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Clinical Outcomes and Adherence to Medications Measured by Claims Data in Patients With Diabetes

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Abstract

OBJECTIVE—Although poor medication adherence may contribute to inadequate diabetes control, ways to feasibly measure adherence in routine clinical practice have yet to be established. The present study was conducted to determine whether pharmacy claims–based measures of medication adherence are associated with clinical outcomes in patients with diabetes.

RESEARCH DESIGN AND METHODS—The study setting was a large, integrated delivery and financial system serving the residents of southeastern Michigan. The study population consisted of 677 randomly selected patients aged ≥ 18 years with a diagnosis of diabetes, hypercholesterolemia, and hypertension and who filled at least one prescription for either an antidiabetic, lipid-lowering, or antihypertensive drug in each of the 3 study years (1999–2001). The main outcome measures were HbA_{1c}, LDL cholesterol levels, and blood pressure.

RESULTS—Nonadherent patients had both statistically and clinically worse outcomes than adherent patients. Even after adjusting for demographic and clinical characteristics, nonadherence was significantly associated with HbA_{1c} and LDL cholesterol levels. A 10% increase in nonadherence to metformin and statins was associated with an increase of 0.14% in HbA_{1c} and an increase of 4.9 mg/dl in LDL cholesterol levels. Nonadherence to ACE inhibitors was not significantly associated with blood pressure.

CONCLUSIONS—Claims-based measures of medication adherence are associated with clinical outcomes in patients with diabetes and may therefore prove to be useful in clinical practice. More research is needed on methods to introduce claims-based adherence measurements into routine clinical practice and how to use these measurements to effectively improve adherence and health outcomes in chronic care management.

Abbreviations

CMG, continuous measure of medication gaps; SBP, systolic blood pressure

Nonadherence to medications is a common problem in clinical practice, especially among patients with asymptomatic chronic conditions such as diabetes, hypertension, and hypercholesterolemia (1–4). A recent meta-analysis has showed that the average adherence in patients with diabetes is 67.5%, which is lower than that found among many conditions (5). Also, recently, a specific systematic review on adherence to medications for diabetes showed

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A table elsewhere in this issue shows conventional and Système International (SI) units and conversion factors for many substances.

that average adherence to oral hypoglycemic agents ranged from 36 to 93% (6). In general, poor adherence to medications has been shown to be associated with the development of complications, disease progression, avoidable hospitalizations, premature disability, and death. In total, the costs associated with poor medication adherence are estimated to approach \$100 billion per year (7). Despite these known consequences, adherence rates have remained unchanged since the 1970s (1,8).

Nonadherence is the result of a complex interaction among the social environment, the patient, and the health providers (9). Adherence to medications is not routinely measured in clinical practice, and a gold standard that can be easily implemented, even for research purposes, does not exist (1,5,10–12). Yet claims data have been shown to be a useful source of adherence information with some degree of both predictive and convergent validity, and its use in routine clinical practice appears both feasible and sustainable, especially in comparison with using other methods of monitoring adherence such as electronic monitoring (13,14). Although some self-report measurements of adherence have similar validity to claims-based measurements, their sensitivity is low, especially when used by clinicians (11–13).

There is now building evidence that pharmacologically treating patients with diabetes to improve their metabolic profile (glycemia and cholesterol) and blood pressure is cost-effective (15,16). Thus, there is reason to believe that improving medication adherence might be cost-effective as well (7), especially among patients with diabetes, where the prevalence of hypercholesterolemia and hypertension is higher than in the general population (17), and their cardiovascular risk factor control is far from optimal (18).

In the present study, we examined the association between claims-based measures of medication adherence and clinical outcomes such as HbA_{1c}, blood pressure, and LDL cholesterol in a sample of diabetic patients.

