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Circulating level of growth-differentiation factor 15 and the functional outcome after acute ischemic stroke: a systematic review and meta-analysis



Yulang Wu¹, Yude Wei¹, Jinrong He¹ and Fengkun Zhou^{1*}

Abstract

Growth-differentiation factor 15 (GDF-15) is a cytokine involved in cellular stress responses and inflammation. This meta-analysis evaluates the association between circulating GDF-15 levels and functional outcomes in patients with acute ischemic stroke (AIS). A comprehensive search of Medline, Web of Science, Embase, Wanfang, and CNKI was conducted up to July 15, 2024. Observational studies with longitudinal follow-up that measured GDF-15 levels within 24 h of stroke onset and reported functional outcomes, defined as a modified Rankin Scale (mRS) score of \geq 2, were included. Odds ratios (OR) with 95% confidence intervals (CI) were used to quantify associations. Heterogeneity was evaluated using I² statistics, and a random-effects model was used to pool the results by incorporating the influence of heterogeneity. Ten studies involving 4,231 patients were included. The pooled OR indicated that high circulating GDF-15 levels were associated with a significantly higher risk of poor functional outcomes at 3 months (OR: 2.60, 95% CI: 1.95 to 3.46, p < 0.001). Sensitivity analyses by excluding one study at a time did not significantly change the results. Subgroup analyses revealed stronger associations in studies with GDF-15 cutoff values < 1200 ng/L as compared to \geq 1200 ng/L, and in those defining poor outcomes as mRS \geq 3 as compared to those \geq 2. In conclusion, elevated circulating GDF-15 levels are associated with worse functional outcomes following AlS. These findings support the potential use of GDF-15 as a prognostic biomarker in stroke patients. Further research is warranted to confirm these results and explore clinical applications.

Keywords Ischemic stroke, Growth-differentiation factor 15, Functional outcome, Predictor, Meta-analysis

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Introduction

Acute ischemic stroke (AIS) is a major global health concern and a leading cause of disability and death [1, 2]. It occurs when a blockage in a cerebral artery impairs blood flow to a part of the brain, leading to neuronal injury and subsequent neurological deficits [3]. According to the Global Burden of Disease Study, ischemic stroke accounts for approximately 11% of all deaths and 5% of disability-adjusted life years (DALYs) worldwide [4]. Despite advancements in acute stroke management and treatment, a substantial proportion of stroke survivors experience poor functional outcomes [5]. The modified Rankin Scale (mRS) is commonly used to assess functional outcomes post-stroke, with scores ranging from 0 (no symptoms) to 6 (death) [6]. A score of ≥ 2 indicates some degree of disability, while a score of ≥ 3 reflects significant impairment [6].

Identifying new risk factors for poor functional outcomes after AIS is critical for improving patient prognosis and tailoring treatment strategies [7]. Biomarkers, in particular, hold promise for predicting stroke severity and outcomes [8]. Growth-differentiation factor 15 (GDF-15) is a cytokine involved in cellular stress responses, inflammation, and tissue repair [9]. Elevated GDF-15 levels have been associated with adverse outcomes in various cardiovascular diseases, suggesting its potential role in stroke prognosis [9–11]. GDF-15 functions by modulating inflammatory responses and apoptosis [11, 12], which are crucial in stroke pathology. High levels of GDF-15 might reflect extensive neuronal damage and heightened inflammatory activity [13], contributing to poor functional recovery.

Recent observational studies have explored the relationship between circulating GDF-15 levels and functional outcomes following AIS [14–23]. However, a quantitative summary for the association remains lacking. In addition, the specific impact of GDF-15 levels on stroke outcomes and the influence of study characteristics, such as the methods for measuring GDF-15, cutoff values used, and patient demographics, remain inadequately explored, underscoring the need for a comprehensive meta-analysis. By aggregating data from multiple studies, this meta-analysis aims to summarize the existing evidence, clarify the robustness of the association between GDF-15 levels and poor functional outcomes, and investigate how different study characteristics might influence the observed outcomes.

Methods

The authors followed the recommendations set forth in PRISMA 2020 [24, 25] and the Cochrane Handbook for Systematic Reviews and Meta-analyses [26] during the entire meta-analysis process, including the study's design, data acquisition, statistical evaluation, and interpretation

of the findings. The protocol of the meta-analysis has been registered at PROSPERO with the identification code CRD42024581753. A PRISMA 2000 Checklist has been included in the Supplementary Material 1.

Database search

To identify studies relevant to the aim of the meta-analvsis, we searched Medline, Web of Science, Embase, Wanfang and China National Knowledge Infrastructure (CNKI) utilizing comprehensive search terms involving (1) "growth differentiation factor 15" OR "macrophage inhibitory cytokine 1" OR "prostate differentiation factor" OR "GDF-15" OR "GDF 15" OR "MIC-1"; and (2) "stroke" OR "transient ischemic stroke" OR "TIA" OR "cerebral infarction" OR "cerebrovascular infarction". The detailed search strategy for each of the included database is detailed in Supplementary Material 2. The search was limited to clinical research conducted on human subjects. We included only full-length articles published in English or Chinese in peer-reviewed journals. Additionally, we manually reviewed the references of relevant original and review articles to identify any additional pertinent studies. The literature search covered publications from the inception of the databases up to July 15, 2024.

Inclusion and exclusion criteria

The inclusion criteria for eligible studies were as follows: (1) observational studies with longitudinal follow-up published as full-length articles, including prospective and retrospective cohort studies, nested case-control studies, and post-hoc analyses of clinical trials; (2) enrollment of patients with AIS, regardless of the primary treatment method; (3) measurement of circulating GDF-15 levels at enrollment (within 24 h of admission after stroke onset) as the exposure factor; (4) reporting of the risk of poor functional outcomes during follow-up, defined as a modified Rankin Scale (mRS) score of ≥ 2 [27]; and (5) comparison of the relative risk of poor functional outcomes after IS between patients with high and low baseline GDF-15 levels. We did not apply restriction for the follow-up duration of the studies. Grey literatures such as preprints, conference abstracts, and unpublished data were not included because these literatures were generally not peer-reviewed, and inclusion of these data may affect the reliability of the meta-analysis results.

The exclusion criteria were: (1) cross-sectional or casecontrol studies; (2) studies involving patients with hemorrhagic stroke; (3) studies that either did not measure or report baseline circulating GDF-15 or analyzed GDF-15 only as a continuous variable; (4) studies that did not report the incidence of poor functional outcomes based on the mRS score during follow-up; and (5) studies published as conference abstracts, unpublished data, reviews, or editorials. For studies with overlapping populations, the one with the largest sample size was chosen for inclusion in the meta-analysis.

Study quality assessment and data collection

The literature search, study identification, quality evaluation, and data extraction were conducted independently by two authors (YL W and YD W). In instances of disagreement, the corresponding author was consulted to achieve consensus. The quality of the included studies was assessed using the Newcastle-Ottawa Scale (NOS) [28], which examines three main domains: selection of the study population, control of confounding factors, and measurement and analysis of outcomes. NOS scores range from 1 to 9, with a score of 9 representing the highest quality. Data collected from each study for subsequent analysis included study-specific information (author, year, country, and design), patient demographics (sample size, age, and sex) at admission, methods for GDF-15 measurement, criteria for determining GDF-15 cutoff values, the actual cutoff values used, follow-up period, definition of poor functional outcomes, the number of patients experiencing poor outcomes during follow-up, and the variables accounted for when analyzing the association between circulating GDF-15 levels and poor functional outcomes.

Statistical analysis

The association between circulating GDF-15 levels and poor functional outcomes after AIS was evaluated using odds ratios (OR) with corresponding 95% confidence intervals (CI), comparing patients with high versus low GDF-15 levels at admission. ORs and their standard errors were calculated from 95% CIs or p-values, followed by logarithmic transformation to stabilize variance. Study heterogeneity was assessed with the Cochrane Q test and I^2 statistics [29], with an I^2 value greater than 50% indicating significant statistical heterogeneity. The results were synthesized using a random-effects model to account for the potential impact of heterogeneity [26]. Sensitivity analyses, in which one dataset was excluded at a time, were performed to evaluate the stability of the findings. Predefined subgroup analyses were also conducted to explore how specific study characteristics might influence the results, with median values of continuous variables serving as cutoffs for subgroup definitions. Publication bias was examined using funnel plots, assessing visual symmetry [30], and confirmed with an Egger's regression test [30]. In addition, We also performed univariate meta-regression to evaluate whether the outcome of the meta-analysis could be significantly modified by study characteristics such as publication year, sample size, mean age of the patients, proportion of men, cutoff value of GDF-15, cutoff mRS for defining a poor functional outcome, and the study quality score in NOS [26]. A p-value < 0.05 indicates significance. Statistical analyses were performed using RevMan (Version 5.1; Cochrane Collaboration, Oxford, UK) and Stata software (version 12.0; Stata Corporation, College Station, TX).

Results

Study inclusion

The study inclusion process is illustrated in Fig. 1. In summary, an initial search of five databases yielded 616 potentially relevant records, of which 185 were excluded due to duplication. The titles and abstracts of the remaining records were then screened, leading to the exclusion of 409 studies, primarily because they did not align with the objectives of the meta-analysis. The full texts of the remaining 22 records were reviewed by two independent authors, resulting in the exclusion of 12 studies for reasons detailed in Fig. 1. Ultimately, ten studies [14–23] were deemed appropriate for inclusion in the quantitative analyses.

