



# Trimethoprim-sulfamethoxazole and the risk of a hospital encounter with hyperkalemia: a matched population-based cohort study

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## ABSTRACT

**Background.** Trimethoprim-sulfamethoxazole (TMP-SMX) can cause hyperkalemia by reducing renal potassium excretion. We assessed the risk of hyperkalemia after initiating TMP-SMX versus amoxicillin and determined if this risk is modified by a patient's baseline kidney function [estimated glomerular filtration rate (eGFR)].

**Methods.** We conducted a population-based cohort study in Ontario, Canada involving adults  $\geq 66$  years of age newly treated with TMP-SMX ( $n = 58\,999$ ) matched 1:1 with those newly treated with amoxicillin (2008–2020). The primary outcome was a hospital encounter with hyperkalemia defined by a laboratory serum potassium value  $\geq 5.5$  mmol/L within 14 days of antibiotic treatment. Secondary outcomes included a hospital encounter with acute kidney injury (AKI) and all-cause hospitalization. Risk ratios (RRs) were obtained using a modified Poisson regression.

**Results.** A hospital encounter with hyperkalemia occurred in 269/58 999 (0.46%) patients treated with TMP-SMX versus 80/58 999 (0.14%) in those treated with amoxicillin [RR 3.36 [95% confidence interval (CI) 2.62–4.31]]. The absolute risk of hyperkalemia in patients treated with TMP-SMX versus amoxicillin increased progressively with decreasing eGFR (risk difference of 0.12% for an eGFR  $\geq 60$  ml/min/1.73 m<sup>2</sup>, 0.42% for eGFR 45–59, 0.85% for eGFR 30–44 and 1.45% for eGFR  $< 30$ ; additive interaction  $P < .001$ ). TMP-SMX versus amoxicillin was associated with a higher risk of a hospital encounter with AKI [RR 3.15 (95% CI 2.82–3.51)] and all-cause hospitalization [RR 1.43 (95% CI 1.34–1.53)].

**Conclusions.** The 14-day risk of a hospital encounter with hyperkalemia was higher in patients newly treated with TMP-SMX versus amoxicillin and the risk was highest in patients with a low eGFR.

**Keywords:** acute kidney injury, chronic kidney disease, hyperkalemia, trimethoprim, trimethoprim-sulfamethoxazole

## INTRODUCTION

Trimethoprim-sulfamethoxazole (TMP-SMX) is a common antibiotic used to treat bacterial infections of the urinary tract, skin and soft tissues [1–4]. TMP-SMX is primarily excreted by the kidneys and can cause hyperkalemia through inhibition of amiloride-sensitive channels in the renal collecting ducts, reducing potassium excretion by the kidneys [5, 6].

Large observational studies report a higher risk of hyperkalemia in patients prescribed TMP-SMX versus other antibiotics. However, hyperkalemia in these studies was assessed solely with administrative database diagnosis codes [7–9] or with a combination of diagnosis codes and a primary care record showing an elevated potassium level [10]. The sensitivity of diagnosis codes to identify hyperkalemia is  $< 15\%$ , raising concerns about accurate outcome ascertainment [11]. Prior studies also did not adequately examine the risk of TMP-SMX-induced hyperkalemia in patients with chronic kidney disease, a growing segment of the population at risk for adverse drug events from renally excreted drugs.

We conducted a population-based cohort study of older adults who received an outpatient prescription for TMP-SMX versus amoxicillin. The primary outcome was a hospital encounter (emergency department visit or a hospital admission) with hyperkalemia assessed with laboratory measurements of serum potassium. Secondary outcomes were a hospital encounter with acute kidney injury (AKI), all-cause hospitalization and all-cause mortality. Prespecified subgroup analyses were conducted by the baseline category of estimated glomerular filtration rate (eGFR), sex, evidence of a urine

## KEY LEARNING POINTS

### What is already known about this subject?

- Trimethoprim-sulfamethoxazole (TMP-SMX) is an antibiotic that can cause hyperkalemia by reducing potassium excretion by the kidneys.
- Previous observational studies reporting a higher risk of hyperkalemia in patients prescribed TMP-SMX versus other antibiotics used administrative diagnosis codes to assess hyperkalemia.
- Administrative diagnosis codes for hyperkalemia are insensitive, raising concerns about accurate outcome ascertainment.

### What this study adds?

- After controlling for potential confounders, new treatment with oral TMP-SMX versus amoxicillin was associated with a 3-fold higher risk of a hospital encounter with hyperkalemia.
- Outcome of hyperkalemia was ascertained using laboratory serum potassium values, which is reflective of current clinical practice.
- The risk of hyperkalemia associated with TMP-SMX was assessed across four estimated glomerular filtration rate (eGFR) categories and the risk was highest in patients with a low eGFR.

### What impact this may have on practice or policy?

- Given the growing number of adults with chronic kidney disease globally and their predispositions to electrolyte disorders and adverse drug outcomes due to poor renal clearance, the study findings have the potential to inform antibiotic prescribing.

culture before the dispense date and whether the TMP-SMX dose was appropriate for the patient's level of eGFR [4, 5].

## MATERIALS AND METHODS

### Study design and setting

We designed a matched new-user, active-comparator, retrospective, population-based cohort study in Ontario, Canada (2008–2020) using provincial linked administrative healthcare databases housed at ICES ([ices.on.ca](http://ices.on.ca)). Ontario residents have universal access to outpatient and inpatient healthcare services. Residents  $\geq 65$  years of age additionally have universal coverage for prescription drugs [12]. Emigration from the province would be the only reason for lost follow-up and the rate is  $<0.5\%$ /year [13]. The use of data in this study was authorized under Section 45 of Ontario's Personal Health Information Protection Act, which does not require review by a research ethics board. We have reported this observational study following the Strengthening the Reporting of Observational studies in Epidemiology and REporting of studies Conducted using Observational Routinely collected Data specific to pharmacoepidemiological research guidelines (Supplementary Table 1) [14, 15].

### Data sources

We ascertained study drug exposure, covariates and outcomes using linked administrative healthcare databases. We ascertained vital information from the Registered Persons Database, which contains demographic information on all Ontario residents. We obtained data on prescription records from the Ontario Drug Benefit (ODB) database, which contains highly accurate records on medication dispensing (the overall error rate is  $<1\%$ ) [12]. Prescriber data were ascertained from the ICES Physician Database and the ODB database.

