RESEARCH

Open Access

Multistate Markov chain modeling for child undernutrition transitions in Ethiopia: a longitudinal data analysis, 2002–2016

Check for updates

Getnet Bogale Begashaw^{1,2*}, Temesgen Zewotir³ and Haile Mekonnen Fenta^{1,4,5}

Abstract

Background The use of the multistate Markov chain model is a valuable tool for studying child undernutrition. This allows us to examine the trends of children's transitions from one state to multiple states of undernutrition.

Objectives In this study, our objective was to estimate the median duration for a child to first transition from one state of undernutrition to another as well as their first recurrence of undernutrition and also to analyze the typical duration of undernourishment. This involves understanding the central tendency of these transitions and durations in the context of longitudinal data.

Methods We used a longitudinal dataset from the Young Lives cohort study (YLCS), which included approximately 1997 Ethiopian children aged 1–15 years. These children were selected from five regions and followed through five survey rounds between 2002 and 2016. The surveys provide comprehensive health and nutrition data and are designed to assess childhood poverty. To analyze this dataset, we employed a Markov chain regression model. The dataset constitutes a cohort with repeated measurements, allowing us to track the transitions of individual children across different states of undernutrition over time.

Results The findings of our study indicate that 46% of children experienced concurrent underweight, stunting, and wasting (referred to as USW). The prevalence of underweight and stunted concurrent condition (US) was 18.7% at baseline, higher among males. The incidence density of undernutrition was calculated at 22.5% per year. On average, it took 3.02 months for a child in a wasting state to transition back to a normal state for the first time, followed by approximately 3.05 months for stunting and 3.89 months for underweight. It is noteworthy that the median duration of undernourishment among children in the US (underweight and stunted concurrently) state was 48.8 months, whereas those concurrently underweight and wasting experienced a median of 45.4 months in this state. Additionally, rural children (HR = 1.75; 95% CI: 1.53–1.97), those with illiterate fathers (HR = 1.50; 95% CI: 1.38–1.62) and mothers (HR = 1.45; 95% CI: 1.02–3.29), and those in households lacking safe drinking water (HR = 1.70; 95% CI: 1.26–2.14) or access to cooking fuel (HR = 1.95; 95% CI: 1.75–2.17) exhibited a higher risk of undernutrition and a slower recovery rate.

Conclusions This study revealed that rural children, especially those with illiterate parents and households lacking safe drinking water but cooking fuels, face an increased risk of undernutrition and slower recovery.

*Correspondence: Getnet Bogale Begashaw Getnetbogale145@gmail.com; getnetbogale@dbu.edu.et Full list of author information is available at the end of the article



© The Author(s) 2024. **Open Access** This article is licensed under a Creative Commons Attribution-NonCommercial-NoDerivatives 4.0 International License, which permits any non-commercial use, sharing, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if you modified the licensed material. You do not have permission under this licence to share adapted material derived from this article or parts of it. The images or other third party material in this article are included in the article's Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit http://creativecommons.org/licenses/by-nc-nd/4.0/.

Keywords Lifetime Undernourished Period, Median First Passage Time, Median Recurrence Time, Transition Intensities, Transition Probability, Young Lives Data

Introduction

The prevalence of undernutrition has shown a decline in Asia from 33% to 27%, in Latin America from 33% to 22%, and in Europe from 20% to 16% [1–3]. However, chronic hunger continues to be an issue in Africa, impacting approximately 50% of children under the age of five. Notably, half of all undernourished children are found in seven specific countries: Nigeria, the Democratic Republic of the Congo (DRC), Ethiopia, India, Pakistan, Bangladesh, and China [2, 4, 5].

The under5 mortality rate is a global concern, with sub-Saharan Africa (sSA) having the highest rate of 1 child out of 13 deaths before the fifth birthday [6]. sSA and central and southern Asia accounted for more than 80% of the under5 deaths in 2019, with half occurring in five countries: Nigeria, India, Pakistan, the Netherlands, and Ethiopia [7]. DRC has a high stunting rate (42% of children under five) in sSA, with undernutrition being the primary cause of nearly half of these deaths [8]. Child undernutrition is a serious health concern in Ethiopia, a developing nation in Africa's Horn [7, 9, 10]. Moreover, a lack of nutrients has an impact on a child's growth, development, and overall health. For several years, Ethiopia has struggled with high rates of child undernutrition [11].

Child undernutrition is a significant issue that has severe consequences, such as increased morbidity, mortality risk, impaired cognitive development, reduced school performance, and decreased productivity in adulthood [7, 10, 12]. Factors contributing to this risk include inadequate dietary intake, poor feeding practices, low household income, food insecurity, limited healthcare access, inadequate sanitation, and lack of education. This long-term issue perpetuates poverty and underdevelopment in communities and nations [9, 12].

The World Health Organization (WHO) [13] defines severe thinness in children as a body mass index (BMI) less than 16 kg/m2 or a Z score less than -3, indicating low undernourishment. Undernourished individuals had Z scores between -3 and -2 (SD) or a BMI between 16 kg/ m^2 and 18.49 kg/m².

Ethiopia's child undernutrition is a significant issue due to high rates of stunting (38-57%), wasting (10-17%), and underweight (22-38%) in children under five years of age [14-16].

Traditional statistical models, such as multivariate logistic regression [17], spatial [18], longitudinal [19] and many others, have limitations in predicting future disease states based on the current state and fail to

account for the complexity of state transitions. They typically do not account for crucial factors such as the duration of stay, periods of undernourishment, and key metrics like Median First Passage Time (the median time until the first transition into a specific state) and Median First Recurrence Time (the median time until returning to a specific state after leaving it). These limitations highlight the need for advanced models that can better capture the dynamic nature of state transitions in predicting disease progression.

However, the multistate Markov chain model improves disease prediction by quantifying transition probabilities and identifying median transit and recurrence times required for first-time transitions between states. Additionally, it considers influencing factors, which improves risk factor identification and enables targeted interventions to improve child nutritional outcomes [20, 21]. It accurately estimates transition probabilities, aiding in understanding patient progress, intervention effectiveness, and developing nutritional status strategies, especially in developing countries.

In recent years, the study of child undernutrition has increasingly focused on understanding the dynamics of state transitions. Key to this understanding are metrics such as median first passage time, which provides insights into how quickly children transition into undernourished states for the first time [22–26], and median recurrence time, which helps to understand the average time until a child returns to a previous health state after leaving it [27–29]. Additionally, the duration of undernourishment indicates the length of time children remain in these vulnerable states [30]. Together, these measures offer a comprehensive view of the dynamics of undernutrition, enabling researchers to identify critical intervention points and the factors that influence these transitions.

Therefore, this study used a multistate Markov chain model to analyze child undernutrition in Ethiopia, focusing on state transitions (such as normal, underweight, stunted, wasted, and their combinations). It identifies critical periods of transition and the factors influencing them, including socioeconomic, demographic, maternal and caregiving factors. The model estimates transition probabilities between undernutrition states, providing insights into children's health outcomes over time.

The primary goal of this research is to investigate the transitions between different states of nutritional deficiency in children. Specifically, we aim to understand how children move between various nutritional states over time by analyzing the median transit and recurrence times between these states. Additionally, we seek to quantify the median duration that children spend in undernourished states and determine the median duration of these states. We also seek to identify covariates that play significant roles in these transitions. By examining these metrics, we aim to gain insights into the complex patterns of nutritional transitions and to identify key covariates that significantly influence these transitions. This study contributes to enhancing our understanding of the factors affecting child nutrition and informs strategies for intervention and prevention.

This study is the first to examine the median transit and recurrence times between undernourished states, as well as the median duration a child spends in undernourished states, and the identification of covariates significantly associated with these transitions in Ethiopia.

Methods

Data source and study participants

The Young Lives cohort study (YLCS) evaluates the impact of poverty reduction policies and interventions on children's lives in low- and middle-income countries, including Ethiopia. Using an observational cohort design, the cohort included approximately 12,000 boys and girls aged from infancy to adulthood from India, Peru, Vietnam, and Ethiopia. There were two cohorts, a young cohort (aged 1 to 15) and an old cohort (aged 8 to 22) in the survey in all countries. However, for this study, we used only young cohort data collected from Ethiopia [31]. In this cohort, 1997 Ethiopian children were admitted to the study. The children were selected from five regions (Amhara, Oromiya, Tigray, SNNP (Southern Nations, Nationalities, and Peoples' Region), and Addis Ababa CA (city administration)), including urban and rural areas, and followed them longitudinally.

