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## Glutamic acid decarboxylase antibodies are indicators of the course, but not of the onset, of diabetes in middle-aged adults: The Atherosclerosis Risk in Communities Study

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### Abstract

To efficiently examine the association of glutamic acid decarboxylase antibody (GADA) positivity with the onset and progression of diabetes in middle-aged adults, we performed a case-cohort study representing the ~9-year experience of 10,275 Atherosclerosis Risk in Communities Study participants, initially aged 45–64. Antibodies to glutamic acid decarboxylase antibody (GAD<sub>65</sub>) were measured by radioimmunoassay on 580 incident diabetes cases and 544 non-cases. The overall weighted prevalence of GADA positivity ( $\geq 1$  U/ml) was 7.3%. Baseline risk factors, with the exception of smoking and interleukin-6 ( $P \leq 0.02$ ), were generally similar between GADA-positive and -negative individuals. GADA positivity did not predict incident diabetes in multiply-adjusted (HR=1.04; 95% CI: 0.55, 1.96) proportional hazard analyses. However, a small non-significant adjusted risk (HR = 1.29; 95% CI: 0.58, 2.88) was seen for those in the highest tertile ( $\geq 2.38$  U/ml) of positivity. GADA-positive and GADA-negative non-diabetic individuals had similar risk profiles for diabetes, with central obesity and elevated inflammation markers, aside from glucose, being the main predictors. Among diabetes cases at study's end, progression to insulin treatment increased monotonically as a function of baseline GADA level. Overall, being GADA positive increased risk of progression to insulin use almost 10 times (HR=9.9; 95% CI: 3.4, 28.5). In conclusion, in initially non-diabetic middle-aged adults, GADA positivity did not increase diabetes risk, and the overall baseline profile of risk factors was similar for positive and negative individuals. Among middle-aged adults, with the possible exception of those with highest GADA levels, autoimmune pathophysiology reflected by GADA may become clinically relevant only after diabetes onset.

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## Keywords

Diabetes mellitus; Glutamic acid decarboxylase; Auto-immune diseases; Inflammation; Risk factors

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## Introduction

Glutamic acid decarboxylase antibodies (GADA) are present in most patients with autoimmune diabetes at or before diagnosis and are related to insulin secretion abnormalities and to the onset of type 1 diabetes (1). Reported to be present in <5% to greater than 20% of adults diagnosed as presumably having type 2 diabetes (2–7), GADA are also markers of a slowly evolving form of autoimmune diabetes, frequently called latent autoimmune diabetes in adults (LADA)(4;8;9). Most studies to date have evaluated individuals at or after diagnosis of diabetes, and little attention has been directed to the question of whether the presence of GADA predicts the development of diabetes in middle-aged adults.

Subclinical inflammation has been found to predict the development of diabetes with onset in middle age (10). Although this association is in great part related to obesity, it is logical to inquire whether it also reflects the fact that some incident, middle-aged cases may have a subclinical, chronic autoimmune disease affecting  $\beta$  cells.

The objectives of this study are to characterize the importance of GADA positivity in a representative sample of initially non-diabetic middle-aged individuals enrolled in the Atherosclerosis Risk in Communities (ARIC) Study. Specifically, we aimed to describe the frequency and correlates of GADA positivity; to evaluate whether GADA positivity predicts the development of diabetes, and to evaluate whether risk factors (including inflammation markers) for diabetes and the early clinical course of diabetes are similar in GADA-positive and GADA-negative individuals.

## Subjects and Methods

In 1987–89 the ARIC study recruited a population-based cohort of 15,792 men and women 45 to 64 years of age from four U.S. communities (11). All subjects were invited to return to three clinic visits, at approximately 3-year intervals, at which incident diabetes was ascertained. Human subjects research review committees at the involved institutions approved the study, and all participants gave written informed consent.

