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# Combined hormonal versus nonhormonal versus progestin-only

contraception in lactation (Review)
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#### [Intervention Review]

# Combined hormonal versus nonhormonal versus progestin-only contraception in lactation

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#### **ABSTRACT**

#### **Background**

Postpartum contraception improves the health of mothers and children by lengthening birth intervals. For lactating women, contraception choices are limited by concerns about hormonal effects on milk quality and quantity and passage of hormones to the infant. Ideally, the contraceptive chosen should not interfere with lactation or infant growth. Timing of contraception initiation is also important. Immediately postpartum, most women have contact with a health professional, but many do not return for follow-up contraceptive counseling. However, immediate initiation of hormonal methods may disrupt the onset of milk production.

# **Objectives**

To determine the effects of hormonal contraceptives on lactation and infant growth

#### Search methods

We searched for eligible trials until 2 March 2015. Sources included the Cochrane Central Register of Controlled Trials (CENTRAL), PubMed, POPLINE, Web of Science, LILACS, ClinicalTrials.gov, and ICTRP. We also examined review articles and contacted investigators.

# Selection criteria

We sought randomized controlled trials in any language that compared hormonal contraception versus another form of hormonal contraception, nonhormonal contraception, or placebo during lactation. Hormonal contraception includes combined or progestin-only oral contraceptives, injectable contraceptives, implants, and intrauterine devices.

Trials had to have one of our primary outcomes: breast milk quantity or biochemical composition; lactation initiation, maintenance, or duration; infant growth; or timing of contraception initiation and effect on lactation. Secondary outcomes included contraceptive efficacy while breastfeeding and birth interval.

# Data collection and analysis

For continuous variables, we calculated the mean difference (MD) with 95% confidence interval (CI). For dichotomous outcomes, we computed the Mantel-Haenszel odds ratio (OR) with 95% CI. Due to differing interventions and outcome measures, we did not aggregate the data in a meta-analysis.



#### **Main results**

In 2014, we added seven trials for a new total of 11. Five reports were published before 1985 and six from 2005 to 2014. They included 1482 women. Four trials examined combined oral contraceptives (COCs), and three studied a levonorgestrel-releasing intrauterine system (LNG-IUS). We found two trials of progestin-only pills (POPs) and two of the etonogestrel-releasing implant. Older studies often lacked quantified results. Most trials did not report significant differences between the study arms in breastfeeding duration, breast milk composition, or infant growth. Exceptions were seen mainly in older studies with limited information.

For breastfeeding duration, two of eight trials indicated a negative effect on lactation. A COC study reported a negative effect on lactation duration compared to placebo but did not quantify results. Another trial showed a lower percentage of the LNG-IUS group breastfeeding at 75 days versus the nonhormonal IUD group (reported P < 0.05) but no significant difference at one year.

For breast milk volume, two older studies indicated lower volume for the COC group versus the placebo group. One trial did not quantify results. The other showed lower means (mL) for the COC group, e.g. at 16 weeks (MD -24.00, 95% CI -34.53 to -13.47) and at 24 weeks (MD -24.90, 95% CI -36.01 to -13.79). Another four trials did not report any significant difference between the study groups in milk volume or composition with two POPs, a COC, or the etonogestrel implant.

Seven trials studied infant growth; one showed greater weight gain (grams) for the etonogestrel implant versus no method for six weeks (MD 426.00, 95% CI 58.94 to 793.06) but less compared with depot medroxyprogesterone acetate (DMPA) from 6 to 12 weeks (MD -271.00, 95% CI -355.10 to -186.90). The others studied POPs, COCs versus POPs, or an LNG-IUS.

#### **Authors' conclusions**

Results were not consistent across the 11 trials. The evidence was limited for any particular hormonal method. The quality of evidence was moderate overall and low for three of four placebo-controlled trials of COCs or POPs. The sensitivity analysis included six trials with moderate quality evidence and sufficient outcome data. Five trials indicated no significant difference between groups in breastfeeding duration (etonogestrel implant insertion times, COC versus POP, and LNG-IUS). For breast milk volume or composition, a COC study showed a negative effect, while an implant trial showed no significant difference. Of four trials that assessed infant growth, three indicated no significant difference between groups. One showed greater weight gain in the etonogestrel implant group versus no method but less versus DMPA.

#### PLAIN LANGUAGE SUMMARY

#### Hormonal and nonhormonal birth control during breastfeeding

Birth control for women who are breastfeeding is important worldwide. Delaying the next pregnancy improves the health of women and children. Each year, millions of women decide whether to use birth control after having a baby. The decision includes the birth control type and when to start using it. Researchers and health care providers debate these issues. Some worry that hormones could affect the breast milk and therefore the baby's growth. Ideally, the birth control would not affect the type or amount of breast milk or the baby's growth. Identifying the best time to start birth control is also important. When monthly cycles return is uncertain, and the woman could get pregnant again.

Combined birth control methods contain the hormones estrogen and progestin. Other types of birth control contain only progestin or no hormones. We looked at whether combined birth control or methods with only progestin affect breastfeeding more than other methods. We ran computer searches for randomized trials of birth control used during breastfeeding until 2 March 2015. These trials compared hormonal methods to other hormonal methods or to placebo ('dummy' method). We also looked at reference lists to find trials. For the initial review, we wrote to researchers to find other studies.

We included 11 studies with a total of 1482 women. These trials looked at many methods: pills, an implant, the injectable 'Depo,' and a hormonal intrauterine device (IUD). Some older reports did not have much data. Most trials showed no major difference due to hormonal birth control use. Two of eight trials noted less breastfeeding among women using hormonal birth control. One was a combined pill with few results and the other a hormonal IUD. In one study, the implant group infants gained more weight than those in the no-method group but less weight than infants in the 'Depo' group. Two trials noted that a combined pill had a negative effect on breast milk volume or content. One report did not have much data. The other showed lower volume for combined pill users than for women taking pills with only progestin.

We found little information on any specific birth control method, with usually two studies per method. Results were not consistent across all trials. The data were of moderate quality overall. The results of better quality showed little effect on breastfeeding or infant growth.



#### BACKGROUND

#### **Description of the condition**

Contraception for women who are breastfeeding is a public health issue of global importance. According to early Demographic and Health Surveys (DHS), nearly two-thirds of women in their first postpartum year have an unmet need for family planning (WHO 2013). More recent DHS indicate that unmet need for modern contraceptive methods is 32% among married women in general, and one-third to one-half of unmarried women have an unmet need for such methods (Westhoff 2012). Each year more than 100 million women make decisions about beginning or resuming contraception after childbirth (Tsui 1997). These decisions include the choice of contraceptive method and the time at which to begin use. For women who are breastfeeding, the choice and timing of hormonal contraception may influence both lactation and infant growth.

Breastfeeding has well-established health benefits. It provides the infant with complete nutrition up to six months, a safe food source, and immunological defense against infectious diseases (USAID 2009; AAP 2012). Breastfeeding has economic benefits from conserving funds that would be spent on milk substitutes (AAP 2012). Lactation is associated with reducing a woman's risk of type 2 diabetes and ovarian and breast cancer (Ip 2007). Infants exclusively breastfed for the first six months appear to have better health outcomes than those partially breastfed as of three or four months of age (Kramer 2012).

Breastfeeding influences the need for and timing of postpartum contraception. An interval of anovulation occurs after delivery, and the length of time until ovulation resumes depends on breastfeeding patterns, biological variation, nutrition, geography, culture, and socioeconomic factors (Knijff 2000). Lactation itself can be an effective form of temporary contraception, known as the lactational amenorrhea method (LAM) (Van der Wijden 2003). For LAM to be effective, the woman must fully (or nearly fully) breastfeed, have no menstrual bleeding, and be within six months of delivery (Kennedy 2011; K4 Health 2014). Return of menstruation and ovulation can be unpredictable in breastfeeding women. Therefore, the timing of contraception initiation is important. Ideally, the contraceptive method chosen should not interfere with lactation.

### **Description of the intervention**

Contraception after childbirth improves the health of mothers and infants by lengthening birth intervals. Women are more likely to report births or pregnancies as unintended when they occur within an interval of 24 months or less (Tsui 1997). Preventing such unintended pregnancies reduces health risks and helps avoid associated financial and psychological costs. A longer birth interval of 18 to 27 months decreases the risk of major maternal complications including death, third-trimester bleeding, puerperal endometritis, and anemia (Conde-Agudelo 2000; WHO 2005). To reduce risk for poor maternal and infant health outcomes, the World Health Organization has recommended waiting 24 months before attempting the next pregnancy (WHO 2005).

Hormonal contraception includes combined methods that contain both estrogen and progestin as well as progestin-only methods. Combined hormonal contraceptives include combined

oral contraceptives (COCs), combined injectables, and the combined vaginal ring and transdermal patch. Progestin-only methods include progestin-only pills (POPs), the injectable depot medroxyprogesterone acetate (DMPA), levonorgestrel and etonogestrel implants, and the levonorgestrel intrauterine system. A progesterone vaginal ring was developed for use during lactation and is available in some Latin American countries (RamaRao 2013).

#### How the intervention might work

Hormonal contraceptives, especially those containing estrogen, may impair lactation through their effect on prolactin, the hormone responsible for production of milk. During pregnancy, prolactin levels rise and they peak at delivery. However, during pregnancy, both estrogen and progesterone block the effect of prolactin on the breasts. After delivery, levels of both estrogen and progesterone drop markedly, and, without their inhibitory effects, prolactin initiates milk production. Infant suckling stimulates more prolactin, which then sustains milk production. Breast engorgement and full milk secretion start three to four days after delivery, when estrogen and progesterone have sufficiently cleared from the maternal circulation (Speroff 2004; Pang 2007).

Choices of contraception may be limited for lactating women because of concerns about potential negative hormonal effects on quality and quantity of milk, passage of hormones to the infant, and infant growth and development. Some studies have found negative effects on lactation from COCs but not from progestinonly contraception (Kapp 2010a; Kapp 2010b; Kennedy 2011). However, those studies had various ways of measuring effects on milk production, yielded inconsistent results, and generally did not show negative effects on infants. Theoretical concerns about milk production focus on the early postpartum period, i.e. the onset of lactogenesis (CDC 2011; Gurtcheff 2011). Gonadal steroids may also impact milk supply in established lactation, as suggested by evidence of slower growth among children who continue to breastfeed when their mothers conceive another pregnancy (Bohler 1996). To determine whether hormonal contraception affects milk production, studies ideally would quantify the rate of exclusive or full breastfeeding, as well as the supplementation that infants receive, among women using different contraceptive methods. In the setting of mixed feeding, comparisons of infant growth are difficult to interpret because mothers can compensate for differences in milk supply by formula supplementation.

Due to concerns about the effect on lactation, combined hormonal contraceptives are considered category 4 for breastfeeding women up to six weeks postpartum and category 3 for six weeks to six months (WHO 2009; CDC 2011). Category 3 means that the theoretical or proven risks usually outweigh the advantages of using the method. The method is not usually recommended unless more appropriate methods are unavailable or unacceptable (WHO 2009). Most progestin-only methods are considered category 3 for less than six weeks postpartum (WHO 2009). For the first month postpartum in the United States, they are considered category 2, i.e. the advantages of using the method generally outweigh the risks (CDC 2011).

#### Why it is important to do this review

Despite potential adverse effects of COCs on lactation, many women prefer this method (Erwin 1994). Combined oral contraceptives have many benefits, including familiarity with the



method, effectiveness, safety, reversibility, excellent cycle control, a decrease in menstrual cramps and pain, decreased days of bleeding and amount of blood loss, and other noncontraceptive benefits. Other hormonal methods, including the progestin-only pill, may not offer all of these advantages (Raymond 2011). Progestin-only pill use is estimated at only 0.4% in the United States, according to an analysis of data from the National Survey of Family Growth (Hall 2012). The percentage may be much higher for breastfeeding women, but the sample size was very small for postpartum women. Some women quit breastfeeding early so they can start the COC (Erwin 1994).

Examining the impact of long-acting reversible contraception (LARC) on lactation is also important. The levonorgestrel-releasing intrauterine system (LNG-IUS) and the etonogestrel-releasing implant (ETG implant) provide highly effective birth control; both are progesterone-only methods.

Clinical recommendations must be evidence-based if women are to make informed choices concerning contraception while breastfeeding. This updated review examined the effects of combined hormonal contraceptives and progestin-only contraceptives on lactation and infant growth.

#### **OBJECTIVES**

To determine the effects of hormonal contraceptives on lactation and infant growth

#### **METHODS**

#### Criteria for considering studies for this review

#### **Types of studies**

All randomized controlled trials (RCTs) reported in any language that compared hormonal contraception during lactation versus other hormonal contraception, nonhormonal contraception, or placebo.

#### **Types of participants**

Breastfeeding women of any age or parity who desired contraception

#### Types of interventions

Any form of hormonal contraception compared with another form of hormonal contraception, nonhormonal contraception, or placebo. Combined hormonal contraceptives include oral or injectable methods, vaginal rings and transdermal patches. Progestin-only contraceptives include oral and injectable methods, subdermal implants, and hormonal intrauterine devices.

# Types of outcome measures

For the 2015 update, we separated the original list of outcomes into primary and secondary outcomes. Trials must have reported on at least one of the primary outcomes, which focus on the effect of the hormonal contraceptive on lactation.

#### **Primary outcomes**

- · Quantity of milk
- Biochemical analysis of milk composition

- Initiation, maintenance and duration of lactation (any or fully breastfeeding)
- · Infant growth
- Timing of contraception initiation and its effects on lactation

#### Secondary outcomes

- Efficacy of contraceptive method while breastfeeding (pregnancy)
- · Birth interval

#### Search methods for identification of studies

#### **Electronic searches**

We searched for eligible RCTs until 2 March 2015. Sources included the Cochrane Central Register of Controlled Trials (CENTRAL), PubMed, POPLINE, Web of Science, and LILACS. We also searched for recent trials via ClinicalTrials.gov and the search portal of the International Clinical Trials Registry Platform (ICTRP). Search strategies are shown in Appendix 1. The initial review and the 2005 and 2008 updates also included EMBASE (Appendix 2).

#### Searching other resources

For the initial review, we began with review articles. We also examined reference lists of relevant articles and book chapters to seek publications comparing different forms of contraception in breastfeeding women and their effects on lactation. We contacted other investigators in the field to find publications that might have been missed, including unpublished reports. For the 2015 update, we also examined reference lists of reviews and relevant articles for other trials.

