Association between dietary patterns and adiposity from 4 to 7 years of age

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Abstract

Objective: The present study aimed to evaluate the association of 4-year-old children's dietary patterns with adiposity at 7 years, according to child's sex, using a conceptual model.

Design: Prospective cohort study. Diet was assessed using an FFQ. Age- and sex-specific BMI standard deviation scores (Z-scores) were defined according to the WHO. Fat mass percentage (FM%), fat mass index (FMI) and waist-to-height ratio (WHtR) were also considered, converted into Z-scores using sex-specific means and standard deviations of the current sample. Dietary patterns were identified by latent class analysis and their association with adiposity was estimated by linear regression models.

Setting: Population-based birth cohort Generation XXI (Porto, Portugal, 2005–2006). Subjects: Children $(n 3473)$ evaluated at both 4 and 7 years of age.

Results: Three dietary patterns were identified: high in energy-dense foods (EDF); low in foods typically consumed at main meals and intermediate in snacks (Snacking); and higher in vegetables and fish and lower in EDF (Healthier, reference). The EDF dietary pattern at 4 years of age was positively associated with later BMI only in girls $(\beta = 0.075, 95\% \text{ CI } 0.009, 0.140, P\text{-interaction} = 0.046)$. The EDF dietary pattern was also associated with other adiposity indicators only in girls (FMI: $\beta = 0.071$, 95 % CI 0.000, 0.142; WHtR: $\beta = 0.094$, 95 % CI 0.023, 0.164). Snacking was not significantly associated with any marker of adiposity in either girls or boys.

Keywords Pre-school children Dietary patterns Adiposity Sexual dimorphism Cohort study

Conclusions: Although dietary patterns and adiposity persisted across the two ages in both sexes, EDF at 4 years of age increased adiposity at 7 years of age only in girls.

Although the prevalence of childhood obesity might be reaching a plateau in some countries^{$(1,2)$ $(1,2)$}, it has generally increased in the last four decades and remains high^{([3](#page-8-0))}. In Europe, the prevalence of overweight in children below 10 years of age is higher in southern countries and differences by sex have been reported^{[\(4\)](#page-8-0)}. In Portugal^{([5](#page-8-0))}, considering either Cole and Lobstein's^{[\(6\)](#page-8-0)} or the WHO's^{([7](#page-8-0))} criteria, prevalence of overweight among 7-year-olds is higher in girls than in boys.

Research on dietary determinants of childhood obesity has focused mainly on specific nutrients and foods, but this approach does not account for the complexity of the human diet. Analysis of dietary patterns considers both interactions between foods and clustered dietary behaviours[\(8](#page-8-0)–[10\)](#page-8-0) and may better capture the overall effect of dietary intake, including total energy intake, on adiposity, helping to better understand the relationship between diet and disease. Besides energy intake, the effect of nutrients on adiposity may also be better captured by overall measures of dietary intake, and both energy and nutrient intakes might have sex-specific effects on adiposity as suggested in a recent study showing that the effect of energy intake was independent of macronutrient distribution in girls, but not in boys $^{(11)}$ $^{(11)}$ $^{(11)}$.

Prospective studies examining the association of empirically derived dietary patterns practised at pre-school age with adiposity are scarce^{$(12,13)$ $(12,13)$} and have not found an association from 3 to 7 years or from 5 to 9 years of age. Methodological aspects in the identification of dietary patterns, such as subjective cut-off points on underlying dimensions, may not allow identification of all dietary patterns in a data set. Hence, using a distinct methodology that avoids subjective decisions, such as latent class analysis^{(14) (14)}, might be advantageous to explore unobserved variables (i.e. latent classes) that distinguish homogeneous clusters of individuals (e.g. dietary patterns) in a data set.

In children, sexual dimorphism is observed from an early age in both total and regional adiposity $(15-17)$ $(15-17)$ $(15-17)$ $(15-17)$ and hormone biology (i.e. insulin and leptin). Yet, prospective studies have not reported formal examination of an interaction of child's sex on the association between dietary patterns of pre-school children and adiposity. In addition, considering sexual dimorphism in BMI trajectories – with girls showing older age and lower BMI at the infancy peak and earlier adiposity rebound^{$(18,19)$ $(18,19)$} – and keeping in mind that sex differences in the susceptibility to environmental factors such as peer influence and restrictive parenting practices have been reported $^{(20)}$ $^{(20)}$ $^{(20)}$, distinct prospective associations of dietary patterns and adiposity in girls and boys could be possible.

