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The Analgesic Effect of Transcranial Direct Current Stimulation in Fibromyalgia. A Systematic Review, Meta-Analysis, and Meta-Regression of Potential Influencers of Clinical Effect

Paulo E. P. Teixeira, PT, MSc^{a,b,c}, Kevin Pacheco-Barrios, MD^{b,c,d}, Luis Castelo Branco, MD, MPH^{b,c}, Paulo S. de Melo, MD^{b,c}, Anna Marduy, MD^{b,c}, Wolnei Caumo, MD, PhD^{b,e,f,g,h}, Stefania Papatheodorou, PhDⁱ, Julie Keysor, PT, PhD^a, Felipe Fregni, MD, PhD^{b,c,i}

^aMGH Institute of Health Professions, Boston, MA, USA

^bNeuromodulation Center and Center for Clinical Research Learning, Spaulding Rehabilitation Hospital and Massachusetts General Hospital, Charlestown, MA, USA.

^cHarvard Medical School, Boston, MA, USA.

^dUniversidad San Ignacio de Loyola, Vicerrectorado de Investigación, Unidad de Investigación para la Generación y Síntesis de Evidencias en Salud. Lima, Peru.

^ePost-Graduate Program in Medical Sciences, School of Medicine, Universidade Federal do Rio Grande do Sul (UFRGS), Brazil.

^fLaboratory of Pain and Neuromodulation at Hospital de Clínicas de Porto Alegre (HCPA), Porto Alegre, Brazil.

^gPain and Palliative Care Service at HCPA, Brazil.

^hDepartment of Surgery, School of Medicine, UFRGS, Brazil.

ⁱHarvard T. H. Chan School of Public Health, Boston, MA, USA.

Abstract

Background: There is tentative evidence to support the analgesic effect of transcranial direct current stimulation (tDCS) in fibromyalgia (FM), with large variability in the effect size encountered in different clinical trials. Understanding the source of the variability and exploring how it relates to the clinical results could characterize effective neuromodulation protocols and ultimately guide care in FM pain. The primary objective of this study was to determine the effect

CORRESPONDING AUTHOR POSTAL ADDRESS: Neuromodulation and Clinical Research Learning Center, 79/96 13th Street, 1st Floor, Suite 105, Charlestown, MA 02129, Tel (office): +1 (617) 952-6160, Mobile: +1 (617) 548-6111, pteixeira2@mg.harvard.edu.

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of tDCS in FM pain as compared to sham tDCS. The secondary objective was to explore the relationship of methodology, population, and intervention factors and the analgesic effect of tDCS in FM.

Materials and Methods: For the primary objective, a systematic review was conducted according to PRISMA guidelines. Randomized clinical trials (RCTs) investigating tDCS as an intervention for FM pain were searched at Medline, EMBASE, and the Web Of Science. Studies were excluded if they used cross-over designs, or if they did not use tDCS as an intervention for pain or did not measure clinical pain. Analysis for the main outcome was performed using a random-effects model. Risk of bias and Evidence certainty were assessed for all studies using Cochrane Risk of Bias and GRADE tools. For the secondary objective, a meta-regression was conducted to explore methodology, population, and intervention factors potentially related to the effect size.

Results: Sixteen RCTs were included. Six studies presented a high risk of bias. Significant reduction in pain scores were found for FM (SMD=1.22; 95% CI: 0.80 to 1.65, p-value<0.001). Subgroup analysis considering tDCS as a neural target revealed no differences between common neural sites. Meta-regression revealed that the duration of the tDCS protocol in weeks was the only factor associated with the effect size (ES), where protocols that lasted four weeks or longer reported larger ES than shorter protocols.

Conclusion: Results suggest an analgesic effect of tDCS for FM. tDCS protocols that last four weeks or more may be associated with larger effect sizes. Definite conclusions are inadequate given the large heterogeneity and limited quality of evidence of the included studies.

Keywords

Systematic review; Fibromyalgia; Transcranial Direct Current Stimulation; Meta-Analysis; Meta-Regression

INTRODUCTION

Five million people suffer from fibromyalgia (FM)¹ in the United States. Annual treatment costs exceeds 29 billion dollars, with a large amount spent treating secondary symptoms². Widespread chronic pain is the main symptom and chief complaint from FM while fatigue, impaired cognition, disrupted sleep and other somatic complaints are also common^{3,4}. The pathophysiology of pain in FM is complex and known to be associated with central sensitization, a hyperalgesia or excessive response to painful stimulation secondary to altered nociception at the central nervous system (CNS)⁵. In FM and other musculoskeletal conditions, chronic pain results from an imbalance between excitatory and inhibitory pain pathways. Recent evidence suggests that pain inhibitory pathways are affected in FM. For example, patients with FM demonstrate lower conditioned pain modulation (CPM), a standardized assessment that can test pain inhibitory efficacy. Indeed, the rate of patients with FM reporting pain facilitation during CPM has been found to be significantly increased compared with that of the controls (41.7% vs 21.2%)⁶. Based on the current literature, there is evidence to support that for some patients with FM, the endogenous pain regulatory system is impaired⁷⁻⁹. There is also evidence of central nervous system changes in FM that are associated with a deficit in the inhibitory control, namely abnormal cortical excitability

as expressed by decreased short intracortical inhibition and facilitation and increased resting motor threshold¹⁰ assessed by Transcranial Magnetic Stimulation. This is thought to reflect an inhibitory deficit that is associated with altered thalamic anatomy and activity¹¹, which may result in abnormal thalamocortical circuits, as evidenced by the association between central pain and thalamic dysrhythmia^{12–14}. Therefore, efforts are being undertaken to develop novel interventions that target these inhibitory deficits underlying pain maintenance mechanisms in patients with FM, as well as for other chronic pain conditions associated with a deficit in the regulatory pain system.

Targeted non-pharmacological therapies that modulate cortical excitability and modify activity of the CNS that modulate the descending pain inhibitory system are new research directions that could affect analgesia. Non-invasive brain stimulation (NIBS) techniques, such as transcranial direct current stimulation (tDCS), are safe, cost-effective, and powerful neuromodulatory tools, which have provided pain relief in several chronic pain conditions such as FM¹⁵, knee osteoarthritis^{16,17}, acute post-stroke pain¹⁸ and chronic pain after spinal cord injury¹⁹, among others, with absent or minimal side effects^{20,21}. Currently, there is tentative evidence to support the analgesic effect of tDCS as compared to sham in FM, with large variability in the effect size (ES) encountered in different clinical trials²². In fact, this variability has raised questions about the validity of tDCS as an effective treatment option for pain in FM. Understanding the source of the variability and explore how it relates to the clinical results may characterize effective neuromodulation protocols and ultimately guide care in FM pain. For example, some clinical trials have compared different protocols of tDCS in an attempt to identify the protocol that produces better responses with a smaller number of sessions, consequently, minimizing the occurrence of adverse effects and reducing the total time and cost of the therapeutic intervention²³. Other trials have investigated different tDCS montage, or electrode positioning, protocols in an attempt to determine the best short-term analgesic effect²⁴. This supports that such factors may be potential sources of variability in effect. In addition to the intervention, different factors related to the study population, such as gender, and the presence of higher levels of psychosocial comorbidities including anxiety and depression, can predict outcomes following other interventions for FM pain²⁵. Lastly, study methodology characteristics, such as inadequately powered studies, may also contribute to this observed variability in effect; and it is unknown to what degree each of these factors is contributing to it. To the best of our knowledge, no research studies, including systematic reviews, have directly explored specific factors that may be related to the analgesic effect of tDCS in FM, possibly contributing to this response variability. Identifying factors associated with tDCS clinical outcomes in FM may improve the development of current intervention protocols and help advance the design of future clinical trials for other chronic musculoskeletal pain conditions that share similar pain processing mechanisms. Although there are previous systematic reviews that have investigated the effect of tDCS on FM pain and could have explored these questions, recent clinical trials^{22,23,26} with conflicting results have not been included in these. In addition, recent reviews did not consider the ES calculated separately from different active tDCS intervention groups used in multi-arm intervention trials. Both of these aspects may relevantly affect a meta-analysis result.

To that end, the current study had the primary objective of determining the effect of tDCS in FM pain as compared to sham tDCS using a systematic review with meta-analysis. We hypothesize that tDCS will show a significant analgesic effect on FM pain as compared to sham tDCS. The secondary objective was to explore the relationship methodology, population, and intervention factors and the analgesic effect of tDCS in FM using meta-regression analysis. We hypothesize that factors such as the presence of higher levels of depression and anxiety, and the number of tDCS sessions will be significantly associated with the analgesic effect of tDCS in FM.

