

## Effects of silymarin use on liver enzymes and metabolic factors in metabolic dysfunction-associated steatotic liver disease: a systematic review and meta-analysis

Adnan Malik MD<sup>1</sup>, Muhammad Malik MRCP<sup>2</sup>, Shahbaz Qureshi MD<sup>1</sup>

### ABSTRACT

**BACKGROUND:** Fatty liver disease comprises a wide range of related liver disorders affecting mainly people who drink no or minimal amounts of alcohol. Silymarin is a member of the *Carduus marianum* family that has been used for centuries to treat different diseases. There is little evidence supporting its efficacy in humans. **OBJECTIVES:** To evaluate the effects of Silymarin in patients with non alcoholic fatty liver disease (NAFLD) or recently renamed metabolic dysfunction-associated steatotic liver disease (MASLD). **METHODS:** We searched PubMed, SCOPUS, Web of Science, and Cochrane Library for relevant clinical trials assessing the use of silymarin in patients with NAFLD. A risk of bias assessment was performed using Cochrane's risk of bias tool. We included the following outcomes: alanine aminotransferase (ALT), aspartate aminotransferase (AST),  $\gamma$ -glutamyl transferase (GGT), total cholesterol (TC), triglyceride (TG), high-density lipoprotein (HDL) (mg/dL), degree of fibrosis resolution, low-density lipoprotein (LDL), and HOMA-IR. We analyzed continuous data using mean difference (MD) and relative 95% confidence interval (CI). **RESULTS:** We included nine clinical trials. We found that silymarin significantly reduced the levels of ALT (MD= -17.12 [-28.81, -4.43]), (P < 0.004), AST (MD= -12.56 [-19.02, -6.10]), (P < 0.0001) and TG (MD = -22.60 [-23.83, -21.38]) (p < 0.00001). It also improved HDL (MD= 2.13 [1.60, 2.66]), (P < 0.01)). There was no significant difference regarding GGT (P=0.07), TC (P= 0.52), LDL (P= 0.06), HOMA-IR (P= 0.06) and BMI (p=0.1). One study reported significant improvement in the degree of fibrosis (P = 0.023). **CONCLUSION:** Silymarin treatment significantly reduces biochemical and transaminase levels in patients with MASLD.

**KEYWORDS:** MASLD; NAFLD; silymarin

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## INTRODUCTION

Non-alcoholic fatty liver disease (NAFLD) is when excess fat is stored in the liver (1). This fat buildup is not caused by heavy alcohol use (2). NAFLD is the most common cause of chronic liver disease as it affects up to 30% of the general population (3). NAFLD is commonly associated with metabolic syndrome, obesity, diabetes, and hyperlipidemia (4).

Recently, the experts advocated renaming non-alcoholic fatty liver disease (NAFLD) metabolic dysfunction-associated steatotic liver disease (MASLD) (5). Nearly 80% of patients with metabolic syndrome have MASLD (6). MASLD is classified into two types: simple fatty liver and non-alcoholic steatohepatitis (NASH) (7). People only develop one type of MASLD, although sometimes people with one form are later diagnosed with the other form of MASLD (8). NASH can lead to liver cirrhosis and liver cancer (9). MASLD is one of the cardiovascular risk factors, which also contributes to mortality in these patients (10). Both environmental and genetic factors are playing an important role in the development of MASLD and its progression. First-degree relatives of patients with MASLD are at higher risk than the general population (11). MASLD can affect people of any age, including children. About 10% of United States children ages 2 to 19 years have MASLD. However, MASLD may be found in people of any age (12).

Laboratory liver tests are defined as tests used in the evaluation of patients with hepatic dysfunction (13). ALT is found in the kidney, heart, muscle, and at a higher concentration in the liver compared with other tissues of the body (14). Marked elevations of ALT levels are observed most often in diseases that primarily affect hepatocytes, such as viral hepatitis, ischemic liver injury, and toxin-induced liver damage (15). AST is found in the highest concentration in the heart compared with other tissues of the body (14). Elevated mitochondrial AST is seen in extensive tissue necrosis during myocardial infarction and also in chronic liver diseases (16).

Silymarin is a member of the *Carduus marianum* family that has been used for centuries for the treatment of different diseases (17). Silymarin has been used as a hepatoprotective drug. Studies in rodents have confirmed that silymarin has very low toxicity, which supports its history as a safe medication in hepatic diseases (18). Silymarin has

antifibrotic, immunomodulating, anti-inflammatory effects, and anti-oxidant properties by scavenging free radicals and increasing the glutathione concentrations (19). Silymarin can be used in hepatitis and hepatic cirrhosis, and mushroom poisoning because of its action (20).

There is no sufficient data regarding the effect of Silymarin on MASLD. Therefore, we performed this systemic review and meta-analysis to evaluate the effect of silymarin on biopsy-proven MASLD.

## METHODS

This meta-analysis was performed following the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) (21) and the guidelines reported in the Cochrane Handbook for Systematic Reviews of Interventions (22).

### Literature search

We searched four databases: Web of Science, SCOPUS, Cochrane CENTRAL, and PubMed, from inception until October 2022. We followed this search strategy with no restriction on time or languages: (silymarin OR marianum OR legalon OR silibinin OR silybin) AND (steatohepatitis).

### Eligibility criteria

We included all the studies that have the following criteria: (i) **Population:** patients with non-alcoholic fatty liver diseases, (ii) **Intervention:** Silymarin regardless of the dose and the mode of administration, (iii) **Comparator:** any control, (iv) **Outcomes:** Alanine transferase (IU/L) (ALT), aspartate transferase (AST) (IU/L), and  $\gamma$ -glutamyl transferase (GGT) (IU/L) as primary outcomes. The secondary outcomes were body mass index (BMI) (kg/m<sup>2</sup>), total cholesterol (TC) (mg/dL), triglyceride (TG) (mg/dL), high-density lipoprotein (HDL) (mg/dL), low-density lipoprotein (LDL) (mg/dL), and HOMA-IR. (v) **Study design:** We included clinical trials characteristics mentioned above. Our exclusion criteria were: (i) studies that did not report data or measures for our selected outcomes, or (ii) that with no available full-text.

### Screening of results

We exported the results of the search into Endnote X8.0.1 (Build 1044), with the removal of duplicates automatically by computer. After that, we screened the studies manually in two steps; first, title and abstract screening, then full-text screening for the preliminary included studies in the first step.

## Data extraction and analysis

After the screening step, we extracted the data from the selected studies and categorized the data into three main groups: (1) baseline and demographic data of patients in each study, including age, sample size, gender, BMI, levels of ALT, AST, GGT, TC, TG, HDL, LDL, and HOMA-IR score. (2) Data for analysis including outcome values of alanine aminotransferase (ALT) (IU/L), aspartate aminotransferase (AST) (IU/L),  $\gamma$ -glutamyl transferase (GGT) (IU/L), body mass index (BMI) ( $\text{kg}/\text{m}^2$ ), total cholesterol (mg/dL), triglyceride (mg/dL), HDL (mg/dL), LDL (mg/dL), degree of fibrosis resolution and HOMA-IR. In addition to the previous two categories, we extracted the data about the seven domains assessing the risk of bias according to Cochrane's risk of bias (23).

