

1 The Emergence of Non-Linear Evolutionary Trade-offs and the 2 Maintenance of Genetic Polymorphisms

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7 Pareto front

8

9 Abstract

10 Evolutionary models of quantitative traits often assume trade-offs between beneficial and
11 detrimental traits, requiring modelers to specify a function linking costs to benefits. The choice of
12 trade-off function is often consequential; functions that assume diminishing returns (accelerating
13 costs) typically lead to single equilibrium genotypes, while decelerating costs often lead to
14 evolutionary branching. Despite their importance, we still lack a strong theoretical foundation to
15 base the choice of trade-off function. To address this gap, we explore how trade-off functions
16 can emerge from the genetic architecture of a quantitative trait. We developed a multi-locus
17 model of disease resistance, assuming each locus had random antagonistic pleiotropic effects
18 on resistance and fecundity. We used this model to generate genotype landscapes and
19 explored how additive versus epistatic genetic architectures influenced the shape of the trade-
20 off function. Regardless of epistasis, our model consistently led to accelerating costs. We then
21 used our genotype landscapes to build an evolutionary model of disease resistance. Unlike
22 other models with accelerating costs, our approach often led to genetic polymorphisms at
23 equilibrium. Our results suggest that accelerating costs are a strong null model for evolutionary
24 trade-offs and that the eco-evolutionary conditions required for polymorphism may be more
25 nuanced than previously believed.

26

27 1. Introduction

28 From life-history to foraging to disease resistance, genetic trade-offs are at the heart of many
29 questions in evolutionary biology. In mathematical models, trade-offs between beneficial and
30 deleterious traits are often necessary to maintain balancing selection [1]. Without an intrinsic
31 downside, there is nothing preventing quantitative traits from evolving towards their maximum.
32 For example, models of disease resistance typically assume the evolution of increased
33 resistance carries a cost to either host mortality or fecundity [2,3]. Such trade-offs could emerge
34 from either physiological constraints, or pleiotropic effects of the mutations affecting the focal
35 trait.

36 Theoretical models of evolutionary processes have shown that particular assumptions
37 about the shape of trade-off function, or how one quantitative trait scales with another can have
38 major implications for evolutionary outcomes [3–7]. Disease resistance is particularly

39 emblematic of trade-off function dependent evolution: Boots and Haraguchi [3] found when
40 fecundity costs scale faster than resistance benefits (referred to as accelerating, or convex cost
41 functions), evolution favours a single intermediate host genotype, whereas decelerating (also
42 referred to as concave) costs lead to the coexistence of resistant and susceptible hosts.
43 Accelerating costs leading to a single optimal genotype while decelerating costs lead to genetic
44 polymorphisms is common outcome of models of quantitative traits with ecological feedbacks
45 [5]. Similar patterns have been shown in predator behaviour models [8,9], life-history evolution
46 models [10] and disease resistance evolution models [3,4]. The shape of trade-off functions has
47 also been shown to determine evolutionary outcomes *in vivo*. By manipulating fecundity-survival
48 trade-offs in *Escherichia coli*, Maharjan et al. were able to experimentally validate the results of
49 theoretical models showing that changes in the shape of trade-off functions can indeed
50 determine evolutionary outcomes [7].

51 Despite the abundance of evidence demonstrating the importance of trade-off functions,
52 we understand their consequences far more than the biological processes that shape trade-off
53 functions. While trade-off functions depict genotypic variation as a one-to-one relationship
54 between quantitative traits, natural variation is two-dimensional. To address this, the concept of
55 the Pareto front is a useful bridge [11]. The Pareto front is defined as the set of all phenotypes,
56 such that improving performance in one trait can only be accomplished through a decrease in
57 performance in another trait. For example, if we consider a trade-off between disease resistance
58 and fecundity, the Pareto front represents the most fecund phenotypes for each level of
59 resistance (Fig 1). Theory predicts that evolution should select for genotypes close to the Pareto
60 front, with genetic polymorphism oriented along the front [12]. Mapping the curvature of the
61 Pareto front can be used as a strategy to identify trade-off functions [13–16], thus understanding
62 how genetic factors shape the Pareto front could be valuable for understanding genetic trade-
63 offs.

