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The efect of aperiodic components in distinguishing Alzheimer's disease from frontotemporal dementia

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Abstract Distinguishing between Alzheimer's disease (AD) and frontotemporal dementia (FTD) presents a clinical challenge. Inexpensive and accessible techniques such as electroencephalography (EEG) are increasingly being used to address this challenge. In particular, the potential relevance between aperiodic components of EEG activity and these disorders has gained interest as our understanding evolves. This study aims to determine the diferences in aperiodic

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activity between AD and FTD and evaluate its potential for distinguishing between the two disorders. A total of 88 participants, including 36 patients with AD, 23 patients with FTD, and 29 healthy controls (CN) underwent cognitive assessment and scalp EEG acquisition. Neuronal power spectra were parameterized to decompose the EEG spectrum, enabling comparison of group diferences in diferent components. A support vector machine was employed to assess the impact of aperiodic parameters on the diferential diagnosis. Compared with the CN group, both the AD and FTD groups showed varying degrees of increased alpha power (both periodic and raw power) and theta alpha power ratio. At the channel level, theta power (both periodic and raw power) in the frontal regions was higher in the AD group compared to the FTD group, and aperiodic parameters (both exponents and ofsets) in the frontal, temporal, central, and parietal regions were higher in the AD group than in the FTD group. Importantly, the inclusion of aperiodic parameters led to improved performance in distinguishing between the two disorders. These fndings highlight the signifcance of aperiodic components in discriminating dementia-related diseases.

Keywords Alzheimer's disease ·

Frontotemporal dementia · Diferential diagnosis · Electroencephalography · Aperiodic · Machine learning

Background

Alzheimer's disease (AD) and frontotemporal dementia (FTD) are the two most prevalent causes of dementia [\[1](#page-14-0)]. AD is characterized by amnesia, fuency aphasia, and visuospatial difficulties, while FTD is characterized by changes in personality and behavior [\[2](#page-14-1)]. However, both diseases often present with overlapping clinical symptoms, and FTD can exhibit diferent patterns of brain atrophy (i.e., temporal and frontal variants), making diagnosis challenging [\[3](#page-15-0)]. Neuropsychological tests have yielded inconclusive or even contradictory results [\[4](#page-15-1), [5\]](#page-15-2). Neuroimaging techniques, such as positron emission tomography, single photon emission computed tomography, and functional magnetic resonance imaging, have been used to identify afected brain regions to improve diagnostic accuracy [\[6](#page-15-3)[–8](#page-15-4)]. Nevertheless, their high cost and limited availability have been a constraint. Electroencephalography (EEG), on the other hand, has gained increased interest in clinical practice and research due to its low cost and wide availability [[9\]](#page-15-5).

Many studies have consistently found an overall slowing of EEG rhythm in patients with AD, evidenced by increased power in slow rhythm (i.e., delta and theta bands) and decreased power in fast rhythm (i.e., alpha and beta bands) $[10-12]$ $[10-12]$, resulting in an increased theta alpha power ratio (TAR), which has shown promise in diferentiating patients with AD from healthy controls [[13\]](#page-15-8). In contrast, there have been fewer quantitative EEG studies of patients with FTD with controversial results, such as no increase in delta and theta band power while alpha band power increases [\[14](#page-15-9)], or an increase in theta power while alpha power remains unchanged [\[15](#page-15-10)]. It is worth noting that previous power spectrum analyses have focused on narrowband power, which measures relative or absolute power changes across specifc frequency bands.

In recent years, the dominant scale-free "1/*f*" components of the power spectrum, known as the aperiodic components, have attracted much attention [\[16](#page-15-11)[–19](#page-15-12)]. Previous studies considered "1/*f*" activity as neural noise and often removed it from analyses to emphasize brain oscillations [[17\]](#page-15-13). However, recent studies have shown that these aperiodic components are not unstructured noise but have unique functional signifcance that indexes the excitation/inhibition (E/I) balance of neuronal populations $[20]$ $[20]$. It is affected by age $[21, 22]$ $[21, 22]$ $[21, 22]$, state of consciousness $[23, 23]$ $[23, 23]$ $[23, 23]$ [24\]](#page-15-18), and disease, including attention-deficit/hyperactivity disorder, schizophrenia, and Parkinson's disease [\[25](#page-15-19)[–27](#page-15-20)]. Notably, the aperiodic components have been found to be excellent predictors of higher-order cognitive tasks, such as cognitive processing speed and language learning [[28,](#page-15-21) [29](#page-15-22)]. To date, no studies have investigated the effect of aperiodic components in dementia-related diseases. Therefore, we employed a recently proposed method of parameterizing neuronal power spectra to analyze scalp EEG of patients with AD and FTD [\[30](#page-16-0)], exhaustively describing all components of the EEG power spectrum and examining the efect of aperiodic components in distinguishing between these two disorders, aiming to improve diagnostic accuracy afordably and conveniently.

