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# State-Level Anti-Immigrant Sentiment and Policies and Health Risks in US Latino Children

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## **Abstract**

**Background:** Although systemic inequities, broadly defined, are associated with health disparities in adults, there is a dearth of research linking contextual measures of exclusionary policies or prejudicial attitudes to health impairments in children, particularly among Latino populations. In this study, we examined a composite measure of systemic inequities in relation to the co-occurrence of multiple health problems in Latino children in the U.S.

**Methods:** Participants included 17,855 Latino children ages 3 to 17 years from the National Survey of Children's Health (2016–2020). We measured state-level systemic inequities using a factor score that combined an index of exclusionary state policies towards immigrants and aggregated survey data on prejudicial attitudes towards immigrants and Latino individuals. Caregivers reported on three categories of child health problems: common health difficulties in

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Dr. Slopen conceptualized the study, conducted the data analysis, interpreted the results, wrote the first draft of the manuscript, and integrated the critical contributions of all co-authors.

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All authors approved the final manuscript as submitted and agree to be accountable for all aspects of the work.

the past year, current chronic physical health conditions, and current mental health conditions. For each category, we constructed a variable reflecting zero, one, or two or more conditions.

**Results:** In models adjusted for sociodemographic covariates, interpersonal discrimination, and state-level income inequality, systemic inequities were associated with 1.13 times the odds of a chronic physical health condition (95% confidence interval (CI): 1.02, 1.25) and 1.24 times the odds of two or more mental health conditions (95% CI: 1.06, 1.45).

**Conclusions:** Latino children residing in states with higher levels of systemic inequity are more likely to experience mental health or chronic physical health conditions relative to those in states with lower levels of systemic inequity.

## **Table of Contents Summary:**

Systemic inequities, assessed at the state-level, are associated with physical and mental health conditions among Latino children.

#### Introduction

Latino/Latina/Hispanic children (referred to as "Latino children" herein) comprise approximately one-quarter of all children under the age of 18 in the U.S.¹ Latino children fare worse than non-Latino White children (referred to as "White children" herein) across several common health conditions, including respiratory illnesses, <sup>2, 3</sup> overweight and obesity, <sup>4, 5</sup> insufficient sleep, <sup>6–8</sup> and heightened levels of inflammation. <sup>9</sup> Systemic inequities related to structural or cultural racism, discrimination, xenophobia, and stigma (referred to as "systemic inequities" herein) limit access to power and resources for members of marginalized or stigmatized groups <sup>10–16</sup> and may contribute to these disparities. <sup>17, 18</sup> Such inequities can take many forms, including pervasive prejudicial attitudes and rhetoric directed toward racial or ethnic minorities, exclusionary laws designed to exclude individuals from various activities within society (e.g., related to education, health care, etc.), and criminalizing immigration policies. <sup>13</sup>

Consistent with research on the effects of discrimination on health in adults. 10, 19 most research on child health has focused on individual experiences of discrimination, <sup>20</sup> despite repeated calls to assess systemic or structural influences. 10, 21-25 Moreover, the majority of studies that have examined state-level measures of systemic inequities (e.g., anti-immigrant policies) in relation to Latino health have focused on adults<sup>26–28</sup> or perinatal outcomes,<sup>29–33</sup> with less known about children and youth. Prior studies have shown that harsh immigrant policies are associated with poor mental health<sup>26</sup> and reduced use of preventive health care and public assistance among Latino adults, 34, 35 as well as food insecurity among Latino immigrant families. <sup>33, 36</sup> Furthermore, restrictive immigration laws, <sup>30, 32</sup> sociopolitical events with relevance to immigration policies (e.g., the 2016 U.S. election), <sup>37, 38</sup> and enforcement actions<sup>39</sup> have each been linked to adverse birth outcomes among children born to Latina mothers. Although research on state-level systemic inequities and child and adolescent health outcomes is sparse, recent evidence found that Latino adolescents in states with greater systemic inequities (measured via a composite index of state immigration policies and aggregate social attitudes towards immigrant and Latino populations) had smaller hippocampal volumes, a brain region associated with chronic stress exposure. 40 This

work is complemented by research on adolescent responses to immigration actions,  $^{41, 42}$  including a study documenting elevated worry and behavioral withdrawal among Latino adolescents with vulnerable versus more secure family immigrant statuses.  $^{41}$ 