64 If we assume a set of pleiotropic alleles have independent, additive contributions to a
65 beneficial and detrimental trait, then low levels of the beneficial trait should be achievable using
66 only the most cost-effective alleles. However, this might not be possible for higher levels of the
67 beneficial trait, meaning evolution must have to rely on costlier alleles. This is one mechanism
68 that could produce accelerating costs, although this prediction relies on strongly simplifying
69 genetic assumptions, principally the absence of epistatic interactions between loci. If beneficial
70 epistatic interactions between multiple alleles are only realized once multiple pleiotropic alleles
71 are fixed, the benefits of subsequent mutations could be magnified, leading to decelerating
72 costs.

73 To test the prediction that strictly additive genetics produces accelerating costs, while
74 epistasis could produce decelerating costs, we developed an allelic model of the evolution of
75 quantitative disease resistance. We assumed that disease resistance was determined by a
76 series of discrete haploid loci, where each locus can have two possible alleles: one with
77 antagonistic pleiotropic effects on fecundity and host resistance and one with no effects on
78 either. With this framework, we generated genotype distributions and investigated the degree to
79 which epistasis can influence the shape of the Pareto front. Next, we incorporated our genotype
80 distribution model into an evolutionary model of disease resistance, where mutation allows
81 hosts to move between genotypes. With this model, we asked whether epistatically induced
82 changes in trade-off functions can result in a shift from a single dominant genotype to the
83 maintenance of genetic polymorphism, mirroring patterns seen in previous models of

84 quantitative disease resistance [3]. Unlike previous models [3,5], our approach requires no initial
85 assumptions about trade-off functions.

86

87 2. Simulating Pareto Fronts

88 To simulate Pareto fronts, we developed a model of quantitative pathogen resistance
89 (henceforth referred to as the discrete random loci model) which assumes that host resistance
90 is determined by a fixed number, n , of haploid loci. Each locus has two possible alleles: a
91 neutral allele which has no effect on the host phenotype, and an active allele which
92 pleiotropically increases host resistance (benefits) and reduces host fecundity (costs). Given a
93 sample of allelic effects, we can then define the set G of all possible genotypes as $G = \{0,1\}^n$,
94 with each genotype $\mathbf{g}_i \in G$ being a vector of length n . For each locus, a 0 represents the neutral
95 allele while a 1 represents the active allele. This process can be thought of as flipping switches
96 on a panel with n different switches, where each combination of switch positions produces a
97 unique genotype. For each locus, we assumed that the active allele has costs and benefits
98 sampled from a random exponential distribution. We define the resistance effect vector, \mathbf{r} by
99 $r_i \sim \text{Exp}(\lambda_b)$ and the fecundity cost vector \mathbf{c} by $c_i \sim \text{Exp}(\lambda_c)$, where $i \leq n$, λ_c represents the cost
100 variance and λ_b represents the benefit variance. We initially assumed active alleles at multiple
101 loci had additive effects for both resistance and fecundity. With this assumption, we define
102 disease transmission, $\beta_{\mathbf{g}_i}$ for a given genotype \mathbf{g}_i as the normalized sum of all the active alleles
103 for that genotype (Eqn. 1, subtracted from 1 to convert resistance to transmission). This value is
104 then multiplied by β_0 , the baseline level of transmission.

$$105 \quad \beta_{\mathbf{g}_i} = \left(1 - \frac{\langle \mathbf{g}_i, \mathbf{r} \rangle}{\sum_n r_i}\right) \beta_0 \quad (1)$$

106 With this normalization, $\beta_{\mathbf{g}_i}$ ranges from 0 to β_0 . The total fecundity cost of each genotype, $\delta_{\mathbf{g}_i}$
107 are defined similarly, where fecundity is normalized to range from 0.2 (the baseline deathrate,
108 see Section 3 below) to 1.