Methods

Participants

We reanalyzed data that were collected from previously published studies [\[31](#page-16-1)[–34](#page-16-2)]. This dataset consists of EEG data from 36 patients with AD, 23 patients with FTD, and 29 healthy controls (CN). Cognitive and neuropsychological status was evaluated using the international Mini-Mental State Examination (MMSE), which ranges from 0 to 30, with lower MMSE scores indicating more severe cognitive decline. The duration of the disease was measured in months, and the median value was 25 with an interquartile range of 24–28.5 months. Regarding the AD groups, no comorbidities related to dementia were reported. All patients were diagnosed by an experienced team of neurologists at the 2nd Department of Neurology of AHEPA General Hospital of Thessaloniki.

EEG acquisition

Recordings were acquired from the 2nd Department of Neurology of AHEPA General Hospital of Thessaloniki. A Nihon Kohden EEG 2100 clinical device with 19 scalp electrodes (Fp1, Fp2, F7, F3, Fz, F4, F8, T3, C3, Cz, C4, T4, T5, P3, Pz, P4, T6, O1, and O2) according to the 10–20 international system, and 2 reference electrodes (A1 and A2) placed on the mastoids for impedance check were used for recording, as specifed in the manual of the device. Each recording was conducted according to the clinical protocol, with participants in a sitting position and their eyes closed. Before the initialization of each recording, the skin impedance value was checked to ensure it was below 5 kΩ. The sampling rate was 500 Hz with a 10 uV/mm resolution. The recording montages were anterior–posterior bipolar and referential montages using Cz as the common reference. The recordings were received under the range of the following parameters of the amplifer: sensitivity, 10 uV/ mm; time constant, 0.3 s; and high-frequency filter at 70 Hz. Each recording lasted approximately 13.5 min for the AD group (min=5.1, max=21.3), 12 min for the FTD group (min=7.9, max = 16.9), and 13.8 min for the CN group (min=12.5, max=16.5). In total, 485.5 min of AD, 276.5 min of FTD, and 402 min of CN recordings were collected.

Signal processing

All data preprocessing steps and subsequent analyses were performed using MATLAB (Version 2020a, MathWorks, Inc., Natick, MA, USA) and Python (Version 3.8). First, the EEG data were bandpass fltered using *pop_frws* function in the EEGLAB toolbox (<http://www.sccn.ucsd.edu/eeglab/>) with a Hamming window with lower and upper cutoff frequencies of 0.5 and 45 Hz, respectively [[35\]](#page-16-3). Subsequently, an adaptive method named Artifact Subspace Reconstruction was used to remove occasional largeamplitude artifacts $[36]$ $[36]$, and an independent component analysis was performed to decompose constant fxed source noise [[37\]](#page-16-5). The data were then manually inspected for periods of artifact, which were marked for exclusion. After artifact removal, the data were re-referenced to a common average reference due to unknown locally confned spatial origins of aperiodic activity.

Data analysis

To avoid EEG instability at the start of the experiment and subject fatigue at the later stages of the experiment, EEG data of a length of 3 min from 2 min after the start of the experiment were selected in each subject for subsequent analysis. Power spectral density (PSD) estimation was performed using the Welch method with a 0.5-s Hamming window, 50% overlap, and a 0.5-Hz resolution. Then, the FOOOF Python package was used to separate periodic and aperiodic components from the spectral results [[30\]](#page-16-0). The power spectral density P(*f*) for each frequency *f* is expressed as: $P(f) = L(f) + \sum_{n} G_n(f)$.

Where the $P(f)$ is a combination of the aperiodic components $L(f)$ and Gaussians $G_n(f)$.

The periodic components were parameterized as a mixture of Gaussian distributions: $G_n(f) = \alpha * \exp\left(\frac{-(f-c)^2}{2*\omega^2} \right)$ $\frac{(f-c)^2}{2\ast\omega^2}$. With α as the height of the peak, over and above the aperiodic components, c as the center frequency of the peak, ω as the width of the peak, and f as the array of frequency values.

The aperiodic components were also parameterized using a Lorentzian function as: $L(f) = b - \log(k + f^{\chi})$. With b as the broadband "offset," k as the "knee," and χ as the "exponent" of the aperiodic ft.

The full model was described using these periodic and aperiodic parameters, the goodness of ft was estimated by comparing each fit with the original PSD in terms of the mean absolute error and the *R*2 of the ft. To avoid overftting, *peak_width_limits* were set at 1–12, *ftted frequency range* was 2–40 Hz, and *aperiodic_mode* was chosen as "fixed" (fitted without a knee parameter) considering the narrow frequency range. For further analysis, all the models meet the condition that R^2 value is higher than 0.95 to exclude the potential infuence of noise on the results. After completing the parameterization of the power spectra, the ftted spectra were used to subtract the aperiodic components to obtain periodic oscillatory components. At last, we use *bandpower* function to calculate the power values, which means calculating the average power in a frequency interval by integrating the power spectral density estimate. We investigated the band-limited power of raw spectra, periodic, and aperiodic components in three frequency bands traditionally used in EEG analysis: theta (4–8 Hz), alpha $(8-12 \text{ Hz})$, and beta $(12-30 \text{ Hz})$. Our analyses were limited to absolute power because relative power measurements were artificially affected by normalization.