RESEARCH DESIGN AND METHODS

The study was conducted in a large, salaried, multispecialty physician group practicing in the mid-western U.S. The study sample was limited to insured patients with prescription drug coverage continuously enrolled in an affiliated health maintenance organization between 1 January 1999 and 31 December 2001. Other inclusion criteria included age ≥ 18 years as of 31 December 1999; diagnosed (at least one ICD-9 code) with diabetes, hypertension, and dyslipidemia during the period of 1999–2001; and at least one prescription drug claim for either an antidiabetic, lipid-lowering, or antihypertensive drug in each of the study years (1999–2001). Patients using insulin were excluded because there is not a feasible method to measure adherence to injectables from claims data, and unmeasured adherence to insulin could confound the effects of adherence to the oral medications. Among patients meeting these criteria, a random sample of 677 patients was drawn.

Automated health plan data and medical group administrative and clinical data were used to identify patients for sample inclusion and to compile data on dates of enrollment, demographics (age, sex, and race), laboratory testing results (plasma HbA_{1c} and LDL cholesterol), and prescription drug use. Information on race/ethnicity available within this system is recorded in eight categories by the medical group front desk staff or hospital admitting staff. Data from a sample of 2,443 group model health plan patients aged 18–45 years with a visit to primary care in 2000 indicate that this race categorization is consistent with self-reported race 95% of the time ($\kappa = 0.87$) (19). BMI and blood pressure measurements were obtained by medical record abstraction by accredited record technicians who were unaware of the patients' medication adherence.

Prescription drug claims data were used to compute a continuous measure of medication gaps (CMG), a measure of nonadherence (14). This index indicates the proportion of days with gaps in medication refills over the days in the observation period. A patient was classified as nonadherent when the percentage of CMG was greater than 20%. Nonadherence was measured for three classes of drugs: metformin, 3-hydroxy-3-methylglutaryl coenzyme A reductase inhibitors (stat-ins), and ACE inhibitors. Within each drug class, CMG indexes were computed for those patients who filled at least one prescription per year ($n = 308$ for metformin, $n = 287$ for statins, and $n = 384$ for ACE inhibitors) and for the period comprised between the first prescription claim after 1 January 1999 and the last prescription available or 31 December 2001, whichever occurred first. Clinical outcomes were evaluated as the average of all measurements of HbA_{1c}, LDL cholesterol, and blood pressure available during the study period (1 January 1999 to 31 December 2001). Institutional review board approval was obtained for all aspects of this study.

Correlation between laboratory values (HbA_{1c} and LDL cholesterol) or blood pressure measurements and CMG indexes was estimated using Pearson and Spearman correlation coefficients. Multivariable linear regression models were used to adjust for sociodemographic and clinic characteristics (BMI, number of total drugs for each outcome, and number of outcome measurements available). Based on a referee suggestion, a separate model with the difference between the last and the first measurement during the study period was included for each out-come (HbA_{1c}, LDL cholesterol, and blood pressure). Those models adjusted for the first outcome measurement available during the study period as well. SPSS version 11.0 was used for all statistical analysis, and significance was set at $P < 0.05$.

RESULTS

Mean age was 64 years (11). A total of 53% of the patients were women, 53% were Caucasian, 41% African American, and 5% were of another race/ethnicity. Mean (\pm SD) outcome levels were 8.0 ± 1.4 for HbA_{1c}, 116.9 ± 30.7 for LDL cholesterol, 33.1 ± 6.8 for BMI, and 138.5 ± 12.8 and 80.0 ± 7.2 for systolic (SBP) and diastolic blood pressure, respectively. Median (mean \pm SD) measures of nonadherence to metformin, statins, and ACE inhibitors were 16.8 (21.6 \pm 18.9), 13.2 (18.4 \pm 17.7), and 7.3 (13.7 \pm 16.1), respectively. Prevalences of nonadherence were 43, 36, and 23% for metformin, statins, and ACE inhibitors, respectively (Table 1).

The correlation coefficients between CMG scores and outcomes (HbA_{1c}, LDL cholesterol, and blood pressure) ranged from 0.15 to 0.32 and were highest for LDL cholesterol and lowest for SBP. Average levels of outcomes were significantly higher in the nonadherent than adherent group (Table 2).