Considering the paucity of evidence, that eating habits start to be established in early years tracking into later $life^(21,22)$ $life^(21,22)$ $life^(21,22)$ $life^(21,22)$ and that the evidence base for intervention at pre-school age is scarce^{([23\)](#page-9-0)}, prospective studies on pre-school children's dietary patterns and adiposity, taking into account child's sex, are particularly warranted. Hence, the aim of the present study was to evaluate if dietary patterns of 4-year-old children were associated with adiposity at 7 years of age, using a conceptual model to examine hypothesized sex-specific prospective associations.

Methods

Subjects

The present study was based on the population-based birth cohort Generation XXI (Porto, Portugal, 2005–2006), described elsewhere^{([24\)](#page-9-0)}. At child's age of 4 years, the entire cohort was invited to a re-evaluation conducted as previously outlined^{$(25,26)$ $(25,26)$ $(25,26)$} and, at 7 years, a new follow-up of the whole cohort took place (81 % participation). The present study considered a sub-sample of singleton children with information on body composition measured at both ages $(n3854)$. After exclusion of children with diseases that might influence dietary intake (e.g. cerebral palsy or food allergy, $n 17$), with missing information on dietary intake ($n\,46$ and $n\,81$ at 4 and 7 years, respectively) and those with incomplete information on potential confounders (exercise, n 12; maternal weigh/height, n 225), 3473 children were included in final analyses.

Dietary intake

Dietary intake was evaluated by an FFQ adapted from a previous questionnaire validated for the general adult population^{(27) (27)}, shortened for application in the current birth cohort. The FFQ was applied by trained staff to the child's main caregiver at both waves and queried about how many times, on average, children consumed thirty-five food items during the previous 6 months. Details on response options and conversion into daily frequencies are described elsewhere^{([26\)](#page-9-0)}. In a sub-sample of 2373 children with complete 3 d food diaries, the validity and reliability of the FFQ were assessed by Pearson's correlation coefficients and intraclass correlation coefficients for key food groups as measured by both methods. For foods most frequently consumed, weak-tomoderate correlations and fair-to-moderate agreement were observed. Significant positive moderate Pearson's correlations were found for vegetable soup $(r=0.54)$, $P < 0.001$), fruit $(r=0.42, P < 0.001)$, milk $(r=0.46,$ $P < 0.001$) and yoghurts ($r = 0.48$, $P < 0.001$). Intraclass correlation coefficients varied from 0·54 (95 % CI 0·51, 0·56) for vegetable soup to 0·17 (95 % CI 0·11, 0·32) for juices. In addition, 3 d food diaries showed that there were no severe outliers $(3 \times$ interquartile range) in daily energy intake in either girls or boys.

As detailed earlier^{[\(28](#page-9-0))}, sixteen food items/groups (i.e. milk, yoghurt, meat/eggs, sweets, soft drinks, fruit, vegetable soup, vegetables, cheese, fish, processed meat, rice/pasta/potatoes, bread, crisps, pizza/burger, salty pastry) were used to identify dietary patterns at 4 years. Due to low variability in frequency of consumption of some foods, items/groups were categorized into 1st quintile, 2nd to 4th quintiles combined and 5th quintile or into two categories (1/d, 2/d) for foods eaten only once or twice daily (e.g. vegetable soup). Because, at 4 years of age, the cut-off points for consumption categories were very similar in girls and boys, the cut-off points of the whole cohort were used to identify dietary patterns. The same sixteen items/groups and cut-off points observed at 4 years were defined at 7 years of age. This approach was defined *a priori* to examine if children changed dietary pattern from 4 to 7 years, rather than to identify new patterns at 7 years. To obtain the dietary patterns at 7 years of age, we calculated the posterior probability of children belonging to each latent class from the parameters estimated by the latent class model at 4 years, considering the food group variables from 7 years.

