# Dietary intake and serum levels of copper and zinc and risk of hepatocellular carcinoma: A matched case-control study

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## Abstract

Background: Copper and zinc are involved in the development of multiple malignancies; yet, epidemiological evidence on hepatocellular carcinoma (HCC) is limited. This study aimed to investigate the association between dietary intake and serum levels of copper and zinc with the risk of HCC.

Methods: A total of 434 case-control pairs matched for sex and age  $(\pm 1 \text{ year})$  were included in this study. Cases with newly diagnosed HCC were from the Guangdong Liver Cancer Cohort (GLCC) study, and healthy controls were from the Guangzhou Nutrition and Health Study (GNHS). A semi-quantitative 79-item food frequency questionnaire (FFQ) was used to assess habitual dietary intakes of copper and zinc. Serum levels of copper and zinc were measured by using inductively coupled plasma mass spectrometry. The copper (Cu)/ zinc (Zn) ratio was computed by dividing copper levels by zinc levels. Conditional logistic regression models were performed to calculate the odds ratio (OR) and 95% confidence intervals (CI) for per 1 standard deviation increase (per-SD increase) in copper and zinc levels.

Results: Higher dietary intake ( $OR_{perSD increase} = 0.65, 95\%$  CI: 0.44, 0.96,  $P_{trend} = 0.029$ ) and serum levels of zinc ( $OR_{perSD increase} = 0.11$ , 95% CI: 0.04, 0.30,  $P_{trend}$  <0.001) were both associated with a lower risk of HCC. Subgroup analyses showed that the inverse association was only pronounced in men but not in women ( $P_{\text{interaction}} = 0.041$  for dietary zinc intake and 0.010 for serum zinc levels). Serum copper levels (OR<sub>perSD increase</sub> = 2.05, 95% CI: 1.39, 3.03,  $P_{\text{trend}} = 0.020$ ) and serum Cu/Zn ratio (OR<sub>perSD increase</sub> = 6.53, 95% CI: 2.52, 16.92,  $P_{\text{trend}} < 0.001$ ) were positively associated with HCC risk, while dietary copper intake and dietary Cu/Zn ratio were not associated with HCC risk.

Conclusion: Zinc may be a protective factor for HCC, especially among men, but the effects of copper on HCC risk are not clear.

Keywords: Copper; Zinc; Hepatocellular carcinoma risk; Case-control studies; Dietary intake; Serum

# Introduction

Liver cancer is the seventh most commonly diagnosed cancer and the second leading cause of cancer death worldwide in 2020.<sup>[1]</sup> The number of new cases and deaths in China were 431,383 and 412,216, respectively.<sup>[2]</sup> China accounts for about half of the total number of new cases and deaths of liver cancer in the world alone.<sup>[3,4]</sup> Hepatocellular carcinoma (HCC), the most predominant type of primary liver cancer, is insidious and highly malignant with a 5-year survival rate ranging between 5% and 30%.<sup>[5]</sup> To prevent the occurrence of HCC, the most effective initiative is to reduce its burden.<sup>[6]</sup> Accumulating evidence has suggested that dietary risks, which are important modifiable lifestyle factors for deaths,<sup>[7]</sup> play an important role in the development of HCC.[8]

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Copper (Cu) and zinc (Zn) are essential micronutrients derived from a wide variety of animal foods, legumes, grain germ, etc.<sup>[9]</sup> They are important components of metalloenzymes involved in tissue respiration, energy metabolism, and antioxidant processes, such as Cu/Znsuperoxide dismutase (Cu/Zn-SOD).<sup>[10,11]</sup> Cu/Zn-SOD is a powerful intracellular antioxidant enzyme that is essential for scavenging reactive oxygen species, but the homeostatic imbalance of copper and zinc will affect Cu/ Zn-SOD activity, causing oxidative damage to biomolecules such as DNA, proteins, and lipids, thus promoting the development of cancer.<sup>[12,13]</sup> In addition, copper and zinc have their own physiological effects in carcinogenesis. Copper is a limiting factor in cancer progression such as tumor growth, angiogenesis, and metastasis.<sup>[14]</sup> Zinc has important functions such as maintaining immune function and cell structural stability and regulating cell

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differentiation and proliferation, DNA/RNA synthesis, and repair.<sup>[15]</sup>

A case-control study also found that serum copper levels were significantly higher and serum zinc levels were significantly lower in HCC patients than in healthy subjects.<sup>[16]</sup> And there was significant copper accumulation in surgically resected HCC tissue compared with para-cancerous tissue in HCC patients.<sup>[17]</sup> These findings suggest that imbalance in copper-zinc homeostasis may be closely related to the incidence of HCC. Nevertheless, there are few epidemiological studies regarding copper and zinc and the risk of HCC. A nested case-control study from the European Prospective Investigation into Cancer and Nutrition (EPIC), including 106 pairs of HCC patients and healthy controls, found that higher serum zinc levels and lower copper-to-zinc ratio are associated with a lower risk of HCC.<sup>[18]</sup> However, two prospective cohort studies from China reported null association between dietary copper and zinc intakes and the risk of liver cancer.<sup>[19]</sup> To our knowledge, limited studies have examined the role of copper and zinc, measured both in diets and blood samples, in the development of HCC.

