

RESEARCH PAPER

Current smoking and the risk of non-fatal myocardial infarction in the WHO MONICA Project populations

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Background: Cohort studies have shown that smoking has a substantial influence on coronary heart disease mortality in young people. Population based data on non-fatal events have been sparse, however.

Objective: To study the impact of smoking on the risk of non-fatal acute myocardial infarction (MI) in young middle age people.

Methods: From 1985 to 1994 all non-fatal MI events in the age group 35–64 were registered in men and women in the WHO MONICA (multinational monitoring of trends and determinants in cardiovascular disease) project populations (18 762 events in men and 4047 in women from 32 populations from 21 countries). In the same populations and age groups 65 741 men and 66 717 women participated in the surveys of risk factors (overall response rate 72%). The relative risk of non-fatal MI for current smokers was compared with non-smokers, by sex and five year age group.

Results: The prevalence of smoking in people aged 35–39 years who experienced non-fatal MI events was 81% in men and 77% in women. It declined with increasing age to 45% in men aged 60–64 years and 36% in women, respectively. In the 35–39 years age group the relative risk of non-fatal MI for smokers was 4.9 (95% confidence interval (CI) 3.9 to 6.1) in men and 5.3 (95% CI 3.2 to 8.7) in women, and the population attributable fractions were 65% and 55%, respectively.

Conclusions: During the study period more than half of the non-fatal MIs occurring in young middle age people can be attributed to smoking.

Although young people may acknowledge the well documented fact^{1,2} that cigarette smoking is a major risk factor for coronary heart disease (CHD) they (and even physicians) may think that this is only a concern in older age. But is this really true?

Several studies have shown acute deleterious effects of smoking on blood clotting, coronary vasomotor tone, and coronary endothelial function.^{3–6} Thus, in current smokers acute myocardial infarction (MI) may occur due to mechanisms other than coronary atherosclerosis and may suddenly affect a previously healthy heart. Also, there may be an interaction between smoking and hereditary coagulation defects such as the Leiden V mutation.⁷ These factors may explain why several studies have found a high prevalence of current smoking in young MI victims.^{8,9} There are, however, sparse population based data on smoking and MI in young people.

The World Health Organization established the MONICA (multinational monitoring of trends and determinants in cardiovascular disease) project to investigate trends and determinants of cardiovascular disease.¹⁰ From the mid 1980s to the mid 1990s, coronary events and acute coronary care were monitored in 32 populations in 21 countries for the age range 35–64 years. At least two CHD risk factor surveys were conducted on random samples of people in these populations during the same time period. To study the impact of smoking on the risk of MI in young people, we have used the MONICA survey data on smoking and contrasted them with the prevalence of smoking in MI patients in the MONICA populations.

METHODS

The methods of the MONICA project and main results have been published.^{11–14} Details of the methods, including

population descriptions and the MONICA manual, are also available at the MONICA WWW pages (<http://www.ktl.fi/publications/monica/index.html>). In short, the project involved monitoring coronary events, acute coronary care, and CHD risk factors over about 10 years among men and women aged 35–64 years who were residents of well defined geographical areas.

Study populations

Due to limited availability of data, the detailed composition of the study populations has varied for different MONICA publications.^{11–14} In this paper we have used the same 31 populations as in the paper by Tunstall-Pedoe and colleagues¹⁴ with an additional population that did not have data spanning enough years for trend analyses to be included in that paper. The 32 study populations were mainly in Europe, but there was one in China, two in Australia, one in New Zealand, and two in North America. Descriptions of the populations and the data collection periods are shown in table 1.

Data collection

All suspected coronary events in the study populations were monitored continuously from mid 1980s to mid 1990s. Non-fatal events and deaths were classified by standard diagnostic criteria as definite MI, possible MI, possible coronary death, unclassifiable death, or not MI.¹¹

Acute coronary care (ACC) data, including smoking status, were collected as a separate data component, and linked to all

Abbreviations: ACC, acute coronary care; CHD, coronary heart disease; CI, confidence interval; ISIS, international studies of infarct survival; MI, myocardial infarction; MONICA, multinational monitoring of trends and determinants in cardiovascular disease; WHO, World Health Organization