Summary of study characteristics

Table 1 presents the overview of the characteristics of the included studies. Overall, four prospective cohort studies [14, 15, 18, 19], five retrospective cohort studies [16, 17, 21–23], and one post-hoc analysis of clinical study [20] were included in the meta-analysis. These studies were reported from 2011 to 2022 and performed in Germany, China, France, and Korea. Overall, 4231 patients with AIS were included, with mean ages varying from 58.4 to 74.6 years, and the proportion of men ranging from 39.0 to 64.1%. The circulating level of GDF-15 was measured all within 24 h after stroke onset, with enzymelinked immunosorbent assay in seven of the included studies [17–23]. For the other three studies [14–16], other methods such as immunoradiometric assay, electrochemiluminescence immunoassay, and radioimmunoassay were used, respectively. The cutoff values for defining a high circulating GDF-15 level were derived via receiver operating characteristic (ROC) curve analysis in three studies [14, 21, 23], via median of GDF-15 in five studies [15–17, 19, 22], via the third tertile [18] and the upper limits of healthy individual [20] in the another two studies. The cutoff values for defining a high GDF-15 varying from 493 to 2088 ng/L. The patients were all followed for 3 months among the included studies. The risk of poor functional outcome at 3 months after stroke were reported in all of the included studies, which were defined as the mRS ≥ 2 in four studies [14–17] and ≥ 3 in six studies [18-23]. Accordingly, 1267 (29.9%) of the included patients had poor functional outcome 3 months after AIS. A multivariate analysis was performed in five studies when the association between circulating GDF-15 and poor functional outcome after AIS was evaluated [14, 15, 19, 20, 22], whereas a univariate analysis was



Fig. 1 PRISMA flowchart of study identification

performed in another five studies [16–18, 21, 23]. The NOS of the included studies were six to nine stars, suggesting overall moderate to good study quality (Table 2).

Results of the meta-analysis and sensitivity analysis

The pooled results of ten studies using a random-effects model suggested that compared to AIS patients with a low circulating GDF-15 at admission, patients with a high circulating GDF-15 were associated with an increased risk of poor functional outcome at 3 months (OR: 2.60, 95% CI: 1.95 to 3.46, p<0.001; Fig. 2A) with moderate statistical heterogeneity (I²=65%). Further analysis

excluding one study at a time consistently demonstrated similar results (OR: 2.38 to 2.84, all p < 0.05).

Results of subgroup analyses and meta-regression analyses

Further subgroup analyses showed similar results in studies from Asia and Europe (OR: 2.94 versus 1.99, p for subgroup difference=0.18; Fig. 2B), in prospective and retrospective studies (OR: 2.23 versus 2.89, p for subgroup difference=0.39; Fig. 2C), in studies with the mean age of the patients < and \geq 65 years (OR: 3.16 versus 2.14, p for subgroup difference=0.22; Fig. 3A), in studies with

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|--|---|---|-------------------------------------|--|-----------------------------------|---|--|---|--|--|--|--|--|
| young | country | nesign | ple size | mean age (years) | (%) | GDF-15 measurement | memous of GDF-15 measurement | Metnods for determining GDF-15 cutoff | GDF-15 | rollow- up duration (months) | of poor functional outcome | wo. or pauents with poor functional outcome | variables aujus cardenary |
| Worth- mann [14] | Germany | PC | 57 | 74 | 47 | At admission | IRMA | ROC curve analysis | 1630 ng/l | m | mRS: 2–6 | 19 | Age and sex |
| Groschel [15] | Germany | DC | 264 | 70.3 | 55.3 | At admission | ECLIA | Median | 1191 ng/l | Ω | mRS: 2–6 | 113 | Age, sex, NIHSS at admis- sion, TOAST classification, SBP, DM, CAD, HDL-C, and thrombolysis |
| Geng [16] | China | RC | 102 | 71.3 | 58.8 | At admission | RIA | Median | 1580 ng/l | e | mRS: 2–6 | 29 | None |
| Lu [17] | China | RC | 120 | 67.8 | 60.8 | Within 24 h of stroke onset | ELISA | Median | 1594 ng/l | m | mRS: 2–6 | 65 | None |
| Yin 2019 | China | Post-hoc analysis | 3066 | 62.3 | 64.1 | Within 24 h of admission | ELISA | UL of healthy individual in previous studies | 1200 ng/l | m | mRS: 3–6 | 762 | Age, sex, current smoking, alcohol consumption, and NIHSS at admission |
| Breniere [18] | France | РС | 173 | 74.6 | 51 | At admission | ELISA | T3 versus T1-2 | 2088 ng/l | m | mRS: 3–6 | 92 | None |
| Dong [19] | China | РС | 83 | 62.4 | 63.9 | At admission | ELISA | Median | 864 ng/l | m | mRS: 3–6 | 42 | Age, sex, comorbidities, and NIHSS at admission |
| Jeong [22] | Korea | RC | 62 | 66.2 | 39 | Within 24 h of admission | ELISA | Median | 493 ng/l | ε | mRS: 3–6 | 20 | Age, sex, comorbidities, eGFR, and timing of recanalization |
| Cai [21] | China | RC | 148 | 58.4 | 62.2 | Within 24 h of admission | ELISA | ROC curve analysis | 826.5 ng/l | m | mRS: 3–6 | 61 | None |
| Li [23] | China | RC | 156 | 58.4 | 55 | Within 24 h of stroke onset | ELISA | ROC curve analysis | 802.1 ng/l | e | mRS: 3–6 | 64 | None |
| GDF-15, grc ROC, Receiv CAD, coron | wth different /er operating - ary artery dise | iation factor 1 characteristic; ase; DM, diab | 5; PC, pro UL, uppe etes mell | ospective er limit; T, t litus; HDL- | cohort; tertile; m C, high- | RC, retrospective co nRS, modified Rankir density lipoprotein | hort; IRMA, immun Score; NIHSS, Natic cholesterol; eGFR, e | oradiometric assay; onal institutes of hea stimated glomerula | ECLIA, electr alth stroke sca ir filtrating ra | ochemilumir ale; TOAST, Tr te; RIA, radioi | iescence immul ial of ORG 10,17. immunoassay | noassay; ELISA, enz) 2 in Acute Stroke Th | /me-linked immunosorbent assay; erapy; SBP, systolic blood pressure; |

 Table 1
 Characteristics of the included studies

Table 2 Study quality evaluation via the Newcastle-Ottawa Scale

| Study | Representa- tiveness of the exposed cohort | Selection of the non- exposed cohort | Ascertain- ment of exposure | Outcome not pres- ent at baseline | Control for age and sex | Control for other con- founding factors | Assess- ment of outcome | Enough long follow- up duration | Adequacy of follow- up of cohort | Total |
|--------------------------|---|---|-----------------------------------|--|-------------------------------|--|-------------------------------|---------------------------------------|---|-------|
| Worthmann [14] | 1 | 1 | 1 | 1 | 1 | 0 | 1 | 1 | 1 | 8 |
| Groschel [15] | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 9 |
| Geng [<mark>16</mark>] | 0 | 1 | 1 | 1 | 0 | 0 | 1 | 1 | 1 | 6 |
| Lu [17] | 0 | 1 | 1 | 1 | 0 | 0 | 1 | 1 | 1 | 6 |
| Yin [<mark>20</mark>] | 0 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 8 |
| Breniere [18] | 1 | 1 | 1 | 1 | 0 | 0 | 1 | 1 | 1 | 7 |
| Dong [19] | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 9 |
| Jeong [22] | 0 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 8 |
| Cai [<mark>2</mark> 1] | 0 | 1 | 1 | 1 | 0 | 0 | 1 | 1 | 1 | 6 |
| Li [23] | 0 | 1 | 1 | 1 | 0 | 0 | 1 | 1 | 1 | 6 |

| | | | | Odds Ratio | | Odd | s Ratio | |
|-----------------------------------|--------------------------------------|------------|---------------------|--------------------|------|----------|---|----|
| Study or Subgroup | log[Odds Ratio] | SE | Weight | IV, Random, 95% C | | IV, Rand | om, 95% Cl | |
| Worthmann 2011 | 0.94 | 0.33 | 9.3% | 2.56 [1.34, 4.89] | | | | |
| Groschel 2012 | 0.41 | 0.16 | 14.5% | 1.51 [1.10, 2.07] | | | | |
| Geng 2015 | 0.53 | 0.35 | 8.5% | 1.70 [0.84, 3.44] | | | | |
| Lu 2015 | 1.02 | 0.37 | 8.1% | 2.80 [1.33, 5.89] | | | | |
| Yin 2019 | 0.59 | 0.12 | 15.5% | 1.81 [1.41, 2.32] | | | - | |
| Breniere 2019 | 0.98 | 0.34 | 9.0% | 2.69 [1.38, 5.24] | | | | |
| Dong 2019 | 1.21 | 0.37 | 8.2% | 3.38 [1.62, 7.04] | | | | |
| Jeong 2020 | 1.59 | 0.60 | 4.3% | 4.93 [1.49, 16.30] | | | · · · · | _ |
| Cai 2020 | 1.43 | 0.26 | 11.1% | 4.22 [2.51, 7.10] | | | | |
| LI 2022 | 1.47 | 0.25 | 11.4% | 4.39 [2.65, 7.27] | | | | |
| Total (95% CI) | | | 100.0% | 2.60 [1.95, 3.46] | | | • | |
| Heterogeneity: Tau ² = | 0.12: Chi ² = 25.65. df = | 9 (P = 0.0 | (002) : $ ^2 = 0$ | 35% | + | + | <u> </u> | |
| Test for overall effect: | Z = 6.53 (P < 0.00001) | - (| ,, . | | 0.05 | 0.2 | 1 5 | 20 |