Classifcation and machine learning algorithms

The library for support vector machine (Version 3.25, [https://www.csie.ntu.edu.tw/~cjlin/libsvm/\)](https://www.csie.ntu.edu.tw/~cjlin/libsvm/) package for MATLAB (Version 2020a, MathWorks, Inc., Natick, MA, USA) was adopted $[38]$; we assessed the ability of each feature and their combinations in distinguishing between patients with AD and FTD using a support vector machine (SVM) classifer (see Fig. [9a](#page-12-0) for pipeline schematic diagram). Specifcally, each feature mentioned above was averaged across channels for each participant. First, we oversampled the samples with the SMOTE (Synthetic Minority Oversampling Technique) algorithm to address the sample imbalance. The SMOTE algorithm operates by selecting a random sample from the minority class and then identifying its k-nearest neighbors. Instead of simply replicating the chosen sample, SMOTE generates new data points along the line segments connecting the original sample to these k-nearest neighbors. This strategy is employed to address the issue of overftting and enhance the quality of synthetic samples, thus contributing to a more efective and balanced dataset [\[39\]](#page-16-7). Next, all participants were randomly assigned to a training set and a test set with a ratio of 7:3, and the value of each feature was normalized in each group. The hyperparameters of each classifer were optimized with inner fvefold cross-validation. Then, an SVM model with a radial basis function kernel was trained using the whole training set with the optimal hyperparameters. We subsequently repeated the whole process 100 times using diferent randomized training and testing sets to ensure the robustness of our results, classifcation efectiveness evaluation was obtained using the receiver operating characteristic curve and area under curve (AUC). Finally, we obtained 100 AUCs for each feature combination and compared them among diferent feature combinations.

Statistical analysis

All statistical analyses were performed with SPSS 25.0 software (SPSS Inc., Chicago, IL, USA), and the significance level was set at α =0.05. The Shapiro–Wilk method was used to test the normality of the data distribution. Since EEG parameters were not normally distributed in all groups, variables were log-transformed before comparison between groups. For channel-averaged group comparisons, EEG parameters for the three groups were analyzed using one-way analysis of variance (ANOVA) and post hoc tests with Bonferroni correction for multiple comparisons. Channel-bychannel group comparisons were performed using Student's *t*-test, and the resulting *p*-values were corrected for multiple comparisons using the false discovery rate (FDR) procedure across 19 channels. Spearman correlation analysis was used to test the correlation between EEG parameters and MMSE scores, and the FDR procedure was used to correct for multiple comparisons. AUC values obtained for diferent feature combinations were compared using one-way ANOVA and post hoc tests (Bonferroni correction).

Results

Demographic

All three groups had similar age distributions (see Table [1\)](#page-3-0). In the AD group, a higher proportion of participants were female, whereas more FTD participants were male. Consistent with the disease characteristics, the AD group exhibited more severe impairment

	$AD(N=36)$	$CN (N=29)$	$FTD(N=23)$	Group differences
Male to female	12:24	18:11	14:9	χ^2 = 6.8, p = 0.034 ^a $p(AD, CN) = 0.021$ $p(AD, FTD) = 0.038$ p (CN, FTD) = 0.930
Age	66.4(7.9)	67.9(5.4)	63.7(8.2)	F $(2, 85) = 2.218$, $p = 0.115^b$
MMSE	17.8(4.5)	30(0)	22.2(2.6)	$t_{57} = -4.3, p < 0.001^c$

Table 1 Demographic and clinical variables, mean (standard deviation)

AD Alzheimer's disease, *CN* healthy control, *FTD* frontotemporal dementia, *MMSE* Mini-Mental State Examination

a Chi-square test AD, CN, FTD

b One-way ANOVA AD, CN, FTD

c Student's *t*-test AD, FTD

in overall cognitive performance (MMSE score) compared to the FTD group ($t_{57} = -4.3$, $p < 0.001$).

Neuronal power spectrum parameterization

Using the method of parameterizing neural power spectral, we separated the raw power into periodic and aperiodic components, with the periodic components represented by periodic power and the aperiodic components represented by offsets and exponents (as illustrated in Fig. [1](#page-4-0), see Fig. [2](#page-4-1) for all ftting curves in three groups). ANOVAs showed the following (see Table [2](#page-5-0) and Fig. [3\)](#page-6-0).

Group diferences in theta power were observed in the raw power analysis (F $(2, 85) = 3.25$, $p = 0.044$), but post hoc multiple comparisons revealed no signifcant diferences among the three groups. In the alpha band, consistent group diferences were observed in both raw power and periodic power analyses. Specifcally, alpha power was reduced in the AD (raw power, mean=4.98, $SD = 0.38$; periodic power, mean=0.40, $SD = 0.32$) and FTD (raw

Fig. 1 Schematic diagram of separating the power spectra. Raw power (left), periodic fits (middle), and aperiodic fits (right) from an example participant, where raw power is equal

to periodic fts plus aperiodic fts. (Spectrogram (top) and time–frequency diagram (bottom))

