medRxiv preprint doi: https://doi.org/10.1101/2024.11.01.24316562; this version posted November 2, 2024. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted medRxiv a license to display the preprint in

perpetuity. It is made available under a CC-BY 4.0 International license .

Positive and negative affect, related mental health traits, and cognitive 1 performance: shared genetic architecture and potential causality 2 3 Chloe Slaney^{1,2}, Naoise Mac Giollabhui³, Peter J. van der Most⁴, Ensor R. Palacios¹, Harold Snieder⁴, Michel Nivard¹, Gibran Hemani¹, Catharina A. Hartman⁵*, Golam M. Khandaker^{1,2,6,7}*. 4 5 ¹ Medical Research Council Integrative Epidemiology Unit, Bristol Medical School, University of 6 7 Bristol, UK. 8 ²Centre for Academic Mental Health, Population Health Sciences, Bristol Medical School, 9 University of Bristol, UK. 10 ³ Depression Clinical & Research Program, Department of Psychiatry, Massachusetts General 11 Hospital, USA. 12 ⁴Department of Epidemiology, University of Groningen, University Medical Center Groningen, 13 the Netherlands. 14 ⁵Interdisciplinary Center Psychopathology and Emotion Regulation, Department of Psychiatry, 15 University of Groningen, University Medical Center Groningen, the Netherlands. 16 ⁶NIHR Bristol Biomedical Research Centre, University Hospitals Bristol and Weston NHS 17 Foundation Trust and University of Bristol, UK. 18 ⁷Avon and Wiltshire Mental Health Partnership NHS Trust, Bristol, UK. 19 20 *Joint Senior Author 21 22 Corresponding author: Dr Chloe Slaney, MRC Integrative Epidemiology Unit, University of 23 Bristol, Oakfield House, Oakfield Grove, Bristol, BS8 2BN, Email: chloe.slaney@bristol.ac.uk. 24 25 Word count: 26 Abstract: 323 27 Manuscript: 5,691 (includes tables/figure titles/footnotes) 28 Keywords: Positive Affect; Negative Affect; Depression; Anxiety; Cognition; General Cognitive 29 Ability; Lifelines Cohort; GWAS; Genetic Correlation; Mendelian randomization. 30 31 32 33 34 35 NOTE: This preprint reports new research that has not been certified by peer review and should not be used to guide clinical practice.

medRxiv preprint doi: https://doi.org/10.1101/2024.11.01.24316562; this version posted November 2, 2024. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted medRxiv a license to display the preprint in

perpetuity. It is made available under a CC-BY 4.0 International license .

36

Abstract

37 Altered affect and cognitive dysfunction are transdiagnostic, burdensome, and pervasive features of many psychiatric conditions which remain poorly understood and 38 39 have few efficacious treatments. Research on the genetic architecture of these 40 phenotypes and causal relationships between them may provide insight into their aetiology and comorbidity. Using data from the Lifelines Cohort Study, we conducted 41 genome-wide association studies (GWAS) on positive and negative affect and four 42 cognitive domains (working memory, reaction time, visual learning and memory, 43 44 executive function). Using publicly available large GWAS on related - albeit distinctphenotypes (depression, anxiety, wellbeing, general cognitive ability [GCA]) we 45 conducted genetic correlation and Mendelian randomization (MR) analyses to examine 46 genetic overlap and causal relationships. We identified one genome-wide hit ($p < 5x10^{-8}$) 47 for reaction time, and many loci with suggestive associations ($p < 5x10^{-6}$; N range= 11-20 48 independent hits) for other phenotypes. For most phenotypes, gene mapping and 49 tissue expression analysis of suggestive hits from the GWAS showed increased gene 50 expression in brain tissue compared to other tissues. As predicted, negative affect is 51 genetically correlated with mental health phenotypes (depression r_g =0.51; anxiety 52 r_{e} =0.70; wellbeing r_{e} = -0.71) and cognitive domains are genetically correlated with GCA 53 and brain volume ($r_g \le 0.66$). Genetic correlations between negative and positive affect 54 suggest that they are dissociable constructs ($r_g = -0.18$) with negative affect having 55 higher genetic overlap with GCA than positive affect (r_g =-0.19 vs -0.06). This could 56 57 indicate that negative affect has a higher shared neural basis with GCA than positive 58 affect and/or GCA and negative affect may exhibit causal relationships. MR analyses suggest potential causal effects of higher GCA on reduced negative affect, reduced risk 59 of depression and anxiety, and higher wellbeing, but little impact on positive affect. We 60 also report evidence for potential causal effects of depression and lower wellbeing on 61 reduced GCA. Taken together, these results suggests that GCA may be a valid target for 62 negative affect (but not positive affect) and depression and wellbeing may be valid 63 64 targets for GCA.

65

1. Introduction 67

68 Positive and negative affect reflect the extent to which a person feels positive or negative emotions (e.g., excited, interested; nervous, distressed), respectively. Positive 69 70 and negative affect can be experienced in a state (temporary) or trait (stable) manner 71 (Watson et al., 1988) and are distinct transdiagnostic features of many health conditions, particularly depression and anxiety. Another critical yet often overlooked 72 transdiagnostic feature of many health conditions is cognitive dysfunction (Colwell et 73 74 al., 2022; Hilton et al., 2024). Despite altered affect and cognitive dysfunction being 75 burdensome and pervasive features of many conditions, they remain poorly understood and have few efficacious treatments (Colwell et al., 2022; Rush et al., 2006). 76

Understanding the genetic architecture of mental health and cognitive 77 78 phenotypes, and their relation to each other, can provide aetiological and biological 79 insights which may facilitate identification of novel treatment targets (Carey et al., 2021; 80 Minikel et al., 2024). Although observational studies have found associations between mental health and cognition (Dam et al., 2021; Rock et al., 2014; Zainal & Newman, 81 82 2021; although see Ball et al., 2024;), causality remains unclear (Suddell et al., 2023). 83 Reported associations may reflect: (1) poorer mental health causing poorer cognition, (2) poorer cognition causing poorer mental health, (3) confounding via shared risk 84 factors (e.g., stress, sedentary behaviour) (Mac Giollabhui, 2021). Crucially, studies 85 testing causal relationships between more narrowly defined mental health phenotypes 86 (e.g., affect) and cognitive domains (e.g., memory, attention) are needed (Chavez-87 88 Baldini et al., 2023). Evidence of bidirectional causality would strengthen the case that targeting poor mental health (either diagnostic conditions or transdiagnostic 89 phenotypes) may prevent and treat cognitive dysfunction, and vice versa. Two 90 91 approaches that can shed light on the aetiology, comorbidity, and/or causal 92 relationships between phenotypes are genetic correlations (using all genome-wide variants) (Bulik-Sullivan, Finucane, et al., 2015) and Mendelian randomization (MR) 93 (typically using genome-wide association study hits) (Davey Smith & Ebrahim, 2003). 94 MR can test for causality given certain assumptions are met (see Supplementary 95 Methods). It does this by using genetic variants robustly associated with the exposure 96 97 as a proxy for it. The properties of genetic variants (random assignment from parents, 98 fixed at conception) mean that they are less likely to be associated with confounders 99 and overcome issues of reverse causality (Davey Smith & Ebrahim, 2003; Sanderson et al., 2022). These methods provide powerful tools given the availability of large genome-100 wide association studies (GWAS) on well-defined phenotypes. 101

102 To date, GWAS on mental health phenotypes have largely been conducted on 103 diagnostic categories such as depression and anxiety (Carey et al., 2021). However, these conditions are heterogeneous, with diagnosed individuals showing diverse 104 105 symptom profiles. For example, a diagnosis of depression requires \geq 5/9 *diverse* symptoms to be present within a two-week period, one of which must be low mood or 106 107 anhedonia (Regier et al., 2013). In addition, symptoms are often not specific to a condition, meaning that GWAS on a condition may in reality capture a broader 108

phenotype. One potential approach to reducing phenotypic heterogeneity while 109 addressing the overlap with other conditions, is to focus on more narrowly defined traits 110 that more closely map onto biological systems (positive and negative affect, specific 111 cognitive domains) (Wendt et al., 2020). In line with the Research Domain Criteria 112 framework (Cuthbert, 2014; Insel et al., 2010), these are often transdiagnostic. GWAS 113 on these more narrowly defined transdiagnostic traits may potentially facilitate insights 114 115 into aetiology, comorbidity, and novel therapeutics.