# **Data collection and analysis**

2003: For the initial review, the authors read titles and abstracts from the searches to assess whether trials appeared to meet the inclusion criteria. We retrieved the full text when necessary. The authors verified that included references were satisfactory and reviewed others that could have met the inclusion criteria. They resolved disagreements by consensus. They sought additional information from investigators of the original included trials. Two investigators responded to questions about randomization methods and blinding (Miller 1970; WHO 1984).

2005 to 2010: The authors reviewed titles and abstracts from the database searches to determine whether trials appeared to be eligible. We retrieved the full text when necessary to determine whether the trial met the inclusion criteria.

<u>2014</u>: We describe below the data collection and analysis methods used in the current version of this review.

### **Selection of studies**

We assessed for inclusion all titles and abstracts identified during the literature searches with no language limitation.

#### **Data extraction and management**

Two authors independently abstracted the data. One author entered the data into Review Manager (RevMan 2014), and a second author verified the accuracy. We resolved discrepancies through discussion. In Characteristics of included studies, we focused on



primary and secondary outcomes for this review, which may not include all outcomes from each study.

#### Assessment of risk of bias in included studies

We examined studies for methodological quality, according to recommended principles (Higgins 2011). We considered factors such as study design, methods used to generate the randomization sequence, allocation concealment, blinding, and losses to follow-up and to early discontinuation. We also examined the methods used for outcome assessment.

#### Measures of treatment effect

For continuous variables, we computed the mean difference (MD) with 95% confidence interval (CI) using a fixed-effect model. RevMan uses the inverse variance approach. For dichotomous outcomes, we calculated the Mantel-Haenszel odds ratio (OR) with 95% CI using a fixed-effect model. When multivariate analysis was conducted, we presented the results as reported by the trial investigators.

#### Dealing with missing data

We wrote to investigators to request missing data, such as sample sizes for analysis and actual numbers for outcomes presented in figures. However, we limited our requests to studies less than 10 years old, unless a report was produced within the past five years. Investigators are unlikely to have access to data from 10 years ago.

#### **Data synthesis**

To assess evidence quality and to address confidence in the effect estimates, we applied principles from GRADE (Grades of Recommendation, Assessment, Development and Evaluation) (Balshem 2011; Higgins 2011). When meta-analysis is not viable because of varied interventions, a Summary of findings table is not feasible. In this review, experimental and comparison interventions differed across the included trials. Outcomes assessed also varied, e.g. breast milk composition, reports of breastfeeding, or change in infant weight. Because of the heterogeneity of interventions and outcomes, we did not conduct a formal GRADE assessment with an evidence profile and Summary of findings table (Guyatt 2011).

Our quality assessment was based on the quality of evidence from the individual studies, which could be rated as high, moderate, low, or very low. We considered the evidence from RCTs to be of high quality initially, then downgraded it for each of the following: (1) no information on randomization sequence generation or allocation concealment, or one was clearly inadequate; (2) no blinding; (3) follow-up less than 8 weeks for infant growth or less than 12 weeks for breastfeeding; (4) losses greater than 20%; and (5) information missing on both blinding and losses.

# **Sensitivity analysis**

We examined a subgroup of trials that provided evidence of moderate or high quality and reported sufficient outcome data. Most of the older trials did not quantify results, limiting the interpretation of effect. Even recent trials might not have reported sufficient detail for interpreting the outcome data presented.

#### RESULTS

#### **Description of studies**

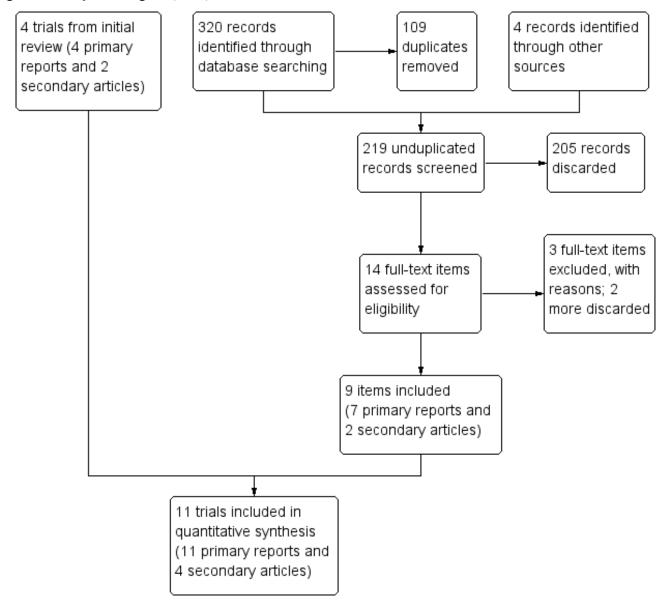
#### Results of the search

The initial search in 2003 identified 50 articles as potentially eligible for inclusion. Seven were included and 43 were excluded, as indicated below. Searches in 2005, 2008, and 2010 did not yield any eligible studies.

In 2014, we identified a study that was eligible but was not include earlier. Therefore, we ran searches starting from September 2001. This revised search, completed in 2015, produced 215 unduplicated references (Figure 1). Duplicates removed totaled 109 (93 identified electronically and 16 by hand). With four citations identified from other sources, we had 219 unduplicated references. We reviewed the full text of 14 items (9 primary and 5 secondary). From recent clinical trial listings, we obtained 15 unduplicated trials. We located a conference abstract for a completed trial, which we included, and we listed another trial in Ongoing studies. In this review, we included seven new primary reports plus two secondary articles. We excluded two new primary reports plus one secondary article. We also discarded two more secondary articles related to an included trial; the abstracts indicated they did not meet our criteria.



Figure 1. Study flow diagram (2014).



#### **Included studies**

A total of 11 trials met our inclusion criteria. The seven trials added in 2014 included three recent reports (Gurtcheff 2011; Espey 2012; Dutta 2013), a recent abstract (Stuart 2014), and three older reports missed during earlier searches (Heikkilä 1982; Shaamash 2005; Brito 2009).

The 11 trials included a total of 1482 women. Sample sizes ranged from 20 to 320 with a median of 80. Only four provided information on sample size calculations. Outcomes of focus were breastfeeding (Shaamash 2005; Gurtcheff 2011; Espey 2012) and infant growth (Shaamash 2005); one was designed to assess change in the woman's weight (Brito 2009).

# Contraceptive methods examined

Seven trials studied just progestin-only methods.

- Two compared progestin-only pills (POPs) versus placebo (Giner Velazquez 1976; Dutta 2013).
- Three studied a progestin-only intrauterine systems (IUS); two compared a hormonal IUS versus a nonhormonal IUD (Heikkilä 1982; Shaamash 2005), and one examined different insertion times for the same hormonal IUS (Stuart 2014).
- Two focused on a progestin-only implant; one compared the implant versus no contraceptive for the first six weeks followed by a progestin-only injectable (Brito 2009), and the other examined different insertion times for the same implant (Gurtcheff 2011).

Four trials examined COCs.

- Two studied a COC versus placebo (Semm 1966; Miller 1970).
- Two compared a COC versus a POP (WHO 1984; Espey 2012).



#### **Outcomes assessed**

- For breast milk, two studies examined quantity (Semm 1966; Miller 1970) and four reported on quantity and composition (Giner Velazquez 1976; WHO 1984; Gurtcheff 2011; Dutta 2013).
- Seven trials assessed any or full (exclusive) breastfeeding (Miller 1970; Heikkilä 1982; Shaamash 2005; Brito 2009; Gurtcheff 2011; Espey 2012; Stuart 2014).
- Seven studies reported infant growth (Giner Velazquez 1976; Heikkilä 1982; WHO 1984; Shaamash 2005; Brito 2009; Espey 2012; Dutta 2013).

#### **Excluded studies**

The initial review excluded 43 reports. Currently, we list articles as 'excluded' if the full text was needed to determine eligibility; otherwise, we 'discard' the abstract. Of the original 43 studies, 14 did not require full-text review. We removed those 14 studies from 'excluded' to shorten the list, and listed them for reference in this update (Appendix 3). Most of the remaining articles had been

excluded because communication with investigators indicated the trials were not RCTs, or because the method of participant allocation was unclear. In addition, Drury 1986 and Gellen 1984 were subgroup analyses that examined similar outcomes from WHO 1984 and were dropped from consideration.

In 2014, we divided the outcomes into primary and secondary and required one of the primary outcome measures. An earlier included trial was no longer eligible (Were 1997). The study had no outcomes addressing lactation; no pregnancy occurred in either group. We excluded three additional trials (Rodrigues da Cunha 2001; Chen 2011; Shaaban 2013) for design issues or for lack of our primary outcomes (Characteristics of excluded studies).

#### Risk of bias in included studies

The included trials span a publication period of nearly 50 years; earlier reports typically have more reporting limitations than later ones. The quality of evidence is discussed below and is summarized later (Table 1). In addition, Figure 2 presents the risk of bias for each study, and Figure 3 summarizes the risk of bias for all included trials.



Figure 2. Risk of bias summary: review authors' judgements about each risk of bias item for each included study.

	Random sequence generation (selection bias)	Allocation concealment (selection bias)	Blinding of participants and personnel (performance bias)	Incomplete outcome data (attrition bias)
Brito 2009	•	?	•	•
Dutta 2013	?	?	?	?
Espey 2012	•	•	•	?
Giner Velazquez 1976	?	?	•	?
Gurtcheff 2011	•	•	•	?
Heikkilä 1982	?	?	•	?
Miller 1970	?	?	•	•
Semm 1966	?	?	?	?
Shaamash 2005	•	•	?	?
Stuart 2014	•	•	•	•
WHO 1984	•	•	•	



Figure 3. Risk of bias graph: review authors' judgements about each risk of bias item presented as percentages across all included studies.



Four of the recent trials mentioned some type of computer-generated sequence (Shaamash 2005; Brito 2009; Gurtcheff 2011; Espey 2012), as did one earlier report (WHO 1984). Four of the five early trials and a recent preliminary report did not specify the method used to generate the random sequence (Semm 1966; Miller 1970; Giner Velazquez 1976; Heikkilä 1982; Dutta 2013). Additionally, Miller 1970 stratified the 'randomization' by gender of the infant and showed a disparity in baseline characteristics (primiparas) unlikely to result from a random process. Heikkilä 1982 also appeared to have an imbalance by parity, but the difference reportedly was not significant.

#### Allocation

Reporting of allocation concealment varied. One trial used pharmacy-blinded pill packages (Espey 2012). Two used sealed, sequentially numbered, opaque envelopes (Shaamash 2005; Gurtcheff 2011). We determined through communication with an investigator that another also used sealed, opaque envelopes (WHO 1984). The remaining trials did not report the method of allocation concealment.

#### **Blinding**

Four trials of oral contraceptives (OCs) had placebos; two used identically labeled placebos (Semm 1966; Miller 1970). Miller 1970 and Giner Velazquez 1976 mentioned double-blinding but did not describe the specifics. Written communication from Miller 1970 indicated that participants and clinicians were kept unaware of participant treatment assignments. Dutta 2013 did not mention blinding.

Two trials compared combined and progestin-only oral contraceptives. Espey 2012 used pharmacy-blinded pill packages. Written correspondence from an investigator in WHO 1984 (Tankeyoon) indicated that participants and clinicians were kept unaware of treatment assignments.

The two implant studies were not blinded. Brito 2009 was identified as an open trial of different contraceptive methods. Gurtcheff 2011 noted that blinding was not feasible because implants were inserted at different time points.

Of the three LNG-IUS studies, Heikkilä 1982 noted that participants were unaware of which IUD was inserted, Stuart 2014 was listed as open label, and Shaamash 2005 did not mention blinding.

#### Incomplete outcome data

Four trials did not mention any losses to follow-up, exclusions, or discontinuations (Semm 1966; Giner Velazquez 1976; Shaamash 2005; Dutta 2013). Heikkilä 1982 had no loss to follow-up. Three trials had losses less than 20% (Miller 1970; Brito 2009; Gurtcheff 2011).

Losses were greater than 20% in two trials (WHO 1984; Espey 2012). In WHO 1984, the disposition of participants in the randomized arms was unclear. One table indicated that 50 participants in each arm completed the study, yet tables with outcome data showed data for 57 and 58 participants in the study arms at 24 weeks (trial completion) (WHO 1984). Losses at 16 weeks were 22% for the COC group and 14% for the POP group. At 24 weeks, from 32% to 34% of participants in each randomized arm were not included in the analysis. The investigators stated that participants who discontinued or were lost to follow-up were analyzed via noncompeting risk life-table procedures, and at least one participant was excluded after randomization.

### **Effects of interventions**

#### Progestin-only pills (POPs) versus placebo

These two studies did not provide sufficient data for analysis. The older trial did not provide actual values for the outcomes and the newer trial report did not have sample sizes for analysis or information on losses. Both studies noted no significant differences between the study groups in breast milk volume and composition and in infant growth.

- o Giner Velazquez 1976 (n = 20) compared a POP containing norethindrone 350 μg versus a placebo for 14 days, starting 48 hours postpartum. Results were presented in figures without actual numbers. The investigators reported no significant differences between the study groups in milk volume and composition, nor in infant growth.
- In Dutta 2013 (n = 400), a POP containing desogestrel 75 µg was compared with placebo, starting six weeks postpartum. This 'preliminary report' stated the intervention was provided for six months. Results were presented in tables without time frames for assessments or specific sample sizes for analysis. The groups were not significantly different for amount or composition of breast milk nor for percent of infants with 'normal growth' (weight, length, and head circumference)



(reported P  $\geq$  0.15). Pregnancy was 0.5% in the POP group and 4% for the placebo group (reported P = 0.018).

# Progestin-only IUS versus nonhormonal IUD or different insertion times

Two studies compared a levonorgestrel-releasing intrauterine system (LNG-IUS) versus a nonhormonal intrauterine device (IUD). Neither provided sample sizes for analysis; results are presented as reported by the investigators.

