

Maintaining Continuity of Care for Nursing Home Residents: Effect of States' Medicaid Bed-Hold Policies and Reimbursement Rates

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Objective. Recent public concern in response to states' intended repeal of Medicaid bed-hold policies and report of their association with higher hospitalization rates prompts examination of these policies in ensuring continuity of care within the broader context of Medicaid policies.

Data Sources/Study Design. Minimum Data Set assessments of long-stay nursing home residents in April–June 2000 linked to Medicare claims enabled tracking residents' hospitalizations during the ensuing 5 months and determining hospital discharge destination. Multinomial multilevel models estimated the effect of state policies on discharge destination controlling for resident, hospitalization, nursing home, and market characteristics.

Results. Among 77,955 hospitalizations, 5,797 (7.4 percent) were not discharged back to the baseline nursing home. Bed-hold policies were associated with lower odds of transfer to another nursing home (AOR = 0.55, 95 percent CI 0.52–0.58) and higher odds of hospitalization (AOR=1.36), translating to 9.5 fewer nursing home transfers and 77.9 more hospitalizations per 1,000 residents annually, and costing Medicaid programs about \$201,311. Higher Medicaid reimbursement rates were associated with lower odds of transfer.

Conclusions. Bed-hold policies were associated with greater continuity of NH care; however, their high cost compared with their small impact on transfer but large impact on increased hospitalizations suggests that they may not be effective.

Key Words. Medicare, multilevel models, hospitalizations, Minimum Data Set (MDS), relocation

The Omnibus Budget Reconciliation Act, or Nursing Home Reform Act of 1987, mandated that nursing homes adopt a bed-hold policy during acute hospitalization of residents. For a fee, nursing homes reserve the resident's bed

during acute hospitalizations. The goals of this policy were twofold: to prevent facilities from discharging “difficult” or expensive patients, and to encourage continuity of residence for the nursing home client (Taylor 2004). Some state Medicaid programs assumed the responsibility for such payments for eligible nursing home residents in various ways, while other states did not.

Recent budget difficulties have led several states to decrease or eliminate bed-hold payments (Maryland 1999; Williams 2003; Correira 2004). Massachusetts has been in a flux regarding its bed-hold policies, changing the number of days and the rates paid several times, discontinuing payment for any bed-hold in 2003, and reinstating it 2 years later. The most attention to the flux in bed-hold policies was garnered in Florida, home to the second largest population of Americans older than 65, where the Medicaid program made \$23 million in bed-hold payments annually in 2004 and 2005, roughly the equivalent of 4 days a year per Medicaid nursing home resident (Florida 2005). As statewide nursing home occupancy was under 90 percent, it was argued that bed-hold payments were superfluous, and could be eliminated (Freeman 2004; Nohlgren 2004). A similar argument was made recently in Pennsylvania where policy makers were considering the removal of bed-hold payments (Rotstein 2006).

Nursing home administrators counter that occupancy rates vary by region and by facility, and that elimination of bed-hold payments would increase hospital discharges to facilities other than the resident’s original “home,” with potential consequences of relocation stress, including disorientation and dissatisfaction (Freeman 2004; Nohlgren 2004).

Historically, nursing home residents have taken legal action to prevent the closure of a substandard facility in order to avoid relocation, clearly implying social value to familiar surroundings (Cohen 1986). There is some literature suggesting that “relocation trauma” has a lasting negative effect on nursing home residents (Friedman et al. 1995; Capezuti et al. 2006; Laughlin et al. 2007). For example, change of venue alone has transiently increased fall rates (Friedman et al. 1995). Relocation may also lead to the duplication of tests and increased likelihood of medical errors (Boockvar et al. 2004; Coleman et al. 2004; Ma et al. 2004), both of which may be associated with

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morbidity and cost. A recent Presidential Council of Bioethics report emphasizes that hospitalizations at the end of life, a time at which more comfort and care are needed, are particularly traumatic to nursing home residents and families, exposing them to the possibility of relocation not only to a hospital, but to other nursing homes (Kass 2005).