Anthropometric data

Child's weight and height were measured at both ages as previously described^{(26) (26)}. Age- and sex-specific BMI standard deviation scores (BMI Z-scores) were computed according to the WHO^{$(7,29)$ $(7,29)$ $(7,29)$}. Fat mass was assessed by tetrapolar bioelectric impedance (BIA 101 Anniversary®; Akern, Florence, Italy) according to standard procedures^{(30) (30)} and equipment outputs were converted into fat-free mass (31) (31) from which fat mass percentage (FM%) and fat mass index (FMI; kg/m²) were derived. Waist circumference was measured at the umbilicus $level^{(32)}$ $level^{(32)}$ $level^{(32)}$ and waist-to-height ratio (WHtR) was calculated. To improve comparability, FMI, FM% and WHtR were converted into Z-score units

using the sex-specific means and standard deviations observed in the sample.

Potential confounders

Potential confounders were selected based on both previous literature and statistical methods, as described in the next section.

Maternal BMI and completed schooling years at child's age of 4 years were included as continuous variables in all analyses. Maternal work status was categorized into not working v . working and maternal smoking status was categorized into smoking v . no.

Child's exact age was considered in years. Any breastfeeding was categorized into $\langle 6 \ v. \ \rangle$ months. Regular practice of structured physical exercise was considered as a dichotomous variable (no v . yes) and screen time was considered as >2 v. \leq 2 h/d.

Statistical analysis

Considering the categorical nature of the response items in the FFQ and their asymmetric distribution, children's dietary patterns were identified by latent class analysis^{[\(14\)](#page-8-0)}. This methodology is a person-centred approach, used to distinguish homogeneous clusters of individuals from a sample[\(14,](#page-8-0)[33\)](#page-9-0). In order to avoid selection of spurious latent classes (type I error) that is more common when using powerful tests such as the Lo–Mendell–Rubin likelihood ratio test, in the present study the number of dietary patterns was defined according to the Bayesian information criterion, which is a criterion that performs well in large sample sizes^{(34) (34) (34)}. Individuals were assigned to each pattern according to the highest probability of class membership and probabilities of choosing each item response (e.g. consumption categories), conditionally on class membership, can be interpreted based on item profiles in each cluster^{$(33,35)$ $(33,35)$}. As energy intake can be an intermediate step in the pathway between diet and adiposity, dietary patterns were not adjusted for energy.

The association of dietary patterns with adiposity was assessed by linear regression coefficients (β) and respective 95 % CI. Four separate models (BMI, FM%, FMI and WHtR) were fitted. Confounding factors were assessed in each model and those not associated with the outcome or that did not change the association of interest were not considered in the final models (maternal age, work status and smoking habits). For theoretical reasons, children's practice of physical exercise, daily screen time and breast-feeding were included in all models. Furthermore, as energy intake was considered a possible intermediate step in the pathway between dietary patterns and adiposity, models were not adjusted for it. The interaction of child's sex on the association with adiposity was examined by a product term added into fully adjusted models.

In order to assess relationships of diet and adiposity across time, an analysis based on a theoretical concept (Fig. 1) was also conducted, including three components: (i) a crossed analysis of exposure at time 1 on outcome at

Fig. 1 Prospective relationships were analysed considering associations of: (a) dietary patterns at 4 years with adiposity at 7 years (results in [Table 3\)](#page-5-0); (b) adiposity at 4 years with later dietary patterns (results in [Table 4](#page-5-0)); (c) dietary patterns across ages (results in [Table 5](#page-6-0)); and (d) adiposity across time (results in [Table 6](#page-6-0))

time 2 (i.e. the main association of interest, denoted (a) in Fig. 1) and of outcome at time 1 on exposure at time 2 (denoted (b) in Fig. 1); (ii) an analysis of the exposure across time (denoted (c) in Fig. 1); and (iii) an analysis of the outcome across the two ages (denoted (d) in Fig. 1).