In order to permit an efficient use of ARIC frozen biologic specimens, a case-cohort design was chosen to investigate our objectives. All specimens were stored at  $-70^{\circ}\text{C}$  within 90 min of venipuncture. Before sampling, we excluded 2,018 participants with prevalent diabetes, 95 members of minority ethnic groups with small numbers, 853 not returning to any follow-up visit, 26 having no valid diabetes determination at follow-ups, 7 with restrictions on stored plasma use, 12 with missing baseline anthropometrics and 2,506 participants in previous ARIC case-control studies involving cardiovascular disease for whom stored plasma was either previously exhausted or held in reserve. This resulted in a final sample of 10,275 individuals (75% of those in the full cohort without diabetes at baseline), 1155 of whom (11.2%) developed diabetes during follow-up. From these 10,275 eligible members of the cohort, we selected and measured analytes on ethnicity-stratified random samples of cases of incident diabetes and of eligible members of the full cohort (a total 1,198 individuals). A few of the incident cases of diabetes overlapped with the cohort sample, and a few were selected only via the cohort sample. Of those sampled, we excluded 45 for incomplete fasting (<8 h) or for not having values for all covariates, 27 identified as non-cases but having 2h glycemia  $\geq 11.1$  mmol/L during the oral

glucose tolerance test performed at the last follow up visit, and 2 individuals lacking information about antibodies for GADA, leaving a total of 1,124 subjects for analysis.

At baseline and at follow-up visits, glucose was measured by a hexokinase method, while fasting serum insulin was measured by a nonspecific radioimmunoassay. The cross-reactivity with pro-insulin was 33%. Waist girth was measured at the umbilical level, and hip circumference at the maximum hip girth. The waist-to-hip ratio (WHR) was defined as the ratio between these two measurements. A family history of diabetes was defined as a report of diabetes in either parent. The definitions and methods for other baseline measurements (height, weight, smoking status, systolic blood pressure, triglycerides, HDL cholesterol, insulin, white cell count, and fibrinogen) have been previously reported (12).

Antibodies to glutamic acid decarboxylase (GAD<sub>65</sub>) were measured in duplicate by radioimmunoassay using human recombinant I<sup>125</sup> according to manufacturer's protocol (Kronus, Boise, ID, USA). The mean of duplicate measurements was used. The reliability coefficient (the ratio of between-person to total variance obtained analyzing replicate pairs of baseline samples drawn at one sitting on a subset of subjects) was 0.98. Subjects were considered positive for GADA if the value was 1 U/ml or higher.

The acute-phase markers C-reactive protein (CRP), orosomucoid, sialic acid, and fibrinogen, as well as interleukin (IL)-6 and white cell count, were chosen to characterize chronic systemic inflammation. White cell count was determined at each center with Coulter counters. The remaining determinations were performed at central laboratories: fibrinogen at baseline in citrated samples using thrombin-time titration, and the other analytes on plasma specimens frozen at baseline. IL-6 levels were measured by high-sensitivity enzyme-linked immunosorbent assay (ELISA; R&D systems, Minneapolis, MN, USA); sialic acid by enzymatic determination using a colorimetric assay (13) (Roche Diagnostics, Indianapolis, IN, USA); orosomucoid utilizing an immunoturbidimetric technique (Kamiya Biomedical, Seattle, WA, USA); and high sensitivity CRP (hsCRP - Latex II, Denka Seikan, Tokyo, Japan) on an Hitachi 911 automated analyzer (Indianapolis, IN, USA). An additional analyte, leptin, was measured in duplicate by direct sandwich ELISA (Linco Research, Inc., St. Charles, Missouri, MO, USA) on the case cohort sample.

As in previous reports (14;15), we created a score to indicate low-grade systemic inflammation ranging from 0 to 6, attributing 1 point to a value greater than the median of the cohort sample for each of the six inflammation markers – IL-6, CRP, orosomucoid, sialic acid, white cell count and fibrinogen.

Diabetes was defined on the basis of 1) a reported physician diagnosis, 2) use of anti-diabetes medications, 3) a fasting ( $\geq 8$  h) glucose  $\geq 7.0$  mmol/L, or 4) a non-fasting glucose of  $\geq 11.1$  mmol/L. For incident diabetes ascertained on the basis of a glucose value, incident date was estimated by linear interpolation using glucose values at the ascertaining visit and the previous one. The fasting glucose at ascertainment for subjects who had been told they had diabetes or who were on diabetic medication may have been affected by their knowledge of their diabetic status (and in some cases was even below 7.0 mmol/L). For these subjects, the time to reach 7.0 mmol/L was estimated by using their fasting glucose at the earlier visit and a slope estimated using information from all diabetic subjects who had been unaware of their status. Hypertension was defined as a blood pressure  $\geq 140/90$  mmHg or reported use of anti-hypertensive medication.

