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Antibiotics for prelabour rupture of membranes at or near term (Review)

Wojcieszek AM, Stock OM, Flenady V

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[Intervention Review]

Antibiotics for prelabour rupture of membranes at or near term

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ABSTRACT

Background

Prelabour rupture of the membranes (PROM) at or near term (defined in this review as 36 weeks' gestation or beyond) increases the risk of infection for the woman and her baby. The routine use of antibiotics for women at the time of term PROM may reduce this risk. However, due to increasing problems with bacterial resistance and the risk of maternal anaphylaxis with antibiotic use, it is important to assess the evidence addressing risks and benefits in order to ensure judicious use of antibiotics. This review was undertaken to assess the balance of risks and benefits to the mother and infant of antibiotic prophylaxis for PROM at or near term.

Objectives

To assess the effects of antibiotics administered prophylactically to women with PROM at 36 weeks' gestation or beyond, on maternal, fetal and neonatal outcomes.

Search methods

We searched the Cochrane Pregnancy and Childbirth Group's Trials Register (31 July 2014).

Selection criteria

All randomised trials that compared outcomes for women and infants when antibiotics were administered prophylactically for prelabour rupture of the membranes at or near term, with outcomes for controls (placebo or no antibiotic).

Data collection and analysis

Two review authors independently extracted the data and assessed risk of bias in the included studies. Additional data were received from the investigators of included studies.

Main results

This update includes an additional two studies involving 1801 women, giving a total of four included studies of 2639 women. Whereas the previous version of this review showed a statistically significant reduction in endometritis with the use of antibiotics, no such effect was shown in this update (average risk ratio (RR) 0.34, 95% confidence interval (Cl) 0.05 to 2.31). No differences were shown on the primary outcome measures of probable early-onset neonatal sepsis (average RR 0.69, 95%; Cl 0.21 to 2.33); definite early-onset neonatal sepsis (average RR 0.57, 95% Cl 0.08 to 4.26); maternal infectious morbidity (chorioamnionitis and/or endometritis) (average RR 0.48, 95% Cl 0.20 to 1.15); stillbirth (RR 3.00, 95% Cl 0.61 to 14.82); and perinatal mortality (RR 1.98, 95% Cl 0.60 to 6.55), though the number of cases in the control group for these outcomes was low. There were no cases of neonatal mortality or serious maternal outcome in the studies assessed. Caesarean section was increased with the use of antibiotics (RR 1.33, 95% Cl 1.09 to 1.61) as was duration of maternal stay in hospital



(mean difference (MD) 0.06 days, 95% CI 0.01 to 0.11), largely owing to one study of 1640 women where repeat caesarean section, increased baseline hypertension and pre-eclampsia were evident in the antibiotic group, despite random allocation and allocation concealment.

Subgroup analyses by timing of induction (early induction versus late induction) showed no difference in either probable or definite earlyonset neonatal sepsis in the early induction group (RR 1.47, 95% CI 0.80 to 2.70 and RR 1.29, 95% CI 0.48 to 3.44, respectively) or the late induction group (RR 0.14, 95% CI 0.02 to 1.13 and RR 0.16, 95% CI 0.02 to 1.34, respectively), although there were trends toward reduced probable and definite early-onset neonatal sepsis in the late induction group. A test for subgroup differences confirmed a differential effect of the intervention on probable early-onset neonatal sepsis between the subgroups ($Chi^2 = 4.50$, df = 1 (P = 0.03), I² = 77.8%). No difference in maternal infectious morbidity (chorioamnionitis and/or endometritis) was found in either subgroup, though again there was a trend towards reduced maternal infectious morbidly in the late induction group (average RR 0.34, 95% CI 0.08 to 1.47). No differences were shown in stillbirth or perinatal mortality. The quality of the evidence for the primary outcomes using GRADE was judged to be low to very low.

Authors' conclusions

This updated review demonstrates no convincing evidence of benefit for mothers or neonates from the routine use of antibiotics for PROM at or near term. We are unable to adequately assess the risk of short- and long-term harms from the use of antibiotics due to the unavailability of data. Given the unmeasured potential adverse effects of antibiotic use, the potential for the development of resistant organisms, and the low risk of maternal infection in the control group, the routine use of antibiotics for PROM at or near term in the absence of confirmed maternal infection should be avoided.

PLAIN LANGUAGE SUMMARY

Antibiotics for rupture of membranes when a pregnant women is at or near term but not in labour

Background

Sometimes the protective bag of fluid around an unborn baby (the membranes) break when the baby is due without the onset of labour (regular uterine contractions). This is called PROM or prelabour rupture of the membranes. When this happens there is a risk of infection entering the womb (uterus) and affecting the mother and her baby. Newborn infections are rare but have the potential to cause serious harm requiring neonatal intensive care. Giving a pregnant woman antibiotics when she has PROM may reduce the risk of infections for the woman and her baby. Most women spontaneously start labour within 24 hours, so delaying induction of labour and waiting for spontaneous onset of labour (expectant management) may be a possibility. Another treatment for term PROM is to induce labour with oxytocin or prostaglandins. Women are often given antibiotics to prevent infection, but there are concerns about possible side-effects of antibiotics, and that overuse of antibiotics can cause resistance to antibiotics so that they become less effective.

Our review questions

Do antibiotics given to women with PROM when they are at or near term (more than 36 weeks' gestation) but not in labour reduce the risk of infection for the baby and the mother? Are there adverse effects from the antibiotics?

What the studies showed

This review included four randomised controlled studies involving 2639 pregnant women at 36 weeks' gestation or more. The evidence showed that routine antibiotics for term PROM did not reduce the risk of infection for pregnant women or their babies when compared to the control group which received a placebo or no antibiotics. There was not enough strong evidence about other outcomes including death, allergic reactions for the woman or complications for the baby, which rarely occurred in the included studies. The quality of the evidence using GRADE was judged to be low to very low.

Overall

The conclusions from this review are limited by the low number of women who developed an infection across the studies overall. There is not enough information in this review to assess the possible side-effects from the use of antibiotics for women or their infants, particularly for any possible long-term harms. Because we do not know enough about side-effects and because we did not find strong evidence of benefit from antibiotics, they should not be routinely used for pregnant women with ruptured membranes prior to labour at term, unless a woman shows signs of infection.

SUMMARY OF FINDINGS

Summary of findings for the main comparison. Any antibiotic compared with placebo or no antibiotic (subgrouped by timing of induction of labour) for prelabour rupture of membranes at or near term

Any antibiotic compared with no antibiotic (subgrouped by timing of induction of labour) for prelabour rupture of membranes at or near term

Population: pregnant women with prelabour rupture of membranes at or near term

Settings: maternity hospitals in Spain, Egypt, Chile, Portugal

Intervention: any antibiotic

Comparison: placebo or no antibiotic

Outcomes	Illustrative comparative risks* (95% CI)		Relative effect (95% CI)	No of Partici- pants	Quality of the evidence	Comments
	Assumed risk	Corresponding risk	- (33 /0 Cl)	(studies)	(GRADE)	
	Control	Any antibiotic compared with no antibiotic (subgrouped by timing of induction of labour)	-			
Probable early-onset neona- tal sepsis - Early induction of labour	Study population		RR 1.47 - (0.8 to 2.7)	1640 (1 study)	⊕⊕⊝⊝ low ¹	
	21 per 1000	30 per 1000 (17 to 56)	(0.8 to 2.7)	(I study)	low -	
	Moderate					
	21 per 1000	31 per 1000 (17 to 57)				
Probable early-onset neonatal sepsis - Late induction of labour	Study population		RR 0.14 (0.02 to 1.13)	838 (2 studies)	⊕⊝⊝⊝ very low ^{2,3}	
sepsis - Late induction of labour	17 per 1000	2 per 1000 (0 to 19)	(0.02 (0 1.13)	(2 studies)		
	Moderate					
	10 per 1000	1 per 1000 (0 to 11)				
Definite early-onset neona- tal sepsis - Early induction of	Study population		RR 1.29 (0.48 to 3.44)	1640 (1 study)	⊕⊕⊝⊝ low 1	
labour	9 per 1000	11 per 1000	- (0.10 (0 3.11)	(I Study)		

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		(4 to 29)			
	Moderate				
	9 per 1000	12 per 1000 (4 to 31)			
Definite early-onset neonatal sepsis - Late induction of labour			RR 0.16 (0.02 to 1.34)	838 (2 studies)	⊕000 very low ^{2,3}
	15 per 1000	2 per 1000 (0 to 20)	- (0.02 10 1.34)		
	Moderate				
	8 per 1000	1 per 1000 (0 to 11)			
Maternal infectious morbidity (chorioamnionitis and/or en-	Study population		RR 1.15 (0.64 to 2.08)	1640 (1 study)	⊕⊕⊙⊙ low 1
dometritis) - Early induction of labour	24 per 1000	28 per 1000 (16 to 51)	(0.01 to 2.00)	(0.04.002.00) (130009)	1010 -
	Moderate				
	24 per 1000	28 per 1000 (15 to 50)			
Maternal infectious morbidity (chorioamnionitis and/or en- dometritis) - Late induction of labour	Study population	population		838 (2 studies)	⊕ooo very low ^{2,3}
	70 per 1000	24 per 1000 (6 to 103)		(0.08 to 1.47) (2 studies) very (
	Moderate				
	109 per 1000	37 per 1000 (9 to 160)			

*The basis for the **assumed risk** (e.g. the median control group risk across studies) is provided in footnotes. The **corresponding risk** (and its 95% confidence interval) is based on the assumed risk in the comparison group and the **relative effect** of the intervention (and its 95% CI). **CI:** Confidence interval; **RR:** Risk ratio;

GRADE Working Group grades of evidence

High quality: Further research is very unlikely to change our confidence in the estimate of effect.

Moderate quality: Further research is likely to have an important impact on our confidence in the estimate of effect and may change the estimate.

Low quality: Further research is very likely to have an important impact on our confidence in the estimate of effect and is likely to change the estimate.

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Trusted evidence. Informed decisions. Better health. Very low quality: We are very uncertain about the estimate.

¹ Wide confidence interval crossing the line of no effect and less than the optimal information size.

² Unclear and high risk of bias for allocation concealment.

³ Wide confidence interval crossing the line of no effect, few events and small sample size.

Summary of findings 2. Any antibiotic compared with placebo or no antibiotic for prelabour rupture of membranes at or near term

Any antibiotic compared with no antibiotic for prelabour rupture of membranes at or near term

Population: pregnant women with prelabour rupture of membranes at or near term Settings: maternity hospitals in Spain, Egypt, Chile, Portugal Intervention: any antibiotic

Comparison: placebo or no antibiotic

Outcomes	Illustrative comp Assumed risk	oarative risks* (95% CI) Corresponding risk	Relative effect 95% CI)	No of Partici- pants (studies)	Quality of the evidence (GRADE)	Comments
	Control	Any antibiotic compared with no antibi- otic	-			
Stillbirth	Study population		RR 3 (0.61 to 14.82)	1906 (3 studies)	$\oplus \oplus \odot \odot$ low 1	
	2 per 1000	6 per 1000 (1 to 31)	(0.01 10 14.02)	(3 studies)	low -	
	Moderate					
	0 per 1000	0 per 1000 (0 to 0)				
Perinatal mor- tality	Study population	1	RR 1.98 - (0.60 to 6.55)	2639 (4 studies)	⊕⊕⊙© low ²	
taity	3 per 1000	6 per 1000 (2 to 20)				
	Moderate					

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Neonatal mor- tality	See comment	See comment	Not estimable	1906 (3 studies)	See comment	There were no cases of neonatal mortality in the trials assessed.
Serious mater- nal outcome	See comment	See comment	Not estimable	1906 (3 studies)	See comment	There were no cases of se- rious maternal outcome in the trials assessed.
***	accuracy rick (a +	he meetice control group viels corece studies)	is provided in feetne	tes. The correspo	nding risk (and its 9 ^r	5% confidence interval) is
based on the assu		he median control group risk across studies) parison group and the relative effect of the i				

² Wide confidence interval crossing the line of no effect, few events.

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BACKGROUND

Description of the condition

Prelabour rupture of membranes (PROM) is defined as rupture of membranes prior to the onset of labour (Duff 1998). Approximately 8% of pregnant women at term experience PROM (Cammu 1990), but the decision as to how term PROM should be managed clinically remains controversial, and there is wide variation in practice with no clear consensus on what constitutes optimal treatment (Marowitz 2007). Although for the majority of women labour will start spontaneously within 24 hours following term PROM, up to 4% of women will not experience spontaneous onset of labour within seven days. While, traditionally, 'term' is defined as gestations of 37 to 41 weeks inclusive, for the purposes of this review infants born at 36 weeks' gestation have been included in the definition of term PROM, as many neonatal outcomes at these gestations are similar to those of gestations greater than 36 weeks (Marshall 2002; Neerhoff 1999).

Description of the intervention

Induction of labour is one strategy intended to reduce infectious morbidity associated with term PROM. There is some evidence that planned management (usually by induction) reduces the risk of infectious maternal morbidity and the number of infants requiring admission to neonatal intensive care (Dare 2006). On the basis of this evidence, clinicians may offer prompt induction with oxytocin or prostaglandins for term PROM (Hannah 1996; Mozurkewich 2006). Others utilise a policy of delayed induction of labour or expectant management (awaiting spontaneous onset of labour), although there remains much discussion regarding the optimum latent period for expectant management (Ezra 2004; Ingemarsson 1996). It is in situations where the latency period from membrane rupture to birth is prolonged that prelabour antibiotics may be beneficial. However, due to increasing problems with bacterial resistance (ACOG 2011; Edwards 2001; Lin 1999) and the rare but potentially life-threatening risk of maternal anaphylaxis with antibiotic use (ACOG 2011; Heim 1991; Jao 2006), it is essential to ensure that antibiotics are used judiciously and when indicated.

How the intervention might work

The reasons for term PROM are not clearly understood. However, subclinical ascending infection is thought to be influential and has been detected in up to one-third of women with term PROM (Romero 1992). There is also evidence of changes that happen within and between the cells of the fetal membranes (Moore 2006; Reti 2007). Despite the antibacterial properties of amniotic fluid, there is an increased risk of infection for the woman and her infant following term PROM (Newton 1993). Neonatal infections in the term population are rare occurrences (2% to 4%), but have the potential for causing mortality or serious morbidity (including the need for neonatal intensive care and mechanical ventilation). Prophylactic antibiotics can be effective against many (but not all) potential pathogens (ACOG 2011). Therefore, the routine use of antibiotics for women at the time of term PROM may reduce the risk of infection for the woman and her baby.