Table 1 MONICA populations and periods of data collection for each population

Country	Population	Mean population aged 35–64 years (1000s)	Periods of the risk factor surveys	Comparison years for non-fatal definite acute myocardial infarctions
Australia	Newcastle	148	5/83–12/83; 6/88–11/89; 6/94–12/94	85; 7/88–12/89; 93
	Perth	380	5/83–11/83; 6/89–12/89; 5/94–11/94	84; 89; 93
Belgium	Ghent/Charleroi	162	2/85–7/87; 4/90–4/92	9/86–7/87; 5/91–3/92
Canada	Halifax	108	9/85–11/88; 5/95–11/95	85–88; 93
China	Beijing	288	9/84–11/85; 9/88–10/89; 9/93–10/93	84–85; 88–89; 93
Czech Republic	Czech Republic	175	3/85–11/85; 3/92–12/92	1/86–5/87; 1/91–11/91
Denmark	Glostrup	134	8/86–4/87; 2/91–3/92	87; 91
Finland	FINMONICA	241	1/87–4/87; 1/92–3/92	9/86–12/86; 9/92–12/92
France	Lille	340	6/86–2/89; 6/95–11/96	10/86–6/87; 89; 94
	Strasbourg	32	1/85–8/87; 3/95–4/97	85; 93
	Toulouse	322	5/85–2/87; 10/88–5/91; 12/94–7/96	86; 7/89–12/89; 90–91 (odd months); 92; 7/93–12/93
Germany	Augsburg	217	10/84–5/85; 10/89–6/90; 10/94–7/95	85; 89–90; 94
	Bremen	217	5/84–11/84; 5/88–11/88; 5/91–1/92	85; 88; 91
	East Germany	46	1/88–11/88; 9/93–12/94	89; 93
	Rhein-Neckar	233	9/83–7/87	85; 88
Iceland	Iceland	75	6/83–11/83; 6/93–4/94	1/82–9/83; 91–92
Italy	Area Brianza	345	4/86–3/87; 5/89–7/90; 9/93–11/94	6/86–1/87; 89–90; 93–94
	Friuli	377	1/86–9/86; 3/89–12/89; 3/94–10/94	86; 89; 93
Lithuania	Kaunas	150	12/86–6/87; 2/92–5/93	87; 92
New Zealand	Auckland	299	1/93–3/94	4/91–9/91
Poland	Tarnobrzeg	199	6/83–11/84; 5/87–11/88; 6/92–7/93	1/86–5/86; 8/87–12/88; 92–93
	Voivodship			
	Warsaw	206	12/83–1/85; 1/88–1/89; 1/93–12/93	86; 88; 93
Russia	Moscow	312	2/84–10/86; 2/88–1/89; 1/92–3/95	86; 89; 92–93
	Novosibirsk	232	5/85–3/86; 5/88–4/89; 5/94–6/95	86; 88–89; 93
Spain	Catalonia	388	4/86–7/88; 10/90–5/92; 6/94–5/96	4/86–12/87; 90–92; 94
Sweden	Gothenburg	153	2/85–11/86; 2/90–5/91	4/86–6/87; 91
	Northern Sweden	194	1/86–4/86; 1/90–4/90; 1/94–4/94	11/86–6/87; 90; 94
Switzerland*	Switzerland	407	10/84–5/86; 10/88–4/90; 10/92–6/93	86; 90; 7/92–12/93
UK	Belfast	157	10/83–9/84; 9/86–12/87; 10/91–12/92	85; 88; 91–92
	Glasgow	131	2/86–7/86; 1/92–9/92; 2/95–10/95	86; 92; 94
USA	Stanford	98	5/79–5/80; 5/85–6/86; 6/89–6/90	81; 1/85–8/86; 89–90
Yugoslavia	Novi Sad	115	9/88–4/89; 9/94–2/95	88–89; 94–95

* Men only.

coronary event records for most of the study period. Several centres collected ACC data continuously from the beginning to the end of the study. Others collected ACC data for a fixed period at the beginning of the study but after 1989 almost all centres collected these data continuously. Full details of the methods used in the event and acute coronary care registration are given in the MONICA manual and in the quality assessment reports (Quality assessment of coronary event registration data in the WHO MONICA Project. (January 1999). URL: www.ktl.fi/publications/monica/coreqa/coreqa.htm, URN: NBN:fi-fe 19991072; Quality assessment of acute coronary care data in the WHO MONICA Project. URL: www.ktl.fi/publications/monica/accqa/accqa.htm, URN: NBN:fi-fe 19991081).

