

Montelukast in paediatric asthma and allergic rhinitis: a systematic review and meta-analysis

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Shareable abstract (@ERSpublications)

This is the first systematic review focused on the impact of montelukast on children and adolescents with asthma, from the patients' perspective, measured with patient-reported outcome measures (PROMs) https://bit.ly/3YwqfBb

Cite this article as: Mayoral K, Lizano-Barrantes C, Zamora V, et al. Montelukast in paediatric asthma and allergic rhinitis: a systematic review and meta-analysis. Eur Respir Rev 2023; 32: 230124 [DOI: 10.1183/16000617.0124-2023].

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Received: 21 June 2023 Accepted: 10 Aug 2023

Abstract

Background: We aim to assess the impact of montelukast on paediatric patients with asthma/allergic rhinitis, measured using patient-reported outcome measures, compared with other treatments or placebo. **Methods:** Protocol registration CRD42020216098 (www.crd.york.ac.uk/PROSPERO). MEDLINE and Embase databases were used to conduct the search. Two authors independently selected studies and extracted data, and a third reviewer resolved discrepancies. Meta-analyses were constructed to estimate the standardised mean difference (SMD) using a random-effects model.

Results: Out of 3937 articles identified, 49 studies met the inclusion criteria, mostly randomised clinical trials (sample sizes: 21–689 patients). The SMD of change pooled estimators for the global, mental and physical domains of health-related quality of life were not statistically significant. For daytime and night-time symptoms scores, the SMD (95% CI) was in favour of inhaled corticosteroids (-0.12, -0.20–-0.05 and -0.23, -0.41–-0.06, respectively). The pooled estimator for global asthma symptoms was better for montelukast when compared with placebo (0.90, 0.44–1.36).

Conclusions: The synthesis of the available evidence suggests that, in children and adolescents, montelukast was effective in controlling asthma symptoms when compared with placebo, but inhaled corticosteroids were superior in controlling symptoms, especially at night-time. These findings of our systematic review concur with current guidelines for asthma treatment.

Introduction



Asthma and allergic rhinitis affect >300 million people globally [1, 2], comprising 10–30% of all adults and up to 40% of children. Montelukast, a leukotriene receptor antagonist (LTRA), was approved by the United States Food and Drug Administration (FDA) in 1998 for the treatment of asthma and in 2002 for

allergic rhinitis [3]. Currently, asthma treatment guidelines [4, 5] include LTRAs in the category of other controller options (in monotherapy or in combination with inhaled corticosteroids (ICS)), highlighting their limited indications and the lack of evidence for efficacy or safety. The guideline for allergic rhinitis [6] limits montelukast only as a treatment option for patients who are not effectively treated with alternative therapies.

In 2009, the FDA requested a review of the clinical trials with montelukast performed by Merck KGaA, whose results indicated that suicidality and behaviour-related adverse experiences were infrequent and similar to those seen in control subjects [7, 8]. Nonetheless, several cases were notified between 2014 and 2018 as suspected adverse neuropsychiatric events [9]. In 2019, a European Medicines Agency review [10] identified some cases in which there had been a delay in recognising neuropsychiatric events as possible adverse drug reactions, and the United Kingdom Medicines and Healthcare Products Regulatory Agency activated a warning including nightmares, depression, insomnia, aggression, anxiety and abnormal or changed behaviour. In 2020, the FDA required a boxed warning advising healthcare providers to avoid prescribing montelukast for patients with mild symptoms [11]. The majority of systematic reviews evaluating montelukast adverse events were performed before the release of the FDA's warning [12–14]. To the best of our knowledge, after the release there was only one systematic review aimed at identifying montelukast adverse drug reactions reported in the medical literature [15], which found a wide range of suspected reactions, predominantly gastrointestinal and neuropsychiatric, and a study reviewing reports of suspected adverse drug reactions to montelukast in the World Health Organization's database (VigiBase) [16], which found 1118 reports of nightmares; two-thirds of which concerned children aged 5–10 years.

The impact of montelukast adverse reactions could be captured through patient-reported outcomes measures (PROMs), which cover a variety of constructs from the patient's perspective, including symptoms and health-related quality of life (HRQoL) among others [17]. Asthma and rhinitis guidelines recommend the use of PROMs to monitor symptom control and to measure the impact of the disease on daily activities and HRQoL [4, 5, 18, 19].

We found five systematic reviews evaluating the efficacy of montelukast or LTRA specifically on paediatric asthma patients [13, 20–23], confirming the superiority of ICS in symptom control [20, 21, 24], symptom scores [13], symptom-free days [24], night-time awakening [24] and HRQoL [24]. Other systematic reviews compared LTRA with placebo [12], or as an addition to usual care [23]. PROMs were secondary end-points in most of these systematic reviews and no meta-analyses were constructed with them. We identified three systematic reviews including adolescents and adults with allergic rhinitis, which compared montelukast with placebo, antihistamines and intranasal corticosteroid [25–27], all considering PROMs as primary and secondary outcomes (symptoms and HRQoL, respectively), but none included a meta-analysis with paediatric patients.

There are few systematic reviews in children and adolescents evaluating montelukast with PROMs, especially for HRQoL in asthma patients. Our primary aim was to evaluate the impact of montelukast on paediatric patients with asthma and/or allergic rhinitis, measured with PROMs, when compared with other treatments or with placebo, through a systematic review and meta-analysis. In addition, we considered a synthesis of the evidence about adverse drug events and withdrawals in the selected studies.

Methods

Protocol and registration

We conducted a systematic review of the literature on the impact of montelukast in children and adolescents with asthma and/or allergic rhinitis, measured with PROMs (https://www.crd.york.ac.uk/PROSPERO identifier CRD42020216098), following the Preferred Reporting Items for Systematic Reviews and Meta-Analyses methodological standards [28].

