REVIEW

Does repetitive transcranial magnetic stimulation have a benefcial efect on improving unilateral spatial neglect caused by stroke? A meta‑analysis

RuixuanLin¹ · Jack Jiaqi Zhang¹ © · Lingling Zhong¹ · Sofina S. Y. Chan¹ · Patrick W. H. Kwong¹ · Lukas Lorentz² · Usman Jawed Shaikh² · Tommy L. H. Lam³ · David M. A. Mehler^{4,5} · Kenneth N. K. Fong¹

Received: 18 June 2024 / Revised: 23 July 2024 / Accepted: 27 July 2024 / Published online: 28 August 2024 © The Author(s) 2024

Abstract

This review aimed to assess the efect of repetitive transcranial magnetic stimulation (rTMS) in improving post-stroke unilateral spatial neglect (USN) using a meta-analysis. Further, we aimed to identify any association between rTMS parameters, patient demographics, and treatment efect sizes using subgroup analyses and meta-regression. A literature search was conducted through four databases from inception to March 6, 2024, to retrieve all relevant controlled trials investigating the efects of rTMS on symptoms of USN in post-stroke patients. Overall, rTMS signifcantly improved post-stroke USN, as measured by the line bisection test (Hedges' $g = -1.301$, $p < 0.0001$), the cancelation test (Hedge's $g = -1.512$, $p < 0.0001$), and the Catherine Bergego Scale (Hedges' $g = -0.770$, $p < 0.0001$), compared to sham stimulation. Subgroup analysis found that generally larger efect sizes following excitatory rTMS across several outcome measures, indicating that excitatory rTMS on the ipsilesional hemisphere may be more efective than inhibitory rTMS on the contralesional hemisphere in ameliorating neglect symptoms. Meta-regression analysis of the line bisection test showed a signifcant diference in the chronicity of stroke patients, suggesting that rTMS may be more efective for USN in patients at the acute stage (within 3 months since stroke) than in those at the post-acute stage $(p=0.035)$. In conclusion, rTMS appears to be effective in promoting recovery from post-stroke USN. Excitatory protocols and early intervention may enhance recovery outcomes for neglect behaviors in post-stroke survivors.

Keywords Stroke · Unilateral spatial neglect · Repetitive transcranial magnetic stimulation · Meta-analysis · Rehabilitation

Introduction

Unilateral spatial neglect (USN) is a neuropsychological disorder manifested as an attention defcit to visual, auditory, or proprioceptive stimuli from the contralesional hemifeld and is often observed in post-stroke patients [\[1](#page-11-0)], especially in those with right hemispheric lesions [[11](#page-12-0)]. USN is also a strong predictor of neurological recovery after stroke, as recovery of cognitive and motor functions in post-stroke patients with USN is worse than in those without USN [[9,](#page-12-1) [13](#page-12-2)]. USN is caused by disruption of the cortical attention network owing to brain injury, which consists of the dorsal attention network (DAN) and ventral attention network (VAN) [\[12](#page-12-3)]. The DAN includes the posterior parietal cortex (PPC) and frontal eye felds, which guide visual–spatial and visual-driven attentions. The VAN includes the right inferior frontal gyrus and right temporoparietal junction, which reorient attention to stimuli-driven covert visual-spatial attention. Damage to either area was associated with the occurrence of USN [\[10](#page-12-4)]. Researchers have proposed a model of spatial attention disturbance in patients with neglect, suggesting that they have an attentional bias toward the side of the space opposite the brain lesion. This bias is caused by an imbalance between the two attention-directing processes controlled by the right and left hemispheres. In a healthy brain, the competition between both hemispheres is mediated by inhibitory connections across the midline of the brain [\[5](#page-12-5)]. However, according to the imbalance hypotheses, disruption of long-range projections due to a brain lesion results in an imbalance that leads to a shift in attention and gaze toward the side of the brain lesion [\[29](#page-12-6)].

