

# Maternal Cigarette Smoking and Oral Clefts: A Population-Based Study

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**Abstract:** Analyses of 1984 data from the Maryland Birth Defects Reporting and Information System indicate that mothers of infants with oral clefts (cleft lip with or without cleft palate; and cleft palate) smoked more during pregnancy than mothers of infants with other defects (odds ratio OR of 2.56 and 2.39, respectively). There was a dose-response relation between the daily amount smoked and the risk of clefting. Adjustment for available confounding variables did not account for the association between smoking and oral clefts. (*Am J Public Health* 1987; 77:623-625.)

## Introduction

Although it is generally accepted that cigarette smoking during pregnancy increases the risk of intrauterine growth retardation, prematurity, perinatal mortality, and spontaneous abortion,<sup>1-3</sup> controversy still exists as to whether or not cigarette smoking increases the risk of congenital malformations.<sup>1</sup> While several studies have reported positive associations between smoking and specific malformations,<sup>4-8</sup> particularly oral clefts,<sup>9-11</sup> others did not find such relationships.<sup>12-21</sup>

In this study, we examined the association between cigarette smoking and oral clefts using data from the Maryland Birth Defects Reporting and Information System (BDRIS). We discuss the importance of grouping malformations into homogeneous categories<sup>22-25</sup> in etiologic studies.

## Methods

The Maryland BDRIS is a birth defects monitoring program that has been in operation since September 1983. Details on the system can be found elsewhere.<sup>26</sup> Briefly, BDRIS ascertains infants (live births or fetal deaths more than 500 g or 20 weeks) born with any one of 12 sentinel defects: anencephaly, spina bifida, hydrocephalus, cleft lip with or without cleft palate (CL), cleft palate (CP), esophageal atresia, anal/rectal atresia, hypospadias, reduction and other deformities of upper limbs, reduction and other deformities of lower limbs, congenital dislocation of the hip, and Down syndrome. Cases are ascertained at birth and reported by hospitals on a special form which includes demographics, obstetric variables, and information concerning prenatal illnesses and exposures. Smoking history during pregnancy is usually obtained by an obstetric nurse directly from the mother. Two questions are asked: 1) Did you smoke at any

time during pregnancy?; and 2) if yes, how many cigarettes a day? (1-5, 6-10, 11-20, >20). No questions are asked about pre-pregnancy smoking or about smoking status by trimester of pregnancy. Details on the completeness of reporting of the system and descriptive epidemiology of sentinel defects for 1984 are reported elsewhere.<sup>26</sup>

We investigated the 1984 completed BDRIS data. Clefts were divided into CL and CP, and by whether the clefting is the only abnormality or whether it exists with other malformations. The frequency of maternal smoking was compared between the cleft groups and a control group that consisted of all other defects combined, except for Down syndrome. This latter group was excluded because of the reported negative association between smoking and Down syndrome.<sup>27</sup> Although other defects are presented in aggregate, the frequency of smoking was examined among each defect group and was found to be homogeneous. A dose-response effect of smoking on the risk of CL and CP was evaluated using data on the reported amount of smoking (grouped into 0, 1-10, 11-20, >20 cigarettes per day). A linear trend in the effect of smoking on the odds of clefting was tested using the weighted regression model described by Rothman.<sup>28</sup> In further analyses, to examine whether the association between smoking and oral clefts is confounded by demographic variables (such as race, sex, residence, maternal age, parity, etc., or by prenatal events such as drug intake, environmental exposures, and illnesses), stratified analysis was used to examine the effects of smoking on the risk of oral clefts within different levels of other variables (such as among Whites and Blacks separately). Adjusted odds ratios were obtained using the Mantel-Haenzel procedure.<sup>29</sup>

## Results

During 1984, the Maryland BDRIS ascertained 28 cases of cleft lip with or without cleft palate (CL) and 26 cases of cleft palate (CP) for rates of 5.0 and 4.6 per 10,000 live births, respectively. In addition, there were 209 cases of other sentinel defects (excluding Down syndrome). Only mothers with known smoking status during pregnancy are included in the analysis (CL: 27/28, CP: 26/26, other defects: 198/209).

