

1 **FLEXIBILITY IN GENE COEXPRESSION AT DEVELOPMENTAL AND EVOLUTIONARY**  
2 **TIMESCALES**

3

4 Eva K Fischer<sup>1\*</sup>, Youngseok Song<sup>2</sup>, Wen Zhou<sup>3</sup>, Kim L Hoke<sup>5</sup>

5

6 <sup>1</sup> Department of Neurobiology, Physiology and Behavior, University of California Davis, Davis, CA

7 95616, USA

8 <sup>2</sup> Department of Statistics, West Virginia University, Morgantown, WV 26506, USA

9 <sup>3</sup> Department of Biostatistics, School of Global Public Health, New York University, New York, NY

10 10003, USA

11 <sup>5</sup> Department of Biology, Colorado State University, Fort Collins, CO 80523, USA

12

13 \*Author for correspondence:

14 Eva K Fischer

15 Department of Neurobiology, Physiology and Behavior

16 University of California Davis

17 One Shields Avenue

18 Davis, CA 95616

19 **ABSTRACT**

20 The explosion of next-generation sequencing technologies has allowed researchers to move from  
21 studying single genes, to thousands of genes, and thereby to also consider the relationships  
22 within gene networks. Like others, we are interested in understanding how developmental and  
23 evolutionary forces shape the expression of individual genes, as well as the interactions among  
24 genes. To this end, we characterized the effects of genetic background and developmental  
25 environment on brain gene coexpression in two parallel, independent evolutionary lineages of  
26 Trinidadian guppies (*Poecilia reticulata*). We asked whether connectivity patterns among genes  
27 differed based on genetic background and rearing environment, and whether a gene's  
28 connectivity predicted its propensity for expression divergence. In pursuing these questions, we  
29 confronted the central challenge that standard approaches fail to control the Type I error and/or  
30 have low power in the presence of high dimensionality (i.e., large number of genes) and small  
31 sample size, as in many gene expression studies. Using our data as a case study, we detail  
32 central challenges, discuss sample size guidelines, and provide rigorous statistical approaches  
33 for exploring coexpression differences with small sample sizes. Using these approaches, we find  
34 evidence that coexpression relationships differ based on both genetic background and rearing  
35 environment. We report greater expression divergence in less connected genes and suggest this  
36 pattern may arise and be reinforced by selection.

## 37 INTRODUCTION

38 Genes neither act nor evolve in isolation. Rather, genes are members of physically and  
39 functionally interacting networks. The nature of these interactions influences the degree to which  
40 changes in gene sequence and gene expression influence higher-level phenotypes, and therefore  
41 the extent to which sequence and expression changes are constrained at developmental and  
42 evolutionary timescales. On the one hand, genes with many interaction partners (i.e., those that  
43 occupy central ‘hub’ positions in a network) may be targets of developmental switches and  
44 evolutionary selection because they most strongly influence network output and higher-level  
45 phenotypes (Chateigner *et al.*, 2020; Friedman *et al.*, 2020). Alternatively, expression changes in  
46 highly connected genes may be constrained by pleiotropic effects imposed by their many  
47 connections (Jeong *et al.*, 2001; Hahn & Kern, 2005), and developmental and evolutionary  
48 changes may therefore be more prevalent in peripheral genes that have lower connectivity and  
49 presumably lower pleiotropic loads (Kim *et al.*, 2007; Mähler *et al.*, 2017). These ideas are of long-  
50 standing interest but have historically been difficult to test because physical interaction networks  
51 were well-characterized in only very few species (e.g. protein networks in yeast (Jeong *et al.*,  
52 2001; Hahn *et al.*, 2004; Jovelin & Phillips, 2009)) and simultaneously surveying expression in  
53 large numbers of genes was challenging, if not impossible. The proliferation of next generation  
54 sequencing technologies, and specifically RNA-sequencing (hereafter RNAseq), has removed  
55 these constraints from a technical perspective. Yet, analytical and statistical methods lag behind  
56 our ability to generate big data and – despite falling costs – the number of biological replicates  
57 remains small in many gene expression studies (Fischer *et al.* 2021). These factors impact  
58 network and gene coexpression analyses in particular.

59         Although network and coexpression analyses of RNAseq datasets remain less common  
60 than gene-wise differential expression analyses, studies in this area provide intriguing – albeit  
61 conflicting – results. Recent studies have shown that centrality in coexpression networks is  
62 negatively correlated with divergence in gene expression (Warnefors & Kaessmann, 2013; Mähler

63 *et al.*, 2017; Kuo *et al.*, 2023) as well as gene sequence evolution (Josephs *et al.*, 2017; Masalia  
64 *et al.*, 2017; Harnqvist, 2021). Conversely, genes in peripheral positions show greater magnitude  
65 expression divergence (Mähler *et al.*, 2017) and signatures of positive selection (Kim *et al.*, 2007).  
66 Collectively, these findings suggest evolutionary constraints imposed on genes with high network  
67 centrality and evolutionary flexibility in genes at the network periphery. Yet this conclusion is in  
68 opposition with evidence for a bias toward changes in the expression of and selection on genes  
69 with high centrality (Koubkova-Yu *et al.*, 2018; Chateigner *et al.*, 2020; Friedman *et al.*, 2020;  
70 Rennison & Peichel, 2022). Evidence for changes in central genes include examples for positive  
71 selection on genes with more interaction partners in human protein networks (Luisi *et al.*, 2015)  
72 and high centrality in coexpression networks of genes associated with lung cancer (Wachi *et al.*,  
73 2005). Together, these latter examples are more consistent with a contrasting hypothesis that  
74 central genes better predict phenotypic variation and are therefore targets of selection.

75         The contradictory evidence for how gene centrality influences gene expression and  
76 sequence evolvability begs the question of what leads to these opposing patterns. One eventuality  
77 is that the above alternatives are not mutually exclusive, but complementary. Indeed, an extensive  
78 survey of stress responses in *Arabidopsis thaliana* found expression differences associated with  
79 drought stress in peripheral genes but cold stress in central genes (Des Marais *et al.*, 2017).  
80 Similarly, some evidence suggests that changes in gene sequence versus gene expression may  
81 be favored for physiological versus morphological traits (Warnefors & Kaessmann, 2013) and that  
82 selection for sequence versus expression changes could act more strongly at central versus  
83 peripheral network positions. The authors of (Hämälä *et al.*, 2020) suggest a ‘goldilocks’  
84 phenomenon, in which intermediate levels of pleiotropy facilitate evolution while excessive  
85 pleiotropy inhibits evolution and insufficient pleiotropy makes selection inefficient (Hämälä *et al.*,  
86 2020). If these seemingly contradictory observations – across taxa as well as within species – are  
87 indeed driven by interactions within gene regulatory networks, then disentangling how interactions

88 among genes shape development and evolution requires understanding at what timescales and  
89 to what extent interactions among genes themselves change.

90 If the relationships among genes are flexible, then the degree of pleiotropy and its  
91 presumed consequences are not fixed. Understanding how changes in coexpression influence  
92 phenotypes is of keen interest for both basic and biomedical research (Gysi & Nowick, 2020;  
93 Stanford *et al.*, 2020). Evolutionary biologists are interested in whether and how changes in  
94 coexpression influence species' propensity for adaptation to novel and changing environments  
95 and whether changes in coexpression patterns can explain distinct genetic mechanisms  
96 underlying convergent phenotypes (Hu *et al.*, 2016; Koubkova-Yu *et al.*, 2018; Yu *et al.*, 2020b).  
97 Behavioral biologists are increasingly asking how 'rewiring' of coexpression networks drives  
98 behavioral changes in health and disease. For example, (Bloch *et al.*, 2021) found changes in  
99 coexpression associated with mating behavior in guppies, and meta-analyses implicated changes  
100 in gene coexpression patterns associated with behavioral disorders in humans (Gaiteri *et al.*,  
101 2014). Changes in coexpression patterns are also of broader biomedical interest as they have  
102 been documented as a feature of diverse cancers (Wachi *et al.*, 2005; Anglani *et al.*, 2014).

103 The above studies highlight both the interest in and the potential of coexpression analyses.  
104 This demand is being met by a growing collection of tools for (differential) gene coexpression  
105 analysis (Wang *et al.*, 2017; Chowdhury *et al.*, 2020; Tommasini & Fogel, 2023). Yet while these  
106 software packages make advanced network analyses accessible, they do not eliminate the  
107 statistical limitations of these approaches. These limitations arise primarily from the combination  
108 of small sample sizes and high-dimensional data (tens of thousands of genes) emblematic of  
109 transcriptomic studies. While sample sizes have increased as sequencing costs have decreased,  
110 per group sample sizes commonly remain less than ten. Pooling samples across experimental  
111 groups or multiple studies can bring the total experimental sample size into the range  
112 recommended for network analyses (e.g., N=20 by (Langfelder & Horvath, 2008; Ballouz *et al.*,  
113 2015)). However, when explicitly asking if coexpression relationships differ between two or among

114 a few experimental groups, small per group samples remain a problem. In brief, inference from  
115 coexpression analyses is problematic without sufficient sample sizes and power, but also when  
116 samples are pooled across experimental groups that can differ in coexpression structure. This  
117 leaves researchers trapped between an experimental rock and hard place.