Fig. 2 Average PSDs in three groups, including linear scale (left) and log–log scale (right) (95% confidential interval was shown in shadow area). AD, Alzheimer's disease; CN, healthy control; FTD, frontotemporal dementia; PSD, power spectral density

	AD	CN	FTD	Group differences
Theta raw power	5.19 [5.06, 5.33]	4.96 [4.82, 5.10]	4.99 [4.78,5.19]	$F(2, 85) = 3.25, p = 0.044$ $p (AD, CN) = 0.068$ $p(AD, FTD) = 0.174$ p (CN, FTD) = 1.0
Alpha raw power	4.98 [4.85, 5.11]	5.33 [5.16, 5.50]	4.99 [4.80,5.18]	$F(2, 85)=6.61, p=0.002$ $p(AD, CN) = 0.004$ $p(AD, FTD) = 1.0$ p (CN, FTD) = 0.013
Beta raw power	16.70 [16.38,17.01]	17.18 [16.88, 17.55]	16.93 [16.30, 17.55]	$F(2, 85) = 1.71, p = 0.187$
Theta periodic power	0.28 [0.23, 0.32]	0.24 [0.16, 0.31]	0.20 [0.13, 0.27]	$F(2, 85) = 1.51, p = 0.228$
Alpha periodic power	0.40 [0.29, 0.51]	0.81 [0.68, 0.94]	0.45 [0.32, 0.58]	$F(2, 85) = 14.63, p < 0.001$ p (AD, CN) < 0.001 $p(AD, FTD) = 1.0$ p (CN, FTD) < 0.001
Beta periodic power	0.43 [0.31, 0.55]	0.91 [0.73, 1.08]	0.48 [0.35, 0.61]	$F(2, 85) = 14.46, p < 0.001$ p (AD, CN) < 0.001 $p (AD, FTD) = 1.0$ p (CN, FTD) < 0.001
Offsets	1.02 [0.84, 1.20]	0.76 [0.63, 0.90]	0.73 [0.49, 0.98]	$F(2, 85) = 3.323, p = 0.041$ $p (AD, CN) = 0.109$ $p(AD, FTD) = 0.089$ p (CN, FTD) = 1.0
Exponents	1.31 [1.17, 1.46]	1.12 [1.02, 1.23]	1.06 [0.85, 1.27]	$F(2, 85) = 0.512, p = 0.048$ $p(AD, CN) = 0.193$ $p(AD, FTD) = 0.071$ p (CN, FTD) = 1.0
TAR	1.05 [1.02, 1.07]	0.93 [0.92, 0.95]	1.00 [0.97, 1.03]	$F(2, 85) = 22.43, p < 0.001$ p (AD, CN) < 0.001 $p(AD, FTD) = 0.053$ p (CN, FTD) = 0.001

Table 2 Group comparison of channel-averaged EEG parameters

Mean (95% confdence interval) of diferent EEG parameters. Group comparisons were performed using ANOVAs followed by post hoc tests, Bonferroni-corrected for multiple comparisons

AD Alzheimer's disease, *CN* healthy control, *FTD* frontotemporal dementia, *TAR* theta alpha power ratio

power, mean=4.99, SD=0.44; periodic power, mean=0.45, $SD = 0.31$) groups compared to the CN (raw power, mean=5.33, $SD = 0.45$; periodic power, mean=0.81, $SD=0.34$) group, with no significant diference observed between the two disease groups $(p=1.0)$. In the beta band, no group differences were observed in the raw power analysis, but the periodic power analysis revealed that the AD (mean=0.43, $SD = 0.34$) and FTD (mean=0.48, $SD = 0.30$) groups reduced compared to the CN group (mean=0.90, SD=0.45). Both offsets (F $(2, 85) = 3.323$, $p = 0.041$) and exponents (F $(2, 85) = 0.512$, $p = 0.048$) showed signifcant group diferences at the signifcance level we set, but post hoc multiple comparisons revealed no signifcant diferences among the three groups in the aperiodic parameters. Finally, raw theta/alpha power ratio (TAR) showed signifcant group diferences (F $(2, 85) = 22.43$, $p < 0.001$), with an increased ratio observed in the AD (mean=1.05, $SD = 0.08$) and FTD (mean=1.00, $SD = 0.07$) groups compared to the CN group (mean=0.93, $SD=0.04$), but no signifcant diference was observed between the two disease groups $(p=0.053)$.

Considering the unknown origin of the aperiodic components, to exclude the potential infuence of channel averaging on the results, we plotted topographies of EEG parameters on all 19 electrodes to illustrate the spatial patterns of EEG parameters in three groups and conducted pairwise group comparisons. Corresponding results of the statistical

(a) Raw power

Fig. 3 Channel-averaging results of the parameterized spectrum. **a** Group comparison of raw power. **b** Group comparison of periodic power. **c** Group comparison of aperiodic parameter. **d** Group comparison of θ/α power ratio. In each box plot, the center line corresponds to the median of the sample; the upper

comparisons between groups are shown in Supplementary Tables S1 and S2.

