

# Planning and monitoring of placebo-controlled survival trials: comparison of the triangular test with usual interim analyses methods

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**Aims** For ethical and economic reasons, interim analysis of phase III clinical trials is essential. This study was conducted to compare the efficiency of two interim analysis procedures, which could be used to allow early termination of a clinical trial.

**Methods** We made a *post hoc* application of two interim analysis methods (Lan & DeMets with O'Brien–Flemming modification, and the triangular test according to Whitehead) by using individual patient data from four published placebo-controlled survival trials. We determined the date the trial would have been stopped had each method been used, and we estimated consequent results in terms of events and patient numbers included in the trial, the duration of the trial, and on treatment effect.

**Results** The triangular test provided the lowest number of events required to reach a conclusion of the trials while providing an accurate estimate of experimental treatment effects. The triangular test thus indicated the smallest number of patients that would have been enrolled, and the shortest trial duration. The difference between the methods was most important with a detrimental effect of experimental treatment: the number of required events was reduced by 75% and the trial duration was shortened by 48% with the triangular test compared to the Lan & DeMets method.

**Conclusions** Stopping a trial early must depend on the clinical context. It is most important to stop a placebo-controlled trial as soon as possible when the experimental treatment can be shown deleterious. In such a situation the triangular test is more appropriate than the Lan & DeMets method. When a treatment effect is no different from, or better than, placebo the triangular test is also superior but the importance of premature termination of the trial in such cases has to be balanced against the inevitable reduction of information that the trial can provide.

**Keywords:** Controlled clinical trials, clinical trials data monitoring committees, trial design, interim analysis, sequential methods, triangular test.

## Introduction

Phase III clinical trials evaluate the benefit of a new treatment. To plan a clinical study, which will achieve this with a specified power, several parameters such as the rate of events in the control group and the expected benefit from the experimental treatment need to be esti-

mated [1]. Because of the approximate nature of such estimates in the trial design it is necessary that data are subjected to frequent analysis to ensure the validity and the safety of an on-going trial [2]. Such monitoring provides opportunities to re-evaluate assumptions for the sample size [3], and also gives an early estimate of the experimental treatment effect. Interim findings could lead to an early termination of the trial and thus prevent unnecessary exposure of patients to inferior treatment, or avoid delays in identifying a superior treatment with a benefit greater than expected. Thus several recent trials have been terminated prematurely for either benefit (HOPE study [4]), or harm of an experimental therapy (ALLHAT

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study [5]). However, because of the increase of the false-positive  $\alpha$  error rate induced by repeated analyses of accumulating data [6], interim results must be interpreted using a specific stopping rule [7]. Several procedures for interim analysis have been proposed [8–11] but methods of monitoring a trial are not widely agreed. In large-scale clinical trials interim safety and efficacy evaluation is usually based on the ‘ $\alpha$ -spending function’ as described by Lan & DeMets [10] with a conventional conservative boundary type. This approach applies the rate of spending the  $\alpha$  error during the trial until its planned end. The ‘stopping rules’ are then based on the level of observed  $P$ -value at the time of interim analysis.

With the triangular test [11] several assumptions, including the  $\alpha$ -error rate, are used to define the rules that will guide the monitoring of the trial. Although some phase III clinical trials [12, 13] had been planned using the triangular test, this method remains widely under-used. However, mathematical simulations [14, 15] have shown that the triangular test seems to be a highly effective method of reducing the sample size and duration of a trial. How effective the triangular test is in the real conditions for clinic trial remains unclear.