118 Like others, we are interested in using transcriptomic analyses to understand the biological  
119 basis of complex phenotypes, and specifically in exploring changes in individual genes as well as  
120 the interactions among genes. To this end, we characterized the effects of genetic background  
121 (high-predation versus low-predation populations) and developmental environment (rearing with  
122 and without predator chemical cues) on brain gene coexpression in two parallel, independent  
123 evolutionary lineages of Trinidadian guppies (*Poecilia reticulata*). In Trinidad, downstream, high-  
124 predation fish have repeatedly and independently colonized upstream, low-predation  
125 environments (Gilliam *et al.*, 1993; Barson *et al.*, 2009; Willing *et al.*, 2010; Fraser *et al.*, 2015),  
126 leading to parallel adaptive changes in life-history, morphology, and behavior (Reznick *et al.*,  
127 1990, 2001; Endler, 1995; Reznick, 1997; Magurran, 2005). In other words, each river drainage  
128 represents a naturally replicated experiment demonstrating parallel phenotypic adaptation.  
129 Recent studies have used laboratory breeding designs to disentangle genetic from environmental  
130 influences, demonstrating that both evolutionary history with predators and developmental  
131 experience with predators shape life history (Torres Dowdall *et al.*, 2012), morphology (Torres-  
132 Dowdall *et al.*, 2012; Fischer *et al.*, 2013; Ruell *et al.*, 2013; Handelsman *et al.*, 2014), physiology  
133 (Handelsman *et al.*, 2013; Fischer *et al.*, 2014), and behavior (Huizinga *et al.*, 2009; Torres-  
134 Dowdall *et al.*, 2012; Fischer *et al.*, 2016b). Using this breeding (Figure 1), we asked (1) whether  
135 connectivity patterns among genes differed among groups, and (2) whether connectivity  
136 influenced a gene's propensity for expression divergence.

137 We previously demonstrated genetic and developmental differences in brain gene  
138 expression (Fischer *et al.*, 2021). Here, we were interested in testing the hypothesis that – in  
139 addition to expression changes in single genes – the relationships among genes are targets of

140 developmental and evolutionary processes. In pursuing these questions, we confronted the  
141 statistical challenges that standard approaches may fail to control Type I error and/or have low  
142 power when dimensionality is high (i.e., large number of genes) and sample size is small. From  
143 a biological perspective, we find evidence for flexibility in coexpression relationships based on  
144 both genetic background and rearing environment, suggesting that changes in the interactions  
145 among genes are associated with phenotypic divergence at developmental and evolutionary  
146 timescales. From a technical perspective, we present a case study for those interested in  
147 (differential) coexpression with small per group sample sizes. We discuss key challenges, set  
148 clear sample size guidelines to control Type I error while maintaining power, and provide rigorous  
149 statistical approaches for exploring coexpression differences even with small sample sizes that  
150 can be readily implemented for similar coexpression analyses in other studies.

151

152

## 153 **RESULTS**

### 154 *Preliminary analyses*

155 Guppies used in this study were second generation lab-born fish from unique family lines  
156 established from wild-caught high-predation (HP) and low-predation (LP) populations in the Aripo  
157 and Quare river drainages in the Northern Range mountains of Trinidad. At birth, we split second-  
158 generation siblings into rearing environments with (pred+) or without (pred-) predator chemical  
159 cues, and they remained in these environments until the completion of the experiment (Figure 1)  
160 (as in (Fischer *et al.*, 2016b)). In brief, each drainage therefore consists of a 2x2 factorial design  
161 that distinguishes genetic from developmental effects of predation. Pair-wise comparisons of  
162 biological relevant are: (1) HP pred- vs LP pred-, an experiment comparing populations reared in  
163 an environment lacking predator cues to identify genetic differences between populations; (2) HP  
164 pred+ vs HP pred-, to identify environmentally induced changes mimicking the situation in which  
165 high-predation fish colonize low-predation environments, i.e., “ancestral plasticity”; (3) LP pred-

166 vs LP pred+, to identify environmentally induced changes comparable to the situation in which  
167 low-predation fish are washed downstream and a measure of whether ancestral plasticity is  
168 maintained in the derived population; (4) HP pred+ vs LP pred+, to identify genetic differences  
169 when fish are raised with environmental cues of predation. We also compared the same  
170 experimental groups across drainages (e.g. HP pred+ in Aripo drainage vs HP pred+ in Quare  
171 drainage) to understand differences associated with parallel, independent evolutionary lineages.

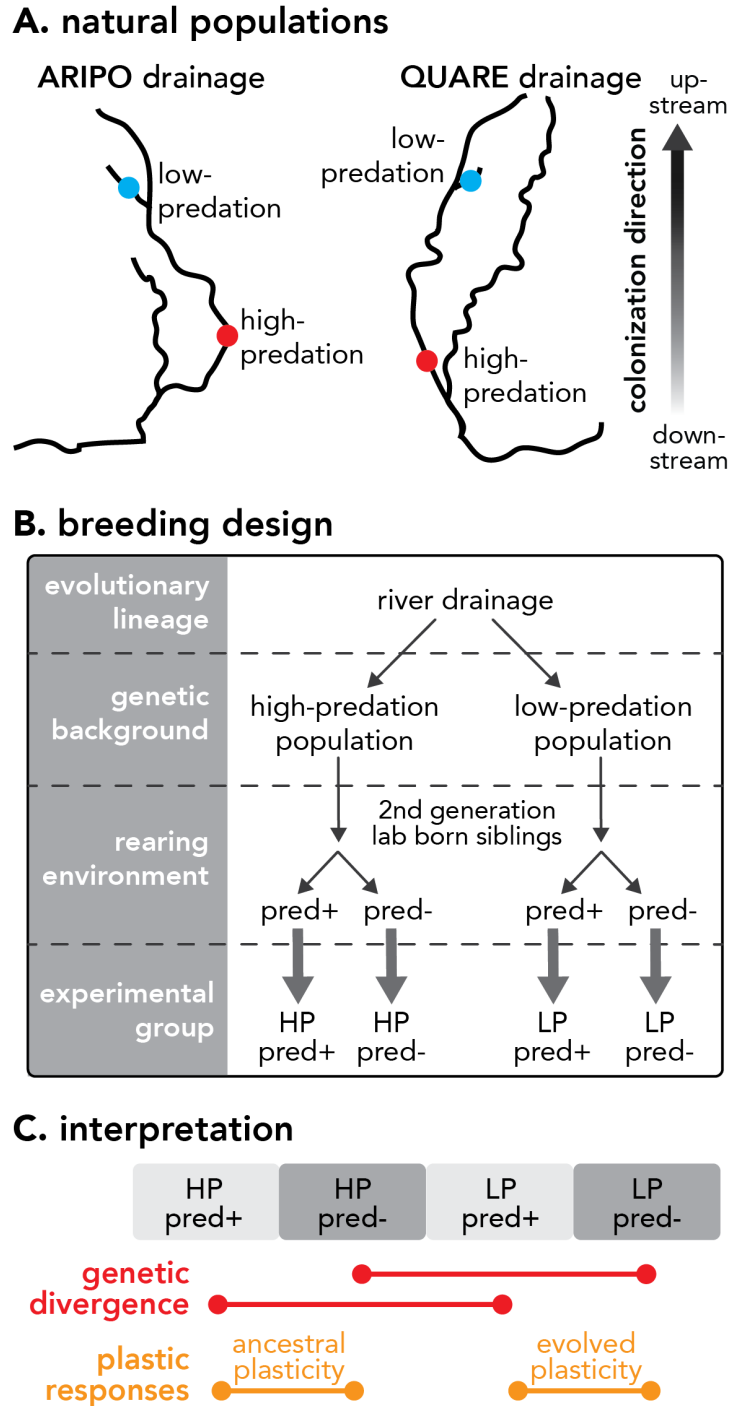
172

173 Our initial approach was to characterize and explore gene coexpression using the popular  
174 Weighted Gene Correlation Network Analysis (WGCNA) package in R (Langfelder & Horvath,  
175 2008). We calculated module preservation scores using the methods implements in WGCNA  
176 which combine a number of difference preservation statistics to calculate a summary preservation  
177 score (Langfelder *et al.*, 2011). We found that, in both drainages, ~50% of gene coexpression  
178 modules were not preserved across experimental groups (Supplemental Materials A). In other  
179 words, we identified substantial differences in network structure between groups, suggesting that  
180 the common practice of reconstructing coexpression networks by pooling samples across groups  
181 may not be valid. In brief, our preliminary analyses using WGCNA underscored the need for a  
182 statistical method to discern network differences across groups when dealing with small group  
183 sizes (N=10-12 in our case). We sought to address these issues through the alternative statistical  
184 approaches detailed in the Methods and Supplemental Materials B. We present results from the  
185 most promising approaches below.

