



Original Article

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Sleep Duration and the Risk of Type 2 Diabetes: A Community-Based Cohort Study with a 16-Year Follow-up

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Background: We aimed to investigate the moderating effects of obesity, age, and sex on the association between sleep duration and the development of diabetes in Asians.

Methods: We analyzed data from a cohort of the Korean Genome and Epidemiology Study conducted from 2001 to 2020. After excluding shift workers and those with diabetes at baseline, 7,407 participants were stratified into three groups according to sleep duration: ≤5 hours/night, >5 to 7 hours/night (reference), and >7 hours/night. The Cox proportional hazards analyses were used to calculate the hazard ratios (HRs) and 95% confidence intervals (CIs) for incident type 2 diabetes mellitus (T2DM). Subgroup analyses were performed according to obesity, age, and sex.

Results: During 16 years of follow-up, 2,024 cases of T2DM were identified. Individuals who slept ≤ 5 h/night had a higher risk of incident diabetes than the reference group (HR, 1.17; 95% CI, 1.02 to 1.33). The subgroup analysis observed a valid interaction with sleep duration only for obesity. A higher risk of T2DM was observed in the ≤ 5 hours/night group in non-obese individuals, men, and those aged < 60 years, and in the > 7 hours/night group in obese individuals (HRs were 1.34 [95% CI, 1.11 to 1.61], 1.22 [95% CI, 1 to 1.49], and 1.18 [95% CI, 1.01 to 1.39], respectively).

Conclusion: This study confirmed the effect of sleep deprivation on the risk of T2DM throughout the 16-year follow-up period. This impact was confined to non-obese or young individuals and men. We observed a significant interaction between sleep duration and obesity.

Keywords: Diabetes mellitus; Sleep deprivation; Sleep duration; Cohort studies; Sleep; Diabetes mellitus, type 2; Obesity

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INTRODUCTION

The number of adults with diabetes worldwide is expected to reach 783 million by 2045 [1]. Besides an unhealthy diet, smoking, and exercise [2], sleep habits have been noticed due to their metabolic influence on various diseases [3]. Because of modern lifestyles, voluntary sleep restriction due to societal demands has become common [4]. The average duration of sleep dropped from to 8- to 9-hour per night in 1960 to 7-hour per night in 1995 [5]. Consequently, sleep insufficiency is regarded as a global health problem [6].

Prior publications suggested that the quantity of sleep was related to the future risk of type 2 diabetes mellitus (T2DM), but they were inconclusive. Although several meta-analyses have reported a U-shaped association between sleep duration and T2DM risk [7,8], other studies have indicated limited significance whether in short or long sleep [9-12]. This discrepancy could be related to the varied definitions of diabetes and the classification of sleep duration. Additionally, differences in the effect of sleep insufficiency on T2DM risk according to ethnicity have been recognized [7,13]. Most Asian studies investigating sleep insufficiency and T2DM risk have several weak points, including a small number of study subjects, a short follow-up period, and a lack of biochemical laboratory information, resulting in inconsistent findings [14-19].

To explain the relationship between sleep insufficiency and incident diabetes, many investigators have attempted to consider the role of body mass index (BMI) and obesity [11,19,20]. Sleep insufficiency is an important contributor not only to diabetes [21,22] but also to obesity [23-26]. Additionally, obese status itself is associated with an increased risk of T2DM [27]. However, the impact of BMI or obesity on sleep characteristics is variable [11,19,20]. Moreover, as is well known, the deleterious effect of obesity in Asians was found in lower BMI values than in other ethnicities [28], a reliable study for Asians is needed. In addition, previous studies had a weakness in the lack of consideration of co-existing sleep-related disorders, such as sleep apnea [9], and more research is needed to specify the group that needs to adjust sleep time intensively.

Regarding other characteristics, although few studies have explored the effect of adjusting for age or sex [10,12,17,19,29], no one has yet concluded as to the specific population susceptible to sleep insufficiency. We have already demonstrated that the presence of obstructive sleep apnea is associated with impaired glucose metabolism only in non-obese individuals [30]. This study aimed to explore the effect of sleep duration on fu-

ture T2DM risk, with the precise identification of diabetes and other sleep-related disorders. In addition, we tested whether obesity status, age group, and sex can be moderators of this relationship in an ongoing cohort study of Asians with an accumulation of follow-up of more than 20 years.

