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## Weight Gain in Older Adolescent Females: The Internet, Sleep, Coffee, and Alcohol

Catherine S. Berkey, ScD<sup>1</sup>, Helaine R.H. Rockett, MS RD<sup>1</sup>, and Graham A. Colditz, MD DrPH<sup>1,2</sup>

<sup>1</sup>Channing Laboratory, Department of Medicine, Brigham & Women's Hospital and Harvard Medical School

<sup>2</sup>Alvin J. Siteman Cancer Center, Washington University School of Medicine

### Abstract

**Objectives**—To examine whether excessive recreational Internet time, insufficient sleep, regular coffee consumption, or alcoholic beverages promote weight gain.

**Study design**—A longitudinal cohort of >5000 girls (Growing Up Today Study), from all over the US and aged 14–21yrs, returned surveys in 2001 reporting typical past-year recreational Internet time, sleep, coffee (with caffeine) and alcohol consumption. We estimated correlations among these four exposures. Each girl also reported her height and weight in 2000 and again in 2001. Multivariate models investigated associations between 1-year change in body mass index (BMI) and same-year exposures, adjusted for adolescent growth/development, activity and inactivity.

**Results**—The exposures were highly ( $p < 0.0001$ ) correlated with each other, except for coffee with Internet time ( $p > 0.50$ ). More Internet time, more alcohol, and less sleep were all associated ( $p < 0.05$ ) with same-year BMI increases. Females, aged 18+yrs, who slept  $\leq 5$ hrs/night ( $p < 0.01$ ) or who consumed alcohol 2+servings/week ( $p < 0.07$ ) gained more BMI from 2000 to 2001. For females in weight-promoting categories of all exposures, this translates to nearly four extra pounds gained over one year. We found no evidence that drinking coffee promotes weight gain.

**Conclusions**—Older girls may benefit from replacing recreational Internet time with sleep and by avoiding alcohol.

### Keywords

body mass index; BMI; girls; inactivity; beverages; overweight; obesity; longitudinal

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Increases in the prevalence of adolescent overweight during recent decades are extensively documented,<sup>1</sup> as are associated health and social consequences.<sup>2–3</sup> The rapid rise in prevalence suggests environmental factors are responsible.<sup>4</sup> Physical activity has declined, sedentary activities have increased, and widespread changes in dietary patterns have taken place. Admittedly, genes play an important role in the development of obesity,<sup>5</sup> but for any particular individual, his or her genes are not modifiable whereas some environmental factors are, though with difficulty.

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**Correspondence:** Dr. Catherine Berkey, Channing Laboratory, 181 Longwood Ave, Boston MA 02115 (Catherine.Berkey@channing.harvard.edu) Phone: 617-525-0845, FAX: 617-525-2008.  
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A review of longitudinal studies of weight gain and childhood inactivity included TV/video watching and playing videogames, not Internet time or sleep.<sup>6</sup> In a cross-sectional study of adolescent girls, those who spent more time on the computer for email, writing and surfing the Internet were more likely to be overweight.<sup>7</sup> Sleep may be displaced by Internet use; cross-sectional studies of children showed associations between less sleep and overweight.<sup>8-9</sup> A longitudinal study found that British 3 yr-olds with short sleep duration were 45% more likely to be obese at age 7yr,<sup>10</sup> and short sleep duration among US third graders was associated with being overweight in sixth grade.<sup>11</sup> Coffee may be partly responsible for reduced sleep time, and thus weight gain, though a study of adolescent girls found no significant association between coffee/tea intakes and BMI.<sup>12</sup> Alcoholic beverage intakes rise greatly throughout adolescence, at the same time that coffee use is increasing. A cross-sectional study of 16-17yo girls found significant correlation between high alcohol intake and percent body fat.<sup>13</sup>

We estimated the correlations among these exposures in adolescent girls, and investigated jointly their associations with BMI change over time. To our knowledge, no longitudinal studies of adolescents have studied any of these exposures with regard to excessive weight gain.