186



187



**Figure 1.** Conceptual overview and interpretation. (A) Natural populations in different river drainages represent distinct evolutionary lineages. (B) Overview of laboratory breeding design disentangling genetic and environmental effects. (C) Interpretation of comparisons of interest between experimental groups resulting from 2x2 breeding design shown in (B). HP = high-predation, LP = low-predation, pred+ = reared with predator chemical cues, pred- = reared without predator chemical cues. Modified from Fischer et al. 2021.

188 *Changes in coexpression networks based on genetics and environment*

189 To overcome problems associated with very small per group sample sizes, yet high-dimensional  
190 data, we used random projection-based tests to compare covariance structures between  
191 experimental groups. We considered the set of all genes (DE and non-DE) that passed filtering  
192 criteria (Aripo: 13,446; Quare: 14,379). We found significant differences in covariance structures  
193 between high-predation and low-predation fish reared with predators (HP pred+ vs LP pred+) in  
194 both drainages (Figure 2; Table 1). Analysis of the Quare dataset found a marginally significant  
195 difference between high-predation fish reared with and without predators (HP pred+ vs HP pred).  
196 We also compared the covariance structures between the same treatment groups across  
197 drainages. Here, we found significant differences in all comparisons (Table 2). In short, when  
198 considering the collection of all genes, we found evidence for changes in gene coexpression  
199 based on evolutionary lineage (drainage), genetic background (population), and rearing  
200 environment.

201

**Table 1.** Approximated p-values from random projection tests comparing covariance structure for all genes (DE and non-DE) between treatment groups.

	<b>ARIPO drainage</b>	<b>QUARE drainage</b>
HP pred+ vs HP pred-	1.0000	0.0639
HP pred+ vs LP pred+	0.0149	<0.0001
HP pred- vs LP pred -	1.0000	0.9568
LP pred+ vs LP pred-	0.1258	0.9943

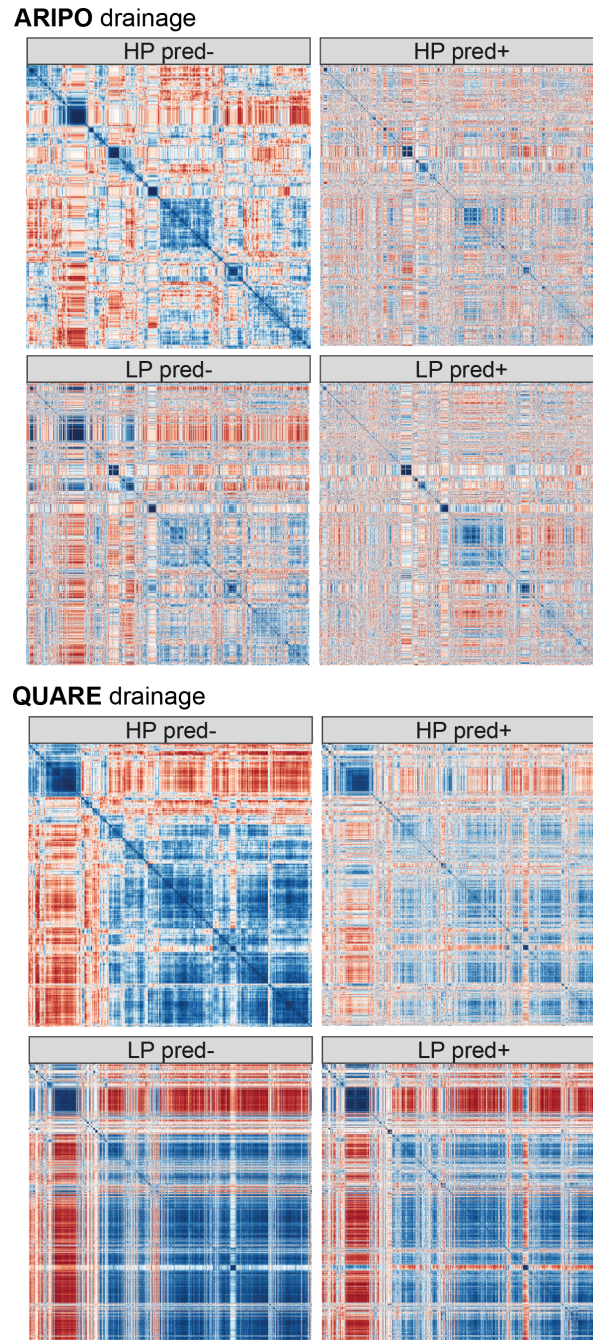
202

**Table 2.** Approximated p-values from random projection tests comparing covariance structure across drainages.

<b>ARIPO vs QUARE</b>	<b>p-value</b>
HP pred-	<0.0001
HP pred+	0.0001
LP pred -	0.0001
LP pred+	0.0107

203

204



**Figure 2.** Visualization of coexpression differences between experimental groups and evolutionary lineages. Correlation heatmaps provide a visual approximation of statistical differences based on genetic background (high- vs low-predation), rearing environment (with (pred+) or without (pred-) predators), and evolutionary lineage (AriPO and Quare drainage). Gene order is determined by hierarchical clustering of the high-predation pred- group, meaning that the same position in two heatmaps represents the correlation of identical pairs of genes. For ease of visualization and computation, only the 1,000 most variable genes are shown.

205 *Coexpression networks and differential expression*



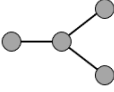


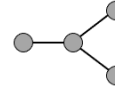
206 In addition to changes in coexpression among genes, we were interested in how interactions  
207 among genes might interface with a gene's propensity to be differentially expressed at  
208 developmental and evolutionary timescales. To explore this question, we compared coexpression  
209 networks between differentially expressed (DE) and non-differentially expressed (non-DE) genes.  
210 We performed these comparisons separately for each treatment group, given group-level  
211 differences in covariance structures (see above) and DE status being inherently related to  
212 differences in gene expression among treatment groups.

213 The comparison involved two steps: (i) reconstruction of the coexpression network using  
214 gene-wise correlations, and (ii) comparing two networks of different sizes. To achieve (i), we first  
215 tested whether the correlation between each pair of genes was zero, while controlling the FDR  
216 using the method from (Cai & Liu, 2016), as detailed in Supplemental Section 4.3. Using this  
217 method, an edge is drawn between any two genes (nodes) with nonzero correlation, forming the  
218 coexpression network. To assess the constructed network's sensitivity to different FDR levels, we  
219 compared network summary plots at multiple FDR cutoffs ( $\alpha = 0.01, 0.05, 0.1$ ). If two networks  
220 are distinct, their summary plots will differ (Maugis *et al.*, 2017). The network summary plots  
221 (Figure S6) suggest that the correlation-based coexpression network is relatively insensitive to  
222 different FDR levels. Therefore, we used a coexpression network with  $\alpha = 0.05$  for all subsequent  
223 analyses.

224 A major concern in comparing DE versus non-DE networks, which has been largely  
225 overlooked in literature, is that the different collections of genes in these two sets (e.g., DE genes  
226 are by definition a small subset of all genes) making the two corresponding networks have  
227 different numbers of unmatchable nodes [60]. To address this, we adopted the network  
228 comparison test proposed by (Shao *et al.*, 2022), which accommodates networks of different sizes  
229 by analyzing their network moments for specific motifs using the difference of two subgraph  
230 densities adjusted for their edge densities. Specifically, we used the v-shape (subgraphs with

231 three nodes and two edges), triangle (subgraphs with three nodes and three edges), and 3-star  
232 (subgraphs with four nodes and three edges) (see visualizations in Figure S9, Table 3) to compare  
233 the correlation-based coexpression networks. In network sciences, these subgraphs reflect  
234 connectivity and clusterability [60, 61, 62]. Apart from HP pred+ and HP pred- in the Aripo dataset,  
235 we observe that all non-DE networks had higher subgraph densities than their DE network  
236 counterparts (Figure S7 & S8; Table S8). Following this overall trend, we compared the different  
237 subgraph types described above. The Aripo dataset had fewer overall differences, especially for  
238 the 3-star subgraph type (Table S9). In the Quare dataset, DE versus non-DE networks differed  
239 in nearly all groups for all subgraph types (Table S9). To further compare subgraph densities in  
240 DE and non-DE networks, we conducted the same test with one-sided alternatives that DE gene  
241 network subgraphs are more or less connected than non-DE gene subgraphs. The results show  
242 that in the Aripo drainage, non-DE networks generally have higher sparsity-adjusted v-shape  
243 densities, while in the Quare drainage, the DE network in HPP has higher adjusted v-shape and  
244 triangle densities, but DE networks in LPP and LPNP have lower adjusted triangle and 3-star  
245 densities than their NDE counterparts (Table 3).

