



# Risk of Incident New-Onset Arterial Hypertension After COVID-19 Recovery: A Systematic Review and Meta-analysis

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## Abstract

**Introduction** Arterial Hypertension (HT) has been described as a common comorbidity and independent risk factor of short-term outcome in COVID-19 patients. However, data regarding the risk of new-onset HT during the post-acute phase of COVID-19 are scant.

**Aim** We assess the risk of new-onset HT in COVID-19 survivors within one year from the index infection by a systematic review and meta-analysis of the available data.

**Methods** Data were obtained searching MEDLINE and Scopus for all studies published at any time up to February 11, 2023, and reporting the long-term risk of new-onset HT in COVID-19 survivors. Risk data were pooled using the Mantel–Haenszel random effects models with Hazard ratio (HR) as the effect measure with 95% confidence interval (CI). Heterogeneity among studies was assessed using  $I^2$  statistic.

**Results** Overall, 19,293,346 patients (mean age 54.6 years, 54.6% males) were included in this analysis. Of them, 758,698 survived to COVID-19 infection. Over a mean follow-up of 6.8 months, new-onset HT occurred to 12.7 [95% CI 11.4–13.5] out of 1000 patients survived to COVID-19 infection compared to 8.17 [95% CI 7.34–8.53] out of 1000 control subjects. Pooled analysis revealed that recovered COVID-19 patients presented an increased risk of new-onset HT (HR 1.70, 95% CI 1.46–1.97,  $p < 0.0001$ ,  $I^2 = 78.9\%$ ) within seven months. This risk was directly influenced by age ( $p = 0.001$ ), female sex ( $p = 0.03$ ) and cancer ( $p < 0.0001$ ) while an indirect association was observed using the follow-up length as moderator ( $p < 0.0001$ ).

**Conclusions** Our findings suggest that new-onset HT represents an important post-acute COVID-19 sequelae.

**Keywords** COVID-19 · Long COVID · Arterial hypertension · Follow-up

## 1 Introduction

Since the beginning of COVID-19 outbreak, arterial hypertension (HT) was recognized as one of the most common comorbidities in COVID-19 patients as well as an independent predictor of short-term mortality and severe disease

[1–4]. Recent analyses have demonstrated an increased risk of cardiovascular sequelae after COVID-19 recovery compared to the general population not exposed to SARS-CoV-2 infection [5–9]. However, data regarding the risk of development new-onset HT as a post-acute COVID-19 sequelae remain scant. Aim of the present manuscript is to assess the risk of new-onset HT in COVID-19 recovered patients by performing a systematic review and meta-analysis of the available data.

## 2 Methods

### 2.1 Study Design

This study followed the Preferred Reporting Items for Systematic Reviews and Meta-analyses (PRISMA) reporting

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guideline (Supplementary file 1) [10]. Data were obtained searching MEDLINE and Scopus for all studies published at any time up to February 11, 2023, and reporting the risk of new-onset HT in COVID-19 recovered patients diagnosed between 4 months (minimum follow-up length of revised investigations) and a maximum of 12 months post discharge (maximum follow-up length of revised studies) after the infection. The diagnosis of new-onset HT was performed after reviewing the clinical records of enrolled subjects using the ICD-10 codes I10-I16. In the revised manuscripts, this group of patients were matched and compared to contemporary cohorts, defined as subjects who did not experience the SARS-CoV-2 infection and developed a pericarditis in the same follow-up period.

## 2.2 Data Extraction and Quality Assessment

The selection of studies included in our analysis was independently conducted by two authors (M.Z., C.B.) in a blinded fashion. Any discrepancies in study selection were resolved by consulting a third author (G.R.). The following MeSH terms were used for the search: “Arterial Hypertension” AND “COVID-19 sequelae” OR “Arterial Hypertension” AND “COVID-19”. Moreover, we searched the bibliographies of the target studies for additional references. Specifically, inclusion criteria were: (i) studies enrolling subjects with previous confirmed COVID-19 infection (ii) providing the hazard ratio (HR) and relative 95% confidence interval (CI) for the risk of new-onset HT after the infection compared to contemporary control cohorts. Conversely, case reports, review articles, abstracts, editorials/letters, and case series with less than 10 participants were excluded. Data extraction was independently conducted by two authors (M.Z., G.R.). For all the reviewed investigations we extracted, when provided, the number of enrolled patients, the mean age, the gender, the prevalence of cardiovascular comorbidities such as arterial hypertension (HT), diabetes mellitus (DM), chronic obstructive pulmonary disease (COPD), chronic kidney disease (CKD), obesity, pre-existing heart failure (HF), cerebrovascular disease and the length of follow-up. The quality of included studies was graded using the Newcastle-Ottawa quality assessment scale (NOS) [11].