In all study populations two risk factor surveys were carried out near the beginning and at the end of the period of event monitoring. In most populations there was a third survey in the middle of the study period. Random samples of the general population (independent of coronary events) were selected from the best available sampling frames and were generally stratified by sex and 10 year age groups. Self reported cigarette smoking was obtained from participants in the risk factor surveys using a standard questionnaire on history of smoking.

Quality assurance

Quality assurance was an integral part of the WHO MONICA project. The internal consistency of the data was checked continuously at the MONICA data centre. In case of inconsistencies the centres were asked to verify the data. Test case series were circulated periodically, and quality assessment reports were produced periodically. The quality

assessment reports are available at the MONICA WWW pages.

Data analyses

Smoking status was obtained for all patients who experienced non-fatal MI events, and smoking status in the referent population was obtained from the population surveys making this a case based study.¹⁵ Records with missing information on smoking were excluded from the analyses.

Only non-fatal MIs were considered in the analyses since data on smoking status were often missing for fatal events. Smoking status for each person was coded as current smoker or non-smoker. Current smoking was defined as any tobacco smoking on a daily basis within three months of the onset of the event. Data on past smoking were not collected in the ACC records.

From the survey data the smoking status of each participant was categorised as current daily cigarette smoker or non-smoker. In the survey questionnaire items about pipe and cigar smoking did not specifically ask about daily smoking but were included in the definition of a smoker to achieve comparability with the ACC data. It should be noted therefore that the definition of current smoker for survey participants was slightly different to the definition used for cases, which was any daily smoking in the past three months. It is possible therefore that some people who would have been classified as current smokers under the case definition were included as non-smokers among the controls. However, in the MONICA database we were able to ascertain whether subjects in the control population had quit smoking in the previous six months, and only 1% of the population had. It is reasonable to assume that only half of these quit in the three

Table 2 Risk of a non-fatal definite AMI for current smokers compared with non-smokers, prevalence of smoking in the general population, and the population attributable risk of smoking (age standardised and aggregated over 32 WHO MONICA populations)

Sex	Age group (years)	Non-fatal definite MI					Population			Population attributable risk (%)	p Value*	Relative risk (95% CI)
		Smokers (n)	Non-smokers (n)	Unknown (n)	Smokers (n)	Non-smokers (n)	Unknown (n)	Smokers (n)	Non-smokers (n)			
Men	35-39	572	138	35	5217	5865	15	64.7	0.195	4.9 (3.9 to 6.1)		
	40-44	1146	356	81	4766	6038	17	59.3	<0.001	4.3 (3.5 to 5.3)		
	45-49	1774	776	119	4884	6373	18	48.8	<0.001	3.2 (2.7 to 3.7)		
	50-54	2072	1282	212	4430	6646	18	40.5	<0.001	2.7 (2.4 to 3.1)		
	55-59	2459	2078	269	4248	7108	23	27.2	<0.001	2.0 (1.8 to 2.3)		
Women	60-64	2285	2747	361	3365	6695	15	19.0	0.033	1.7 (1.6 to 1.9)		
	35-39	58	17	7	3174	8195	12	54.6	0.981	5.3 (3.2 to 8.7)		
	40-44	109	61	10	2788	8341	13	48.8	0.392	4.8 (3.3 to 6.9)		
	45-49	182	116	28	2350	9169	19	47.9	0.754	5.5 (4.2 to 7.1)		
	50-54	307	266	47	1957	9287	17	39.8	0.449	4.8 (4.0 to 5.9)		
55-59	439	559	102	1622	9382	12	27.7	0.209	3.6 (3.0 to 4.3)			
60-64	578	1015	146	1419	8937	23	22.3	0.169	3.1 (2.6 to 3.6)			

*Q statistic for test of heterogeneity (see methods).

months before the survey and therefore it is unlikely that the difference definitions led to any significant increase in the estimates of relative risk.