### Statistical analysis

Statistical analysis was based on our case-cohort sampling design, weights being defined as the inverse of the ethnicity-specific sampling fractions, thus permitting statistical estimation

and inference relevant to the entire cohort. In the cohort random sample, we did weighted analysis of covariance to compute adjusted means and proportions of socio-demographic variables and risk factors. Tertiles of GADA levels among GADA-positive subjects, used in some of the analyses, were computed in the cohort random sample. We tested differences in clinical characteristics adjusted for baseline levels using a similar weighted covariance analysis. Unadjusted and adjusted relative risks of developing diabetes and of progressing to insulin use once diabetes was present were estimated in proportional hazards models fit using SUDAAN to account for the case-cohort design with stratified sampling both from the whole cohort and from the incident diabetes cases (16). Statistical significance of differences in hazard ratios between GADA-positive and -negative individuals was evaluated by Wald tests of the interaction terms in similar models. Analyses were performed using SAS (17) and SUDAAN (18).

## Results

After the exclusions, the cohort random sample comprised 639 individuals, the incident diabetes random sample 558 subjects. With overlap, the total sample consisted of 1124 subjects, 580 who went on to develop incident diabetes and 544 who did not. The median (interquartile range) follow-up to time of onset in those who developed diabetes was 3.0 (1.7 – 5.9) years and to censoring for those who did not was 8.9 (8.7 – 9.0) years. Of incident case subjects, 498 (86%) were ascertained by a fasting glucose  $\geq 7.0$  mmol/L, 5 (1%) by a nonfasting glucose  $\geq 11.1$  mmol/L, and, independently of their glucose values, 125 (22%) by reported physician diagnosis or medication use. The characteristics of this case-cohort sample have been previously described (14). Among GADA-positive and -negative participants, 38.6 and 37.3% were male, 32.3 and 21.7% had parental history of diabetes, and 30.2 and 22.0% were African-American, respectively. Mean  $\pm$  SEM ages were  $51.6 \pm 0.66$  and  $52.6 \pm 0.23$ , and mean BMI's  $27.2 \pm 0.80$  kg/m<sup>2</sup> and  $27.3 \pm 0.21$  kg/m<sup>2</sup>, respectively. None of these differences was statistically significant ( $P > 0.05$ ).

For cases who returned to the final ARIC clinic visit, weighted mean (95% CI) follow up after estimated date of diabetes onset was 4.8 (3.8 – 5.8) years for those who were GADA positive and 4.9 (4.7 – 5.2) for those who were GADA negative.

Among the subjects of the cohort random sample, 52 were GADA positive and 587 were GADA negative, the weighted prevalence of GADA positivity being 7.3%. Among those who went on to develop diabetes, 18 (5.9%; 95% CI = 3.3, 8.6%) whites and 23 (8.3%; 95% CI = 5.1, 11.6%) African-Americans were positive for GADA antibodies. Among GADA-positive subjects, most (55%) had low antibody levels ( $< 2$  U/ml), and only 10% had values greater than 10 U/ml.

Few differences in clinical characteristics were seen according to the presence of GADA among participants of the cohort random sample. Current smoking was less common among GADA-positive individuals ( $P=0.02$ ). Among the markers of inflammation, only IL-6 showed a significant difference, being slightly lower among GADA-positive individuals ( $P=0.01$ ). However, when inflammation was assessed by an aggregate score, no difference was apparent – inflammation score (mean  $\pm$  SEM) of  $2.6 \pm 0.24$  for those GADA-positive subjects vs  $3.0 \pm 0.08$  for GADA-negative subjects ( $P=0.15$ ). Statistical significance of GADA –inflammation marker associations was unchanged after adjusting for differences in smoking prevalence. Additionally, no differences were seen between GADA-positive and -negative individuals in age, gender, race, family history of diabetes, prevalence of hypertension, obesity indices, fasting insulin, glucose, homa insulin sensitivity or beta cell function, triglycerides, HDLcholesterol, leptin, or adiponectin.