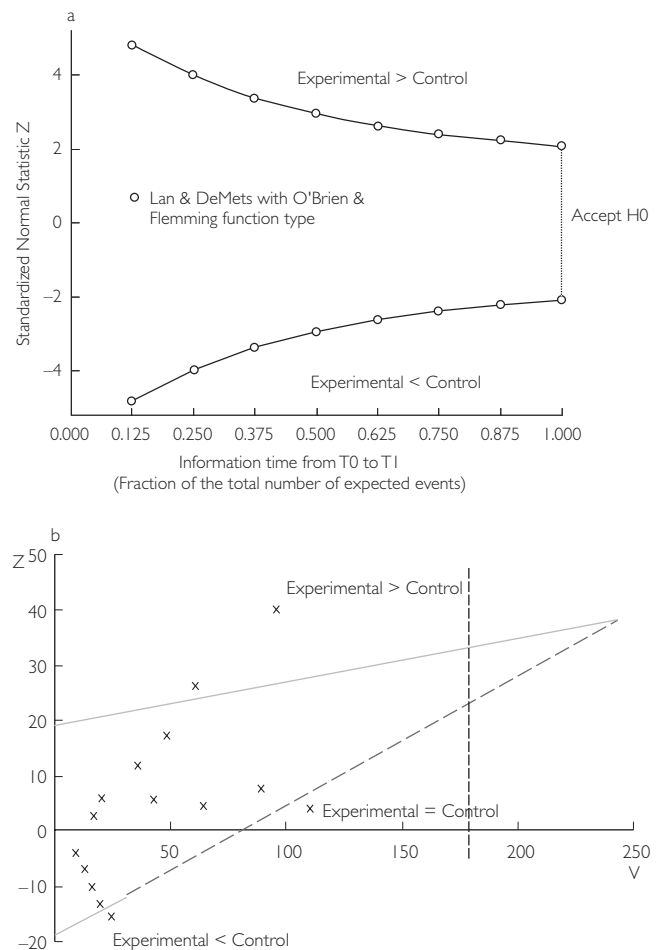
Therefore in order to evaluate the potential value of the triangular test and the planning and monitoring of phase III placebo-controlled survival trials in comparison with other methods, we compared the triangular test and the Lan & DeMets method retrospectively to four published trials using their original data.

**Methods**

*Sequential procedures*

*Repeated significance testing approach according to Lan & DeMets.* In this approach, a nominal significance level  $\alpha' < \alpha$  is used at each analysis in order to maintain the predetermined overall significance level  $\alpha$ . Various adjustment strategies have been proposed. Lan and DeMets developed the ‘ $\alpha$ -spending function’ [10], which defines a continuous increasing function characterizing the rate the  $\alpha$  error is ‘spent’ according to the progress of the trial (Figure 1a). This progress is defined according to the number of expected events at the time of trial design. The calculations of nominal significance levels depend upon the predetermined two-sided overall  $\alpha$  error rate, the total number of analyses, and the boundary shape parameter, which defines the likelihood of stopping at the early stage as opposed to the late stages of the trial. In this study, the boundary shape parameter is set to zero corresponding to the popular and validated O’Brien and Fleming modifications.

*Boundaries approach: the triangular test.* The triangular test, developed by Whitehead from initial work of Wald and



**Figure 1** (a) Two-sided graphical representation of the critical values according to the  $\alpha$ -spending function with O’Brien–Flemming-type. The global  $\alpha$  is fixed to 0.05 and 8 analyses were planned. (b) Triangular design with the three possible conclusions according to the crossing boundary. The vertical dotted line represents the information equivalent to the fixed sample design. It plays no role in the procedure and is only reported here for comparison.

Anderson, uses a sequential plan defined by two perpendicular axes corresponding to two several statistics  $Z$  and  $V$ . The statistic  $Z$  (vertical axis) summarizes the current difference between the experimental treatment and control. The statistic  $V$  (horizontal axis) represents the quantity of information accumulated since the beginning of the trial [11]. It reflects both the cumulated number of patients and the number of observed events, and it will thus increase as the study progresses.

