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New-onset obstructive airway disease following COVID-19: a multicenter retrospective cohort study

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Abstract

Background The study assessed the association between COVID-19 and new-onset obstructive airway diseases, including asthma, chronic obstructive pulmonary disease, and bronchiectasis among vaccinated individuals recovering from COVID-19 during the Omicron wave.

Methods This multicenter retrospective cohort study comprised 549,606 individuals from the U.S. Collaborative Network of TriNetX database, from January 8, 2022, to January 17, 2024. The hazard of new-onset obstructive airway diseases between COVID-19 and no-COVID-19 groups were compared following propensity score matching using the Kaplan–Meier method and Cox proportional hazards model.

Results After propensity score matching, each group contained 274,803 participants. Patients with COVID-19 exhibited a higher risk of developing new-onset asthma than that of individuals without COVID-19 (adjusted hazard ratio (aHR), 1.27; 95% CI, 1.22–1.33; $p < 0.001$). Stratified analyses by age, SARS-CoV-2 variant, vaccination status, and infection status consistently supported this association. Non-hospitalized individuals with COVID-19 demonstrated a higher risk of new-onset asthma (aHR, 1.27; 95% CI, 1.22–1.33; $p < 0.001$); however, no significant differences were observed in hospitalized and critically ill groups. The study also identified an increased risk of subsequent bronchiectasis following COVID-19 (aHR, 1.30; 95% CI, 1.13–1.50; $p < 0.001$). In contrast, there was no significant difference in the hazard of chronic obstructive pulmonary disease between the groups (aHR, 1.00; 95% CI, 0.95–1.06; $p = 0.994$).

Conclusion This study offers convincing evidence of the association between COVID-19 and the subsequent onset of asthma and bronchiectasis. It underscores the need for a multidisciplinary approach to post-COVID-19 care, with a particular focus on respiratory health.

Keywords Asthma, Bronchiectasis, COPD, COVID-19, Obstructive airway diseases

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Background

Over the past 3 years, severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2) has infected more than 773 million individuals globally [1]. Despite coronavirus disease (COVID-19) being responsible for over seven million fatalities, the case fatality rate stood at a mere 0.9% [1], suggesting that more than 99% of patients with COVID-19 have survived the acute phase of SARS-CoV-2 infection. Nevertheless, a significant number of COVID-19 survivors develop post-COVID-19 conditions, also termed long COVID, which present a grave threat to global health [2–4]. Reports have documented a broad range of persistent symptoms linked to post-COVID-19 conditions, impacting the cardiovascular and pulmonary systems, gastrointestinal tract, neurocognitive functions, psychological well-being, and musculoskeletal system [3–7].

The respiratory system is frequently involved in post-COVID-19 conditions, as demonstrated by clinical and radiologic findings, along with pulmonary function tests [8, 9]. Predominantly, studies have concentrated on the restrictive ventilatory defect, which is marked by lung fibrosis, a diminished diffusion capacity of the lung for carbon monoxide, and reduced lung volume [8–11]. In addition, residual abnormalities on computed tomography for patients with COVID-19 are not uncommon. One meta-analysis showed that fibrotic-like changes had the highest event rates of 0.44 and 0.38 during both short-term (1–6 months) and long-term (12–24 months) follow-up periods [12]. Moreover, patients with severe COVID-19 exhibited significantly higher rates of various abnormalities, including bronchiectasis, fibrotic-like changes, and reticulation, at long-term follow-ups compared to those in the non-severe subgroup [12]. Conversely, although prior research has identified a link between respiratory virus infections and the subsequent onset of asthma [13–15], the evaluation of obstructive airway disease risk following COVID-19 has been limited [16–18].

Kim et al. using the Korean National Health Insurance claim-based database demonstrated that COVID-19 could be associated with a higher risk of new-onset asthma (adjusted hazard ratio (aHR), 2.14; 95% CI, 1.88–2.45) [16]. However, the study was conducted between 2020 and June 30, 2021, in South Korea and focused specifically on asthma. Their findings were thus limited to the Asian population and the non-Omicron wave. Since then, the virulence, epidemiology, and disease severity of the currently prevalent Omicron variant have changed significantly from those of the early strains [19, 20]. An updated investigation among non-Asian populations is also needed to provide a broader understanding of the association between COVID-19 and obstructive airway

diseases. Therefore, we conducted this study using the TriNetX network to assess the association between COVID-19 and new-onset obstructive airway diseases, particularly asthma, following COVID-19 during the Omicron wave.

