

# 1 Early life adversity has sex-dependent effects on 2 survival across the lifespan in rhesus macaques

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35 Abstract

36 Exposure to adversity during early life is linked to lasting detrimental effects on  
37 evolutionary fitness across many taxa. However, due to the challenges of collecting  
38 longitudinal data, especially in species where one sex disperses, direct evidence from  
39 long-lived species remains relatively scarce. Here we test the effects of early life  
40 adversity on male and female longevity in a free-ranging population of rhesus  
41 macaques (*Macaca mulatta*) at Cayo Santiago, Puerto Rico. We leveraged six decades  
42 of data to quantify the relative importance of ten forms of early life adversity for 6,599  
43 macaques (3,230 male, 3,369 female), with a smaller sample size (N=299) for one form  
44 of adversity (maternal social isolation) which required high-resolution behavioral data.  
45 We found that individuals who experienced more early life adversity died earlier than  
46 those who experienced less adversity. Mortality risk was highest during early life,  
47 defined as birth to four years old, suggesting acute survival effects of adversity, but  
48 heightened mortality risk was also present in macaques who survived to adulthood.  
49 Females and males were affected differently by some forms of adversity, and these  
50 differences might be driven by varying energetic demands, female philopatry, and male  
51 dispersal. By leveraging data on thousands of macaques collected over decades, our  
52 results show that the fitness consequences of early life adversity are not uniform across  
53 individuals but vary as a function of the type of adversity, timing, and social context, and  
54 thus contribute to our limited but growing understanding of the evolution of early life  
55 sensitivities in long-lived species.

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## 58 Significance Statement

59 Exposure to early life adversity, even when conditions subsequently improve, can have  
60 profound and persistent consequences for human health. Negative effects of early life  
61 adversity appear widespread across the animal kingdom. To date, however, direct  
62 evidence from long-lived species is relatively scarce due to the difficulties of collecting  
63 data from early life till death. We leverage six decades of observations on thousands of  
64 free-ranging male and female rhesus macaques to examine the complex ways early life  
65 adversity impacts survival. Our results suggest that the type of adversity and life history  
66 factors intersect to impact immediate and downstream survival. By studying early life  
67 adversity across environments, cultures, contexts, and species, we can better  
68 understand the evolutionary underpinnings of early life sensitivities.

69

## 70 Introduction

71

72 Exposure to adversity such as food shortages and social isolation during early life can  
73 result in long-term health and evolutionary fitness consequences in a wide range of  
74 species such as insects, birds, fish, reptiles, and mammals (Cooper & Kruuk, 2018;  
75 Eyck et al., 2019; Lea & Rosenbaum, 2020; Lu et al., 2019). For example, female red  
76 deer (*Cervus elaphus*) that face high resource competition during development exhibit  
77 accelerated senescence in adulthood (Nussey et al., 2007). Early life adversity in  
78 female baboons (*Papio sp.*) is associated with reduced fecundity and poorer offspring  
79 survival (Lange et al., 2023; Lea et al., 2015; Patterson et al., 2021; Tung et al., 2016,  
80 2023; Weibel et al., 2020; Zippel et al., 2019). Organisms are hypothesized to adjust

81 their developmental trajectories in response to early life adversity in order to improve  
82 immediate survival (Lea et al., 2017; Lea & Rosenbaum, 2020; Patterson, Petersen et  
83 al., 2023), but such adjustments may lead to these detrimental outcomes in adulthood.

84 A small but growing number of studies have tested the long-term impacts of early life  
85 adversity, but further research on how different forms of early life adversity shape the  
86 timing of fitness consequences in a variety of species, populations, and contexts is  
87 needed to better understand the evolution of early life sensitivities to adversity.

88

89 Survival, which is a prerequisite for reproductive success and an important aspect of  
90 evolutionary fitness, has been linked to early life adversity in a number of species.

91 Given the difficulty of measuring fitness directly, a common approach is to quantify  
92 lifetime reproductive success (LRS), which is the total number of offspring produced

93 over an individual's lifetime (Clutton-Brock 1988). Adult lifespan is the biggest

94 contributor to LRS in long-lived species like roe deer (*Capreolus capreolus*), baboons,

95 rhesus macaques (*Macaca mulatta*) (Blomquist, 2009), and gorillas (*Gorilla beringei*

96 *beringei*) (Clutton-Brock et al., 1983; Kjellander et al., 2004; Rhine et al., 2000; Robbins

97 et al., 2011; Van de Walle et al., 2022). Females who live longer have a longer

98 reproductive span, and are able to produce more offspring. Long-term studies have

99 shown that early life adversity is associated with reduced survival in adult female

100 baboons and hyenas (*Crocuta crocuta*) (Gicquel et al., 2022; Lange et al., 2023;

101 Strauss et al., 2020; Tung et al., 2016). In populations characterized by high mortality

102 rates prior to reproductive maturity, LRS based on the total number of offspring *reaching*

103 *reproductive maturity* is the best proxy for fitness (e.g., Alif et al. 2022). Death prior to

104 maturity has severe fitness costs for the organism who fails to reproduce, but also for  
105 the organism's parents. Male and female gorillas exposed to early life adversity  
106 experience reduced survival prior to reproductive maturity, but do not experience  
107 survival costs after maturity (Morrison et al., 2023). As such, adult survival patterns in  
108 gorillas differ from those in yellow baboons and spotted hyenas, but pre-reproductive  
109 survival patterns are not yet available across species for comparison. More studies  
110 across different species are thus needed to both identify the fitness consequences of  
111 early life sensitivities and to draw comparisons and better understand the evolutionary  
112 pressures which shape the developmental responses to adversity and ultimately  
113 influence fitness across the lifespan.

114

115 The fitness consequences of early life adversity might vary in a sex-dependent manner  
116 due to differences in life history strategies. During adverse early life conditions, the sex  
117 with more energetically demanding traits is predicted to be more susceptible to  
118 nutritional constraints (Clutton-Brock, 1994; Clutton-Brock et al., 1985). Further, during  
119 adverse conditions, parents are predicted to reduce investment in more energetically  
120 costly offspring, thus exacerbating consequences of adversity especially during the  
121 period of care (Clutton-Brock et al., 1985; Trivers & Willard, 1973). In many species,  
122 males are considered more energetically costly given faster growth and larger body size  
123 compared to females. When male fitness is largely determined by access to mates via  
124 competitive ability, males should also invest in costly developmental processes like play  
125 and motor skill development (Lonsdorf, 2017). In support of these predicted differences,  
126 among red deer (*Cervus elaphus*), maternal death prior to weaning was linked to higher

127 mortality risk among male compared to female offspring (Andres et al., 2013). Sex-  
128 dependent effects of early life adversity are challenging to study because many species  
129 are characterized by sex-biased dispersal such that pre- or post-dispersal data are  
130 typically missing for individuals of the dispersing sex. More studies are thus needed that  
131 can follow both males and females from birth till death to investigate how life history and  
132 parental investment strategies shape fitness consequences and developmental  
133 responses to early life adversity.

134

135 Here we leverage a large historical dataset of the free-ranging rhesus macaques  
136 (*Macaca mulatta*) of Cayo Santiago to advance our understanding about the magnitude,  
137 form, timing, and sex-dependence of early life adversity effects. Complete demographic  
138 records extend back to the 1960s for thousands of male and female macaques at the  
139 Cayo Santiago field site. In this population, abusive maternal care behavior is linked to  
140 differences in HPA function in juveniles, the presence of a competing younger sibling is  
141 linked to reduced survival during juvenility, and exposure to hurricanes and high  
142 population density during early life are linked to life history trade-offs in adulthood (Lee  
143 et al., 2019; Luevano et al., 2022; Petrullo et al., 2016). We examine the effects of ten  
144 forms of potential early life adversity on sex-specific mortality risk across early life, and  
145 separately, across adulthood. By examining mortality across the lifespan, we can  
146 identify the full extent of variation in the fitness consequences of early life adversity. If  
147 mortality consequences are severe in early life, this might influence survival patterns  
148 observed in adulthood (e.g., survivorship bias), and it would be important to consider  
149 why some individuals can survive to reproductive debut while others do not.

150

151 We predict that rhesus macaques exposed to greater amounts of early life adversity will  
152 have increased mortality risk. We predict mortality risks will be more severe during the  
153 first four years of life when individuals are still growing and adversity is more recent.  
154 However, we also predict heightened mortality risk will persist into adulthood among  
155 those who survive past four years of age. Given male life history strategies in this  
156 species (Hoffman et al., 2008) that prioritize costly traits like faster growth, larger body  
157 size, and motor skill development (Kulik et al., 2015; Schwartz & Kemnitz, 1992;  
158 Turcotte et al., 2022), we predict that early life adversity will exert larger effects on  
159 males than females.

160

## 161 Methods

### 162 Study site and population

163 We studied a free-ranging population of rhesus macaques living on Cayo Santiago, a  
164 15.2-ha island off the southeastern coast of Puerto Rico. The current population of  
165 individually recognized ~1,700 rhesus macaques are the descendants of 409 monkeys  
166 that were transported from India to the island in 1938. This population is managed by  
167 the Caribbean Primate Research Center (CPRC) of the University of Puerto Rico.  
168 Monkeys are fed monkey chow daily and water catchments provide *ad libitum* access to  
169 drinking water. The monkeys live in naturally forming multi-male, multi-female social  
170 groups characterized by dominance hierarchies and male dispersal. Monkeys mate with  
171 multiple partners and breed seasonally. This species is characterized by sexual

172 dimorphism with males exhibiting larger body mass and canine length than females  
173 (Schwartz & Kemnitz, 1992; Turcotte et al., 2022). Males queue for dominance rank,  
174 have large testes, and experience indirect male-male competition (Higham &  
175 Maestriperi, 2014; Kimock et al., 2019, 2022). Male mortality is highest during the  
176 mating season, consistent with the notion that males prioritize investment in mating  
177 effort (Higham & Maestriperi, 2014; Hoffman et al., 2008). Females prioritize investment  
178 in gestation and lactation, and face the highest mortality risk during the birth season  
179 (Hoffman et al., 2008). The island is free of predators and there is no regular veterinary  
180 intervention, so the primary causes of death are illness and injury (Pavez-Fox et al.,  
181 2022).