METHODS

Study design

We analyzed data from the Ansan and Ansung cohort of the Korean Genome and Epidemiology Study (KoGES), an ongoing prospective cohort study started in 2001 ("Exam 1"), whose design, aims, and detailed descriptions have been previously reported [30-32]. The participants, aged 40 to 69 years, were followed up with biennial examinations addressing demographic characteristics, medical history, lifestyle habits, and sleep-related factors.

Among the 5,012 participants of the Ansan Cohort and the 5,018 participants of the Ansung Cohort who attended the KoG-ES from June 2001 to January 2003 ("Exam 1," hereafter referred to as "Baseline"), 2,630 participants were excluded as follows: the presence of diabetes or missing information about diabetes at baseline (n=1,311); missing information for BMI or waist circumference (WC) at baseline (n=13); missing data for a questionnaire about sleep habits (n=277); shift worker at baseline (n=126); no follow-up examination after baseline (n=735). Finally, 7,407 participants were included in this study (Fig. 1).

Before participation, the participants provided written informed consent for the use of their data. The study was reviewed and approved by the Institutional Review Board of the Korea University Ansan Hospital (2006AS0045) and conducted in accordance with the tenets of the Helsinki Declaration of 1975.

Assessment of sleep habit

Information on habitual sleep duration, the presence of snoring, and sleep apnea was collected from the questionnaires with respect to usual weekdays and nights [33]. The question about sleep duration was as follows: "How many hours did you usually sleep during the last month?"

Habitual snoring was defined as snoring ≥4 days/week. A previous study demonstrated substantial reliability for the question of snoring [34]. The presence of sleep apnea was identified from the answer to the questionnaire "Has anyone who sleeps with you ever said that you stopped breathing while snoring?" indicating a witness to sleep apnea.

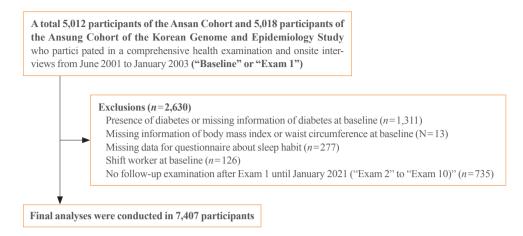


Fig. 1. Flowchart of the selection of participants for this study.

Study outcome

The study endpoint was incident T2DM until January 2020 ("Exam 10"), defined by levels of fasting plasma glucose \geq 126 mg/dL, 2-hour plasma glucose \geq 200 mg/dL during the 75 g oral glucose tolerance test (OGTT), glycated hemoglobin (HbA1c) levels of \geq 6.5%, or a medical history of antidiabetic agents provided for those considered to have T2DM [35]. The follow-up period was defined as the period from the baseline to the first diagnosis of diabetes or the date of the last examination.

Measurements of covariates

Covariate measures were derived from a comprehensive health examination and questionnaire-based interviews at each visit. Smoking status and alcohol consumption were classified as never, former, or current, respectively. Regular exercise was defined as exercise for >30 minutes at least three times a week during the previous month.

The BMI of the participants was calculated as their weight in kilograms divided by the square of their height in meters. WC was measured at the narrowest part between the lower rib and iliac crest in the standing position, and the average of three repeated measurements was calculated. Blood pressure (BP) was measured by trained nurses with the patient in a sitting position using standardized sphygmomanometers after 5 minutes of rest. The average of repeated measurements with a 30-second interval was used.

Venous blood sampling was conducted in the morning (8:00 AM to 9:00 AM) after an overnight fast of at least 8 hours. Plasma glucose, serum total cholesterol, triglycerides high-density lipoprotein cholesterol, and low-density lipoprotein cholesterol levels were measured using an autoanalyzer (ADVIA 1650, Siemens, Tarrytown, NY, USA). Insulin levels were measured

using an immunoradiometric assay kit (INS-IRMA Kit, Bio-Source, Nivelles, Belgium), with a Packard γ counter system. Insulin resistance was estimated using the homeostasis model assessment of insulin resistance (HOMA-IR) and calculated using the following formula: HOMA-IR=fasting insulin (IU/mL)×fasting glucose (mmol/L)/22.5 [36].