Why it is important to do this review

This review aims to address the balance of risks and benefits to the mother and infant of antibiotic prophylaxis for prelabour rupture

of the membranes at or near term. The review will also explore the differential effects of antibiotics and induction of labour.

Although Group B Streptococcus (GBS) is the most common cause of serious neonatal infection in the first seven days of life, this review does not address the role of intrapartum antibiotic GBS prophylaxis as this is a separate clinical question from prelabour antibiotic usage. The role of intrapartum antibiotic prophylaxis for GBS is addressed in another Cochrane review (Ohlsson 2014).

OBJECTIVES

To assess the effects of antibiotics administered prophylactically to women with prelabour rupture of the membranes at 36 weeks' gestation or beyond, on maternal, fetal and neonatal outcomes.

METHODS

Criteria for considering studies for this review

Types of studies

All randomised trials that compared outcomes for women and infants when antibiotics were administered prophylactically for prelabour rupture of the membranes at or near term, with outcomes for controls (placebo or no antibiotic). Studies using a quasi-random allocation method and cross-over design were excluded. Cluster-randomised trials were eligible for inclusion.

Types of participants

Women with spontaneous rupture of the fetal membranes prior to the onset of regular uterine contractions at 36 weeks' gestation or beyond.

Types of interventions

Any antibiotics, administered as prophylaxis, by any route, to women at 36 weeks' gestation or beyond, with prelabour rupture of the membranes.

Types of outcome measures

In this review we aimed to examine the effect of prophylactic antibiotics on clinically important outcome measures of maternal and infant morbidity and mortality, and maternal adverse effects. We also explored cost effectiveness and health service utilisation.

Primary outcomes

Primary outcomes were chosen to be most representative of the clinically important measures of effectiveness and complications. We assessed the primary outcome of maternal infectious morbidity (chorioamnionitis and/or endometritis) collectively, and explored its individual components as secondary outcomes. Primary outcomes were:

- 1. probable early-onset neonatal sepsis;
- 2. definite early-onset neonatal sepsis;
- maternal infectious morbidity (chorioamnionitis and/or endometritis);
- 4. stillbirth;
- 5. perinatal mortality;
- 6. neonatal mortality;



7. serious maternal outcome (defined as death, cardiac arrest, respiratory arrest, anaphylaxis, admission to intensive care unit).

Secondary outcomes

Secondary outcomes include other measures of morbidity and mortality, effectiveness, complications, health service utilisation and cost effectiveness.

Maternal outcomes

- 1. Chorioamnionitis (either suspected or proven);
- 2. endometritis;
- 3. caesarean section;
- 4. operative vaginal birth;
- 5. internal fetal monitoring;
- 6. epidural analgesia;
- 7. postpartum haemorrhage;
- 8. postpartum pyrexia;
- 9. postpartum septicaemia;
- 10.wound infection;
- 11.postpartum antibiotic usage;
- 12.breastfeeding on discharge from hospital;
- 13.adverse effects.

Neonatal outcomes

- 1. Neonatal meningitis;
- 2. neonatal pneumonia;
- 3. Apgar score less than seven at five minutes;
- 4. admission to neonatal special care nursery;
- 5. admission to neonatal intensive care unit;
- 6. antibiotic usage;
- 7. respiratory distress syndrome; and
- 8. use of mechanical ventilation.

Cost effectiveness

Health service utilisation

- 1. Duration of maternal stay in hospital (days);
- 2. duration of neonatal stay in hospital (days);
- 3. duration of neonatal stay in intensive care unit (days).

A priori subgroup analyses

- 1. Nulliparae;
- early induction of labour (less than 12 hours from rupture of membranes);
- 3. late induction of labour (at 12 hours or greater from rupture of membranes).

Outcome definitions

Suspected or proven chorioamnionitis: uterine infection prior to birth of the baby diagnosed on clinical signs, including pyrexia with or without a positive culture result or haematological signs of infection.

Endometritis: clinical signs of uterine infection following labour and birth.

Maternal pyrexia: maternal temperature of 38 degrees Celsius or higher.

Postpartum septicaemia: maternal positive blood culture in the presence of pyrexia following birth of the baby.

Early onset sepsis: definite or probable infection within the first seven days of life.

Definite infection: positive culture from a normally sterile site.

Probable infection: clinical signs and blood count suggestive of infection and a possible causative organism identified (i.e. gastric aspirate, urine antigen).

Search methods for identification of studies

The following methods section of this review is based on a standard template used by the Cochrane Pregnancy and Childbirth Group.

Electronic searches

We searched the Cochrane Pregnancy and Childbirth Group's Trials Register by contacting the Trials Search Co-ordinator (31 July 2014).

The Cochrane Pregnancy and Childbirth Group's Trials Register is maintained by the Trials Search Co-ordinator and contains trials identified from:

- 1. monthly searches of the Cochrane Central Register of Controlled Trials (CENTRAL);
- 2. weekly searches of MEDLINE;
- 3. weekly searches of Embase;
- 4. handsearches of 30 journals and the proceedings of major conferences;
- 5. weekly current awareness alerts for a further 44 journals plus monthly BioMed Central email alerts.

Details of the search strategies for CENTRAL, MEDLINE and Embase, the list of handsearched journals and conference proceedings, and the list of journals reviewed via the current awareness service can be found in the 'Specialized Register' section within the editorial information about the Cochrane Pregnancy and Childbirth Group.

Trials identified through the searching activities described above are each assigned to a review topic (or topics). The Trials Search Coordinator searches the register for each review using the topic list rather than keywords.

We did not apply any language restrictions.

Searches carried out in the previous version of the review are listed in Appendix 1.

Data collection and analysis

For the methods used when assessing the studies identified in a previous version of this review, *see* Flenady 2002.

For this update, the following methods were used for assessing the two studies that were identified as a result of the updated search. Where required, information pertaining to the two previous studies was updated according to methods outlined in the *Cochrane* Handbook for Systematic Reviews of Interventions (Higgins 2011).



Selection of studies

Two review authors independently assessed for inclusion all the potential studies identified as a result of the search strategy. We resolved any disagreement through discussion or, if required, we consulted the third review author.

Data extraction and management

At least two review authors extracted the data using the agreed form. We resolved discrepancies through discussion or, if required, we consulted the third review author. Data were entered into Review Manager software (RevMan 2014) and checked for accuracy.

When information regarding any of the above was unclear, we contacted authors of the original reports to provide further details.

Assessment of risk of bias in included studies

Two review authors independently assessed risk of bias for each study using the criteria outlined in the *Cochrane Handbook for Systematic Reviews of Interventions* (Higgins 2011). Any disagreement was resolved by discussion or by involving a third assessor.

(1) Random sequence generation (checking for possible selection bias)

We described for each included study the method used to generate the allocation sequence in sufficient detail to allow an assessment of whether it should produce comparable groups.

We assessed the method as:

- low risk of bias (any truly random process, e.g. random number table; computer random number generator);
- high risk of bias (any non-random process, e.g. odd or even date of birth; hospital or clinic record number);
- unclear risk of bias.

(2) Allocation concealment (checking for possible selection bias)

We described for each included study the method used to conceal allocation to interventions prior to assignment and assessed whether intervention allocation could have been foreseen in advance of, or during recruitment, or changed after assignment.

We assessed the methods as:

- low risk of bias (e.g. telephone or central randomisation; consecutively numbered sealed opaque envelopes);
- high risk of bias (open random allocation; unsealed or nonopaque envelopes, alternation; date of birth);
- unclear risk of bias.

(3.1) Blinding of participants and personnel (checking for possible performance bias)

We described for each included study the methods used, if any, to blind study participants and personnel from knowledge of which intervention a participant received. We considered that studies were at low risk of bias if they were blinded, or if we judged that the lack of blinding unlikely to affect results. We assessed blinding separately for different outcomes or classes of outcomes.

We have assessed the methods as:

low, high or unclear risk of bias for participants;

low, high or unclear risk of bias for personnel.

(3.2) Blinding of outcome assessment (checking for possible detection bias)

We described for each included study the methods used, if any, to blind outcome assessors from knowledge of which intervention a participant received. We assessed blinding separately for different outcomes or classes of outcomes.

We have assessed methods used to blind outcome assessment as:

low, high or unclear risk of bias.

(4) Blinding of outcome assessment (checking for possible detection bias)

We described for each included study the methods used, if any, to blind outcome assessors from knowledge of which intervention the participant received.

We assessed the methods as:

- low risk of bias (e.g. blinding, and unlikely that the blinding could have been broken, or no blinding or incomplete blinding, but outcome unlikely to be influenced);
- high risk of bias (e.g., no blinding, incomplete or broken blinding, and outcome likely to be influenced);
- unclear risk of bias.

(5) Incomplete outcome data (checking for possible attrition bias due to the amount, nature and handling of incomplete outcome data)

We described for each included study, and for each outcome or class of outcomes, the completeness of data including attrition and exclusions from the analysis. We stated whether attrition and exclusions were reported and the numbers included in the analysis at each stage (compared with the total randomised participants), reasons for attrition or exclusion where reported, and whether missing data were balanced across groups or were related to outcomes. Where sufficient information was reported, or could be supplied by the study authors, we planned to re-include missing data in the analyses which we undertook.

We assessed methods as:

- low risk of bias (e.g. no missing outcome data; missing outcome data balanced across groups);
- high risk of bias (e.g. numbers or reasons for missing data imbalanced across groups; 'as treated' analysis done with substantial departure of intervention received from that assigned at randomisation);
- unclear risk of bias.

(6) Selective reporting (checking for reporting bias)

We described for each included study how we investigated the possibility of selective outcome reporting bias and what we found.

We assessed the methods as:



- low risk of bias (where it is clear that all of the study's prespecified outcomes and all expected outcomes of interest to the review have been reported);
- high risk of bias (where not all the study's pre-specified outcomes have been reported; one or more reported primary outcomes were not pre-specified; outcomes of interest are reported incompletely and so cannot be used; study fails to include results of a key outcome that would have been expected to have been reported);
- unclear risk of bias.

(7) Other bias (checking for bias due to problems not covered by (1) to (6) above)

We described for each included study any important concerns we had about other possible sources of bias.

Assessment of the body of evidence - The GRADE approach

For this update we also evaluated the quality of the evidence using the GRADE approach (Schunemann 2009). The GRADE approach uses five considerations (study limitations, consistency of effect, imprecision, indirectness and publication bias) to assess the quality of the body of evidence for specific outcomes. The evidence can be downgraded from 'high quality' by one level for serious (or by two levels for very serious) limitations, depending on assessments for risk of bias, indirectness of evidence, serious inconsistency, imprecision of effect estimates or potential publication bias. In this review we used the GRADE approach to assess the seven primary outcomes, as follows:

- 1. probable early-onset neonatal sepsis;
- 2. definite early-onset neonatal sepsis;
- maternal infectious morbidity (chorioamnionitis and/or endometritis);
- 4. stillbirth;
- 5. perinatal mortality;
- 6. neonatal mortality;
- 7. serious maternal outcome (defined as death, cardiac arrest, respiratory arrest, anaphylaxis, admission to intensive care unit).

'Summary of findings' table

We used GRADE Profiler (GRADE 2008) to import data from Review Manager 5.3 (RevMan 2014) in order to create a 'Summary of findings' table. A summary of the intervention effect and a measure of quality according to the GRADE approach is presented in the 'Summary of findings' table for each of the above outcomes. Due to high heterogeneity in the analysis of the first three primary outcomes (probable early-onset neonatal sepsis; definite earlyonset neonatal sepsis; maternal infectious morbidity), a 'Summary of findings' table was produced presenting the results for these three outcomes subgrouped by timing of induction of labour (early induction versus late induction).

Measures of treatment effect

Dichotomous data

For dichotomous data, we presented results as summary risk ratio with 95% confidence intervals.

Continuous data

We used the mean difference if outcomes were measured in the same way between studies. We planned to use the standardised mean difference (SMD) to combine studies that measured the same outcome, but used different methods. However, SMD was not used in this update.

Unit of analysis issues

Cluster-randomised trials

We did not identify any cluster-randomised trials for inclusion in this review, but we may include trials of this type in future updates.

If cluster-randomised trials are included in future reviews, we plan to include cluster-randomised trials in the analyses along with individually-randomised trials. Their sample sizes will be adjusted using the methods described in the Handbook (Higgins 2011) using an estimate of the intra-cluster correlation co-efficient (ICC) derived from the trial (if possible), or from another source. If ICCs from other sources are used, we will report this and conduct sensitivity analyses to investigate the effect of variation in the ICC. If we identify both cluster-randomised trials and individuallyrandomised trials, we plan to synthesise the relevant information. We consider it reasonable to combine the results from both if there is little heterogeneity between the study designs and the interaction between the effect of intervention and the choice of randomisation unit is considered to be unlikely. We will also acknowledge heterogeneity in the randomisation unit and perform a sensitivity analysis to investigate the effects of the randomisation units.

Other unit of analysis issues

Cross-over trials

We excluded cross-over designs as these are unlikely to be a valid study design for Pregnancy and Childbirth reviews.

Multiple pregnancies

We did not identify any eligible studies that included multiple pregnancies in this review update, but we may include studies of this type in future updates.

When multiple pregnancies are included in studies, wherever possible, analyses should be adjusted for clustering to take into account the non-independence of babies from the same pregnancy (Gates 2004). Treating babies from multiple pregnancies as if they are independent, when they are more likely to have similar outcomes than babies from different pregnancies, will overestimate the sample size and give confidence intervals that are too narrow. Each woman can be considered a cluster in multiple pregnancy, with the number of individuals in the cluster being equal to the number of fetuses in her pregnancy. Analysis using cluster-trial methods allows calculation of risk ratio and adjustment of confidence intervals. Usually this will mean that the confidence intervals become wider. Although this may make little difference to the conclusion of a trial, it avoids misleading results in those trials where the difference may be substantial.

If multiple pregnancies are included in future updates of this review, we will adjust for clustering in the analyses wherever possible, and to use the inverse variance method for adjusted



analyses, as described in the *Cochrane Handbook for Systematic Reviews of Interventions* (Higgins 2011) and in Yelland 2011.