The mainstream rehabilitative intervention methods for post-stroke USN include prism adaptation (PA), visual scanning, optokinetic stimulation, and mirror therapy [\[45](#page-13-0)]. All these rehabilitative interventions activate the VAN via a bottom–up approach. Contrastingly, non-invasive brain stimulation promotes USN recovery by adjusting the balance of both brain hemispheres via top–down modulation [[26\]](#page-12-7). Transcranial magnetic stimulation (TMS), the most commonly used non-invasive brain stimulation technique in post-stroke rehabilitation, uses the principle of electromagnetic induction to generate an electric current that acts on the neurons of the cerebral cortex, thereby afecting the metabolism and neuroelectric activity of the brain [[31](#page-12-8)]. Repetitive transcranial magnetic stimulation (rTMS) has multifaceted efects on the nervous system. It infuences the release of neurotransmitters and neurotrophic factors [[36\]](#page-12-9) and alters the functional connections between diferent brain regions [\[24](#page-12-10)]. Typically, high-frequency rTMS (HF-rTMS) and intermittent theta-burst stimulation (iTBS) induce an increase in excitability of the stimulated cortex, whereas low-frequency rTMS (LF-rTMS) and continuous theta-burst stimulation (cTBS) inhibit excitability in the stimulated cortex [\[30\]](#page-12-11). In clinical trials, rTMS is frequently used to treat USN by regulating the interhemispheric balance of attention system located in the bilateral parietal areas, through either inhibitory rTMS to the unafected hemisphere or applying excitatory rTMS to the afected hemisphere, thus achieving a balance between the hemispheres to reduce the severity of USN. According to our previous review, rTMS yielded the largest efect size among interventions tailored to unilateral neglect [[45](#page-13-0)].

Previous systematic reviews and meta-analyses have reported the positive efect of rTMS in improving USN in post-stroke patients [[22,](#page-12-12) [26](#page-12-7), [35](#page-12-13), [44,](#page-13-1) [48](#page-13-2)]. However, the possible efficacy-related modulators on the treatment effects were not yet well explored in the previous reviews [[22,](#page-12-12) [26](#page-12-7), [35](#page-12-13), [44,](#page-13-1) [48\]](#page-13-2), which can be investigated by conducting meta-regression and subgroup analyses based on the metaanalytic results. Here, our review aimed to assess the efect of diferent rTMS protocols in improving post-stroke USN using a meta-analysis. Further, we aimed to identify any association between rTMS parameters, patient demographics, and treatment efect sizes using subgroup analyses and meta-regression.

Methods

Data sources and search strategy

The protocol of this review was registered in PROSPERO (CRD42024553592). Randomized controlled trials (RCTs) or nonrandomized controlled trials on the impact of rTMS in patients with USN after stroke were retrieved from Pub-Med, EMBASE, MEDLINE, and Web of Science databases from inception until March 6, 2024. The search strategy was developed as follows: (transcranial magnetic stimulation OR theta-burst stimulation) AND (unilateral neglect OR unilateral spatial neglect OR visuospatial neglect OR visuomotor neglect OR behavioral inattention OR hemispatial neglect) AND (stroke OR cerebrovascular accident OR cerebral infarction OR cerebral hemorrhage).

Inclusion criteria

The inclusion criteria were developed in accordance with the PICOS framework: *Population (P):* adult patients (>18 years old) with post-stroke USN. Intervention (I): using any rTMS protocols (HF-rTMS, LF-rTMS, cTBS, and iTBS). *Comparison (C):* sham stimulation or no rTMS. *Outcomes (O):* Studies must provide outcomes that reflect the severity of USN or the severity of USN afecting activities of daily living. The line bisection test (LBT) and star cancelation test (SCT) are the standard paper–pencil tests most used to assess the severity of USN. The Catherine Bergego Scale (CBS) is a standardized behavioral test of USN. These three tests were selected as primary outcomes of the current meta-analysis. When the LBT was not available, a subtest (bisection test) of the behavioral inattention test (BIT) was used in the meta-analysis. If the SCT was not available, the cancelation subtest (CT) of the BIT and the Albert test (AT), which are similar in nature to the SCT, were used for analysis. *Study design:* RCT or nonrandomized controlled trials.