Associations of adiposity measures across time were assessed by linear regression models. Associations of adiposity at 4 years with later dietary pattern, as well as of patterns across the two ages, were assessed by OR and respective 95% CI estimated by multinomial logistic regression models. In linear regression models homoscedasticity of the error and normality of error distribution were assessed, while in multinomial logistic regression models goodness-of-fit was assessed by Pearson's χ^2 tests.

Statistical analyses were conducted using the statistical software package IBM SPSS Statistics version 22 and Mplus version 5.2 (Muthén & Muthén, Los Angeles, CA, USA), and a significance level of 5 % was considered.

Results

Three dietary patterns were identified at 4 years of age [\(Table 1\)](#page-3-0): (i) a pattern characterized by higher proportions of children within the 5th quintile for sweets, soft drinks, salty pastry and processed meat (named 'Energy-Dense Foods' (EDF)); (ii) a pattern lower in foods typically eaten at main meals (e.g. vegetables on a plate, fish, meat and eggs, rice/ pasta/potatoes) and intermediate in snacks (named 'Snacking'); and (iii) a dietary pattern higher in healthy foods and lower in unhealthier ones (named 'Healthier' and used as reference category). At 4 years, 43·7 % of girls and 44·5 % of boys followed the EDF dietary pattern, while at 7 years [\(Table 2\)](#page-4-0) proportions increased to 48·8 and 52·7 %, respectively. In contrast, proportions of children following the Snacking pattern decreased from 4 to 7 years of age in both sexes. Likewise, class membership in the Healthier dietary pattern decreased.

Table 1 Dietary patterns at 4 years of age: proportions of children in lower and higher consumption categories by sex (n 3473)*, Generation XXI prospective cohort study (Porto, Portugal, 2005–2006)

LC, latent class; EDF, Energy-Dense Foods dietary pattern.

*Proportions were compared by χ^2 test considering Bonferroni's correction for post hoc multiple comparisons. Unlike superscript letters indicate significant differences between dietary patterns at a significance level of 5 %.

†In order to avoid redundancy, reference category (2nd to 4th quintiles, combined) is not presented.

Prospective relationships between dietary patterns and adiposity are presented according to the conceptual model depicted in [Fig. 1](#page-2-0). The association of dietary patterns at 4 years of age with adiposity 3 years later (association (a) in [Fig. 1](#page-2-0); [Table 3\)](#page-5-0) was modified by child's sex. The interaction term was statistically significant in the BMI Z-score model $(P=0.046)$ and was consistent in other models (FMI, $P = 0.053$; WHtR, $P = 0.063$). Hence, we

chose to show final models stratified by child's sex. In crude models (model 1, [Table 3\)](#page-5-0), girls following the EDF dietary pattern at 4 years of age, when compared with those following the Healthier pattern, had significantly higher BMI Z-scores (β = 0·142; 95 % CI 0·029, 0·256), FMI $(\beta = 0.108; 95\% \text{ CI } 0.008, 0.209)$ and WHtR $(\beta = 0.136; 95\%$ CI 0·036, 0·237), while in boys these associations were not significant. After adjustment for maternal (education, BMI)

Table 2 Dietary patterns at 7 years of age: proportions of children in lower and higher consumption categories by sex (n 3473)*, Generation XXI prospective cohort study (Porto, Portugal, 2005–2006)

LC, latent class; EDF, Energy-Dense Foods dietary pattern.

*Proportions were compared by χ^2 test considering Bonferroni's correction for post hoc multiple comparisons. Unlike superscript letters indicate significant differences between dietary patterns at a significance level of 5 %.

†In order to avoid redundancy, reference category (2nd to 4th quintiles, combined) is not presented.

and child's characteristics (any breast-feeding, physical exercise, screen time, exact age and adiposity at 4 years of age; model 2, [Table 3\)](#page-5-0), these associations were attenuated (BMI Z-score: $\beta = 0.075$; 95% CI 0.009, 0.140; FMI: β =0·071; 95% CI 0·000, 0·142; WHtR: β =0·094; 95% CI 0·023, 0·164), but retained statistical significance.

Associations of adiposity measures at 4 years of age with dietary patterns at 7 years of age are shown in [Table 4](#page-5-0) (association (b) in [Fig. 1](#page-2-0)). In both crude and adjusted models, none of the adiposity measures at 4 years of age was associated with subsequent dietary patterns in girls or boys.