In terms of raw power, signifcant diferences were found between the AD and CN groups in all three bands but in diferent brain regions (see Fig. [4](#page-7-0)). Specifcally, diferences in the theta band were concentrated in the frontotemporal region, while diferences

and lower borders of the boxes indicate the 25th and 75th percentiles, respectively; the length of the whiskers represents the range of the data. $*p < 0.05$, $**p < 0.01$. Corresponding results of the statistical comparisons between groups are shown in Table [2](#page-5-0)

in the beta band were observed in the parietooccipital region. In contrast, signifcant diferences in the alpha band were observed in almost the entire brain. The FTD and CN groups showed diferences in both alpha (almost the whole brain) and beta (parietal (P3 and Pz), temporal (T5), and bilateral occipital (O1 and O2)) bands. Only theta power in the frontal regions **Fig. 4** Scalp topography and group comparison for raw power in three groups. **a** Theta raw power. **b** Alpha raw power. **c** Beta raw power. All test statistics were rescaled and normalized to the percentage of the most signifcant test statistic for better visualization, such that 100 would refect the most signifcant test statistic. All electrodes with signifcant test statistics (FDR-corrected $p < 0.05$) were marked with red stars. AD, Alzheimer's disease; CN, healthy control; FTD, frontotemporal dementia; FDR, false discovery rate

(F3, F4, F7, and Fz) difered between the AD and FTD groups, while no significant differences were found in the alpha and beta bands of all channels.

After removing the aperiodic components, we analyzed only periodic power and found signifcant diferences in all three frequency bands between the AD and CN groups (see Fig. [5](#page-8-0)). Specifcally, in the theta band, diferential channels were in frontal regions (Fz) and bilateral temporal regions (T3 and T4). In the alpha and beta bands, the AD group **Fig. 5** Scalp topography and group comparison for periodic power in three groups. **a** Theta periodic power. **b** Alpha periodic power. **c** Beta periodic power. All test statistics were rescaled and normalized to the percentage of the most signifcant test statistic for better visualization, such that 100 would refect the most signifcant test statistic. All electrodes with signifcant test statistics (FDR-corrected $p < 0.05$) were marked with red stars. AD, Alzheimer's disease; CN, healthy control; FTD, frontotemporal dementia; FDR, false discovery rate

showed higher periodic power in all channels compared to the CN group. Similar results were found in the comparison of the FTD and CN groups, with the FTD group showing overall higher periodic power in the alpha and beta bands compared to the CN group.

Additionally, compared to the FTD group, the AD group had higher theta power in the frontal (F3 and Fz) and temporal (T5) regions, but no diferences were found between the two groups in the fast EEG rhythm (alpha and beta band).

The channel-by-channel comparisons of aperiodic parameters revealed that the AD group exhibited higher offsets and exponents compared to the CN group in the frontal, temporal, parietal, and occipital regions, whereas no signifcant diferences were observed in the aperiodic parameters between the FTD and CN groups across all channels (see Fig. [6](#page-9-0)). Notably, in the frontal (F3 and F4), temporal (T3),

FTD group. Finally, we compared the TAR results of all channels (see Fig. [7](#page-10-0)). Compared to the CN group, the TAR of the AD and FTD groups was increased at the whole brain level. Specifcally, we observed higher TAR in the frontal (F3, Fz, and F4), temporal (T3, T5, and T6), central (Cz), and occipital (O2) regions in the AD group compared to the FTD group.

central (C3), and parietal (P3 and Pz) regions, the AD group showed higher offsets and exponents than the

Correlation relationship analyses

We examined the possible correlations between EEG parameters and overall cognitive impairment (MMSE score) (see Fig. [8\)](#page-10-1). First, the aperiodic parameters had a signifcant and high positive correlation between each other in all three groups (AD: offsets–exponents, $\rho = 0.92$, $p_{FDR} < 0.001$; CN: offsets–exponents, ρ =0.81, *p* $_{FDR}$ <0.001; FTD: offsets–exponents, ρ =0.85, *p*_{FDR}<0.001). Also, there was a low but signifcant positive correlation between theta periodic power and ofsets in all three groups (AD: theta periodic power–offsets, $\rho = 0.45$, $p_{FDR} < 0.01$; CN: theta periodic power–offsets, $\rho = 0.47$, $p_{FDR} < 0.01$; FTD: theta periodic power–offsets, $\rho = 0.55$, $p_{FDR} < 0.01$), while the low positive correlation between periodic theta power and exponents was found only between the AD and FTD groups (AD: theta periodic

Fig. 6 Scalp topography and group comparison for aperiodic parameters in three groups. **a** Aperiodic offsets. **b** Aperiodic exponents. All test statistics were rescaled and normalized to the percentage of the most signifcant test statistic for better visualization, such that 100 would refect the most signifcant test statistic. All electrodes with signifcant test statistics (FDR-corrected $p < 0.05$) were marked with red stars. AD, Alzheimer's disease; CN, healthy control; FTD, frontotemporal dementia; FDR, false discovery rate