182

183 During the study period (1960-2021), observers monitored and recorded demographic  
184 events daily. These records include births, deaths, sex, matriline, matriline rank,  
185 maternal identification, sires when genetic data were available, and group emigration  
186 and immigration events. A genetic pedigree is available for much of the population  
187 (Widdig et al., 2016). Daily total rainfall and mean maximum temperature data were  
188 pulled from the NOAA station in Rio Piedras, Puerto Rico. Over a 61-year period (1959-  
189 2020), data were not recorded by this NOAA station for 21% of days. Rather than  
190 removing a large portion of data, we imputed missing rainfall and temperature data  
191 using the 'mice' package in R (Buuren & Groothuis-Oudshoorn, 2011). This study  
192 includes 6,599 individuals for which there are complete data available. We had data  
193 covering the entire lifespan—birth to death—from 2,513 macaques. The remaining 4,086  
194 macaques were either alive at the time of this study or were removed from the island



195 prior to natural death as a result of population management (i.e., right-censored  
196 samples).

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198

199 Early life adversities

200

201 We used historical demographic records to assess individual exposure to early life  
202 adversity (Tung et al., 2016). We consider ten forms of potential early life adversity  
203 based on previous research on this population and other species, as described below.  
204 In choosing time periods of exposure for each form of adversity, we followed Tung and  
205 colleagues 2016:

206

207 **Maternal loss:** Maternal death increases offspring mortality in humans and other  
208 mammals (Cayo Santiago macaques: Blomquist, 2013; red deer: Andres et al., 2013;  
209 Asian elephants (*Elephas maximus*): Lahdenperä et al., 2016; humans: Sear & Mace,  
210 2008; chimpanzees (*Pan troglodytes*): Stanton et al., 2020; yellow baboons: Tung et al.,  
211 2016). Following previous work (Tung et al., 2016; Zippel et al., 2021), we consider an  
212 individual to experience maternal loss if their mother died (including natural death  
213 (N=1,165) and permanent removal from the population (N=299)) before the individual  
214 reached 4 years of age. This four year period includes the period during which young  
215 macaques are nutritionally and socially dependent on their mothers. While a mother's  
216 removal from the population had to occur while the offspring was alive (i.e., prior to  
217 offspring death if they have died) to be considered an adversity, this was not a

218 requirement for maternal death because an imminent maternal death is linked to  
219 offspring mortality risk—an association likely explained by poor maternal condition  
220 (Zipple et al., 2021). If maternal loss occurred after the offspring reached four years old,  
221 we did not consider the offspring to have experienced this source of early life adversity.  
222 Maternal loss was measured as a binary variable: experienced maternal loss or did not  
223 experience loss.

224 **Competing sibling:** The presence of a close in age younger sibling represents a  
225 source of competition over maternal resources and is associated with higher mortality  
226 risk (Cayo Santiago macaques: Lee et al., 2019; yellow baboons: Tung et al., 2016). We  
227 considered a sibling to be a competitor if the sibling was born within 355 days, which  
228 represented the bottom quartile of interbirth intervals in our sample. Last born offspring  
229 and individuals which died before their sibling was born did not experience this  
230 adversity. The presence of competing siblings was measured as a binary variable.

231 **Group size:** High group size and high population density are indicative of more  
232 competition and are associated with reductions in fecundity (Cayo Santiago macaques:  
233 Luevano et al., 2022; red deer: Clutton-Brock et al., 1982, 1983; meerkats (*Suricata*  
234 *suricatta*): Clutton-Bock 2008). We used group size as a proxy for within-group  
235 competition. Demographic records were used to construct group composition over the  
236 study period. Group size was defined as the number of adults ( $\geq 4$  years of age) of both  
237 sexes in an individual's social group on the day that individual was born (Tung et al.,  
238 2016), and was included in our models as a continuous variable (range: 2-222  
239 individuals).

240 **Primiparity:** The high energetic demands on first time mothers can result in negative  
241 outcomes for offspring such as increased mortality risk (Rhesus macaques: Bercovitch  
242 et al., 1998; Blomquist, 2013; Nuñez et al., 2015; vervet monkeys (*Chlorocebus*  
243 *sabaeus*): Fairbanks & Mcguire, 1995; mantled howler monkeys (*Alouatta palliata*):  
244 Glander, 1980; Asian elephants: Mar et al., 2012; olive baboons (*Papio anubis*): Smuts  
245 & Nicolson, 1989). We used a binary measurement for primiparity: first born or not first  
246 born.

247 **Matriline rank:** Dominance rank mediates access to food and is linked to survival,  
248 fecundity, and offspring growth (Cayo Santiago macaques: Blomquist et al., 2011; Weiß  
249 et al., 2016; yellow baboons: Altmann & Alberts, 2005; olive baboons: Garcia et al.,  
250 2009; chacma baboons (*Papio ursinus*): Johnson, 2003). Matrilineal dominance  
251 hierarchies for a given social group and year are recorded as categorical – high, middle,  
252 low – based on data from Donald S. Sade and John D. Berard, who recorded dyadic  
253 agonistic interactions (e.g., threats, displacements, submissive behaviors) across the  
254 year to calculate matrilineal dominance matrices (Lee et al., 2019; Missakian, 1972).

255 **Kin network:** Among prime aged adult females at Cayo Santiago, the presence of more  
256 maternal kin is linked to better survival in any given year (Brent et al., 2017). We  
257 measured an individual's maternal kin network size at birth as the number of living  
258 females over 4 years of age with a relatedness coefficient of at least 0.063. This  
259 relatedness coefficient was chosen because 0.063 represents the threshold at which  
260 macaques in this population can recognize kin via vocalizations (Rendall et al., 1996),  
261 and this threshold was used in previous work showing a positive association between

262 the number of relatives present and adult survival (Brent et al., 2017). Kin network size  
263 was included as a continuous variable (range: 1-21 individuals).

264 **Maternal social connectedness:** Greater social connectedness is associated with  
265 better survival and better offspring survival (Cayo Santiago macaques: Brent,  
266 Heilbronner, et al., 2013; Ellis et al., 2019; yellow & chacma baboons: Archie et al.,  
267 2014; Silk et al., 2003, 2010). We used behavioral data collected during 10-min focal  
268 animal samples on adults in several social groups from 2010-2017 (details provided in  
269 the SOM). To measure maternal social connectedness, we calculated a composite  
270 sociality index (CSI) using the affiliative social behaviors, approaches and grooming.  
271 For each mother in each year, we tabulated the rate of approaches (approaches to and  
272 from other adult females / hours observed) and the rate of grooming bouts (number of  
273 grooming bouts given and received / hours observed). A mother's approach and  
274 grooming rates were divided by the mean rate for all adult females in each social group  
275 in each year. These standardized approach and grooming rates were added together  
276 and divided by 2 (the number of behaviors) to create the CSI for each mother. Here, we  
277 followed Tung and colleagues (2016): for each offspring in our analyses, we averaged  
278 their mother's composite sociality index for the first two years of life.

279 **Rainfall:** More rainfall is indicative of greater food and water availability, and is linked to  
280 greater fecundity and better survival (primates: Campos et al., 2017; yellow baboons:  
281 Lea et al., 2015; gelada monkeys (*Theropithecus gelada*): Sloan et al., 2022). However,  
282 because food and water are provisioned and due to the negative effects of tropical  
283 storms at Cayo Santiago, low rainfall might not be as relevant or have negative

284 consequences in this population. Here, we used total rainfall across the first year of life  
285 (range: 1,021.4-3,157.1 mm) (Tung et al., 2016).

286 **Temperature:** Higher temperatures are linked to fecundity, cognition, and mortality  
287 (Western Australian magpies (*Cracticus tibicen dorsalis*): Blackburn et al., 2022;  
288 southern pied babbler (*Turdoides bicolor*): Bourne et al., 2020; dairy cattle (*Bos taurus*):  
289 Polsky & von Keyserlingk, 2017; gelada monkeys: Sloan et al., 2022). Here, we  
290 averaged mean maximum daily temperatures across the first year of life (range: 85.12-  
291 89.89 F) (Tung et al., 2016).

292 **Hurricanes:** Exposure to major hurricanes is linked to female reproductive strategies,  
293 demographic roles, and immunological aging (Cayo Santiago macaques: Diaz et al.,  
294 2023; Luevano et al., 2022; Watowich et al., 2022). We recorded individual exposure to  
295 any of the 3 major hurricanes that had major impacts on Cayo Santiago (Hugo on  
296 September 18, 1989, Georges on September 21, 1998, and Maria on September 20,  
297 2017) during the first year of life. Hurricane exposure was not included in previous  
298 studies of early life adversity, so we chose the first year of life as our window of  
299 exposure to align with our other weather variables, rainfall and temperature. If  
300 individuals were exposed to hurricanes when they were over one year old, we did not  
301 consider them to have experienced this adversity as *early life* adversity.