Dealing with missing data

For included studies, levels of attrition were noted. In future updates, if more eligible studies are included, the impact of including studies with high levels of missing data in the overall assessment of treatment effect will be explored by using sensitivity analysis.

For all outcomes, analyses were carried out, as far as possible, on an intention-to-treat basis, i.e. we attempted to include all participants randomised to each group in the analyses. The denominator for each outcome in each trial was the number randomised minus any participants whose outcomes were known to be missing.

Assessment of heterogeneity

We assessed statistical heterogeneity in each meta-analysis using the Tau², I² and Chi² statistics. We regarded heterogeneity as substantial if an I² was greater than 30% and either the Tau² was greater than zero, or there was a low P value (less than 0.10) in the Chi² test for heterogeneity. If we identified substantial heterogeneity (above 30%), we planned to explore it by prespecified subgroup analysis.

Assessment of reporting biases

In future updates, if there are 10 or more studies in the metaanalysis we will investigate reporting biases (such as publication bias) using funnel plots. We will assess funnel plot asymmetry visually. If asymmetry is suggested by a visual assessment, we will perform exploratory analyses to investigate it.

Data synthesis

We carried out statistical analysis using the Review Manager software (RevMan 2014). We used fixed-effect meta-analysis for combining data where it was reasonable to assume that studies were estimating the same underlying treatment effect: i.e. where studies were examining the same intervention, and the studies' populations and methods were judged sufficiently similar.

If there was clinical heterogeneity sufficient to expect that the underlying treatment effects differed between studies, or if substantial statistical heterogeneity was detected, we used random-effects meta-analysis to produce an overall summary if an average treatment effect across studies was considered clinically meaningful. The random-effects summary was treated as the average range of possible treatment effects and we discussed the clinical implications of treatment effects differing between studies. If the average treatment effect was not clinically meaningful, we did not combine studies. Where we used random-effects analyses, the results were presented as the average treatment effect with 95% confidence intervals, and the estimates of Tau² and I².

Subgroup analysis and investigation of heterogeneity

If we identified substantial heterogeneity among primary outcomes, we investigated it using subgroup analyses. We considered whether an overall summary was meaningful, and if it was, we used random-effects analysis to produce it. The following subgroup analyses were planned for the primary outcomes:

- nulliparae;
- early induction of labour (less than 12 hours from rupture of membranes);
- late induction of labour (at 12 hours or greater from rupture of membranes).

We assessed subgroup differences by interaction tests available within RevMan (RevMan 2014). We reported the results of subgroup analyses quoting the Chi² statistic and P value, and the interaction test I² value.

Sensitivity analysis

We did not conduct sensitivity analyses in this update due to the low number of included studies. In future updates we plan to carry out sensitivity analyses to explore the effect of risk of bias assessed by random sequence generation, concealment of allocation, and blinding of participants and personnel, with poor-quality studies (including those assessed as high or unknown risk of bias) being excluded from the analyses in order to assess whether this made any difference to the overall result.

RESULTS

Description of studies

Results of the search

Eight studies in all were identified for possible inclusion in this review, with four studies being excluded for the reasons outlined below. This updated review includes an additional two studies involving 1801 women (Nabhan 2014; Passos 2012), giving a total of four included studies of 2639 women. The previous version included two studies and 838 women.

Included studies

Two of the included studies compared antibiotics versus no antibiotics (Cararach 1998; Passos 2012) and two compared antibiotics with placebo (Nabhan 2014; Ovalle 1998). Cararach 1998 received partial support from the Spanish Ministry of Health, Nabhan 2014 and Ovalle 1998 received in-house support, and Passos 2012 received no funding. Declaration of conflicts of interest was provided for Nabhan 2014, Ovalle 1998 and Passos 2012, and all authors reported no conflicts.

Participants

The population of women in the included studies were homogenous. The gestation of women enrolled was 36 weeks or greater for Cararach 1998 and Nabhan 2014, 37 weeks or greater for Passos 2012, and 37 to 42 weeks for Ovalle 1998. All studies excluded women with multiple pregnancy and major obstetric complications and had stipulated criteria for diagnosis of membrane rupture.

Some differences were apparent in management protocols and types of antibiotics used as outlined below.

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Management protocols

Cararach 1998 and Ovalle 1998 had systematic approaches to routine maternal cervicovaginal cultures on admission. Ovalle 1998 also conducted amniocentesis for culture of amniotic fluid.

Induction of labour was undertaken in all studies however, the timing of induction differed. Cararach 1998 employed a policy of induction of labour for all women not in labour after 12 hours of membrane rupture. In Ovalle 1998, induction of labour was undertaken within 24 hours of membrane rupture and in an unknown proportion an earlier threshold was used (details are unclear). In this study all women undergoing induction of labour were induced after 12 hours (late induction). Nabhan 2014 employed a policy of immediate induction, ensuring the latency between rupture of membranes and labour was less than 12 hours. In Passos 2012, whether induction of labour or caesarean section was performed was determined at the discretion of the attending physicians (no further details given). All studies used Intravenous oxytocin as the method of induction of labour.

All studies employed management protocols which involved attempts to avoid or minimise vaginal examinations. Routine antibiotic prophylaxis at the time of caesarean section was given in Cararach 1998, Nabhan 2014 and Passos 2012, but inconsistently in Ovalle 1998.

Nabhan 2014 reported that no Group B streptococcal cultures were taken and no routine intrapartum Group B streptococcal prophylaxis was given, as per the hospital protocol for spontaneous rupture of membranes. Group B streptococcal cultures were taken in Passos 2012 and women with a positive or unknown culture between 35 and 37 weeks were excluded from the study. Neither Cararach 1998 nor Ovalle 1998 included a description of a policy for prevention of neonatal early onset Group B streptococcal disease. Ovalle 1998 administered routine antibiotics to neonates of mothers with clinical chorioamnionitis or positive maternal admission cultures (Group B streptococcal, haemophilus influenzae or chlamydia trachomatis). As per hospital policy, Nabhan 2014 prescribed routine prophylactic antibiotics to all neonates where maternal risk factors were present. In Passos 2012, antibiotic therapy was initiated only following early-onset neonatal infection, which was diagnosed, in this study, with or without a positive blood culture. Cararach 1998 did not describe the use of neonatal antibiotics for maternal risk factors.

Antibiotics

Cararach 1998 used intravenous ampicillin (or intramuscular erythromycin for women with penicillin allergy) with intramuscular gentamicin, Nabhan 2014 used parenteral ampicillin/sulbactam, Ovalle 1998 used intravenous cefuroxime and clindamycin for 48 hours then oral cefuroxime and clindamycin for a further 24 hours, and Passos 2012 used intravenous ampicillin with gentamicin.

Outcomes

The outcomes of maternal infectious morbidity (chorioamnionitis and/or endometritis) and early-onset neonatal sepsis (probable and definite) were assessed in all included studies. Maternal infection (chorioamnionitis and/or endometritis) was well defined in all studies. All studies diagnosed early-onset neonatal sepsis according to well defined clinical criteria, with or without positive blood culture. although no cases of neonatal morbidity were reported in Ovalle 1998.

Excluded studies

Four studies were excluded: Lebherz 1963 was excluded as the antibiotic (tetracycline) is now contraindicated in pregnancy and Brelje 1966 and Gordon 1974 because a quasi-random method of allocation was employed. A further study was excluded because additional information on methods of randomisation and allocation to treatment was not available (Walss Rodriguez 1988).

Risk of bias in included studies

The quality of the included studies ranged from fair to high. Cararach 1998 was of low quality and Ovalle 1998 and Passos 2012 were of moderate quality. Nabhan 2014 was of very high quality, with low risk of bias on all domains.

Please see Figure 1 and Figure 2 for a summary of 'Risk of bias' assessments. Please refer to the table of Characteristics of included studies for details about each study.

Figure 1. 'Risk of bias' graph: review authors' judgements about each risk of bias item presented as percentages across all included studies.

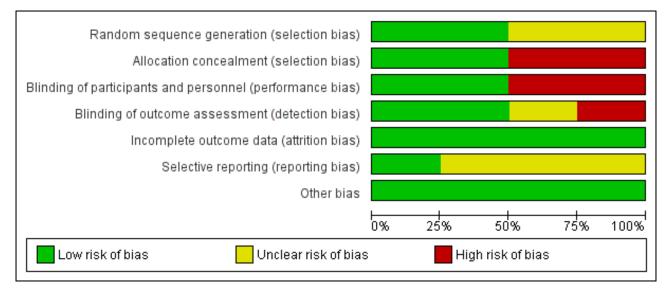
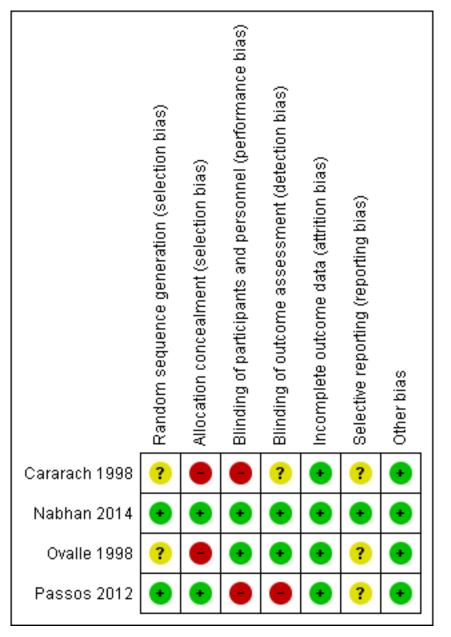




Figure 2. 'Risk of bias' summary: review authors' judgements about each risk of bias item for each included study.



Allocation

Adequate sequence generation was achieved and judged as low risk for selection bias in two of the four studies (Nabhan 2014; Passos 2012), both of which used computer-generated random number lists. Risk of bias was unclear in Cararach 1998 and Ovalle 1998. Cararach 1998 used a 'randomisation list in each participating centre' but provided no further details and Ovalle 1998 used a 'preestablished allocation code' but, again, provided no further details. Allocation concealment was achieved in Nabhan 2014 and Passos 2012, but not in Cararach 1998 or Ovalle 1998.

Blinding

The intervention was blinded to participants and personnel in Ovalle 1998 and Nabhan 2014 and judged as low risk for performance bias owing to the use of a placebo for the control group. No placebo was used in Cararach 1998 or Passos 2012 and we therefore judged these studies as high risk for performance bias. Nabhan 2014 undertook blinded assessment of outcomes, as did Ovalle 1998 for all outcomes and Cararach 1998 for neonatal outcomes, with disclosure of allocation only in cases of neonatal sepsis. We judged the risk of detection bias as high for Passos 2012, who undertook only partial blinding of outcomes assessors (only one of the three outcome assessors was blind to group allocation during the process).

Incomplete outcome data

All studies reported an intention-to-treat analysis. Cararach 1998 and Ovalle 1998 reported complete follow-up. Some attrition occurred in Nabhan 2014 due to incomplete data collection, and in Passos 2012 due to missing data and protocol violations, however,

attrition was low and comparable between groups in both studies. All studies were judged as low risk for attrition bias.

Selective reporting

The potential for selective reporting was unclear in Cararach 1998, Ovalle 1998 and Passos 2012 as no protocols were available for review. A prospectively-registered protocol was available for Nabhan 2014 and all outcomes were reported as expected (low risk for reporting bias).

Other potential sources of bias

We found no evidence of other bias in the studies.

Effects of interventions

See: Summary of findings for the main comparison Any antibiotic compared with placebo or no antibiotic (subgrouped by timing of induction of labour) for prelabour rupture of membranes at or near term; Summary of findings 2 Any antibiotic compared with placebo or no antibiotic for prelabour rupture of membranes at or near term

Comparisons and subgroup analyses were undertaken as follows.

- Comparison 1: any antibiotic compared with placebo or no antibiotic.
- Comparison 2: any antibiotic compared with placebo or no antibiotic subgrouped by timing of induction of labour: early induction (less than 12 hours from rupture of membranes) versus late induction (at 12 hours or greater from rupture of membranes).

We were unable to conduct the remaining planned subgroup analysis due to unavailability of data.

Comparison 1. Any antibiotic compared with placebo or no antibiotic

The results of four studies comparing the use of antibiotics with no use of antibiotics for women with term prelabour rupture of membranes (PROM) (Cararach 1998; Nabhan 2014; Ovalle 1998; Passos 2012), involving a total of 2639 women, are included in this review.

Primary outcomes

Early-onset neonatal sepsis (probable and definite)

A moderate degree of statistical heterogeneity was evident for both of these outcome measures. However, upon exploration of the possible reasons for the heterogeneity by examining clinical features of the studies, we considered an overall summary was meaningful using a random-effects analysis. Comparing antibiotics to placebo or no antibiotics, no difference was shown in probable early-onset neonatal sepsis (average risk ratio (RR) 0.69, 95% confidence interval (Cl) 0.21 to 2.33; babies = 2639; studies = four; Tau² = 0.71; Chi² = 5.34, df = 2 (P = 0.07); I² = 63%) (Analysis 1.1), or definite early onset neonatal sepsis (average RR 0.57, 95% Cl 0.08 to 4.26; babies = 2639; studies = four; Tau² = 1.51; Chi² = 3.14, df = 1 (P = 0.08); I² = 68%) (Analysis 1.2). Cases of definite early onset neonatal sepsis were low (less than 1% of all neonates), and there were no cases in two of the four studies (Ovalle 1998; Passos 2012).

Maternal infectious morbidity (chorioamnionitis and/or endometritis) (Analysis 1.3)

A moderate degree of statistical heterogeneity was also evident for this outcome measure. Upon exploration of the possible reasons for the heterogeneity by examining clinical features of the studies, we again considered an overall summary was meaningful using a random-effects analysis. No difference in maternal infectious morbidity (chorioamnionitis and/or endometritis) was shown (average RR 0.48, 95% Cl 0.20 to 1.15; women = 2639; studies = four; Tau² = 0.48; Chi² = 9.27, df = 3 (P = 0.03); I² = 68%).

Stillbirth (Analysis 1.4)

No difference in stillbirth was shown when comparing antibiotics with placebo or no antibiotics (RR 3.00, 95% CI 0.61 to 14.82; babies = 1906; studies = three). There were no cases of stillbirth in two of the three studies reporting this outcome (Ovalle 1998; Passos 2012).