Fig. 8 Spearman's correlation between the severity of overall cognitive impairment (MMSE scores) and EEG parameterization results in three groups. **a** AD. **b** CN. **c** FTD. Statistically signifcant groups are marked by black asterisks on the lower left ellipse (FDR-corrected *p*-value, $* p < 0.05$, $* p < 0.01$, ****p*<0.001), and the transparent number on the upper right

represents the correlation coefficient value. Also, the color of the ellipse maps the correlation coefficient value. AD, Alzheimer's disease; CN, healthy control; FTD, frontotemporal dementia; TAR, theta alpha power ratio; MMSE, Mini-Mental State Examination; FDR, false discovery rate

power–exponents, $\rho = 0.43$, $p_{FDR} < 0.01$; FTD: theta periodic power–exponents, $\rho = 0.53$, $p_{FDR} < 0.01$). We also observed consistent signifcant positive correlations between alpha and beta periodic power (AD: alpha periodic power–beta periodic power, $\rho = 0.64$, *p FDR*<0.001; CN: alpha periodic power–beta periodic power, $\rho = 0.48$, $p_{FDR} < 0.01$; FTD: alpha periodic power–beta periodic power, $\rho = 0.45$, $p_{FDR} < 0.05$), and consistent signifcant negative correlations between TAR and periodic power (alpha and beta band) in all three groups (AD: TAR–beta periodic power, $\rho = -0.83$, $p_{FDR} < 0.001$, TAR–beta periodic power, *ρ* = −0.54, *p*_{*FDR}* < 0.001; CN: TAR–beta peri-</sub> odic power, *ρ*= −0.70, *p FDR*<0.001, TAR–beta periodic power, $\rho = -0.53$, $p_{FDR} < 0.01$; FTD, TAR–beta periodic power, *ρ*= −0.64, *p FDR*<0.001, TAR–beta periodic power, $\rho = -0.54$, $p_{FDR} < 0.01$). Notably, exponents were negatively correlated with MMSE scores in the AD and FTD groups (AD: exponents—MMSE, *ρ* = −0.13, *p* = 0.02; FTD: exponents–MMSE: $\rho = -0.24$, $p = 0.03$). However, all correlations between aperiodic parameters and MMSE scores were not signifcant after correction for FDR multiple comparisons (all p_{FDR} >0.05).

SVM analysis for diagnostic discrimination

Based on the results above, we chose certain EEG parameters of AD and FTD groups for machine learning classifcation, ANOVA tests and multiple comparison corrections were performed on the AUC values obtained for all feature combinations. All statistical results can be found in Supplementary Table 2 and 3, partially signifcantly diferent comparative results are presented in the Fig. [9](#page-12-0)b.

Compared with using theta periodic power alone, we obtained signifcantly higher AUCs using the combination of theta periodic power and aperiodic parameters (theta periodic power vs. offsets and theta periodic power, $p < 0.001$; theta periodic power vs. exponents and theta periodic power, $p < 0.001$; theta periodic power vs. offsets and exponents and theta periodic power $p < 0.001$). Similarly, we obtained signifcantly higher AUCs using the combination of TAR and aperiodic parameters compared with using TAR alone (TAR vs. offsets and TAR, $p < 0.001$; TAR vs. exponents and TAR, $p < 0.001$; TAR vs. offsets and exponents and TAR, $p < 0.001$).

Our results indicate that when using a single EEG parameter as a feature, the aperiodic parameters offsets (AUC, mean= 0.66 , SD= 0.12) and exponents (AUC, mean= 0.66 , SD= 0.13) showed the best classifcation performance. However, when multiple parameters were used as features, the combination of aperiodic parameters (offsets and exponents) and the periodic parameter (theta periodic power) resulted in the best classifcation performance (AUC, mean=0.73, $SD=0.12$). Surprisingly, the combination of all four parameters did not show the best performance (AUC, mean $= 0.71$, SD $= 0.12$).

Discussion

In this study, we aimed to explore changes in diferent components of scalp EEG between patients with AD and TD by spectral decomposition and investigate the potential of aperiodic components in distinguishing between these two disorders. Specifcally, our results revealed that aperiodic parameters (offsets and exponents) and theta power were signifcantly higher in the AD group compared to the FTD group at the channel level. Moreover, combining aperiodic parameters improved the discriminative ability of EEG features in diferentiating AD from FTD. These results suggest that aperiodic parameters could serve as a promising biomarker for the identifcation and distinction of these dementia-related diseases.

The initial analysis revealed that both the AD and FTD groups exhibited varying degrees of increased EEG low-frequency raw power compared to the CN group, as evidenced by an increase in slow rhythms (theta) and a decrease in fast rhythms (alpha and beta), or an increase in TAR, which is consistent with previous studies on scalp EEG in patients with AD [\[10](#page-15-6), [12\]](#page-15-7). Importantly, we found that this trend persisted even after the removal of the aperiodic components. Our correlation results indicate a robust association between the TAR and both alpha and beta periodic power, which aligns with Donoghue et al.'s fndings, suggesting that TAR predominantly refects alpha power [[40\]](#page-16-8). However, it is essential to recognize that the band ratio measure is a broad, non-specifc metric that amalgamates various potential spectral variations. This raises questions about the interpretability of TAR, indicating that it may not be a straightforward indicator of EEG slowing. This underscores the importance of incorporating parameterized neural power spectra in future studies to gain a more nuanced and precise understanding of underlying neural dynamics. In contrast to a study by Passant and colleagues who reported well-preserved alpha power in patients with FTD [[15\]](#page-15-10), our results showed reduced alpha power in the FTD group. We suggest that our results better refect reality since our FTD group consisted of a larger sample of patients with similar ages and symptom severity.

