

BMJ Open Where are the inequalities in colorectal cancer care in a country with universal healthcare? A systematic review and narrative synthesis

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ABSTRACT

Objective Patients diagnosed with colorectal cancer living in more deprived areas experience worse survival than those in more affluent areas. Those living in more deprived areas face barriers to accessing timely, quality healthcare. These barriers may contribute to socioeconomic inequalities in survival. We evaluated the literature for any association between socioeconomic group, hospital delay and treatments received among patients with colorectal cancer in the UK, a country with universal healthcare.

Design MEDLINE, EMBASE, CINAHL, CENTRAL, SCIE, AMED and PsycINFO were searched from inception to January 2023. Grey literature, including HMIC, BASE and Google Advanced Search, and forward and backward citation searches were conducted. Two reviewers independently reviewed titles, abstracts and full-text articles. Observational UK-based studies were included if they reported socioeconomic measures and an association with either hospital delay or treatments received. The QUIPS tool assessed bias risk, and a narrative synthesis was conducted. The review is reported to Preferred Reporting Items for Systematic Reviews and Meta-Analyses 2020.

Results 41 of the 7209 identified references were included. 12 studies evaluated 7 different hospital intervals. There was a significant association between area-level deprivation and a longer time from first presentation in primary care to diagnosis. 32 studies evaluated treatments received. There were socioeconomic inequalities in surgery and chemotherapy but not radiotherapy.

Conclusion Patients with colorectal cancer face inequalities across the cancer care continuum. Further research is needed to understand why and what evidence-based actions can reduce these inequalities in treatment. Qualitative research of patients and clinicians conducted across various settings would provide a rich understanding of the complex factors that drive these inequalities. Further research should also consider using a causal approach to future studies to considerably strengthen the interpretation. Clinicians can try and mitigate some potential causes of colorectal cancer inequalities, including signposting to financial advice and patient transport schemes.

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STRENGTHS AND LIMITATIONS OF THIS STUDY

- ⇒ The searches were extensive—conducted across eight databases, supplemented with citation searching and hand-searching websites.
- ⇒ The search strategy was validated.
- ⇒ The inclusion of non-peer-reviewed literature was a key strength.
- ⇒ Due to heterogeneous methods, meta-analysis was not possible.

INTRODUCTION

Colorectal cancer is the second most common cause of cancer-related death in the UK.¹ Survival has improved since the 1990s but lags behind comparable countries.² There are also survival gradients within countries, including those with universal healthcare, such as the UK and Australia.³ In particular, patients living in more deprived areas experience significantly worse survival outcomes.^{1,3} Healthcare systems can contribute to these inequalities, as treatment differences likely compound differential outcomes across populations.²

Timely diagnosis and treatment are also essential, with delays associated with worse outcomes. The Aarhus statement suggested a framework for measuring these delays, categorising the patient journey into patient, doctor and system intervals.⁴ Specifically, the system interval was defined as the period from primary care-initiated investigations or referral to the commencement of treatment.⁴ Socioeconomic circumstances can impact this interval and yet is comparatively under-researched.

Existing inequalities have been exacerbated by the COVID-19 pandemic, with vulnerable patient groups disproportionately affected by suboptimal care.⁵ The evolution of precision medicine and the development of new

technologies and surgical approaches will likely worsen existing inequalities, a process described as the ‘inverse equity law’.⁶ Worryingly, disparities in access to precision oncology are already well documented.⁷ Understanding where inequalities are in the pathways of care for patients with colorectal cancer is essential to inform policy and identify areas of further research to target evidence-based action.

We evaluated the literature for any association between socioeconomic group, system interval and treatment among patients with colorectal cancer in the UK. By focusing exclusively on studies conducted within a single country with a universal healthcare system, our systematic review homogenised the healthcare infrastructure, policy and patient population, ensuring a more interpretable analysis of disparities in cancer care with greater scope for policy impact.

METHODS

This systematic review was registered with PROSPERO (CRD42022347652). The review is reported according to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses 2020 statement (online supplemental appendix S1).⁸

Patient and public involvement

This study was discussed with Involve Hull, a patient and public involvement group affiliated with the author’s institution. The review was considered necessary by all members of the group.

Eligibility criteria

Published and grey literature observational studies were considered for inclusion if relevant outcomes of patients with a primary diagnosis of colorectal cancer (International Classification of Diseases Tenth Revision C18-C20) in the UK were reported.

Outcomes were only included if they had been analysed by a measure of socioeconomic status (eg, an area-based measure such as the Index of Multiple Deprivation or individual measures such as occupation). The relevant outcomes were defined as follows:

- ▶ The association between socioeconomic status and the length of the system interval, as defined by the Aarhus statement.⁴ Any part of the system interval could have been measured.
- ▶ Receipt of cancer-directed treatment (defined as receipt of surgery, chemotherapy or radiotherapy). Studies evaluating palliative or supportive care only were excluded.

Information sources

The following bibliographic databases were searched from inception to 26 January 2023: MEDLINE, EMBASE, AMED, PsycINFO, CINAHL, CENTRAL and Science Citation Index Expanded.

The grey literature was searched using HMIC, BASE, NICE Evidence Search and Google Advanced Search on 26 January 2023. In addition, 12 websites were systematically hand-searched, and backward and forward citation searches were conducted on 30 March 2023 (details in online supplemental appendix S2).

Search strategy

The search strategies are listed in online supplemental appendix S3. The search strategy was developed and validated in conjunction with SG, an information specialist (details in online supplemental appendix S4). BAP-S and another reviewer (MHS or KS) independently screened all titles and abstracts against the predetermined eligibility criteria. The full texts of eligible titles and abstracts were obtained and independently screened for inclusion. Conflicts were resolved by consensus.

Data collection process

One researcher (BAP-S) extracted information from the included studies, collating the relevant data onto a data extraction form. A second author (KS) checked the extracted data, and discrepancies were reconciled by consensus. The data items and effect measures that were sought for extraction are detailed in online supplemental appendix S5.

Study risk of bias assessment

Two researchers (BAP-S and KS) independently evaluated the study risk of bias against domains adapted from the Quality in Prognosis Studies (QUIPS) tool.⁹ Each domain was judged to have a high, moderate or low risk of bias, with the evaluations collated onto a pre-prepared form (online supplemental appendix S6).