Two straight lines intersect and define an asymmetric triangular area with two external stopping boundaries. At each interim analysis, the two statistics  $Z$  and  $V$  are calculated from the accumulated data and define a point reported on the sequential plan. The consecutive points plot a sample path going from the left to the right. The study is continued as long as the sample path stays within the two boundaries and is stopped when one of the

boundaries is crossed. The conclusion of the study depends on, which boundary is crossed (Figure 1b): experimental better than control for the upper boundary, experimental nondifferent or inferior to control for the lower boundary. The equations of each boundary are specific for a trial and depend on prespecified  $\alpha$  (two-sided) and  $\beta$  error rates, on the expected benefit of the experimental treatment over control and on the expected rate of events in the control group of the study. Furthermore, after each interim analysis, boundaries are adjusted thus defining a continuation region with a shape of a Christmas tree. Such adjustment maintains the  $\alpha$  error risk at its prespecified level when study starts.

### Reanalysed clinical trials

Four published phase III clinical trials were reanalysed in this study. Table 1 shows their main characteristics: their duration (calculated between the first inclusion and the final analysis, which lead to stopping the trial), and the final results on the difference between the experimental treatment effect and control. All these trials are recent multicentre randomized, double-blind and placebo-controlled trials. For all of them, all-cause mortality was the primary end-point. It was expressed as survival time (censored criterion).

These trials represent a sample of the different conclusions that a phase III clinical trial may reach. One trial (CIBIS II [16]) demonstrated a benefit of experimental treatment, one (EMIAT [17]) showed no significant difference between treatments, one (PRIME II [18]) was stopped because of a deleterious effect of the experimental treatment and one (CIBIS I [19]) was underpowered.

For each trial, the initial hypothesis on the two-sided  $\alpha$  error rate and the power ( $1-\beta$ ), the expected benefit of the experimental treatment over control and the expected rate of events in the control group were observed from the protocol (Table 2). Using this information, the sequential design according to the two studied methods was rebuilt for each trial.

*Interim analyses planning.* For each method, interim analyses were performed according to the total number of expected deaths in a trial. With the triangular test, analyses started after 10% of the number of expected deaths (minimum 40 events) then each 5% with a minimum of

20 events between each analyses. With the Lan & DeMets method, analyses were performed after 25%, 37.5%, 50%, 62.5%, 75%, 87.5% and 100% according to Lan & DeMets recommendations [20].

*Re-analysis of trials.* The individual data for each trial were set on SAS (SAS Institute, Cary, USA) or EXCEL (Microsoft Corp.) files and included for each patient the date of inclusion, the treatment code and the event code with its date if occurred. Thus, the real progress of each trial was rebuilt and the two interim analyses procedures were applied. For each method, the date the trial would have been stopped was then determined.

### Statistical analysis

Statistical analyses were computed using the SAS v6.12 software. The monitoring and the sequential analyses using the triangular test were performed with the Planning and Evaluation of Sequential Trials (PEST 4.0) statistical software [21]. The Lan & DeMets method was performed with the EAST software [22].

Differences in survival between treatments were assessed using the log-rank test. Hazard ratios and 95%CI were calculated with Cox's proportional hazards regression model at the date of stopping decision provided by each methods. For the triangular test, PEST 4.0 software provides corrected hazard ratios taking into account the sequential nature of analyses [11].

## Results

Figure 2 shows the design and the monitoring of each trial with the triangular test. The number of events and patients enrolled at the stopping time for each method are reported in Table 3. The triangular test provides the shortest sizes and durations for all the re-analysed trials except for the underpowered CIBIS I study.