109 Beyond purely additive allele interactions, we also explored how non-additive epistasis
110 can influence the shape of the Pareto front. Here, a random subset of all active allele pairs is
111 considered to have an epistatic interaction. For every interacting pair of alleles, epistasis either
112 increases or decreases the combined effect of both alleles on the host resistance. We
113 considered two forms of epistasis: first-order, where pairs of alleles have an epistatic interaction
114 and second-order, where triplets of alleles have an epistatic interaction. There are $\binom{n}{k}$ unique
115 loci pairs, which could possibly have an epistatic interaction, where n is the number of loci and
116 $k = 2$ for first-order epistasis or $k = 3$ for second-order epistasis. We randomly assigned a fixed
117 proportion of these pairs and triples. We then assigned each pair an epistatic interaction with
118 probability p_1 for first-order epistasis and p_2 for second-order epistasis. Next, we modified the
119 cumulative effect of each loci pair (i, j) on resistance to $\theta_1(r_i + r_j)$ where $\theta_1 \sim N(1, \sigma_1^2)$. We
120 implemented second-order epistasis similarly, by setting the cumulative effect of three given
121 alleles to $\theta_2(r_i + r_j + r_k)$ where $\theta_2 \sim N(1, \sigma_2^2)$. For all simulations with epistasis, the normalization
122 step occurs after the epistatic effects are introduced.

123 We ran three series of simulations: one with no epistasis (Fig. 2A), one with only first-
124 order epistasis (Fig. 2B), and one with both first and second order epistasis (Fig. 2C). For all
125 simulations, we set the number of loci, n , to 9. Based on GWAS studies, this is a small but
126 plausible number of loci [17–19]. Alternative models with either 5 or 13 loci did not have a
127 qualitatively different effect on the Parent front (Fig. S1). For each epistasis treatment, we ran
128 100 random instantiations, and computed the Pareto front as the average optimal fecundity for
129 every level of resistance (see the orange line, Fig 2A-C for an example of a single instantiation,
130 Fig 2D-E for the average). Regardless of epistasis, the resulting Pareto had clear accelerating
131 costs (Fig 2). However, with second-order epistasis, the trade-off function had a reduced
132 curvature relative to other scenarios (Fig. 2C).

133

134 3. Evolutionary Dynamics

135 To determine whether randomly generated genotype distributions drive similar evolutionary
136 outcomes to standard models with trade-off functions, we built an evolutionary model on top of
137 our randomly generated genotypes distributions. Given that we found accelerating costs when
138 generating Pareto fronts, we predicted that this model would produce a single equilibrium
139 genotype. Our model uses time-separated mutation and selection steps, similar to adaptive
140 dynamics [20]. In classic adaptive dynamics models, mutations are introduced into populations
141 at equilibrium, and this process is iterated until a final evolutionary equilibrium is reached. While
142 adaptive dynamics models assume that new mutants differ from their parental generation by a
143 small phenotypic value given by a trade-off function, our implementation assumes that new
144 mutants differ from their progenitors by a single allele at a given locus. We used the discrete
145 random loci model as the basis for the phenotype of each genotype, where a single mutation
146 does not necessarily correspond to a small phenotypic change. Furthermore, instead of
147 assuming a smooth trade-off function, trade-offs are generated by a random process and are
148 inherently non-smooth.

149 For a given instantiation of random of allelic effects, each simulation begins with 100
150 uninfected hosts from the completely susceptible genotype, ($\mathbf{g}_0 = (0, \dots, 0)$) and 10 infected
151 hosts. We assumed that hosts reproduce asexually. Furthermore, we assume a sterilizing,
152 density-dependent disease without recovery, such that infection results in a total loss of
153 fecundity without induced mortality. We then computed numerical solutions, from $t = 0$ to $t =$
154 1000, so that the hosts can reach the ecological equilibrium. At this point, we introduced
155 mutation by taking 5% of all extant hosts and reassigning them to genotypes which differ from
156 their progenitors by one allele. Analogous to adaptive dynamics, we then ran the simulation to
157 ecological equilibrium again, and iteratively introduced new mutations. We ran the simulations
158 for a total of 15 mutational iterations to reach evolutionary equilibrium, using the same genotype
159 distribution parameters in Fig. 2. Since any two genotypes can differ by at most n loci, n
160 mutational steps are sufficient for all possible genotypes to be reached. Since the shortest path
161 to a particular genotype might not be evolutionary feasible, we include extra mutational
162 iterations to allow for evolutionary equilibrium. The equations governing these dynamics are
163 given below (Eqns. 2-3).