Exclusion criteria

This meta-analysis included only articles written in English language. Articles that did not use a scale measuring the degree of USN as an outcome, studies without any parallel control group (self-control, cohort studies, case–control studies, cross-sectional studies, animal experiments, expert consensus, conference abstracts, meta-analyses, or reviews), studies that did not use any form of therapeutic TMS protocols, and duplicate publications were excluded from consideration.

Data extraction

Endnote version 21 was used to manage the citations. The extracted data included the frst author, publication year, number of participants, mean age, sex, nature of lesion, stroke duration, treatment protocol, combined intervention, Class of studies, outcome measures, pre- and post-treatment means, mean change score, and standard deviations for outcome measures.

All included studies were classifed into four classes based on the criteria used in the clinical guideline by Lefaucheur et al. [\[34\]](#page-12-14). If the mean and SD were not provided in the study, but the median and range of the score were available, this formulation was used: mean = $(a+2m+b)/4$ (where a is the smallest value, b is the largest value, and m is the median); the standard deviation was calculated using the interquartile range/4 [[23\]](#page-12-15).

Quality and risk‑of‑*bias* **assessment**

The PEDro scale was used to evaluate the reporting quality of the methodology [\[19\]](#page-12-16). The scale consists of 11 items, with the frst item assessing the external validity of a study and the remaining items assessing internal validity. The PEDro Scale was scored from 0 to 10, with higher scores indicating higher research quality. A total score falling within the range of zero to three indicates low methodological quality, while a score of four-to-fve accounts for moderate quality. When the score is between six and eight the quality is considered good. A score of nine or ten suggests excellent quality [[15](#page-12-17)]. Two authors (RL and LZ) independently rated the PEDro scores for each study, and any disagreements were resolved through discussion with a senior author (JZ).

Statistical analysis

Comprehensive meta-analysis (CMA) software (version 3.0) was used for the meta-analysis. The authors were contacted by email when the required data were missing. In case no response could be received from the authors, a data digitizer was used to extract the data obtained in the form of graphs. The data used in this analysis were continuous variables. The change score (postminus pre) and its standard deviation were used to compute the pooled efect size in the form of Hedge's *g* [[39](#page-13-3)]. The 95% confdence interval (CI) was calculated, and the significance level was set to $\alpha = 0.05$. Efect sizes measured by Hedge's *g* values of 0.15, 0.40, and 0.75 are interpreted as indicating small, medium, and large effects, respectively [\[3](#page-11-1)]. The q -statistics and I^2 indices were used to evaluate the heterogeneity of each efect size. The random-efects model was used in all meta-analyses because of the signifcant clinical and statistical heterogeneity among the studies [[2\]](#page-11-2). Due to the lack of follow-up data and different follow-up times in many of the included articles, we only focused on the efect of rTMS post-intervention in this meta-analysis. Subgroup analysis was performed according to the diferent rTMS protocols, Class of studies, and the time since stroke onset. We used 3 months as the cutoff, because the spontaneous biological recovery of USN is dominant within the frst 3 months after stroke [[33,](#page-12-18) [37](#page-12-19)], and this subgroup analysis was carried out to explore the possible diferential efects between TMS applied on top of the spontaneous biological recovery trend post-stroke (within 3 months) and TMS applied when spontaneous biological recovery is largely diminished (over 3 months). Meta-regression was performed to investigate the association between various predictors and efect sizes (Hedges' g). Predictors included in the meta-regression were mean age, sex (expressed as the percentage of female patients), the percentage of ischemic stroke patients, the baseline severity (baseline test score), and the rTMS parameters (the total number of applied pulses, the number of sessions, and the applied pulses each session). Sensitivity analysis was performed using the leave-one-out method to test the robustness of signifcant fndings. Publication bias was investigated using funnel plots and the Egger's regression test, with a significant level of $p=0.1$.