As shown in Table 1, cigarette smoking during pregnancy was associated with clefting (odds ratios of 2.56 and 2.39 for CL and CP, respectively). Table 2 shows that, with

**TABLE 1—Maternal Cigarette Smoking during Pregnancy and Oral Clefts, Maryland Birth Defects Reporting and Information System, 1984**

	Cleft lip ± cleft palate (N = 27)	Cleft palate (N = 26)	Other defects* (N = 198)
No. Smokers	15	14	65
Odds Ratio	2.56	2.39	1.0
95% Confidence intervals	1.13-5.78	1.04-5.45	—

\*Other defects include: anencephaly, spina bifida, hydrocephalus, esophageal atresia, anal/rectal atresia, limb reduction deformities, hypospadias, congenital dislocation of the hip. Mothers with unknown smoking status are excluded.

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**TABLE 2—Odds Ratios for the Effect of Maternal Cigarette Smoking on Oral Clefts, by Amount Smoked, Maryland Birth Defects Reporting and Information System, 1984**

Amount Smoked	Cleft lip ± Cleft palate (N = 25)			Cleft palate (N = 25)		
	N	Crude Odds Ratio (95% CI)	Predicted* Odds Ratio (95% CI)	N	Crude Odds Ratio (95% CI)	Predicted* Odds Ratio (95% CI)
0	12	1.0	1.0	12	1.0	1.0
1-10	4	1.17 (.4-3.8)	1.68 (.8-3.3)	7	2.04 (.7-5.6)	2.00 (1.1-3.7)
11-20	6	4.16 (1.4-12.6)	3.04 (1.0-9.5)	4	2.77 (.8-9.6)	4.00 (1.6-9.9)
21+	3	4.75 (1.1-20.8)	4.50 (1.2-16.3)	2	3.17 (.6-17.0)	6.01 (2.2-16.5)

\*Predicted odds ratios are obtained by fitting a weighted regression line to case-control odds as described by Rothman.<sup>29</sup>  
 For CL: odds ratio = 1 + .14 X (X = number cigarettes smoked)  
 For CP: odds ratio = 1 + .20 X (X = number cigarettes smoked)

**TABLE 3—Odds Ratio for the Effects of Maternal Cigarette Smoking on Oral Clefts, by the Presence of Associated Malformations, Maryland Birth Defects Reporting and Information System, 1984**

Associated Malformations	Cleft lip ± Cleft palate (N = 27)			Cleft palate (N = 26)		
	N	Odds Ratio	95% CI	N	Odds Ratio	95% CI
None	21	3.33	1.3-8.4	17	1.82	.67- 4.9
Multiple	6	1.02	.2-5.7	9	4.09	.99-16.9

increasing amounts of smoking, there is a more pronounced association between smoking and clefts for both CL and CP.

When oral clefts were divided into those that occur isolated (about two-thirds of the cases) and those that occur in conjunction with other defects (Table 3), the association between smoking and CL holds only for the isolated CL but not the multiple defect group (odds ratios of 3.33 and 1.02, respectively). For CP, there is a tendency for the reverse to occur (but not statistically significant).

As shown in Table 4, adjustment for a series of potential confounding demographics and perceived exposure variables did not appreciably change the magnitude of association between smoking and clefts. Other available variables such as maternal epilepsy, anticonvulsant use, diabetes, cancer, and emotional disturbances occurred too infrequently to be used for adjustment. Because of the small number of cases in this series, these variables were adjusted one at a time using the Mantel-Haenzel procedure, rather than in a multivariate model.