Risk of bias assessments informed the narrative synthesis, with greater weight given to studies with a lower risk of bias. A study’s evidence was considered ‘strong’ if there were no high risk of bias categories, ‘moderate’ if there was a high risk of bias in one category and ‘weak’ if there were two or more categories at high risk of bias. However, studies were not excluded based on this.

Synthesis methods

A narrative synthesis was conducted, according to the synthesis without meta-analysis in systematic reviews reporting guideline.¹⁰ An overall assessment of the association between socioeconomic status and each outcome was made, considering the consistency and strength of supporting evidence from each study. Coefficients were extracted based on multi-variable models. Given the inherent methodological heterogeneity, diverse patient populations, varying measures of deprivation and significant statistical heterogeneity observed across the included studies, a meta-analysis was deemed inappropriate as it could yield misleading or oversimplified results. While a meta-analysis was not conducted, forest plots were generated to visually illustrate the observed outcomes in individual studies.

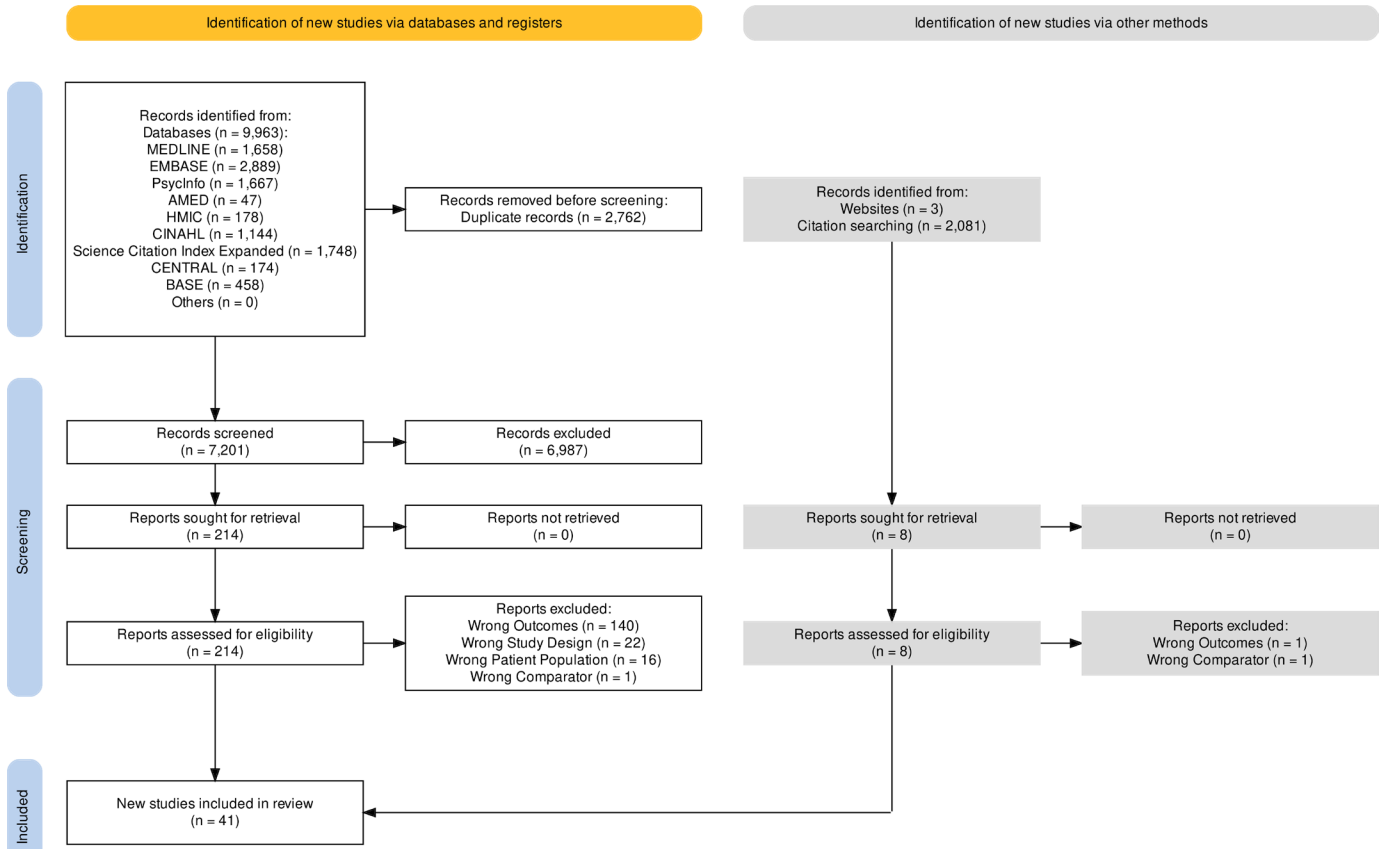


Figure 1 Preferred Reporting Items for Systematic Reviews and Meta-Analyses flow diagram of included studies.

RESULTS

Study selection

The database searches yielded 7201 studies, 214 of which were retrieved for full-text screening. An additional six studies were identified from the grey literature. Overall, 41 studies were included (figure 1).¹¹

Study characteristics

The characteristics of the included studies are summarised in online supplemental appendix S7. The system interval was examined in 12 studies, with 7 different time points evaluated, summarised in figure 2.^{12–23} In total, 15 studies reported the receipt of surgery,^{19 20 24–36} 7 studies evaluated surgical variation,^{37–43} 14 studies reported the receipt of chemotherapy,^{19 20 24–27 44–51} 7 reported the receipt of radiotherapy^{19 20 25–27 43 52} and 2 reported the receipt of any treatment.^{17 46}

In total, 32 of the 41 studies adjusted or stratified for at least one other factor.^{12–26 32–41 44–49 51} The remaining nine studies provided unadjusted rates.^{27–31 42 43 50 52}

Risk of bias in studies

Assessments of the risk of bias are summarised in figure 3 and online supplemental appendix S6. The domain most at risk of bias was study confounding, with 16 studies at high risk of bias.^{13 27–31 39–43 47–50 52} Although some of these studies conducted adjusted analyses, important factors such as stage were unaccounted for.