Variables in the study

The data provider used WHO standards to standardize Z scores for each observation and defined children's anthropometric status as underweight, stunted, or wasted as outcome variables. We proceeded with the analysis after cleaning and making necessary adjustments to the dataset. Therefore, the state variables are shown in Table 1 and are classified as follows: normal, underweight only, stunted only, wasted only, underweight and stunted, underweight and wasted, stunted and wasted, and underweight, stunted and wasted concurrently [32]. Children who have a height-for-age Z-score (HAZ), a weight-forheight z-score (WHZ), and weight for age z-score (WAZ) which is below two are defined as having stunting, wasting and underweight respectively [33–35].

The following figure clearly demonstrates the study's covariates and their connection with the state variable (Fig. 1). However, it is important to note that certain variables widely discussed in literature, such as breastfeeding duration, parental smoking habits, access to prenatal care, exposure to indoor air pollution, maternal employment status, and availability of childcare services, were not included in the dataset.

Statistical analysis

This paper utilized a multistate Markov chain model with a set of discrete nutritional states $S = \{s_1, s_2, s_3, s_4, s_5, s_6, s_7, s_8\}$ observed at different discrete time points (t = 2002, 2006, 2009, 2013, and 2016) [36].

The Markov chain model is mathematically explained by considering a child's nutritional state at time t (X_t) and transition probability $(p_{ij}=\mathbf{P}(X_{t+1}=s_j | X_t=s_i))$, where $\mathbf{P}=[p_{ij}], 1 \le i, j \le 8$.

In the transition probability matrix **P**, which is an 8×8 square matrix, each element p_{ij} represents nonnegativity, and the row sums to one while transitioning from nutritional state s_i to s_i .

$$p_{ij} = \frac{Number of transitions from state i to state j}{Total number of transitions from state i}$$
(1)

$$P = \begin{bmatrix} p_{11} & p_{12} & \dots & p_{18} \\ p_{21} & p_{22} & \dots & p_{28} \\ p_{31} & p_{32} & \dots & p_{38} \\ \vdots & \vdots & \ddots & \vdots \\ p_{81} & p_{82} & \dots & p_{88} \end{bmatrix}$$
(2)

The undernutrition status (N, U, S, W, US, UW, SW, and USW) of the children at each time point was recorded as an outcome variable in the dataset called the state (Fig. 2). A child's condition can be classified into one state if they fall into the normal, underweight, stunted, or wasted category; two states if they experience both conditions simultaneously; and three states if they are simultaneously affected by all three conditions. Therefore, subjects with multiple states at a time are treated using the multistate Markov chain model.

The undernutrition status (N, U, S, W, US, UW, SW, and USW) of the children at each time point was recorded as an outcome variable in the dataset called the state (Fig. 2). A child's condition can be classified into one state if they fall into the underweight, stunted, or wasted categories; into two states if they experience both conditions simultaneously; and into three states if they are simultaneously affected by all three conditions (underweight, stunting, and wasting). The inclusion

Group	Description of the group	Definition	Underweight	Stunting	Wasting	2002	2006	2009	2013	2016
z	Normal	AII WAZ, HAZ, & WHZ >-2SD	No	No	No	46.2%	58.1%	54.1%	39.6%	49.9%
n	Underweight only	WAZ <-25D	Yes	No	No	2.0%	3.5%	6.5%	4.5%	2.5%
S	Stunting only	HAZ <-25D	No	Yes	No	16.5%	15.1%	4.1%	1.6%	1.0%
M	Wasting only	WHZ <-2SD	No	No	Yes	3.2%	3.0%	6.8%	6.6%	7.3%
US	Underweight and stunting	WAZ & HAZ <-2SD	Yes	Yes	No	18.7%	14.7%	13.6%	13.1%	10.2%
NN	Underweight and wasting	WHZ & WAZ <-2SD	Yes	No	Yes	6.1%	4.1%	11.0%	20.5%	14.6%
SW	Stunting and wasting	HAZ & WHZ < -2SD	No	Yes	Yes	0.0%	0.0%	0.1%	0.2%	0.1%
NSW	Underweight, stunting, and wasting	All WHZ, WAZ, & HAZ < -2SD	Yes	Yes	Yes	7.4%	1.5%	3.9%	13.9%	14.3%
Underweight	U + US + UW + USW					34.2%	23.8%	35.0%	52.0%	41.6%
Stunting	S+US+SW+USW					42.6%	31.3%	21.7%	28.8%	25.6%
Wasting	W + UW + SW + USW					16.7%	8.6%	21.8%	41.2%	36.3%

<u>a</u> .
8
Ĕ
击
.⊆
Ð
E
Ţ
Ş
Ó
5
itio
utr
ЪС
der
ğ
Ľ
ild
5
f
ification and percentage of child undernutrition over time i
ğ
Jta
rcer
2
ă
p
ar
uo
Ę
G
sifi
ass
Ū
-
-
able
~

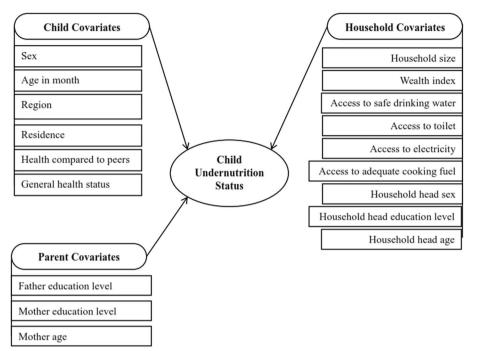


Fig. 1 Classifications of covariates in the study associated with child undernutrition

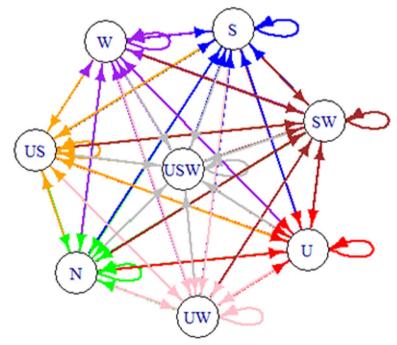


Fig. 2 State transition diagram

of the Normal (N) state reflects both the overall health status of children in the dataset and serves as a reference category for comparing undernutrition states. This approach enables a comprehensive assessment of nutritional outcomes and health dynamics within the studied population. Therefore, subjects with multiple states at a time are treated using the multistate Markov chain model.

Median First Passage Time (MFPT) and Median First Recurrence Time (MFRT)

The median is particularly useful in scenarios where the distribution of times may be skewed or when outliers could disproportionately influence the mean. The Median First Passage Time (MFPT), denoted as \tilde{m}_{ij} , represents the median time it takes for a child to transition from nutritional state s_i to state s_j for the first time. This can be computed by identifying the point at which 50% of the transition times from s_i to s_j occur, effectively capturing the typical experience of a child navigating nutritional states.

$$P(T_{ij} \le \widetilde{m}_{ij}) = 0.5 \tag{3}$$

where, $P(T_{ij} \le \tilde{m}_{ij})$ represents the probability that the time to transition from state s_i to s_j is less than or equal to the median value \tilde{m}_{ij} .

Similarly, the Median First Recurrence Time (MFRT), denoted as \tilde{r}_i quantifies the median time required for a child to return to a specific nutritional state s_i after its initial visit. This measure accounts for the recurrence dynamics, reflecting the state's stability and the likelihood of relapse into previous nutritional states. The calculation involves determining the median of the times taken to return to state s_i across all relevant transitions.

$$P(T_{ii} \le \tilde{r}_i) = 0.5 \tag{4}$$

where, $P(T_{ii} \le \tilde{r}_i)$ is the probability that the time to return to state s_i is less than or equal to the median recurrence time \tilde{r}_i .

Lifetime Undernourished Period (LUP)

The LUP measures the duration spent in undernourished states (U, S, W, US, UW, SW, and USW) over time, considering all possible sequences leading to undernourished periods. It is calculated by summing the probabilities of

- $P(X_t = m | X_{t-1} \neq m)$ is the probability of transitioning to an undernourished state at time t, given that the nutritional state at time t-1 is not the undernourished state *m*.
- δ(*t*, *m*) is an indicator function that equals 1 if the nutritional state at time t is *m* and 0 otherwise.

Markov regression

The multistate Markov chain model is a statistical technique that analyzes transitions between undernutrition states over time, providing a comprehensive understanding of child undernutrition. It captures the dynamics and an interrelationship among undernutrition states; an estimate transition probability between different combinations of undernutrition states and identifies patterns and determinants of transitions. One key advantage of the model is its ability to estimate transition probabilities between different combinations of under nutritional states, identifying patterns and determinants of transitions [36–39].

