Testicular cancer risk and maternal parity: a population-based cohort study

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Summary The aim was to study, in a population-based cohort design, whether first-born sons run a higher risk of testicular cancer than later born sons; to investigate whether this difference in risk was affected by birth cohort, age of the son, maternal age, interval to previous delivery and other reproductive factors; and, finally, to evaluate to what extent changes in women's parity over time might explain the increasing incidence of testicular cancer. By using data from the Civil Registration System, a database was established of all women born in Denmark since 1935 and all their children alive in 1968 or born later. Sons with testicular cancer were identified in the Danish Cancer Registry. Among 1 015 994 sons followed for 15 981 967 person–years, 626 developed testicular cancer (443 non-seminomas, 183 seminomas). Later born sons had a decreased risk of testicular cancer (RR = 0.80, 95% CI = 0.67–0.95) compared with first-born sons. The RR was 0.79 (95% CI = 0.64–0.98) for non-seminomas and 0.81 (95% CI = 0.58–1.13) for seminomas. There was no association between testicular cancer risk and overall parity of the mother, maternal or paternal age at the birth of the son, or maternal age at first birth. The decreased risk of testicular cancer among later born sons was not modified by age, birth cohort, interval to the previous birth, sex of the first-born child, or maternal age at birth of the son or at first birth. The increased proportion of first-borns from birth cohort 1946 to birth cohort 1969 only explained around 3% of an approximated two-fold increase in incidence between the cohorts. Our data document a distinctly higher risk of testicular cancer in first-born compared with later born sons and suggest that the most likely explanation should be sought among exposures in utero. The increase in the proportion of first-borns in the population has only contributed marginally to the increase in testicular cancer incidence.

Keywords: birth order; cohort study; epidemiology; oestrogen; maternal age; parity; risk factor; testicular cancer

The incidence rate of testicular cancer in Denmark is among the highest in the world (IARC, 1992) and the risk has increased three-fold during the last half-century (Møller, 1993; Danish Cancer Society, 1996). Cryptorchidism is one of the few established risk factors (Forman et al, 1990; United Kingdom Testicular Cancer Study Group, 1994), but the main causes of testicular cancer remain unknown. Studies of familial occurrence of testicular cancer have suggested a genetic component (Forman et al, 1992; Heimdal et al, 1996; Westergaard et al, 1996), but environmental factors are likely to play an important role (Adami et al, 1994). The specific age pattern of testicular cancer, with a marked peak in incidence in young adult life, as well as the association with cryptorchidism suggest that causal factors may be operating early in life, perhaps in utero.

Several case–control studies (Depue et al, 1983; Swerdlow et al, 1987; Prener et al, 1992; Møller and Skakkebæk, 1996), but not all (Morrison, 1976; Henderson et al, 1979; Brown et al, 1986; Akre et al, 1996), have reported an increased risk of testicular cancer in first-born compared with later born sons. If such an association exists it suggests a modifying effect of environmental factors on the risk of testicular cancer (MacMahon and Pugh, 1970; Khoury et al, 1993), factors that are likely to be operational during fetal life

or early childhood. Higher levels of oestrogens have been measured during first compared with subsequent pregnancies (Bernstein et al, 1986; Panagiotopoulou et al, 1990; Key et al, 1996) and a higher risk of testicular cancer among first-born sons has been interpreted as supporting the hypothesis that high levels of endogenous oestrogens during critical stages of fetal development may increase the risk of testicular cancer (Bernstein et al, 1986; Panagiotopoulou et al, 1990; Prener et al, 1992).

In the present study, we used a large population-based cohort design to study testicular cancer and maternal parity and also to study other possible risk factors. Furthermore, we considered how much of the increase in the incidence of testicular cancer between two birth cohorts might be explained by changes in parity.

MATERIALS AND METHODS

Cohort

Data from the Danish Civil Registration System were used to generate a population-based database of women born in Denmark and their live-born children. All live-born children and new residents in Denmark are recorded in the Civil Registration System (Malig, 1995) and assigned a unique personal identification number that contains information on the date of birth and gender of the person. Individual information is kept under this identification number in all national registers, enabling identity-secure linkage of information between registers. The Civil Registration System was established on 1 April 1968, when all persons living in Denmark were registered. A link to the parents was established at that time for children living at the same address as their parents,

and links between parents and live-born children have been recorded in the register ever since. The registration system also includes continuously updated information on name, address and vital status. We established a population-based database by extracting data recorded in the Civil Registration System on all women born in Denmark during the period 1 April 1935 to 31 March 1978. We also included all their children who were alive on 1 April 1968 or later, and who were born before 1 January 1993. It was not possible to distinguish children who were adopted from children who had a link to their biological parents. However, only around 1% of children are adopted by women in Denmark (Danmarks Statistik, 1960–85). The male children in the above database constituted the study cohort.