Against this background, this study was designed to investigate the association between dietary intake, serum levels of copper and zinc with the risk of HCC in a 1:1 case-control study.

# **Methods**

# Study population

This 1:1 matched case-control study was conducted in Guangdong, China. HCC cases were from the Guangdong Liver Cancer Cohort (GLCC). A prospective cohort study of patients aged 18-80 years with incident liver cancer, was established in September 2013 at Sun Yatsen University Cancer Center (SYSUCC). In this study, we limited our analyses to cases with newly diagnosed HCC (C22.0 as per the International Statistical Classification of Diseases and Related Health Problems, Tenth Revision [ICD-10], codes), who had not received antitumor therapy at recruitment and who had data on both dietary intake and serum levels of copper and zinc. Cases were excluded if they had the following conditions: (1) other serious chronic diseases (including heart failure, liver failure, renal failure, physical or mental disability); (2) significant changes in dietary habits or lifestyles within the previous 5 years (i.e., changing from omnivorous diets to vegetarian diets, abstaining recruited from smoking and alcohol drinking). Healthy controls were recruited from the Guangzhou Nutrition and Health Study (GNHS). The GNHS, established in 2008, is a communitybased prospective cohort study. The inclusion and exclusion criteria for cases and controls were the same (except for the diagnosis of HCC). The recruitment and enrollment procedures of the cohorts have been previously described in detail.<sup>[20,21]</sup>

# Ethical approval

The GLCC and GNHS were approved by the Ethics

Committee of the School of Public Health at Sun Yatsen University (Nos.013[2017] and 048[2018]). Written informed consent was obtained from all participants.

Cases and controls were matched 1:1 according to sex and age ( $\pm$ 1 year), and a total of 434 case-control pairs were included in the present analysis. The study protocol was approved by the institutional review board of the School of Public Health at Sun Yat-sen University. Written informed consent was provided by each participant.

# **Dietary assessment**

A semi-quantitative 79-item food frequency questionnaire (FFQ) was used to assess dietary intake during the previous year before HCC diagnosis for cases or prior to the interview for controls. Each food item had five predefined qualitative responses, namely "daily," "weekly," "monthly," "annually," and "not eaten." Participants were asked to report the frequency and intake amount of each food consumed. Photographs of common food portion sizes were provided to help participants to quantify their consumption. Individual food intake was translated into grams per day. Daily dietary intakes of total energy, copper, zinc, and other nutrients were then calculated by multiplying the daily consumption of each food item by the nutrient content obtained from the China Food Composition Table 2009.<sup>[22]</sup> Copper and zinc intakes were adjusted for total energy intake (i.e., 1922 kcal/day for men and 1688 kcal/day for women) by the residual method.<sup>[23]</sup> The Healthy Eating Index 2015 (HEI-2015)<sup>[24]</sup> was also calculated to reflect the diet quality of participants, with higher scores indicating better adherence to dietary guidelines. The validity and reproducibility of the FFQ had been confirmed by six non-consecutive, 3-day dietary records at intervals of 2 months within 1 year and two FFQs administered 1 year apart among 61 healthy women recruited from Guangdong.<sup>[25]</sup>

# Laboratory assays

Blood samples were collected on the second day of the hospital admission for HCC cases and at the first followup for controls. Serum samples were separated, aliquoted, and stored in  $-80^{\circ}$ C freezers. To minimize technical variability, 100 µL of serum samples was assayed in batches at the KingMed Diagnostics Laboratory (Guangzhou, China) by blinded laboratory personnel between December 2016 and March 2017. Serum copper and zinc concentrations were quantified by using inductively coupled plasma mass spectrometry (Agilent 7700x ICP-MS spectrometer, Agilent Technologies, Germany). Masked, replicate, and quality control samples were interspersed among the samples. The intra-assay coefficients of variance for serum copper and zinc were 7.4% and 6.6%, respectively.

Routine laboratory parameters including seropositivity for hepatitis B surface antigen (HBsAg) were analyzed to a standardized protocol at the Clinical Laboratory of SYSUCC for cases and the KingMed Diagnostics Laboratory for controls.