Perinatal mortality (Analysis 1.5)

All studies reported data on perinatal mortality, and in two studies (Ovalle 1998; Passos 2012) there were no cases. Overall, no difference was detected (RR 1.98, 95% CI 0.60 to 6.55; babies = 2639; studies = four).

Neonatal mortality

There were no cases of neonatal mortality in three studies reporting this outcome (Nabhan 2014; Ovalle 1998; Passos 2012).

Serious maternal outcome (defined as death, cardiac arrest, respiratory arrest, anaphylaxis, admission to intensive care unit)

There were no cases of serious maternal outcome in the three trials reporting this outcome (Nabhan 2014; Ovalle 1998; Passos 2012).

Secondary outcomes

For the woman

Caesarean section was increased with the use of antibiotics when compared to placebo or no antibiotics (RR 1.33, 95% Cl 1.09 to 1.61; number needed to treat to harm (NNTH) 20, 95% Cl 11 to 74; women = 1906; studies = three (Analysis 1.10), as was duration of maternal stay in hospital (mean difference (MD) 0.06 days, 95% Cl 0.01 to 0.11; women = 1906; studies = three) (Analysis 1.28).

Two studies comparing antibiotics with placebo reported data on postpartum antibiotic usage. Due to extreme heterogeneity (Tau² = 1.32; $Chi^2 = 14.47$, df = 1 (P = 0.0001); $I^2 = 93\%$), we did not combine the data from these studies. One study of 1640 women (Nabhan 2014) reported no difference in postpartum antibiotic usage (RR 1.11, 95% CI 0.60 to 2.04), whereas the other study of 105 women (Ovalle 1998) reported an increase in postpartum antibiotic usage (RR 5.95, 95% CI 3.22 to 10.99, NNTH 2, 95% CI 1 to 3) (Analysis 1.18). The same two studies also reported data on postpartum pyrexia. Again, due to substantial heterogeneity (Tau² = 1.79; Chi² = 4.11, df = 1 (P = 0.04); I^2 = 76%) we did not combine these data. The Nabhan 2014 study reported no difference in postpartum pyrexia (RR 0.97, 95% CI 0.58 to 1.61), whereas the Ovalle 1998 study reported a reduction in postpartum pyrexia (RR 0.11, 95% CI 0.01 to 0.88; number needed to treat to benefit (NNTB) 7, 95% CI 8 to 53) (Analysis 1.15).



No differences were shown on the remaining secondary maternal outcomes as outlined below. For the outcomes of chorioamnionitis, endometritis, and postpartum haemorrhage, a moderate degree of statistical heterogeneity was evident, but upon exploration of the possible reasons for the heterogeneity by examining clinical features of the studies, we again considered an overall summary was meaningful using a random-effects analysis:

- chorioamnionitis (suspected or proven) (average RR 0.65, 95% CI 0.34 to 1.26; women = 2639; studies = four; Tau² = 0.17; Chi² = 5.07, df = 3 (P = 0.17); I² = 41%) (Analysis 1.8);
- endometritis (average RR 0.34, 95% CI 0.05 to 2.31; women = 2639; studies = four; Tau² = 2.06; Chi² = 6.74, df = 3 (P = 0.08); I² = 55%) (Analysis 1.9);
- operative vaginal birth (RR 0.95, 95% CI 0.63 to 1.44; women = 1906; studies = three) (Analysis 1.11);
- internal fetal monitoring (no cases, two studies);
- epidural analgesia (RR 1.00, 95% CI 0.98 to 1.02; women = 266; studies = two) (Analysis 1.13);
- postpartum haemorrhage (average RR 1.11, 95% CI 0.14 to 8.93; women = 1906; studies = three; Tau² = 1.51; Chi² = 2.23, df = 1 (P = 0.14); l² = 55%) (Analysis 1.14);
- postpartum septicaemia (no cases, three studies);
- wound infection (RR 0.79, 95% CI 0.36 to 1.72; women = 1906; studies = three) (Analysis 1.17);
- adverse effects (RR 2.93, 95% CI 0.12 to 71.63; women = 2639; studies = four) (Analysis 1.19).

No studies reported data on breastfeeding on discharge from hospital.

For the neonate

One study of 105 neonates (Ovalle 1998) reported reduced neonatal stay in hospital with the use of antibiotics (MD -0.90, 95% CI -1.34 to -0.46) (Analysis 1.29).

No differences were shown on the remaining secondary neonatal outcomes as outlined below. For the outcome of neonatal antibiotic usage a moderate degree of statistical heterogeneity was evident. Upon exploration of the possible reasons for the heterogeneity by examining clinical features of the studies, we considered an overall summary was meaningful using a random-effects analysis:

- neonatal meningitis (RR 0.33, 95% CI 0.03 to 3.11; babies = 2639; studies = four) (Analysis 1.20);
- neonatal pneumonia (RR 0.33, 95% CI 0.01 to 7.96; babies = 2639; studies = four) (Analysis 1.21);
- Apgar score less than seven at five minutes (RR 1.65, 95% CI 0.81 to 3.36; babies = 2639; studies = four) (Analysis 1.22);
- admission to neonatal special care nursery (average RR 0.32, 95% Cl 0.06 to 1.74; babies = 266; studies = two; Tau² = 0.77; Chi² = 1.97, df = 1 (P = 0.16); l² = 49%) (Analysis 1.23);
- admission to neonatal intensive care unit (RR 1.23, 95% CI 0.82 to 1.85; babies = 1906; studies = three) (Analysis 1.24);
- antibiotic usage (average RR 0.55, 95% Cl 0.15 to 1.97; babies = 1906; studies = three; Tau² = 0.87; Chi² = 6.55, df = 2 (P = 0.04); l² = 69%) (Analysis 1.25);
- respiratory distress syndrome (no cases, two studies);
- use of mechanical ventilation (RR 1.21, 95% CI 0.50 to 2.91; babies = 2639; studies = four) (Analysis 1.27);

 duration of neonatal stay in intensive care unit (MD 0.05 days, 95% CI -0.09 to 0.19; babies = 1640; studies = one) (Analysis 1.30).

Cost effectiveness

No data on cost effectiveness were available.

Comparison 2. Any antibiotic compared with placebo or no antibiotic subgrouped by timing of induction of labour (early induction versus late induction)

Three studies involving 2478 women were included in the subgroup analysis of timing of induction of labour (Cararach 1998; Nabhan 2014; Ovalle 1998).

Early-onset neonatal sepsis (probable and definite)

No difference in either probable or definite early-onset neonatal sepsis was shown in the early induction group (RR 1.47, 95% CI 0.80 to 2.70; babies = 1640; studies = one and RR 1.29, 95% CI 0.48 to 3.44; babies = 1640; studies = one, respectively) (Analysis 2.1 and Analysis 2.2). There were trends toward reduced probable and definite early-onset neonatal sepsis in the late induction group, but neither result reached statistical significance (RR 0.14, 95% CI 0.02 to 1.13; babies = 838; studies = two, and RR 0.16, 95% CI 0.02 to 1.34; babies = 838; studies = two, respectively). A test for subgroup differences was significant (Chi² = 4.50, df = 1 (P = 0.03), l² = 77.8%), suggesting a differential effect of the intervention on probable early-onset neonatal sepsis between the early and late induction groups.

Maternal infectious morbidity (chorioamnionitis and/or endometritis)

No difference in maternal infectious morbidity (chorioamnionitis and/or endometritis) was shown in the early induction group (average RR 1.15, 95% Cl 0.64 to 2.08; women = 1640; studies = one). There was a trend toward reduced maternal infectious morbidity in the late induction group (average RR 0.34, 95% Cl 0.08 to 1.47; babies = 838; studies = two; Tau² = 0.70; Chi² = 2.16, df = 1 (P = 0.14); l² = 54%) (Analysis 2.3), but this result was not statistically significant.

Stillbirth

No difference in stillbirth was shown in one study of 1640 women in the early induction group (RR 3.00, 95% CI 0.61 to 14.82) (Analysis 2.4). No data were available for the late induction subgroup.

Perinatal mortality

No difference in perinatal mortality was shown in either the early induction or late induction groups (RR 3.00, 95% CI 0.61 to 14.82; babies = 1640; studies = one, and RR 0.98, 95% CI 0.14 to 6.89; babies = 838; studies = two, respectively) (Analysis 2.5).

DISCUSSION

Summary of main results

This review included four studies of 2639 women comparing antibiotics for women with term prelabour rupture of membranes with placebo or no antibiotics. The previous version of this review included two studies of 838 women. Whereas the previous version of this review showed a statistically significant reduction in endometritis, no such effect was shown in this update after an additional two studies of 1801 women. No differences were shown on the primary outcome measures of

early-onset neonatal sepsis (probable and definite), maternal infectious morbidity (chorioamnionitis and/or endometritis), stillbirth, perinatal mortality, neonatal mortality, and serious maternal outcome, though the number of cases in the control group for these outcomes was low.

Caesarean section was increased with the use of antibiotics, as was duration of maternal stay in hospital. These results appear largely driven by the Nabhan 2014 study, in which 165 women in the antibiotic group underwent caesarean section (20%), compared to 122 women in the placebo group (approximately 15%). Repeat caesarean section was significantly more common in the antibiotic group compared to the placebo group in this study. The larger incidence of caesarean section in the antibiotic group may also be partially accounted for by the increased hypertension and pre-eclampsia in the antibiotic group in this study.

Due to heterogeneity we did not combine the data from two studies measuring postpartum antibiotic usage and postpartum pyrexia. A small study of 105 women (Ovalle 1998) reported an increase in postpartum antibiotic usage and a reduction in postpartum pyrexia, whereas a larger study of 1640 women (Nabhan 2014) reported no differences on either outcome. The increased use of postpartum antibiotic usage in Ovalle 1998 appears to be explained by the study protocol for antibiotic provision, and the reduction in postpartum pyrexia in Ovalle 1998 is likely owing to the observed reduction in maternal infectious morbidity in this study.

No difference in any measures of maternal morbidity or neonatal morbidity were shown, although one small study showed a reduction in neonatal nursery stay. No data on cost effectiveness were available.

With the additional data available in this update we were able to introduce a second comparison of antibiotics versus placebo or no antibiotics subgrouped by timing of induction, whereby we compared early induction (defined as induction less than 12 hours from rupture of membranes) with late induction (defined as induction at 12 hours or greater from rupture of membranes). Three studies involving 2478 women were included in this subgroup analysis. No difference in either probable or definite earlyonset neonatal sepsis was shown in either subgroup, although there were trends toward reduced probable and definite earlyonset neonatal sepsis in the late induction group. Similarly, no difference in maternal infectious morbidity (chorioamnionitis and/ or endometritis) was found in either subgroup analysis, though there was a trend towards reduced maternal infectious morbidly in the late induction group. A test for subgroup differences confirmed a differential effect of the intervention on probable early-onset neonatal sepsis between the early and late induction groups. Given these data, it appears that it is the difference in the timing of induction and the period of latency between membrane rupture and birth that is likely to have the greatest impact on infectious morbidity, with the least effect of antibiotic treatment seen in the Nabhan 2014 study where the period between admission and birth was shortest, and greater effect of antibiotic treatment seen in the remaining studies in these analyses where the period of latency was longer.

No differences were shown in stillbirth or perinatal mortality, and there were no cases of serious maternal outcome or neonatal mortality.

Overall completeness and applicability of evidence

The four studies included in this review showed some heterogeneity. The main source of clinical heterogeneity was in the timing of induction of labour, which we have attempted to address via subgroup analyses. Nonetheless, the conclusions from this review are limited by rare event rates across the primary outcomes. This review is underpowered to detect differences in stillbirth, perinatal mortality, and neonatal mortality and it is unlikely that future studies will be able to rectify this problem due to the sheer numbers of women and neonates that would be required to complete such clinical trials. A further limitation to the applicability of the evidence in this review is regarding the use or method of routine screening for Group B Streptococcus, which was included in only one (Passos 2012) of the included studies. The review also provides little information on possible short- and long-term adverse effects from the use of antibiotics including allergic reaction and anaphylaxis for women (ACOG 2011), and the development of resistant organisms among neonates such as gram-negative organisms (Edwards 2001; Stoll 2002). Due to unavailability of data, we are unable to assess not only the direct economic impact of routine use of antibiotics for PROM at term, but also the economic impact of treating the antibiotic-resistant infections that may result. Due to the low number of included studies, we were also not able to investigate the likelihood of reporting biases such as publication bias.

Quality of the evidence

Overall risk of bias in this review was fair to high. Allocation to groups was not concealed in two of the four studies (Cararach 1998; Ovalle 1998), and in a third study (Passos 2012) both performance and detection bias was judged to be high risk. The largest and most recent study (Nabhan 2014) was of particularly high quality, showing low risk of bias on all domains assessed. Using the GRADE approach, quality of evidence for each of the primary outcomes yielding a point estimate (probable early-onset neonatal sepsis; definite early-onset neonatal sepsis; maternal infectious morbidity; stillbirth; and perinatal mortality) was judged to be low to very low. We downgraded evidence for each of these outcomes due to risk of bias and imprecision of results (see Summary of findings for the main comparison; Summary of findings 2).

Potential biases in the review process

We are aware that the review process itself is subject to bias, and we took steps to minimise bias. At least two review authors carried out data extraction and assessed risk of bias independently; however, a different review team may not have made identical decisions.

Agreements and disagreements with other studies or reviews

In their retrospective cohort study of 3841 women at a singlesite in the United States, Tran 2007 and colleagues examined the relationship between duration of membrane rupture and maternal infection, finding that the risk chorioamnionitis and endometritis was increased at 12 hours and 16 hours respectively post membrane rupture. The study included women with singleton pregnancy and cephalic presentation, with PROM diagnosed at 37 weeks' gestation or later. Women with multiple pregnancies and obstetric complications were excluded and definitions for chorioamnionitis and endometritis were similar to those in this review. The study also found that the risk of postpartum

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haemorrhage was increased with durations of membrane rupture, but we were unable to investigate this in the current review due to unavailability of data. The Cochrane Review on *Planned early birth versus expectant management (waiting) for prelabour rupture of membranes at term (37 weeks or more)* (Dare 2006) showed that planned management (usually by induction of labour) reduced chorioamnionitis and endometritis and, similar to this review, no difference in neonatal infection was shown. It is important to note that the NICE guideline on intrapartum care states that induction of labour is appropriate approximately 24 hours after rupture of membranes (NICE guideline 55). As in this update, the guideline also states that maternal and neonatal antibiotics should not be administered in the absence of signs of maternal infection.

