The association between perihaematomal oedema and functional outcome after spontaneous intracerebral haemorrhage: A systematic review and meta-analysis

European Stroke Journal 2023, Vol. 8(2) 423–433 © European Stroke Organisation 2023 \bullet $\left($

DOI: 10.1177/23969873231157884 Article reuse guidelines: [sagepub.com/journals-permissions](https://uk.sagepub.com/en-gb/journals-permissions) journals.sagepub.com/home/eso

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Abstract

Purpose: Perihaematomal oedema (PHO) formation has gained increasing interest as a therapeutic target after spontaneous intracerebral haemorrhage (ICH). Whether PHO contributes to poor outcome is unclear. We aimed to determine the association between PHO and outcome in patients with spontaneous ICH.

Method: We searched five databases up to 17 November 2021 for studies of ≥10 adults with ICH reporting the presence of PHO and outcome. We assessed risk of bias, extracted aggregate data and used random effects meta-analysis to pool studies that reported odds ratios (OR) with 95% confidence intervals (CI). Primary outcome was poor functional outcome defined as modified Rankin Scale score of 3–6 at 3months. Additionally, we assessed PHO growth and poor outcome at any time of follow-up. We prospectively registered the protocol in PROSPERO (CRD42020157088).

Findings: We identified 12,968 articles, of which we included 27 studies (*n*=9534). Eighteen studies reported an association between larger PHO volume and poor outcome, six a neutral result and three an inverse relationship. Larger absolute PHO volume was associated with poor functional outcome at 3months (OR per mL increase of absolute PHO 1.03, 95% CI 1.00–1.06, ^{β} 44%, four studies). Additionally, PHO growth was associated with poor outcome (OR 1.04, 95% CI 1.02–1.06, *I* 2 0%, seven studies).

Discussion: In patients with spontaneous ICH, larger PHO volume is associated with poor functional outcome at 3months. These findings support the development and investigation of new therapeutic interventions targeting PHO formation to evaluate if reduction of PHO improves outcome after ICH.

Keywords

Intracerebral haemorrhage, perihaematomal oedema, systematic review, meta-analysis

Date received: 4 January 2023; accepted: 27 January 2023

Introduction

Spontaneous intracerebral haemorrhage (ICH) is the second most common subtype of stroke, affecting more than 3 million people worldwide each year.¹ One-month case fatality rate is approximately 40%.² Apart from the effect of stroke unit care and early control of elevated blood pressure that may be beneficial, there are no treatments with proven benefit.3,4

The mechanisms leading to brain injury in ICH are complex and can be divided into two main categories.⁵ Primary brain injury results from the immediate disruptive mass effect caused by the haematoma and occurs within the first hours after ICH.⁵ Subsequently, the local tissue damage stimulates the release of inflammatory

factors, blood-brain-barrier breakdown, the activation of microglia and influx of circulating inflammatory cells.⁶ These secondary processes result in secondary brain

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injury and the development of perihaematomal oedema (PHO). Development of PHO may be detrimental via enhancement of the harmful mass effect but toxic dysregulation of the local osmotic gradient has also been suggested.7 PHO is considered a quantifiable radiological marker of secondary brain injury⁸ and has been used as an outcome measure in previous clinical studies targeting secondary brain injury after ICH.^{9,10} However, whether PHO affects clinical outcome after spontaneous ICH remains controversial as previous reports have shown conflicting results.^{8,11,12}

We aimed to systematically review the literature and meta-analyse studies that investigated the association between PHO and outcome in adults with spontaneous ICH.

Methods

Search strategy and study selection

We searched MEDLINE, Embase, Cochrane library, clinicaltrials.gov and International Standard Randomised Controlled Trial Number Register (ISRCTN) up to 17 November 2021 for published prospective or retrospective observational cohort studies, case control studies and randomised controlled trials in human adults that investigated the association between PHO and outcome ≥ 30 days after symptom onset. We performed an electronic search strategy consisting of different combinations of the terms for ICH AND (perihaematomal) oedema (Supplemental Material). There were no restrictions on language or publication date. Studies had to comprise at least 10 patients to be included. PHO could be reported as absolute PHO volume (aPHO), relative PHO volume

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(rPHO = \frac{aPHO}{ICH volume})
$$
 or oedema extension distance

$$
(\text{OED} = \sqrt[3]{\frac{PHO \text{ volume} + ICH \text{ volume}}{\frac{4}{3}\pi}} - \sqrt[3]{\frac{ICH \text{ volume}}{\frac{4}{3}\pi}})^{13}
$$

Imaging modalities to measure PHO could be computed tomography (CT) or magnetic resonance imaging (MRI). The following studies were excluded: conference abstracts, studies regarding ICH secondary to an underlying macrovascular cause identified by brain imaging, studies including solely children (<18 years) and studies that reported in-hospital outcome measures only, because outcome of patients with ICH can improve after (longer) time.¹⁴ The study protocol was prospectively registered with the International Prospective Register of Systematic Reviews (CRD42020157088).