Results

Study selection

The search strategy identifed 466 records in the database. After screening the records and according to the inclusion and exclusion criteria, 18 articles that met the inclusion criteria were included in the systematic review [[4](#page-12-20), [6](#page-12-21)[–8,](#page-12-22) [16,](#page-12-23) [17](#page-12-24), [21](#page-12-25), [25](#page-12-26), [27](#page-12-27), [28](#page-12-28), [32](#page-12-29), [40](#page-13-4)[–42](#page-13-5), [46](#page-13-6), [47](#page-13-7), [49](#page-13-8)]. Two studies were excluded from the meta-analysis, because the data were not available; therefore, 16 studies were included in the metaanalysis [\[4](#page-12-20), [6](#page-12-21)[–8](#page-12-22), [17](#page-12-24), [25,](#page-12-26) [27,](#page-12-27) [28,](#page-12-28) [32,](#page-12-29) [40–](#page-13-4)[42](#page-13-5), [46](#page-13-6), [47](#page-13-7), [49](#page-13-8)]. The characteristics of the included studies are summarized in Table [1](#page-3-0). The flowchart of the study selection is shown in Fig. [1.](#page-6-0)

Methodological quality assessment

The quality of the included articles was rated using the PEDro scale (Table S1). The mean score of the 18 articles

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Fig. 1 Flowchart of literature search

was 8.61, ranging from 5 to 10. This indicated that the included articles were of moderate-to-high quality.

rTMS protocols

Most studies $(n=9)$ used LF-rTMS, four of which used the 1-Hz frequency for the P3 based on the EEG 10–20 system [\[7](#page-12-30), [8](#page-12-22), [27](#page-12-27), [47\]](#page-13-7) and the other two also used the same frequency while applying it to the contralesional angular gyrus $[25]$, or P5 $[28, 46]$ $[28, 46]$ $[28, 46]$ $[28, 46]$. Other three studies also chose to use LF-rTMS, but they applied diferent frequencies: two studies used 0.5 Hz rTMS to the P3 [[40,](#page-13-4) [49\]](#page-13-8). One study applied 0.9 Hz rTMS to the P5 [[28\]](#page-12-28). High-frequency rTMS (10 Hz) was used in two studies, separately acting on the P3 [[47\]](#page-13-7) or P4 [\[27](#page-12-27)]. cTBS was used in nine studies, the stimulation target was set at the P3 in eight studies [[6,](#page-12-21) [16,](#page-12-23) [21](#page-12-25), [32](#page-12-29), [41,](#page-13-9) [42](#page-13-5), [47\]](#page-13-7), and one study chose the P5 as the stimulation target [[17\]](#page-12-24). Only one study [\[4](#page-12-20)] used the iTBS protocol for the left DLPFC, which was localized by the F5 channel in the EEG 10–20 system.

Fig. 2 Meta-analysis for LBT

Study name		treatment Outcome	Statistics for each study			Hedges's g and 95% CI
Excitatory protocosis			Hedges's g	Lower limit	Upper limit	
Yang et al. 2015	HF rTMS	LBT	-0.886	-1.769	-0.002	
Kim et al. 2013	HF rTMS	LBT	-3.220	-4.592	-1.849	
Cao et al. 2016	iTBS	LBT	-1.853	-3.092	-0.613	
			-1.906	-3.241	-0.572	
Inibitory protocol						
Koch et al. 2012	cTBS	BIT-LB	-0.713	-1.581	0.155	
Fu et al. 2015	cTBS	LBT	-0.591	-1.450	0.268	
Vatanparas et al. 2023 cTBS		LBT	-0.514	-1.513	0.485	
Yang et al. 2015a	cTBS	LBT	-0.888	-1.793	0.017	
Yang et al. 2015b	LF rTMS	LBT	-0.886	-1.791	0.019	
Kim et al. 2018	LF rTMS	LBT	-0.353	-1.200	0.493	
Cha et al. 2015	LF rTMS	LBT	-1.843	-2.993	-0.692	
Cha et al. 2016	LF rTMS	LBT	-2.240	-3.138	-1.342	
Kim et al. 2013 a	LF rTMS	LBT	-2.890	-4.180	-1.599	
Song et al. 2009	LF rTMS	LBT	-1.126	-2.191	-0.060	
Zhang et al. 2013	LF rTMS	LBT	-1.483	-2.274	-0.692	
			-1.171	-1.602	-0.741	
Random-effects model			-1.301	-1.718	-0.884	
Heterogeneity: I ² =60.654%						0.00 -8.00 -4.00 4.00 8.00
Test for overall effect: $Z = -6.117$ (P<0.001)						
						control TMS