$$164 \quad \dot{S}_i = S_i(b - \delta_i - \mu - \gamma N - \beta_i I) \quad (2)$$

165

$$\dot{I} = I \left(\sum_i \beta_i S_i - \mu \right) \quad (3)$$

166 Here, S_i denotes the abundance of uninfected host genotype i , and I denotes the number of
167 infected hosts and N represents the total number of hosts, both susceptible and infected. The
168 host resistance and costs of resistance are given by β_i and δ_i respectively. Demographics are
169 controlled by the birthrate, b the deathrate, μ , and the coefficient of density-dependent growth,
170 γ . We considered three cases: no epistasis, first-order epistasis, and first and second-order
171 epistasis (Fig 3).

172 To test whether epistasis affected equilibrium host genetic diversity, we ran 100
173 simulations for each epistasis treatment. We then calculated the host genetic diversity at
174 equilibrium using the Shannon index, H , where $H = -\sum_i p_i \ln(p_i)$, with p_i being the proportion of
175 each genotype. Here, $H = 0$ indicates a monomorphic population, and $H > 0$ indicates a
176 polymorphic population. As most simulations resulted in either one or two host genotypes, we
177 used a non-parametric Kruskal-Wallis test to test whether the epistasis treatment produced
178 significant differences in host diversity.

179 We expected diversity would be lowest in the purely additive model, since the Pareto
180 front in the model was strongly accelerating, and in classic adaptive dynamics models only
181 decelerating cost functions lead to stable polymorphisms. However, we found that certain model
182 instantiations were able to maintain polymorphisms, even for purely additive models (Fig 3A-B).
183 For simulations without epistasis, 30% had a polymorphism with at least 2 genotypes having
184 equilibrium abundance greater than 5 (to ensure polymorphisms were not solely maintained by
185 new mutants), while 37% were polymorphic for first-order epistasis and 43% for second-order
186 epistasis. Epistasis did not significantly affect the host's equilibrium genetic diversity ($p = 0.54$).
187 We did not observe these polymorphisms when we ran adaptive dynamics simulations with
188 equivalent parameters (Fig. S2).

189

190 4. Discussion

191 The discrete random loci model demonstrates how selection acting on a random assortment of
192 mutations can generate non-linear cost functions. Our approach bridges fitness landscape
193 models, such as the $N - k$ model [21], and adaptive dynamics models [20] to test how genetic
194 processes can define trade-off functions and enable polymorphism via ecological feedbacks.
195 Our first key result is that accelerating cost function can emerge naturally from a process of
196 random pleiotropic mutations followed by selection. We found that cost curves were most
197 strongly accelerating when loci were purely additive. Epistasis could only blunt this trend but
198 could not produce linear or decelerating costs. Our second key result is that allowing for a multi-
199 locus mutational process instead of a fixed trade-off function changes evolutionary predictions,
200 resulting in more polymorphic outcomes. In single locus adaptive dynamics models,
201 accelerating cost functions generally lead to stable, monomorphic populations [5]. However, we
202 found that even though our mutation model generated Pareto fronts with an accelerating cost
203 curve, the evolution and maintenance of stable polymorphism was common. This result
204 contradicts the findings of classical adaptive dynamics models [3,5], suggesting that other eco-
205 ecological factors beyond the shape of trade-off functions can drive genetic polymorphisms.