### *Trial showing benefit of the experimental treatment: CIBIS II*

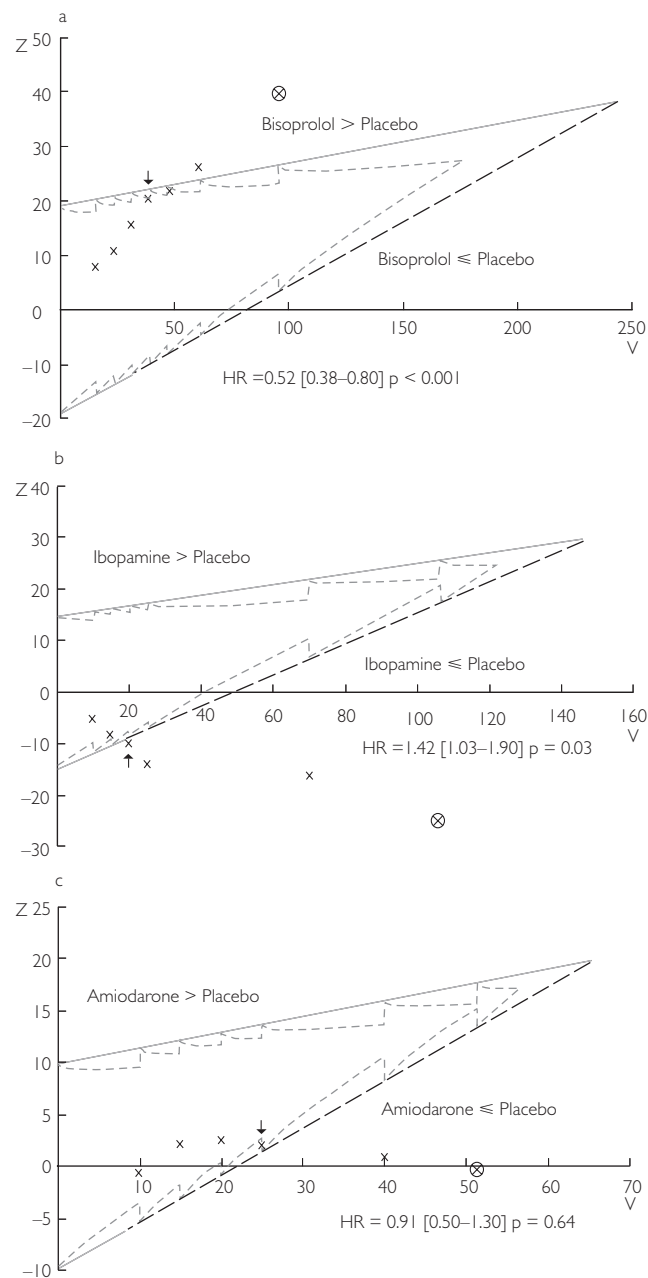
With the triangular test, the fourth interim analysis (25% of the expected death) reached the upper boundary, which would have led to stopping the trial prematurely

**Table 1** Published characteristics of the re-analysed clinical trials.

Study	Trial duration (days)	No. of patients	Hazard Ratio (%95CI)	p
CIBIS II Bisoprolol vs placebo in chronic heart failure	829	2647	0.66 (0.54–0.81)	<0.0001
EMIAT Amiodarone vs placebo in post-MI	1795	1486	0.99 (0.76–1.31)	0.96
PRIME II Ibopamine vs placebo in chronic heart failure	1066	1906	1.26 (1.04–1.53)	0.017
CIBIS I Bisoprolol vs placebo in chronic heart failure	1442	641	0.80 (0.56–1.15)	0.8

**Table 2** Initial hypothesis used for the rebuilding of sequential designs.

Study	$\alpha$ (two-sided)	Power (1- $\beta$ )	Expected rate of events in control group	Hazard ratio for the difference between treatments
CIBIS II	0.05	0.95	Annual mortality rate: 0.112	0.75
EMIAT	0.05	0.8	2-year mortality rate: 0.15	0.65
PRIME II	0.05	0.8	Annual mortality rate: 0.2	0.75
CIBIS I	0.05	0.9	2-year mortality rate: 0.37	0.67