Patient comorbidities were ascertained using the Canadian Institute for Health Information Discharge Abstract Database (CIHI-DAD) and the Ontario Health Insurance Plan (OHIP) database. The CIHI-DAD contains information on diagnoses and procedures during acute inpatient encounters. The OHIP database contains claims information for all physician services. Emergency department visits and hospitalization data were ascertained from the National Ambulatory Care Reporting System and CIHI-DAD, respectively. Laboratory measurements for serum potassium and creatinine were obtained from the Ontario Laboratories Information System, a provincial electronic repository of laboratory results from hospitals and community and public health laboratories [16]. These datasets were linked using unique encoded identifiers and analyzed at ICES.

### Study cohort

The study cohort comprised adults  $\geq 66$  years of age with a new outpatient prescription for TMP-SMX or amoxicillin between 1 January 2008 and 17 December 2020. The prescription dispense date served as the date of cohort entry (i.e. the index date). We restricted the cohort to patients with baseline measurements of both serum potassium and creatinine taken in the 365- to 7-day period before the index date. Serum creatinine measurements were standardized using isotope dilution mass spectrometry. We grouped patients by four prespecified eGFR categories:  $\geq 60$ , 45–59, 30–44 and  $<30$  ml/min/1.73 m<sup>2</sup> [17]. We calculated eGFR with the Chronic Kidney Disease Epidemiology Collaboration (CKD-EPI) equation using the patient's most recent serum creatinine measurement obtained in the 365- to 7-day period before the index date [18]. To ensure we were studying new antibiotic prescriptions, we excluded those who received TMP-SMX or amoxicillin in the 180-day period before the index date and

excluded patients who received other (nonstudy) antibiotics in the 180 days leading up to and including the index date [19]. To ensure study antibiotics were initiated in an outpatient setting, we excluded those with a hospital discharge or an emergency department visit in the 2 days before or on the index date. We also excluded patients who received potassium binders or had evidence of kidney failure (defined as receipt of chronic dialysis or a kidney transplant) in the 180-day period before the index date and excluded those with evidence of hyperkalemia (defined as a laboratory serum potassium value  $\geq 5.5$  mmol/L) in the 365-day period before the index date.

### Exposure and comparator groups

The exposure group comprised outpatients newly treated with oral TMP-SMX. The comparator group comprised outpatients newly treated with oral amoxicillin (amoxicillin alone or in combination with clavulanic acid). Amoxicillin was chosen as an active comparator because it is commonly prescribed to treat similar infections as TMP-SMX but is not associated with hyperkalemia [7, 10]. Patients in the exposed and comparator groups were matched 1:1 on *a priori* selected baseline characteristics as described in the Statistical analysis section.

### Outcomes

The primary outcome was a hospital encounter (an emergency department visit or a hospital admission) with hyperkalemia, defined by a laboratory serum potassium value  $\geq 5.5$  mmol/L. Three prespecified secondary outcomes were a hospital encounter with AKI, all-cause hospitalization and all-cause mortality. AKI was defined as an increase in serum creatinine concentration from the baseline value by either  $\geq 0.3$  mg/dl ( $\geq 26.5$   $\mu$ mol/L) or  $\geq 50\%$  or receipt of acute dialysis. As with serum potassium, the baseline serum creatinine value was the most recent measurement obtained in the 365- to 7-day period before the index date. Coding definitions for all-cause hospitalization and all-cause mortality are provided in Supplementary Table 2.

Patients were followed for 14 days after the index date to assess all study outcomes. This time frame reflects the typical antibiotic prescription duration and the period of acute exposure [8–10]. Hyperkalemia, when it occurs, is typically observed 5–10 days after TMP-SMX initiation [20, 21].

### Baseline characteristics

Patient characteristics, including demographics, comorbidities, coprescriptions, healthcare use and laboratory measurements, were obtained from our study databases as described above. Coding definitions for these characteristics are provided in Supplementary Table 3. We examined comorbidities in the 5-year period before the index date, coprescriptions in the 180-day period before the index date and healthcare visits and laboratory measurements in the 365-day period before the index date; for laboratory measurements, the most recent measurement before the index date was used. Data on

diagnostic tests such as urine cultures, chest X-rays, sputum collections, wound swabs and blood smears were examined in the 7-day period as well as 365-day period before the index date.

### Statistical analysis

We used both standard and propensity score matching to address confounding in this study. We matched new users of TMP-SMX and amoxicillin 1:1 on the following variables: eGFR category ( $\geq 60$ , 45–59, 30–44 or  $< 30$  ml/min/1.73 m<sup>2</sup>), sex, presence of a urine culture in the 7 days up to and including the index date and the logit of the propensity score for the predicted probability of newly receiving TMP-SMX [within  $\pm 0.2$  standard deviations (SD)]. We matched on sex and receipt of a urine culture to increase the probability that patients in the two groups received the study antibiotics for similar reasons and to facilitate subgroup analyses by sex and potential indication. The propensity score was derived from a multivariable logistic regression model with 146 measured baseline characteristics chosen *a priori*, including demographics, comorbidities, coprescriptions, healthcare use and laboratory measurements (see Supplementary Table 4 for a list of variables) [22]. Comorbidities and coprescriptions that may affect a patient's potassium level were carefully considered and included in the logistic model [23–25].

We summarized continuous variables as means and SDs or medians and interquartile ranges (IQRs; 25th–75th percentile) as appropriate and categorical variables as number and percentages. We used standardized differences to evaluate between-group differences on baseline characteristics in both the unmatched and matched cohorts [26]. Standardized differences  $> 10\%$  were considered meaningful. We estimated the risks of the primary and secondary outcomes in the matched cohort in both relative and absolute terms using risk ratios (RRs) and risk differences (RDs), respectively. Absolute risk was additionally expressed as the number needed to harm (1/RD). We estimated RRs and 95% confidence intervals (CIs) for the outcomes using modified Poisson regression models, using generalized estimating equations to account for the correlation induced by matching.

We tested for additive and multiplicative interaction for the primary outcome of hyperkalemia in the following four prespecified subgroups: eGFR category ( $\geq 60$ , 45–59, 30–44 and  $< 30$  ml/min/1.73 m<sup>2</sup>, sex (female or male), presence of a urine culture (absent or present) and appropriate TMP-SMX dose reduction for the level of eGFR (no or yes). A single-strength TMP-SMX tablet contains 80 mg TMP and a double-strength TMP-SMX tablet contains 160 mg TMP [5]. An appropriate daily dose for a patient's level of kidney function was defined as a TMP dose  $\leq 320$  mg for an eGFR  $\geq 30$  ml/min/1.73 m<sup>2</sup>, a TMP dose  $\leq 160$  mg for an eGFR of 15–29 ml/min/1.73 m<sup>2</sup> and a TMP dose  $\leq 80$  mg for an eGFR  $< 15$  ml/min/1.73 m<sup>2</sup> [4, 5].