## METHODS

Established in 1996, the Growing Up Today Study (GUTS) includes 16 771 boys and girls from all 50 states who are children of Nurses' Health Study II (NHSII) participants. The study, approved by Human Subjects Committees at Harvard School of Public Health and Brigham and Women's Hospital, is described elsewhere.<sup>14</sup> Mothers provided informed consent, and their children assented by completing baseline questionnaires. The cohort, aged 9-14 years in 1996, returned follow-up questionnaires annually through 2001. Girls' responses to one or more of these follow-ups were 95.5%; more relevant to this analysis is that 70% returned the 2001 survey that asks about all four exposures. A total of 5502 girls returned surveys in both 2000 and 2001, the source of data for these analyses.

Children reported their heights and weights in 2000 and again in 2001. Our questionnaire provided specific measuring instructions but suggested they seek assistance; their mothers (nurses) biennially self-report their own heights and weights for NHSII. We assessed relative weight status by computing body mass index ( $BMI = \text{weight}/\text{height}^2$ , ( $\text{kg}/\text{m}^2$ )). The validity of self-reported BMI is supported by National Longitudinal Study of Adolescent Health analyses that found a correlation of 0.92 between BMI computed from measured values and from self-reports by youth in grades 7-12.<sup>15</sup> Our outcome measure is change in relative body weight,  $BMI_{2001} - BMI_{2000}$ , divided by the time interval (to the month) between survey return dates. After excluding outliers, we had 5,036 girls with annualized BMI changes from 2000 to 2001. Analyses of BMI changes are superior to analyzing differences between BMI Z-scores or percentiles.<sup>16</sup>

### Recreational Internet Time

In 2001, we asked the participants how many hours/week they typically spent on Internet/Computers, not including schoolwork; our 7 response options ranged from none to 31+hours/week. (The preceding question regarded computer/video game time). A study on self-reported computer use in Minnesota middle school children demonstrated adequate reliability (a week apart, test with re-test correlation=0.60) and validity (test with seven-day log, correlation=0.39).<sup>17</sup>

### Sleep

Our 2001 survey asked “On a typical night when you have school or work the next day, how many hours of sleep do you get?” Eight response options ranged from “less than 5 hours” to “11 or more hours”. Validation studies on adolescents indicate they can validly self-report sleep time.<sup>18</sup>

### Coffee

Average past year coffee (“not decaf”) use was obtained by the 2001 survey. Eight response options ranged from Never to >3 cups/day. Self-reported intakes for beverages in general, and coffee specifically, had high validity in a study on adult women.<sup>19</sup>

### Alcohol

The 2001 survey inquired about typical past year intakes of beer, wine, and liquor. A review article, on the validity of adolescent self-reports of risky behaviors, concluded that the privacy of self-administered questionnaires produces higher, supposedly more valid, reported rates of alcohol use.<sup>20</sup> Studies that incorporated (bogus) saliva tests along with self-reports suggested that adolescents provided valid responses. Test-retest reliability levels were generally high for alcohol, but it is difficult to do validation studies of long-term alcohol use due to the lack of biochemical measures.<sup>20</sup> The validation study we cited for coffee also supported high validity for alcoholic beverages.<sup>19</sup>

### Statistical Analyses

We computed Pearson correlations among the four exposures collected on our 2001 survey. Any girl who reported that she was a smoker (>100 cigarettes lifetime, and at least one within the past month) or who was pregnant (currently or in past year) was excluded, from these correlations and from all models outlined below, leaving N=4427 girls for analysis of BMI change.

For analyzing relative weight change from 2000 to 2001, we fitted multivariate linear regression models on the 4 exposures simultaneously (N=4427 girls). All models were estimated using generalized estimating equations to allow for correlations among the small number of sisters in the cohort.<sup>21</sup> We used a continuous measure of each exposure (typical past-year cups/day of coffee, etc.) reported in 2001 to estimate linear associations with BMI change from 2000 to 2001. In a separate model, we used a categorical version of each exposure, in order to observe and estimate any nonlinear trends. Both models adjusted for adolescent growth and development by including age (age-18.0) and age<sup>2</sup>, menstrual status (2000 and 2001), height growth from 2000 to 2001, and prior BMI (2000). Fully adjusted models further included past-year physical activity and TV/videos/computer-games,<sup>14</sup> reported in 2001. Finally, these fully-adjusted models were fit to only the oldest (18+ yrs) girls (N=1424).