**Figure 2** Results of the monitoring of clinical trials using the triangular test: (a) CIBIS 2; (b) PRIME 2; (c) EMIAT. The black arrow points the stopping time provided by this method. The estimation of the experimental treatment effect (hazard ratio and *P*-value) at this time is reported on each design. The circled cross corresponds to the report on the design of the real final analysis. HR = hazard ratio.

for a significant benefit of bisoprolol over placebo (Figure 2). A total number of 157 events would have been necessary to reach this conclusion with the triangular test instead of 314 (50% of expected deaths) with the Lan & DeMets method. The early stopping provided by the triangular test would have reduced the number of enrolled patients by 9% compared to the Lan & DeMets method. Despite the lower number of events with the triangular test, the estimation of the bisoprolol effect on mortality (Figure 2) remains correct compared to the published result. The trial duration of the CIBIS II study would have been reduced by 30% with the triangular test compared to the Lan & DeMets method: 517 days vs 738 days, respectively.

*Trial showing no significant differences between experimental and control: EMIAT*

Because of the low number of expected events (about 200 deaths), the first interim analysis was performed after 40 deaths (20%) then each 20 deaths (10%) with the triangular test. At the fourth interim analysis (50% of expected death), the lower boundary is crossed (Figure 2) in its dashed part and the trial would have been stopped for lack of difference between amiodarone and placebo on mortality in postmyocardial infarction. 931 days would have been necessary to reach this conclusion. The hazard ratio for the effect of amiodarone stays consistent with the value later published (Figure 2). Had the trial been stopped then 35% less patients would have been enrolled. Unlike the triangular test, the Lan & DeMets method does not provide stopping rules for the early acceptance of the null hypothesis (meaning the lack of significant difference between treatments). Thus, no early termination of this trial could have been observed using the Lan & DeMets method.

*Trial showing a deleterious effect of the experimental treatment: PRIME II*

With the triangular test, the third interim analysis (20% of the expected death) crosses the lower boundary in its solid part and would stop the trial because of inferiority of ibopamine compared with control on mortality in chronic heart failure (hazard ratio on Figure 2). The Lan

**Table 3** Number of events and of enrolled patients at the time of stopping decision according to each interim analyses procedures.

	Triangular test		Lan & DeMets	
	Events (Exp/Control)	No. of patients	Events (Exp/Control)	No. of patients
CIBIS II	157 (58/99)	2416	314 (129/185)	2647
EMIAT	100 (48/52)	960	–	–
PRIME II	80 (49/31)	522	340 (187/153)	1570
CIBIS I	120 (53/67)	641	120 (53/67)	641

Exp, experimental.

& DeMets method would have also provided an earlier detection of such detrimental effect but need a four fold higher number of events than the triangular test to demonstrate this effect. Thus, the number of enrolled patients was reduced by 66% (522 *vs* 1570) and the trial duration by 48% (486 days *vs* 950 days) with the triangular test compared to the Lan & DeMets method.

#### Underpowered trial: CIBIS I

The lower-than-expected mortality rate largely reduced the power of the study: the mortality rate in the placebo group was expected to be between 36 and 38% but was actually only 21%. Using both the interim analysis methods, no early termination of the study would have been advised. In this under-powered study, the triangular test would have lead to continuation of the study at the time of final analysis (20% of the expected number of deaths) as no boundary had been crossed. The Lan & DeMets method based on an information fraction related to the expected deaths would have lead to the same conclusion.

#### Report of later observations in the triangular test

In figure2, the points following the first crossing of a boundary (stopping time) provided by the triangular test (black arrow) correspond to the further information provided by the real data (the circled cross represents the real final analysis). In all trials, extra information confirms the consistency of the earlier decision. In these trials, the sample path keeps the same direction and all further points remain in the same side of triangle boundaries.