Results of studies reporting variations in the system interval Referral to first-seen interval

Three studies evaluated the referral to first-seen interval.^{13 15 18} Two studies estimated the odds of being seen by a specialist within 2 weeks of referral; one demonstrated reduced unadjusted odds (OR 0.80, 95% CI 0.70 to 0.91),¹⁸ while there was no significant association in the other (OR 0.95, 95% CI 0.87 to 1.03) after adjusting for age, stage and site (colon vs rectal) (online supplemental appendix S8).¹⁵

Another study used generalised linear modelling to estimate the association between occupation and the number of days to see a specialist after referral, adjusting for age, marital status and ethnicity.¹³ This study reported no significant association ($p > 0.05$).¹³ Overall, the evidence was inconclusive for an association between deprivation and the referral to first-seen interval (table 1 and online supplemental appendix S8).

First-seen to diagnosis interval

One study estimated the association between occupation and the number of days from the first hospital appointment to communication of diagnosis.¹³ A significant association was demonstrated ($p = 0.028$), but no magnitude or direction of effect was provided. The evidence was, therefore, inconclusive (table 1 and online supplemental appendix S8).

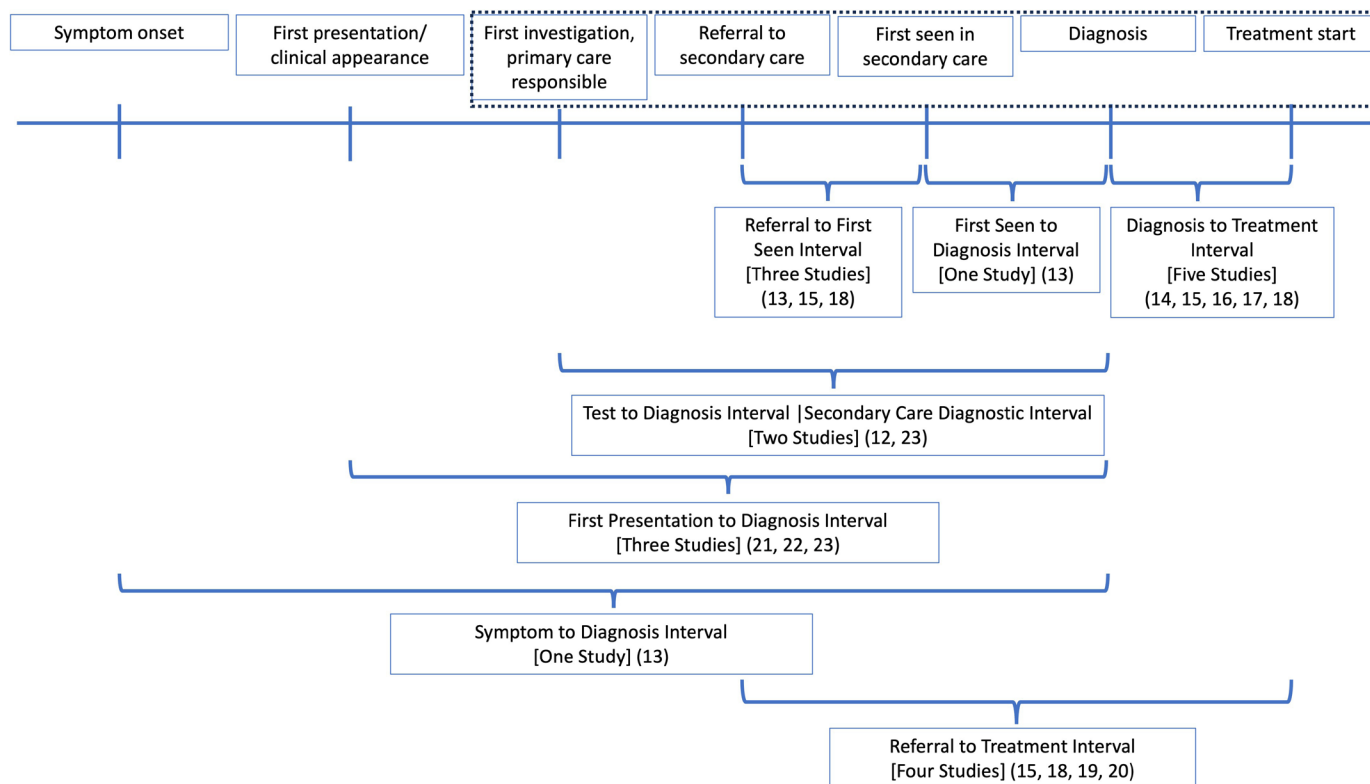


Figure 2 Time intervals evaluated in the included studies. The blue dotted line indicates the system interval defined by the Aarhus statement. Studies that included any aspect of this system interval were included, even if the interval commenced before the system interval defined here.

Diagnosis to treatment interval

Five studies evaluated the diagnosis to treatment interval.^{14–18} Two estimated the number of days from diagnosis to major surgery, adjusting for stage, sex, age, grade and morphology.^{14,16} There was no significant impact of deprivation on the length of the diagnosis to treatment interval demonstrated in these two studies (coefficient 0.99, 95% CI 0.97 to 1.02)¹⁴ (coefficient 0.21, 95% CI –0.55 to 0.98) (online supplemental appendix S8).¹⁶

Two studies evaluated the likelihood of commencing treatment within 31 days from the date a treatment plan was agreed on.^{15,18} One study demonstrated increased unadjusted odds (OR 1.28, 95% CI 1.14 to 1.44),¹⁸ while the other presented reduced adjusted odds of patients from the most deprived areas commencing treatment within 31 days (OR 0.91, 95% CI 0.84 to 0.98) (online supplemental appendix S8).¹⁵

Another study calculated the likelihood of treatment for the most deprived quintile across several time points. They demonstrated reduced adjusted odds of treatment within 1 week (OR 0.78, 95% CI 0.72 to 0.84), 1 month (OR 0.84, 95% CI 0.78 to 0.90) and 2–3 months (OR 0.91, 95% CI 0.85 to 0.98) but non-reduced odds at 4–6 months (OR 1.07, 95% CI 0.96 to 1.18) after the first contact with the health system (online supplemental appendix S8).¹⁷

Overall, the evidence for an association between deprivation and length of the diagnosis to treatment interval was inconclusive (table 1 and online supplemental appendix S8).