Results

Characteristics of study subjects

Among the vaccinated individuals without pre-existing asthma, COPD, or bronchiectasis, the study included 274,820 enrollees in the COVID-19 group and 2,214,332 in the no-COVID-19 group (Fig. 1). Prior to PSM, the average age of patients with COVID-19 was 55.5 years, compared to 54.8 years for those without. The COVID-19 group was 57.4% female, while the no-COVID-19 group was 58.9% female; 64% of the COVID-19 group and 67.5% of the no-COVID-19 group were Caucasians. Age, sex, and race did not differ significantly between the groups (all standardized differences < 0.1). However, the COVID-19 group showed a higher prevalence of comorbidities, including DM, dyslipidemia, hypertension, ischemic heart disease, heart failure, cerebrovascular disease, CKD, neoplasm, anemia, and systemic connective tissue disease, as well as allergic rhinitis, compared to their counterparts (Table 1). Additionally, a greater percentage of the COVID-19 group had histories of tobacco use/nicotine dependence, a BMI of ≥ 30 kg/m², and socioeconomic status (SES)- or psychosocial-related health hazards. Regarding medication, a higher proportion of the COVID-19 group had received systemic glucocorticoids, antihistamines, beta-blockers, aspirin, and NSAIDs within 1 year preceding the index date (Table 1).

After PSM, each group consisted of 274,803 patients. Characteristics for all the considered covariates, including demographics, comorbidities, and medications, were balanced between the two groups with standardized differences < 0.1 (Table 1). The group with COVID-19 was designated as the study group, whereas the group without COVID-19 was designated as the control group.

Primary outcome

During the follow-up period, 4216 new asthma cases were diagnosed in the study group, compared with 3352 in the control group. The study group had a significantly higher risk of developing new-onset asthma than the control group (aHR, 1.27; 95% CI, 1.22–1.33; $p < 0.001$; see Table 2). Additionally, Kaplan–Meier analysis showed a lower probability of event-free survival for the study group relative to the control group (log-rank $p < 0.001$; see Fig. 2).

Stratified analyses by age group, SARS-CoV-2 variant, vaccine status, and infection status consistently demonstrated similar trends. Patients with COVID-19 exhibited

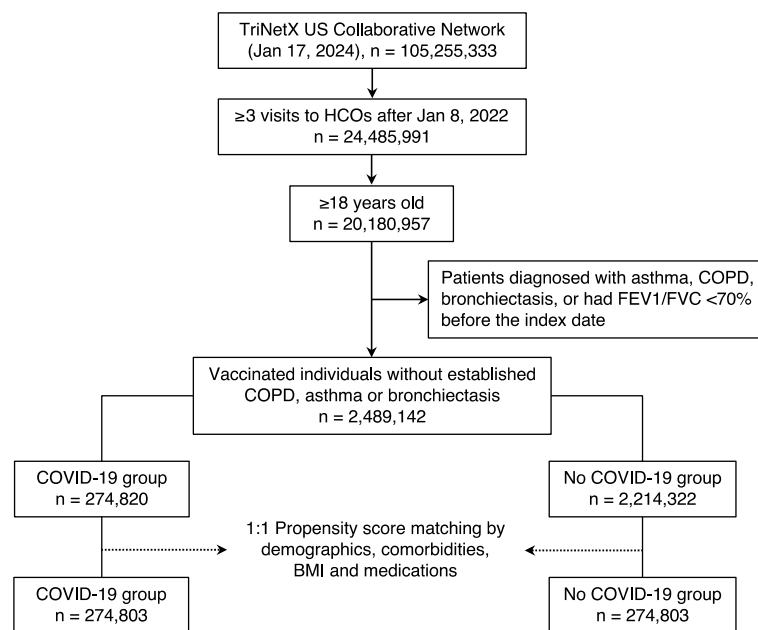


Fig. 1 Flowchart of cohort construction. BMI body mass index, COVID-19 coronavirus disease 2019, FEV1 forced expiratory volume in the first second, FVC forced vital capacity, HCO healthcare organization, COPD chronic obstructive pulmonary disease

a higher risk of new-onset asthma, as illustrated in Fig. 3. Notably, there were no significant differences in effect size estimates across the analyses, except for those based on disease severity (p for interaction = 0.011). A significantly higher risk of new-onset asthma was observed only in non-hospitalized patients with COVID-19, whereas no significant differences were observed between hospitalized and critically ill groups (Fig. 3).