After adjusting for age, sex, race, BMI, total number of active drugs for each outcome (antihypertensive drugs for blood pressure, antidiabetic oral agents for HbA_{1c}, and lipid-lowering drugs for LDL cholesterol), and number of outcome measurements, nonadherence was significantly and positively associated with both HbA_{1c} and LDL cholesterol but not SBP (Table 3). The CMG score for ACE inhibitors was not significantly associated with diastolic blood pressure either (data not shown). A 10% increase in CMG percentage score (worse adherence) for metformin and statins was associated with an increase of 0.14% in HbA_{1c} and an increase of 4.9 mg/dl in LDL cholesterol, respectively. Interactions between adherence and the other significant predictors included in the linear regression models were tested, but none were statistically significant ($P > 0.10$). The proportion of outcome variability explained by the three different outcome models ranged from 17 (SBP model) to 24% (HbA_{1c} model). The LDL cholesterol model explained 22% of the variability. The separate models with the difference between last and first outcome measurement, as the dependent variable did not

change the results and nonadherence indexes for metformin and statins, were still significantly associated with clinical outcomes (Table 4).

CONCLUSIONS

The introduction of nonadherence information into clinical practice has the potential to improve both adherence and clinical outcomes. However, before introducing nonadherence information in routine clinical practice, it is critical to ensure that such information can be obtained economically and is associated with meaningful clinical outcomes. The present study provides evidence that claims-based measurements of nonadherence are associated with relevant intermediate outcomes in patients with diabetes. These associations are clinically and statistically significant for HbA_{1c} and LDL cholesterol. The association between ACE inhibitor nonadherence and blood pressure measurements was no longer significant in the multivariable models. However, nonadherence levels for ACE inhibitors were low. Only 23% of the patients on ACE inhibitors were nonadherent. Also, the average number of drugs for blood pressure among patients treated with ACE inhibitors was higher than the average number of drugs among patients treated with either metformin or statins. When the number of drugs for blood pressure was introduced in the model, the variable was strongly associated with blood pressure and the nonadherence co-efficient for ACE inhibitors was not statistically significant when compared with the one in the model that did not include the total number of drugs for blood pressure ($\beta = 0.09$, $P = 0.046$). Therefore, in this population of patients treated with ACE inhibitors and low nonadherence levels, treatment complexity was a much stronger predictor of blood pressure than nonadherence. The amount of HbA_{1c} variability explained by the multivariable linear models including CMG scores is moderate ($R^2 = 24\%$). However, it is slightly greater than the one ($R^2 \sim 10\%$) explained by other studies looking at overall HbA_{1c} predictors (20,21). Those studies included clinical and sociodemographic variables as predictors. Apart from adherence, other factors (duration and severity of disease, lifestyle, and self-management skills), which are not available from automated data sources, are important predictors of HbA_{1c}. However, even when such studies (21) have had some of these measurements available, they have only been able to explain $\sim 10\%$ of the variability. Our findings are consistent with those of Schectman, Nadkarni, and Voss (22), who evaluated the association between adherence to diabetes medications (measured using pharmacy prescription refill data) and HbA_{1c}. After adjusting for clinical and sociodemographic variables, they found that for each 10% increase in adherence, HbA_{1c} decreased by 0.16% ($P < 0.0001$). We are not aware of any other studies among patients with diabetes that look at the associations of claims-based measurements of adherence to lipid-lowering drugs and hypotensive agents with lipid levels and blood pressure, respectively. From the clinical point of view, our results highlight the potential use of claims-based measurements of nonadherence in designing and implementing interventions to improve both adherence and relevant clinical outcomes.