302

303 Several approaches have been used to conceptualize, process, and analyze early life  
304 adversity data. Some studies of early life adversity use broad, cumulative measures of  
305 adversity, while others focus on different forms of adversity separately (Gunnar, 2020;  
306 Smith & Pollak, 2020). Empirical evidence suggests that the accumulation of multiple

307 adversities is a better predictor of adult outcomes than any particular form of adversity,  
308 but there is also evidence that specific forms of adversity lead to different outcomes  
309 (Gicquel et al., 2022; Gunnar, 2020; Tung et al., 2016). Here, we examine cumulative  
310 adversity measures and then examine individual forms of adversity separately. To  
311 construct a cumulative early life adversity index, we summed individuals' exposure to  
312 different forms of adversity. Previous studies typically relied on binary scores for each  
313 form of adversity. Here, we use continuous measures of adversity when feasible. For  
314 purposes of the cumulative index, continuous measures (i.e., group size, kin network  
315 size, high temperature, and rainfall) were normalized so values range from zero to one.  
316 For binary measures (i.e., maternal loss, being a first born, presence of a competing  
317 sibling, and hurricane exposure), individuals were assigned a value of one if they  
318 experienced a given form of adversity and a value of zero if they did not experience that  
319 form of adversity. Those born into high ranking matriline were assigned a zero, mid  
320 ranking matriline were assigned 0.5, and low ranking matriline were assigned a value  
321 of one. As such, each variable ranged from zero to one and were summed together into  
322 a cumulative index to represent the total exposure to early life adversity. Our main  
323 cumulative early life adversity index could range from 0-9 because it included nine  
324 variables: maternal loss, presence of a competing younger sibling, high group size,  
325 primiparity, low matrilineal dominance rank, small kin network, hurricane exposure, high  
326 temperature, and low rainfall. Maternal social connectedness is not included because it  
327 was derived only for a subset of our data.

328

329 Data analysis

330 To determine if early life adversity predicts survival, we used survival models. The  
331 outcome variable was age at death. Individuals who were either still alive at the end of  
332 the study or removed from the island for population control were right-censored. We ran  
333 models on the full sample of all ages (N=6,599) but right-censored to four years old to  
334 examine early life mortality, and we ran models on a subsample of individuals who  
335 survived beyond four years of age (N=2,866) to examine mortality across adulthood.  
336 Early life adversity predictor variables were modeled two ways: 1) cumulative index  
337 models include all forms of adversity summed together into one variable, and 2)  
338 multivariate models include each form of adversity modeled as individual predictor  
339 variables. We ran separate models to examine the survival effects of maternal social  
340 connectedness during early life because our focal behavioral data does not span the  
341 entire study period (N=299 early life survival; N=101 adult survival). Models included  
342 early life adversity index, sex, and an interaction term between sex and early life  
343 adversity. Models also included a varying intercept for birth year and maternal  
344 identification. The different forms of adversity we examined were not correlated, but  
345 high temperatures in the first year of life were highly correlated with birth year (Table  
346 S1).

347

348 We first used Cox survival models, but the proportional hazards assumption in the Cox  
349 model was violated in three out of four models. Specifically, early life and adult  
350 multivariate models and the early life cumulative index model violated assumptions.  
351 Individual sex, maternal loss, high temperatures, group size, kin network size,  
352 matrilineal rank, hurricane exposure, and four of the sex interaction terms violated

353 proportional hazard assumptions (cox.zph function in R package, “survival”:  $p < 0.05$ ).

354 Instead, we fit Accelerated Failure Time (AFT) survival models with a Weibull  
355 distribution. The presence of a competing younger sibling is time-dependent since  
356 individuals cannot experience this exposure unless they survive till a given age, i.e.,  
357 until it’s biologically possible for the mother to give birth again. To include this variable,  
358 we would need to include it as a time-varying variable in a Cox proportional hazard  
359 model. As such, we were unable to test how the presence of a competing younger  
360 sibling affects survival in early life and we excluded this variable from the cumulative  
361 index for the early life survival model. We could, however, examine this in adulthood  
362 since all individuals in the sample survived to adulthood and the presence of a  
363 competing younger sibling is not time-varying.

364

365 Genetics can contribute to the effects of early life adversity. For example, individuals  
366 experiencing maternal loss might have shorter lifespans due to genes shared by both  
367 the mother and offspring. To estimate to what extent variance in survival is explained by  
368 genetics and early life adversity, we account for pedigree in a subsample of the data for  
369 which we had complete pedigree information (N=923 individuals during early life; N=307  
370 adults). To do so, we used the animal model and incorporated the relationship  
371 covariance matrix as a random effect (Wilson et al., 2010). The models with and without  
372 pedigree produced similar results (Table S2). The effects of pedigree on survival were  
373 substantial, but accounting for genetic relatedness in the model did not diminish the  
374 effects of early life adversity on survival in early life (with pedigree:  $\beta = -0.33 \pm 0.12$ ;  
375 without pedigree:  $\beta = -0.31 \pm 0.11$ ) or adulthood (with pedigree:  $\beta = -0.12 \pm 0.04$ ; without



376 pedigree:  $\beta = -0.12 \pm 0.04$ ). The full pedigreed sample is smaller than our main dataset  
377 because paternity is unknown for many animals earlier in the study. Because the effects  
378 of early life adversity were unaffected by pedigree inclusion and because the sample  
379 size for pedigree inclusion is much smaller, we have presented the larger set of data  
380 without pedigree in the main text.

381  
382 Models were run with the brms package (v 2.16.3) in R (v 4.1.2 ) (Bürkner, 2017). All  
383 continuous predictor variables were transformed to a mean of 0 and a standard  
384 deviation of 1. All models are Bayesian, and we used weakly informative priors for fixed  
385 effects, setting the mean to zero and the standard deviation to one. To produce more  
386 accurate predictions for age at death, we used more regularizing priors for the intercept  
387 (a mean of 1 and standard deviation of 0.1 for the early life survival models, and a mean  
388 of 12 and standard deviation of 0.4 for the adult survival models). Specifically, our  
389 analyses contain a high proportion of right-censored cases, which can lead to model  
390 predictions that overestimate life expectancy (Alam et al., 2022). We use credible  
391 intervals to determine whether the effect of a variable is substantial or not. If the 85%  
392 credible interval for an effect does not overlap with zero, the effect is considered  
393 substantial. When the vast majority of the 85% credible interval does not span zero, but  
394 there is some overlap, we consider the model to be uncertain about the effect. The code  
395 and data used can be found here: [https://github.com/skpatter/ELA\\_Survival\\_Macaques](https://github.com/skpatter/ELA_Survival_Macaques)

396

397 Results

398 *Cumulative early life adversity is associated with reduced survival during early life and*  
399 *during adulthood*

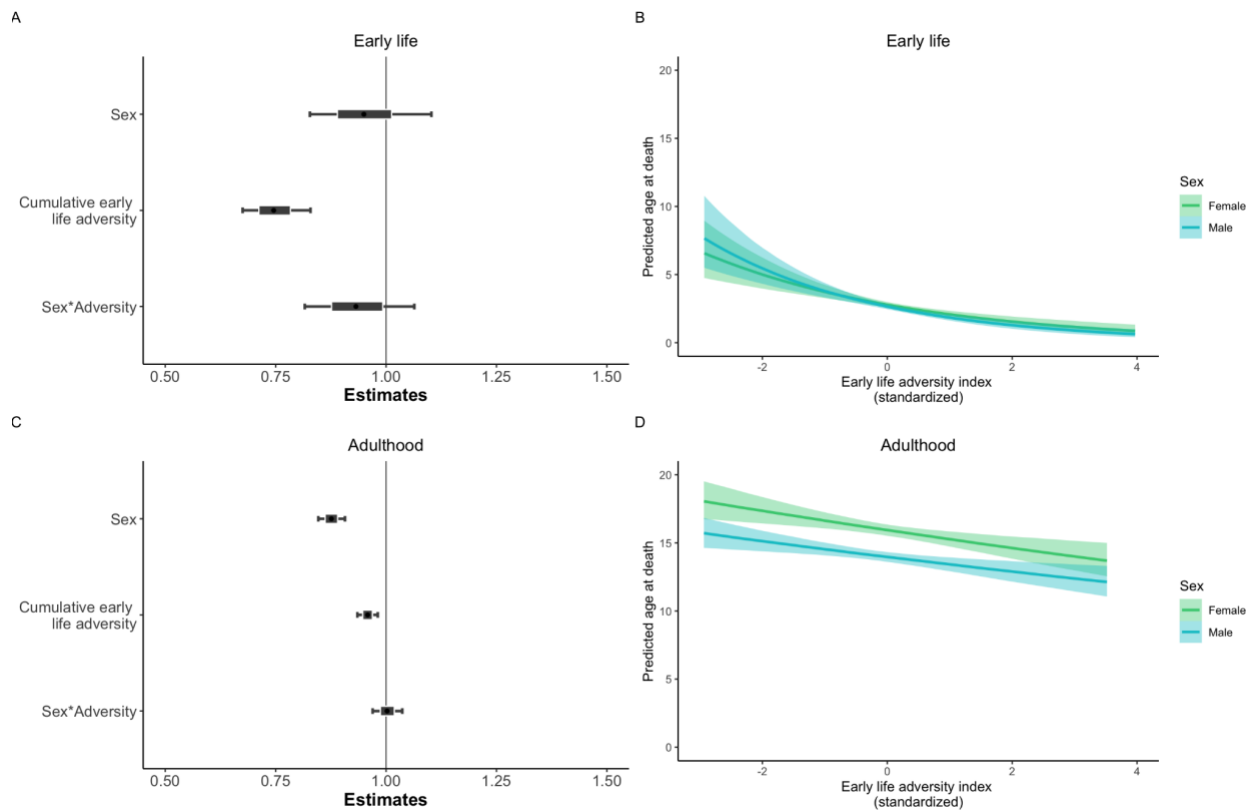
400

401 Individuals who experienced more cumulative early life adversity had higher mortality  
402 during early life ( $\beta=-0.29\pm 0.07$ ; Figure 1, Table S3). There were no clear differences in  
403 mortality for males versus females during early life ( $\beta=-0.05\pm 0.10$ ), and there was no  
404 evidence that early life adversity differentially affected mortality risk as a function of sex  
405 during early life ( $\beta=-0.07\pm 0.09$ ; Figure 1, Table S3). Adults who experienced more  
406 cumulative early life adversity had shorter lives than adults with less early life adversity  
407 ( $\beta=-0.04\pm 0.02$ ; Figure 1, Table S3). Among adults, females lived longer than males ( $\beta=-$   
408  $0.13\pm 0.02$ ), and there was no evidence that cumulative early life adversity differentially  
409 affected mortality risk among males and females ( $\beta=0.00\pm 0.02$ ; Figure 1; Table S3).  
410 Between those who experienced the least and the most amount of cumulative early life  
411 adversity in our sample, these effects translate to a 4.78-year difference in average life  
412 expectancy among adult females and a 3.94-year difference in average life expectancy  
413 among adult males.