Although the rationale for bed-hold policies was to facilitate residents' return to their original nursing home, this effect has never been demonstrated. Hospital discharge destination is more complex than bed availability alone, and may be affected by other patients, nursing homes, market, and state factors, as has been shown for other transitions (Intrator et al. 2004, 2007). For specific diagnoses, and for sicker residents, discharge destinations other than nursing homes must be considered. Some patients or families may choose a different facility. For some patients, hospitalization may indicate that their condition is too severe to be managed in the originating nursing home, and indeed, possibly in any nursing home. Residents discharged with rehabilitation needs, such as following a hospitalization for hip fracture, stroke, or psychiatric diagnoses, may have limited facilities that can more appropriately serve them. Facilities with more available beds, advanced clinical capabilities, or higher staffing levels may be more likely to readmit their hospitalized residents. Within a given market, the availability of alternative nursing homes may affect discharge locale. Moreover, differences in the hospitalization rates of nursing home residents have been reported to vary with state policies (Intrator and Mor 2004; Nohlgren 2004). In particular, residents in states with any bed-hold payment policies were reported to have higher odds of being hospitalized (Intrator et al. 2007). Higher Medicaid reimbursement rates have been associated with fewer hospitalizations (Intrator et al. 2007) possibly because higher reimbursement allows facilities to invest in infrastructure and staffing that facilitate care within the nursing home environment as demonstrated in several papers (Intrator et al. 2005; Feng et al. 2008). The purpose of this paper was therefore to study the intended effect of bed-hold policies within the context of today's nursing home market, in light of the reported unintended consequences of associated increased hospitalization rates, while controlling for Medicaid reimbursement rates.

METHODS

Data Sources

We obtained resident characteristics from the Minimum Data Set (MDS), a federally mandated assessment for all nursing home residents that includes

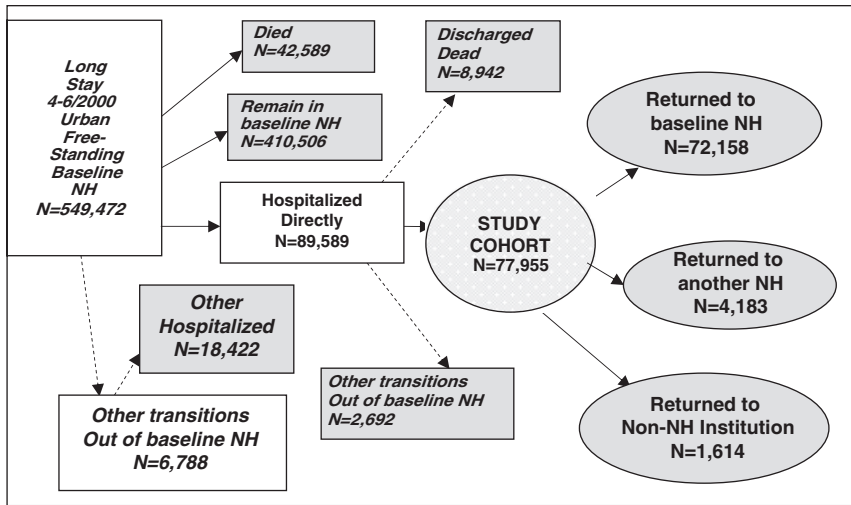
approximately 400 data elements, including demographics, diagnoses, treatments and measures of both physical and cognitive functions (Morris et al. 1990; Hawes et al. 1995). We used the Centers for Medicare and Medicaid Services' (CMS's) Standard Analytic Files for part A claims for inpatient hospital, skilled nursing facility (SNF), hospice, home health, and outpatient services to determine other locations of care using the Residential History File methodology (Intrator et al. 2003). The CMS Online Survey Certification and Reporting (OSCAR) system provided nursing home data. We aggregated OSCAR data and used the Area Resource File to obtain market data (Stambler 1988). State Medicaid policies for the period of our study were available from a survey of the 48 contiguous states (Grabowski et al. 2004). We matched all data at the level of the resident, and performed our evaluation at the level of the hospitalization.

