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# Association of continuous renal replacement therapy downtime with fuid balance gap and clinical outcomes: a retrospective cohort analysis utilizing EHR and machine data

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### **Abstract**

**Background** Fluid balance gap (FBgap—prescribed vs. achieved) is associated with hospital mortality. Downtime is an important quality indicator for the delivery of continuous renal replacement therapy (CRRT). We examined the association of CRRT downtime with FBgap and clinical outcomes including mortality.

**Methods** This is a retrospective cohort study of critically ill adults receiving CRRT utilizing both electronic health records (EHR) and CRRT machine data. FBgap was calculated as achieved minus prescribed fuid balance. Downtime, or percent treatment time loss (%TTL), was defined as CRRT downtime in relation to the total CRRT time. Data collection stopped upon transition to intermittent hemodialysis when applicable. Linear and logistic regression models were used to analyze the association of %TTL with FBgap and hospital mortality, respectively. Covariates included demographics, Sequential Organ Failure Assessment (SOFA) score at CRRT initiation, use of organ support devices, and the interaction between %TTL and machine alarms.

**Results** We included 3630 CRRT patient-days from 500 patients with a median age of 59.5 years (IQR 50–67). Patients had a median SOFA score at CRRT initiation of 13 (IQR 10–16). Median %TTL was 8.1% (IQR 4.3–12.5) and median FBgap was 17.4 mL/kg/day (IQR 8.2-30.4). In adjusted models, there was a significant positive relationship between FBgap and %TTL only in the subgroup with higher alarm frequency (6+alarms per CRRT-day)  $(\beta = 0.87$  per 1% increase, 95%CI 0.48–1.26). No association was found in the subgroups with lower alarm frequency (0–2 and 3–5 alarms). There was no statistical evidence for an association between %TTL and hospital mortality in the adjusted model with the interaction term of alarm frequency.

**Conclusions** In critically ill adult patients undergoing CRRT, %TTL was associated with FBgap only in the subgroup with higher alarm frequency, but not in the other subgroups with lower alarms. No association between %TTL and mortality was observed. More frequent alarms, possibly indicating unexpected downtime, may suggest compromised CRRT delivery and could negatively impact FBgap.

**Keywords** Fluid management, CRRT, Mortality, AKI, Downtime

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#### **Background**

Acute kidney injury (AKI) is common in critically ill adults admitted to the intensive care unit (ICU) [[1\]](#page-8-0). Critically ill patients with severe AKI and positive fuid balance or fuid accumulation that do not respond to diuretics are often treated with continuous renal replacement therapy (CRRT) with the goal of mitigating poor outcomes that are associated with fuid accumulation [[2](#page-8-1)–[7](#page-8-2)]. Managing positive fuid balance with CRRT requires prescription of a fuid removal rate according to fuid balance goals and adjusting these goals based on achieved fuid balance and the patient's clinical course [[8\]](#page-8-3). It has recently been shown that despite provider best eforts, there is frequently a gap between the prescribed fuid balance and what is efectively achieved. Moreover, this fuid balance gap (FBgap) has been shown to be associated with hospital mortality [[9](#page-8-4)]. There is paucity of data about what contributes to the FBgap. The Acute Disease Quality Initiative (ADQI) 2019 consensus conference identifed (among other factors) downtime as a key quality indicator in CRRT treatment [[10](#page-8-5)]. However, the impact downtime has on fuid management and clinical outcomes has not been thoroughly investigated  $[11-14]$  $[11-14]$  $[11-14]$  $[11-14]$  $[11-14]$ .

We hypothesized that percent treatment time loss (%TTL), in other words CRRT machine downtime during a patient's CRRT treatment course, would be positively associated with FBgap—higher downtime, higher FBgap. We aimed to describe how %TTL relates to FBgap and relevant clinical outcomes after adjusting for patient-related factors. Finally, we set out to describe patient and machine factors that may be implicated in %TTL.