- Heikkilä 1982 (n = 80) compared an LNG-IUS (30 µg) versus the Nova T IUD. Insertion was done at six weeks postpartum with a range from 29 to 56 days. Sample sizes were lacking in most tables.
  - o The source of breastfeeding data was unclear (record or recall); one analysis date (75 days after insertion) did not correspond to clinic visit times (3, 6, and 12 months after insertion). The investigators did not mention whether women were asked to record days of breastfeeding as they did for bleeding and spotting. The percentage of women breastfeeding at 75 days was lower in the LNG-IUS group than in the Nova T group (Analysis 1.1). However, the groups did not differ significantly in mean days of breastfeeding during the study (Analysis 1.2).
  - Women were asked to have their infants weighed and measured monthly and to record those numbers on a special card; results were shown in figures without actual numbers.
     Reportedly, infant growth did not differ between the two study groups.
  - A secondary article noted that no pregnancies occurred during the study.
- Shaamash 2005 (n = 320) compared the LNG-IUS (20 µg) versus the CuT 308A IUD. The article did not provide sample sizes for analysis or information on losses. Means and standard deviations (SDs) or standard errors (SEs) are presented as available for the 6- and 12-month assessments. The groups did not differ significantly for mean infant weight or length or in the Kaplan-Meier rates for full breastfeeding or continued use of the IUS or IUD (Analysis 2.1 to Analysis 2.4). Sample size for the study was based on detecting differences in breastfeeding rate and infant weight gain. No pregnancy occurred in either group.
- In Stuart 2014 (n = 35), the LNG-IUS was inserted either within 48 hours of delivery or four to eight weeks after delivery. The trial was stopped early because expulsion rates met a priori stopping rules. Only 35 women were randomized; 190 had been planned. At the six-month study visit, the groups were not significantly different for any breastfeeding, but the sample size was much smaller than was intended (Analysis 3.1).

# Progestin-only implant versus no method or delayed method

Two studies examined the etonogestrel-releasing (ETG) implant. Brito 2009 compared the implant versus no contraceptive for the first six weeks, then DMPA initiated at postpartum week six. In Gurtcheff 2011, both groups received the ETG implant, either by early insertion (one to three days postpartum) or standard insertion (four to eight weeks postpartum).

 In Brito 2009 (n = 40), the sample size calculation was based on the women's weight change and a metabolic measure, which are not outcomes for our review. Women with a body mass index
 ≥ 30 kg/m² were excluded from the trial, potentially affecting generalizability. The study groups did not differ significantly in proportions of women fully breastfeeding at 6 or 12 weeks postpartum (Analysis 4.1). None of the participants completely ceased breastfeeding during the study. Those who were not fully breastfeeding had started supplemental feeding. For infant weight, mean change (grams) by six weeks was greater in the ETG implant group than in the no-contraceptive group (MD 426.00, 95% CI 58.94 to 793.06; Analysis 4.2). The wide CI indicates imprecision in the estimate; the standard deviations were large. Mean change between 6 and 12 weeks was lower for the ETG implant group than for the group that received DMPA (MD -271.00, 95% CI -355.10 to -186.90; Analysis 4.2). The report did not provide the actual infant weights.

• For Gurtcheff 2011 (n = 69), the main outcomes were lactation failure and time to lactogenesis stage II (copious milk secretion). Sample size was based on establishing non-inferiority in those outcomes. Non-inferiority margins were a 15% increase in lactation failure and an additional eight hours to lactogenesis II in the early insertion group, as assessed through the 95% CI. Study groups were not significantly different for lactation failure (Analysis 6.1), mean time to lactogenesis (Analysis 6.2), or mean creamatocrit of breast milk at six weeks (Analysis 6.3). See Characteristics of included studies for explanations of these outcomes. However, our results differed from the non-inferiority results for time to lactogenesis; our 95% CI exceeded the prespecified margin of an additional eight hours (MD -0.90, 95% CI -9.90 to 8.10). For breastfeeding (full or any), the investigator provided the actual numbers for analysis. Early and standard insertion groups did not differ significantly for full breastfeeding (Analysis 6.4) or for any breastfeeding (Analysis 6.5) at any time point. About half of study participants were fully breastfeeding at two weeks.

# Combined oral contraceptive versus placebo

Findings from Miller 1970 and Semm 1966 were conflicting. Results were presented in figures without actual numbers. Neither trial quantified the outcomes, making interpretation difficult.

- Miller 1970 (n = 50) examined a COC containing norethindrone 1 mg plus mestranol 80 μg versus placebo. The investigators noted inhibitory effects on milk volume and lactation duration from COC use. They indirectly measured milk volume by assessing the subjective need for supplemental infant feeds and infant weight as a proxy for milk adequacy. Only general estimates were given for the effects of the COC on lactation duration.
- Semm 1966 (n = 100) compared a COC containing lynestrenol 2.5 mg plus mestranol 75 μg versus placebo. The investigators found no differences in milk volume or lactation initiation during the first 10 days postpartum. Hormone doses were larger than those in currently marketed COCs; therefore generalizability is limited.

#### Combined versus progestin-only oral contraceptives

The two studies in this group included an older multisite trial of OCs containing different progestins (WHO 1984) and a recent trial that compared OCs containing the same progestin (Espey 2012). The report of WHO 1984 provided data on milk volume, but did not have sufficient data to analyze infant growth in this review. Also, data on milk composition were presented by center, not for the full trial. Espey 2012 examined infant growth and breastfeeding continuation.



- WHO 1984 (n = 171) compared a COC containing levonorgestrel 150 μg plus ethinyl estradiol 30 μg versus a POP containing norgestrel 75 μg. Breast milk volume was determined by pump expression using standardized procedures. The participants breastfed their infants in the morning, waited two hours, and pumped milk while simultaneously nursing from the other breast for 20 minutes. The process was repeated two hours later, using the opposite breast for pumping. The 'average' amount was then presented. Mean milk volume (mL) was lower for the COC group than for the POP group at 9 weeks (MD -17.80, 95% CI -28.80 to -6.80), 16 weeks (MD -24.00, 95% CI -34.53 to -13.47), and 24 weeks (MD -24.90, 95% CI -36.01 to -13.79) (Analysis 7.1). Declines began after study initiation at six weeks postpartum and continued throughout the trial. From week 6 through week 24, average milk volume for COC users declined by 42% versus 12% among POP users. The randomized groups were not significantly different in infant weight change; sample sizes were not available for analysis (Analysis 7.2). Biochemical composition of breast milk was presented by center, not for the total sample. Differences within centers were small and inconsistent. Increases in milk lipid among combined contraceptive users in one center are of unknown clinical significance.
- Espey 2012 (n = 197) compared a COC containing norethindrone 1 mg plus ethinyl estradiol 35 µg versus a POP containing norethindrone 350 μg. The trial was powered to assess differences in continuation of breastfeeding. The investigator provided means and standard deviations for actual change and percentage change in infant weight and length. The study groups did not differ significantly for mean change in infant weight (Analysis 8.1) or in infant length (Analysis 8.2) from week two to week eight. The groups did not differ significantly at two or eight weeks for supplementing with formula. Results for the Cox proportional hazard regression indicate that the intervention group was not statistically associated with breastfeeding discontinuation (Analysis 8.3). The regression included the covariates of OC history and breastfeeding history due to baseline differences. However, formula supplementation and concerns about milk supply were associated with breastfeeding discontinuation (reported P = 0.033 and 0.005, respectively). No pregnancies were reported by eight weeks.

#### DISCUSSION

#### **Summary of main results**

Most trials did not show or report significant differences between study arms in breastfeeding duration, breast milk composition, or infant growth (Table 2). The few exceptions were seen mainly in older studies with limited reporting.

• Breastfeeding continuation was studied in eight trials. One older study noted a negative effect of a combined oral contraceptive (COC) on lactation duration but did not quantify results. An early trial of a levonorgestrel-releasing intrauterine system (LNG-IUS) showed a lower percentage of the LNG-IUS group breastfeeding at 75 days but no significant difference at one year. The other five trials indicated no significant differences between the study arms. Of those five, two studied insertion times for the etonogestrel-releasing (ETG) implant, one compared a COC versus a progestin-only pill (POP), and two examined the

- LNG-IUS. Of the latter, one compared the LNG-IUS versus a nonhormonal intrauterine device (IUD); the other compared insertion times and was stopped early because of expulsion rates.
- Milk volume or composition was examined in six trials. Two older studies indicated conflicting effects of a COC on milk volume, but neither quantified results. Another older trial showed lower volume for the COC versus POP group. Two placebo-controlled trials reported no significant effect of a POP on breast milk volume or composition, although one did not quantify results, and the other did not provide sample sizes for analysis. An ETG implant study showed no significant effect of early versus standard insertion.
- Infant growth was assessed in seven trials. One showed greater
  infant weight gain for the ETG implant group than for the nomethod group during the first six weeks. The implant group
  had less weight gain from 6 to 12 weeks when compared to
  those given depot medroxyprogesterone acetate (DMPA). The
  other six studied POPs only, COCs versus POPs, or the LNG-IUS,
  and indicated no significant differences between groups. Trial
  reports did not present the amount of supplemental feeding
  provided to the infants.
- Pregnancy was self-reported. One trial showed a lower percentage in the POP group. Three had no pregnancy in either group; they examined an LNG-IUS or a COC and POP.

#### Sensitivity analysis

This analysis was restricted to six studies that provided sufficient data and moderate-quality evidence (Table 3). Two examined the LNG-IUS, with one comparing it with a nonhormonal IUD and one examining insertion times. Two trials studied insertion times for the ETG implant, and two examined a COC versus a POP. Five of the six trials were published since 2005.

- Lactation duration: The five trials that examined breastfeeding duration indicated no significant differences between the study groups.
- Breast milk volume or composition: One study showed a negative effect of the COC versus a POP on milk volume (mL expressed) through 24 weeks postpartum (mean difference (MD) -24.90, 95% confidence interval (CI) -36.01 to -13.79), and a recent implant trial showed no significant difference between insertion times.
- Infant growth: Three trials indicated no significant differences in weight or length. One trial showed greater weight gain in the ETG implant group compared with the no-method group at six weeks (MD 426.00, 95% CI 58.94 to 793.06) but less gain compared with the DMPA group at 12 weeks (MD -271.00, 95% CI -355.10 to -186.90).

# Overall completeness and applicability of evidence

Most studies assessed breastfeeding duration (full or any) or breast milk volume or composition as well as infant growth. Four examined self-reported pregnancy. Time frames for contraceptive initiation in the experimental group varied: one to three days after delivery in five trials (implant, LNG-IUS, and OCs); two weeks postpartum in two trials (OCs); and approximately six weeks postpartum in four trials (OCs or IUS). Follow-up ranged from 10 days (oldest trial) to one year (more recent studies).



Studies were conducted in Asia, Europe, North Africa, and the Americas. Five trials published before 1985 included a multicenter trial conducted in Hungary and Thailand and four other trials from Finland, Germany, Mexico, and the United States. Six newer trials, published from 2005 to 2014, were conducted in Brazil, Egypt, India, and the United States (three studies).

Inclusion and exclusion criteria varied among trials. Most studies were limited to healthy women with birth at 37 weeks gestation or later. Some trials included any woman intending to initiate breastfeeding, whereas others were limited to women who planned to breastfeed exclusively (Brito 2009) or for a minimum duration (Miller 1970; Stuart 2014), had previously breastfed successfully (WHO 1984), or were exclusively breastfeeding at the time of randomization (Shaamash 2005).

The evidence was limited for any particular hormonal method. Four trials examined COCs: two versus placebo and two versus a POP. Three of the four studies were at least 30 years old and had limited reporting. Older studies used hormonal preparations and doses that may be applicable to contemporary practice. Seven trials focused on progestin-only methods. Two studied a POP versus placebo; the recent one was a preliminary report. Two trials compared ETG implant insertion times. Three examined an LNG-IUS (one had an older IUS); two compared the LNG-IUS versus a nonhormonal IUD, and one compared insertion times. We did not find any RCTs of a vaginal ring or a transdermal patch among lactating women. A trial of emergency contraception as a backup for the lactational amenorrhea method did not include data on our primary outcomes (Shaaban 2013).

This review examined infant growth rather than development or other health outcomes. Two included trials assessed infant health issues at 12 months. In Heikkilä 1982, the LNG-IUS and Nova T groups did not differ significantly for incidence of otitis media and respiratory infection as reported by the mothers. Shaamash 2005 assessed 19 infant development items, grouped as gross motor, vision and fine motor, hearing and language, self-help skills, and social skills (WHO 1986). No significant difference was noted between the LNG-IUS and CuT 380A groups in time to pass any test. An earlier non-randomized study of progestin-only methods used the same 19 items (WHO 1994). The investigators found a few differences that were mainly site-specific but no consistent trends. The progestin-only methods apparently had no negative effect on infant development. A systematic review of mainly observational studies came to the same conclusion (Kapp 2010a). In a review of combined oral contraceptives, the limited evidence indicated no adverse health outcomes in infants (Kapp 2010b).

#### Quality of the evidence

The quality of reporting varied over time. The earlier publications often provided little detail on methodology and provided insufficient data on results. Standards for publishing trials were developed in the late 1990s, and the CONSORT statement was widely adopted in 2010 (Schulz 2010). We summarized the quality of evidence from the individual trials (Table 1). Overall, the quality was considered moderate. The lower-quality evidence, often due to limited reporting, came from three of the four placebo-controlled trials of oral contraceptives.

Half of the trials provided inadequate information on randomization and allocation concealment. Two trials had high

losses to follow-up (WHO 1984; Espey 2012). Loss to follow-up greater than 20% threatens trial validity (Strauss 2005). Two trials with moderate-quality evidence provided insufficient or inconsistent data (Miller 1970; Heikkilä 1982). One trial was terminated early because expulsion rates met a priori stopping rules (Stuart 2014). The resulting sample size was much smaller than was planned, so the trial was underpowered to detect differences between groups in breastfeeding outcomes.

# Agreements and disagreements with other studies or reviews

Two systematic reviews of hormonal contraceptives and lactation included studies of various designs (Kapp 2010a; Kapp 2010b). Progestin-only methods did not appear to negatively affect lactation (Kapp 2010a). The evidence was limited and inconsistent for COCs and breastfeeding (Kapp 2010b). For infant growth and health, the evidence for progestin-only methods was consistent though limited but did not appear to show an adverse effect. The evidence for COCs was considered inadequate at the time for assessing the effect on infant health. Another systematic review found limited evidence regarding medroxyprogesterone use and breastfeeding at less than six weeks postpartum (Brownell 2012). Only three studies of limited quality were eligible and none adjusted for potential confounding. A non-randomized study of hormonal contraceptives and lactation examined a COC, the LNG-IUS, the etonogestrel implant, and a nonhormonal IUD (CuT 380A) initiated at 42 days postpartum (Bahamondes 2013). Participants had previously breastfed and were willing to breastfeed exclusively and on demand for the study duration. The investigators reported no significant differences between the study groups from 42 days to 63 days postpartum in infant milk intake, infant weight or height, or breastfeeding (also assessed at six months).

