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Genetics and population analysis

# MOIRE: A software package for the estimation of allele frequencies and effective multiplicity of infection from polyallelic data

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# Abstract

**Motivation:** Malaria parasite genetic data can provide insight into parasite phenotypes, evolution, and transmission. However, estimating key parameters such as allele frequencies, multiplicity of infection (MOI), and within-host relatedness from genetic data is challenging, particularly in the presence of multiple related coinfecting strains. Existing methods often rely on single nucleotide polymorphism (SNP) data and do not account for within-host relatedness.

**Results:** We present MOIRE (Multiplicity Of Infection and allele frequency REcovery), a Bayesian approach to estimate allele frequencies, MOI, and within-host relatedness from genetic data subject to experimental error. MOIRE accommodates both polyallelic and SNP data, making it applicable to diverse genotyping panels. We also introduce a novel metric, the effective MOI (eMOI), which integrates MOI and within-host relatedness, providing a robust and interpretable measure of genetic diversity. Extensive simulations and real-world data from a malaria study in Namibia demonstrate the superior performance of MOIRE over naive estimation methods, accurately estimating MOI up to 7 with moderate sized panels of diverse loci (e.g. microhaplotypes). MOIRE also revealed substantial heterogeneity in population mean MOI and mean relatedness across health districts in Namibia, suggesting detectable differences in transmission dynamics. Notably, eMOI emerges as a portable metric of within-host diversity, facilitating meaningful comparisons across settings when allele frequencies or genotyping panels differ. Compared to existing software, MOIRE enables more comprehensive insights into within-host diversity and population structure.

Availability: MOIRE is available as an R package at https://eppicenter.github.io/moire/. Contact: mm@maxmurphy.dev

Supplementary information: Supplementary data are available at *Bioinformatics* online.

# 1 Introduction

2 Genetic data can be a powerful source of information for 3 understanding malaria parasite phenotype and transmission 4 dynamics, providing insight into population structure and 5 connectivity, and thereby informing control and elimination 6 efforts. However, analysis is complicated in malaria due to 7 the presence of multiple coinfecting, genetically distinct strains. 8 More specifically, genetically distinct strains may share the 9 same alleles at genetic loci, rendering the actual number of 10 strains contributing a particular allele unknown and making it

difficult to estimate fundamental statistics such as population 11 allele frequencies and multiplicity of infection (MOI). Standard 12 methods to address this either naively estimate allele frequencies 13 and MOI without considering the total number of strains 14 contributing an allele (Roh et al., 2019; Tessema et al., 2019; 15 Pringle et al., 2019), or completely ignore polyclonal samples 16 during analysis. Naive estimation without accounting for strain 17 count contribution results in biased estimates of allele frequencies 18 and MOI, leading to meaningful systematic biases in summary 19 statistics. For example, naive estimation of allele frequencies 20

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21 without consideration of strain composition from polyclonal 22 samples results in a consistent overestimation of heterozygosity, leading to potentially faulty inference about population diversity. 23 Additionally, naive estimation offers no principled way to address 24 genotyping error beyond heuristics, further biasing estimates of 25 diversity in ways that depend on choices made during initial 26 interpretation of genotyping data. Alternatively, considering only 27 monoclonal samples is potentially problematic, as this may require 28 a substantial number of samples to be discarded when collected 29 from regions where multiple infection is the rule rather than 30 the exception. Further, the monoclonal subset of samples are 31 fundamentally different from the larger population of interest, as 32 they preclude the possibility of within-host relatedness between 33 34 strains. This ignores a potentially important source of information about transmission dynamics, as within-host relatedness may 35 be indicative of co-transmission or persistent local transmission 36 (Wong et al., 2017; Nkhoma et al., 2020; Wong et al., 2018). 37

To address these issues and make full use of available data, 38 Chang et al. (2017) developed a Bayesian approach (THE REAL 39 40 McCOIL) to estimate allele frequencies and MOI in the context of polygenomic infections from single nucleotide polymorphism 41 (SNP) based data. More recently, coiaf (Paschalidis et al., 42 2023) and SNP-Slice (Ju et al., 2023) have been developed to 43 further improve computational efficiency and resolving power. 44 Briefly, coiaf takes user provided allele frequencies and SNP read 45 count data and applies an optimization procedure to estimate 46 either discrete or continuous values for MOI. SNP-Slice also 47 takes SNP read count data and uses a non-parametric Bayesian 48 approach to simultaneously estimate phased strain identity and 49 within-host strain composition. While Paschalidis et al. suggest 50 within-host relatedness as a possible explanation for continuous 51 values of MOI, and the method by Ju et al. may provide 52 a way to interrogate within-host relatedness through phased 53 54 strain composition, none of these methods directly consider or estimate within-host relatedness. Further, these methods are all 55 tailored to SNP based data and are unable to accommodate more 56 diverse polyallelic loci, such as microsatellites, which have been 57 widely used in population genetic studies (Anderson *et al.*, 2000; 58 Tessema et al., 2019; Roh et al., 2019; Pringle et al., 2019). 59 Other methods that infer within-host relatedness (Zhu et al., 60 2019), in contrast, rely on whole genome sequencing (WGS) 61 data. WGS based approaches, however, frequently have poor 62 63 sensitivity for detecting minority strains and low density infections (Tessema et al., 2022). In recent years, the declining cost of DNA 64 65 sequencing and development of high throughput, high diversity, 66 targeted sequencing panels have made polyallelic data even more 67 attractive for genomic based studies of malaria (Tessema et al., 2022; LaVerriere et al., 2022; Kattenberg et al., 2023). Genetic 68 analysis methods leveraging polyallelic loci have the potential 69 for substantially increased resolving power over their SNP based 70 counterparts, particularly in the context of related polyclonal 71 infections in malaria (Taylor et al., 2019; Inna Gerlovina et al., 72 2022). Unfortunately, there are limited tools available to analyze 73 these types of data. 74