## RESULTS

Nearly all participants, daughters of NHSII nurses, are white (95%). At baseline, 12.7% were overweight (85<sup>th</sup>–95<sup>th</sup> percentile on CDC BMI charts), 4.8% were obese (>95<sup>th</sup> percentile), and 4.7% were very lean (<5<sup>th</sup> percentile). In 2001, when data were collected on our 4 exposures, 16% of the girls were aged 14–15yr, 38% were 16–17yr, 34% were 18–19yr, and 12% were 20–21yr. At that time, 52% of girls were spending 1–5 hours/week on the computer for recreation (34%, even more hours), only 31.3% were sleeping 8 or more hours/night, 36% drank coffee, and 35% drank alcohol. The mean annual BMI gains, from year 2000 (when girls were aged 13–20yr) to 2001, declined with age, except for age 18 to 19yr (+0.47kg/m<sup>2</sup>), which

may reflect weight increases as girls begin college. Beginning at age 18yr, mean alcohol intakes (0.22 servings/day) exceeded mean coffee intakes (0.15 servings/day).

Correlations among the 4 exposures were highly significant, except for coffee with Internet time ( $R=+0.009$ ,  $p=0.54$ ), demonstrating the importance of joint multivariate analyses of these exposures in relation to weight gain. The other correlations (all  $P<0.0001$ ) were: Internet with sleep  $R=-0.093$ , Internet with alcohol  $R=+0.086$ , sleep with coffee  $R=-0.102$ , sleep with alcohol  $R=-0.116$ , and coffee with alcohol  $R=+0.111$ .

Girls spending more time on the Internet had significantly greater BMI increases during the same year (Table I, top row), whether Internet time was measured categorically (6–10hrs/wk and 16+hrs/week) or continuously (hours/day). When we further included past year physical activity and TV/videos/computer-games in the model, the Internet effect was weaker. Adjusting for TV may be over-control as girls watch television at the same time as checking email, Instant Messaging or surfing the Internet. For girls aged 18+ years, the Internet effects were further weakened and not statistically significant. Older girls are likely also spending substantial non-recreational time on the computer for college or employment (data not collected), which may have diminished our estimates.

Associations between sleep and BMI gain were summarized similarly (Table II). Estimates shown in the top row were obtained from the same two multivariate models that provided estimates for Table I, Table III, and Table IV. A significant inverse trend was present even after adjusting for physical activity and TV/videos/computer-games. The strongest effects were on the older girls (18+). Those who slept 5 or fewer hours/night gained significantly more BMI ( $+0.322$  kg/m<sup>2</sup>) during the year compared with those sleeping 8 hours.

None of the coffee models (Table III) provided evidence that drinking coffee promotes weight gain. In fact, though not significant, the bulk of the evidence (see 5+cups/week and continuous model estimates) suggested protection against weight gain.

Girls who typically consumed 2+servings/week of alcoholic beverages gained significantly more weight than those who consumed the least (Table IV). Statistically significant linear trends were observed, as well. For the older girls (18+yrs), the estimated effect ( $\beta =+0.115$ ) for the 2+servings/week group, though only marginally significant ( $p<0.07$ ), was larger than the significant effect ( $\beta =+0.108$ ) from the full group of girls (both older and younger than 18yr).

To illustrate the effect of modifying all 4 exposures simultaneously, we compare 2 hypothetical girls who are identical (age 19, same prior BMI, same physical activity and same TV/videos/computer-games) except in the four exposures. One girl had no recreational Internet time, slept 8 hours nightly, consumed no alcohol, and drank a cup of coffee daily. The other girl was on the Internet 6–10 hours/week, slept 5 hours/night, drank no coffee but regularly consumed 2+ alcohol servings/week. The latter girl gained (according to the categorical exposures model, bottom of Table I, Table II, Table III, and Table IV) +0.65 more BMI over 1 year. For a 19-yr-old of average height and weight, this represents a nearly 4 pound weight gain.

