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### Prospective analysis of plasma amyloid beta and postoperative delirium in the Interventions for Postoperative Delirium: Biomarker-3 study

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#### Abstract

**Background:** The effect of postoperative delirium on the amyloid cascade of Alzheimer's dementia is poorly understood. Using early postoperative plasma biomarkers, we explored whether surgery and delirium are associated with changes in amyloid pathways.

**Methods:** We analysed data from 100 participants in the Interventions for Postoperative Delirium: Biomarker-3 (IPOD-B3) cohort study in the USA (NCT03124303 and NCT01980511), which recruited participants aged >65 yr undergoing non-intracranial surgery. We assessed the relationship between the change in plasma amyloid beta ratio (AβR; Aβ42:Aβ40) and delirium incidence (defined by the 3-Minute Diagnostic Confusion Assessment Method) and severity (quantified by the Delirium Rating Scale-Revised-98, the study's primary outcome). We also tested the relationship between plasma amyloid beta and intraoperative variables.

**Results:** Across all participants, the plasma  $A\beta R$  increased from the preoperative period to postoperative Day 1 (Wilcoxon P<0.001). However, this increase was not associated with delirium incidence (Wilcoxon P=0.22) or peak severity after adjusting for confounders (log[incidence rate ratio]=0.43; P=0.14). Postoperative Day 1 change in plasma  $A\beta R$  was not associated with postoperative Day 1 change in plasma tau, neurofilament light, or inflammatory markers (interleukin [IL]-1 $\beta$ , IL-1Ra, IL-2, IL-4, IL-6, IL-8, IL-10, and IL-12), or with operative time or low intraoperative arterial pressure. **Conclusions:** Perioperative changes in plasma amyloid do not appear to be associated with postoperative delirium. Our findings do not support associations of dynamic changes in amyloid with postoperative delirium. **Clinical trial registration:** NCT03124303 and NCT01980511.

Keywords: Alzheimer's disease; amyloid beta; anaesthesia; delirium; dementia; neurocognitive disorders; surgery

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#### Editor's key points

- A relationship between postoperative delirium and the amyloid cascade of Alzheimer's disease has been suggested but is poorly understood.
- This study used early postoperative plasma biomarkers to explore whether surgery and delirium are associated with changes in amyloid pathways.
- An increase in the plasma amyloid beta ratio from preoperative baseline to postoperative Day 1 was observed but was not associated with the incidence or severity of postoperative delirium.
- These findings do not support associations of changes in plasma amyloid with postoperative delirium; further studies are required to validate other potential biomarkers.

Surgery and anaesthesia have been associated with perioperative neurocognitive disorders (NCDs), including short-term (postoperative delirium and delayed neurocognitive recovery)<sup>1</sup> and long-term (postoperative mild or major neurocognitive disorder)<sup>2,3</sup> manifestations. Postoperative delirium may play a causal role in accelerating long-term cognitive decline<sup>4,5</sup>; however, confounding factors in this relationship mean inferences of causation are tenuous. Moreover, not all studies have observed poorer overall long-term cognition in patients with postoperative delirium.<sup>6</sup> Investigation of this topic is hampered by uncertainty regarding the neuropathological mechanisms that could underlie this relationship.

The recent amyloid–tau–neurodegeneration (ATN) model provides a pathogenic framework of Alzheimer's disease (AD).<sup>7</sup> This model argues that accumulation of toxic amyloid beta (A $\beta$ ) species ('A') and intracellular hyperphosphorylated tau ('T'), with accompanying neurodegeneration ('N'), are pathological hallmarks of AD. Although designed to make no inferences of causality or temporal sequences of these events, the ATN framework implies a deterministic model that relies heavily on the amyloid hypothesis. The amyloid hypothesis posits that the pathophysiological cascade that results in AD is initially triggered by pathogenic A $\beta$  peptide aggregation.<sup>8</sup> Preclinical studies have shown that general anaesthetic exposure is associated with increased A $\beta$  production and aggregation,<sup>9</sup> possibly implicating general anaesthetics (especially volatile agents<sup>10</sup>) in the amyloid hypothesis.

Walker and colleagues<sup>11</sup> recently highlighted the need for a greater understanding of the biological link between postoperative delirium and long-term cognitive changes, particularly with respect to the amyloid hypothesis as a possible explanation. Previously, we have shown associations of delirium with markers of neurodegeneration (plasma neurofilament light [NfL]<sup>12,13</sup> and tau<sup>13,14</sup>), which have been replicated in other studies.<sup>15,16</sup> Herein, we extend our biomarker analyses focusing on the amyloid hypothesis.