#### **Methods**

#### **Study design and participants**

We conducted a single-center, retrospective cohort study of adult patients ( $\geq$  18 years) admitted to all ICUs with AKI who received CRRT at the University of Kentucky Albert B. Chandler Hospital between August 2017 and April 2021. We excluded patients for whom ICU admission and discharge records were unavailable, those missing data (body weight at admission, Sequential Organ Failure Assessment (SOFA) score, fluid balance, and machine parameters), and those with end stage kidney disease or kidney transplant. Body weight values below the 0.5th percentile or above the 99.5th percentile were considered outliers and treated as missing. To address outliers in FBgap, we used Cook's distance based on a linear regression model that included both FBgap and %TTL as the dependent and independent variable, respectively [[15\]](#page-8-8). See additional materials regarding distribution of FBgap (Additional Fig. [1](#page-1-0)). Patients with Cook's distance values exceeding four times the mean Cook's distance of the cohort were excluded from the analysis. CRRT was performed with continuous venovenous hemodiafltration according to institution standards with regional citrate most commonly used for anticoagulation. This study was approved by the Institutional Review Board (IRB) from the University of Kentucky (17–0444-P1G). Given the observational nature of this investigation, informed consent was waived by the Institutional Review Board.

#### **Data sources**

We utilized two data sources: electronic health record (EHR) and CRRT machine data. We obtained data cards containing machine parameters from the CRRT



<span id="page-1-0"></span>**Fig. 1** Consort diagram. RRT—continuous renal replacement therapy, ICU—intensive care unit, SOFA—sequential organ failure assessment, EHR electronic health record, FBgap—fuid balance gap, %TTL—treatment time lost

machines through manual extraction. We matched EHR data with CRRT machine data in a phased approach, utilizing variables in descending order of priority, such as the medical record number, patient name, encounter identifcation number, CRRT start and stop dates/times, death date/time (if applicable), body weight, and total machine fuid removal per day. For patients who could not be matched through these criteria, we employed a further matching process using the remaining available variables from the CRRT data and dates with information available in the patient record. These variables were cross-referenced against known values in the clinical record. To avoid the use of inaccurate information, only patients whose machine data and EHR matched with a timestamp concordance rate of 90% or higher were included. For a more detailed description of the matching process, refer to the additional material (Additional Table 2).

#### **Primary independent and dependent variables**

The primary independent variable, %TTL, was calculated based on data from CRRT machines. %TTL represents the ratio of CRRT downtime (therapy is paused) to the total CRRT time when the machine is on (except when priming). As %TTL is a ratio, it is adjusted for total time on CRRT. The analysis would have excluded any intentional CRRT stops. Similarly, once CRRT was permanently stopped, such as in patients with kidney function recovery or planned transition to intermittent hemodialysis data collection stopped. Data regarding specifc causes of downtime were unavailable with the exception of frequency of CRRT alarms. Alarms from the CRRT machine, a component of downtime calculation, were categorized as catheter alarms (access and return alarms) and flter alarms (transmembrane pressure [TMP], flter pressure [FP], and clotting alarms). The frequency of these alarms, along with the number of flter changes and the effluent flow rate (CRRT dose), were extracted from machine data. The daily average alarms and number of each type of alarm were calculated. Total alarms per day were categorized into three groups:  $0-2$ ,  $3-5$ , and  $6+$ , to ensure a close to equal distribution of patients within each group, guided by prior data distribution knowledge at the study site. Details on alarm processing is available in the additional materials (Additional Table 2).

The primary outcome of interest, FBgap, represents the deviation of the actual net fuid balance from the pre-established fuid balance goal, adjusted by the number of CRRT days and the patient's admission body weight. FBgap was calculated using the formula: (actual patient's net fuid balance – patient's fuid balance goal prescribed during CRRT)/total CRRT days/body



<span id="page-2-0"></span>**Fig. 2** Relationship between FBgap and %TTL stratifed by total alarms. FBgap, fuid balance gap; %TTL, percent treatment time loss; CCI, Charlson comorbidity index; SOFA, sequential organ failure assessment; CRRT, continuous renal replacement therapy; MV, mechanical ventilation; IABP, intra-aortic balloon pump; VAD, ventricular assist device; ECMO, extracorporeal membrane oxygenation. Adjusted FBgap was calculated based on the multivariate linear regression model with FBgap as the dependent variable and %TTL, along with age, sex, race, CCI, SOFA score at CRRT initiation, use of MV at CRRT initiation, use of circulatory support devices (any of IABP, VAD, ECMO) at CRRT initiation, and %fuid overload as independent variables