## DISCUSSION

Adolescent girls who, between 2000 and 2001, spent more recreational time on the Internet, slept less, or consumed more alcohol, gained more BMI during that year. Although the estimated effects were small, they may accumulate over time and become clinically important as these behaviors are maintained, or more likely magnified, as girls approach adulthood. We found no evidence that coffee consumption promoted weight gain.

Internet time likely promotes weight gain through increased sedentary time, reducing total energy expenditure. Our findings are consistent with a cross-sectional study in which adolescent girls who spent more recreational time on the Internet were more likely to be overweight.<sup>7</sup> The Internet can also be beneficial, as in Internet-aided programs for weight management in children.<sup>22</sup>

Our findings regarding sleep are consistent with longitudinal studies on younger children in which short sleep duration preceded overweight.<sup>10–11</sup> Insufficient sleep might result in daytime fatigue and reduced physical activity, with fewer calories expended during the day, though Patel found no evidence to support this in adult women.<sup>23</sup> Sleep restriction has metabolic effects that may predispose humans to weight gain and to diabetes.<sup>24</sup> A longitudinal study of adults also showed short sleep duration preceding onset of overweight.<sup>23</sup> A recent paper recommended more sleep for children to prevent obesity.<sup>25</sup>

These analyses did not support our hypothesis, that coffee promotes weight gain in adolescent girls, but instead were consistent with a cross-sectional study finding no association.<sup>12</sup> Because we had no information available on the use of cream, milk, or sugar in coffee, this affects the validity of inferences we can make regarding coffee and weight gain in our cohort. This may partly explain why our findings did not confirm adult studies indicating that coffee consumption may prevent weight gain.<sup>26–28</sup>

Liquid (beverage) calories may not be as well compensated for in the overall diet, by reduced subsequent food/beverage intakes, as calories from food. Thus, beverage intakes (including beer, wine and liquor) are more likely to result in higher total energy intakes, and weight gain. A prospective analysis of women showed the heaviest drinkers gained the most weight,<sup>29</sup> and a longitudinal study of older Danish women found that higher alcohol intakes were associated with 5-yr increases in waist circumference.<sup>30</sup> A recent cross-sectional analysis of 16–17yo girls found significant correlation between high alcohol intake and percent body fat.<sup>13</sup> Our results extend to the literature the positive association between weight gain over time and alcohol consumption in adolescent girls.

A major strength of this analysis is the measurement, on a large cohort of girls from all over the US, of height and weight in two consecutive years. Longitudinal observational studies such as ours cannot determine causality as validly as randomized controlled trials, but our study design is superior to cross-sectional studies, where associations may represent reverse causality. Though we controlled for many potential confounders in our models, some residual and unmeasured confounding may remain. We cannot exclude the possibility of incomplete adjustment of some covariates, or confounding through variables not considered, such as employment-related physical activity or inactivity. We omitted from these analyses girls who were pregnant or smokers. The major limitation of our study was the necessity to collect data by self-report on mailed questionnaires, but with our large cohort of participants residing all over the US, alternatives were not feasible. Errors in reporting sleep, Internet time, and coffee are likely to be non-differential with respect to weight gain, so that any resulting bias should produce underestimates of the strength of true associations. Regarding alcohol consumption, assurances that we will not release information to parents should help maintain accuracy. But to the extent that some participants may under-report alcohol intake, this also should bias estimates toward the null, obscuring real relationships between alcohol and weight gain. Although our cohort is not representative of US girls, associations among factors within our cohort should still be valid. Because all participants are daughters of nurses, this reduces confounding by socioeconomic and other unmeasured factors, as well as enhances the accuracy of the information provided. But the racial and ethnic makeup of our cohort (95% white) is clearly a limitation to generalizing our findings.

If future studies confirm our findings, that more sleep, less recreational Internet time, and less alcohol may prevent excessive weight gain or promote weight loss, then adolescents could safely and inexpensively use this information to manage their body weight. Aside from preventing obesity, these findings may encourage youth to make lifestyle changes that provide educational and other health benefits, as well.

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## Abbreviations

BMI, body mass index; GUTS, Growing Up Today Study; NHSII, Nurses' Health Study II.