We did not review the evidence for venous thromboembolism (VTE) risk, which was a factor in developing medical eligibility criteria for hormonal contraceptive use by postpartum women (WHO 2009; CDC 2011). Physiological risk of VTE is high in the postpartum period, especially for the first three weeks, and declines by six weeks' postpartum (WHO 2010). However, direct evidence is lacking for risk of VTE with combined hormonal contraception.

# **AUTHORS' CONCLUSIONS**

#### Implications for practice

We found limited information on any particular hormonal method and lactation. The six trials of higher quality examined the levonorgestrel-releasing IUS, the etonogestrel-releasing implant, and two COCs versus progestin-only pills. Of those six trials, an older one showed a negative effect of the COC on breast milk volume. A later study within those six indicated greater infant weight gain with the implant versus no contraceptive but less gain compared with the injectable DMPA. Volume of supplementation was not assessed.

As noted earlier, because of concerns about the effect on lactation, combined hormonal contraceptives are considered category 4 for breastfeeding women up to six weeks postpartum, and category 3 for six weeks to six months (WHO 2009; CDC 2011). Category 3 means that the risks outweigh the advantages. Progestin-only methods are considered category 3 for less than six weeks' postpartum in global guidelines (WHO 2009), and category 2 for the



first month postpartum in the USA (CDC 2010). The latter means that the advantages generally outweigh the risks.

# Implications for research

The body of evidence regarding the effects of hormonal contraceptives on lactation has grown considerably. Six RCTs were published after a 20-year gap. The 11 trials in this review examined a range of contraceptive methods; consequently, evidence was limited for any specific method. Most early studies focused on oral contraceptives, often compared with placebo. Most of the later trials examined the levonorgestrel-releasing intrauterine system or the etonogestrel-releasing implant, with several comparing insertion times. The overall quality of evidence was moderate. Five

of the six newer trials provided moderate-quality data, as did one older multisite trial.

Further research on progestin-only methods and lactation would be beneficial, especially because some are long-acting methods. The field could benefit from additional research into the effects of initiation time on lactation and infant health. To more accurately assess those outcomes, future research should assess the mother's breastfeeding intention, duration of full and any breastfeeding, and amount of supplemental feeding.

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# CHARACTERISTICS OF STUDIES

**Characteristics of included studies** [ordered by study ID]

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\* Indicates the major publication for the study

Methods	Time frame and location: recruitment from September 2007 to July 2008; São Paulo, Brazil
	Described as a pilot study. Further details on methods were provided in later publication (Brito 2012)
	Sample size calculation and outcomes of focus: women's body weight and resistance to activated protein C (not outcomes for this review); difference of 1 SD in first 6 weeks; 20 per group due to potential loss during study; power not reported
Participants	General with N and source: 40 women in prenatal care program
	Inclusion criteria: 18 and 35 years old, exclusively breastfeeding and desires long-acting contraception after delivery
	Exclusion criteria: smoker, alcoholic, or user of recreational drugs; body mass index (BMI) (kg/m²) ≥ 30 or systemic disease (diabetes mellitus, cardiovascular disease, liver disease, thyroid disease, autoimmune disease); history (personal or family) of thromboembolic events; alterations in hepatic enzymes or allergy to local anesthetics (xylocaine)
Interventions	Treatment: etonogestrel-releasing implant, early insertion (24 to 48 hours after delivery)
	Comparison: no contraceptive for 6 weeks after delivery, then depot medroxyprogesterone acetate at week 6 postpartum
	Timing: as noted above
Outcomes	Measures: infant weight gain; maintenance of exclusive lactation to week 12 postpartum
	Assessments: 6 and 12 weeks
Notes	



# Brito 2009 (Continued)

# Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Computerized; 1:1 ratio
Allocation concealment (selection bias)	Unclear risk	According to communication with investigator, randomized groups were stored in sealed envelopes by one investigator and only opened by another investigator.
Blinding of participants and personnel (perfor- mance bias) All outcomes	High risk	Open; ethical reasons
Incomplete outcome data (attrition bias) All outcomes	Low risk	No losses or discontinuations, according to study flow chart

# **Dutta 2013**

Bias	Authors' judgement Support for judgement
Risk of bias	
Notes	Investigator provided some design information; we were unable to obtain additional data for analysis and interpretation.
	Assessments: 8, 12, 16, and 24 weeks; 1 year
	(Normal infant growth defined as weight doubled by 4 months and tripled by 1 year, height 50 cm at birth increased to 62 cm by 6 months and 75 cm by 1 year, head circumference increased by 2 cm per month until l year.)
Outcomes	Measures: milk volume and composition, pregnancy, normal infant growth (see below); time frames and samples sizes for analysis not specified
	Timing: starting 6 weeks after delivery
	Comparison: placebo
Interventions	Treatment: oral contraceptive (OC) containing desogestrel 75 $\mu g$ for 6 months
	Exclusion criteria: family history or past history of thromboembolism
	Inclusion criteria: normal lipid profile, blood sugar levels, and thyroid stimulating hormone
Participants	General with N and source: 400 lactating women from clinic
	Sample size calculation and outcome of focus: no mention
	Limited details about methods; published as a preliminary report
Methods	Time frame and location: conducted from January 2010 through June 2011 in Kalyani, Nadia, West Be gal (India)



Dutta 2013 (Continued)		
Random sequence generation (selection bias)	Unclear risk	Interventions were provided "in a random sequence." Communication with investigator indicated that this was an RCT.
Allocation concealment (selection bias)	Unclear risk	No mention
Blinding of participants and personnel (perfor- mance bias) All outcomes	Unclear risk	No mention
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	No mention
Espey 2012		
Methods	Time frame and locat (USA)	ion: conducted from January 2005 to June 2008 in Albuquerque, New Mexico
	for 80% power to deteratio = 2). Assumed 50	on and outcome of focus: breastfeeding continuation at 8 weeks; 120 participants ect continuation difference of 35% in COC group versus 60% in POP group (hazard 0% breastfeeding at 8 weeks. With potential 20% loss to follow-up, 150 needed; insee of high loss between enrollment and randomization.
Participants	General with N and so	ource: 197 postpartum women in hospital after delivery
		e 15 to 45 years; delivered at University of New Mexico Hospital; intended to to use oral contraceptives as family planning method
	bolism, uncontrolled	edical contraindication to combined pills, including history of venous thromboem- hypertension, or complex migraine headaches; preterm birth (< 37 weeks); new- ional age (< 2500 g) or large for gestational age (> 4500 g); newborn had major con-
Interventions		d oral contraceptive (COC) containing norethindrone 1 mg plus ethinyl estradiol and placebo for 7 days
	Comparison: OC cont	aining norethindrone 350 μg daily
	Timing: starting 2 wee	eks postpartum
Outcomes		ling continuation; infant weight, length, and head circumference (in figures with- self-reported pregnancy
	Assessments: phone of months postpartum	call 3 to 7 weeks postpartum and 4 to 6 months postpartum; hospital visit 2
Notes	Investigator provided	I means and SD for change in infant weight and length.
Risk of bias		

Generated in randomly-permuted blocks of 6.

Low risk

Random sequence genera-

tion (selection bias)



Espey 2012 (Continued)		
Allocation concealment (selection bias)	Low risk	Pharmacy-blinded packages; pills placed in red capsules; identical monthly pill dispensers
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Blinding of participants and study personnel
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Loss to follow-up: COC 23% (15/64); POP 22% (14/63)
		Women who stopped breastfeeding before 8 weeks were discontinued from the study.
		Infant growth was not obtained at 8 weeks: COC 12.5% (8/64); POP 14% (9/63).

# **Giner Velazquez 1976**

Methods	Time frame and location: no time frame; Mexico				
	Limited details about methods				
	Sample size calculation and outcome of focus: no mention; small sample size limits power				
Participants	General with N and source: 20 women selected immediately after birth				
	Inclusion criteria: age 18 to 36 years; adequate health and nutrition during pregnancy				
	Exclusion criteria: not specified				
Interventions	Treatment: OC containing norethindrone 350 μg				
	Comparison: placebo				
	Timing: starting 48 hours postpartum				
Outcomes	Measures: average milk production; biochemical composition of milk; infant weights; results primarily in figures without actual numbers				
	Assessments: 14 days				
Notes	Information from report translation				

# Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Divided randomly (12 treatment and 8 control)
Allocation concealment (selection bias)	Unclear risk	Not specified
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Double blind; no detail



**Giner Velazquez 1976** (Continued)

Incomplete outcome data (attrition bias)
All outcomes

Unclear risk

No mention of discontinuation, loss to follow-up or exclusions

# **Gurtcheff 2011**

Methods	Time frame and location: January to October 2009 in Salt Lake City, Utah (USA)
	Described as non-inferiority trial
	Sample size calculation and outcomes of focus:
	• lactation failure: assuming 5% failure in both groups, n = 27 per group for 80% power and 15% non-inferiority margin;
	• hours to lactogenesis II: assuming mean $54\pm12$ , n = 34 per group for $80\%$ power, and non-inferiority margin of $8$ additional hours
Participants	General with N and source: 69 healthy peripartum women
	Inclusion criteria: healthy; intended to breastfeed
	Exclusion criteria: onset of lactogenesis before randomization; hemorrhage requiring transfusion; severe pregnancy-induced hypertension; prolonged hospitalization; coagulopathy; liver disease; undiagnosed genital bleeding; other contraindication to etonogestrel implant insertion; taking inducers of hepatic enzymes
Interventions	Etonogestrel implant with different insertion times
	Treatment: early insertion, 1 to 3 days postpartum
	Comparison: standard insertion, 4 to 8 weeks postpartum
	Timing: See above
Outcomes	Measures: lactation failure (assessed 3 times daily in hospital and at least daily after discharge; diagnosed by 120 hours); time to lactogenesis stage II (copious milk secretion); breastfeeding, full or any (shown in figure without actual numbers); creamatocrit for fat and energy content of milk sample from 6-week visit
	Assessments: 2 and 6 weeks, 3 and 6 months postpartum
Notes	Investigator provided data for breastfeeding continuation (full or any).

# Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Computer-generated random numbers in varying blocks of 2, 4, 6
Allocation concealment (selection bias)	Low risk	Sequentially numbered, opaque, sealed envelopes
Blinding of participants and personnel (perfor- mance bias) All outcomes	High risk	Not blinded; considered not feasible



# **Gurtcheff 2011** (Continued)

Incomplete outcome data (attrition bias)
All outcomes

Unclear risk

Loss to follow-up: early insertion 17% (6/35); standard insertion 18% (6/34)

# Heikkilä 1982

Methods	Time frame and location: no time frame; conducted in Helsinki, Finland	
	Some methods from secondary article	
	Sample size calculation and outcome of focus: not specified	
Participants	General with N and source: 80 women, delivered at State Maternity Hospital	
	Inclusion criteria: breastfeeding and amenorrheic	
	Exclusion criteria: not specified	
Interventions	Treatment: levonorgestrel-releasing IUD (30 μg)	
	Comparison: Nova T intrauterine device (IUD) (copper releasing)	
	Timing: inserted 6 weeks postpartum; range 29 to 56 days in one report, and 32 to 56 days in another	
Outcomes	Measures: breastfeeding duration (no sample sizes for analysis); infant growth measured outside of study and presented in figure without actual numbers	
	Assessment: unclear source and time frame; 75 days after insertion for analysis of breastfeeding data; secondary report notes recording of bleeding and spotting on cards and clinic visits at 3, 6, 12 months	
Notes	Third group was added later (not randomized); received levonorgestrel-releasing IUD (10 μg)	

### Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	"Divided at random" into 2 groups
Allocation concealment (selection bias)	Unclear risk	Not specified
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Participants did not know which IUD they received (30 μg/day LNG-IUD or Nova T).
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Loss to follow-up: none Discontinuation (expulsion or removal): 4 LNG-IUD 30 µg; 4 Nova T. Also, 2 excluded from study due to uterine perforation at insertion (group not specified).  Sample sizes for analysis not specified or inconsistent across reports.

# **Miller 1970**

Methods	Time frame and location: September 1967 to June 1968 in Iowa City, Iowa (USA)



Miller 1970 (Continued)	
	Limited detail about methods
	Sample size calculation and outcome of focus: no mention
Participants	General with N and source: 50 women delivering healthy term infants at city hospital.
	Inclusion criteria: desire to nurse for 3 months and use oral contraceptives while nursing
	Exclusion criteria: none specified
Interventions	Treatment: COC containing norethindrone 1 mg plus mestranol 80 μg daily for 21 days
	Comparison: identically-labeled placebos for 21 days
	Timing: postpartum day 14; all received COC starting 6 weeks postpartum
Outcomes	Measures: lactation duration, milk volume production. Results presented in figures without actual numbers; graphic variations not identified.
	Assessments: 3 weeks, 5 weeks, and 3 months postpartum
Notes	

# Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Not specified
Allocation concealment (selection bias)	Unclear risk	Not specified
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Participants and investigators blinded
Incomplete outcome data (attrition bias) All outcomes	Low risk	Loss to follow-up: 1 in each arm; 1 discontinuation in placebo arm

# **Semm 1966**

Time frame and location: no time frame; Munich, Germany Limited detail about methods	
General with N and source: 100 women delivering with a definite desire to lactate	
Inclusion criteria: women in "child-bed"; with definite desire to lactate	
Exclusion criteria: not specified	
Treatment: COC containing lynestrenol 2.5 mg plus mestranol 75 μg daily	
Comparison: identically labeled placebos	