Our results revealed signifcantly higher theta power in the AD group compared to the FTD group, which is partially consistent with a previous quantitative EEG study [\[41](#page-16-9)], inconsistent in that our results difered only in a few channels, whereas Caso et al. found diferences in regions other than the central region. This discrepancy may be attributed to the vulnerability of averaged over diferent regions of interest to the extreme values of individual channels. Overall, our results indicate that the theta power, including both periodic and raw power, in the frontal regions of the EEG was signifcantly elevated in the AD group compared to the FTD group under

Fig. 9 Schematic diagram of the SVM model and AUC results. **a** Pipeline for training and evaluation of SVM models. Feature combinations were extracted from the total dataset (AD group, $N=36$; FTD group, $N=36$ (data over-sampling)), then normalized and divided into training and test sets in the ratio of 7:3. Hyperparametric optimization was performed in the training set using inner fvefold cross-validation to obtain the best SVM model, and the test set was used to evaluate the performance of the optimized SVM model. Finally, the whole procedure was repeated 100 times using diferent random training and test sets. Posterior probabilities were calculated to obtain ROC curves and AUC values. **b** Group comparison of AUC values in diferent feature combinations. In each violin plot, the white dot in the center corresponds to the mean of the sample; the upper and lower boundaries of the boxes indicate the 25th percentile and 75th percentile, respectively; the length of the whiskers represents the range of the stated data. $* p < 0.001$. Violin diagrams with the top 3 AUC values are represented in a darker color. Only partially signifcantly diferent comparative results are presented in the fgure, corresponding results of all statistical comparisons between different feature combinations are shown in Supplementary Table S3. AD, Alzheimer's disease; FTD, frontotemporal dementia; SVM, support vector machine; ROC, receiver operating characteristic; AUC, area under curve; O, offsets; E, exponents; T, theta raw power; TAR, theta alpha power ratio

age-matched conditions. These results suggest that varying degrees of EEG slowing may be indicative of distinct pathophysiological processes in diferent neurodegenerative diseases.

Both computational models and experimental evidence have linked aperiodic activity to an underlying E/I balance of neuronal population [[20\]](#page-15-14), where an increase in inhibition (e.g., the use of GABAergic drugs such as propofol) leads to a steepening of the power spectrum (i.e., an increase in exponents) [[24,](#page-15-18) [42\]](#page-16-10), whereas an increase in excitation (e.g., the use of 5-hydroxytryptamine reuptake inhibitors such as escitalopram) leads to a fattening of the power spectrum (i.e., a decrease in exponents) [\[43](#page-16-11)]. Additionally, there is growing evidence that alterations in neuronal E/I balance may be a potential cause of neural network dysfunction in AD [\[44](#page-16-12), [45\]](#page-16-13). Our results indicate higher offsets and exponents in the AD group compared to the FTD group, with the largest diference observed in the temporal region (channel T3). Given the good age match between the two groups, it is unlikely that the observed changes in aperiodic parameters are due to age-related diferences [\[21](#page-15-15), [46\]](#page-16-14). The higher exponents observed in the AD group may suggest a higher degree of inhibition in the neuronal population, but the results obtained from macroscopic-scale analysis difer from those observed at the microscopic scale. At the synaptic level, Aβ plaques cause local alterations in dendrites and cytosol in contact with $\mathbf{A}\beta$, leading to local loss of dendritic spines and GABAergic interneuron terminals [[47–](#page-16-15)[49\]](#page-16-16), which can promote the presence of overactive neurons that may cause epileptiform activity. Conversely, neurons with intracellular neurofbrillary tangles (NFTs), another characteristic marker of AD, exhibit signifcant loss of dendritic spines and consequent synaptic loss in all dendritic regions [\[50](#page-16-17)]. Thus, while Aβ plaques may primarily lead to neuronal hyperexcitability, NFTs may be more associated with a reduction in neuronal activity and connectivity[[44\]](#page-16-12), which could explain the contradictory results at the micro- and macro-scale. Overall, we inferred the altered E/I balance associated with AD and FTD patients from macroscopic-scale electrophysiological recordings.