**Table 3.** Network comparisons of DE vs non-DE genes within treatment groups and across datasets. Comparisons of sparsity-adjusted subgraph densities tested the alternatives that DE gene networks had smaller or larger subgraph density than non-DE networks. P-values from one-sided alternative tests are reported for the v-shape, triangle, and 3-star subgraph types.

	ARIPO drainage			QUARE drainage		
	<u>v-shape</u>	<u>triangle</u>	<u>3-star</u>	<u>v-shape</u>	<u>triangle</u>	<u>3-star</u>
						
<b>DE &lt; non-DE</b>						
HP pred-	0.0051	0.8611	0.6149	<0.0001	0.9523	1.0000
HP pred+	0.0676	0.2330	0.5027	<0.0001	<0.0001	0.0665
LP pred -	0.0056	0.1700	0.2650	1.0000	1.0000	1.0000
LP pred+	<0.0001	0.0152	0.0749	0.4844	1.0000	1.0000
<b>DE &gt; non-DE</b>						
HP pred-	0.9948	0.1389	0.3742	1.0000	0.0482	<0.0001
HP pred+	0.9358	0.7676	0.5013	1.0000	1.0000	0.9347
LP pred -	0.9938	0.8277	0.7304	<0.0001	<0.0001	<0.0001
LP pred+	1.0000	0.9850	0.9252	0.5122	<0.0001	<0.0001

246

247

## 248 DISCUSSION

249 Our goal in this study was to understand how genetic background and rearing environment shape  
 250 relationships among genes. We previously characterized expression changes at the level of  
 251 individual genes (Fischer *et al.*, 2021), and here we were interested in exploring changes in  
 252 coexpression patterns among genes. Our findings suggest that coexpression patterns are flexible  
 253 at evolutionary and developmental timescales. Exciting from a biological perspective, exploring  
 254 these questions presents statistical challenges, in particular for RNAseq datasets characterized  
 255 by high-dimensionality and small sample sizes. We discuss the implications of our work from both  
 256 angles.

257 Gene expression studies remain plagued by small per-group samples sizes and high  
 258 dimensionality. Network construction is far from trivial, if not problematic, under these conditions,  
 259 especially when network structure – and not just network expression level – differs among  
 260 experimental groups. In our own study, we had an overall sample size of N=98 individuals, well

261 above the recommendation of  $N=30$  for network construction. However, this total sample size  
262 includes samples from two drainages and four experimental groups, and – based on our analyses  
263 here and preliminary analyses using the WGCNA package – we found evidence that network  
264 structure differs between experimental groups and even more strongly between drainages. These  
265 differences are of key biological interest as they suggest that expression relationships among  
266 genes (i.e., network structure) are subject to developmental plasticity and evolutionary change.  
267 However, if network structure differs across experimental groups, then networks must be  
268 constructed separately for each experimental group to avoid construction of ‘average’ networks  
269 that can obscure differences of (biological) interest and lead to biased conclusions (Zhao *et al.*,  
270 2014; Shojaie, 2021; Li *et al.*, 2022; Sai Li & Li, 2023). To take an extreme example, if two genes  
271 have opposing correlations of the same magnitude in two groups, the average correlation across  
272 groups will be zero. Thus, it is the per-group sample size that is most important for network  
273 construction and comparison when gene coexpression patterns are of interest.

274 While our per-group sample size of  $N=10-15$  is relatively large for an RNAseq study, it is  
275 below the recommended threshold for network construction, such as the minimum sample of  
276  $N=20$  suggested for RNAseq analyses by (Ballouz *et al.*, 2015). As we illustrate in the  
277 Supplemental Materials, these sample sizes are surprisingly inadequate for recently developed  
278 high-dimensional statistical tests thought to be robust against high-dimensionality, to control Type  
279 I error, and to maintain power. Indeed, from our simulation experiments, most common methods  
280 require  $N>50$  to retain the generally accepted nominal significance levels of 0.05 and satisfactory  
281 power exceeding 0.8. Importantly, the potential misinterpretations resulting from these  
282 shortcomings are not systematic (i.e., directionally biased) and therefore difficult to predict. As a  
283 growing number of studies consider how interactions among genes shape phenotypic differences  
284 across timescales, we present our work as a case study to increase awareness of these  
285 limitations, present complementary statistical approaches to those commonly used, and in hopes  
286 that others will consider these issues in experimental design and analysis.

287           We first investigated whether gene coexpression patterns differed based on genetic  
288 background and rearing environment. Using pairwise expression covariance tests, we identified  
289 differences in coexpression patterns between fish reared with and without predators, fish adapted  
290 to high- versus low-predation environments, and fish from distinct evolutionary lineages.  
291 Differences in coexpression structures in both drainages were most pronounced between high-  
292 and low-predation fish reared with predators (HP pred+ vs LP pred+). This comparison represents  
293 the ancestral population adapted to life with predators (HP pred+) versus the derived low-  
294 predation population adapted to predator and suddenly re-exposed to predator cues (e.g., as  
295 when fish are washed downstream; LP pred+). Fish adapted to a low-predation life are expected  
296 to be poorly equipped deal with the sudden stressors of predation. Indeed, we previously found  
297 HP pred+ fish to be behaviorally least variable and LP pred+ fish to be behaviorally most variable  
298 (Fischer *et al.*, 2016b). In light of findings here, we suggest that these unpredictable behavioral  
299 patterns could arise from disruptions in gene coexpression networks in the brain.

300           In addition to differences within drainages, we found more evidence for coexpression  
301 differences in the Quare as compared to the Aripo drainage, and that differences were ubiquitous  
302 when comparing between the two drainages. We suggest these patterns arise – at least in part –  
303 from the extent of genetic divergence between populations: high- and low-predation populations  
304 in the Quare drainage show greater genetic (Willing *et al.*, 2010) and gene expression (Fischer *et*  
305 *al.*, 2021) divergence than those in the Aripo drainage, and the two drainages represent distinct  
306 evolutionary lineages (Willing *et al.*, 2010). The importance of genetic background in shaping  
307 evolutionary trajectories is highlighted by our previous work demonstrating distinct underlying  
308 mechanisms associated with parallel phenotypic adaptation in guppies from distinct evolutionary  
309 lineages (Fischer *et al.*, 2016a, 2021). Similar mechanistic flexibility has also been demonstrated  
310 in other systems (Cordero *et al.*, 2018; Jacobs *et al.*, 2020), including those known for parallel  
311 phenotypic evolution (e.g. (Laporte *et al.*, 2015; Hanson *et al.*, 2017; Bolnick *et al.*, 2018)). Our  
312 findings here extend these observations from the expression of individual genes to coexpression



313 patterns among genes, suggesting that alternative gene expression network configurations can  
314 give rise to shared organism-level phenotypes.

315         What are the potential consequences of changes in gene coexpression relationships at  
316 developmental and evolutionary timescales? If pleiotropic interactions among genes constrain  
317 their propensity for change, then flexibility in these relationships could reduce pleiotropic load  
318 (e.g., (Wang *et al.*, 2010; Pavlicev & Wagner, 2012; Pavličev & Cheverud, 2015). In other words,  
319 genes may be more able to change in expression and drive phenotypic change if their interactions  
320 with other genes can be altered to minimize off-target effects. Conversely, flexibility in  
321 coexpression relationships might improve the ability of underlying gene expression networks to  
322 buffer higher-level phenotypes through homeostatic change (e.g., (Fischer *et al.*, 2016a; Badyaev,  
323 2018; Hoke *et al.*, 2019)). Importantly, either scenario implies that the relationships among genes  
324 may themselves be targets of selection. Alternatively, changes in gene coexpression could  
325 represent transcriptional noise if these changes do not amount to selectable differences at the  
326 network and/or organismal level. While this last scenario is less interesting from an adaptationist  
327 perspective, such ‘neutral’ changes may nonetheless have consequences for evolutionary  
328 trajectories, for example by giving rise to cryptic variation that is revealed under novel  
329 environmental conditions (West-Eberhard, 2003; McGuigan & Sgrò, 2009; Paaby & Rockman,  
330 2014). In brief, all three alternatives highlight that coexpression relationships can change at  
331 developmental and evolutionary timescales with consequences in the short and long term.