1

| Study or Subgroup | log[Odds Patio] | SE | Weight | Odds Ratio | Odds Ratio |
|-----------------------------------|---------------------------------------|-------------|--------------------------------------|---------------------|----------------------|
| 1.2.1 Asian | log[ouds Ratio] | 02 | Weight | IV. Random, 3576 Of | TV. Randolli. 35% Of |
| Geng 2015 | 0.53 | 0.35 | 8.5% | 1.70 [0.84, 3.44] | + |
| Lu 2015 | 1.02 | 0.37 | 8.1% | 2.80 [1.33, 5.89] | |
| Yin 2019 | 0.59 | 0.12 | 15.5% | 1.81 [1.41, 2.32] | - |
| Dong 2019 | 1.21 | 0.37 | 8.2% | 3.38 [1.62, 7.04] | |
| Jeong 2020 | 1.59 | 0.60 | 4.3% | 4.93 [1.49, 16.30] | |
| Cai 2020 | 1.43 | 0.26 | 11.1% | 4.22 [2.51, 7.10] | |
| LI 2022 | 1.47 | 0.25 | 11.4% | 4.39 [2.65, 7.27] | |
| Subtotal (95% CI) | | | 67.2% | 2.94 [2.01, 4.29] | • |
| 1.2.2 European | | | | | |
| Worthmann 2011 | 0.94 | 0.33 | 9.3% | 2.56 [1.34, 4.89] | |
| Groschel 2012 | 0.41 | 0.16 | 14.5% | 1.51 [1.10, 2.07] | |
| Breniere 2019 | 0.98 | 0.34 | 9.0% | 2.69 [1.38, 5.24] | |
| Subtotal (95% CI) | 0.00 | 0101 | 32.8% | 1.99 [1.32, 3.01] | • |
| Heterogeneity: Tau ² = | 0.06; Chi ² = 3.74, df = 2 | 2 (P = 0.15 | 5); ² = 46 ⁴ | % | |
| Test for overall effect: | Z = 3.28 (P = 0.001) | | ,, | | |
| Total (95% CI) | | | 100.0% | 2.60 [1.95, 3.46] | • |
| Heterogeneity: Tau ² = | 0.12; Chi ² = 25.65, df = | 9 (P = 0.0 | 002); l ² = 6 | 35% | |
| Tool for a second set | 7 - 0 F0 (D + 0 00004) | | | | 0.05 0.2 1 5 |

Test for overall effect: Z = 6.53 (P < 0.00001) Test for subgroup differences: Chi² = 1.83 df = 1 (P = 0.18) l² = 45.3%

| Test for subdroub differences: $Chi^{\mu} = 1.83$. df = 1 (P = 0.18). $i^{\mu} = 45$. | 5% |
|---|----|
| | |
| Odds Ratio | |

| | | | | Odda Datia | Odda Datia |
|-----------------------------------|---------------------------------------|-------------|--------------------------|--------------------|--------------------|
| o | La rolla Daria | 05 | 147-1-1-4 | Udds Ratio | Udds Ratio |
| Study or Subgroup | logiOdds Ratioj | SE | weight | IV, Random, 95% CI | IV. Random, 95% CI |
| 1.3.1 Prospective | | | | | |
| Worthmann 2011 | 0.94 | 0.33 | 9.3% | 2.56 [1.34, 4.89] | |
| Groschel 2012 | 0.41 | 0.16 | 14.5% | 1.51 [1.10, 2.07] | - |
| Breniere 2019 | 0.98 | 0.34 | 9.0% | 2.69 [1.38, 5.24] | |
| Dong 2019 | 1.21 | 0.37 | 8.2% | 3.38 [1.62, 7.04] | |
| Subtotal (95% CI) | | | 41.0% | 2.23 [1.48, 3.36] | - |
| Heterogeneity: Tau ² = | 0.09; Chi ² = 6.27, df = 3 | 3 (P = 0.10 |); l ² = 525 | % | |
| Test for overall effect: | Z = 3.86 (P = 0.0001) | | | | |
| 1.3.2 Retrospective of | or post-hoc | | | | |
| Geng 2015 | 0.53 | 0.35 | 8.5% | 1.70 [0.84, 3.44] | + |
| Lu 2015 | 1.02 | 0.37 | 8.1% | 2.80 [1.33, 5.89] | |
| Yin 2019 | 0.59 | 0.12 | 15.5% | 1.81 [1.41, 2.32] | - |
| Jeong 2020 | 1.59 | 0.60 | 4.3% | 4.93 [1.49, 16.30] | |
| Cai 2020 | 1.43 | 0.26 | 11.1% | 4.22 [2.51, 7,10] | |
| LI 2022 | 1.47 | 0.25 | 11.4% | 4.39 [2.65, 7.27] | |
| Subtotal (95% CI) | | | 59.0% | 2.89 [1.89, 4.42] | • |
| Heterogeneity: Tau ² = | 0.18: Chi ² = 17.41. df = | 5(P = 0.0) | $(04): ^2 = 7$ | /1% | |
| Test for overall effect: | Z = 4.87 (P < 0.00001) | | | | |
| | , | | | | |
| Total (95% CI) | | | 100.0% | 2.60 [1.95, 3.46] | • |
| Heterogeneity: Tau ² = | 0.12; Chi ² = 25.65, df = | 9 (P = 0.0 | 002); l² = 6 | 35% [†] | 0.05 0.2 1 5 |
| Test for overall effect: | Z = 6.53 (P < 0.00001) | | | , | 1.00 0.2 1 5 |
| Test for subaroup diffe | erences: Chi ² = 0.73. df | = 1 (P = 0 | .39), l ² = (| 0% | |

Fig. 2 Forest plots for the meta-analysis of the association between circulating GDF-15 level at admission and functional outcome after AIS; A, overall meta-analysis; B, subgroup analysis according to study country; and C, subgroup analysis according to study design

the proportion of men<and \geq 60% (OR: 2.52 versus 2.79, p for subgroup difference=0.76; Fig. 3B), and in studies with the cutoff of GDF-15 defined by ROC analysis and the medians (OR: 3.80 versus 2.25, p for subgroup difference=0.07; Fig. 4A). Interestingly, the subgroup analysis suggested that the association between a high circulating GDF-15 and the risk of poor functional outcome after AIS was stronger for studies with cutoff of GDF-15<1200 ng/L as compared to those \geq 1200 ng/L (OR: 3.52 versus 1.90, p for subgroup difference=0.003; Fig. 4B), which

substantially explained the source of heterogeneity. In addition, a stronger association was observed in studies with poor functional outcome define as mRS \geq 3 compared to those as mRS \geq 2 (OR: 3.17 versus 1.85, p for subgroup difference=0.04; Fig. 5A). Subsequently, consistent results were obtained for studies with univariate and multivariate analysis (OR: 3.03 versus 2.34, p for subgroup difference=0.33; Fig. 5B). Finally, the univariate meta-regression analyses did not show that study characteristics such as publication year, sample size, mean age