Identification of cases

Cases of germ cell testicular cancer were identified by linkage with the Danish Cancer Registry, which has recorded nearly all cases of cancer in Denmark since 1943 (Storm, 1991). Patients with testicular cancer in the cohort that had been diagnosed before 1978, when the *International Classification of Diseases for Oncology* (ICD-O) (World Health Organization, 1976) was introduced in the registry, were re-evaluated on the basis of the original information from clinicians and pathologists and assigned an ICD-O code. A total of 646 patients with testicular cancer were identified. Among these, 20 were excluded (15 with non-germ cell testicular cancer and five with unknown histology) leaving 626 patients with germ cell testicular cancer (ICD-O morphology codes 9060–9102 and topography codes 186.0 or 186.9) for subsequent analysis.

Data analysis

The sons in the cohort were each followed for a first diagnosis of testicular cancer from 1 April 1968, or date of birth, whichever occurred last, until the date of death (1.5% of the cohort), emigration (1.3%), disappearance (0.015%), or 31 December 1992 (97.1%), whichever occurred first. A log-linear Poisson regression model (Breslow and Day, 1987) was used to investigate the possible effect on the risk of testicular cancer of the following co-variates: parity of the mother at birth of the son; overall parity of the mother; interval from birth of the son to the previous delivery to the mother; sex of the first-born child, maternal and paternal age at birth of the son, maternal age at first birth, and multiple birth. Parity of the mother at the son's birth was defined as 1 plus the previous number of deliveries that had resulted in live-born children. Overall parity of the mother was defined as the cumulated number of deliveries resulting in live-born children and was calculated as a time-dependent variable. The analyses were performed for all testicular cancers combined, as well as for seminomas (i.e. pure seminomas) and non-seminomas separately. All analyses were adjusted for age of the son (0–14, 15–19, 20–24, 25–29 and ≥30 years) and birth cohort of the son (1950–57, 1958–62, 1963–67, 1968–92). Adjustments were furthermore carried out for maternal age at birth of the son, maternal age at first birth, parity of the mother at birth of the son, and overall parity of the mother.

All analyses were performed using the SAS procedure GENMOD (SAS Institute, 1996). Ratios of incidence rates for testicular cancer were used as a measure of the relative risk (RR). Two-tailed *P*-values were obtained from likelihood ratio tests and

95% confidence intervals (CI) were calculated by the use of Wald's test. Trends were estimated as the slope when the categorical variables of interest were treated as quantitative variables. The numerical value assigned to a given category was chosen as the median of the distribution of person–years within the category. The log-linear assumptions were checked by likelihood ratio tests.

The proportion of first-born children in Denmark has increased from 32.5% in 1946 to 38.0% in 1969 and to 47.0% in 1990 (Danmarks Statistik, 1973; Danmarks Statistik, 1992). We estimated how much higher the incidence of testicular cancer in the 1946 birth cohort would have been had there been the same proportion of first-borns as in the birth cohort of 1969 or 1990 (Breslow and Day, 1980). Furthermore, we estimated how much of the relative difference in testicular cancer incidence between the birth cohorts of 1946 and 1969 might be explained by differences in the birth order distribution in the two birth cohorts (the relative attributable risk) (Breslow and Day, 1980).

RESULTS

The cohort of 1 015 994 sons was followed for the occurrence of testicular cancer for 15 981 967 person–years. A total of 626 sons had a first diagnosis of testicular cancer (443 non-seminomas and 183 seminomas) during the follow-up period.

Sons born at second or later deliveries were at lower risk of testicular cancer (RR = 0.80; 95% CI = 0.67–0.95) than first-born sons (Table 1). Adjustment for overall parity did not affect the risk estimate (RR = 0.80; 95% CI = 0.66–0.97). The risk among later born sons was similarly decreased for non-seminomas (RR = 0.79; 95% CI = 0.64–0.98) and seminomas (RR = 0.81; 95% CI = 0.58–1.13), but this association was only statistically significant for non-seminomas.