Test to diagnosis interval/secondary care diagnostic interval (SCDI)

One study evaluated the SCDI, defined as the period between the date of the first interaction with secondary care and the date of diagnosis.¹² This study evaluated the factors associated with an interval greater than the median, adjusting for sex, age, stage, comorbidities, ethnicity, route to diagnosis and additional diagnostic tests.¹² The odds of a longer interval were not significantly increased for patients from the most deprived quintile (OR 1.07, 95% CI 1.00 to 1.13) (online supplemental appendix S8).

Another study evaluated the time from the first investigation to cancer diagnosis.²³ The authors conducted quantile regression, adjusting for age, comorbidities, sex, test type and symptom category, focusing on the median and 75th centiles.²³ There was no significant association between deprivation and interval length (coefficient 0.7, 95% CI –2.7 to 4.1) (online supplemental appendix S8).

Overall, there was no evidence of a prolonged SCDI or test-to-diagnosis interval for patients from the most deprived background (table 1 and online supplemental appendix S8).

First presentation to diagnosis interval

Three studies evaluated the time from the first symptom or feature of colorectal cancer in primary care records to diagnosis.^{21–23} One study demonstrated an association between deprivation and a longer interval in two of three econometric analyses (pre-to-post

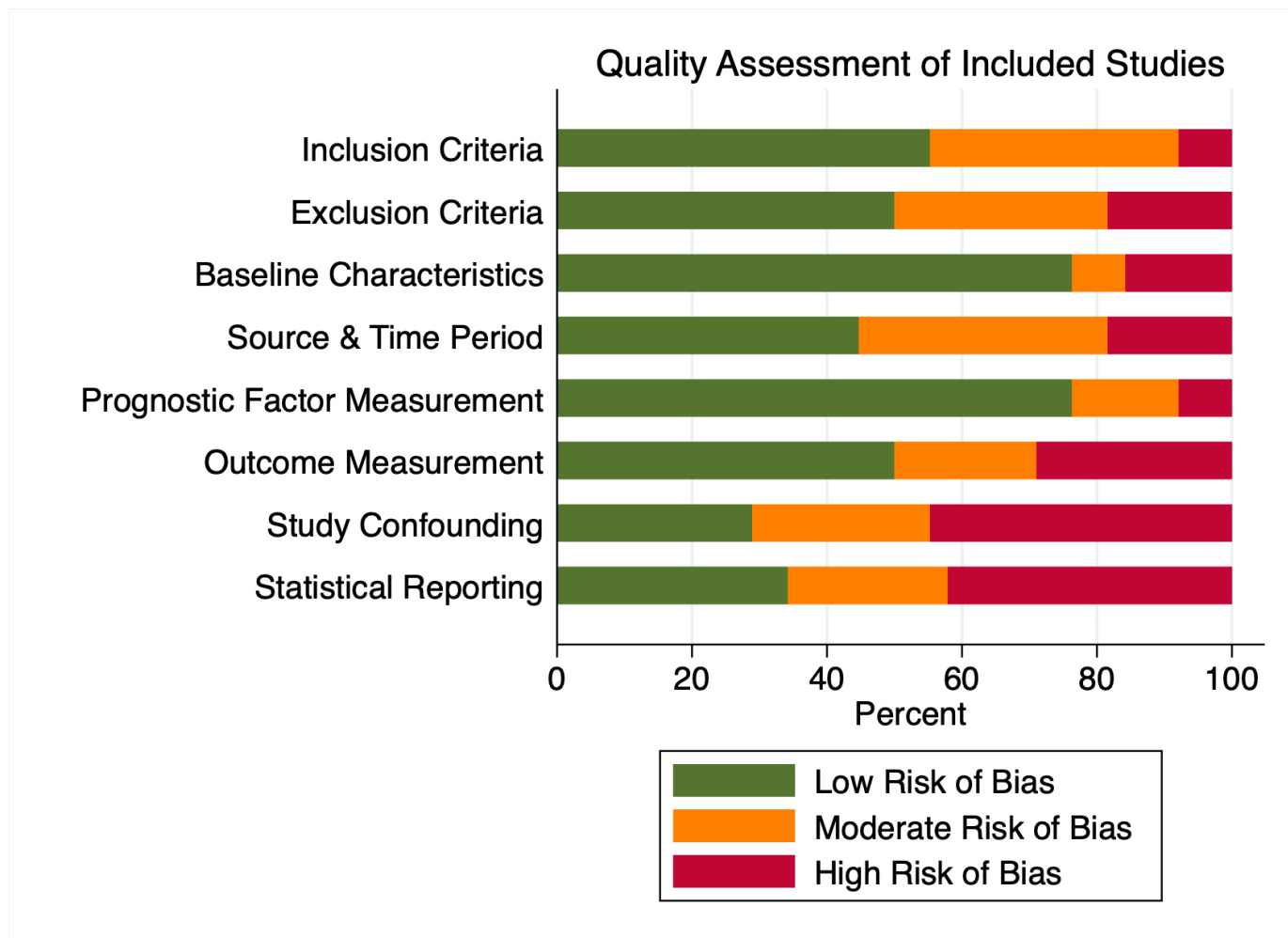


Figure 3 Risk of bias in the included studies. For each element the proportion of studies with high, moderate and low risk of bias is illustrated.

difference-in-differences 95% CI -0.03 to 0.2 and $p=0.147$ or event-study difference-in-differences 95% CI 0.002 to 0.136 and $p=0.043$ or semiparametric varying-coefficient analysis significance stated but not reported).²¹ The other two studies conducted quantile regression, focusing on the median and 75th centiles, adjusting for age, comorbidities, sex and type of symptom.^{22,23} Both studies demonstrated an association between the most deprived quintile and a longer first presentation to diagnosis interval for patients with colon cancer (eg, adjusted median interval of 204 vs 126 days, $p=0.04$).²² Meanwhile, there was no such association among patients with rectal cancer,²³ possibly reflecting that patients with rectal cancer are more likely to present with localising symptoms (online supplemental appendix S8).