### Additional analyses

The primary outcome of hyperkalemia was further assessed using more severe definitions: a serum potassium level

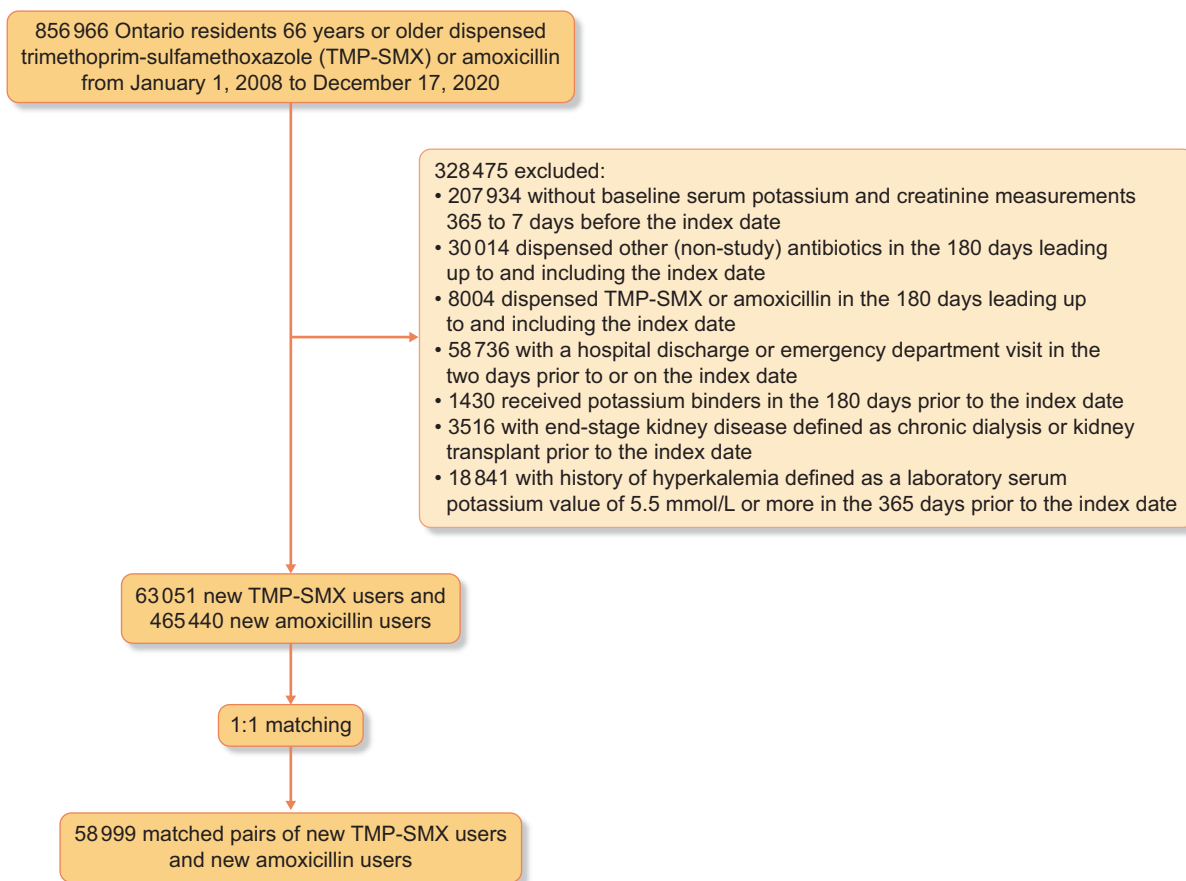


Figure 1: Cohort derivation.

$\geq 6.0$  mmol/L and  $\geq 6.5$  mmol/L. We calculated the *E*-value to evaluate the potential impact of unmeasured confounding on the risk estimate [27]. Briefly, an *E*-value indicates the strength of association that an unmeasured confounder needs to have with both the exposure and outcome to entirely explain the observed association between the two variables [28]. For example, an *E*-value of 6 would suggest that an observed association could be explained away by an unmeasured confounder that was associated with both the exposure and the outcome by an RR of 6-fold each. The higher the *E*-value the less likely the chance the observed association is due to unmeasured confounding. We performed all analyses with SAS version 9.4 (SAS Institute, Cary, NC, USA). In all outcome analyses we interpreted two-tailed *P*-values  $< .05$  as statistically significant.

## RESULTS

### Study cohort

Cohort selection is presented in Fig. 1. We identified 63 051 new TMP-SMX users and 465 440 new amoxicillin users between 2008 and 2020. Before matching, TMP-SMX users were on average older; more likely to be female; more likely to have a history of dementia, urinary tract infection, pneumonia and urinary retention; and were more likely to have coprescriptions for loop diuretics and angiotensin II

receptor blockers (Table 1 and Supplementary Table 4). TMP-SMX users had a lower mean eGFR than amoxicillin users, a higher Charlson comorbidity score and a greater number of coprescriptions. The full set of 146 characteristics is shown in Supplementary Table 4. A total of 58 999 TMP-SMX users were successfully matched 1:1 with 58 999 amoxicillin users (93.6% of TMP-SMX users). After matching, all characteristics were well balanced between the two groups (all standardized differences  $< 10\%$ ; Table 1 and Supplementary Table 4).

The median prescription duration for TMP-SMX was 7 days (IQR 5–7) and for amoxicillin it was 7 days (IQR 7–10). When examined by eGFR category, the median daily dose of TMP-SMX was 320 mg (IQR 320–320) in patients in the three highest eGFR categories ( $\geq 60$ , 45–59 and 30–44 ml/min/1.73 m<sup>2</sup>) and it was 320 mg (IQR 160–320) in patients with an eGFR  $< 30$  ml/min/1.73 m<sup>2</sup> (Table 2).

### Primary outcome: a hospital encounter with hyperkalemia

The primary outcome results are presented in relative and absolute terms in Table 3. The risk of a hospital encounter with hyperkalemia was higher in patients treated with TMP-SMX [269/58 999 (0.46%)] versus amoxicillin [80/58 999 (0.14%)]; the RR was 3.36 (95% CI 2.62–4.31) and the RD was 0.32% (CI 0.26–0.38).