## REFERENCES

- Ogden CL, Carroll MD, Curtin LR, McDowell MA, Tabak CJ, Flegal KM. Prevalence of overweight and obesity in the United States, 1999–2004. *JAMA* 2006;295:1549–1555. [PubMed: 16595758]
- Gortmaker SL, Must A, Perrin JM, Sobol AM, Dietz WH. Social and economic consequences of overweight in adolescence and young adulthood. *N Engl J Med* 1993;329:1008–1012. [PubMed: 8366901]
- Must A, Jacques PF, Dallal GE, Bajema CJ, Dietz WH. Long-term morbidity and mortality of overweight adolescents: a follow-up of the Harvard Growth Study of 1922 to 1935. *N Engl J Med* 1992;327:1350–1355. [PubMed: 1406836]
- Hill JO, Peters JC. Environmental contributions to the obesity epidemic. *Science* 1998;280:1371–1374. [PubMed: 9603719]
- Gesta S, Bluher M, Yamamoto Y, Norris AW, Berndt J, Kralisch S, et al. Evidence for a role of developmental genes in the origin of obesity and body fat distribution. *Proc Natl Acad Sci USA* 2006;25(103):6676–6681. [PubMed: 16617105]
- Must A, Tybor DJ. Physical activity and sedentary behavior: a review of longitudinal studies of weight and adiposity in youth. *Int J Obes* 2005;29:S84–S96.
- Kautiainen S, Koivusilta L, Lintonen T, Virtanen SM, Rimpela A. Use of information and communication technology and prevalence of overweight and obesity among adolescents. *Int J Obes (Lond)* 2005;29:925–933. [PubMed: 15925961]
- Knutson KL. Sex differences in the association between sleep and body mass index in adolescents. *J Pediatr* 2005;147:830–834. [PubMed: 16356441]
- Chaput JP, Brunet M, Tremblay A. Relationship between short sleeping hours and childhood overweight/obesity: results from the 'Quebec en Forme' Project. *Int J Obes (Lond)* 2006;30:1080–1085. [PubMed: 16534525]
- Reilly JJ, Armstrong J, Dorosty AR, Emmett PM, Ness A, Rogers I, et al. Early life risk factors for obesity in childhood: cohort study. *BMJ* 2005;330:1357–1363. [PubMed: 15908441]
- Lumeng JC, Somashekar D, Appugliese D, Kaciroti N, Corwyn RF, Bradley RH. Shorter sleep duration is associated with increased risk for being overweight at ages 9 to 12 years. *Pediatrics* 2007;120:1020–1029. [PubMed: 17974739]
- Striegel-Moore RH, Thompson D, Affenito SG, Franko DL, Obarzanek E, Barton BA, et al. Correlates of beverage intake in adolescent girls: the National Heart, Lung, and Blood Institute Growth and Health Study. *J Pediatr* 2006;148:183–187. [PubMed: 16492426]