MATERIALS AND METHODS

To determine the analgesic effect of tDCS in FM, a systematic review and meta-analysis of randomized clinical trials (RCTs) was performed. To explore the relationship between the methodology, population, and intervention-related characteristics used in the clinical trials, and the analgesic effect of tDCS in FM, a meta-regression analysis was used to analyze potential sources of ES heterogeneity originating from these factors. This systematic review was conducted according to the preferred reporting items for systematic reviews and meta-analyses (PRISMA) guidelines²⁷.

Search Strategy

Medline, EMBASE, and the Web Of Science databases were searched for potential articles. A broad search strategy was used to guarantee the detection of all potential articles (Appendix 1). Search terms included multiple variants of the terms Transcranial direct current stimulation, pain, and FM, in addition to its Medical Subject Headings (MeSH). The final search strategy was reviewed by a librarian specialized in systematic reviews and the search was performed by one author (PT), in all databases at the same day. The search inception date to all databases was February 3rd, 2022, and was performed with no language restriction. Additionally, a snowball searching method was used on any systematic reviews or randomized trials on the topic as well as verification with senior authors who were (WC, FF) experts in the field to check if any know articles were missed.

Selection criteria

Research articles investigating the effects of tDCS on FM were included if they met the following criteria: (1) utilized tDCS alone or combined with any other therapy as a treatment tool for FM pain, including exercise, cognitive-behavioral or other neurostimulation interventions (2) measured clinical pain as a primary or secondary outcome via visual analog or numeric scale or another clinical self-reported pain tool that measured pain perception; (3) were RCTs (not including cross-over designs). Journal articles were excluded if: (1) did not use sham tDCS as a comparison group; (2) did not measure clinical pain as an outcome measure; (3) animal studies; (4) oral or poster presentation abstracts.

Article Screening

The Covidence web-based systematic review software was used to organize and screen the search results. As a first level of screening, titles and abstracts were assessed to determine article eligibility. Then, full-text articles were reviewed for the inclusion and exclusion

criteria. Two independent assessors (PT, KP), trained in clinical trial methodology and evidence synthesis, screened, and selected the final articles. Any non-unanimous criteria judgment among the two assessors would require a full-text review by each assessor to confirm their decision. If after full-text review, disagreement still existed, the full-text was reviewed by a senior assessor (FF) for an ultimate decision.

Data Extraction

Three authors performed the data extraction and entered it into a systematized spreadsheet that contained all the pre-determinant variables. Data on the factors related to the study methodology, population and intervention were extracted to create the variables used in the analysis (Table 1). If any study reported more than one pain outcome measure, any pain measurement used as a primary outcome of the study was prioritized and used in the meta-analysis. If pain was a secondary outcome of the study, then any continuous, ordinal pain outcome type, such as the visual analog scale or the numerical rating scale was considered before any categorical rating scale. This hierarchy was based on the common use of the visual and numerical rating scales by trialists, and to facilitate comparison among studies. If pain data was collected but not completely reported, the authors reached out to the first author of each study have the complete data necessary for the meta-analysis.

As part of the data extraction, mean pain scores and respective variability estimates were collected from the included studies to be used in the meta-analysis. The web-based software WebPlotDigitizer (version 4.5) was used to extract data when articles did not include specific values in tables or in text. The primary outcome used was the mean difference in pain score between baseline and the primary time-point to which the study's sample size was calculated for.

Risk of Bias Assessment

Two independent reviewers (PM, AM) trained in clinical trials methodology used the Revised Cochrane risk-of-bias tool for randomized trials (RoB 2)²⁸⁻³⁰ to assess and report the risk of bias of the included studies. The tool is organized into different domains of bias that are related to trial design, conduct and reporting. Bias due to deviations from the intended intervention was judged based on both the assignment and the adherence to intervention. Overall bias was categorized into low risk, some concerns and high risk. Assessors rated bias independently and crosschecked results with each other. Any conflicting results were discussed to reach an ultimate decision.

Evidence certainty

Evidence certainty and quality of the evidence were assessed by two independent assessors (PT, KP) using the 'Grading of Recommendations Assessment, Development, and Evaluation' (GRADE) methodology. A Summary of findings table was created.

Publication Bias

Publication bias was explored using visual analysis of the funnel plot asymmetry and tested for significance by the Egger's test for bias and the Begg's test for small study effect with continuity corrected p-value to confirm the presence publication bias. If bias was present,

adjustment was performed using the Duval and Tweedie³¹ nonparametric “trim and fill” method of accounting for publication bias in meta-analysis, to re-calculate the overall ES estimate.

Statistical Analysis

Main effect—Analysis of the analgesic effect of tDCS on pain in FM was reported as a pooled effect of all selected trials. A random-effects model was used with weights based on the DerSimonian and Laird method using standardized mean differences (SMD) as it was expected that the included studies used different pain scale measurements. Results are presented using SMD with their respective 95% confidence intervals. If a clinical trial had more than one active tDCS intervention protocol being investigated, ES was calculated for all separate intervention arms and normalized by the correlation value of 0.55 to account for the assumption of independent observations³². As the articles did not provide data on group correlations for the pain outcome at baseline, the value of 0.55 was used according to previous reference³³. A sensitivity analysis using correlation values of 0.25, 0.50, and 0.75 was also performed. Between-study statistical heterogeneity was examined using the I^2 and Cochran’s Q test. Values of I^2 greater than 50% were considered evidence for substantial heterogeneity. Lastly, descriptive statistics were used to summarize the methodological, population and intervention characteristics collected from the included studies.

Subgroup Analysis: As it was expected that studies would vary in time intervals for measuring the pain outcome score, a sensitivity analysis was planned a priori for short and moderate-term follow-up data. To that end, a short-term and moderate-term follow-up result was defined using a four-week timepoint cutoff. The 4-week cutoff was chosen based on the median of all the follow-up used in all studies. Thus, if a pain outcome score was reported less than four weeks post intervention, the ES calculated for this timepoint was used to characterize a short-term result; and scores reported after four weeks would be used as moderate term results.

Additional subgroup analysis was planned a priori considering the placement of the active electrode location for the tDCS intervention. In addition, an analysis excluding the studies which fell outside of the publication bias funnel plot was also performed, as well as an analysis considering the use of any combined intervention together with tDCS. Lastly, a subgroup analysis based on the results of risk of bias analysis comparing the classifications of “high-risk” of bias studies versus “low-risk” and “some concerns” was performed.

Meta-Regression—Data on the factors related to the study methodology, population and intervention were used to explore potential sources of ES heterogeneity and a meta-regression analysis was used for both continuous and categorical variables. To best select the variables, meta-regression models were created following a model building process intended to lower the chances of type 1 error with multiple testing. First, univariate models were used to determine the potential most significant covariates, and to eliminate non-significant variables, diminishing the number of variables for the following stages of the analysis. For that, a relaxed regression coefficient significance p-value of $p < 0.2$ was used. The chances of type 1 error with multiple independent testing was lowered by having less variables

to build the final models³⁴. Secondly, the significant variables were included in an initial multivariate model. Then, using a backward elimination-based approach, the variable with the largest non-significant p-value was excluded from the model until a final model with only significant coefficients ($p < 0.05$) variables was achieved.

Observations with values greater than three standard deviations above or below the mean scores were considered outliers and excluded as per the Empirical Rule³⁴. For all meta-regression models, the assumption of linearity was assessed by visual scatterplot analysis with a superimposed regression line. Homoscedasticity was checked by comparison of scatterplots of the standardized predicted values and standardized residuals³⁶. Normality of the residuals was checked using the Shapiro-Wilk normality test³⁷. Durbin Watson estimate, tolerance values, Cook distance, variance inflation factor and Standardized DFBeta values were analyzed for regression diagnostics such as multicollinearity and influential cases. Due to the exploratory nature of the meta-regression analysis, the conservative approach to correct the significance level for multiple testing was not performed. All tests were two-tailed, using a $p < 0.05$ threshold for significance, and were performed using Stata V17 BE Statistical Software (StataCorp. 2021. College Station, TX: StataCorp LLC.)

RESULTS

The result of our search strategy revealed 418 potential articles, while one study was identified using the snowball strategy. After duplicates were removed, 314 articles were considered for the first level of screening. After all levels of screening and selection, 16 articles were selected. Figure 1 shows the CONSORT diagram for the selection of the final articles.