206 While some studies have demonstrated accelerating costs in experimental evolution
207 studies [35–37], quantifying the relationship between traits and their costs is difficult. Costs can
208 manifest in many different ways, potentially via specific ecological contexts [38], meaning that
209 recreating the context in which costs manifest can be impractical if not impossible. For disease
210 resistance, detecting any costs can be difficult, yet alone mapping costs to resistance levels with
211 sufficient resolution to define a cost curve [39]. Despite this uncertainty, accelerating costs are a
212 common assumption in adaptive dynamics models [5]. Our results suggest that this is a
213 reasonable null model for trade-offs between quantitative traits, while decelerating costs might
214 require more justification.

215 Contrary to our initial predictions, our model resulted in accelerating costs even with
216 second-order epistasis. While decelerating costs might be more likely with third or even fourth
217 order epistasis, such interactions are plausible but likely less frequent [22]. Decelerating costs
218 might also emerge from evolvability constraints. In this case, even when the Pareto depicts
219 accelerating costs, evolution may be unable able to track the front, instead following a path of
220 decelerating costs. For example, the initial cost of evolutionary innovations may be reduced by
221 compensatory mutations which can only emerge later [23]. Evolutionary trajectories may also
222 depend on stepwise mutations at a single locus [24], as well as recombination, which are not
223 included in our model. These simplifying assumptions in our model make it easier for evolution
224 to reach all genotypes, potentially removing mechanisms that lead to a broader range of trade-
225 off functions. Decelerating costs could also result from physiological constraints, such as
226 allometric scaling laws. However, for a trait like quantitative resistance, it is not clear that such
227 physiological constraints would have a greater role in defining trade-offs than additive genetic
228 variance.

229 Our evolutionary model of disease resistance differs from traditional adaptive dynamics
230 approaches in several important ways. First, individual mutations do not necessarily result in
231 small phenotypic changes. First, similar polymorphisms can emerge from models with only two
232 alleles at a single locus [2]. Antonovics and Thrall found that when one host genotype is highly
233 resistant, it allows for the coexistence of more susceptible genotypes by reducing the
234 prevalence of infection. Since our model has the potential for single alleles with large effects,
235 the same mechanism as in single locus models could produce polymorphisms. Secondly, as the
236 Pareto fronts generated from my model are non-smooth, small perturbations from a purely
237 accelerating cost function may result in small regions where costs grow at a decelerating rate
238 relative to resistance. Such deviations can be seen in Fig. 2, where each curve has areas where
239 it does not reflect the overall accelerating cost pattern. Since trade-offs in vivo are unlikely to be
240 perfectly smooth [29], locally decelerating costs could be a plausible mechanism for maintaining
241 genetic diversity in natural systems. This influence of both the smoothness of the trade-off curve
242 could be tested by introducing perturbations into trade-off functions in adaptive dynamics
243 models known to produce a single continuously stable strategy to see if that strategy remains
244 stable with perturbation. If small regions of decelerating costs are responsible for polymorphism
245 in our model, then this should be reflected through adaptive dynamics as well.

246 With the advent of modern genomics, many assumptions of our model are increasingly
247 testable [25]. Our assumption that allelic effects are exponentially distributed is supported by
248 population genetics theory and GWAS studies [26–28]. While quantifying the frequency and
249 magnitude of epistatic interactions between many loci is difficult, combinatorial approaches to
250 mapping out fitness landscapes can be illuminating [29–31]. Depending on the trait and model

251 system, the frequency of epistasis is highly variable [25,32]. Furthermore, studies in yeasts and
252 bacteria have found that higher-order epistasis is nearly as prevalent as pairwise epistasis, and
253 that epistatic interactions occur between roughly 10% of mutation triplets [22,33,34]. While the
254 prevalence of epistatic interactions is highly species and phenotype dependent, what we do
255 know suggests that our implementation is a reasonable first approach.