414

415 **Fig 1.** Model effects of cumulative early life adversity and sex on survival during early  
416 life (A) and adulthood (C). The outer bars show the 85% credible intervals, the inner box  
417 shows the 50% credible intervals, and the black circle in the middle shows the median  
418 of the posterior distribution. Model predictions are shown for the effect of cumulative  
419 early life adversity on lifespan in early life (B) and adulthood (D). Green predictions

420 represent females and blue predictions represent males. The solid lines show the  
421 median estimates and the shaded region shows the 85% credible intervals.



422

423

424 *Several forms of early life adversity are associated with reduced survival, and some*  
425 *effects are sex-dependent*

426

427 Individuals who **lost their mother** during the first four years of life had a higher mortality

428 risk during early life ( $\beta=-0.33\pm 0.14$ ) and adulthood ( $\beta=-0.06\pm 0.04$ ) than those who did

429 not lose their mother (Figure 2, Table S4). Maternal loss had a larger negative effect on

430 sons than daughters in both early life ( $\beta=-0.32\pm 0.20$ ) and adulthood ( $\beta=-0.05\pm 0.06$ ;

431 Figure 3, Table S4). **Higher temperatures** during the first year of life were associated

432 with higher mortality rises in early life ( $\beta=-0.06\pm 0.24$ ) and adulthood ( $\beta=-0.02\pm 0.02$ ), but

433 the models were uncertain about these effects (Figure 2, Table S4). The effect of high  
434 temperatures on survival was moderated by sex (Figure 3; Table S4). Higher  
435 temperatures during the first year of life were more strongly associated with reduced  
436 survival among males than females during early life ( $\beta=-0.23\pm0.09$ ), but in adulthood,  
437 only females experienced this survival cost ( $\beta=0.02\pm0.03$ ).

438

439 **First-born** offspring had elevated mortality risk during early life ( $\beta=-0.36\pm0.16$ ). In  
440 contrast, individuals born to primiparous mothers had better survival in adulthood than  
441 those born to multiparous mothers, although the model was uncertain about this effect  
442 ( $\beta=0.06\pm0.05$ ). Effects of maternal primiparity were not moderated by sex (early life:  
443  $\beta=0.09\pm0.22$ ; adulthood:  $\beta=-0.04\pm0.07$ ). Macaques born into **low ranking matriline**  
444 had a higher mortality risk during early life ( $\beta=-0.44\pm0.17$ ) and adulthood ( $\beta=-0.16\pm0.05$ )  
445 than those born into high ranking matriline. Matrilineal rank was more strongly  
446 associated with female survival than male survival during both periods of life (early life:  
447  $\beta=0.47\pm0.22$ ; adulthood:  $\beta=0.04\pm0.06$ ; Figure 3; Table S4). We also treated matrilineal  
448 rank as an ordinal variable and found similar results (Table S4).

449

450 **Smaller maternal kin networks** at birth were associated with higher early life mortality  
451 risk, especially for males ( $\beta=0.05\pm0.07$ ; sex interaction:  $\beta=0.14\pm0.10$ ). Smaller maternal  
452 kin networks at birth were associated with better survival for adult males, but reduced  
453 survival for adult females ( $\beta=0.01\pm0.02$ ; sex interaction:  $\beta=-0.05\pm0.03$ ; Figure 3; Table  
454 S4). The presence of **competing younger siblings** was associated with higher  
455 mortality risk, but the model was uncertain about this effect ( $\beta=-0.03\pm0.04$ ). Although

456 the model was uncertain, a competing sibling had a slightly larger effect on females  
457 ( $\beta=0.06\pm0.06$ ). We were unable to examine survival effects during early life given time-  
458 varying issues.

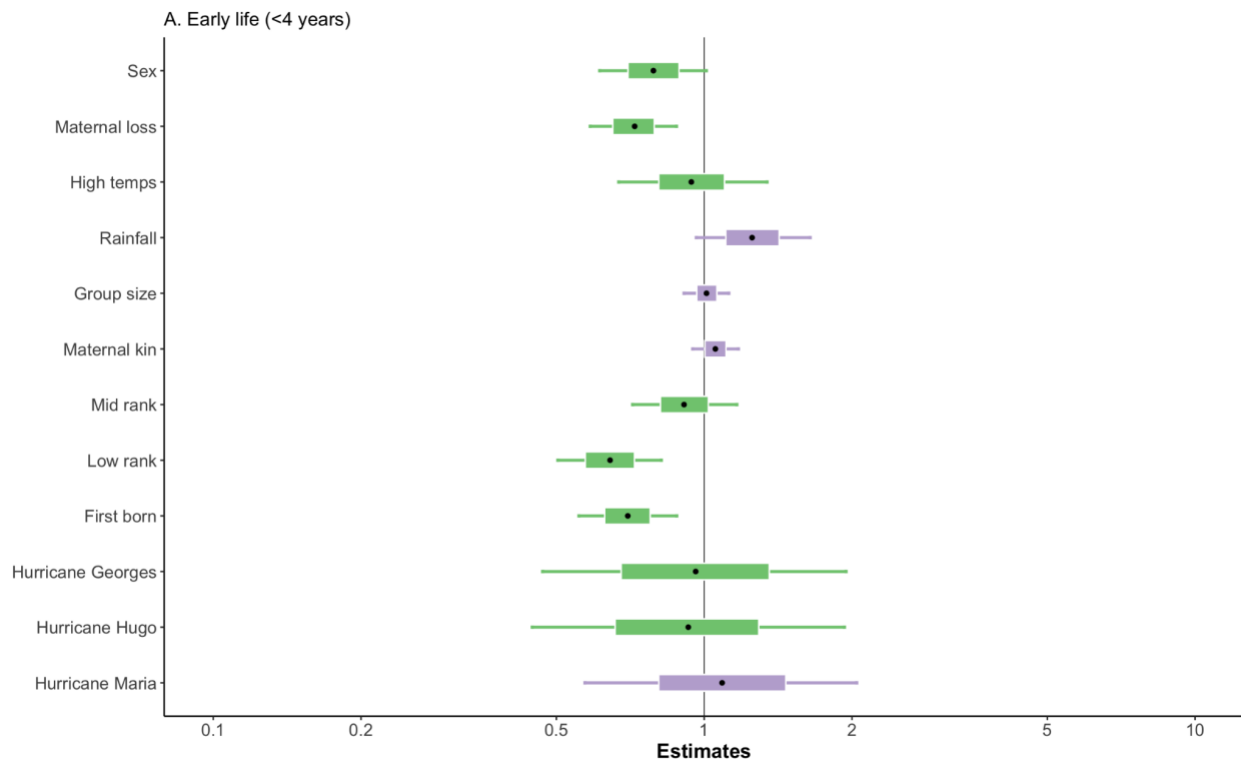
459

460 Low **rainfall** was associated with reduced survival during early life, but the model was  
461 uncertain about this estimate ( $\beta=0.23\pm0.18$ ). No effect of rainfall was found during  
462 adulthood ( $\beta=-0.01\pm0.02$ ; Figure 2; Table S4). No effect of **group size** was observed  
463 during early life ( $\beta=0.01\pm0.07$ ), and while the model was uncertain, it seems adults born  
464 into larger groups exhibited reduced survival ( $\beta=-0.02\pm0.02$ ; Figure 2; Table S4). No  
465 effect of major **hurricanes** was observed during early life (Georges:  $\beta=-0.04\pm0.50$ ;  
466 Hugo:  $\beta=-0.08\pm0.50$ ; Maria:  $\beta=0.08\pm0.44$ ) or adulthood (Georges:  $\beta=-0.02\pm0.10$ ; Hugo:  
467  $\beta=-0.04\pm0.10$ ; Figure 2; Table S4). The model was uncertain about the effects of  
468 maternal social isolation; individuals born to **socially isolated mothers** seemed to have  
469 higher mortality during early life than those born to more socially connected mothers  
470 ( $\beta=0.15\pm0.17$ ; Table S5), and males were more affected by this than females  
471 ( $\beta=0.29\pm0.25$ ). No effect of maternal social isolation was observed among adults ( $\beta=-$   
472  $0.04\pm0.02$ ; Table S5). Model comparisons revealed no substantial difference between  
473 models constructed with the cumulative early life adversity index versus those  
474 constructed with each form of early life adversity separately (Table S6).