Efect of rTMS on unilateral spatial neglect

Line bisection test

Eleven studies with 14 units of analysis were included in this meta-analysis of LBT [[4](#page-12-20), [7,](#page-12-30) [8](#page-12-22), [17,](#page-12-24) [27](#page-12-27), [28,](#page-12-28) [32](#page-12-29), [40\]](#page-13-4) (Fig. [2](#page-7-0)). An overall improvement in the LBT score was found in the rTMS group compared with the control group (Hedges' $g = -1.301$, $p < 0.0001$, $I^2 = 60.65\%$), and the overall signifcance was robust to the leave-one-out sensitivity analysis (Hedge's g from -0.884 to -1.718 , which consistently indicated a large efect size in our sensitivity analysis). Subgroup analyses showed that the excitatory and inhibitory rTMS protocols both signifcantly improved USN. The pooled effect size for excitatory rTMS stroke acute was

numerically larger than the inhibitory rTMS (excitatory: Hedge's $g = -1.906$, $p = 0.005$, $I^2 = 75.05\%$; inhibitory: Hedges' $g = -1.171$, $p < 0.0001$, $I^2 = 56.33\%$). There was sign of publication bias according to the signifcant results of Egger's test ($p = 0.010$) (Figure S1). Univariate meta-regression analysis (Fig. [3\)](#page-7-1) showed that the post-stroke period of patients (acute or post-acute) was a signifcant predictor of the efect size, and rTMS appeared to be more efective for patients in the acute phase $(p=0.035)$ than those in the post-acute phase. Other predictors were not signifcant in the meta-regression. The subgroup analysis based on the classes of studies shows no signifcant diference in the efect sizes between subgroups ($Q=1.07$, $p=0.586$). Table [2](#page-8-0) summarizes the results of the univariate meta-regression.

Fig. 3 Between-group diferences in the efect sizes of LBT in acute and postacute patients

scores

Table 2 Results of meta-

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* *p*<0.05; ***p*<0.01; ****p*<0.001

Heterogeneity: I²= 66.315 Test for overall effect: Z=-6.344 (P<0.001)

Fig. 4 Meta-analysis for CT

Cancelation test

Eleven studies with 14 units of analysis were included in the meta-analysis of the cancelation test scores [[4,](#page-12-20) [8](#page-12-22), [17](#page-12-24), [27,](#page-12-27) [28,](#page-12-28) [32](#page-12-29), [40,](#page-13-4) [42](#page-13-5), [46](#page-13-6), [47,](#page-13-7) [49\]](#page-13-8) (Fig. [4\)](#page-8-1). The results indicated that the rTMS group showed a signifcant improvement in scores compared with the control group (Hedge's $g = -1.512$, $p < 0.0001$, $I^2 = 66.32\%$). The overall effect was still robust in the leave-one-out sensitivity analysis (Hedge's *g* from – 1.979 to – 1.405, which consistently indicated a large effect size in our sensitivity analysis). There was a sign of publication bias according to the signifcant results of

TMS

Control

the efect sizes of rTMS in CT scores

Table 3 Results of metaregression of moderators for

Egger's test $(p=0.020)$ (Fig. S2). The subgroup analysis showed that the excitatory rTMS group had a larger efect size than the inhibitory rTMS subgroup, and both were statistically signifcant (excitatory: Hedge's *g*=– 2.497, $p = 0.001$, $I^2 = 69.22\%$; inhibitory: Hedges' $g = -1.305$, $p < 0.0001$, $I^2 = 61.68\%$). The subgroup analysis based on the classes of studies shows no signifcant diference in the effect sizes between subgroups ($Q = 2.18$, $p = 0.336$). Table [3](#page-9-0) summarizes the results of the univariate meta-regression.