75 We present here a new Bayesian approach, Multiplicity Of 76 Infection and allele frequency REcovery from noisy polyallelic 77 data (*MOIRE*), that, like *THE REAL McCOIL*, enables the 78 estimation of allele frequencies and MOI from genomic data 79 that are subject to experimental error. In addition, MOIRE 80 estimates and accounts for within-host relatedness of parasites,

a common occurrence due to the inbreeding of parasites serially 81 co-transmitted by mosquitoes (Nkhoma et al., 2020, 2012). 82 Critically, MOIRE takes as input genetic data of arbitrary 83 diversity, allowing for estimation of allele frequencies, MOI, 84 and within-host relatedness from polyallelic as well as biallelic 85 data. MOIRE is able to fully utilize polyallelic data, yielding 86 joint estimates of allele frequencies, sample specific MOIs and 87 within-host relatedness along with probabilistic measures of 88 uncertainty. We demonstrate through simulations and applications 89 to empirical data the ability of MOIRE to leverage a variety of 90 polyallelic markers. Polyallelic markers can greatly improve jointly 91 estimating sample MOI, within-host relatedness, and population 92 allele frequencies, resulting in reduced bias and increased power 93 for understanding population dynamics from genetic data. We also 94 introduce a new metric of diversity, the effective MOI (eMOI), a 95 continuous value that combines estimates of the true MOI and 96 the degree of within-host relatedness in a single sample, providing 97 an interpretable quantity that is comparable across genotyping 98 panels and transmission settings. We contrast this with the within-99 host infection fixation index,  $F_{WS}$ , a frequently used metric 100 of within-host diversity and signal of inbreeding and population 101 sub-structure (Manske et al., 2012; Auburn et al., 2012), and 102 demonstrate the inherent shortcomings of  $F_{WS}$  as a non-portable 103 metric. 104

#### Methods

### A model of infection and observation

Consider observed genetic data  $X = (X_1, \ldots, X_n)$  from *n* samples 107 indexed by i, where each  $X_i$  is a collection of vectors indexed 108 by l of possibly differing length, representing the varying number 109 of alleles possible at each locus, e.g. polyallelic loci. Each vector 110 is binary, with 1 representing the allele was observed or 0 111 representing the allele went unobserved at locus l for sample i. 112 From this data, we wish to estimate MOI for each individual 113  $(\boldsymbol{\mu} = [\mu_1, \dots, \mu_n])$ , within host relatedness  $(\boldsymbol{r} = [r_1, \dots, r_n])$ , 114 defined as the average proportion of the genome that is identical by 115 descent across all strains, individual specific genotyping error rates 116  $(\epsilon^+ = [\epsilon_1^+, \dots, \epsilon_n^+]$  and  $\epsilon^- = [\epsilon_1^-, \dots, \epsilon_n^-])$ , and population allele 117 frequencies at each locus ( $\boldsymbol{\pi} = [\pi_1, \ldots, \pi_l]$ ). Similar to Chang *et al.* 118 (2017), we applied a Bayesian approach and looked to estimate the 119 posterior distribution of  $\mu, r, \epsilon^+, \epsilon^-$  and  $\pi$  as 120

$$P(\boldsymbol{\mu}, \boldsymbol{r}, \boldsymbol{\epsilon}^+, \boldsymbol{\epsilon}^-, \boldsymbol{\pi} | X) \propto \prod_{i=1}^n P(X_i | \boldsymbol{\mu}_i, r_i, \boldsymbol{\epsilon}_i^+, \boldsymbol{\epsilon}_i^-, \boldsymbol{\pi}) P(\boldsymbol{\mu}, \boldsymbol{r}, \boldsymbol{\epsilon}^+, \boldsymbol{\epsilon}^-, \boldsymbol{\pi})$$
(1)

where we assumed independence between samples.