To further examine differences in clinical characteristics according to GADA levels, within the limits imposed by a sample of 84 GADA-positive individuals, we compared these characteristics across tertiles of GADA-positive individuals, cutoff points for these tertiles being 1.37 and 2.38 U/ml. Though no difference was seen in obesity measures overall, those with GADA values in the highest tertile had somewhat lower BMI,  $26.1 \pm 1.23 \text{ kg/m}^2$ , and waist circumference,  $92.4 \pm 3.77 \text{ cm}$ , than those GADA-negative ( $27.3 \pm 0.21 \text{ kg/m}^2$  and  $94.9 \pm 0.53 \text{ cm}$ ,  $P=0.35$  and  $0.55$ , respectively). Lowest values for IL-6 ( $2.0 \pm 0.22 \text{ pg/ml}$ ) and the inflammation score ( $2.5 \pm 0.38$ ) were seen in the 3<sup>rd</sup> GADA tertile. White cell count ( $5.3 \pm 0.25$  vs.  $5.9 \pm 0.03 \times 10^9/\text{L}$  in GADA negatives) and orosomucoid ( $78.2 \pm 3.85$  vs.  $87.3 \pm 0.84 \text{ mg/dl}$  in GADA negatives) were also notably lower in this tertile.

Next we evaluated the association of GADA with the development of diabetes using proportional hazard models. The presence of GADA was not associated with the development of diabetes in unadjusted analyses (HR = 0.93; 95% CI = 0.59, 1.48), nor when adjusted for age, gender, ethnicity, family history of diabetes, glucose, body mass index, waist-to-hip ratio, insulin, hypertension and inflammation score (HR = 1.04; 95% CI = 0.55, 1.96). The association was also absent, when analyzed separately, in men and women, blacks and whites, younger and older subjects (45–54 and 55–64 years), and smokers and non-smokers. GADA-positive lean individuals (BMI < 25 kg/m<sup>2</sup>) had a somewhat larger, but not statistically significant risk of developing diabetes compared to GADA-negative lean individuals (HR = 1.23; 95% CI = 0.47, 3.25). Individuals in the highest tertile of GADA (~3% of the cohort random sample) had no greater risk in crude analyses (HR = 1.08, 95% CI = 0.21, 5.49), nor in the fully adjusted model (HR = 1.45; 95% CI = 0.39, 5.42).

To compare the risk factor profile for developing diabetes with and without GADA, we next modeled incident diabetes, separately for GADA-positive and GADA-negative individuals. Table 1 (top) shows that the magnitude of risk is generally similar for each group in univariate analyses, the inflammation score being the only variable demonstrating greater risk in GADA-positive individuals. Since there were only 40 GADA-positive cases, we performed multivariable modeling starting with all available risk factors for diabetes and using backward elimination until 4 variables remained, 3 being statistically significant (WHR, baseline fasting glucose and inflammation score, Table 1, bottom). Modeling with these same risk factors for GADA-negative subjects produced smaller hazard ratios for these three principal risk factors, statistically significantly so for WHR (interaction  $P=0.03$ ). Having a leptin value above the sex-specific median (4.43 ng/mL for males and 17.03 ng/mL for females) was not a risk factor in GADA-positive cases, either in crude or adjusted (HR = 1.54; 95% CI = 0.52, 4.54) analyses. However, leptin, in crude analyses, was highly associated with developing of diabetes among the 53 GADA-positive women (HR = 6.26, 95%CI = 1.61, 24.33), though not statistically so after adjustment for WHR (HR = 2.57, 95%CI = 0.59, 11.18).

To compare clinical progression of GADA-positive and -negative incident cases, we determined case characteristics at the last ARIC follow-up visit (Table 2). Among ascertained cases of diabetes present at this final visit, a somewhat greater percentage of those GADA-positive (59%) than those GADA-negative (49%) subjects reported having received a clinical diagnosis of diabetes (or to be currently using anti-diabetes medication) ( $P=0.32$ ). Anti-diabetes medication use was reported by 30% of GADA-positive and 28% of GADA-negative cases. However, among medication users at Visit 4, 62% of those GADA positive vs. only 9% of those GADA negative were using insulin. Thus, GADA positivity notably increased risk of progressing to insulin use once diabetes was present (HR=9.9; 95%CI=3.4, 28.5), ( $P<0.001$ ). Furthermore, progression to insulin treatment increased with greater baseline GADA values (9.1, 20.0, and 34.6% for those in the 1st, 2nd and 3rd GADA tertiles, vs. 2.4% in those GADA negative; Figure 1).