Overall, three robust studies provided evidence that patients from the most deprived quintile experienced a longer first presentation to diagnosis interval (table 1 and online supplemental appendix S8).

Symptom to diagnosis interval

One study estimated the effect of occupation on the time between a patient's first symptom and diagnosis.¹³

No significant effect was demonstrated, adjusting for ethnicity, age, marital status and sex ($p>0.05$) (table 1 and online supplemental appendix S8).¹³

Referral to treatment interval

Four studies evaluated the time from referral to treatment.^{15 18–20} Two studies demonstrated no significant association between deprivation and the likelihood of commencing treatment within 62 days of referral (range of ORs 1.02 – 1.07).^{18 19} Another study demonstrated reduced odds of patients commencing treatment within 62 days of referral, adjusted for age, stage, referral interval and first treatment received (OR 0.82 , 95% CI 0.74 to 0.91) (online supplemental appendix S8).¹⁵

Meanwhile, one study estimated HRs for the time between referral and first treatment, adjusting for stage, distance and presentation.²⁰ There was no significant association between deprivation and time to treatment (HR 1.24 , 95% CI 0.93 to 1.67) (online supplemental appendix S8).

Overall, the association between deprivation and this interval was inconclusive (table 1 and online supplemental appendix S8).

Table 1 Narrative synthesis—assessment of the relationship between deprivation, the system interval and treatment received

Specific outcome reported	Overall assessment/conclusion	Studies, n (subjects, n)	Studies demonstrating adverse effect of deprivation	Studies demonstrating protective effect of deprivation	Studies demonstrating no impact of deprivation	Further information
Referral to first-seen interval	Inconclusive impact of deprivation on the length of the referral to first-seen interval	3 (86 644)	1 strong ¹⁸	–	1 strong ¹⁵ 1 weak ¹³	Online supplemental appendix S8: Results of studies reporting variations in the system interval
First-seen to diagnosis interval	Inconclusive impact of deprivation on the length of the first-seen to diagnosis interval	1 (15 891)	–	1 weak ¹³	–	
Referral to treatment interval	Inconclusive impact of deprivation on the length of the referral to treatment interval	4 (69 892)	1 strong ¹⁵	–	1 strong ¹⁸ 2 weak ^{19 20}	
Diagnosis to treatment interval	Inconclusive impact of deprivation on the length of the diagnosis to treatment interval	5 (292 502)	1 strong ¹⁵ 1 moderate ¹⁷	1 strong ¹⁸	2 strong ^{14 16}	
Test to diagnosis/secondary care diagnostic interval (SCDI)	No impact of deprivation on the length of the test to diagnosis/SCDI	2 (68 794)	–	–	2 strong ^{12 23}	
First presentation to diagnosis interval	Deprivation associated with increased length of the first presentation to diagnosis interval	3 (at least 6951)	3 strong ^{*21–23}	–	1 strong ^{*23}	
Symptom to diagnosis interval	Inconclusive impact of deprivation on the length of the symptom to diagnosis interval	1 (15 891)	–	–	1 weak ¹³	
Likelihood of receipt of surgery	Strong evidence for reduced surgery with increasing deprivation	11 (374 869)	2 strong ^{*24 36} 1 moderate ²⁷ 4 weak ^{26 28 30 31}	1 strong ²⁵	1 strong ^{*36} 3 weak ^{19 20 29}	Online supplemental appendix S9: Results—likelihood of receipt of surgery
Likelihood of receipt of liver resection	Strong evidence for reduced liver resection with increasing deprivation	3 (285 194)	3 strong ^{32–34}	–	–	Online supplemental appendix S9: Results—likelihood of receipt of surgery
Likelihood of receipt of pulmonary resection	No impact of deprivation on likelihood of pulmonary resection	1 (80 869)	–	–	1 strong ³⁵	Online supplemental appendix S9: Results—likelihood of receipt of surgery
Likelihood of receipt of APER	Strong evidence for increased likelihood of APER versus AR with increasing deprivation	6 (128 946)	1 strong ³⁷ 4 weak ^{39–42}	–	1 weak ³⁸	Online supplemental appendix S11: Results—likelihood of surgical variation

Continued

Table 1 Continued

Specific outcome reported	Overall assessment/conclusion	Studies, n (subjects, n)	Studies demonstrating adverse effect of deprivation	Studies demonstrating protective effect of deprivation	Studies demonstrating no impact of deprivation	Further information
Likelihood of receipt of TPE	No impact of deprivation on likelihood of TPE versus PPE with increasing deprivation	1 (120)	–	–	1 weak ⁴³	Online supplemental appendix S11: Results — likelihood of surgical variation
Likelihood of receipt of chemotherapy	Strong evidence for reduced chemotherapy with increasing deprivation	13 (251 862)	4 strong ^{24 25 44 45} 2 moderate ^{*27 47} 5 weak ^{*19 26 46 48 50}	–	1 moderate ^{*27} 3 weak ^{*20 46 49}	Online supplemental appendix S13: Results — likelihood of receipt of chemotherapy
Likelihood of receipt of combination chemotherapy	Strong evidence for reduced use of combination chemotherapy with increasing deprivation	1 (8750)	1 strong ⁵¹	–	–	Online supplemental appendix S13: Results — likelihood of receipt of chemotherapy
Likelihood of receipt of radiotherapy	No impact of deprivation on likelihood of radiotherapy	7 (79 053)	–	1 moderate ²⁷ 1 weak ⁵²	1 strong ²⁵ 4 weak ^{19 20 26 43}	Online supplemental appendix S15: Results — likelihood of receipt of radiotherapy
Likelihood of receipt of any treatment	Moderate evidence for reduced any treatment with increasing deprivation	2 (90 138)	1 moderate ¹⁷ 1 weak ⁴⁶	–	–	Online supplemental appendix S16: Results — likelihood of receipt of any treatment

*Studies represented in more than one column due to different conclusions depending on the underlying cancer type (colon vs rectal cancer).^{23 27 36 46}
 APER, abdominoperineal resection; AR, anterior resection; TPE, total pelvic exenteration.