All records of potentially eligible studies were imported into Covidence (covidence.org). Two of four authors (MC, LS, FS, NS) independently screened all abstracts and two authors (MC, LS) assessed full texts to identify studies that met the predefined inclusion criteria. Disagreements were resolved by a third author (FS or NS). When two studies used overlapping cohorts, the study with the largest number of patients that best matched the inclusion criteria was included in our primary analysis. This process was repeated for the secondary analyses.

Data extraction

Two authors (LS and MC) independently performed the methodologic quality assessment of the included studies using the Newcastle-Ottawa Quality scale (NOS) for cohort studies and for case control studies, with 0 points reflecting the highest risk of bias and 9 reflecting the lowest risk of bias. Studies scoring 0–3 points were considered of poor quality, with 4–6 points reflecting fair quality and 7–9 points good quality. Using a prespecified structured data extraction form, two authors (LS and MC) extracted the following data from all included studies: first author, year of publication, in- and exclusion criteria, baseline characteristics of the included subjects (age, sex, blood pressure on admission, medical history, medication use), ICH imaging characteristics (imaging modality, location, ICH volume, time since symptom onset), PHO parameters (aPHO, rPHO, OED, PHO growth, modality and timing of imaging), the number of patients with a certain functional outcome, the used definition of functional outcome and the timing of follow-up. We extracted this information for patients with good and poor outcome separately, to exclude patients without follow up. Discrepancies were resolved by discussion and, if necessary, by a third reviewer (FS) in a consensus meeting.

Outcomes

The primary outcome was poor functional outcome at 3-month follow-up, defined as a dichotomised modified Rankin Scale (mRS) score of 3–6. Secondary outcomes were a mRS score 3–6 at any time of follow-up, a mRS score 4–6 at 3months and death at any time of follow-up. For the primary and secondary analyses, we combined all studies that measured PHO by means of aPHO, rPHO and OED at any timepoint.

Data analysis

We reported the effects of PHO on outcome in all included studies descriptively. We pooled reported odds ratio's (ORs) with 95% confidence intervals (CI) for primary and secondary outcomes in a generic inverse-variance based random-effects method meta-analysis. When both unadjusted and adjusted ORs were available we included only the adjusted OR in the meta-analysis. Additionally, the association between PHO growth and any kind of reported outcome at any time of follow-up was analysed.

To maximise the number of studies contributing to our analyses, we also calculated standardised mean differences

Figure 1. Study selection flow chart.

An overview of the 27 excluded studies in the last step is provided in the Supplemental Material.

(SMD) between patients with a poor and patients with a good outcome for studies reporting crude aPHO measures for each outcome group, to determine whether results would support the results of the pooled OR for the primary outcome. When studies provided median and interquartile values (IQR), we approximated the sample mean and standard deviation (SD) following a standard method.15 Metaanalysis of SMDs was performed using a random effects model.

We assessed heterogeneity with the I-squared statistic (I^2) and categorised heterogeneity as follows: 0% –40% heterogeneity that might not be important; 30%–60%, moderate heterogeneity; 50%–90%, substantial heterogeneity; and 75%–100%, considerable heterogeneity.

We aimed to investigate the following prespecified potential modifying factors by means of a meta-regression analysis if at least 10 studies in a meta-analysis had these data available: age, (systolic) blood pressure, the use of antithrombotic/antiplatelet agents or the use of statins. We constructed funnel plots to assess potential publication bias.

We used R and R-studio version 3.6.2 with packages 'rmeta' and 'metafor' for all statistical analysis.

Results

We identified 12,968 references of which 309 studies were assessed for eligibility. After full text screening of 309 studies, we extracted 54 studies of which we included a total of 27 studies in our analyses with a total of 9534 patients (Figure 1; Supplemental Table 1).16–42 Using the NOS scale, 11 studies were classified as of fair quality and 16 studies were deemed of high quality. Median risk of bias score was seven (IQR 6–8; Supplemental Table 2). The most common reason for possible bias was the lack of information on premorbid functional status of the included patients. The main study characteristics are summarised in Table 1.