weight, expressed in mL/kg/day. Additional Fig. [2](#page-2-0) illustrates the distribution of patient's FB prescribed and FB achieved. FB goal was determined by clinicians as "no fuid removal in 24 h", "net negative X liters in 24 h", or "net even matching intake and output in 24 h" and was aggregated during the total period of CRRT for analysis. Notably, if fuid removal was not clinically indicated a FB goal was still documented. Similarly, if a patient had increased native urine output, this would be included in the net FB calculation and thus accounted for in their FBgap. These goals were set according to clinicians' goals of fuid management and adjusted on average once or twice a day. Nevertheless, the FB goal could be adjusted more frequently if clinically necessary as the bedside nurse has a CRRT dedicated flowsheet that compiled all fuid management parameters every hour. This flowsheet accounted for the FB goal per hour (ex. FB goal net negative 50 ml/h for a FB goal net negative of 1.2 L in 24 h), the defcit from the prior hour, and the suggested target to set the machine for the next hour. These data were documented in the electronic CRRT flowsheet and then extracted from the Electronic Data Warehouse for validation and analysis. The FB goals were determined by consultant nephrologists in collaboration with ICU physicians, based on the institutional standard of care.

#### **Other variables: patient characteristics and process variables**

Patient demographics (age, sex, race, body weight, survival status, baseline estimated Glomerular Filtration Rate (eGFR), SOFA score [\[16](#page-8-9)], and International Classifcation of Diseases-code-based comorbidities) were extracted from the EHR. The Charlson Comorbidity Index (CCI) was calculated based on recorded comorbidities  $[17]$  $[17]$ . The use of CRRT, mechanical ventilation (MV), intra-aortic balloon pump (IABP), ventricular assist devices (VAD), and extracorporeal membrane oxygenation (ECMO) was also extracted from EHR. ICU and device free days were computed as the total number of days without each respective intervention within a 28-day period starting from the initiation of CRRT. If a patient died within 28 days, the free days were counted as zero. Fluid overload (%FO) was defned at CRRT initiation by the following formula: (Total Fluid Input in  $L -$ Total Fluid Output in L) \*100/Weight (Admission) in Kg.

#### **Statistical analysis**

The analyses in this study were divided into two major parts, and all statistical analyses and fgures were generated using R Studio version 4.2.2.

#### *Primary analysis*

To elucidate the relationship between %TTL and FBgap, a comprehensive linear regression analysis was conducted with multiple model specifcations. Specifcally, models were ftted both with and without adjustment for patient-specifc factors, and with and without identifed interactions. To identify these interactions, we utilized the Feasible Solution Algorithm (FSA) [[18](#page-8-11)], which identifed potentially novel two-way interactions between %TTL and patient/machine factors that contributed to variations in FBgap. This interaction identification analysis aimed to investigate how the main independent variable (%TTL) diferentially afected FBgap across various subgroups, thereby elucidating which group(s) are most susceptible to adverse efects associated with %TTL. In essence, this analysis provides a nuanced understanding of how %TTL infuences FBgap in distinct patient and machine contexts, ultimately informing strategies for mitigating potential negative outcomes.

#### *Secondary analysis*

We reported the characteristics of the cohort by %TTL tertiles, as well as by survival status at discharge. To explore the factors associated with %TTL, we used a linear regression model to examine patient factors, such as age, sex, race, CCI, SOFA score at CRRT initiation, use of MV at CRRT initiation, use of circulatory support devices (including IABP, VAD, ECMO) at CRRT initiation, and %FO at CRRT initiation. We also investigated the association between machine alarms, including access pressure, return pressure, transmembrane pressure, flter pressure, and clotting alarms, with %TTL.

#### *Multiplicity*

The alpha ( $\alpha$  = 0.05) spending in this study relates to the models outlined in the primary analysis section above. In total, the four planned models contained the main independent variable (%TTL) either with or without adjustment, and with or without an interaction term. To address the issue of multiplicity, we used the Bonferroni correction to adjust the signifcance level for these four models. Within these models, we identifed all terms involving %TTL (main or interaction terms). This number of terms, T, was used to adjust the cutoff for statistical significance by dividing 0.05 (α) by the number of terms. We also report the results of exploratory analyses, including logistic regression models for hospital mortality and linear regression models for ICU-, CRRT-, and MV-free days, without adjustments for multiple testing, as they are not part of our alpha spending aims in this study.