In this study, we first conducted GWAS of negative and positive affect (assessed 116 117 using the Positive and Negative Affect Schedule [PANAS]) and cognitive task performance in several domains (executive function, working memory, visual learning 118 and memory, and reaction time). Second, we conducted gene mapping and tissue 119 expression analysis to gain greater insight into associated genetic variants. Third, we 120 121 investigated genetic overlap between mental health (negative affect, positive affect, 122 depression, anxiety, wellbeing) and cognitive (general cognitive ability [GCA], specific cognitive domains) phenotypes using genetic correlations. Fourth, we tested evidence 123 of potential causality between mental health and cognitive phenotypes using MR 124 analyses. This is the first GWAS conducted on positive and negative affect using well-125 126 validated measures of these phenotypes. To achieve this, we use a large Dutch population-based cohort, the Lifelines Cohort. The application of MR to test evidence of 127 causal relationships between affect and cognition is critical for improving our 128 129 conceptual understanding of these phenotypes.

130 We hypothesised that: (i) follow-up analyses on all GWAS show increased gene expression in brain tissue compared to other tissues; (ii) affect and cognitive measures 131 in the Lifelines Cohort genetically correlate with related (albeit distinct) phenotypes 132 133 from external GWAS. Specifically, negative affect positively correlates with 134 depression/anxiety and negatively correlates with wellbeing (reverse direction for positive affect), and higher cognitive performance on all domains correlates with GCA; 135 (iii) there are bidirectional causal relationships between mental health and cognitive 136 137 phenotypes. 138 139

140

141

142

143

144

145

Figure 1. Study Objectives. 147



148

2. Methods and Materials 149

150

151 2.1. Description of the Lifelines Cohort

Lifelines is a multi-disciplinary prospective population-based cohort study 152 examining in a unique three-generation design the health and health-related behaviours 153 of 167,729 persons living in the north of the Netherlands. It employs a broad range of 154 investigative procedures in assessing the biomedical, socio-demographic, behavioural, 155 physical and psychological factors which contribute to the health and disease of the 156 general population, with a special focus on multi-morbidity and complex genetics. 157 Participants were recruited between 2006-2013 via their GP (49%), participating family 158 members (38%), and self-registration on the Lifelines website (13%) (Scholtens et al., 159 160 2015). Exclusion criteria for GP recruitment were: insufficient knowledge of Dutch language, severe psychiatric or physical illness, limited life expectancy (<5 years) 161 (Scholtens et al., 2015). Baseline data included approximately: 15,000 children (0-17 162 years), 140,000 adults (18-65 years), 12,000 elderly individuals (65+ years). Following 163 baseline, participants completed follow-up questionnaires every 1.5 years and study 164 visits every 5 years (1st follow-up visit [2014-2017], 2nd follow-up visit [2019-2023]) 165 (Scholtens et al., 2015; Sijtsma et al., 2022). In this study, we included participants ≥18 166 years who were genotyped on a GWAS array and excluded participants who have 167 conditions with a significant cognitive sequel: Alzheimer's disease, other dementia, 168

169 epilepsy, multiple sclerosis, Parkinson's disease, and stroke.

170

171 2.2. Assessment of Positive and Negative Affect

The Positive and Negative Affect Scale (PANAS) was used. PANAS assesses 172 173 positive and negative affect using two separate sub-scales, each containing 10 items (example items: excited, upset, nervous) (Crawford & Henry, 2004; Watson et al., 1988). 174 Participants rate the extent they experienced each item during the last four weeks on a 175 176 five-point scale (ranging from 'not at all' to 'extremely'). The outcome is the summed 177 score on each subscale, which ranges from 10-50 (higher values reflect higher positive or negative affect). The scales have high internal consistency (a=.87) and moderate test-178 179 retest reliability over 8-weeks (positive r=.58; negative r=.48) (Watson et al., 1988). 180

100

181 2.3. Assessment of Cognitive Performance

182 Ruff Figural Fluency Test (RFFT) was used, which is a valid and reliable measure 183 of executive functioning (Ross, 2014). The task consists of five parts, each containing 35 184 identical five-dot patterns. Participants draw as many unique designs as possible within 185 one minute by connecting dots in different patterns (Kuiper et al., 2017). The primary 186 outcome is total number of unique designs.

Four tasks from the CogState Test Battery were administered in Lifelines, which 187 188 tap into specific cognitive domains: visual learning & memory (one card learning task), reaction time/attention (identification task and detection task) and working memory 189 (one-back task). As the detection and identification task both assess reaction time, we 190 include only the identification task in this study which has a larger sample size. We used 191 192 the recommended primary outcomes for all tasks. All outcomes were transformed 193 (reaction time [ms] on the identification task was log10 transformed, and accuracy on the one-back and one card learning task were arcsine transformed). On both memory 194 tasks, higher values reflect better memory; on the reaction time task higher values 195 reflect poorer performance (i.e., slower reaction time). See Supplementary Methods for 196 197 further information on dataset preparation.

198

199 2.4. Genetics Data and Quality Control

200 Genotyping was conducted using three chip arrays in three subsets of the 201 Lifelines cohort: (i) Illumina CytoSNP-12 Bead Chip v2 array (N=17,033), (ii) Infinium Global Screening Array (GSA) Beadchip-24 v1.0 (N=38,030), (iii) FinnGen Thermo Fisher 202 Axiom[®] custom array (Affymetrix; N=29,166). See Supplementary for quality control 203 (QC) and imputation procedures conducted by Lifelines. Following Lifelines QCs, 204 205 73,086 participants were potentially eligible for this study (CytoSNP N=14,942; GSA N=31,810; Affymetrix N=26,334). We applied additional QCs, removing: (i) duplicates 206 207 (individuals genotyped on >1 chip) and 1st-degree relatives between chips, (ii) non-European individuals, and (iii) genetic outliers, see Supplementary Figure S1. Following 208

medRxiv preprint doi: https://doi.org/10.1101/2024.11.01.24316562; this version posted November 2, 2024. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted medRxiv a license to display the preprint in

perpetuity. It is made available under a CC-BY 4.0 International license .

additional QCs, 58,713 participants with genetics data were included in this study 209 (CytoSNP N=7,632; GSA N=24,975; Affymetrix N=26,106). 210

211

2.5. GWAS Summary Statistics for Related Mental Health and Cognitive Phenotypes 212

To examine genetic overlap and potential causality between affect and cognitive 213 performance, we used summary statistics from large publicly available GWAS of 214 depression (N Cases=294,322; N Controls=741,438) (Als et al., 2023), anxiety (N 215 216 Cases=7,016; N Controls=14,745)(Otowa et al., 2016), wellbeing (N=2,370,390) 217 (Baselmans et al., 2019), general cognitive ability (GCA; N=373,617)(Lam et al., 2021), and brain volume (N=47,316) (Jansen et al., 2020). See Supplementary Methods and 218 Table S1-S2 for further details on these GWAS. 219 220

Statistical Analyses 221 2.6.

2.6.1. GWAS of Negative Affect, Positive Affect, and Cognitive Performance in Lifelines 222

We conducted GWAS (negative affect score, positive affect score, and cognitive 223 224 performance score on each task) in each genotyped subset separately using REGENIE, which accounts for relatedness (Mbatchou et al., 2021). Prior to conducting GWAS, 225 within each subset, we standardised all outcomes (mean=0; SD=1) and used PLINK 226 (Purcell et al., 2007) to clean the genetics data for REGENIE Step 1. We included 227 228 variants that met the following criteria: call rate (0.95), Hardy-Weinberg equilibrium (1e-6), minor allele count (100), minor allele frequency (0.01), not multi-allelic; and 229 230 restricted to individuals with low missingness (0.95). In each GWAS, covariates included age, sex, and top 10 genetic principal components. We excluded poorly imputed 231 232 variants (INFO score < 0.8). For each outcome, we meta-analysed GWAS across the 233 three subsets using STDERR model in METAL and applied genomic corrections (Willer et al., 2010). GWAS quality was inspected using GWAS Inspector 1.6.4.0 (Ani et al., 2021). 234 The standard p-value threshold of 5x10⁻⁸ was used to determine genome-wide 235 significance (for suggestive significance: p-value <5x10⁻⁶). Variants with MAF < 0.01 were 236 excluded from all follow-up analyses. 237

2.6.2. Tissue Specificity of Prioritised Genes 238

We used gene and tissue mapping to explore associated variants in each GWAS. 239 240 We used an online platform (FUMA) which integrates resources to annotate, prioritize, 241 and visualise the summary statistics (Watanabe et al., 2017). The SNP2GENE function 242 was used to annotate SNPs and prioritise genes at each locus using positional mapping. Using these prioritised genes, we used GENE2FUNC to investigate tissue 243 specificity of differentially expressed gene (DEG) sets using GTEx v8 data on 54 non-244 245 diseased tissue types (N≤838 adults). For further details, see Supplementary Methods.