Dietary patterns tracked across the two ages (association (c) in [Fig. 1](#page-2-0); results in [Table 5\)](#page-6-0). The EDF pattern at 4 years significantly increased the odds of children following the same pattern at 7 years in both girls $OR = 4.48$; 95 % CI 3·51, 5·70) and boys (OR=4·70; 95 % CI 3·68, 5·98). Table 3 Association of dietary patterns at 4 years of age with adiposity at 7 years of age by sex (n 3473)*, Generation XXI prospective cohort study (Porto, Portugal, 2005–2006)

β, regression coefficient; EDF, Energy-Dense Foods dietary pattern; FM%, fat mass percentage; FMI, fat mass index; WHtR, waist-to-height ratio; Ref., reference category.

*Results correspond to association (a) in [Fig. 1.](#page-2-0) Statistically significant associations are highlighted in bold.

†All adiposity measures are in Z-score units. BMI Z-scores were defined according to the WHO's age- and sex-specific standards. FM%, FMI and WHtR were converted into Z-scores using the sex-specific means and standard deviations observed in the study sample. ‡Crude.

§Adjusted for maternal (education, BMI) and child (any breast-feeding, physical exercise, screen time, exact age and each adiposity measure at 4 years) characteristics.

Table 4 Association of adiposity measures at 4 years of age with dietary patterns at 7 years of age by sex (n 3473)*, Generation XXI prospective cohort study (Porto, Portugal, 2005–2006)

EDF, Energy-Dense Foods dietary pattern; FM%, fat mass percentage; FMI, fat mass index; WHtR, waist-to-height ratio.

*Results correspond to association (b) in [Fig. 1.](#page-2-0)

†All adiposity measures are in Z-score units. BMI Z-scores were defined according to the WHO's age- and sex-specific standards. FM%, FMI and WHtR were converted into Z-scores using the sex-specific means and standard deviations observed in the study sample.

‡Reference category (healthier dietary pattern at 7 years) not shown in order to avoid redundancy. §Crude.

||Adjusted for maternal (education, BMI) and child (any breast-feeding, physical exercise practice, daily screen time, exact age and dietary pattern at 4 years) characteristics.

Table 5 Association of dietary patterns at 4 years of age with dietary patterns at 7 years of age by sex (n 3473)*, Generation XXI prospective

EDF, Energy-Dense Foods dietary pattern; Ref., reference category.

*Results correspond to association (c) in [Fig. 1](#page-2-0). Statistically significant associations are highlighted in bold.

†Reference category (healthier dietary pattern at 7 years) not shown in order to avoid redundancy.

‡Model 1 is adjusted for maternal (education, BMI) and child (any breast-feeding, physical exercise practice, daily screen time and exact age) characteristics. §Model 2 includes variables in Model 1 and is further adjusted for BMI Z-score at child's age of 4 years.

Likewise, the Snacking pattern persisted in time in both sexes.

Adiposity measures also tracked across the two ages (association (d) in [Fig. 1](#page-2-0); results in Table 6). In girls, each increase of 1 Z-score unit in BMI at 4 years significantly increased later BMI by 0·829 (95 % CI 0·802, 0·857) Z-score units; in boys, an increase of 1 Z-score unit in BMI at 4 years was associated with a significant increase of 0·851 (95 % CI 0·817, 0·885) Z-score units in BMI at 7 years of age.

Discussion

Present findings indicate that dietary patterns as well as adiposity measures persist from 4 to 7 years in both sexes, and that an EDF dietary pattern high in sweets, sugarsweetened beverages, salty pastry and processed meat followed at pre-school age is significantly positively associated with measures of total and central adiposity in 7-year-old girls, but not in boys.