Building on this nascent literature, the present study examined associations between state-level, systemic inequities and the number of reported health problems in Latino children. Co-occurrence of problems is an understudied aspect of child health, <sup>43–46</sup> with potential implications for costs and quality of life for parents and children, health and earnings across the life course, as well as approaches to prevention. Systemic inequities may affect a broad range of children's health conditions—and the clustering of conditions—via chronic stress and associated disruptions to the child's stress-response system, <sup>47, 48</sup> or through the deprivation of resources needed to support healthy development (e.g., lack of access to safe and secure neighborhoods and schools or affordable, nourishing food). <sup>49</sup> These postulated pathways of chronic stress and deprivation of resources suggest potential shared mechanisms for mental and physical health problems among children.

Based on prior studies, <sup>26–28, 40</sup> we hypothesized that greater systemic inequities would be associated with increased reported health problems among Latino children. We include both U.S.-born and foreign-born Latino children in our study based on: (a) quantitative research showing that restrictive immigration policies are associated with poor mental health among Latino adults<sup>26</sup> and birth outcomes<sup>32</sup> regardless of personal immigration history, and (b) qualitative research demonstrating few differences between US-born and foreign-born Latino adults in perceptions of vulnerability or psychological distress related to immigration enforcement activities. <sup>50</sup> These findings are likely due to several factors, including that 40 percent of Latino adults live in households with mixed immigration status, <sup>51</sup> thus making immigration-related policies salient to a large portion of Latino families.

## **Methods**

#### Sample

We used data from 17,855 Latino children from the National Survey of Children's Health (NSCH) (2016–2020), a cross-sectional, nationally representative, weighted probability sample of non-institutionalized children from birth through age 17. Each year, randomly selected households across the U.S. are mailed an invitation to complete a household screener and child-level questionnaire via a secured website or on paper. The paper and web instruments are available in both English and Spanish, and additional language support is available via telephone. Parents or guardians familiar with the child's health and medical care are the respondents. After completing the screener, a single child from each home is randomly selected as the focal child. Details on design, administration, and completion rates are available at <a href="https://www.childhealthdata.org/learn-about-the-nsch/NSCH">https://www.childhealthdata.org/learn-about-the-nsch/NSCH</a>. Our analytic sample included children ages 3 years of age and older given the health outcomes on which we focus. Children from Washington, DC—notably, considered to be a "sanctuary city" were excluded from our analytic sample because data needed to construct the state-level systemic inequities score were unavailable.

#### **Measures**

Child health.—We assessed the number of problems within and across three distinct dimensions of child health, following Jackson and colleagues' approach with the 2016 NSCH.<sup>53</sup> Caregiver respondents reported on: (a) health difficulties in the past twelve months (six items: eating or swallowing because of a health condition; digesting food, including stomach or intestinal problems, constipation, or diarrhea; repeated or chronic physical pain, including headaches or other back and body pain; toothaches; bleeding gums; and decayed teeth or cavities); (b) provider-diagnosed, current, chronic physical conditions (six items: allergies; asthma; blood disorders; diabetes; heart condition; and arthritis); and (c) providerdiagnosed, current mental health conditions (four items: depression; Attention-Deficit/ Hyperactivity Disorder (ADHD); anxiety problems; and behavioral or conduct problems). We examined the extent of problems within each health dimension by constructing a threelevel variable to reflect zero, one, or two or more. We also created a four-level variable to indicate the co-occurrence of problems across dimensions (i.e., no health problems and problems within one, two, or all three dimensions). These outcomes are designed to indicate the pervasiveness of problems within and across multiple dimensions of mental and physical health, for both chronic and temporary conditions. <sup>53</sup> As a secondary analysis, we examined outcomes individually as well.

**Systemic inequities.**—We operationalized systemic inequities via a factor score developed using data-driven methods and used in prior research.<sup>40</sup> We included measures of aggregated public attitudes and social policies—and refer to them together as "systemic"—as both policies and attitudes reflect the broader macro-social context, are highly correlated,<sup>54</sup> and are consistent with conceptual frameworks from minority stress theory and stigma research.<sup>55, 56</sup>