The present study has several limitations. First, patients taking insulin were excluded. Because of the limitations of claims data to measure adherence to insulin, we decided to measure adherence to antidiabetic oral drugs only. However, had data on insulin use been included in the HbA_{1c} model, the estimate of the association between nonadherence and HbA_{1c} could have been different, and the study's external validity would have increased. Nevertheless, the study by Schectman, Nadkarni, and Voss (22) adjusted for insulin use, and the association reported between adherence to oral agents and HbA_{1c} was almost identical to the one we are reporting. It is also noticeable that this consistency between study results holds even after considering that, while we only measured nonadherence to metformin, Schectman, Nadkarni, and Voss (22) measured and averaged adherence for all oral antidiabetic agents. Second, we were not able to control for other variables that are associated with the outcomes analyzed, such as lifestyle behaviors (diet and exercise), self-management skills (23), and other medical treatments. Some of these variables are also correlated with adherence and could have

confounded the associations observed in our study. Third, we measured both nonadherence and outcomes during the same period (1999–2001). Therefore, we cannot clearly establish the temporal sequence of the estimated associations. However, nonadherence was positively associated with changes in both HbA_{1c} and LDL cholesterol. Also, baseline adherence is the best predictor of future adherence (24), and when we used the last outcome measurement available instead of the average of all measurements available, the results were similar (data not shown).

Claims-based measurements of adherence to medications are not a perfectly valid method. Patients may obtain medications from alternative sources (via dual insurance coverage, samples, etc.) or may temporarily stop a medication because of a physician recommendation or side effects. Also, filling a prescription does not necessarily imply that the medication has been taken. On the other hand, there is no perfect method to measure adherence to medications (5) and claims-based measurements of adherence are not expensive to generate, are available in most managed care systems, and could be integrated within existing information technology (IT) clinical systems. When compared with electronic monitoring, Choo et al. (13) showed that adherence levels determined from pharmacy dispensing records correlate more closely with quantity than with timing of doses. Although not as robust as electronic monitoring, pharmacy refill rates for all antiretroviral medications were also associated with virologic response, and the highest specificity was attained when both the Medical Event Monitoring System and pharmacy refills were used in combination (25).

The consequences of making adherence information available to patients and physicians are unknown. Thus, although some have argued that it would be desirable to introduce adherence measurements into routine clinical practice and that its introduction is feasible at a sustainable cost using administrative data sources (26,27), adherence information could potentially lead to confrontation between clinicians and their patients (28,29). In this sense, especially in chronic diseases like diabetes, even the concept of adherence has been questioned and other terms like self-management have been proposed to better describe the patient-provider relationship (29). Also, recent evidence suggests that giving adherence information alone to clinicians does not improve adherence rates. Instead, adherence rates improved only when physicians received both nonadherence information and training on how to use such information (30,31).

The Chronic Care Model proposed by Bodenheimer, Wagner, and Grumbach (32,33) and the concept of productive physician-patient interactions offer a model of how adherence-to-medications data could be used by both patients and clinicians to improve health outcomes. In a patient-centered model, clinicians do not judge or overreact to patients' final decisions regarding medication adherence. Rather, clinicians try to facilitate an atmosphere where patients feel comfortable enough to report their worries, side effects, and reservations regarding prescribed medications and use empathy, positive reinforcement, and support to motivate the patients toward adherence self-management. The model also provides a framework in which other non-physician health professionals could be involved in monitoring and improving adherence and in which patients could also access the data to self-monitor adherence (self-efficacy). Also consistent with this model would be the use of informatics tools that are accessible to both patients and providers and allow patients to provide feedback data (updated information on drugs being used, remaining pills, over-the-counter drug use, side effects, etc.) to the system to further refine adherence measurements. In light of this potential, the absence of initiatives to implement outpatient computerized prescription systems in the U.S. is surprising (34). A computerized prescription system would connect clinicians IT systems to pharmacy data and would represent a clear improvement over the use of claims data. Data linkage via automated reminders would allow clinicians to monitor and improve adherence in

real time and could also contribute to medication safety monitoring in the outpatient setting (35).

Claims-based measures of adherence to medications are associated with health outcomes in patients with diabetes, thus making them relevant measures for intervention. More research is needed on methods to introduce adherence measurements into routine clinical practice and how to use these measurements to effectively improve adherence and health outcomes in chronic care management. Unless feasible adherence measurements are introduced in routine clinical practice, the status quo of low adherence levels and the unavailability of effective tools to improve medication adherence is likely to remain unchallenged.