Secondary outcomes

We also identified that patients with COVID-19 had a higher risk of bronchiectasis (aHR, 1.30; 95% CI, 1.13–1.50, $p < 0.001$, Table 2). There was a trend towards an increased risk of bronchiectasis in patients with critical COVID-19 compared to those not hospitalized, but no statistically significant between-subgroup heterogeneity was observed for bronchiectasis as evidenced by the interaction tests (all p values for interaction > 0.05) and the largely overlapping confidence intervals (Fig. 4). On the other hand, there was no significant difference in the incidence of COPD between the groups (aHR, 1.00; 95% CI, 0.95–1.06, $p = 0.994$, Table 2).

Sensitivity analysis

Our sensitivity analysis using different cutoff landmarks (2, 3, and 6 months after the index date) to define new-onset obstructive airway diseases and the start of follow-up, showed consistent results. These results suggested that COVID-19 is associated with an increased risk of

new-onset asthma and bronchiectasis, but not COPD (Additional file 1: Table S4).

Regarding the negative and positive control outcomes, we observed a positive association between COVID-19 and symptoms characteristic of the post-COVID-19 condition, as expected, compared to the no-COVID-19 group. No significant associations were observed with the negative control outcomes, as depicted in Additional file 1: Fig. S1.

Discussion

This large-scale study evaluated the respiratory sequelae of COVID-19, including asthma, COPD, and bronchiectasis. We found a potential link between COVID-19 and an increased risk of new-onset asthma. Specifically, patients with COVID-19 had a 27% higher risk of developing asthma during the follow-up period compared to those without COVID-19. Further analyses, stratified by age, SARS-CoV-2 variant, vaccination status, and infection status, as well as sensitivity analyses using different landmarks, consistently supported these findings. In summary, our study presents compelling evidence regarding the association between COVID-19 and the subsequent development of asthma, highlighting the potential long-term respiratory impacts of SARS-CoV-2 infection and enhancing our knowledge of the disease's progression beyond the acute phase.

Stratified analysis by disease severity revealed an intriguing nuance. Although the overall risk of new-onset