**Table 1: Baseline characteristics of new TMP-SMX users and amoxicillin users before and after matching (2008–2020).**

Characteristics	Before matching			After matching		
	TMP-SMX (n = 63 051)	Amoxicillin (n = 465 440)	Standardized difference, %	TMP-SMX (n = 58 999)	Amoxicillin (n = 58 999)	Standardized difference, %
Age (years), mean (SD)	78.33 (8.6)	75.07 (7.6)	40	78.29 (8.6)	78.49 (8.8)	2
Female, n (%)	43 169 (68.5)	250 699 (53.9)	30	40 103 (68.0)	40 103 (68.0)	0
Year of cohort entry <sup>b</sup> , n (%)						
2008–2011	828 (1.2)	3243 (0.7)	5	768 (1.2)	517 (0.9)	3
2012–2014	8048 (12.8)	53 326 (11.5)	4	7548 (12.8)	6895 (11.7)	3
2015–2017	27 123 (43.0)	210 009 (45.1)	4	25 302 (42.9)	26 634 (45.2)	5
2018–2020	27 052 (42.9)	198 862 (42.7)	0	25 381 (43.0)	24 953 (42.3)	1
Rural residence, n (%)	10 051 (15.9)	48 131 (10.3)	17	9148 (15.5)	9152 (15.5)	0
Neighborhood income quintile <sup>c</sup> , n (%)						
1 (lowest income)	14 593 (23.1)	92 240 (19.8)	2	13 666 (23.2)	13 782 (23.4)	0
2	13 232 (21.0)	95 111 (20.4)	8	12 408 (21.0)	12 480 (21.2)	0
3	12 237 (19.4)	92 687 (19.9)	1	11 616 (19.7)	11 535 (19.6)	0
4	11 454 (18.2)	88 510 (19.0)	1	10 710 (18.2)	10 654 (18.1)	0
5 (highest income)	11 324 (18.0)	95 836 (20.6)	2	10 599 (18.0)	10 548 (17.9)	0
Long-term care residence, n (%)	11 497 (18.2)	22 417 (4.8)	43	10 398 (17.6)	10 289 (17.4)	1
Prescriber specialty <sup>d</sup> , n (%)						
General/family medicine	49 919 (79.2)	247 076 (53.1)	57	47 073 (79.8)	47 379 (80.3)	1
Nurse practitioner	3665 (5.8)	10 743 (2.3)	18	3277 (5.6)	3247 (5.5)	0
Urology	2829 (4.5)	2090 (0.4)	27	2235 (3.8)	1987 (3.4)	2
Internal medicine	429 (0.7)	2447 (0.5)	3	414 (0.7)	445 (0.8)	1
Nephrology	179 (0.3)	428 (0.1)	4	174 (0.3)	169 (0.3)	0
Dental	10 (0.0)	33 416 (7.2)	39	10 (0.0)	11 (0.0)	4
Other	3997 (6.3)	15 782 (3.4)	14	3795 (6.4)	4287 (7.3)	4
Comorbidities in prior 5 years <sup>e</sup> , n (%)						
Skin and soft tissue infection	27 516 (43.6)	181 789 (39.1)	9	25 676 (43.5)	25 796 (43.7)	0
Diabetes mellitus	17 797 (28.2)	132 318 (28.4)	0	16 772 (28.4)	16 977 (28.8)	1
Cancer	10 756 (17.1)	69 191 (14.9)	6	10 007 (17.0)	10 172 (17.2)	1
Urinary tract infection	10 680 (16.9)	26 789 (5.8)	36	9624 (16.3)	9291 (15.7)	2
Congestive heart failure	9369 (14.9)	51 309 (11.0)	12	8879 (15.0)	9132 (15.5)	1
Joint infection (septic arthritis)	9289 (14.7)	65 841 (14.1)	2	8680 (14.7)	8716 (14.8)	0
Pneumonia	4802 (7.6)	19 402 (4.2)	14	4464 (7.6)	4554 (7.7)	0
Chronic liver disease	2804 (4.4)	23 311 (5.0)	3	2657 (4.5)	2743 (4.6)	0
Other infection	23 013 (36.5)	172 556 (37.1)	1	21 733 (36.8)	22 167 (37.6)	2
Charlson Comorbidity Index, mean (SD) <sup>f</sup>	0.9 (1.5)	0.5 (1.2)	23	0.8 (1.5)	0.9 (1.5)	1
Coprescriptions in prior 180 days <sup>g</sup> , n (%)						
Glucocorticoid	17 243 (27.3)	120 428 (25.9)	3	16 186 (27.4)	16 443 (27.9)	1
Beta blockers	16 452 (26.1)	113 776 (24.4)	4	15 538 (26.3)	15 709 (26.6)	1
Calcium channel blocker	16 404 (26.0)	121 195 (26.0)	0	15 513 (26.3)	15 891 (26.9)	1
Angiotensin-converting enzyme inhibitor	12 243 (19.4)	99 717 (21.4)	5	11 607 (19.7)	11 958 (20.3)	2
Angiotensin II receptor blocker	9134 (14.5)	87 499 (18.8)	12	8762 (14.9)	9140 (15.5)	2
Loop diuretics	9068 (14.4)	40 814 (8.8)	18	8511 (14.4)	8695 (14.7)	1
Thiazide diuretics	7774 (12.3)	55 730 (12.0)	1	7283 (12.3)	7237 (12.3)	0
Nonsteroidal anti-inflammatory drug (excluding aspirin)	5958 (9.4)	45 955 (9.9)	2	5540 (9.4)	5418 (9.2)	1
Prednisone	3237 (5.1)	17 974 (3.9)	6	3030 (5.1)	3162 (5.4)	1
Antiarrhythmic agent	2167 (3.4)	12 656 (2.7)	4	2071 (3.5)	2211 (3.7)	1
Chemotherapeutic drugs	1416 (2.2)	8130 (1.7)	4	1332 (2.3)	1300 (2.2)	1
Tacrolimus	163 (0.3)	1237 (0.3)	0	151 (0.3)	161 (0.3)	0
Antiretroviral therapy	63 (0.1)	632 (0.1)	0	62 (0.1)	63 (0.1)	0
Cyclosporine	27 (0.0)	166 (0.0)	6	26 (0.0)	33 (0.1)	4
Healthcare visits in prior year, median (IQR)						
General/family medicine visits	9 (5–15)	7 (4–12)	–	9 (5–15)	9 (5–15)	–
Hospitalizations	0 (0–1)	0 (0–1)	–	0 (0–1)	0 (0–1)	–
Emergency department visits	0 (0–2)	0 (0–1)	–	0 (0–2)	0 (0–2)	–
Diagnostic tests in prior 7 days, n (%)						
Urine culture	32 463 (51.5)	38 160 (8.2)	107	29 183 (49.5)	29 183 (49.5)	0
Wound swab	2788 (4.4)	5521 (1.2)	19	2435 (4.1)	2664 (4.5)	2
Chest X-ray	1907 (3.0)	21 721 (4.7)	9	1879 (3.2)	2041 (3.5)	2
Sputum collection	79 (0.1)	318 (0.1)	0	76 (0.1)	87 (0.1)	0
Blood smear	29 (0.0)	108 (0.0)	19	29 (0.0)	26 (0.0)	2