Studies were performed between 2006 and 2022 in seven different countries, 10 in Brazil^{23,24,26,38–44}, two in Belgium^{45,46}, and individual studies in Egypt⁴⁷, Norway⁴⁸, United States¹⁵, and Spain⁴⁹. A total of 813 subjects were included in the analysis. Three studies^{45,46,48} used the pain numerical rating scale, two studies^{24,39} used the visual numerical scale, one study⁴⁰ used the Short-Form 36 Health Questionnaire (SF36) pain subscale, while all other studies used the pain visual analog scale (10 studies). Only one multi-arm intervention study used the cathodal as the active polarity²⁴. The two milliamperes (mA) current intensity was the most used. Other current intensities included 1.5mA^{45,46} and 1mA⁴³. Three studies applied tDCS in combination with other intervention^{39,41,44}. Active electrode location varied mostly between the primary motor cortex region (M1) (11 studies)^{15,23,24,39,40,42–44,47–49} or the dorsolateral prefrontal cortex (DLPFC) (eight studies^{15,26,38,41,42,45,46,49}). Some studies used both M1 and DLPFC as different active tDCS groups in the same study. The reference electrode locations varied and are listed on Table 2. Table 2 summarizes the characteristics of the selected studies used in the analysis.

Risk of Bias Assessment

Our analysis detected varying levels of bias among the selected studies. The percentage agreement between the assessors were 95%. The percentage of studies in the classifications of high-risk, some concerns and low risk were respectively 37.5%, 37.5%, and 25%. The

domains more frequently characterized as “some concerns” and “high risk” of bias were the ones evaluating the randomization (D1), deviations from intended interventions (D2), and missing outcome data (D3). Many of the studies did not properly report the methodological domains of the risk of bias tool, being characterized as “no informed”. We provide the risk of bias assessment of the included RCTs in Appendix 2.

Evidence Certainty

The GRADEpro GDT software⁵⁰ was used to rate the quality of the evidence according to the GRADE criteria for the pain intensity outcome at short and moderate term follow-ups. Single randomized studies (with under 300 participants) had quality downgraded twice as it was considered inconsistent and potentially imprecise (i.e., if wide 95% CI). Presence of unclear or high risk of bias was also downgraded once according to the criteria. A “Summary of findings” table (Table 3) was created to present the result and the criteria explanation is shown in the table.

Main Effect Analysis

The meta-analysis for the main effect of active tDCS versus sham tDCS for pain in FM showed a positive result favoring the active tDCS intervention with a pooled ES of $d=1.22$ (95% CI: 0.80, 1.65, p -value<0.001, p -value < 0.001), which is considered to be large based on suggested standards⁵¹. This ES corresponds to a 41% change, which is above the threshold for the minimal clinically important difference for pain in FM (34.2%)⁵². Large heterogeneity was found among studies ($I^2=83.6\%$). Figure 2 shows the forest plot for the included studies.

The publication bias analysis revealed an asymmetric distribution on the funnel plot suggesting a publication bias in the analysis (Appendix 3). A significant Egger’s test for bias was found $t = 2.70$, p -value = 0.012. Begg’s test for small study effect was statistically significant (continuity corrected p -value = 0.022) which provides evidence of publication bias. After adjustment for publication bias, the overall ES estimate was 1.161 (95% C.I. 0.51, 1.81) and still significant.

Subgroup Analysis

Short-Term versus Moderate Term Effects—As defined a priori, a sensitivity analysis according to short-term versus moderate term effects was performed. For the short-term effects of tDCS on pain, two studies^{40,43} did not collect pain outcomes before four weeks, and seven studies did not have moderate term data on pain^{15,23,24,38,41,45,46}. The median timepoint for the short-term outcome results was three weeks, while for moderate term was 10 weeks. The pooled ES estimates were respectively 0.81 [95% CI: 0.43, 1.19, p -value = 0.000; $I^2=78.5\%$; $\text{Chi}^2(23) = 107.21$, $p = 0.000$; $\tau^2 = 0.66$], and 1.69 [95% CI: 0.84, 2.54, p -value = 0.000; $I^2=91.0\%$; $\text{Chi}^2(9) = 100.3$, $p = 0.000$; $\tau^2 = 1.59$]. The results suggest that moderate term effects of tDCS on pain seem to be larger than short-term effects. The forest plots are offered as a supplemental material.

Active Electrode Location—The sensitivity analysis for the main effect of tDCS versus sham tDCS for pain in FM considering the placement of the active electrode location

showed no difference when M1 location was compared directly to DLPFC location. The pooled ES estimates for studies which used M1 and DLPFC were respectively 1.25 [95% CI: 0.52, 1.98, p-value = 0.000; $I^2=88\%$; $\text{Chi}^2(12) = 99.71$, $p = 0.000$; $\tau^2 = 1.51$], and 1.43 [95% CI: 0.73, 2.14, p-value = 0.000; $I^2=81.7\%$; $\text{Chi}^2(7) = 38.17$, $p = 0.000$; $\tau^2 = 0.78$] (Supplemental Material). The overlapping of the 95%CI shows that there is no advantage of using M1 or DLPFC as compared with each other.

Risk of Bias Subgroups—A total of 6 studies were classified with “high-risk” of bias (Appendix 2), while 10 studies were classified as others (“low risk” or “some concerns”). The effect sizes were respectively 1.53 [95% CI: 0.47, 2.59) and 1.07 [95% CI: 0.65, 1.49). The forest plot of this comparison analysis is provided as a supplemental material.

Presence of Combined Intervention—Three studies used a combined intervention with tDCS^{39,41,44} while the other 13 studies did not. The effect sizes were respectively 1.70 [95% CI: 1.24, 2.16) and 1.16 [95% CI: 0.69, 1.63). Two studies^{39,44} used an exercise-based type of intervention, while one study⁴¹ used a cognitive-behavioral based therapy. As for the timing of the combination, two studies^{41,44} combined concurrently with the tDCS stimulation, while one study³⁹ used immediately after tDCS stimulation. The forest plot of this comparison analysis is provided as a supplemental material.

Meta-Regression—No multivariate models were significant. As FM is a disease which occurs predominantly in women gender variability not sufficient to be properly investigated as an independent variable in the meta-regression. Only one meta-regression model showed significance in univariate analysis for the variable of the duration of the tDCS intervention protocol in weeks. The direction of the regression coefficient indicated that the longer the duration of the tDCS intervention protocol in weeks, the larger the effect size ($R^2 = 31.3\%$). Figure 3 shows model information and regression plot.

After visual analysis of the regression plot, we noticed that studies with protocols lasting approximately 4–5 weeks presented higher effect sizes as compared to less than 4–5 weeks. In attempt to facilitate the clinical application of this information we have explored this data by dichotomizing the duration of the protocol variables using different cut-offs (1, 2, 3, 4 and 5 weeks) and tested for significance using meta-regression. The result was significant only for the 4-week cut-off with a regression coefficient $\beta = 2.35$ (standard error: 0.95, $p=0.021$, 95% C.I: 0.39, 4.30) and adjusted $R^2 = 22.1\%$. All other baseline variables were not significant, and results are shown on Appendix 4. Although the variables that represented the total number of tDCS sessions, and the tDCS session duration in minutes are shown to be significant, the p-values were above 0.05 after influential point adjustment.

DISCUSSION

Our meta-analysis included 16 RCTs with 26 different active tDCS protocol groups that were compared to sham tDCS. The field required a meta-analysis update as trials which have been published in the last five years, with varied ES, were not included in past reviews^{22,23,26}. In addition, to the best of our knowledge, this is the first meta-analysis in the field that used the ES calculated separately for all active tDCS intervention groups used

in multi-arm intervention trials. It is imperative to consider all variations of tDCS protocols as the variability in ES may be explained by differences in protocols. This updated review also included the exploration of factors related to ES variability using meta-regression analysis and revealing that the duration of the intervention protocol in weeks seemed to be a significant factor to explain the variability in the effect of tDCS on pain for FM.

Our results for the main effect of tDCS on FM pain showed a significant analgesic effect (pooled ES of $d=1.22$ (95% CI: 0.80, 1.65). Although this ES is larger than previously reported, our finding agree with a previous systematic review on the topic²² that reported medium to large ES (SMD= -0.66), thus supporting the analgesic effect. Chaturvedi et. al. performed a systematic review also in the same topic, however no meta-analysis was made, or ES was reported. A potential reason for the ES differences of our results as compared to previous reviews may be the fact that our review included recent clinical trials with conflicting results that were not included in preceding reviews and meta-analyses. In addition, our results were calculated considering the ES calculated separately from different active tDCS intervention groups used in multi-arm intervention trials, a methodology detail that has also not been used in previous reviews. Such factors may relevantly affect a meta-analysis result. In addition, we have explored specific factors that may be related to the analgesic effect of tDCS in FM using rigorous methodology.