The factor score was based on three measures. First, a state-level summary index reflecting restrictiveness or supportiveness of state policies (related to health services, private sector employment, business licensing, rental housing access, higher education access, driver's license access, immigration policy enforcement, non-English language use, identification requirements, and discrimination prohibition) towards immigrants as of 2016, with a positive point awarded for each of the items, and a negative point awarded if the state explicitly prohibited the item.<sup>57</sup> Second, we used survey responses from the American National Election Study (ANES) to a "feelings thermometers" (i.e., a measurement technique where participants report their feelings towards a target on a scale ranging from 0 (extremely cold or negative feelings) to 100 (extremely warm or positive feelings)) reflecting attitudes towards Latino individuals (pooled, 1996 to 2016). Third, we used survey responses from the ANES on a "feelings thermometer" reflecting attitudes towards immigrants (pooled, 2004 to 2016). Responses on the ANES feelings thermometers were standardized for all respondents and then aggregated at the state-level. All three components were reverse scored so that higher ratings represented higher levels of structural inequity.

For the ANES feeling thermometer measures, survey years were pooled to maximize the number of respondents per state and to minimize measurement error; this approach is supported by research showing the stability of states relative to each other in terms of their

residents' attitudes towards marginalized groups (e.g., racial minorities and women) over 30 years. <sup>58, 59</sup> We included attitudes and policies related to immigration for this measure, despite the fact that only a third of Latino individuals in the U.S. are foreign-born, <sup>60</sup> because of the mixed status of many Latino households <sup>51</sup> and because non-Latino individuals in the U.S. often conflate immigrant identity with Latino identity. <sup>61</sup> See Appendix 2 for a table describing the component measures.

The model-based factor score was constructed for each state using exploratory factor analysis, with all three measures coded with higher values reflecting higher levels of systemic inequity. We have displayed the distribution of factor scores across states (see Figure 1). Appendix 3 presents the scores for each state. The continuous factor score ranged from -1.75 to +1.76, representing the state's relative standing on the latent factor of systemic inequities for Latino children, with higher values reflecting higher levels of systemic inequity.

**Covariates.**—Caregivers reported on children's ethnicity. We selected covariates to be consistent with prior research<sup>53</sup> and constructed both minimally and fully adjusted models recognizing that some of the covariates could be on the causal pathway. Our basic set of covariates included child's age and sex, survey year, family immigration history, mother's age at child's birth, and state-level Gini Index, to control for other macro-level characteristics related to income inequality (see Table 1 for variable categories). Our extended set of covariates additionally included highest education level in household, income-to-needs ratio (using the multiple imputed values provided by Census), caregiver report that the child ever resided in an unsafe neighborhood, health insurance status, caregiver self-rated health index (i.e., sum of single-item self-reports of physical and mental health), an index reflecting number of social services received (i.e., cash assistance, Special Supplemental Nutrition Program for Women, Infants, and Children (WIC), Supplemental Nutrition Assistance Program (SNAP), and free or reduced-price school lunch), and caregiver report of the child's personal experience of racism.

**Analysis**—First, we display the social, demographic, and health characteristics for each group. Second, we used generalized estimating equations (GEE)<sup>62</sup> and multinomial logistic regression to estimate the relative risk of exhibiting a single health problem and multiple problems within each of the three health dimensions, and across health dimensions, using systemic inequities as the independent variable. We selected multinomial regression because we conceptualized the co-occurrence of more than one condition as a discrete, qualitative outcome, rather than a count (consistent with prior research using a similar set of outcomes<sup>53</sup>). This approach accounts for the complex sample design and for correlations among children who reside in the same state. We present odds ratios, 95% confidence intervals (CIs), and *p*-values from Wald's tests, which describe the significance of the association between the exposure and multi-category outcome, collectively (i.e., both categories against zero).<sup>63</sup> As sensitivity analyses, we examined each health outcome individually and tested for effect modification by sex.

Descriptive statistics and models were generated using SUDAAN 11.0.3, and we weighted the results to represent non-institutionalized U.S. children. To handle missing data, analyses

used six imputed data sets. The results using the imputed data are nearly identical to those using complete case data.

## Results

## Demographic characteristics.

Nearly one-third of children (29.50%) lived in households below the federal poverty level, and over one in ten (10.56%) did not have health insurance (see Table 1). Slightly over half of the children (54.30%) were either born outside of the U.S. or had a parent born outside the U.S. Table 2 presents the distributions for number of problems within and across three distinct dimensions of child health. See Appendix 1 for prevalence of each specific health outcome.