Measures of PHO differed between studies (Supplemental Table 3). Twenty-two studies measured aPHO, ^{16-19,21,22,24,26-40} with the study mean aPHO ranging from 0.94 to 42.6mL. Six studies reported rPHO,^{19,20,22,30,33,37} three studies evaluated OED,22,28,40 eight studies assessed PHO gro wth.21,23,25,29,30,36,41,42 Ten of the 27 included studies reported on multiple measures of PHO.19,21,22,28–30,33,36,37,40 Seventeen studies had an average interval between symptom onset and imaging of 24h or less, $16-19,21,25-34,37,41$ eight studies

Table 1. Included studies.

(Continued)

Figure 2. Estimates of the association between absolute perihaematomal oedema volume and poor clinical outcome (mRS score 3–6) at 3months follow-up.

N: total participants; *n:* participants with poor outcome; OR: odds ratio; 95% CI: 95% confidence interval.

All studies reported an OR adjusted for at least age, ICH volume and clinical condition on admission described as GCS or NIHSS. Murthy et al. 30 adjusted for an additional five factors, Urday et al.³⁸ for two additional factors, Iglesias-Rey et al.²⁴ for nine and Huan et al.²² for three additional factors (Supplemental Table 4).

performed imaging between 25 and 72h post onset,20,22,23,35,36,38,40,42 one study had an interval longer than $72h^{24}$ and one study assessed peak PHO volume between 1 and 12 days after onset.³⁹ Twenty-six of 27 studies used CT as their primary imaging modality, the other assessed PHO with MRI.³⁷ Outcome measures and follow-up duration were various among studies (Table 1). A total of 20 studies assessed outcome at 3months after ICH, with poor outcome defined as mRS score 3–6 in 11 of these studies,17–20,22–24,30,34,38,41 mRS score $4-6$ in seven studies^{25,27,29,31,36,39,42} two reporting death as primary outcome.16,26 Six studies assessed clinical outcome at a different intervals, varying between 1 and 12months post ictus, two of which investigated mRS score $3-6^{28,37}$ four that assessed death.^{21,32,33,40} Lastly, one study reported clinical outcome only as 'worse mRS'.35

Primary outcome. In 18 studies (*n*=7711 patie nts)17,18,21–24,27,29–32,34–36,39–42 a larger volume of PHO (aPHO, rPHO, OED and/or growth) was associated with a higher risk of any type of poor outcome (mRS 3–6, mRS 4–6 and death) at any time, while six studies (*n*=1633 patie nts)16,25,28,33,37,38 reported a neutral result and three studies $(n=190 \text{ patients})^{19,20,26}$ found an association between larger PHO volume and a lower risk of poor outcome. The pooled OR of the four studies reporting the association between PHO measured at any timepoint (all aPHO), and mRS score 3–6 at 3months was 1.03 (95% CI 1.00–1.06, *p*=0.036, Figure 2), indicating a 3% increase in the odds of poor functional outcome increases for each mL of aPHO.³⁸ In all four studies included in this pooled effect size the OR had been adjusted for multiple factors (Supplemental Table 4), including at least age, ICH volume and severity of ICH by either GCS score or NIHSS score.^{22,24,30,38} Heterogeneity among these four studies was moderate $(I² 44%)$. Quality assessment showed an intermediate^{22,30} to $low^{24,38}$ risk of bias in these studies. There was no evidence for publication bias (Supplemental Figure 1).

Secondary outcomes. Two studies^{28,37} reported on the association between PHO and mRS score 3 and 6 at either 637 or 12 months after ICH.²⁸ PHO was measured as either rPHO³⁷ or aPHO.28 Notably, the study investigating the mRS at 12months, presented an OR adjusted for ICH volume, age, ICH location, IVH and GCS score,²⁸ but the study investigating the mRS at 6months only provided an unadjusted OR (Supplemental Table 4). 37 Combining these two studies^{28,37} with the studies included the primary analysis,22,24,30,38 we found a pooled OR of 1.02 (95% CI 1.00–1.04, $p=0.1$, $I^2 = 76\%$ for the association between PHO and mRS 3 and 6 at any time of follow-up.