Cardiovascular disease (CVD) was defined as a self-reported history of myocardial infarction, congestive heart failure, coronary artery disease, or peripheral vascular disease. Prescribed medications for diabetes, hypertension, dyslipidemia, and CVD were reported.

Statistical analysis

Data are presented as the mean \pm standard deviation, number (%), or median (interquartile range). The baseline characteristics of the participants were compared according to the presence of T2DM and sleep duration using the analysis of variance for continuous variables and the chi-square test for categorical variables. The skewed variables were log-transformed for analysis. Based on sleep duration, study participants were stratified into three groups as follows: (1) \leq 5 hours night (n=881), (2) >5 to 7 hours/night (n=4,060), and (3) >7 hour/night (n=2,466).

To evaluate the risk of diabetes according to the three groups of sleep duration, we performed a Cox proportional hazards analysis with a >5 to 7 hour/night group as the reference group. The risk is shown as a hazard ratio (HR) and Bonferroni-adjusted 95% confidence interval (CI). We adjusted for various cardiometabolic factors: age, sex, BMI, study site (Ansan or Ansung), alcohol consumption, smoking, regular exercise, presence of hypertension or CVD, use of antihyperlipidemic drugs, fasting glucose levels, HOMA-IR, habitual snoring, and sleep apnea.

To explore the factors modifying the association between sleep

Table 1. Baseline Characteristics of the Study Participants according to Sleep Duration

| Characteristic | \leq 5-hr/night (n =881) | >5 to 7-hr/night (<i>n</i> =4,060) | >7-hr/night (<i>n</i> =2,466) | P value ^a |
|----------------------------------|-------------------------------|-------------------------------------|--------------------------------|----------------------|
| Age, yr | 52.4±9 ^b | 50.7±8.4 | 52.6±9.0° | < 0.001 |
| Male sex | 349 (39.6) ^{b,d} | 1,928 (47.5) | 1,189 (48.2) | < 0.001 |
| WC, cm | 81.8±9.2 | 81.7±8.7 | 82.6±8.5° | < 0.001 |
| BMI, kg/m ² | 24.6±3.2 | 24.5±3.1 | 24.4 ± 3.0 | 0.174 |
| Systolic BP, mm Hg | 120.4 ± 19.0 | 119.4 ± 17.7 | $121.6 \pm 18.0^{\circ}$ | < 0.001 |
| Diastolic BP, mm Hg | 79.7 ± 11.9^{d} | 79.3 ± 11.3 | $80.8 \pm 11.3^{\circ}$ | < 0.001 |
| Fasting glucose, mg/dL | 82.5 ± 8.4 | 82.7±8.5 | 82.8 ± 8.6 | 0.692 |
| Glycated hemoglobin, % | 5.6±0.3 | 5.5±0.3 | 5.5 ± 0.3 | 0.059 |
| Total cholesterol, mg/dL | 190.2±35.2 | 189.4 ± 34.3 | 190±33.4 | 0.713 |
| Triglyceride, mg/dL | 132.2 (100–188) | 126.8 (94.6–178) | 136 (100–187) ^c | < 0.001 |
| HDL-C, mg/dL | 45.2 ± 10.7^d | 45.1 ± 10.1 | 44.6±9.9 | 0.096 |
| LDL-C, mg/dL | 113.5±33 | 114.2±31.8 | 114±31.4 | 0.841 |
| HOMA-IR | 1.4 (1-1.9) | 1.4 (1-1.9) | 1.4 (1–2) | 0.387 |
| Creatinine, mg/dL | 0.6 ± 0.2^{b} | 0.6 ± 0.2 | 0.6 ± 0.2 | 0.002 |
| eGFR, mL/min/1.73 m ² | 110.2 ± 11.6 | 110.9 ± 12.2 | 110±12.6° | 0.012 |
| Current drinker | 366 (41.5) ^{b,d} | 1,975 (48.7) | 1,188 (48.2) | 0.001 |
| Current smoker | 182 (20.7) ^d | 960 (23.7) | 673 (27.3)° | < 0.001 |
| Regular exercise | 243 (27.6) ^d | 995 (24.5) | 381 (15.5) ^c | < 0.001 |
| Antihypertensive drug | 88 (10) | 339 (8.4) | 248 (10.1)° | 0.042 |
| Antihyperlipidemic drug | 2 (0.2) | 19 (0.5) | 6 (0.2) | 0.265 |
| Cardiovascular disease | 16 (1.8) | 79 (2) | 58 (2.4) | 0.459 |
| Habitual snoring | 121 (13.7) | 561 (13.8) | 331 (13.4) | 0.902 |
| Sleep apnea | 99 (11.3) | 524 (13) | 305 (12.5) | 0.364 |