Dietary patterns identified in the current study are comparable to those identified in other reports. Among pre-school children from the $UK^{(36,37)}$ $UK^{(36,37)}$ $UK^{(36,37)}$ three dietary patterns were identified at 3 and 4 years of age: a 'traditional' pattern characterized by meat and vegetables; a 'junk' pattern characterized by high-fat processed foods (comparable to the present EDF pattern); and a 'health conscious' pattern rich in vegetarian-style foods, cheese and fish (comparable to the current Healthier pattern). Additionally, snacking dietary patterns – partly comparable to the current Snacking pattern – were also identified at 3 years of age in children from the $UK^{(36)}$ $UK^{(36)}$ $UK^{(36)}$ and in 5–12–year-old children from Colombia^{[\(38\)](#page-9-0)}. Two other studies identified dietary patterns comparable to the current EDF pattern; a 'Processed and fast foods' pattern was identified at multiple time points among French children^{[\(39\)](#page-9-0)} and a 'Snacky' pattern was identified in Greek 4 -year-olds^{(40) (40)}.

Table 6 Association of adiposity measures across the two ages by sex (n 3473)*, Generation XXI prospective cohort study (Porto, Portugal, 2005–2006)

β, regression coefficient; FM%, fat mass percentage; FMI, fat mass index; WHtR, waist-to-height ratio.

*Results correspond to association (d) in [Fig. 1.](#page-2-0) Statistically significant associations are highlighted in bold.

†All adiposity measures are in Z-score units. BMI Z-scores were defined according to the WHO's age- and sex-specific standards. FM%, FMI and WHtR were converted into Z-scores using the sex-specific means and standard deviations observed in the study sample.

‡Model 1 is adjusted for maternal (education, BMI) and child (any breastfeeding, physical exercise practice, daily screen time and exact age) characteristics.

§Model 2 is further adjusted for dietary pattern at child's age of 4 years.

Tracking of dietary patterns is also comparable to reports from other studies^{$(13,39)$ $(13,39)$ $(13,39)$ $(13,39)$}. In the current sample, the odds of staying in the EDF pattern across the two ages were considerably higher than the odds of changing into the Snacking pattern and vice versa. Still, some interchangeability between these two patterns was observed and some children changed from the Snacking into the EDF pattern (girls, 42 %; boys 45 %). Keeping in mind that the 'snack' pattern identified among 3-year-olds from the $UK^{(22)}$ $UK^{(22)}$ $UK^{(22)}$ was not identified at 4, 7 or 9 years of age, it is possible that these 'snack'/'snacking' patterns are still reflecting a transition to the family diet and that – with increasing age – children tend to change into other patterns.

Considering the association of dietary patterns with adiposity, the 'junk' dietary pattern identified in 3-yearolds from the $UK^{(12)}$ $UK^{(12)}$ $UK^{(12)}$ was not associated with obesity at 7 years of age and the 'energy-dense, low-fibre, high-fat' pattern practised at 5 years was not significantly associated with fat mass at 7 years, although this pattern practised at 7 years was significantly positively associated with adiposity at 9 years of $age^{(13)}$ $age^{(13)}$ $age^{(13)}$. As these studies did not report sex-specific associations, comparison with the present findings is hampered. Regarding other ages, in a previous study conducted in our cohort from 2 to 4 years of age^{([41](#page-9-0))}, we did not find evidence of effect modification by sex on the association between consumption of EDF and BMI. In contrast, among adolescents from the USA, slight sex differences have been reported: a 'starchy foods' pattern was negatively associated with overweight in boys, while no association was found in girls^{(42) (42)}.

The positive association between the EDF dietary pattern and adiposity found in girls may be explained by the sexual dimorphism in adiposity that emerges before the onset of puberty^{$(17,43)$ $(17,43)$}, with girls showing higher relative body fat and abdominal subcutaneous adipose tissue compared with boys^{([44\)](#page-9-0)}. An earlier adiposity rebound reported in girls $⁽¹⁸⁾$ $⁽¹⁸⁾$ $⁽¹⁸⁾$ could also explain the present findings.</sup>

Considering the possibility of distinct BMI trajectories and distinct susceptibilities to certain environmental factors in girls and boys, prospective associations could – hypothetically – differ according to sex. Hence, the conceptual model presented in [Fig. 1](#page-2-0) was used in order to examine if such differences were observed. Yet there was no evidence of distinct prospective associations.