#### Systemic inequities and number of health outcomes.

The multinomial models to estimate odds ratios for each health dimension, and across dimensions, were similar when adjusting for a basic set of covariates and when additionally adjusting for a more comprehensive set of social and demographic covariates (see Appendix 4 for values from the basic and fully-adjusted models). Figure 2 displays the odds ratios from the fully-adjusted models only.

In minimally adjusted models, systemic inequities were significantly associated with the number of chronic physical health conditions and number of mental health problems. Specifically, a one-unit increase in systemic inequities was associated increased odds of one (adjusted odds ratio (AOR)=1.13, 95% confidence interval (CI): 1.02, 1.25) or two or more (AOR=1.20, 95% CI: 1.00, 1.45) chronic physical health conditions, and two or more mental health conditions (AOR=1.27, 95% CI: 1.09, 1.48) (see Appendix 4). In models that additionally adjusted for demographic characteristics that could be on the causal pathway, systemic inequities were significantly associated with increased odds of having a chronic physical health condition (AOR=1.13, 95% CI: 1.02, 1.25) and increased odds of having two or more mental health conditions (AOR=1.24, 95% CI: 1.06, 1.45). No associations were evident for the outcome of health difficulties in the past 12 months. Although an increasing relationship between systemic inequities and the number of health dimensions with one or more health problems was suggested by the point estimates, with AORs increasing in magnitude as the number of dimensions with 1+ health problems increased, associations were not significant at *p*<.05.

In sensitivity analyses, we disaggregated our health categories to further examine the results (see Table 3): in fully-adjusted models, anxiety problems (AOR=1.24, 95% CI: 1.06, 1.46), ADHD (AOR: 1.20, 95%: 1.03, 1.40), and allergies (AOR=1.15, 95% CI: 1.04, 1.27) were each associated with systemic inequities at the p < .05 threshold. Finally, we did not find evidence for effect modification based on child's sex (p-values>0.05) across the four primary outcomes.

## **Discussion**

This study examined whether Latino children who reside in states with higher levels of systemic inequities experience a greater co-occurrence of health problems relative to children who live in states with lower levels of inequity. We used nationally representative data from the NSCH linked to a state-level measure of systemic inequities, generated from aggregated public opinion data about Latino groups and immigrant populations as well as both exclusionary and inclusive policies towards immigrants.

As hypothesized based on prior research, <sup>26–28, 32, 64, 65</sup> systemic inequities were associated with a greater co-occurrence of mental health conditions and the occurrence of chronic physical health conditions among Latino children, even after adjusting for a broad set of child and family characteristics and individual experiences of discrimination. Of note, while the observed associations are small in magnitude, research suggests that small effects can be meaningful when scaled across populations, as is the case in our measure of structural inequalities. <sup>66, 67</sup> For both chronic physical and mental health conditions, we observed a graded relationship where the estimated associations were larger as the number of health conditions increased. This pattern is consistent with conceptualizations of systemic inequities as a broad, generalizable, risk factor.

Our results reinforce and build on prior studies of personally-experienced racism and child health, <sup>20, 68–70</sup> as well as restrictive immigration policies and adult <sup>26–28</sup> and perinatal health, <sup>29, 30, 33, 39</sup> in several ways. First, we use a recent, large nationally representative sample of children, which improves the generalizability of our results and the ability to study variation in systemic inequities across states. Second, drawing on evidence that immigration policies and anti-immigrant sentiment are interconnected, <sup>54</sup> our measure of systemic inequities combines both aggregated social attitudes and policies, thereby improving construct validity. Third, our analyses consider the co-occurrence of health problems, an understudied aspect of child development that has relevance for health equity research, <sup>43</sup> which has rarely been studied in relation to structural contexts in childhood.

There are also limitations to consider in interpreting these study results. First, state-level analyses of systemic inequities are appropriate given the many important legislative activities at that level, but they offer a conservative test because more proximal environments are likely to exert stronger associations. Thus, more localized aspects of place-based inequities and protective factors that influence child health and development should be studied, 71–73 as there is often substantial heterogeneity within states in terms of social climates surrounding Latino populations (e.g., differential enforcement of immigration policies). Also related to our exposure measure, the index of state policies reflects policies in place in 2016, 57 and our measures of prejudicial attitudes pool across many years, up to 2016. Although our approach is supported by research showing stability in the rank ordering of state-level attitudes towards marginalized groups, 58, 59 future studies might benefit from examining time-varying measures of systemic inequities. We also recognize that, despite the aforementioned strengths of using a factor score, one of the limitations of this approach is that there is not a direct interpretation of a one-unit change in this continuous measure because it combines interrelated components of systemic inequities.