Subject Sample

We used MDS data from the second quarter of calendar year 2000 from free-standing urban nursing homes to identify a cohort of all long-stay (>90 days) residents. This established the nursing home of origin for each of 549,472 residents.

We matched residents to Medicare claims to identify the first acute hospitalization over the ensuing 5 months ($N = 97,141$). We limited our analysis to 5 months to avoid hospitalizations that extended beyond the calendar year 2000. When hospitalization at one hospital was immediately followed by hospitalization at another, we considered this as a single hospitalization. We excluded hospitalizations in long-term hospitals or rehabilitation facilities ($N = 6,067$), and hospitalizations not definitively originating from the baseline nursing home ($N = 1,485$). Hospitalizations with discharges due to death were excluded as mortality was not relevant to the purpose of this study ($N = 8,942$). Finally, hospitalizations with no matching resident, nursing home, or county information were excluded, and ambiguous discharge locations were censored ($N = 2,692$). The final cohort included 77,955 residents served in 8,652 facilities in 812 counties in the 48 contiguous U.S. states. Figure 1 presents the framework for this paper in which the first hospitalization of long-stay residents was identified and followed for hospital discharge destination. Sample selection is denoted on the left part of the flow chart, and study outcomes on the right.

Figure 1: Flow chart of study sample selection and outcomes. Rectangular nodes relate to sample selection, among which those with a darker background represent cases omitted from analyses. Oval nodes represent outcomes. The circular node in the middle represents the final cohort. Of 549,472 long-stay residents of free-standing urban nursing homes, in the second quarter of 2000, 77,955 were hospitalized directly from their original nursing home and entered the study. Of cohort members, 72,158 were discharged back to the original nursing home from the hospital; 4,183 were discharged to another nursing home; and 1,614 were discharged to another institution.



Outcomes

We used the MDS assessment dates, Medicare claims, and discharge location information to determine each resident’s first location within 7 days after hospital discharge using the Residential History File methodology (Intrator et al. 2003). We considered three possibilities for discharge destination: the original nursing home, a different nursing home, and a nonnursing home institutional provider (including long-term hospitals, rehabilitation facilities or home hospice-care). Figure 1 presents these outcomes.

Variable Definitions

The main research question in this study was to examine the effect of two state policies on the post-hospital discharge location of long-stay nursing home res-

idents. Any bed-hold policy was measured by a state-specific indicator. Medicaid per-diem reimbursement was measured as a continuous variable of the average Medicaid per-diem rate (total Medicaid payments to nursing homes divided by total bed-days paid for), standardized at its mean among the 48 states (\$103.30), and in steps of \$10 (1/2 SD). We note that bed-hold policies included maximum number of days paid within a prespecified period of time, the proportion of the regular per-diem paid, and whether there were occupancy standards that facilities were to meet in order to be paid for bed-hold. Table 1 lists the three state-bed-hold parameters: specific maximal days, the period during which those days were counted, and the proportion of the Medicaid per-diem rate paid. There was no correlation among any of these parameters and the per-diem reimbursement rate. Equivalent Reimbursement Days (ERD), a product of the maximum annual days for bed-hold and proportion of rate paid, incorporating the three parameters of bed-hold policies to a standardized measure, was also not correlated with Medicaid per-diem.

Confounders to the investigated relationship between bed-hold policies and discharge destination are described in Table 2. These included measures of inpatient acuity from the inpatient claims: inpatient length of stay, several categories of primary diagnoses, Elixhauser comorbidity index (Elixhauser et al. 1998), and an indicator of intensive care unit (ICU) use. Resident level controls relating to resident condition or potential preferences were obtained from the MDS. Characteristics of the discharging nursing home obtained from the OSCAR data included measures of structure, markers of facility revenue, low occupancy (< 85 percent) as an indicator of ready bed supply for a return the availability of a dementia unit and a rehabilitation unit (which may potentially offer opportunities for returning to the baseline facility), and two measures of nursing home casemix acuity of the admitted and long-stay residents¹ (Fries et al. 1994; Feng et al. 2006). Finally, we controlled for nurse and medical staffing levels, at the county level, we controlled for nursing home competition (Nyman 1989), assuming that return to the baseline nursing home would be more likely in competitive markets (Grabowski 2002; Mukamel et al. 2005; Castle et al. 2007), and average hospital occupancy rate in the county, indicating a pressure to discharge residents quicker, potentially leading to more transitions. Local wage index was used to control for geographical differences in the cost of living.