	Timing: postpartum da	y 1	
Outcomes	Measures: lactation initiation and milk volume yield; results in figure without actual numbers  Assessments: daily for 10 days; women also weighed babies at home in following 3 weeks, but data not shown		
Notes			
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Unclear risk	Distribution of 2 types of tablets was "purely random"	
Allocation concealment (selection bias)	Unclear risk	Not specified	
Blinding of participants and personnel (perfor- mance bias) All outcomes	Unclear risk	Not specified	
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Loss to follow-up: no mention, nor any mention of discontinuation or exclusions	
Shaamash 2005			
shaamash 2005 Methods	Time frame and location	on: conducted from June 2001 to June 2003 in Assiut, Egypt	
		on: conducted from June 2001 to June 2003 in Assiut, Egypt	
	Sample size calculation • Breastfeeding rate:	n and outcomes of focus: 151 per group for 80% power to detect difference of 15 per 100 women ales and 56 females for 90% power to detect 400 gram difference at 12 months,	
	<ul><li>Sample size calculation</li><li>Breastfeeding rate:</li><li>Infant weight: 56 m assuming 20% attrit</li></ul>	n and outcomes of focus: 151 per group for 80% power to detect difference of 15 per 100 women ales and 56 females for 90% power to detect 400 gram difference at 12 months,	
Methods	<ul> <li>Sample size calculation</li> <li>Breastfeeding rate:</li> <li>Infant weight: 56 m assuming 20% attrit</li> <li>General with N and sou</li> <li>Inclusion criteria: gave</li> </ul>	n and outcomes of focus: 151 per group for 80% power to detect difference of 15 per 100 women ales and 56 females for 90% power to detect 400 gram difference at 12 months, ion	
Methods	<ul> <li>Sample size calculation</li> <li>Breastfeeding rate:</li> <li>Infant weight: 56 m assuming 20% attrit</li> <li>General with N and sou</li> <li>Inclusion criteria: gave</li> </ul>	n and outcomes of focus:  151 per group for 80% power to detect difference of 15 per 100 women ales and 56 females for 90% power to detect 400 gram difference at 12 months, ion  Irce: 320 exclusively breastfeeding women; university hospital clinic birth to healthy term baby with no health problem; had no contraindication to er IUD contraceptives; planning to breastfeed for at least 1 year	
Methods	Sample size calculation  Breastfeeding rate:  Infant weight: 56 m assuming 20% attrit  General with N and sou  Inclusion criteria: gave the use of LNG or copp  Exclusion criteria: not s	n and outcomes of focus:  151 per group for 80% power to detect difference of 15 per 100 women ales and 56 females for 90% power to detect 400 gram difference at 12 months, ion  Irce: 320 exclusively breastfeeding women; university hospital clinic birth to healthy term baby with no health problem; had no contraindication to er IUD contraceptives; planning to breastfeed for at least 1 year	
Methods  Participants	Sample size calculation  Breastfeeding rate:  Infant weight: 56 m assuming 20% attrit  General with N and sou  Inclusion criteria: gave the use of LNG or copp  Exclusion criteria: not s	n and outcomes of focus:  151 per group for 80% power to detect difference of 15 per 100 women ales and 56 females for 90% power to detect 400 gram difference at 12 months, cion  arce: 320 exclusively breastfeeding women; university hospital clinic birth to healthy term baby with no health problem; had no contraindication to er IUD contraceptives; planning to breastfeed for at least 1 year specified  trel-releasing IUS (20 µg)	
Participants	Sample size calculation  Breastfeeding rate:  Infant weight: 56 m assuming 20% attrit  General with N and sou Inclusion criteria: gave the use of LNG or copp  Exclusion criteria: not so	n and outcomes of focus:  151 per group for 80% power to detect difference of 15 per 100 women ales and 56 females for 90% power to detect 400 gram difference at 12 months, ion  Irce: 320 exclusively breastfeeding women; university hospital clinic birth to healthy term baby with no health problem; had no contraindication to er IUD contraceptives; planning to breastfeed for at least 1 year specified  trel-releasing IUS (20 µg)  380A IUD	
Methods  Participants	Sample size calculation  Breastfeeding rate:  Infant weight: 56 m assuming 20% attrit  General with N and sou Inclusion criteria: gave the use of LNG or copp  Exclusion criteria: not s  Treatment: levonorges  Comparison: Copper T  Timing: inserted 6 to 8  Measures: breastfeedin	n and outcomes of focus:  151 per group for 80% power to detect difference of 15 per 100 women ales and 56 females for 90% power to detect 400 gram difference at 12 months, ion  Irce: 320 exclusively breastfeeding women; university hospital clinic birth to healthy term baby with no health problem; had no contraindication to er IUD contraceptives; planning to breastfeed for at least 1 year specified trel-releasing IUS (20 µg)  380A IUD	
Methods  Participants  Interventions	Sample size calculation  Breastfeeding rate:  Infant weight: 56 m assuming 20% attrit  General with N and sou Inclusion criteria: gave the use of LNG or copp  Exclusion criteria: not s  Treatment: levonorges  Comparison: Copper T  Timing: inserted 6 to 8  Measures: breastfeedin	n and outcomes of focus:  151 per group for 80% power to detect difference of 15 per 100 women ales and 56 females for 90% power to detect 400 gram difference at 12 months, ion  17ce: 320 exclusively breastfeeding women; university hospital clinic birth to healthy term baby with no health problem; had no contraindication to er IUD contraceptives; planning to breastfeed for at least 1 year specified trel-releasing IUS (20 µg)  380A IUD  weeks postpartum  19 (full, partial); infant growth (weight, length, head circumference, mid-arm cirinfold); pregnancy (presumably self-reported)	



# Shaamash 2005 (Continued)

Notes

Risk	of	bias
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Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Randomization scheme was computer generated.
Allocation concealment (selection bias)	Low risk	Sealed opaque envelopes; opened in sequential order at time of enrollment
Blinding of participants and personnel (perfor- mance bias) All outcomes	Unclear risk	No mention
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	No information on losses or sample sizes for analysis

#### Stuart 2014

Stuart 2014	
Methods	Time frame and location: March 2012 to June 2013; Chapel Hill, NC (USA)
	RCT, phase 4
	Sample size calculation and outcome of focus: no information. Planned to enroll 190; randomized 35. Study stopped early; expulsion rate met a priori stopping rules.
Participants	General with N and source: 35 women, 18 to 45 years old; Women's Hospital of North Carolina
	Inclusion criteria (for antepartum enrollment):
	Pregnant and ≥ 24 weeks of estimated gestational age; intent to breastfeed for at least 6 months; plan to use LNG-IUS postpartum; anticipation of vaginal delivery; HIV negative; intention to stay in Chapel Hill area for at least 6 months after birth; no medical or personal conditions that in the judgment of study staff preclude participation; no allergies to any component of LNG-IUS; no known uterine anomalies; fluent in English  No history of ectopic pregnancy, carcinoma of the breast, acute liver disease or liver tumor (benign or malignant); uterine or cervical neoplasia or unresolved abnormal pap smear; active pelvic inflammatory disease; hypersensitivity to component of LNG-IUS; genital bleeding of unknown etiology; solid organ transplantation
	Additional eligibility criteria for trial entry (assessed postpartum):  No endometritis or chorioamnionitis; membranes ruptured < 24 hours before delivery; no fever ≥ 38°C during intrapartum or postpartum period; did not receive medication other than pitocin or misoprosto to control postpartum bleeding; did not have blood loss > 750 mL intrapartum, blood transfusion for postpartum hemorrhage, third or fourth degree laceration at delivery; infant > 35 weeks gestation as determined by physical exam; weight at least 2727 grams; singleton birth; not in intensive care nursery; not diagnosed with condition that precludes long-term feeding
	Exclusion criteria: no other mention
Interventions	Levonorgestrel-releasing IUS (20 $\mu$ g) with different insertion times
	Treatment: inserted within 48 hours of delivery



Stuart 2014 (Continued)	Comparison: inserted 4 to 8 weeks after delivery	
Outcomes	Measures: any breastfeeding	
	Assessments: 6 months	
Notes	Information from published conference abstract, ClinicalTrials.gov listing, and communication wi vestigator	
	Full report to be submitted for publication by mid-September, according to investigator.	

# Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence genera- tion (selection bias)	Low risk	1:1 allocation
tion (selection bias)		Computer generated by someone not involved with study activities, according to communication with investigator
Allocation concealment (selection bias)	Low risk	Sequentially numbered opaque envelopes, according to communication with investigator
Blinding of participants and personnel (perfor- mance bias) All outcomes	High risk	Open label
Incomplete outcome data (attrition bias) All outcomes	Low risk	Six-month analysis for LNG-IUS continuation included all women randomized.

# **WHO 1984**

VHO 1984			
Methods	Time frame and location: no time frame; conducted in Szeged, Hungary; Bangkok, Thailand; Khon Kaen, Thailand		
	Multicenter trial		
	Sample size calculation and outcome of focus: no information		
Participants	General with N and source: 171 women delivering and chose to use OCs		
	Inclusion criteria: age 20 to 35 years; 2 to 4 live births; prior successful breastfeeding (3 months); hemoglobin 10 to 12 g/dL; normal gestation and uncomplicated singleton delivery (2700 g to 3700 g); no breast abnormalities; desiring to use hormonal contraception during lactation		
	Exclusion criteria: none specified		
Interventions	Treatment: COC containing levonorgestrel 150 μg plus ethinyl estradiol 30 μg		
	Comparison: OC containing norgestrel 75 μg		
	Timing: initiated at 6 weeks postpartum (± 3 days)		
	Study length: 24 weeks		
Outcomes	Measures:		



#### WHO 1984 (Continued)

- 1984 report (Tankeyoon): average breast milk volume; change in infant weight (no sample sizes for analysis); study withdrawal due to inadequate milk supply or infant growth (withdrawal by center, not total)
- 1986 report (Sas): milk composition (lipids and fatty acids); by center, not total
- 1988 report (WHO): milk composition (fat, nitrogen, caloric content, lactose, and osmolality) and infant weight; by center, not total

Assessments: 6 (baseline), 9, 12, 16, 20, 24 weeks postpartum

Notes 341 women were recruited; only 171 women were randomized. Others chose depot medroxyprogesterone acetate (DMPA) or nonhormonal methods.

#### Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Computer-generated sequence from WHO
Allocation concealment (selection bias)	Low risk	Sealed, opaque envelopes, according to communication with investigator
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Double-blind; pills packaged identically
Incomplete outcome data (attrition bias)	High risk	Loss to follow-up: unclear due to inconsistent reporting across and within reports
Alloutcomes		COC: 9% at 9 weeks, 22% at 16 weeks, 34% at 24 weeks
		POP: 2% at 9 weeks, 14% at 16 weeks, 32% at 24 weeks

COC: combined oral contraceptive

IUD: intrauterine device

LNG-IUS: levonorgestrel-releasing intrauterine system

OC: oral contraceptive

POP: progestin-only contraceptive WHO: World Health Organization

# **Characteristics of excluded studies** [ordered by study ID]

Study	Reason for exclusion
Abdel-Aleem 1996	Not a randomized controlled trial
Barsivala 1973	Unclear whether participant allocation was random
Bassol 2002	Outcomes not applicable for this review
Betrabet 1987	Per correspondence with the investigator (Betrabet), study participants received personal choice of intervention, not random allocation.
Bhatia 1987	Not a randomized controlled trial
Borglin 1971	Study described as a "carefully controlled trial." No evidence of randomization



Study	Reason for exclusion				
Chen 2011	Secondary analysis; trial of immediate versus delayed insertion of levonorgestrel-releasing intrauterine system (LNG-IUS)				
	Interim contraception varied for delayed-insertion group until LNG-IUS was inserted at 6 to 8 weeks postpartum; included condoms, depot medroxyprogesterone acetate (DMPA), abstinence, and none. Women were not randomized to specific method during those first few weeks.				
Croxatto 1982	Per correspondence with the investigator, study participants received personal choice of intervention, not random allocation.				
Croxatto 1983	Per correspondence with the investigator, study participants received personal choice of intervention, not random allocation.				
Diaz 1983	Per correspondence with the investigator, study participants received personal choice of intervention, not random allocation.				
Diaz 1985	Per correspondence with the investigator, study participants received personal choice of intervention, not random allocation.				
Diaz 1997	Not a randomized controlled trial				
Drury 1986	Subgroup analysis of WHO 1984				
Gellen 1984	Subgroup analysis of WHO 1984				
Gupta 1974	Not a randomized controlled trial				
Hefnawi 1970	Unclear whether participant allocation was random				
Kader 1969	Unclear whether participant allocation was random				
Kader 1975	Unclear whether participant allocation was random				
Kaern 1967	Unclear whether participant allocation was random				
Kamal 1969a	Unclear whether participant allocation was random				
Kamal 1969b	Unclear whether participant allocation was random				
Kamal 1970	Unclear whether participant allocation was random				
Karim 1971	Not a randomized controlled trial				
Koetsawang 1972a	Unclear whether participant allocation was random				
Koetsawang 1972b	Unclear whether participant allocation was random				
Peralta 1983	Per correspondence with the investigator, all trials in the series were not randomized. Participants chose their intervention.				
Rodrigues da Cunha 2001	Not an RCT; choice of contraceptive				
Seth 1977	Unclear whether participant allocation was random				
Shaaban 2013	Does not have any of our primary outcomes.				



Study	Reason for exclusion
Sinchai 1995	Not a randomized controlled trial
Toaff 1969	Trial does not include regularly marketed combined oral contraceptive. Intervention on postpartum days 1 to 5 further limited its usefulness.
Were 1997	No relevant outcome when outcomes were divided into primary and secondary in 2014; trial did not address lactation. Further, losses were > 85% in each arm.
Zanartu 1976	Not a randomized controlled trial

# **Characteristics of ongoing studies** [ordered by study ID]

# **Turok 2014**

1010K 2014				
Trial name or title	BLIS - Breastfeeding Levonorgestrel IUD Study			
Methods	Time frame and location: January 2014 to December 2015 in Albuquerque, NM, and Salt Lake City, UT (USA)			
	Randomized controlled trial, non-inferiority, phase 4; open label			
	Sample size calculation and outcome of focus: no mention			
Participants	General with N: 317 pregnant women			
	Inclusion criteria: healthy; 18 to 40 years old; pregnant; intention to breastfeed; desire for LNG-IUD (levonorgestrel-releasing intrauterine device) as method of contraception; delivered healthy term infant (37 weeks gestation)  Exclusion criteria: chorioamnionitis; obstetric complications including transfusion, severe pregnancy-induced hypertension, prolonged hospitalization, coagulopathy, liver disease, undiagnosed genital bleeding, or other relative contraindication to LNG-IUD insertion (known or suspected pregnancy, uterine cavity; abnormality, known, suspected, or history of breast cancer; or hypersensitivity to any of the components in the LNG-IUD)			
Interventions	Levonorgestrel-releasing IUD			
	Treatment: immediately after delivery of baby and placenta (within 10 minutes)			
	Comparison: 4 to 6 weeks after delivery			
	Timing: as above			
Outcomes	Measures: breastfeeding continuation and exclusivity; lactogenesis stage 2 (first 5 days after birth)			
	Assessments: 5 days (lactogenesis stage 2); 8 and 26 weeks (breastfeeding)			
Starting date	January 2014; estimated primary completion December 2015			
Contact information	Contact: Maria Masters; maria.masters@hsc.utah.edu; 801-213-2286			
	PI: David K Turok, MD; University of Utah, Department of Obstetrics and Gynecology			
Notes				



#### DATA AND ANALYSES

# Comparison 1. LNG-IUS (30 µg) versus nonhormonal IUD (Nova T)

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
1 Breastfeeding continuation at 75 days			Other data	No numeric data
2 Mean days of breastfeeding (over 12 months)			Other data	No numeric data

# Analysis 1.1. Comparison 1 LNG-IUS (30 $\mu g$ ) versus nonhormonal IUD (Nova T), Outcome 1 Breastfeeding continuation at 75 days.