Results of studies reporting treatment inequalities

Results of studies reporting likelihood of receipt of primary surgery

The outcome of interest was primary surgery in 11 studies, here defined as resection of the tumour.^{19 20 24–31 36} Five studies clearly defined the outcome as a tumour resection,^{25 27–29 36} while the received surgical procedure was not identified in the other six studies (online supplemental appendix S9).^{19 20 24 26 30 31}

Across seven studies, adjustment was made for different factors: age,^{19 20 24–26 29 36} stage,^{19 20 24–26 36} sex,^{19 24–26 29 36} comorbidity,^{24 25 36} site (colon vs rectum),^{19 25 36} distance or time to hospital,^{20 26} year of diagnosis,^{24 36} region¹⁹ and histology, grade and presentation.³⁶ Meanwhile, four studies provided only rates of patients receiving surgery (online supplemental appendix S9).^{27 28 30 31}

Six studies presented reduced odds of surgery for patients from the most deprived background (range of ORs 0.32–0.99).^{24 26–28 30 31} One study presented increased odds of *not* receiving surgery among the most deprived patients with rectal cancer (OR 1.35, 95% CI 1.22 to 1.49) but no significant association among patients with colon cancer (OR 0.96, 95% CI 0.87 to 1.07).³⁶ Meanwhile, three studies demonstrated no association (range of ORs 0.52–0.88).^{19 20 29}

One study revealed a higher likelihood of surgery for patients from the most deprived background (OR 1.63, 95% CI 1.17 to 2.26).²⁵ Additionally, the study reported

increased odds of surgery in older age groups. These findings, which were unexpected, were confirmed by consulting the author. However, it is important to note that this analysis was based on regional data from a historical cohort of colorectal cancers diagnosed between 1997 and 2004. While the reported methodology appears robust, the results of this small study are opposed to other studies (see figure 4) and cautious interpretation is required.

Figure 4 displays a forest plot, which provides an overview of the findings from multiple studies investigating the likelihood of undergoing surgery for colorectal cancer. The plot reveals that a majority of studies considering primary surgery (10/12) indicate a decrease in the likelihood of surgical intervention among patients belonging to the most deprived group. Overall, the evidence strongly supports the hypothesis that patients from the most deprived group are less likely to receive surgery (table 1 and online supplemental appendix S9).

Results of studies reporting likelihood of receipt of surgery for oligometastatic disease

Four studies examined the receipt of surgery in presumed oligometastatic disease, all adjusted for age, stage, comorbidity and site (colon vs rectal).^{32–35} Three studies examined the receipt of liver resection, demonstrating significantly reduced odds of resection for patients from

Forest plot demonstrating odds of receipt of surgery

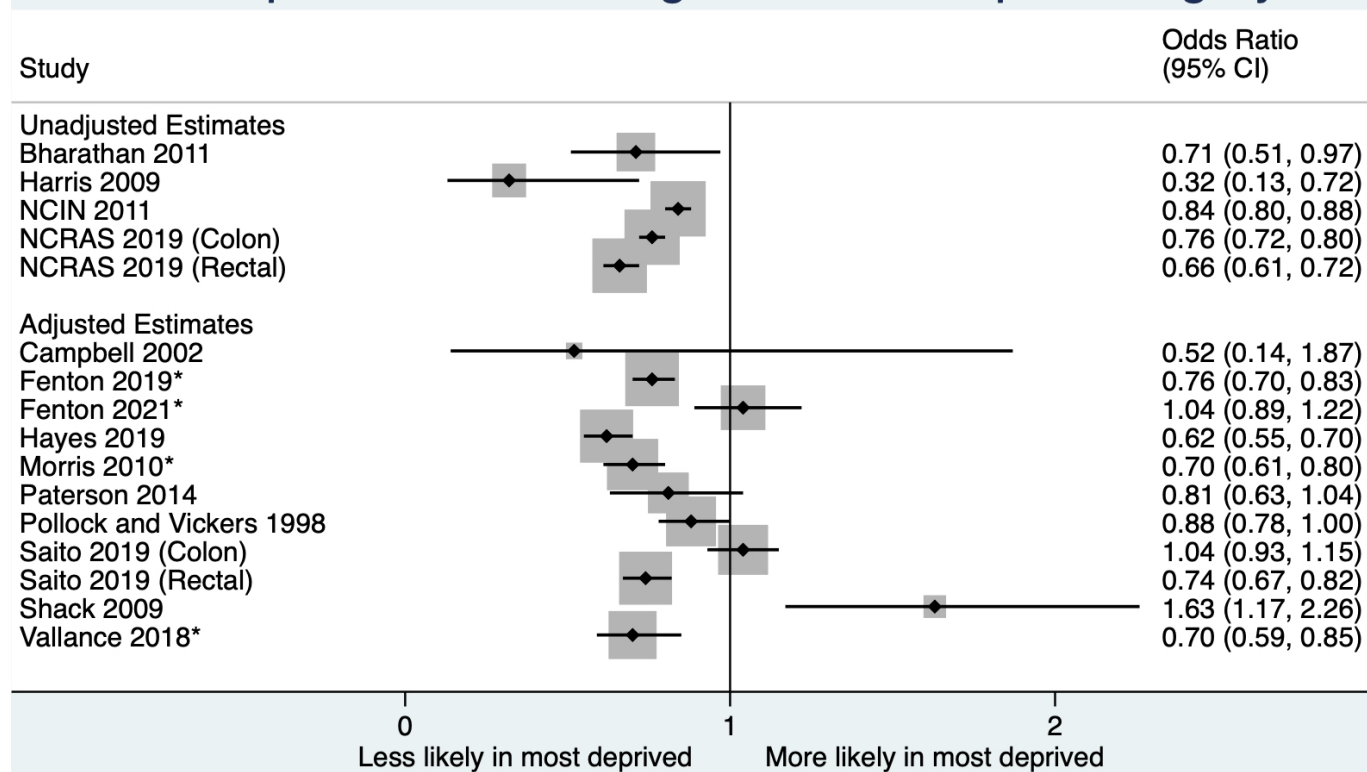


Figure 4 Forest plot demonstrating the odds of receipt of surgery in the most deprived versus the least deprived patient group.

the most deprived group (range of ORs 0.70–0.76).^{32–34} One study examined the receipt of pulmonary resection, with no significant association demonstrated between deprivation and the likelihood of resection (OR 1.04, 95% CI 0.89 to 1.22) (table 1 and online supplemental appendix S9).³⁵ and Figure 4 displays a forest plot, providing an overview of the findings from these studies, each highlighted with an asterisk.