256 In natural populations, traits like quantitative pathogen resistance often have a high
257 degree of genetic variability, thus theoretical models must reflect how this variation is
258 maintained. Our model demonstrates that without very strong epistasis, or a clear physiological
259 mechanism, accelerating costs might be the most realistic cost function for most evolutionary
260 trade-offs. However, unlike models with smooth trade-off functions, our model shows that even
261 these accelerating cost functions can lead to polymorphic outcomes. Stochastic, jagged trade-
262 off functions may therefore be an important driver of genetic variation.

263

264 Ethics

265 This work did not require ethical approval from a human subject or animal welfare committee.

266

267 Data accessibility

268 The python code used to generate all data and figures in this article is available on GitHub at
269 <https://github.com/svhulse/cost-model>.

270

271 Declaration of AI use

272 We have not used AI-assisted technologies in creating this article.

273

274 Authors' contributions

275 S.V.H.: conceptualization, formal analysis, software, visualization, writing - original draft. E.L.B.:
276 funding acquisition, supervision, writing – review and editing.

277

278 Conflict of interest declaration

279 We declare we have no competing interests.

280

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288

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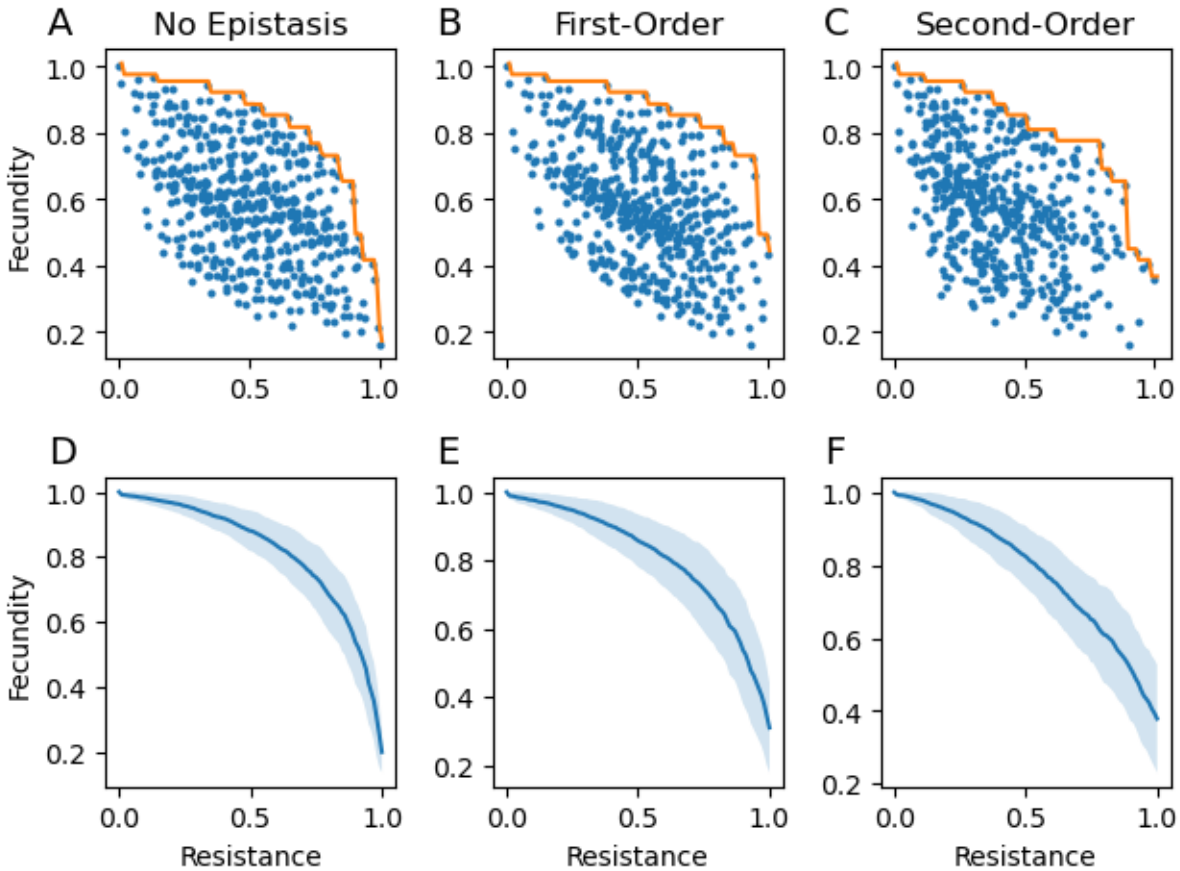
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388 **Figure 1:** Schematic of how Pareto Fronts can define cost functions for (A): linear costs, (B):
389 decelerating costs and (C): accelerating costs. Each dot represents a host genotype. The lighter
390 red dots represent possible genotypes that would be removed by selection. The dashed blue
391 line represents the Pareto front, where the phenotype space beyond is inaccessible to evolution.