2.6.3. Genetic Correlations 246

Linkage Disequilibrium Score Regression (LDSR) (Bulik-Sullivan et al., 2015; 247 Bulik-Sullivan et al., 2015) was used to estimate genetic correlations of mental health 248

and cognitive phenotypes. This included phenotypes from Lifelines (positive and 249 negative affect, cognitive performance), and phenotypes from large external GWAS 250 (depression, anxiety, wellbeing, GCA, brain volume). We used precomputed LD scores 251 calculated using 1000G European data. As low imputation guality may confound LDSR 252 we filtered to HapMap3 panel SNPs which tend to be well-imputed in most studies 253 (Bulik-Sullivan, Finucane, et al., 2015). If GWAS did not have a sample size column for 254 255 SNPs, we assumed the same sample size for all SNPs. We checked all phenotypes had 256 heritability z-scores > 4 in line with standard procedure (Bulik-Sullivan et al., 2015). For LDSR cleaning filters applied, see (Bulik-Sullivan et al., 2015). 257

258

2.7. Bidirectional Mendelian Randomization 259

MR was performed to investigate potential causal effect of mental health 260 phenotypes on cognition and vice versa. This was done in R v4.2.2 using TwoSampleMR 261 (Hemani et al., 2017) and CAUSE v1.2.0.0335 (Morrison et al., 2020). 262

We applied the following criteria to identify genetic instruments: (1) ρ < 5x10⁻⁸ (if 263 not available, a lenient $p < 5x10^{-6}$ threshold was used) (2) independent ($r^2 < 0.01$, kb = 264 265 1000; based on European clustering in the 1000 genomes reference panel using *ld_clump()* in the *ieugwasr* package (Hemani et al., 2024)), (3) minor allele frequency 266 (MAF) ≥1%. Any SNPs not available in the outcome GWAS were excluded. For primary 267 analyses, we used the Inverse-Variance Weighted method (>1 SNP available) or Wald 268 269 ratio (if only 1 SNP was available). Where multiple genetic variants were available, we conducted sensitivity analyses using different MR methods which have different 270 assumptions regarding the validity of the genetic instruments: MR-Egger (Bowden et al., 271 272 2015), weighted median (Bowden et al., 2016), weighted mode (Hartwig et al., 2017).

273 We also applied Steiger filtering to assess whether genetic variants have stronger associations with exposures than outcomes (Hemani et al., 2017). Given the possibility 274 275 of correlated pleiotropy, we additionally applied causal analysis using summary effect estimates (CAUSE) which accounts for correlated and uncorrelated pleiotropy 276 277 (Morrison et al., 2020). For CAUSE, we included genome-wide variants pruned using 278 default criteria (r²=0.01, p-value=0.003) based on European 1000 genomes reference panel. Given the possibility of population-level confounding (e.g., dynastic effects), we 279 280 also conducted within-sibship MR in the MR-base platform (Hemani et al., 2018) using publicly available within-sibship GWAS on wellbeing, depressive symptoms, and 281 cognitive ability (Howe et al., 2022); see Supplementary methods for more detail. We 282 also checked for heterogeneity (Cochran's Q-statistic) and pleiotropy (Egger intercept). 283 For further details on assumptions of different MR methods used here, see 284 Supplementary methods. 285

286 In our primary analysis, we included phenotypes which have well-powered GWAS as exposures (>1 SNP available: depression, wellbeing, GCA). In secondary 287 analysis, we included phenotypes which have less well-powered GWAS as exposures (1 288

medRxiv preprint doi: https://doi.org/10.1101/2024.11.01.24316562; this version posted November 2, 2024. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted medRxiv a license to display the preprint in

perpetuity. It is made available under a CC-BY 4.0 International license .

- SNP or lenient p-value criteria: anxiety, negative affect, positive affect) to investigate 289 whether similar patterns of effects are observed for these phenotypes. 290
- 3. Results 291
- 292
- GWAS of Negative Affect, Positive Affect, and Cognitive Performance in the 293 3.1. Lifelines Cohort 294

No SNPs met genome-wide significance threshold ($p < 5x10^{-8}$), except one SNP for 295 reaction time (rs2920287, MAF=0.04, p=1.907e-08). Visual inspection of the locus zoom 296 plot of the region around rs2920287 shows many nearby genes, with PSCA being in 297 closest proximity (Figure S4). Follow-up analyses of this SNP using the GWAS Catalog 298 did not show associations with any traits. Lowering the threshold to $p < 5x10^{-6}$ yielded 299 associated SNPs for all phenotypes (Table 1). For Manhattan plots, see Figures S2-S3. 300

- Table 1. GWAS Results of Positive Affect, Negative Affect, and Cognitive Performance in 301
- the Lifelines cohort. 302

Phenotype	N	Hits (p<5e-08)	Hits Clumped (1000kb, r ² =0.01)	Hits (p<5e-06)	Hits Clumped (1000kb, r²=0.01)
Positive affect	57,946	0	0	67	20
Negative affect	57,946	0	0	73	15
Executive function	36,563	0	0	108	12
Visual learning and memory	36,783	0	0	139	12
Reaction time	35,729	1	1	176	19
Working memory	36,349	0	0	56	11

303 Note: Variants with MAE < 0.01 excluded.

304

3.2. **Tissue Specificity of Prioritised Genes** 305

306 We used a lenient p-value threshold ($p < 5 \times 10^{-6}$) to identify SNPs for follow-up analysis in FUMA. For all phenotype's, prioritised genes were upregulated across 307 multiple tissue types, particularly brain tissue (see Supplementary Figures S5-S6 and 308 Tables S18-S23). Following FUMAs default Bonferroni corrections for multiple 309 comparisons, there was significant upregulation in brain tissue for positive affect 310 311 (substantia nigra) and visual learning and memory (cerebellum and cerebellar hemispheres) compared to other tissue types. For visualisation of tissue specificity of 312 prioritised genes, see Figures S5-S6. 313

- 314
- 315

medRxiv preprint doi: https://doi.org/10.1101/2024.11.01.24316562; this version posted November 2, 2024. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted medRxiv a license to display the preprint in perpetuity.

polpotally	•
It is made available under a CC-BY	4.0 International license .

316	3.3.	Genetic Correlations between Affect and Mental Health Phenotypes
317	3.3.1.	Negative Affect is Genetically Correlated with Mental Health Phenotypes
318		Negative affect score was genetically correlated with depression (r_{g} = 0.51
319	[SE=0.	05], $p=2.80E-28$), wellbeing (r_g = -0.71 [SE=0.05], $p=8.65E-41$), and anxiety (r_g =
320	0.70 [S	SE=0.18], $p=0.0002$). Positive affect showed weaker genetic correlations with
321	depres	ssion (rg= -0.11 [SE=0.04], p=0.005), wellbeing (rg= 0.30 [SE=0.05], p=2.38E-09),
322	with lit	ttle evidence for anxiety (r_g = -0.16 [SE=0.17], p =0.33).
323	3.3.2.	Cognitive Domains are Genetically Correlated with GCA
324		Cognitive performance in Lifelines tasks were genetically correlated with the
325	larges	t GWAS for GCA: executive function (r_g =0.66 [SE=0.07], p =1.39E-23), visual
326	learnir	ng & memory (rg=0.54 [SE=0.05], p=6.90E-29), working memory (rg=0.53 [SE=0.06],
327	p=1.23	3E-18), and reaction time (r_g = -0.39 [SE=0.06], p =1.48E-12).
328	3.3.3.	Compared to Positive Affect, Negative Affect has Stronger Genetic Correlation

329 *with GCA*.

There was a weak negative genetic correlation between positive and negative affect scores (r_g = -0.18 [SE=0.08], p=0.016). There was stronger evidence of genetic correlations between GCA and negative affect (r_g = -0.19 [SE=0.04], p=7.56E-06) compared to GCA and positive affect (r_g = -0.06 [SE=0.04], p=0.18).

334 For all results, see Figure 2 and Supplementary Table S3.

335

Figure 2. Genetic correlations between Lifelines phenotypes (negative affect, positive
affect, cognitive performance) and external phenotypes (depression, anxiety, wellbeing,
GCA, brain volume) using LDSR.