**Table 1: (Continued.)**

Characteristics	Before matching			After matching		
	TMP-SMX (n = 63 051)	Amoxicillin (n = 465 440)	Standardized difference, %	TMP-SMX (n = 58 999)	Amoxicillin (n = 58 999)	Standardized difference, %
Laboratory measurements <sup>h</sup>						
eGFR (ml/min/1.73 m <sup>2</sup> ), mean (SD) <sup>i</sup>	66 (19)	70 (18)	19	66 (19)	66 (19)	0
eGFR categories (ml/min/1.73 m <sup>2</sup> ), n (%)						
≥60	40 393 (64.1)	335 855 (72.2)	17	37 048 (62.8)	37 048 (62.8)	0
45–59	13 251 (21.0)	83 466 (17.9)	8	12 734 (21.6)	12 734 (21.6)	0
30–44	7171 (11.4)	35 508 (7.6)	13	7011 (11.9)	7011 (11.9)	0
<30	2236 (3.5)	10 611 (2.3)	7	2206 (3.7)	2206 (3.7)	0
Serum potassium (mmol/L), mean (SD)	4.4 (0.4)	4.4 (0.4)	10	4.4 (0.4)	4.4 (0.4)	0

<sup>a</sup>The difference between the groups divided by the pooled SD; a value >10% is interpreted as a meaningful difference.

<sup>b</sup>The prescription dispense date was the date of cohort entry (i.e. the index date).

<sup>c</sup>Income was categorized into fifths of average neighborhood income on the index date.

<sup>d</sup>Information on prescriber specialty was available for 61 028 (96.8%) TMP-SMX users and 311 982 (67.0%) amoxicillin users before matching and 56 978 (96.6%) TMP-SMX users and 57 525 (97.6%) amoxicillin users after matching.

<sup>e</sup>Comorbidities were assessed in the 5-year period before the index date. Information on hyperaldosteronism is not presented due to the limited number of patients with this comorbidity.

<sup>f</sup>Charlson Comorbidity Index was calculated based on hospitalization data during the 5 years preceding the index date. For each patient, the index considers hospitalizations with the comorbidities of interest (acute myocardial infarction, congestive heart failure, peripheral vascular disease, cerebrovascular disease, dementia, chronic lung disease, rheumatic disease, peptic ulcer disease, mild and moderate/severe liver disease, diabetes mellitus with and without complications, hemiplegia/paraplegia, renal disease, cancer and metastatic solid tumor and acquired immunodeficiency syndrome/human immunodeficiency virus). It assigns a point score (1, 2, 3 or 6) for each comorbidity and sums them to generate an overall score of disease burden. The final risk scores range between 0 and 13, with higher values associated with higher mortality. Patients without a history of hospitalization received a score of 0.

<sup>g</sup>Coprescriptions were examined in the 180-day period before the index date.

<sup>h</sup>Laboratory measurements were examined in the 365-day period prior to the index date; the most recent value before the index date was used.

<sup>i</sup>eGFR was calculated based on an individual's most recent serum creatinine measurement found in the 365 to 7 days prior to the index date, using the CKD-EPI equation:  $141 \times \min([\text{serum creatinine concentration in } \mu\text{mol/L}/88.4]/\kappa, 1)^{\alpha} \times \max([\text{serum creatinine concentration in } \mu\text{mol/L}/88.4]/\kappa, 1)^{-1.209} \times 0.993^{\text{Age}} \times 1.018$  (if female)  $\times 1.159$  (if African American);  $\kappa = 0.7$  if female and 0.9 if male;  $\alpha = -0.329$  if female and  $-0.411$  if male; min = the minimum of serum creatinine concentration/ $\kappa$  or 1; max = the maximum of serum creatinine concentration/ $\kappa$  or 1. Information on race was not available in our data sources and all patients were assumed not to be of African Canadian race/ethnicity; African Canadians comprised <5% of the population of Ontario in 2006.

**Table 2: Daily dose of TMP across eGFR categories.**

eGFR category (ml/min/1.73 m <sup>2</sup> )	n	Mean (SD)	Median (IQR)	Patients receiving a higher than appropriate daily TMP dose, n (%)
≥60	37 048	313.80 (111.49)	320 (320–320)	515 (1.4)
45–59	12 734	304.46 (80.59)	320 (320–320)	165 (1.3)
30–44	7011	284.53 (101.37)	320 (300–320)	97 (1.4)
<30	2206	252.08 (86.64)	320 (160–320)	1379 (62.5)

<sup>a</sup>A TMP-SMX tablet contains 80 mg TMP and a TMP-SMX double-strength tablet contains 160 mg TMP. An appropriate dose for a patient's eGFR was defined as a daily TMP dose ≤320 mg/day for eGFR ≥30 ml/min/1.73 m<sup>2</sup>, ≤160 mg/day for eGFR 15–29 and ≤80 mg/day for eGFR <15.

### Secondary outcomes

The secondary outcome results are presented in relative and absolute terms in Table 3. The risk of a hospital encounter with AKI was higher in patients treated with TMP-SMX [1328/58 999 (2.25%)] versus amoxicillin [422/58 999 (0.72%)]; the RR was 3.15 (CI 2.82–3.51) and the RD was 1.54% (CI 1.40–1.67). The risk of all-cause hospitalization was higher in patients treated with TMP-SMX [2264/58 999 (3.84%)] versus amoxicillin [1581/58 999 (2.68%)]; the RR was 1.43 (CI 1.34–1.53) and the RD was 1.16% (CI 0.96–1.36). There was no significant difference in the risk of all-cause mortality between patients treated with TMP-SMX versus amoxicillin [RR 1.09 (CI 0.96–1.25)].

### Subgroup analyses

The results of the subgroup analyses are presented in Table 4. The absolute risk of hyperkalemia in patients treated

with TMP-SMX versus amoxicillin increased progressively with decreasing eGFR [at an eGFR ≥60 ml/min/1.73 m<sup>2</sup>, the RD was 0.12% (CI 0.06–0.18); at an eGFR of 45–59, the RD was 0.42% (CI 0.27–0.56); at an eGFR of 30–44, the RD was 0.85% (CI 0.58–1.11); and at an eGFR <30, the RD was 1.45% (CI 0.80–2.12); additive interaction  $P < .0001$ ]. The relative risk of hyperkalemia in patients treated with TMP-SMX versus amoxicillin was not significantly modified by the eGFR category (multiplicative interaction  $P = .13$ ).