Changes in the amyloid cascade that might accompany postoperative delirium are not well understood. Of two small human positron emission tomography (PET) imaging studies, one observed no correlation between cerebral amyloid burden and postoperative delirium,<sup>17</sup> whereas our pilot study reported a positive association with delirium severity.<sup>18</sup> The pitfall of these neuroimaging studies is that the static 'snapshot' of imaging might fail to track important short-term changes, which could influence the pathological trajectory longer term. Although low preoperative CSF A<sup>β</sup>42 has been associated with postoperative delirium in one study,<sup>19</sup> another study found no such relationship.<sup>20</sup> A recent small study by our group reported that CSF and plasma amyloid beta ratio (ABR; AB42:AB40) increased from the preoperative to postoperative period, but peak postoperative change in CSF and plasma  $A\beta R$  was not associated with peak delirium severity or delirium incidence.<sup>13</sup> However, an adequately powered sample was not obtained, and hence, we undertook a larger sample based on our previous observation<sup>13</sup> that CSF and plasma values for change in Aβ40 (Spearman  $\rho{=}0.929;$  P=0.007) and A\beta42 (Pearson r=0.793; P=0.033) on postoperative Day 1 (POD1) were correlated in the 13 patients with CSF samples. We focus on the plasma A $\beta$ R over A $\beta$ 40 or A $\beta$ 42 levels, as the former has been shown in meta-analysis to be more closely associated with cognitive decline.<sup>21</sup> Our approach is consistent with other studies using plasma A $\beta$ R in chronic settings.<sup>22,23</sup>

We hypothesise a causal pathway that anaesthesia/surgery alters the amyloid pathway (detected by fluctuations in plasma  $A\beta$ ) that is proportional to delirium severity with potential implications for long-term cognition. We contend that a strong correlation between perioperative changes in plasma amyloid and delirium, which is not explained by known confounders, would provide evidence for acceleration of the amyloid cascade, which might explain links between delirium and dementia. Our specific aims were to (i) investigate the relationship between plasma  $A\beta R$  and delirium incidence and severity, (ii) confirm the postoperative increase in plasma  $A\beta R$ in a larger cohort than our previous study, and (iii) establish covariates that could explain such an increase in plasma  $A\beta R$ 

#### Methods

We report outcomes from participants in the Interventions for Postoperative Delirium: Biomarker-3 (IPOD-B3) study (NCT03124303 and NCT01980511), which is an ongoing prospective observational cohort study in the USA enrolling patients undergoing non-intracranial surgery  $\geq$ 65 yr of age. The University of Wisconsin–Madison Institutional Review Board provided ethical approval for the study (2015-374). Further details on this cohort are described elsewhere.<sup>14,24</sup> The 14 participants in our previous study<sup>13</sup> with data for postoperative change in plasma A $\beta$ R were also included in this study. Strengthening the Reporting of Observational Studies in Epidemiology (STROBE) guidelines were used to report our findings.<sup>25</sup>

#### Outcomes

The primary outcome of this study was POD1 change in  $A\beta R$ relative to peak delirium severity, as defined by the Delirium Rating Scale-Revised-98 (DRS-R-98).<sup>26</sup> Consistent with our previous work,<sup>12</sup> POD1 change was preferred over peak postoperative change, as the latter is biased towards patients with longer postoperative stays.<sup>27</sup> Secondary outcomes included POD1 change in  $A\beta R$  based on delirium incidence, which was assessed using the 3-Minute Diagnostic Confusion Assessment Method (3D-CAM)<sup>28</sup> or CAM for the ICU (CAM-ICU)<sup>29</sup> (if intubated). Assessments were performed twice daily from postoperative Days 1-4, between 05:00-10:00 and 16:00-20:00 supplemented with a chart-based review.<sup>30</sup> Preoperative baseline cognitive testing was performed using the Montreal Cognitive Assessment (MoCA),<sup>31</sup> Trail Making Test B (TMTB),<sup>32</sup> and Controlled Oral Word Association Test (COWAT).<sup>33</sup> Intraoperative data, including operation time in minutes, arterial pressure, and blood loss, were obtained from the medical

record. Secondary analyses were performed for all outcomes using Aβ40 and Aβ42, and using the peak postoperative change in Aβ40, Aβ42, and AβR. The AβR was calculated by dividing the plasma concentration of Aβ42 by that of Aβ40. We emphasise the results of Aβ42 over Aβ40 because the former is thought to be the more pathogenic amyloid species.<sup>34</sup> We also analysed the correlation between postoperative change in amyloid and operative time, intraoperative arterial pressure, and anaesthetic dose to assess the relationship to intraoperative variables. Age-adjusted median sevoflurane (AMS) concentration was calculated by dividing the median sevoflurane concentration (in minimum alveolar concentration [MAC] units) by  $2.03*10^{(-0.00301*[age minus 40]).^{35}$ 

#### Biomarker collection and analysis

Venous blood was collected in ethylenediaminetetraacetic acid tubes before surgery and each morning (04:00–11:00) from POD1 to 4 as close as possible to the time of delirium assessment and stored at  $-80^{\circ}$ C. Samples were also taken at long-term follow-up, at POD90 and POD365. We sent samples to the University of Gothenburg for analysis using an ultrasensitive single-molecule array<sup>36</sup> immunoassay for quantification of plasma A $\beta$ 40 and A $\beta$ 42 according to the manufacturer (Quanterix, Billerica, MA, USA). Blood samples were analysed by laboratory technicians who were blinded to clinical details.