Additional characteristics of GADA-positive and -negative cases at Visit 4 are presented in Table 2. Comparison of means across all categories revealed no statistically significant differences overall, perhaps due to the very small size of GADA tertile case groups. Nonetheless, at study conclusion, those in the highest tertile ( $\geq 2.38$  U/ml) of GADA at baseline had waist girth 11 cm smaller and BMI 2.7 kg/m<sup>2</sup> less than GADA-negative individuals. Gains in waist and weight from baseline were likewise 4 cm and 3 kg less for those in the highest tertile of GADA. Similar differences in final anthropometric values were seen among those with GADA values between 1.37 and 2.38 U/ml, but not for those with GADA values less than 1.37 U/ml. Among cases not reporting medication use, GADA-positive and GADA-negative individuals had virtually identical values for fasting and 2-h glucose, fasting insulin, triglycerides and HDL cholesterol (data not shown).

## Discussion

Our results demonstrate that approximately 7% of middle-aged ARIC participants were positive for GADA, a value similar to those of previous population-based reports. Those positive and negative had similar levels of most diabetes risk factors. Although a non-significant tendency toward increased risk was seen with highest levels of GADA, antibody positivity was not a risk factor for the development of diabetes among ARIC Study participants. Inflammation, central obesity and high baseline glucose, known to predict type 2 diabetes, were the principal diabetes risk factors for both GADA-positive and -negative cases. Important differences in obesity measures and insulin use were found in diabetes cases with higher GADA levels an estimated 4.8 years after ascertainment, a monotonic increase in insulin use being seen with increasing baseline GADA levels.

We know of no other published study of associations with GADA status in initially non-diabetic subjects. Previously published studies, including population-based ones (19) describe associations with already prevalent diabetes. Two groups, the Diabetes Prevention Project (DPP) (20) and the Botnia Study (21), have recently reported in abstract form that autoantibodies do predict incident diabetes. However both, like our study, did not find an association with a single measurement of GADA. In the DPP, the association was largely explained by the presence, considerably less frequent, of the IA-2 antibody. In the Botnia Study, an association was seen only in younger (<45 year old) subjects with repeated positivity (~1/3 of all those positive) or in those in the upper quartile of GADA positivity. Both the DPP and the Botnia Study extended enrollment to younger participants than did ARIC. Neither presented data exclusively for middle-aged and elderly individuals, age ranges in which the bulk of diabetes cases are detected.

Given what is currently believed about the autoimmune origin of LADA, the relative similarity between GADA-positive and -negative subjects at ARIC baseline, when coupled with absence of risk associated with GADA positivity, is surprising. Previous studies, such as that of the UK Prospective Diabetes Study (UKPDS) (22), have suggested that GADA positivity identifies a subtype of adult onset diabetes of distinct etiology, leaner and with less pancreatic reserve than those with type 2 diabetes. Here too, differences in study populations may explain the apparent discrepancies. The UKPDS enrolled patients aged between 25 and 65 years shortly after clinical diagnosis. By contrast, ARIC incident cases were all over 45, only 22% being ascertained because of clinical diagnosis; and only 59% of GADA-positive cases being clinically diagnosed by the end of follow-up. If one assumes that 10 years is the average lag between the onset of diabetes and its clinical diagnosis, and that some additional lag existed between clinical diagnosis and enrollment in the UKPDS, our GADA-positive cases were ascertained, as a group, more than 5 years earlier in the natural history of their disease than those of the UKPDS. In fact, in the UKPDS, a study of much greater size, differences in the clinical profile among older (>55 years of age) GADA-positive and -negative patients were less than among younger

ones (22). That those GADA-positive in ARIC did not have higher levels of inflammation markers provides further evidence that an autoimmune pancreatitis is not the source of subclinical inflammation characteristic of the pre-diabetic state in adults. That central obesity and subclinical inflammation were important risk factors for diabetes onset for these individuals suggests a pathophysiology for GADA-positive cases quite similar to that of type 2 diabetes. In the majority of GADA-positive individuals who developed diabetes, the autoimmune pancreatitis denoted by GADA, instead of precipitating disease onset, may rather lead to a more rapid pancreatic failure once islet cells are stressed by the diabetic milieu. Indeed, our study, like others, clearly suggests a rapid progression to insulin use in a large fraction of GADA-positive (>1 U/ml) cases.

In this regard, leptin has been suggested in an animal study (23) to accelerate autoimmune diabetes in females. Though crude analysis showed a large risk for high-leptin values in women, our sample of GADA-positive women is too small to permit adequate multivariable investigation of this complex issue.