Results of studies reporting likelihood of surgical variation

Seven studies evaluated variations in surgery.^{37–43} Six reported rates or odds of abdominoperineal resection (APER) or anterior resection (AR).^{37–42} Five studies adjusted for variables, including age,^{37–40} sex,^{37–41} stage,^{37,38} year of diagnosis or resection,^{37–41} surgeon workload^{37,38} and admission type.^{37–40} Online supplemental appendix S10 displays a forest plot, providing an overview of the findings from these studies. Five of the seven studies demonstrated that APER was significantly more likely than AR for patients from the most deprived areas (range of ORs 1.37–1.64) (table 1 and online supplemental appendix S11).^{37,39–42}

Meanwhile, one study of 120 patients presented unadjusted rates of total pelvic exenteration (TPE) compared with partial pelvic exenteration.⁴³ There was a non-significant association between deprivation and the unadjusted odds of TPE (OR 1.75, 95% CI 0.55 to 5.68) (table 1 and online supplemental appendix S11).

Results of studies reporting likelihood of receipt of chemotherapy

In total, 13 studies examined whether patients received any chemotherapy,^{19,20,24–27,44–50} 11 of which conducted adjusted analyses.^{19,20,24–26,44–49} Six studies evaluated the use of adjuvant chemotherapy.^{24,44,45,49–51} Two studies evaluated the use of palliative chemotherapy.^{24,46} Meanwhile, the intent of chemotherapy was unknown in the remaining seven studies.^{19,20,25–27,47,48}

Online supplemental appendix S12 displays a forest plot, providing an overview of the findings from the studies. Eight studies demonstrated reduced adjusted odds of chemotherapy for patients from the most deprived group (range of ORs 0.44–0.99).^{19,24–26,44,45,47,48} One study demonstrated reduced adjusted odds for patients from the most deprived group with colon (OR 0.45, 95% CI 0.27 to 0.77) but not rectal cancer (OR 0.73, 95% CI 0.36 to 1.50).⁴⁶ Two studies did not show a significant association between deprivation and receipt of chemotherapy (range of ORs 0.49–2.13) (online supplemental appendix S13).^{20,49}

Meanwhile, two studies presented unadjusted rates.^{27,50} One demonstrated reduced odds of chemotherapy for the most deprived patients with colorectal cancer (OR 0.31, 95% CI 0.09 to 0.91).⁵⁰ The other demonstrated reduced odds of chemotherapy for the most deprived patients with colon (OR 0.85, 95% CI 0.81 to 0.89) but not rectal cancer (OR 1.03, 95% CI 0.95 to 1.11) (online supplemental appendix S13).²⁷

One study examined the receipt of combination versus single-agent chemotherapy, adjusting for age, sex, ethnicity, tumour size, lymph node yield and year of diagnosis.⁵¹ However, no adjustment was made for comorbidity. Patients from the most deprived area had significantly reduced odds of receiving combination chemotherapy (OR 0.50, 95% CI 0.42 to 0.59) (online supplemental appendix S13).⁵¹

Five of the six studies evaluating the use of adjuvant chemotherapy demonstrated inequalities.^{24,44,45,50,51} Meanwhile, both studies evaluating the use of palliative chemotherapy demonstrated similar inequalities.^{24,46} Overall, the evidence strongly supports the hypothesis that patients from the most deprived group are less likely to receive chemotherapy or combination adjuvant chemotherapy (table 1 and online supplemental appendix S13).

Results of studies reporting likelihood of receipt of radiotherapy

Seven studies reported receipt of radiotherapy by socioeconomic group.^{19,20,25–27,43,52} Two studies evaluated the use of neoadjuvant radiotherapy.^{19,43} One study evaluated patterns of preoperative and postoperative radiotherapy.⁵² The intent of radiotherapy was unknown in four studies.^{20,25–27}

Three studies conducted analyses that adjusted for important factors, including age,^{20,25,26} stage,^{20,25,26} sex,^{25,26} distance or journey time,^{20,26} tumour site (colon vs rectum)²⁰ and comorbidity.²⁵ None of these studies demonstrated a significant association between deprivation group and radiotherapy (range of ORs 0.85–0.99). Online supplemental appendix S14 presents a forest plot, providing an overview of the findings from these studies. The remaining four studies reported unadjusted rates of radiotherapy.^{19,27,43,52} Two of these studies demonstrated increased odds of radiotherapy for patients from the most deprived group (range of ORs 1.33–1.39).^{27,52} The other two studies looked at rates of neoadjuvant radiotherapy specifically and did not show a significant association between deprivation and odds of treatment (range of ORs 1.00–1.15) (online supplemental appendix S15).^{19,43}

Overall, there was no evidence to support an association between socioeconomic status and receipt of radiotherapy (table 1 and online supplemental appendix S15). This conclusion may depend on the intent of radiotherapy and would, therefore, have been stronger if all outcomes were differentiated by intent (eg, neoadjuvant or palliative).

Results of studies reporting receipt of any treatment

Two studies evaluated the likelihood of any treatment by deprivation quintile, adjusting for age,^{17,46} sex⁴⁶ and stage.^{17,46} It was assumed this meant receiving surgery, radiotherapy or chemotherapy. However, these outcomes needed to be more clearly defined. For the most socioeconomically deprived quintile, both studies reported significantly reduced odds of any treatment within 6 months of diagnosis⁴⁶ or 6 months of the first contact with the NHS (range of ORs 0.54–0.87) (table 1 and online supplemental appendix S16).¹⁷

DISCUSSION

Main findings

This is the first systematic review to evaluate what is already known about the relationship between socioeconomic status, the system interval and the treatment that patients with colorectal cancer receive.