13. Vagstrand K, Barkeling B, Forslund HB, Elfhag K, Linne Y, Rossner S, et al. Eating habits in relation to body fatness and gender in adolescents – results from the ‘SWEDES’ study. *Eur J Clin Nutr* 2007;61:517–525. [PubMed: 17006444]
14. Berkey CS, Rockett HRH, Gillman MW, Colditz GA. One year changes in activity and in inactivity among 10 to 15 year old boys and girls: Relationship to change in BMI. *Pediatrics* 2003;111:836–843. [PubMed: 12671121]
15. Goodman E, Hinden BR, Khandelwal S. Accuracy of teen and parental reports of obesity and body mass index. *Pediatrics* 2000;106:52–58. [PubMed: 10878149]
16. Berkey CS, Colditz GA. Adiposity in adolescents: Change in actual BMI works better than change in BMI Z Score for longitudinal studies. *Ann Epidemiol* 2007;17:44–50.
17. Schmitz KH, Harnack L, Fulton JE, Jacobs DR Jr, Gao S, Lytle LA, et al. Reliability and validity of a brief questionnaire to assess television viewing and computer use by middle school children. *J Sch Health* 2004;74:370–377. [PubMed: 15656264]
18. Wolfson A, Carskadon M, Acebo C, Seifer R, Fallone G, Labyak S, et al. Evidence for the validity of a sleep habits survey for adolescents. *Sleep* 2003;26:213–216. [PubMed: 12683482]
19. Salvini S, Hunter DJ, Sampson L, Stampfer MJ, Colditz GA, Rosner B, et al. Food-based validation of a dietary questionnaire: the effects of week-to-week variation in food consumption. *Int J Epidemiol* 1989;18:858–867. [PubMed: 2621022]
20. Brener N, Billy J, Grady W. Assessment of factors affecting the validity of self-reported health-risk behavior among adolescents: evidence from the scientific literature. *Journal of Adolescent Health* 2003;33:436–457. [PubMed: 14642706]
21. Diggle, PJ.; Liang, KY.; Zeger, SL. *Analysis of Longitudinal Data*. Oxford: Clarendon Press; 1994.
22. Williamson DA, Walden HM, White MA, York-Crowe E, Newton RL Jr, Alfonso A, et al. Two-year Internet based randomized controlled trial for weight loss in African-American girls. *Obesity (Silver Spring)* 2006;14:1213–1243.
23. Patel SR, Malhotra A, White DP, Gottlieb DJ, Hu FB. Association between reduced sleep and weight gain in women. *Am J Epidemiology* 2006;164:947–954.
24. Currie A, Cappuccio FP. Sleep in children and adolescents: A worrying scenario. Can we understand the sleep deprivation—obesity epidemic? *Nutr, Metab & Cardio Dis* 2007;17:230–222.
25. Taheri S. The link between short sleep duration and obesity: should we recommend more sleep to prevent obesity. *Arch Dis Childhood* 2006;91:881–884. [PubMed: 17056861]
26. Westerterp-Plantenga M, Diepvens K, Joosen AM, Berube-Parent S, Tremblay A. Metabolic effects of spices, teas, and caffeine. *Physiol Behav* 2006;89:85–91. [PubMed: 16580033]
27. Lopez-Garcia E, van Dam RM, Rajpathak S, Willett WC, Manson JE, Hu FB. Changes in caffeine intake and long-term weight change in men and women. *Am J Clin Nutr* 2006;83:674–680. [PubMed: 16522916]
28. Greenberg JA, Boozer CN, Geliebter A. Coffee, diabetes, and weight control. *Am J Clin Nutr* 2006;84:682–693. [PubMed: 17023692]
29. Wannamethee SF, Field AE, Colditz GA, Rimm EB. Alcohol intake and 8-year weight gain in women: a prospective study. *Obes Res* 2004;12:1386–1396. [PubMed: 15483203]
30. Halkjaer J, Tjonneland A, Thomsen BL, Overvad K, Sorensen TI. Intake of macronutrients as predictors of 5-y changes in waist circumference. *Am J Clin Nutr* 2006;84:789–797. [PubMed: 17023705]

Past-yr recreational Internet use (2001 report) and its estimated association ( $\beta$ , se) with BMI change ( $\text{kg/m}^2$ ), 2000 to 2001. Models of categorical Internet time appear in middle 5 columns, and continuous Internet models are on right. Models adjust for age, age<sup>2</sup>, menarche (2000–2001), height growth (2000–2001), 2000 BMI.

TABLE 1

Categories of Time for Internet Use	None (referent)	1–5hrs/wk	6–10hrs/wk	11–15hrs	16+hrs/wk	Continuous Internet Time $\beta$ (se) ( $\text{kg/m}^2$ per hr/day)
<b>Adjust for:</b>						
Sleep, Coffee, Alcohol	0	+077 (.047)	+164 <sup>**</sup> (.056)	+061 (.069)	+198 <sup>*</sup> (.082)	+045 <sup>*</sup> (.021)
<b>Further adjust for:</b>						
Past Year Physical Activity, TV/ Videos, Games	0	+069 (.048)	+150 <sup>**</sup> (.056)	+042 (.070)	+157 <sup>#</sup> (.084)	+034 (.022)
<b>Only Girls Aged 18+</b> (height growth and menarche omitted)	0	+011 (.085)	+116 (.097)	+023 (.123)	+107 (.135)	+038 (.035)