## 2.3 Data Synthesis and Analysis

Continuous variables were expressed as mean while categorical variables were presented as numbers and relative percentages. New-onset HT risk data were pooled using the Mantel–Haenszel random effects models with Hazard ratio (HR) as the effect measure with 95% confidence interval (CI). Heterogeneity among studies was assessed using Higgins and Thomson  $I^2$  statistic. Specifically, the  $I^2$

values correspond to the following levels of heterogeneity: low (< 25%), moderate (25–75%) and high (> 75%). The presence of potential publication bias was verified by visual inspection of the funnel plot. Due to the low number of the included studies (< 10), small-study bias was not examined as our analysis was underpowered to detect such bias. However, a predefined sensitivity analysis (leave-one-out analysis) was performed removing 1 study at the time, to evaluate the stability of our results regarding the risk of new-onset HT. To further appraise the impact of potential baseline confounders, a meta-regression analysis was also performed. All meta-analyses were conducted using Comprehensive Meta-Analysis software, version 3 (Biostat, USA).

## 3 Results

### 3.1 Search Results and Included Studies

A total of 2497 articles were obtained using our search strategy. After excluding duplicates and preliminary screening, 230 full-text articles were assessed for eligibility. Among them, 225 studies were excluded for not meeting the inclusion criteria, leaving 5 investigations fulfilling the inclusion criteria (Fig. 1) [12–16].

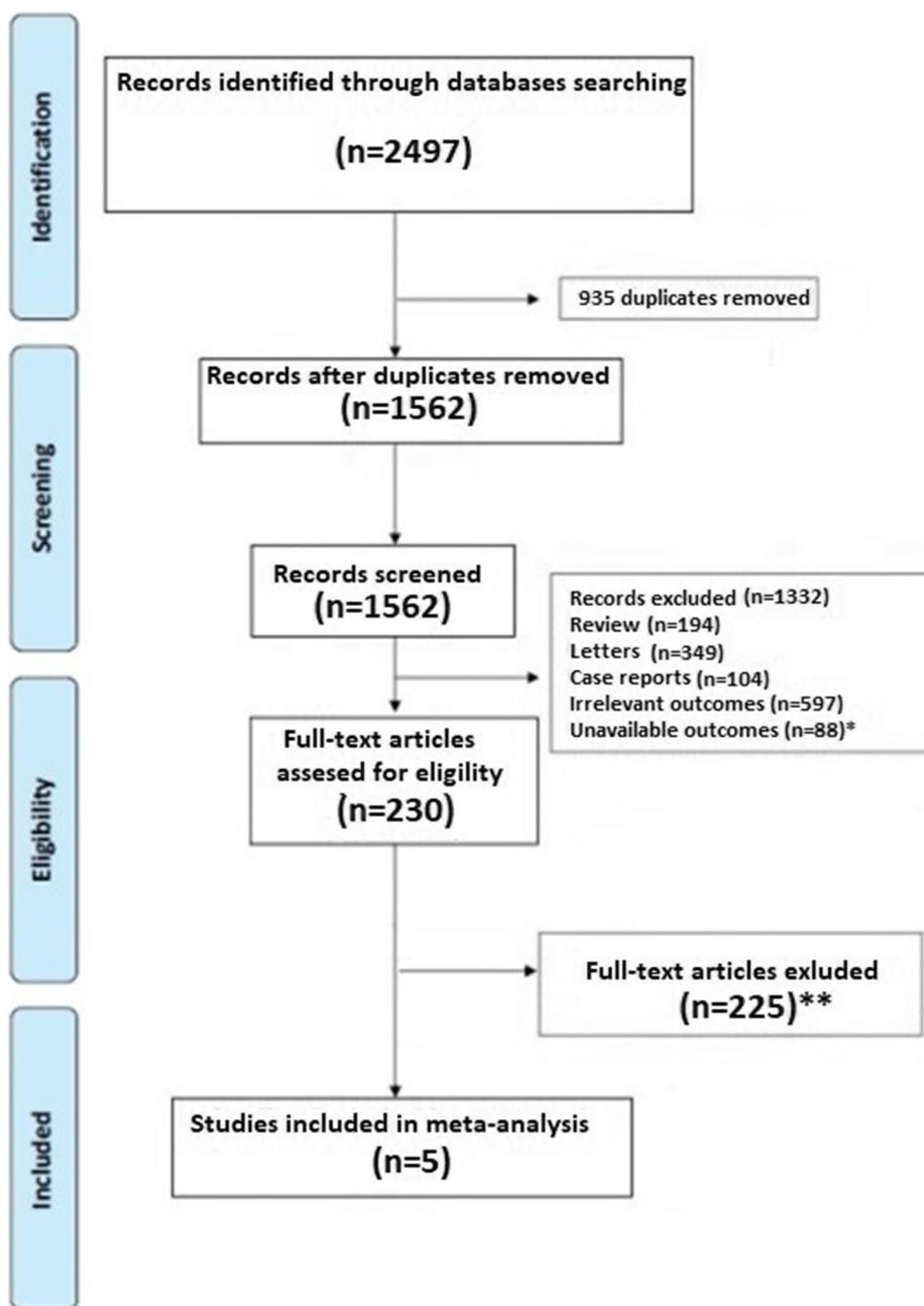
### 3.2 Characteristics of the Population and Quality Assessment