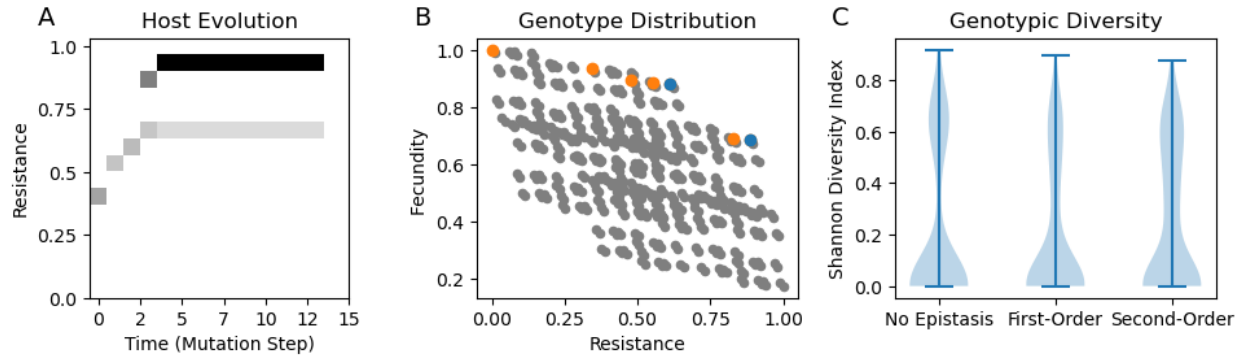
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394 **Figure 2:** Distributions of genotype resistances and costs. A-C: Genotype distributions showing
395 the fecundity and resistance level for all host genotypes for an instantiation with no epistasis (A),
396 first-order epistasis (B), and first and second order epistasis (C). The clustering observed here
397 is a byproduct of the additive loci: each additional locus effectively copies and shifts the
398 distribution without it, leading to the observed patchiness when individual loci have large effects.
399 For each, the orange line represents the Pareto front. D-F: Simulated Pareto front averaged
400 over 100 instantiations for no epistasis (D), first-order epistasis (E), and first and second order
401 epistasis (F). The light blue region represents values within one standard deviation of the
402 average fecundity for a particular level of resistance. Parameters: $\lambda_b = 0.1, \lambda_c = 0.1, p_1 =$
403 $0.3, p_2 = 0.3, \sigma_1^2 = 0.2, \sigma_2^2 = 0.2$.

404



405

406 **Figure 3:** Evolutionary dynamics of disease resistance, using the discrete random loci model A:
407 Phenotypic changes in resistance over the course of a simulation with no epistasis. B:
408 Distribution of host genotypes with genotypes that had large populations at some point in
409 evolutionary time in orange. Blue dots represent genotypes present at the end of the simulation,
410 while orange dots are genotypes that were present at previous ecological equilibrium but not the
411 final equilibrium. Panels A and B correspond to the same simulation. C: Equilibrium genetic
412 diversity across simulations with 100 different allelic instantiations for each epistasis treatment.
413 There was no significant difference in the Shannon diversity across treatments. Parameters:
414 $\beta_0 = 0.005, \mu = 0.2, \gamma = 0.001$, the parameters for the trait distributions are the same as in Fig.
415 2.