Catherine Bergego scale

In total, six articles with eight units of analysis were included in this meta-analysis of CBS [[6](#page-12-21), [21,](#page-12-25) [25](#page-12-26), [27,](#page-12-27) [28,](#page-12-28) [46](#page-13-6)] (Fig. [5\)](#page-10-0). The rTMS group showed a signifcant improvement in the CBS score compared with the control group (Hedges' $g = -0.770$, $p < 0.0001$, $I^2 = 35.40\%$), and this overall signifcance remained robust in the leaveone-out sensitivity analysis (Hedge's *g* from – 0.379 $to -1.161$, which showed a medium-to-large effect size in our sensitivity analysis). There was a sign of publication bias according to the signifcant results of Egger's test $(p=0.010)$ (Fig. S3). Subgroup analysis revealed that both excitatory and inhibitory rTMS protocols led to signifcant improvements in the behavioral test results of USN. However, the efect size for excitatory rTMS was numerically larger than that for inhibitory rTMS (Excitatory: Hedges' $g = -2.215$, $p = 0.002$, $I^2 = 0.00\%$; inhibitory: Hedges' $g = -0.618$, $p < 0.0001$, $I^2 = 0.00\%$). The subgroup analysis based on the classes of studies shows no signifcant difference in the effect sizes between subgroups $(Q=4.87,$ $p=0.431$). Table [4](#page-10-1) summarizes the results of the univariate meta-regression.

Combined intervention

In total, 11 out of 18 studies included in this review utilized a combination of conventional rehabilitation and rTMS. However, other studies have opted for different approaches, such as smooth pursuit eye movement training [[21\]](#page-12-25), robot therapy for simultaneous visual scanning and limb activation [[28\]](#page-12-28), and sensory cueing using a wearable device [[46](#page-13-6)]. Additionally, two studies employed PA [\[41,](#page-13-9) [42\]](#page-13-5). All these interventions demonstrated effectiveness when combined with rTMS to improve USN.

Level of recommendation

We summarized the level of recommendation of the efficacy of different rTMS protocols in treatment of poststroke neglect, following the definition of Lefaucheur et al. [[34\]](#page-12-14). The level of recommendation in the efficacy of LF-rTMS has now reached level A, with two Class I studies [\[46,](#page-13-6) [47](#page-13-7)], six Class II studies [\[7,](#page-12-30) [8,](#page-12-22) [25,](#page-12-26) [27](#page-12-27), [28,](#page-12-28) [49](#page-13-8)], and one Class III study [[40\]](#page-13-4). The level of recommendation in the efficacy of cTBS has now reached Level B, with one Class I study [\[47\]](#page-13-7), five Class II studies [[6](#page-12-21), [17,](#page-12-24) [32,](#page-12-29) [42](#page-13-5)], and three Class III studies [\[16,](#page-12-23) [21](#page-12-25), [41](#page-13-9)]. The level of recommendation in the efficacy of HF-rTMS has now reached level C, with one Class I study [\[47\]](#page-13-7) and one Class II study [[27](#page-12-27)]. No recommendation for iTBS can be made as only one experiment [\[4](#page-12-20)] was available. The two excitatory proposals (HF-rTMS and iTBS) have not received high recommendation levels due to the limited number of studies at this stage; however, the numerically larger effect size observed in meta-analyses to some extent can support that the excitatory protocols have great potential in the treatment of post-stroke USN.

Fig. 5 Meta-analysis for CBS

Table 4 Results of metaregression of moderators for the efect sizes of rTMS in CBS scores

Discussion

Our study found that (1) rTMS was signifcantly efective in improving post-stroke USN compared with the control group. (2) Excitatory rTMS appears to be more efective than inhibitory rTMS, with a numerically larger efect size than that of inhibitory protocols; however, a few studies have utilized excitatory rTMS and the level of recommendation is thereby lower when compared to inhibitory protocols. (3) rTMS seems to efectively improve neglect behaviors during daily activities in post-stroke patients. (4) A signifcant diference was found between the chronicity of stroke patients and the efect size of rTMS in the LBT meta-regression, indicating that the timing of stroke may be a factor influencing the efficacy of rTMS, with patients in the acute phase (within the frst 3 months) potentially benefting more from the non-invasive brain stimulation therapy.