AUTHORS' CONCLUSIONS

Implications for practice

This review demonstrates no convincing maternal or neonatal benefits of the routine use of antibiotics for prelabour rupture of membranes (PROM) at or near term in the absence of confirmed clinical infection. It is likely that the modest benefit for maternal endometritis seen in the previous version of this review can be largely accounted for by the late use of induction of labour in both included studies, where the duration of membrane rupture was more prolonged (related to either a policy of expectant management or a delay in induction greater than 24 hours). The low rate of maternal infection in the control population of this updated review (less than 5%) further reinforces that administration of antibiotics to women with PROM should be restricted to those who develop clinical indications for antibiotic treatment.

Implications for research

Antibiotics may not offer any benefit under a clinical policy of immediate or early induction of labour, where the duration of membrane rupture is minimised. For women being managed expectantly or with a policy of delayed induction, there may be a role for antibiotics, but further well-designed randomised controlled trials are needed. These trials should utilise blinding of the intervention and need to be adequately sized to address clinically important maternal and neonatal outcomes. All future trials should consider including a cost analysis to determine the economic impact of routine antibiotics for PROM at term.

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CHARACTERISTICS OF STUDIES

Characteristics of included studies [ordered by study ID]

Cararach 1998

Methods Multicentre randomised trial. Participants 733 women at 36 weeks' gestation or more with singleton pregnancy, MR duration less than 12 hrs and absence of uterine contractions. Setting: 11 hospitals in Spain. Exclusion criteria: fetal death or anomaly, placenta praevia, placental abruption, fetal distress, chorioamnionitis, indication for elective CS, allergy to penicillin and erythromycin. Interventions Intervention group: antibiotics on admission following vaginal and endocervical culture. IV ampicillin 1 g every 6 hrs and IM gentamicin 80 mg every 8 hrs or IM erythromycin 500 mg every 6 hrs for women with penicillin allergy. Control group: no treatment. Admission cultures as for intervention group. Outcomes Maternal: mode of birth, chorioamnionitis, endometritis, postpartum antibiotics. Neonatal: Apgar score, umbilical artery and vein pH values, early neonatal infection (i.e. sepsis, meningitis, or pneumonia), respiratory complications, intracranial haemorrhage. Data were requested and received for: maternal adverse drug reaction, neonatal mechanical ventilation, perinatal death. Notes Management of all women enrolled: single vaginal examination until initiation of labour. Vaginal and endocervical culture on admission. IOL with IV oxytocin after 12 hrs of MR in the absence of regular uterine contractions. Routine antibiotic prophylaxis for women undergoing CS. Partial support received via a grant from the 'Fondo de Investigaciones Sanitarias' (FIS), Spanish Ministry of Health. Declaration of conflicts of interest not included. **Risk of bias** Bias **Authors' judgement** Support for judgement Random sequence genera-Unclear risk "Randomisation list in each participating centre". No details provided. tion (selection bias) Allocation concealment High risk Open list used, no concealment of allocation. (selection bias) Blinding of participants High risk Unblinded.

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and personnel (perfor-

mance bias)

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* Indicates the major publication for the study



Blinding of outcome as- sessment (detection bias) All outcomes	Unclear risk	Blinding of outcome assessment for neonatal outcomes only. Unblinding in cases of neonatal sepsis.
Incomplete outcome data (attrition bias) All outcomes	Low risk	Complete follow-up.
Selective reporting (re- porting bias)	Unclear risk	No protocol available. Intracranial haemorrhage not reported.
Other bias	Low risk	None evident.

Nabhan 2014

Methods	Randomised controlled trial between June 2009 and January 2011. Parallel group design.				
Participants	1640 women at 36 weeks' gestation or more with prelabour rupture of membranes and singleton preg- nancy. Setting: Labour and delivery ward, Department of Obstetrics, Ain Shams University, Cairo, Egypt. Exclusion criteria: rupture of membranes for more than 12 hrs, chorioamnionitis, meconium-stained liquor on admission, fetal anomalies, maternal rheumatic valve disease.				
Interventions	Intervention group: single dose 1500mg parenteral ampicillin/sulbactam. Control group: placebo (solvent without active ingredient). Both groups: no GBS cultures taken and no routine intrapartum GBS prophylaxis given according to hospital protocol. All women were managed on admission either by induction of labour or CS.				
Outcomes	Neonatal: primary outcome EONS. EONS defined as neonatal infection within the first 7 days of life. Definite infection defined as above plus 1 of: (1) positive blood, CSF, urine, or tracheal aspirate culture; (2) positive CSF Gram stain; (3) positive blood, CSF, or urine antigen detection test; or (4) positive chest radiograph confirming neonatal pneumonia. Probable EONS defined by clinical signs of infection plus either a complete blood count suggestive of infection or a CSF with an elevated leukocyte count, a high protein concentration, or a low glucose concentration. Additional neonatal outcomes included new- born pneumonia, fetal or early neonatal death (excluding CA), Apgar < 7 at 5 mins, admission to NICU, need for and duration of mechanical ventilation, and duration of NICU stay. Maternal outcomes were chorioamnionitis, endometritis, mode of birth, PPH, postpartum pyrexia, ADRs, duration of hospital stay.				
Notes	Intention-to-treat analyses used.				
	Employed a policy of immediate induction, ensuring the latency between rupture of membranes and labour was less than 12 hrs.				
	No GBS cultures taken and no routine intrapartum GBS prophylaxis as per hospital protocol.				
	Intramaural support only, no external funding. Authors reported no conflicts of interest.				
Risk of bias					

Bias	Authors' judgement	Support for judgement
Random sequence genera- tion (selection bias)	Low risk	Computer-generated simple randomisation procedure (1:1 allocation ratio).

Nabhan 2014 (Continued)

Allocation concealment (selection bias)	Low risk	An obstetrician on the labour ward enrolled participants. For each participant, the obstetrician obtained the next randomisation number by telephone.
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Participants and (most) personnel were blinded to the group assignment (an- tibiotics or placebo was administered by a labour ward nurse).
Blinding of outcome as- sessment (detection bias) All outcomes	Low risk	Staff assessing outcomes were blinded. Only the labour ward nurse who ad- ministered antibiotics or placebo was aware of group allocation, this nurse did not assess outcomes.
Incomplete outcome data (attrition bias) All outcomes	Low risk	62 participants in total were lost to follow-up due to incomplete collection of data (attrition rates acceptable: 2.6% in intervention and 5% in control group). Intention-to-treat analyses used.
Selective reporting (re- porting bias)	Low risk	Prospectively registered protocol available and outcomes reported as expect- ed.
Other bias	Low risk	None evident.

Ovalle 1998

Methods	Single-centre randomised trial between August 1990 and December 1993.
Participants	105 women at 37-42 weeks' gestation with a singleton pregnancy, duration of MR less than 12 hrs and no labour.
	Setting: San Borja Arriarán Hospital, Chile. Exclusion criteria: previous CS, malpresentation, fetal distress, fetal malformation, chorioamnionitis, antibiotics given within 30 days.
Interventions	Intervention: antibiotics on admission following cervicovaginal and amniotic fluid culture. IV clin- damycin 600 mg every 6 hrs and IV cefuroxime 750 mg every 8 hrs for 48 hrs then oral cefuroxime 250 mg every 12 hrs and clindamycin 300 mg every 6 hrs for a further 24 hrs.
	Control: placebo following admission cultures. No details provided.
Outcomes	Maternal: chorioamnionitis, endometritis. Neonatal: Apgar score < 7 at 5 mins, neonatal morbidity. Data were requested and received for: neonatal sepsis, pneumonia, meningitis, neonatal mechanical ventilation, admission to SCN, maternal and neonatal length of stay, perinatal death, maternal adverse drug reaction.
Notes	Management of all women enrolled: no digital vaginal examination until active labour. Cervicovaginal culture and culture of amniotic fluid by amniocentesis on admission. IOL with IV oxytocin within 24 hrs of MR if no labour. Antibiotic prophylaxis given to some women un- dergoing CS. Neonatal antibiotics routinely given when chorioamnionitis present or positive maternal admission cultures.
	No funding details provided. Declaration of conflicts of interest not included.
Risk of bias	
Bias	Authors' judgement Support for judgement

Ovalle 1998 (Continued)

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Random sequence genera- tion (selection bias)	Unclear risk	Pre-established allocation code. No other details given.
Allocation concealment (selection bias)	High risk	Pharmacist allocation using a list of random numbers. No central randomisa- tion, or sequentially numbered drug containers or opaque, sealed envelopes.
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Placebo-controlled trial.
Blinding of outcome as- sessment (detection bias) All outcomes	Low risk	Placebo-controlled trial. The treating physicians were independent from those in charge of the study and were unaware of group allocation.
Incomplete outcome data (attrition bias) All outcomes	Low risk	Complete follow-up.
Selective reporting (re- porting bias)	Unclear risk	No protocol available.
Other bias	Low risk	None evident.

Passos 2012

Methods	Randomised controlled trial between October 2008 and January 2012.
Participants	161 women at 37 weeks' gestation or more with singleton pregnancy, vertex presentation, ruptured membranes for less than 12 hrs, and a negative GBS culture between 35 and 37 weeks. Setting: Department of Obstetrics, Gynecology, and Reproductive Medicine, Lisbon University Hospi- tal. Exclusion criteria: active labour, meconium liquor, no GBS culture, indication for GBS antibiotic pro- phylaxis, or contraindication to expectant management (e.g. fetal distress).
Interventions	Intervention group: 1 g IV ampicillin every 6 hrs and 240 mg IV gentamicin daily. Control group: no treatment.
Outcomes	Maternal: primary outcome was maternal infection (chorioamnionitis or puerperal endometritis). Ad- ditional outcomes were time from PROM to admission, time from PROM to initiation of spontaneous labour, time from PROM to IOL, time from rupture of membranes to birth; mode of birth and CS for fetal distress. Neonatal: primary outcomes was neonatal infection rate (EONS, meningitis, and pneumonia). Addi- tional outcomes were Apgar score at 1 and 5 mins, birthweight, early-onset neonatal infections and death.
	Data were requested and received for operative birth, epidural analgesia, PPH, postpartum septi- caemia, wound infection, maternal adverse effects, Apgar scores, admission to SCN and NICU, EONS, neonatal use of antibiotics and mechanical ventilation.
Notes	Intention-to-treat analysis used, women were excluded where outcome data were not available.
	IOL or CS performed at the discretion of the attending physicians (no details given).
	Antibiotic therapy for neonates initiated only following diagnosis of EONS.
	No funding details provided. Authors reported no conflicts of interest.

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Passos 2012 (Continued)

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence genera- tion (selection bias)	Low risk	Participants were 'randomly assigned to the antibiotic group or to the control group, according to a computer-generated randomisation list'.
Allocation concealment (selection bias)	Low risk	Sequentially numbered opaque envelopes used.
Blinding of participants and personnel (perfor- mance bias) All outcomes	High risk	Unblinded.
Blinding of outcome as- sessment (detection bias) All outcomes	High risk	2 obstetric and 1 paediatric investigators reviewed data to confirm diagnoses of maternal and neonatal infections. Only the paediatric investigator was blind to group allocation during the process.
Incomplete outcome data (attrition bias) All outcomes	Low risk	5 participants were excluded due to missing clinical data. 6 were excluded due to protocol violations. Intention-to-treat analyses used.
Selective reporting (re- porting bias)	Unclear risk	No protocol available. Trial was registered via ClinicalTrials.gov in 2012 (no prospective registration).
Other bias	Low risk	None evident.

ADR: adverse drug reaction CA: congenital abnormality CS: caesarean section CSF: cerebrospinal fluid EONS: early-onset neonatal sepsis GBS: group B streptococcus hrs: hours IM: intramuscular IOL: induction of labour IV: intravenous mins: minutes NICU: neonatal intensive care unit MR: membrane rupture PPH: postpartum haemorrhage PPROM: prelabour rupture of the membranes SCN: special care nursery

Characteristics of excluded studies [ordered by study ID]

Study	Reason for exclusion
Brelje 1966	Quasi-random allocation used.
Gordon 1974	Quasi-random allocation used.
Lebherz 1963	Antibiotic used no longer recommended for use in pregnancy.



Study	Reason for exclusion
Walss Rodriguez 1988	Further information on method of randomisation and allocation to treatment was requested but is not yet available.