332         To begin to address the above alternatives, we asked whether connectivity influenced a  
333 gene’s propensity for expression divergence. This question has been considered at various levels,  
334 including in protein-protein interactions (Kim *et al.*, 2007; Luisi *et al.*, 2015), gene expression  
335 networks (Oldham *et al.*, 2006; Bloch *et al.*, 2021), gene sequence changes (Josephs *et al.*, 2017;  
336 Makinen *et al.*, 2018), and the relationship between expression and sequence divergence  
337 (Harrison *et al.*, 2012; Warnefors & Kaessmann, 2013; Chateigner *et al.*, 2020). Overall, we found  
338 that networks of non-DE genes were denser than networks of DE genes. This suggests that DE

339 genes are less well-connected than their non-DE counterparts, supporting the idea that differential  
340 expression may be facilitated by lower connectivity and decrease pleiotropy. Comparing across  
341 lineages, we observed more widespread differences in network connectivity in the Quare as  
342 compared to the Aripo drainage. As high- and low-predation populations in the Quare drainage  
343 are more genetically diverged than those in the Aripo (Willing *et al.*, 2010), this observations  
344 suggests that greater genetic divergence may increase network sparseness, suggesting that  
345 selection can reshape network structure. Taken together, our findings support greater expression  
346 divergence in less connected genes and suggest this pattern could facilitate changes in network  
347 topology if there is feedback between greater magnitude expression change and decreased  
348 network connectivity.

349

### 350 *Conclusions*

351 Understanding how underlying genetic architecture shapes the maintenance and evolution of  
352 complex traits is a fundamental goal of biological research. Over the past two decades, the  
353 explosion of next-generation sequencing technologies has allowed us to move beyond the genetic  
354 scale – considering one or a few genes or loci – to genomic scales – considering thousands to  
355 tens of thousands of genes or loci. Among the key advances afforded by these approaches, are  
356 the ease of conducting broadscale, exploratory studies; the opportunity to characterize underlying  
357 mechanisms in non-model species; and the ability to consider genes in the context of their  
358 interactions. As a growing number of studies consider how interactions among genes shape  
359 phenotypic differences across timescales, we provide a case study to increase awareness of  
360 limitations and provide suggestions for analysis. Our findings provide intriguing evidence of  
361 extensive coexpression flexibility at multiple timescales in a species known for rapid adaptation.  
362 The generality of these phenomena and their consequences for adaptation will be revealed by  
363 more studies with larger sample sizes and new statistical approaches. Understanding whether  
364 and how relationships among genes change at developmental and evolutionary timescales has

365 consequences for our understanding of how underlying mechanisms shape flexibility and  
366 robustness in higher order phenotypes, how animals adapt to novel and changing environments,  
367 and how behavior and physiology are regulated in health and disease.

368

369

## 370 **METHODS**

### 371 *Fish collection and rearing*

372 Samples here are the same as those described in (Fischer *et al.*, 2021). Briefly, guppies used in  
373 this study were second generation lab-born fish from unique family lines established from wild-  
374 caught high-predation (HP) and low-predation (LP) populations in the Aripo and Quare river  
375 drainages in the Northern Range mountains of Trinidad. At birth, we split second-generation  
376 siblings into rearing environments with (pred+) or without (pred-) predator chemical cues, and  
377 they remained in these environments until the completion of the experiment (Figure 1) (as in  
378 (Fischer *et al.*, 2016b)). Guppies were individually housed in 12:12 hour light cycle and fed a  
379 measured food diet once daily. All experimental methods were approved by the Colorado State  
380 University Animal Care and Use Committee (Approval #12-3818A).

381

### 382 *Tissue collection and processing*

383 We collected brain tissue from mature males in the groups described above 10 minutes after  
384 lights on in the morning. We extracted RNA from whole brains using the Qiagen RNeasy Lipid  
385 Tissue Mini Kit (Qiagen, Germany) and constructed a sequencing library for each individual using  
386 the NEBNext Ultra RNA Library Prep Kit for Illumina (New England Biolabs, Massachusetts,  
387 USA). For the Aripo dataset, 40 samples (N=10 per group) were pooled with unique barcodes  
388 into eight samples per sequencing library and each library was sequenced on a single lane. For  
389 the Quare dataset, 60 samples (N=12-16 per group) were pooled into three sequencing libraries  
390 with 20 samples per pool and each library was sequenced in two separate lanes. Libraries were

391 sequenced as 100bp paired-end reads on an Illumina HiSeq 2000 at the Florida State University  
392 College of Medicine Translational Science Laboratory (Tallahassee, Florida) in May 2014 (Aripo  
393 dataset) and January 2016 (Quare dataset).

394

#### 395 *Differential expression analysis*

396 We reported results of differential expression analyses in another study (Fischer *et al.*, 2021), and  
397 we use the resulting differential expression status (DE versus non-DE) as an important criterion  
398 in analyses performed here (see below). Briefly, we normalized read counts using DESeq2 (Love  
399 *et al.*, 2014) and performed differential expression analysis using the lme4 package in R  
400 ([github.com/lme4](https://github.com/lme4)). We used generalized linear mixed models with a negative binomial link to  
401 accommodate our experimental design and data type, including family and sampling week as  
402 random effects. Population of origin (HP / LP), rearing environment (pred+ / pred-), and their  
403 interaction were included as fixed effects. We used the residuals in the covariance and correlation  
404 analyses described below to adjust for fixed and random effects included in model fitting. To  
405 compare differentially expressed and non-differentially expressed genes, we adjusted p-values  
406 for multiple hypothesis testing using a direct approach for FDR control (Storey, 2002) as  
407 implemented in the fdrtool package in R (Strimmer, 2008). We considered transcripts differentially  
408 expressed (DE) if the adjusted p-value was  $<0.05$ , and all other genes non differentially expressed  
409 (non-DE).

410

#### 411 *Challenges from small sample, high dimensional data*

412 In wanting to understand differences in covariance structure and changes in network architecture  
413 between our experimental groups, we faced two fundamental challenges. First, controlling Type  
414 I error becomes difficult when comparing large networks or covariance structures with limited  
415 sample sizes, often leading to spurious discoveries and unreliable results. Also, detecting true  
416 effects for separating complicated networks or covariance structures in ultra-high dimensional

417 datasets is difficult due to low statistical power; a common phenomenon in coexpression analysis  
418 of RNAseq studies. While there are commonly accepted methods for two sample covariance  
419 tests, they either fail to control the Type I error rate with extremely small sample size (e.g.,  $N < 25$   
420 per group) or have substantially low power (see Supplemental Materials for simulations) (Li &  
421 Chen, 2012; Cai *et al.*, 2013; Chang *et al.*, 2017; Yu *et al.*, 2020a). Second, comparison of multiple  
422 networks is nontrivial, particularly when the networks are of different sizes or have unmatching  
423 nodes (Tang *et al.*, 2017; Agterberg *et al.*, 2020; Qi *et al.*, 2024). For gene coexpression analysis  
424 these issues apply, for example, when sample sizes vary between groups and are fairly small due  
425 to the limitation of experiment constraints, or when comparing subsets of genes of interest that  
426 vary in size (e.g., comparing non-DE vs DE gene sets, or differently sized coexpression modules)  
427 (Agterberg *et al.*, 2020; Alyakin *et al.*, 2024; Jin *et al.*, 2024; Qi *et al.*, 2024).

428 To highlight these challenges, we first present simulation studies using existing high-  
429 dimensional methods designed for valid inference on comparing large covariance structures with  
430 controlled Type I error rate and reasonable power (Li & Chen, 2012; Cai *et al.*, 2013; Chang *et*  
431 *al.*, 2017; Yu *et al.*, 2020a), before presenting our method and examining the real data set (see  
432 below). We summarize the outcomes of the simulation studies and exploratory comparisons here  
433 and refer the interested reader to additional details provided in the Supplemental Information.

434 We found that Type I error rates for existing tests were uncontrolled for small sample sizes  
435 of  $N < 50$  per group, even when the number of genes was relatively small (250 genes, orders of  
436 magnitude smaller than what is typical for RNAseq analysis) (Figure S1). In addition to  
437 uncontrolled Type I error, the existing methods were substantially underpowered for small sample  
438 sizes. Specifically, the empirical power was overall low ( $< 0.25$ ) for sample sizes  $N < 30$  per group  
439 (Figure S2). These issues plagued our dataset, which is representative of most RNAseq studies  
440 exploring connections between gene expression and behavior (per group samples  $N < 10$ ,  
441  $\sim 20,000+$  genes). Importantly, these issues are not resolved by subsampling the data to include  
442 a smaller number of genes (Figure S3), an approach commonly deployed by network analysis

443 packages (e.g., filtering for the 5,000-8,000 most variable genes in WGCNA or the approach of  
444 (Qiu *et al.*, 2021)).