Eligibility criteria

We consider as inclusion criteria: studies including children and adolescents (aged <18 years) with a diagnosis of asthma or allergic rhinitis, comparing montelukast with other treatments or placebo, with or without randomisation, and with at least one end-point measured with a validated PROM. No limitation of language, region or year of publication was considered.

Studies were excluded if samples were composed of patients with multimorbidity, case—control, a longitudinal design without a comparison group, comparison groups with nonpharmacological intervention, data not stratified for patients aged <18 years, with a small sample size ($n \le 30$) or not published as original articles.

Information sources and search

The search was conducted in Ovid MEDLINE and Ovid Embase databases on 1 June 2022, and updated last on 31 January 2023. We searched for both Medical Subject Headings (MeSH) and text word terms including "asthma", "rhinitis", "leukotriene antagonists" and "patient-reported outcome measures". The detailed search strategy can be found in the supplementary material (supplementary figure S1). In addition, a manual search was performed of the reference list of selected articles and other published systematic reviews.

Selection process

Two members (K. Mayoral, V. Zamora) independently reviewed titles and abstracts using the Covidence software (www.covidence.org). A pilot test was conducted to standardise criteria among reviewers. Two pairs of members (K. Mayoral, C. Lizano-Barrantes; V. Zamora, C. Miret) reviewed the articles' full text, to select the articles for data extraction. Disagreements in all phases were resolved through discussion, with the participation of a third party (M. Ferrer).

Data collection process and data items

Data extraction was carried out by one researcher (K. Mayoral, C. Lizano-Barrantes or V. Zamora) with independent verification performed by a statistician (A. Pont). We designed a pre-defined data collection form in Microsoft Excel with the information to be extracted: author and year of publication, study design, follow-up period, sample size and age, diagnosis and severity, intervention and control groups, PROMs administered, type of end-point and central tendency and dispersion statistics, adverse events and withdrawals. Authors were contacted *via* email to request any PROM information not reported in the manuscript.

Study risk-of-bias assessment

To assess the risk of bias we used RoB 2 version 2 [29], or the Risk of Bias in Non-randomized Studies of Interventions (ROBINS)-I [30] tool, according to the study design. Both tools have a common set of domains of bias: deviations from intended interventions, missing outcome data, measurement of the outcome and selection of the reported result. For the nonrandomised studies, the bias due to confounding, selection of participants and classification of interventions was added and, for the randomised controlled trials (RCTs), the bias arising from the randomisation process. The risk of bias arising from each domain is classified for the RoB 2 as "low", "high" or "some concerns"; and for ROBINS-I, "serious", "moderate", "low", or "no information". We used the Robvis tool [31] for presenting the risk-of-bias results in graphics.

Summary measures

The primary outcome was defined as the difference between treatment groups on the mean change of the PROM score. As studies differed in the instrument to gather PROMs and algorithms to construct scores also differed among instruments, PROMs results were standardised by applying the *escalc* function in R-4.2.0 [32]. Either the difference between treatment groups on the mean change of the PROM score or the mean change of each group was used to estimate the standardised mean difference (SMD). Otherwise, we calculated the mean change through a basic subtraction of the means at baseline and at follow-up evaluations for each group and the pooled standard deviation. Standardising PROM results in units of standard deviations allows them to be directly comparable, and to interpret the magnitude (effect size) of the SMD [33] as small for 0.2 sp., moderate for 0.5 sp and large for 0.8 sp, regardless of the instrument used to gather PROMs.

Synthesis of results

Aggregated data were described as part of the general narrative synthesis. Meta-analyses for quantitative syntheses were constructed to test the impact of montelukast on PROMs, by performing subgroup analyses according to type of asthma (considering severity or chronicity). A random-effects model (DerSimonian–Laird method) was employed due to expecting variation in study populations, diseases and data collection. Heterogeneity among studies was evaluated using the I^2 statistic and categorised as 25–50% low, 50–75% moderate and >75% high [34]. To explore a possible publication bias, a funnel plot was planned when the number of studies pooled is \geqslant 10. The statistical analysis software programme used was RevMan version 5.4 [35].

Additional analyses

Sensitivity analyses by risk of bias were carried out: studies rated as high risk of bias using the tool RoB 2 or the ROBINS-I were excluded as they were considered a potential source of heterogeneity.

Results

Study selection

Figure 1 shows that 3937 articles were identified, 3250 titles and abstracts were reviewed and a complete reading of 265 articles was carried out. Of them, 215 were excluded mostly due to "only adult patients" (32.5%), "other outcomes" (23.3%) and "small sample size (n<30)" (15.8%). Finally, of the 50 articles that fulfilled the inclusion criteria, 28 did not report sufficient data to be included in the meta-analysis (supplementary table A) and 22 did [36–57]. Data from five articles [43, 44, 50, 53, 57] were obtained by contacting the sponsor (a pharmaceutical company) or authors.