Given that the observed genetic data are experimentally derived, 123 they are subject to some rate of false positives where an allele 124 is erroneously called as present, and false negatives where an 125 allele is erroneously called as absent. To address this issue, we 126 augmented our model with a latent true genetic state Y, reflecting 127 the true presence or absence of alleles at each locus for each 128 individual. Augmenting our model with this latent state allowed 129 us to incorporate and model the uncertainty around measurement 130 of genetic data separately from the uncertainty around the true 131 genetic state, as expressed in the following factorization: 132

$$P(X_{i}|Y_{i},\mu_{i},r_{i},\epsilon_{i}^{+},\epsilon_{i}^{-},\pi) = P(X_{i}|Y_{i},\epsilon_{i}^{+},\epsilon_{i}^{-})P(Y_{i}|\mu_{i},r_{i},\pi) \quad (2) \quad 133$$

We assumed a prior in which the MOI of each individual was 134 independent of the MOI of other individuals, relatedness was 135

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136 independent across individuals, error rates were independent 137 across individuals, and allele frequencies were independent across 138 loci and without linkage disequilibrium, yielding the following

factorization: 139

<sup>140</sup> 
$$P(\boldsymbol{\mu}, \boldsymbol{r}, \boldsymbol{\epsilon}^+, \boldsymbol{\epsilon}^-, \boldsymbol{\pi}) = \prod_{i=1}^n P(\mu_i) P(r_i) P(\boldsymbol{\epsilon}_i^+) P(\boldsymbol{\epsilon}_i^-) \prod_{j=1}^l P(\pi_l) \quad (3)$$

Details of the likelihood and prior distributions are provided in 141 the supplementary material (section 1 and section 2), as well as 142 practical considerations when using MOIRE with real world data 143 144 (section 3).

#### Effective MOI 145

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By estimating MOI and within-host relatedness, we can estimate 146 a continuous metric of genetic diversity within a host, the effective 147 MOI (eMOI), which we define as: 148

$$eMOI = 1 + (1 - r)(\mu - 1)$$
(4)

One interpretation of the effective MOI is the expected number 150 of distinct alleles at a locus with infinite diversity, i.e. a locus 151 where heterozygosity is 1 (see supplementary section 6 for a formal 152 derivation). In the case of no within-host relatedness, this is 153 simply the MOI. However, when there is within-host relatedness, 154 155 the effective MOI is the MOI weighted by the probability that a given strain is unrelated to all other strains within the host, 156 157 and ranges from 1 to  $\mu$ . This value better reflects the true genetic diversity within a host than the MOI alone, and allows 158 for comparison and differentiation of genetic diversity across hosts 159 with the same MOIs. We also note that eMOI is likely to be more 160 identifiable than MOI or within-host relatedness alone because it 161 is a one-dimensional combination of the two estimated parameters 162 with synergistic properties around precision. As MOI increases, 163 precision around estimates of within-host relatedness also increases 164 as there are more observations available to inform within-host 165 relatedness. As MOI decreases, precision around the estimate of 166 within-host relatedness decreases, however the contribution to the 167 168 estimate of eMOI from within-host relatedness also decreases, and thus the overall precision of eMOI is maintained. 169

#### Inference and Implementation 170

We fit our model to observed genetic data using a Markov 171 Chain Monte Carlo (MCMC) approach using the Metropolis-172 Hastings algorithm with a variety of update kernels. Details of 173 sampling and implementation are described in the supplementary 174 material (section 5). MOIRE is implemented as an R package 175 and is available with tutorials and usage guidance at https: 176 //eppicenter.github.io/moire/. All sampling procedures were 177 178 implemented using Rcpp (Eddelbuettel and Francois, 2011) for efficiency. Substantial effort was placed on ease of use and 179 limiting the amount of tuning required by the user by leveraging 180 adaptive sampling methods. We provide weak default priors for 181 all parameters and recommend that users only modify priors if 182 they have strong prior knowledge about the parameters, such 183 as experimentally derived estimates of false positive and false 184 negative rates using samples with known parasite compositions 185 and densities. All analysis conducted in this paper was done 186 using MOIRE with default priors and settings, using 40 parallel 187

tempered chains for 5000 burn-in steps, followed by 10,000 samples which were thinned to 1000 total samples.

# Results

# Estimation of multiplicity of infection, within-host relatedness, and allele frequencies

We simulated collections of 100 samples under varied combinations 193 of population mean MOI, average within-host relatedness, false 194 positive and false negative rates, and different genotyping panels 195 (details of our simulation procedure may be found in the 196 supplement section 4). Individual MOIs were drawn from zero 197 truncated Poisson (ZTP) distributions with rate parameters 1, 3, 198 and 5, resulting in mean MOIs of 1.58, 3.16, and 5.03 respectively. 199 Within-host relatedness was simulated from settings with low, 200 moderate, and high relatedness. False positive and false negative 201 rates were varied from 0 to 0.1. We first simulated synthetic 202 genomic loci with prespecified diversity: 100 SNPs, 30 loci with 5 203 alleles (moderate diversity), 30 loci with 10 alleles (high diversity), 204 and 30 loci with 20 alleles (very high diversity) with frequencies 205 drawn from the uniform Dirichlet distribution. We also assessed 206 potential real world performance of MOIRE by simulating data for 207 5 currently used genotyping panels from 12 regional populations 208 characterized by the MalariaGEN Pf7 dataset (Abdel Hamid et al., 209 2023) as described in the supplementary material (section 7, 210 Supplementary Figure 4). Genetic loci were selected according to 211 a 24 SNP panel (Daniels et al., 2008), a 101 SNP panel (Chang 212 et al., 2019), and 3 recently developed amplicon sequencing panels 213 consisting of 128 (LaVerriere et al., 2022), 165 (Aranda-Diaz and 214 Neubauer Vickers, 2022), and 233 (Kattenberg et al., 2023) diverse 215 microhaplotypes respectively. Like the fully synthetic simulations, 216 these simulations were varied over a range of MOI and within-217 host relatedness, however error rates were fixed at moderate false 218 positive and false negative rates of .01 and .1 respectively for 219 the purposes of computational feasibility due to the extensive 220 number of simulations required. We chose these levels as we 221 believe they are reflective of the most likely situation of higher 222 levels of false negatives and relatively low rates of false positives 223 from a typical bioinformatics pipeline. We then ran MOIRE and 224 calculated summary statistics of interest on the sampled posterior 225 distributions. 226