#### **Power analysis**

Our primary outcome of POD1 change in A $\beta$ R and peak delirium severity was adjusted for possible confounders: age, sex, and cognitive baseline. Based on a four-factor general linear model, 99 participants would be required to show a moderate effect size (Cohen F<sub>2</sub>=0.15) with 90% power at the 0.05 significance level. To ensure this sample size was reached and allow correspondence to the prior analyses of tau and NfL, we sent 114 subject samples for analysis, of which 105 were successfuly assayed for baseline amyloid data.

#### Statistical analysis

All analyses were performed in R (R Studio 2022.02.1 build 461, base R 4.1.3; https://posit.co/). All biomarker and TMTB data were normalised by log<sub>10</sub> transformation. In time course plots, biomarkers were further normalised to baseline. Changes in each biomarker were calculated by subtracting the baseline value from the value on the postoperative day of interest. For all analyses, outliers were identified using Cook's distance with a threshold of >4 $\mu$  to justify exclusion. The Shapiro–Wilk test, skewness, and visual inspection of histograms and boxplots were used to assess for normality. For unpaired, nonnormal data, Spearman correlation methods were used to compare continuous outcomes, and Mann-Whitney/Wilcoxon rank sum tests with continuity correction were used for dichotomous delirium incidence. Fisher's exact test was used to compare proportions. Preoperative to postoperative (paired) change in amyloid was assessed using a Wilcoxon signed rank test. Linear regression with peak DRS-R-98 as the dependent variable used a Poisson distribution family. This was the most suitable model given DRS-R-98 is a count variable with significant rightward skew. We performed a sensitivity analysis using a Gaussian distribution, given this was the method used for our power analysis. Predictors for multivariable models were selected based on previous evidence of their influence on delirium severity or plasma A $\beta$  concentration. Predictors were added using forced entry methods. Bayesian information criteria (BIC) were used to assess model fit. A P-value of <0.05 was used as the threshold for statistical significance. No adjustments were made for multiple comparisons.

#### Results

A total of 101 participants in the IPOD-B3 study had documented delirium assessments and baseline plasma amyloid levels. One patient experienced acute alcohol withdrawal and was excluded from our analysis, leaving 100 subjects (Supplementary Fig 1). The baseline characteristics of participants in this study are provided in Table 1; 35 of the 100

Table 1 Baseline patient characteristics. AUC, area under the curve; CI, confidence interval; COWAT, Controlled Oral Word Association Test; IQR, inter-quartile range; MoCA, Montreal Cognitive Assessment; NSQIP-D, National Surgical Quality Improvement Program-risk of death; NSQIP-SC, National Surgical Quality Improvement Program-serious complications; TMTB, Trail Making Test B. \*Median (IQR); *n* (%). <sup>†</sup>Standardised mean difference.

Characteristic	Overall, N=100*	Delirium		Difference <sup>†</sup>	95% CI <sup>†</sup>
		Yes, N=35*	No, N=65*		
Age (yr) Sex, n (%)	71(67—75)	70 (66–74)	71 (67–76)	-0.40 0.20	-0.81 to 0.02 -0.21 to 0.61
Female	45 (45)	18 (51)	27 (42)		
Male	55 (55)	17 (49)	38 (58)		
NSQIP-D	1.0 (0.2–3.7)	3.4 (1.6–5.4)	0.6 (0.2–1.9)	0.66	0.24-1.1
NSQIP-SC	14 (7–27)	27 (19–36)	10 (7-16)	1.3	0.85-1.7
Operating time (min)	318 (206–446)	453 (400-579)	259 (180–350)	1.4	0.94-1.8
Blood loss (ml)	500 (150-2350)	3000 (1000-5675)	320 (100-700)	1.1	0.62-1.5
Log <sub>10</sub> blood pressure AUC 10%	4.88 (4.54-5.27)	4.92 (4.46-5.39)	4.87 (4.56-5.11)	0.12	-0.29 to 0.53
Surgery type, n (%)	. ,	. ,	. ,	0.76	0.34-1.2
General	10 (10)	4 (11)	6 (9.2)		
Orthopaedic	36 (36)	6 (17)	30 (46)		
Urological	9 (9.0)	2 (5.7)	7 (11)		
Vascular	45 (45)	23 (66)	22 (34)		
Baseline TMTB (s)	84 (60–121)	96 (72–150)	81 (57–114)	0.41	-0.05 to 0.87
Baseline MoCA	24 (23–26)	24 (23–26)	24 (23–26)	-0.18	-0.65 to 0.29
Baseline COWAT	32 (24–39)	27 (18–32)	32 (27-42)	-0.81	-1.3 to -0.34
Length of hospital stay (days)	4 (2–9)	9 (7–14)	3 (2–5)	0.79	0.36-1.2