Study characteristics

The 22 articles (21 studies) included in some of the meta-analyses (table 1) [36–57] were published between 1998 and 2015, and all assessed asthma, except for one study evaluating allergic rhinitis [54] and another one evaluating patients with both conditions [53]. All studies were performed only in paediatric samples except for four including adults, but with data stratified by age [43–45, 53]. Sample size ranged from 21 to 689, and the follow-up period was from 2 to 52 weeks. There were seven studies using ICS in monotherapy as the comparator group [36–42], five studies using diverse treatments [43–47] and nine

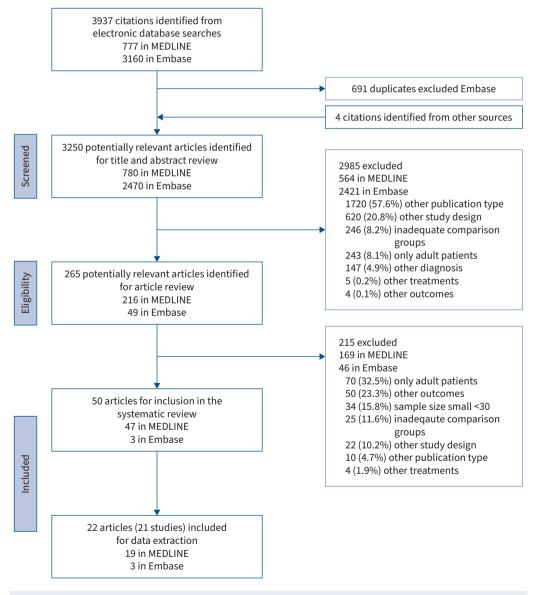


FIGURE 1 Preferred Reporting Items for Systematic Reviews and Meta-Analyses flowchart of the systematic literature review.

EUROPEAN RESPIRATORY REVIEW

TABLE 1 Characteristics of the	ne included studies							
First author [reference], publication year; study funding	Study design; follow-up period	Sample size; age range (mean)	Diagnosis	Intervention or control group	PROM administered	PROM as primary, secondary or tertiary outcome	End-point measure	Adverse events reported
LTRA versus ICS								
BUKSTEIN [36], 2003; Merck & Co, Clinical Development, US Human Health	Observational; 12 months	Total n=104 6–15 years (9.6 years)	Mild persistent asthma	I: montelukast (10 mg) C: fluticasone (176 μg)	ATAQ	Secondary	Mean± sb	No
STELMACH [37], 2005 [#] ; Medical University of Lodz, Poland	RCT; 6 months	Total n=51 6–18 years (11.8 years)	Newly diagnosed atopic asthma	I: montelukast (5 or 10 mg) C: budesonide (400 or 800 μg)	Asthma symptom score (0–9) (daytime/night-time/ β_2 -agonist use)	Secondary	Mean ±seм	No
García-García [38], 2005; Merck & Co	RCT; 12 months	Total n=541 6–14 years (median 9 years)	Mild persistent asthma	I: montelukast (5 mg) C: fluticasone (200 μg)	PATAQ overall	Tertiary	Mean change (95% CI)	Yes
Ostrom [39], 2005 [#] ; GlaxoSmithKline Inc	RCT; 12 weeks	Total n=342 6–12 years (9.35 years)	Chronic asthma	I: montelukast (5 mg) C: fluticasone (100 μg)	Daytime symptom score (0–5) Night-time symptom score (0–3) Satisfaction with treatment ⁺	Not specified	Mean±seм Mean change±seм	Yes
Szefler [40], 2007; AstraZeneca LP	RCT; 52 weeks	Total n=312 2–8 years (4.65 years)	Mild persistent asthma	I: montelukast (5 mg) C: budesonide (500 μg)	CHQ-PF50: general/ activities CHSA: physical/activities/ emotional Asthma symptom score (am/pm)	Secondary	Mean±sb Mean±sem Adjusted mean change±sem	Yes
OLSZOWIEC-CHLEBNA [41], 2010; Medical University of Lodz, Poland	RCT; 6 months	Total n=60 5–18 years (8.53 years)	Newly diagnosed asthma	I: montelukast (5 or 10 mg) C: budesonide (400 and 800 µg)	ATAQ	Not specified	Mean±sb	No
Parisi [42], 2022; University of Catania, Italy	RCT; 12 weeks	Total n=42 2–6 years (3.5 years)	Recurrent wheezing	I: montelukast (4 or 5 mg) C: beclomethasone (200 µg)	TRACK	Primary	Mean±sd	Yes
LTRA versus other treatment								
Malmstrom [43], 1999 [¶] ; Merck Research Laboratories	RCT; 12 weeks	Total n=895 >15 years Paediatric n=45 15–18 years	Chronic asthma	I: montelukast (10 mg) C: beclomethasone (400 μg) C: placebo	Daytime asthma symptoms Global evaluation question [†] AQLQ [†]	Primary	Mean±sp Mean change±sp	Yes
		•						Continued

EUROPEAN RESPIRATORY REVIEW

TABLE 1 Continued								
First author [reference], publication year; study funding	Study design; follow-up period	Sample size; age range (mean)	Diagnosis	Intervention or control group	PROM administered	PROM as primary, secondary or tertiary outcome	End-point measure	Adverse events reported
Williams [44], 2001 ⁴ ; Merck Research Laboratories	RCT; 16 weeks	Total n=1055 15–85 years Paediatric n=245 6–14 years	Chronic asthma	I: montelukast (5 or 10 mg) C: beclomethasone (300 μg) C: placebo	PAQLQ: activity/symptom/ emotional	Not specified	Mean Adjusted mean change (95% CI)	No
PETERS [45], 2007; GlaxoSmithKline, American Lung Association	RCT; 16 weeks	Total n=500 >15 years Paediatric n=245 6–14 years	Mild persistent controlled asthma	I: montelukast (5 or 10 mg) C: fluticasone (200 µg) C: fluticasone (100 µg) + salmeterol (50 µg)	Mini AQLQ ASUI ⁺ ACQ ⁺	Secondary	Adjusted mean±sb (95% CI)	No
Máspero [46], 2008; GlaxoSmithKline	RCT; 12 weeks	Total n=548 6–14 years (9.3 years)	Persistent uncontrolled asthma	I: montelukast (5 mg) C: salmeterol/ fluticasone (100/ 200 µg)	PAQLQ Asthma symptom score (0–5)⁺	Secondary	Adjusted mean change±seм	No
Bérubé [47], 2014; Merck Frosst Canada Ltd	Observational; 12 weeks	Total n=328 2–14 years (6.92 years)	Uncontrolled asthma	I: montelukast (4 or 5 mg) C: montelukast (4 or 5 mg) ICS (different doses)	ACQ	Primary and secondary	Mean±sb (95% CI) Mean change±sb	No
LTRA versus placebo								
Knorr [48], 1998; Merck & Co Inc Becker [49], 2004; Merck & Co Inc	RCT; 8 weeks	Total n=336 6–14 years (median 11 years)	Mild asthma	I: montelukast (5 mg) C: placebo	PAQLQ: activity/symptom/ emotional Daytime symptom score (0–5) Night-time symptom score (0–5)	Primary	Mean±sb Mean change (95% CI)	No
K _{NORR} [50], 2001 [¶] ; Merck Research Laboratories	RCT; 12 weeks	Total n=689 2–5 years (3.6 years)	Mild to severe persistent asthma	I: montelukast (4 mg) C: placebo	Daytime-symptom score (0–5) Overnight asthma symptoms	Secondary	Mean change (95% CI)	Yes
Stelmach [51], 2002; Medical University of Lodz, Poland	RCT; 6 weeks	Total n=32 6–18 years (13.5 years)	Mild to moderate asthma	I: montelukast (5 or 10 mg) C: placebo	Asthma symptom score (0–9) (daytime/night-time/ β ₂ -agonist use)	Secondary	Mean±sb (95% CI)	No