Markov regression using transition probabilities is a method that considers the current-state response as an additional covariate alongside usual risk factors and expresses conditional probability as a function of both factors (1-normal, 2-underweight only, 3-stunting only, 4-wasted only, 5-underweight and stunted, 6-underweight and wasted, 7- stunted and wasted, and 8-underweight, stunted and wasted).

$$Pr(y_{i(t+1)} = 1 | y_{it} = 1) = x_{l(t+1)}/\beta_{1}$$

$$Pr(y_{i(t+1)} = 1 | y_{it} = 2) = x_{l(t+1)}/\beta_{2}$$

$$Pr(y_{i(t+1)} = 1 | y_{it} = 3) = x_{l(t+1)}/\beta_{3}$$

$$Pr(y_{i(t+1)} = 1 | y_{it} = 4) = x_{l(t+1)}/\beta_{4}$$

$$Pr(y_{i(t+1)} = 1 | y_{it} = 5) = x_{l(t+1)}/\beta_{5}$$

$$Pr(y_{i(t+1)} = 1 | y_{it} = 6) = x_{l(t+1)}/\beta_{6}$$

$$Pr(y_{i(t+1)} = 1 | y_{it} = 7) = x_{l(t+1)}/\beta_{7}$$

$$Pr(y_{i(t+1)} = 1 | y_{it} = 8) = x_{l(t+1)}/\beta_{8}$$
(6)

The model can be written as follows:

logit
$$Pr(Y_{i(t+1)} = 1 | Y_{it} = y_{it}) = x_{l(t+1)} \beta_0 + y_{it} x_{l(t+1)} \omega$$
 so that $\beta_1 = \beta_0 + \omega$

transitioning to an undernourished state at each time point [30].

$$LUP = \sum_{i \neq j} \left[P(X_t = m | X_{t-1} \neq m) * \delta(t, m) \right]$$
 (5)

where,

- *X_t* is the nutritional state at time t.
- *m* represents the undernourished state.

This logistic model considers the current state of child undernutrition y_{it} and the interaction between y_{it} and risk factors, allowing us to determine whether a specific risk factor affects response probability.

With a first-order Markov model, the contribution to the likelihood for the ith subject can be written as

$$L_{i}(y_{i1},...,y_{int}) = f(y_{i1}) \prod_{j=2}^{n_{i}} f(y_{it} | \mathcal{H}_{it})$$
(8)

The conditional distribution of Y_{it} in a Markov model of order q is

$$f(\mathcal{Y}_{it} \mid \mathcal{H}_{it}) = f\left(\mathcal{Y}_{it} \mid \mathcal{Y}_{it-1}, \dots, \mathcal{Y}_{it-q}\right)$$
(9)

Therefore, the likelihood contribution for the ith subject becomes

$$f(y_{i1},\ldots,y_{iq})\prod_{j=q+1}^{n_1}f((y_{it}\mid y_{it-1},\ldots,y_{it-q}))$$
(10)

The logistic case does not determine $f(y_{i1}, \ldots, y_{iq})$ from the GLM assumption, and the full likelihood is unavailable. An alternative is to estimate β and α by maximizing the conditional likelihood.

$$\prod_{i=1}^{m} f(y_{iq+1}, \dots, y_{in_i} \mid y_{i1}, \dots, y_{iq}) = \prod_{i=1}^{m} \prod_{t=q+1}^{n_i} f(y_{it} \mid t)$$
(11)

In addition, multistate Markov chain models are governed by transition intensity functions, which represent the instantaneous incidence rate of moving from one state to another state at time t [37].

$$\lambda_{jk} = \lim_{\Delta t \to 0} \frac{P\{Y(t + \Delta t) = k \mid Y(t) = j\}}{\Delta t}, j \neq k \quad (12)$$

where Y(t) is the state occupied at time t. This transition intensity is the (j, k) entry of the transition intensity matrix denoted by Λ , the rows of which sum to zero. The model is time homogeneous, which means that $\lambda_{jk}(t) = \lambda_{jk}$ or all t. The diagonal entries of Λ are defined by convention as the sum of off diagonal entries with negative magnitude [37]. When the time is homogeneous, $w_q = t_q - t_{q-1}$, q = 1, 2, 3, ..., myields the log likelihood

$$L(\theta) = \sum_{q=1}^{m} \sum_{j,k=1}^{r} n_{jkq} \log p_{jk} w_q$$
(15)

The data analysis utilized two prominent R packages, msm [40] and markovchain [41].

Results

The study showed that US was prevalent at a baseline level of 18.7%, with male children having a greater incidence of undernutrition (19.9%) than female children (17.3%). In rural areas, the prevalence was 11.6%, while in urban areas, it was greater (22.5%). The prevalence of US, UW, and USW was found to be greater for children with illiterate fathers (22.5%, 8.0%, and 9.6%, respectively), and mothers (22.3%, 7.0%, and 9.8%, respectively) than for those with literate families (Table 2).

Similarly, in Table 2: children with unsafe drinking water in their household experienced higher levels of US (23.2%) than did those with safe water at baseline (14.7%). Likewise, children without toilet access in households at baseline were more affected by US (23.5%), UW (7.7%) and USW (9.4%) than were those with toilet access (10.7%, 3.3% and 4.1%, respectively). In addition, children in households without electricity access were more affected by the US (22.8%), the UW (7.8%) and the USW (9.4%) than were those with electricity access (11%, 3%, and 3.7%, respectively). Finally, children in households with cooking fuel access at baseline had higher rates of

	$\Gamma - (\lambda_{12} + \lambda_{13} + \dots + \lambda_{18})$	λ_{12}		λ_{18} 7	
	1	$-(\lambda_{21}+\lambda_{23}+\cdots+\lambda_{28})$		n	(13)
$\Lambda =$	λ ₃₁	λ_{32}	•••	λ_{38}	(10)
	:	:	۰.	:	
	· ·	·	•	•	
	λ_{81}	λ_{82}		$-(\lambda_{81}+\lambda_{82}+\cdots+\lambda_{87}) \rfloor$	

The model includes other risk factors; hence, the model is $\log (\lambda_{jk}) = \beta_{0jk} + \sum_{l=1}^{L} \beta_{ljk}$ where l is the number of risk factors (l=1, 2, ..., L).

For a given random sample of n individuals, the likelihood function for θ is calculated using the number of individuals in state j at t-1 and k at t by the distribution of individuals among states at t₀;

$$L(\theta) = \prod_{q=1}^{m} \left\{ \prod_{j,k=1}^{r} p_{jk} (t_{q-1,t_q}) n_{jkq} \right\}$$
(14)

experiencing US (19.6%), UW (6.6%), and USW (8%) than did those without cooking fuel (8.1%, 0.6%, and 1.2%, respectively). The zero values in the SW column indicate that there were no children concurrently affected by stunting and wasting at baseline.

The study revealed changes in the prevalence of various nutritional states over the years. The study revealed that the prevalence of normal nutritional status increased from 922 in 2002 to 1110 in 2006 and 1019 in 2009 but decreased to 741 in 2013 and 906 in 2016. The lowest incidence of underweight was 39 at baseline, followed by 46 in 2016, 67 in 2006, 84 in 2013, and 123 in 2009. The highest incidence of UW was 383 in 2013, and the

Table 2 Prevalence of child nutritional status at baseline (2002) across sc	ociodemographic characteristics
---	---------------------------------