**Discussion**

This study has some limitations. Aside from the small sample size, these include:

- the inability to refine the timing of smoking during pregnancy into exposure that occurred during the vulnerable embryological time in the first trimester. Since this misclassification problem is nondifferential among infants with clefts and other defects, its net effect will be to bias the association towards the null;
- underreporting of clefts to the Maryland BDRIS. As shown elsewhere,<sup>26</sup> about 20 per cent of sentinel defects were not reported to the BDRIS in 1984. This underreport-

**TABLE 4—Odds Ratios for the Effect of Maternal Cigarette Smoking on Oral Clefts, Adjusted for Potential Confounding Variables, Maryland Birth Defects Reporting and Information System, 1984**

Variables	Cleft lip ± Cleft palate (N = 27)		Cleft palate (N = 26)	
	Odds Ratio	95% CI	Odds Ratio	95% CI
Race				
White/Black	2.58	1.1-6.3	2.23	.5-9.4
Sex				
Males/Females	2.55	1.1-6.1	2.67	.7-8.8
Residence				
Baltimore/Other	2.52	1.0-6.2	2.33	.6-8.5
Maternal age (years)				
>30/≤30	2.58	1.0-6.6	2.27	.6-8.0
Parity				
≤2/>2	2.53	1.0-6.2	2.15	.6-8.3
Perceived maternal exposures				
No/Yes	2.44	1.0-6.0	2.28	.6-9.3
Use of birth control				
No/Yes	2.44	1.0-6.0	2.56	1.6-11.0
Drug use				
No/Yes	2.40	.8-6.5	2.30	.6-9.3
URI during pregnancy				
No/Yes	2.63	1.1-6.4	2.23	.6-8.3
Nausea/vomiting				
No/Yes	2.68	1.1-6.4	1.87	.5-7.0

Mothers with unknown status for a variable were excluded.

ing is due, in part, to the loss of cases to District of Columbia area hospitals, but there is no reason to believe it is related to cigarette smoking during pregnancy;

- the possibility of unmeasured confounding variables (for example, a factor not routinely collected via BDRIS is maternal alcohol consumption during pregnancy); and
- the absence of a normal control group.

Although there is much debate in the literature about the appropriate control group in birth defects studies,<sup>30</sup> the use of abnormal controls can be defended here on the basis of two arguments: First, the frequency of smoking among mothers of infants with other defects is similar to the frequency of smoking among females of childbearing age in general, and is within the range of reported smoking frequency among pregnant women in particular.<sup>1</sup> Second, oral clefts stand out as the most consistently found association with smoking,<sup>9-11</sup> and is supported by our inability to find meaningful differ-

ences in the frequency of smoking among mothers of infants with other defects when broken down into specific types of defects. Nevertheless, these findings should be viewed as conservative in the sense that if smoking increases the risk of some other defects, the odds ratios measured here will tend toward unity and against finding an association.

Many previous studies of smoking and birth defects have not refined cases into embryologically and pathogenetically homogeneous categories. Since it is highly unlikely that any putative teratogen can cause an increased risk of all malformations, a negative association between smoking and all defects combined<sup>12,16,19</sup> is not illuminating. Furthermore, because of cumulating evidence from epidemiologic,<sup>23,31</sup> family,<sup>24,32</sup> and embryologic studies<sup>22,25</sup> that the same clinical phenotype (such as cleft palate) may be related to different etiologic and embryologic mechanisms, it is not revealing to report negative associations between smoking and defects of major organ systems (such as cardiovascular musculoskeletal).<sup>16,17</sup> Even studies of specific defects such as oral clefts do not usually separate cases into cleft lip and cleft palate, and into isolated cleft and those with multiple abnormalities<sup>9,18</sup> despite the epidemiologic heterogeneity between the two groups of clefts.<sup>31</sup> The presence of associated defects may point to a chromosomal or Meadlian etiology.<sup>23,32</sup>

If smoking is related to isolated CL but not CL associated with other anomalies, as suggested here, lumping all clefts together for analysis, or with major organ structures (such as craniofacial), may dilute the magnitude of the association between smoking and the specific defect group.

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