## Data analysis

We used Review Manager Software (RevMan 5.4.1) to perform our analysis. We conducted the analysis of continuous outcomes using mean difference (MD) and standard deviations (SD), and relative 95% confidence interval (CI) under the Inverse variance method.

## Quality assessment

We evaluated the quality of this systematic review and meta-analysis using the Grading of Recommendations Assessment, Development, and Evaluation (GRADE) guidelines (24). According to the Cochrane risk of bias (ROB) tool for clinical trials, we performed the risk of bias (ROB) for the included studies. The tool depends on the following domains for assessment of the risk of bias: (1) proper randomization, (2) blinding allocation of the included patients into each group, (3) blinding of patients only (single-blinding), blinding of both personnel and participants (double-blinding), or not blinding at all, (4) attrition bias, (5) selection bias (outcomes reported matches with that of the protocol or not), (6) awareness of the outcome assessor (whether blinded or not), (7) other bias. We assessed the total risk of bias for the studies as well.

## RESULTS

### Summary of included studies

Figure 1 shows a PRISMA flow diagram of our literature search. In our study, we performed an analysis of 820 patients from nine studies (25–33). A total of 412 patients were allocated to receive

silymarin, and 408 patients entered the control group. The mean age of the recipient in the treatment group was  $44.8 \pm 8.38$  years, while that of the control group was  $44.2 \pm 8.09$ . The mean body mass index of the patients in the silymarin group was  $27.888 \pm 1.955$ , while that of the control group was  $28.27 \pm 2.815$ . Table 1–3 shows a detailed summary of the included participants, their demographic data, liver enzymes and complete lipid profile.

### Results of risk of bias assessment

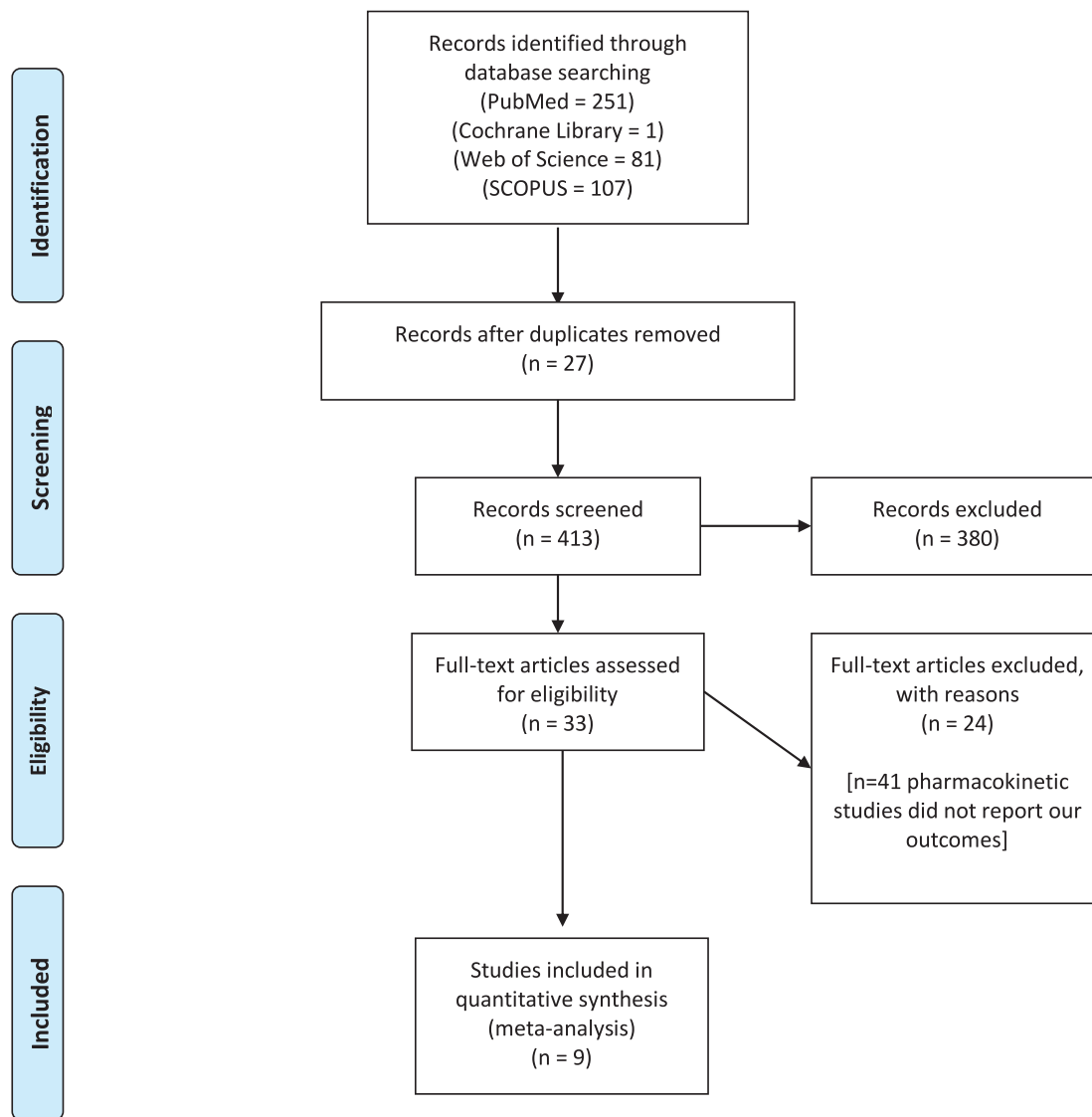
The result of the risk of bias assessments yielded an overall low risk of bias, according to Cochrane's tool. Regarding randomization, all studies were at low risk of randomization, except Kheong et al. and Masoodi et al. (25,30) were non-randomized trials. As for the allocation concealment, six studies (24,26–29,32) reported adequate allocation concealment; therefore, they were put to a low risk of bias. Two studies (25,31) did not report enough data about allocation concealment, thus put to an unclear risk of bias. One study (30) reported no allocation concealment. The majority of the included studies were blinded, and four studies (25,30–32) were not blinded to the participants and personnel. Four studies (24,26,27,29) were at low risk of blinding of outcome assessment. Two studies (25,28) were not blinded to outcome assessment, and three studies (30–32) did not report enough data about blinding of outcome assessment, thus put to an unclear risk of bias. The remaining domains of the Cochrane tool were all at low risk of bias. A summarized illustration of the risk of bias of included trials, Figure 2.

### Analysis of Outcomes

#### ALT

All studies (24–32) reported the ALT. The overall mean difference showed that there was a significant difference between both groups (MD =  $-17.12$  [ $-28.81, -4.43$ ]), ( $P < 0.004$ ). Pooled analysis was heterogeneous ( $P < 0.00001$ );  $I^2 = 99\%$  as shown in Figure 3. We solved the heterogeneity by performing a subgroup analysis according to the duration of drug administration.

