# Review

## Probiotics for pediatric antibiotic-associated diarrhea: a meta-analysis of randomized placebo-controlled trials

Bradley C. Johnston, Alison L. Supina, Sunita Vohra

### Abstract

**Background:** Antibiotic treatment is known to disturb gastrointestinal microflora, which results in a range of clinical symptoms — most notably, diarrhea. This is especially important in children, for whom antibiotics are prescribed frequently. Although meta-analyses have been conducted to evaluate the ability of probiotics to prevent antibioticinduced diarrhea in the general population, little is known about which probiotic strains and doses might be of most benefit to children. Our objective in this study was to assess the efficacy of probiotics (of any specified strain or dose) for the prevention of antibiotic-associated diarrhea in children and to assess adverse events associated with the use of probiotics when coadministered with antibiotics to children.

**Methods:** A comprehensive search was performed of the major electronic databases (e.g., CENTRAL, MEDLINE, EMBASE, CINAHL, AMED) from their inception to January 2005. We also contacted experts and searched registries and meeting abstracts for additional relevant articles. Randomized controlled trials that compared probiotic treatment with placebo or no treatment, involving pediatric subjects less than 19 years of age were included. Two reviewers independently applied eligibility criteria and assessed the studies for methodological quality. Data were independently extracted by 2 reviewers and analyzed via the standard Cochrane methodology.

**Results:** Six studies were included (total n = 707 patients). The combined results, analyzed with a per-protocol method that reported on the incidence of diarrhea during antibiotic treatment, showed significant benefit for the use of probiotics over placebo (relative risk [RR] 0.43, 95% confidence interval [CI] 0.25–0.75, I<sup>2</sup> = 70.1%). In contrast, results from intention-to-treat analysis were nonsignificant overall (RR 1.01, 95% CI 0.64–1.61). Subgroup analysis on 4 studies that provided at least 5 billion single-strain colony-forming units (CFUs) daily (range 5.5–40 × 10<sup>9</sup> *Lactobacillus GG, L. sporogens* or *Saccharomyces boulardii*) showed strong evidence with narrow CIs for the preventative effects of probiotics for antibiotic-associated diarrhea (RR 0.36, 95% CI 0.25–0.53, I<sup>2</sup> = 3.5%). No serious adverse events were reported.

**Interpretation:** The potential protective effects of probiotics to prevent antibiotic-associated diarrhea in children do not withstand intention-to-treat analysis. Before routine use is recommended, further studies (with limited losses of sub-

jects to follow-up) are merited. Trials should involve those probiotic strains and doses with the most promising evidence (i.e., *Lactobacillus GG, L. sporogens* or *S. boulardii* at  $5-40 \times 10^9$  CFUs daily).

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In 2003, the Canadian Paediatric Society released prescription information from administrative databases for over 1 000 000 pediatric claimants. Among the 20 most frequently prescribed drugs, 14 were antibiotics; 76% of children in the study population had been given a least 1 prescription for antibiotics in 1999–2000.<sup>1</sup>

Almost all antibiotic treatments may disturb the colonization resistance of gastrointestinal flora and cause a range of clinical symptoms, most notably diarrhea. Clostridium difficile diarrhea, however, which is among the most serious of the adverse events related to antibiotic-associated diarrhea (AAD), occurs most often in older, immunocompromised adults who have been admitted to hospital.<sup>2</sup> In the general population, AAD varies in incidence from 5% to 62%, and in timing, from at the initiation of therapy to as long as 2 months after the end of treatment.<sup>2–4</sup> Among children who receive broad-spectrum antibiotics, reported incidences of diarrhea vary from 11% to 62%.<sup>4–6</sup> Although the World Health Organization (WHO) defines AAD as 3 or more abnormally loose bowel movements in a 24-hour period, definitions used in pediatric and adult trials have varied from 1 to 3 abnormally loose stools per 24–48 hours.<sup>7,8</sup>