³³⁹ 340

GCA=General Cognitive Ability; EF=Executive Functioning; Memory=Visual Learning and Memory; Neg Affect=Negative Affect (PANAS); Pos Affect=Positive Affect (PANAS);

342 3.4. Potential Causal Effects: Results from Mendelian Randomization 343 3.4.1. Effect of Mental Health Phenotypes on Cognition

344 In primary analyses, across all MR methods there was consistent evidence for a potential causal effect of depression on lower GCA (IVW estimate: -0.14 [95% CI= -0.19 345 to -0.09], p=0.00000006) and of one SD increase in wellbeing on higher GCA (IVW) 346 estimate: 0.30 [95% CI= 0.13 to 0.46], p=0.0004) (Figure 3). These effects remained after 347 applying Steiger Filtering (Figure 4; Supplementary Tables S6-S9) and were supported by 348 349 MR CAUSE (Supplementary Table S24-S25). There was also evidence for a potential 350 causal effect of one SD increase in wellbeing on better executive functioning (IVW estimate: 0.24 [95% CI= 0.03 to 0.44], p=0.024), and weak evidence of an effect on 351 memory (IVW estimate: 0.19 [95% CI= -0.02 to 0.41], p=0.076). However, these effects 352 were not consistent across MR methods and were not supported following Steiger 353 354 filtering (Figure 4; Supplementary Tables S6-S9).

In secondary analyses, there was evidence for a potential causal effect of one SD increase in negative affect on poorer memory (IVW estimate: -0.15 [95% CI= -0.27 to -0.03], p=0.012) and one SD increase in positive affect on better executive functioning (IVW estimate: 0.16 [95% CI= 0.02 to 0.30], p=0.024). However, these effects were not consistent across MR methods, see Figures 4 and Supplementary Tables S10-S13.

For other mental health and cognitive phenotypes, there was little evidence of
causality based on IVW estimates (Figures 4; Supplementary Tables S6-S14).
Confidence intervals (CIs) using within-sibship GWAS were very large with imprecise
estimates (Figure S16).

There was evidence of heterogeneity, with the strongest evidence in analyses of depression and wellbeing on GCA, see Supplementary Table S15. There was little evidence of horizontal pleiotropy (based on Egger Intercept; $ps \ge 0.14$), except for weak evidence in MR analyses of depression on GCA (p=0.084) and executive function (p=0.048); see Supplementary Table S16.

369

370 3.4.2. Effect of GCA on Mental Health Phenotypes

There was evidence for a potential causal effect of one SD increase in GCA on 371 lower negative affect (IVW estimate: -0.11 [95% CI= -0.16 to -0.05], p=0.0002), reduced 372 373 risk of depression (OR: 0.88 [95% CI=0.83 to 0.93], p=0.00003), and anxiety (OR: 0.64 374 [95% CI=0.51 to 0.81], p= 0.0002), and higher wellbeing (IVW estimate: 0.02 [95% CI= 0.01 to 0.04], p=0.011) (Figure 3; Supplementary Table S4). However, effects were not 375 consistent across all MR methods: there was more consistent evidence for GCA on 376 negative affect (3/4 methods provide evidence supporting causality) compared to 377 378 depression (2/4), anxiety (2/4), and wellbeing (1/4).

Following Steiger Filtering, there was still evidence for a potential causal effect of one SD increase in GCA on lower negative affect (IVW estimate: -0.07 [95% CI= -0.11 to -0.02], *p*=0.007) but weaker evidence for one SD increase in GCA on reduced risk of

- anxiety (IVW estimate: -0.24 [95% CI= -0.50 to 0.02], p=0.072), although CIs were large 382
- (see Figure 4 and Supplementary Table S5). For GCA on depression and wellbeing, 383
- results were unaltered as no variants were removed in Steiger Filtering. MR CAUSE 384
- support that the data are consistent with a causal effect of GCA on negative affect, 385
- depression, and wellbeing (See Supplementary Tables S24-S25). Although MR CAUSE 386
- suggests the data for GCA on anxiety fit the causal model better than the null or sharing 387 model, this did not meet conventional p-value criteria (p>0.05). In within-sibship MR, CI 388
- 389 were very large with imprecise estimates (Figure S16).
- 390 There was strong evidence of heterogeneity in MR analyses testing effects of GCA on mental health phenotypes, except for GCA on anxiety (p=0.34); see 391 Supplementary Table S15. There was little evidence of horizontal pleiotropy based on 392 MR Egger Intercept ($ps \ge .15$) (Supplementary Table S16). 393
- 394

395

397 Figure 3. Mendelian Randomization Analyses Testing Evidence of Potential Causality between Mental Health and Cognition Phenotypes. 398



- 400 Note: Bars reflect 95% CIs. Axes differ to ensure inclusion of CIs for all analyses. For binary phenotypes (anxiety, depression), beta reflects log(OR). 401

399

Figure 4. Steiger Filtered Mendelian Randomization Analyses Testing Evidence of Potential Causality between Mental Health and Cognition Phenotypes.



Note: Bars reflect 95% Cls. Axes differ to increase visibility of results in different analyses; CI of MR-Egger (cognition on anxiety) is absent on graph due to very

large CI. For binary phenotypes (anxiety, depression), beta reflects log(OR).

4. Discussion 411

We conducted genome-wide association studies (GWAS) on negative and 412 positive affect using a well-validated scale (PANAS) and four cognitive domains 413 (executive function, working memory, visual learning and memory, reaction time) in the 414 Lifelines Cohort. We identified one genome-wide hit ($p < 5x10^{-8}$) for reaction time, and 415 many loci with suggestive associations ($p < 5 \times 10^{-6}$) for other phenotypes. As predicted, 416 gene mapping and tissue expression analysis of suggestive hits show higher gene 417 418 expression in brain tissue compared to other tissues for most phenotypes; negative 419 affect is genetically correlated with mental health phenotypes (depression $r_g=0.51$; anxiety r_g =0.70; wellbeing r_g = -0.71) and cognitive domains are genetically correlated 420 with GCA and brain volume ($r_g \le 0.66$). Genetic correlations between negative and 421 422 positive affect suggest that they are dissociable constructs ($r_g = -0.18$); with negative 423 affect having higher genetic overlap with GCA than positive affect (r_g =-0.19 vs -0.06). 424 Importantly, MR results suggest evidence of potential causal effects of higher GCA on reduced negative affect, reduced risk of depression and anxiety, and higher wellbeing 425 426 (with the most robust result being higher GCA on reduced negative affect), but little impact of GCA on positive affect. We also report evidence for potential causal effects of 427 428 depression and lower wellbeing on reduced GCA. Taken together, these results suggests 429 that GCA may be a valid target for reducing negative affect, but not for increasing positive affect; and depression and wellbeing may be valid targets for GCA. 430

4.1. Potential Causality between Mental Health and Cognitive Phenotypes 431

432 Across different MR methods, we found evidence of a potential causal effect of higher GCA on many mental health outcomes (reduced negative affect, reduced risk of 433 434 depression and anxiety, increased wellbeing). There was also evidence of potential 435 causal effects of depression and wellbeing on GCA. Whilst we did not observe evidence 436 for negative affect and anxiety on GCA, given the consistent direction of effect, this may 437 be due to these GWAS being smaller and having decreased statistical power when used as exposures. Our findings complement recent studies using MR analyses and 438 complementary designs (e.g., within-sibship analyses) reporting evidence of causality 439 440 between mental health phenotypes (depression, wellbeing, and/or anxiety) and 441 constructs related to cognition [e.g., educational attainment (Demange et al., 2024), education duration (Van De Weijer et al., 2024)]). 442

However, few studies have directly tested causality between mental health and 443 444 cognition. Using MR analyses in the ALSPAC cohort, Suddell et al. (2023) tested causality between mental health conditions (depression, anxiety) and specific cognitive 445 domains (response inhibition, working memory, emotion recognition) but reported that 446 estimates were imprecise likely due to limited statistical power (Suddell et al., 2023). 447 Marchi et al. (2024) used multivariable MR to test the causal effect of GCA and poverty 448 on several mental health conditions (including depression and anxiety) (Marchi et al., 449 2024). After adjusting for poverty, they found evidence that higher GCA may causally 450 reduce risk of depression and anxiety. Through inclusion of additional GWAS (negative 451 affect, positive affect, larger depression GWAS), our findings lend support to these 452

findings and additionally suggest: (1) higher GCA may also cause lower negative affect
(but not influence positive affect), (2) evidence of bidirectional effects (i.e., higher risk of
depression and lower wellbeing may also cause poorer GCA). Taken together, these MR
findings suggest a potential bidirectional causal relationship between mental health
and cognition.