Fig 1. (a, c, e) Time plots of Aβ40, Aβ42, and Aβ ratio across postoperative Days 1–4 and (b, d, f) change from baseline on postoperative Day 1. Wilcoxon signed rank test was used to determine the postoperative Day 1 change in each biomarker (paired data). Based on Cook's distances, two, three, and three outliers were excluded for Aβ40, Aβ42, and AβR, respectively. Red triangles indicate participants diagnosed with postoperative delirium, and blue circles represent those without delirium. The red and blue lines represent smoothed LOESS regression curves for participants with and without delirium, respectively. Aβ, amyloid beta; AβR, amyloid beta ratio; DRS, Delirium Rating Scale; LOESS, locally estimated scatterplot smoothing; POD, postoperative day.

subjects (35%) were diagnosed with postoperative delirium. Consistent with our prior report on this cohort,<sup>12</sup> greater intraoperative blood loss and longer operation times were observed in subjects with delirium. Higher National Surgical Quality Improvement Program-risk of death (NSQIP-D) and -serious complications (NSQIP-SC) scores were also noted in the delirium group. In the delirium group, the mean baseline COWAT scores were lower, and the mean baseline TMTB scores were qualitatively higher than in the no-delirium group. The median length of stay was longer in patients with delirium, and the distribution of length of stay showed a significant rightward skew (Supplementary Fig 2).

#### Time course of plasma $A\beta$ levels

Time course plots of plasma A $\beta$ 40, A $\beta$ 42, and A $\beta$ R concentrations (normalised to baseline) over the first 4 postoperative days are shown in Fig 1 with smoothed LOESS (locally estimated scatterplot smoothing) regression curves for subjects with and without delirium. Although there was no significant change in plasma Aβ40 (Wilcoxon P=0.15; Fig 1b) or Aβ42 (Wilcoxon P=0.45; Fig 1d) from baseline to POD1, there was an increase in plasma  $A\beta R$  across this time period (Wilcoxon P<0.001; Fig 1f). This finding remained significant when excluding the 14 participants who overlapped from our previous study (Supplementary Fig 3). Analyses using peak postoperative change in  $A\beta$  also showed significant increases in A $\beta$ 40, A $\beta$ 42, and A $\beta$ R from baseline (Supplementary Fig 4). Comparison of the plasma CSF time course of  $A\beta$  shows strong similarity in trajectory between the two compartments (Supplementary Fig 5), consistent with our previous correlations.<sup>13</sup> A generalised linear model was constructed to investigate this postoperative rise in  $A\beta R$ , which suggested that age, sex, NSQIP-D, and area under the curve (AUC) of sevoflurane dose did not explain the POD1 rise in A $\beta$ R (Table 2).

### Primary outcome: association of $A\beta R$ with postoperative peak delirium severity

There was no statistically significant correlation between peak DRS-R-98 and POD1 change in Aβ40 (Spearman's  $\rho$ =-0.067; P=0.54; Fig 2a), Aβ42 ( $\rho$ =0.0085; P=0.94; Fig 2c), or the primary outcome, AβR ( $\rho$ =0.13; P=0.23; Fig 2e). Similarly, there was no difference in patients with and without delirium in terms of POD1 change in Aβ40 (Wilcoxon P=0.54; Fig 2b), Aβ42 (Wilcoxon P=0.78; Fig 2d), or AβR (Wilcoxon P=0.22; Fig 2f).

Poisson regression of peak DRS-R-98 with POD1 plasma A $\beta$  change after adjusting for age, sex, and baseline TMTB is described in Table 3. A decrease in plasma A $\beta$ 40 on POD1 was associated with more severe delirium in unadjusted (log[incidence rate ratio {IRR}]=-0.31; P<0.001) and adjusted (log [IRR]=-0.22; P=0.023) models. Conversely, POD1 change in plasma A $\beta$ 42 was not associated with peak delirium severity in unadjusted (log[IRR]=-0.19; P=0.29) or adjusted (log[IRR]=-0.13; P=0.49) models. An increase in POD1 plasma A $\beta$ R was associated with higher delirium severity in unadjusted (log

[IRR]=0.57; P<0.001) but not adjusted (log[IRR]=0.43; P=0.14) models. In the sensitivity analysis using a Gaussian distribution, POD1 change in plasma A $\beta$ R was not associated with peak DRS-R-98 in unadjusted or adjusted models (Supplementary Table 1).