Variables	Category	N	U	S	w	US	UW	SW	USW	Total
	5 1	n(%)	n(%)	n(%)	n(%)	n(%)	n(%)	n(%)	n(%)	n
Gender	Male	433 (41.4)	17 (1.6)	195 (18.6)	38 (3.6)	208 (19.9)	65 (6.2)	0 (0)	91 (8.7)	1047
	Female	489 (51.6)	22 (2.3)	134 (14.1)	26 (2.7)	164 (17.3)	56 (5.9)	0 (0)	56 (5.9)	947
Residence	Urban	413 (59.2)	12 (1.7)	128 (18.3)	23 (3.3)	81 (11.6)	17 (2.4)	0 (0)	24 (3.4)	698
	Rural	509 (39.3)	27 (2.1)	201 (15.5)	41 (3.2)	291 (22.5)	104 (8)	0 (0)	123 (9.5)	1296
Father's Edu	Illiterate	407 (38.9)	21 (2)	166 (15.9)	33 (3.2)	235 (22.5)	84 (8)	0 (0)	100 (9.6)	1045
	Literate	516 (54.4)	18 (1.9)	163 (17.2)	31 (3.3)	137 (14.4)	37 (3.9)	0 (0)	47 (5)	949
Mother's Edu	Illiterate	474 (39.8)	22 (1.8)	193 (16.2)	36 (3)	266 (22.3)	84 (7)	0 (0)	117 (9.8)	1192
	Literate	448 (55.9)	17 (2.1)	136 (17)	28 (3.5)	106 (13.2)	37 (4.6)	0 (0)	30 (3.7)	802
Access to Safe Drink Water	No	364 (39.4)	23 (2.5)	139 (15)	28 (3)	215 (23.2)	77 (8.3)	0 (0)	79 (8.5)	925
	Yes	558 (52.2)	16 (1.5)	190 (17.8)	36 (3.4)	157 (14.7)	44 (4.1)	0 (0)	68 (6.4)	1069
Access to Toilet	No	466 (37.6)	28 (2.3)	207 (16.7)	35 (2.8)	291 (23.5)	96 (7.7)	0 (0)	116 (9.4)	1239
	Yes	456 (60.4)	11 (1.5)	122 (16.2)	29 (3.8)	81 (10.7)	25 (3.3)	0 (0)	31 (4.1)	755
Access to Electricity	No	498 (38.7)	27 (2.1)	205 (15.9)	42 (3.3)	294 (22.8)	100 (7.8)	0 (0)	121 (9.4)	1287
	Yes	424 (60)	12 (1.7)	124 (17.5)	22 (3.1)	78 (11)	21 (3)	0 (0)	26 (3.7)	707
Access to Cooking Fuel	No	815 (44.7)	34 (1.9)	290 (15.9)	60 (3.3)	358 (19.6)	120 (6.6)	0 (0)	145 (8)	1822
	Yes	107 (62.2)	5 (2.9)	39 (22.7)	4 (2.3)	14 (8.1)	1 (0.6)	0 (0)	2 (1.2)	172
Overall		922 (46.2)	39(1.96)	329(16.5)	64(3.2.)	372(18.7)	121(6.1)	0(0)	147(7.4)	1994

Table 3 Prevalence of child nutritional status at each time-point

Variables	2002 Ye	ar	2006 Yea	r	2009 Yea	r	2013 Ye	ar	2016 Year	
	n	%	n	%	n	%	n	%	n	%
Undernutritic	on State									
Ν	922	46.2	1110	58.1	1019	54.1	741	39.6	906	49.9
U	39	2	67	3.5	123	6.5	84	4.5	46	2.5
S	329	16.5	289	15.1	77	4.1	30	1.6	18	1
W	64	3.2	57	3	129	6.8	123	6.6	133	7.3
US	372	18.7	282	14.7	256	13.6	246	13.1	185	10.2
UW	121	6.1	78	4.1	207	11	383	20.5	265	14.6
SW	0	0	0	0	1	0.1	4	0.2	2	0.1
USW	147	7.4	29	1.5	73	3.9	261	13.9	260	14.3

lowest was 78 in 2006. The prevalence of both stunted and wasted children was extremely low. Similarly, the prevalence of both stunted and wasted children was extremely low. Finally, the prevalence of children who were USW increased from 147 thousand in 2002 to 260 thousand in 2016 (Table 3).

Table 4 revealed that almost half 2269(48.3%) of male children were normal, 177(49.3%) were underweight, 443 (59.6\%) were stunted, 255(50.4\%) were wasted, 752(56.1\%) were US, 597 (56.6\%) were UW, 3(42.9\%) were SW, and 502(65.2\%) were USW while 182(50.7\%) of females were underweight, 300(40.4\%) stunted, 251(49.6\%) wasted, 589(43.9\%) US, 457(43.4\%) UW,

4(57.1%) SW, and 268(34.8%) were USW. The majority of children in the normal, underweight and stunted states were from the SNNP region (25.4%, 28.7%, and 23.3%, respectively), while the majority wasted children were from the Amhara region (115; 22.7%). Similarly, the majority of children affected by US were diagnosed with SNNP 361 (26.9%). However, the majority of the children affected by UW, SW, or USW were from the Amhara region (286; 27.1%), (4; 57.1%), and (237; 30.8%), respectively.

The study showed that the probability of children being in a normal state at time t + 1 was 0.68. The probability of transitioning from underweight to normal was 0.31, while the probability of transitioning from

Variables	Category	Ν	U	S	W	US	UW	SW	USW
		n(%)	n(%)	n(%)	n(%)	n(%)	n(%)	n(%)	n(%)
Gender	Male	2269 (48.3)	177 (49.3)	443 (59.6)	255 (50.4)	752 (56.1)	597 (56.6)	3 (42.9)	502 (65.2)
	Female	2429 (51.7)	182 (50.7)	300 (40.4)	251 (49.6)	589 (43.9)	457 (43.4)	4 (57.1)	268 (34.8)
Region	Tigray	884 (18.8)	77 (21.4)	135 (18.2)	108 (21.3)	275 (20.5)	265 (25.1)	2 (28.6)	174 (22.6)
	Amhara	713 (15.2)	82 (22.8)	155 (20.9)	115 (22.7)	292 (21.8)	286 (27.1)	4 (57.1)	237 (30.8)
	Oromiya	1021 (21.7)	62 (17.3)	168 (22.6)	77 (15.2)	315 (23.5)	157 (14.9)	0 (0.0)	139 (18.1)
	SNNP	1191 (25.4)	103 (28.7)	173 (23.3)	101 (20.0)	361 (26.9)	245 (23.2)	1 (14.3)	197 (25.6)
	Addis Ababa CA	889 (18.9)	35 (9.7)	112 (15.1)	105 (20.8)	98 (7.3)	101 (9.6)	0 (0.0)	23 (3.0)
Residence	Urban	2041 (43.4)	93 (25.9)	274 (36.9)	234 (46.2)	308 (23.0)	282 (26.8)	2 (28.6)	144 (18.7)
	Rural	2657 (56.6)	266 (74.1)	469 (63.1)	272 (53.8)	1033 (77)	772 (73.2)	5 (71.4)	626 (81.3)
Mother's level of education	Illiterate	1884 (40.1)	176 (49.0)	407 (54.8)	202 (39.9)	782 (58.3)	538 (51.0)	3 (42.9)	435 (56.5)
	Literate	2814 (59.9)	183 (51.0)	336 (45.2)	304 (60.1)	559 (41.7)	516 (49.0)	4 (57.1)	335 (43.5)
Father's level of education	Illiterate	1328 (28.3)	128 (35.7)	316 (42.5)	149 (29.4)	579 (43.2)	349 (33.1)	1 (14.3)	256 (33.2)
	Literate	3370 (71.7)	231 (64.3)	427 (57.5)	357 (70.6)	762 (56.8)	705 (66.9)	6 (85.7)	514 (66.8
Access to safe drinking water	Yes	1627 (34.6)	165 (46.0)	285 (38.4)	173 (34.2)	659 (49.1)	506 (48.0)	4 (57.1)	380 (49.4)
	No	3071 (65.4)	194 (54.0)	458 (61.6)	333 (65.8)	682 (50.9)	548 (52.0)	3 (42.9)	390 (50.6)
Access to sanitation	Yes	1490 (31.7)	133 (37.0)	388 (52.2)	142 (28.1)	663 (49.4)	358 (34.0)	3 (42.9)	305 (39.6)
	No	3208 (68.3)	226 (63.0)	355 (47.8)	364 (71.9)	678 (50.6)	696 (66.0)	4 (57.1)	465 (60.4)
Access to electricity	Yes	1949 (41.5)	221 (61.6)	429 (57.7)	199 (39.3)	888 (66.2)	605 (57.4)	4 (57.1)	491 (63.8)
	No	2749 (58.5)	138 (38.4)	314 (42.3)	307 (60.7)	453 (33.8)	449 (42.6)	3 (42.9)	279 (36.2)
Access to cooking fuels	Yes	4082 (86.9)	336 (93.6)	672 (90.4)	439 (86.8)	1271 (94.8)	998 (94.7)	7 (100.0)	741 (96.2)
-	No	616 (13.1)	23 (6.4)	71 (9.6)	67 (13.2)	70 (5.2)	56 (5.3)	0 (0.0)	29 (3.8)
Total		4698	359	743	506	1341	1054	7	770

underweight to a UW state was 0.25. Similarly, the probability of recovering from stunting to normal was 0.47, and a higher transition of 0.23 was to the SW state. Finally, children in a wasting state at time t had a probability of 0.46 of recovering to the normal state and had a higher probability of transitioning to UW (0.20) (Fig. 3).

Children who were facing US simultaneously had a 0.24 chance of recovering to normal, with a higher probability of 0.35 remaining in the same state at time t+1. Similarly, children in the UW state at time t had a 0.29 chance of recovering and 0.30 chance of remaining in the same state at time t+1. Likewise, children who were affected by SW had a probability of 0.29 moving to a normal state, with a greater likelihood of 0.29 moving to both the UW and the US. Finally, children in all USW states concurrently had a 0.24 chance of recovering, with a decreased transition to SW. Generally, the transitions from any state to concurrent stunting and wasting states were relatively fewer (near zero) (Fig. 3).