We estimated allele frequencies, heterozygosity, MOI, within-host 227 relatedness, and eMOI using the mean or median of the posterior 228 distribution output by MOIRE. It should be noted that within-229 host relatedness is only defined for polyclonal infections, so the 230 posterior distribution of within-host relatedness is conditional 231 on the MOI being greater than 1. We contrasted these with 232 naive estimates of allele frequency and MOI by assuming that an 233 observed allele was contributed by a single strain, and estimated 234 MOI as equal to the second-highest number of alleles observed 235 across loci. We calculated ground truth allele frequencies using 236 the true number of strains contributing each allele. 237

Under moderate false positive and false negative rates of 0.01 238 and 0.1 respectively, MOIRE accurately recovered parameters of 239 interest across a range of genotyping panels, population MOI, 240 and within-host relatedness (Figure 1, Table 1). Allele frequencies 241 estimated by MOIRE were unbiased across genotyping panels 242 (Figure 1B), leading to unbiased estimates of heterozygosity 243 (Figure 1C). Naive estimation exhibited substantial bias that 244

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Panel	Source	Heterozygosity		Allele Freqs.		MOI		Relatedness	eMOI
		MOIRE	Naive	MOIRE	Naive	MOIRE	Naive	MOIRE	MOIRE
100 SNP	Synthetic	0.01 (.95)	0.05	0.02(.95)	0.06	1.72(.85)	3.61	0.20 (.70)	0.37(.77)
Moderate Div.	Synthetic	0.01 (.99)	0.04	0.01 (.96)	0.02	1.55(.88)	2.53	0.14(.77)	0.17(.91)
High Div.	Synthetic	0.01 (.99)	0.02	0.01(.91)	0.01	1.29(.86)	1.87	0.11(.75)	0.12(.89)
Very High Div.	Synthetic	0.02(.60)	0.01	$0.01 \ (.82)$	0.01	1.02(.86)	1.28	0.10(.77)	0.10(.85)
24 SNP	Daniels et al. (2008)	0.01 (.90)	0.04	0.02(.90)	0.05	1.95(.81)	3.66	0.21 (.75)	0.45(.86)
101 SNP	Chang <i>et al.</i> (2019)	0.01 (.90)	0.04	0.02(.95)	0.05	1.77(.85)	3.62	0.20(.71)	0.36(.79)
AMPLseq	LaVerriere et al. (2022)	0.01 (.97)	0.05	0.01(.94)	0.02	1.32(.88)	1.86	0.12(.71)	0.14(.88)
$MaD^4HatTeR$	Aranda-Diaz and Neubauer Vickers (2022)	0.01 (.98)	0.05	0.01 (.95)	0.03	1.24(.90)	1.88	0.13(.68)	0.12(.88)
AmpliSeq	Kattenberg et al. (2023)	0.02(.94)	0.06	$0.01 \ (.93)$	0.02	1.34(.83)	1.56	0.12(.67)	0.12(.88)

Table 1. Mean absolute deviation (MAD) of estimates of MOI, heterozygosity, within-host relatedness, and eMOI across simulations using synthetic (top) and real-world (bottom) genotyping panels. The MAD of estimates of MOI were calculated by taking the mean of the MAD for each stratum of true MOI between 1 and 10. MOI Within-host relatedness accuracy is only considered for samples with a true MOI > 1. Coverage rates of 95% credible intervals are shown in parentheses for estimates by MOIRE.



Fig. 1: True vs. estimated values of parameters across panels of varying genetic diversity. Panel A summarizes the distribution of heterozygosity across each panel used. Each symbol represents the estimated value of the parameter for a single simulated dataset, with the true value of the parameter on the x-axis and the estimated value on the y-axis. Simulations were pooled across mean MOIs and levels of relatedness. False positive and false negatives rates were fixed to 0.01 and 0.1 respectively. Opacity was set to accommodate overplotting, except in the case of withinhost relatedness where opacity reflects the estimated probability that a sample is polyclonal, calculated as the posterior probability of the sample MOI being greater than 1, as individual withinhost relatedness is only defined for samples with MOI greater than 1. MOIRE accurately recovered parameters of interest with increasing accuracy as panel diversity increased, while naive estimation exhibited substantial bias where such estimators exist.

varied with respect to the true allele frequency. Rare alleles tended
to be overestimated and common alleles underestimated, leading
to inflated estimates of heterozygosity.