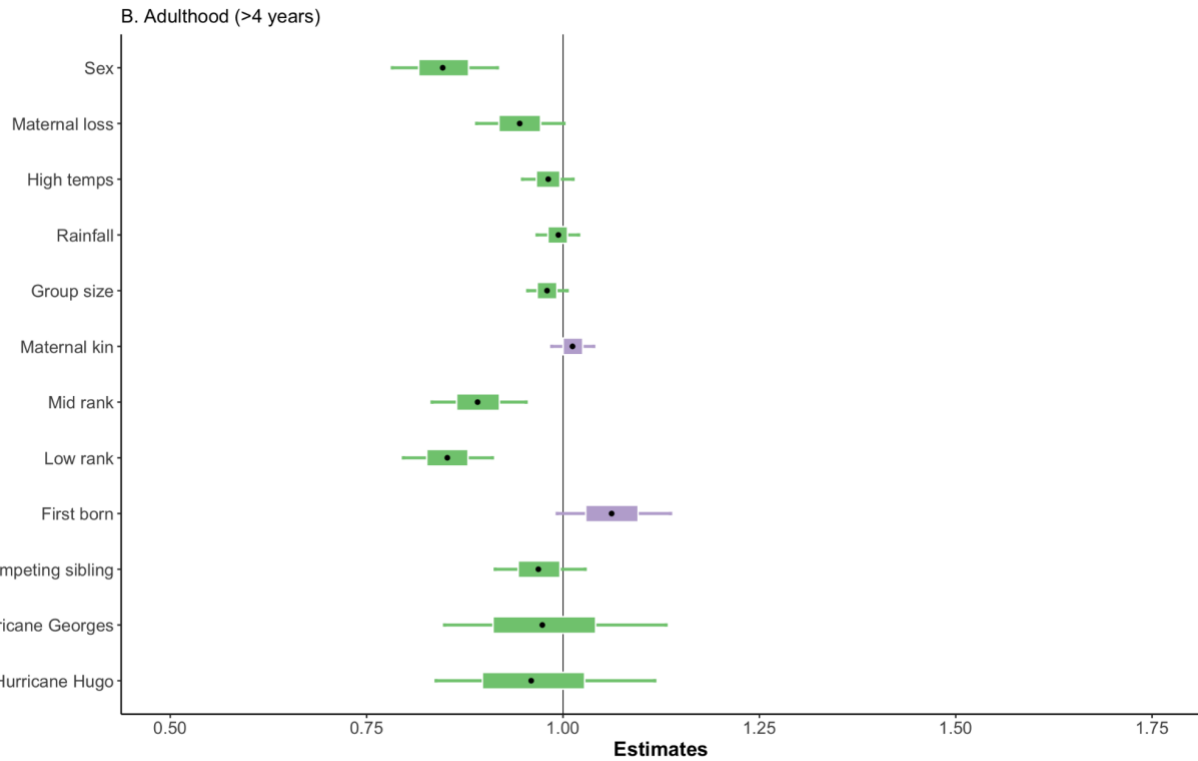
475

476 **Fig2.** Model effects of sex and the forms of early life adversity on survival during early  
477 life (A) and adulthood (B). The outer bars show the 85% credible intervals, the inner  
478 boxes show the 50% credible intervals, and the black circles in the middle show the

479 medians of the posterior distributions. Green shading represents negative effect sizes,  
480 meaning that the variable is associated with shorter lifespans, and purple shading  
481 represents positive effect sizes, meaning that the variable is associated with longer  
482 lifespans. For interaction effects, see Figure S1.



483



484

485

486 **Fig 3.** Interactions between sex and three forms of early life adversity on adult survival.

487 (A) and (E) Predicted relationship between maternal kin network size at birth and

488 survival for females (green) and males (blue). (B) and (F) Predicted relationship

489 between high temperatures during the first year of life and survival for females (green)

490 and males (blue). (C) and (G) Predicted relationship between maternal loss during the

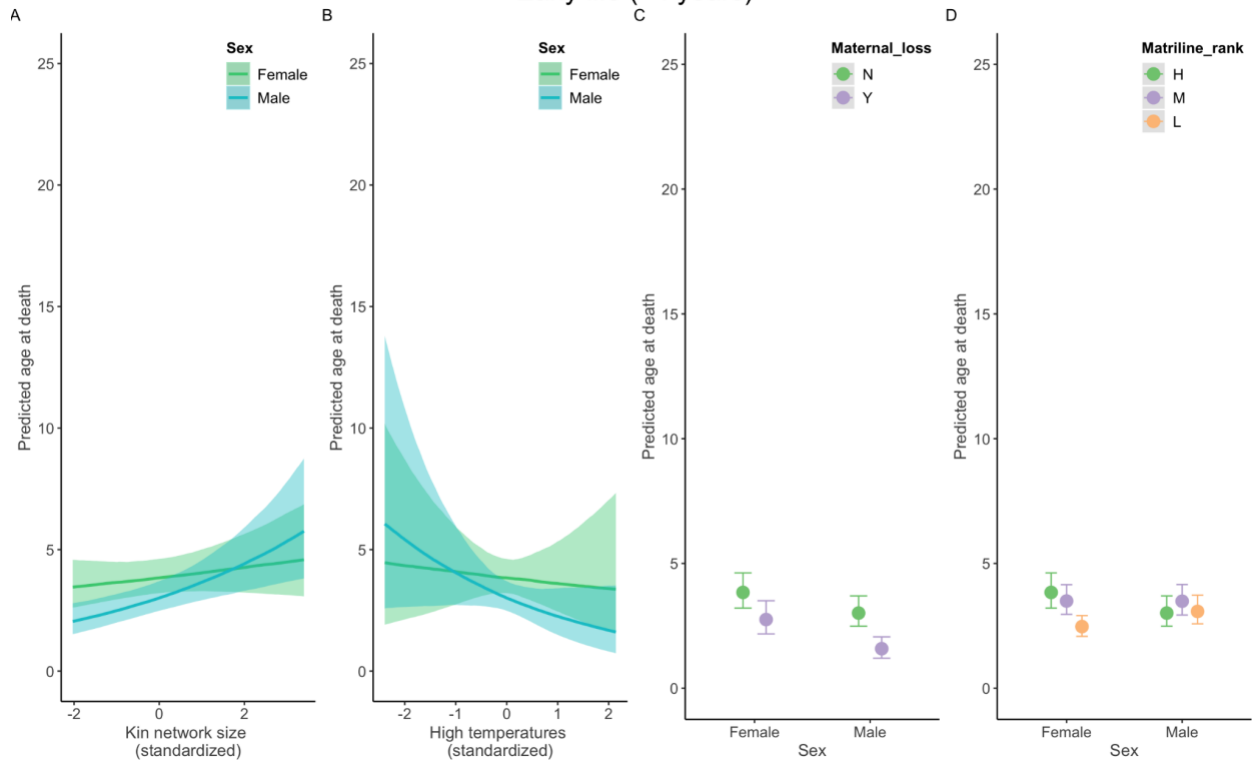
491 first four years of life and survival for males versus females. The circles show the

492 median estimate, and the bars show the 85% credible intervals. (D) and (H) Predicted

493 relationship between matrilineal rank and survival for males versus females. The circles

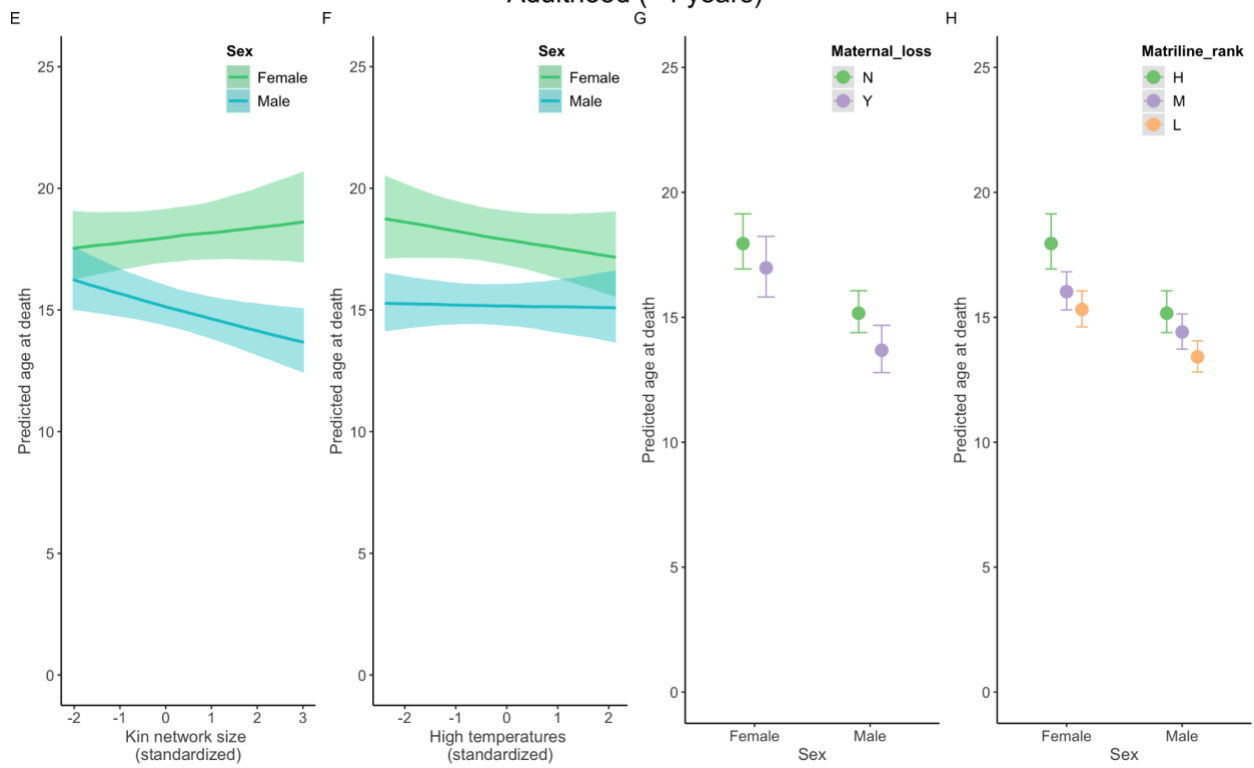
494 show the median estimate, and the bars show the 85% credible intervals.