#### **Results**

#### **Cohort characteristics and clinical outcomes**

We identifed 908 patients admitted to the ICU who received CRRT during the study period. A total of 500 patients accounting for 3,630 patient–CRRT days were included in the final analysis (Fig.  $1$ ). The cohort had a median age of 59.5 years (Interquartile range [IQR]: 50.0, 67.3) and consisted of 305 males (61.0%), mostly white (89.8%). The median age differed based on %TTL tertile. The median baseline eGFR was  $41.6$  mL/min/1.73m<sup>2</sup> (IQR: 18.5, 74.8), the median CCI was 5 (IQR: 3, 7), and the median SOFA score at CRRT initiation was 13  $(IQR: 11, 16)$  (Table [1\)](#page-4-0). Cohort characteristics by mortality can be found in additional materials (Additional Table 3). The median total alarms per day was  $2.67$  (IQR: 1.50, 4.72), with catheter alarms being more frequent than filter alarms. The median FBgap was 17.37 mL/kg/ day (IQR: 8.22, 30.39). The association between patient and machine factors with FBgap can be found in additional materials (Additional Table 4). The median %TTL was 8.14% (IQR: 4.34, 12.45). The overall mortality was 54.2%. Median ICU length of stay was 10 days (IQR: 5, 19) and the median CRRT duration was 5 days (IQR: 3, 9) (Table [2\)](#page-5-0).

#### **Relationship between clinical and machine factors and %TTL**

In adjusted models with %TTL as the dependent variable, the β coefficient was  $-0.05$  (95% Confidence Interval

#### <span id="page-4-0"></span>**Table 1** Cohort characteristics stratifed by %TTL tertiles



pressure; FP, flter pressure; FBgap, fuid balance gap; IQR, interquartile range \*The tertile cutofs for %TTL values are as follows: lowest:<5.63%, middle: 5.63–10.76%, highest:≥10.77%

%TTL, percent treatment time loss; ICU, intensive care unit; SOFA, sequential organ failure assessment; eGFR, estimated glomerular fltration rate; TMP, transmembrane

†The number of each alarm is the average per day

[CI] −0.09 to −0.003) for age and −2.39 (95%CI −4.09 to  $-0.69$ ) for MV use at CRRT initiation. The β coefficient for %FO was 0.18 (95%CI 0.10–0.25), for access alarms 1.61 (95%CI 1.25–1.98), for return alarms 0.51 (95%CI 0.14–0.88), and for flter pressure alarms 1.75 (95%CI 0.33–3.46) (Table [3](#page-5-1)). We found no statistical evidence for association of sex, race, CCI, SOFA score at CRRT initiation, use of cardiac support device at CRRT initiation, TMP alarms, and clotting alarms with %TTL.

#### **Relationship between %TTL and FBgap**

Based on FSA, a signifcant interaction between %TTL and the total number of alarms was observed, suggesting that the association of %TTL with FBgap is infuenced by machine alarms. Therefore, considering the variation in the distribution of %TTL according to the number of alarms, we decided to categorize the total number of alarms into three groups  $(0-2, 3-5, 3)$  and  $6+$ alarms per CRRT-day) for further analysis (Additional Fig. 3). The

#### <span id="page-5-0"></span>**Table 2** Clinical and process outcomes stratifed by %TTL tertiles



The outcome of free days indicates the number of days within the 28-day period after CRRT initiation during which there was no exposure to each respective factor. If the patient died within the 28 days, the number of free days is treated as 0

ICU, intensive care unit; CRRT, continuous renal replacement therapy; MV, mechanical ventilation; IQR, interquartile range

#### <span id="page-5-1"></span>**Table 3** Association of patient and machine factors with %TTL



Linear regression model with %TTL as the dependent variable as well as the factors shown in the table as independent variables was performed

%TTL, percent treatment time loss; CCI, Charlson comorbidity index; SOFA, sequential organ failure assessment; CRRT, continuous renal replacement therapy; MV, mechanical ventilation; TMP, transmembrane pressure; FP, flter pressure; FBgap, fuid balance gap; IQR, interquartile range

association of %TTL with FBgap varied across the total alarms subgroups (Fig. [2](#page-2-0)).

In adjusted models with the Bonferroni correction, there was a negative association between %TTL and FBgap without stratifcation by the total alarms subgroups (Table [4\)](#page-6-0). However, a signifcant positive association was only found in the subgroup with a high number of alarms ( $\geq$ 6), with a β coefficient of 0.72 (95% CI 0.34 to 1.09) in the unadjusted model and a β coefficient of 0.87 (95% CI 0.48 to 1.26) in the adjusted model with the Bonferroni correction (Table [4](#page-6-0)).