## **Covariates**

The same structured questionnaire was used to collect the following information from cases and controls through face-to-face interviews: socio-demographic characteristics (i.e., sex, age); lifestyle habits (i.e., diet, smoking status, alcohol drinking status); and relevant diseases (i.e., a history of fatty liver disease, a history of diabetes). Smokers or alcohol drinkers were defined as participants who smoked at least one cigarette per day or drank alcohol at least once a week continuously for at least 6 months. Anthropometric measurements (i.e., height and body weight) were collected by a standard procedure. Body mass index (BMI, kg/m<sup>2</sup>) was then calculated by dividing weight (kg) by height squared (m<sup>2</sup>).

# Statistical analysis

The Cu/Zn ratio was computed by dividing copper levels by zinc levels. Paired student's t-test, Wilcoxon signed-rank test, or McNemar  $\chi^2$  test were used to compare differences in baseline characteristics between cases and controls. Spearman's rank correlation coefficients were estimated between dietary intake and serum levels of copper or zinc. Cases and controls were divided into three groups according to the tertiles of copper and zinc levels in the control group. Conditional logistic regression models were used to calculate the odds ratio (OR) and 95% confidence intervals (95% CIs). The trend test was performed by assigning the median value for each group as a continuous variable in the regression models. We also calculated OR and 95% CIs for one standard deviation increase (per-SD increase) in copper and zinc levels. Model 1 was adjusted for age and total energy intake; model 2 was additionally adjusted for BMI, smoking status, alcohol drinking status, and HEI-2015 based on model 1; and model 3 was additionally adjusted for HBsAg, history of fatty liver disease, and history of diabetes based on model 2. The interaction of copper and zinc with other established risk factors for HCC including sex, age, smoking status, alcohol drinking status, HBsAg, history of fatty liver disease, and history of diabetes was explored using the multiplicative scale. Stratified analyses were further performed using a binary logistic regression model for all factors except sex and age.

The statistical analyses were performed in SPSS (IBM, Armonk, NY, USA), version 26.0. All tests were twosided, and a value of P < 0.05 was considered statistically significant.

## **Results**

#### **Baseline characteristics**

A total of 434 case-control pairs were included in this study. Men accounted for 83.2% (361/434); the mean  $\pm$  standard deviation (SD) age was 59.8  $\pm$  7.6 years. The baseline characteristics of the case and control groups are shown in Table 1. The case group had a higher proportion of smokers, alcohol drinkers, HBsAg (+), and a history of fatty liver disease but lower HEI-2015 scores than the control group (all *P* <0.001). No significant differences in BMI, total energy intake, and the

proportion with a history of diabetes were found between cases and controls.

As shown in Table 2, dietary intakes of copper (median: 1.56 mg/day *vs.* 2.10 mg/day), zinc (median: 12.3 mg/day *vs.* 13.0 mg/day), and their ratio (median: 0.13 *vs.* 0.16) were significantly lower in cases than in controls (all P < 0.001). Regarding serum levels, serum copper (median: 982 µg/L *vs.* 874 µg/L) and Cu/Zn ratio (median: 1.13 *vs.* 0.79) were higher, while serum zinc was lower (median: 871 µg/L *vs.* 1098 µg/L) in cases than in controls (all P < 0.001). There was significantly inverse correlation between serum copper levels and dietary copper intake ( $r_{\text{spearman}} = -0.101$ , P = 0.003), whereas serum zinc levels were positively correlated with dietary zinc intake ( $r_{\text{spearman}} = 0.171$ , P < 0.001).

## Association of dietary copper and zinc intake with HCC risk

The associations of dietary intake of copper and zinc and their ratios with HCC risk are shown in Table 3. Higher dietary zinc intake was associated with a lower risk of HCC after adjustment for age and total energy intake (model 1: OR [95% CI] Tertile 3 vs. Tertile 1 = 0.39 [0.27, 0.56],  $P_{\text{trend}}$  <0.001). The inverse association remained significant after further adjustment for BMI, smoking, alcohol drinking, HEI-2015 (model 2: OR [95% CI] <sub>Tertile 3 vs. Tertile 1</sub> = 0.33 [0.17, 0.63],  $P_{\text{trend}} = 0.001$ ), as well as HBsAg, history of fatty liver disease, and history of diabetes (model 3: OR [95% CI] Tertile 3 vs. Tertile 1 = 0.26  $[0.08, 0.81], P_{trend} = 0.029$ ). Per-SD increase in dietary zinc intake was associated with a 35% decrease in the risk of HCC (OR [95% CI]  $_{per-SD increase} = 0.65$  [0.44, 0.96]). An inverse association with HCC risk was also observed for dietary copper intake and dietary Cu/Zn ratio in model 1, however, the association became non-significant after additional adjustment for other risk factors.