MOI was also well estimated by MOIRE, with accuracy increasing 248 substantially in the presence of more diverse loci (Figure 1D). 249 In the context of SNPs, MOIRE recovered MOI accurately up 250 to approximately 4 strains, and then began to exhibit limited 251 ability to resolve. More diverse panels enabled greatly improved 252 resolving power, allowing for the accurate recovery of MOI 253 up to approximately 7 strains. Naive estimation substantially 254 underestimated MOI in comparison, due in part to the limited 255 capacity of low diversity loci to discriminate MOI, as well as 256 the presence of related strains that deflate the observed number 257 of distinct alleles. This bias was particularly prominent for low 258 diversity markers such as SNPs which can only resolve up to 2 259 strains. 260

MOIRE was generally able to recover within-host relatedness, 261 particularly for moderate and high diversity markers in the 262 context of high relatedness (Figure 1E). SNP based panels had 263 difficulty resolving individual level within-host relatedness and 264 were sensitive to the uniform prior. It should be noted that in 265 the circumstance that a monoclonal infection has an inferred MOI 266 greater than 1, MOIRE will likely classify these infections with 267 very high relatedness (Figure 1E). This is due to the presence of 268 false positives that MOIRE will sometimes infer as an infection 269 consisting of highly related strains rather than being explained 270 by observation error. Therefore, within-host relatedness should be 271 interpreted in the context of the probability of the infection being 272 polyclonal. A more robust metric is eMOI, since it is a metric of 273 diversity that integrates MOI and within-host relatedness. 274

MOIRE recovered eMOI with high accuracy under all conditions using polyallelic panels (Figure 1F). SNP panels exhibited a larger degree of bias at higher eMOI, but still performed relatively well for eMOI of up to 4. This demonstrates that while identifiability of MOI or within-host relatedness may be challenging in some situations, eMOI is a reliably identifiable quantity when estimated using highly polymorphic markers. 281

All simulations were also conducted without any relatedness 282 present. MOIRE was still able to accurately recover allele 283 frequencies, heterozygosity, and MOI, indicating that minimal 284 bias or uncertainty are introduced by attempting to estimate 285 relatedness (Supplementary Figure 1). 286

These patterns held across the range of false positive and false 287 negative rates simulated with the fully synthetic simulations. 288

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Allele frequencies and heterozygosity remained well estimated by MOIRE across settings, however bias was elevated for individual level estimates of MOI, within-host relatedness, and eMOI when false positive rates were increased and panel diversity was low.

<sup>293</sup> Increased false negative rates did not result in any additional bias

<sup>294</sup> within the range of tested values (Supplementary Figure 2).

### 295 Population inference

MOIRE is a probabilistic approach providing a full posterior 296 297 distribution over model parameters, allowing estimation of credible intervals for model parameters as well as functions thereof. While 298 sample level parameters estimated by the model are useful, it may 299 also be useful to estimate population level summary statistics 300 for reporting and comparison purposes. We thus calculated the 301 posterior distribution of population level summaries of interest, 302 303 such as mean MOI, mean within-host relatedness, and mean eMOI. We note that mean within-host relatedness is defined only 304 for samples with MOI greater than 1, therefore the posterior 305 distribution of mean within-host relatedness was calculated across 306 samples with MOI greater than 1 at each iteration of the MCMC 307 algorithm. MOIRE accurately estimated these quantities across 308 a range of conditions (Supplementary Figure 3), with the best 309 performance seen for polyallelic data. 310

Population mean MOI was accurately estimated across all panels, 311 with improved precision at lower levels (Supplementary Figure 3A, 312 Table 1). Credible interval (CI) coverage in general was poor, 313 likely due to the challenge of identifiability in conjunction with 314 within-host relatedness. SNP panels were largely unable to resolve 315 population level mean within-host relatedness and exhibited poor 316 CI coverage and substantial sensitivity to the uniform prior 317 specification due to the low relative information contained in these 318 markers. Polyallelic panels in contrast had improved precision as 319 more diverse panels were used, although CI coverage was also 320 poor due to persistent sensitivity to the uniform prior as indicated 321 by slightly overestimating within-host relatedness below .5 and 322 underestimating within-host relatedness above .5. 323

Population mean eMOI was remarkably accurate for low and 324 medium mean MOI when using SNP based panels, with bias 325 only becoming apparent at higher mean MOI (Supplementary 326 Figure 3C, Table 1). Polyallelic panels had substantially improved 327 precision across a wide range of values, further demonstrating that 328 while population mean within-host relatedness or mean MOI may 329 be challenging to identify, mean eMOI remains a highly identifiable 330 quantity when genetic markers with sufficient diversity are used. 331