Values are expressed as mean ± standard deviation, number (%), or median (interquartile range).

WC, waist circumference; BMI, body mass index; BP, blood pressure; HDL-C, high-density lipoprotein cholesterol; LDL-C, low-density lipoprotein cholesterol; HOMA-IR, homeostasis model assessment of insulin resistance; eGFR, estimated glomerular filtration rate.

duration and the risk of diabetes, we performed a subgroup analysis after stratifying the participants according to sex, age <60 and \geq 60 years, and obesity, defined as a BMI of \geq 25 kg/m² [37].

In addition, we conducted a sensitivity analysis in participants in the same sleep duration group at the baseline and final visit before the study outcome to explore the effect of persistent sleep duration. As the questionnaire regarding sleep habits was included only in Exam 1 in the Ansung cohort, we conducted this sensitivity analysis in the Ansan cohort, excluding participants with incident diabetes in Exam 2 performed from January 2003 to February 2005. Data from a total of 1,772 participants were extracted. We repeated the above-mentioned Cox proportional hazards analyses to calculate the HR and 95% CI of incident diabetes, using the group that slept persistently >5 to 7 hours/ night as a reference category.

Statistical analyses were performed using SAS version 9.1 for Windows (SAS Institute Inc., Cary, NC, USA). P values were corrected using Tukey's method because of multiple testing, and a P<0.05 was considered to be statistically significant.

RESULTS

Table 1 shows the baseline characteristics of the participants according to sleep duration. Individuals in the >5 to 7 hours/night group were younger and had a higher proportion of men and drinkers than those in the ≤5-hour/night group, whereas the

^aP values were derived from analysis of variance for continuous variables and chi-square tests for categorical variables. P values were corrected using post hoc Tukey's method due to multiple testing. The right-skewed variables were log-transformed for the analysis; bP<0.05, in comparison between \leq 5 and >5 to 7 hours; $^{\circ}P < 0.05$, comparison between >5 and 7 vs. >7 hours; $^{\circ}P < 0.05$, compared between ≤ 5 and >7 hours.

| | | Multivariate adjusted HR (95% CI) | | D 0 1 1 1 1 |
|-----------------------------|-------------------|-----------------------------------|------------------|----------------------------------|
| Sleep duration, hr/day | Incident diabetes | Model 1 | Model 2 | — P for interaction ^a |
| Total $(n=7,407)$ | | | | |
| $\leq 5 (n=881)$ | 272 (30.87) | 1.19 (1.04–1.36) | 1.17 (1.02–1.33) | |
| >5 to 7 ($n=4,060$) | 1,096 (27) | 1 (Ref) | 1 (Ref) | |
| >7 (<i>n</i> =2,466) | 656 (26.6) | 1.01 (0.91–1.11) | 1.03 (0.93–1.14) | |
| Women $(n=3,941)$ | | | | 0.174 |
| $\leq 5 (n=532)$ | 149 (28.01) | 1.14 (0.95–1.37) | 1.12 (0.94–1.35) | |
| >5 to 7 ($n=2,132$) | 536 (25.14) | 1 (Ref) | 1 (Ref) | |
| >7 (<i>n</i> =1,277) | 342 (26.78) | 1.12 (0.97–1.29) | 1.1 (0.96–1.27) | |
| Men (<i>n</i> =3,466) | | | | |
| $\leq 5 (n=349)$ | 123 (35.24) | 1.25 (1.03–1.53) | 1.22 (1-1.49) | |
| >5 to 7 ($n=1,928$) | 560 (29.05) | 1 (Ref) | 1 (Ref) | |
| >7 (n=1,189) | 314 (26.41) | 0.91 (0.79–1.05) | 0.97 (0.84-1.11) | |
| Age $<60 (n=5,645)$ | | | | 0.130 |
| $\leq 5 (n=630)$ | 185 (29.37) | 1.21 (1.03–1.42) | 1.18 (1.01–1.39) | |
| >5 to 7 ($n=3,257$) | 855 (26.25) | 1 (Ref) | 1 (Ref) | |
| >7 (n=1,758) | 467 (26.56) | 1.06 (0.95–1.19) | 1.09 (0.97–1.22) | |
| Age \geq 60 (n =1,762) | | | | |
| $\leq 5 (n=251)$ | 87 (34.66) | 1.19 (0.93–1.53) | 1.2 (0.94–1.54) | |