The absolute risk of hyperkalemia in patients treated with TMP-SMX versus amoxicillin was higher in males than females (additive interaction  $P = .0013$ ); however, the relative risk was not significantly modified by sex (multiplicative interaction  $P = .24$ ). Neither the absolute risk nor the relative risk of hyperkalemia in patients treated with TMP-SMX versus amoxicillin was modified in patients with

**Table 3: Risk of a hospital encounter with hyperkalemia and other outcomes within 14 days of initiating TMP-SMX versus amoxicillin.**

Outcome	Event, n (%)		Risk difference, % (95% CI)	Number needed to harm (95% CI)	RR (95% CI)
	TMP-SMX (n = 58 999)	Amoxicillin (n = 58 999)			
Primary outcome					
Hyperkalemia <sup>a</sup>	269 (0.46)	80 (0.14)	0.32 (0.26–0.38)	313 (263–385)	3.36 (2.62–4.31)
Secondary outcomes					
AKI <sup>b</sup>	1328 (2.25)	422 (0.72)	1.54 (1.40–1.67)	65 (60–71)	3.15 (2.82–3.51)
All-cause hospitalization	2264 (3.84)	1581 (2.68)	1.16 (0.96–1.36)	86 (74–104)	1.43 (1.34–1.53)
All-cause mortality	467 (0.79)	427 (0.72)	0.07 (–0.03–0.17)	Not significant	1.09 (0.96–1.25)

<sup>a</sup>Hyperkalemia was defined as a serum potassium level  $\geq 5.5$  mmol/L.

<sup>b</sup>AKI was defined as an increase in serum creatinine concentration from the baseline value of  $\geq 0.3$  mg/dl ( $\geq 26.5$   $\mu$ mol/L) or an increase of  $\geq 50\%$ , or receipt of acute dialysis. The baseline serum creatinine was the most recent value 365 to 7 days prior to the index date.

<sup>c</sup>Risk differences and RRs were estimated within the matched cohort.

evidence of a urine culture (additive and multiplicative interaction  $P > .05$ ). The absolute risk of hyperkalemia in patients treated with TMP-SMX versus amoxicillin was higher in those prescribed a higher-than-recommended TMP-SMX dose for their eGFR than those prescribed an appropriate dose (additive interaction  $P = .0052$ ); however, the relative risk was not significantly modified (multiplicative interaction  $P = .90$ ).

### Additional analyses

The risk of hyperkalemia associated with TMP-SMX versus amoxicillin was similar using more severe definitions of hyperkalemia at a serum potassium level  $\geq 6.0$  mmol/L and  $\geq 6.5$  mmol/L (Supplementary Table 5). The calculated  $E$ -value for the RR and lower CI for the association between TMP-SMX and the primary outcome of hyperkalemia was 6.18 and 4.68 (Supplementary Fig. 1).

## DISCUSSION

In this real-world cohort study of older adults, we found that new treatment with TMP-SMX versus amoxicillin was associated with a higher risk of a hospital encounter with hyperkalemia within 14 days. The difference in the risk of hyperkalemia in patients treated with TMP-SMX versus amoxicillin increased progressively across lower eGFR categories, from 0.12% (1/834) among patients with an eGFR  $\geq 60$  ml/min/1.73 m<sup>2</sup> to 1.45% (1/69) among those with an eGFR  $< 30$  ml/min/1.73 m<sup>2</sup>. TMP-SMX versus amoxicillin use was also associated with a higher risk of a hospital encounter with AKI and all-cause hospitalization within 14 days. The difference in the risk of hyperkalemia with TMP-SMX versus amoxicillin use was higher in males than females and in those who received a higher-than-recommended daily dose of TMP for their level of eGFR [4, 5].

Our matched cohort of 117 998 older adults newly treated with TMP-SMX or amoxicillin provides robust estimates for the risk of a hospital encounter with hyperkalemia. The observed risk is supported by the known putative mechanism of reduced renal potassium excretion with TMP-SMX [6]. Our findings confirm the results of large cohort studies that examined the risk of hyperkalemia associated with TMP-

SMX [7–9]. Whereas other studies primarily used database diagnosis codes to define hyperkalemic events (which have low sensitivity), we defined hyperkalemia using elevated serum potassium concentrations from laboratory measurements [11]. This is a more accurate way to define hyperkalemia and likely explains the 2-fold greater 14-day incidence of hyperkalemia in TMP-SMX users in our study compared with prior studies [10]. Similarly, our study used increasing concentrations of serum creatinine to define AKI rather than database codes, which likely explains the 5-fold higher 14-day incidence of AKI in the TMP-SMX users in our study compared with other studies [10, 29].

As with all renally cleared drugs, clinicians can consider alternate, non-renally cleared antibiotics guided by local microbiologic data when treating patients with chronic kidney disease. When TMP-SMX is prescribed to patients with diminished kidney function, strategies to mitigate renal dosing errors, such as featuring the information from this study within computerized order-entry systems with dosing prompts, should also be considered [30, 31].

Our study has several strengths. We conducted a population-based cohort study in the most populous province of Canada, a region with universal health insurance coverage. We included all older adults with valid study antibiotic prescriptions. We had an appropriately short duration of follow-up of 14 days to best capture outcome events relevant to the acute antibiotic exposure [20, 21]. Using laboratory measurements, we overcame the limitation of relying on diagnosis codes to ascertain study outcomes, which may substantively underestimate the true event rate. Our previous validation of International Classification of Diseases, Tenth Revision codes for hyperkalemia and AKI showed these codes have poor sensitivity when measured against laboratory-based definitions [11, 29].