## Secondary analysis: association of peak change and baseline amyloid with delirium severity

Analyses using peak change in plasma A $\beta$  showed a statistically significant association of change in plasma A $\beta$ 40 with peak DRS-R-98 ( $\rho$ =0.26; P=0.013), and the change in A $\beta$ 40 was higher in delirious patients (Wilcoxon P=0.0024). There were no statistically significant associations between peak plasma A $\beta$ 42 or A $\beta$ R and delirium incidence or severity (Supplementary Fig 6). Peak DRS-R-98 was not correlated with preoperative plasma A $\beta$ 40, A $\beta$ 42, or A $\beta$ R, and preoperative concentrations of these biomarkers were not different in subjects with postoperative delirium compared with those without delirium (Supplementary Fig 7).

Poisson regression showed that an increase in peak plasma A $\beta$ 40 and A $\beta$ 42 explained peak delirium severity in adjusted and unadjusted models. A $\beta$ R was significant in the unadjusted model but not in the model adjusted for age, sex, and baseline TMTB (Supplementary Table 2).

### Secondary outcome: association of perioperative variables with amyloid

There was no relationship between AUC low arterial pressure (<10% of baseline) and plasma A $\beta$ 40, A $\beta$ 42, and A $\beta$ R on either POD1 or peak change analyses (Supplementary Fig 8). Similarly, there was no correlation between operation time and peak or POD1 change in plasma A $\beta$ 40, A $\beta$ 42, and A $\beta$ R, with the exception of peak change in A $\beta$ 40, which was positively with operation time correlated (ρ=0.26; P=0.012: Supplementary Fig 9). There was no significant correlation between POD1 change in plasma  $A\beta R$  and POD1 changes in plasma inflammatory markers (interleukin [IL]-1β, IL-1Ra, IL-2, IL-4, IL-6, IL-8, IL-10, and IL-12) (Supplementary Fig 10). Plasma A $\beta$ 40 was negatively correlated with POD1 change in plasma IL-2 and IL-4 (Supplementary Fig 11), whereas POD1 change in Aβ42 was negatively correlated with POD1 change in plasma IL-1β, IL-2, IL-4, IL-6, and IL-12 (Supplementary Fig 12).

A total of 70 participants received sevoflurane for maintenance for anaesthesia, whereas the remaining 30 did not receive a volatile agent. There was no difference in POD1 change in plasma A $\beta$ 40, A $\beta$ 42, and A $\beta$ R in those with an AMS

Table 2 Linear regression predicting change in AβR from preoperative to POD1. Two outliers were excluded based on Cook's distances for peak DRS-R-98 ~ POD1 change in plasma AβR. AβR, amyloid beta ratio; AIC, Akaike's information criteria; AUC, area under the curve; BIC, Bayesian information criteria; CI, confidence interval; df, degrees of freedom; DRS-R-98, Delirium Rating Scale-Revised-98; NSQIP-D, National Surgical Quality Improvement Program-risk of death; POD, postoperative day; sE, standard error. Number of observations=65; log-likelihood=36.8; AIC=-61.5; BIC=-48.5; residual df=60. \*P<0.05; \*\*P<0.01; \*\*\*P<0.01.

Characteristic	Beta	SE	95% CI	P-value
(Intercept)	0.1003	0.266	-0.4220 to 0.6227	0.71
Age	-0.0008	0.004	-0.0078 to 0.0061	0.81
Sex: male	0.0383	0.036	-0.0332 to 0.1097	0.30
NSQIP-D	0.0063	0.007	-0.0071 to 0.0196	0.36
Log <sub>10</sub> AUC sevoflurane	-0.0077	0.028	-0.0626 to 0.0471	0.78



Fig 2. Peak DRS-R-98 to POD1 change in plasma amyloid univariate (a, c, e) correlation plots and (b, d, f) delirium boxplots. Based on Cook's distances for peak DRS-R-98 ~ amyloid univariate regression, six, seven, and three outliers were excluded for Aβ40, Aβ42, and AβR, respectively. Spearman methods were used for univariate correlation. Red triangles indicate participants diagnosed with postoperative delirium, and blue circles represent those without delirium. Black lines represent linear regression smoothing for all participants. Aβ, amyloid beta; AβR, amyloid beta ratio; DRS, Delirium Rating Scale-Revised-98; POD, postoperative day.

concentration of 0.2–1.0 MAC compared with those with AMS >1.0 MAC (Supplementary Fig 13). Similarly, excluding those with AMS <0.2, there was no correlation between AUC of sevoflurane and delirium incidence or severity (Supplementary Fig 14).