Overall, 19,293,346 patients (mean age 54.6 years, 54.6% males) were included in this analysis [12–16]. Among them 758,698 had confirmed COVID-19 infection. The general characteristics of the studies included are presented in Table 1. Although the demographic characteristics and concomitant comorbidities were not systematically recorded in all investigations, the cohorts mainly consisted of middle-aged patients. The mean length of follow-up was 6.8 months. Over the follow-up period, incident pericarditis occurred to 12.7 [95% CI 11.4–13.5] out of 1000 patients and to 8.17 [95% CI 7.34–8.53] out of 1000 patients in survived to COVID-19 infection and in control subjects, respectively. Quality assessment showed that all studies were of moderate-high quality according to the NOS scale (Table 1) [11].

### 3.3 Risk of New-Onset Arterial Hypertension After COVID-19 Recovery

During the follow-up period, recovered COVID-19 patients presented an increased risk of new-onset HT (HR: 1.70, 95% CI 1.46–1.97,  $p < 0.0001$ ,  $I^2 = 78.9\%$ ) compared to subjects who did not experience COVID-19 infection but developed a pericarditis over the same period (Fig. 2). The funnel plot is presented in Supplemental File 2. The sensitivity analysis

**Fig. 1** Flow diagram of selected studies for the meta-analysis according to the Preferred reporting items for systematic reviews and meta-analyses (PRISMA). \*Articles excluded because not provided data on new-onset arterial hypertension; \*\*Articles excluded because not provided Hazard ratio for new-onset arterial hypertension



confirmed yielded results reporting an HR ranging between 1.62 (95% CI 1.53–1.71,  $p < 0.0001$ ,  $I^2: 65.6\%$ ) and 1.80 (95% CI 1.56–2.07,  $p < 0.0001$ ,  $I^2: 71.8\%$ ), indicating that the obtained results were not driven by any single study.

### 3.4 Meta-regression for the Risk of New-Onset Arterial Hypertension

A meta-regression analysis showed a significant direct relationship for the risk of new-onset HT using age ( $p = 0.001$ ), female sex ( $p = 0.03$ ) and cancer ( $p < 0.0001$ ) as moderators,

while an indirect association was observed when the follow-up length ( $p < 0.0001$ ) was adopted as moderating variables (Table 2).

## 4 Discussion

Our meta-analysis, based on a large population of more than 19 million people, showed that COVID-19 recovery subjects had an additional 70% risk of developing new-onset HF within 7 months from the acute infection. As demonstrated