To our knowledge, only one previous study⁵³ attempted to explore intervention related factors such as number of sessions, weekly frequency, and protocol duration, or different tDCS montages (M1 versus DLPFC). Hou⁵³ and colleagues investigated the effects of different non-invasive brain stimulation treatments, including tDCS and repetitive transcranial magnetic stimulation (rTMS) as add-on treatment approaches to FM. Although the study included only five tDCS trials, the results suggested that the M1 location produced larger effect sizes as compared to DLPFC ($p < 0.0001$). Our subgroup analysis comparing studies with M1 and DLPFC stimulation sites did not show differences in the effect of tDCS on pain. We believe the multidimensional nature of pain in FM could be related to this finding and although the mechanisms would be different, both strategies would affect central pain processing centers. In fact, targeting M1 would lead to motor cortex excitability modulation, subsequently affecting components of sensory pain processing, ultimately facilitating the descending pain inhibitory system⁵⁴. Meanwhile, targeting the DLPFC would lead to modulation of the cognitive-affective components of pain^{55,56} due to its connections with the limbic system structures. Upcoming mechanistic studies evaluating outcomes related to pain processing, such as quantitative sensory testing, could help clarify the effects of tDCS on different stimulation sites⁵⁷. Conditioned pain modulation, for example, is believed to be a marker of the descending pain inhibitory system⁵⁸ that is potentially more influenced by M1 than DLPFC stimulation⁵⁹. Comparably, the effects of DLPFC stimulation on mood, fatigue, and attention, commonly impaired in patients with FM⁶⁰, could also be assessed and later correlated to its effects on pain. One important consideration is that pain measurement as indexed by VAS is an unspecific measurement that is poorly correlated with mechanistic pain measurements^{8,61}.

It is possible that the lack of differences in effect between M1 and DLPFC stimulation sites could be simply due to the lack of focality of the tDCS stimulation. Therefore, an

aspect that is sometimes viewed as a disadvantage of the technique^{62,63} could in fact lead to positive results in this case. Targeting M1 or DLPFC would lead, therefore, to simultaneous modulating of several pain processing and affective/cognitive pathways. Here, again, future study designs comparing the two stimulation sites and their effects on different dimensions of pain would help address this potential explanation. Hou⁵³ et. al. also explored factors related to treatment regimen such as total number of sessions, frequency, and protocol duration, however, no significant linear relationship was found. Despite the fact that our initial analysis showed that the total number of sessions and the duration of the tDCS session in minutes were significantly influencing the effect size, such variables were not relevant after controlling for influential observations.

We found that tDCS protocols that lasted four weeks or more, reported larger effect sizes than shorter protocols. These results persisted after adjusting for the number of sessions in the meta-regression model, suggesting an essential role of the session distribution across weeks more so than only the total number of sessions. Our finding is aligned with a previous mechanistic study⁶⁴ that showed that the broader repetition interval of M1 tDCS (more than 24 hours) produces the strongest cortical excitability changes. Similarly, a previous tDCS trial on patients with FM reported during a longitudinal analysis of analgesic effects, that the highest clinically significant differences compared to sham tDCS, were achieved after the fourth week of stimulation (approximately after 15 sessions)⁶⁵. These time-dependent phenomena are not exclusive to tDCS studies on pain populations; in fact, recent evidence-based guidelines on tDCS for neuropsychiatric conditions⁶⁶ reported similar patterns for stroke recovery and depression management.

The potential underlying physiological explanation is the time-dependency of neuroplasticity. Previous studies^{67,68} have suggested that tDCS effects depend on the engagement of NMDA-related networks within the brain target; thus, the magnitude and duration of tDCS-induced plasticity must be related to stimulation duration and timing⁶⁹. However, as it has been reported in several animal models, NMDA-related long-term potentiation and depression have a saturation effect (homeostatic plasticity) that prevents destabilization of synaptic function; this process depends on protein synthesis regulation, which requires a time window to be executed⁷⁰⁻⁷³. Therefore, it can be argued that some stabilization of plasticity takes place during that prolonged period between stimulation sessions, which could facilitate the cumulative gains of plasticity.

Nevertheless, we need to consider that only some studies performed protocols with a duration longer than four weeks; hence, our findings could be underpowered or driven by few influential studies and require further exploration. But it is clear that session distribution and specific timing of repetition intervals are important for optimizing the cumulative effects of tDCS protocols. As suggested by Pacheco-Barríos et al.⁷⁴, there is a need for more extended protocols with an optimized architecture based on neuroplasticity physiology. For instance, including an induction phase with daily sessions (which produces an optimal target engagement^{64,75}) for the first one or two weeks, followed by booster sessions across time^{76,77} to maximize cumulative plasticity and synaptic stabilization. Further studies must test methodological alternatives to increase the feasibility of these approaches, such as the inclusion of digital health methods and home-based stimulation protocols^{78,79}.

Despite supporting evidence on the emotional and affective impact of FM and its effect on pain perception^{80,81}, our hypothesis that the presence of higher levels of depression and anxiety, or other psychosocial related factors would be significantly associated with the analgesic effect of tDCS in FM was not confirmed as it could not be adequately tested with our data. To explore this, we extracted data on baseline levels of depression, anxiety, and pain catastrophizing. Although depression scores were reported in 14 of the 16 studies, there was small variability on levels of depression as most studies excluded subjects with severe depression. Only few studies have provided information on anxiety and pain catastrophizing, as well as other potential key factors such as the use of medications and sleep quality scores. Therefore, the available data was insufficient to construct stable meta-regression models for many of the population related variables, and model significance was likely influenced by the low number of observations. Consequently, it is not possible, or adequate, to make conclusions on the relevance of these population related factors on their influence on the analgesic effect of tDCS on pain. This finding points out the lack of consistency and structure in reporting population related factors. Future study designs are encouraged to include information and variability of such factors so its relation to tDCS clinical effect can be appropriately explored and used to diminish response variability in the field of neuromodulation.

One potentially relevant finding is related to the subgroup analysis based on the presence of a combined intervention used with tDCS. Although this comparison was not statistically significant, the three studies which used a combined intervention had a pooled ES of 1.70 [95% CI: 1.24, 2.16) and two out of the three studies used an exercise-based intervention. The studies that did not use other combined interventions had an ES of 1.16 [95% CI: 0.69, 1.63). This finding supports the use of tDCS in combination with exercises for FM pain. Research has showed that the combination of tDCS and other therapies, such as exercise, are believed to increase clinical effects on pain, as compared to tDCS alone. The evidence supports the mechanism of increasing the brain responsiveness to the corticomotor benefits of exercise or other interventions as for example increased cortical excitability, improved motor control, enhanced muscle activation⁸², and by adding effects on the pain system function.

Limitations

Our review and meta-regression exploration were performed using 16 RCTs with a total of 813 subjects included in the analysis. Although our analysis suggested that our review is based on low to very low evidence quality, included trials were mainly penalized for imprecision due to small sample size (under 300 participants). We agree that small and low quality RCTs may indeed present compromised results which might not represent the true effect of the intervention, or the true population characteristics. For example, small trials may have the randomization assignment compromised and suffer from potential confounding. We also acknowledge that this limitation is understandable given the challenges with recruitment and adherence in RCTs assessing NIBS in this population⁸³. Our results did not find any significant multivariate meta regression models. This finding should be cautiously interpreted as it is likely that potentially existing multivariate models may have been underpowered. In addition, our meta-regression analysis

had an exploratory nature, and model building process was performed accordingly. Also, there is a possible influence of the risk of bias on the pooled results as the individual ES may not represent the true effect due to the study biases. Among the limitation, the comparison of different measures to assess clinical pain; the male gender not properly represented making the results generalizable only to the female population, the lack of studies using alternative cortical targets to assess the effect of tDCS, and the impossibility to assess the effect of medication should be noted and considered before any conclusion is attempted. Too, the studies selected did not included more recent approaches of transcranial electrical stimulation, such as focal-tDCS or transcranial random noise stimulation (trNS), which perhaps may be able to exert stronger stimulation effects over the same cortical targets with reported therapeutical effects on pain⁸⁴. Lastly, the sham tDCS montage could have been a limiting factor for comparison between active and sham groups for the multi-arm intervention trials. The authors caution that all inferences about any results should consider these limitations and be interpreted accordingly.