Our study also has limitations. As with any observational study, our findings remain subject to residual confounding, including confounding by indication. To minimize these confounding effects, we employed a new-user, active-comparator design and used propensity score matching to balance comparison groups on baseline health, including on microbiology and radiologic investigations [32]. Moreover, the observed  $E$ -value of 6.18 for the estimated RR and 4.68 for the lower CI suggest

**Table 4: Risk of a hospital encounter with hyperkalemia within 14 days of initiating TMP-SMX versus amoxicillin: subgroup analysis by eGFR categories, sex, presence of urine culture before antibiotic prescription and appropriate dose adjustment for kidney function.**

Subgroup	eGFR (ml/min/1.73 m <sup>2</sup> )	Exposure	n	Event, n (%)	Risk difference, % (95% CI)	Additive interaction P-value	RR (95% CI)	Multiplicative interaction P-value
	≥60	TMP-SMX	37 048	81 (0.22)	0.12 (0.06–0.18)	<.0001	2.25 (1.52–3.33)	.13
		Amoxicillin	37 048	36 (0.10)	Reference		Reference	
	45–59	TMP-SMX	12 734	69 (0.54)	0.42 (0.27–0.56)		4.31 (2.50–7.43)	
		Amoxicillin	12 734	16 (0.13)	Reference		Reference	
	30–44	TMP-SMX	7011	76 (1.08)	0.85 (0.58–1.11)		4.47 (2.66–7.52)	
		Amoxicillin	7011	17 (0.24)	Reference		Reference	
	<30	TMP-SMX	2206	43 (1.95)	1.45 (0.80–2.12)		3.91 (2.02–7.58)	
		Amoxicillin	2206	11 (0.50)	Reference		Reference	
Sex	Female	TMP-SMX	40 103	148 (0.37)	0.24 (0.18–0.31)	.0013	2.96 (2.15–4.07)	.24
		Amoxicillin	40 103	50 (0.12)	Reference		Reference	
	Male	TMP-SMX	18 896	121 (0.64)	0.48 (0.35–0.61)		4.03 (2.70–6.02)	
		Amoxicillin	18 896	30 (0.16)	Reference		Reference	
Urine culture within 7 days before or on the day of antibiotic initiation	Absent	TMP-SMX	29 816	152 (0.51)	0.36 (0.27–0.45)	.22	3.38 (2.43–4.70)	.97
		Amoxicillin	29 816	45 (0.15)	Reference		Reference	
	Present	TMP-SMX	29 183	117 (0.40)	0.28 (0.20–0.36)		3.34 (2.29–4.88)	
		Amoxicillin	29 183	35 (0.12)	Reference		Reference	
Appropriate dose reduction per kidney function <sup>a</sup>	No	TMP-SMX	2156	35 (1.62)	1.17 (0.55–1.78)	.0052	3.50 (1.73–7.07)	.90
		Amoxicillin	2156	10 (0.46)	Reference		Reference	
	Yes	TMP-SMX	56 843	234 (0.41)	0.29 (0.23–0.35)		3.34 (2.56–4.36)	
		Amoxicillin	56 843	70 (0.12)	Reference		Reference	

<sup>a</sup>A TMP-SMX tablet contains 80 mg TMP and a TMP-SMX double-strength tablet contains 160 mg TMP. An appropriate dose for a patient's eGFR was defined as a daily TMP dose ≤320 mg/day for eGFR ≥30 ml/min/1.73 m<sup>2</sup>, ≤160 mg/day for eGFR 15–<30 and ≤80 mg/day for eGFR <15.

<sup>b</sup>Risk differences and RRs were estimated within the matched cohort.



that substantial unmeasured confounding would be needed to nullify the observed association or its 95% CI [28].

Our study only included adults  $\geq 66$  years of age and thus the generalizability of the study results to younger patients is uncertain. Older adults are of a population of interest for investigating adverse drug events, as they are predisposed to such outcomes [33]. Also, the follow-up serum potassium measurements in this study were done in routine care rather than as part of a research protocol where all participants would have a measurement at a fixed time point in follow-up.

In conclusion, treatment with TMP-SMX associates with a 3-fold higher risk of a hospital encounter with hyperkalemia in older adults. This risk is greater among patients with a lower eGFR and those who received a higher-than-recommended daily dose of TMP-SMX. Our findings have the potential to inform prescribing practice in clinical settings that warrant consideration of TMP-SMX use, especially among patients with diminished kidney function.

## SUPPLEMENTARY DATA

Supplementary data are available at *ndt* online.

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## AUTHORS' CONTRIBUTIONS

Y.J.H., F.T.M. and A.X.G. conceived and designed the study. All authors contributed to the acquisition, analysis or interpretation of data. Y.J.H. drafted the manuscript. All authors contributed to critical revision of the manuscript for important

intellectual content. E.M. performed statistical analysis. All authors read and approved the submitted manuscript.

## DATA AVAILABILITY STATEMENT

The data set from this study is held securely coded from ICES. Although data sharing agreements prohibit ICES from making the data set publicly available, access may be granted to those who meet prespecified criteria for confidential access, available at <https://www.ices.on.ca/DAS>.

## CONFLICT OF INTEREST STATEMENT

All authors have declared no competing interests. Results presented in this article have not been published in whole or part.

## REFERENCES

1. Gupta K, Hooton TM, Naber KG *et al*. International clinical practice guidelines for the treatment of acute uncomplicated cystitis and pyelonephritis in women: a 2010 update by the Infectious Diseases Society of America and the European Society for Microbiology and Infectious Diseases. *Clin Infect Dis* 2011;**52**:e103–20. doi:10.1093/cid/ciq257.
2. Drekonja DM, Rector TS, Cutting A *et al*. Urinary tract infection in male veterans: treatment patterns and outcomes. *JAMA Intern Med* 2013;**173**:62. doi:10.1001/2013.jamainternmed.829.
3. Stevens DL, Bisno AL, Chambers HF *et al*. Practice guidelines for the diagnosis and management of skin and soft tissue infections: 2014 update by the Infectious Diseases Society of America. *Clin Infect Dis* 2014;**59**:e10–52. doi:10.1093/cid/ciu296.
4. May DB. Trimethoprim-sulfamethoxazole: an overview. In: *UpToDate*. Hooper D.C., Mitty J. (Eds.). 2021. [https://www.uptodate.com/contents/trimethoprim-sulfamethoxazole-an-overview?search=tmptomx&source=search\\_result&selectedTitle=1~150&usage\\_type=default&display\\_rank=1#H7](https://www.uptodate.com/contents/trimethoprim-sulfamethoxazole-an-overview?search=tmptomx&source=search_result&selectedTitle=1~150&usage_type=default&display_rank=1#H7) (19 November 2021, date last accessed).
5. U.S. Food and Drug Administration. Bactrim™ sulfamethoxazole and trimethoprim DS (double strength) tablets and tablets USP. June 2013. [https://www.accessdata.fda.gov/drugsatfda\\_docs/label/2013/017377s068s073lbl.pdf](https://www.accessdata.fda.gov/drugsatfda_docs/label/2013/017377s068s073lbl.pdf) (19 November 2021, date last accessed).
6. Muto S, Tsuruoka S, Miyata Y *et al*. Effect of trimethoprim-sulfamethoxazole on Na and K<sup>+</sup> transport properties in the rabbit cortical collecting duct perfused in vitro. *Nephron Physiol* 2006;**102**:51–60. doi:10.1159/000089682.
7. Lam N, Weir MA, Juurlink DN *et al*. Hospital admissions for hyperkalemia with trimethoprim-sulfamethoxazole: a cohort study using health care database codes for 393,039 older women with urinary tract infections. *Am J Kidney Dis* 2011;**57**:521–3. doi:10.1053/j.ajkd.2010.11.006.
8. Antoniou T, Gomes T, Juurlink DN *et al*. Trimethoprim-sulfamethoxazole-induced hyperkalemia in patients receiving inhibitors of the renin-angiotensin system: a population-based study. *Arch Intern Med* 2010;**170**:1045. doi:10.1001/archinternmed.2010.142.
9. Antoniou T, Gomes T, Mamdani MM *et al*. Trimethoprim-sulfamethoxazole induced hyperkalaemia in elderly patients receiving spirinolactone: nested case-control study. *BMJ* 2011;**343**:d5228. doi:10.1136/bmj.d5228.
10. Crellin E, Mansfield KE, Leyrat C *et al*. Trimethoprim use for urinary tract infection and risk of adverse outcomes in older patients: cohort study. *BMJ* 2018;**360**:k341. doi:10.1136/bmj.k341.
11. Fleet JL, Shariff SZ, Gandhi S *et al*. Validity of the International Classification of Diseases 10th revision code for hyperkalaemia in elderly patients at presentation to an emergency department and at hospital admission. *BMJ Open* 2012;**2**:e002011. doi:10.1136/bmjopen-2012-002011.
12. Levy AR, O'Brien BJ, Sellors C *et al*. Coding accuracy of administrative drug claims in the Ontario Drug Benefit database. *Can J Clin Pharmacol* 2003;**10**:67–71.