Diagnostic and treatment delays

There were seven intervals evaluated. The evidence for system delays was generally inconclusive, given substantial heterogeneity in methods and outcomes. However, there was substantial evidence that the first presentation to diagnosis interval was longer for patients from the most deprived background, depending on the underlying site. The underlying reasons require further elucidation using qualitative studies. This would help us understand the extent to which these delays are driven by patient or healthcare factors and how these can be addressed. Possible causes include missed appointments due to competing demands such as employment or care responsibilities.^{53 54} Other reasons might include complex transport and travel arrangements causing difficulties in attending appointments.^{53 54}

Surgery in the management of colorectal cancer

There was strong evidence for inequalities in primary surgery. However, most studies had limitations; few adjusted for stage, most combined colon and rectal cancers and many included patients diagnosed before 2010.

There was also strong and consistent evidence that patients from the most deprived areas were less likely to undergo a liver resection and were more likely to undergo an APER than AR. APER is associated with a worse quality of life and is generally considered less preferable if a less deforming surgery is possible.

Despite adjustment, socioeconomic inequalities were frequently observed. This suggests the presence of uncaptured factors such as comorbidity or frailty. There may also have been variations in access to specialist care, financial and employment factors, patient choice, health-seeking behaviours and health literacy, all of which warrant further investigation.^{55–57}

Chemotherapy in the management of colorectal cancer

There was strong evidence that patients from more deprived areas were less likely to receive chemotherapy or combination adjuvant chemotherapy. Trust in clinicians, financial and employment factors, social support, adequate communication and provision of information are critical in influencing the use of chemotherapy.^{58–61} These, among other uncaptured factors such as comorbidity or frailty, could be responsible for the observed inequalities.

Radiotherapy in the management of rectal cancer

There was no evidence that patients from more deprived areas were less likely to receive radiotherapy. The absence of observed inequalities could reflect the nature of this

outpatient treatment and the availability of patient transport. This is compared with, for example, surgery, which necessitates hospital admission and prolonged time away from work and social support. A lung cancer study similarly demonstrated a greater likelihood of radiotherapy but a reduced likelihood of surgery among less affluent patients.⁶²

Strengths and weaknesses

This systematic review identified many studies and employed a robust methodology. The process of identifying search terms was thorough, and the search was validated. The searches were extensive, conducted across eight databases, supplemented with citation searching and a thorough examination of the grey literature. These additional search methods identified six studies.^{27 28 35 36 44 52} Inclusion of non-peer-reviewed literature was also a key strength of this review.^{25 27 28 36}

The included studies were, however, heterogeneous in the methodology and populations studied. Out of 41 studies, only 15 included patients diagnosed after 2010.^{12 14 18 21–23 27 32 33 35 36 43–45 51} Of the six studies evaluating the system interval in patients diagnosed since 2010, four demonstrated some inequalities.^{18 21–23} Meanwhile, seven out of the nine studies that evaluated inequalities in treatments among patients diagnosed after 2010 demonstrated the presence of inequalities.^{27 32 33 36 44 45 51} Therefore, although most studies included patients from over a decade ago, inequalities persisted in recent cohorts despite a national focus on reducing inequalities.

Another limitation was that studies frequently analysed colorectal cancer as a single disease despite differences in presentation and management. Significantly, no study used causal inference approaches, exemplified by an absence of reported directed acyclic graphs.⁶³ The methods used could have introduced a bias known as the ‘table 2 fallacy’, whereby estimates from regression models are mistakenly interpreted.⁶³ Using a causal approach to future studies would considerably strengthen the interpretation and, thus, meaningfully impact policy.⁶⁴

Implications for policy and practice

Due to significant heterogeneity across studies, we could not firmly conclude whether patients from more deprived backgrounds systematically experience longer system intervals. However, COVID-19 detrimentally impacted cancer diagnostic activity for most patients, especially those in deprived areas.⁵ It is important to ensure measures are in place to monitor the system interval for patients most at risk of delays.⁵

There was strong evidence of socioeconomic inequalities in surgery and chemotherapy. Some inequalities may partly be due to wording in clinical guidelines. For example, the National Institute for Health and Care Excellence advises that primary surgery for colorectal cancer is ‘offered’ (a strong recommendation); the same guideline advises liver resection be ‘considered’ (less certain benefit).⁶⁵ Similarly, adjuvant chemotherapy can

be estimated to reduce the risk of death in stage III disease by 10%–15%. However, there is a significant risk of long-term toxicity. Patients must carefully weigh the potential harms and benefits of these less strongly recommended treatments. Shared decision-making is vital. Inequalities will result when some patients experience better shared decision-making and can cover the costs of additional treatment, such as time off work.⁶⁶

Clinicians can mitigate some of the effects of deprivation. Such strategies may include referring patients for pre rehabilitation, tailored communication and ensuring patients are aware of appropriate financial support and transport schemes.⁶⁶

Further studies are needed to evaluate for inequalities in novel treatments. In the era of precision oncology and an ever-increasing armamentarium of novel treatments, the marginal benefits of new therapies must not just be experienced by the most affluent. A prostate cancer study exemplified this, demonstrating that patients from more deprived backgrounds living at greater distances from specialist centres were significantly less likely to receive robotic prostatectomy.⁶⁷ If we accept the benefit of newer surgical technology and techniques, such as robotic surgery, these should be available for all patients no matter where they live.

Future research

Further research evaluating the whole of the system interval is needed. Further research should also aim to understand why deprivation is associated with a reduced likelihood of chemotherapy and surgery. In particular, observational research of recent cohorts should use causal inference. Beyond this, qualitative research will be of great value in gaining a richer insight into the causes and drivers of these inequalities.

CONCLUSIONS

Despite a healthcare system that provides free healthcare at the point of access, there were unexplained socioeconomic inequalities in surgery, chemotherapy and aspects of the system interval. Further research is needed to understand the variations in treatment between socioeconomic groups.

Differences in patient selection for treatment have been linked with worse colorectal cancer survival within and between countries, with evidence of improved outcomes when care is aligned with optimal pathways.⁶⁸ Eliminating inequalities could narrow survival gaps within and between countries. These findings will interest policymakers, clinicians and researchers worldwide, as inequalities in cancer care and outcomes of different socioeconomic groups have been recognised across healthcare jurisdictions.

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