#### Breastfeeding continuation at 75 days

Study	LNG-IUS	Nova T	Reported P
Heikkilä 1982	56%	79%	<.05

# Analysis 1.2. Comparison 1 LNG-IUS (30 $\mu$ g) versus nonhormonal IUD (Nova T), Outcome 2 Mean days of breastfeeding (over 12 months).

#### Mean days of breastfeeding (over 12 months)

Study	LNG-IUS (mean <u>+</u> SD)	Nova T (mean <u>+</u> SD)	Reported P
Heikkilä 1982	197 <u>+</u> 141	208 <u>+</u> 104	not significant (NS)

# Comparison 2. LNG-IUS (20 μg) versus nonhormonal IUD (CuT 380A)

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Reported mean infant weight (g)			Other data	No numeric data
2 Reported mean infant length (cm)			Other data	No numeric data
3 Reported net cumulative rate for full breast- feeding			Other data	No numeric data
4 Reported net cumulative rate for continuation of IUS or IUD			Other data	No numeric data

# Analysis 2.1. Comparison 2 LNG-IUS (20 $\mu$ g) versus nonhormonal IUD (CuT 380A), Outcome 1 Reported mean infant weight (g).

#### Reported mean infant weight (g)

Study	Time frame	LNG-IUS (mean <u>+</u> SD)	CuT 380A (mean <u>+</u> SD)	Reported P
Shaamash 2005	6 months	7225 <u>+</u> 810	7157 <u>+</u> 759	NS
Shaamash 2005	12 months	9284 <u>+</u> 423	9183 <u>+</u> 421	NS



# Analysis 2.2. Comparison 2 LNG-IUS (20 μg) versus nonhormonal IUD (CuT 380A), Outcome 2 Reported mean infant length (cm).

Reported mean infant length (cm)

Study	Time frame	LNG-IUS (mean <u>+</u> SD)	CuT 380A (mean <u>+</u> SD)	Reported P
Shaamash 2005	6 months	65.1 <u>+</u> 3.2	64.4 <u>+</u> 3.0	NS
Shaamash 2005	12 months	72.6 <u>+</u> 3.2	72.3 <u>+</u> 3.9	NS

# Analysis 2.3. Comparison 2 LNG-IUS (20 $\mu$ g) versus nonhormonal IUD (CuT 380A), Outcome 3 Reported net cumulative rate for full breastfeeding.

Reported net cumulative rate for full breastfeeding

Study	Time frame	LNG-IUS (mean <u>+</u> SE)	CuT 380A (mean <u>+</u> SE)
Shaamash 2005	6 months	16.5 <u>+</u> 3.4	19.9 <u>+</u> 3.6
Shaamash 2005	12 months	0.3 <u>+</u> 0.4	0.0 <u>+</u> 0.0

# Analysis 2.4. Comparison 2 LNG-IUS (20 $\mu$ g) versus nonhormonal IUD (CuT 380A), Outcome 4 Reported net cumulative rate for continuation of IUS or IUD.

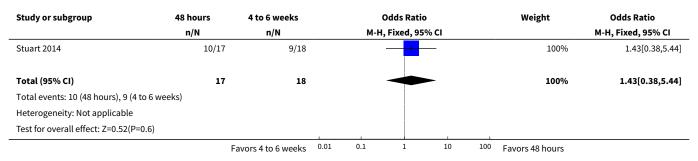
Reported net cumulative rate for continuation of IUS or IUD

Study	Time frame	LNG-IUS (mean <u>+</u> SE)	CuT 380A (mean <u>+</u> SE)
Shaamash 2005	6 months	90.6 <u>+</u> 2.3	93.5 <u>+</u> 2.0
Shaamash 2005	12 months	89.3 <u>+</u> 2.5	90.9 <u>+</u> 2.3

# Comparison 3. LNG-IUS insertion time after delivery: 48 hours versus 4 to 6 weeks

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
1 Any breastfeeding at 6 months	1	35	Odds Ratio (M-H, Fixed, 95% CI)	1.43 [0.38, 5.44]

# Analysis 3.1. Comparison 3 LNG-IUS insertion time after delivery: 48 hours versus 4 to 6 weeks, Outcome 1 Any breastfeeding at 6 months.

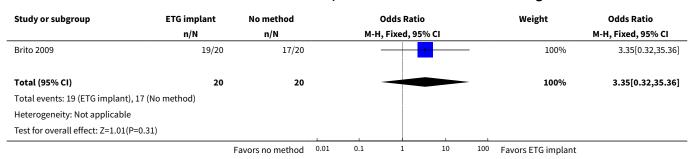




# Comparison 4. Etonogestrel implant (at 24 to 48 hours) versus no method until 6 weeks

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
1 Exclusive breastfeeding	1	40	Odds Ratio (M-H, Fixed, 95% CI)	3.35 [0.32, 35.36]
2 Mean change in infant weight (g)	1	40	Mean Difference (IV, Fixed, 95% CI)	426.0 [58.94, 793.06]

# Analysis 4.1. Comparison 4 Etonogestrel implant (at 24 to 48 hours) versus no method until 6 weeks, Outcome 1 Exclusive breastfeeding.



# Analysis 4.2. Comparison 4 Etonogestrel implant (at 24 to 48 hours) versus no method until 6 weeks, Outcome 2 Mean change in infant weight (g).

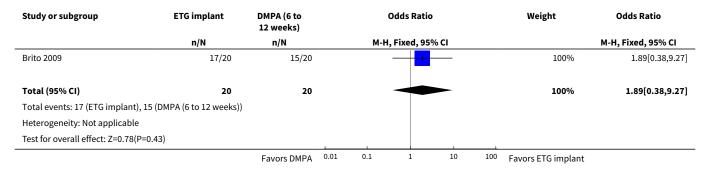
Study or subgroup	ETG	implant	No	method		M	ean Difference	Weight	<b>Mean Difference</b>
	N	Mean(SD)	N	Mean(SD)		ı	ixed, 95% CI		Fixed, 95% CI
Brito 2009	20	1461 (621)	20	1035 (562)				100%	426[58.94,793.06]
Total ***	20		20				-	100%	426[58.94,793.06]
Heterogeneity: Not applicable									
Test for overall effect: Z=2.27(P=0.02)								-1	
			Favo	rs no method	-1000	-500	0 500 10	<sup>00</sup> Favors E	TG implant

# Comparison 5. ETG implant (at 24 to 48 hours) versus DMPA from 6 to 12 weeks

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
1 Exclusive breastfeeding	1	40	Odds Ratio (M-H, Fixed, 95% CI)	1.89 [0.38, 9.27]
2 Mean change in infant weight (g)	1	40	Mean Difference (IV, Fixed, 95% CI)	-271.0 [-355.10, -186.90]



# Analysis 5.1. Comparison 5 ETG implant (at 24 to 48 hours) versus DMPA from 6 to 12 weeks, Outcome 1 Exclusive breastfeeding.



# Analysis 5.2. Comparison 5 ETG implant (at 24 to 48 hours) versus DMPA from 6 to 12 weeks, Outcome 2 Mean change in infant weight (g).

Study or subgroup	ETG	implant		IPA (6 to weeks)		Mean Dif	ference		Weight	Mean Difference
	N	Mean(SD)	N	Mean(SD)		Fixed, 9	5% CI			Fixed, 95% CI
Brito 2009	20	366 (89)	20	637 (170)	-				100%	-271[-355.1,-186.9]
Total ***	20		20		•				100%	-271[-355.1,-186.9]
Heterogeneity: Not applicable										
Test for overall effect: Z=6.32(P<0.000	01)									
				Favors DMPA	-400 -20	0 0	200	400	Favors ETG	implant

# Comparison 6. Etonogestrel implant: early versus standard insertion

Outcome or subgroup ti- tle	No. of studies	No. of partici- pants	Statistical method	Effect size
1 Lactation failure	1	69	Odds Ratio (M-H, Fixed, 95% CI)	3.18 [0.13, 80.79]
2 Mean time to lactogenesis stage II	1	69	Mean Difference (IV, Fixed, 95% CI)	-0.90 [-9.90, 8.10]
3 Mean creamatocrit of breast milk at 6 weeks	1	69	Mean Difference (IV, Fixed, 95% CI)	0.70 [-0.76, 2.16]
4 Breastfeeding fully	1		Odds Ratio (M-H, Fixed, 95% CI)	Subtotals only
4.1 At 6 weeks	1	65	Odds Ratio (M-H, Fixed, 95% CI)	1.23 [0.46, 3.28]
4.2 At 3 months	1	63	Odds Ratio (M-H, Fixed, 95% CI)	1.24 [0.45, 3.45]
4.3 At 6 months	1	61	Odds Ratio (M-H, Fixed, 95% CI)	0.88 [0.28, 2.74]
5 Breastfeeding, any	1		Odds Ratio (M-H, Fixed, 95% CI)	Subtotals only
5.1 At 6 weeks	1	65	Odds Ratio (M-H, Fixed, 95% CI)	3.08 [0.55, 17.18]



Outcome or subgroup ti- tle	No. of studies	No. of partici- pants	Statistical method	Effect size
5.2 At 3 months	1	63	Odds Ratio (M-H, Fixed, 95% CI)	3.13 [1.00, 9.77]
5.3 At 6 months	1	57	Odds Ratio (M-H, Fixed, 95% CI)	1.42 [0.50, 4.03]

## Analysis 6.1. Comparison 6 Etonogestrel implant: early versus standard insertion, Outcome 1 Lactation failure.

Study or subgroup	Early	Standard			Odds Ratio	•		Weight	Odds Ratio
	n/N	n/N		M-H	I, Fixed, 95	% CI			M-H, Fixed, 95% CI
Gurtcheff 2011	1/34	0/35		_		-		100%	3.18[0.13,80.79]
Total (95% CI)	34	35		_				100%	3.18[0.13,80.79]
Total events: 1 (Early), 0 (Standard)									
Heterogeneity: Not applicable									
Test for overall effect: Z=0.7(P=0.48)									
		Favors early	0.01	0.1	1	10	100	Favors standard	

# Analysis 6.2. Comparison 6 Etonogestrel implant: early versus standard insertion, Outcome 2 Mean time to lactogenesis stage II.

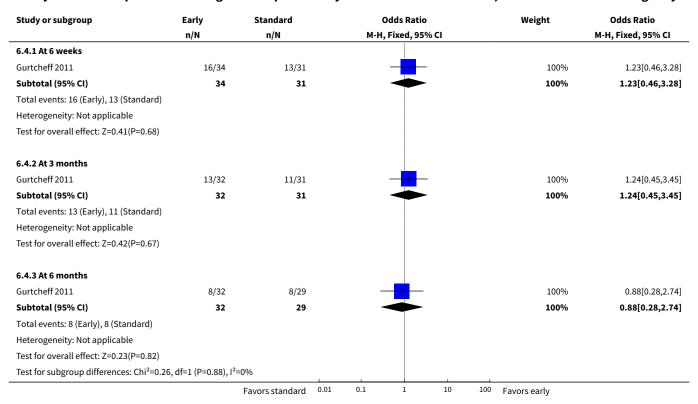
Study or subgroup		Early	St	andard		Mea	n Difference	e		Weight	Mean Difference
	N	Mean(SD)	N	Mean(SD)		Fix	ed, 95% CI				Fixed, 95% CI
Gurtcheff 2011	34	64.3 (19.6)	35	65.2 (18.5)						100%	-0.9[-9.9,8.1]
Total ***	34		35			-				100%	-0.9[-9.9,8.1]
Heterogeneity: Tau <sup>2</sup> =0; Chi <sup>2</sup> =0	, df=0(P<0.0001	.); I <sup>2</sup> =100%									
Test for overall effect: Z=0.2(P	=0.84)										
				Favors early	-20	-10	0	10	20	Favors standar	d

# Analysis 6.3. Comparison 6 Etonogestrel implant: early versus standard insertion, Outcome 3 Mean creamatocrit of breast milk at 6 weeks.

Study or subgroup		Early	St	andard		Mea	n Difference		Weight	Mean Difference
	N	Mean(SD)	N	Mean(SD)		Fix	ed, 95% CI			Fixed, 95% CI
Gurtcheff 2011	34	7.5 (3)	35	6.8 (3.2)					100%	0.7[-0.76,2.16]
Total ***	34		35						100%	0.7[-0.76,2.16]
Heterogeneity: Not applicable										
Test for overall effect: Z=0.94(P=0.3	5)									
			Fa	vors standard	-5	-2.5	0 2.5	5	Favors early	



Analysis 6.4. Comparison 6 Etonogestrel implant: early versus standard insertion, Outcome 4 Breastfeeding fully.



Analysis 6.5. Comparison 6 Etonogestrel implant: early versus standard insertion, Outcome 5 Breastfeeding, any.

Study or subgroup	Early	Standard	Odds Ratio	Weight	Odds Ratio
	n/N	n/N	M-H, Fixed, 95% CI		M-H, Fixed, 95% CI
6.5.1 At 6 weeks					
Gurtcheff 2011	32/34	26/31	<del>-                                     </del>	100%	3.08[0.55,17.18]
Subtotal (95% CI)	34	31		100%	3.08[0.55,17.18]
Total events: 32 (Early), 26 (Standard)					
Heterogeneity: Not applicable					
Test for overall effect: Z=1.28(P=0.2)					
6.5.2 At 3 months					
Gurtcheff 2011	26/32	18/31	<del></del>	100%	3.13[1,9.77]
Subtotal (95% CI)	32	31		100%	3.13[1,9.77]
Total events: 26 (Early), 18 (Standard)					
Heterogeneity: Not applicable					
Test for overall effect: Z=1.96(P=0.05)					
6.5.3 At 6 months					
Gurtcheff 2011	15/28	13/29	— <del></del>	100%	1.42[0.5,4.03]
Subtotal (95% CI)	28	29		100%	1.42[0.5,4.03]
Total events: 15 (Early), 13 (Standard)					
Heterogeneity: Not applicable					
Test for overall effect: Z=0.66(P=0.51)					
Test for subgroup differences: Chi <sup>2</sup> =1.19	, df=1 (P=0.55), I <sup>2</sup> =	:0%			
		Favors standard 0.01	0.1 1 10 1	.00 Favors early	



# Comparison 7. COC (levonorgestrel 150 μg + EE 30 μg) versus POP (norgestrel 75 μg)

Outcome or subgroup ti- tle	No. of studies	No. of partici- pants	Statistical method	Effect size
1 Mean breast milk volume (mL)	1		Mean Difference (IV, Fixed, 95% CI)	Subtotals only
1.1 At 9 weeks	1	161	Mean Difference (IV, Fixed, 95% CI)	-17.80 [-28.80, -6.80]
1.2 At 16 weeks	1	140	Mean Difference (IV, Fixed, 95% CI)	-24.0 [-34.53, -13.47]
1.3 At 24 weeks	1	115	Mean Difference (IV, Fixed, 95% CI)	-24.90 [-36.01, -13.79]
2 Mean change in infant weight (g)			Other data	No numeric data

# Analysis 7.1. Comparison 7 COC (levonorgestrel 150 $\mu g$ + EE 30 $\mu g$ ) versus POP (norgestrel 75 $\mu g$ ), Outcome 1 Mean breast milk volume (mL).