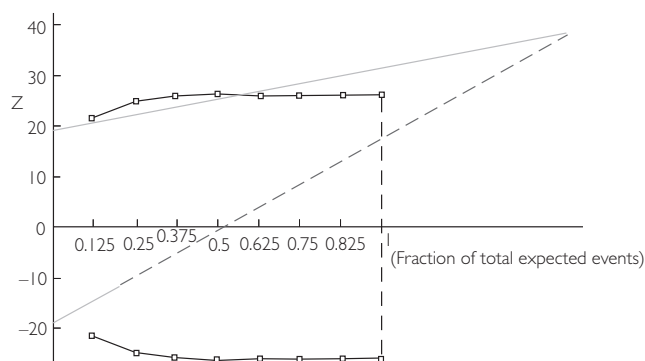
## Discussion

This study addressed the potential advantages and limits of the sequential approach for the monitoring of phase III placebo-controlled survival trials, and was based on a retrospective analysis of four trials using their individual data. Our results show that the triangular test method provides an early determination of the experimental treatment effect whatever it may be. However, the

importance of such a premature detection of treatment effect depends on the actual results observed.

The largest reduction in sample size and trial duration was detected when the experimental treatment did not show any benefit compared with placebo, and this is a major advantage of the triangular test. In survival trials there is a need to use formal interim analysis procedures, which monitor both efficacy and safety of the new therapy [23]. It is also desirable to stop the trial as soon as it is clear that the experimental treatment is not better, and even more so if it appears detrimental. The two-sided but asymmetric conception of the triangular design and the lower boundary double  $\times$  meaning [24] (Figure1b) appear optimally adapted for this.

The very early detection of the deleterious effect of a new treatment compared to placebo provided by the triangular test (as in the PRIME II study) is of considerable practical and ethical importance. The Lan & DeMets method, which is widely used in large-scale clinical trial (e.g. GUSTO V study [25]), provides symmetric boundaries for early detection of treatment differences. Our results show that the detection of a deleterious effect in the PRIME II study would have occurred with this method about one year later than with the triangular test. However, ethical reasons justify asymmetric stopping rules, since early termination for safety reasons should require less secure evidence than efficacy. Peto's rule [26], which assigns a fixed minimized  $\alpha'$  (at 0.001 for efficacy) for each interim analysis was used in the PRIME II study, where  $\alpha'$  was adjusted to 0.01 for safety. Compared with this well-accepted adjustment technique, the triangular test would have led to an earlier stopping of this trial. In contrast with the ' $\alpha$ -spending function' approach, the stopping rules provided by the triangular test are not only based on the  $P$ -value but also include the expected value of the experimental treatment benefit. Thus, at the very early time of stopping decision allowed by the triangular test in the PRIME II study, the  $P$ -value for the difference between treatment is 0.03 but the experimental treatment clearly showed no benefit. The difference between stopping rules for detrimental



**Figure 3** Graphical comparison of both interim analysis methods reported on same scales. The horizontal axis represents the information time according to the expected number of expected events and the vertical axis, the current difference between treatments. The dotted (light grey) and broken (----) lines draw the triangular test boundaries and the black lines (with  $\square$ ) represent the boundaries according to the Lan & DeMets method.

effect according to each method is well shown in the Figure 3 where both methods are represented on same scales for a same trial. One could wonder if the investigators would have really stopped the trial at the early date supported by the triangular test with so few patients, so few events and a  $P$ -value of 0.03? In the case of the PRIME II study later observations proved the triangular test to be right.

Similarly, the earliest detection of a lack of efficacy of the experimental treatment in the EMIAT trial was provided by the triangular test. While several trials have been stopped early for harm from experimental therapy [5, 18], early stopping for lack of difference between the experimental and the control treatments remains rare [12]. In contrast with the triangular test, the Lan & DeMets method does not provide rules to stop a trial early when there is a lack of difference between treatments. A design, which allows early stopping for accepting the null hypothesis can be however, implemented with the EAST software [22] and its application to the EMIAT study would have provided the same stopping time as with the triangular test (data not shown). However, one has to keep in mind that the boundary for lack of difference between treatments depends on the expected benefit for the experimental therapy and that interpretation of such results has to be cautious. The EMIAT study failed to demonstrate an expected 35% reduction of mortality in the amiodarone group but amiodarone was found to reduce the total mortality by 13% in the Amiodarone Trials Meta-Analysis, which included 6553 patients [27]. Thus the EMIAT study was underpowered to demonstrate this lower benefit, and the early stopping of the trial would have reduced this power even more. Under such circumstances the use of the triangular