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References

1. Dunbar-Jacob J, Mortimer-Stephens MK. Treatment adherence in chronic disease. *J Clin Epidemiol* 2001;54:S57–S60. [PubMed: 11750211]
2. Haynes RB, McDonald H, Garg AX, Mon-tague P. Interventions for helping patients to follow prescriptions for medications (Review). *Cochrane Database Syst Rev* 2002:CD000011. [PubMed: 12076376]
3. Sabaté E. Adherence to long-term therapies: evidence for action. Geneva, World Health Org. [article online]. Accessed 5 March 2004. Available from http://www.who.int/chronic_conditions/adherencereport/en/
4. Roter DL, Hall JA, Merisca R, Nordstrom B, Cretin D, Svarstad B. Effectiveness of interventions to improve patient compliance: a meta-analysis. *Med Care* 1998;36:1138–1161. [PubMed: 9708588]
5. DiMatteo MRP. Variations in patients' adherence to medical recommendations: a quantitative review of 50 years of research. *Med Care* 2004;42:200–209. [PubMed: 15076819]
6. Cramer JA. A systematic review of adherence with medications for diabetes. *Diabetes Care* 2004;27:1218–1224. [PubMed: 15111553]
7. Hughes DA, Bagust A, Haycox A, Walley T. The impact of non-compliance on the cost-effectiveness of pharmaceuticals: a review of the literature. *Health Econ* 2001;10:601–615. [PubMed: 11747044]
8. Walker EA, Usher JA. Understanding and enhancing adherence in adults with diabetes. *Curr Diab Rep* 2003;3:141–148. [PubMed: 12728640]
9. Kidd KE, Altman DG. Adherence in social context. *Control Clin Trials* 2000;21:184–187.
10. Ellis S, Shumaker S, Sieber W, Rand C. Adherence to pharmacological interventions: current trends and future directions: the Pharmacological Intervention Working Group. *Control Clin Trials* 2000;21 (Suppl 5):S218–S225.
11. Turner BJ, Hecht FM. Improving on a coin toss to predict patient adherence to medications. *Ann Intern Med* 2001;134:1004–1006. [PubMed: 11352702]
12. Vitolins MZ, Rand CS, Rapp SR, Ribisl PM, Sevick MA. Measuring adherence to behavioral and medical interventions. *Control Clin Trials* 2000;21:188–194.
13. Choo PW, Rand CS, Inui TS, Lee ML, Cain E, Cordeiro-Breault M, Canning C, Platt R. Validation of patient reports, automated pharmacy records, and pill counts with electronic monitoring of adherence to antihypertensive therapy. *Med Care* 1999;37:846–857. [PubMed: 10493464]
14. Steiner JF, Prochazka AV. The assessment of refill compliance using pharmacy records: methods, validity, and applications. *J Clin Epidemiol* 1997;50:105–116. [PubMed: 9048695]
15. Gray A, Clarke P, Farmer A, Holman R. Implementing intensive control of blood glucose concentration and blood pressure in type 2 diabetes in England: cost analysis (UKPDS 63) (Letter). *BMJ* 2002;325:860. [PubMed: 12386035]