# Association of serum levels of copper and zinc with HCC risk

The associations of serum levels of copper and zinc and their ratio with HCC risk are shown in Table 4. After adjusting for potential confounders, higher serum zinc levels were associated with a lower risk of HCC (model 3: OR [95% CI] Tertile 3 *vs.* Tertile 1 = 0.02 [0.002, 0.13],  $P_{\text{trend}} < 0.001$ ). For each SD increase in serum zinc levels, the risk of HCC was decreased by 89% (model 3: OR [95% CI] per-SD increase = 0.11 [0.04, 0.30]). However, serum copper levels (model 3: OR [95% CI] per-SD increase = 2.05 [1.39, 3.03]) and serum Cu/Zn ratio (model 3: OR [95% CI] per-SD increase = 6.53 [2.52, 16.92]) were positively associated with HCC risk.

# Stratified analyses

Supplementary Tables 1 and 2, http://links.lww.com/CM9/ B614 show a significant interaction of dietary zinc intake and serum zinc levels with sex, respectively ( $P_{\text{interaction}} = 0.041$ ; 0.010). Higher zinc intake (OR<sub>Tertile 3 vs. Tertile 1</sub> = 0.54, 95% CI [0.32, 0.92],  $P_{\text{trend}} = 0.038$ ) or higher serum zinc levels (OR<sub>Tertile 3 vs. Tertile 1</sub> = 0.06, 95% CI [0.01, 0.29],  $P_{\text{trend}} < 0.001$ ) were associated with a lower risk of HCC in men, while no significant association was observed in

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| Table 1: Baseline characteristics of HCC cases and matched controls. |                        |                               |                    |          |  |  |
|--|------------------------|-------------------------------|--------------------|----------|--|--|
| Characteristics  | HCC cases<br>(n = 434) | Matched controls<br>(n = 434) | Statistics         | P values |  |  |
| Sex  |                        |                               | $0^{\dagger}$      | 1.000    |  |  |
| Women  | 73 (16.8)              | 73 (16.8)                     |                    |          |  |  |
| Men  | 361 (83.2)             | 361 (83.2)                    |                    |          |  |  |
| Age (years)  | $59.8 \pm 7.6$         | $59.9 \pm 7.5$                | $1.98^*$           | 0.048    |  |  |
| BMI (kg/m <sup>2</sup> )   | $22.8 \pm 3.2$         | $22.8 \pm 2.8$                | $0.35^{*}$         | 0.727    |  |  |
| Smoking  |                        |                               | $117.44^{\dagger}$ | < 0.001  |  |  |
| No   | 169 (38.9)             | 327 (75.3)                    |                    |          |  |  |
| Yes  | 265 (61.1)             | 107 (24.7)                    |                    |          |  |  |
| Alcohol drinking   |                        |                               | $73.36^{\dagger}$  | < 0.001  |  |  |
| No   | 253 (58.3)             | 367 (84.6)                    |                    |          |  |  |
| Yes  | 181 (41.7)             | 67 (15.4)                     |                    |          |  |  |
| BCLC stage   |                        |                               | -                  | -        |  |  |
| 0  | 41 (9.4)               | _                             |                    |          |  |  |
| А  | 165 (38.0)             | -                             |                    |          |  |  |
| В  | 42 (9.7)               | -                             |                    |          |  |  |
| С  | 186 (42.9)             | -                             |                    |          |  |  |
| D  | -                      | -                             |                    |          |  |  |
| HBsAg (+)  |                        |                               | $448.00^\dagger$   | < 0.001  |  |  |
| No   | 66 (15.2)              | 366 (84.3)                    |                    |          |  |  |
| Yes  | 368 (84.8)             | 51 (11.8)                     |                    |          |  |  |
| Missing  | -                      | 17 (3.9)                      |                    |          |  |  |
| History of fatty liver disease                                       |                        |                               | $14.34^{\dagger}$  | < 0.001  |  |  |
| No   | 339 (78.1)             | 375 (86.4)                    |                    |          |  |  |
| Yes  | 95 (21.9)              | 52 (12.0)                     |                    |          |  |  |
| Missing  | -                      | 7 (1.6)                       |                    |          |  |  |
| History of diabetes  |                        |                               | $1.63^{\dagger}$   | 0.257    |  |  |
| No   | 378 (87.1)             | 390 (89.9)                    |                    |          |  |  |
| Yes  | 56 (12.9)              | 44 (10.1)                     |                    |          |  |  |
| HEI-2015   | $55.5 \pm 7.0$         | $65.1 \pm 6.3$                | $21.01^{*}$        | < 0.001  |  |  |
| Total energy intake (kcal/day)                                       | $1912 \pm 610$         | $1951 \pm 585$                | $0.97^{*}$         | 0.334    |  |  |

Values are shown as n (%) or mean  $\pm$  standard deviation. <sup>\*</sup>t value, <sup>†</sup>chi-squared value. BCLC stage: Barcelona Clinic Liver Cancer stage; BMI: Body mass index; HBsAg: Hepatitis B surface antigen; HCC: Hepatocellular carcinoma; HEI-2015: Healthy Eating Index 2015. –: Not available.