Study or subgroup		coc		POP	Mean Difference	Weight	<b>Mean Difference</b>
	N	Mean(SD)	N	Mean(SD)	Fixed, 95% CI		Fixed, 95% CI
7.1.1 At 9 weeks							
WHO 1984	78	57.6 (31.8)	83	75.4 (39.2)	-	100%	-17.8[-28.8,-6.8]
Subtotal ***	78		83		<b>→</b>	100%	-17.8[-28.8,-6.8]
Heterogeneity: Not applicable							
Test for overall effect: Z=3.17(P=0)							
7.1.2 At 16 weeks							
WHO 1984	67	46.5 (21.3)	73	70.5 (40.2)	-	100%	-24[-34.53,-13.47]
Subtotal ***	67		73		•	100%	-24[-34.53,-13.47]
Heterogeneity: Not applicable							
Test for overall effect: Z=4.47(P<0.0	0001)						
7.1.3 At 24 weeks							
WHO 1984	57	40.5 (22.7)	58	65.4 (36.6)	_ <b></b>	100%	-24.9[-36.01,-13.79]
Subtotal ***	57		58		•	100%	-24.9[-36.01,-13.79]
Heterogeneity: Not applicable							
Test for overall effect: Z=4.39(P<0.0	0001)						
Test for subgroup differences: Chi <sup>2</sup> =	=0.95, df=1	(P=0.62), I <sup>2</sup> =0%					
				Favors POP	-50 -25 0 25	50 Favors COC	

# Analysis 7.2. Comparison 7 COC (levonorgestrel 150 $\mu$ g + EE 30 $\mu$ g) versus POP (norgestrel 75 $\mu$ g), Outcome 2 Mean change in infant weight (g).

Mean change	in	infant	weight (g)
mean change	••••	munc	weight (g)

:	Study Visit interval (postpartum)	сос	POP
WHO 1984		Mean <u>+</u> standard error (SE)	Mean <u>+</u> SE
WHO 1984	6 to 9 weeks	561.9 <u>+</u> 24.4	623.5 <u>+</u> 23.5

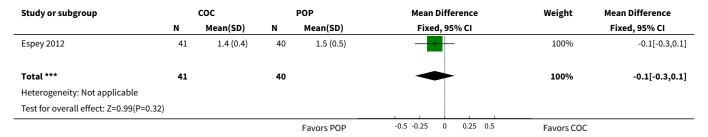


	Mean cha	nge in infant weight (g)		
Study	Visit interval (postpartum)	сос	POP	
WHO 1984	9 to 12 weeks	523.0 <u>+</u> 24.3	531.9 <u>+</u> 23.2	
WHO 1984	20 to 24 weeks	354.6 <u>+</u> 19.9	395.7 <u>+</u> 19.8	

# Comparison 8. COC (norethinodrone 1 mg + EE 35 μg) versus POP (norethinodrone 350 μg)

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
1 Mean change in infant weight (kg) from week 2 to 8	1	81	Mean Difference (IV, Fixed, 95% CI)	-0.10 [-0.30, 0.10]
2 Mean change in infant length (cm) from week 2 to week 8	1	81	Mean Difference (IV, Fixed, 95% CI)	0.27 [-0.48, 1.02]
3 Breastfeeding discontinuation by 6 months			Other data	No numeric data

# Analysis 8.1. Comparison 8 COC (norethinodrone 1 mg + EE 35 $\mu$ g) versus POP (norethinodrone 350 $\mu$ g), Outcome 1 Mean change in infant weight (kg) from week 2 to 8.



# Analysis 8.2. Comparison 8 COC (norethinodrone 1 mg + EE 35 $\mu$ g) versus POP (norethinodrone 350 $\mu$ g), Outcome 2 Mean change in infant length (cm) from week 2 to week 8.

Study or subgroup		coc		POP	Mean Difference	Weight	Mean Difference
	N	Mean(SD)	N	Mean(SD)	Fixed, 95% CI		Fixed, 95% CI
Espey 2012	41	5.6 (1.8)	40	5.3 (1.7)	<del>-   -</del>	100%	0.27[-0.48,1.02]
Total ***	41		40			100%	0.27[-0.48,1.02]
Heterogeneity: Not applicable							
Test for overall effect: Z=0.7(P=0.48)				_		_	
				Favors POP	-1 -0.5 0 0.5 1	Favors COC	



# Analysis 8.3. Comparison 8 COC (norethinodrone 1 mg + EE 35 $\mu$ g) versus POP (norethinodrone 350 $\mu$ g), Outcome 3 Breastfeeding discontinuation by 6 months.

## Breastfeeding discontinuation by 6 months

Study	Variable	Reported hazard ratio	Reported 95% CI	Reported P value
Espey 2012	Group: COC versus POP	1.42	0.76 to 2.65	.270
Espey 2012	Supplementing: yes versus no	2.81	1.09 to 7.23	.033
Espey 2012	Milk concerns: yes versus no	2.07	1.37 to 5.91	.005

## **ADDITIONAL TABLES**

# Table 1. Evidence quality

Study	Comparison groups	Inade- quate random- ization and allo- cation con- cealment	No blind- ing	Follow-up: < 8 weeks for infant growth; < 12 weeks for lactation	Loss to fol- low-up > 20%	Quality of evidence <sup>a</sup>
Progestin-onl	y pills (POPs) versus placebo					
Giner Ve- lazquez 1976	<ul><li>POP norethindrone</li><li>Placebo</li></ul>	-1		-1	No infor- mation	Low
Dutta 2013	<ul><li>POP desogestrel</li><li>Placebo</li></ul>	-1	No infor- mation		No infor- mation	Low
Levonorgestr	el-releasing intrauterine system (LNG-IUS)					
Heikkilä 1982	<ul><li>LNG-IUS (30 μg)</li><li>Nova T IUD</li></ul>	-1				Moderate
Shaamash 2005	<ul> <li>LNG-IUS (20 μg)</li> <li>CuT 380A IUD</li> </ul>		No infor- mation		No infor- mation	Moderate
Stuart 2014	<ul> <li>LNG-IUS (20 μg), insertion times</li> <li>Within 48 hours</li> <li>4 to 8 weeks</li> </ul>		-1			Moderate
Etonogestrel-	releasing (ETG) implant, early versus stand	ard insertion				
Brito 2009	<ul> <li>ETG implant, early insertion (1 to 2 days)</li> <li>No method 6 weeks, then DMPA</li> </ul>		-1			Moderate
Gurtcheff 2011	<ul> <li>ETG implant, postpartum insertion time</li> <li>Early (1 to 3 days)</li> <li>Standard (4 to 8 weeks)</li> </ul>		-1			Moderate



Table 1. Evid	dence quality (Continued)					
Semm 1966	<ul><li>COC lynestrenol + mestranol</li><li>Placebo</li></ul>	-1	No infor- mation	-1	No infor- mation	Very low
Miller 1970	<ul><li>COC norethindrone + mestranol</li><li>Placebo</li></ul>	-1				Moderate
Combined or	al contraceptive versus progestin-only p	oill				
WHO 1984	<ul><li>COC levonorgestrel + EE</li><li>POP norgestrel</li></ul>				-1	Moderate
Espey 2012	<ul><li>COC norethindrone + EE</li><li>POP norethindrone</li></ul>				-1	Moderate
Overall quali	ty of evidence					Moderate

aGrade levels were high, moderate, low, or very low. RCTs were downgraded by one level for each of the following: (1) no information on randomization sequence generation or allocation concealment, or one was clearly inadequate; (2) no blinding; (3) follow-up < 8 weeks for infant growth or < 12 weeks for lactation; (4) loss to follow-up > 20%; (5) information missing for both blinding and losses.

41	44-
Library	Cochrane

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Table 2.	Outcome	summary
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Study	Comparison groups <sup>a</sup>	N	Postpartum initiation	Lactation <sup>b</sup>	Breast milk <sup>c</sup>	Infant growth <sup>d</sup>	Pregnancy
Progestin-only	y pill (POP) versus placebo						
Giner Ve- lazquez 1976	<ul><li>POP norethindrone</li><li>Placebo</li></ul>	20	48 hours		NS <sup>e</sup> ; not quantified	NS <sup>e</sup> ; not quantified	
Dutta 2013	<ul><li>POP desogestrel</li><li>Placebo</li></ul>	400	6 weeks		NSe	NSe	POP < place- bo <sup>e</sup>
Levonorgestre	el-releasing intrauterine system (LNG-IUS)						
Heikkilä 1982	<ul><li>LNG-IUS (30 μg)</li><li>Nova T IUD</li></ul>	80	6 weeks	BFe: LNG-IUS < Nova T IUD		NS <sup>e</sup> ; not quantified	None
Shaamash 2005	<ul><li>LNG-IUS (20 μg)</li><li>CuT 380A IUD</li></ul>	320	6 to 8 weeks	NSe		NSe	None
Stuart 2014	LNG-IUS (20 µg), insertion times  • ≤ 48 hours  • 4 to 8 weeks	35	48 hours vs 4 to 8 weeks	NS			
Etonogestrel-r	releasing (ETG) implant: early versus standard in	sertion					
Brito 2009	<ul> <li>ETG implant, early insertion (24 to 48 hours)</li> <li>No method 6 weeks, then DMPA</li> </ul>	40	24 to 48 hours vs 6 weeks	NS		Weight: 6 weeks, implant > no method; 12 weeks, DMPA > implant	
Gurtcheff 2011	<ul><li>ETG implant insertion</li><li>Early (1 to 3 days)</li><li>Standard (4 to 8 weeks)</li></ul>	69	1 to 3 days vs 4 to 8 weeks	NS	NS		
Combined ora	l contraceptive (COC) versus placebo						
Semm 1966	<ul><li>COC lynestrenol + mestranol</li><li>Placebo</li></ul>	100	1 day	NS <sup>e</sup> ; not quantified	NS <sup>e</sup> ; not quantified		

Table 2. Out	come summary (Continued)						
Miller 1970	<ul><li>COC norethindrone + mestranol</li><li>Placebo</li></ul>	50	14 days	BFe: COC < placebo; un- clear, not quantified	Volume <sup>e</sup> : COC < placebo; unclear, not quantified,		
Combined ve	rsus progestin-only oral contraceptive						
WHO 1984	COC levonorgestrel + EE	171	6 weeks		Volume:	NSe	
	POP norgestrel				COC < POP		
Espey 2012	COC norethindrone + EE	197	2 weeks	NS		NS	None

 $^{o}$ Table shows direction of significant differences (reported or analyzed); NS = not significantly different.

bInitiation, failure, breastfeeding (BF) fully or any.

<sup>c</sup>Time to lactogenesis II, volume, composition.

dWeight or length.

<sup>e</sup>Reported; insufficient data for analysis in this review.

• POP norethindrone



Table 3.	Sensitivity	/ analv	/sis
----------	-------------	---------	------

Study <sup>a</sup>	Comparison groups	N	Postpar- tum initia- tion	Lactation <sup>b</sup>	Breast milk <sup>c</sup>	Infant growth
Levonorgest	rel-releasing intrauterine syste	m (LNG-IUS)	)			
Shaamash 2005	<ul><li>LNG-IUS (20 μg)</li><li>CuT 380A IUD</li></ul>	320	6 to 8 weeks	NS <sup>d</sup> : full BF (Analysis 2.3)		NSd: weight (Analysis 2.1) and length Analysis 2.2)
Stuart 2014	LNG-IUS (20 µg) insertion times  • ≤ 48 hours  • 4 to 8 weeks	35	48 hours vs 4 to 8 weeks	NS: any BF (Analysis 3.1)		
Etonogestre	l-releasing (ETG) implant: early	versus stan	dard insertion			
Brito 2009	<ul> <li>ETG implant, early insertion (1 to 2 days)</li> <li>No method for 6 weeks, then DMPA</li> </ul>	40	24 to 48 hours vs 6 weeks	NS: full BF (Analy- sis 4.1)		Weight: 6 weeks, ETG implant > no method (Analysis 4.2); 12 weeks, DMPA > ETG implant (Analysis 5.2)
Gurtcheff 2011	ETG implant, postpartum insertion  Early (1 to 3 days)  Standard (4 to 8 weeks)	65	1 to 3 days vs 4 to 8 weeks	NS: failure (Analysis 6.1); full BF (Analysis 6.4); any BF (Analysis 6.5)	NS: lactoge- nesis II (Analysis 6.2); cream- atocrit (Analysis 6.3)	
Combined (C	COC) versus progestin-only (POI	P) oral contr	aceptive			
WHO 1984	<ul><li>COC levonorgestrel + EE</li><li>POP norgestrel</li></ul>	171	6 weeks		Volume: COC < POP (Analysis 7.1)	NS <sup>d</sup> : weight (Analysis 8.3)
Espey 2012	<ul><li>COC norethindrone + EE</li><li>POP norethindrone</li></ul>	179	2 weeks	NS: BF (Analysis 8.3)		NS: weight (Analysis 8.1) and length (Analysis 8.2)

<sup>&</sup>lt;sup>a</sup>Includes studies with moderate-quality evidence (Table 1) and sufficient outcome data (Table 2). Direction of significant differences shown (reported or analyzed); NS = not significantly different.

bInitiation; failure; breastfeeding (BF) full, any, or discontinuation.

<sup>&</sup>lt;sup>c</sup>Time to lactogenesis II; milk volume or composition.

dReported; insufficient data for analysis in this review.