test could warn the investigators that the objective of the trial is unlikely to be reached. However, the decision to stop such a trial should remain the responsibility of an independent data committee, which will take into account the clinical context.

On the other hand the value of the earliest possible detection of experimental treatment benefit is not necessarily an over-riding priority. The CIBIS II study was prematurely stopped using a Peto's rule [26]. However, the triangular test would have provided an earlier stopping time than was indicated with either the Peto's rule or the Lan & DeMets method. This would have led to an important reduction in the information about the mechanism of the benefit and about secondary endpoints. A trial that has demonstrated benefit from a new treatment will not be repeated. It will inevitably lead to wide-scale use of the new drug, and as much information as possible must be derived from the trial about the drugs efficacy and safety in patient subgroups. Trials showing an unexpected benefit from an experimental treatment must not therefore be stopped early simply on statistical grounds: such practical problems have been previously discussed [28, 29].

In our study, interim analyses were planned according to the total number of expected deaths. The frequency of analysis seems to have no influence on the statistical properties of the triangular test [15]. The planning of interim analyses after each 5% of the expected deaths was chosen for pragmatic reasons including feasibility and validity. Interim analyses were less frequent with the Lan & DeMets method as it has been showed that the times for these would be chosen independently of the behaviour of the sample path [20]. The application of identical interim analysis planning for both methods would not have changed our results.

The stopping rules for the triangular test were based here on a single crossing of a boundary. The consistent crossing of a boundary by several analyses before a stopping decision would be required as in the HOPE study [4]. Such a requirement would not have changed our conclusions.

Several other methods have been described for the 'repeated significance approach'. We chose here the Lan & DeMets method because of its ' $\alpha$ -spending function' flexibility [30] and its wide and accepted use in large-scale clinical trials. Several types of ' $\alpha$ -spending function' have been described ranging from the conservative O'Brien-Flemming type to the more aggressive Pocock type. We chose the O'Brien-Flemming variety because of the lower risk of  $\alpha$ -error rate inflation if underlying assumptions were incorrect or if there was an increase in the frequency of the analyses [15, 31].

Our study inevitably has limitations. We made a retrospective application of interim analyses methods to

real clinical trials. These trials are finished, published, and would not have used the triangular test when they were conducted. Our results are therefore not critical of those trials, but aim to improve planning and methodology of trials in future. An essential prerequisite for a successful application of such methods lies in the assurance of an accurate and swift reporting of events. The retrospective application of methods of interim analysis in our work was based on complete and validated data and this is seldom possible when interim analyses are performed. Planning data management becomes more important if interim analyses are made more efficient. First, the minimization of delays is necessary whatever procedure is used [32]. Second, there is no need for complete reports and only essential details such as treatment code, prognostic factors, date of inclusion and date of death required. Such data could be processed separately and rapidly communicated to the safety committee.

Finally, the decision to stop a study early involves ethical, administrative, economic and not just statistical, consideration. Statistics should be considered as a tool, which could trigger the procedure to stop a trial but statistics cannot supplant a trial safety committee.

In conclusion, the sequential approach based on the triangular test appears to be a useful methodological tool to assist in the planning and monitoring of placebo-controlled survival trials. Its major advantage lies in the early detection of lack of benefit of the experimental treatment compared to placebo, especially when the experimental treatment is detrimental. However, the use of this or any other approach must depend on the clinical context and the relevance of early trial terminations.

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