Finally, the probability of children being in a state varied, with normal children having the highest probability (0.68), followed by US (0.35), UW (0.30), USW (0.28), wasting (0.19), stunting (0.15), and underweight children (0.08) (Fig. 3). Figure 3 includes the normal state, providing a comprehensive view of nutritional states among the children in the study. Understanding its transition probabilities is crucial for assessing nutritional outcomes and serves as a baseline for comparing transitions into and out of states like underweight, stunting, wasting, and their combinations (US, UW, SW, and USW).

There was a greater incidence of US in male children (16%) than in female children (14.6%). In rural areas, the cumulative incidence of UWs was 10.3%, while in urban areas, it was lower (6.5%). Children in households without safe drinking water experienced greater incidences of US, UW, SW and USW (15.7%, 9.2%, 0.1% and 6.7%, respectively) than did those with safe drinking water. Similarly, children without access to toilets in households face greater incidents of stunting, US and USW (13.7%, 20.7%, and 7.4%, respectively) than do those who have access to toilets. Similarly, children in households without access to electricity experienced higher incidences of underweight, stunting, US, UW, and USW (4.2%, 11.7%, 20%, 10% and 8.2%, respectively) than did the other children. Finally, children in families without cooking fuel access were more likely to be underweight or stunted and to have experienced incidents of US and UW (3.7%,

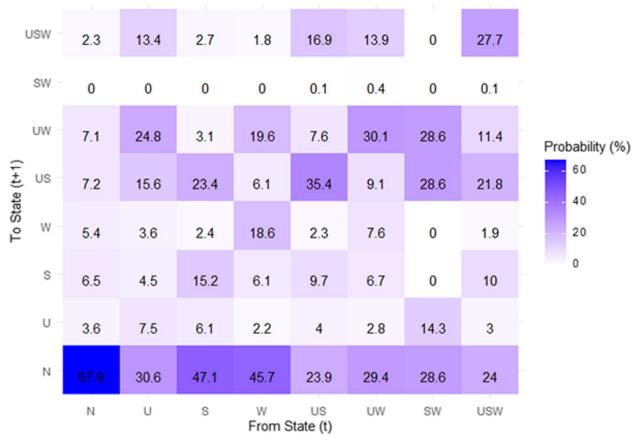


Fig. 3 State transition probability (%) matrix. NB: the diagonal elements representing the probability of remaining in a state and off-diagonal elements showing transition probabilities between states

Table 5 Cu	umulative incidence	of child undernutritior	n at the end of the surve	ey in 2016
------------	---------------------	-------------------------	---------------------------	------------

Variables	Category	N	U	S	W	US	UW	SW	USW	Total
		n(%)	n(%)	n(%)	n(%)	n(%)	n(%)	n(%)	n(%)	
Gender	Male	7140 (47.0)	516 (3.4)	1853 (12.2)	684 (4.5)	2431 (16.0)	1443 (9.5)	0 (0)	1124 (7.4)	15191
	Female	6813 (53.9)	493 (3.9)	1176 (9.3)	581 (4.6)	1846 (14.6)	1049 (8.3)	0 (0)	670 (5.3)	13641
Residence	Urban	6219 (61.0)	275 (2.7)	1132 (11.1)	581 (5.7)	989 (9.7)	663 (6.5)	0 (0)	347 (3.4)	10195
	Rural	8275 (44.4)	764 (4.1)	1976 (10.6)	727 (3.9)	3448 (18.5)	1920 (10.3)	0 (0)	1510 (8.1)	18637
Mother's Edu	Illiterate	6293 (43.3)	552 (3.8)	1715 (11.8)	567 (3.9)	2791 (19.2)	1410 (9.7)	0 (0)	1177 (8.1)	14534
	Literate	8207 (57.4)	500 (3.5)	1358 (9.5)	743 (5.2)	1644 (11.5)	1158 (8.1)	0 (0)	672 (4.7)	14298
Father's Edu	Illiterate	4785 (43.0)	423 (3.8)	1380 (12.4)	456 (4.1)	2215 (19.9)	1046 (9.4)	0 (0)	824 (7.4)	11129
	Literate	9719 (54.9)	620 (3.5)	1735 (9.8)	850 (4.8)	2213 (12.5)	1522 (8.6)	18 (0.1)	1027 (5.8)	17703
Access to Safe Drink Water	No	4472 (36.8)	413 (3.4)	1021 (8.4)	413 (3.4)	1908 (15.7)	1118 (9.2)	12 (0.1)	814 (6.7)	12153
	Yes	10,675 (64.0)	634 (3.8)	2202 (13.2)	951 (5.7)	2502 (15.0)	1451 (8.7)	0 (0)	1034 (6.2)	16679
Access to Toilet	No	5308 (42.5)	450 (3.6)	1711 (13.7)	425 (3.4)	2585 (20.7)	1087 (8.7)	0 (0)	924 (7.4)	12490
	Yes	9184 (56.2)	588 (3.6)	1405 (8.6)	882 (5.4)	1847 (11.3)	1487 (9.1)	0 (0)	931 (5.7)	16342
Access to Electricity	No	6720 (42.1)	670 (4.2)	1868 (11.7)	591 (3.7)	3193 (20.0)	1596 (10.0)	0 (0)	1309 (8.2)	15963
	Yes	7773 (60.4)	373 (2.9)	1248 (9.7)	721 (5.6)	1235 (9.6)	965 (7.5)	0 (0)	540 (4.2)	12869
Access to Cooking Fuel	No	12,812 (48.7)	973 (3.7)	2815 (10.7)	1184 (4.5)	4236 (16.1)	2473 (9.4)	0 (0)	1815 (6.9)	26308
	Yes	1689 (60.3)	64 (2.3)	289 (10.3)	140 (5.0)	193 (6.9)	101 (3.6)	0 (0)	325 (11.6)	2801
Overall		14,495(34.9)	14,495(34.9)	1046(2.5)	3110(7.5)	1314(3.2)	4433(10.7)	2569(6.2)	13(0.0)	41475

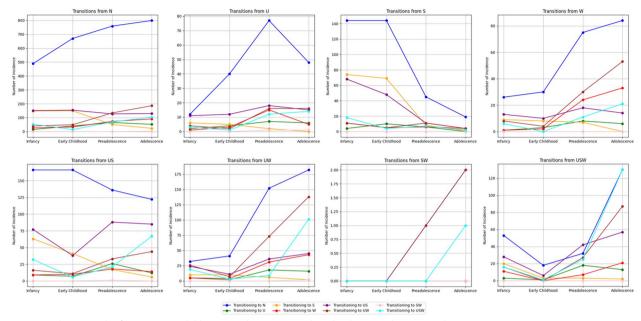


Fig. 4 Incidence trends of transitions in child nutritional status across their developmental stages in Ethiopia

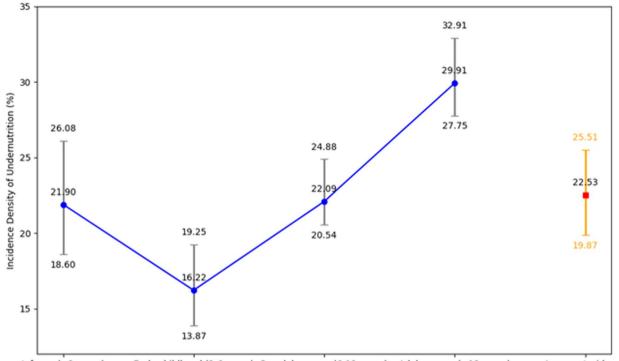
10.7%, 16.1%, and 9.4%, respectively) than were those with access to cooking fuel (Table 5).

The incidence trends of transitions between the eight distinct nutritional statuses of children across their developmental stages are displayed in Fig. 4. This figure tracks how children move from one nutritional status to another at each stage of development, with incidence rates reported per 100 children. The study finds varying rates of undernutrition prevalence, with the highest incidence observed in adolescence (29.91 per 100 children) and the lowest in early childhood (16.22 per 100 children). For instance, in infancy (<2 years), the incidence rate is 21.90 per 100 children, while in early childhood (2-6 years) it decreases to 16.22 per 100 children. Conversely, preadolescence (6-12 years) shows an increase to 22.09 per 100 children, and adolescence (>12 years) exhibits the highest incidence at 29.91 per 100 children (Fig. 5). These figures provide a comprehensive overview of undernutrition incidence across different developmental stages, highlighting notable variations in vulnerability among children (Figs. 4 and 5).