Table 1 Baseline characteristics

Variables	Before matching			After matching		
	COVID-19 (n = 274,820)	No COVID-19 (n = 2,214,322)	Std. diff.	COVID-19 (n = 274,803)	No COVID-19 (n = 274,803)	Std. diff.
Demographics						
Age (mean ± SD)	55.5 ± 18.5	54.8 ± 18.8	0.037	55.5 ± 18.5	55.8 ± 18.3	0.014
Female	161,968 (58.9)	1,271,264 (57.4)	0.031	161,954 (58.9)	161,511 (58.5)	0.003
White	185,452 (67.5)	1,417,970 (64.0)	0.073	185,438 (67.5)	185,767 (67.6)	0.003
African American	31,209 (11.4)	291,711 (13.2)	0.055	31,208 (11.4)	31,720 (11.1)	0.006
Asian	21,816 (7.9)	184,713 (8.3)	0.015	21,816 (7.9)	21,761 (7.7)	0.001
Comorbidities						
Diabetes mellitus	47,704 (17.4)	208,628 (9.4)	0.235	47,691 (17.4)	47,293 (17.1)	0.004
Dyslipidemia	105,886 (38.5)	460,850 (20.8)	0.395	105,870 (38.5)	110,171 (38.3)	0.032
Hypertension	107,739 (39.2)	494,876 (22.4)	0.371	107,723 (39.2)	110,598 (39.3)	0.021
Ischemic heart diseases	27,825 (10.1)	109,008 (4.9)	0.198	27,815 (10.1)	26,617 (10.1)	0.015
Heart failure	12,402 (4.5)	41,265 (1.9)	0.151	12,395 (4.5)	10,877 (4.4)	0.027
Peripheral vascular disease	6,295 (2.3)	22,709 (1.0)	0.099	6,289 (2.3)	5,680 (2.2)	0.015
Cerebrovascular diseases	14,040 (5.1)	52,719 (2.4)	0.144	14,032 (5.1)	12,813 (5.5)	0.021
Dementia	3,446 (1.3)	10,705 (0.5)	0.083	3,443 (1.3)	3,131 (1.1)	0.010
Alzheimer's disease	1,938 (0.7)	7,280 (0.3)	0.053	1,937 (0.7)	1,825 (0.0)	0.005
Chronic kidney disease	23,891 (8.7)	86,687 (3.9)	0.198	23,884 (8.7)	22,449 (8.8)	0.019
Inflammatory liver diseases	3,907 (1.4)	12,246 (0.6)	0.088	3,904 (1.4)	3,328 (1.1)	0.018
Liver cirrhosis	3,057 (1.1)	11,227 (0.5)	0.068	3,057 (1.1)	2,608 (1.1)	0.016
Neoplasms	61,041 (22.2)	272,071 (12.3)	0.265	61,026 (22.2)	61,983 (22.2)	0.008
Anemia	25,937 (9.4)	85,755 (3.9)	0.225	25,928 (9.4)	23,833 (9.9)	0.027
Systemic CTDs	7,369 (2.7)	27,394 (1.2)	0.104	7,365 (2.7)	6,643 (2.2)	0.017
Rheumatoid arthritis	4,177 (1.5)	15,896 (0.7)	0.076	4,170 (1.5)	3,873 (1.1)	0.009
Atopic dermatitis	1,509 (0.6)	6,073 (0.3)	0.043	1,509 (0.6)	1,346 (0.0)	0.008
Allergic rhinitis	19,379 (7.1)	69,874 (3.2)	0.178	19,371 (7.1)	19,381 (7.7)	0.000
Tobacco use	11,697 (4.3)	38,852 (1.8)	0.147	11,688 (4.3)	10,181 (4.4)	0.028
Nicotine dependence	12,913 (4.7)	60,134 (2.7)	0.105	12,908 (4.7)	12,503 (4.4)	0.007
Alcohol-related disorders	5,919 (2.2)	20,583 (0.9)	0.099	5,913 (2.2)	5,185 (2.2)	0.019
SES and psychosocial-related health hazards	9,816 (3.6)	28,838 (1.3)	0.148	9,805 (3.6)	8,539 (3.3)	0.026
HIV	2,797 (1.0)	14,075 (0.6)	0.042	2,797 (1.0)	2,501 (1.1)	0.011
BMI ≥30 kg/m ²	39,113 (14.2)	168,388 (7.6)	0.214	39,100 (14.2)	40,247 (14.1)	0.012
Medications						
Glucocorticoids	88,008 (32.0)	371,437 (16.8)	0.361	88,992 (32.0)	88,114 (32.3)	0.001
Antihistamines	55,601 (20.2)	217,363 (9.8)	0.295	55,587 (20.2)	55,187 (20.2)	0.004
Beta blockers	53,298 (19.4)	258,025 (11.7)	0.215	53,287 (19.4)	52,790 (19.1)	0.005
Aspirin	26,372 (9.6)	112,237 (5.1)	0.174	26,366 (9.6)	24,916 (9.9)	0.018
	25,176 (9.2)	109,365 (4.9)	0.165	25,166 (9.2)	24,976 (9.9)	0.002

BMI Body mass index, *COVID-19* Coronavirus disease 2019, *CTD* Connective tissue disease, *HIV* Human immunodeficiency virus, *NSAID* Non-steroidal anti-inflammatory drugs, *SD* Standard deviation, *SES* Socioeconomic status, *Std. diff* Standardized difference

asthma was significantly higher in the COVID-19 group, this association was particularly pronounced in non-hospitalized individuals. The absence of a significant difference in hospitalized and critically ill groups might be partly due to a smaller patient count in these categories. Progression to viral pneumonia could indicate more

severe inflammation of the lower airway and could be associated with the risk of subsequent obstructive airway diseases, but its occurrence, the chronological relationship with SARS-CoV-2 infection, and the etiologies could not be ascertained in the current database, thereby precluding further analysis on this variable. However, the

Table 2 Adjusted hazard ratios of primary and secondary outcomes

Outcomes	No. of patients with outcome		Adjusted hazard ratio (95% CI)	p-value
	COVID-19 (n = 274,803)	No COVID-19 (n = 274,803)		
Primary outcome				
Asthma	4216	3352	1.27 (1.22–1.33) [*]	<0.001
Secondary outcomes				
Bronchiectasis	438	345	1.30 (1.13–1.50) [*]	<0.001
COPD	2180	1964	1.00 (0.95–1.06) [*]	0.994

CI Confidence interval, COPD Chronic obstructive pulmonary disease

^{*} Proportionality (Schoenfeld test $p > 0.05$)

interaction test suggested a significant difference, indicating a complex relationship between COVID-19 disease severity and the risk of obstructive airway diseases that warrants further exploration.