#### **Relationship between %TTL and process outcomes**

Adjusted models did not show statistical evidence for an association of %TTL with in-hospital mortality (odds ratio

per 1% increase in %TTL: 0.99, 95% CI 0.96–1.03). Furthermore, we found no statistical evidence that %TTL was associated with ICU-free days (β coefficient:  $-0.13$ , 95% CI − 0.27 to 0.003), CRRT-free days (β coefficient: − 0.11, 95% CI – 0.26 to 0.05), or MV-free days ( $\beta$  coefficient: – 0.10, 95% CI − 0.26 to 0.06) (Additional Table 5). In the survivor-only analysis, the estimates were as follows: ICU-free days ( $\beta$  coefficient:  $-0.26$ , 95% CI  $-0.49$  to  $-0.03$ ), CRRTfree days (β coefficient:  $- 0.22$ , 95% CI  $- 0.47$  to 0.04), and MV-free days (β coefficient:  $-$  0.17, 95% CI  $-$  0.44 to 0.09) (Additional Table 5).

#### **Discussion**

We conducted a retrospective cohort study with the novel approach of combining EHR and CRRT machine data to investigate the relationship between %TTL



#### <span id="page-6-0"></span>**Table 4** Association of %TTL (independent variable) with FBgap (dependent variable)

The β coefficient represents the change in FBgap for each 1% increase in %TTL

Model 1: A linear regression with FBgap as the dependent variable and %TTL as the independent variable, as well as the interaction term between %TTL and total alarms

Model 2: A linear regression with FBgap as the dependent variable and %TTL as the independent variable, without including the interaction term. In addition, a subgroup analysis was performed for each category of total alarms (0–2, 3–5, 6+)

Model 3: A linear regression with FBgap as the dependent variable and %TTL as the main independent variable, including the interaction term between %TTL and total alarms. Other covariates included age, sex, race, CCI, SOFA score at CRRT initiation, use of MV at CRRT initiation, use of circulatory support devices (any of IABP, VAD, ECMO) at CRRT initiation, and %fuid overload as independent variables

Model 4: A linear regression with FBgap as the dependent variable and %TTL as the main independent variable, without including the interaction term. The covariates included the same as in Model 3. In addition, a subgroup analysis was performed for each category of total alarms (0–2, 3–5, 6+)

CI, confdence interval; FBgap, fuid balance gap; %TTL, percent treatment time loss; CCI, Charlson comorbidity index; SOFA, sequential organ failure assessment; CRRT, continuous renal replacement therapy; MV, mechanical ventilation; IABP, intra-aortic balloon pump; VAD, ventricular assist device; ECMO, extracorporeal membrane oxygenation

\* After applying the Bonferroni correction, the signifcance threshold was set to *p*<0.005, as 10 models were used

(i.e. CRRT downtime) and FBgap (i.e. the gap between achieved and goal fuid balance) in critically ill adults. Against our hypothesis, %TTL was negatively associated with FBgap in the overall cohort (i.e. higher %TTL, lower FBgap). However, given the effect modification of total number of alarms on the relationship of %TTL with FBgap, the only statistically signifcant relationship that prevailed in stratifed analyses was that higher %TTL was positively associated with FBgap in patients with  $\geq 6$ alarms per CRRT-day. This suggests that downtime has a stronger impact on fuid management in patients with frequent alarms. One possible explanation for this is that downtime unrelated to alarms are more likely to be anticipated (ex. planned filter change for off unit procedure) and later compensated for clinically. Conversely, when the number of alarms is high, signifcant treatment time loss suggests that downtime is unexpected, leading to greater deviations in fuid management.

Prior literature has described some contributing factors to CRRT downtime, such as flter clotting, catheter malfunction, procedures, transfers, and patient mobilization [[11,](#page-8-6) [14](#page-8-7), [19](#page-8-12)]. Predictable downtimes may allow for adjustments in CRRT dose and goal setting, while unexpected issues like clotting or access failure may signifcantly impact CRRT deliverables, including fuid management.

Future investigations should describe factors contributing to downtime, distinguishing between predictable and unpredictable factors. By doing so, we can aim to minimize unexpected factors and the downtime they cause, potentially leading to more optimal CRRT delivery.

The total downtime in this study was about 6.4 h over the entire CRRT course. Prior literature reports downtime in hours per day anywhere from 3.0 to 5.3 h. Compared to previous studies our %TTL was lower (8.14% vs. 22%)  $[11-14]$  $[11-14]$  $[11-14]$ . Downtime may vary significantly between study sites as diferent institutions may have diferent CRRT protocols, policies, and stafng. Furthermore, this study used data directly extracted from the CRRT machine to quantify downtime. This may have made this study more accurate in determining downtime and running time of CRRT, possibly resulting in the identifed diferences, and thus improving external validity of this study.