16. CDC Diabetes Cost-effectiveness Group. Cost-effectiveness of intensive glycemic control, intensified hypertension control, and serum cholesterol level reduction for type 2 diabetes. *JAMA* 2002;287:2542–2551. [PubMed: 12020335]
17. Egede LE, Zheng D. Modifiable cardiovascular risk factors in adults with diabetes: prevalence and missed opportunities for physician counseling. *Arch Intern Med* 2002;162:427–433. [PubMed: 11863475]
18. Saydah SH, Fradkin J, Cowie CC. Poor control of risk factors for vascular disease among adults with previously diagnosed diabetes. *JAMA* 2004;291:335–342. [PubMed: 14734596]
19. Kolodner K, Lipton R, Elston Lafata J, Leotta C, Liberman J, Chee E, Moon C. Pharmacy and medical claims data identified migraine sufferers with high specificity but modest sensitivity. *J Clin Epidemiol* 2004;57:959–969.
20. Zhang Q, Safford M, Ottenweller J, Hawley G, Repke D, Burgess JF, Dhar S, Cheng H, Naito H, Pogach LM. Performance status of health care facilities changes with risk adjustment of HbA_{1c}. *Diabetes Care* 2000;23:919–927. [PubMed: 10895841]
21. Nichols GA, Hillier TA, Javor K, Brown JB. Predictors of glycemic control in insulin-using adults with type 2 diabetes. *Diabetes Care* 2000;23:273–277. [PubMed: 10868850]
22. Schectman JM, Nadkarni MM, Voss JD. The association between diabetes, metabolic control, and drug adherence in an indigent population. *Diabetes Care* 2002;25:1015–1021. [PubMed: 12032108]
23. Heisler M, Smith DM, Hayward RA, Krein SL, Kerr EA. How well do patients' assessments of their diabetes self-management correlate with actual glycemic control and receipt of recommended diabetes services? *Diabetes Care* 2003;26:738–743. [PubMed: 12610031]
24. DiMatteo MR, Sherbourne CD, Hays RD, Ordway L, Kravitz RL, McGlynn EA, Kaplan S, Rogers WH. Physicians' characteristics influence patients' adherence to medical treatment: results from the Medical Outcomes Study. *Health Psychol* 1993;12:93–102. [PubMed: 8500445]
25. Farley J, Hines S, Musk A, Ferrus S, Tepper V. Assessment of adherence to antiviral therapy in HIV-infected children using the Medication Event Monitoring System, pharmacy refill, provider assessment, caregiver self-report, and appointment keeping. *J Acquir Immune Defic Syndr* 2003;33:211–218. [PubMed: 12794557]
26. Burnier M, Brunner HR. Impact on clinical outcomes. In *Compliance in Healthcare and Research* Burke LE, Ockene IS, Eds. Armonk, New York, Futura, 2001, p.299–316
27. Applegate WB. Elderly patients' adherence to statin therapy. *JAMA* 2002;288:495–497. [PubMed: 12132982]
28. Rand CS, Sevick MA. Ethics in adherence promotion and monitoring. *Control Clin Trials* 2000;21 (Suppl 5):S241–S247.
29. Glasgow RE, Anderson RM. In diabetes care, moving from compliance to adherence is not enough: something entirely different is needed. *Diabetes Care* 1999;22:2090–2092. [PubMed: 10587854]
30. Schectman JM, Schorling JB, Nadkarni MM, Voss JD. Can prescription refill feedback to physicians improve patient adherence? *Am J Med Sci* 2004;327:19–24. [PubMed: 14722392]
31. Rosen MI, Rigsby MO, Salahi JT, Ryan CE, Cramer JA. Electronic monitoring and counseling to improve medication adherence. *Behav Res Ther* 2004;42:409–422. [PubMed: 14998735]
32. Bodenheimer T, Wagner EH, Grumbach K. Improving primary care for patients with chronic illness. *JAMA* 2002;288:1775–1779. [PubMed: 12365965]
33. Bodenheimer T, Wagner EH, Grumbach K. Improving primary care for patients with chronic illness: the chronic care model. Part 2. *JAMA* 2002;288:1909–1914. [PubMed: 12377092]
34. Schiff GD, Rucker TD. Computerized prescribing: building the electronic infrastructure for better medication usage. *JAMA* 1998;279:1024–1029. [PubMed: 9533503]
35. Richardson WC, Berwick DM, Bisgard JC. The Institute of Medicine report on medical errors (Letter). *N Engl J Med* 2000;343:663–664. [PubMed: 10979811]