Our findings are consistent with those reported by Kim et al., who used the Korean National Health Insurance claim-based database. They found that 1.6% of the COVID-19 cohort and 0.7% of the matched cohort developed new-onset asthma, with an aHR of 2.14 (95% CI, 1.88–2.45) [16]. However, unlike the Korean study [16] conducted during 2020–2021, our analyses were based on data collected after 2022 and feature a more ethnically diverse population from the TriNetX platform. Consequently, our results provided updated insights during the

Omicron wave and are likely to be more generalizable and relevant to the current context.

Respiratory tract viral infections, including respiratory syncytial virus and measles, have been identified as potential causes of bronchiectasis [21], but the link between SARS-CoV-2 and post-infectious bronchiectasis is not well established. Although a meta-analysis by Guinto et al. [22] reported a 16.8% prevalence of bronchiectasis (95% CI, 9.10–26.1%) based on imaging studies during post-COVID-19 follow-up, our study is the first to demonstrate an increased risk of subsequent bronchiectasis following COVID-19, with an aHR of 1.30 (95% CI, 1.13–1.50). These findings suggest that bronchiectasis may be a respiratory sequela of COVID-19. Respiratory viruses were found more frequently in bronchiectasis exacerbations and the proposed mechanisms included disturbances in host-defense responses, heightened inflammation, and changes in bacterial virulence [23], which could be involved in the vicious cycle model of development and progression of bronchiectasis, for which the primary insult is often unknown [24]. Our findings accordingly supported that clinicians should be vigilant for this potential complication after COVID-19.

The current study aligns with previous research [16, 22] highlighting the respiratory system's vulnerability to post-COVID-19 conditions. Notably, the increased incidence of new-onset asthma and bronchiectasis observed in the study group suggests a potential link between SARS-CoV-2 infection and the development of obstructive airway diseases. This represents a contradiction with the results of previous studies [25, 26] which primarily focused on restrictive ventilatory defects associated with lung fibrosis. However, further investigation is crucial to

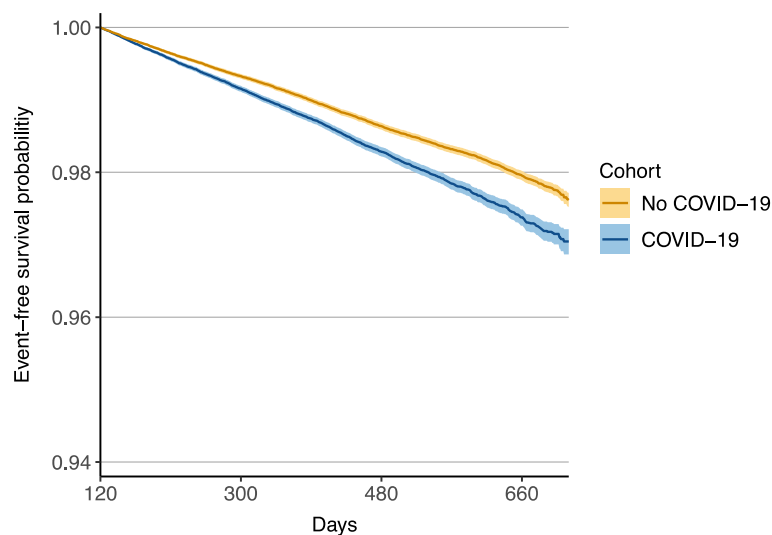


Fig. 2 Kaplan–Meier curves of event-free survival for new-onset asthma (log-rank test $p < 0.001$)