Accordingly, the study revealed that a child initially in a USW takes approximately a median of 13.90 months to make the first transition to the normal state, 24 months or two years to underweight, and 21.50 months to stunt. The transition period for a child initially in US, UW, and USW to experience SW for the first time is relatively lengthy. According to Table 6, initially, individuals spent a median of 24 months or two years transitioning from a healthy state to underweight, followed by 22 months to stunted, 17.9 months to USW, 17.5 months to wasted, 13.75 months to US, and 9.5 months to UW. The zero values (not exactly zero but near to zero) in the SW column indicate that the median time to transition from another nutritional state to SW is extremely short or negligible. Lastly, underweight children had a median recurrence period of 23.4 months, followed by 19.8 months for stunted, 14.5 months for wasted, 8 months for US, 6.5 months for UW, and 9.8 months for USW states (Table 6).

Furthermore, the study showed that underweight children had a shorter lifetime undernourished period (LUP) of approximately 15.3 months, while those with stunting had a LUP of nearly 23 months. Those with wasting had a median LUP of 30.4 months. The combined category of underweight and stunted children (US) had a higher LUP of approximately 48.8 months, while those with USW had a median LUP of approximately 45.4 months. However, a child experiencing SW had a LUP of nearly five months, implying a shorter duration of undernutrition in this state (Table 7).

Additionally, in Table 8: males had a median LUP of approximately 89.7 months, while females had a lower median LUP of approximately 75.4 months. Regionally, children living in the Amhara region had the highest median LUP, nearly 94.5 months; followed by those living in the Tigray region, 87.2 months; and those living in the SNNP region, 82.3 months. In the Oromia region, the median LUP was approximately 78.1 months. Finally, the Addis Ababa City Administration had the lowest median



 Infancy (<2 years)</th>
 Early childhood (2-6 years)
 Preadolescence (6-12 years)
 Adolescence (>12 years)
 Average incidence

 Fig. 5
 Trend of incidence density of undernutrition by child developmental stages with 95% confidence intervals in Ethiopia
 Average incidence

Table 6 Median first passage time and median first recurrence time of child undernutrition in Ethiopia

		To time (t+1)							
		Median First Pa	assage Time and	Median First Re	currence Time	(95% Cl) (in mor	iths)		
		N	U	S	W	US	UW	SW	USW
From time	N	11.5 (10.8, 12.2)	24.0 (19.5, 28.5)	22.0 (18.0, 26.0)	17.5 (13.0, 22.0)	13.8 (11.5, 15.0)	9.5 (7.5, 11.5)	0	17.9 (15.8, 20.0)
(t)	U	12.8 (12.0, 13.5)	23.4 (18.0, 29.0)	22.5 (19.0, 26.0)	18.5 (14.0, 23.0)	11.5 (9.5, 13.0)	7.5 (5.5, 9.5)	0	14.1 (11.5, 16.5)
	S	12.5 (11.8, 13.2)	23.9 (19.0, 28.5)	19.8 (15.0, 24.0)	18.8 (15.0, 22.0)	10.6 (8.0, 12.0)	10.0 (8.0, 12.0)	0	16.5 (14.0, 19.0)
	W	12.6 (11.9, 13.3)	24.5 (20.0, 29.0)	23.5 (19.0, 28.0)	14.5 (10.0, 18.0)	13.5 (11.0, 16.0)	7.5 (5.5, 9.5)	0	17.0 (14.0, 20.0)
	US	13.2 (12.0, 14.0)	24.4 (19.0, 29.5)	20.5 (16.0, 25.0)	19.5 (16.0, 23.0)	8.0 (6.0, 10.0)	9.4 (7.0, 11.0)	41 (39, 43)	12.8 (10.5, 15.0)
	UW	13.0 (12.0, 14.0)	24.8 (20.0, 29.5)	22.6 (18.0, 27.0)	17.0 (13.0, 21.0)	12.6 (10.5, 14.0)	6.5 (4.5, 8.5)	43 (40, 46)	13.5 (11.0, 16.0)
	SW	13.6 (12.5, 14.5)	20.5 (16.0, 25.0)	23.5 (20.0, 27.0)	19.0 (15.0, 23.0)	10.7 (8.0, 12.5)	6.0 (4.0, 8.0)	0	15.5 (13.0, 18.0)
	USW	13.9 (13.0, 14.8)	24.0 (19.0, 29.0)	21.5 (17.0, 25.0)	19.2 (15.0, 23.0)	9.5 (7.0, 11.0)	8.8 (6.5, 10.5)	34 (31, 37)	9.8 (7.0, 12.0)

NB: The diagonal entries in the above table represent MFRT, indicating the median time taken for children to return to a specific state after initially transitioning away from it. The off-diagonal entries denote MFPT, showing the median time for children to transition from one nutritional state to another for the first time during the study period

 Table 7
 Median and IQR of the Lifetime Undernourished Period

 (LUP) of a children in Ethiopia

-		
States	Median LUP (months)	Inter Quartile Range (IQR) of LUP
N	83.6	10.5
U	15.3	5.0
S	23.0	7.5
W	30.4	6.0
US	48.8	8.0
UW	45.4	4.5
SW	5.2	2.0
USW	48.1	9.0

Table 8 Median and IQR of the Lifetime Undernourished Period
(LUP) in Ethiopia

Variables	Category	Median LUP (months)	IQR of LUP (months)	
Gender	Male	89.7	12.5	
	Female	75.4	10.4	
Region	Tigray	87.2	15.8	
	Amhara	94.5	8.7	
	Oromiya	78.1	6.5	
	SNNP	82.3	9.3	
	Addis Ababa CA	55.7	5.5	
Residence	Urban	60.4	7.6	
	Rural	89.3	11.3	
Father's level of education	Illiterate	52.3	9.5	
	Literate	56.9	7.1	
Mother's level of education	Illiterate	68.5	6.4	
	Literate	57.3	5.5	
Access to safe drinking	Yes	48.3	6.3	
water	No	52.1	5.7	
Access to toilet	Yes	45.2	4.8	
	No	43.8	3.5	
Access to electricity	Yes	40.7	8.4	
	No	67.5	10.2	
Access to cooking fuels	Yes	25.1	5.5	
	No	70.8	12.1	

LUP among the regions, occurring at 55.7 months. Similarly, rural children had the highest median LUP of approximately 89.3 months. Children from illiterate mothers had a greater LUP 68.5 months than did those from literate mothers. Children from households without access to safe water, electricity, or cooking fuel had higher LUPs (52.1 months, 67.5 months, and 70.8 months, respectively (see Table 8).

The Markov regression analysis revealed that a previous state of undernutrition was the main risk factor for current undernutrition. Other risk factors include rural residence, child's health compared to peers, education of mother and father, access to safe drink water, and cooking fuel. Children from rural areas were 70% (95% CI: 1.40, 2.06) more likely to be undernourished than were those from urban residences. Compared to their peers, children with worse health conditions were 80% (95% CI: 1.40, 2.30) more likely to be undernourished. Children with illiterate fathers (OR = 1.85: 1.45, 2.42) and mothers (OR = 1.40; 95% CI: 1.24, 1.56) had greater odds of wasting than did those with literate mothers. Children who did not drink safe water had greater odds (OR = 1.60; 95% CI=1.21, 2.04) of experiencing undernutrition. Children with access to cooking fuel were 50% more likely to be undernourished, irrespective of previous undernutrition, than were those without access to cooking fuel (OR = 1.50; 95% CI: 1.40, 1.61). These findings highlight the importance of addressing undernutrition in children's lives (Table 9).

The Markov regression analysis using transition intensity revealed that children in rural areas with undernutrition had a 1.75 (95% CI: 1.53, 1.67) greater risk of being undernourished and a 0.63% (95% CI: 1.57, 1.69) slower recovery from undernutrition to normal status than did those in urban areas. This finding suggested that the probability of a child recovering from undernutrition to normal was less than the probability of progression. Children with illiterate fathers (HR = 1.50; 95% CI: 1.38, 1.62) and mothers (HR = 1.45; 95% CI: 1.02, 3.29) transitioned faster from normal to undernutrition, while those with illiterate mothers 1.81 (95% CI: 1.01, 3.24) recovered to normal. Additionally, children in households without safe drinking water (HR = 1.70; 95% CI: 1.26, 2.14) were more likely to transition from normal to undernutrition than were those with access to safe water. Children with access to cooking fuel also had a faster (HR = 1.95; 95% CI: 1.75, 2.17) transition from normal to undernutrition, while those with fuel-based families had a slower recovery rate. These findings highlight the importance of accessing safe drinking water and cooking fuel for children's health (Table 10). However, due to the complexity and volume of transition intensity data across multiple states (U, S, W, US, UW, SW, and USW), these categories were combined into a single Undernourished category in Table 10 for clarity and simplicity of presentation.