Seven studies $(n=2793 \text{ patients})^{25,27,29,31,36,39,42}$ reported on the influence of PHO and mRS score 4–6 at 3months. Three of these studies $(n=1398)^{25,36,42}$ assessed PHO growth as their primary analysis, the other four $(n=1395)^{27,29,31,39}$ measured aPHO. One study $(n=59$ patients)²⁷ did not provide an OR but reported that larger PHO on day 3 after admission was associated with mRS 4–6. Of the six studies presenting an OR, four studies presented an OR adjusted for at least ICH volume, $25,36,39,42$ one study³¹ presented an OR that was adjusted for several factors but not for ICH volume and one study²⁹ presented an unadjusted OR (Supplemental Table 4). In addition, one of

Author(s), Year	N	n	SMD, 95% CI	Weight	SMD [95% CI]
Delgado et al, 2006	78	48		12.82%	0.68 [0.21, 1.15]
Tsai et al, 2013	47	29		10.68%	0.27 [-0.32, 0.86]
Gupta et al, 2014	44	20			10.48% -0.52 [-1.12, 0.08]
Rodriguez-Luna et al, 2016	322	188	$\overline{}$	16.95%	0.66 [$0.41, 0.91$]
Iglesias-Rey et al, 2018	887	513	۰.	18.57%	0.72 [0.58, 0.86]
Chen et al, 2019	131	77	$\overline{}$	14.97%	0.61 [0.25, 0.96]
Huan et al, 2021	159	77	$\overline{}$	15.52%	0.91 [0.58 , 1.23]
Heterogeneity					100.00% 0.54 [0.25, 0.83]
$I^2 = 82$, $\tau^2 = 0.11$, $p = 0.0036$				Overall effect; $p = 0.00025$	
			-1.5 -0.5 0.5 1.5		
	Larger PHO in good outcome			Larger PHO in poor outcome	

Figure 3. Standardised mean difference in absolute perihaematomal oedema volume between poor clinical outcome (mRS score 3–6) and good outcome (mRS score 0–2) at 3months follow-up.

N: total participants; *n*: participants with poor outcome; SMD: standardised mean difference; 95% CI: 95% confidence interval.

these studies $(n=292 \text{ patients})^{39}$ reported that peak aPHO volume was associated with a decreased odds of an mRS 0–3 (aOR 0.984, 95% CI 0.973–0.994, *p*=0.002) after adjustment for age, ICH volume by location, IVH and NIHSS. Two studies $(n=1044 \text{ patients})^{29,31}$ reported an OR (adjusted in one study, 31 Supplemental Table 4), for the association between aPHO and mRS score 4 and 6 after 3months resulting in a pooled OR of 1.06 (95% CI 0.97– 1.17, $p=0.21$) with considerable heterogeneity (l^2 93%).

A total of 10 studies (*n*=2936 patients) reported on the association between PHO volume (aPHO, rPHO, PHO growth or OED) and death.^{16,19,21,26,30,32–34,36,40} Five studies provided only descriptive results.16,19,32,34,36 Three of these studies (*n*=874 patients) reported a significantly higher aPHO in patients that died within one^{32,36} or three³⁴ months after ICH while two studies $(n=69 \text{ patients})^{16,19}$ reported no statistically significant association between aPHO and death at 3months. The other five studies provided an adjusted OR and used aPHO as their primary metric with different intervals between ICH and mortality assessment: 1 month in two studies, $2^{1,33}$ 3 months in two studies^{26,30} and 6months in one study.40 All five reported ORs were adjusted for multiple factors but at least for ICH volume (Supplemental Table 4). One study reported an OR per 100cc increase in aPHO, which we have transformed to an OR per mL increase.²⁶ We found a pooled OR of 1.02 for death at any time of follow-up (95% CI 0.99–1.05, *p*=0.14, I^2 =96%, five studies, Supplemental Figure 2).

Meta-analysis of SMDs in the seven studies (*n*=1668 patients) reporting aPHO for mRS score 3–6 versus mRS score 0–2 at 3months revealed a significant difference in mean aPHO, with a higher aPHO in patients with a poor outcome (SMD 0.54, 95% CI 0.25–0.83, *p*=<0.001,

 I^2 =82%; Figure 3). Five of the seven studies carried an intermediate risk of bias^{18,20,22,34,37} while the other two were of high quality (Supplemental Table 3).^{17,24}