DATA AND ANALYSES

Comparison 1. Any antibiotic compared with placebo or no antibiotic

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size		
1 Probable early-onset neonatal sepsis	4	2639	Risk Ratio (M-H, Random, 95% CI)	0.69 [0.21, 2.33]		
2 Definite early-onset neonatal sepsis	4	2639	Risk Ratio (M-H, Random, 95% CI)	0.57 [0.08, 4.26]		
3 Maternal infectious morbidi- ty (chorioamnionitis and/or en- dometritis)	4	2639	Risk Ratio (M-H, Random, 95% CI)	0.48 [0.20, 1.15]		
4 Stillbirth	Stillbirth 3 1906		Risk Ratio (M-H, Fixed, 95% CI)	3.0 [0.61, 14.82]		
5 Perinatal mortality	4	2639	Risk Ratio (M-H, Fixed, 95% CI)	1.98 [0.60, 6.55]		
6 Neonatal mortality	3	1906	Risk Ratio (M-H, Fixed, 95% CI)	0.0 [0.0, 0.0]		
7 Serious maternal outcome	3	1906	Risk Ratio (M-H, Fixed, 95% CI)	0.0 [0.0, 0.0]		
8 Chorioamnionitis (suspected or proven)	4	2639	Risk Ratio (M-H, Random, 95% CI)	0.65 [0.34, 1.26]		
9 Endometritis	4	2639	Risk Ratio (M-H, Random, 95% CI)	0.34 [0.05, 2.31]		
10 Caesarean section	3	1906	Risk Ratio (M-H, Fixed, 95% CI)	1.33 [1.09, 1.61]		
11 Operative vaginal birth	3	1906	Risk Ratio (M-H, Fixed, 95% CI)	0.95 [0.63, 1.44]		
12 Internal fetal monitoring	2	1745	Risk Ratio (M-H, Fixed, 95% CI)	0.0 [0.0, 0.0]		
13 Epidural analgesia	2	266	Risk Ratio (M-H, Fixed, 95% CI)	1.0 [0.98, 1.02]		
14 Postpartum haemorrhage	3	1906	Risk Ratio (M-H, Random, 95% CI)	1.11 [0.14, 8.93]		
15 Postpartum pyrexia	2		Risk Ratio (M-H, Random, 95% CI)	Subtotals only		
16 Postpartum septicaemia	3	1906	Risk Ratio (M-H, Fixed, 95% CI)	0.0 [0.0, 0.0]		
17 Wound infection	3	1906	Risk Ratio (M-H, Fixed, 95% CI)	0.79 [0.36, 1.72]		

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Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size		
18 Maternal postpartum antibi- otic usage	2		Risk Ratio (M-H, Random, 95% CI)	Subtotals only		
19 Maternal adverse effects	4	2639	Risk Ratio (M-H, Fixed, 95% CI)	2.93 [0.12, 71.63]		
20 Neonatal meningitis	4	2639	Risk Ratio (M-H, Fixed, 95% CI)	0.33 [0.03, 3.11]		
21 Neonatal pneumonia	4	2639	Risk Ratio (M-H, Fixed, 95% CI)	0.33 [0.01, 7.96]		
22 Apgar score < 7 at 5 minutes	4	2639	Risk Ratio (M-H, Fixed, 95% CI)	1.65 [0.81, 3.36]		
23 Admission to neonatal special care nursery	2	266	Risk Ratio (M-H, Random, 95% CI)	0.32 [0.06, 1.74]		
24 Admission to neonatal inten- sive care unit	3	1906	Risk Ratio (M-H, Fixed, 95% CI)	1.23 [0.82, 1.85]		
25 Neonatal antibiotic usage	3	1906	Risk Ratio (M-H, Random, 95% CI)	0.55 [0.15, 1.97]		
26 Respiratory distress syn- drome	2	266	Risk Ratio (M-H, Fixed, 95% CI)	0.0 [0.0, 0.0]		
27 Neonatal mechanical ventila- tion	4	2639	Risk Ratio (M-H, Fixed, 95% CI)	1.21 [0.50, 2.91]		
28 Duration of maternal stay in hospital (days)	3	1906	Mean Difference (IV, Fixed, 95% CI)	0.06 [0.01, 0.11]		
29 Duration of neonatal stay in hospital (days)	1	105	Mean Difference (IV, Fixed, 95% CI)	-0.90 [-1.34, -0.46]		
30 Duration of neonatal stay in intensive care unit (days)	1	1640	Mean Difference (IV, Fixed, 95% CI)	0.05 [-0.09, 0.19]		

Analysis 1.1. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 1 Probable early-onset neonatal sepsis.

Study or subgroup	Treatment	Control	Risk Ratio						Weight	Risk Ratio	
	n/N	n/N		M-	H, Rar	ndom,	95% CI				M-H, Random, 95% Cl
Cararach 1998	1/371	7/362	+			+				20.77%	0.14[0.02,1.13]
Nabhan 2014	25/820	17/820				+				47.71%	1.47[0.8,2.7]
Ovalle 1998	0/55	0/50									Not estimable
Passos 2012	3/78	5/83								31.52%	0.64[0.16,2.58]
Total (95% CI)	1324	1315								100%	0.69[0.21,2.33]
Total events: 29 (Treatment), 29	(Control)										
Heterogeneity: Tau ² =0.71; Chi ² =	5.34, df=2(P=0.07); I ² =62.54	4%									
Test for overall effect: Z=0.59(P=	=0.55)										
	Fa	vours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotic	S



Analysis 1.2. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 2 Definite early-onset neonatal sepsis.

Study or subgroup	Treatment	Control		Risk Ratio						Weight	Risk Ratio	
	n/N	n/N			M-H, Ra	ndom	, 95% CI				M-H, Random, 95% CI	
Cararach 1998	1/371	6/362	•	1						39.74%	0.16[0.02,1.34]	
Nabhan 2014	9/820	7/820						-		60.26%	1.29[0.48,3.44]	
Ovalle 1998	0/55	0/50									Not estimable	
Passos 2012	0/78	0/83									Not estimable	
Total (95% CI)	1324	1315						_		100%	0.57[0.08,4.26]	
Total events: 10 (Treatment), 13	8 (Control)											
Heterogeneity: Tau ² =1.51; Chi ² =	3.14, df=1(P=0.08); l ² =68.14	4%										
Test for overall effect: Z=0.55(P=	-0.58)											
	Fa	vours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotic	S	

Analysis 1.3. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 3 Maternal infectious morbidity (chorioamnionitis and/or endometritis).

Study or subgroup	Treatment	Control		Risk Ratio					Weight	Risk Ratio
	n/N	n/N		M-H, Ra	ndom	, 95% CI				M-H, Random, 95% Cl
Cararach 1998	12/371	21/362			-				33.02%	0.56[0.28,1.12]
Nabhan 2014	23/820	20/820		-					35%	1.15[0.64,2.08]
Ovalle 1998	1/55	8/50	-		-				12.82%	0.11[0.01,0.88]
Passos 2012	2/78	11/83	←	•	-				19.17%	0.19[0.04,0.85]
Total (95% CI)	1324	1315							100%	0.48[0.2,1.15]
Total events: 38 (Treatment), 6	50 (Control)									
Heterogeneity: Tau ² =0.48; Chi ²	² =9.27, df=3(P=0.03); I ² =67.64	1%								
Test for overall effect: Z=1.64(F	P=0.1)									
	Fav	ours antibiotics	0.1	0.2 0.5	1	2	5	10	Favours no antibiotic	S

Analysis 1.4. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 4 Stillbirth.

Study or subgroup	Treatment	Control			Risk Ratio			Weight	Risk Ratio
	n/N	n/N		М-	H, Fixed, 95	% CI			M-H, Fixed, 95% CI
Nabhan 2014	6/820	2/820						100%	3[0.61,14.82]
Ovalle 1998	0/55	0/50				_			Not estimable
Passos 2012	0/78	0/83							Not estimable
Total (95% CI)	953	953						100%	3[0.61,14.82]
Total events: 6 (Treatment), 2 (Control))								
Heterogeneity: Not applicable									
Test for overall effect: Z=1.35(P=0.18)									
	Fa	vours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	

Analysis 1.5. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 5 Perinatal mortality.

Study or subgroup	Treatment	Control		Risk Ratio				Weight	Risk Ratio		
	n/N	n/N			М-Н, Р	ixed, 9	95% CI				M-H, Fixed, 95% CI
Cararach 1998	2/371	2/362	-						-	50.31%	0.98[0.14,6.89]
Nabhan 2014	6/820	2/820			-	_	-		\rightarrow	49.69%	3[0.61,14.82]
Ovalle 1998	0/55	0/50									Not estimable
Passos 2012	0/78	0/83									Not estimable
Total (95% CI)	1324	1315			_					100%	1.98[0.6,6.55]
Total events: 8 (Treatment), 4 (Co	ntrol)										
Heterogeneity: Tau ² =0; Chi ² =0.76,	df=1(P=0.38); I ² =0%										
Test for overall effect: Z=1.12(P=0.	.26)				1						
	F	avours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotics	

Analysis 1.6. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 6 Neonatal mortality.

Study or subgroup	Treatment	Treatment Control			Risk Ratio		Weight	Risk Ratio
	n/N	n/N		M-H	Fixed, 95% C	1	M-H, Fixed, 95% CI	
Nabhan 2014	0/820	0/820						Not estimable
Ovalle 1998	0/55	0/50						Not estimable
Passos 2012	0/78	0/83						Not estimable
Total (95% CI)	953	953						Not estimable
Total events: 0 (Treatment), 0 (Contro	ol)							
Heterogeneity: Not applicable								
Test for overall effect: Not applicable								
	Ea	wours antibiotics	0.01	0.1	1	10 100	Eavours no antibiotics	

Favours antibiotics 0.01 0.1 1 10 100 Favours no antibiotics

Analysis 1.7. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 7 Serious maternal outcome.

Study or subgroup	Treatment	Control		Risk Ratio				Weight	Risk Ratio
	n/N	n/N		M-H	, Fixed, 959	% CI			M-H, Fixed, 95% CI
Nabhan 2014	0/820	0/820							Not estimable
Ovalle 1998	0/55	0/50							Not estimable
Passos 2012	0/78	0/83							Not estimable
Total (95% CI)	953	953							Not estimable
Total events: 0 (Treatment), 0 (Control	l)								
Heterogeneity: Not applicable									
Test for overall effect: Not applicable									
	Fa	vours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	

Analysis 1.8. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 8 Chorioamnionitis (suspected or proven).

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Study or subgroup	Treatment	Control			Ri	sk Rat	io			Weight	Risk Ratio
	n/N	n/N			M-H, Ra	ndom	, 95% CI				M-H, Random, 95% CI
Cararach 1998	12/371	17/362					-			36.04%	0.69[0.33,1.42]
Nabhan 2014	21/820	19/820			_					41.22%	1.11[0.6,2.04]
Ovalle 1998	1/55	4/50	←	•						8.05%	0.23[0.03,1.97]
Passos 2012	2/78	9/83	←	+						14.69%	0.24[0.05,1.06]
Total (95% CI)	1324	1315								100%	0.65[0.34,1.26]
Total events: 36 (Treatment),	49 (Control)										
Heterogeneity: Tau ² =0.17; Chi	² =5.07, df=3(P=0.17); l ² =40.88	3%									
Test for overall effect: Z=1.27(P=0.2)										
	Fa	vours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotic	S

Analysis 1.9. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 9 Endometritis.

Study or subgroup	Treatment	Control				Weight	Risk Ratio				
	n/N	n/N	/N M-H, Random, 95% Cl								M-H, Random, 95% CI
Cararach 1998	0/371	4/362	-							22.06%	0.11[0.01,2.01]
Nabhan 2014	5/820	2/820					-		→	34.22%	2.5[0.49,12.85]
Ovalle 1998	0/55	5/50	←				-			22.44%	0.08[0,1.46]
Passos 2012	0/78	2/83	←	•				_		21.27%	0.21[0.01,4.36]
Total (95% CI)	1324	1315								100%	0.34[0.05,2.31]
Total events: 5 (Treatment), 13	B (Control)										
Heterogeneity: Tau ² =2.06; Chi ²	e6.74, df=3(P=0.08); l ² =55.46	5%									
Test for overall effect: Z=1.1(P=	=0.27)		1								
	Fa	vours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotic	S

Analysis 1.10. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 10 Caesarean section.

Study or subgroup	Treatment	Control		Ri	sk Ratio			Weight	Risk Ratio
	n/N	n/N		M-H, F	ixed, 95°	% CI			M-H, Fixed, 95% CI
Nabhan 2014	165/820	122/820						84.23%	1.35[1.09,1.67]
Ovalle 1998	10/55	7/50			++			5.06%	1.3[0.54,3.15]
Passos 2012	17/78	16/83			+			10.7%	1.13[0.62,2.08]
Total (95% CI)	953	953			•			100%	1.33[1.09,1.61]
Total events: 192 (Treatment),	145 (Control)								
Heterogeneity: Tau ² =0; Chi ² =0.	.3, df=2(P=0.86); l ² =0%								
Test for overall effect: Z=2.82(F	P=0)								
	Fa	vours antibiotics	0.2	0.5	1	2	5	Favours no antibiotics	

Analysis 1.11. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 11 Operative vaginal birth.

Study or subgroup	Treatment	Control			Risk Ratio			Weight	Risk Ratio
	n/N	n/N		M-I	H, Fixed, 95%	CI			M-H, Fixed, 95% Cl
Nabhan 2014	7/820	12/820		-				32.98%	0.58[0.23,1.47]
Ovalle 1998	3/55	2/50			+			5.76%	1.36[0.24,7.83]
Passos 2012	24/78	23/83			+			61.26%	1.11[0.69,1.8]
Total (95% CI)	953	953			•			100%	0.95[0.63,1.44]
Total events: 34 (Treatment),	37 (Control)								
Heterogeneity: Tau ² =0; Chi ² =1	1.63, df=2(P=0.44); l ² =0%								
Test for overall effect: Z=0.24((P=0.81)						1		
	Fa	wours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	

Analysis 1.12. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 12 Internal fetal monitoring.

Study or subgroup	Treatment	Control			Risk Ratio			Weight	Risk Ratio
	n/N	n/N		M-H	l, Fixed, 95	% CI			M-H, Fixed, 95% Cl
Nabhan 2014	0/820	0/820							Not estimable
Ovalle 1998	0/55	0/50							Not estimable
Total (95% CI)	875	870							Not estimable
Total events: 0 (Treatment), 0 (Control))								
Heterogeneity: Not applicable									
Test for overall effect: Not applicable									
	Fa	vours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	

Analysis 1.13. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 13 Epidural analgesia.

Study or subgroup	Treatment	Control			Risk Ratio			Weight	Risk Ratio
	n/N	n/N		M-I	l, Fixed, 95%	CI			M-H, Fixed, 95% CI
Ovalle 1998	55/55	50/50			•			39.51%	1[0.96,1.04]
Passos 2012	78/78	83/83			•			60.49%	1[0.98,1.02]
Total (95% CI)	133	133						100%	1[0.98,1.02]
Total events: 133 (Treatment)	, 133 (Control)								
Heterogeneity: Tau ² =0; Chi ² =0	, df=1(P=1); l ² =0%								
Test for overall effect: Not app	licable								
	Fa	vours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	

Analysis 1.14. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 14 Postpartum haemorrhage.

Study or subgroup	Treatment	Control			Risk Ratio			Weight	Risk Ratio	
	n/N	n/N		м-н,	Random, 9	5% CI			M-H, Random, 95% CI	
Nabhan 2014	35/820	16/820				-		70.85%	2.19[1.22,3.92]	
Ovalle 1998	0/55	0/50							Not estimable	
Passos 2012	0/78	2/83				_		29.15%	0.21[0.01,4.36]	
Total (95% CI)	953	953						100%	1.11[0.14,8.93]	
Total events: 35 (Treatment),	18 (Control)									
Heterogeneity: Tau ² =1.51; Chi	² =2.23, df=1(P=0.14); l ² =55.08	8%								
Test for overall effect: Z=0.1(P	=0.92)						1			
	Fa	vours antibiotics	0.01	0.1	1	10	100	Favours no antibiotic	S	

Analysis 1.15. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 15 Postpartum pyrexia.