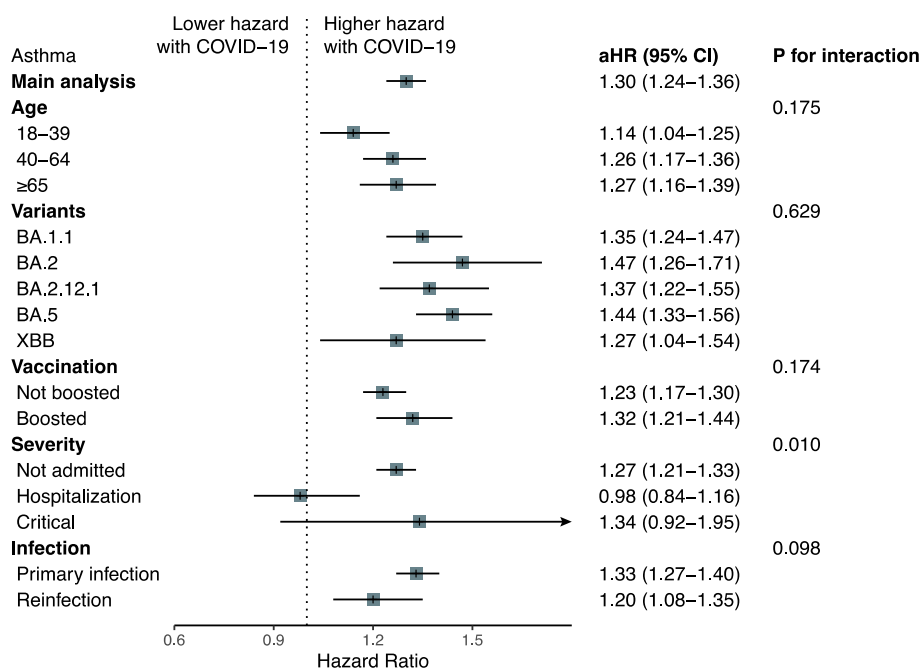


Fig. 3 Subgroup analysis of new-onset asthma between COVID-19 and no-COVID-19 groups

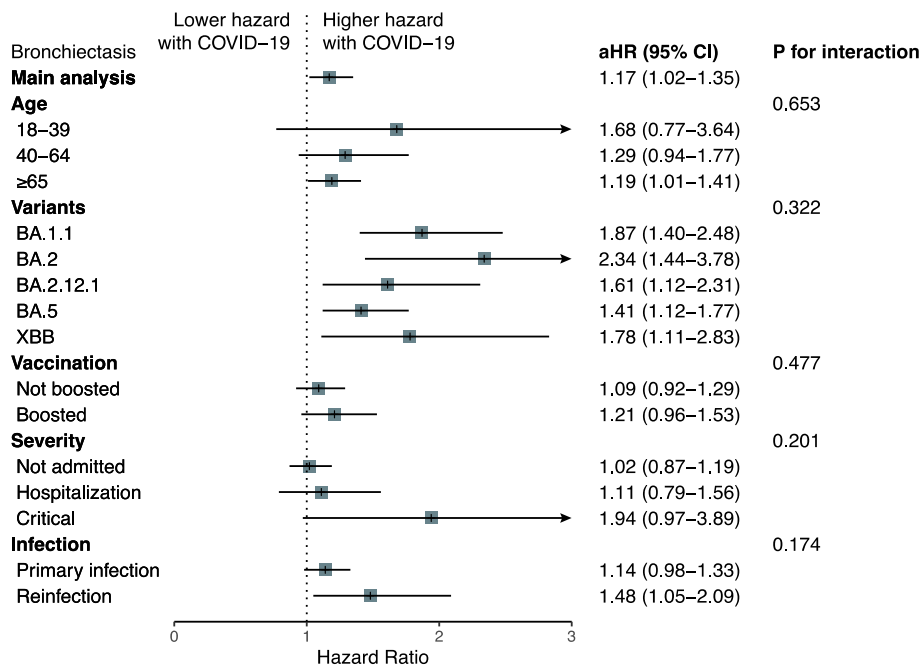


Fig. 4 Subgroup analysis of new-onset bronchiectasis between COVID-19 and no-COVID-19 groups

validate these findings and elucidate the potential mechanisms involved.

This study boasts several strengths. First, the analyses utilized the TriNetX platform, a sizable database, which facilitated the inclusion of a substantial patient cohort.

Second, we conducted numerous stratification analyses and sensitivity tests, with the majority yielding consistent results. Third, although patients with COVID-19 exhibited a higher prevalence of comorbidities such as diabetes mellitus, cardiovascular diseases, and systemic

connective tissue disorders—conditions potentially elevating the risk of obstructive airway diseases—these factors were meticulously adjusted in relation to study outcomes. The application of PSM effectively equalized baseline characteristics between the COVID-19 and control groups, thus bolstering the study's internal validity.

Despite the study's strengths, we must acknowledge certain limitations. The reliance on electronic health records and administrative data raises the possibility of misclassification and underreporting. Furthermore, the study does not investigate potential mechanisms behind the observed associations, which leaves a gap for future research to clarify the pathophysiological connections between SARS-CoV-2 infection and obstructive airway diseases. Long-term patient follow-up is also essential to comprehend the persistence and progression of these respiratory conditions over time. Meanwhile, FEV₁ and disease staging data for asthma or COPD were not documented for most patients in the TriNetX database, thereby introducing potential information bias and precluding further analysis on these variables. Finally, the possibility of residual confounding could not be fully eliminated though we had accounted for numerous clinically significant confounders in our analyses.