Table 2: Dietary intakes and serum levels of copper and zinc of HCC cases and matched controls.

|                                  |                   | Matched controls  |            |          |
|----------------------------------|-------------------|-------------------|------------|----------|
| Characteristics                  | (n = 434)         | (n = 434)         | Statistics | P values |
| Dietary copper intakes (mg/day)* | 1.56 (1.21, 2.04) | 2.10 (1.63, 2.89) | -8.26      | < 0.001  |
| Dietary zinc intakes (mg/day)*   | 12.3 (11.1, 13.6) | 13.0 (12.0, 14.0) | -5.48      | < 0.001  |
| Dietary Cu/Zn ratio <sup>*</sup> | 0.13 (0.10, 0.16) | 0.16 (0.13, 0.23) | -7.15      | < 0.001  |
| Serum copper (µg/L)              | 982 (838, 1200)   | 874 (802, 955)    | -9.26      | < 0.001  |
| Serum zinc (µg/L)                | 871 (773, 980)    | 1098 (968, 1257)  | -14.85     | < 0.001  |
| Serum Cu/Zn ratio                | 1.13 (0.94, 1.42) | 0.79 (0.66, 0.93) | -15.67     | < 0.001  |

Values are shown as median (IQR). <sup>\*</sup>Intakes of copper and zinc were adjusted for sex-specific mean total energy intake (i.e., 1922 kcal/day for men and 1688 kcal/day for women) using the Residual Method. Cu/Zn: Copper/zinc; HCC: Hepatocellular carcinoma; IQR: Interquartile range.

women. There was no significant interaction between zinc and the remaining risk factors, including age, smoking, alcohol drinking, HBsAg (+), a history of fatty liver disease, and a history of diabetes (all  $P_{\text{interaction}} > 0.05$ ). In Supplementary Tables 3 and 4, http://links.lww.com/CM9/B614, there was no significant interaction between dietary intake or serum levels of copper and the stratified factors (all  $P_{\text{interaction}} > 0.05$ ).

# Table 3: OR (95% CI) of the risk of HCC by tertiles of dietary copper and zinc intakes.

| OR (95% CI)            |                           |                     |                   | OR (95% CI)               |                   |
|------------------------|---------------------------|---------------------|-------------------|---------------------------|-------------------|
| Characteristics        | Tertile 1                 | Tertile 2           | Tertile 3         | P <sub>trend</sub> values | per-SD increase   |
| Dietary copper intake* | ≤1.76 mg/day              | >1.76-2.52 mg/day   | >2.52 mg/day      |                           | per 2.02 mg/day   |
| Cases/control          | 273/144                   | 93/145              | 68/145            |                           | 434/434           |
| Model $1^{\dagger}$    | 1.00 (ref <sup>II</sup> ) | 0.31 (0.21, 0.45)   | 0.25 (0.17, 0.37) | < 0.001                   | 0.60 (0.49, 0.74) |
| Model 2 <sup>‡</sup>   | 1.00 (ref <sup>II</sup> ) | 1.08 (0.58, 2.00)   | 0.66 (0.36, 1.22) | 0.140                     | 0.87 (0.67, 1.12) |
| Model 3 <sup>§</sup>   | 1.00 (ref <sup>II</sup> ) | 1.15 (0.40, 3.29)   | 0.94 (0.37, 2.38) | 0.834                     | 0.86 (0.59, 1.26) |
| Dietary zinc intake*   | ≤12.35 mg/day             | >12.35-13.68 mg/day | >13.68 mg/day     |                           | per 1.81 mg/day   |
| Cases/controls         | 229/144                   | 102/145             | 103/145           |                           | 434/434           |
| Model $1^{\dagger}$    | 1.00 (ref <sup>II</sup> ) | 0.38 (0.27, 0.55)   | 0.39 (0.27, 0.56) | < 0.001                   | 0.70 (0.61, 0.81) |
| Model 2 <sup>‡</sup>   | 1.00 (ref <sup>II</sup> ) | 0.23 (0.11, 0.46)   | 0.33 (0.17, 0.63) | 0.001                     | 0.68 (0.54, 0.86) |
| Model 3 <sup>§</sup>   | 1.00 (ref <sup>II</sup> ) | 0.19 (0.06, 0.63)   | 0.26 (0.08, 0.81) | 0.029                     | 0.65 (0.44, 0.96) |
| Dietary Cu/Zn ratio*   | ≤0.14                     | >0.14-0.20          | >0.20             |                           | per 0.19          |
| Cases/controls         | 267/145                   | 99/144              | 68/145            |                           | 434/434           |
| Model $1^{\dagger}$    | 1.00 (ref <sup>II</sup> ) | 0.32 (0.22, 0.47)   | 0.25 (0.17, 0.37) | < 0.001                   | 0.70 (0.57, 0.86) |
| Model 2 <sup>‡</sup>   | 1.00 (ref <sup>ll</sup> ) | 1.09 (0.59, 2.05)   | 0.65 (0.35, 1.23) | 0.110                     | 0.95 (0.74, 1.23) |
| Model 3 <sup>§</sup>   | 1.00 (ref <sup>ll</sup> ) | 0.86 (0.30, 2.47)   | 0.88 (0.36, 2.16) | 0.799                     | 0.97 (0.68, 1.39) |