# 0.05<P<0.10

\* P<0.05

\*\* P<0.01



Usual sleep (hours/night), reported in 2001, and its estimated association ( $\beta$ , se) with BMI change ( $\text{kg}/\text{m}^2$ ) from 2000 to 2001. Models of categorical sleep time appear in middle five columns, and continuous sleep models are on right. Models adjust for age, age<sup>2</sup>, menarche (2000–2001), height growth (2000–2001), and 2000 BMI.

TABLE II

Categories of Sleep	$\leq 5$ hrs	6 hours	7 hours	8 hours (referent)	9+ hours	Continuous Sleep $\beta$ (se) ( $\text{kg}/\text{m}^2$ per hr/night)
<b>Adjust for:</b> Internet, Coffee, Alcohol	+133 <sup>#</sup> (.076)	+059 (.046)	-025 (.039)	0	-028 (.073)	-.034 <sup>*</sup> (.017)
<b>Further adjust for:</b> Past Year Phys Act, TV/videos, Games	+140 <sup>#</sup> (.075)	+061 (.047)	-020 (.039)	0	-023 (.073)	-.034 <sup>*</sup> (.016)
<b>Only Girls Aged 18+</b> (height growth and menarche omitted)	+322 <sup>**</sup> (.112)	-021 (.076)	-049 (.067)	0	-134 (.121)	-.076 <sup>**</sup> (.026)

<sup>#</sup> 0.05 < P < 0.10

<sup>\*</sup> P < 0.05

<sup>\*\*</sup> P < 0.01

Past-yr coffee intake (not decaf, 2001 report) and its estimated association ( $\beta$ , se) with BMI change ( $\text{kg}/\text{m}^2$ ), 2000 to 2001. Models of categorical coffee intake appear in middle five columns, and continuous coffee models are on right. Models adjust for age, age<sup>2</sup>, menarche (2000–2001), height growth (2000–2001), and 2000 BMI.

TABLE III

Categories of Coffee Intake	None (referent)	1–3cups/mo	1/wk	2–4/wk	5+cups/wk	Continuous Coffee $\beta$ (se) ( $\text{kg}/\text{m}^2$ per cup/day)
<b>Adjust for:</b> Internet, Sleep, Alcohol	0	+033 (.044)	–006 (.061)	+078 (.077)	–089 (.071)	–035 (.056)
<b>Further adjust for:</b> Past Year Physical Activity, TV/videos/games	0	+030 (.044)	–008 (.061)	+070 (.077)	–088 (.072)	–031 (.056)
<b>Only Girls Aged 18+</b> (height growth and menarche omitted)	0	–022 (.071)	–051 (.092)	+169 (.120)	–097 (.095)	–032 (.068)

# 0.05<P<0.10

\* P<0.05

\*\* P<0.01

**TABLE IV**

Past-yr alcohol consumption (2001 report) and its estimated association ( $\beta$ , se) with BMI change ( $\text{kg}/\text{m}^2$ ), 2000 to 2001. Models of categorical alcohol intake appear in middle three columns, and continuous alcohol models appear on far right. Models adjust for age, age<sup>2</sup>, menarche (2000–2001), height growth (2000–2001), and 2000 BMI.

Categories of Alcohol Intake	0–3 servings/mo (referent)	1/wk	2+/week	Continuous Alcohol $\beta$ (se) ( $\text{kg}/\text{m}^2$ per serv/day)
<b>Adjust for:</b> Internet, Sleep, Coffee	0	+0.050 (.063)	+0.114* (.047)	+0.113* (.048)
<b>Further adjust for:</b> Past Year Physical Activity, TV/ videos/Games	0	+0.047 (.064)	+0.108* (.048)	+0.111* (.048)
<b>Only Girls Aged 18+</b> (height growth and menarche omitted)	0	+0.036 (.086)	+0.115# (.063)	+0.064 (.054)

# 0.05 < P < 0.10

\* P < 0.05

\*\* P < 0.01