445

446 *Our new high-dimensional covariance comparison*

447 From the above simulations, it is clear that existing approaches to compare large covariance  
448 structures fail even when using only a small subset of genes. To overcome these issues, we  
449 employed the random projection (Wu & Li, 2020) to develop a new two-sample covariance  
450 comparison method (Supplemental Information). This new test can control Type I error (Figure  
451 S4) and have satisfactory power (Figure S5), even with very small per group sample sizes and a  
452 large number of genes.

453 We conduct our newly developed random projection-based tests on residual from the  
454 generalized linear mixed model described above and in (Fischer *et al.*, 2021). We focused on the  
455 pairwise comparisons of biological interest (Figure 1c). Within each drainage, we compared (1)  
456 HP pred- vs LP pred-, an experiment comparing populations reared in an environment lacking  
457 predator cues to identify genetic differences between populations; (2) HP pred+ vs HP pred-, to  
458 identify environmentally induced changes mimicking the situation in which high-predation fish  
459 colonize low-predation environments, i.e., “ancestral plasticity”; (3) LP pred- vs LP pred+, to  
460 identify environmentally induced changes comparable to the situation in which low-predation fish  
461 are washed downstream and a measure of whether ancestral plasticity is maintained in the  
462 derived population; (4) HP pred+ vs LP pred+, to identify genetic differences when fish are raised  
463 with environmental cues of predation. We also compared the same experimental groups across  
464 drainages (e.g. HP pred+ in Aripo drainage vs HP pred+ in Quare drainage) to understand  
465 differences associated with parallel, independent evolutionary lineages. For both within and  
466 between drainage comparison, we considered the four comparisons jointly to control family-wise  
467 error rate.

468

## 469 *Correlation network comparisons*

470 We used correlation network analyses to examine whether DE versus non-DE genes differed in  
471 their relationships with other genes. A challenge with this analysis was the lack of consensus to  
472 define networks, in addition to the fact that derived networks will usually have very different sizes  
473 and unmatched nodes (i.e., non-DE genes far outnumber DE genes and a single gene is  
474 inherently only in one category). We constructed adjacency matrices using the correlations of  
475 residuals from the aforementioned generalized linear mixed model. From these adjacency  
476 matrices, we identified overall network structure and the frequency of specific subgraphs (i.e.,  
477 small motifs or subnetworks with two to seven nodes and one to six edges) for each experimental  
478 group within each drainage. We then compared network pairs of interest to examine their  
479 structural differences, using a network summary plot and two-sample network tests based on  
480 relative frequencies of different subgraphs adjusted for the sparsity of networks (additional details  
481 in Supplemental Information) (Maugis *et al.*, 2017; Shao *et al.*, 2022).

482

483

## 484 **ACKNOWLEDGEMENTS**

485 We thank the members of the Colorado State University Guppy Group for fish care and help with  
486 tissue collection and processing. This work was supported by the National Science Foundation  
487 DDIG-1311680 (to EKF), RCN IOS-1256839 (to EKF), IOS-1354755 (to KLH), [IOS 1922701 \(to](#)  
488 [WZ\)](#), US Department of Energy DE-SC0018344 (to WZ), and [NIH R01GM144961](#) (to WZ).

489

490

## 491 **DATA ACCESSIBILITY**

492 Raw sequencing reads are available through the NCBI SRA repository (PRJNA601479). R code  
493 for statistical analyses will be made available on GitHub  
494 (<https://github.com/EnigmaSong/GeneFlexibilityStudy>) upon publication.

495

496

## 497 **AUTHOR CONTRIBUTIONS**

498 EKF and KLH conceived of the study; EKF collected samples and performed molecular work,  
499 gene expression mapping, transcript abundance estimation, and preliminary differential  
500 expression analyses; YS and WZ devised and performed statistical analyses with input from EKF  
501 and KLH; EKF and YS performed data visualization; EKF wrote the manuscript with contributions  
502 from all authors.

503

504

## 505 **REFERENCES**

- 506 Agterberg, J., Tang, M. & Priebe, C.E. 2020. Nonparametric two-sample hypothesis testing for  
507 random graphs with negative and repeated eigenvalues. *arXiv*. 2012.09828.
- 508 Alyakin, A.A., Agterberg, J., Helm, H.S. & Priebe, C.E. 2024. Correcting a Nonparametric Two-  
509 sample Graph Hypothesis Test for Graphs with Different Numbers of Vertices with  
510 Applications to Connectomics. *Appl Netw Sci*. 9: 1.
- 511 Anglani, R., Creanza, T.M., Liuzzi, V.C., Piepoli, A., Panza, A., Andriulli, A., *et al.* 2014. Loss of  
512 connectivity in cancer co-expression networks. *PLoS One*. 9: 0087075.
- 513 Badyaev, A. V. 2018. Evolutionary transitions in controls reconcile adaptation with continuity of  
514 evolution. *Semin Cell Dev Biol*. 88: 36–45.
- 515 Ballouz, S., Verleyen, W. & Gillis, J. 2015. Guidance for RNA-seq co-expression network  
516 construction and analysis: Safety in numbers. *Bioinformatics*. 31: 2123–2130.
- 517 Barson, N.J., Cable, J. & Van Oosterhout, C. 2009. Population genetic analysis of microsatellite  
518 variation of guppies (*Poecilia reticulata*) in Trinidad and Tobago: Evidence for a dynamic  
519 source-sink metapopulation structure, founder events and population bottlenecks. *J Evol*  
520 *Biol*. 22: 485–497.



- 521 Bloch, N.I., Corral-López, A., Buechel, S.D., Kotrschal, A., Kolm, N. & Mank, J.E. 2021. Different  
522 mating contexts lead to extensive rewiring of female brain coexpression networks in the  
523 guppy. *Genes Brain Behav.* 20: 1–11.
- 524 Bolnick, D.I., Barrett, R.D.H., Oke, K.B., Rennison, D.J. & Stuart, Y.E. 2018. (Non)Parallel  
525 evolution. *Annu Rev Ecol Evol Syst.* 49: 303–330.
- 526 Cai, T., Liu, W. & Xia, Y. 2013. Two-sample covariance matrix testing and support recovery in  
527 high-dimensional and sparse settings. *J Am Stat Assoc.* 108: 265–277.
- 528 Cai, T.T. & Liu, W. 2016. Large-Scale Multiple Testing of Correlations. *J Am Stat Assoc.* 111:  
529 229–240.
- 530 Chang, J., Zhou, W., Zhou, W.X. & Wang, L. 2017. Comparing large covariance matrices under  
531 weak conditions on the dependence structure and its application to gene clustering.  
532 *Biometrics.* 73: 31–41.
- 533 Chateigner, A., Lesage-Descauses, M.C., Rogier, O., Jorge, V., Leplé, J.C., Brunaud, V., *et al.*  
534 2020. Gene expression predictions and networks in natural populations supports the  
535 omnigenic theory. *BMC Genomics.* 21: 1–16.
- 536 Chowdhury, H.A., Bhattacharyya, D.K. & Kalita, J.K. 2020. (Differential) Co-Expression Analysis  
537 of Gene Expression: A Survey of Best Practices. *IEEE/ACM Trans Comput Biol Bioinform.*  
538 17: 1154–1173.
- 539 Cordero, G.A., Liu, H., Wimalanathan, K., Weber, R., Quinteros, K. & Janzen, F.J. 2018. Gene  
540 network variation and alternative paths to convergent evolution in turtles. *Evol Dev.* 20:  
541 172–185.
- 542 Des Marais, D.L., Guerrero, R.F., Lasky, J.R. & Scarpino, S. V. 2017. Topological features of a  
543 gene co-expression network predict patterns of natural diversity in environmental response.  
544 *Proceedings of the Royal Society B: Biological Sciences.* 284: 20170914.
- 545 Endler, J.A. 1995. Multiple-trait coevolution and environmental gradients in guppies. *Trends*  
546 *Ecol Evol* .10: 22–29.