There was no association between the overall parity of the mother and the RR of testicular cancer (Table 1). Likewise, the risk was not influenced by maternal or paternal age at the birth of the son (Table 1). Nor was there an association between the risk of testicular cancer and maternal age at birth for first-born sons, the RR for first-born sons being 0.97 (95% CI 0.78–1.22) for maternal age at birth <20 years, 1 (reference) for 20–24 years, 0.81 (0.59–1.12) for 25–29 years and 1.14 (0.50–2.59) for ≥30 years. Whereas the RR of non-seminomas was lower for maternal ages at first birth of 25–29 years compared with 20–24 years, the overall risk estimates according to maternal age at first birth were not significantly different (Table 1). Multiple births tended to have a decreased risk of testicular cancer (RR = 0.42; 95% CI = 0.16–1.12; *n* = 4) after adjustment for age of the son, birth cohort of the son, maternal age at birth of the son, and parity of the mother at birth of the son. Exclusion of multiple births from the analyses did not affect the other risk estimates.

Compared with first-borns, the RR of testicular cancer among later born sons tended to be decreased for most intervals to the previous delivery and the RRs among later born sons did not differ significantly from each other (Table 1). The decreased RR of testicular cancer among second- and later born sons was not affected by the sex of the first-born child. The RR of testicular cancer for second-, third-, and fourth- or later born sons was 0.81 (95% CI = 0.64–1.02), 0.87 (95% CI = 0.59–1.29), and 0.64 (95% CI = 0.28–1.44), respectively, for sons from families in which the first-born child was a boy, and 0.74 (95% CI = 0.57–0.95), 0.85 (0.55–1.32) and 0.73 (0.30–1.78), respectively, for sons from families where the first-born child was a girl.

Table 1   Relative risk (RR) of a first diagnosis of testicular cancer in a cohort of 1 016 165 sons born to Danish-born women

	Cohort	All testicular cancers		Non-seminomas		Seminomas	
	Person-years	Cases	RR (95% CI)	Cases	RR (95% CI)	Cases	RR (95% CI)
Parity[a] of mother at birth of son[b]							
1	7 972 276	398	1 ref.	277	1 ref.	121	1 ref.
2	5 590 265	165	0.78 (0.65–0.95)	120	0.78 (0.63–0.98)	45	0.78 (0.55–1.13)
3	1 860 303	52	0.87 (0.64–1.19)	37	0.83 (0.58–1.20)	15	0.90 (0.52–1.57)[c]
4+	559 123	11	0.68 (0.37–1.27)	9	0.73 (0.37–1.46)	2	
P for trend			P = 0.03		P = 0.06		P = 0.36
1	7 972 276	398	1 ref.	277	1 ref.	121	1 ref.
2+	8 009 691	228	0.80 (0.67–0.95)	166	0.79 (0.64–0.98)	62	0.81 (0.58–1.13)
			P = 0.01		P = 0.03		P = 0.21
Overall parity of mother[d]							
1	2 371 288	57	1.16 (0.87–1.56)	45	1.27 (0.91–1.78)	12	0.89 (0.48–1.66)
2	7 453 348	242	1 ref.	176	1 ref.	66	1 ref.
3	4 292 150	216	1.10 (0.91–1.33)	153	1.11 (0.88–1.39)	63	1.09 (0.76–1.55)
4	1 358 354	78	1.04 (0.79–1.37)	48	0.94 (0.66–1.32)	30	1.24 (0.78–1.96)
5+	506 827	33	1.08 (0.73–1.60)	21	1.01 (0.61–1.66)	12	1.13 (0.59–2.18)
P for trend			P = 0.92		P = 0.53		P = 0.34
Maternal age at birth of son (years)[e]							
<20	1 863 190	130	0.97 (0.79–1.20)	88	1.00 (0.78–1.29)	42	0.92 (0.63–1.35)
20–24	6 746 494	344	1 ref.	243	1 ref.	101	1 ref.
25–29	5 239 978	127	0.84 (0.67–1.04)	90	0.73 (0.56–0.95)	37	1.25 (0.82–1.89)
≥30	2 132 304	25	0.93 (0.60–1.45)	22	0.90 (0.56–1.45)	3	0.99 (0.29–3.36)
P for trend							
Paternal age at birth of son (years)[e]							
<20	417 733	30	1.10 (0.75–1.62)	25	1.43 (0.93–2.19)	5	0.52 (0.21–1.28)
20–24	4 040 550	204	1 ref.	135	1 ref.	69	1 ref.
25–29	6 063 189	243	1.12 (0.92–1.35)	175	1.17 (0.93–1.47)	68	1.02 (0.72–1.44)
≥30	5 059 463	126	1.01 (0.79–1.27)	92	1.01 (0.76–1.35)	34	1.00 (0.65–1.54)
Missing	401 030	23	–	16	–	7	–
Maternal age at first birth (years)[e]							
<20	3 837 310	200	0.97 (0.81–1.16)	141	1.04 (0.84–1.28)	59	0.82 (0.59–1.13)
20–24	8 484 786	360	1 ref.	252	1 ref.	108	1 ref.
25–29	3 129 796	59	0.74 (0.56–0.98)	43	0.68 (0.49–0.94)	16	0.97 (0.56–1.68)[f]
≥30	530 074	7	1.08 (0.50–2.31)	7	1.21 (0.56–2.59)	0	–
Interval from the son's birth to the previous delivery of the mother (years)[b]							
No previous deliveries	7 972 276	398	1 ref.	277	1 ref.	121	1 ref.
<2	1 774 323	75	0.78 (0.61–1.01)	52	0.78 (0.57–1.05)	23	0.81 (0.51–1.27)
2 to <3	2 072 615	58	0.71 (0.54–0.95)	44	0.74 (0.54–1.03)	14	0.64 (0.36–1.12)
3 to <4	1 581 352	47	0.92 (0.67–1.26)	31	0.82 (0.56–1.21)	16	1.22 (0.70–2.12)
4 to <5	970 719	27	1.00 (0.67–1.51)	19	0.94 (0.58–1.53)	8	0.71 (0.34–1.49)[g]
≥5	1 610 682	21	0.68 (0.42–1.10)	20	0.83 (0.50–1.38)	1	