Second, the NSCH has several limitations, such as reliance on caregiver report of providerdiagnosed mental health conditions; given the disparities in specialized mental health services for Latino children, <sup>74, 75</sup> these outcomes are likely to be underestimates. Related, the survey was administered via mail and online, which could exclude families without permanent mailing addresses or reliable internet access, which represent some of the highest risk populations. Furthermore, our analysis is not inclusive of all relevant child outcomes (e.g., we could not include provider-diagnosed overweight or obesity since it was not asked in 2016 and 2017); and, our ability to explore potential within-group interactions by country of origin, age, geography, duration of time in the United States among the subset of children born outside of the US, and other child and family characteristics (e.g., interpersonal experiences of racism) was limited by insufficient sample sizes. We also were unable to account for how long a child lived in a state at the time of the survey, which could introduce measurement error, and our cross-sectional design prohibits causal inferences and examination of both unique and shared mechanisms, all of which represent important areas for future investigation. Finally, while interpreting these results it is important to keep in mind that odds ratios are overestimates of risk for common outcomes (i.e., >10 percent<sup>76</sup>).

## **Implications**

Childhood health provides a foundation for well-being across the life course, including promotion of school attendance and performance, 77–79 reducing risk for substance abuse. 80, 81 and positive health and socioeconomic attainment in adulthood. 82-84 Accordingly, our results and related studies have implications for a wide range of healthpromoting policies, particularly in the face of persistent structural inequities related to racism, xenophobia, and punitive approaches to immigration. Policy statements from the American Academy of Pediatrics and other reports have called on pediatricians to play a more active role in educating the public about the adverse effects of systemic racism experienced by children of color and immigrant families. 17, 85–88 This study underscores the importance of addressing the health impacts of state laws as well as the effects of public attitudes that perpetuate racist and/or anti-immigrant sentiments, all of which influence access to opportunities and resources that promote healthy development.<sup>89</sup> Previous research has demonstrated that inclusive immigrant policies can be protective for educational attainment, 90, 91 labor market outcomes, 92 and other measures of socioeconomic wellbeing, 93 which directly influence the resources available to minoritized children. Although studies of the potential benefits of inclusive immigrant policies for child health outcomes are limited, one quasi-experimental study of children whose mothers received protection from deportation via the United States' Deferred Action for Childhood Arrivals program (i.e., determined based on their birth date) reported 50 percent fewer diagnoses of anxiety and adjustment disorders compared to children who did not receive this protection.<sup>94</sup> Further investigation is needed to identify policies, administrative practices, and localized programs that are most effective in advancing health equity. In addition, pediatricians working with Latino children and children in immigrant families should be cognizant of major changes to immigrant-related policies or highly visible discriminatory events and can advocate for strategies to minimize structural or cultural racism, including the removal of exclusionary policies.

## Conclusion

This study begins to address significant gaps in the empirical literature on the harmful consequences of discriminatory policies and prejudicial social contexts on children's health. Beyond the need for a strong pediatric voice in educating policymakers and the general public about this threat to child wellbeing, a deeper understanding of the causal mechanisms that explain these findings is essential for moving beyond documenting the consequences of structural inequities and towards accelerating the development of more effective strategies to prevent, reduce, and/or mitigate their harmful effects.

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## **Appendix**

## **Appendix**

Appendix 1.

Specific health outcomes in children (*n*=17,855)

	%	(S.E.)
Health difficulties, past 12 months		
Eating or swallowing problems	1.41	(0.18)
Digesting food, including stomach or intestinal problems	8.74	(0.45)
Chronic or repeated physical pain	8.63	(0.46)
Toothaches	4.75	(0.36)
Bleeding gums	3.11	(0.28)
Decayed teeth or cavities	15.53	(0.63)
Any dental problem	18.28	(0.67)
Current chronic health conditions		
Allergies	16.93	(0.57)
Asthma	8.46	(0.44)
Blood disorders	0.21	(0.07)
Diabetes	0.26	(0.05)
Heart conditions	0.94	(0.11)
Arthritis	0.12	(0.02)
Current mental health conditions		
Depression	2.93	(0.24)
Anxiety problems	6.53	(0.35)
ADHD	6.75	(0.37)
Behavioral or conduct problems	6.03	(0.35)