Recovery from neglect after stroke depends on neuroplasticity [\[18](#page-12-31)]. Spontaneous biological recovery, which is dominant with the frst three months after stroke, signifcantly contributes to the spontaneous recovery from USN. Our analysis suggested that the timing of rTMS intervention may be a potential factor infuencing the outcome of USN after stroke. Our analysis showed that rTMS delivered within the frst three months post-stroke demonstrated a signifcantly stronger efect on facilitating the recovery from USN, compared with rTMS applied after the frst three months. The results indicated that rTMS may have an add-on efect on spontaneous biological

recovery from USN in post-stroke survivors. However, this signifcant diference was only observed in the meta-analysis of LBT, perhaps due to the diferences among the neglect measures [[20\]](#page-12-32). LBT necessitates the correct perception of the size of a single stimulus, while CT depends on a normal visual search within an array of various stimuli [\[14\]](#page-12-33). CBS is related to daily activities; therefore, it requires higher cognitive function. This fnding may suggest that the add-on beneft from rTMS on spontaneous recovery from USN is more specifc to basic attention to a single stimulus.

Functional connectivity between the bilateral attention systems is assumably disrupted in post-stroke patients with USN [[38\]](#page-12-34). Inhibitory protocols applied to the contralesional attention system can suppress the interhemispheric inhibition from the contralesional to the ipsilesional attention system, therefore facilitating recovery from USN. However, its efect depends on intact interhemispheric connectivity [\[37](#page-12-19)]. The efect of inhibitory rTMS may be therefore limited in patients with an injured corpus callosum, although we were unable to perform a quantitative analysis based on the integrity of the corpus callosum, because this information was not usually reported in the previous trials. In contrast, excitatory rTMS can promote the activation of adjacent functional areas through the parieto-frontal attention network in the afected hemisphere, such as the temporoparietal junction and the inferior frontal gyrus [\[27\]](#page-12-27). Besides, excitatory rTMS on the same hemisphere may not only promote the recovery of visual-spatial attention, but also potentially facilitate the recovery of other non-spatial functions, such as alertness and novelty detection [\[43](#page-13-10)]. Therefore, the excitatory protocol might yield a more consistent treatment response in USN.

Limitations

This study also has some limitations. First, the rTMS protocols used in the included studies varied in treatment duration, target points, frequency, etc., which may have led to high heterogeneity. Second, in this meta-analysis, diferent scaling methods used for the same test were not considered for grouping. For example, LBT can be performed using both fve- and one-line methods. Additionally, due to the unavailability of information on lesion location and specifc types of USN for all participants, further analyses could not be conducted. Third, there is a risk of publication bias according to our analysis. In addition, the small sample size of the included studies, as well as the exploratory nature and lack of follow-up data, may downgrade the level of evidence in this feld. Future pre-registered studies using large sample sizes are needed to verify the current fndings. Finally, our meta-analyses were limited to published aggregate data. Mega-analyses using individual data will allow further investigation but require data sharing, for instance through large consortia such as ENIGMA.

Conclusion

rTMS has shown promise as a potential treatment for facilitating recovery from post-stroke USN. Furthermore, generally larger efect sizes following excitatory rTMS across several outcome measures suggest that excitatory rTMS on the ipsilesional hemisphere may be more effective than inhibitory rTMS on the contralesional hemisphere in ameliorating neglect symptoms. Early delivery of rTMS treatment may yield a more favorable recovery outcome in neglect behaviors in post-stroke patients by accelerating the spontaneous recovery from USN within the frst 3 months.

Supplementary Information The online version contains supplementary material available at<https://doi.org/10.1007/s00415-024-12612-w>.

Author contributions Study objective: JZ and RL. Literature search: RL, JZ, and LZ. Data extraction: RL, JZ, and LZ. Methodological quality assessment: RL, JZ, SC, and LZ. Critical review and approval of manuscript: JZ, SC, PK, LL, US, TL, DM, and KF. All authors read and approved the fnal manuscript.

Funding Open access funding provided by The Hong Kong Polytechnic University. This study was partially supported by the Start-Up Fund for RAPs under the Strategic Hiring Scheme (P0048866) to JZ.

Data availability Data supporting the fndings of this study are available from the corresponding author on reasonable request.

Declarations

Conflicts of interest None.

Ethics approval Not applicable. The current study is a review.

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