EUROPEAN RESPIRATORY REVIEW

irst author [reference], publication year; study funding	Study design; follow-up period	Sample size; age range (mean)	Diagnosis	Intervention or control group	PROM administered	PROM as primary, secondary or tertiary outcome	End-point measure	Adverse events reported
STELMACH [52], 2002#; Medical University of Lodz, Poland	RCT; 8 weeks	Total n=49 9–15 years (11.9 years)	Moderate atopic asthma	I: montelukast (5 or 10 mg) C: placebo	Asthma symptom score (0–9) (daytime/night-time/ β ₂ -agonist use) PAOLO ⁺	Not specified	Mean±sb (95% CI)	No
PHILIP [53], 2004 [¶] ; Merck Research Laboratories	RCT; 2 weeks	Total n=831 15–85 years Paediatric n=102 15–18 years	Asthma and seasonal allergic rhinitis	l: montelukast (10 mg) C: placebo	Daily rhinitis symptoms (0–3) Daytime symptom score (0–3) Night-time symptom score (0–3) ROLO+	Primary	Mean change±sp	No
CHEN [54], 2006; Chung Shan Medical University Hospital, Taiwan	RCT; 12 weeks	Total n=40 2–6 years (4.4 years)	Rhinitis	I: montelukast (5 mg) C: placebo	PRQLQ Symptom score	Not specified	Mean±sd Mean change±sd	No
Spahn [55], 2006; [#] Merck & Co Inc	RCT; 8 weeks	Total n=21 9–18 years (13.1 years)	Mild to moderate asthma	I: montelukast (5 or 10 mg) C: placebo	Daytime symptom score (0–3) Night-time symptom score (0–3)	Not specified	Mean±seм	Yes
STELMACH [56], 2015; Medical University of Lodz, Poland. National Science Centre	RCT; 30 weeks	Total n=76 6–14 years (10.7 years)	Allergic asthma	I: montelukast (5 mg) C: placebo	ACT	Not specified	Mean±sb	No
SAHINER [57], 2021; Merck Sharp and Dohme Company	RCT; 4 weeks	Total n=46 6–18 years (11 years)	Asthma	I: montelukast (10 mg) C: placebo	ACT	Secondary	Mean±sp	No

PROM: patient-reported outcome measure; LTRA: leukotriene receptor antagonist; ICS: inhaled corticosteroid; I: intervention group; C: control group; ATAQ: Asthma Therapy Assessment Questionnaire for children; RCT: randomised controlled trial; PATAQ: Pediatric Asthma Therapy Assessment Questionnaire; CHQ-PF50: Child Health Questionnaire Parent Form-50; CHSA: Children's Health Survey for Asthma; TRACK: Test for Respiratory and Asthma Control in Kids; AQLQ: Asthma Quality of Life Questionnaire; PAQLQ: Pediatric Asthma Quality of Life Questionnaire; ASUI: Asthma Symptom Utility Index; ACQ: Asthma Control Questionnaire; RQLQ: Rhinoconjunctivitis Quality of Life Questionnaire; PRQLQ: Pediatric Rhinoconjunctivitis Quality of Life Questionnaire; ACT: Asthma Control Test. #: citations identified from other reviews; *1: data obtained after contacting sponsor; *: PROM information not available.

studies (10 articles) comparing it only with placebo [48–57]. All administered PROMS were disease-specific, except for a RCT which used the generic questionnaire Child Health Questionnaire Parent Form-50 (CHQ-PF50) [40]. The constructs collected were symptoms (differentiating between daytime and night-time, or considered globally) and HRQoL.

Authors' judgements about each risk-of-bias item for each included study

Figure 2 shows that, out of 20 RCTs, 12 had a low risk of bias [37, 39, 41, 43–46, 48, 50, 52, 54, 57], six had high risk [38, 40, 42, 52, 55, 56] and one study presented some concerns [53] (figure 2a). The most frequent domain causing downgrading was "bias in measurement of the outcome", because the assessment of the outcome was probably being influenced by knowledge of the intervention received by patients [38, 40, 42, 52, 55, 56], and no information was reported in the articles. Figure 2b shows two nonrandomised studies [36, 47] which presented moderate and serious risk of bias.