## Appendix 2.

## Structural inequity related to Latino ethnicity

Component	Source	Year(s) of data
Illegal immigrants feelings thermometer $^{I}$	American National Election Survey	2004, 2008, 2012, 2016
Hispanics feelings thermometer $^{I}$	American National Election Survey	1996, 2000, 2002, 2004, 2008, 2012, 2016
Index for the following 10 policies: Access to health services for immigrants; cooperation by state or local law enforcement with federal immigration enforcement; use of non-English language; immigrant employment in a broad segment of the private sector; restricts immigrant access to a broad class of business licenses; access to rental housing for immigrants; prohibits discrimination based on citizenship or immigration status; provides identification to access relevant services or opportunities; access to higher education for immigrants; access to driver's licenses for immigrants <sup>2</sup>	Wake Forest Immigration Law Coding Project	2016

Note: All components were reverse scored so that higher ratings represented higher levels of structural inequity.

## Appendix 3.

## Systemic inequity factor score for each state

Alabama	1.192609415
Alaska	1.762166823
Arizona	0.09262617
Arkansas	0.362811591
California	-1.746000735
Colorado	-0.764800338
Connecticut	-0.594610506
Delaware	-1.131580132
Florida	-0.534851602
Georgia	0.674533941
Hawaii	-0.053595288
Idaho	-0.68201278
Illinois	0.379332559
Indiana	0.647283447
Iowa	0.489941806
Kansas	-0.052408919
Kentucky	0.103890173
Louisiana	-0.090151361
Maine	-0.594149601
Maryland	-0.160155081
Massachusetts	-0.122937467

<sup>&</sup>lt;sup>1</sup>A "feelings thermometer" is a visual analog scale; respondents are asked to rank their views of the subject on a scale from "cold," which indicates cold or negative feelings, to "hot" which indicates warm, or positive feelings, scored on a range from 0 to 100. National scores were standardized and then aggregated at the state-level.

<sup>&</sup>lt;sup>2</sup>Score was computed by awarding a positive point for the presence of a given item; a negative point was awarded if there was a law to explicitly prohibit the item.

Michigan	-0.277416171
Minnesota	-0.253823636
Mississippi	0.585269372
Missouri	0.881863619
Montana	-0.002640792
Nebraska	1.069464245
Nevada	0.247743259
New Hampshire	-0.160049548
New jersey	-1.1168013
New Mexico	-1.196954002
New York	0.340094743
North Carolina	0.54052828
North Dakota	0.312859548
Ohio	0.446518179
Oklahoma	0.127157003
Oregon	-0.624085962
Pennsylvania	0.39290517
Rhode island	-0.637706676
South Carolina	0.316076419
South Dakota	0.509147968
Tennessee	0.698644606
Texas	-1.111367825
Utah	-0.264389678
Vermont	-0.492881888
Virginia	-0.454053702
Washington	-0.797251392
West Virginia	0.638810582
Wisconsin	0.383011156
Wyoming	0.721386309

Note: Darker shade reflects higher values of the state-level systemic inequity score

## Appendix 4.

Multinomial models estimating associations between systemic inequities and health problems in Latino children (n=17,855) ages 3-17 years

	Model 1		Model 2	
	AOR	(95% CI)	AOR	(95% CI)
Health difficulties				
0	1.00		1.00	
1	0.95	(0.85, 1.06)	0.94	(0.84, 1.05)
2+	1.00	(0.88, 1.15)	0.99	(0.86, 1.13)
Wald p-value	0.67		0.55	
Chronic physical health conditions				

	Model 1		Model 2	
	AOR	(95% CI)	AOR	(95% CI)
0	1.00		1.00	
1	1.13	(1.02, 1.25)	1.13	(1.02, 1.24)
2+	1.20	(1.00, 1.45)	1.20	(0.99, 1.44)
Wald p-value	0.02		0.02	
Mental health conditions				
0	1.00		1.00	
1	1.08	(0.92, 1.26)	1.07	(0.92, 1.24)
2+	1.27	(1.09, 1.48)	1.24	(1.06, 1.45)
Wald p-value	0.01		0.03	
# of health dimensions with 1+ condition				
0	1.00		1.00	
1	1.04	(0.95, 1.14)	1.04	(0.95, 1.13)
2	1.12	(0.98, 1.28)	1.11	(0.98, 1.27)
3	1.19	(0.97, 1.47)	1.15	(0.93, 1.42)
Wald p-value	0.18		0.32	

Notes: AOR=adjusted odds ratios. Model 1 is adjusted for child's age, sex, survey year, family immigration history, mother's age at child's birth, and state income inequality. Model 2 is adjusted for covariates in Model 1, in addition to personal experience of racism, household income, highest education in household, social service use index, neighborhood safety, insurance status, caregiver health index, family structure, and state income inequality. OR=odds ratio; CI=confidence interval. Values for Model 2 correspond to the values presented in Figure 2.