RR, relative risk; CI, confidence interval. [a]Note definition of parity in Data analysis. [b]Adjusted for age, birth cohort, and parity of the mother at birth of son. [c]Adjusted for age, birth cohort, and maternal age at birth of the son. '3+. [d]Adjusted for age, birth cohort, and parity of the mother at birth of the son. ≥25 years. [e]Adjusted for age, birth cohort, and parity of the mother at birth of son. [f]Adjusted for age, birth cohort, and parity of the mother at birth of the son. ≥25 years. [g]Adjusted for age, birth cohort, and maternal age at birth of the son. ≥4 years.

The RR of testicular cancer was particularly low for later born sons aged 30 years or more (RR = 0.35; 95% CI = 0.17–0.72) compared with first-born sons of similar age (Table 2). However, the risk among later born sons compared with first-borns tended to be lower for most age categories and the RRs among later born sons did not differ significantly from each other for testicular cancer overall or for non-seminomas or seminomas. Neither was the decreased risk among later born sons significantly modified by maternal age at birth of the son, maternal age at first birth, or birth cohort of the son, although the association was particularly strong for those born after 1968 (Table 2).

The 20% decreased risk of testicular cancer in later born compared with first-born sons corresponds to an increased risk of testicular cancer of 25% in first-born compared with later born sons. With this RR estimate we calculated that if the 1946 birth cohort had had the same birth order distribution as the 1969 or 1990 birth cohort, the incidence of testicular cancer in the 1946 cohort would have been 1.3% and 3.4% higher respectively. With an approximated twofold increase in the incidence of testicular cancer for the 1969 birth cohort compared with the 1946 cohort and under the assumption that our finding of a 25% higher risk of testicular cancer in first-born compared with later born sons has remained unchanged over time, the changed birth order distribution was estimated to explain 2.6% of the increase in incidence between the two birth cohorts.

DISCUSSION

Overall, we found a significantly decreased risk of testicular cancer among sons born at second or later deliveries compared with first-born sons. This decreased risk among later born sons of approximately 20% was similar for both non-seminomas and seminomas. It was not modified by other factors studied, i.e. interval to the previous delivery, sex of the first-born child, maternal age at birth of the son, maternal age at first birth, age of the son or birth cohort.

The proportion of first-born children in Denmark has increased considerably during the last half century (Danmarks Statistik, 1973, 1992). However, the change in the birth order distribution between the 1946 and 1969 birth cohort was found to explain less than 3% of an approximated twofold increase in incidence of testicular cancer between the two birth cohorts. This estimate was made under the assumption that the 20% lower risk of testicular cancer in later born compared with first-born sons remained unchanged over time. Our result is in contrast to calculations made by Prener et al (1992), who on the basis of their case–control study of 183 case patients estimated that between 15% and 20% of this remarkable increase in the incidence of testicular cancer might be explained by changes in the birth order distribution of young men (Prener et al, 1992).