The POD1 change in both plasma A $\beta$ 40 and A $\beta$ 42 was positively correlated with the POD1 change in plasma tau, whereas there was no correlation between A $\beta$ R and tau. The POD1 change in plasma NfL was not correlated with the POD1 change in any plasma amyloid biomarker (Supplementary Fig 15).

Characteristic	Aβ4	0				Αβ4:	2				AβR				
	z	Log(IRR)	se	95% CI	P-value	z	Log(IRR)	se	95% CI	P-value	z	Log(IRR)	se	95% CI	P-value
Unadjusted analysis (Intercept)	8	2.4	0.033	2.4–2.5	<0.001***	28	2.4	0.033	2.3–2.5	<0.001***	88	2.4	0.033	2. <u>4</u> —2.5	<0.001***
POD1 change in	85	-0.31	0.092	-0.48 to -0.13	<0.001***	84	-0.19	0.18	-0.55 to 0.16	0.29	88	0.57	0.152	0.26-0.86	<0.001***
plasma amyloid Adjusted analysis															
(Intercept)	77	2.97	0.559	1.88 - 4.07	<0.001***	77	3.14	0.56	2.049 - 4.24	$< 0.001^{***}$	79	3.36	0.540	2.30 - 4.42	<0.001***
POD1 change in plasma amyloid	77	-0.22	0.100	-0.42 to -0.029	0.023*	77	-0.13	0.20	-0.523 to 0.24	0.49	79	0.43	0.294	-0.13 to 1.02	0.14
Age	77	-0.010	0.008	-0.026 to 0.0057	0.21	77	-0.013	0.008	-0.029 to 0.0028	0.11	79	-0.016	0.008	-0.031 to -0.0007	0.042*
Sex: male	77	-0.18	0.070	-0.32 to -0.042	0.010*	77	-0.15	0.070	-0.287 to -0.012	0.034*	79	-0.11	0.068	-0.25 to 0.020	0.094
TMTB baseline	77	0.0023	0.001	0.0009-0.004	<0.001***	77	0.0025	0.001	0.001 - 0.004	<0.001***	79	0.0026	0.001	0.0013-0.0039	<0.001***

We explored the short-term relationship between surgery, postoperative delirium, and plasma  $A\beta$  levels. As in our previous study,<sup>13</sup> plasma  $A\beta R$  increased from baseline to POD1. However, after adjusting for covariates, this increase did not appear to be associated with delirium incidence or peak severity. Postoperative plasma amyloid concentrations do not appear to correlate sufficiently with delirium to implicate amyloid changes in the link between delirium and accelerated cognitive decline that has been reported in other studies. Hence, despite no long-term cognitive outcomes being reported, our results do not support a pathophysiological chain linking postoperative delirium with long-term postoperative cognitive decline involving the amyloid hypothesis of AD.

The absence of a significant change in plasma A $\beta$ 42 from before surgery to postoperative Day 1 supports other studies without delirium analyses showing no change in CSF Aβ42 postoperatively.<sup>37–39</sup> One study showed no change in CSF A $\beta$ 42 from baseline to up to 24 h postoperatively in 39 participants,<sup>31</sup> which was supported by another study with 103 participants showing no change in CSF A $\beta$ 42 out to 6 weeks postoperatively.<sup>38</sup> Furthermore, a PET study of 40 patients undergoing cardiac surgery showed no association between cortical amyloid burden and cognitive function at 6 weeks postoperatively, no difference in cortical amyloid burden between patients undergoing surgery and controls not undergoing surgery at 6 weeks postoperatively, and no association between cortical amyloid burden at 6 weeks or 1 yr postoperatively and cognitive function at 1 or 3 yr postoperatively.<sup>40</sup> Another PET study in 313 participants showed no differences in the odds of elevated brain amyloid between participants with previous surgical hospitalisation compared with those without.<sup>41</sup> Although no direct inferences about delirium can be made, the normal distribution of data in the aforementioned studies provides evidence against a subpopulation (including those with postoperative delirium) that might have larger fluctuations in CSF Aβ42 in the perioperative period. However, these studies do not report the changes in CSF A $\beta$ R with surgery, which have been shown to be more strongly associated with CSF tau and cognitive decline than A $\beta$ 42<sup>42</sup>; it was for this reason that we focused on the A $\beta$ R. Furthermore, our prior analyses of  $\mbox{PET}^{18}$  and  $\mbox{CSF}^{13}$  focused on continuous relationships with delirium severity that are not dependent on subgroups, so we believe that further investigations are required to elucidate links of amyloid disease, delirium, and longer-term cognitive changes.