The period in which information on smoking status was collected for coronary events varied between centres, but most of the 32 centres involved in this analysis contributed data from the middle of the 1980s to the middle of the 1990s.

Statistical analysis

Data were analysed separately for men and women and for each five year age group from 35-39 years to 60-64 years. The age standardised prevalence of smoking in each centre was estimated by the direct method using the world standard population weights of 6, 6, 6, 5, 4 and 4 for the five year age groups from 35-39 years to 60-64 years.¹⁶

The relative risk of MI for current smokers compared with non-smokers was estimated using the random effects model proposed by DerSimonian and Laird,¹⁷ and tests of heterogeneity were conducted using the Q statistic with the mean value obtained by the Mantel-Haenszel method.

The proportion of non-fatal definite MIs that occurred as a result of exposure to smoking was estimated using the attributable fraction (AF) as described in Rothman and Greenland¹⁸: $AF = [P \times (RR - 1)] / [P \times (RR - 1) + 1]$, where P = prevalence of smoking in the general population, and RR = relative risk of non-fatal definite MI in the population.

We examined the issue of missing data among cases and non-response rates among controls using graphical techniques and sensitivity analyses. Scatter plots and correlation coefficients were used to test for an association between the percentage of missing data among cases and prevalence of smoking among cases as well as the association between non-response rates and the prevalence of smoking among controls. Forest plots were used to look for an association between centre specific estimates of relative risk and potential sources of bias such as non-response rates, missing data, and prevalence of smoking. Sensitivity analyses were conducted by assuming all cases with missing data were smokers and secondly by assuming all cases with missing data were non-smokers. Similarly sensitivity analyses were conducted assuming that all non-responders to the population surveys were current smokers and then assuming they were non-smokers.

RESULTS

In the risk factor surveys the proportion of missing data on smoking was only 0-1%; however, response rates varied considerably (range 50-90%). In the ACC data for non-fatal definite MIs, information on smoking was fairly complete: the median amount of missing data for men was 3.5% (range 1-36%), for women 7.5% (range 0-36%). Centres in Poland, Russia, Denmark, and the Czech Republic had high levels of missing information on smoking (15-36%) for people with MI, but excluding these centres from the analysis made little difference to the results.

The age standardised prevalence of current smoking (excluding records with missing information on smoking) from the risk factor surveys varied from 18% (New Zealand-Auckland) to 65% (China-Beijing) in men and from 4% (Russia-Novosibirsk and Lithuania-Kaunas) to 44% (Denmark-Glostrup and Scotland-Glasgow) in women. The age standardised prevalence of current smoking in people with non-fatal MIs, obtained from the ACC data (excluding records with missing information on smoking), varied from 52% (Lithuania-Kaunas) to 84% (China-Beijing) in men and from 4% (Russia-Novosibirsk) to 85% (Denmark-Glostrup) in women. People who had a non-fatal MI were far more likely to be current smokers than people in the general community (figs 1 and 2).

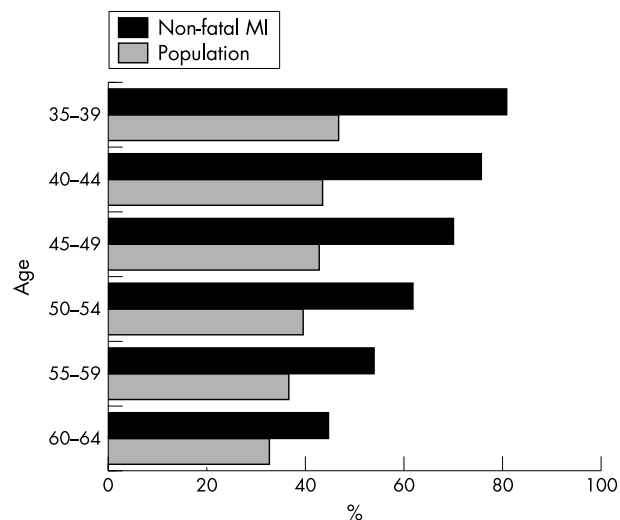


Figure 1 Prevalence of smoking in MONICA populations in non-fatal MI events and in the population. Men aged 35–64 years. Records with missing data on smoking were excluded.