| | | | | Odds Ratio | 0 | das Ratio | |
|--|---|---|---|---|---------------|------------------------------------|--|
| Study or Subgroup | log[Odds Ratio] | SE | Weight | IV, Random, 95% C | I IV, Ra | andom, 95% Cl | |
| 1.4.1 < 65 years | | | | | | | |
| Yin 2019 | 0.59 | 0.12 | 15.5% | 1.81 [1.41, 2.32] | | | |
| Dong 2019 | 1.21 | 0.37 | 8.2% | 3.38 [1.62, 7.04] | | | |
| Cai 2020 | 1.43 | 0.26 | 11.1% | 4.22 [2.51, 7.10] | | | |
| LI 2022 | 1.47 | 0.25 | 11.4% | 4.39 [2.65, 7.27] | | | |
| Subtotal (95% CI) | | | 46.2% | 3.16 [1.85, 5.41] | | | |
| Heterogeneity: Tau ² = | 0.23; Chi ² = 15.72, df = | 3 (P = 0.0 | 01); l² = 8 | 31% | | | |
| Test for overall effect: | Z = 4.19 (P < 0.0001) | | | | | | |
| 1.4.2 ≥ 65 years | | | | | | | |
| Worthmann 2011 | 0.94 | 0.33 | 9.3% | 2.56 [1.34, 4.89] | | | |
| Groschel 2012 | 0.41 | 0.16 | 14.5% | 1.51 [1.10, 2.07] | | | |
| Geng 2015 | 0.53 | 0.35 | 8.5% | 1.70 [0.84, 3.44] | | + | |
| Lu 2015 | 1.02 | 0.37 | 8.1% | 2.80 [1.33, 5.89] | | | |
| Breniere 2019 | 0.98 | 0.34 | 9.0% | 2.69 [1.38, 5.24] | | — - | |
| Jeong 2020 | 1.59 | 0.60 | 4.3% | 4.93 [1.49, 16.30] | | | |
| Subtotal (95% CI) | | | 53.8% | 2.14 [1.56, 2.93] | | • | |
| Heterogeneity: Tau ² = | 0.05; Chi ² = 7.49. df = 5 | 5 (P = 0.19 |); ² = 339 | 6 | | | |
| Test for overall effect: | Z = 4.73 (P < 0.00001) | | ,, | | | | |
| | , | | | | | | |
| Total (95% CI) | | | 100.0% | 2.60 [1.95, 3.46] | | • | |
| Heterogeneity: Tau ² = | 0.12; Chi ² = 25.65, df = | 9 (P = 0.0 | 02); I ² = 6 | 5% | | | |
| | | | | | 0.00 0.2 | 1 5 | |
| Test for overall effect: | Z = 6.53 (P < 0.00001) | | | | | | |
| Test for overall effect: Test for subaroup diffe | Z = 6.53 (P < 0.00001) erences: Chi ² = 1.51. df | = 1 (P = 0. | 22). I² = 3 | 3.8% | | | |
| Test for overall effect: Test for subaroup diffe | Z = 6.53 (P < 0.00001) erences: Chi ² = 1.51. df | = 1 (P = 0. | 22). I² = 3 | 3.8% Odds Ratio | 0 | dds Ratio | |
| Test for overall effect: Test for subaroup diffe Study or Subgroup | Z = 6.53 (P < 0.00001) erences: Chi ² = 1.51. df <u>log[Odds Ratio]</u> | = 1 (P = 0. SE | 22). I² = 3 Weight | 3.8% Odds Ratio <u>IV, Random, 95% C</u> | 0 I IV, Ra | dds Ratio andom, 95% Cl | |
| Test for overall effect: Test for subaroup diffe Study or Subgroup 1.5.1 Men < 60% | Z = 6.53 (P < 0.00001) erences: Chi ² = 1.51. df log[Odds Ratio] | = 1 (P = 0. SE | 22). I ² = 3 Weight | 3.8% Odds Ratio IV, Random, 95% C | O I IV, Ra | dds Ratio andom, 95% Cl | |
| Test for overall effect: Test for subaroup diffe Study or Subgroup 1.5.1 Men < 60% Worthmann 2011 | Z = 6.53 (P < 0.00001) erences: Chi ² = 1.51. df log[Odds Ratio] 0.94 | = 1 (P = 0. <u>SE</u> 0.33 | 22). I ² = 3 <u>Weight</u> 9.3% | 3.8% Odds Ratio <u>IV, Random, 95% C</u> 2.56 [1.34, 4.89] | O I IV, Ra | dds Ratio andom, 95% Cl | |
| Test for overall effect: Test for subaroup diffe Study or Subgroup 1.5.1 Men < 60% Worthmann 2011 Groschel 2012 | Z = 6.53 (P < 0.00001) erences: Chi ² = 1.51. df log[Odds Ratio] 0.94 0.41 | = 1 (P = 0. <u>SE</u> 0.33 0.16 | 22). I ² = 3 <u>Weight</u> 9.3% 14.5% | 3.8% Odds Ratio <u>IV, Random, 95% C</u> 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] | O I IV, Ra | dds Ratio andom, 95% Cl | |
| Test for overall effect: Test for subaroup diffe Study or Subgroup 1.5.1 Men < 60% Worthmann 2011 Groschel 2012 Geng 2015 | Z = 6.53 (P < 0.00001) erences: Chi ² = 1.51. df log[Odds Ratio] 0.94 0.41 0.53 | = 1 (P = 0. <u>SE</u> 0.33 0.16 0.35 | 22). I ² = 3 <u>Weight</u> 9.3% 14.5% 8.5% | 3.8% Odds Ratio <u>IV, Random, 95% C</u> 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84] 3.44] | O I IV, Ra | dds Ratio andom, 95% Cl | |
| Test for overall effect: Test for subaroup diffe Study or Subgroup 1.5.1 Men < 60% Worthmann 2011 Groschel 2012 Geng 2015 Breniere 2019 | Z = 6.53 (P < 0.00001) erences: Chi ² = 1.51. df log[Odds Ratio] 0.94 0.41 0.53 0.98 | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 | 22). l ² = 3 <u>Weight</u> 9.3% 14.5% 8.5% 9.0% | 3.8% Odds Ratio IV, Random, 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] | O I IV, Ra | dds Ratio andom, 95% Cl | |
| Test for overall effect: Test for subaroup diffe Study or Subgroup 1.5.1 Men < 60% Worthmann 2011 Groschel 2012 Geng 2015 Breniere 2019 Jeong 2020 | Z = 6.53 (P < 0.00001) erences: Chi ² = 1.51. df log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 | 22). I ² = 3 <u>Weight</u> 9.3% 14.5% 8.5% 9.0% 4.3% | 3.8% Odds Ratio IV, Random, 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] | O I IV, Ra | dds Ratio andom, 95% Cl | |
| Test for overall effect: Test for subaroup diffe Study or Subgroup 1.5.1 Men < 60% Worthmann 2011 Groschel 2012 Geng 2015 Breniere 2019 Jeong 2020 LL 2022 | Z = 6.53 (P < 0.00001) erences: Chi ² = 1.51. df log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 | 22). I ² = 3 <u>Weight</u> 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% | 33.8% Odds Ratio IV. Random. 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65 7 27] | O I IV. Ra | Idds Ratio Indom, 95% Cl | |
| Test for overall effect: Test for subaroup diffe Study or Subgroup 1.5.1 Men < 60% Worthmann 2011 Groschel 2012 Geng 2015 Breniere 2019 Jeong 2020 LI 2022 Subtotal (95% CI) | Z = 6.53 (P < 0.00001) erences: Chi ² = 1.51. df log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 | 22). I ² = 3 Weight 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% | 3.8% Odds Ratio IV. Random. 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] | O I IV. Ra | dds Ratio andom, 95% Cl | |
| Test for overall effect: Test for subaroup diffe Study or Subgroup 1.5.1 Men < 60% Worthmann 2011 Groschel 2012 Geng 2015 Breniere 2019 Jeong 2020 LI 2022 Subtotal (95% CI) Heterogeneity: Tau ² = | Z = 6.53 (P < 0.00001) erences: Chi ² = 1.51. df log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 0 18: Chi ² = 15 56. df = | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 | 22). l ² = 3 Weight 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 908): l ² = 6 | 33.8% Odds Ratio IV. Random. 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% | O I IV. Ra | dds Ratio andom, 95% CI | |
| Test for overall effect: Test for subaroup diffe Study or Subgroup 1.5.1 Men < 60% Worthmann 2011 Groschel 2012 Geng 2015 Breniere 2019 Jeong 2020 LI 2022 Subtotal (95% CI) Heterogeneity: Tau ² = Test for overall effect: | Z = 6.53 (P < 0.00001) prences: Chi ² = 1.51. df log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 0.18; Chi ² = 15.56, df = Z = 4.21 (P < 0.0001) | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 | 22). l ² = 3 Weight 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 08); l ² = 6 | 33.8% Odds Ratio IV, Random, 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% | O I IV. Ra | dds Ratio andom, 95% Cl | |
| Test for overall effect: Test for subaroup diffe Study or Subaroup diffe 1.5.1 Men < 60% Worthmann 2011 Groschel 2012 Geng 2015 Breniere 2019 Jeong 2020 LI 2022 Subtotal (95% CI) Heterogeneity: Tau ² = Test for overall effect: | Z = $6.53 (P < 0.00001)$ prences: Chi ² = $1.51. df$ 0.94 0.94 0.41 0.53 0.98 1.59 1.47 0.18; Chi ² = $15.56, df = Z = 4.21 (P < 0.0001)$ | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 | 22). l ² = 3 Weight 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 108); l ² = 6 | 33.8% Odds Ratio IV. Random. 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% | O I IV, Ra | odds Ratio | |