Whilst our results could indicate a bidirectional causal effect, other explanations 458 (which violate MR assumptions) may account for these results including: (1) correlated 459 460 pleiotropy and (2) population-level confounding. Correlated pleiotropy occurs when 461 genetic variants used in MR analyses affect both the exposure and outcome via a shared heritable factor (Morrison et al., 2020). An example of how this could occur here 462 is through genetic variants impacting brain-related processes (e.g., synaptic plasticity) 463 which directly affect both cognition and mental health. Whilst we did include several 464 465 MR methods which allow for some correlated pleiotropy (e.g., weighted median, weighted mode, MR CAUSE) (Morrison et al., 2020), these methods could still give 466 biased estimates (and lead to incorrect conclusions regarding causality) if the majority 467 of instruments exhibit correlated pleiotropy. Another factor which may account for 468 these results is population-level confounding (e.g., assortative mating, dynastic effects) 469 470 (Brumpton et al., 2020). Whilst we are interested in direct effects from GWAS (i.e., 471 genetic variants effect on phenotypic variation), for many phenotypes GWAS will also pick up indirect effects (e.g., dynastic effects: parental genotype affecting offspring 472 473 phenotype via environmental factors). This is problematic as it violates MR assumptions 474 and could lead to incorrect conclusions (Brumpton et al., 2020). One approach 475 proposed to overcome this is to use within-family GWAS (Brumpton et al., 2020; Howe et al., 2022). Howe et al. (2022) found that GWAS estimates for some phenotypes, 476 477 including cognitive ability and depressive symptoms, were attenuated when using within-sibship GWAS compared to population-level GWAS (Howe et al., 2022). To try to 478 address this, we also conducted MR using within-sibship GWAS (depressive symptoms, 479 wellbeing, and cognitive ability) to test whether this impacted our findings. 480 Unfortunately, confidence intervals (CIs) were very large with imprecise estimates 481 (Supplementary Methods and Figure S16). This is unsurprising given the much smaller 482 483 sample size of within-sibship GWAS compared to population-level GWAS and highlights 484 the need for larger within-family GWAS on complex phenotypes like cognitive ability and depressive symptoms. 485

486 4.2. Additional Insights from Genetic Analyses on Negative and Positive Affect, and 487 Cognitive Performance.

Genetic correlation analyses reveal key insights into the genetic architecture of these phenotypes. First, as predicted, there was moderate genetic overlap between negative affect and mental health phenotypes (depression r_g =0.51; anxiety r_g =0.70; wellbeing r_g = -0.71), and between cognitive domains and cognition-related phenotypes (GCA, brain volume; $r_g \le 0.66$). This suggests that although these phenotypes are related, they are not interchangeable but rather have partly distinct genetic components. Second, as expected based on phenotypic correlations, genetic medRxiv preprint doi: https://doi.org/10.1101/2024.11.01.24316562; this version posted November 2, 2024. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted medRxiv a license to display the preprint in

perpetuity. It is made available under a CC-BY 4.0 International license .

correlation between negative and positive affect suggest they are dissociable 495 constructs as opposed to being opposite ends of the same spectrum ($r_g = -0.18$) 496 (Watson et al., 1988). Third, compared to positive affect, negative affect has higher 497 genetic overlap with GCA (r_g =-0.19 vs -0.06). This could indicate that negative affect has 498 a higher shared neural basis with GCA than positive affect and/or GCA and negative 499 affect may exhibit causal relationships (as suggested by our MR results). 500

Despite the negative affect GWAS having a much smaller sample size than the 501 502 depression GWAS (N=57,946 versus N=1,035,760), when used as an outcome, we 503 found more consistent evidence that GCA may play a causal role in negative affect 504 across MR methods. Speculatively, this may be because the GWAS on negative affect consists of a more homogeneous phenotype which may increase statistical power and 505 impact effect estimates (Manchia et al., 2013). This highlights the importance of future 506 507 studies conducting GWAS on more homogeneous phenotypes (Nagel et al., 2018). 508 Whilst we focused on transdiagnostic features of positive and negative affect and specific cognitive domains, there is a need for GWAS on other transdiagnostic features 509 [e.g., sleep disturbances, anhedonia, hot cognition (Roiser & Sahakian, 2013)]. It is 510 likely that advances will also be gained by parsing heterogeneity using other 511 512 approaches. For example, GWAS on depressed patients with specific characteristics [e.g., immune-metabolic depression (Milaneschi et al., 2020)]. Research focusing on 513 improving the validity of subtypes within and across psychiatric conditions will be 514 necessary for advancing our understanding of these conditions (Hammen, 2018). 515

516 4.3. Limitations

517 Limitations of this study must be considered when interpreting the results. First, smaller sample sizes for some GWAS resulted in a lack of genome-wide significant 518 variants ($p < 5x10^{-8}$) and/or larger CIs in MR analyses (positive and negative affect, 519 520 specific cognitive domains, anxiety). Consortia combining data from several large 521 datasets are necessary to provide well-powered GWAS on these phenotypes; our GWAS in the Lifelines Cohort will provide a useful contribution to this endeavour. Second, MR 522 523 estimates lifetime effect of an exposure (e.g., depression) on an outcome (e.g., GCA) 524 (Sanderson et al., 2022). Whilst this study is informative for understanding lifetime risk, it is unclear what time periods would be best to intervene on. This requires either a 525 randomised controlled trial (RCT; which would be expensive and time consuming) or MR 526 with large GWAS on exposures and outcomes at specific ages in the lifespan (Power et 527 al., 2023). Third, cognitive performance is highly related to other socioeconomic 528 529 phenotypes (e.g., education, socioeconomic status). Future studies testing 530 independent effects and interactions between these phenotypes on mental health using other methods (e.g., multivariable MR) would be useful (see Marchi et al., 2024). 531 Fourth, as discussed above, observed bidirectional causal relationships in MR analyses 532 533 between mental health and GCA could instead be due to violation of MR assumptions. Triangulating results from standard MR analyses with other methods (e.g., within-534 sibship MR) may help to increase confidence in conclusions drawn. Fifth, many GWAS 535 536 use data from large population-based cohorts which are less representative of some

populations (e.g., less affluent people) which may hinder generalizability of the findings. 537 Sixth, for continuous variables (e.g., GCA, negative affect, wellbeing), our study cannot 538 shed light on whether these relationships are nonlinear. As currently available nonlinear 539 MR approaches have provided implausible results (Wade et al., 2023), there is a need 540 for other methods to be used to characterise the shape of these relationships (see 541 Pines et al., 2024). Seventh, we focus on a subset of phenotypes, future studies should 542 543 expand this to provide insight into other mental health and cognitive phenotypes which 544 may show different relationships (e.g., schizophrenia, hot cognition) (Danhauer et al., 545 2013). Finally, many GWAS on psychiatric conditions do not exclude people with comorbidities. For example, the depression GWAS includes UK Biobank which defines 546 depression based on the following question: "Have you ever seen a general practitioner 547 (or psychiatrist) for nerves, anxiety, tension, or depression?". This may result in many 548 individuals with anxiety being characterised as having depression and makes it 549 challenging to conduct subsequent analyses testing genetic overlap/causality between 550 different conditions. 551

552 4.4. Implications

553 Our GWAS on positive and negative affect and cognitive domains in the Lifelines Cohort provide valuable resources which may facilitate insights into aetiology, 554 comorbidity, and causal risk factors for these phenotypes. We found evidence of 555 potential causal relationships between mental health phenotypes (negative affect, 556 depression, anxiety, wellbeing) and GCA. This may suggest that strategies targeting poor 557 mental health may prevent/treat cognitive dysfunction, and vice versa. However, to 558 increase confidence in this finding, triangulation using other methods which have 559 different strengths/limitations to MR, and consideration of highly related phenotypes 560 561 (e.g., education), are needed. Additionally, GWAS on other transdiagnostic phenotypes 562 are necessary to enable clearer insights into potential causal relationships. If multiple lines of evidence support causality, careful consideration of potential interventions 563 (e.g., age to intervene, length of intervention, whether interventions targeting 564 depression to reduce poorer GCA [or vice versa] could also have their own direct effect 565 of GCA) would be necessary. This could have important implications for clinical 566 practice (e.g., targeting depression may help prevent/treat cognitive impairments in 567 health conditions such as dementia; cognitive remediation therapy may help 568 569 prevent/treat depression). Considering the broader literature (Demange et al., 2024; Marchi et al., 2024), policy changes targeting factors impacting GCA [e.g., education 570 571 (Anderson et al., 2020)] would potentially be promising for reducing future mental health challenges in the general population. 572