The first subgroup (6 months) included five studies (26–28,30,32), the mean difference showed a significant difference between both groups (MD =  $-30.12$  [ $-36.42, -23.77$ ]), ( $p < 0.01$ ). Pooled analysis was homogeneous ( $p = 0.16$ );  $I^2 = 38\%$  as shown in Figure 3.



**Figure 1:** PRISMA flow diagram of our literature search

The second subgroup (2 months) included two studies (25,31) the mean difference showed that there was a significant difference between both groups (MD = -11.19 [-12.96, -9.42]), ( $p < 0.01$ ). Pooled analysis was homogeneous ( $p = 0.63$ );  $I^2 = 0\%$  as shown in Figure 3C. The third subgroup (3 months) had two studies (24,29) the mean difference showed that there was no significant difference between both groups (MD = 1.10 [-15.57, 17.76]) ( $p < 0.90$ ). Pooled analysis was homogeneous ( $p = 0.1$ );  $I^2 = 64\%$ , as shown in Figure 3.

### AST

All studies (24–32) reported the AST. The overall mean difference showed that there was a significant difference between both groups (MD= -12.56 [-19.02, -6.10]), ( $P < 0.0001$ ). Pooled analysis was

heterogeneous ( $p < 0.00001$ );  $I^2 = 96\%$  as shown in Figure 4. We solved the heterogeneity by subgroup analysis according to the duration of drug administration. The first subgroup (6 months) contains five studies (26–28,30,32), the mean difference showed that there was a significant difference between both groups (MD = -21.90 [-25.62, -18.17]), ( $p < 0.00001$ ). Pooled analysis was homogeneous ( $p = 0.25$ );  $I^2 = 26\%$  as shown in Figure 4. The second subgroup (2 months) contains two studies (25,31), the mean difference showed that there was a significant difference between both groups (MD = -8.15 [-9.89, -6.40]), ( $p < 0.00001$ ). Pooled analysis was homogeneous ( $p = 0.33$ );  $I^2 = 0\%$  as shown in Figure 4. The third subgroup (3 months) contains two studies (24,29). The mean difference showed that there was no significant difference

**Table 1:** Shows a detailed summary of the included participants, their demographic data. BMI, age, and male number

Study ID	BMI (kg/m <sup>2</sup> )				Age (years)				Male	
	Mean		SD		Mean		SD		Number	
	Silymarin		Control		Silymarin		Control		Silymarin	Control
Wah Kheong 2017	30	0.4	31	4.6	49.6	12.7	50.1	10.2	24	22
Cacciapuoti 2013	26.7	1.67	26.7	1.67	44	3.2	44	3.2	40	40
Hajiaghamohammadi 2012	27.44	1.65	27.44	1.65	32.62	6.4	32.62	6.4	14	14
Hashemi 2009	26.75	2.65	27.8	3.75	39.28	11.117	39	10.7	28	29
Masoodi 2013	29.04	3.66	29.18	3.32	48.42	6.75	48.32	5.45	31	31
Jelodar 2015	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR
Solhi 2014	27.4	1.7	27.5	1.9	43.6	8.3	39.36	10.5	19	19
Frățilă 2019	NR	NR	NR	NR	NR	10.2	56.08	10.2	25	25
Taghvaei	NR	NR	NR	NR	NR	NR	NR	NR	14	13

Data are represented as mean (SD). BMI = body mass index; NR = not reported.

**Table 2:** Shows a detailed summary of the included participant's ALT, AST, GGT levels

Study ID	Alanine aminotransferase (ALT) (IU/L)				Aspartate aminotransferase (AST) (IU/L)				GGT (U/L)			
	Mean		SD		Mean		SD		Mean		SD	
	Silymarin		Control		Silymarin		Control		Silymarin		Control	
Wah Kheong 2017	93.3	53.3	80.3	54.8	55.3	31.1	49	28.9	98.7	52.6	87.7	70.4
Cacciapuoti 2013	109.48	4.4	109.48	4.4	72.39	8.4	72.39	8.4	45.51	1.2	45.51	1.2
Hajiaghamohammadi 2012	78.73	19.71	78.37	19.71	56	11.07	56	11.07	NR	NR	NR	NR
Hashemi 2009	113.54	50.92	104.54	41.82	71.42	66.5	73.02	39.62	NR	NR	NR	NR
Masoodi 2013	84.06	6.65	74.48	4.45	71.94	4.56	62.94	4.45	NR	NR	NR	NR
Jelodar 2015	52.12	6.46	52.12	6.46	48.73	7.09	48.73	7.09	NR	NR	NR	NR
Solhi 2014	91.3	21.3	84.6	23.3	62.8	10.5	70.4	18.9	NR	NR	NR	NR
Frățilă 2019	67.35	58.83	67.35	58.83	75.22	58.13	75.22	58.13	150.75	113.78	150.75	11378
Taghvaei	69.9	27.06	88.55	23.58	56.67	30.5	60.2	27.8	NR	NR	NR	NR

Data are represented as mean (SD). ALT = alanine aminotransferase, AST = aspartate aminotransferase, NR = not reported

between both groups (MD = -3.07 [-13.22, -7.08]), ( $p = 0.55$ ). Pooled analysis was heterogeneous ( $p = 0.08$ );  $I^2 = 68\%$ , as shown in [Figure 4](#).

### GGT

Three studies (29,30,32) reported the GGT. The overall mean difference showed that there was no

significant difference between both groups (MD = -(MD= -13.20 [-34.65, - 80]), ( $P < 0.23$ ). Pooled analysis was heterogeneous ( $p = 0.06$ ;  $I^2 = 65\%$ ) as shown in [Figure 5A](#). We solved the heterogeneity by the exclusion of [Wah Kheong 2017](#) et al. (29) ( $p = 0.18$ );  $I^2 = 43\%$ . The pooled analysis after exclusion showed that there was no significant

**Table 3:** Shows a detailed summary of the included participant's TC, TG, HDL, and LDL

Study ID	Total cholesterol (TC) (mg/dL)				Triglyceride (TG) (mg/dL)				HDL (mg/dL)				LDL (mg/dL)			
	mean	SD	mean	SD	mean	SD	MEAN	SD	mean	SD	mean	SD	mean	SD	mean	SD
	silymarin		control		silymarin		control		silymarin		control		silymarin		control	
<b>Wah Kheong 2017</b>	204.95	42.54	181.75	34.8	159.43	35.43	124	53.14	46.4	11.6	46.4	11.6	123.74	46.4	112.14	34.8
<b>Cacciapuoti 2013</b>	205.7	9.3	205.7	9.3	178.4	4.1	178.4	4.1	43.6	2.1	43.6	2.1	157.4	4.3	157.4	4.3
<b>Hajjaghahammadi 2012</b>	191.68	33.81	191.68	33.81	254.18	56.04	254.18	56.04	NR	NR	NR	NR	NR	NR	NR	NR
<b>Hashemi 2009</b>	235.18	59.25	216.18	52.12	281.48	116.66	261.32	102.02	40.58	5.19	41.06	6	163.3	49.69	135.36	47.45
<b>Masoodi 2013</b>	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR
<b>Jelodar 2015</b>	242.88	16.2	242.88	16.2	232.75	48.09	232.75	48.09	NR	NR	NR	NR	NR	NR	NR	NR
<b>Solhi 2014</b>	195.7	34.4	193	35.7	252.2	52.8	248.4	53.2	NR	NR	NR	NR	NR	NR	NR	NR
<b>Frățilă 2019</b>	181.75	45.14	181.75	45.14	125	56.35	125	56.35	NR	NR	NR	NR	NR	NR	NR	NR
<b>Taghvaei</b>	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR

Data are represented as mean(SD). HDL = high density lipoprotein; LDL = low density lipoprotein; TC = total cholesterol; TG = triglyceride; NR = not reported

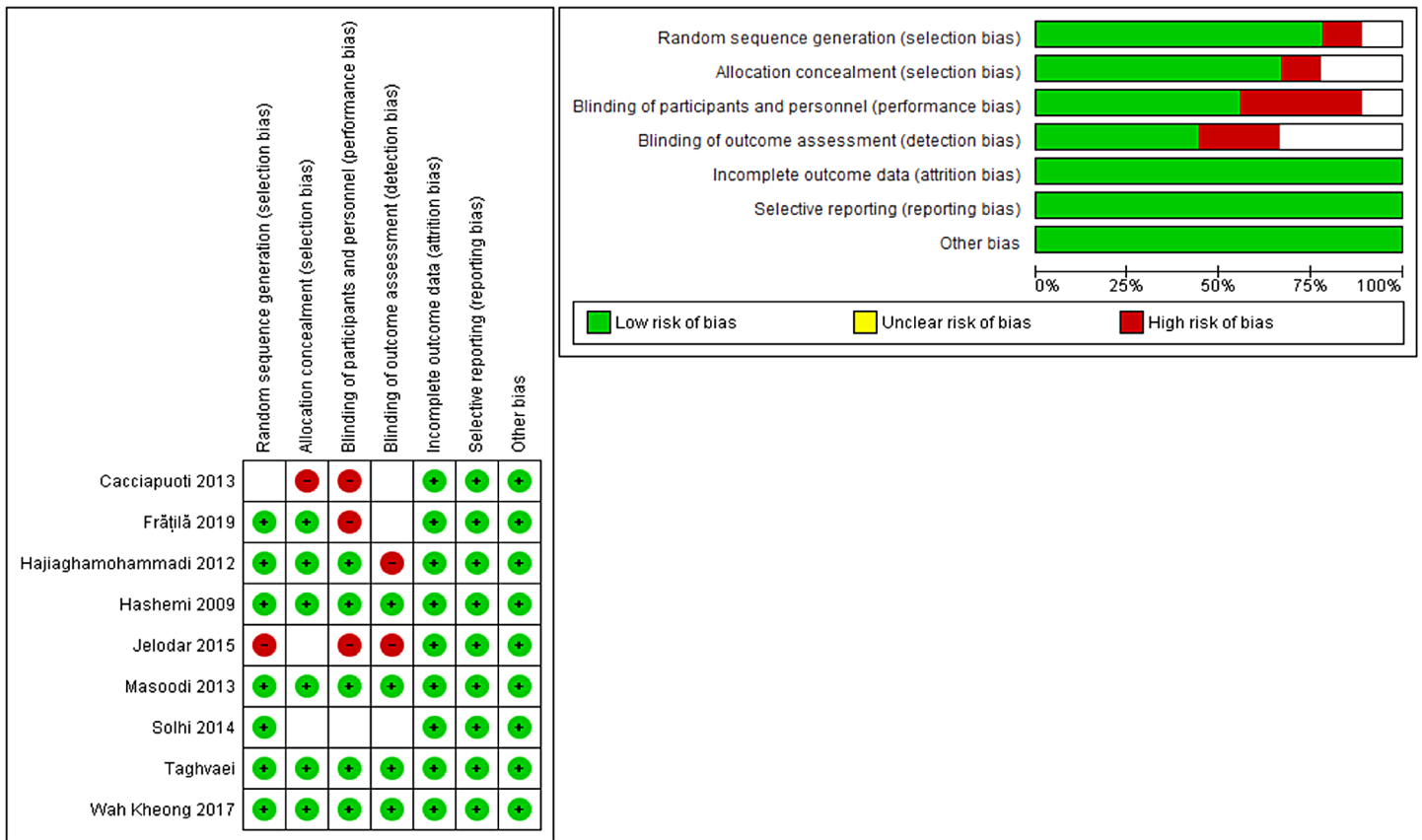


Figure 2: Risk of bias graph and summary

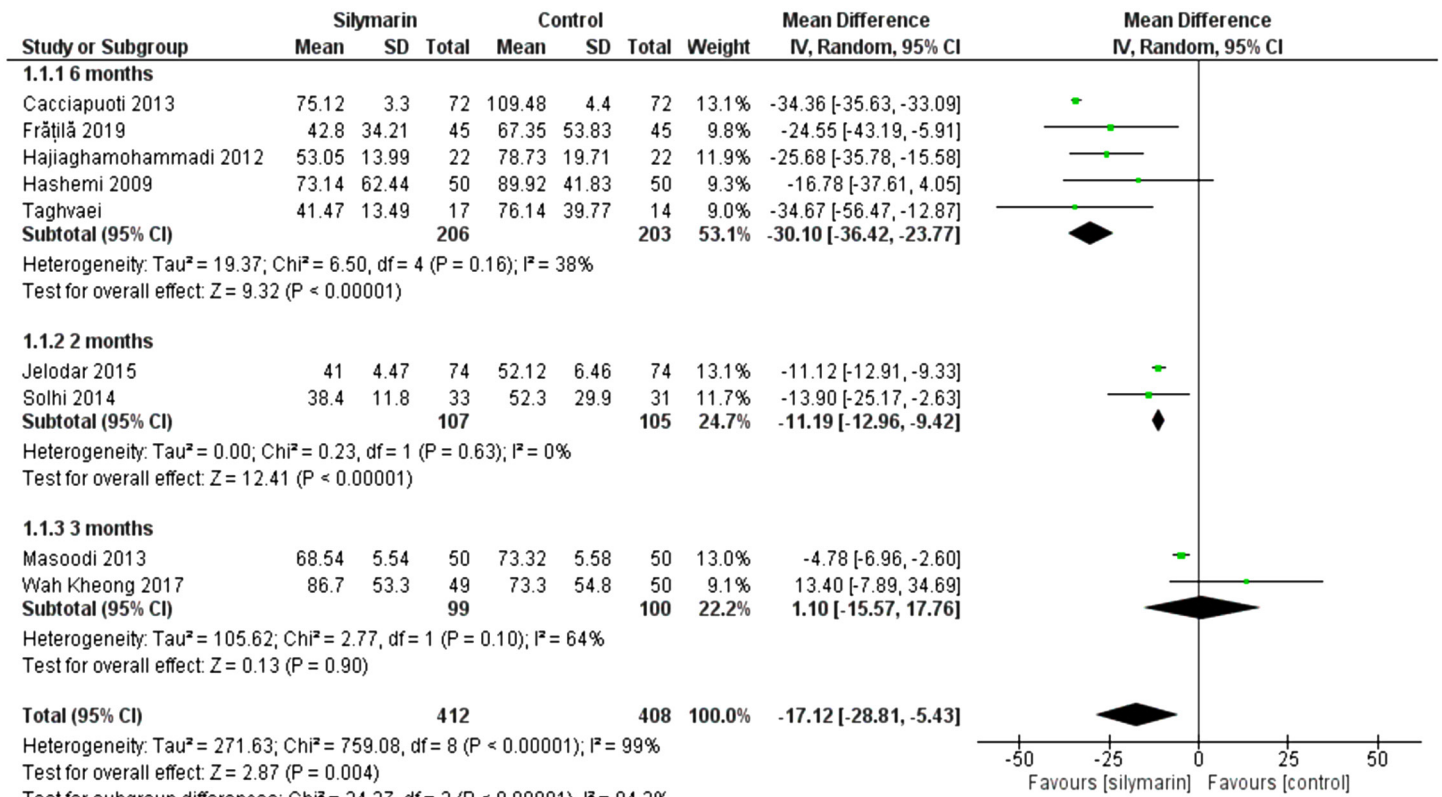


Figure 3: Analysis of ALT outcome

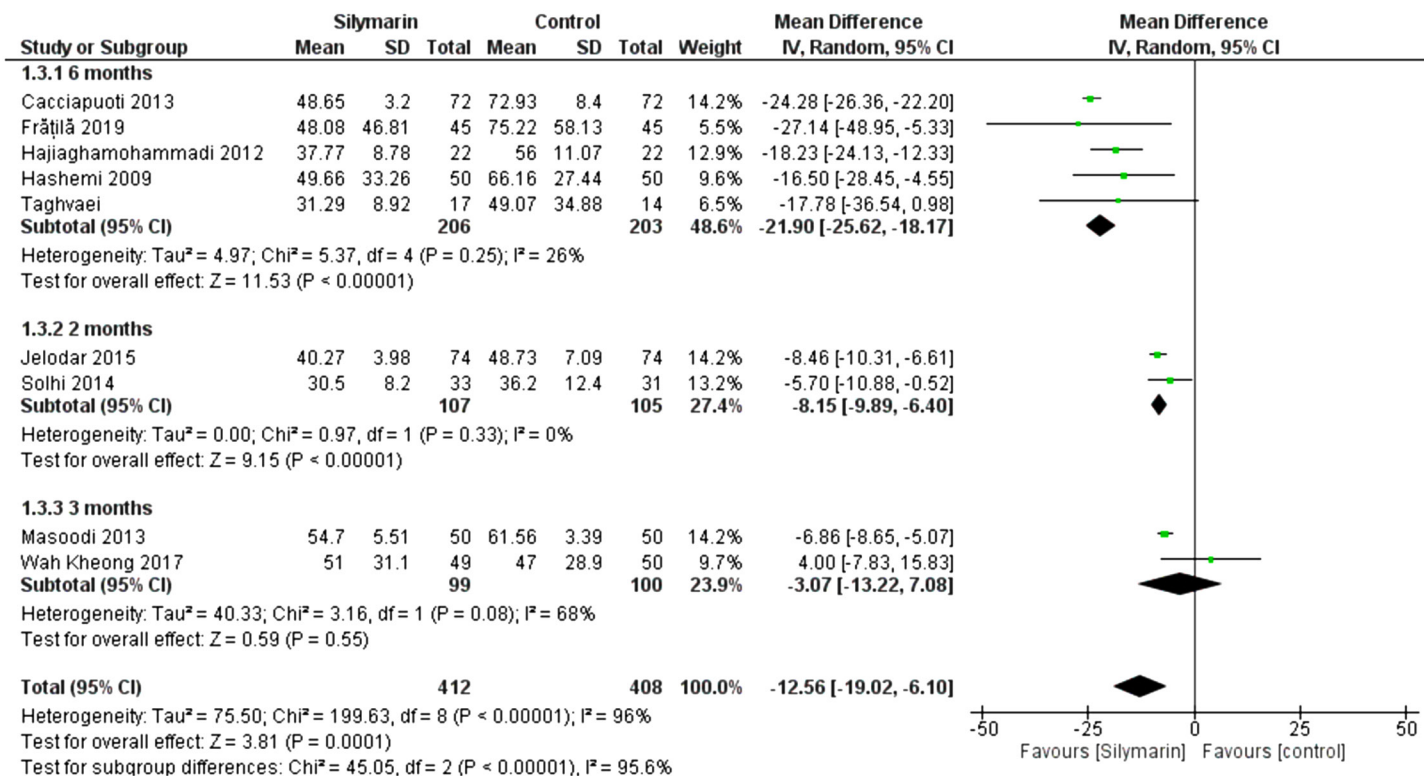