- 547 Fischer, E.K., Ghalambor, C.K. & Hoke, K.L. 2016a. Can a Network Approach Resolve How  
548 Adaptive vs Nonadaptive Plasticity Impacts Evolutionary Trajectories? *Integr Comp Biol*.  
549 56: 877–888.
- 550 Fischer, E.K., Ghalambor, C.K. & Hoke, K.L. 2016b. Plasticity and evolution in correlated suites  
551 of traits. *J Evol Biol*. 29: 991–1002.
- 552 Fischer, E.K., Harris, R.M., Hofmann, H.A. & Hoke, K.L. 2014. Predator exposure alters stress  
553 physiology in guppies across timescales. *Horm Behav*. 65: 165–172.
- 554 Fischer, E.K., Hauber, M.E. & Bell, A.M. 2021a. Back to the basics? Transcriptomics offers  
555 integrative insights into the role of space, time and the environment for gene expression  
556 and behaviour. *Biology Letters*. 2021;17:20210293.
- 557 Fischer, E.K., Soares, D., Archer, K.R., Ghalambor, C.K. & Hoke, K.L. 2013. Genetically and  
558 environmentally mediated divergence in lateral line morphology in the Trinidadian guppy  
559 (*Poecilia reticulata*). *Journal of Experimental Biology*. 216: 3132–3142.
- 560 Fischer, E.K., Song, Y., Hughes, K.A., Zhou, W. & Hoke, K.L. 2021. Non-parallel transcriptional  
561 divergence during parallel adaptation. *Mol Ecol*. 30: 1516–1530.
- 562 Fraser, B.A., Künstner, A., Reznick, D.N., Dreyer, C. & Weigel, D. 2015. Population genomics of  
563 natural and experimental populations of guppies (*Poecilia reticulata*). *Mol Ecol*. 24: 389–  
564 408.
- 565 Friedman, D.A., York, R.A., Hilliard, A.T. & Gordon, D.M. 2020. Gene expression variation in the  
566 brains of harvester ant foragers is associated with collective behavior. *Commun Biol*. 3.  
567 Springer US.
- 568 Gaiteri, C., Ding, Y., French, B., Tseng, G.C. & Sibille, E. 2014. Beyond modules and hubs: The  
569 potential of gene coexpression networks for investigating molecular mechanisms of  
570 complex brain disorders. *Genes Brain Behav*. 13: 13–24.
- 571 Gilliam, J.F., Fraser, D.F. & Alkins-Koo, M. 1993. Structure of a Tropical Stream Fish  
572 Community : A Role for Biotic Interactions. *Ecology*. 74: 1856–1870.

- 573 Gysi, D.M. & Nowick, K. 2020. Construction, comparison and evolution of networks in life  
574 sciences and other disciplines. *J R Soc Interface*. 17.
- 575 Hahn, M.W., Conant, G.C. & Wagner, A. 2004. Molecular Evolution in Large Genetic Networks:  
576 Does Connectivity Equal Constraint? *J Mol Evol*. 58: 203–211.
- 577 Hahn, M.W. & Kern, A.D. 2005. Comparative genomics of centrality and essentiality in three  
578 eukaryotic protein-interaction networks. *Mol Biol Evol*. 22: 803–806.
- 579 Hämälä, T., Gorton, A.J., Moeller, D.A. & Tiffin, P. 2020. Pleiotropy facilitates local adaptation to  
580 distant optima in common ragweed (*Ambrosia artemisiifolia*). *PLoS Gene.t* 16: 1–23.
- 581 Handelsman, C.A., Broder, E.D., Dalton, C.M., Ruell, E.W., Myrick, C.A., Reznick, D.N., *et al.*  
582 2013. Predator-induced phenotypic plasticity in metabolism and rate of growth: Rapid  
583 adaptation to a novel environment. *Integr Comp Biol*. 53: 975–988.
- 584 Handelsman, C.A., Ruell, E.W., Torres-Dowdall, J. & Ghalambor, C.K. 2014. Phenotypic  
585 Plasticity Changes Correlations of Traits Following Experimental Introductions of  
586 Trinidadian Guppies (*Poecilia reticulata*). *Integr Comp Biol*. 54: 794–804.
- 587 Hanson, D., Hu, J., Hendry, A.P. & Barrett, R.D.H. 2017. Heritable gene expression differences  
588 between lake and stream stickleback include both parallel and antiparallel components.  
589 *Heredity*. 119: 339–348.
- 590 Harnqvist SE, Grace CA, Jeffares DC. Variables Influencing Differences in Sequence  
591 Conservation in the Fission Yeast *Schizosaccharomyces pombe*. *J Mol Evol*. 2021;89.
- 592 Harrison, P.W., Wright, A.E. & Mank, J.E. 2012. The evolution of gene expression and the  
593 transcriptome-phenotype relationship. *Semin Cell Dev Biol*. 23: 222–229.
- 594 Hoke, K.L., Adkins-Regan, E., Bass, A.H., McCune, A.R. & Wolfner, M.F. 2019. Co-opting evo-  
595 devo concepts for new insights into mechanisms of behavioural diversity. *Journal of*  
596 *Experimental Biology*. 222: jeb190058.

- 597 Hu, G., Hovav, R., Grover, C.E., Faigenboim-Doron, A., Kadmon, N., Page, J.T., *et al.* 2016.  
598 Evolutionary conservation and divergence of gene coexpression networks in *Gossypium*  
599 (Cotton) seeds. *Genome Biol Evol.* 8: 3765–3783.
- 600 Huizinga, M., Ghalambor, C.K. & Reznick, D.N. 2009. The genetic and environmental basis of  
601 adaptive differences in shoaling behaviour among populations of Trinidadian guppies,  
602 *Poecilia reticulata*. *J Evol Biol.* 22: 1860–1866.
- 603 Jacobs A, Carruthers M, Yurchenko A, Gordeeva N V., Alekseyev SS, Hooker O, *et al.*  
604 Parallelism in eco-morphology and gene expression despite variable evolutionary and  
605 genomic backgrounds in a Holarctic fish. *PLoS Genetics.* 2020;16:e1008658.
- 606 Jeong, H., Mason, S.P., Barabási, A.L. & Oltvai, Z.N. 2001. Lethality and centrality in protein  
607 networks. *Nature.* 411: 41–42.
- 608 Jin, X., Chan, K., Barnett, I. & Ghosh, R.P. 2024. Two-Sample Hypothesis Testing for Large  
609 Random Graphs of Unequal Size. *arXiv.* 2024.11133.
- 610 Josephs, E.B., Wright, S.I., Stinchcombe, J.R. & Schoen, D.J. 2017. The relationship between  
611 selection, network connectivity, and regulatory variation within a population of *Capsella*  
612 *grandiflora*. *Genome Biol Evol.* 9: 1099–1109.
- 613 Jovelin, R. & Phillips, P.C. 2009. Evolutionary rates and centrality in the yeast gene regulatory  
614 network. *Genome Biol.* 10:R35.
- 615 Kim, P.M., Korbelt, J.O. & Gerstein, M.B. 2007. Positive selection at the protein network  
616 periphery: Evaluation in terms of structural constraints and cellular context. *Proc Natl Acad*  
617 *Sci.* 104: 20274–20279.
- 618 Koubkova-Yu, T.C.T., Chao, J.C. & Leu, J.Y. 2018. Heterologous Hsp90 promotes phenotypic  
619 diversity through network evolution. *PLoS Biol.* 16: e2006450.
- 620 Kuo, H.C., Yao, C. Te, Liao, B.Y., Weng, M.P., Dong, F., Hsu, Y.C., *et al.* 2023. Weak gene–  
621 gene interaction facilitates the evolution of gene expression plasticity. *BMC Biol.* 21: 1–20.

- 622 Langfelder, P. & Horvath, S. 2008. WGCNA: An R package for weighted correlation network  
623 analysis. *BMC Bioinformatics*. 9: 1–13.
- 624 Langfelder, P., Luo, R., Oldham, M.C. & Horvath, S. 2011. Is my network module preserved and  
625 reproducible? *PLoS Comput Biol*. 7: e1001057.
- 626 Laporte, M., Rogers, S.M., Dion-Côté, A.M., Normandeau, E., Gagnaire, P.A., Dalziel, A.C., et  
627 al. 2015. RAD-QTL mapping reveals both genome-level parallelism and different genetic  
628 architecture underlying the evolution of body shape in lake Whitefish (*Coregonus*  
629 *clupeaformis*) species pairs. *G3: Genes, Genomes, Genetics*. 5: 1481–1491.
- 630 Li, J. & Chen, S.X. 2012. Two sample tests for high-dimensional covariance matrices. *Ann Stat*.  
631 40: 908–940.
- 632 Li, S., Cai, T.T. & Li, H. 2022. Transfer Learning in Large-Scale Gaussian Graphical Models with  
633 False Discovery Rate Control. *J Am Stat Assoc*. 118: 2171–2183.
- 634 Love, M.I., Huber, W. & Anders, S. 2014. Moderated estimation of fold change and dispersion  
635 for RNA-seq data with DESeq2. *Genome Biol*. 15: 1–21.
- 636 Luisi, P., Alvarez-Ponce, D., Pybus, M., Fares, M.A., Bertranpetit, J. & Laayouni, H. 2015.  
637 Recent positive selection has acted on genes encoding proteins with more interactions  
638 within the whole human interactome. *Genome Biol Evol*. 7: 1141–1154.
- 639 Magurran, A.E. 2005. *Evolutionary ecology: the Trinidadian guppy*. Oxford University Press,  
640 Oxford.
- 641 Mähler, N., Wang, J., Terebieniec, B.K., Ingvarsson, P.K., Street, N.R. & Hvidsten, T.R. 2017.  
642 Gene co-expression network connectivity is an important determinant of selective  
643 constraint. *PLoS Genet*. 13: 1–33.
- 644 Makinen, H., Savilampi, T., Papakostas, S., Leder, E., Vøllestad, L.A. & Primmer, C.R. 2018.  
645 Modularity facilitates flexible tuning of plastic and evolutionary gene expression responses  
646 during early divergence. *Genome Biol Evol*. 10: 77–93.