### Early life (<4 years)



495

### Adulthood (>4 years)



496

497



## 498 Discussion

499 Our findings indicate that early life adversity shapes both early life survival and adult  
500 survival in free-ranging rhesus macaques. Individuals experiencing more cumulative  
501 early life adversity lived shorter lives than those with less adversity. The effect size of  
502 early life adversity on mortality risk was larger in the first four years of life than  
503 adulthood, but risks were also elevated in adulthood. Strong effects on early life  
504 mortality risk are consistent with the notion of an overall greater vulnerability during  
505 development (Walasek et al., 2022; West-Eberhard, 2003). Given the fitness costs of  
506 dying prior to reproduction, our results demonstrate that the effects of early life adversity  
507 prior to maturity have major fitness ramifications and the full consequences of early life  
508 adversity are likely to be larger than predicted in previous studies focused on adult  
509 fitness.

510

511 The various forms of potential adversity we measured did not contribute equally to  
512 survival with maternal-related adversities exhibiting the largest effects. Maternal death  
513 in the first four years of an individual's life and low matrilineal rank were associated with  
514 higher mortality risk in early life and adulthood. The lasting effects of these maternal-  
515 related adversities are unsurprising given similar consequences in other mammalian  
516 species (e.g., Stanton et al., 2020; Strauss et al., 2020; Tung et al., 2016; Zippel et al.,  
517 2019), as well as consequences of parental-related hardships in humans (Fields et al.,  
518 2021; Glover et al., 2018; O'Donnell et al., 2014; Reid et al., 2018; Thayer & Kuzawa,  
519 2014). Other maternal effects were also linked to mortality, but the effects were not as  
520 strong or straightforward. Survival advantages were observed among offspring born to

521 more socially connected mothers, but there was considerable variation in this effect and  
522 it did not persist into adulthood. Future work is needed to examine how an individual's  
523 own social connectedness in adulthood potentially interacts with the maternal social  
524 network they experienced during early life to shape survival. Small maternal kin  
525 networks, a proxy for low social support, predicted reduced survival, but effects were  
526 sex-dependent. From previous analyses, we know the presence of a competing  
527 younger sibling increases mortality risk during early life (Lee et al., 2019), and these  
528 effects appear to persist into adulthood, at least for females. Offspring born to first time  
529 mothers were more likely to die during early life than those who were not the first born.  
530 However, while our model estimates were uncertain, among those who survived into  
531 adulthood, survival odds were likely better for those born to primiparous than  
532 multiparous mothers. Given the strong negative effects of primiparity on offspring  
533 survival in the first four years of life, the higher survival odds of adults who were first  
534 borns could reflect survivorship bias. Alternatively, confounding variables that were not  
535 included in the analyses such as the presence of the grandmother, mother, and other  
536 kin across adulthood might also contribute to the patterns observed among adults.  
537  
538 Other forms of adversity, which were not directly related to the maternal environment,  
539 also contributed to survival. High temperatures generally predicted reduced survival, but  
540 this effect was sex-dependent. High temperatures are not typically included in measures  
541 of adversity for primates, but identifying potential health consequences and the  
542 underlying mechanisms is of interest given rising global temperatures (Hondula et al.,  
543 2015). While low rainfall seemed to have some negative effects on survival during early

544 life, we found no substantial relationship between rainfall and adult survival. High group  
545 size at birth was not associated with survival during early life, but seemed to have some  
546 negative effects on adult survival. Rainfall and group size might have limited effects on  
547 survival because drinking water and food are provisioned in the study population. While  
548 the macaques still compete over access to food and water resources, competition is  
549 likely reduced compared to wild populations.

550

551 Consistent with previous analyses in this population (Luevano et al., 2022; Morcillo et  
552 al., 2020), we did not find substantial impacts of early life hurricane exposure on  
553 survival. This is surprising given that macaques in this population exposed to Hurricane  
554 Maria showed divergent immune cell gene regulation, suggestive of accelerated aging  
555 (Watowich et al., 2022). Exposure to major hurricanes also led to greater heterogeneity  
556 in reproductive strategies and longevity, and macaques might reduce fertility as a  
557 strategy to prioritize survival odds (Diaz et al., 2023; Morcillo et al., 2020). Given our  
558 results showing heat effects on mortality and the fact that temperatures increase  
559 following hurricanes on Cayo Santiago (Testard et al., 2021; Watowich et al., 2022),  
560 hurricanes might affect macaques indirectly via factors such as deforestation, shade  
561 scarcity, and heat. Further, given the recency of Hurricane Maria and our small sample  
562 of individuals exposed to Maria in this study's dataset, we are currently limited in our  
563 ability to analyze survival outcomes for this most recent hurricane event. Potential  
564 impacts of Hurricane Maria may also have been socially buffered - macaques on Cayo  
565 Santiago adjusted their social networks after Hurricane Maria (Testard et al., 2021) and  
566 built new social connections, which may buffer negative impacts.

567

568 The survival effects of some forms of early life adversity were sex-dependent. Studies of  
569 early life adversity in long-lived animals have typically only been assessed in one sex,  
570 the non-dispersing sex (e.g., Gicquel et al., 2022; Patterson et al., 2022; Tung et al.,  
571 2016), but studies that have examined males and females produced mixed results.  
572 While male red deer are more negatively impacted by maternal death than female deer  
573 (Andres et al., 2013), there do not appear to be any sex-based differences in  
574 susceptibility to the survival costs of cumulative early life adversity in gorillas (Morrison  
575 et al., 2023), though males and females might vary in their responses to specific types  
576 of adversity. We predicted that early life adversity would have greater effects on  
577 mortality risk in males than females because of sex differences in life history strategies  
578 with males typically prioritizing more energetically intensive processes (Higham &  
579 Maestriperi, 2014; Hoffman et al., 2008; Schwartz & Kemnitz, 1992; Turcotte et al.,  
580 2022). We found some evidence in support of this prediction. During early life, male  
581 survival was more negatively affected by three forms of adversity: small maternal kin  
582 networks, high temperatures, and maternal loss. In adulthood, males continued to suffer  
583 greater costs of early maternal loss, perhaps reflecting the long-term costs of severe  
584 energetic constraints during early life. Males might be more affected by these  
585 adversities than females prior to reproductive maturity due to their energetically costly  
586 developmental trajectories and/or due to maternal decisions to reduce investment in  
587 energetically costly offspring during harsh environments (Clutton-Brock, 1994; Clutton-  
588 Brock et al., 1985; Trivers & Willard, 1973). Further research is needed to investigate

589 how effects of early life adversity might be moderated or mediated by developmental  
590 trajectories and parental investment strategies.

591

592 We also found evidence contrary to our predictions about sex-dependent effects: in  
593 adulthood, females were more affected by several forms of early life adversity than  
594 males. These findings are consistent with studies on later life mortality in humans which  
595 suggest that women tend to be more susceptible to early life adversity than men  
596 (Johnson et al., 2020; Lee & Ryff, 2019). Differences between men and women could  
597 be due to survivorship bias, underreporting of symptoms among men, confounds of  
598 societal gender biases and inequities, or due to underlying biology (Johnson et al.,  
599 2020). Given the complexity of potential factors contributing to effects in humans, it can  
600 be useful to turn to simpler animal models. Experimental work on mice also showed that  
601 early life adversity resulted in increased anxiety and depressive behaviors in female  
602 mice but not males (Bondar et al., 2018; Goodwill et al., 2019). Our results suggest  
603 species-specific social patterns and underlying biology might contribute to sex-  
604 dependent patterns associated with early life adversity. First, adult females were more  
605 affected by matriline rank than adult males likely because males disperse (Weiß et al.,  
606 2016), female dominance hierarchies are fairly stable across time (Blomquist et al.,  
607 2011), and females typically inherit dominance rank via their matriline. That this sex  
608 difference was already apparent in early life might reflect variation in social networks  
609 and social priorities between immature males and females (Amici et al., 2019). Second,  
610 being born into large maternal kin networks had a positive effect on adult female  
611 survival but a negative effect on adult male survival. Given dispersal, males might not

612 receive any immediate benefits of kin support in adulthood and thus only experience the  
613 long-term costs associated with earlier competition, consistent with the idea that  
614 individuals face tradeoffs between benefits of kin support and costs of competition with  
615 kin (Croft et al., 2017). Third, males were more susceptible than females to high  
616 temperatures during early life, but females were more susceptible in adulthood. Sex-  
617 dependent differences in sensitivity to heat are not well understood in humans, as  
618 findings have suggested both greater and lesser susceptibility to heat stroke in women  
619 versus men (Giersch et al., 2022). Experimental studies with mice showed females, but  
620 not males, exposed to exertional heat stroke exhibited delayed myocardial dysfunction,  
621 potentially influencing long-term cardiovascular health (Laitano et al., 2020). Future  
622 studies are needed to examine how body size, physiology, cardiovascular health, and  
623 energetic expense patterns are linked to temperature fluctuations, hurricane exposures,  
624 and mortality across ages in this population.