This study also noted %TTL was not associated with hospital mortality or process outcomes. Similar to our results, a prior retrospective observational study compared clinical outcomes between patients with<20% downtime ( $n = 88$ ) to those with > 20% downtime ( $n = 44$ ). Despite some diferences in clinical indicators, such as blood urea nitrogen and serum creatinine, they found no

significant difference in 28-day mortality  $[14]$  $[14]$ . This study also did not fnd any relationship between %TTL and process outcomes (ex. MV or ICU free days). Although the 2019 ADQI consensus recommended a goal of<10% downtime as a reasonable quality improvement measure, validation of proposed cutoffs is needed  $[10]$  $[10]$ .

This study has many strengths including a large sample size with data from 500 patients with over 3,000 CRRT patient-days. In addition, this study introduces novelty in utilizing multimodal data—both EHR and CRRT machine data. Because we used data directly extracted from the CRRT machine our results may refect higher accuracy when determining CRRT downtime. Nonetheless, this study also has limitations. First, it is a single center observational study. This study involved a population of about 90% white patients limiting extrapolation to other populations. We only accounted for select high yield clinical and machine variables, and therefore residual confounding by unmeasured variables is possible. It relied heavily on matching algorithms between EHR and CRRT machine data, which poses potential systematic errors in data mining. We attempted to overcome this by including only those patients with>90% multimodal data matches. We also performed a 10% manual validation with the goal of identifying a systematic error if one existed. Also innate to our methodology of utilizing EHR data, some important data were unable to be systematically incorporated. For example, pre-ICU fuid resuscitation data were not available. For this reason, our %FO at CRRT initiation may not accurately refect the fuid status of every patient. We also do not have data on etiology of CRRT downtime beyond alarm information. Since we were unable to obtain data on the specifc indications for CRRT, the study population includes patients that received CRRT for purposes other than fluid management. Finally, while we followed previous research by focusing on the overall FB gap during the entire CRRT period [[9\]](#page-8-4), this approach may overlook the fuctuations in daily gaps. We are considering evaluating the trajectory of FB gap in future research. Relatedly, determining fuid needs, especially in a critically ill population, are often imprecise and frequently changing. Fluid balance goals may be prone to variability and subjectivity and are typically not guided by evidence or consensus which limits the reliability of FB gap as an established CRRT metric.

Despite these limitations, this study lays the groundwork for future work which should look at better assessment of patient fuid balance and more accurate fuid balance prescriptions. Additional research should be dedicated to elucidating specifc causes of downtime and the clinical impact of each type as well as precision medicine-based approaches that decrease unnecessary alarms and downtime during CRRT. This study should also be replicated across other centers and other populations.

#### **Conclusion**

This study found the association between %TTL and FBgap was signifcant only in patients with a high number of CRRT alarms ( $\geq 6$  alarms per CRRT-day). Although this fnding suggests that downtime due to unexpected factors (i.e. treatment interruptions due to recurrent alarms) may have a stronger impact on fuid management, the mechanisms behind this efect modifcation are not fully explored, and further studies are needed to understand the clinical implications. This study did not fnd an association between %TTL and hospital mortality or other process outcomes, such as ICU-free days, CRRT-free days, or MV-free days. This lack of association raises questions about the signifcance of %TTL as a clinical rather than a quality indicator of CRRT.

#### **Abbreviations**



#### **Supplementary Information**

The online version contains supplementary material available at [https://doi.](https://doi.org/10.1186/s40560-024-00772-w) [org/10.1186/s40560-024-00772-w.](https://doi.org/10.1186/s40560-024-00772-w)

Additional file 1.

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#### **Author contributions**

CB participated in methodologic development and writing of the manuscript. TT participated in statistical analysis and writing of the manuscript. JL participated in statistical analysis. LL, SR, and SC participated in data collection and processing. AT and WB-S provided critical feedback and consultation to all aspects of the work. JN led the project including study design, data collection and management, statistical analysis and manuscript drafting. All authors have read and approved the manuscript.

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#### **Availability of data and materials**

The data sets used and/or analyzed during the current study are available from the corresponding author on reasonable request.

#### **Declarations**

#### **Ethics approval and consent to participate**

The study was approved by the Institutional Review Board (IRB) from the University of Kentucky (17-0444-P1G). Given the observational nature of this investigation, informed consent was waived by the IRB.

#### **Consent for publication**

Not applicable.

#### **Competing interests**

The authors declare that they have no competing interests.

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