Table 1
Sociodemographic, nonadherence, and clinic characteristics

Characteristics	
<i>n</i>	677
Age (years) (means ± SD)	63.9 ± 10.6
Women (%)	53.2
Race (%)	
Caucasian	53.3
African American	41.4
Other	5.3
Clinical outcomes [means ± SD (<i>n</i>)]	
HbA _{1c} (%)	8.0 ± 1.4 (675)
LDL cholesterol (mg/dl)	116.9 ± 30.7 (659)
Blood pressure (mmHg)	138.5 ± 12.8/80.0 ± 7.2 (674)
BMI (kg/m ²)	33.1 ± 6.8 (560)
Treatments during each study year [% (<i>n</i>)]	
Metformin	45.5 (308)
Statins	42.4 (287)
ACE inhibitors	56.7 (384)
Number of drugs used [means (median)]	
Oral antidiabetic drugs (among metformin users)	2.1 (2)
Lipid-lowering drugs (among statin users)	1.1 (1)
Antihypertensive drugs (among ACE inhibitor users)	2.6 (2)
Nonadherence indices (CMG percentages) [*] [median (means ± SD)]	
Metformin	16.8 (21.6 ± 18.9)
Statins	13.2 (18.4 ± 17.7)
ACE inhibitors	7.3 (13.7 ± 16.1)
Prevalence of nonadherence (%) [†]	
Metformin	43
Statins	36
ACE inhibitors	23

* CMG indicates the proportion of days with gaps in medication refills over the days in the observation period.

[†] Nonadherence is defined as CMG >20%.

Table 2
Association between nonadherence levels and outcomes

Characteristics/outcome of interest	HbA _{1c}	LDL cholesterol	SBP
<i>n</i>	308	287	384
Pearson's correlation (CMG) *	0.25 [†]	0.32 [†]	0.16 [†]
Spearman's correlation (CMG) *	0.21 [†]	0.30 [†]	0.15 [†]
Outcome levels (means ± SD)			
Nonadherent patients [‡]	8.5 ± 1.6 [§]	124 ± 34.1 [§]	143 ± 12.9 [§]
Adherent patients	8.0 ± 1.2	103 ± 28.1	138 ± 12.4

* HbA_{1c} levels were correlated with nonadherence to metformin, LDL levels with nonadherence to statins, and SBP levels with nonadherence to ACE inhibitors. CMG indicates the proportion of days with gaps in medication refills over the days in the observation period.

[†] $P < 0.01$.

[‡] Nonadherence is defined as CMG > 20%.

[§] $P < 0.01$ for differences between nonadherent and adherent patients.

Multivariable linear regression: adjusted relationship of nonadherence (CMG indices) and changes in clinical outcomes between the last and first outcome available

Table 4

Predictor	Change in HbA _{1c} model* (adjusted R ² = 49%) (n = 257)			Change in LDL model* (adjusted R ² = 43%) (n = 215)			Change in SBP model* (adjusted R ² = 43%) (n = 310)		
	β	SE (β)	P	β	SE (β)	P	β	SE (β)	P
Intercept	7.837	1.251	0.000	-2.158	24.529	0.930	84.979	12.440	0.000
CMG for metformin [†] (%)	0.015	0.005	0.006	0.317	0.110	0.004	-0.005	0.053	0.926
Age	-0.041	0.011	0.000	0.245	0.210	0.244	0.135	0.090	0.135
African American vs. Caucasian	-0.003	0.207	0.987	8.273	4.425	0.063	-4.946	1.770	0.006
Men vs. women	0.146	0.203	0.474	-1.162	4.032	0.773	-3.130	1.728	0.071
BMI	-0.011	0.016	0.485	0.353	0.343	0.304	0.071	0.126	0.573
First HbA _{1c}	-0.795	0.050	0.000	-0.578	0.050	0.000	-0.652	0.048	0.000
Total number of anti-diabetic drugs	0.299	0.150	0.047	10.935	5.172	0.036	0.089	0.753	0.906
Number of HbA _{1c} tests	0.029	0.040	0.471	0.180	0.902	0.842	-2.140	1.308	0.103

* A positive parameter estimate indicates the outcome worsened (increased) with increments in independent variable.

[†] CMG indicates the proportion of days with gaps in medication refills over the days in the observation period.