Discussion

One study showed that the prevalence of stunting and underweight was greater than 20% [42]. The present study reported that the incidences of stunting and US were above 16.5% and 18.7%, respectively, which were

Variables	Category	OR	Robust SE	95% CI		P value
				LCL	UCL	
Residence	Urban ^a					
	Rural	1.70	0.110	1.395	2.059	0.010
Previous state of Undernutrition		92.00	0.280	50.500	160.00	< 0.001
Residence*Previous State		1.85	0.107	1.480	2.360	0.015
Child's health compared to peers	Same ^a	1.00				
	Better	1.55	0.450	0.620	3.940	0.950
	Worse	1.80	0.135	1.400	2.300	0.003
Previous state of Undernutrition		102.00	0.290	57.04	182.41	< 0.001
Child's health*Previous State		1.15	0.075	1.010	1.320	0.750
Father's level of education	Illiterate	1.85	0.130	1.450	2.420	0.004
	Literate ^a	1.00				
Previous state of Undernutrition		90.00	0.055	80.201	100.04	< 0.001
level of education *Previous State		1.70	0.250	1.030	2.830	0.040
Mother's level of education	Illiterate	1.40	0.055	1.240	1.560	0.010
	Literate ^a	1.00				
Previous state of Undernutrition		85.00	0.295	45.04	150.81	< 0.001
level of education *Previous State		1.25	0.030	1.170	1.340	0.030
Access to safe drinking water	Yes ^a					
	No	1.60	0.140	1.210	2.040	0.015
Previous state of Undernutrition		95.00	0.120	75.50	115.04	< 0.001
Safe drink water *Previous State						
Access to cooking fuel	Yes	1.50	0.002	1.400	1.610	0.020
-	No ^a					
Previous state of Undernutrition		105.00	0.310	58.15	190.74	< 0.001
Cooking fuel *Previous State						

Table 9 Factors influencing childhood undernutrition risk in Ethiopia

^a Denotes the reference category

*The asterisks represent interaction effects rather than statistical significance

similar to the findings of previous studies. Furthermore, the study revealed that male children have a longer median lifetime undernourished period (89.7 months) than female children (75.4 months). The present study showed that male children had a greater risk of undernutrition than female children did.

This finding was similar to the finding from the NHFS-3 data [43]. A case–control study from Bangladesh found no significant difference in all undernutrition states (underweight, stunting, or wasting) among male and female children [44]. The present study revealed that urban children have better health status and immunization rates, while rural children have a greater risk of undernutrition. This finding is supported by a Ugandan study that revealed that residence significantly impacts nutritional health [45].

The Markov chain principle states that the current state is dependent on the previous state, allowing for the calculation of the first median transition time. This time is used to estimate the median duration for underweight, stunted or wasted children to reach normalcy and to identify the factors affecting this time. This type of time has not been reported in longitudinal studies. The present study reported the probability of transition from one state at the previous time to the current state of undernutrition, providing valuable insights for appropriate intervention. This information is crucial for understanding the progression of undernourished children and their needs.

In many studies involving multistate Markov chain modeling with follow-up times, mean and standard deviation (SD) are commonly reported as summary statistics [31, 46, 47]. However, since follow-up times are typically skewed, the use of mean and SD may not accurately represent the data's distribution. In this study, we opted for the median and IQR as more appropriate summary statistics, as they provide a clearer picture of central tendency and variability while reducing the impact of outliers. This

Variables	Category	From time (t)	To time (t + 1)			
			Normal		Undernourished	
			Hazard Ratio	95% CI	Hazard Ratio	95% CI
Sex of child	Male	Normal			1.75	1.53, 1.67
		Undernutrition	1.63	1.57, 1.69		
	Female ^a	Normal			1.00	
		Undernutrition	1.00			
Residence	Urban	Normal			1.42	1.40, 1.44
		Undernutrition	0.42	0.40, 0.44		
	Rural ^a	Normal			1.00	
		Undernutrition	1.00			
Father's level of education	Illiterate	Normal			1.50	1.38, 1.62
		Undernutrition	1.34	0.59, 2.91		
	Literate ^a	Normal			1.00	
		Undernutrition	1.00			
Mother's level of education	Illiterate	Normal			1.45	1.02, 3.29
		Undernutrition	1.81	1.01, 3.24		
	Literate ^a	Normal			1.00	
		Undernutrition	1.00			
Access to safe drinking water	Yes ^a	Normal			1.00	
		Undernutrition	1.00			
	No	Normal			1.70	1.26, 2.14
		Undernutrition	0.35	0.33, 0.37		
Access to cooking fuel	Yes	Normal			1.95	1.75, 2.17
		Undernutrition	0.64	0.50, 0.84		
	No ^a	Normal			1.00	
		Undernutrition	1.00			

Table 10 Transition intensity analysis of childhood undernutrition risk factors in Ethiopia

^a Denotes the reference category

approach offers a more accurate representation of followup durations in skewed data.

The study showed that children in the USW state required an average of 14.63 months to transition to normal health, those in the stunting state took an average of 13.05 months, and those in the underweight state took an average of 13.89 months. However, a study by [37] revealed that it takes an average of 2.73 years (about 33 months) to transition from a wasting state to normal undernutrition and 1.97 years (about 24 months) to transition from stunting to normal undernutrition.

This study examined the risk factors for undernutrition on the Markovian property. The authors found that area of residence where the children lived in rural areas, having worse health conditions than their peers, illiterate parents, lack of safe drinking water, and access to cooking fuels were significant factors associated with wasting, regardless of previous undernutrition. The interaction effect of these covariates with the previous state was also statistically significant. This study highlights the importance of addressing these factors to improve nutritional outcomes.

The study revealed that children in rural areas with undernutrition face a greater risk of undernutrition and slower recovery from undernutrition than urban children. Likewise, children with illiterate fathers and mothers experienced faster transitions from normal to undernutrition. In addition, these findings also indicate that children living in households without safe drinking water but with access to cooking fuels are at a greater risk of transitioning from a normal to an undernutrition state rather than from undernutrition to a normal state.

Conclusions

In conclusion, our study highlights on critical factors influencing transition of child undernutrition in Ethiopia. Children facing USW take an average of 14.63 months to return to normal health, with male children taking a relatively longer time to become undernourished. Our findings underscore the increased risk among children in rural areas, those with illiterate parents, and those lacking access to safe drinking water and cooking fuel, who experience worse health conditions compared to their peers. Addressing these disparities is crucial for improving child health outcomes. Public health efforts should focus on targeted nutritional interventions, enhancing access to basic amenities, and promoting educational programs to empower families. These initiatives, supported by robust policy frameworks, are essential for reducing the duration of undernutrition and recurrence of child undernutrition and fostering healthier communities in Ethiopia.

Supplementary Information

The online version contains supplementary material available at https://doi. org/10.1186/s12874-024-02399-9.

Supplementary Material 1.

Acknowledgements

The authors express their gratitude to the Young Lives Study teams for providing access to the data files.

Authors' contributions

GBB: Contributed to the conceptualization, conducted the formal analysis, shaped the methodology, performed the data analysis, engaged in the writing of the report, and authored the original draft.TZ: Provided insights into conceptualization and guided methodology, reviewed and edited the manuscript, and contributed to the data analysis. HMF: Contributed to the methodology, played a key role in creating the original draft, and conducted the data analysis.

Funding

The authors have no support or funding to report.

Data availability

The dataset used in this study was obtained from the Young Lives Study. Access to the data can be obtained either by completing the form available at https://www.younglives.org.uk/use-our-data-form, selecting the dataset "Young Lives: Rounds 1-5 constructed files, 2002-2016" (used in this study), or by creating a user account through the https://ukdataservice.ac.uk/, subject to their terms and conditions. Additionally, the survey questionnaires for each round (Rounds 1-5) are available through the following link: https://www. younglives.org.uk/round-1-questionnaires. By adjusting the round number in the URL or navigating through the menu on the Young Lives website, users can access the questionnaires for each respective round.

Declarations

Ethics approval and consent to participate

Ethics approval for this study was not required since the data are secondary and available in the public domain.

Consent for publication

Not applicable.

Competing interests

The authors declare no competing interests.

Author details

¹Department of Statistics, College of Science, Bahir Dar University, Bahir Dar, Ethiopia. ²Department of Data Science, College of Natural and Computational Science, Debre Berhan University, Debre Berhan, Ethiopia. ³School of Mathematics, Statistics and Computer Science, College of Agriculture, Engineering and Science, University of KwaZulu-Natal, Durban, South Africa. ⁴Center for Environmental and Respiratory Health Research (CERH), Research Unit of Population Health, University of Oulu, Oulu, Finland. $^{\rm 5}{\rm Biocenter}$ Oulu, University of Oulu, Oulu, Finland.