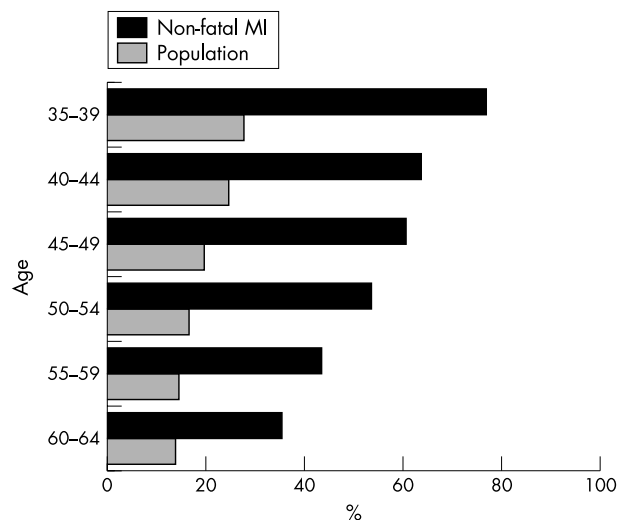


Figure 2 Prevalence of smoking in MONICA populations in non-fatal MI events and in the population. Women aged 35–64 years. Records with missing data on smoking were excluded.

The risk of a non-fatal MI in current smokers was notably higher than the risk in non-smokers. In men, the relative risk ranged from 4.9 (95% confidence interval (CI) 3.9 to 6.1) in the age group 35–39 years to 1.7 (95% CI 1.6 to 1.9) in the age group 60–64 years, and in women from 5.3 (95% CI 3.2 to 8.7) to 3.1 (95% CI 2.6 to 3.6), correspondingly (table 2). In the age group 35–39 years, approximately 65% of non-fatal MIs among men and 55% among women could be attributed to smoking (table 2).

The sensitivity analysis produced predictable results. When all cases with missing data were assumed to be current smokers the relative risk estimates increase by approximately 0.3 for men and 0.4 for women but maintaining the same gradient across age groups. Similarly, the same gradient was maintained when all non-responders to the population surveys were assumed to be current smokers, although the impact on the estimates of relative risk was greater. For the youngest to oldest age groups, the relative risk estimates for men were 3.0, 2.6, 2.0, 1.6, 1.2, and 1.0, respectively. Except for the estimate in the oldest age group all estimates were significantly different from unity. For women the estimates of relative risk were 2.9, 2.4, 2.6, 2.0, 1.3, and 1.0 and again all but the 1.0 were significant.

There was significant heterogeneity in the risk estimates within most age and sex groups. However, forest plots showed that non-response rates, percentage of missing data, and the prevalence of smoking were not associated with the estimates of relative risk.

DISCUSSION

The data from this large population based study confirmed the high prevalence of smoking in young MI patients (about 80% of men and women aged 35–39 years were smokers), and the very substantially increased risk of MI in young smokers, compared with non-smokers. In men and women aged 35–39 years, the risk of MI for those who smoked was five times higher than the risk for those who did not smoke. About 50% of MIs in men and women younger than 50 years were attributable to smoking and therefore potentially preventable. These findings for non-fatal definite MI events are in line with the observations on non-fatal events in the large case–control study in which the survivors of ISIS-3 and ISIS-4 studies (international studies of infarct survival) were cases.⁹ That study showed that the risk of non-fatal MI was

five times higher in smokers than non-smokers in the age group 30–49 years. In the British male doctors cohort study reported by Doll and Peto,¹⁹ the risk of CHD death in heavy smokers was 15 times higher than in non-smokers in the age group < 45 years and three times higher in the age group 45–54 years during 20 years follow up.

The data on MI occurring below the age 35 were sparse. The risk in the young (below 35 years) could therefore not be assessed reliably.

In this study, it was not possible to obtain complete data on smoking for people who died, and therefore only data for cases with non-fatal MI were used. This might lead to spurious association if smokers were protected from dying, as several clinical trials that include only hospitalised cases have indicated (for example, Barbash *et al*²⁰). However, sudden death is more common in smokers, compared with non-smokers,²¹ and population based studies show that the risk of death during an acute event (including out-of-hospital and hospital deaths) is similar for smokers and non-smokers.^{22–23} Thus, it is reasonable to assume that the risk of non-fatal events associated with smoking is representative of the risk of all events.