Our observation that the plasma ABR increases after surgery suggests a small increase in A $\beta$ 42 and decrease in A $\beta$ 40 that were not large enough on their own to reach statistical significance. It is unclear for how long the short-term increase in plasma A $\beta$ R that we observed persists and if there is any impact on long-term trends in plasma ABR, brain amyloid deposition, or cognitive function. Longitudinal studies have shown that plasma  $A\beta R$  decreases proportionally with age and increasing cerebral amyloid deposition,  $^{\rm 22,43}$  and this decrease is associated with greater cognitive decline in  $\mathrm{many}^{44,45}$  but not all<sup>46</sup> studies. It is possible that the increase in plasma  $A\beta R$ relates to worsening AD pathology that is independent of any effect on postoperative delirium, although the links drawn here would be speculative. Moreover, if general anaesthesia were increasing accumulation of the more pathogenic  $A\beta 42$ species in the brain, we would expect a decrease, not an increase, in the plasma  $A\beta R$  after surgery.

The observed negative correlation between plasma A $\beta$ 40 and delirium severity could represent greater CNS A $\beta$ 40 deposition in patients with more severe delirium. However, this is unlikely to carry pathological significance. Despite the concentration of A $\beta$ 40 in the CNS being several-fold higher than of A $\beta$ 42,<sup>47</sup> the latter forms the major (and sometimes only) component of brain amyloid plaques,<sup>48</sup> and even small increases in the A $\beta$ R induce greater neurotoxicity.<sup>49</sup> Perhaps a more parsimonious explanation is that plasma A $\beta$ 40 has shown an inverse correlation with IL-1 $\beta$  in other studies<sup>50</sup> (although it did not reach statistical significance in our cohort [P=0.079]), which suggests the association of A $\beta$ 40 with delirium severity merely reflects its inverse relationship with inflammation, a putative driver of delirium severity.

The increase in CSF and plasma  $A\beta R$  observed in our study could represent increased amyloid clearance from the CNS by microglia; expression of IL-1 $\beta$  is increased in surgery, and stimulation by IL-1 $\beta$  has been shown to promote microglial clearance of  $A\beta$ .<sup>51</sup> Although this hypothesis and other provocative neuropathological explanations are attractive, they are difficult to support. First, we observed no correlation between plasma inflammatory markers and POD1 change in plasma A $\beta$ R. Second, plasma A $\beta$  is a less reliable indicator of AD compared with CSF  $A\beta$ , <sup>52,53</sup> as  $A\beta$  is also produced by non-CNS organs (e.g. liver and kidney),<sup>54</sup> in which homeostasis is often perturbed by surgery, and other unknown factors contributing to the production, clearance, and equilibration of A $\beta$  might also play a role. The relationship between CSF and plasma A $\beta$  is much more poorly understood in the perioperative setting compared with non-perioperative settings. As such, inferences regarding CNS amyloid based on plasma amyloid measurements should be drawn with caution until more is known about the relationship between amyloid concentrations in the two compartments. Changes in plasma  $A\beta$ have also been associated with white matter hyperintensities, lacunar infarcts, and hypoxic brain injury after cardiac arrest,  $^{55-57}$  meaning perioperative plasma  $A\beta$  could be influenced by ischaemia. Together, the aforementioned observations suggest plasma  $A\beta$  could fluctuate in the perioperative period independent of any change in cerebral amyloid burden that is of pathophysiological significance to delirium and AD. Third, preoperative  $A\beta$  was not correlated with delirium incidence or severity in our cohort; however, our study was likely underpowered for this secondary outcome. Finally, we are not aware of a rationale for the relative increase in plasma  $A\beta 42$ over Aβ40 in the perioperative period. Although differential cleavage of the amyloid precursor protein in central or peripheral tissues could be induced by anaesthesia or surgery, or Aβ42 could be preferentially cleared from plaques in the brain, preclinical studies have shown no obvious signal for this.<sup>9,10,58</sup>

We did not find evidence to support  $A\beta$  as an explanatory link for any causative relationship between short-term and long-term postoperative cognitive decline. These findings are supported by a recent cross-sectional study that observed an association between past surgery and cortical thinning on PET imaging in areas typically implicated in AD but no association between past surgery and cerebral amyloid deposition.<sup>59</sup> Considering the reported association between surgery and accelerated cognitive decline in some<sup>59,60</sup> but not all<sup>61</sup> highquality studies, current evidence suggests a harmful effect of surgery on neurodegeneration via a mechanism that is independent of cerebral  $A\beta$  deposition. This is supported by a study of CSF biomarkers showing no impact of surgery on CSF  $A\beta42$ at 6 weeks postoperatively and no association between changes in A $\beta$ 42 and postoperative cognitive function, both of which were also true for CSF tau and pTau<sub>181</sub>.<sup>38</sup> However, these findings stand in contrast to our own observations, where there were strong associations for CSF and plasma tau and pTau<sub>181</sub>.<sup>13</sup> Further research is required to reconcile these differences.