Results of syntheses

Figure 3 shows forest plots of studies comparing HRQoL of patients with asthma treated with ICS in monotherapy or combined with long-acting β -agonists (LABAs) *versus* montelukast. Figure 3a shows the forest plot of the SMD of change for the global HRQoL, which included six studies with a sample size of 656 participants for the intervention group, and 535 for the control group (ICS in monotherapy); the pooled estimator was -0.06 (95% CI -0.14-0.02). The subgroup analysis was statistically significant (p<0.001), showing that, in patients with mild persistent asthma and unspecified type of asthma, the change in HRQoL was in favour of ICS, but there was one study in favour of montelukast including patients with mild persistent controlled asthma.

Similarly, the HRQoL forest plot of the two studies comparing montelukast with ICS/LABA (figure 3b) was in favour of the latter group without statistically significant differences: overall estimator -0.13 (95% CI -0.28–0.02) with a moderate heterogeneity (I²=69%). Both studies were carried out among patients with persistent asthma [45, 46]. Figure 3c and d shows forest plots of mental and physical domains constructed with only two studies providing data at domain level (n=298 intervention, n=158 control), which used ICS in monotherapy as the comparator group. Pooled estimators were not statistically significant and heterogeneity was high in both cases.

Figure 4a and b shows forest plots of studies comparing patients with asthma treated with ICS *versus* montelukast in terms of daytime (n=386 intervention, n=376 control) and night-time (n=363 intervention, n=363 control) symptoms, which showed a statistically significant difference in favour of ICS. When assessing the symptoms globally (figure 4c), there were no differences between montelukast and ICS (-0.04, 95% CI -0.55-0.47). The subgroup analysis by type of asthma was statistically significant (p<0.001), and heterogeneity was high (I^2 >90%) for the three meta-analyses. The change was statistically significant in favour of montelukast in a study among newly diagnosed atopic asthma patients, while favouring ICS in a study among patients with recurrent wheezing.

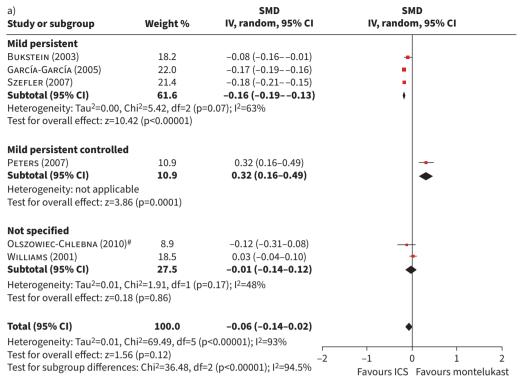
Meta-analyses of studies comparing montelukast with placebo are shown in figure 5. Pooled estimators of daytime, night-time symptoms and global symptoms scores (figure 5a, b and c) showed a statistically significant difference in favour of montelukast: 0.11 (95% CI 0.05–0.18), 0.15 (95% CI 0.09–0.21) and 0.90 (95% CI 0.44–1.36). The heterogeneity was high, except for the night-time symptoms score ($I^2 = 38\%$). The subgroup analysis by asthma severity was not statistically significant in any of these three meta-analyses.

Sensitivity analyses excluding studies with a high risk of bias are shown in supplementary figure S2. Pooled estimators obtained in these sensitivity analyses were in the same direction and similar magnitude to the meta-analyses constructed with all the studies.

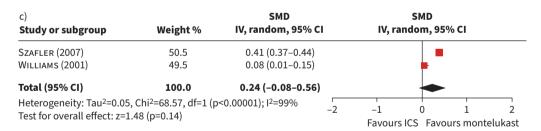
Six studies [38–40, 47, 48, 50] provided information on adverse drug events and study withdrawals, and three other studies [37, 46, 55] provided information only on study withdrawals (supplementary table B). Most studies reporting adverse drug events concluded that those were not significantly different for montelukast when compared to other treatments [38–40, 48, 50], while some studies reported withdrawals more frequently in the montelukast group [38–40, 46, 48]. Asthma exacerbations, headaches and upper respiratory tract infections were among the most commonly reported events [38–40, 48, 50]. After having requested them from the authors, one study [50] provided adverse effects for 60 participants, the most frequent being headache (n=35), irritability (n=7), nervousness (n=4), behaviour disturbance (n=4) and drug overdose (n=4, presenting thirst, mydriasis and somnolence). There was one study [47] that reported both adverse events and withdrawals in general, but montelukast was included in both the intervention and



FIGURE 2 a) Risk-of-bias summary for randomised controlled trials using the revised tool for risk of bias (RoB 2) in randomised trials; b) risk of bias of nonrandomised studies assessed using the Risk of Bias in Non-randomized Studies of Interventions (ROBINS-I) tool.

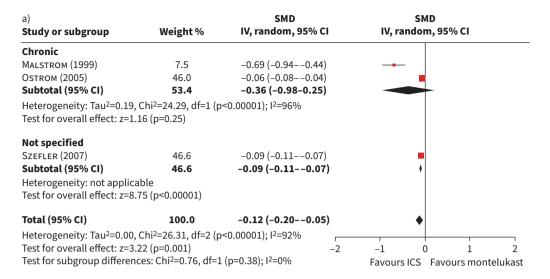


b) Study or subgroup	SMD roup Weight % IV, random, 95			SMD lom, 95% (:1	
Máspero (2008)	65.4	-0.07 (-0.090.06)				
PETERS (2007)	34.6	-0.23 (-0.410.06)	-	-		
Total (95% CI)	100.0	-0.13 (-0.28-0.02)	•	•		
Heterogeneity: Tau ² =0.01 Test for overall effect: z=1	-2 -1 Favours ICS/LAB	0 A Favours	1 monte	 2 elukast		



d) Study or subgroup	Weight %	SMD IV, random, 95% CI		IV, ra			
SZAFLER (2007) WILLIAMS (2001)	52.6 47.4	-0.09 (-0.120.06) 0.06 (-0.02-0.13)					
Total (95% CI)	100.0	-0.02 (-0.16-0.12)			•		
Heterogeneity: Tau ² =0.01 Test for overall effect: z=0	=0.0002); I ² =93%	-2	-1 Favour	0 s ICS Favo	1 ours mont	2 elukast	