#### 332 Metric stability across genetic backgrounds

Population metrics of genetic diversity enable researchers to make 333 comparisons across space and time, and to answer questions 334 relating to differences in transmission dynamics. In order for a 335 metric to be useful for these purposes, it must be sensitive to 336 changes in transmission dynamics while remaining insensitive to 337 other factors that vary and may confound interpretation, such as 338 the genotyping panel used, or the local allele frequencies for a 339 given panel. For example, if we were to compare two populations 340 that exhibit the same transmission dynamics, we would want the 341 metric to be the same, uninfluenced by differing population allele 342 frequencies. It would be even better if the metric is insensitive to 343 the genotyping panel used, allowing for comparisons across studies 344 that are independent of the technology utilized. 345



rig. 2. Comparison of mean endot to other summary measures of diversity across varying levels of within-host relatedness. For each level of relatedness (low and high), we simulated 100 infections with a mean MOI of 1.51 and 3.16, for a total of 400 infections across 4 conditions. Keeping the MOI and relatedness fixed for each sample, we varied the genetic diversity of the panel used to genotype each sample. We then calculated the mean eMOI from MOIRE, mean MOI using the naive estimator, and mean  $F_{WS}$  using a naive estimate of allele frequencies for each simulation to assess the sensitivity of each metric to varying the genetic diversity of the panel. True mean eMOI and mean MOI are fixed values within levels of within-host relatedness and are annotated by dashed lines. Mean  $F_{WS}$  is not fixed within levels of within-host relatedness and MOI because it is a function of the genetic diversity of the panel.

To explore the performance of eMOI across varying transmission 346 settings, we simulated 100 samples with MOI drawn from a ZTP 347 distribution with either  $\lambda = 1$  or  $\lambda = 3$ . For each sample, we 348 then simulated either low or high within-host relatedness. For each 349 individual level simulation, we then observed simulated genetics 350 parameterized by each of the 12 regional populations previously 351 described using the 5 genotyping panels, followed by the previously 352 described observation process with false positive and false negative 353 rates of .01 and .1 respectively. We then fit MOIRE on each 354 simulation independently. 355

For each simulation, we calculated mean eMOI, mean naive MOI, 356 and the within-host infection fixation index  $(F_{WS})$  (Roh et al., 357 2019; Manske et al., 2012), a frequently used metric of within-358 host diversity that relates genetic diversity of the individual 359 infection to diversity of the parasite population. Mean MOI was 360 calculated using the second-highest number of observed alleles, 361 and  $F_{WS}$  used the observed genetics, assuming all alleles were 362 equifrequent within hosts, and naive estimates of allele frequencies 363 to estimate heterozygosity. For these metrics to be most useful in 364 characterizing transmission dynamics, they should be the same for 365 all simulations with the same degree of within-host relatedness and 366 mean MOI, no matter the panel used nor the genetic background 367 of the population. We found that mean eMOI was stable across all 368 genetic backgrounds using microhaplotype based panels, yielding 369 accurate estimates of mean eMOI despite substantial variability in 370 local diversity of alleles, as shown by heterozygosity, and differing 371

372 genomic loci (Figure 2A). Interestingly, while the SNP panels 373 exhibited reduced precision and downward bias as expected, they were consistently biased with respect to the true eMOI, even across 374 different panels. This suggests that SNP panels, while limited 375 in resolving power, may still have utility in estimating relative 376 ordering of eMOI. These results also demonstrate that eMOI may 377 be readily used and compared across transmission settings and 378 is relatively insensitive to other factors such as heterozygosity 379 that may vary across settings. In contrast, mean naive MOI and 380  $F_{WS}$  were sensitive to genetic background and genotyping panel 381 in confounded ways. Mean naive MOI, only useful with polyallelic 382 markers, exhibits an inherent upward bias as mean heterozygosity 383 increases that is most severe at higher mean MOI. This bias 384 385 also varied with the genotyping panel used, making it difficult to interpret and compare across settings (Figure 2B).  $F_{WS}$  is 386 also sensitive to genetic background and panel used, exhibiting 387 an upward trend as heterozygosity increases and a bias that varies 388 across panels. This is inherent to the construction of the metric, as 389 it is coupled to an estimate of the true heterozygosity of genetic loci 390 being used (Figure 2C). This simulation demonstrates limitations 391 in the utility of  $F_{WS}$  as a metric of within-host diversity for a 392 population as it is inherently uncomparable across settings due to 393 its high sensitivity to varying genetic background and genotyping 394 panel used. Mean eMOI, in contrast, is a stable metric of genetic 395 diversity that is insensitive to genetic background and genotyping 396 panel, and is thus readily comparable across settings. 397

# 398 Application to a study in Northern Namibia

We next used MOIRE to reanalyze data from a previously 399 conducted study carried out in northeastern Namibia consisting 400 of 2585 samples from 29 health facilities across 4 health districts 401 genotyped at 26 microsatellite loci (Tessema et al., 2019). We ran 402 MOIRE across samples collected from each of the 4 health districts 403 independently. Running MOIRE in this way implies that we are 404 assuming that all samples from each health district come from 405 a shared population with the same allele frequencies. We then 406 407 calculated summary statistics of interest on the sampled posterior distributions. 408