The term probiotic refers to a product or preparation containing viable, defined microorganisms in numbers thought to be sufficient to alter the host's microflora (by implantation or colonization) and thereby exert beneficial effects.9 The rationale behind probiotic administration is based on reinoculation with specific probiotic strains to normalize unbalanced indigenous microflora. Two meta-analyses of trials of probiotics and prevention of AAD in the general population have been completed, with results in favour of probiotic coadministration with antibiotics (odds ratio<sup>8</sup> 0.37, 95% confidence interval [CI] 0.26-0.53; risk ratio10 [RR] 0.40, 95% CI 0.28-0.57); a third review, however, found insufficient evidence for the routine clinical use of probiotics to prevent or treat C. difficile-associated diarrhea in adults.<sup>11</sup> Safety does not appear to be a concern in healthy individuals, although serious infections (e.g., bacteremia, meningitis, endocarditis, pneumonia and deep abdominal abscesses) have been reported in neonates and in severely debilitated and immunocompromised patients.<sup>12–20</sup>

We conducted a systematic review of randomized controlled trials of probiotics (of any specified strain or dose) coadministered with antibiotics (any agent) to assess their safety and efficacy in the prevention of pediatric AAD.

#### Methods

A comprehensive search was conducted for studies of probiotics, irrespective of publication status or language. The databases searched included ISI's Web of Science (1945– 2005), MEDLINE (1966–2005), EMBASE (1980–2005), CIN-AHL (1982–2005), AMED (1985–2005), the Cochrane Central Register of Controlled Trials (CENTRAL, 2005, issue 1) and the Cochrane Complementary Medicine Field's Register of Controlled Trials. Conference proceedings and dissertation abstracts were searched through Ovid's Dissertation Abstracts (1980–2005), the Conference Papers Index (1982– 2005) and the OCLC Online Computer Library Center (1992–2005). We also searched the "grey" literature (i.e., studies not published or with limited distribution), including, among other venues, databases housed by the Chalmers Research Group (PedCAM), CISCOM (Centralised In-





formation Service for Complementary Medicine) and the Cochrane Inflammatory Bowel Disease Review Group; bibliographies of randomized controlled trials and review articles; American Gastroenterological Association meeting abstracts; and ongoing trials, located through the Current Controlled Trial Register. Inclusion criteria for our review were that the study be a randomized controlled trial; involve pediatric subjects (o–18 years of age); use any dose of a specified probiotic of any strain, administered in conjunction with any antibiotic and compared with placebo; and report incidence outcomes for diarrhea (as defined by the study's authors) or adverse events.

Searches were first screened by 2 independent reviewers (B.C.J., A.L.S.) who looked at titles of papers and, when available, abstracts. The full text of each article selected was then retrieved and independently assessed by each of the 2 reviewers for inclusion according to prespecified selection criteria and for methodological quality. Each included study was evaluated with the (previously validated) 5-point Jadad scale<sup>21</sup> to assess its methods of randomization and double-blinding and the numbers of withdrawals and dropouts. Concealment of allocation was assessed as adequate, inadequate or unclear according to trial-design methodology described by Schulz and coauthors.<sup>22</sup> Interrater reliability was assessed for both quality scales by using K statistics; disagreements were re-

solved by consensus.

Using a standardized data extraction form, the same 2 reviewers independently extracted data items, including methodological design, definition and diagnostic criteria for diarrhea, and inclusion and exclusion criteria for participants.

#### Results

From the primary electronic databases, we identified a total of 520 studies (EMBASE 277, MEDLINE 156, CENTRAL 38, Web of Science 32, CINAHL 17, AMED 0). Thirty additional citations were identified from the grey literature (Fig. 1). Using prespecified inclusion criteria, the independent reviewers identified 17 trials for full review. Upon full text review, 6 relevant studies (5 in English,<sup>7,23-26</sup> 1 in Italian<sup>4</sup>) were included. Appendix 1 presents details on the 11 studies excluded.<sup>27-38</sup> Interrater agreement for inclusion and exclusion was very good (estimated  $\kappa$  0.70).<sup>27</sup> The 6 included studies enrolled a total of 836 participants (408 for treatment and 428 as control subjects, which totalled 707 after losses to followup). In each trial, participants were administered probiotics at the onset of antibiotic treatment; daily probiotic treatment was continued until the completion of the antibiotic course (7-14 days). All 6 studies provided data on the incidence of pediatric diarrhea. A detailed summary of the studies included can be found in Table 1.