Previous research has suggested a potential relationship between aperiodic activity and cognitive status [[28,](#page-15-21) [29](#page-15-22)]. However, we did not observe a signifcant correlation between aperiodic parameters and MMSE scores in either the AD or FTD groups in our study. One possible explanation for this discrepancy is that aperiodic parameters may refect disease-specifc pathological changes, such as selective degeneration of the frontal or temporal cortex in patients with FTD [\[51](#page-16-18)], in addition to cognitive performance. Nonetheless, we did observe a trend of negative correlation between aperiodic parameters and cognition performance in the FTD group, indicating that larger exponents (steeper power spectrum) were associated with poorer overall cognitive performance. This fnding is consistent with a recent study that examined the efects of anesthesia and surgery on cognitive performance, which found that a reduction in the slope of the spectrum after anesthesia (fatter power spectrum) was associated with better language learning [[52\]](#page-16-19). It should be noted that the neuropsychological assessment protocol used in this study contained more detailed and comprehensive subitems compared to the MMSE. Therefore, another explanation for the lack of signifcant correlation between aperiodic parameters and MMSE scores may be due to the limited sensitivity of the MMSE in assessing subtle cognitive changes. Future studies may need to consider more sensitive and specifc neurocognitive assessment protocols to better capture the relationship between aperiodic activity and cognitive status in AD and FTD patients. Moreover, our study did not observe any signifcant correlation between other EEG parameters and MMSE scores, which contradicts some existing research. For instance, Garn et al. found a signifcant positive correlation between relative delta power during closed-eye resting state and MMSE scores [\[53](#page-16-20)], while Meghdadi et al.'s study revealed a significant negative correlation between TAR and MMSE scores [\[13](#page-15-8)]. Although our results indicated higher TAR values in AD and FTD patients compared to healthy controls, aligning with the trend observed in Meghdadi et al.'s study, we believe that these difering results may be attributed to limitations in sample size and the characteristics of the study population. Future investigations should aim to expand the sample size, include a more diverse range of subjects with varying disease severity, and provide a more comprehensive evaluation of the relationship between EEG parameters and MMSE scores.

Furthermore, we observed a weak correlation between aperiodic parameters and periodic theta power, suggesting that they provide distinct information regarding physiological or pathological activity. This indicates that aperiodic parameters and periodic theta power could serve as complementary measures for understanding the disease, refecting different neuronal mechanisms underlying these distinct types of neural activity [\[54](#page-17-0)]. To test this hypothesis, we utilized SVM to classify the two diseases based on diferent feature combinations. We found that the use of aperiodic parameters yielded higher AUC values compared to the use of only periodic theta power. However, the combination of both aperiodic offsets and exponents did not signifcantly improve the classifcation performance, likely due to their high positive correlation, consistent with a correlation-based feature selection approach [\[55](#page-17-1)]. Interestingly, the feature combinations that included aperiodic parameters resulted in higher AUC values compared to individual parameters as features (see Supplementary Table 3), indicating that they provide non-redundant information. Specifcally, the feature combinations that included ofsets, exponents, and periodic theta power yielded the best classifcation performance. Overall, these results highlight the distinct physiological signifcance of aperiodic activity in dementia-related diseases and demonstrate the utility of aperiodic parameters in distinguishing AD from FTD.

One potential limitation of this study is the disparity in the gender composition between the AD and FTD groups. Although the impact of gender on aperiodic activity has not been reported, future research should strive to replicate these fndings with gender-matched cohorts. Furthermore, we did not diferentiate FTD by clinical phenotype, which may result in potential selection bias. Lastly, high-density EEG data should be included in further research to facilitate source analysis and accurately identify the brain regions responsible for the origin of aperiodic components.

Conclusions

Increasing evidence suggests that aperiodic components in neurophysiological signals have physiological signifcance but have not been studied in dementia-related diseases. In this study, we revealed potential diferences in neural oscillatory and non-oscillatory dynamics between AD and FTD patients, with aperiodic components exhibiting better

performance in distinguishing between AD and FTD. Our fndings provide important new insights for the precise identifcation and distinction of these two dementia-related diseases.

Abbreviations *AD***: Alzheimer's disease;** *ANOVA***: Analysis of variance;** *AUC***: Area under curve;** *CN***: Healthy controls;** *EEG***: Electroencephalography;** *E/I***: Excitation/inhibition;** *FDR***: False discovery rate;** *FTD***: Frontotemporal dementia;** *MMSE***: Mini-Mental State Examination;** *NFTs***: Neurofbrillary tangles;** *PSD***: Power spectral density;** *SMOTE***: Synthetic Minority Oversampling Technique;** *SVM***: Support vector machine;** *TAR***: Theta alpha power ratio**

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Author contribution Z.W. and W.Z. provided the concept and design; Z.W., A.L., J.Y., P.W., and Y.B. analyzed and explained the data; J.Z. and S.X. guided the data processing in this study; Z.W. and W.Z. wrote the main manuscript; W.Z., B.G., and J.Z. supervised the study. All authors read and approved the fnal manuscript.

Data availability The datasets analyzed during the current study are available in the ds004504 repository, [https://github.](https://github.com/OpenNeuroDatasets/ds004504) [com/OpenNeuroDatasets/ds004504.](https://github.com/OpenNeuroDatasets/ds004504) The custom codes used for the current study are available at [https://github.com/dcgggg/](https://github.com/dcgggg/ds004504_Aperdiodic) [ds004504_Aperdiodic](https://github.com/dcgggg/ds004504_Aperdiodic).

Declarations

Ethical approval Ethical approval was already approved for data acquisition in the previous work [\[32,](#page-16-21) [33\]](#page-16-22).

Confict of interest The authors declare no competing interests.

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