Study or subgroup	Treatment	Control		Risk Ratio						Weight	Risk Ratio
	n/N	n/N			M-H, Ra	ndom	, 95% CI				M-H, Random, 95% CI
Nabhan 2014	28/820	29/820			_		_			0%	0.97[0.58,1.61]
Ovalle 1998	1/55	8/50	-			-	1			0%	0.11[0.01,0.88]
	Fav	ours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotics	5

Analysis 1.16. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 16 Postpartum septicaemia.

Study or subgroup	Treatment	Control			Risk Ratio			Weight	Risk Ratio
	n/N	n/N		M-H	I, Fixed, 95	% CI			M-H, Fixed, 95% Cl
Nabhan 2014	0/820	0/820							Not estimable
Ovalle 1998	0/55	0/50							Not estimable
Passos 2012	0/78	0/83							Not estimable
Total (95% CI)	953	953							Not estimable
Total events: 0 (Treatment), 0 (Control)								
Heterogeneity: Not applicable									
Test for overall effect: Not applicable						I.			
	Fa	vours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	

Analysis 1.17. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 17 Wound infection.

Study or subgroup	Treatment	Control			Risk Ratio			Weight	Risk Ratio
	n/N	n/N		M-H	l, Fixed, 95	% CI			M-H, Fixed, 95% Cl
Nabhan 2014	11/820	14/820						100%	0.79[0.36,1.72]
Ovalle 1998	0/55	0/50							Not estimable
Passos 2012	0/78	0/83							Not estimable
	Fa	vours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	

Antibiotics for prelabour rupture of membranes at or near term (Review)

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Study or subgroup	Treatment	Control			Risk Ratio			Weight	Risk Ratio
	n/N	n/N		М-	H, Fixed, 95	% CI			M-H, Fixed, 95% CI
Total (95% CI)	953	953			•			100%	0.79[0.36,1.72]
Total events: 11 (Treatment), 14 (Cont	trol)								
Heterogeneity: Not applicable									
Test for overall effect: Z=0.6(P=0.55)									
	Fa	wours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	

Analysis 1.18. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 18 Maternal postpartum antibiotic usage.

Study or subgroup	Treatment	Control		Risk Ratio			Weight	Risk Ratio
	n/N	n/N	M-	H, Random, 9	5% CI			M-H, Random, 95% CI
Nabhan 2014	21/820	19/820		-+			0%	1.11[0.6,2.04]
Ovalle 1998	55/55	8/50					0%	5.95[3.22,10.99]
	Fa	vours antibiotics	0.05 0.2	1	5	20	Fougurs no ontibiotic	~

Favours antibiotics 0.05 0.2 1 5 20 Favours no antibiotics

Analysis 1.19. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 19 Maternal adverse effects.

Study or subgroup	Treatment	Control		R	isk Rati	io		Weight	Risk Ratio
	n/N	n/N		М-Н, Р	ixed, 9	5% CI			M-H, Fixed, 95% Cl
Cararach 1998	1/371	0/362					_	100%	2.93[0.12,71.63]
Nabhan 2014	0/820	0/820				_			Not estimable
Ovalle 1998	0/55	0/50							Not estimable
Passos 2012	0/78	0/83							Not estimable
Total (95% CI)	1324	1315					-	100%	2.93[0.12,71.63]
Total events: 1 (Treatment), 0 (Control	l)								
Heterogeneity: Not applicable									
Test for overall effect: Z=0.66(P=0.51)									
	Fa	wours antibiotics	0.002	0.1	1	10	500	Favours no antibiotics	

Analysis 1.20. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 20 Neonatal meningitis.

Study or subgroup	Treatment	Control			Ri	sk Rat	io			Weight	Risk Ratio
	n/N	n/N			M-H, F	ixed, 9	95% CI				M-H, Fixed, 95% CI
Cararach 1998	1/371	3/362	€		+	_		-		100%	0.33[0.03,3.11]
Nabhan 2014	0/820	0/820				ĺ					Not estimable
Ovalle 1998	0/55	0/50									Not estimable
Passos 2012	0/78	0/83									Not estimable
Total (95% CI)	1324	1315						-		100%	0.33[0.03,3.11]
Total events: 1 (Treatment), 3 (Control)					İ					
Heterogeneity: Not applicable											
	Fa	vours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotics	;



Study or subgroup	Treatment n/N	Control n/N			Ris M-H, Fi	sk Ra ixed,				Weight	Risk Ratio M-H, Fixed, 95% CI
Test for overall effect: Z=0.97(P=0.33)											
		Favours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotics	

Analysis 1.21. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 21 Neonatal pneumonia.

Study or subgroup	Treatment	Control	Control Risk Ratio			Weight	Risk Ratio				
	n/N			М-Н, F	ixed, 9	95% CI				M-H, Fixed, 95% Cl	
Cararach 1998	0/371	1/362	←						_	100%	0.33[0.01,7.96]
Nabhan 2014	0/820	0/820									Not estimable
Ovalle 1998	0/55	0/50									Not estimable
Passos 2012	0/78	0/83									Not estimable
Total (95% CI)	1324	1315								100%	0.33[0.01,7.96]
Total events: 0 (Treatment), 1 (Control))										
Heterogeneity: Not applicable											
Test for overall effect: Z=0.69(P=0.49)											
	Fa	vours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotics	

Analysis 1.22. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 22 Apgar score < 7 at 5 minutes.

Study or subgroup	Treatment	Control			Ri	sk Rat	io			Weight	Risk Ratio
	n/N	n/N		M-H, Fixed, 95% Cl							M-H, Fixed, 95% CI
Cararach 1998	5/371	5/362		-		-		-		41.96%	0.98[0.28,3.34]
Nabhan 2014	15/820	7/820				-	+			58.04%	2.14[0.88,5.23]
Ovalle 1998	0/55	0/50									Not estimable
Passos 2012	0/78	0/83									Not estimable
Total (95% CI)	1324	1315								100%	1.65[0.81,3.36]
Total events: 20 (Treatment), :	12 (Control)										
Heterogeneity: Tau ² =0; Chi ² =1	.03, df=1(P=0.31); I ² =2.88%										
Test for overall effect: Z=1.39(F	P=0.17)										
	Fa	vours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotics	

Analysis 1.23. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 23 Admission to neonatal special care nursery.

Study or subgroup	Treatment	Control		Risk Ratio		Weight	Risk Ratio
	n/N	n/N	М-Н, І	Random, 95% C	l		M-H, Random, 95% Cl
Ovalle 1998	1/55	8/50				40.78%	0.11[0.01,0.88]
Passos 2012	3/78	5/83				59.22%	0.64[0.16,2.58]
Total (95% CI)	133	133				100%	0.32[0.06,1.74]
	Fa	vours antibiotics	0.01 0.1	1 1	0 100	Favours no antibiotic	S



Study or subgroup	Treatment	Control			Risk Ratio			Weight	Risk Ratio
	n/N	n/N		м-н,	Random, 9	5% CI			M-H, Random, 95% CI
Total events: 4 (Treatment), 1	3 (Control)								
Heterogeneity: Tau ² =0.77; Chi	² =1.97, df=1(P=0.16); l ² =49.	16%							
Test for overall effect: Z=1.32(P=0.19)								
	F	avours antibiotics	0.01	0.1	1	10	100	Favours no antibioti	cs

Analysis 1.24. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 24 Admission to neonatal intensive care unit.

Study or subgroup	Treatment	Control			Ri	sk Rat	io			Weight	Risk Ratio
	n/N	n/N			M-H, F	ixed, 9	5% CI				M-H, Fixed, 95% Cl
Nabhan 2014	43/820	37/820					_			92.72%	1.16[0.76,1.78]
Ovalle 1998	0/55	0/50									Not estimable
Passos 2012	6/78	3/83					+		_	7.28%	2.13[0.55,8.22]
Total (95% CI)	953	953				-	•			100%	1.23[0.82,1.85]
Total events: 49 (Treatment), 4	10 (Control)										
Heterogeneity: Tau ² =0; Chi ² =0	.7, df=1(P=0.4); I ² =0%										
Test for overall effect: Z=1.01(F	P=0.31)										
	Fa	vours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotics	

Analysis 1.25. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 25 Neonatal antibiotic usage.

Study or subgroup	Treatment	Control	Risk Ratio					Weight	Risk Ratio	
	n/N	n/N		м-н,	Random, 9	5% CI			M-H, Random, 95% CI	
Nabhan 2014	43/820	37/820						46.76%	1.16[0.76,1.78]	
Ovalle 1998	1/55	10/50		•				22.16%	0.09[0.01,0.69]	
Passos 2012	3/78	5/83						31.08%	0.64[0.16,2.58]	
Total (95% CI)	953	953						100%	0.55[0.15,1.97]	
Total events: 47 (Treatment),	52 (Control)									
Heterogeneity: Tau ² =0.87; Chi	² =6.55, df=2(P=0.04); l ² =69.44	8%								
Test for overall effect: Z=0.92(P=0.36)									
	Fa	vours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	5	

Analysis 1.26. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 26 Respiratory distress syndrome.

Study or subgroup	Treatment	Control		Risk Ratio			Weight	Risk Ratio
	n/N	n/N	м	-H, Fixed, 95%	5 CI			M-H, Fixed, 95% CI
Ovalle 1998	0/55	0/50						Not estimable
Passos 2012	0/78	0/83						Not estimable
Total (95% CI)	133	133						Not estimable
	Fa	vours antibiotics	0.01 0.1	1	10	100	Favours no antibiotics	3



Study or subgroup	Treatment	Control	Risk Ratio			Weight	Risk Ratio		
	n/N	n/N		M-H	I, Fixed, 95	5% CI			M-H, Fixed, 95% CI
Total events: 0 (Treatment), 0 (Contr	rol)								
Heterogeneity: Not applicable									
Test for overall effect: Not applicable	e								
		Favours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	

Analysis 1.27. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 27 Neonatal mechanical ventilation.

Study or subgroup	Treatment	Control			Ri	sk Rat	io			Weight	Risk Ratio	
	n/N	n/N			M-H, F	ixed, 9	95% CI				M-H, Fixed, 95% CI	
Cararach 1998	3/371	4/362			-	•				44.75%	0.73[0.16,3.25]	
Nabhan 2014	8/820	5/820				_	•			55.25%	1.6[0.53,4.87]	
Ovalle 1998	0/55	0/50									Not estimable	
Passos 2012	0/78	0/83									Not estimable	
Total (95% CI)	1324	1315								100%	1.21[0.5,2.91]	
Total events: 11 (Treatment),	9 (Control)											
Heterogeneity: Tau ² =0; Chi ² =0	0.68, df=1(P=0.41); l ² =0%											
Test for overall effect: Z=0.43(P=0.67)											
	Fa	vours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotics		

Analysis 1.28. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 28 Duration of maternal stay in hospital (days).

Study or subgroup	Tre	eatment	c	ontrol		Me	an Difference		Weight	Mean Difference
	N	Mean(SD)	Ν	Mean(SD)		Fi	xed, 95% CI			Fixed, 95% CI
Nabhan 2014	820	0.5 (0.6)	820	0.5 (0.5)					98.02%	0.06[0.01,0.11]
Ovalle 1998	55	3.7 (1.1)	50	3.6 (1.7)	-			\rightarrow	0.78%	0.1[-0.45,0.65]
Passos 2012	78	2.6 (1.7)	83	2.7 (1.1)	←				1.19%	-0.1[-0.55,0.35]
Total ***	953		953						100%	0.06[0.01,0.11]
Heterogeneity: Tau ² =0; Chi ² =	0.51, df=2(P=0.7	7); I ² =0%								
Test for overall effect: Z=2.35	(P=0.02)									
			Favou	urs antibiotics	-0.2	-0.1	0 0.1	0.2	Favours no	antibiotics

Analysis 1.29. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 29 Duration of neonatal stay in hospital (days).

Study or subgroup	Tre	eatment	c	ontrol	Mean Difference	Weight	Mean Difference
	Ν	Mean(SD)	Ν	Mean(SD)	Fixed, 95% CI		Fixed, 95% CI
Ovalle 1998	55	3 (1.2)	50	3.9 (1.1)	+	100%	-0.9[-1.34,-0.46]
Total ***	55		50		•	100%	-0.9[-1.34,-0.46]
Heterogeneity: Not applicable							
			Favou	urs antibiotics -10	-5 0 5	¹⁰ Favours no	antibiotics

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Study or subgroup	Т	reatment	Control			Me	an Differer	ıce		Weight	Mean Difference
	Ν	Mean(SD)	Ν	Mean(SD)	Fixed, 95		ixed, 95% (95% CI			Fixed, 95% CI
Test for overall effect: Z=3.99(P<0.0001)				_							
			Favo	urs antibiotics	-10	-5	0	5	10	Favours no a	antibiotics

Analysis 1.30. Comparison 1 Any antibiotic compared with placebo or no antibiotic, Outcome 30 Duration of neonatal stay in intensive care unit (days).