The overall pooled results with random-effects per-protocol analysis showed that the use of probiotics produced a significant reduction in the incidence of AAD (Fig. 2). The number needed to treat (NNT) in order to prevent 1 case of diarrhea was found to be 6 patients (NNT 5.74, 95% CI 4.36–13.08). Statistical heterogeneity, however, was moderate ( $I^2 = 70\%$ ;  $I^2$ represents the percentage of variability due to between-study variability).<sup>39</sup> A sensitivity analysis that incorporated extremecase intention to treat (ITT)<sup>40</sup> showed nonsignificant results overall (Fig. 3).

None of the 6 studies reported a definition of what constituted an adverse event. Three<sup>23,24,26</sup> of the 4 studies that did monitor for adverse events reported that none occured. The fourth<sup>7</sup> reported 14 minor adverse events (including rash, gas, vomiting, increased phlegm and chest pain) but not the group (treatment or control) in which they occurred.

Jadad quality scores for the 6 studies are shown in Table 1. The mean score of 3.3 indicated good quality overall,<sup>21</sup> whereas results for concealment of treatment allocation were mixed.<sup>22</sup> Since there was full agreement between the 2 reviewers for both the Jadad and Shulz quality scales,  $\kappa$  statistics were not calculated.

Two *a priori* subgroups were analyzed with respect to the primary outcome, incidence of diarrhea: probiotic strain and probiotic dose. Two of the 6 trials administered identical probiotic strains — *Lactobacillus GG (L. casei* spp *rhamnos-us)*.<sup>25,26</sup> The summary statistics from these 2 studies (n = 307) provided statistical significance for a protective effect (RR = 0.29) of the *Lactobacillus GG* strain (Fig. 4).

The dose-based subgroup analyses compared primary outcomes for daily doses of  $5 \times 10^9$  CFUs and above with those for

lesser doses. The 4 studies that provided children with 5.5–40 × 10° bacteria or yeast cells per day showed evidence for the preventative effects of probiotics (RR 0.36, 95% CI 0.25– 0.53),<sup>4,24–26</sup> whereas those for the single study that provided a substantially lower daily dose,  $2 \times 10^{\circ}$  probiotic bacteria,<sup>7</sup> were nonsignificant (Fig. 5). A  $\chi^2$  test for potential dose-related heterogeneity (4 studies providing  $\geq 5 \times 10^{\circ}$  CFUs/d v. 1 study providing <  $5 \times 10^{\circ}$  CFUs/d) was significant (P = 0.002) and may explain the moderate statistical heterogeneity ( $I^2 = 70\%$ ) observed for the the incidence of diarrhea<sup>41</sup> (Fig. 2, Fig. 5).

#### Interpretation

This systematic review of probiotic trials, which was focused on pediatric patients with AAD, incorporated 3 trials<sup>4,23,24</sup> not included in previous meta-analyses of probiotics for AAD in the general population.<sup>8,10</sup> Yet another review evaluated probiotics for the prevention and treatment of *C. difficile*–associated diarrhea in adults.<sup>11</sup> Of the 6 pediatric trials we included, 2 assayed stool samples for *C. difficile*.<sup>24,25</sup> Since 26%–50% of AAD may be attributed to *C. difficile*,<sup>2</sup> future trials should better address the efficacy of probiotics in preventing pediatric AAD caused by enteropathogens, particularly *C. difficile*.

Per-protocol pooled results of the 6 studies we reviewed yielded favourable results for the efficacy of probiotics (RR 0.43, 95% CI 0.25–0.75); treatment of 6 patients (NNT 5.7, 95% CI 4.4–13.1) should prevent 1 case of diarrhea. Our findings are consistent with those of 2 earlier meta-analyses involving the general population: one by D'Souza and colleagues,<sup>8</sup> who used a per-protocol analysis (odds ratio 0.37, 95% CI 0.26–0.53), and another by Cremonini and co-authors<sup>10</sup> (RR 0.40, 95% CI 0.27–0.57), who excluded trials with losses to follow-up of 15% or more.