Figure 4: Analysis of AST outcome

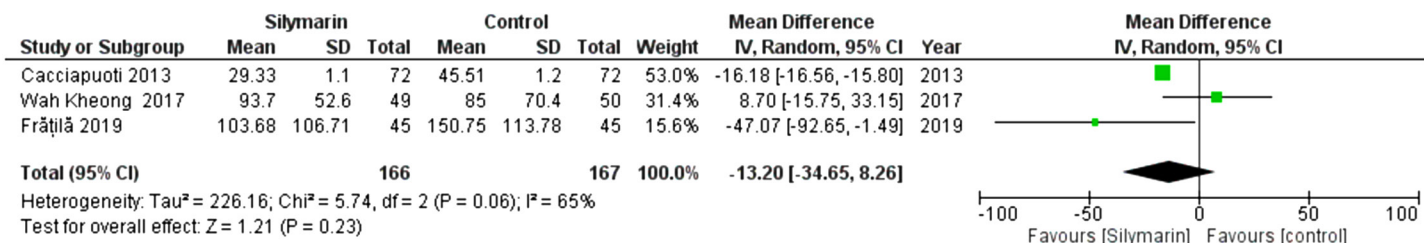


Figure 5A: Analysis of GGT Outcome

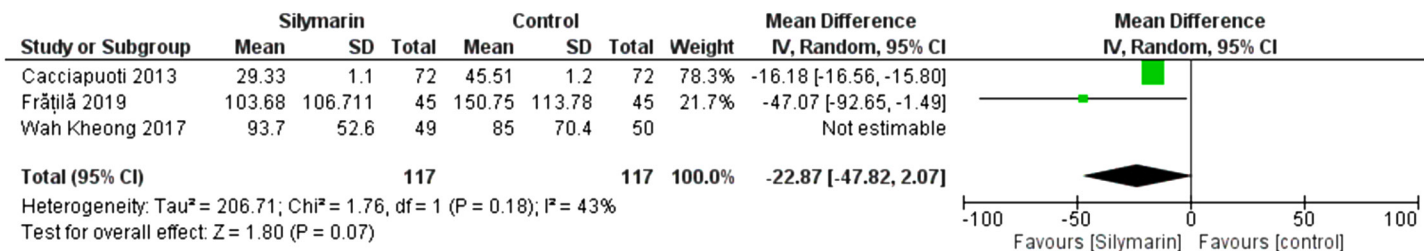


Figure 5B: Analysis of GGT Outcome – Rule one out analysis

difference between both groups (MD= -22.87 [-47.82, 2.07]) (P < 0.07). Figure 5B illustrates the analysis after the exclusion of one study.