It should be highlighted that adjustment for BMI at 4 years of age diminished the magnitude of the association between the EDF dietary pattern and adiposity observed in girls. This suggests that, in girls, the EDF dietary pattern has a direct effect on later adiposity, but also an indirect effect mediated through adiposity at the age of 4 years. In a sensitivity analysis, we fitted models excluding children who were overweight at age 4 years and results were similar (data not shown), giving strength to the current findings. If the observed effect modification by sex is true, the present results may be reflecting the higher propensity to store body fat and higher risk of obesity reported in females^{$(15,16,20)$ $(15,16,20)$ $(15,16,20)$ $(15,16,20)$ $(15,16,20)$}. However, given that males and females are both susceptible to obesity, and that girls may have an earlier adiposity rebound, it is possible that the EDF pattern is positively associated with adiposity in boys at later stages.

Evaluation of distinct measures of adiposity is a strength of the present study. Since BMI is not able to distinguish between lean and fat mass, and the accuracy of FM% depends on height and cannot be assessed independently of free-fat mass, it is important to use an index adjusted for height (i.e. $FMI^{(45)}$ $FMI^{(45)}$ $FMI^{(45)}$. Since these measures do not indicate regional fat distribution, it was also crucial to assess a measure of central fatness (i.e. WHtR) as abdominal adiposity is associated with increased metabolic compli-cations in children^{[\(46](#page-9-0))}.

The use of latent class analysis to identify dietary patterns is an advantage, as this methodology is especially well suited for categorical variables asymmetrically distributed, common in FFQ. Latent class analysis focuses on relationships between individuals and describes how the probabilities of a set of observed categorical variables (e.g. FFQ items) vary across individuals, finding the smallest number of unobserved variables (i.e. latent classes) that distinguish homogeneous clusters of individuals (e.g. patterns). It has the advantage of avoiding subjective cut-off points on underlying dimensions, as it adds classes stepwise until the model fits the data well and the classification is provided directly by the $model⁽¹⁴⁾$ $model⁽¹⁴⁾$ $model⁽¹⁴⁾$.

Restricting the analysis of dietary patterns at 7 years to the consumption categories observed at 4 years deserves discussion. This approach was defined *a priori* to assess if children switched between dietary patterns from 4 to 7 years of age, rather than to identify new patterns at 7 years. Considering that median probabilities of class membership supported an acceptable classification of individuals into each latent class at both ages (data not shown), this approach was deemed appropriate.

Besides the prospective design and children's age frame, another strength of the current study was its sufficient statistical power to assess the interaction of child's sex on the association between diet and adiposity. In addition, as children were part of a large population-based cohort regularly followed, several potential confounding factors could be assessed.

Assessment of dietary intake by FFQ could be argued as a limitation. However, comparison with 3 d food diaries supported a reasonable validity and reliability of the FFQ data. Also, the EDF pattern was significantly higher in energy than the Snacking and Healthier patterns (girls, 6765 v. 6304 and 6576 kJ; boys, 7175 v. 6735 and 6910 kJ).

Inclusion of a sub-sample could have introduced some bias into the analysis. A previous study conducted in this $\text{cohort}^{(47)}$ $\text{cohort}^{(47)}$ $\text{cohort}^{(47)}$ managed missing data using full-information maximum likelihood estimation and comparison of this model with the model excluding missing values suggested that substantial bias was not expected (Cohen's κ = 0.85). Hence, it is not expected that alternative methods to manage missing data would result in different conclusions. In addition, as Cohen's effect size values are not high (<0.35) , the slight differences between this sample and the remaining cohort are likely due to large sample size rather than to substantial differences between participants^{[\(48](#page-9-0))}.

Children's dietary patterns and adiposity 1981

Conclusion

Children practising unhealthier dietary patterns at the age of 4 years were more likely to follow the unhealthier dietary patterns 3 years later. Likewise, higher adiposity persisted across the two ages in both girls and boys. In contrast, a dietary pattern high in EDF was consistently positively associated with adiposity 3 years later only in girls. Future studies should further examine if sex modifies the association between diet and adiposity in distinct populations at different ages, with particular focus on underlying metabolic pathways. Present findings suggest that interventions to prevent unhealthy dietary habits should focus on early dietary patterns and that girls in particular may need early attention.

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