Table 1: Characteristics of studies included in the meta-analysis						
Study	No of participants (loss to FU); age range*	Probiotic (antibiotic) interventions with dosages, if reported	Definition of diarrhea for purposes of incidence	Study quality: Jadad score† (allocation concealment)		
Arvola 1999 <sup>25</sup>	167 (48) 2 wk-12 yr	<i>Lactobacillus GG</i> , 20-40 billion CFU/d for 7-14 d (antibiotics not specified)	$\geq$ 3 watery or loose stools per day for $\geq$ 2 consecutive days	2 (unclear)		
Jirapinyo 2002 <sup>23</sup>	18 (0) 1-36 mo	L. acidophilus, Bifidobacterium infantis 7 d of probiotic treatment, dose not stated (cefprozil, ampicillin, gentamycin, cloxacillin)	Not reported	2 (unclear)		
Kotowski 2005 <sup>24</sup>	269 (23) 5 mo-15 yr	Saccharomyces boulardii 10 billion CFU/d for 7-9 days (a mixture of antibiotics)	$\geq$ 3 loose or watery stools daily for $\geq$ 48 h, during or $\leq$ 2 wk after therapy	5 (adequate)		
LaRosa 2003⁴	120 (10) Mean age 6.6 yr	<i>L. sporogens</i> 5.5 billion CFU/d for 10 d (amoxicillin, amoxicillin + clavulanate, cephalosporin, erythromycin)	≥ 2 liquid stools over 24 h during study period	4 (unclear)		
Tankanow 1990 <sup>7</sup>	60 (22) 5 mo-6 yr	L. acidophilus, L. bulgaricus 3 billion CFU/d for 10 days (amoxicillin)	≥ 1 abnormally loose stool daily throughout entire study period	3 (adequate)		
Vanderhoof 1999 <sup>26</sup>	202 (14) 4-12 yr	Lactobacillus GG for 7-10 d, 10 billion for children < 12 kg, 20 billion for $\ge$ 12 kg (amoxicillin, cefprozil, clarithomycin)	$\geq$ 2 liquid stools per day, twice or more during study period	4 (adequate)		

Note: FU = follow-up, CFU = colony-forming units.

\*Unless otherwise specified.

+Score < 3 indicates a study of poor quality; 5 indicates a study of maximal quality.

We also analyzed incidence results for AAD with a more conservative extreme-case ITT analysis, the results of which were nonsignificant. A sensitivity analysis suggested that 2 studies with extreme losses to follow-up  $(37\%^7 \text{ and } 29\%^{25})$ accounted for why results did not withstand ITT analysis. In 1 study,25 these losses may have been related to feasibility issues; a 3-month follow-up with families, for example, was difficult to accomplish (Dr. Taina Arvola, personal communication, 2005). By acceptance that noncompliance and protocol deviations are likely to occur in real clinical situations, ITT analysis preserves the integrity of randomization, enhances the external validity of the results and (by controlling for the removal of noncompliers from the analysis) dissuades overoptimistic statements of the efficacy of an intervention.<sup>42</sup> For this reason, it is the preferred analytic approach for pragmatic trials of effectiveness.43 However, ITT analysis is not robust if rates of loss to follow-up are high. Since our review involved the meta-analysis of explanatory rather than pragmatic trials, and since 3 trials<sup>4,7,25</sup> of the 6 examined had high losses to fol-

Study	Study gro	oup, <i>n/N</i> Control	Risk ratio (RR), random, with 95% confidence interval (CI)	Weight %	RR, random (95% CI)
Tankanow Arvola Vanderhoof Jirapinyo LaRosa Kotowska	10/15 3/61 7/93 3/8 14/48 4/119	16/23 9/58 25/95 8/10 31/50 22/127		22.12 11.09 16.98 14.73 21.50 13.58	0.96 (0.61-1.50) 0.32 (0.09-1.11) 0.29 (0.13-0.63) 0.47 (0.18-1.21) 0.47 (0.29-0.77) 0.19 (0.07-0.55)
Total events $\chi^2 = 16.72 \ (p z score 2.99)$	41/344 = 0.005), I (p = 0.003)	111/363 <sup>2</sup> = 70.1%	0.1 0.2 0.5 1 2 Favours treatment Favour	100.00 5 s placebo	0.43 (0.25-0.75) ר 10

**Fig 2:** Per-protocol analysis of pediatric incidence of antibiotic-associated diarrhea. Although the difference in overall effect (i.e., the z score) was statistically significant, so too was the heterogeneity ( $\chi^2$ , here with 5 degrees of freedom. Note also that higher values of I<sup>2</sup> indicate greater heterogeneity).