| Test for overall effect: Test for subaroup different different for subaroup different different for subaroup different differ | Z = $6.53 (P < 0.00001)$ prences: Chi ² = $1.51. df$ <u>log[Odds Ratio]</u> 0.94 0.41 0.53 0.98 1.59 1.47 0.18; Chi ² = $15.56, df = Z = 4.21 (P < 0.0001)$ | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 | 22). l ² = 3 Weight 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 108); l ² = 6 | 33.8% Odds Ratio IV. Random. 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% | O I IV, Ra | odds Ratio | |
| Test for overall effect: Test for subgroup different different different different for subgroup different for subgroup different for subgroup different for subgroup different diff | Z = $6.53 (P < 0.00001)$ prences: Chi ² = $1.51. df$ log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 0.18; Chi ² = $15.56, df = Z = 4.21 (P < 0.0001)$ 1.02 | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 0.37 | 22). l ² = 3 Weight 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 108); l ² = 6 8.1% | 33.8% Odds Ratio IV. Random. 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% 2.80 [1.33, 5.89] | O I IV, Ra | odds Ratio | |
| Test for overall effect: Test for subaroup different different for subaroup different different for subaroup different di | Z = $6.53 (P < 0.00001)$ prences: Chi ² = $1.51. df$ log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 0.18; Chi ² = $15.56, df = Z = 4.21 (P < 0.0001)$ 1.02 0.59 | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 0.37 0.12 | 22). l ² = 3 <u>Weight</u> 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 108); l ² = 6 8.1% 15.5% | 33.8% Odds Ratio IV. Random. 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] | O I IV, Ra | odds Ratio | |
| Test for overall effect: Test for subgroup different different different for subgroup different different for subgroup different for subgroup different for subgroup different di d | Z = 6.53 (P < 0.00001) prences: Chi ² = 1.51 . df log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 0.18; Chi ² = 15.56 , df = Z = 4.21 (P < 0.0001) 1.02 0.59 1.21 | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 0.37 0.12 0.37 | 22). l ² = 3 <u>Weight</u> 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 108); l ² = 6 8.1% 15.5% 8.2% | 33.8% Odds Ratio IV. Random, 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] 3.38 [1.62, 7.04] | O I IV, Ra | odds Ratio andom, 95% CI | |
| Test for overall effect: Test for subgroup different for subgroup d | Z = $6.53 (P < 0.00001)$ prences: Chi ² = $1.51. df$ log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 0.18; Chi ² = $15.56, df = Z = 4.21 (P < 0.0001)$ 1.02 0.59 1.21 1.43 | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 0.37 0.12 0.37 0.26 | 22). ² = 3 <u>Weight</u> 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 108); ² = 6 8.1% 15.5% 8.2% 11.1% | 3.8% Odds Ratio IV, Random, 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] 3.38 [1.62, 7.04] 4.22 [2.51, 7.10] | O I IV, Ra | odds Ratio | |
| Test for overall effect: Test for subgroup different different for subgroup different for | Z = 6.53 (P < 0.00001) prences: Chi ² = 1.51. df log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 0.18; Chi ² = 15.56, df = Z = 4.21 (P < 0.0001) 1.02 0.59 1.21 1.43 | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 0.37 0.12 0.37 0.26 | 22). I ² = 3 Weight 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 008); I ² = 6 8.1% 15.5% 8.2% 11.1% 42.9% | 33.8% Odds Ratio IV, Random, 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] 3.38 [1.62, 7.04] 4.22 [2.51, 7.10] 2.79 [1.73, 4.50] | O I IV, Ra | odds Ratio | |
| Test for overall effect: Test for subaroup different for subaroup d | Z = $6.53 (P < 0.00001)$ prences: Chi ² = $1.51. df$ log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 0.18; Chi ² = $15.56, df =$ Z = $4.21 (P < 0.0001)$ 1.02 0.59 1.21 1.43 0.16; Chi ² = $10.04, df =$ | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 0.37 0.12 0.37 0.26 3 (P = 0.0 | 22). ² = 3 Weight 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 108); ² = 6 8.1% 15.5% 8.2% 11.1% 42.9% 12); ² = 70 | 33.8% Odds Ratio IV. Random. 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] 3.38 [1.62, 7.04] 4.22 [2.51, 7.10] 2.79 [1.73, 4.50] | O I IV, Ra | odds Ratio | |
| Test for overall effect: Test for subgroup different different different different for subgroup different different different for subgroup different differ | Z = 6.53 (P < 0.00001) prences: Chi ² = 1.51. df log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 0.18; Chi ² = 15.56, df = Z = 4.21 (P < 0.0001) 1.02 0.59 1.21 1.43 0.16; Chi ² = 10.04, df = Z = 4.19 (P < 0.0001) | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 0.37 0.12 0.37 0.26 3 (P = 0.0 | 22). l ² = 3 Weight 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 108); l ² = 6 8.1% 15.5% 8.2% 11.1% 42.9% 12); l ² = 70 | 33.8% Odds Ratio IV. Random. 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] 3.38 [1.62, 7.04] 4.22 [2.51, 7.10] 2.79 [1.73, 4.50] 3% | O I IV. Ra | odds Ratio | |
| Test for overall effect: Test for subgroup 1.5.1 Men < 60% Worthmann 2011 Groschel 2012 Geng 2015 Breniere 2019 Jeong 2020 Ll 2022 Subtotal (95% Cl) Heterogeneity: Tau ² = Test for overall effect: 1.5.2 Men $\ge 60\%$ Lu 2015 Yin 2019 Dong 2019 Cai 2020 Subtotal (95% Cl) Heterogeneity: Tau ² = Test for overall effect: Total (95% Cl) | Z = 6.53 (P < 0.00001) prences: Chi ² = 1.51. df log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 0.18; Chi ² = 15.56, df = Z = 4.21 (P < 0.0001) 1.02 0.59 1.21 1.43 0.16; Chi ² = 10.04, df = Z = 4.19 (P < 0.0001) | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 0.37 0.12 0.37 0.26 3 (P = 0.0 | 22). l ² = 3 Weight 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 108); l ² = 6 8.1% 15.5% 8.2% 11.1% 42.9% 12); l ² = 70 100.0% | 33.8% Odds Ratio IV. Random. 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] 3.38 [1.62, 7.04] 4.22 [2.51, 7.10] 2.79 [1.73, 4.50] 0% 2.60 [1.95, 3.46] | O I IV, Ra | odds Ratio andom, 95% Cl | |
| Test for overall effect: Test for subgroup 1.5.1 Men < 60% Worthmann 2011 Groschel 2012 Geng 2015 Breniere 2019 Jeong 2020 Ll 2022 Subtotal (95% Cl) Heterogeneity: Tau ² = Test for overall effect: 1.5.2 Men \ge 60% Lu 2015 Yin 2019 Dong 2019 Cai 2020 Subtotal (95% Cl) Heterogeneity: Tau ² = Test for overall effect: Total (95% Cl) Heterogeneity: Tau ² = | Z = 6.53 (P < 0.00001) prences: Chi ² = 1.51. df log[Odds Ratio] 0.94 0.41 0.53 0.98 1.59 1.47 0.18; Chi ² = 15.56, df = Z = 4.21 (P < 0.0001) 1.02 0.59 1.21 1.43 0.16; Chi ² = 10.04, df = Z = 4.19 (P < 0.0001) 0.12; Chi ² = 25.65, df = | = 1 (P = 0. SE 0.33 0.16 0.35 0.34 0.60 0.25 5 (P = 0.0 0.37 0.12 0.37 0.26 3 (P = 0.0 9 (P = 0.0 | 22). $l^2 = 3$ Weight 9.3% 14.5% 8.5% 9.0% 4.3% 11.4% 57.1% 108); $l^2 = 6$ 8.1% 15.5% 8.2% 11.1% 42.9% 12); $l^2 = 70$ 100.0% 102); $l^2 = 6$ | 33.8% Odds Ratio IV. Random. 95% C 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 1.70 [0.84, 3.44] 2.69 [1.38, 5.24] 4.93 [1.49, 16.30] 4.39 [2.65, 7.27] 2.52 [1.64, 3.87] 38% 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] 3.38 [1.62, 7.04] 4.22 [2.51, 7.10] 2.79 [1.73, 4.50] 0% 2.60 [1.95, 3.46] | | odds Ratio | |