625

626 Our findings complement a body of work examining the effects of early life adversity in  
627 humans. Studies have shown that early life adversity is associated with poorer health  
628 and reduced longevity in humans (e.g., Barker et al., 2002; Deighton et al., 2018;  
629 Gluckman et al., 2008). However, this research faces challenges such as the  
630 prevalence of confounding variables like smoking and job insecurity, the difficulty of  
631 disentangling effects of different adverse experiences which tend to co-occur, and a  
632 reliance on retrospective surveys, which are prone to recall and reporting bias. Animal  
633 studies can help overcome these challenges because non-human species are  
634 characterized by simpler systems with fewer confounding variables and less clustering

635 of different types of adversity, and non-human animals tend to have shorter lifespans,  
636 which allows researchers to observe individuals from birth to death (Dettmer & Chusyd,  
637 2023; Patterson et al., 2023; Snyder-Mackler et al., 2020). Research on early life  
638 adversity in other species, especially natural populations, can shed light on the  
639 evolutionary pressures shaping early life sensitivities and provide the opportunity to  
640 disentangle confounding and correlated environmental factors. In our study, we  
641 measured ten forms of adversity and age at death in a large sample of males and  
642 females, and were able to show that the form of adversity, socio-sexual context, and  
643 other biological factors interact to shape the timing and severity of consequences.  
644 Natural populations of non-human animals can prove valuable not only for improving  
645 our understanding of the evolutionary pressures that shape developmental plasticity, life  
646 history strategies, and early life sensitivities, but also for better contextualizing findings  
647 in humans and informing future research in humans.

648

649 Our study demonstrates that exposure to early life adversity increases mortality risk in  
650 male and female rhesus macaques. Lower odds of surviving to reproductive age  
651 indicates that early life adversity can have major fitness ramifications for both an  
652 organism and their parents. Reduced life expectancy among those who survive to  
653 adulthood, suggests that early life adversity can have persisting fitness costs and long-  
654 term health consequences. Adversities related to the maternal social and nutritional  
655 environment generally had the largest impacts on offspring survival. We also found sex-  
656 dependent effects of early life adversity in this study, which are likely driven by social  
657 system characteristics (i.e., female philopatry) and sex-based variation in energetic

658 demands. Further research into the biological mechanisms underlying these survival  
659 patterns are needed to better understand how early life adversity impacts fitness and  
660 health. Importantly, variation in model estimates and predictions convey that while early  
661 life adversity can have negative consequences, such effects are not definitive. Social  
662 connections and behavioral adjustments (Campos & Fedigan, 2009; Testard et al.,  
663 2021) should be investigated as potential contributors to resilience.

664

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671

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1013

## 1014 Supplementary Materials

1015

### 1016 Social behavior

1017 Behavioral data were collected continuously using 10-min focal animal samples on  
1018 handheld computers (Brent, MacLarnon, et al., 2013). From 2010 to 2021, 17 observers  
1019 conducted focal samples on adults in 6 social groups (groups F, HH, KK, R, S, and V).  
1020 During focal samples, observers recorded activity state (e.g., resting, traveling, feeding)  
1021 and social interactions with other adults. For social interactions, observers recorded  
1022 type of behavior (e.g., grooming, vocal grunts, approaches, contact aggression, threats,  
1023 displacements), the identity of the partner, and whether the focal or partner initiated the  
1024 interaction. Behaviors were recorded as instantaneous point occurrences and durational  
1025 states. Focal animals were selected randomly and data were balanced such that  
1026 animals were sampled roughly equally across times of day and across the study period.

1027 Animals included in our analyses were focal followed for a mean of 5.72 hours per study  
 1028 year (range: 3.17-10.33 hours).

1029

1030 To measure maternal sociality, we calculated a composite sociality index (CSI) using  
 1031 the affiliative social behaviors, approaches and grooming. For each mother in each  
 1032 year, we tabulated the rate of approaches (count of approaches to and from other adult  
 1033 females divided by the number of hours the mother was focalled) and the rate of  
 1034 grooming bouts (count of grooming bouts given to and received from adult females  
 1035 divided by the number of hours the mother was focalled). We calculated the mean  
 1036 approach and grooming rates for all adult females in each social group in each year  
 1037 because these metrics vary across groups and across time. A mother's approach and  
 1038 grooming rates were divided by the mean rate for a given group-year. These  
 1039 standardized approach and grooming rates were added together and divided by 2 (the  
 1040 number of behaviors) to create the CSI for each mother.

1041

1042 Table S1. Correlations among model variables.

	Temperature	Rainfall	Kin network size	Maternal loss	Primiparity	Hurricane	Matriline rank	Competing sibling	Year
Group size	0.20	-0.04	0.06	0.00	-0.03	0.10	-0.10	0.01	0.29
Temperature		-0.06	0.06	0.02	-0.07	0.02	-0.03	-0.01	0.64
Rainfall			-0.05	0.04	-0.02	0.02	0.05	0.02	0.06
Kin network size				0.01	-0.20	0.08	-0.13	0.00	0.06
Maternal loss					-0.06	0.01	0.04	-0.05	0.05
Primiparity						0.00	-0.02	-0.17	-0.07
Hurricane							-0.05	-0.04	0.10
Matriline rank								0.02	-0.04
Competing sibling									0.14

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1045 Table S2. Models including and excluding pedigree produce similar predicted effects of  
 1046 early life adversity on survival  
 1047

Early life									
Model including pedigree					Model excluding pedigree				
	Estimate	Est.Error	Q2.5	Q97.5		Estimate	Est.Error	Q2.5	Q97.5
Intercept	2.67	0.28	2.19	3.27	Intercept	3.19	0.23	2.80	3.67
ELA Index (continuous, limited)	-0.33	0.12	-0.58	-0.10	ELA Index (continuous, limited)	-0.31	0.11	-0.53	-0.11
Pedigree Intercept	2.19	0.22	1.68	2.59					

Adults									
Model including pedigree					Model excluding pedigree				
	Estimate	Est.Error	Q2.5	Q97.5		Estimate	Est.Error	Q2.5	Q97.5
Intercept	2.63	0.04	2.56	2.72	Intercept	2.63	0.04	2.57	2.71
ELA Index	-0.12	0.04	-0.20	-0.03	ELA Index	-0.12	0.04	-0.21	-0.04
Pedigree Intercept	0.14	0.09	0.01	0.34					

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Table S3. Models of cumulative early life adversity on survival in early life and adulthood

Early life					Adults				
	Estimate	Est.Error	Q2.5	Q97.5		Estimate	Est.Error	Q2.5	Q97.5
Intercept	1.03	0.05	0.93	1.12	Intercept	2.77	0.02	2.73	2.80
Sex (male)	-0.05	0.10	-	0.15	Sex (male)	-0.13	0.02	-	-0.08
Cumulative early life adversity index	-0.29	0.07	-	-0.15	Cumulative early life adversity index	-0.04	0.02	-	-0.01
Sex*Cumulative early life adversity index	-0.07	0.09	-	0.12	Sex*Cumulative early life adversity index	0.00	0.02	-	0.05

Birth year intercept	3.37	0.33	2.78	4.04	Birth year intercept	0.03	0.02	0.00	0.07
Maternal ID intercept	0.78	0.09	0.60	0.96	Maternal ID intercept	0.07	0.03	0.01	0.13

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Table S4. Survival as a function of individual forms of early life adversity

	Early life				Adults				
	Estimate	Est.Error	Q2.5	Q97.5	Estimate	Est.Error	Q2.5	Q97.5	
Intercept	1.30	0.13	1.10	1.60	Intercept	2.90	0.04	2.80	3.00
Sex (male)	-0.24	0.17	-0.58	0.11	Sex (male)	-0.17	0.05	-0.27	-0.06
Maternal loss	-0.33	0.14	-0.61	-0.06	Maternal loss	-0.06	0.04	-0.14	0.02
High temperatures	-0.06	0.24	-0.52	0.42	High temperatures	-0.02	0.02	-0.07	0.02
Rainfall	0.23	0.18	-0.12	0.59	Rainfall	-0.01	0.02	-0.04	0.03
Group Size	0.01	0.07	-0.13	0.15	Group Size	-0.02	0.02	-0.06	0.02
Kin network	0.05	0.07	-0.09	0.20	Kin network	0.01	0.02	-0.02	0.05
Matriline rank (mid)	-0.09	0.17	-0.44	0.24	Matriline rank (mid)	-0.12	0.05	-0.21	-0.03
Matriline rank (low)	-0.44	0.17	-0.78	-0.13	Matriline rank (low)	-0.16	0.05	-0.25	-0.07
Primiparous	-0.36	0.16	-0.66	-0.05	Primiparous	0.06	0.05	-0.03	0.16
Competing sibling					Competing sibling	-0.03	0.04	-0.11	0.05
Hurricane Georges	-0.04	0.50	-1.00	0.94	Hurricane Georges	-0.02	0.10	-0.21	0.18
Hurricane Hugo	-0.08	0.50	-1.00	0.91	Hurricane Hugo	-0.04	0.10	-0.22	0.17
Hurricane Maria	0.08	0.44	-0.75	0.93	Hurricane Maria				
Sex*Maternal loss	-0.32	0.20	-0.71	0.06	Sex*Maternal loss	-0.05	0.06	-0.15	0.07
Sex*High temperatures	-0.23	0.09	-0.41	-0.05	Sex*High temperatures	0.02	0.03	-0.04	0.08
Sex*Rainfall	-0.04	0.09	-0.23	0.13	Sex*Rainfall	-0.01	0.03	-0.06	0.04
Sex*Group Size	-0.03	0.10	-0.23	0.17	Sex*Group Size	0.01	0.03	-0.04	0.06
Sex*Kin network	0.14	0.10	-0.05	0.32	Sex*Kin network	-0.05	0.03	-0.10	0.00