FIGURE 3 Forest plots of studies comparing patients with asthma treated with inhaled corticosteroids (ICS) in monotherapy or as a combined treatment with long-acting β -agonists (LABAs) *versus* montelukast in terms of health-related quality of life (HRQoL). a) Global scores of HRQoL for monotherapy; b) global scores of HRQoL for combined therapy; c) mental and d) physical dimension scores of HRQoL for monotherapy. SMD: standardised mean difference; IV: inverse variance. #: study including more than one dose of ICS; 400 μg budesonide was used.



b) Study or subgroup	Weight %	SMD IV, random, 95% CI	IV, rand			
Оѕтком (2005)	49.9	-0.32(-0.350.30)				
Szefler (2007)	50.1	-0.14 (-0.160.12)	l			
Total (95% CI)	100.0	-0.23 (-0.410.06)	•			
Heterogeneity: Tau ² =0.02 Test for overall effect: z=2		p<0.00001); I ² =99% -2	-1 Favours IC	0 S Favou	1 Irs mont	2 elukast

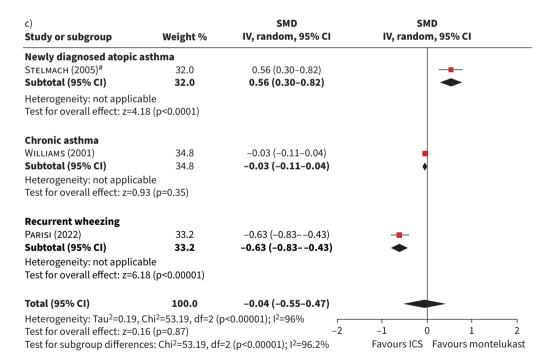
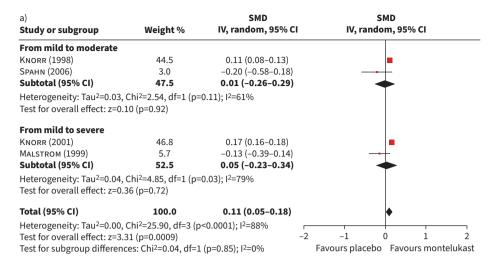
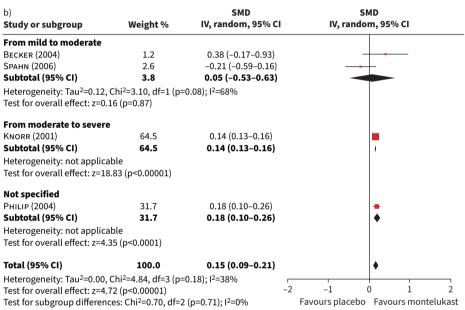


FIGURE 4 Forest plot of studies comparing patients with asthma treated with inhaled corticosteroids (ICS) *versus* montelukast in terms of symptoms. a) Daytime symptom score; b) night-time symptom score; c) global symptom score. SMD: standardised mean difference; IV: inverse variance. #: study including more than one dose of ICS; 400 µg budesonide was used.





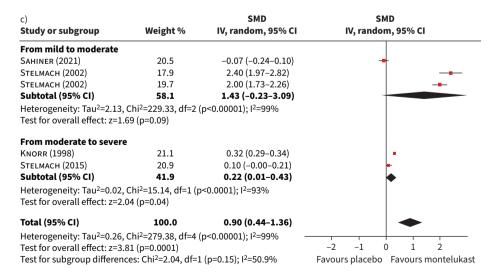


FIGURE 5 Forest plot of studies comparing patients with asthma treated with placebo *versus* montelukast in terms of symptoms. a) Daytime symptom score; b) night-time symptom score; c) global symptom score. SMD: standardised mean difference; IV: inverse variance.

control groups: nightmares and sleep terror (n=6), abdominal pain (n=5), insomnia (n=2) and headache (n=2). Another study [38] reported one death in the montelukast group, although it was not considered to be related to the drug.

Discussion

In this systematic review, we identified 50 articles (49 studies) evaluating the impact of montelukast on PROMs when compared with other treatments or placebo in children and adolescents with asthma/allergic rhinitis. The 21 studies finally included in some of the meta-analyses allowed for the construction of pooled estimators of montelukast impact on HRQoL (global, mental and physical domains) and symptoms (daytime, night-time and globally) with high heterogeneity. Results were in favour of ICS without statistically significant differences, in the global, mental and physical HRQoL mean changes. For daytime and night-time symptom scores, the change was significant and in favour of ICS. The incidence of adverse events and withdrawals was only provided by six and nine of these studies, respectively, showing infrequent adverse events, which were similar to those seen in the control group, and only one study reported neuropsychiatric events, but withdrawals were slightly more frequent in the montelukast group.

In our systematic review, of the 49 studies that fulfilled inclusion criteria, 11 (22%) included PROMs as a primary outcome, 21 (43%) as a secondary outcome, one (2%) as a tertiary outcome and 16 (33%) did not specify. Furthermore, of the 21 studies included in the meta-analysis, six did not include complete data for all the PROMs administered [39, 43, 45, 46, 52, 53]. The 28 studies whose data could not be extracted for meta-analysis were predominantly RCTs in asthma patients comparing montelukast with ICS [58–67] or placebo [64, 68–75] (supplementary table A). Reasons for noninclusion in the meta-analyses were results not being segregated for children and adults [59, 60, 64, 66, 67, 69, 72–74, 76–82], or PROM data not fully reported [58, 61–63, 65, 68, 70, 71, 75, 83–85]. These latter 12 studies with missing information on PROMs, which remained unavailable after contacting authors, suggest the poor relevance given to these end-points. This is in line with reviews [86, 87] highlighting the small proportion of asthma clinical trials including PROMs: only 20 out of 300 published RCTs [86], and only 27% of 96 736 registered trials [87]. There is growing evidence that PROM results are frequently omitted and not reported, and data are presented with suboptimal reporting standards [88, 89]. Issues such as lack of PROMs training, poor understanding of the purpose of PROMs assessment and questionnaire selection [90] should be contemplated to avoid research waste.