Fig. 3 Forest plots for the subgroup analysis of the association between circulating GDF-15 level at admission and functional outcome after AIS; **A**, subgroup analysis according to the mean age of the patients; and **B**, subgroup analysis according to the proportion of men

of the patients, proportion of men, cutoff value of GDF-15, cutoff mRS for defining a poor functional outcome, or the study quality score in NOS could significantly affect the association between circulating GDF-15 at admission and the risk of poor functional outcome after AIS (p all >0.05; Table 3). However, publication year and NOS scores have relatively higher adjusted R^2 (27.4% and 28.3%, respectively), indicating they may explain more of the between-study heterogeneity compared to other factors. These findings should be interpreted cautiously, acknowledging the lack of statistical significance.

| Δ | | | | | Odds Ratio | Odds Ratio | |
|------------|---|---|--|--|---|----------------------------------|---|
| <u> </u> | Study or Subgroup | log[Odds Ratio] | SE | Weight | IV, Random, 95% CI | IV, Random, 95% Cl | |
| | 1.6.1 ROC analysis de | erived | | | | | |
| | Worthmann 2011 | 0.94 | 0.33 | 12.5% | 2.56 [1.34, 4.89] | | |
| | Cai 2020 | 1.43 | 0.26 | 14.4% | 4.22 [2.51, 7.10] | _ | |
| | LI 2022 | 1.47 | 0.25 | 14.7% | 4.39 [2.65, 7.27] | | |
| | Subtotal (95% CI) | | | 41.6% | 3.80 [2.77, 5.22] | | |
| | Heterogeneity: Tau ² = 0 | 0.00; Chi² = 1.90, df = 2 | 2 (P = 0.39 |); I² = 0% | | | |
| | Test for overall effect: 2 | Z = 8.29 (P < 0.00001) | | | | | |
| | 1.6.2 Median | | | | | | |
| | Groschel 2012 | 0.41 | 0.16 | 17.5% | 1.51 [1.10, 2.07] | | |
| | Geng 2015 | 0.53 | 0.35 | 11.7% | 1.70 [0.84, 3.44] | + | |
| | Lu 2015 | 1.02 | 0.37 | 11.2% | 2.80 [1.33, 5.89] | | |
| | Dong 2019 | 1.21 | 0.37 | 11.3% | 3.38 [1.62, 7.04] | | |
| | Jeong 2020 | 1.59 | 0.60 | 6.6% | 4.93 [1.49, 16.30] | | |
| | Subtotal (95% CI) | | | 58.4% | 2.25 [1.48, 3.42] | | |
| | Heterogeneity: Tau ² = 0 | 0.11; Chi² = 8.01, df = 4 | (P = 0.09 |); l² = 50% | % | | |
| | Test for overall effect: 2 | Z = 3.78 (P = 0.0002) | | | | | |
| | Total (95% CI) | | | 100.0% | 2.82 [1.94, 4.10] | • | |
| | Heterogeneity: Tau ² = 0 | 0.18; Chi² = 21.71, df = | 7 (P = 0.0 | 03); I² = 6 | \$8% | | ł |
| | Test for overall effect: 2 | Z = 5.41 (P < 0.00001) | | | 0.0 | JS 0.2 I S 20 | , |
| | Test for subaroup differ | rences: Chi ² = 3.84. df : | = 1 (P = 0. | (07) , $ ^2 = 7$ | 74 0% | | |
| | | | | | 4.070 | | |
| R | | | | | Odds Ratio | Odds Ratio | |
| B | Study or Subgroup | log[Odds Ratio] | SE | Weight | Odds Ratio IV, Random, 95% Cl | Odds Ratio IV, Random, 95% Cl | |
| B | Study or Subgroup 1.7.1 GDF-15 cutoff < | log[Odds Ratio] 1200 ng/l | SE | Weight | Odds Ratio IV, Random, 95% Cl | Odds Ratio IV, Random, 95% Cl | |
| B | <u>Study or Subgroup</u> 1.7.1 GDF-15 cutoff < Geng 2015 | log[Odds Ratio] 1200 ng/l 0.53 | SE 0.35 | Weight 8.5% | Odds Ratio IV. Random, 95% CI 1.70 [0.84, 3.44] | Odds Ratio IV, Random, 95% Cl | |
| B _ | Study or Subgroup 1.7.1 GDF-15 cutoff < Geng 2015 Dong 2019 | log[Odds Ratio] 1200 ng/l 0.53 1.21 | SE 0.35 0.37 | Weight 8.5% 8.2% | Odds Ratio IV, Random, 95% Cl 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] | Odds Ratio IV, Random, 95% Cl | |
| B_ | Study or Subgroup 1.7.1 GDF-15 cutoff < Geng 2015 Dong 2019 Jeong 2020 | log[Odds Ratio] 1200 ng/l 0.53 1.21 1.59 | SE 0.35 0.37 0.60 | Weight 8.5% 8.2% 4.3% | Odds Ratio IV. Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] | Odds Ratio IV, Random, 95% Cl | |
| B _ | Study or Subgroup 1.7.1 GDF-15 cutoff < Geng 2015 Dong 2019 Jeong 2020 Cai 2020 | log[Odds Ratio] 1200 ng/l 0.53 1.21 1.59 1.43 | SE 0.35 0.37 0.60 0.26 | Weight 8.5% 8.2% 4.3% 11.1% | Odds Ratio <u>IV, Random, 95% CI</u> 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] | Odds Ratio IV, Random, 95% Cl | |
| B_ | Study or Subgroup 1.7.1 GDF-15 cutoff < Geng 2015 Dong 2019 Jeong 2020 Cai 2020 LI 2022 | log[Odds Ratio] 1200 ng/l 0.53 1.21 1.59 1.43 1.47 | SE 0.35 0.37 0.60 0.26 0.25 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% | Odds Ratio IV, Random, 95% Cl 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] | Odds Ratio IV, Random, 95% Cl | |
| B | Study or Subgroup 1.7.1 GDF-15 cutoff < | log[Odds Ratio] 1200 ng/l 0.53 1.21 1.59 1.43 1.47 | SE 0.35 0.37 0.60 0.26 0.25 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% 43.6% | Odds Ratio IV, Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] 3.52 [2.47, 5.00] | Odds Ratio IV, Random, 95% CI | |
| B_ | Study or Subgroup 1.7.1 GDF-15 cutoff < Geng 2015 Dong 2019 Jeong 2020 Cai 2020 LI 2022 Subtotal (95% CI) Heterogeneity: Tau ² = | log[Odds Ratio] 1200 ng/l 0.53 1.21 1.59 1.43 1.47 0.05; Chi ² = 5.60, df = 4 | SE 0.35 0.37 0.60 0.26 0.25 ↓ (P = 0.23 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% 43.6% 3); ² = 29° | Odds Ratio IV, Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] 3.52 [2.47, 5.00] % | Odds Ratio IV, Random, 95% Cl | |
| B_ | Study or Subgroup 1.7.1 GDF-15 cutoff < Geng 2015 Dong 2019 Jeong 2020 Cai 2020 LI 2022 Subtotal (95% CI) Heterogeneity: Tau ² = Test for overall effect: 2 | log[Odds Ratio] 1200 ng/l 0.53 1.21 1.59 1.43 1.47 0.05; Chi ² = 5.60, df = 4 Z = 7.01 (P < 0.00001) | SE 0.35 0.37 0.60 0.26 0.25 4 (P = 0.23 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% 43.6% 3); ² = 29° | Odds Ratio IV. Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] 3.52 [2.47, 5.00] % | Odds Ratio IV, Random, 95% CI | |
| B | Study or Subgroup 1.7.1 GDF-15 cutoff < Geng 2015 Dong 2019 Jeong 2020 Cai 2020 LI 2022 Subtotal (95% CI) Heterogeneity: Tau ² = Test for overall effect: 2 1.7.2 GDF-15 cutoff ≥ | log[Odds Ratio] 0.53 1.21 1.59 1.43 1.47 0.05; Chi ² = 5.60, df = 4 Z = 7.01 (P < 0.00001) ≥ 1200 ng/l | SE 0.35 0.37 0.60 0.26 0.25 4 (P = 0.23 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% 43.6% 3); l² = 29° | Odds Ratio IV, Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] 3.52 [2.47, 5.00] % | Odds Ratio IV, Random, 95% Cl | |
| B_ | Study or Subgroup 1.7.1 GDF-15 cutoff < | log[Odds Ratio] 0.53 1.21 1.59 1.43 1.47 0.05; Chi ² = 5.60, df = 4 Z = 7.01 (P < 0.00001) ≥ 1200 ng/l 0.94 | SE 0.35 0.37 0.60 0.26 0.25 ↓ (P = 0.23 0.33 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% 43.6% 3); I² = 29° 9.3% | Odds Ratio IV, Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] 3.52 [2.47, 5.00] % | Odds Ratio IV, Random, 95% CI | |
| B | Study or Subgroup1.7.1 GDF-15 cutoff < | log[Odds Ratio] 0.53 1.21 1.59 1.43 1.47 0.05; Chi ² = 5.60, df = 4 Z = 7.01 (P < 0.00001) ≥ 1200 ng/l 0.94 0.41 | SE 0.35 0.37 0.60 0.26 0.25 ↓ (P = 0.23 0.33 0.16 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% 43.6% 3); I² = 29° 9.3% 14.5% | Odds Ratio IV, Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] 3.52 [2.47, 5.00] % 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] | Odds Ratio IV, Random, 95% CI | |
| B_ | Study or Subgroup 1.7.1 GDF-15 cutoff < Geng 2015 Dong 2019 Jeong 2020 Cai 2020 LI 2022 Subtotal (95% CI) Heterogeneity: Tau ² = Test for overall effect: 3 1.7.2 GDF-15 cutoff ≥ Worthmann 2011 Groschel 2012 Lu 2015 | log[Odds Ratio] 0.53 1.21 1.59 1.43 1.47 0.05; Chi ² = 5.60, df = 4 Z = 7.01 (P < 0.00001) ≥ 1200 ng/l 0.94 0.41 1.02 | SE 0.35 0.37 0.60 0.26 0.25 4 (P = 0.23 0.33 0.16 0.37 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% 43.6% 3); I² = 29° 9.3% 14.5% 8.1% | Odds Ratio IV, Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] 3.52 [2.47, 5.00] % 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 2.80 [1.33, 5.89] | Odds Ratio IV, Random, 95% CI | |
| B | Study or Subgroup 1.7.1 GDF-15 cutoff < Geng 2015 Dong 2019 Jeong 2020 Cai 2020 LI 2022 Subtotal (95% CI) Heterogeneity: Tau ² = Test for overall effect: 3 1.7.2 GDF-15 cutoff ≥ Worthmann 2011 Groschel 2012 Lu 2015 Yin 2019 | log[Odds Ratio] 0.53 1.21 1.59 1.43 1.47 0.05; Chi ² = 5.60, df = 4 Z = 7.01 (P < 0.00001) ≥ 1200 ng/l 0.94 0.41 1.02 0.59 | SE 0.35 0.37 0.60 0.26 0.25 4 (P = 0.23 0.33 0.16 0.37 0.12 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% 43.6% 3); I² = 29° 9.3% 14.5% 8.1% 15.5% | Odds Ratio IV, Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] 3.52 [2.47, 5.00] % 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] | Odds Ratio IV, Random, 95% CI | |
| B | Study or Subgroup 1.7.1 GDF-15 cutoff < Geng 2015 Dong 2019 Jeong 2020 Cai 2020 LI 2022 Subtotal (95% CI) Heterogeneity: Tau ² = 1 Test for overall effect: 3 1.7.2 GDF-15 cutoff ≥ Worthmann 2011 Groschel 2012 Lu 2015 Yin 2019 Breniere 2019 | log[Odds Ratio] 0.53 1.21 1.59 1.43 1.47 0.05; Chi ² = 5.60, df = 4 Z = 7.01 (P < 0.00001) ≥ 1200 ng/l 0.94 0.41 1.02 0.59 0.98 | SE 0.35 0.37 0.60 0.26 0.25 4 (P = 0.23 0.33 0.16 0.37 0.12 0.34 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% 43.6% 8); I² = 29° 9.3% 14.5% 8.1% 15.5% 9.0% | Odds Ratio IV, Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] 3.52 [2.47, 5.00] % 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] 2.69 [1.38, 5.24] | Odds Ratio IV, Random, 95% CI | |
| B | Study or Subgroup 1.7.1 GDF-15 cutoff < | log[Odds Ratio] 0.53 1.21 1.59 1.43 1.47 0.05; Chi ² = 5.60, df = 4 Z = 7.01 (P < 0.00001) ≥ 1200 ng/l 0.94 0.41 1.02 0.59 0.98 | SE 0.35 0.37 0.60 0.26 0.25 4 (P = 0.23 0.33 0.16 0.37 0.12 0.34 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% 43.6% 8); I ² = 29° 9.3% 14.5% 8.1% 15.5% 9.0% 56.4% | Odds Ratio IV, Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] 3.52 [2.47, 5.00] % 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] 2.69 [1.38, 5.24] 1.90 [1.53, 2.35] | Odds Ratio IV, Random, 95% CI | |
| B | Study or Subgroup1.7.1 GDF-15 cutoff < | log[Odds Ratio] 0.53 1.21 1.59 1.43 1.47 0.05; Chi ² = 5.60, df = 4 Z = 7.01 (P < 0.00001) ≥ 1200 ng/l 0.94 0.41 1.02 0.59 0.98 0.01; Chi ² = 5.01, df = 4 | SE 0.35 0.37 0.60 0.26 0.25 4 (P = 0.23 0.33 0.16 0.37 0.12 0.34 4 (P = 0.29 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% 43.6% 3); ² = 29° 9.3% 14.5% 8.1% 15.5% 9.0% 56.4% 9); ² = 20° | Odds Ratio IV. Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] 3.52 [2.47, 5.00] % 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] 2.69 [1.38, 5.24] 1.90 [1.53, 2.35] % | Odds Ratio IV, Random, 95% CI | |
| B _ | Study or Subgroup1.7.1 GDF-15 cutoff < | log[Odds Ratio] 0.53 1.21 1.59 1.43 1.47 0.05; Chi² = 5.60, df = 4 Z = 7.01 (P < 0.00001) ≥ 1200 ng/l 0.94 0.41 1.02 0.59 0.98 0.01; Chi² = 5.01, df = 4 Z = 5.86 (P < 0.00001) | SE 0.35 0.37 0.60 0.26 0.25 4 (P = 0.23 0.33 0.16 0.37 0.12 0.34 4 (P = 0.29 | Weight 8.5% 8.2% 4.3% 11.1% 11.4% 43.6% 8); I² = 29° 9.3% 14.5% 8.1% 15.5% 9.0% 56.4% 9); I² = 20° | Odds Ratio IV. Random, 95% CI 1.70 [0.84, 3.44] 3.38 [1.62, 7.04] 4.93 [1.49, 16.30] 4.22 [2.51, 7.10] 4.39 [2.65, 7.27] 3.52 [2.47, 5.00] % 2.56 [1.34, 4.89] 1.51 [1.10, 2.07] 2.80 [1.33, 5.89] 1.81 [1.41, 2.32] 2.69 [1.38, 5.24] 1.90 [1.53, 2.35] % | Odds Ratio IV, Random, 95% CI | |