We compared our results to the naive estimation conducted in 409 the original study and found that overall relative ordering of 410 mean MOI was maintained, with Andara and Rundu exhibiting 411 the highest MOI, Zambezi the lowest, and Nyangana in between, 412 consistent with contemporary estimates of transmission intensity 413 (Tessema et al., 2019). However, similar to our simulations, naive 414 estimation substantially underestimated mean MOI across health 415 districts compared to MOIRE (Figure 3A and C). Individual 416 within-host relatedness was estimated to be very high across 417 sites (IQR: .61-.91) with no differences between sites (Figure 3B). 418 This suggests substantial inbreeding which may be indicative of 419 420 persistent local transmission, consistent with the original findings by Tessema et al. (2019) We also found that heterozygosity 421 across loci estimated by MOIRE was generally lower (IQR: 422 .55 - .85), consistent with the previously described simulations 423 demonstrating that naive estimation overestimates heterozygosity, 424 and that previously detected statistically significant differences in 425 heterozygosity between the Zambezi region and the other three 426 regions may have been an artifact of biased estimation (Figure 3D). 427

428 We also ran MOIRE independently across each of the 29 health 429 facilities, excluding 2 health facilities from the Zambezi region



3: Estimated MOI, relatedness, eMOI Fig. and heterozygosity in Northern Namibia. MOIRE was run on data from 2585 samples from 29 clinics genotyped at 26 microsatellite loci, subset across four health districts. Each point represents the posterior mean or median for each sample or locus level parameter. The black circle represents the population mean with 95% credible interval for each health district and the black triangle indicates the naive estimate where applicable. In the case of eMOI (C), the naive estimate is simply the MOI. Opacity was used to accommodate overplotting in A, C and D, however opacity in B is reflective of the posterior probability of a particular sample being polyclonal to emphasize that an observation's contribution to the posterior distribution of mean within-host relatedness is weighted by its probability of being polyclonal. This is due to the fact that mean within-host relatedness is only defined for samples with MOI greater than 1, and thus the posterior distribution of within-host relatedness was calculated by taking the mean withinhost relatedness across samples with MOI greater than 1 at each iteration of the MCMC algorithm. Therefore, the opacity of each point in B is reflective of the contribution of that sample to the posterior distribution of mean within-host relatedness.

due to low total number of samples (n = 9 in each). Stratifying 430 by health facility revealed substantial heterogeneity in mean MOI, 431 within-host relatedness, and consequently eMOI, also consistent 432 with the findings by Tessema et al. (2019) (Figure 4). Interestingly, 433 Tessema et al. (2019) identified Rundu district hospital as having 434 exceptionally high within-host diversity as measured by  $F_{WS}$ , 435 which was posited to be due to a large fraction of the patients 436 having traveled or resided in Angola. We found that Rundu district 437 hospital had the highest mean eMOI and greatest spread across 438 observations (mean = 4.3 [95% CI: 4.18 - 4.4], IQR = 4.88). This 439 was mainly driven by much higher mean MOI (7 [95% CI: 6.5-440 (7.5], and low mean within-host relatedness (.47 [95% CI: .43 -441 0.51]). The combination of high MOI and relatively low within-442 host relatedness, translating into high population mean eMOI, 443

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Fig. 4: Estimated MOI, relatedness, eMOI and heterozygosity in Northern Namibia, stratified by health facility. MOIRE was run independently on data from each health facility. Two health facilities from the Zambezi region were excluded due to only having 9 samples present in each subset. Health facilities are plotted in geographic order from West to East. Plotting conventions are the same as in Figure 3.

suggests that samples collected here reflect a parasite population
experiencing less inbreeding and more superinfection, which may
be indicative of higher transmission intensity in tandem with a
larger effective population size.

## 448 Discussion

Translating *Plasmodium* genetic data from naturally acquired 449 infections into meaningful insights about population genetics or 450 malaria transmission dynamics often begins with estimation of 451 allele frequencies and MOI. We demonstrated through simulation 452 that naive estimation introduces substantial biases, rendering 453 estimation unreliable and uncomparable between settings. 454 In particular, naive estimation systematically overestimates 455 measures of allelic diversity such as heterozygosity and 456 systematically underestimates MOI. State-of-the-art methods 457 previously available to more accurately estimate individual level 458 459 MOI and population allele frequencies only allow for SNP based data, and fail to directly consider within-host relatedness as 460 an important biologic factor (Chang et al., 2017; Paschalidis 461 et al., 2023; Ju et al., 2023). MOIRE fills these important 462 gaps, demonstrating both the power and necessity of polyallelic 463 data to obtain precise estimates of these key parameters for 464 understanding of parasite population structure and dynamics. The 465 R package implementing MOIRE provides a user-friendly interface 466 for researchers to easily leverage SNP and polyallelic data to 467 estimate these individual and population diversity metrics which 468 are fundamental for many downstream analyses and often of direct 469 470 interest themselves.