This study was designed as a historic prospective population-based cohort study based on mandatorily reported registry information. Compared with problems often relevant to case–control studies, our data were not influenced by, for example, selection bias, recall bias or overmatching. There was a potential risk of misclassification of parity of the mother at birth of the son as only children alive in 1968 (when the Civil Registration system was established) and those born after that time were considered. However, this potential misclassification would only tend to dilute a difference in risk of testicular cancer between first-born and later born sons, and thus the approximately 20% reduced risk among later borns is likely to be a conservative estimate.

The fact that neither birth cohort, interval to the previous delivery, overall parity of the mother nor other factors studied influenced the 20% lower risk in later born sons suggests that the mechanism behind this phenomenon may be related to factors that are influential during life in utero. It has been hypothesized that high levels of endogenous oestrogens during critical stages of development in utero may increase the risk of cryptorchidism and testicular cancer (Henderson et al, 1979; 1983), and at least three studies have found some evidence for increased oestrogen levels in first compared with later pregnancies (Bernstein et al, 1986; Panagiotopoulou et al, 1990; Key et al, 1996). Other studies also tend to support this hypothesis (McLachlan et al, 1975; Depue, 1984; Brown et al, 1986; Depue et al, 1983; Bernstein et al, 1988).

Cryptorchidism has been found to be associated with being first born. One large Swedish study including 2424 cryptorchid boys found a significantly increased risk of cryptorchidism among first births (Hjertkvist et al, 1989) compared with later births, and a suggestion of a similar association was found in two other studies (Swerdlow et al, 1983; Møller and Skakkebæk, 1996), although some smaller studies failed to support this observation (Beard et al, 1984; Depue, 1984; Berkowitz et al, 1995). However, the prevalence of cryptorchidism in patients with testicular cancer is only around 10% (United Kingdom Testicular Cancer Study Group, 1994; Schottenfeld, 1996) and therefore not likely in itself to account for the observed association between parity of the mother at birth and the risk of testicular cancer in the male offspring. Rather, our findings would be in line with the hypothesis that cryptorchidism and testicular cancer may be associated through shared risk factors (Henderson et al, 1979).

Nevertheless, the evidence of a link between endogenous hormone levels in utero and testicular cancer remains indirect, and the consistent finding of an increased risk of testicular cancer among first-born sons could also have its explanation in other environmental factors that differ between first-born and later born children. First-born children may in general be exposed to infections in the environment at a later age than later born children (Fox et al, 1970; MacMahon, 1992). Epidemiological features of testicular cancer such as the age pattern with a peak in young adult life and the association with high socioeconomic status reported in several countries (Swerdlow et al, 1991; Schottenfeld, 1996) have been interpreted as supportive of the hypothesis that late exposure to a common infectious agent might increase the risk of testicular cancer (Newell et al, 1984). However, it should be recalled that the overall evidence for an infectious aetiology in testicular cancer is limited (Swerdlow, 1993).

We did not have information on socioeconomic status of the sons. However, three studies from Denmark did not find any clear association between testicular cancer and socioeconomic status, perhaps because Denmark is a rather egalitarian country (Davies et al, 1990; Prener et al, 1992; Møller and Skakkebæk 1996). Moreover, it is unlikely that socioeconomic status alone could explain an increased risk among first-born sons compared with all later births independent of overall parity of the mother and maternal age at first birth.

It has recently been proposed that a decline in maternal age at first birth might explain part of the observed increase in testicular cancer incidence (Henderson et al, 1997). However, we found no association between maternal age at first birth and the risk of testicular cancer, either for the whole cohort or for the first-borns.

Multiple pregnancies differ from singleton pregnancies in several respects, including hormone levels, birth weight, maternal age and parity. However, exclusion of multiple births had no effect on the results. Contrary to previous suggestions (Depue et al, 1983;

Table 2. Relative risk (RR) of a first diagnosis of testicular cancer in sons born at their mother's first delivery (reference) compared with sons born at second or later deliveries, in a cohort of 1 015 994 sons born to Danish-born women. By age of son, maternal age at birth of the son, maternal age at first birth, and birth cohort of the son