<sup>\*</sup>Intakes of copper and zinc were adjusted for total energy intake (i.e., 1922 kcal/day for men and 1688 kcal/day for women) using the residual method. <sup>†</sup>Model 1: Adjusted for age and total energy intake. <sup>‡</sup>Model 2: Adjusted for covariates in model 1 + BMI, smoking status, alcohol drinking status, and HEI-2015. <sup>§</sup>Model 3: Adjusted for covariates in model 2 + HBsAg, history of fatty liver disease, and history of diabetes. <sup>ll</sup>"ref" refers to the inclusion of copper and zinc indicators as categorical variables in the model, using Tertile 1 as a reference to calculate the OR (95% CI) for Tertile 2 and Tertile 3, respectively. BMI: Body mass index; CI: Confidence interval; Cu/Zn: Copper/zinc; HBsAg: Hepatitis B surface antigen; HCC: Hepatocellular carcinoma; HEI-2015: Healthy Eating Index 2015; OR: Odds ratio; SD: Standard deviation.

# Table 4: OR (95% CI) of the risk of HCC by tertiles of serum copper and zinc levels.

|                      |                          | OR (95% CI)        |                      |                           | OR (95% CI)        |  |
|----------------------|--------------------------|--------------------|----------------------|---------------------------|--------------------|--|
| Characteristics      | Tertile 1                | Tertile 2          | Tertile 3            | P <sub>trend</sub> values | per-SD increase    |  |
| Serum copper levels  | ≤825 µg/L                | >825–921 µg/L      | >921 µg/L            |                           | per 122 µg/L       |  |
| Cases/controls       | 100/144                  | 74/146             | 260/144              |                           | 434/434            |  |
| Model 1 <sup>*</sup> | 1.00 (ref <sup>§</sup> ) | 0.76 (0.51, 1.14)  | 2.51 (1.78, 3.55)    | < 0.001                   | 1.56 (1.40, 1.73)  |  |
| Model 2 <sup>†</sup> | 1.00 (ref <sup>§</sup> ) | 0.75 (0.38, 1.47)  | 1.48 (0.82, 2.68)    | 0.100                     | 1.33 (1.14, 1.55)  |  |
| Model 3 <sup>‡</sup> | 1.00 (ref <sup>§</sup> ) | 0.68 (0.19, 2.38)  | 2.57 (0.90, 7.32)    | 0.020                     | 2.05 (1.39, 3.03)  |  |
| Serum zinc levels    | ≤1020 µg/L               | >1020–1186 µg/L    | >1186 µg/L           |                           | per 404 µg/L       |  |
| Cases/controls       | 361/144                  | 57/146             | 16/144               |                           | 434/434            |  |
| Model 1 <sup>*</sup> | 1.00 (ref <sup>§</sup> ) | 0.10 (0.06, 0.17)  | 0.03 (0.01, 0.06)    | < 0.001                   | 0.05 (0.03, 0.58)  |  |
| Model 2 <sup>†</sup> | 1.00 (ref <sup>§</sup> ) | 0.16 (0.07, 0.37)  | 0.03 (0.01, 0.09)    | < 0.001                   | 0.09 (0.04, 0.20)  |  |
| Model 3 <sup>‡</sup> | 1.00 (ref <sup>§</sup> ) | 0.43 (0.13, 1.44)  | 0.02 (0.002, 0.13)   | < 0.001                   | 0.11 (0.04, 0.30)  |  |
| Serum Cu/Zn ratio    | ≤0.71                    | >0.71-0.87         | >0.87                |                           | per 0.21           |  |
| Cases/controls       | 17/144                   | 56/145             | 361/145              |                           | 434/434            |  |
| Model 1 <sup>*</sup> | 1.00 (ref <sup>§</sup> ) | 3.04 (1.50, 6.14)  | 24.86 (12.34, 50.07) | < 0.001                   | 4.05 (3.03, 5.40)  |  |
| Model 2 <sup>†</sup> | 1.00 (ref <sup>§</sup> ) | 7.27 (2.43, 21.80) | 26.06 (9.16, 74.12)  | < 0.001                   | 3.97 (2.56, 6.16)  |  |
| Model 3 <sup>‡</sup> | 1.00 (ref <sup>§</sup> ) | 9.17 (1.41, 59.65) | 65.42 (8.90, 480.83) | < 0.001                   | 6.53 (2.52, 16.92) |  |