**Fig 3:** Incidence of antibiotic-associated diarrhea — intention-to-treat analysis. The analysis showed a nonsignificant difference between probiotics and placebo (z score) and statistically significant heterogeneity.

low-up (18%–37%), the validity of ITT analysis in this review can be questioned. We chose instead to focus on the per-protocol analysis, as other authors<sup>8,11</sup> in this field have done.

The statistical heterogeneity we found was notable: moderate (70%) in the per-protocol analysis (Fig. 2) and large (78%) in the ITT analysis (Fig. 3). Nonetheless, the quantification of heterogeneity is one component of a wider investigation of variability across studies, perhaps the most important being diversity in clinical trials, which we explored via subgroup analysis.<sup>39</sup> Our subgroup analyses (based on perprotocol analysis) explored the potential reasons for statistical heterogeneity, including dose and strain.

No dose-ranging studies have been reported to determine the minimal effective dose of a probiotic in preventing AAD. *A priori*, because we wanted to separate trials of subtherapeutic doses from the others, we chose for subgroup analysis daily doses of 5 billion CFUs and above versus lesser doses. Although dosage recommendations on the labels of probiotic products available in Canadian health food stores varies from

> I billion to 40 billion CFUs per day, doses at the lower end of this range might not colonize the intestine.<sup>44</sup> Since clinical effectiveness has been demonstrated at 5.5 billion CFUs daily, we chose 5 billion CFUs per day as our cut-off.<sup>4</sup>

> Our subgroup analyses provided preliminary evidence that probiotic dose may be responsible for the clinical and thus statistical heterogeneity observed ( $I^2 = 76\%$ ). A perprotocol subgroup analysis of 4 studies that provided  $\geq$  5 billion CFUs daily to study participants  $(5.5-40 \times 10^9 \text{ single-strain probiotic})$ CFUs per day of *Lactobacillus GG*, L. sporogens or Saccharomyces *boulardii*) showed strong evidence for the preventative effects of probiotics (RR = 0.36, 95% CI 0.25-0.53,  $I^2 = 3.5\%$ ). Our  $\chi^2$  test for potential dose-related heterogeneity (Fig. 5) showed significance and may explain the moderate statistical heterogeneity observed (Fig. 2). When included trials that used the same probiotic intervention underwent subgroup analysis for probiotic strain (Fig. 4), the 2 trials of Lactobacillus GG (10–40  $\times$  10<sup>9</sup> CFUs/d) demonstrated almost no statistical heterogeneity with pooled results, which significantly favoured probiotics (RR 0.29, 95% CI 0.15-0.57).<sup>25,26</sup> Since heterogeneity dropped substantially when data on identical strains (Lactobacillus GG. Fig. 4) and on similar doses were

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combined (Fig. 5), we speculated that the differences in probiotic strain and dose were plausible reasons for the observed heterogeneity of the pooled results.

With respect to safety, no trials reported a serious adverse event. Our meta-analysis found no significant differences in the incidence of any adverse events between treatment and control groups, although no trial defined *a priori* what was considered to be an adverse event/reaction.

Two additional issues are worthy of further consideration during review of pediatric trials of AAD: the effect of age and the definition of AAD. Whereas the 2 meta-analyses done previously included both children and adults, with the majority of included trials involving adults, our review was restricted to the pediatric population; the included patients ranged in age from 2 weeks to 15 vears.<sup>8,10</sup> The difficulty in accurately measuring frequency or consistency of diarrhea in diapered infants versus in children and adolescents may make it difficult to detect differences. For example, infants not only have bowel movements more frequently,45 but their stools may also be looser in consistency than those of older children, which makes infant stools more difficult to distinguish from diarrhea. Furthermore, it can be problematic to detect the frequency and consistency of stools accurately in diapered infants. Although randomized trials may evenly distribute infants between treatment and control groups so as to eliminate this potential confounder, the considerable differences in bowel habits of children at various ages may obscure the effect of probiotics in this population, and should be taken into consideration when planning pediatric trials.46

Second, the definitions of AAD in included studies varied, including "1 or more abnormally loose bowel movements per day throughout the study period,"<sup>7</sup> as opposed to "at least 3 watery or loose stools per day for at least 2 consecutive days."<sup>25</sup> One trial<sup>23</sup> did not provide a definition at all. A survey of clinicians and