Sex*Matriline rank (mid)	0.24	0.21	-0.17	0.67	Sex*Matriline rank (mid)	0.06	0.06	-0.07	0.19
Sex*Matriline rank (low)	0.47	0.22	0.04	0.90	Sex*Matriline rank (low)	0.04	0.06	-0.09	0.15
Sex*Primiparous	0.09	0.22	-0.34	0.52	Sex*Primiparous	-0.04	0.07	-0.16	0.10
Sex*Competing sibling					Sex*Competing sibling	0.06	0.06	-0.06	0.18
Sex*Hurricane Georges	-0.05	0.37	-0.78	0.68	Sex*Hurricane Georges	0.07	0.14	-0.21	0.35
Sex*Hurricane Hugo	-0.15	0.40	-0.95	0.63	Sex*Hurricane Hugo	0.08	0.14	-0.20	0.36
Sex*Hurricane Maria	0.00	0.42	-0.85	0.85	Sex*Hurricane Maria				
Birth year intercept	3.40	0.34	2.80	4.10	Birth year intercept	0.03	0.02	0.00	0.08
Maternal ID intercept	0.78	0.09	0.60	0.95	Maternal ID intercept	0.05	0.03	0.00	0.12

1053 **Note** when matrilineal rank is coded as an ordinal variable, similar results are found:

1054 Early life: Rank =  $\beta = -0.28 \pm 0.09$ ; Rank\*Sex =  $\beta = 0.31 \pm 0.12$

1055 Adulthood: Rank =  $\beta = -0.08 \pm 0.02$ ; Rank\*Sex =  $\beta = 0.02 \pm 0.03$

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Table S5. Maternal social connectedness and offspring survival in early life and adulthood

	Early life				Adults				
	Estimate	Est. Error	Q2.5	Q97.5	Estimate	Est. Error	Q2.5	Q97.5	
Intercept	0.93	0.16	0.62	1.25	Intercept	2.38	0.25	1.99	2.97
Sex (male)	0.22	0.24	-0.26	0.70	Sex (male)	0.15	0.19	-0.21	0.56
Maternal social connectedness	0.15	0.17	-0.19	0.49	Maternal social connectedness	-0.02	0.12	-0.26	0.24
Sex*Maternal social connectedness	0.29	0.25	-0.19	0.81	Sex*Maternal social connectedness	-0.00	0.19	-0.40	0.38
Birth year intercept	2.18	0.72	1.21	3.94	Birth year intercept	0.20	0.17	0.01	0.63
Maternal ID intercept	0.66	0.34	0.04	1.30	Maternal ID intercept	0.20	0.16	0.01	0.59

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Table S6. Model comparisons using the loo\_compare function in the “brms” R package

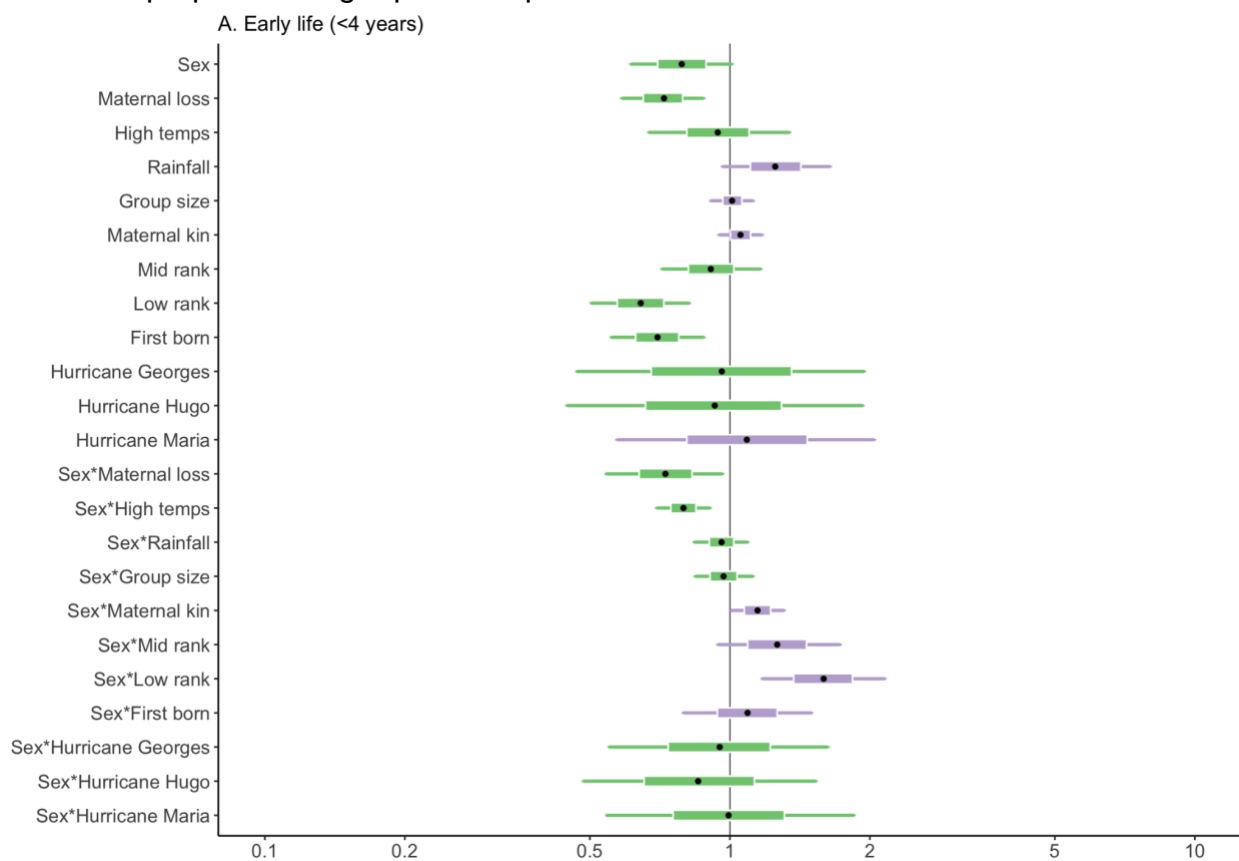
Early life survival			Adult survival		
Model	elpd_diff	se_diff	Model	elpd_diff	se_diff

Index	0.0	0.0	Index	0.0	0.0
Multivariate	-1.0	4.4	Multivariate	-3.3	5.9

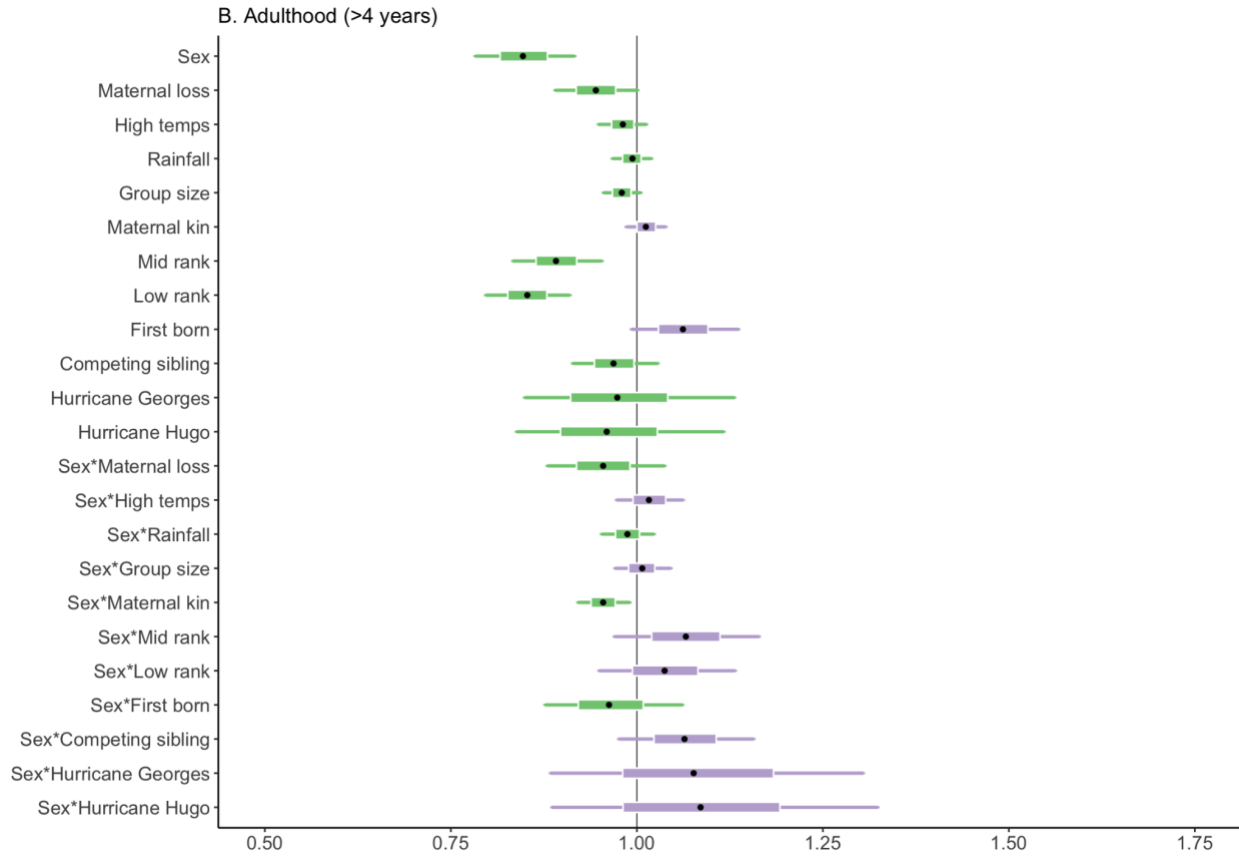
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1062 Figure S1. Model effects of sex and the forms of early life adversity on survival during  
 1063 early life (A) and adulthood (B). The outer bars show the 85% credible intervals, the  
 1064 inner boxes show the 50% credible intervals, and the black circles in the middle show  
 1065 the medians of the posterior distributions. Green shading represents negative effect  
 1066 sizes and purple shading represents positive effect sizes.



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