## **Abbreviations:**

**NSCH** National Survey of Children's Health

**ADHD** attention deficit hyperactivity activity disorder

**ANES** American National Election Study

**OR** odds ratio

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## What's Known on This Subject

Prior studies have documented that systemic inequities, including harsh immigrant policies, are associated with poor mental health among Latino adults and adverse birth outcomes for Latino infants. Few studies have examined state-level measures of systemic inequities and children's health.

## What This Study Adds

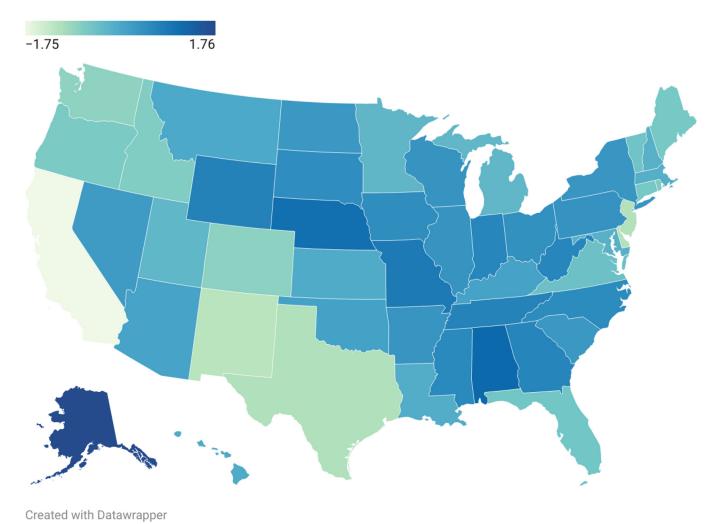
Systemic inequities—including exclusionary state policies and prejudicial attitudes—are associated with the occurrence of multiple physical and mental health conditions among Latino children over and above individual and family characteristics, highlighting the importance of considering macro-level social determinants of child health.

## What's Known on This Subject

Systemic inequities, including harsh immigration policies, are associated with poor mental health among Latino adults and adverse birth outcomes for Latino infants. Yet, few studies have examined associations between state-level measures of systemic inequities and children's health.

## What This Study Adds

Systemic inequities—including exclusionary state policies and prejudicial attitudes—are associated with the occurrence of physical and mental health conditions among Latino children over and above individual and family characteristics, highlighting the importance of considering macro-level social determinants of child health.



**Figure 1.**Distribution of systemic inequities for Latino populations across the United States **Note**: Darker shade reflects higher values of the state-level systemic inequity score.

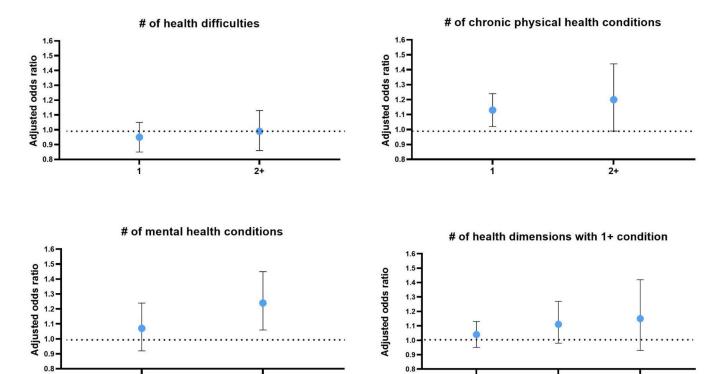


Figure 2. Adjusted odds ratios to display the relationship between state-level systemic inequities and health problems among Latino children (n=17,855) 3–17 years of age

**Notes:** Models are adjusted for child's age, sex, personal experience of racism, survey year, household income, highest education in household, family immigration history, social service use index, mother's age at child's birth, neighborhood safety, insurance status, caregiver health, family structure, and state income inequality. See Appendix 4 for numeric values corresponding to estimates presented in the figure and Wald p-values.