1 (Ref)

0.9(0.75-1.1)

1.34 (1.11–1.61)

0.9(0.78-1.04)

1.06 (0.88-1.29)

1.12 (0.98-1.28)

1 (Ref)

1 (Ref)

Values are expressed as number (%). Model 1 was adjusted for age, sex, study site, body mass index, alcohol consumption, smoking, regular exercise, and fasting glucose levels; Model 2 was adjusted for age, sex, study site, body mass index, alcohol consumption, smoking, regular exercise, homeostasis model assessment-estimated insulin resistance, presence of hypertension or cardiovascular disease, use of antihyperlipidemic drugs, presence of habitual snoring, and sleep apnea.

HR, hazard ratio; CI, confidence interval; BMI, body mass index.

^aThe interaction P values between the sleep duration group and the corresponding characteristic were adjusted by age, sex, and study site.

241 (30.01)

189 (26.69)

144 (28.86)

519 (21.77)

300 (20.08)

128 (33.51)

577 (34.43)

356 (36.63)

>7-hour/night group consisted of those who were older, had higher BP and triglyceride levels and had a higher proportion of smokers and those who were taking antihypertensive drugs.

During the 16 years (interquartile range, 7.8 to 17.8) of median follow-up periods, 2,024 patients with T2DM were identified. In the results of the multivariate adjusted Cox analysis (Table 2), individuals who slept \leq 5 hours/night had a higher risk of incident diabetes than those who slept >5 to 7 hours/night. However, there was no impact of >7 hours/night of sleep on the risk of T2DM.

1 (Ref)

0.92 (0.76-1.12)

1.34 (1.11–1.61)

0.91 (0.79-1.06)

1.03 (0.85-1.25)

1.13 (1.01-1.33)

1 (Ref)

1 (Ref)

0.002

In the subgroup analysis, no valid interaction was observed at 60 years of age or by sex (Table 2). A marked increased risk of T2DM in ≤5 hours/night of sleep was found in men and individuals aged less than 60 years. However, in the case of obesity

>5 to 7 (n=803)

BMI $< 25 \text{ kg/m}^2 (n=4,377)$

BMI \geq 25 kg/m² (n=3,030)

>7 (n=708)

 $\leq 5 (n=499)$

>7 (n=1,494)

 $\leq 5 (n=382)$

>7 (n=972)

>5 to 7 (n=2,384)

>5 to 7 (n=1,676)

Table 3. Hazard Ratios and 95% Confidence Intervals for Incident Diabetes according to Persistent Habitual Sleep Duration