471 By estimating within-host relatedness, we also have introduced 472 a new metric of diversity—eMOI—a continuous metric that 473 integrates within-host relatedness and MOI, providing the first 474 portable metric of within-host diversity. This metric is highly 475 identifiable and robust to varying genetic backgrounds, and thus 476 readily comparable across settings and genotyping technologies. 477 We demonstrated that eMOI is a more stable metric of genetic diversity than naive MOI or  $F_{WS}$ , and is insensitive to other 478 factors that may vary across settings such as allele frequencies 479 of given genetic markers. Further, by decomposing the genetic 480 state of an infection into components of within-host relatedness 481 and the number of distinct strains present, we have enabled 482 the characterization of these quantities independently, which 483 may be of interest in their own right. For example, within-host 484 relatedness may be of interest in the context of understanding the 485 role of inbreeding and co-transmission in the parasite population 486 (Wong et al., 2022; Nkhoma et al., 2020), and the number of 487 distinct strains may be of interest in the context of understanding 488 superinfection dynamics. 489

While we have demonstrated the utility of polyallelic data, MOIRE 490 is still compatible with SNP based data and can offer benefits 491 over other approaches. When using SNP based panels, eMOI 492 is still well characterized up to moderate levels, and while the 493 reduced capacity of SNPs generally results in biased estimates, 494 the estimates recovered reflect changes in within-host relatedness 495 yet are stable across genetic backgrounds. Thus, these data may be 496 useful for comparing relative ordering of eMOI across settings and 497 providing inference. In contrast, existing analytical approaches are 498 likely to be sensitive to model misspecification by not considering 499 within-host relatedness and varying genetic backgrounds, and may 500 be biased in ways that are difficult to interpret and compare across 501 settings. 502

We also note that while increasing the number of loci genotyped 503 is always beneficial, the largest gains in recovering estimates of 504 interest are through using sufficiently diverse loci. Our simulations 505 demonstrate that, even with a modest number of very diverse 506 loci such as our synthetic simulations using 30 loci, eMOI can be 507 recovered with a high accuracy and precision. Marginal increases in 508 complexity of incorporating several highly diverse loci, for example 509 in the context of drug resistance monitoring, may be outweighed 510 by the substantial insights obtained from jointly understanding 511 transmission dynamics, population structure, and drug resistance 512 through increased accuracy of estimating resistance marker 513 allele frequency. Modern amplicon sequencing panels have been 514 developed precisely for these contexts, combining high diversity 515 targets with comprehensive coverage of known resistance markers 516 (LaVerriere et al., 2022; Aranda-Diaz and Neubauer Vickers, 2022; 517 Kattenberg et al., 2023). 518

MOIRE provides a powerful tool for leveraging polyallelic data 519 to understand malaria epidemiology, and there are multiple 520 avenues for future work to further improve inference. First, 521 the observation model does not currently fully leverage the 522 information in sequencing based data where the actual number of 523 reads may be available. This may provide additional information, 524 e.g. to inform false positive rates by considering the number of 525 reads attributable to an allele, as well as false negative rates by 526 considering the total number of reads at a locus which may be 527 indicative of sample quality. Second, we currently consider only a 528 single, well mixed, background population parameterized by allele 529 frequencies at each locus. However, it may be the case that there 530 are multiple distinct populations with their own allele frequencies, 531 and that the observed data is a mixture of these populations. This 532 may be particularly relevant in the context of malaria transmission 533 where there may be multiple distinct populations of parasites 534 circulating in a region. Future work may consider a mixture 535 model over allele frequencies, where the number of populations 536

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537 is a priori specified or determined through data adaptive nonparametric Bayesian modeling, and thereby identify population substructure. Alternatively, a spatially explicit approach may be 530 feasible that would model the allele frequencies as a function 540 of geographic location, potentially enabling resolving geographic 541 origin of parasites within observed infections. Third, MOIRE 542 currently assumes independence of loci. In the case of locus 543 dependence where there is some amount of linkage disequilibrium, 544 we would expect estimates of allele frequencies and sample specific 545 eMOI to still be consistent if there is not a systematic bias in loci 546 towards regions of high or low within-host relatedness. 547

In summary, MOIRE enables the use of polyallelic data to 548 estimate allele frequencies, MOI, and within-host relatedness, and 549 provides a new metric of genetic diversity, the eMOI. We have 550 demonstrated that eMOI has improved utility, interpretability, 551 and stability across simulated transmission settings than existing 552 metrics of within-host diversity such as  $F_{WS}$ . Furthermore, 553 we demonstrated the utility of MOIRE through simulation and 554 reanalysis of previously collected data, and have provided an R 555 package to enable researchers to easily leverage polyallelic data 556 to make inferences about malaria population dynamics. MOIRE 557 also serves as a fundamental building block for future work, 558 as it provides a principled approach to jointly estimate allele 559 frequencies, MOI, and within-host relatedness from polyallelic 560 data, which can be used as a basis for more complex modeling 561 of population dynamics. These methods may also be of utility 562 for other pathogens where superinfection is common, such as 563 schistosomiasis or filarial diseases (Aemero et al., 2015; Hedtke 564 et al., 2020). 565

## 566 Competing interests

567 No competing interest is declared.

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