Subcategory Study group, n/N and study Treatment Control	Risk ratio (RR), random, with 95% confidence interval (CI)	Weight %	RR, random (95% CI)
Lactobacillus GGArvola $3/61$ $9/58$ Vanderhoof $7/93$ $25/95$ Subtotal $10/154$ $34/153$ Heterogeneity: $\chi^2 = 0.02$ ( $p = 0.89$ Overall effect:z score 3.59 ( $p < 0$	), I <sup>2</sup> = 0 .001)	11.09 16.98 28.07	0.32 (0.09-1.11) 0.29 (0.13-0.63) 0.29 (0.15-0.57)
<i>L. acidophilus</i> and <i>L. bulgaricus</i> Tankanow 10/15 16/23 z score 0.19 (p = 0.85)	-	22.12	0.96 (0.61-1.50)
<i>L. acidophilus</i> and <i>Bifidobacteriu</i> Jirapinyo 3/8 8/10 z score 1.57 (p = 0.12)	n infantis	14.73	0.47 (0.18-1.21)
<i>L. sporgens</i> and fructo-oligosaccha LaRosa 14/48 31/50 z score 3.01 (p = 0.003)	arides	21.50	0.47 (0.29-0.77)
<b>Saccharomyces boulardii</b> Kotowska 4/119 22/127 z score 3.10 (p = 0.002)		13.58	0.19 (0.07-0.55)
Total events 41/344 111/363 $\chi^2$ = 16.72 ( <i>p</i> = 0.005), I <sup>2</sup> = 70.1% z score 2.99 ( <i>p</i> = 0.003)	0.1 0.2 0.5 1 2 Favours treatment Favour	100.00 5 rs placebo	0.43 (0.25-0.75) - 10

**Fig 4:** Incidence of antibiotic-associated diarrhea, by the probiotic strain administered. Both the overall heterogeneity  $(I^2)$  and the difference between probiotics and placebo (z score) were statistically significant. For the combined results of 2 trials of *Lactobacillus GG*, the heterogeneity was nonsignificant, with a significant difference in overall effect compared with placebo.

Subcategory and study	Study gro Treatment	up, <i>n/N</i> Control	Risk ratio (RR), ra 95% confidence i	andom, with interval (CI)	Weight %	RR, random (95% CI)
≥ 5 billion ba	acteria					
Arvola	3/61	9/58	← ■		13.85	0.32 (0.09-1.11)
Vanderhoof	7/93	25/95		_	20.09	0.29 (0.13-0.63)
LaRosa	14/48	31/50			24.46	0.47 (0.29-0.77)
Kotowska	4/119	22/127	• • • • • • • • • • • • • • • • • • •	-	16.57	0.19 (0.07-0.55)
Subtotal	28/321	87/330			74.96	0.36 (0.25-0.53)
< 5 billion ba Tankanow z score 0.19	acteria 10/15 (p = 0.85)	16/23		+	25.04	0.96 (0.61-1.50)
Total events $\chi^2 = 16.89 (p z score 2.63)$	38/336 = 0.002), I <sup>2</sup> (p = 0.008)	103/353 <sup>2</sup> = 76.3%			100.00	0.42 (0.22-0.80)
			0.1 0.2 Favours treat	0.5 1 2 tment Favour	5 rs placebo	10

**Fig 5:** Incidence of antibiotic-associated diarrhea, by probiotic dose. For the 4 studies administering daily doses of at least  $5 \times 10^{9}$  colony-forming units (CFUs), the overall effect differed significantly from that of placebo, with minimal heterogeneity. When all data were pooled, however, heterogeneity also became statistically significant.

parents may be helpful to document what is considered a clinically meaningful reduction in AAD and the point at which parents and clinicians would be likely to consider the coadministration of a probiotic.

This systematic review has several strengths. Compared with previous reviews of probiotics for AAD in the general population, our search strategy was more comprehensive. Although we had too few trials reporting on the incidence of pediatric diarrhea (n = 6) to properly analyze for publication bias,<sup>47</sup> several strategies were implemented to control this. We conducted explicit searches in multiple databases, and considered any trial irrespective of publication format (e.g., abstracts) or language. We also conducted *a priori* subgroup (probiotic strain, dose) and sensitivity analyses (per-protocol v. ITT) to further explore our hypotheses, and looked for heterogeneity as a result of pooling.