CONCLUSION

Our study provides an updated systematic review and meta-analysis with meta-regression methodology on the clinical analgesic effect of tDCS in FM. The main effect of active tDCS versus sham tDCS for clinical pain in FM was confirmed to have a positive result favoring the active tDCS, with a large ES. Our findings also suggest that tDCS protocols that last four weeks or more may be associated with larger effect sizes than shorter protocols, and time-dependency neuroplasticity is discussed as a potential underlying physiological cause. The offered conclusions should be interpreted with caution due to the risk of bias and low levels of certainty found in our analysis and considering the lack of robust studies. Lastly, future trials must be thorough to methodological details such as randomization process, deviations from intended interventions, missing data, and reporting results to minimize bias.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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Appendix

Appendix

Appendix 1.

Search strategy.

Database	Search Terms	Results
MEDLINE PubMed	TDCS "noninvasive brain stimulation"[Title/Abstract] OR "non-invasive brain	18,246

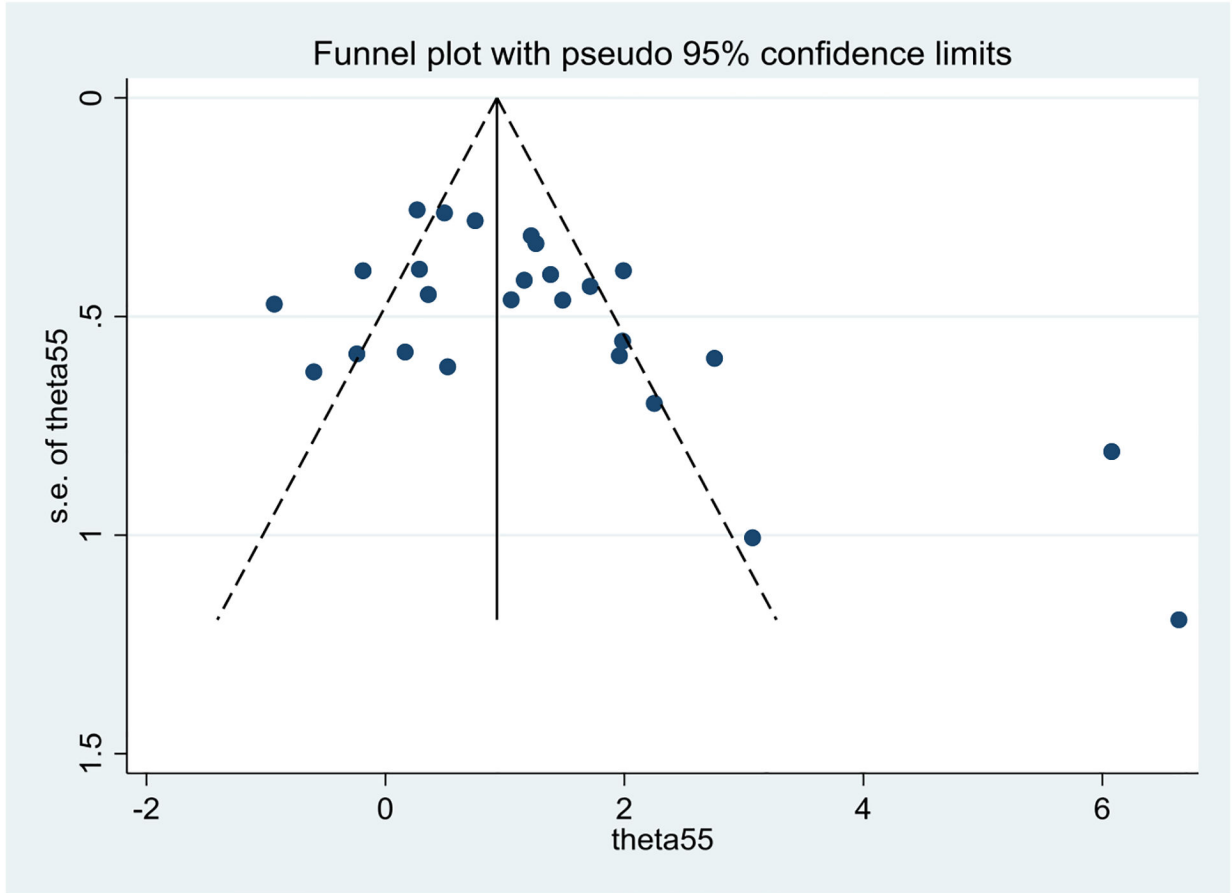
Database	Search Terms	Results
	stimulation"[Title/Abstract] OR "neuromodulation"[Title/Abstract] OR "NIBS"[Title/Abstract] OR "motor cortex stimulation"[Title/Abstract] OR "Transcranial Electric Stimulation"[Title/Abstract] OR "transcranial direct current stimulation"[MeSH Terms] OR "transcranial direct current stimulation"[Title/Abstract] OR "tdcs"[All Fields] OR "transcranial direct current stimulation"[All Fields]	
	Fibromyalgia: "Fibromyalgia"[MeSH Terms] OR "Fibromyalgia"[All Fields] OR "fatigue syndrome, chronic"[MeSH Terms] OR "chronic fatigue syndrome"[All Fields] OR "Myofascial Pain Syndromes"[MeSH Terms] OR "Myofascial Pain Syndrome"[All Fields] OR "Trigger Points"[MeSH Terms] OR "Trigger Points"[All Fields] OR "Trigger Point Pain"[All Fields] OR "Fibromyalgia"[All Fields] OR "Widespread chronic pain"[All Fields] OR "widespread pain"[All Fields]	27,761
	TDCS and FIBROMYALGIA	109
EMBASE	TDCS 'transcranial direct current stimulation'/exp OR 'noninvasive brain stimulation':ab,ti OR 'noninvasive brain stimulation'/exp OR 'non-invasive brain stimulation':ab,ti OR 'transcranial direct current stimulation':ab,ti OR 'neuromodulation':ab,ti OR 'neuromodulation'/exp OR 'nibs':ab,ti OR 'tdcs':ab,ti OR 'transcranial direct current stimulator'	59,921
	FIBROMYALGIA 'fibromyalgia'/exp OR 'fibromyalgia':ab,ti OR 'myofascial pain'/exp OR 'myofascial pain syndrome'/exp OR 'myofascial pain':ab,ti OR 'trigger point pain' OR 'trigger point':ab,ti OR 'chronic fatigue syndrome'/exp OR 'chronic fatigue syndrome':ab,ti OR 'chronic fatigue':ab,ti OR 'widespread chronic pain' OR 'widespread chronic pain':ab,ti	45,271
	TDCS and FIBROMYALGIA	490
	EMBASE FILTER by Study Type: clinical trial + randomized controlled trial + randomized controlled trial topic	130
WEB OF SCIENCE	TDCS transcranial direct current stimulation OR noninvasive brain stimulation OR noninvasive brain stimulation OR non-invasive brain stimulation OR transcranial direct current stimulation OR neuromodulation OR neuromodulation OR nibs OR tdcs	32,311
	FIBROMYALGIA Fibromyalgia OR chronic fatigue syndrome OR Myofascial Pain Syndrome OR Myofascial Pain OR Trigger Points OR trigger point pain OR Fibromyalgia OR chronic fatigue syndrome OR Widespread chronic pain OR Widespread pain	65,205
	TDCS and FIBROMYALGIA	267
	WeB of Science Filter - Document Type Articles	179

Appendix

Study ID		D1	D2	D3	D4	D5	Overall	
Bin Yoo	2018	!	!	+	+	+	!	+
Brietzke	2019	+	+	+	+	+	+	!
Caumo	2022	+	+	+	+	+	+	-
De Melo	2020	+	-	-	+	+	-	
Fagerlund	2015	+	!	+	+	+	!	D1 Randomization Process
Fregni	2006	!	+	+	+	+	!	D2 Deviations from intended interventions
Jales Junior	2015	!	+	+	+	+	!	D3 Missing outcome data
Khedr	2017	+	!	-	+	+	-	D4 Measurement of the outcome
Matias	2022	+	!	-	+	+	-	D5 Selection of the reported result
Mendonça	2011	!	+	+	+	+	!	
Mendonça	2016	+	+	+	+	+	+	
Riberto	2011	-	+	+	+	+	-	
Samartin-Veiga	2021	+	+	+	+	+	+	
Santos	2018	+	!	+	!	+	!	
To	2017	!	!	-	!	+	-	
Valle	2009	!	!	+	+	-	-	

Appendix 2.
Risk of bias assessment of the included studies.