\*Model 1: Adjusted for age and total energy intake. <sup>†</sup>Model 2: Adjusted for covariates in model 1 + BMI, smoking status, alcohol drinking status, and HEI-2015. <sup>‡</sup>Model 3: Adjusted for covariates in model 2 + HBsAg, history of fatty liver disease, and history of diabetes. <sup>§</sup>"ref" refers to the inclusion of copper and zinc indicators as categorical covariates in the model, using Tertile 1 as a reference to calculate the OR (95% CI) for Tertile 2 and Tertile 3, respectively. BMI: Body mass index; CI: Confidence interval; Cu/Zn: Copper/zinc; HBsAg: Hepatitis B surface antigen; HEI-2015: Healthy Eating Index 2015; HCC: Hepatocellular carcinoma; OR: Odds ratio; SD: Standard deviation.

## Discussion

In this 1:1 matched case-control study, we found that higher dietary intake and serum levels of zinc were associated with a lower risk of HCC, and the inverse association between zinc and HCC risk was only predominant in men. In addition, serum levels of copper and serum Cu/Zn ratio were positively associated with HCC risk, while dietary copper intake and dietary Cu/Zn ratio were not associated with HCC risk in model 3.

In line with our findings, a nested case-control study conducted within the EPIC showed an inverse association between serum zinc levels and HCC risk (n = 106).<sup>[18]</sup> However, the results from the Shanghai Women's and Men's Study observed no significant association between dietary zinc intake and liver cancer risk,<sup>[19]</sup> which were inconsistent with our study. In this study, we found both dietary zinc intake and serum zinc levels were inversely associated with HCC risk, even after additional adjustment for other risk factors for HCC including hepatitis B virus infection. Our findings are supported by several studies with zinc supplementation in patients with chronic liver diseases (CLDs). Randomized controlled trials have showed that zinc supplementation significantly increased serum zinc levels and decreased levels of liver enzymes and oxidative stress in patients with non-alcoholic fatty liver disease.<sup>[26,27]</sup> A retrospective study reported that zinc supplementation maintained liver function and reduced the incidence of HCC in 267 patients with CLDs, particularly in those with post-treatment serum zinc levels  $\geq 700 \ \mu g/L.^{[28]}$ 

In stratified analyses, we also found that the inverse association between dietary intake and serum levels of zinc and the risk of HCC was only pronounced in men but not in women. The exact mechanism for the sex differences remains to be elucidated. Animal studies have shown that sexually dimorphic liver gene expression is modulated by regulator of sex-limitation 1 (Rsl1), one of the Krüppel-associated box zinc finger proteins (KRAB-ZFPs), and thus contribute to the hepatic sex differences in liver disease pathogenesis.<sup>[29,30]</sup> This may partially explain why zinc is a protective factor for HCC in men but not in women.

HBV infection and alcohol drinking are well-established risk factors for HCC. Although the case group had a higher proportion of participants with HBsAg (+) and alcohol drinkers than the control group, the association between serum levels and dietary intake of zinc and HCC risk was not modified by HBV infection and alcohol drinking status in our study. However, the limited number of HBsAg(-) participants and non-alcohol drinkers preclude drawing a firm conclusion.

Copper intake were not observed to be associated with the risk of HCC in this study, which was in accord with the findings from the Shanghai Women's and Men's Study.<sup>[19]</sup> However, we found a positive association between serum copper levels and HCC risk. Given the fact that dietary copper levels were lower while serum copper levels were higher in HCC cases than in healthy controls in our study, we speculate that the increase in serum copper levels in patients with HCC is likely due to the redistribution of hepatic copper caused by the pathological inflammation in the liver rather than excessive copper intake.<sup>[31]</sup> Redistribution of hepatic copper includes excessive activation of ceruloplasmin and release of hepatocyte necrosis, which leads to excessive copper release from the liver into the blood.<sup>[32,33]</sup> Therefore, high serum copper levels may be a consequence of abnormal liver function in patients with HCC, rather than an independent risk factor for HCC. Consistent with the assumption, a nested case-control study within the EPIC showed null association between serum copper levels and HCC risk.<sup>[18]</sup> Nevertheless, our previous study found that higher serum copper levels were associated with a poorer prognosis in HCC patients,<sup>[20]</sup> which may suggest a more important role of copper in the progression of HCC.