Global HRQoL pooled results of our meta-analyses were in favour of ICS in monotherapy or combined with LABAs, but they were not statistically significant, with a negligible magnitude of difference. The study that constituted the subgroup of mild persistent controlled asthma was the only one favouring montelukast, which is consistent with by the Global Initiative for Asthma (GINA)'s recommendation of using montelukast in patients with fewer frequency of symptoms (more than once a month, but less than daily) [4]. The studies included in our systematic review of HRQoL mostly administered asthma-specific questionnaires [36, 38, 40, 41, 43–46, 48, 52–54], similarly to findings of a review of PROMs in asthma [86], where the low use of generic questionnaires was highlighted. In fact, the National Asthma Education and Prevention Program [91] in the United States supported the use of generic HRQoL instruments to capture patients' global health perceptions.

Results for both meta-analyses on mental and physical domains of HRQoL were also in favour of ICS, without statistically significant differences; however, the first one achieves a small magnitude of the difference between treatment groups. Mental health is an important component of health in children with asthma [92, 93]. The USA National Survey of Children's Health [94] reported higher odds ratios for developmental or behavioural problems in children with asthma. Furthermore, considering the lack of screening tools for anxious and depressive symptomatology validated among asthma populations remarked by GINA [4], mental dimensions covered by PROMs could be a valuable proxy, since they could capture mental adverse problems related to drugs or disease. Several studies included in our meta-analyses measured the mental health component of the HRQoL, such as those administering the Children's Health Survey for Asthma (CHSA) [40] (emotional health domain); the CHQ-PF50 (role/social emotional functioning, role/social behavioural functioning, mental health and behaviour domains); the Pediatric Asthma Quality of Life Questionnaire (PAQLQ) [44, 46, 48, 52], the Asthma Quality of Life Questionnaire (AQLQ) [43] or Mini-AQLQ [45] (emotional function domain); the Asthma Therapy Assessment Questionnaire for children [36, 41] and Pediatric Asthma Therapy Assessment Questionnaire [38] (attitude and behaviour domain); and the Pediatric Rhinoconjunctivitis Quality of Life Questionnaire [53, 54] (feeling irritable item). Unfortunately, the majority of the articles identified only included total scores [36, 38, 41, 44, 45, 54], and none provided information at dimension and item level, which may have limited the possibility of finding the real impact of montelukast on mental and physical components of health.

Only three studies provided information on the dimension scores: one study that administered the CHSA [40], and two that administered the PAQLQ [44, 48]. The latter used placebo as the comparison group, and the first one used ICS.

The meta-analysis of global symptoms shows opposite results according to the type of asthma used as inclusion criteria in the studies (newly diagnosed atopic asthma, chronic asthma and recurrent wheezing). Results of daytime and night-time symptom meta-analyses were statistically significant, favouring ICS over montelukast, with only night-time achieving a small magnitude. This is in line with asthma guidelines recommendation of ICS therapy as the step-one treatment for paediatric patients with mild-to-moderate persistent asthma [4, 18].

Furthermore, our results are consistent with those obtained in previous systematic reviews of schoolchildren with mild to moderate asthma [13], preschool children [21] and paediatric patients [20], concluding that those treated with ICS had better symptom control than those with montelukast. Previous systematic reviews evaluating the impact of montelukast through PROMs were predominantly focused on symptom control [13, 20, 21], while there have been very few reviews gathering information on global HRQoL [24], and none were found summarising HRQoL mental and physical domains.

The synthesis of trials using placebo as a comparison group confirmed the efficacy of montelukast in terms of statistically significant improvement in symptoms during daytime, night-time and globally. However, the magnitude of the benefits was only large for symptoms when measured globally, but very small (<0.2~sd) for symptoms measured separately for daytime and night-time. The large magnitude for the global symptoms pooled estimator (0.9~sd) is explained by two [51, 52], of the five studies included. These studies [51, 52] measured asthma symptoms with an overall score covering daytime, night-time and also the use of β_2 -agonists, differing from the other studies which did not include the use of rescue inhalers [48, 56, 57]. Results of studies using placebo as comparison group, without enough information to be included in the meta-analysis, did not show significant differences for HRQoL [70] and symptom score [75] in asthma patients using montelukast, except for one study which obtained a statistically significant higher HRQoL improvement in this group [71]. Another study in children with allergic rhinitis found that montelukast's improvement of HQRL was significantly higher than placebo [68].

In this systematic review, we have not found any PROMs aiming to measure treatment adverse events, and only six studies reporting them [38–40, 47, 48, 50]. Only one of these studies reported neuropsychiatry-related events [50] such as irritability, nervousness, behaviour disturbance, insomnia, dream abnormality, anxiety, bipolar disorder and personality change, but they were more infrequent than in the control group (4.0% *versus* 5.1%). This is consistent with results of the FDA-requested review of behaviour-related adverse experiences in montelukast clinical trials [14], but it is important to consider that some neuropsychiatric events could have been under-reported, since patients/parents may have not associated the treatment with the adverse experience. Conversely, study withdrawals, extracted as a proxy to adverse events, were more frequent in the montelukast group in five out of eight studies [38–40, 46, 48]. This result is consistent with another systematic review [24] which found that LTRA associated with a 24% increased risk of overall withdrawals in paediatric patients, but also with no statistically significant group difference due to adverse effects. Although treatment may occasionally be discontinued for mild or moderate adverse events, withdrawals are especially important because they reflect the ultimate decision of the participant and/or physician to discontinue treatment.