| | Cohort | | All testicular cancers | | | | Non-seminomas | | | | Seminomas | | | |
| | Person–years | | Cases | | RR (95% CI) | | Cases | | RR (95% CI) | | Cases | | RR (95% CI) | |
Parity of mother at birth of son	1	2+	1	2+	1	2+	1	2+	1	2+	1	2+	1	2+
Age of son (years)[a]														
0–14	5 063 890	5 635 261	15	10	1 ref.	0.62 (0.28–1.39)	14	10	1 ref.	0.68 (0.30–1.55)	1	0		
15–19	1 291 014	1 237 759	57	47	1 ref.	0.87 (0.59–1.30)	54	43	1 ref.	0.86 (0.57–1.30)	3	4	1 ref.	0.98 (0.24–3.97)[b]
20–24	907 937	752 903	133	89	1 ref.	0.81 (0.61–1.07)	104	70	1 ref.	0.83 (0.61–1.14)	29	19	1 ref.	0.74 (0.41–1.35)
25–29	508 299	322 283	123	74	1 ref.	0.93 (0.69–1.25)	78	43	1 ref.	0.72 (0.50–1.03)[c]	45	31	1 ref.	1.00 (0.62–1.61)
≥30	201 136	61 485	70	8	1 ref.	0.35 (0.17–0.72)	27	0	1 ref.		43	8	1 ref.	0.54 (0.25–1.17)
Maternal age at birth of son (years)[d]														
<20	1 663 524	199 666	119	11	1 ref.	0.77 (0.41–1.43)	82	6	1 ref.	0.59 (0.26–1.34)	37	5	1 ref.	1.23 (0.48–3.14)
20–24	4 140 397	2 606 097	227	117	1 ref.	0.79 (0.63–0.99)	158	85	1 ref.	0.80 (0.62–1.05)	69	32	1 ref.	0.77 (0.50–1.17)
25–29	1 792 651	3 447 327	46	81	1 ref.	0.84 (0.59–1.21)	31	59	1 ref.	0.90 (0.58–1.39)	15	22	1 ref.	0.78 (0.41–1.48)[c]
≥30	375 704	1 756 600	6	19	1 ref.	0.62 (0.25–1.55)	6	16	1 ref.	0.52 (0.20–1.33)	0	3		
Maternal age at first birth (years)[d]														
<20	1 663 524	2 173 787	119	81	1 ref.	0.77 (0.57–1.02)	82	59	1 ref.	0.72 (0.51–1.02)	37	22	1 ref.	0.87 (0.50–1.49)
20–24	4 140 397	4 344 389	227	133	1 ref.	0.79 (0.63–0.98)	158	94	1 ref.	0.72 (0.55–0.93)	69	39	1 ref.	1.00 (0.67–1.50)
25–29	1 792 672	1 337 124	46	13	1 ref.	0.56 (0.30–1.04)	31	12	1 ref.	0.71 (0.37–1.39)	15	1	1 ref.	0.18 (0.02–1.40)[c]
≥30	375 683	154 392	6	1	1 ref.	0.59 (0.07–4.87)	6	1	1 ref.	0.56 (0.07–4.68)	0	0		
Birth cohort of son[e]														
1950–62	1 612 463	741 255	194	62	1 ref.	0.75 (0.56–1.01)	112	33	1 ref.	0.68 (0.46–1.01)	82	29	1 ref.	0.83 (0.54–1.30)
1963–67	1 959 205	1 910 268	130	112	1 ref.	0.94 (0.72–1.22)	100	84	1 ref.	0.95 (0.71–1.29)	30	28	1 ref.	0.90 (0.53–1.53)
1968–92	4 400 609	5 358 168	74	54	1 ref.	0.63 (0.43–0.91)	65	49	1 ref.	0.67 (0.46–1.00)	9	5	1 ref.	0.43 (0.14–1.32)

RR, relative risk; CI, confidence interval. [a]Adjusted for birth cohort of son, and maternal age at birth of the son. [b]0–19 years. [c]≥25 years. [d]Adjusted for age, and birth cohort of son. [e]Adjusted for age, and maternal age at birth of the son at birth of the son

Swerdlow et al, 1987; Braun et al, 1994; Braun et al, 1995), we found that multiple births tended to have a lower risk of testicular cancer than singletons. This finding, however, was based on a very limited number of cases.

In conclusion, our large cohort study documents the existence of a higher risk of testicular cancer among first-born compared with later born sons. This association was not significantly influenced by a number of potential risk modifiers, including birth cohort, age of the son, maternal age at birth of the son, maternal age at first birth and interval to the previous delivery. Our findings suggest that early environmental exposures are of importance for the development of testicular cancer and are compatible with the hypothesis that exposures in utero may play a central role for this risk difference between first-born and later born males. The increase in the proportion of first-born children during the last half century has only marginally contributed to the observed increase in the incidence of testicular cancer.

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