Received: 29 December 2023 Accepted: 31 October 2024 Published online: 15 November 2024

References

- Khan AA, Bano N, Salam A. Child malnutrition in South Asia: a comparative perspective. South Asian Survey. 2007;14(1):129–45.
- UNICEF. Progress for children: a report card on nutrition. Unicef; 2006. https://data.unicef.org/resources/progress-for-children-areport-card-on-nutrition/.
- Meerman J, Carisma B, Thompson B. Global, regional and subregional trends in undernourishment and malnutrition. SOFA FOA. 2012;1:1–33. https://www.fao.org/fileadmin/user_upload/agn/pdf/SOFA_2013global. regional.subregional.trends.pdf.
- 4. Watkins K. The state of the world's children 2016: a fair chance for every child. 2016. ERIC.
- Rudan I, et al. Epidemiology and etiology of childhood pneumonia. Bull World Health Organ. 2008;86:408-416B.
- Agarwal M, Banerjee A. Prospects for achieving the sustainable development goals. In: ECONOMICS OF G20: A World Scientific Reference Vol 2: How Developing Countries can Achieve Sustainable Development Goals. Singapore: World Scientific; 2020. p. 65–99.
- Black RE, et al. Maternal and child undernutrition and overweight in lowincome and middle-income countries. Lancet. 2013;382(9890):427–51.
- 8. Organization WH. UNICEF-WHO low birthweight estimates: levels and trends 2000–2015. Geneva: World Health Organization; 2019.
- Victora CG, et al. Maternal and child undernutrition: consequences for adult health and human capital. Lancet. 2008;371(9609):340–57.
- 10. Fenta SM, Fenta HM, Ayenew GM. The best statistical model to estimate predictors of under-five mortality in Ethiopia. J Big Data. 2020;7(1):1–14.
- Woldeyohannes M, et al. Ethiopia national food and nutrition survey to inform the Ethiopian national food and nutrition strategy: a study protocol. BMJ Open. 2023;13(4):e067641.
- 12. Lassi ZS, et al. Systematic review of complementary feeding strategies amongst children less than two years of age. Karachi City, Pakistan: Division of Women and Child Health, The Aga Khan University; 2013.
- WHO, WHO, Global strategy on diet, physical activity and health. In: fifty-seventh world health assembly. 2004. https://www.who.int/publi cations/i/item/9241592222.
- Berhe K, et al. Risk factors of stunting (chronic undernutrition) of children aged 6 to 24 months in Mekelle City, Tigray Region, North Ethiopia: an unmatched case-control study. PLoS ONE. 2019;14(6):e0217736.
- Mengistu K, Alemu K, Destaw B. Prevalence of malnutrition and associated factors among children aged 6–59 months at Hidabu Abote District, North Shewa, Oromia Regional State. J nutr disorders ther. 2013;1(1):2161–509.
- Wondimagegn ZT. Magnitude and determinants of stunting among children in Africa: a systematic review. Curr Res Nutr Food Sci J. 2014;2(2):88–93.
- Tesfaw LM, Fenta HM. Multivariate logistic regression analysis on the association between anthropometric indicators of under-five children in Nigeria: NDHS 2018. BMC Pediatr. 2021;21(1):1–13.
- Fenta HM, Zewotir T, Muluneh EK. Spatial data analysis of malnutrition among children under-five years in Ethiopia. BMC Med Res Methodol. 2021;21(1):1–13.
- Takele K, Zewotir T, Ndanguza D. Understanding correlates of child stunting in Ethiopia using generalized linear mixed models. BMC Public Health. 2019;19(1):1–8.
- Zheng X, et al. Multistate Markov model application for blood pressure transition among the Chinese elderly population: a quantitative longitudinal study. BMJ Open. 2022;12(7): e059805.
- Matsena Zingoni Z, et al. Hiv disease progression among antiretroviral therapy patients in Zimbabwe: a multistate Markov model. Front Public Health. 2019;7:326.
- 22. Belan S. Median and mode in first passage under restart. Phys Rev Res. 2020;2(1):013243.

- 23. Warr RL, Woodfield TB. Bayesian nonparametric estimation of first passage distributions in semi-Markov processes. Appl Stoch Model Bus Ind. 2020;36(2):237–50.
- 24. Eisler Z, et al. Diffusive behavior and the modeling of characteristic times in limit order executions. Quant Finance. 2009;9(5):547–63.
- Schwarz W. Small steps at high speed: the diffusion limit. In: Random Walk and Diffusion Models: An Introduction for Life and Behavioral Scientists. Cham: Springer; 2022. p. 55–70.
- Mermoud, G. Model-Based Real-Time Control. In: Stochastic Reactive Distributed Robotic Systems. Springer Tracts in Advanced Robotics, vol 93. Cham: Springer; 2014. https://doi.org/10.1007/978-3-319-02609-1_12.
- Suit H, Wette R, Lindberg R. Analysis of tumor-recurrence times. Radiology. 1967;88(2):311–21.
- Huang R, et al. A retrospective analysis of the risk factors affecting recurrence time in patients with recurrent glioblastoma. Ann Palliat Med. 2021;10(5):5391399–5399.
- Guraya SY. Pattern, stage, and time of recurrent colorectal cancer after curative surgery. Clin Colorectal Cancer. 2019;18(2):e223–8.
- Owoeye SM, Oseni BM, Gayawan E. Estimating lifetime malnourished period and its statistics based on the concept of Markov chain with reward. Heliyon. 2020;6(5):e04073.
- Lives Y. A guide to Young Lives research. Lives Y, editor. Oxford; 2017. https://www.younglives.org.uk/sites/default/files/migrated/ GuidetoYLResearch_0.pdf.
- 32. Fenta HM, Zewotir T, Muluneh EK. Disparities in childhood composite index of anthropometric failure prevalence and determinants across Ethiopian administrative zones. PLoS One. 2021;16(9):e0256726.
- Khan S, Zaheer S, Safdar NF. Determinants of stunting, underweight and wasting among children< 5 years of age: evidence from 2012– 2013 Pakistan demographic and health survey. BMC Public Health. 2019;19:1–15.
- Gupta AK, Borkotoky K. Exploring the multidimensional nature of anthropometric indicators for under-five children in India. Indian J Public Health. 2016;60(1):68–72.
- Kassie GW, Workie DL. Exploring the association of anthropometric indicators for under-five children in Ethiopia. BMC Public Health. 2019;19:1–6.
- Simegnew M. Modeling the Progression of Neonatal Hypothermia Disease Progression Using Data on New Borns at Dilla University Referral Hospital Application of Multi State Hidden Markov Model. 2022, HU.
- 37. Visalakshi J. Estimation of mean transition time using markov model and comparison of risk factors of malnutrition using Markov regression to generalized estimating equations and random effects model in a longitudinal study. The Tamilnadu Dr. Chennai: MGR Medical University; 2012.
- Dong H, et al. Multi-state analysis of the impact of childhood starvation on the healthy life expectancy of the elderly in China. Front Public Health. 2021;9: 690645.
- Moreira C, et al. A multistate Model for analyzing transitions between body mass index categories during childhood: the Generation XXI Birth Cohort Study. Am J Epidemiol. 2019;188(2):305–13.
- Jackson C. Multi-state models for panel data: the msm package for R. J Stat Softw. 2011;38:1–28.
- 41. Spedicato G, Signorelli M. The markovchain Package: A Package for Easily Handling Discrete Markov Chains in R. 2014.
- 42. Kain J, et al. Trends in height and BMI of 6-year-old children during the nutrition transition in Chile. Obes Res. 2005;13(12):2178–86.
- International Institute for Population Sciences (IIPS) and ORC Macro, MEASURE/DHS+, National Family Health Survey (NFHS-2), 1998-99: India, vol. 1. International Institute for Population Sciences; 2000.
- Nahar B, et al. Risk factors associated with severe underweight among young children reporting to a diarrhoea treatment facility in Bangladesh. J Health Popul Nutr. 2010;28(5):476.
- 45. Kikafunda JK, et al. Risk factors for early childhood malnutrition in Uganda. Pediatrics. 1998;102(4):e45–e45.
- Wang W. Multi-state Markov models with fixed and random effects. UCL (University College London); 2022.
- Abner EL. Multistate Markov chains and their application to the Biologically Resilient Adults in Neurological Studies cohort. University of Kentucky; 2013.

Publisher's Note

Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.