The main limitation related to the systematic review process was that we may have failed to identify all articles assessing the impact of montelukast on PROMs in children/adolescents. However, we believe that this has been minimised due to the bibliographic database interface used (Ovid), the sensitive search strategy, and the additional hand search of references. Regarding limitation related to the studies included and pooled results obtained, first, heterogeneity was high (I²>75%) in all meta-analyses comparing montelukast with ICS in monotherapy, and moderate (I²= 69%) when compared with ICS combined with LABAs. Since we hypothesised asthma severity as a potential source of heterogeneity, we performed subgroup analyses according to this variable, but it was statistically significant only in one of the three meta-analyses where it could be tested: the one comparing HRQoL between montelukast and ICS. It is important to remark the lack of uniform criteria across studies for asthma severity and classification types, which limited our subgroup analyses, but many other possible reasons could explain such heterogeneity. Second, the internal validity of the pooled estimator provided by a meta-analysis depends on the quality of primary studies. Results of the sensitivity analyses excluding the 29.4% of studies qualified as high risk of bias were consistent with those obtained by the meta-analyses constructed with all the studies. Third, the follow-up period in most of trials could be considered short to detect montelukast's adverse events, which

could appear after stopping treatment [16]. In fact, 11 studies were conducted over \leq 12 weeks [39, 43, 46–48, 50–55]. Fourth, the ability to generalise our findings in preschool children (\leq 5 years) is limited because only two studies examined this population; likewise in paediatric patients with allergic rhinitis, for which there were five studies and most of them without data to be included in the meta-analyses. Fifth, no funnel plot for publication bias assessment could be constructed because of the low number of studies pooled (\leq 10) in all forest plots. Finally, concerns about the sources of these studies are remarkable, as more than half of the studies included were sponsored by pharmaceutical companies.

Points for clinical practice

- The results of this systematic review suggest that the benefit of ICS on night-time symptoms was statistically higher than that of montelukast, although the magnitude of the difference was small.
- HRQoL did not present any statistically significant differences between montelukast and ICS, and their magnitude was negligible except for the mental domain, which was small.
- Montelukast was effective in controlling asthma symptoms when compared with placebo.
- These findings are in accordance with current guidelines for asthma treatment, which includes LTRAs in the category of other controller options, but remark upon their lack of evidence for safety.

Conclusion

Our results suggest that the benefit of ICS on night-time symptoms was statistically higher than that of montelukast, although the magnitude of the difference was small. HRQoL did not present any statistically significant differences between montelukast and ICS, and their magnitude was small only for the mental domain (negligible for the rest). Compared with placebo, montelukast was effective in controlling asthma symptoms. These findings of our systematic review are in accordance with current guidelines for asthma treatment, which recommend the use of ICS therapy as first step treatment for paediatric patients with mild-to-moderate persistent asthma; and considering LTRAs in the category of other controller options, but taking into account the lack of evidence for safety. It is essential to achieve a broader understanding of the safety of montelukast by long-term post-marketing studies, and improving the collection of adverse events in clinical trials, specifically measuring them through validated PROMs.

Provenance: Submitted article, peer reviewed.

Acknowledgements: The authors would like to acknowledge Raúl Castro (Universitat Autònoma de Barcelona, Barcelona, Spain) for helping with part of the search process, and Aurea Martin (CIBER en Epidemiología y Salud Pública, CIBERESP, Madrid, Spain and Health Services Research Group, Hospital del Mar Research Institute, Barcelona) for her help in the English editing and proofreading process and in finalising this manuscript.

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Authors' contributions: M. Ferrer conceived the research question, contributed to the conception and design of the article, oversaw all aspects, contributed to the statistical analyses, carried out the interpretation of data, and contributed to the writing of the article; K. Mayoral, C. Lizano-Barrantes and V. Zamora participated in the conception of the research question, conducted the literature search and data extraction, as well as the quality assessment, conceptualised and oversaw analyses, and wrote the article; C. Miret and C. Barrufet conducted the literature search and data extraction, as well as the quality assessment; A. Pont contributed to the analysis and provided statistical support and help in the interpretation of data; M.A. Caballero-Rabasco, M. Praena-Crespo, A. Bercedo, L. Valdesoiro-Navarrete and M.T. Guerra participated in the conception of the research question, ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved, and helped in the interpretation of data and drafting of the paper; O. Garin, Y. Pardo and

M.J. Martínez Zapata revised important intellectual content and the draft versions of the manuscript. All the co-authors critically revised the manuscript and approved the final draft before submission, and can attest to the validity and legitimacy of the data in the manuscript and agree to be named as author of the manuscript.

Conflict of interest: All authors have nothing to disclose.

Support statement: Financial support for this study was provided through grants by the Instituto de Salud Carlos III FEDER: Fondo Europeo de Desarrollo Regional (PI15/00449) and DIUE of Generalitat de Catalunya (AGAUR 2021 SGR 00624, 2017 SGR 452). The following researchers have worked on this manuscript while funded by grants: K. Mayoral and V. Zamora (Instituto de Salud Carlos III FEDER: Fondo Europeo de Desarrollo Regional FI16/00071, FI19/00229), K. Mayoral also obtained financial support from CIBER of Epidemiology and Public Health CIBERESP for a short PhD stay abroad, and C. Lizano-Barrantes (University of Costa Rica OAICE-85-2019). The funding agreements ensure the authors' independence in designing the study, interpreting the data, and writing and publishing the report. Funding information for this article has been deposited with the Crossref Funder Registry.

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