We explored only the risk of current smoking, since it was not possible to distinguish between never and past smokers among people who had a non-fatal MI. Since the risk of MI decreases quite rapidly after quitting smoking,²⁴ the bias caused by this misclassification is likely to be small, especially in the young age group. In the British doctors' cohort total mortality was essentially the same in those who stopped smoking before the age of 35 as in non-smokers, but not subsequently.² This indicates that the current study may underestimate the risk in older age groups. On the other hand, survey non-respondents may more often be smokers, causing an underestimation of smoking prevalence in the population and thus inflating the risk estimates. However, in our sensitivity analysis the gradient of risk was maintained and even if all non-responders were current smokers, which is extremely unlikely, the risk of MI among smokers was still statistically significant and biologically important in all but the oldest age group.

The relative risks for women were higher than for men in this study, particularly for the older age groups. One possible interpretation of this finding is that women are more sensitive to smoking than men, an issue of recent

What this paper adds

Several case-control studies and case series have revealed a high prevalence of smoking in young myocardial infarction (MI) patients. Large scale, population based studies examining the true impact of current smoking on the occurrence of MI in young age have, however, not been available thus far.

This large study documents the very high prevalence of current smoking in young MI patients. It also documents that current smoking is the single most important factor behind an acute MI in young patients. Current smoking is the sole or substantial contributory factor in at least 60% of acute MIs in the age group 35–39 years. Thus, every effort should be put to use to make young people realise the true and imminent risks of smoking to prevent them to start smoking, and to help those already caught to quit smoking.

controversy.²⁵ Another possible explanation, which is more consistent with the age related trend in relative risk for men, is that in the absence of high levels of other risk factors for coronary disease smoking is the main determinant of risk.²⁶

It is possible that the interaction between smoking and other factors—for example, hereditary coagulation defects⁷—may be particularly important in the occurrence of MI in young people. Many of these other factors may not be easily treatable. Indeed, smoking cessation programmes will probably be the mainstay of prevention available for such patients.

In conclusion, our data indicate that 50% of non-fatal MIs in men and women younger than 50 years (even more in younger age groups) would be preventable if smoking cessation programmes were successful. Smoking cessation would prevent excess mortality from a host of other diseases as well.² There is a continuing need for public health programmes and anti-smoking campaigns targeted at young people to keep them healthy, and specifically from our results, also to prevent the particular tragedy of heart attack at a young age.

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*See appendix for sites and key personnel of the WHO MONICA Project

APPENDIX

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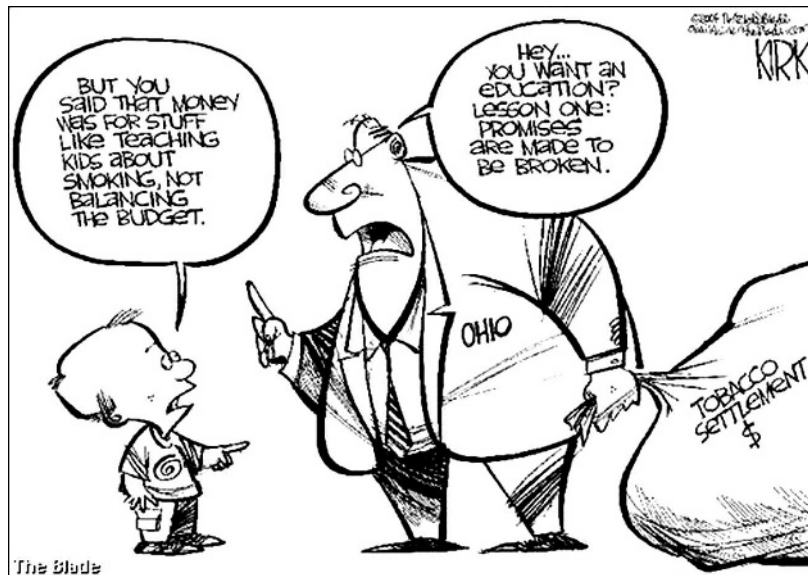
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The lighter side



Ohio is one of dozens of US states to have US MSA (Master Settlement Agreement) funds for general revenues rather than for tobacco control measures.

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