Fig. 4 Forest plots for the subgroup analysis of the association between circulating GDF-15 level at admission and functional outcome after AlS; **A**, subgroup analysis according to the methods for defining the cutoff of GDF-15; and **B**, subgroup analysis according to the cutoff values for defining a high circulating GDF-15 level

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| | | | | Odds Ratio | Ode | ds Ratio | |
|---|---|-------------|-------------------------|--------------------|----------|--------------------|---|
| Study or Subgroup | log[Odds Ratio] | SE | Weight | IV, Random, 95% CI | IV, Ran | <u>dom, 95% Cl</u> | |
| 1.8.1 mRS 2~6 | | | | | | | |
| Worthmann 2011 | 0.94 | 0.33 | 9.3% | 2.56 [1.34, 4.89] | | | |
| Groschel 2012 | 0.41 | 0.16 | 14.5% | 1.51 [1.10, 2.07] | | | |
| Geng 2015 | 0.53 | 0.35 | 8.5% | 1.70 [0.84, 3.44] | | | |
| Lu 2015 | 1.02 | 0.37 | 8.1% | 2.80 [1.33, 5.89] | | | |
| Subtotal (95% CI) | | | 40.4% | 1.85 [1.37, 2.50] | | • | |
| Heterogeneity: Tau ² = | 0.02; Chi ² = 3.71, df = 3 | 3 (P = 0.29 |); l² = 199 | 6 | | | |
| Test for overall effect: | Z = 4.02 (P < 0.0001) | | | | | | |
| 1.8.2 mRS 3~6 | | | | | | | |
| Yin 2019 | 0.59 | 0.12 | 15.5% | 1.81 [1.41, 2.32] | | | |
| Breniere 2019 | 0.98 | 0.34 | 9.0% | 2.69 [1.38, 5.24] | | | |
| Dong 2019 | 1.21 | 0.37 | 8.2% | 3.38 [1.62, 7.04] | | | |
| Jeong 2020 | 1.59 | 0.60 | 4.3% | 4.93 [1.49, 16.30] | | | |
| Cai 2020 | 1.43 | 0.26 | 11.1% | 4.22 [2.51, 7.10] | | | |
| LI 2022 | 1.47 | 0.25 | 11.4% | 4.39 [2.65, 7.27] | | | |
| Subtotal (95% CI) | | | 59.6% | 3.17 [2.09, 4.82] | | • | |
| Heterogeneity: Tau ² = | 0.17; Chi ² = 17.04, df = | 5 (P = 0.0 | 04); l ² = 7 | ′1% | | | |
| Test for overall effect: | Z = 5.42 (P < 0.00001) | | ,, | | | | |
| Total (95% CI) | | | 100.0% | 2.60 [1.95, 3.46] | | • | |
| Heterogeneity: Tau ² = | 0.12; Chi ² = 25.65, df = | 9 (P = 0.0 | 02); l ² = 6 | 5% | | <u> </u> | |
| Test for overall effect: | Z = 6.53 (P < 0.00001) | , | ,, | | 0.05 0.2 | 1 5 | 2 |
| Test for subaroup diffe | erences: Chi² = 4.22. df | = 1 (P = 0. | 04). I² = 7 | 6.3% | | | |
| | | | | Odds Ratio | Ode | ds Ratio | |
| Study or Subgroup | log[Odds Ratio] | SE | Weight | IV, Random, 95% CI | IV, Ran | <u>dom, 95% Cl</u> | |
| 1.9.1 Univariate | | | | | | | |
| Geng 2015 | 0.53 | 0.35 | 8.5% | 1.70 [0.84, 3.44] | | + | |
| Lu 2015 | 1.02 | 0.37 | 8.1% | 2.80 [1.33, 5.89] | | | |
| Breniere 2019 | 0.98 | 0.34 | 9.0% | 2.69 [1.38, 5.24] | | | |
| Jeong 2020 | 1.59 | 0.60 | 4.3% | 4.93 [1.49, 16.30] | | | |
| LI 2022 | 1.47 | 0.25 | 11.4% | 4.39 [2.65, 7.27] | | | |
| Subtotal (95% CI) | | | 11 3% | 3 03 [2 10 4 38] | | • | |
| | | | 41.370 | 0.00 [2.10, 4.00] | | • | |
| Heterogeneity: Tau ² = | 0.05; Chi² = 5.45, df = 4 | 4 (P = 0.24 |); l ² = 279 | % | | | |
| Heterogeneity: Tau ² = Test for overall effect: | 0.05; Chi² = 5.45, df = 4 Z = 5.90 (P < 0.00001) | 4 (P = 0.24 |); l ² = 279 | % | | | |

| Heterogeneity: Tau ² = 0.0 | 05; Chi² = 5.45, df | = 4 (P = 0.24 | l); l² = 27% | | | | |
|---------------------------------------|---------------------|----------------|---------------------------|-------------------|----------|-----|-----|
| Test for overall effect: Z = | = 5.90 (P < 0.0000 | 1) | | | | | |
| 1.9.2 Multivariate | | | | | | | |
| Worthmann 2011 | 0.94 | 0.33 | 9.3% | 2.56 [1.34, 4.89] | | | |
| Groschel 2012 | 0.41 | 0.16 | 14.5% | 1.51 [1.10, 2.07] | | | |
| Yin 2019 | 0.59 | 0.12 | 15.5% | 1.81 [1.41, 2.32] | | | |
| Dong 2019 | 1.21 | 0.37 | 8.2% | 3.38 [1.62, 7.04] | | | |
| Cai 2020 | 1.43 | 0.26 | 11.1% | 4.22 [2.51, 7.10] | | | |
| Subtotal (95% CI) | | | 58.7% | 2.34 [1.62, 3.37] | | • | |
| Heterogeneity: Tau ² = 0. | 12; Chi² = 14.12, d | f = 4 (P = 0.0 | 007); l² = 72 | 2% | | | |
| Test for overall effect: Z | = 4.53 (P < 0.0000 | 1) | · | | | | |
| Total (95% CI) | | | 100.0% | 2.60 [1.95, 3.46] | | • | |
| Heterogeneity: Tau ² = 0. | 12; Chi² = 25.65, d | f = 9 (P = 0.0 |)02); l ² = 65 | 5% | | | |
| To at fair an annual offer at 7 | | 4 | <i>,.</i> | | 0.05 0.2 | 1 5 | -20 |