**HOMA-IR**

Three studies (28–30) reported the HOMA-IR score. The overall mean difference showed that there was no significant difference between both

groups (MD= -0.42 [-0.87, 0.02]), (p = 0.06). Pooled analysis was homogeneous (p = 0.24); I<sup>2</sup> = 31% as shown in Figure 6.

**TC**

Six studies (25,27–30,32) reported the TC. The overall mean difference showed that there was no significant difference between both groups



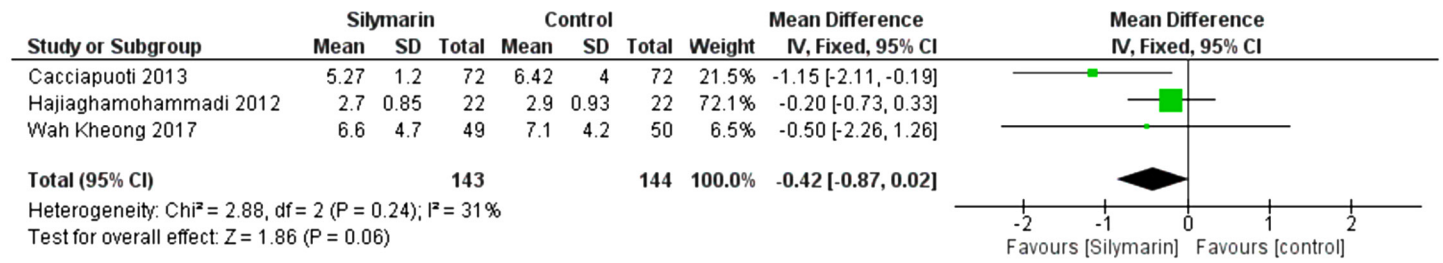


Figure 6: Analysis of HOMA outcome

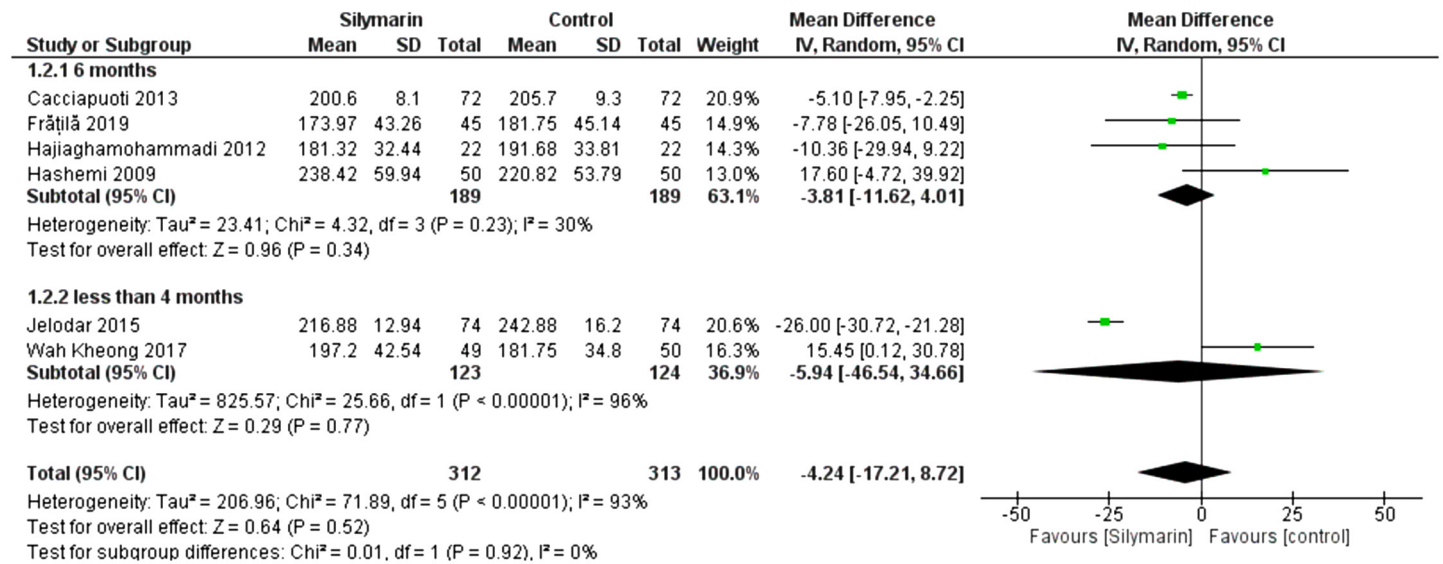


Figure 7: Analysis of TC outcome

(MD = -4.24 [-17.21, -8.72]), ( $p = 0.52$ ). Pooled analysis was heterogeneous ( $p = 0.00001$ );  $I^2 = 93%$  as shown in Figure 7. We solved the heterogeneity by subgroup analysis according to the duration of drug administration. The first subgroup (6 months) contains four studies (27,28,30,32). The mean difference showed that there was no significant difference between both groups (MD = -3.81 [-11.62, -4.01]), ( $p = 0.34$ ). Pooled analysis was homogeneous ( $p = 0.22$ );  $I^2 = 30%$  as shown in Figure 7. The second subgroup (less than 4 months) contains two studies (25,29). The mean difference showed that there was no significant difference between both groups (MD = -5.94 [-46.54, -34.66]), ( $p = 0.77$ ). Pooled analysis was heterogeneous ( $p = 0.00001$ );  $I^2 = 96%$  as shown in Figure 7.

**TG**

Six studies (25,27-30,32) reported the TG. The overall mean difference showed that there was a non-significant difference between both groups (MD= -9.85 [-24.04, 19.90]), ( $P < 0.17$ ). Pooled

analysis was heterogeneous ( $P = 0.0005$ ;  $I^2 = 78%$ ) as shown in Figure 8A. We solved the heterogeneity by the exclusion of Wah Kheong et al. (2017) (29) ( $p = 0.64$ );  $I^2 = 0%$ . The pooled analysis after exclusion showed that there was a significant difference between both groups (MD = -22.60 [-23.83, -21.38]) ( $p < 0.00001$ ). Figure 8B illustrates the analysis after the exclusion of one study.