13. Statistics Canada. Table 17-10-0022-01. Estimates of interprovincial migrants by province or territory of origin and destination, annual. 29 September 2021. <https://www150.statcan.gc.ca/t1/tbl1/en/tv.action?pid=1710002201> (10 January 2022, date last accessed).
14. von Elm E, Altman DG, Egger M *et al*. The Strengthening of Reporting of Observational Studies in Epidemiology (STROBE) statement: guidelines for reporting observational studies. *Ann Intern Med* 2007;**147**: 573–7.
15. Langan SM, Schmidt SA, Wing K *et al*. The reporting of studies conducted using observational routinely collected health data statement for pharmacoepidemiology (RECORD-PE). *BMJ* 2018;**363**:k3532. doi:10.1136/bmj.k3532.
16. Iskander C, McArthur E, Nash DM *et al*. Identifying Ontario geographic regions to assess adults who present to hospital with laboratory-defined conditions: a descriptive study. *CMAJ Open* 2019;**7**:E624–9. doi:10.9778/cmaj.20190065.
17. Levey AS, Eckardt KU, Dorman NM *et al*. Nomenclature for kidney function and disease: report of a Kidney Disease: Improving Global Outcomes (KDIGO) Consensus Conference. *Kidney Int* 2020;**97**:1117–29. doi:10.1016/j.kint.2020.02.010.
18. Levey AS, Stevens LA, Schmid CH *et al*. A new equation to estimate glomerular filtration rate. *Ann Intern Med* 2009;**150**:604–12. doi:10.7326/0003-4819-150-9-200905050-00006.
19. Ray WA. Evaluating medication effects outside of clinical trials: new-user designs. *Am J Epidemiol* 2003;**158**:915–20. doi:10.1093/aje/kwg231.
20. Greenberg S, Reiser IW, Chou SY *et al*. Trimethoprim-sulfamethoxazole induces reversible hyperkalemia. *Ann Intern Med* 1993;**119**:291–5.
21. Alappan R. Hyperkalemia in hospitalized patients treated with trimethoprim-sulfamethoxazole. *Ann Intern Med* 1996;**124**:316. doi:10.7326/0003-4819-124-3-199602010-00006.
22. Austin PC. An introduction to propensity score methods for reducing the effects of confounding in observational studies. *Multivariate Behav Res* 2011;**46**:399–424. doi:10.1080/00273171.2011.568786.
23. Adelborg K, Nicolaisen SK, Hasvold P *et al*. Predictors for repeated hyperkalemia and potassium trajectories in high-risk patients—a population-based cohort study. *PLoS One* 2019;**14**:e0218739. doi:10.1371/journal.pone.0218739.
24. Thomsen RW, Nicolaisen SK, Hasvold P *et al*. Elevated potassium levels in patients with chronic kidney disease: occurrence, risk factors and clinical outcomes—a Danish population-based cohort study. *Nephrol Dial Transplant* 2018;**33**:1610–20. doi:10.1093/ndt/gfx312.
25. Thomsen RW, Nicolaisen SK, Hasvold P *et al*. Elevated potassium levels in patients with congestive heart failure: occurrence, risk factors, and clinical outcomes: a Danish population-based cohort study. *J Am Heart Assoc* 2018;**7**:e008912. doi:10.1161/JAHA.118.008912.
26. Austin PC. Using the standardized difference to compare the prevalence of a binary variable between two groups in observational research. *Commun Stat Simul Comput* 2009;**38**:1228–34. doi:10.1080/03610910902859574.
27. Mathur M, Ding P, Riddell C *et al*. E-value calculator. 2020. <https://www.evalue-calculator.com/> (1 November 2021, date last accessed).
28. VanderWeele TJ, Ding P. Sensitivity analysis in observational research: introducing the E-value. *Ann Intern Med* 2017;**167**:268–74. doi:10.7326/M16-2607.
29. Hwang YJ, Shariff SZ, Gandhi S *et al*. Validity of the international classification of diseases, tenth revision code for acute kidney injury in elderly patients at presentation to the emergency department and at hospital admission. *BMJ Open* 2012;**2**:e001821. doi:10.1136/bmjopen-2012-001821.
30. Tawadrous D, Shariff SZ, Haynes RB *et al*. Use of clinical decision support systems for kidney-related drug prescribing: a systematic review. *Am J Kidney Dis* 2011;**58**:903–14. doi:10.1053/j.ajkd.2011.07.022.
31. Erler A, Beyer M, Petersen JJ *et al*. How to improve drug dosing for patients with renal impairment in primary care - a cluster-randomized controlled trial. *BMC Fam Pract* 2012;**13**:91. doi:10.1186/1471-2296-13-91.
32. Yoshida K, Solomon DH, Kim SC. Active-comparator design and new-user design in observational studies. *Nat Rev Rheumatol* 2015;**11**:437–41. doi:10.1038/nrrheum.2015.30.
33. Evans RS, Lloyd JF, Stoddard GJ *et al*. Risk factors for adverse drug events: a 10-year analysis. *Ann Pharmacother* 2005;**39**:1161–8. doi:10.1345/aph.1E642.

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