Nevertheless, there is also evidence suggesting that cognitive impairments can
persist in remitted depressed individuals (Semkovska et al., 2019). This may appear to
contrast the idea that treating depression may help to prevent/treat cognitive
impairments. Speculatively, this could be because: (1) once depressive symptoms have
decreased, cognitive impairments reduce but require a longer time to observe effects
(RCTs may not have long enough follow-up lengths), (2) reducing depression may help

medRxiv preprint doi: https://doi.org/10.1101/2024.11.01.24316562; this version posted November 2, 2024. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted medRxiv a license to display the preprint in

perpetuity. It is made available under a CC-BY 4.0 International license .

to prevent cognitive impairments, but may not help treat them once they are already 579 experienced (i.e., a 'scar effect'), (3) treating depression may improve cognition in a 580 subset of people (not all individuals), and/or (4) some symptoms of depression when 581 treated may improve cognition (but not all symptoms when treated will improve 582 cognition). There is a need for research testing these different theories to better 583 understand the dynamic relationship between depression and cognition. 584

585 4.5. Conclusions

In summary, we conducted GWAS on transdiagnostic features of many health 586 conditions (positive and negative affect, four specific cognitive domains). We identified 587 one genome-wide hit (p<5x10⁻⁸) for reaction time, and many loci with suggestive 588 associations ($p < 5x10^{-6}$) for other cognitive phenotypes. Follow-up gene mapping and 589 tissue expression analyses of suggestive hits show higher gene expression in brain 590 tissue compared to other tissues for most phenotypes. Genetic correlation analyses 591 show that negative and positive affect are dissociable constructs, with negative affect 592 having higher genetic overlap with GCA than positive affect. Importantly, in MR 593 594 analyses, we found evidence of a potential causal effect of higher GCA on multiple 595 mental health phenotypes (reduced negative affect, depression, and anxiety; and increased wellbeing), with little evidence on positive affect. We also report evidence of 596 potential causal effects of depression and lower wellbeing on reduced GCA. Taken 597 together, as the most robust evidence was for GCA on negative affect, with little effect 598 on positive affect, this suggests that GCA may be a valid target for negative affect (but 599 600 not positive affect) and depression and wellbeing may be valid targets for GCA. Further research testing the relationship between depression and cognition using 601 complementary research designs is warranted, particularly there is a need for studies to 602 603 test different theories we proposed in the discussion.

605 Acknowledgements

- 606 The generation and management of GWAS genotype data for the Lifelines Cohort Study
- 607 is supported by the UMCG Genetics Lifelines Initiative (UGLI). UGLI is partly supported
- 608 by a Spinoza Grant from NWO, awarded to Cisca Wijmenga. The authors wish to
- 609 acknowledge the services of the Lifelines Cohort Study, the contributing research
- 610 centers delivering data to Lifelines, and all the study participants.
- 611 The Genotype-Tissue Expression (GTEx) Project was supported by the <u>Common Fund</u> of
- the Office of the Director of the National Institutes of Health, and by NCI, NHGRI,
- 613 NHLBI, NIDA, NIMH, and NINDS. The final data used for the analyses described in this
- 614 manuscript were obtained from: the GTEx Portal via FUMA on 01/07/2024 and
- 615 18/07/2024.

616 Ethics approval

The general Lifelines protocol has been approved by the UMCG Medical ethicalcommittee under number 2007/152.

619 Funding

- 620 The Lifelines initiative has been made possible by subsidy from the Dutch Ministry of
- Health, Welfare and Sport, the Dutch Ministry of Economic Affairs, the University
- 622 Medical Center Groningen (UMCG), Groningen University and the Provinces in the North
- 623 of the Netherlands (Drenthe, Friesland, Groningen).
- This work was supported by a UK Medical Research Council (MRC) grant to GMK
- 625 (MC_UU_00032/06) which forms part of the Integrative Epidemiology Unit (IEU) at the
- 626 University of Bristol. The grant also supports CS. NMG acknowledges funding support by
- 627 Harvard University's Mind Brain Behavior Interfaculty Initiative and National Institute of
- 628 Mental Health grant K23MH132893. GMK acknowledges additional funding support
- 629 from the Wellcome Trust (201486/Z/16/Z and 201486/B/16/Z), MRC (MR/W014416/1;
- 630 MR/S037675/1; and MR/Z50354X/1), and the UK National Institute of Health and Care
- 631 Research (NIHR) Bristol Biomedical Research Centre (NIHR 203315). GH is supported
- by the MRC (MC_UU_00032/01), and the NIHR Bristol Biomedical Research Centre
- 633 (NIHR 203315). The views expressed are those of the authors and not necessarily those
- 634 of the NIHR or the Department of Health and Social Care, UK.

635 Contributions

CS, GK, CH, NMG and GH conceptualized and designed the study. CS analyzed the data
and drafted the manuscript. All authors advised on the project/analysis and approved
the final version of the manuscript.

639 Conflicts of Interest

- 640 No conflicts of interest were reported.
- 641

64	12

642

643

References

Als, T. D., Kurki, M. I., Grove, J., Voloudakis, G., Therrien, K., Tasanko, E., Nielsen, T. T., 644 Naamanka, J., Veerapen, K., Levey, D. F., Bendl, J., Bybjerg-Grauholm, J., Zeng, B., 645 Demontis, D., Rosengren, A., Athanasiadis, G., Bækved-Hansen, M., Qvist, P., 646 Bragi Walters, G., ... Børglum, A. D. (2023). Depression pathophysiology, risk 647 prediction of recurrence and comorbid psychiatric disorders using genome-wide 648 analyses. Nature Medicine, 29(7), 1832–1844. 649 650 Anderson, E. L., Howe, L. D., Wade, K. H., Ben-Shlomo, Y., Hill, W. D., Deary, I. J., Sanderson, E. C., Zheng, J., Korologou-Linden, R., Stergiakouli, E., Davey Smith, G., 651 Davies, N. M., & Hemani, G. (2020). Education, intelligence and Alzheimer's 652 653 disease: evidence from a multivariable two-sample Mendelian randomization 654 study. International Journal of Epidemiology, 49(4), 1163–1172. Ani, A., Van Der Most, P. J., Snieder, H., Vaez, A., & Nolte, I. M. (2021). GWASinspector: 655 comprehensive quality control of genome-wide association study results. 656 657 Bioinformatics (Oxford, England), 37(1), 129–130. Ball, E. L., Morillo, L., Poyner, E., McIntosh, A. M., & Iveson, M. H. (2024). Cognitive 658 659 ability in early life and risk of depression in adulthood: A systematic review and 660 meta-analysis. Journal of Affective Disorders, 352, 498-508. 661 Baselmans, B. M. L., Jansen, R., Ip, H. F., van Dongen, J., Abdellaoui, A., van de Weijer, M. P., Bao, Y., Smart, M., Kumari, M., Willemsen, G., Hottenga, J. J., Boomsma, D. 662 I., de Geus, E. J. C., Nivard, M. G., & Bartels, M. (2019). Multivariate genome-wide 663 664 analyses of the well-being spectrum. Nature Genetics, 51(3), 445–451. 665 Brumpton, B., Sanderson, E., Heilbron, K., Hartwig, F. P., Harrison, S., Vie, G. Å., Cho, Y., Howe, L. D., Hughes, A., Boomsma, D. I., Havdahl, A., Hopper, J., Neale, M., 666 Nivard, M. G., Pedersen, N. L., Reynolds, C. A., Tucker-Drob, E. M., Grotzinger, A., 667 Howe, L., ... Davies, N. M. (2020). Avoiding dynastic, assortative mating, and 668 669 population stratification biases in Mendelian randomization through within-family 670 analyses. *Nature Communications 2020 11:1, 11(1), 1–13.* Bulik-Sullivan, B., Finucane, H. K., Anttila, V., Gusev, A., Day, F. R., Loh, P. R., Duncan, 671 672 L., Perry, J. R. B., Patterson, N., Robinson, E. B., Daly, M. J., Price, A. L., & Neale, B. 673 M. (2015). An atlas of genetic correlations across human diseases and traits. 674 Nature Genetics 2015 47:11, 47(11), 1236–1241. 675 Bulik-Sullivan, B., Loh, P. R., Finucane, H. K., Ripke, S., Yang, J., Patterson, N., Daly, M. J., Price, A. L., Neale, B. M., Corvin, A., Walters, J. T. R., Farh, K. H., Holmans, P. A., 676 Lee, P., Collier, D. A., Huang, H., Pers, T. H., Agartz, I., Agerbo, E., ... O'Donovan, M. 677 C. (2015). LD Score regression distinguishes confounding from polygenicity in 678 679 genome-wide association studies. Nature Genetics 2015 47:3, 47(3), 291–295. Carey, C., Strong, R., Huang, Y., Aslibekyan, S., Gentleman, R., Smoller, J., Wilmer, J., 680 Robinson, E., & Germine, L. (2021). Shared and Distinct Genetic Influences 681 682 Between Cognitive Domains and Psychiatric Disorder Risk Based on Genome-Wide Data. Biological Psychiatry, 89(9), S45-S46. 683

medRxiv preprint doi: https://doi.org/10.1101/2024.11.01.24316562; this version posted November 2, 2024. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted medRxiv a license to display the preprint in perpetuity.