The lower prevalence of smoking seen in GADA-positive individuals at baseline may be due to random error, as multiple comparisons were made and P values were not far from borderline significance. However, nicotine, whose plasma levels are increased in smokers, has been reported to have anti-inflammatory actions in macrophages and adipocytes (24–26). Its anti-inflammatory effects might protect against autoimmune pancreatitis. In this regard, others have reported that smoking protects against the development of GADA-positive diabetes. Compared with adult Norwegian subjects who had never smoked, the adjusted risk of developing LADA (N=81; RR=0.25; 95% CI=0.11, 0.60) or type 1 diabetes (N=18; RR=0.17; 95% CI=0.04, 0.73), was substantially lower for current (vs never) smokers, whereas the adjusted risk of developing type 2 diabetes was greater (N=738; RR=1.64; 95% CI=1.12, 2.39) for current heavy (vs never) smokers (27). Our data suggest a possible alternative explanation for this reported association – that smoking, by furnishing nicotine, diminishes or abolishes the autoimmune state which leads to GADA positivity, without protecting against the development of diabetes in those who remain GADA positive. Smoking has been repeatedly suggested, if anything, to be a diabetes risk factor (27–29).

Limitations to our study should be noted. The number of GADA-positive individuals and GADA-positive cases of incident diabetes in our sample is relatively small, limiting our ability to describe and contrast clinical characteristics and to perform multivariable modeling. We did not measure additional autoantibodies. In terms of the generalizability of our findings, our GADA-positive cases were considerably obese, as indicated by their waist girth, as a group, of 106.5 cm at the end of follow-up. GADA positivity could possibly be more important in the development of diabetes in middle-aged or elderly populations with less obesity than our ARIC sample. Additional studies on the early natural history of diabetes in middle-aged and elderly GADA-positive individuals are warranted.

In conclusion, in these middle-aged adults, GADA positivity did not increase diabetes risk, and the overall profile of risk factors was similar for positive and negative individuals. These ARIC data suggest that, with the possible exception of those with highest GADA levels, the presumably low-grade pancreatitis reflected specifically by GADA appears to become clinically relevant in middle-aged individuals only after diabetes onset, with more rapid progression to insulin use. Additional prospective studies, including a wider range of antibodies when possible, are necessary to clarify the role of autoimmunity in the onset of diabetes in the middle-aged and elderly.

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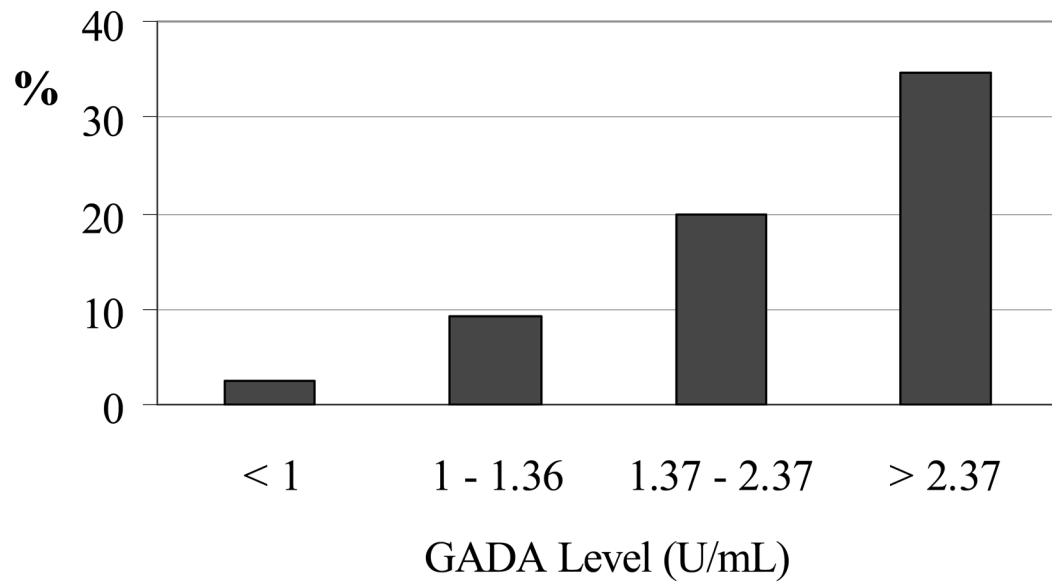
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**Figure 1.** Insulin use at end of follow-up among individuals by presence of and level of glutamic acid decarboxylase antibody (GADA) positivity.