**Table 1** General characteristics of the population reviewed

Authors	Study Design	Sample size n	Cases n	Controls n	Age (years)	Males n (%)	HT n (%)	DM n (%)	COPD n (%)	CKD n (%)	Obesity n (%)	HF n (%)	Cancer n (%)	Cerebrovascular disease n (%)	FW-length (months)	NOS
Cohen et al. [12]	R	2,895,943	133,366	2,762,577	75.7	1,227,545 (42.0)	2,081,772 (72.0)	938,043 (32.7)	578,650 (20) <sup>a</sup>	528,314 (14.0)	478,902 (17.0)	334,654 (12)	418,700 (14.4)	364,782 (13.0)	4	8
Daugherty et al. [13]	R	9,247,505	266,586	8,980,919	42.4	4,640,393 (50.2)	NR	521,699 (5.6)	NR	NR	NR	NR	NR	NR	6	6
Al-Aly et al. [14]	R	5,017,431	33,940	4,983,491	64.9	4,512,549 (89.9)	NR	1,275,370	NR	423,329 (8.4)	NR	NR	371,526 (7.4)	288,266 (5.7)	6	7
Mizrahi et al. [15]	R	1,913,324	320,857	1,792,150	25.0 <sup>e</sup>	1,48,095 (49.4) <sup>c</sup>	22,490 (7.5) <sup>c</sup>	11,412 (3.8) <sup>c</sup>	1519 (0.5) <sup>c</sup>	4756 (1.6) <sup>c</sup>	63,418 (21.1) <sup>c</sup>	NR	6953 (2.3) <sup>c</sup>	NR	6	8
Tisler et al. [16]	R	19,460	3949	15,511	65.1	8900 (45.7)	10,601 (54.4)	3341 (17.1)	2205 (13.9) <sup>a</sup>	870 (4.4)	1496 (7.6)	NR	3998 (20.5)	578 (2.9)	12	8

HT Arterial Hypertension, DM Diabetes Mellitus, COPD chronic obstructive pulmonary disease, CKD Chronic Kidney disease, HF Heart failure, FW Follow-up, NR Not reported, R Retrospective

<sup>a</sup>Defined as Chronic Pulmonary disease

<sup>b</sup>Only DM type 2. R: Retrospective

<sup>c</sup>Referred to propensity matched cohort

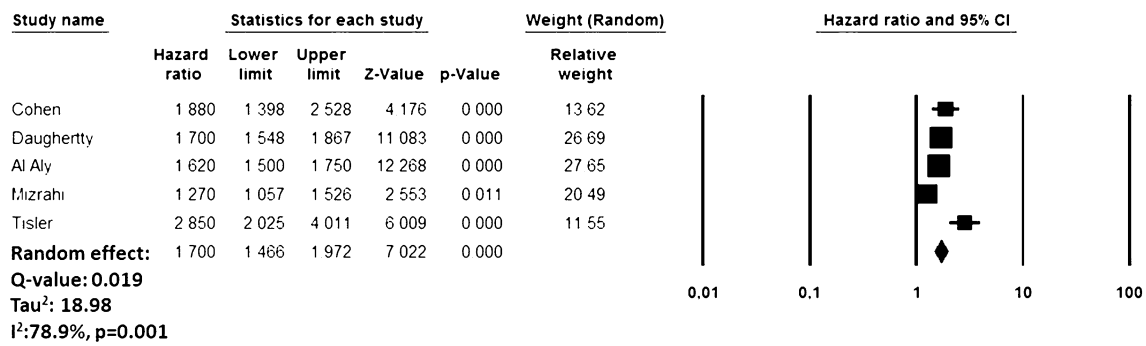


Fig. 2 Forest plot investigating the risk of new-onset arterial hypertension after COVID-19 Infection

Table 2 Meta-regression analysis for the risk of new-onset arterial hypertension after COVID-19 infection

Items	Coeff.	95% CI	p
Age (years)	0.008	0.001 to 0.018	0.001
Females (%)	0.003	0.008 to 0.014	0.03
DM (%)	0.011	- 0.033 to 0.054	0.62
CKD (%)	0.013	- 0.047 to 0.074	0.66
Cancer (%)	0.040	0.021 to 0.059	<0.0001
Follow-up length (months)	0.073	- 0.153 to - 0.005	<0.0001

CI Confidence interval, DM diabetes mellitus, Chronic kidney disease

by the meta-regression, this risk was directly influenced by age, female sex and cancer and resulted higher in the early post-acute phase of the infection. The moderate heterogeneity observed may be also partly explained by the differences in baseline characteristics of the population enrolled, pre-existing cardiovascular risk factors and/or chronic conditions and immunization against COVID-19 infection. However, sensitivity analysis confirmed the robustness of our results.

The revised population had a lowest incidence rate of new-onset HT compared to the previous analyses based on the general population which vary between 15.3 and 47.3 per 1000 people [17–19]. This difference could be partially explained by the fact that the population revised was relatively young, having a median age of 54 years. Furthermore, potential sex and racial differences [20, 21] as well as the absence of a systematic blood pressure screening may have significantly contributed to the lower incidence rate observed. Indeed, in the article reviewed, the new-onset HT was identified using the relative ICD-10 codes. However, our findings are in accordance with those presented by Zhang et al. [9] and Ogungbe et al. [22], which demonstrated an increased risk for essential HT after COVID-19 infection.