Study or subgroup	Tre	eatment	Control		Mean Difference	Weight	Mean Difference
	Ν	Mean(SD)	Ν	Mean(SD)	Fixed, 95% CI		Fixed, 95% CI
Nabhan 2014	820	0.3 (1.5)	820	0.3 (1.3)	-	100%	0.05[-0.09,0.19]
Total ***	820		820		•	100%	0.05[-0.09,0.19]
Heterogeneity: Not applicable							
Test for overall effect: Z=0.72(P=0.	47)						
			Favou	irs antibiotics	-0.5 -0.25 0 0.25 0.5	Favours no	antibiotics

Comparison 2. Any antibiotic compared with placebo or no antibiotic (subgrouped by timing of induction of labour)

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
1 Probable early-onset neonatal sepsis	3		Risk Ratio (M-H, Fixed, 95% CI)	Subtotals only
1.1 Early induction of labour	1	1640	Risk Ratio (M-H, Fixed, 95% CI)	1.47 [0.80, 2.70]
1.2 Late induction of labour	2	838	Risk Ratio (M-H, Fixed, 95% CI)	0.14 [0.02, 1.13]
2 Definite early-onset neonatal sepsis	3		Risk Ratio (M-H, Fixed, 95% CI)	Subtotals only
2.1 Early induction of labour	1	1640	Risk Ratio (M-H, Fixed, 95% CI)	1.29 [0.48, 3.44]
2.2 Late induction of labour	2	838	Risk Ratio (M-H, Fixed, 95% CI)	0.16 [0.02, 1.34]
3 Maternal infectious morbidi- ty (chorioamnionitis and/or en- dometritis)	3		Risk Ratio (M-H, Random, 95% CI)	Subtotals only
3.1 Early induction of labour	1	1640	Risk Ratio (M-H, Random, 95% CI)	1.15 [0.64, 2.08]
3.2 Late induction of labour	2	838	Risk Ratio (M-H, Random, 95% CI)	0.34 [0.08, 1.47]
4 Stillbirth	2		Risk Ratio (M-H, Fixed, 95% CI)	Subtotals only
4.1 Early induction of labour	1	1640	Risk Ratio (M-H, Fixed, 95% CI)	3.0 [0.61, 14.82]
4.2 Late induction of labour	1	105	Risk Ratio (M-H, Fixed, 95% CI)	0.0 [0.0, 0.0]
5 Perinatal mortality	3		Risk Ratio (M-H, Fixed, 95% CI)	Subtotals only

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Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
5.1 Early induction of labour	1	1640	Risk Ratio (M-H, Fixed, 95% CI)	3.0 [0.61, 14.82]
5.2 Late induction of labour	2	838	Risk Ratio (M-H, Fixed, 95% CI)	0.98 [0.14, 6.89]
6 Neonatal mortality	2		Risk Ratio (M-H, Fixed, 95% CI)	Subtotals only
6.1 Early induction of labour	1	1640	Risk Ratio (M-H, Fixed, 95% CI)	0.0 [0.0, 0.0]
6.2 Late induction of labour	1	105	Risk Ratio (M-H, Fixed, 95% CI)	0.0 [0.0, 0.0]
7 Serious maternal outcome	2		Risk Ratio (M-H, Fixed, 95% CI)	Subtotals only
7.1 Early induction of labour	1	1640	Risk Ratio (M-H, Fixed, 95% CI)	0.0 [0.0, 0.0]
7.2 Late induction of labour	1	105	Risk Ratio (M-H, Fixed, 95% CI)	0.0 [0.0, 0.0]

Analysis 2.1. Comparison 2 Any antibiotic compared with placebo or no antibiotic (subgrouped by timing of induction of labour), Outcome 1 Probable early-onset neonatal sepsis.

Study or subgroup	Treatment	Control		Risk Ratio		Weight	Risk Ratio
	n/N	n/N		M-H, Fixed, 95% C	I		M-H, Fixed, 95% CI
2.1.1 Early induction of labour							
Nabhan 2014	25/820	17/820			-	100%	1.47[0.8,2.7]
Subtotal (95% CI)	820	820				100%	1.47[0.8,2.7]
Total events: 25 (Treatment), 17 (Cont	rol)						
Heterogeneity: Not applicable							
Test for overall effect: Z=1.24(P=0.21)							
2.1.2 Late induction of labour							
Cararach 1998	1/371	7/362				100%	0.14[0.02,1.13]
Ovalle 1998	0/55	0/50					Not estimable
Subtotal (95% CI)	426	412				100%	0.14[0.02,1.13]
Total events: 1 (Treatment), 7 (Control)						
Heterogeneity: Not applicable							
Test for overall effect: Z=1.85(P=0.06)							
Test for subgroup differences: Chi ² =4.5	5, df=1 (P=0.03), I ² =7	7.78%					
	Fa	vours antibiotics	0.1 0.2	0.5 1 2	5 1	¹⁰ Favours no antibiotic	S

Analysis 2.2. Comparison 2 Any antibiotic compared with placebo or no antibiotic (subgrouped by timing of induction of labour), Outcome 2 Definite early-onset neonatal sepsis.

Study or subgroup	ly or subgroup Treatment			Control Risk Ratio						Weight	Risk Ratio
	n/N	n/N			М-Н, Р	ixed,	95% CI				M-H, Fixed, 95% CI
2.2.1 Early induction of labour											
Nabhan 2014	9/820	7/820						_		100%	1.29[0.48,3.44]
Subtotal (95% CI)	820	820			-					100%	1.29[0.48,3.44]
	Fa	wours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotics	



Study or subgroup	Treatment	Control			R	isk Rat	io			Weight	Risk Ratio
	n/N	n/N		M-H, Fixed, 95% Cl							M-H, Fixed, 95% CI
Total events: 9 (Treatment), 7 (Control)										
Heterogeneity: Not applicable											
Test for overall effect: Z=0.5(P=0.62)											
2.2.2 Late induction of labour											
Cararach 1998	1/371	6/362		+						100%	0.16[0.02,1.34]
Ovalle 1998	0/55	0/50									Not estimable
Subtotal (95% CI)	426	412								100%	0.16[0.02,1.34]
Total events: 1 (Treatment), 6 (Control)										
Heterogeneity: Not applicable											
Test for overall effect: Z=1.69(P=0.09)											
Test for subgroup differences: Chi ² =3.0)3, df=1 (P=0.08), I ² =	66.95%									
	Fa	wours antibiotics	0.1	0.2	0.5	1	2	5	10	Favours no antibiotics	

Analysis 2.3. Comparison 2 Any antibiotic compared with placebo or no antibiotic (subgrouped by timing of induction of labour), Outcome 3 Maternal infectious morbidity (chorioamnionitis and/or endometritis).

Study or subgroup	Treatment	Control			Risk Ra	tio			Weight	Risk Ratio	
	n/N	n/N	n/N		Randon	n, 95% Cl				M-H, Random, 95% Cl	
2.3.1 Early induction of labour											
Nabhan 2014	23/820	20/820				<u> </u>			100%	1.15[0.64,2.08]	
Subtotal (95% CI)	820	820							100%	1.15[0.64,2.08]	
Total events: 23 (Treatment), 20 (Co	ntrol)										
Heterogeneity: Not applicable											
Test for overall effect: Z=0.46(P=0.64	1)										
2.3.2 Late induction of labour											
Cararach 1998	12/371	21/362							68.36%	0.56[0.28,1.12]	
Ovalle 1998	1/55	8/50	-						31.64%	0.11[0.01,0.88]	
Subtotal (95% CI)	426	412				-			100%	0.34[0.08,1.47]	
Total events: 13 (Treatment), 29 (Co	ntrol)										
Heterogeneity: Tau ² =0.7; Chi ² =2.16,	df=1(P=0.14); I ² =53.69	%									
Test for overall effect: Z=1.45(P=0.15	5)										
Test for subgroup differences: Chi ² =	2.29, df=1 (P=0.13), I ² =	56.36%									
	Fa	vours antibiotics	0.1	0.2 0.	5 1	2	5	10	Favours no antibiotic	s	

Analysis 2.4. Comparison 2 Any antibiotic compared with placebo or no antibiotic (subgrouped by timing of induction of labour), Outcome 4 Stillbirth.

Study or subgroup	Treatment	Control		Risk Ratio				Weight	Risk Ratio
	n/N	n/N		M-H	, Fixed, 95	% CI			M-H, Fixed, 95% CI
2.4.1 Early induction of labour									
Nabhan 2014	6/820	2/820						100%	3[0.61,14.82]
Subtotal (95% CI)	820	820						100%	3[0.61,14.82]
Total events: 6 (Treatment), 2 (Control	l)								
Heterogeneity: Not applicable									
Test for overall effect: Z=1.35(P=0.18)									
	Fa	vours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	

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Study or subgroup T	reatment	Control			Risk Ratio)		Weight	Risk Ratio
	n/N	n/N	M-H, Fixed, 95% Cl					M-H, Fixed, 95% CI	
2.4.2 Late induction of labour									
Ovalle 1998	0/55	0/50							Not estimable
Subtotal (95% CI)	55	50							Not estimable
Total events: 0 (Treatment), 0 (Control)									
Heterogeneity: Not applicable									
Test for overall effect: Not applicable									
Test for subgroup differences: Not applica	able								
	Fa	avours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	

Analysis 2.5. Comparison 2 Any antibiotic compared with placebo or no antibiotic (subgrouped by timing of induction of labour), Outcome 5 Perinatal mortality.

Study or subgroup	Treatment	Control	Risk Ratio	Weight	Risk Ratio
	n/N	n/N	M-H, Fixed, 95% Cl		M-H, Fixed, 95% Cl
2.5.1 Early induction of labour					
Nabhan 2014	6/820	2/820		100%	3[0.61,14.82]
Subtotal (95% CI)	820	820		100%	3[0.61,14.82]
Total events: 6 (Treatment), 2 (Contro)				
Heterogeneity: Not applicable					
Test for overall effect: Z=1.35(P=0.18)					
2.5.2 Late induction of labour					
Cararach 1998	2/371	2/362		100%	0.98[0.14,6.89]
Ovalle 1998	0/55	0/50			Not estimable
Subtotal (95% CI)	426	412		100%	0.98[0.14,6.89]
Total events: 2 (Treatment), 2 (Contro)				
Heterogeneity: Not applicable					
Test for overall effect: Z=0.02(P=0.98)					
Test for subgroup differences: Chi ² =0.7	'6, df=1 (P=0.38), I ² =	0%			
	Fa	vours antibiotics	0.1 0.2 0.5 1 2 5 10	^D Favours no antibiotio	CS

Analysis 2.6. Comparison 2 Any antibiotic compared with placebo or no antibiotic (subgrouped by timing of induction of labour), Outcome 6 Neonatal mortality.

Study or subgroup	Treatment	Control			Risk Ratio			Weight	Risk Ratio
	n/N	n/N		M-H	, Fixed, 959	% CI			M-H, Fixed, 95% CI
2.6.1 Early induction of labour									
Nabhan 2014	0/820	0/820							Not estimable
Subtotal (95% CI)	820	820							Not estimable
Total events: 0 (Treatment), 0 (Control)								
Heterogeneity: Not applicable									
Test for overall effect: Not applicable									
2.6.2 Late induction of labour									
Ovalle 1998	0/55	0/50							Not estimable
Subtotal (95% CI)	55	50							Not estimable
	Fa	vours antibiotics	0.01	0.1	1	10	100	Favours no antibiotics	



Study or subgroup	Treatment	Control			Risk Ratio			Weight	Risk Ratio
	n/N	n/N	M-H, Fixed, 95% CI					M-H, Fixed, 95% CI	
Total events: 0 (Treatment), 0 (0	Control)								
Heterogeneity: Not applicable									
Test for overall effect: Not appli	cable								
Test for subgroup differences: N	lot applicable								
		Favours antibiotics	0.01	0.1	1	10	100	Favours no antibiotic	S

Analysis 2.7. Comparison 2 Any antibiotic compared with placebo or no antibiotic (subgrouped by timing of induction of labour), Outcome 7 Serious maternal outcome.

Study or subgroup T	reatment	Control		I	Risk Ratio		Weight	Risk Ratio
	n/N	n/N	M-H, Fixed, 95% CI					M-H, Fixed, 95% CI
2.7.1 Early induction of labour								
Nabhan 2014	0/820	0/820						Not estimable
Subtotal (95% CI)	820	820						Not estimable
Total events: 0 (Treatment), 0 (Control)								
Heterogeneity: Not applicable								
Test for overall effect: Not applicable								
2.7.2 Late induction of labour								
Ovalle 1998	0/55	0/50						Not estimable
Subtotal (95% CI)	55	50						Not estimable
Total events: 0 (Treatment), 0 (Control)								
Heterogeneity: Not applicable								
Test for overall effect: Not applicable								
Test for subgroup differences: Not application	able					1		
	Fa	vours antibiotics	0.01	0.1	1 10	100	Favours no antibiotics	5

APPENDICES

Appendix 1. Searches carried out in the previous version

The authors conducted a systematic literature search which included electronic databases: the Cochrane Controlled Trials Register (*The Cochrane Library* 2001, Issue 4) and MEDLINE (1965 to 2001), using MeSH headings: pregnancy and childbirth, infant-newborn and the search terms: term, chorioamnionitis, membrane*, rupture*, prelabour, prelabor, ROM, antibiotic*, neonat*, sepsis, early onset sepsis.

The authors also contacted recognised experts and cross referenced relevant material.

WHAT'S NEW

Date	Event	Description
8 September 2014	New citation required and conclusions have changed	Two new trials of 1801 women have been incorporated. The pre- vious reduction in endometritis is no longer evident. A 'Summary of findings' table has been added to this update.
31 July 2014	New search has been performed	Search updated. Two trials identified. Methods updated.



HISTORY

Protocol first published: Issue 4, 1999 Review first published: Issue 3, 2002

Date	Event	Description
23 December 2008	New search has been performed	Search updated. No new trials identified. Two studies previously awaiting classification have been excluded (Gordon 1974; Walss Rodriguez 1988).
28 August 2008	Amended	Converted to new review format.
30 September 2005	New search has been performed	Search updated. No new trials identified.

CONTRIBUTIONS OF AUTHORS

Aleena M Wojcieszek and Owen Stock undertook data extraction and quality assessments for studies included in this review update. Aleena M Wojcieszek led revisions of the review for the current update. All authors assisted with the interpretation and final editing.

Vicki Flenady and James King worked as equal partners in the production of the original review.

DECLARATIONS OF INTEREST

None.

SOURCES OF SUPPORT

Internal sources

- Centre for Clinical Studies Mater Hospital, South Brisbane, Queensland, Australia.
- J P Kelly Research Foundation, Mater Hospital, South Brisbane, Queensland, Australia.
- Department of Perinatal Medicine, Royal Women's Hospital, Melbourne, Victoria, Australia.

External sources

• Department of Health and Ageing, Commonwealth Government, Canberra, Australia.

DIFFERENCES BETWEEN PROTOCOL AND REVIEW

For this update outcome measures were divided into primary and secondary outcomes. Primary outcomes were selected by the authorship team as those most representative of the clinically important measures of effectiveness and complications. Respiratory distress syndrome was added as a secondary outcome. Methods for assessing risk of bias were updated as per Chapter 8 of the *Cochrane Handbook for Reviews of Intervention* (Higgins 2011). These changes include evaluating the domain of blinding separately for *participants and personnel* (performance bias) and for *outcome assessment* (detection bias). The quality of the evidence was re-assessed using the GRADE approach (Schunemann 2009) in order to assess the quality of the body of evidence relating to key outcomes.

INDEX TERMS

Medical Subject Headings (MeSH)

*Antibiotic Prophylaxis; *Fetal Membranes, Premature Rupture; Bacterial Infections [*prevention & control]; Chorioamnionitis [prevention & control]; Risk Assessment; Treatment Outcome

MeSH check words

Female; Humans; Infant, Newborn; Pregnancy