**HDL**

Three studies (27,29,30) reported the HDL. The overall mean difference showed that there was a significant difference between both groups (MD = 2.13 [1.60, 2.66]), ( $p < 0.01$ ). Pooled analysis was homogeneous ( $p = 0.54$ );  $I^2 = 0%$  as shown in Figure 9.

**LDL**

Three studies (27,29,30) reported the LDL. The overall mean difference showed that there was a non-significant difference between both groups (MD= -2.92 [-16.37, 22.21]), ( $P < 0.77$ ). Pooled analysis was heterogeneous ( $p = 0.0008$ ;  $I^2 = 86%$ )

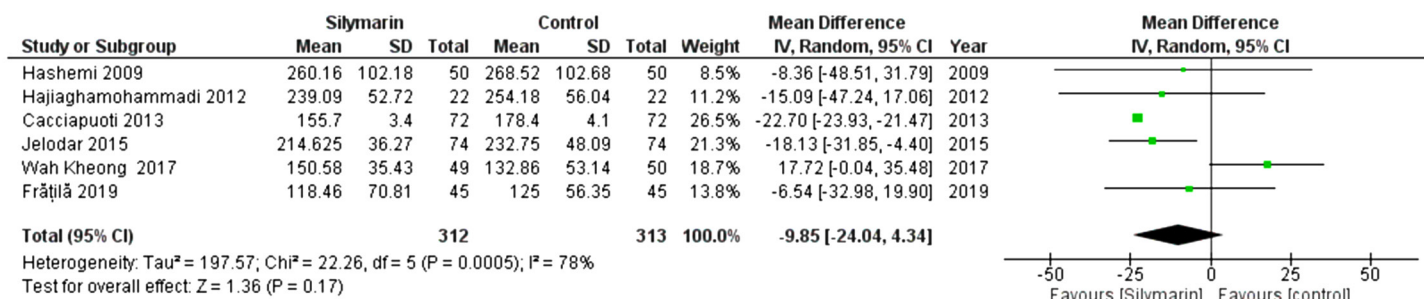


Figure 8A: Analysis of TG outcome

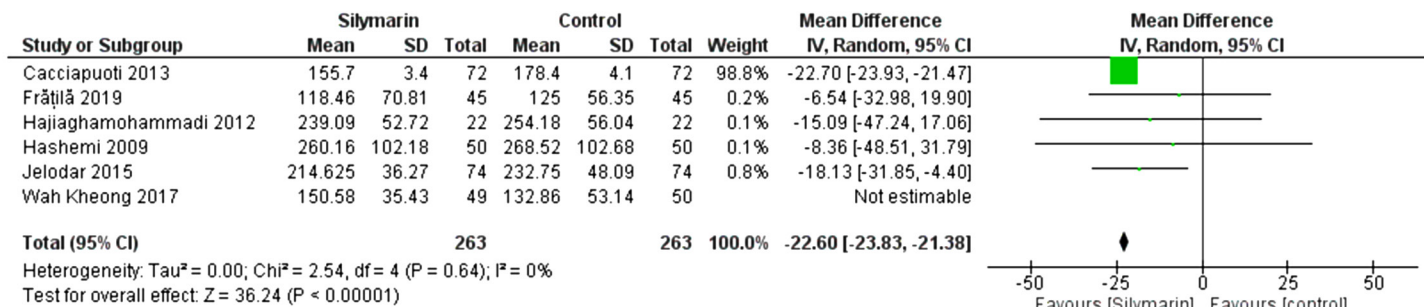


Figure 8B: Analysis of TG outcome -Rule one out analysis

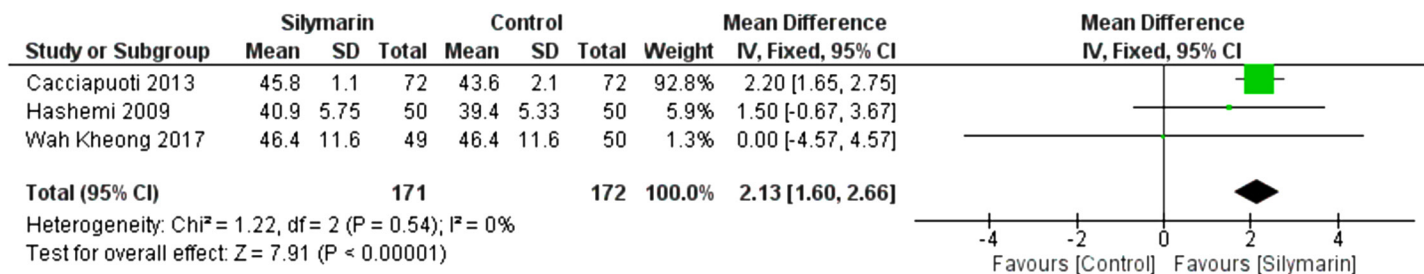


Figure 9: Analysis of HDL outcome

as shown in Figure 10A. We solved the heterogeneity by the exclusion of Cacciapuoti et al. (30,31) ( $p = 0.44$ );  $I^2 = 0\%$ . The pooled analysis after exclusion showed that there was no significant difference between both groups (MD = 11.88 [-0.39, 24.14]) ( $p = 0.06$ ). Figure 10B illustrates the analysis after the exclusion of one study.

### BMI

Five studies (24,27–30) reported the BMI. The overall mean difference of the BMI showed that there was no significant difference between either group (MD = -0.32 [-0.70, 0.06]), ( $p = 0.1$ ). Pooled analysis was homogeneous ( $p = 0.75$ );  $I^2 = 0\%$  as shown in Figure 11.

### Degree of fibrosis resolution

Only one study (30) reported significant improvement in the degree of fibrosis resolution. Their results revealed that higher proportion of patients in the silymarin group had reductions in fibrosis based on histology (reductions of 1 point or more), (22.4%) than the placebo group (6.0%) ( $P = 0.023$ )

### DISCUSSION

In this meta-analysis, we included 412 patients treated with silymarin from nine clinical trials. We found that silymarin significantly reduced AST and ALT levels. The results also showed a significant reduction in the levels of triglyceride,