Seven studies (*n*=4473 patients) assessed the influence of PHO growth on poor functional outcome.^{21,23,25,29,36,41,42} Six out of the seven studies measured aPHO increase while one study assessed increase in OED.²³ There was large variation in the interval between symptom onset, baseline PHO measurement and the timing of the assessment of PHO growth. Four studies measured PHO growth within the first $24h$,^{21,25,29,41} one study between 24 and 48 h³⁶ and two studies between approximately 48 and 72 h after ICH onset.^{23,42} Timing of follow-up and definition of outcome varied between these studies. Studies defined poor outcome as mRS score $3-6^{41}$ or mRS score $4-6^{25,29,36,42}$ at 3 months, mRS score $3-6$ at 12 months²³ or as death at 1 month after ICH.²¹ All studies presented ORs adjusted for several factors but at least for ICH volume (Supplemental Table 4). Meta-analysis of the seven studies reporting PHO growth resulted in a pooled OR of 1.04 (95% CI 1.02–1.06, $p=0.0013$, $I^2=0\%$; Figure 4). Sub-analysis of the four studies measuring PHO growth within the first 24h revealed a pooled OR of 1.04 (95% CI 1.01–1.06, *p*=0.0019, *I* 2=0.13), while the pooled OR of the three studies measuring PHO growth between 24 and 72 h after symptom onset was 1.40 (95% CI 0.94–2.08, *p*=0.098, *I* 2=79%).

There was an insufficient number of studies in any of the meta-analyses to perform a meta-regression analysis.

Discussion

In this systematic review and meta-analysis, we found that in adults with spontaneous ICH, aPHO was associated with

Figure 4. Estimates of the association between perihaematomal oedema growth and any poor outcome at any time of follow-up, stratified by timing of the assessment of growth.

N: total participants; *n*: participants with poor outcome; OR: odds ratio; 95% CI: 95% confidence interval.

*Poor outcome defined as mRS score 3–6 at 3 months after ICH. ^Poor outcome defined as mRS score 4–6 at 3 months after ICH. ~Poor outcome defined as death at 1month after ICH. " Poor outcome defined as mRS score 3–6 at 12months after ICH.

poor outcome at 3months. In addition, PHO growth was associated with poor outcome defined as either mRS 3–6, mRS 4–6 or death.

Secondary brain injury has gained increasing interest as a potential therapeutic target over recent years. However, systematic assessment of the influence of PHO on functional outcome after ICH is hampered by the large variation in methods for PHO measurement and outcome assessment. This heterogeneity also limited a previous systematic review of PHO and outcome, which included 21 articles up to 2016.12 In their meta-analysis consisting of just two studies,30,38 the authors reported a significant association between aPHO measured at 72h after ICH onset and poor functional outcome, defined as mRS 3–6, at 90days (OR 1.02, 95% CI 1.00–1.03, $p=0.007$). In comparison to their work, we were able to include multiple new studies and chose to perform a meta-analysis of aPHO at all available timepoints and functional outcome, resulting in increased statistical power.

We found that for every extra mL of aPHO the odds of poor functional outcome after ICH increases with 3%. PHO develops quickly during the first day and increases over the first week after ICH with a reported peak between 7 and 11days.8,38 In our systematic review the majority of studies measured PHO within 24–72 h of ICH onset but there was insufficient data to perform meta-regression to assess the effect of time between symptom onset and PHO measurement on the association between PHO and outcome. Therefore, the effect of peak PHO volume on functional outcome after ICH remains unclear. Future studies of PHO formation could offer valuable insights into the effect of PHO on clinical outcome by measuring PHO at later timepoints as well.

Considering the progressive development of PHO over time, we also assessed the influence of PHO growth on functional outcome. In contrast to the previous meta-analysis,¹² we found that PHO growth appears to be associated with poor outcome (mRS 3–6, mRS 4–6 and death). This strengthens the hypothesis that PHO formation might be a valuable target to improve outcome after ICH.

Besides the variation in timing of PHO assessment, the studies in this review used a variety of PHO metrics. aPHO volume is traditionally the most frequently used PHO metric. However, it is known that both aPHO and rPHO are strongly correlated with ICH volume. Recently, the OED has been developed as a new PHO metric that is affected less by ICH volume and could reduce the required sample size in clinical trials by as much as 75% .¹³ Only four studies in this review assessed OED in their population, mostly as a secondary measure in addition to aPHO and with different clinical outcome measures. Data was insufficient for separate meta-analysis of OED as PHO metric and functional outcome.