Appendix



Appendix 3.
Funnel plot for publication bias analysis.

Appendix

Appendix 4.

Univariate models for all explored variables.

Study Methodology Related Factors

Variables	n	z2	Adj. R2	β	SE	t	p	95%CI	
								LB	UP
Use of intention-to-treat analysis	22	2.53	8.14	0.30	0.91	0.33	0.75	-1.60	2.19
Size of active intervention group	26	2.05	6.87	0.00	0.04	-0.09	0.93	-0.08	0.07
Size of sham intervention group	26	2.05	6.72	0.00	0.05	0.09	0.93	-0.09	0.10
# of Dropouts in active intervention group	22	2.40	1.88	0.19	0.22	-0.86	0.40	-0.64	0.27
# of Dropouts in sham intervention group	22	2.41	2.48	-0.18	0.21	-0.82	0.42	-0.62	0.27
Dropout % in active intervention group	24	2.04	1.79	-0.05	0.05	-1.02	0.32	-0.15	0.05
Dropout % in sham intervention group	24	2.11	1.54	-0.04	0.05	-0.85	0.41	-0.14	0.06
Blinding of participants	26	2.05	6.89	0.13	1.64	0.08	0.94	-3.24	3.52

Study Methodology Related Factors

Variables	n	z2	Adj. R2	β	SE	t	p	95%CI	
								LB	UP
Blinding of intervention staff	26	1.93	0.47	1.58	1.66	0.95	0.35	-1.85	5.02

Population Related Factors

Variables	Active Group									Sham Group								
	95%CI									95%CI								
	n	z2	Adj. R2	β	SE	t	p	LB	UP	n	z2	Adj. R2	β	SE	t	p	LB	UP
Average age	23	1.83	1.15	-0.06	0.06	-1.04	0.31	-0.19	0.06	23	1.76	4.52	-0.09	0.08	-1.20	0.24	-0.2	
Duration of symptoms (months)	8	2.99	0.64	-0.01	0.01	-1.06	0.33	-0.05	0.02	8	2.95	0.73	-0.01	0.01	-1.10	0.31	-0.0	
Percentage of women	26	1.91	0.36	-0.02	0.02	-0.85	0.40	-0.05	0.02	26	1.80	6.38	-0.02	0.02	-1.25	0.23	-0.0	
Body mass index	7	2.64	17.21	-0.73	0.51	-1.43	0.21	-2.04	0.58	7	3.91	22.42	0.37	0.60	0.61	0.57	-1.1	
Number of subjects with another chronic illness	6	6.42	21.10	0.10	0.19	0.51	0.63	-0.43	0.63	6	6.70	26.46	0.07	0.19	0.35	0.74	-0.4	
Baseline depression level	14	2.48	2.18	-0.81	0.75	-1.09	0.30	-2.44	0.81	14	2.48	2.18	-0.08	0.75	-1.09	0.30	-2.4	
Baseline anxiety level	7	3.33	6.21	-0.76	0.84	-0.90	0.41	-2.93	1.41	7	3.31	5.36	-0.73	0.80	-0.92	0.40	-2.7	
Baseline pain catastrophizing score	5	3.70	33.52	0.17	0.23	0.75	0.51	-0.55	0.89	5	3.86	39.35	-0.19	0.33	-0.58	0.61	-1.2	
Percentage of subjects using opioids [†]																		
Disability score at baseline	14	0.62	7.92	-0.02	0.25	-0.78	0.45	-0.07	0.04	14	0.65	4.07	-0.02	0.23	-0.97	0.35	-0.0	
Sleep quality score at baseline	7	2.90	8.46	0.01	0.08	1.14	0.31	-0.12	0.31	7	3.10	15.95	0.09	0.09	0.99	0.37	-0.1	

[†] Calculation not possible due to low number of observations

Intervention Related Factors

Variables	n	z2	Adj. R2	β	SE	t	p	95%CI	
								LB	UP
Presence of tDCS with a combined intervention [‡]									
Total number of intervention sessions	26	1.45	24.40	0.09	0.03	2.97	0.007 [‡]	0.03	0.15
# of Intervention sessions above/below 8 (median)	26	2.01	4.60	0.56	0.64	0.87	0.39	-0.77	1.89

Intervention Related Factors

Variables	n	z2	Adj. R2	β	SE	t	p	95%CI	
								LB	UP
Frequency of intervention sessions per week	26	2.04	6.30	0.09	0.17	0.52	0.61	-0.27	0.45
Duration of tDCS intervention (in weeks)	26	1.32	31.30	0.28	0.09	2.96	0.007*	0.08	0.47
Patient setting (inpatient/ outpatient/ home)	26	1.69	12.09	-1.99	0.96	-2.06	0.05	-3.98	0.00
Active electrode location	26	2.03	5.91	0.08	0.34	0.23	0.82	-0.63	0.79
Use of a combined intervention with tDCS	26	2.00	4.19	0.46	0.98	0.47	0.64	-1.57	2.49
tDCS duration of session (in minutes)	26	1.34	30.11	0.55	0.18	3.08	0.005 [†]	0.18	0.92

[†]Not significant after influential point adjustment;

* Statistically significant at 0.01 level;

[‡]Calculation not possible due to low number of observations

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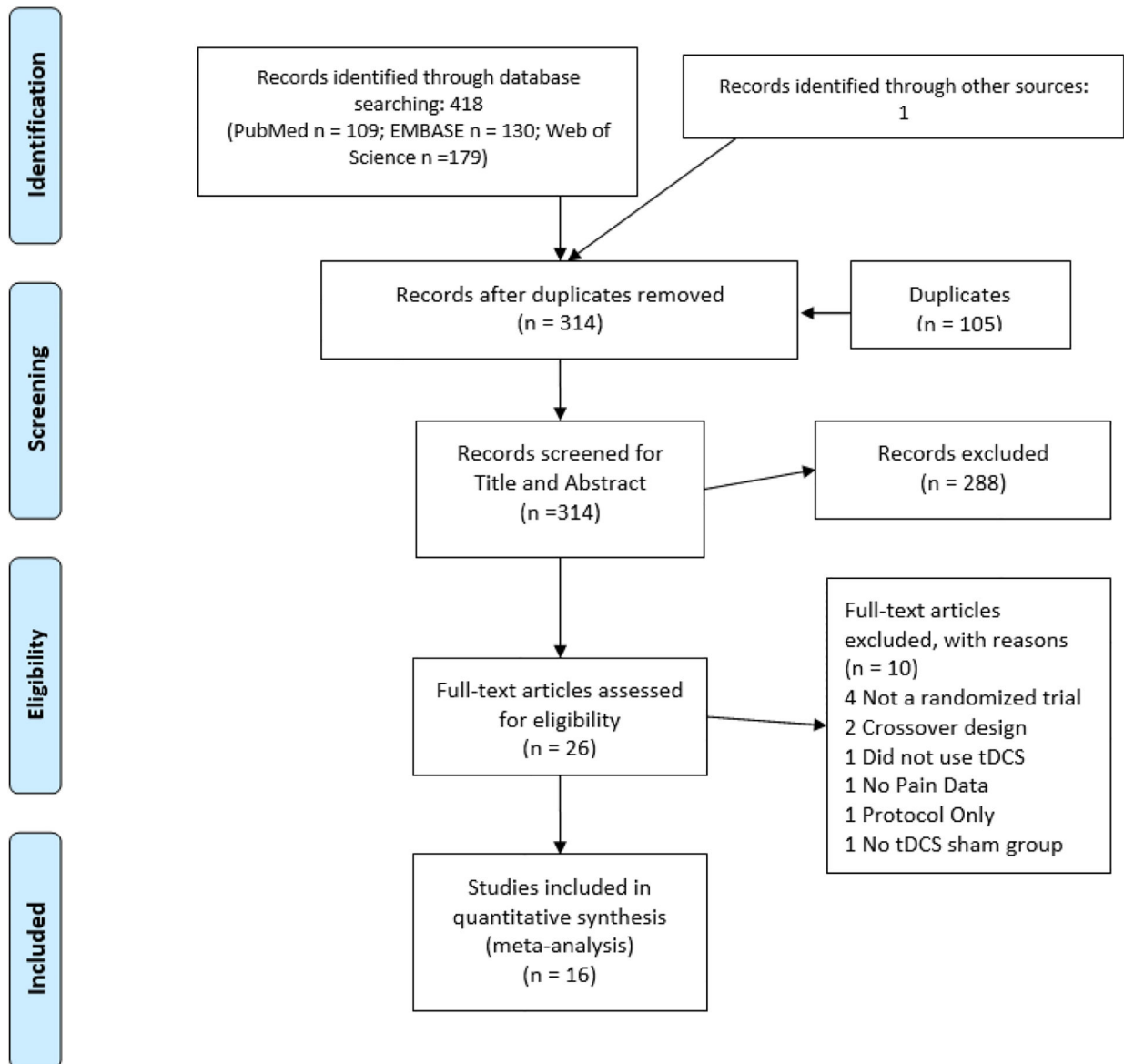


Figure 1.
Flow diagram of study selection.