Unfortunately, the revised studies did not systematically report data regarding the left ventricular ejection fraction or the presence of target organ damage (TOD), limiting the

possibility to further characterize the reported risk. However, to the best of our knowledge, our analysis represents the first attempt to comprehensively assess the risk of new-onset HT in the post-acute phase of COVID-19 subjects. Moreover, it is also true that when the death rate is high from causes other than the disease of interest, the incidence rates of the illness may be overestimated in traditional Kaplan–Meier survival analysis due to existence of competing risks [23].

Currently, pathophysiological mechanisms potentially involved in the onset of new HT remains poorly understood. However, it seems that viral persistence, delayed resolution of inflammation or residual organ damage after the acute phase of infection as well as renin-angiotensin-aldosterone system (RAAS) dysregulation and previous autoimmune conditions, may represent potential contributors, alone or together [24–26]. From our data it seems that female sex influences the risk of new-onset HT sequelae after COVID-19 infection. This difference may be partially explained by the known differences in immune system function between females and males [27]. Indeed, females generally mount more rapid and robust innate and adaptive immune responses, which can protect them from initial infection and severity. Wang et al. [28], analyzing the long-term cardiovascular sequelae in COVID-19 survivors evidenced that women had a higher risk of cardiovascular events, including myocarditis, ischemic cardiomyopathy and atrial fibrillation compared to the sex-matched control subjects.

Unfortunately, most of the studies reviewed did not evaluate or report granular data by sex, which limited sex-specific sub-analysis in our analysis. Similarly, few data were provided regarding the potential association between the severity of the index COVID-19 infection and the subsequent risk of new-onset HT. However, data from Al-Aly et al. [14] and Tisler et al. [16] suggested that this risk significantly increase in patients hospitalized and in those requiring intensive care treatments compared to those not hospitalized. Indeed, the hospitalization setting may indirectly reflect the severity of COVID-19 infection. Moreover, Mizrahi et al. [15] observed



an increased risk of new-onset HT during the recovery period in unvaccinated subjects compared to those immunized.

Our data, based on a large cohort, may be useful for designing strategies to minimize the risk of new HT onset during the post-acute phase after COVID-19 infection, although our results must be considered preliminary and cannot be directly translated into clinical practice regarding the use of antihypertensive drugs. However, our findings may have several implications for daily clinical practice and future research. Indeed, being the risk of new HT onset not limited just to the acute phase of the infection but extends in the medium/long-term period [29], indicating the need to prevent the infection and to consider post COVID-19 patients at future of HT and related cardiovascular consequences.

Changes in HT incidence over time represents an important issue which require urgent actions as increased awareness, adequate screening, and preventive measure to minimize the risk of future cardiovascular disease.

#### 4.1 Limitations

Our study has several limitations related to the observational nature of the studies reviewed and their own limitations with all inherited bias. Potential underestimation could derive from detection bias considering that patients were not systematically screened HT and per se, these conditions is clinically silent. Moreover, sampling bias by the competing risk of death may also have led to underestimation of the real cumulative incidence of new HT cases. Our data must be carefully interpreted considering that the reviewed studies mainly based their search strategy on the revision of hospital admission and screening of ICD-10 codes during the follow-up period; indeed, the utility of such data largely depends on their accuracy and reliability. In fact, the absence of a dedicated screening for HT may have determined some bias in our analysis, evidencing the lower limit of such phenomenon. Furthermore, no data regarding the type of HT, the dipping status and the chronic use of anti-hypertensive drugs limited our ability to perform further sub analyses.

We can neither exclude those geographical differences, also related to the quality of care, may have influenced our findings. Finally, no data regarding the vaccination status of patients enrolled as well as information on variants of the SARS-CoV-2 virus were systematically provided in the investigators reviewed, limiting the possibility to conduct further sub-analyses.

## 5 Conclusions

During the first months after COVID-19 infection, recovered patients have a higher risk to develop HT compared to subjects from the general population, which increase with

aging, male sex and cancer. These findings suggest that HT represents an important post-acute COVID-19 sequelae that might benefit from an adequate primary prevention against SARS-CoV-2 infection and an appropriate follow-up in COVID-19 patients.

**Supplementary Information** The online version contains supplementary material available at <https://doi.org/10.1007/s40292-023-00574-5>.

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#### Declarations

**Conflict of Interest** None of the Authors have conflicts of interest to declare.

**Ethical Approval** This study was performed in line with the principles of the Declaration of Helsinki. The study did not require Ethical approval due to the study design.

**Consent to Participate** Not applicable.

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