| Clean dynation hu/dex | Incident dichetes | Multivariate adjusted HR (95% CI) | | D. C i t |
|---|-------------------|-----------------------------------|------------------|----------------------------------|
| Sleep duration, hr/day | Incident diabetes | Model 1 | Model 2 | — P for interaction ^a |
| Total (n=1,772) | | | | |
| $\leq 5 (n=224)$ | 77 (34.48) | 1.37 (1.06–1.76) | 1.32 (1.02–1.7) | |
| >5 to 7 ($n=1,305$) | 361 (27.66) | 1 (Ref) | 1 (Ref) | |
| >7 (<i>n</i> =243) | 65 (26.75) | 0.97 (0.74–1.26) | 1.07 (0.82–1.4) | |
| BMI $<25 \text{ kg/m}^2 (n=1,048)$ | | | | 0.161 |
| \leq 5 (n =127) | 42 (33.07) | 1.64 (1.16–2.32) | 1.64 (1.16–2.32) | |
| >5 to $7 (n=771)$ | 168 (21.79) | 1 (Ref) | 1 (Ref) | |
| >7 (n=150) | 38 (25.33) | 1.15 (0.8–1.65) | 1.26 (0.88–1.81) | |
| BMI \geq 25 kg/m ² (n =724) | | | | |
| $\leq 5 (n=97)$ | 35 (36.08) | 1.1 (0.76–1.6) | 1.13 (0.78–1.64) | |
| >5 to 7 ($n=534$) | 193 (36.14) | 1 (Ref) | 1 (Ref) | |
| >7 (<i>n</i> =93) | 27 (29.03) | 0.77 (0.51–1.15) | 0.86 (0.57-1.29) | |

Values are expressed as number (%). Model 1 was adjusted for age, sex, study site, body mass index, alcohol consumption, smoking, regular exercise. and fasting glucose levels; Model 2 was adjusted for age, sex, study site, body mass index, alcohol consumption, smoking, regular exercise, homeostasis model assessment-estimated insulin resistance, presence of hypertension or cardiovascular disease, use of antihyperlipidemic drugs, presence of habitual snoring, and sleep apnea.

HR, hazard ratio; CI, confidence interval; BMI, body mass index.

status, there was a significant interaction between sleep duration and obesity status (P=0.02, with adjustment by age, sex, and study site). In the fully adjusted model 2, the ≤ 5 hours/night group showed a higher risk of T2DM than the >5 to 7 hours/ night group in non-obese individuals, whereas the >7 hours/ night group only showed a higher risk in obese individuals.

We also analyzed whether this relationship was still significant in those with persistent short sleep duration (Table 3), and those who slept ≤ 5 hours/night consistently exhibited a higher risk of T2DM than those in the >5 to 7 hours/night group. There was a 64% increased risk of diabetes in those not obese and with persistent short sleep duration; however, the interaction according to obesity status was not significant.

DISCUSSION

The present study confirmed the effect of sleep deprivation on the risk of T2DM throughout a 16-year follow-up period, and the analysis of persistent sleep habits reinforced these findings. This impact was confined to men, individuals aged <60 years, and non-obese individuals. We also suggested a significant interaction between sleep duration and obesity.

Regarding sleep duration and T2DM risk, previous publications on Asians have reported inconsistent findings [14,1620,29]. A study from Japan indicated the predictive value of short sleep (<5.5 hours/day), but this finding was limited to individuals aged less than 60 years [17]. More recently, a nested case-control study in a rural Chinese population indicated that both shorter (≤6 hours/day) and longer (8 to 9 hours/day) sleep durations increased the risk of T2DM [19]. However, after adjusting for lifestyle habits and BMI, this effect was maintained only in women who slept ≤ 6 hours/day.

Several limitations of the previous study design have been reported [14-17]. First, the identification of diabetes was mainly based on the self-report of a physician's diagnosis of diabetes or fasting blood glucose levels [16,18,19]. In this study, however, we identified incident T2DM using 75 g OGTT, HbA1c level, and a history of antidiabetic agent prescription, thereby enhancing the accuracy of the diagnosis. Second, the follow-up period was less than 10 years [14-19] which is not sufficient to demonstrate the long-term effect of sleep insufficiency. Considering that it was previously known that the risk of future diabetes related to poor sleep tends to decrease with the duration of followup [12,22], the results of our study, which has the longest follow-up period among Asian studies, are meaningful. Third, covariates were confined to the questionnaires or lifestyle habits [14-19]. Although a large longitudinal epidemiological study in Korea revealed that short sleep duration could affect the risk of

^aThe interaction P value between groups of sleep duration and obesity status was adjusted by age, sex, and study site.

diabetes by adapting various covariates, the median follow-up period was 22 months [20]. In our study, we could adjust for comprehensive factors, including cardiometabolic parameters and the HOMA-IR. Lastly, few studies have evaluated the impact of other sleep characteristics such as afternoon napping or sleep quality [14,18]. However, we could minimize the bias from other sleep habits by adjusting sleep-related variables (e.g., snoring and objective sleep apnea).