Conclusions

This study provides essential insights into the long-term respiratory consequences of COVID-19, highlighting the need for continued monitoring and care for individuals recovering from SARS-CoV-2 infection. The associations found with new-onset asthma and bronchiectasis underscore the importance of a multidisciplinary approach to post-COVID-19 care, with a particular focus on respiratory issues. Further research is necessary to investigate the underlying mechanisms and to develop therapeutic strategies to mitigate the effects of SARS-CoV-2 on the respiratory system.

Methods

Data source

The present study utilized data from the US Collaborative Network of the TriNetX database, which gathered de-identified patient-level information from electronic health records. A health care organization (HCO) typically referred to an academic healthcare center that compiled data from its associated facilities, including the main and satellite hospitals, and outpatient clinics. The collected data encompassed patient demographics, clinical diagnoses (coded using ICD-10-CM), medical procedures (categorized according to ICD-10-PCS or Current Procedural Terminology), medications (coded based on the Veterans Affairs Drug Classification

System and RxNorm medication codes), laboratory tests (organized with Logical Observation Identifiers Names and Codes), and records of healthcare utilization. The US Collaborative Network contained data from over 100 million patients across 61 HCOs in the United States. Analysis of patient-level data was conducted on the TriNetX platform, and the results were provided to researchers in a summarized format. The TriNetX database performs intensive data preprocessing procedures to minimize missing values and maps the data to a consistent framework. Details regarding the database can be found online [21].

The current study's use of the TriNetX database received ethical approval from the Chi-Mei Hospital's Institutional Review Board (no: 11202–002). We conducted the study in accordance with the Declaration of Helsinki and reported our findings following the Strengthening the Reporting of Observational Studies in Epidemiology (STROBE) guidelines.

Study population

Participants of this study were ≥ 18 -year-olds and visited HCOs ≥ 3 times after January 8, 2022. This date coincides with the emergence of the Omicron variant as the predominant strain in the United States [26], ensuring that the data for our analysis were collected within a comparable timeframe. We limited our inclusion to those vaccinated against COVID-19 to reflect the high vaccination rates among U.S. adults, which had surpassed 90% in U.S. adults [27], and to minimize biases related to health-seeking behavior using vaccination status as a proxy. To avoid misinterpretation, we excluded individuals with pre-existing diagnoses of asthma, chronic obstructive pulmonary disease (COPD), bronchiectasis, or forced expiration in the first second (FEV₁)/forced vital capacity (FVC) $< 70\%$ before the index date. The index date was defined as the date of COVID-19 diagnosis for the COVID-19 group and the date of the first visit to HCOs during the inclusion period for the no-COVID-19 group.

The enrollees were further divided into those diagnosed with COVID-19 (the COVID-19 group) during the study timeframe and those who never had COVID-19 (the no-COVID-19 group). One-to-one propensity score matching (PSM) was conducted, involving 34 variables, including demographics, comorbidities, and medication usage, to balance the two groups. Both groups were followed for a maximum of 2 years or until the date of data analysis on January 17, 2024. Details regarding the codes used to identify demographics, diagnoses, procedures, and medications are provided in Additional file 1: Table S2.

Covariates

In the current analysis, we considered the following variables to balance baseline characteristics between the COVID-19 and non-COVID-19 groups: age, sex, race, diabetes mellitus (DM), dyslipidemia, cardiovascular diseases (hypertension, ischemic heart diseases, heart failure, peripheral vascular disease, cerebrovascular diseases), dementia, chronic kidney disease (CKD), hepatitis, cirrhosis, autoimmune diseases (systemic connective tissue disorders, rheumatoid arthritis), anemia, neoplasms, human immunodeficiency virus (HIV) disease, atopic dermatitis, allergic rhinitis, tobacco use/nicotine dependence, alcohol-related disorders, potential health hazards related to socioeconomic and psychosocial circumstances, and BMI (≥ 30 kg/m²). We also included medications that could alter allergic or inflammatory responses or are potentially associated with bronchospasm (glucocorticoids, antihistamines, beta-adrenergic blockers, aspirin and other non-steroidal anti-inflammatory drugs (NSAIDs)). The codes used to define the covariates are provided in Additional file 1: Table S3.