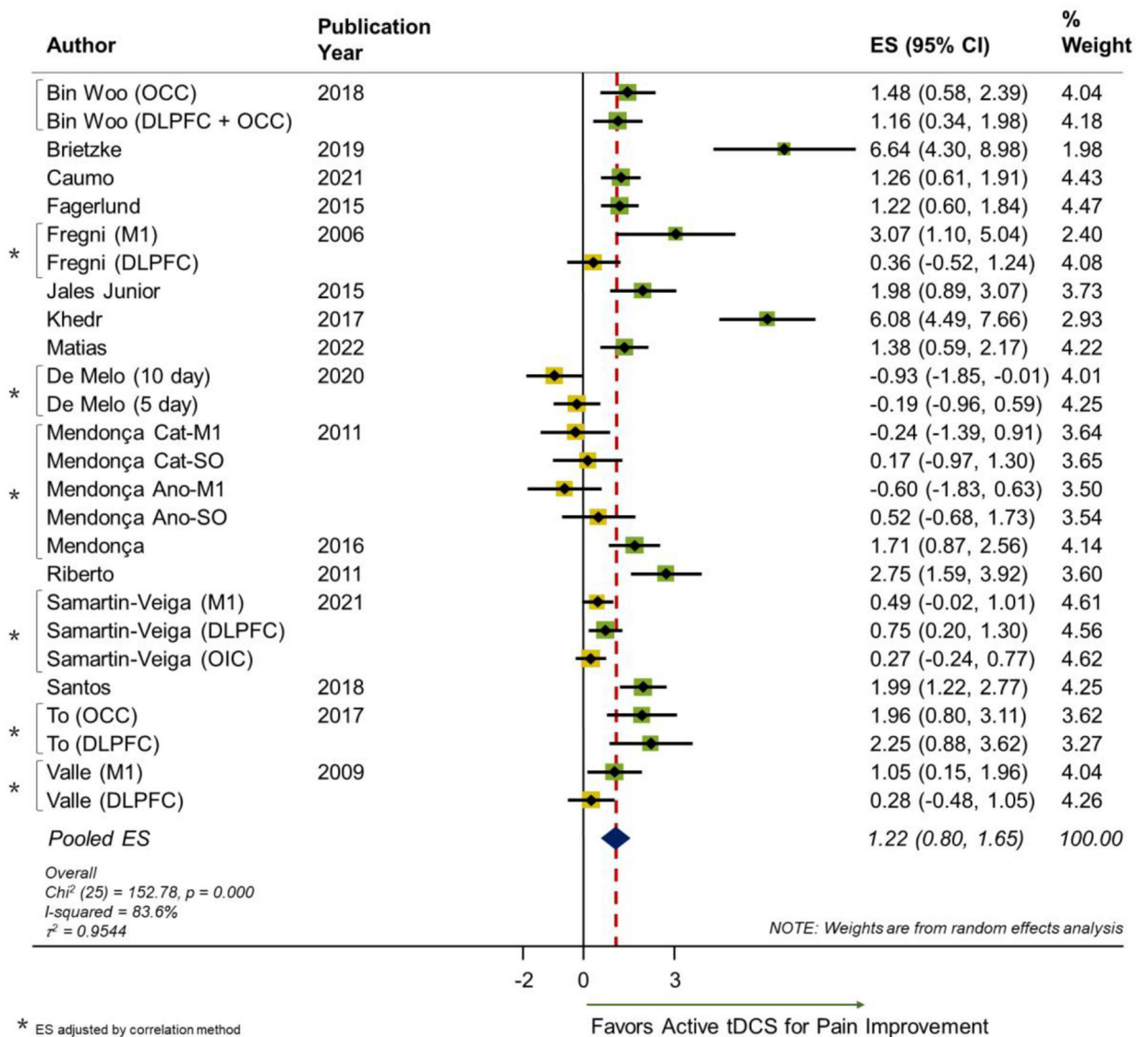


Figure 2.

Forest plot of the main effect of tDCS on pain in FM.

OCC: occipital region; DLPFC: dorsolateral prefrontal cortex; M1: primary motor cortex; OIC: operculo-insular cortex; NRS: numeric rating scale; Cat: Cathodal; Ano: Anodal; SO: Supraorbital; ES: Effect Size; CI: Confidence Interval.

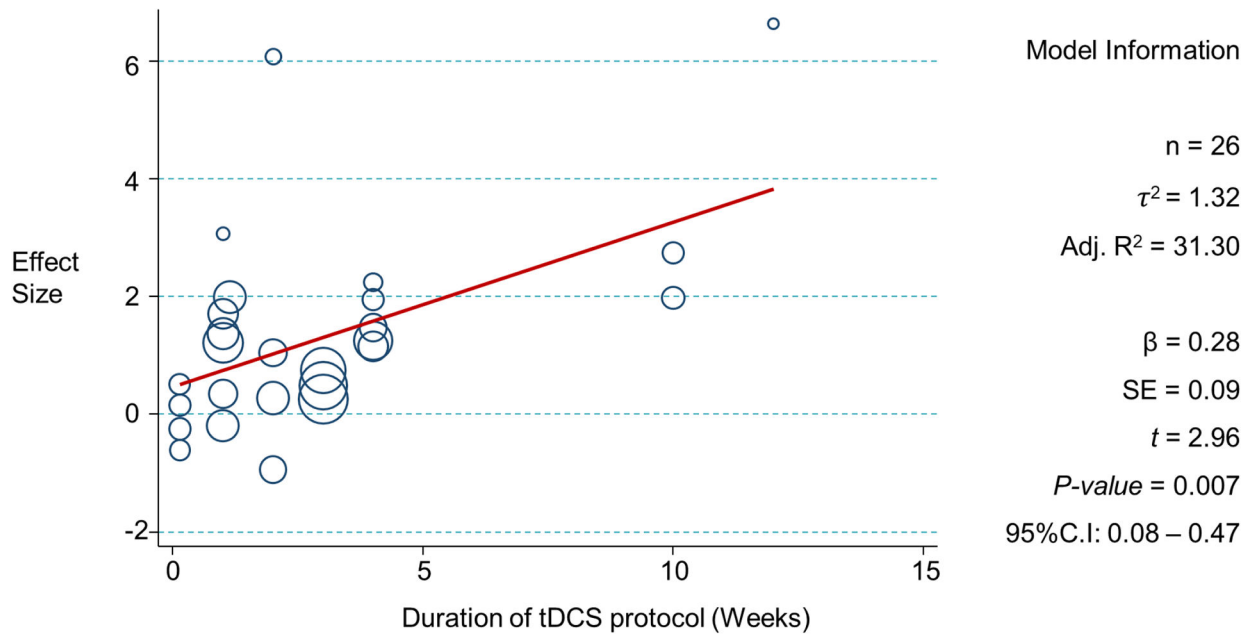


Figure 3. Significant univariate meta-regression model – Duration of tDCS intervention measured in weeks.

n: Number of observations; τ^2 : tau-squared; Adj. R²: adjusted r-squared; β : beta coefficient; SE: standard error; t: t statistic; 95% C.I.: 95% confidence interval

Table 1.

Considered factors related to the study methodology, population and intervention used to create the variables used in the analysis.

Study Methodology	Population	Intervention
- Use of intention-to-treat analysis	- Age	- Use of tDCS with a combined intervention
- Size of intervention groups	- Duration of symptoms	- Type of combined intervention
- Dropouts in intervention groups	- Gender	- Timing of combined intervention
- Disease diagnostic definition criteria	- Race	- Total number of intervention sessions
- Blinding of study subject	- Body mass index	- Frequency of intervention sessions per week
- Blinding of intervention staff	- Number of subjects with another chronic illness	- Duration of tDCS intervention in weeks
	- Baseline depression level	- Patient setting
	- Baseline anxiety level	- Type of tDCS montage
	- Baseline pain catastrophizing score	- Number of stimulation channels
	- Percentage of subjects using opioids	- Stimulation polarity
	- Percentage of subjects using gabapentinoids	- Active electrode location
	- Percentage of subjects using antidepressants	- tDCS current intensity parameter (mA)
	- Percentage of subjects using anticonvulsants	- tDCS session duration
	- Percentage of subjects using analgesics	- tDCS electrode surface area
	- Percentage of subjects using myorelaxants	
	- Disability score at baseline	
	- Sleep quality score at baseline	
	- Baseline pain Level (Active and Sham groups)	

Table 2.