Test for overall effect: Z = 6.53 (P < 0.00001)

Test for subaroup differences: Chi² = 0.97. df = 1 (P = 0.33). $I^2 = 0\%$

Fig. 5 Forest plots for the subgroup analyses of the association between circulating GDF-15 level at admission and functional outcome after AIS; **A**, subgroup analysis according to the definition of poor functional outcome; and **B**, subgroup analysis according to the analytic models (univariate or multivariate)

Table 3 Results of univariate meta-regression analysis

| Variables | OR for the as functional o | ssociation circula utcome after AIS | ting GDF- | 15 and |
|---|-------------------------------|--|-----------|------------------------------|
| | Coefficient | 95% CI | p values | Adjust- ed R ² |
| Publication year | 0.071 | -0.011 to 0.153 | 0.11 | 27.4% |
| Sample size | -0.00015 | -0.00046 to 0.00015 | 0.28 | 2.9% |
| Mean age (years) | -0.036 | -0.090 to 0.017 | 0.16 | 19.5% |
| Men (%) | -0.012 | -0.064 to 0.040 | 0.62 | -11.3% |
| Cutoff value of GDF-15 (ng/L) | -0.00045 | -0.00123 to 0.00033 | 0.22 | 8.4% |
| Cutoff mRS of poor functional outcome | 0.45 | -0.18 to 1.09 | 0.14 | 18.7% |
| NOS | -0.19 | -0.42 to 0.04 | 0.10 | 28.3% |

GDF-15, growth differentiation factor 15; AIS, acute isckemic stroke; OR, odds ratio; CI, confidence interval; mRS, modified Rankin Score; NOS, Newcastle-Ottawa Scale



Fig. 6 Funnel plots for the meta-analysis of the association between circulating GDF-15 level at admission and functional outcome after AIS

Publication bias

Visual inspection of the funnel plots for the meta-analysis examining the relationship between circulating GDF-15 levels and poor functional outcomes after AIS indicates symmetry, suggesting a low risk of publication bias (Fig. 6). This observation is further supported by Egger's regression test results (p=0.45), which also suggest a low risk of publication bias.

Discussion

This meta-analysis synthesizes data from ten studies to examine the relationship between circulating GDF-15 levels and functional outcomes after AIS. Our findings reveal a consistent association between elevated GDF-15 levels and an increased risk of poor functional outcomes, with a pooled OR of 2.60. This significant association underscores the potential utility of GDF-15 as a prognostic biomarker in stroke patients, helping to identify those at higher risk for adverse recovery.

GDF-15 is a member of the transforming growth factor-beta (TGF- β) superfamily and plays a critical role in cellular stress responses and inflammation [31]. Its physiological and pathological functions have been linked to various mechanisms that could explain its association with poor functional outcomes after stroke. One key mechanism involves its role in inflammation [32]. GDF-15 is upregulated in response to tissue injury and inflammation and acts as an anti-inflammatory cytokine. It modulates macrophage activation and inhibits the release of pro-inflammatory cytokines [33]. Elevated GDF-15 levels may indicate a dysregulated inflammatory response, where excessive inflammation or prolonged inflammatory signaling contributes to tissue damage and impaired recovery [34]. Additionally, GDF-15 influences apoptosis and cell survival pathways [35]. It has been shown to modulate apoptosis through the activation of the MAPK (mitogen-activated protein kinase) and PI3K/ Akt (phosphoinositide 3-kinase/protein kinase B) signaling pathways [36]. These pathways are crucial in regulating cell survival and apoptosis, particularly in the context of ischemic injury [37]. Elevated GDF-15 levels might reflect enhanced cell death or apoptosis in brain tissue, contributing to poor functional outcomes. A recent study using a rat model of AIS found that GDF-15 gene expression significantly increased in the ipsilateral cortex and cerebellum, with a smaller increase observed in the contralateral cortex [38]. Additionally, GDF-15 expression was correlated with the neurological deficit score [38]. Additionally, GDF-15 is also involved in tissue repair and remodeling. After ischemic injury, the repair processes are crucial for recovery. The role of GDF-15 in modulating these processes, potentially by influencing fibroblast and endothelial cell functions, could impact the extent of tissue damage and the recovery trajectory [39, 40]. High levels of GDF-15 might signify an imbalance in repair processes, further impairing functional recovery. Studies are warranted in the future to validate these hypotheses.

Our subgroup analyses reveal important insights into the variability of the association between GDF-15 and functional outcomes after AIS. The observed stronger association in studies using cutoff values for GDF-15<1200 ng/L compared to ≥1200 ng/L suggests that lower cutoff values may better capture clinically significant elevations of GDF-15 associated with poorer outcomes. This implies that high GDF-15 levels are more predictive of adverse outcomes when measured against lower thresholds, possibly due to a more pronounced inflammatory or pathological response captured by these values. The variation in the definition of poor functional outcomes also impacts the observed associations. Studies defining poor outcomes as an mRS score \geq 3 showed a stronger association with elevated GDF-15 levels compared to those defining poor outcomes as mRS \geq 2.

This indicates that high GDF-15 levels are more closely linked with severe disability rather than moderate disability. This finding highlights potential role of GDF-15 as a marker for severe functional impairments, providing valuable prognostic information that could guide clinical decision-making.

This meta-analysis has several strengths. We employed a comprehensive search strategy across multiple databases, ensuring a thorough identification of relevant studies. The use of a random-effects model allowed us to account for variability across studies and provided a more accurate estimate of the association between GDF-15 levels and functional outcomes. Only longitudinal observational studies were included, so as to provide a sequential relationship between high GDF-15 and increased risk of poor functional outcome after AIS. Sensitivity analyses further validated the robustness of our findings, while subgroup analyses offered insights into how study characteristics influence results. However, there are limitations to consider. The variability in measurement methods for GDF-15, cutoff values used, and definitions of poor outcomes across studies could affect the generalizability of the results. In addition, influences of concurrent medications and treatments for AIS on the results of the metaanalysis could not be determined because the influences of these factors were generally not reported among the included studies. For example, statins have been wellacknowledged to exert anti-inflammatory effect [41]. It remains to be determined if the association between a high circulating GDF-15 level and the increased risk of poor functional outcome after AIS could be significantly modified by the use of statins. Moreover, the observational nature of the included studies limits the ability to establish causality and may introduce biases. Finally, the quality of studies varied, with some having methodological limitations that could influence the findings.

From a clinical perspective, our findings suggest that elevated circulating GDF-15 levels could be a valuable biomarker for predicting poor functional outcomes in stroke patients. Identifying patients with high GDF-15 levels could help clinicians stratify risk and tailor treatment strategies accordingly. Incorporating GDF-15 into clinical assessments could improve prognosis and guide therapeutic interventions, potentially enhancing patient outcomes. Future research should focus on standardizing measurement methods for GDF-15 and defining optimal cutoff values to improve consistency across studies. Investigating the role of GDF-15 in different stroke subtypes and its interaction with other biomarkers could provide a more comprehensive understanding of its prognostic value. Additionally, prospective studies and clinical trials are needed to validate GDF-15 as a clinical tool and explore its potential therapeutic applications.

Conclusions

This meta-analysis demonstrates a significant association between elevated circulating GDF-15 levels and poor functional outcomes following AIS. The consistent findings across studies highlight the potential role of GDF-15 as a prognostic biomarker. While further research is needed to confirm these results and refine clinical applications, GDF-15 could play a crucial role in improving stroke management and patient outcomes by identifying individuals at higher risk for adverse recovery.

Supplementary Information

The online version contains supplementary material available at https://doi.or g/10.1186/s13005-024-00476-4.

Supplementary Material 1

Supplementary Material 2

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Not applicable.

Author contributions

Yulang Wu and Fengkun Zhou designed the study. Yulang Wu and Yude Wei performed database search, literature review, study identification, study quality assessment, and data collection. Yulang Wu, Jinrong He, and Fengkun Zhou performed statistical analyses and interpreted the results. Yulang Wu drafted the manuscript. All authors revised the manuscript and approved the submission.

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Data availability

No datasets were generated or analysed during the current study.

Declarations

Ethical approval

Ethics approval was not required as this was a systematic review of published studies.

Informed consent

Not applicable.

Competing interests

The authors declare no competing interests.

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