Previous studies have found a positive correlation between serum Cu/Zn ratio and the overall severity of viral hepatitis and liver cirrhosis,<sup>[16,34]</sup> and Cu/Zn ratio may serve as a reference indicator to predict the incidence of HCC.<sup>[35]</sup> We also observed a positive association between serum Cu/Zn ratio and HCC risk, but in consideration of copper not being an independent risk factor for HCC, the positive association may be mainly attributed to zinc alone.

Zinc, a core component of Cu/Zn-SOD, is a key factor in reducing cellular oxidative stress and inhibiting cancer initiation.<sup>[13,36]</sup> It was found that zinc supplementation in healthy elderly subjects increased their serum zinc levels and significantly increased the antioxidant activity of circulating Cu/Zn-SOD, while zinc deficiency at baseline was associated with lower activity of circulating Cu/Zn-SOD.<sup>[37,38]</sup> Zinc deficiency also decreases glutathione levels, which in turn increases hepatic genotoxicity.<sup>[39]</sup> Zinc is particularly important for cellular and humoral immunity in humans. Zinc is involved in the proliferation and differentiation of immune cells, as well as in the regulation of a series of cytokines such as interleukins, interferons, and tumor necrosis factors.[15,40] Zinc deficiency decreases the quantity of granulocytes and natural killer cells as well as the phagocytic capacity of macrophages and also affects the quantity and function of T and B lymphocytes in the thymus and marrow, thereby increasing the risk of various cancers.[41,42] Meanwhile, zinc binds to the oncogene p53 in the form of zinc-binding protein, performing the functions of monitoring cell cycle and DNA repair and inhibiting excessive cell proliferation. Zinc deficiency decreases the level of binding to p53, causes misfolding of p53, and impairs DNA repair functions, thus promoting cancer development.<sup>[39,43]</sup> In addition, zinc is a physiological trigger for metallothioneins (MTs), which have a high affinity for heavy metals and prevent oxidative stress and DNA damage caused by heavy metals.<sup>[36]</sup> Thus, the mechanistic aspect supports our view that high serum zinc levels contribute to the prevention of HCC.

Strengths of this study are that, first, we measured zinc and copper in both diet and serum in the same population, which allows us to explore the link between dietary intake and serum levels. Second, the sample size of HCC cases is relatively large and all HCC cases were newly diagnosed and previously untreated to avoid the confounding effect of anti-tumor therapy on dietary intake and serum levels of copper and zinc. Third, cases and controls were matched by sex and age ( $\pm 1$  year), which increases the statistical power. Fourth, dietary information was collected by the same FFQ, and serum levels of zinc and copper were quantified using ICP-MS at the same laboratory concurrently for cases and controls, which increases comparability of exposures between the groups. Finally, a wide range of covariates was collected, including sex, age, total energy intake, BMI, smoking status, alcohol drinking status, HEI-2015, HBsAg seropositivity, the history of fatty liver disease, and diabetes.

Nevertheless, several limitations of this study are needed to be noticed. First, dietary information within the previous year before HCC diagnosis and blood samples from cases were collected at the diagnosis of HCC. Therefore, the possibility of reverse causality cannot be ruled out due to the case-control design. Second, we have only one measurement of dietary intake and serum levels of copper and zinc, which cannot reflect the long-term nutritional status of copper and zinc in the study population. Third, the FFQ used in this study were only validated in the healthy subjects, and recall bias is inevitable. To minimize the bias, we only included participants whose dietary habits without marked change in the past 5 years and patients with newly diagnosed HCC. In addition, photographs were provided to help participants assess food consumption. Fourth, hepatitis C virus (HCV) infection was not considered in the analysis. However, since the prevalence of HCV infection in China is low,<sup>[44]</sup> its impact on the results may be negligible. Lastly, the study population was all of Asians living in southern China, and caution should be taken when generalizing the findings to different regions and ethnicities.

In summary, higher dietary intakes and serum levels of zinc were associated with a lower risk of HCC, especially in men, suggesting that zinc may protect against HCC. Although serum copper was significantly higher in HCC than in healthy controls, copper may not be an independent risk factor for HCC, as it is likely due to the redistribution of hepatic copper caused by the pathological inflammation of the liver. Large-scale prospective cohort studies or randomized controlled studies are warranted.

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# **Conflicts of interest**

None.

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