Prespecified outcomes

The primary outcome of this study was the hazard of new-onset asthma during the follow-up period, which commenced 4 months after the index date to preclude the confounding effects of post-viral bronchial hyper-reactivity syndrome. Secondary outcomes encompassed the hazard of bronchiectasis and COPD within the same follow-up timeframe. To ascertain the specificity of our results, we incorporated negative control outcomes, including traumatic intracranial injury, skin cancer, schizophrenia, and cataract. Conversely, the positive control outcome was the post-COVID-19 condition, characterized by a collection of symptoms such as fatigue, headache, dizziness, myalgia, sleep disturbance, emotional distress, cognitive impairment, palpitation, shortness of breath, and changes in bowel habits [28]. Additional file 1: Table S4 contains the definitions and codes pertinent to these outcomes.

Statistical analysis

Baseline characteristics for the COVID-19 and non-COVID-19 groups are presented as means with standard deviations (SD) or as counts and percentages. We compared categorical variables using the χ^2 test and assessed continuous variables with the independent two-sample *t*-test. We conducted one-to-one PSM using the greedy nearest neighbor algorithm with a caliper of 0.1 pooled standardized differences to balance baseline characteristics. Variables were considered adequately matched post-PSM if the standardized difference between groups

was less than 0.1. We computed survival probabilities using the Kaplan–Meier method and calculated adjusted hazard ratios (aHR) for the outcomes using the Cox proportional hazards model, including corresponding 95% confidence intervals (CI) and *p*-values. We tested the proportional hazards assumption with the generalized Schoenfeld residuals method. Outcome variables were categorized as present or absent; thus, missingness was not applicable. We excluded cases lost to follow-up to reduce potential biases or inaccuracies from incomplete data.

We conducted sensitivity analyses using different cutoff landmarks (2, 3, and 6 months post-index date) to initiate follow-up and define new-onset asthma. To investigate potential differences in effect sizes among clinically relevant subgroups, we performed pre-specified subgroup analyses based on age (≥ 18 to <40 , ≥ 40 to <65 , or ≥ 65 years), variants (BA.1.1, BA.2, BA.2.12.1, BA.5, or XBB, corresponding to periods when a variant predominated in the U.S. [27]), booster vaccination status, COVID-19 severity (not admitted, hospitalized, or critical, the latter defined by the need for endotracheal intubation, mechanical ventilation, extracorporeal membrane oxygenation, or intensive care unit admission), and infection status (primary infection or reinfection). All tests were two-sided with a significance threshold of 0.05. Statistical analyses were performed using TriNetX analytics tools and R (version 4.2.2; R Foundation for Statistical Computing, Vienna, Austria).

Abbreviations

aHR	Adjusted hazard ratio
BMI	Body mass index
CI	Confidence interval
COPD	Chronic obstructive pulmonary disease
COVID-19	Coronavirus disease 2019
CKD	Chronic kidney disease
DM	Diabetes mellitus
FEV1	Forced expiration in the first second.
FVC	Forced vital capacity
HCO	Health care organization
HIV	Human immunodeficiency virus
ICD-10-CM	International Classification of Diseases, Tenth Revision, Clinical Modification
ICD-10-PCS	International Classification of Diseases, Tenth Revision, Procedure Coding System
NSAID	Non-steroidal anti-inflammatory drugs
PSM	Propensity score matching
SARS-CoV-2	Severe acute respiratory syndrome coronavirus 2
SD	Standard deviation

Supplementary Information

The online version contains supplementary material available at <https://doi.org/10.1186/s12916-024-03589-4>.

Additional file 1: Table S1. Sensitivity analysis for primary and secondary outcomes with different cutoff landmarks to define new-onset obstructive airway diseases and the start of follow-up. Table S2. Demographic, diagnostic, procedural, medication, visit, and laboratory codes used in

the definition of the cohort. Table S3. Demographic, diagnostic, laboratory, and medication codes used in the definition of covariates. Table S4. Diagnostic codes used in the definition of outcomes. FigS1. Results for negative and positive control outcomes.

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Authors' contributions

MHC conceptualized, designed the study performed the data analysis and drafted the manuscript. WH, YWT, WHH, JYW, THL and PYH assisted data correction and created figures. MHC, and CCL contributed to project design and edited the manuscript. MHC and CCL was responsible for the data interpretation. MHC and CCL finalized the manuscript. All authors approved the final version of the manuscript.

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Availability of data and materials

All data generated or analyzed during this study are included in this published article and will be available upon request to CCL.

Data availability

No datasets were generated or analysed during the current study.

Declarations

Ethics approval and consent to participate

The Institutional Review Board of the Chi Mei Medical Center approved the study protocol (no. 11202–002). Written informed consent was not required because TriNetX contains anonymized data.

Consent for publication

None.

Competing interests

None.

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