- 647 Masalia, R.R., Bewick, A.J. & Burke, J.M. 2017. Connectivity in gene coexpression networks  
648 negatively correlates with rates of molecular evolution in flowering plants. *PLoS One* 12: 1–  
649 10.
- 650 Maugis, P.-A.G., Olhede, S.C. & Wolfe, P.J. 2017. Topology Reveals Universal Features for  
651 Network Comparison. *arXiv.1705.05677*.
- 652 McGuigan, K. & Sgrò, C.M. 2009. Evolutionary consequences of cryptic genetic variation.  
653 *Trends Ecol Evol.* 24: 305–311.
- 654 Oldham, M.C., Horvath, S. & Geschwind, D.H. 2006. Conservation and evolution of gene  
655 coexpression networks in human and chimpanzee brains. *Proc Natl Acad Sci.* 103: 17973–  
656 17978.
- 657 Paaby, A.B. & Rockman, M. V. 2014. Cryptic genetic variation: Evolution’s hidden substrate. *Nat*  
658 *Rev Genet.* 15: 247–258.
- 659 Pavličev, M. & Cheverud, J.M. 2015. Constraints Evolve: Context Dependency of Gene Effects  
660 Allows Evolution of Pleiotropy. *Annu Rev Ecol Evol. Syst* 46: 413–434.
- 661 Pavlicev, M. & Wagner, G.P. 2012. A model of developmental evolution: Selection, pleiotropy  
662 and compensation. *Trends Ecol Evol.* 27: 316–322.
- 663 Qi, M., Li, T. & Zhou, W. 2024. Multivariate Inference of Network Moments by Subsampling.  
664 *arXiv.* 2409.01599.
- 665 Qiu, T., Xu, W. & Zhu, L. 2021. Two-sample test in high dimensions through random selection.  
666 *Comput Stat Data Anal.* 160: 107218.
- 667 Rennison, D.J. & Peichel, C.L. 2022. Pleiotropy facilitates parallel adaptation in sticklebacks.  
668 *Mol Ecol.* 31: 1476–1486.
- 669 Reznick, D., Butler IV, M.J. & Rodd, H. 2001. Life-History Evolution in Guppies. VII. The  
670 Comparative Ecology of High- and Low-Predation Environments. *Am Nat.* 157: 126–140.
- 671 Reznick, D.A., Bryga, H. & Endler, J.A. 1990. Experimentally induced life-history evolution in a  
672 natural population. *Nature.* 346: 357–359.

- 673 Reznick, D.N. 1997. Life history evolution in guppies (*Poecilia reticulata*): Guppies as a model  
674 for studying the evolutionary biology of aging. *Exp Gerontol.* 32: 245–258.
- 675 Ruell, E.W., Handelsman, C.A., Hawkins, C.L., Sofaer, H.R., Ghalambor, C.K. & Angeloni, L.  
676 2013. Fear, food and sexual ornamentation: plasticity of colour development in Trinidadian  
677 guppies. *Proceedings of the Royal Society B: Biological Sciences.* 280: 20122019–  
678 20122019.
- 679 Sai Li, T.T.C. & Li, H. 2023. Transfer Learning in Large-Scale Gaussian Graphical Models with  
680 False Discovery Rate Control. *J Am Stat Assoc.* 118: 2171–2183.
- 681 Shao, M., Xia, D., Zhang, Y., Wu, Q. & Chen, S. 2022. Higher-Order Accurate Two-Sample  
682 Network Inference and Network Hashing. *arXiv.* 2208.07573.
- 683 Shojaie, A. 2021. Differential network analysis: A statistical perspective. *Wiley Interdisciplinary*  
684 *Reviews: Computational Statistics.* 2021;13:e1508.
- 685 Stanford, B.C.M., Clake, D.J., Morris, M.R.J. & Rogers, S.M. 2020. The power and limitations of  
686 gene expression pathway analyses toward predicting population response to  
687 environmental stressors. *Evol Appl.* 13: 1166–1182.
- 688 Storey, J.D. 2002. A direct approach to false discovery rates. *J R Stat Soc Series B Stat*  
689 *Methodol.* 64: 479–498.
- 690 Strimmer, K. 2008. fdrtool: A versatile R package for estimating local and tail area-based false  
691 discovery rates. *Bioinformatics.* 24: 1461–1462.
- 692 Tang, M., Athreya, A., Sussman, D.L., Lyzinski, V., Park, Y. & Priebe, C.E. 2017. A  
693 Semiparametric Two-Sample Hypothesis Testing Problem for Random Graphs. *Journal of*  
694 *Computational and Graphical Statistics.* 26: 344-354.
- 695 Tommasini, D. & Fogel, B.L. 2023. multiWGCNA: an R package for deep mining gene co-  
696 expression networks in multi-trait expression data. *BMC Bioinformatics.* 24: 1–15.

- 697 Torres Dowdall, J., Handelsman, C.A., Ruell, E.W., Auer, S.K., Reznick, D.N. & Ghalambor,  
698 C.K. 2012. Fine-scale local adaptation in life histories along a continuous environmental  
699 gradient in Trinidadian guppies. *Funct Ecol.* 26: 616–627.
- 700 Torres-Dowdall, J., Handelsman, C. a, Reznick, D.N. & Ghalambor, C.K. 2012. Local adaptation  
701 and the evolution of of phenotypic plasticity in Trinidadian guppies (*Poecilia reticulata*).  
702 *Evolution.* 66: 3432–3443.
- 703 Torres-Dowdall, J., Handelsman, C.A., Reznick, D.N. & Ghalambor, C.K. 2012. Local adaptation  
704 and the evolution of phenotypic plasticity in Trinidadian guppies (*Poecilia reticulata*).  
705 *Evolution.* 66: 3432–3443.
- 706 Wachi, S., Yoneda, K. & Wu, R. 2005. Interactome-transcriptome analysis reveals the high  
707 centrality of genes differentially expressed in lung cancer tissues. *Bioinformatics.* 21: 4205–  
708 4208.
- 709 Wang, D., Wang, J., Jiang, Y., Liang, Y. & Xu, D. 2017. BFDCA: A Comprehensive Tool of  
710 Using Bayes Factor for Differential Co-Expression Analysis. *J Mol Biol.* 429: 446–453.
- 711 Wang, Z., Liao, B.-Y. & Zhang, J. 2010. Genomic patterns of pleiotropy and the evolution of  
712 complexity. *Proceedings of the National Academy of Sciences.* 107: 18034–18039.
- 713 Warnefors, M. & Kaessmann, H. 2013. Evolution of the correlation between expression  
714 divergence and protein divergence in mammals. *Genome Biol Evol.* 5: 1324–1335.
- 715 West-Eberhard, M.J. 2003. *Developmental plasticity and evolution.* Oxford University Press,  
716 New York.
- 717 Willing, E.M., Bentzen, P., Van Oosterhout, C., Hoffmann, M., Cable, J., Breden, F., *et al.* 2010.  
718 Genome-wide single nucleotide polymorphisms reveal population history and adaptive  
719 divergence in wild guppies. *Mol Ecol.* 19: 968–984.
- 720 Wu, T.-L. & Li, P. 2020. Projected Tests for High-Dimensional Covariance Matrices. *J Stat Plan*  
721 *Inference.* 207: 73–85.



- 722 Yu, X., Li, D. & Xue, L. 2020a. Fisher's Combined Probability Test for High-Dimensional  
723 Covariance Matrices. *arXiv*. 2006.00426.
- 724 Yu, Y., Hu, H., Doust, A.N. & Kellogg, E.A. 2020b. Divergent gene expression networks underlie  
725 morphological diversity of abscission zones in grasses. *New Phytologist* 225: 1799–1815.
- 726 Zhao, S.D., Cai, T.T. & Li, H. 2014. Direct estimation of differential networks. *Biometrika*. 101:  
727 253–268.
- 728