Support for the role of tau in the clinical progression of AD appears to be greater than that for  $A\beta$ ,<sup>62</sup> and we have recently associated plasma tau with postoperative delirium incidence and severity.<sup>14,63</sup> Furthermore, tau resolved in parallel with resolution of delirium symptoms. We observed a positive correlation of change in plasma tau with A $\beta$ 40 and A $\beta$ 42, and no correlation with A $\beta$ R. Given the POD1 change in plasma Aβ40 and Aβ42 was not related to delirium incidence, and there was no correlation between amyloid and NfL, the relationship between tau and amyloid likely arose from processes separate to those occurring in delirium. Aside from tau, the recently proposed probabilistic model of AD also stresses the increased importance of stochastic factors and decreased importance of the amyloid cascade in the pathogenesis of apolipoprotein E (APOE) ε4-unrelated AD (62% of AD cases<sup>64</sup>).<sup>65</sup> This suggests that future studies could consider assessing a broader range of potential pathophysiological pathways. For example, neuroinflammation driven by the surgery-induced peripheral inflammatory response has also emerged as a potentially critical mediator of delirium and postoperative cognitive decline,<sup>66</sup> and inflammation was unrelated to amyloid in our data.

Our study had some limitations. We analysed only one component of the ATN framework, and we did not consider the many other factors implicated in AD pathogenesis. Moreover, the amyloid hypothesis, which was the focus of this study, implies a linear causal chain that begins with an amyloid trigger, an assumption that has come under significant scrutiny in recent years.<sup>67</sup> Another possible reason for our negative finding is missing data in our primary outcome (n=82 prior to outlier exclusion): not all participants with amyloid data underwent baseline cognitive testing. However, it is worth reflecting that we retained 83% power to detect our primary outcome. We also did not report long-term cognitive data; hence, we cannot be certain that delirium was associated with long-term cognitive decline in our cohort. Future adequately powered studies should investigate the link with long-term cognitive decline and  $A\beta R$ ; however, with advances in plasma phosphorylated tau assays,  $A\beta R$  could be superseded by a superior marker.<sup>13</sup> We did not control for APOE ɛ4 status, which is known to affect plasma A $\beta$  concentrations.<sup>68</sup> However, APOE status has been found to not be associated with delirium.<sup>69</sup> Our attrition rate for biomarker collection from POD1 onwards was an expected consequence of patient discharge from hospital (with a median length of stay of 4 days). We focused on POD1 in an effort to reduce bias, the pitfall being that we are limited in commenting on important fluctuations in A $\beta$  that could occur in subsequent days and weeks.

#### Conclusions

We observed an increase in the plasma A $\beta$ R from the preoperative to postoperative period; however, this increase was not associated with the incidence or severity of postoperative delirium. Perioperative fluctuations in the plasma A $\beta$ R therefore appear to be unrelated to severe perioperative changes in cognition. Our findings do not support associations of dynamic changes in amyloid with postoperative delirium.

#### Authors' contributions

Study design: RDS, RCL, RAP (in consultation with MP, CC, and DK).

Supply of assays: HZ, KB Management of biofluid analysis: HZ, KB Data analysis: TP Drafting of paper: TP (with input from JT). All authors provided critical feedback on the paper.

#### **Declarations of interest**

HZ has served at scientific advisory boards or as a consultant for AbbVie, Acumen, Alector, ALZpath, Annexon, Apellis, Artery Therapeutics, AZTherapies, CogRx, Denali, Eisai, Nervgen, Novo Nordisk, Passage Bio, Pinteon Therapeutics, Red Abbey Labs, reMYND, Roche, Samumed, Siemens Healthineers, Triplet Therapeutics, and Wave; has given lectures in symposia sponsored by Cellectricon, Fujirebio, AlzeCure, Biogen, and Roche; and is a co-founder of Brain Biomarker Solutions in Gothenburg AB (BBS), which is a part of the GU Ventures Incubator Program. KB has served as a consultant at advisory boards or at data monitoring committees for Abcam, Axon, BioArctic, Biogen, JOMDD/Shimadzu, Julius Clinical, Lilly, MagQu, Novartis, Ono Pharma, Pharmatrophix, Prothena, Roche Diagnostics, and Siemens Healthineers, and is a co-founder of Brain Biomarker Solutions in Gothenburg AB (BBS), which is a part of the GU Ventures Incubator Program, outside the work presented in this paper. The other authors declare no competing interests that may be relevant to this work.

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#### Appendix A. Supplementary data

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