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Table 1.

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Descriptive characteristics of the sample (n=17,855)

	%	(S.E.)
Child sex (male)	50.90	(0.83)
Age cohort		
3–5 years	18.84	(0.66)
6–11 years	40.38	(0.83)
12–17 years	40.77	(0.81)
No health insurance	10.56	(0.56)
Household income-to-needs ratio		
Less than FPL	29.50	(0.91)
100 – 399% FPL	45.66	(0.87)
400% FPL	24.84	(0.72)
Highest household education		
Less than high school	23.74	(0.83)
High school	26.93	(0.74)
More than high school	49.33	(0.83)
Family Structure		
Two adults, married	58.35	(0.84)
Two adults, unmarried	13.19	(0.61)
Single parent	22.71	(0.69)
Grandparent or Other	5.75	(0.43)
Family Immigration History		
Child or parent born out of US	54.30	(0.83)
Child born in US, parent data missing	8.46	(0.53)
Parent and child born in US	37.24	(0.78)
Count of social services received		
0	39.25	(0.80)
1	31.46	(0.81)
2	19.03	(0.75)
3	8.68	(0.52)
4	1.58	(0.23)
Caregiver health index		
Excellent (score=2–3)	36.42	(0.81)
Good (score=4-6)	55.23	(0.84)
Fair/poor (score=7–10)	8.35	(0.46)
Parent-report, unsafe neighborhood	7.74	(0.45)
Parent reported unfair treatment of child due to race or ethnicity	5.66	(0.34)

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Table 2.

Health characteristics of the sample (n=17,855)

	%	(S.E.)
Health difficulties		
0	70.76	(0.77)
1	19.76	(0.69)
2+	9.48	(0.48)
Chronic physical health conditions		
0	78.22	(0.64)
1	16.86	(0.58)
2+	4.92	(0.32)
Mental health conditions		
0	86.42	(0.52)
1	7.73	(0.44)
2+	5.85	(0.32)
# of health dimensions with 1+ condition		
0	52.89	(0.83)
1	32.42	(0.79)
2	11.89	(0.51)
3	2.80	(0.23)

Note: Health difficulties (past 12 months) include problems (1) eating or swallowing; (2) digesting food, including stomach/intestinal problems; (3) repeated or chronic physical pain; (4) toothaches; (5) bleeding gums; and (6) cavities. Chronic health problems (current) include caregiver report of health care provider's diagnosis of: (1) allergies; (2) asthma; (3) blood disorder; (4) diabetes; (5) heart condition; and (6) arthritis. Mental health disorders (current) include health care provider or educator report of (1) depression; (2) Attention Deficit Hyperactivity Disorder; (3) anxiety problems; and (4) behavioral or conduct problems. Number of health dimensions with 1+ condition is a count of the health outcome categories in which a child had one or more health conditions (range: 0 to 3).

Table 3.

Adjusted odds ratios to describe the relationship between the systemic inequities score and each outcome among Latino children, ages 3 to 17 years

	Systemic Inequities	
	AOR	95% CI
Health difficulties, past 12 months		
Eating or swallowing problems	1.18	(0.84, 1.65)
Digesting food, including stomach or intestinal problems	1.04	(0.90, 1.20)
Chronic or repeated physical pain	0.96	(0.83, 1.11)
Toothaches	1.01	(0.83, 1.23)
Bleeding gums	0.98	(0.79, 1.23)
Decayed teeth or cavities	0.93	(0.82, 1.04)
Any dental problem	0.92	(0.82, 1.03)
Chronic health conditions		
Allergies	1.15	(1.04, 1.27)
Asthma	1.11	(0.97, 1.28)
Mental health conditions		
Depression	1.05	(0.82, 1.35)
Anxiety problems	1.24	(1.06, 1.46)
ADHD	1.20	(1.03, 1.40)
Behavioral or conduct problems	1.14	(0.96, 1.35)

Notes: AOR = Adjusted odds ratios. All models are adjusted for child's age, sex, survey year, household income, highest education in household, family immigration history, social service use index, mother's age at child's birth, neighborhood safety, insurance status, caregiver health, family structure, personal experience of racism, and state income inequality. We did not examine uncommon chronic physical health conditions as separate outcomes (i.e., blood disorders, diabetes, heart conditions, arthritis) due to data limitations resulting from the low prevalence for these conditions.