Crude and adjusted associations of risk factors for incident diabetes according to glutamic acid decarboxylase antibody (GADA) positivity, Atherosclerosis Risk in Communities Study, 1987–98.

Table 1

Risk factors	Standard Deviation		GADA Positive (N=84)		GADA Negative (N=1040)		P value*
	HR	95% CI	HR	95% CI	HR	95% CI	
Crude associations							
Age (years)	5.41	1.67	1.07, 2.60	1.17	1.04, 1.32	0.15	
Body mass index (kg/m <sup>2</sup> )	5.64	1.57	1.03, 2.38	1.82	1.55, 2.13	0.53	
Waist-to-hip ratio	0.08	3.07	1.69, 5.57	1.83	1.54, 2.19	0.12	
Insulin (pmol/L) <sup>a</sup>	0.30	2.59	1.34, 4.99	3.18	2.71, 3.74	0.57	
Male		1.24	0.50, 3.12	1.25	0.97, 1.62	0.97	
African-American		2.00	0.91, 4.41	2.12	1.98, 2.28	0.99	
Family history of diabetes		1.43	0.56, 3.67	1.92	1.45, 2.54	0.60	
Hypertension		2.71	1.08, 6.85	2.54	1.97, 3.28	0.86	
Current smoker		1.55	0.46, 5.23	0.90	0.68, 1.21	0.35	
Glucose (mmol/L)	0.51	4.51	2.73, 7.44	3.46	2.91, 4.12	0.36	
Leptin		2.24	0.91, 5.54	1.69	1.32, 2.17	0.52	
Women only		6.26	1.61, 24.33	3.08	2.19, 4.33	0.31	
Inflammation score	1.79	3.05	1.79, 5.20	1.73	1.52, 1.96	0.04	
Adjusted							
Waist-to-hip ratio	0.08	2.20	1.43, 3.38	1.32	1.10, 1.58	0.03	
Hypertension		0.36	0.13, 1.00	1.00	0.71, 1.42	0.38	
Glucose (mmol/L)	0.51	4.33	2.58, 7.27	3.09	2.54, 3.76	0.30	
Inflammation score	1.79	1.82	1.06, 3.12	1.37	1.15, 1.63	0.31	

\* For continuous variables, the hazard ratio (HR) is expressed for a difference of 1 standard deviation of the variable (in the cohort random sample), with the exception of leptin, which is dichotomized as above or below the sex-specific median value. Adjusted associations were obtained through proportional hazards modeling for the four variables with smallest P-values in backward modeling of GADA-positive cases.

<sup>a</sup> log<sub>10</sub>-transformed.

\* For difference in hazard ratio between GADA-positive and -negative individuals.

Follow-up and changes in clinical characteristics of cases by glutamic acid decarboxylase antibody (GADA) status, Atherosclerosis Risk in Communities Study, 1996–98.

**Table 2**

Variable	GADA Negative (N=450)	GADA positive 1 <sup>st</sup> Tertile (N=14)	GADA positive 2 <sup>nd</sup> Tertile (N=10)	GADA positive 3 <sup>rd</sup> Tertile (N=10)	P*
Waist (cm)					
At Follow-up <sup>a</sup>	110.4 ± 0.70	114.4 ± 4.94	101.8 ± 2.93	99.6 ± 4.98	0.41
Difference from Baseline	5.6 ± 0.38	6.6 ± 3.18	6.5 ± 2.40	1.8 ± 3.19	0.53
Weight (kg)					
At Follow-up <sup>a</sup>	90.2 ± 0.85	95.0 ± 6.08	76.6 ± 4.31	82.3 ± 4.54	0.63
Difference from Baseline	3.0 ± 0.38	3.6 ± 3.39	2.5 ± 2.74	-0.16 ± 1.43	0.69
BMI (kg/m <sup>2</sup> )					
At Follow-up <sup>a</sup>	32.1 ± 0.29	33.1 ± 1.94	28.9 ± 1.28	29.4 ± 1.28	0.58
Difference from Baseline	1.3 ± 0.14	1.4 ± 1.16	1.2 ± 1.09	0.04 ± 0.49	0.63

Data are reported as means ± SEM. BMI = body mass index. Cutoff points for GADA tertiles were 1.37 and 2.38.

\* Overall test comparing means between all groups.

<sup>a</sup> Values at follow-up adjusted for the baseline values through analysis of covariance.