Summary of the included RCTs that investigated the analgesic effect of tDCS on FM.

Paper Year	tDCS characteristics										Disability score at baseline (SD)					
	Country	Active Groups (n)	Sham Group (n)	Active electrode location	Intensity (mA)	Polarity (Active)	Reference Electrode Location *	Session duration (min)	Number of sessions	Sessions / week		Duration of tDCS intervention (weeks)	Type of Combined intervention	Pain outcome used	Use of Mechanistic outcomes	Disability tool
Bin Yoo 2018	Belgium	21	16	OCC DLPFC +OCC	1,5 2	Anode	O2 F3	20	8	2	4	-	NRS	No	FIQ	Active OCC: 55,80 (14,45) Active OCC + DLPFC: 50,38 (14,85) Sham: 56,58 (13,86)
Briefcase 2019	Brazil	10	10	DLPFC	2	Anode	F4	30	60	5	12	-	VAS	Yes	FIQ	Active DLPFC: 71,7 (12) Sham: 63 9 (21,5)
Caumo 2021	Brazil	32	16	DLPFC	2	Anode	F4	20	20	5	4	-	VAS	Yes	FIQ	Active DLPFC: 70,2 (15,2) Sham: 65,88 (15,53)
Fagerlund 2014	Norway	24	24	M1	2	Anode	Fp2	20	5	5	1	-	NRS	No	FIQ	Active M1: 57 (36,54) Sham: 48,41 (37,44)
Fregni 2006	United States	11 11	10	M1 DLPFC	2	Anode	Fp2	20	5	5	1	-	VAS	No	SF36-PF	Active M1: 32,1 (6,30) Active DLPFC: 30,4 (33,50) Sham: 30,3 (25,61)

tDCS characteristics

Paper Year	Country	Active Groups (n)	Sham Group (n)	Active electrode location	Intensity (mA)	Polarity (Active)	Reference Electrode Location *	Session duration (min)	Number of sessions	Sessions / week	Duration of tDCS intervention (weeks)	Type of Combined intervention	Pain outcome used	Use of Mechanistic outcomes	Disability tool	Disability score at baseline (SD)
Jales Junior 2015 ⁴³	Brazil	10	10	M1	1	Anode	Fp2	20	10	1	10	-	VAS	Yes	SF36-PF	Active MI: 48 (16.36) Sham: 31 (23.7)
Khedr 2017 ⁴⁴	Egypt	18	18	M1	2	Anode	extra- cephalic (arm)	20	10	5	2	-	VAS	No	-	-
Mattis 2022	Brazil	17	14	M1	2	Anode	Fp2	20	5	5	1	Exercise-based	VAS	Yes	FIQ	Active MI: 67 (15.76) Sham: 73.44 (11.23)
Melo 2020 ²³	Brazil	13 12	13	M1	2	Anode	NR	20	10 5	5	2 1	-	VAS	Yes	CIRS	-
Mendonça 2011 ³⁵	Brazil	6 6 6	6	M1 APFC M1 APFC	2	Cathodal Anode Anode	C-T Spine Transition	20	1	1	0.14	-	VNS	Yes	-	-
Mendonça 2016 ³⁶	Brazil	15	15	M1	2	Anode	Fp2	20	5	5	1	Exercise-based	VNS	Yes	SF36-PF	Active MI + Exercise: 41 (4.8) Sham + Exercise: 45 (4.1)
Ribeiro 2011 ⁴⁰	Brazil	11	12	M1	2	Anode	Fp2	20	10	1	10	-	SF36 (Pain)	No	FIQ	Active MI: 62.79 (15.03) Sham: 62.6 (18.08)
Samartin-Veiga 2021 ⁴⁹	Spain	34 33 33	30	M1 DLPFC OIC	2	Anode	Fp2 Ear Lobe	20	15	5	3	-	FIQ and VAS	Yes	-	-
Santos 2018 ⁴¹	Brazil	19	20	DLPFC	2	Anode	Fp2	20	8	7	1.14	Cognitive-behavioral based	VAS	Yes	FIQ	Active DLPFC: 63.43 (18.29) Sham: 66.16 (15.31)

tDCS characteristics

Paper Year	Country	Active Groups (n)	Sham Group (n)	Active electrode location	Intensity (mA)	Polarity (Active)	Reference Electrode Location *	Session duration (min)	Number of sessions	Sessions / week	Duration of tDCS intervention (weeks)	Type of Combined intervention	Pain outcome used	Use of Mechanistic outcomes	Disability tool	Disability score at baseline (SD)
To 2017 ⁴⁶	Belgium	15 11	16	OCC DLPFC	1.5	Anode	O2 Fp2	20	8	2	4	-	NRS	No	-	-
Valles 2009 ⁴²	Brazil	14 13	14	M1 DLPFC	2	Anode	Fp2	20	10	5	2	-	VAS	No	FIQ	Active MI: 90.95 (23.94) Active DLPFC: 92.62 (18.29) Sham: 96.28 (28.89)

OCC: occipital region; PFC: prefrontal cortex; DLPFC: dorsolateral prefrontal cortex; M1: primary motor cortex; APFC: anterior prefrontal cortex; OIC: operculo-insular cortex; NRS: numeric rating scale; VAS: visual analog scale; VNS: vagus nerve stimulation; SF36: short form 36; FIQ: fibromyalgia impact questionnaire; CIRTS: cumulative illness rating scale; ma: milli-Ampere; min: minutes; SD: standard deviation;

* according to the EEG 1020 system;

NR: Not Reported

The GRADE “Summary of Findings” table for quality of evidence, magnitude of effect of the interventions at short- and medium to moderate-term follow-up.

Table 3.

Identifying methodology, population, and intervention factors associated with the analgesic effect of tDCS in Fibromyalgia. A Systematic Review, Meta-Analysis, and Meta-Regression			
Patient or population: Patients with diagnosed fibromyalgia			
Settings: Outpatient			
Intervention: Active tDCS with or without a combined intervention			
Comparison: Sham tDCS			
Outcomes	Effect size	Relative and absolute effect (average % improvement (reduction) in pain (95% CIs) in relation to post-treatment score from sham group) ** Where 95% CIs do not cross the line of no effect.	No of participants (studies) Quality of the evidence (GRADE)
Pain intensity – Short Term Effect (< 4 weeks from baseline)			
Pain intensity measured using a visual analog scale or numerical rating scale	SMD 0.81 95%CI (0.43, 1.19)	This equates to a 27%, 95% CI (14% to 40%) reduction in pain intensity, on a 0 to 100 pain intensity scale	770 total. 401 on active and 369 on sham. (14 Studies) ⊕⊕○○ low ¹
Pain intensity – Moderate Term Effect (> 4 weeks from baseline)			
Pain intensity measured using a visual analog scale or numerical rating scale	SMD 1.689 95%CI (0.84, 2.54)	This equates to a 57%, 95% CI (28% to 85%) reduction in pain intensity, on a 0 to 100 pain intensity scale	375 total. 193 on active and 182 on sham. (7 Studies) ⊕○○○ very low ^{2, 3, 4}

GRADE Working Group grades of evidence

High quality: we are very confident that the true effect lies close to that of the estimate of the effect
Moderate quality: we are moderately confident in the effect estimate; the true effect is likely to be close to the estimate of effect, but there is a possibility that it is substantially different
Low quality: our confidence in the effect estimate is limited; the true effect may be substantially different from the estimate of the effect
Very low quality: we have very little confidence in the effect estimate; the true effect is likely to be substantially different from the estimate of effect

¹ Downgraded twice. Single randomized study (with under 300 participants) was considered inconsistent, imprecise (i.e., wide 95%CI) and with potential publication bias, providing low quality of evidence.
² Downgraded once for inconsistency due to heterogeneity.
³ Downgraded once for study limitation due to high or unclear risk of bias.
⁴ Downgraded once for imprecision due to small sample size
 SMD: Standardized Mean Difference; CI: confidence Interval