Few studies have documented the influence of obesity on the association between sleep and diabetes. In multi-ethnic cohort study, the higher risk for T2DM observed in long sleepers was valid only in participants with BMI $\geq 25 \text{ kg/m}^2$ [10]. More recently, a prospective Mexican American cohort study showed the partially valid effect of short sleep on the 3-year-risk of T2DM in patients with a BMI <30 kg/m² [12]. The present study widened prior evidence with a long follow-up period in the Asian populations. We suggest that obesity might modify the significance of sleep insufficiency on the risk of T2DM in the opposite direction.

The causal relationship between short sleep duration and the risk of obesity or T2DM is speculated to involve several possible mechanisms. Insufficient sleep enhances a positive energy balance by increasing energy intake to compensate for the lack of biological processes that normally occur during nighttime sleep, resulting in weight gain [38]. Sleep deprivation can activate the sympathetic nervous system and induce cortisol dysregulation, which is attributed to insulin resistance and visceral adiposity [39]. Short sleep can disturb metabolic processes by upregulating daily levels of ghrelin and insulin resistance and declining daily levels of leptin, which increases appetite and consequently increases the risk of obesity and T2DM [40-42]. In addition, inflammation and oxidative stress, such as tumor necrosis factor-α and interleukin-6, may also play a role [43]. The reasons for the limited significance observed only in nonobese individuals is not well known; however, we can speculate that the mechanisms above have already been activated in obese populations, thereby the impact of short sleep may be more prominent in those not obese since. The results of the analysis of persistent sleep duration confirmed the association between sleep deprivation and increased risk of T2DM.

Among obese individuals, the >7 hours/night group showed a high risk of T2DM. This can be explained by the fact that long periods of sleep may be derived from poor sleep quality, sleep fragmentation, and frequent awakenings [29]. Additionally, long sleep duration can indicate poor metabolic characteristics, such as depression, poor economic status, low physical activity, and obesity [44,45].

In a subgroup analysis, we found an impact of a short night's sleep on the risk of T2DM, particularly in men and individuals aged < 60 years, without significance in the interaction analysis. This is in line with prior evidence suggesting that a prominent association between short or long sleep duration and T2DM risk was confined to young individuals [10,17,29]. Similarly, in prior research, short sleep duration was not associated with hypertension or obesity in the elderly [46,47]. Differences in sleep time, architecture, quality, and psychological factors are thought to contribute to attenuation in advanced age [48]. Finally, only men were found to be sensitive to short sleep duration. Regarding variance in sex, inconsistent results have been observed in previous studies [12,19]. Although we cannot explain the exact mechanisms underlying these findings, the current study suggests that maintaining adequate sleep might be particularly important for young or middle-aged men to prevent the development of future diabetes.

However, we acknowledge several limitations. First, the identification of sleep duration and sleep habits was derived from a questionnaire that may overestimate or underestimate short sleep duration [49]. However, objective evaluation of sleep duration in large populations is impossible due to cost, and a moderate correlation (r=0.47) between self-reported sleep duration and objective time measured by actigraphy was noted [49], and a few prior investigations have documented the usefulness of questionnaire-derived sleep habits to various outcomes [33,50]. Second, although we explored the relationship between sleep habits and sleep duration, we did not consider the changes in other biomarkers during the follow-up period. Finally, we did not study the influence of insomnia or exact changes in sleep duration in every examination, although we analyzed the impact of persistent short sleep duration, defined as the baseline and final visit before the study outcome.

In conclusion, this study indicated that sleep duration ≤ 5 hours was significantly associated with a higher risk of developing diabetes in non-obese or young individuals and men after adjusting for some sleep habits. The impact of sleep duration varies according to obesity status with a valid interaction between sleep duration and obesity.

CONFLICTS OF INTEREST

No potential conflict of interest relevant to this article was reported.

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