perpetuit	у.
It is made available under a CC-B	Y 4.0 International license

684 685 686 687 688 689	Chavez-Baldini, U., Nieman, D. H., Keestra, A., Lok, A., Mocking, R. J. T., De Koning, P., Krzhizhanovskaya, V. V., Bockting, C. L. H., Van Rooijen, G., Smit, D. J. A., Sutterland, A. L., Verweij, K. J. H., Van Wingen, G., Wigman, J. T. W., Vulink, N. C., & Denys, D. (2023). The relationship between cognitive functioning and psychopathology in patients with psychiatric disorders: a transdiagnostic network analysis. <i>Psychological Medicine</i> , 53(2), 476.
690 691 692 693	Colwell, M. J., Tagomori, H., Chapman, S., Gillespie, A. L., Cowen, P. J., Harmer, C. J., & Murphy, S. E. (2022). Pharmacological targeting of cognitive impairment in depression: recent developments and challenges in human clinical research. <i>Translational Psychiatry 2022 12:1, 12</i> (1), 1–16.
694 695 696	Crawford, J. R., & Henry, J. D. (2004). The positive and negative affect schedule (PANAS): construct validity, measurement properties and normative data in a large non-clinical sample. <i>The British Journal of Clinical Psychology</i> , <i>43</i> (Pt 3), 245–265.
697 698 699	Cuthbert, B. N. (2014). The RDoC framework: facilitating transition from ICD/DSM to dimensional approaches that integrate neuroscience and psychopathology. <i>World Psychiatry</i> , <i>13</i> (1), 28–35.
700 701 702 703	Dam, V. H., Stenbæk, D. S., Köhler-Forsberg, K., Ip, C., Ozenne, B., Sahakian, B. J., Knudsen, G. M., Jørgensen, M. B., & Frokjaer, V. G. (2021). Hot and cold cognitive disturbances in antidepressant-free patients with major depressive disorder: a NeuroPharm study. <i>Psychological Medicine</i> , <i>51</i> (14), 2347–2356.
704 705 706 707 708 709	 Danhauer, S. C., Legault, C., Bandos, H., Kidwell, K., Costantino, J., Vaughan, L., Avis, N. E., Rapp, S., Coker, L. H., Naughton, M., Naylor, C., Terracciano, A., & Shumaker, S. (2013). Positive and negative affect, depression, and cognitive processes in the Cognition in the Study of Tamoxifen and Raloxifene (Co-STAR) Trial. <i>Neuropsychology, Development, and Cognition. Section B, Aging, Neuropsychology and Cognition</i>, 20(5), 532–552.
710 711 712	Davey Smith, G., & Ebrahim, S. (2003). "Mendelian randomization": Can genetic epidemiology contribute to understanding environmental determinants of disease? <i>International Journal of Epidemiology</i> , <i>32</i> (1), 1–22.
713 714 715	Demange, P. A., Boomsma, D. I., van Bergen, E., & Nivard, M. G. (2024). Educational attainment and psychiatric diagnoses: a national registry data and two-sample Mendelian randomization study. <i>Nature Mental Health 2024 2</i> :6, <i>2</i> (6), 668–679.
716 717	Hammen, C. (2018). Risk Factors for Depression: An Autobiographical Review. <i>Annual Review of Clinical Psychology, 14</i> (Volume 14, 2018), 1–28.
718 719 720	Hemani, G, Elsworth, B., Palmer, T., & Rasteiro, R. (2024). <i>ieugwasr: Interface to the</i> "OpenGWAS" Database API. R package version 1.0.1, https://mrcieu.github.io/ieugwasr/. https://github.com/MRCIEU/ieugwasr
721 722 723	Hemani, Gibran, Tilling, K., & Davey Smith, G. (2017). Orienting the causal relationship between imprecisely measured traits using GWAS summary data. <i>PLOS Genetics</i> , <i>13</i> (11), e1007081.
724	Hemani, Gibran, Zheng, J., Elsworth, B., Wade, K. H., Haberland, V., Baird, D., Laurin,

725	C., Burgess, S., Bowden, J., Langdon, R., Tan, V. Y., Yarmolinsky, J., Shihab, H. A.,
726	Timpson, N. J., Evans, D. M., Relton, C., Martin, R. M., Davey Smith, G., Gaunt, T.
727	R., & Haycock, P. C. (2018). The MR-base platform supports systematic causal
728	inference across the human phenome. <i>ELife, 7</i> .
729 730 731 732	 Hilton, R. A., Tozzi, L., Nesamoney, S., Kozlowska, K., Kohn, M. R., Harris, A., Clarke, S., & Williams, L. M. (2024). Transdiagnostic neurocognitive dysfunction in children and adolescents with mental illness. <i>Nature Mental Health 2024 2:3</i>, 2(3), 299–309.
733 734 735 736 737	 Howe, L. J., Nivard, M. G., Morris, T. T., Hansen, A. F., Rasheed, H., Cho, Y., Chittoor, G., Ahlskog, R., Lind, P. A., Palviainen, T., van der Zee, M. D., Cheesman, R., Mangino, M., Wang, Y., Li, S., Klaric, L., Ratliff, S. M., Bielak, L. F., Nygaard, M., Davies, N. M. (2022). Within-sibship genome-wide association analyses decrease bias in estimates of direct genetic effects. <i>Nature Genetics 2022 54</i>:5, <i>54</i>(5), 581–592.
738	Insel, T., Cuthbert, B., Garvey, M., Heinssen, R., Pine, D. S., Quinn, K., Sanislow, C., &
739	Wang, P. (2010). Research Domain Criteria (RDoC): Toward a New Classification
740	Framework for Research on Mental Disorders. <i>American Journal of Psychiatry</i> ,
741	167(7), 748–751.
742	Jansen, P. R., Nagel, M., Watanabe, K., Wei, Y., Savage, J. E., de Leeuw, C. A., van den
743	Heuvel, M. P., van der Sluis, S., & Posthuma, D. (2020). Genome-wide meta-
744	analysis of brain volume identifies genomic loci and genes shared with
745	intelligence. <i>Nature Communications 2020 11:1, 11</i> (1), 1–12.
746 747 748 749 750	 Lam, M., Chen, CY., Ge, T., Xia, Y., Hill, D. W., Trampush, J. W., Yu, J., Knowles, E., Davies, G., Stahl, E. A., Huckins, L., Liewald, D. C., Djurovic, S., Melle, I., Christoforou, A., Reinvang, I., DeRosse, P., Lundervold, A. J., Steen, V. M., Lencz, T. (2021). Identifying nootropic drug targets via large-scale cognitive GWAS and transcriptomics. <i>Neuropsychopharmacology</i>, <i>17</i>, 47.
751	Mac Giollabhui, N. (2021). Inflammation and depression: Research designs to better
752	understand the mechanistic relationships between depression, inflammation,
753	cognitive dysfunction, and their shared risk factors. <i>Brain, Behavior, & Immunity -</i>
754	<i>Health, 15</i> .
755 756 757	Manchia, M., Cullis, J., Turecki, G., Rouleau, G. A., Uher, R., & Alda, M. (2013). The impact of phenotypic and genetic heterogeneity on results of genome wide association studies of complex diseases. <i>PloS One</i> , <i>8</i> (10).
758	Marchi, M., Alkema, A., Xia, C., Thio, C. H. L., Chen, L. Y., Schalkwijk, W., Galeazzi, G.
759	M., Ferrari, S., Pingani, L., Kweon, H., Evans-Lacko, S., David Hill, W., & Boks, M. P.
760	(2024). Investigating the impact of poverty on mental illness in the UK Biobank
761	using Mendelian randomization. <i>Nature Human Behaviour 2024</i> , 1–13.
762	Mbatchou, J., Barnard, L., Backman, J., Marcketta, A., Kosmicki, J. A., Ziyatdinov, A.,
763	Benner, C., O'Dushlaine, C., Barber, M., Boutkov, B., Habegger, L., Ferreira, M.,
764	Baras, A., Reid, J., Abecasis, G., Maxwell, E., & Marchini, J. (2021). Computationally
765	efficient whole-genome regression for quantitative and binary traits. <i>Nature</i>
766	<i>Genetics 2021</i> 53:7, 53(7), 1097–1103.