Our review also had a number of limitations. We did not search the European Gastroenterology Week conference proceedings, a potential home for relevant trials; nor did we include trials that assessed probiotics versus standard conventional care for AAD (Appendix 1).<sup>31,32</sup> Although these effectiveness trials may have provided additional relevant data, we were interested in comparing the efficacy of probiotics with placebo via explanatory trials.<sup>48</sup> Finally, some readers may question the pooling of different probiotic strains. In keeping with our justification for including the combinations of probiotic strains that were used in 2 trials included in this review (Tankanow and associates<sup>7</sup> administered both L. acidophilus and L. bulgaricus; Jirapinyo and coworkers23 administered L. acidophilus with Bifidobacterium infantis), we pooled our overall results because the probiotics used share the recommended characteristics of a viable probiotic agent: nonpathogenic properties, the ability to survive transit through the gastrointestinal tract, adherence to intestinal epithelium, colonization in the intestinal tract, production of antimicrobial substances, and a good shelf life in food or powdered form.49

In sum, the potential protective effects of probiotics to prevent pediatric AAD did not withstand ITT analysis. Future studies should engage probiotic strains and doses with the most promising evidence (e.g., Lactobacillus GG, L. sporogens or S. boulardii at  $5-40 \times 10^{9}$  CFUs/d); determine the effect of age (i.e., infants v. older children) and ensure that adverse events are reported and losses to follow-up are limited. Trials would also benefit from a validated primary outcome measure for AAD that is sensitive to change and reflects what stool frequencies and consistencies have importance in the eyes of clinicians, parents and children. The current data are promising, but inconclusive; it is premature to draw a conclusion about the efficacy and safety of probiotics for pediatric AAD until such trials are completed. For clinicians contemplating the use of probiotics in their pediatric practice, the use of Lactobacillus GG, L. sporogens or S. boulardii at a dose of 5–40 billion CFUs per day appears to hold promise as an option for coadministration with antibiotics, and no serious adverse events were reported in the included studies. Nevertheless, safety is better assessed in population-based samples than in the relatively small numbers exposed in randomized controlled trials.

This article has been peer reviewed.

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**Competing interests:** Bradley Johnston has received study product from Culturelle Inc., the manufacturer of *Lactobacillus GG*, for evaluation in pediatric patients with irritable bowel syndrome.

**Contributors:** Bradley Johnston wrote the study protocol, managed and participated in screening and quality assessment of and data extraction from articles, completed the statistical analyses, interpretated results, and drafted the manuscript. Alison Supina assisted with screening, data extraction and quality assessment. Sunita Vohra assisted with design of the methodology, statistical analyses and interpretation of results. All coauthors critically revised the article and approved the version to be published.

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*Correspondence to:* Dr. Sunita Vohra, Director, CARE Program, Stollery Children's Hospital, Associate Professor of Pediatrics, Department of Pediatrics, University of Alberta, Aberhart Centre #1, Rm. 8213, 11402 University Ave. NW, Edmonton AB T6G 2J3; 780 407-3798; fax 780 407-2105; svohra@ualberta.ca

Appendix 1: Studies excluded after full-text review					
Reason for exclusion	Study				
Nonpediatric study population	Thomas 2001 <sup>34</sup> Witsell 1995 <sup>35</sup> Siitonen 1990 <sup>36</sup>				
Study groups not randomly assigned	Seki 2003 <sup>30</sup> Chapoy 1985 <sup>28</sup>				
<ul><li>Probiotic compared to other than placebo:</li><li>A nutritional drink</li><li>Diosmectite, an antidiarrheal drug</li></ul>	Schrezenmeir 2004 <sup>32</sup> Benhamou 1999 <sup>31</sup>				
Acute diarrhea not associated with use of antibiotic therapy	Erdeve 2004 <sup>29</sup> Michielutti 1996 <sup>33</sup>				
Outcomes not particular to incidence of diarrhea or adverse events	Zoppi 2001 <sup>37</sup> Contardi 1991 <sup>38</sup>				

#### LEADERSHIP

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