#### APPENDICES

### Appendix 1. Search 2015

#### PubMed (1 September 2001 to 6 January 2015)

((breastfeeding OR lactation) AND (contraceptive devices, female OR contraceptive agents, female)) NOT (cancer OR cows) Article types: Clinical Trial

## **CENTRAL (2015, Issue 2 (on 2 March 2015))**

Title, Abstract, Keywords: lactat\* or breastfeed\* OR breast-feed\* AND Title, Abstract, Keywords: contracept\* Publication Date from 2001 to 2015 Limited to Trials

## **POPLINE (2001 to 8 July 2014)**

### Two strategies:

Keyword: Contraception AND Keyword: Lactation

Filter by Keyword: Research report

Years: 2001 to 2014

Keyword: Contraception AND Keyword: Breastfeeding AND Keyword: Clinical research

Years: 2001 to 2014

### Web of Science (11 December 2014)

TOPIC: (contracept\*)

AND TOPIC: (lactat\*)

Refined by: WEB OF SCIENCE CATEGORIES: (OBSTETRICS GYNECOLOGY OR PUBLIC ENVIRONMENTAL OCCUPATIONAL HEALTH OR PUBLIC ENVIRONMENTAL OCCUPATIONAL HEALTH SSCI)

AND DOCUMENT TYPES: (ARTICLE OR MEETING ABSTRACT OR REVIEW)

Timespan: 2001-2014

#### LILACS (4 March 2015)

contraceptive agents, female or agentes anticonceptivos femeninos or anticoncepcionais femeninos [Words] and ("LACTATION") or "BREASTFEEDING" [Words]

## ClinicalTrials.gov (14 January 2015)

Condition: breastfeeding OR lactation Intervention: contraception OR contraceptive Study type: Interventional studies First received: from 01 June 2010

#### ICTRP (14 January 2015)

Title: breast-feed\* OR breastfeed\* OR lactat\* Condition: contraception OR contraceptive

Recruitment status: All

Date of registration: from 01 June 2010 to 14 January 2015



#### **Appendix 2. Prior searches**

#### 2010 update

#### MEDLINE (via PubMed) (to 29 October 2014)

((breastfeeding OR lactation) AND (contraceptive devices, female OR contraceptive agents, female)) AND (randomized controlled trial [pt] OR controlled clinical trial [pt] OR random allocation [mh] OR double-blind method [mh] OR single-blind method [mh] OR clinical trial [pt] OR ("clinical trial" [tw]) OR ((singl\* [tw] OR doubl\* [tw] OR trebl\* [tw] OR tripl\* [tw]) AND (mask\* [tw] OR blind\* [tw])) OR ("latin square" [tw]) OR placebos [mh] OR placebo\* [tw] OR random\* [tw] OR research design [mh:noexp] OR follow-up studies [mh] OR prospective studies [mh] OR control\* [tw] OR prospective [tw]) OR volunteer\* [tw]) NOT (cancer OR cows)

#### POPLINE (2006 to 1 November 2010)

(lactating / lactation / breastfeed / breastfeeding / breastfeeding / breastfeeders / breastfield) & contraception & clinical trials

#### CENTRAL (2008 to 29 October 2010)

lactat\* or breastfeed\* in Title, Abstract, or Keywords AND contracept\* in Title, Abstract, or Keywords

#### ClinicalTrials.gov (to 1 November 2010)

(lactat\* or breastfeed\*) AND contracept\*

#### ICTRP (to 1 November 2010)

Condition: lactating OR lactation OR breastfeed OR breastfeeding

Intervention: contraception OR contraceptive

#### Initial review and updates in 2005 and 2008

Strategies for MEDLINE, POPLINE, and CENTRAL were those listed above for 2010. In addition, EMBASE and LILACS were searched as shown below. The 2005 searches were run to 5 April 2005, and the 2008 searches were run to 6 May 2008.

#### **EMBASE**

- 1. contracep?
- 2. contraception!
- 3.1 or 2
- 4. breast(W)milk OR breastmilk
- 5. breastfeed? OR breast(W)feed?
- 6. lactation
- 7.4 OR 5 OR 6
- 8. clinical trial!
- 9. controlled study!
- 10.8 OR 9
- 11. 3 AND 7 AND 10
- 12. 11/human

### **LILACS**

lactation, breastfeeding, contraception/contraceptives, clinical trials

#### Appendix 3. Non-randomized studies (previously listed as excluded)

Full-text review was not needed to determine that these studies were not eligible. In 2014, we classified them as 'discarded' to shorten the list of Excluded studies.

- 1. Bjarnadottir RI, Gottfredsdottir H, Sigurdardottir K, Geirsson RT, Dieben TO. Comparative study of the effects of a progestogen-only pill containing desogestrel and an intrauterine contraceptive device in lactating women. British Journal of Obstetrics and Gynaecology 2001;108:1174-80.
- 2. Chen JH, Wu SC, Shao WQ, Zou MH, Hu J, Cong L, et al. The comparative trial of TCu 380A IUD and progesterone-releasing vaginal ring used by lactating women. Contraception 1998;57:371-9.
- 3. Coutinho EM, Athayde C, Dantas C, Hirsch C, Barbosa I. Use of a single implant of elcometrine (ST-1435), a nonorally active progestin, as a long acting contraceptive for postpartum nursing women. Contraception 1999;59:115-22.



- 4. Danli S, Qingxiang S, Guowei S. A multicentered clinical trial of the long-acting injectable contraceptive Depo Provera in Chinese women. Contraception 2000;62:15-8.
- 5. Dunson TR, McLaurin VL, Grubb GS, Rosman AW. A multicenter clinical trial of a progestin-only oral contraceptive in lactating women. Contraception 1993;47:23-35.
- 6. Hannon PR, Duggan AK, Serwint JR, Vogelhut JW, Witter F, DeAngelis C. The influence of medroxyprogesterone on the duration of breast-feeding in mothers in an urban community. Archives of Pediatrics & Adolescent Medicine 1997;151:490-6.
- 7. Lönnerdal B, Forsum E, Hambraeus L. Effect of oral contraceptives on composition and volume of breast milk. American Journal of Clinical Nutriton 1980;33:816-24.
- 8. Massai MR, Diaz S, Quinteros E, Reyes MV, Herreros C, Zepeda A, et al. Contraceptive efficacy and clinical performance of Nestorone implants in postpartum women. Contraception 2001;64:369-76.
- 9. McCann MF, Moggia AV, Higgins JE, Potts M, Becker C. The effects of a progestin-only oral contraceptive (levonorgestrel 0.03 mg) on breast-feeding. Contraception 1989;40:635-48.
- 10. Reinprayoon D, Taneepanichskul S, Bunyavejchevin S, Thaithumyanon P, Punnahitananda S, Tosukhowong P, et al. Effects of the etonogestrel-releasing contraceptive implant (Implanon) on parameters of breastfeeding compared to those of an intrauterine device. Contraception 2000;62:239-46.
- 11. Virutamasen P, Leepipatpaiboon S, Kriengsinyot R, Vichaidith P, Muia PN, Sekadde-Kigondu CB, et al. Pharmacodynamic effects of depot-medroxyprogesterone acetate (DMPA) administered to lactating women and their male infants. Contraception 1996;54:153-7.
- 12. World Health Organization Task Force for Epidemiological Research on Reproductive Health; Special Programme of Research, Development, and Research Training in Human Reproduction. Progestogen-only contraceptives during lactation: I. infant growth. Contraception 1994;50:35-53.
- 13. Zacharias S, Aguilera E, Assenzo JR, Zanartu J. Effects of hormonal and nonhormonal contraceptives on lactation and incidence of pregnancy. Contraception 1986;33:203-12.
- 14. Zanartu J, Aguilera E, Munoz-Pinto G. Maintenance of lactation by means of continuous low-dose progestogen given post-partum as a contraceptive. Contraception 1976;13:313-8.

#### **FEEDBACK**

#### Combined hormonal versus nonhormonal versus progestin-only contraception

# Summary

The exclusion of the study by Diaz et al (Contraception 1983; 27: 1-11) is not justified. In the review it is stated: "Per correspondence with the author, study participants received personal choice of intervention, not random allocation". In the methods division of this study the authors wrote: "Women requesting an oral contraceptive were assigned at random to the contraceptive pill under study or to an oral placebo on a patient-blind basis. Both pills were offered as a low-dose O.C. with no demonstrated effects upon lactation". I wrote to dr Diaz and asked her why the Cochrane review could come to a different conclusion. She answered me that the person doing the Cochrane review asked her only in general about all the studies on breastfeeding and contraception. In all the other trials women had the free choice of contraceptive method. Because the reviewer did not ask in detail about every trial she forgot to mention that in one trial the treatment was randomized.

In the discussion of the review is stated: "No trial to date has documented an adverse effect of hormonal contraceptives on infant growth". This statement is wrong because in the trial by Diaz et al (1983) the oral contraceptive group showed a significantly lower average absolute weight at days 61 and 91 postpartum and a significantly lower average daily weight increase during the first month of treatment.

In the discussion of the Cochrane review, the review authors criticized several trials because the putative effects of hormonal contraceptives on lactation could have been masked by the influence of supplemental foods. The review authors apparently did not realize that it is impossible to do a month long trial on breastfeeding and forbid the mothers to give supplements to their babies if they consider that the babies get insufficient food.

In the implications for research, the authors of the review suggest that investigators should do a trial of contraceptives versus placebo: "such a trial might enroll women who are not at risk of pregnancy because of a sterilization procedure (i.e. postpartum sterilization or partner vasectomy)."This proposal would imply that researchers would have to ask mothers who were planning to breastfeed their infants during several months to take a pill that could potentially deteriorate their lactation, while this medication could not offer her or her baby any advantage. Is there any ethical committee that would approve such a proposal?



I certify that I have no affiliations with or involvement in any organization or entity with a direct financial interest in the subject matter of my criticisms.

February, 2004.

#### Reply

#### Response to Treffers re:

Dr. Treffers suggests that we incorrectly excluded the report by Diaz et al. (Contraception 1983;27:1-11) and thus reached the wrong conclusion about the effect of combined oral contraceptives on infant growth. We stand by our exclusion for several reasons. First, we contacted Dr. Diaz and were advised by her in writing on April 1, 2002, that randomization had not been used in her study. Second, the report appears to describe a cohort study, with another group (not randomized) having received an intramuscular placebo. Third, the disparity in sample size between the ostensibly randomized groups (oral contraceptive, 103 participants; oral placebo, 79 participants) is unlikely to have resulted from simple randomization. By binomial theorem, the likelihood of getting a difference this large or larger due to chance is 4%. Stated alternatively, one can be 96% sure that randomization did not yield this result. Hence, we conclude that the Diaz 1983 study was not a randomized controlled trial, and our interpretation of the literature stands.

Given the absence of any demonstrable adverse effect of oral contraceptives on infant growth, the age and limited quality of existing studies, and the public health importance of the question, we believe a proper trial is both appropriate and ethical.

#### **Contributors**

Treffers, Pieter

### WHAT'S NEW

Date	Event	Description
2 March 2015	New citation required but conclusions have not changed	Searches updated
4 August 2014	Amended	Summaries of evidence quality (Table 1) and outcomes (Table 2) and a sensitivity analysis (Table 3) added
31 July 2014	New search has been performed	7 trials (Heikkilä 1982; Shaamash 2005; Brito 2009; Gurtcheff 2011; Espey 2012; Dutta 2013; Stuart 2014) and 1 ongoing trial (Turok 2014) added
14 July 2014	Amended	Types of outcome measures separated into primary (effect on lactation) and secondary (contraceptive efficacy). Trials must report a primary outcome; Were 1997 determined to be no longer eligible (Excluded studies)

#### HISTORY

Protocol first published: Issue 1, 2003 Review first published: Issue 2, 2003

Date	Event	Description
17 December 2012	New search has been performed	Study added to 'awaiting classification' (Espey 2012).
2 November 2010	New search has been performed	Searches were updated for MEDLINE, CENTRAL, and POPLINE. New searches were added for ClinicalTrials.gov and ICTRP. No new trials met the inclusion criteria.
7 May 2008	New search has been performed	Searches were updated; no new trials were found.



Date	Event	Description
3 October 2005	New citation required and conclusions have changed	Substantive amendment

#### **CONTRIBUTIONS OF AUTHORS**

#### 2014 revision

LM Lopez conducted the searches and primary data extraction, incorporated the six trials, and revised the review. TW Grey conducted the secondary data extraction for study characteristics and outcomes. M Chen contributed to the methodology and to statistical interpretation, and examined the quality assessment data. AM Stuebe contributed research expertise in lactation to the background and discussion sections. ST Truitt contributed clinical expertise in lactation to the interpretation of results. MF Gallo commented on the overall presentation and interpretation of results.

#### 2005, 2008, and 2010 updates

LM Lopez wrote the plain language summary, reviewed the search results, and edited the text for Cochrane style.

#### 2003 initial review

Sarah T Truitt conceived and designed the review, performed initial database searches and retrieval of papers, wrote the protocol, and wrote to study authors for additional information.

Anna B Fraser (formerly of St John's Clinic, Springfield, Montana) screened search results and retrieved papers, edited the protocol, and wrote the review.

Maria F Gallo designed the database searches, screened search results and retrieved papers, and edited the protocol and the review. David A Grimes (formerly of FHI 360) supervised the review, provided advice, edited the review, screened retrieved papers, and wrote to study authors for additional information.

Kenneth F Schulz (from FHI 360) provided advice, interpreted data, and contributed statistical expertise.

#### **DECLARATIONS OF INTEREST**

Lopez LM, Grey TW, Stuebe AM, Chen M, Truitt ST, Gallo MF: none.

# SOURCES OF SUPPORT

### **Internal sources**

· No sources of support supplied

#### **External sources**

• National Institute of Child Health and Human Development, USA.

2002 through 2014: Support for conducting the review and updates at FHI 360

• US Agency for International Development, USA.

2002 through 2010: Support for conducting the review and updates at FHI 360

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### INDEX TERMS

### **Medical Subject Headings (MeSH)**

Breast Feeding [\*statistics & numerical data]; Child Development [drug effects]; Contraception [\*methods]; Contraceptives, Oral, Combined [\*pharmacology]; Contraceptives, Oral, Hormonal [pharmacology]; Desogestrel [pharmacology]; Lactation [\*drug effects]; Levonorgestrel [pharmacology]; Milk, Human [\*drug effects]; Progestins [\*pharmacology]; Randomized Controlled Trials as Topic

## MeSH check words

Female; Humans; Infant