Strengths of our systematic review and meta-analysis include the comprehensive literature search without restrictions in publication language. This resulted in a high number of identified studies and a total number of included patients twice as high as in a previously published systematic review.12 Moreover, we applied a generic inverse-variance based random effects model in the meta-analyses, minimising imprecision in the pooled OR estimate. In the analysis of the primary outcome, all ORs that were included were at least adjusted for age, ICH volume and clinical examination by either GCS or NIHSS. In addition, we performed an analysis based on calculated SMDs, in order to consider all the available data even in the absence of ORs. This analysis supported the results of the primary analysis, which strengthens the validity of our findings.

This study also has some limitations. First, quality assessment of the included studies by the NOS revealed a risk of bias almost half of the included studies. Second, meta-analysis was hampered by the variation in applied method and timing of measurements of PHO, and by the variation in the type and timing of outcome measures. Third, not all studies provided ORs and therefore not all studies could be included in the meta-analysis. Calculation of the SMDs partly overcame the small number of ORs reported, but an important limitation of this approach is that for some studies statistical approximation of the mean PHO volume and corresponding SD had to be applied, and we were not able to correct these SMDs for other characteristics that influence the outcome, such as ICH volume. Lastly, due to an insufficient number of studies, we were unable to perform meta-regression to examine potential modifying factors such as age, ICH location or admission blood pressure.

Based on our experience in this systematic review we want to emphasise the importance of uniformity in measurements and reporting of PHO, in the timing of PHO assessment, and in the definition of functional outcome, to allow for optimal comparison of treatment effects. Also, a planned individual patient data meta-analysis of the available literature might provide further insight in the relationship between PHO and functional outcome (PROSPERO CRD42021253263 UK).

Our findings support further research to develop treatment strategies aimed at preventing PHO formation. There is a window of opportunity in the early stage after hospital admission when PHO is developing. Treatments targeting secondary brain injury and PHO formation probably need to be administered during the first days after ICH as PHO continues to develop in this early timeframe. When PHO is assessed as outcome, the timing of its measurement is of great importance. Multiple randomised clinical trials aimed at ameliorating PHO formation are ongoing, investigating IL1-Ra anakinra (NCT04834388), atorvastatin (NCT04857632), fingolimod (NCT04088630) and sodium aescinate (NCT05263167).

Conclusion

Our data indicates an association between aPHO and PHO growth with poor outcome after ICH. These findings support the development and investigation of new therapeutic interventions targeting PHO formation to evaluate if reduction of PHO leads to improved outcome after ICH. Uniformity in PHO assessment in future studies is of great importance to compare effectiveness of potential new treatment strategies.

Acknowledgements

None.

Declaration of conflicting interests

The author(s) declared no potential conflicts of interest with respect to the research, authorship, and/or publication of this article.

Funding

The author(s) disclosed receipt of the following financial support for the research, authorship, and/or publication of this article: FS and MC are funded by a senior clinical scientist grant from the Dutch Heart Foundation (grant 2019T060). CK and LS are supported by the Netherlands Cardiovascular Research Initiative, which is supported by the Dutch Heart Foundation, CVON2015- 01: CONTRAST, and the support of the Brain Foundation Netherlands (HA2015.01.06). The collaboration project is additionally financed by the Ministry of Economic Affairs by means of the PPP Allowance made available by the Top Sector Life Sciences & Health to stimulate public-private partnerships (LSHM17016). This work was funded in part through unrestricted funding by Stryker, Medtronic and Cerenovus. Radboud UMC and Erasmus MC received additional unrestricted funding on behalf of CONTRAST, for the execution of the Dutch ICH Surgery Trial pilot study. For the Dutch ICH Surgery Trial, Radboudumc and Erasmus MC received funding from Penumbra Inc. and from ZonMw and 'Zorginstituut' (80-86200-98-25001). NS receives funding from NHS Research Scotland and the UK Stroke Association.

Ethical approval

Not applicable. We conformed to the ICMJE Recommendations for the conduct, reporting, editing and publication of scholarly work in medical journals.

Informed consent

Not applicable.

Guarantor

FHBMS

Contributorship

CK, LS, NS, FS: study conceptualisation and methodology. MC, LS, NS, FS: data extraction, MC and LS data analysis, all authors: interpretation of data. MC, LS, FS: writing original draft, NS, RASS, CK and FS: supervision, all authors: writing-review, editing and approval of the manuscript.

Trial registration

Not applicable.

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Data availability

The full data and extraction forms of this review can be provided on request via the corresponding author.

Supplemental material

Supplemental material for this article is available online.

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