medRxiv preprint doi: https://doi.org/10.1101/2024.11.01.24316562; this version posted November 2, 2024. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted medRxiv a license to display the preprint in

- perpetuity. It is made available under a CC-BY 4.0 International license .
- Milaneschi, Y., Lamers, F., Berk, M., & Penninx, B. W. J. H. (2020). Depression 767 768 Heterogeneity and Its Biological Underpinnings: Toward Immunometabolic 769 Depression. Biological Psychiatry, 88(5), 369-380. 770 Minikel, E. V., Painter, J. L., Dong, C. C., & Nelson, M. R. (2024). Refining the impact of 771 genetic evidence on clinical success. Nature 2024 629:8012, 629(8012), 624–629. 772 Morrison, J., Knoblauch, N., Marcus, J. H., Stephens, M., & He, X. (2020). Mendelian randomization accounting for correlated and uncorrelated pleiotropic effects using 773 774 genome-wide summary statistics. Nature Genetics 2020 52:7, 52(7), 740–747. 775 Nagel, M., Watanabe, K., Stringer, S., Posthuma, D., & Van Der Sluis, S. (2018). Item-776 level analyses reveal genetic heterogeneity in neuroticism. Nature 777 Communications 2018 9:1, 9(1), 1–10. 778 Otowa, T., Hek, K., Lee, M., Byrne, E. M., Mirza, S. S., Nivard, M. G., Bigdeli, T., Aggen, S. 779 H., Adkins, D., Wolen, A., Fanous, A., Keller, M. C., Castelao, E., Kutalik, Z., Der Auwera, S. V., Homuth, G., Nauck, M., Teumer, A., Milaneschi, Y., ... Hettema, J. M. 780 (2016). Meta-analysis of genome-wide association studies of anxiety disorders. 781 782 Molecular Psychiatry, 21(10), 1391–1399. 783 Pines, A., Tozzi, L., Bertrand, C., Keller, A. S., Zhang, X., Whitfield-Gabrieli, S., Hastie, T., Larsen, B., Leikauf, J., & Williams, L. M. (2024). Psychiatric Symptoms, Cognition, 784 and Symptom Severity in Children. JAMA Psychiatry. 785 Power, G. M., Sanderson, E., Pagoni, P., Fraser, A., Morris, T., Prince, C., Frayling, T. M., 786 787 Heron, J., Richardson, T. G., Richmond, R., Tyrrell, J., Warrington, N., Davey Smith, 788 G., Howe, L. D., & Tilling, K. M. (2023). Methodological approaches, challenges, 789 and opportunities in the application of Mendelian randomisation to lifecourse epidemiology: A systematic literature review. European Journal of Epidemiology 790 2023 39:5, 39(5), 501-520. 791 792 Purcell, S., Neale, B., Todd-Brown, K., Thomas, L., Ferreira, M. A. R., Bender, D., Maller, J., Sklar, P., De Bakker, P. I. W., Daly, M. J., & Sham, P. C. (2007). PLINK: A Tool Set 793 for Whole-Genome Association and Population-Based Linkage Analyses. American 794 Journal of Human Genetics, 81(3), 559. 795 Regier, D. A., Kuhl, E. A., & Kupfer, D. J. (2013). The DSM-5: Classification and criteria 796 797 changes. World Psychiatry, 12(2), 92-98. 798 Rock, P. L., Roiser, J. P., Riedel, W. J., & Blackwell, A. D. (2014). Cognitive impairment in 799 depression: A systematic review and meta-analysis. Psychological Medicine, 44(10), 2029–2040. 800 Roiser, J. P., & Sahakian, B. J. (2013). Hot and cold cognition in depression. CNS 801 Spectrums, 18(3), 139-149. 802 Ross, T. P. (2014). The reliability and convergent and divergent validity of the Ruff Figural 803 804 Fluency Test in healthy young adults. Archives of Clinical Neuropsychology: The 805 Official Journal of the National Academy of Neuropsychologists, 29(8), 806–817. Rush, A. J., Trivedi, M. H., Wisniewski, S. R., Nierenberg, A. A., Stewart, J. W., Warden, 806 807 D., Niederehe, G., Thase, M. E., Lavori, P. W., Lebowitz, B. D., McGrath, P. J.,

808 809 810 811	Rosenbaum, J. F., Sackeim, H. A., Kupfer, D. J., Luther, J., & Fava, M. (2006). Acute and longer-term outcomes in depressed outpatients requiring one or several treatment steps: a STAR*D report. <i>The American Journal of Psychiatry</i> , <i>163</i> (11), 1905–1917.
812 813 814	Sanderson, E., Glymour, M. M., Holmes, M. V., Kang, H., Morrison, J., Munafò, M. R., Palmer, T., Schooling, C. M., Wallace, C., Zhao, Q., & Davey Smith, G. (2022). Mendelian randomization. <i>Nature Reviews Methods Primers</i> , <i>2</i> (1), 1–21.
815 816 817 818	 Scholtens, S., Smidt, N., Swertz, M. A., Bakker, S. J. L., Dotinga, A., Vonk, J. M., Van Dijk, F., Van Zon, S. K. R., Wijmenga, C., Wolffenbuttel, B. H. R., & Stolk, R. P. (2015). Cohort Profile: LifeLines, a three-generation cohort study and biobank. International Journal of Epidemiology, 44(4), 1172–1180.
819 820 821 822	Semkovska, M., Quinlivan, L., O'Grady, T., Johnson, R., Collins, A., O'Connor, J., Knittle, H., Ahern, E., & Gload, T. (2019). Cognitive function following a major depressive episode: a systematic review and meta-analysis. <i>The Lancet Psychiatry</i> , 6(10), 851–861.
823 824 825	Sijtsma, A., Rienks, J., Van Der Harst, P., Navis, G., Rosmalen, J. G. M., & Dotinga, A. (2022). Cohort Profile Update: Lifelines, a three-generation cohort study and biobank. <i>International Journal of Epidemiology</i> , <i>51</i> (5), E295–E302.
826 827 828	Suddell, S., Mahedy, L., Skirrow, C., Penton-Voak, I. S., Munafò, M. R., & Wootton, R. E. (2023). Cognitive functioning in anxiety and depression: results from the ALSPAC cohort. <i>Royal Society Open Science</i> , <i>10</i> (8).
829 830 831	Van De Weijer, M. P., Demange, P. A., Pelt, D. H. M., Bartels, M., & Nivard, M. G. (2024). Disentangling potential causal effects of educational duration on well-being, and mental and physical health outcomes. <i>Psychological Medicine</i> , <i>54</i> (7), 1403–1418.
832 833 834	Wade, K. H., Hamilton, F. W., Carslake, D., Sattar, N., Davey Smith, G., & Timpson, N. J. (2023). Challenges in undertaking nonlinear Mendelian randomization. <i>Obesity</i> <i>(Silver Spring, Md.)</i> , <i>31</i> (12), 2887–2890.
835 836 837	Watanabe, K., Taskesen, E., Van Bochoven, A., & Posthuma, D. (2017). Functional mapping and annotation of genetic associations with FUMA. <i>Nature Communications 2017 8:1</i> , 8(1), 1–11.
838 839 840	Watson, D, Clark, L. A., & Tellegen, A. (1988). Development and validation of brief measures of positive and negative affect: the PANAS scales. <i>Journal of Personality and Social Psychology</i> , <i>54</i> (6), 1063–1070.
841 842 843	Watson, David, Clark, L. A., & Tellegen, A. (1988). Development and validation of brief measures of positive and negative affect: the PANAS scales. <i>Journal of Personality and Social Psychology</i> , 54(6), 1063–1070.
844 845 846	Wendt, F. R., Pathak, G. A., Tylee, D. S., Goswami, A., & Polimanti, R. (2020). Heterogeneity and Polygenicity in Psychiatric Disorders: A Genome-Wide Perspective. <i>Chronic Stress</i> , <i>4</i> .
847 848	Willer, C. J., Li, Y., & Abecasis, G. R. (2010). METAL: fast and efficient meta-analysis of genomewide association scans. <i>Bioinformatics</i> , <i>2</i> 6(17), 2190.

- 849 Zainal, N. H., & Newman, M. G. (2021). Depression and worry symptoms predict future executive functioning impairment via inflammation. Psychological Medicine. 850