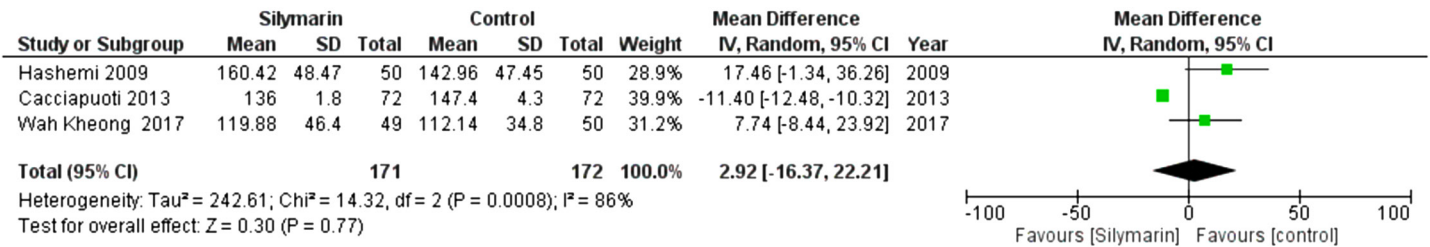


Figure 10A: Analysis of LDL outcome

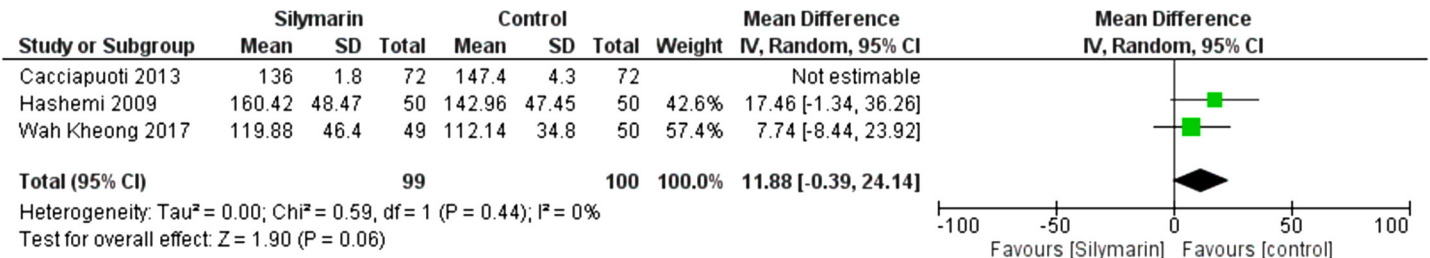


Figure 10B: Analysis of LDL outcome: Rule one out analysis

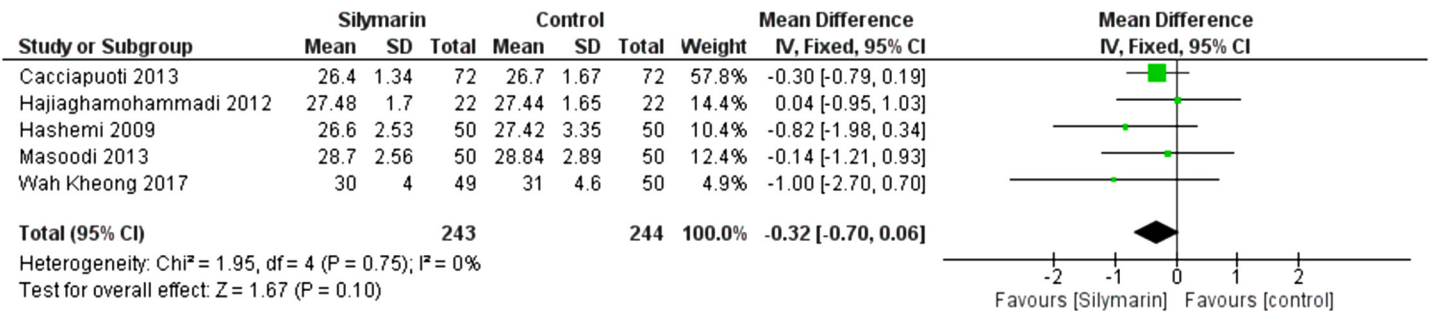


Figure 11: Analysis of BMI outcome

and improvement in HDL. Subgroup analysis demonstrated that more duration of follow up was associated with more reduced levels of liver transaminases.

Overall, our analysis is more comprehensive and up to date than Kalopitas et al.'s study. It adds to the evidence for silymarin's positive effects on liver enzymes and metabolic variables in MASLD (34).

Elevated liver enzymes are the prevailing abnormality in patients with MASLD. MASLD is commonly associated with metabolic syndrome and insulin resistance, which in turn, may elevate the liver enzymes (35). Elevated transaminase levels correlate with the histological state of the liver. However, histological outcomes are relevant as biochemical outcomes in NAFLD for various reasons.

For this reason, liver enzymes could be a non-invasive clue to generalized liver injury. Silymarin,

an insulin sensitizer, significantly reduced the liver enzymes (34). However, it had less significant results in restoring TG and TC levels back to normal. Recent trials performed by Kheong et al. (30), Cacciapuoti et al. (31), Hashemi et al. (28), Taghvaei et al. (27), Jelodar et al. (26), and Dahm et al. (24), found that silymarin showed a significant reduction in the levels of ALT and AST. Previous studies have already proved that reduction in transaminases is associated with reviving of liver function in MASLD patients (35,36). This also predicted a better prognosis and a lower incidence of progression to either liver cirrhosis or hepatocellular carcinoma. Silymarin was also associated with more reduction in ALT level than AST level because the original level of ALT was much higher than AST level in patients with MASLD.

A clinical review by Flora et al. (37) found that silymarin may be associated with a better clinical

prognosis of both acute and chronic hepatitis, drugs, and alcohol-induced hepatitis. A trial by Kheong et al. (30) demonstrated similar results; they found that silymarin may be effective in reducing ultrasonographic changes and the biochemical induced by MASLD. A study by Loguerio et al. (38) found that silymarin added to vitamin E and phospholipids could be associated with improvement in the liver enzymes in patients with chronic liver damage.

Solhi et al. (32) proved that the combination of silymarin, *Phyllanthus niruri*, and choline was associated with better hepatoprotective effects than silymarin alone in hepatic disorders of different inflammation and destruction stages. This hepatoprotective combination had a smaller dose of silymarin with similar results, which may reduce the side effects produced by the drug. Haji-aghahammadi et al. (29) found that silymarin 700 mg three times daily for 48 weeks, a well-tolerated and safe dose, was associated with greater fibrosis improvement compared with placebo. For this reason, silymarin seemed to be useful for fibrosis improvement in biopsy-proven NASH patients.

A trial by Masoodi et al. (25) found that silymarin extract in patients with MAFLD significantly reduced levels of plasma cholesterol, fasting glucose, triglyceride, AST, and ALT, while the plasma levels of folate and B12 did not change significantly. The major strength points of our analysis were that we only included clinical trials, which was an important strength point to ensure the highest evidence according to GRADE. Another strength point was that all the included studies showed low risk of bias. We also conducted the analysis of a good sample size. We tried to solve any inconsistency among studies using appropriate methodologies reported by the Cochrane's handbook (39) including mainly the leave-one-out and subgroup analysis.

## LIMITATIONS AND RECOMMENDATION

The heterogeneity in some outcomes was a major limitation; therefore, we tried to solve the heterogeneity by subgroup analysis. In addition, we have not registered a protocol for our study. Another limitation was that three of our included trials were not blinded, which might have affected the results of the analysis. We conclude that silymarin treatment seems to be significantly effective in the

reduction of biochemical and transaminases levels in patients with MAFLD. In addition, a highly need for future well-designed studies to ensure our outcomes and to examine whether the reduction in transaminase levels corresponds to histological improvement.

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