Multilevel Model

We constructed a hierarchical, four-level multinomial model using *MLwiN* (Centre for Multilevel Modeling, Institute of Education, London, U.K.). In this

Table 1: States' Medicaid Policies and Distribution of Hospital Discharge Destination

State Policies				Hospital Discharge Destination per Resident												
Per-Diem (\$)	% Per-Diem	Max Days	Time Frame	Bed-Hold Policies		Occupancy Requirement	ERD	All States	Total		Baseline NH		Other NH		Non NH	
				N/A	N/A				N	N	N	N	N	N	%	%
103.30	64.8	16.6	N/A	N/A	N/A	N/A	16.0	All States	77,955	72,158	92.5	4,783	5.4	1,614	2.1	0.8
112.54	75	4	Hosp		No	3.6	Alabama	1,322	1,280	96.8	31	2.3	11	0.8		
99.57	100	9	Cal		No	9.0	Arizona	265	238	89.8						
69.36	100	5	Hosp		Yes	6.0	Arkansas	566	506	89.4	49	8.7	11	1.9		
110.27	100	7	Hosp		No	8.4	California	5,556	4,683	84.3	655	11.8	218	3.9		
111.62	0	0	NA		NA	0.0	Colorado	432	398	92.1	22	5.1	12	2.8		
151.59	100	15	Hosp		Yes	18.0	Connecticut	1,533	1,451	94.7	61	4	21	1.4		
117.66	100	14	Cal		No	14.0	Delaware	202	198	98						
113.45	100	8	Hosp		Yes	9.6	Florida	5,188	4,704	90.7	297	5.7	187	3.6		
83.64	100	7	Hosp		No	8.4	Georgia	1,547	1,483	95.9	45	2.9	19	1.2		
116.37	0	0	NA		NA	0.0	Idaho	64	52	81.3						
90.06	75	10	Hosp		Yes	9.0	Illinois	5,038	4,705	93.4	274	5.4	59	1.2		
92.83	50	15	Hosp		No	9.0	Indiana	2,617	2,481	94.8	99	3.8	37	1.4		
85.90	42	10	Hosp		No	5.0	Iowa	671	612	91.2						
83.53	67	10	Hosp		Yes	8.0	Kansas	592	506	85.5	71	12	15	2.5		
100.35	100	14	Mixed		No	16.8	Kentucky	1,031	988	95.8	29	2.8	14	1.4		
68.97	75	7	Hosp		No	6.3	Louisiana	2,120	1,742	82.2	236	11.1	142	6.7		
115.77	100	10	Cal		No	10.0	Maine	123	119	96.7						
122.15	48	15	Hosp		No	8.6	Maryland	1,201	1,144	95.3	46	3.8	11	0.9		
124.47	0	0	Hosp		No	24.0	Massachusetts	3,659	3,469	94.8	126	3.4	64	1.7		
98.87	79	18	Hosp		Yes	0	Michigan	2,696	2,513	93.2	97	3.6	86	3.2		
116.84	100	15	Hosp		No	17.1	Minnesota	854	798	93.4	41	4.8	15	1.8		
90.38	100	12	Mixed		Yes	18.0	Mississippi	397	375	94.5						
91.65	100	10	Mixed		Yes	3.6	Missouri	1,711	1,525	89.1	169	9.9	17	1		
94.04	100	365	Hosp		Yes	365.0	Montana	64	63	98.4						
81.42	100	15	Hosp		Yes	18.0	Nebraska	277	256	92.4						

continued

Table 1. Continued

State Policies				Hospital Discharge Destination per Resident										
Per-Diem (\$)	% Per-Diem	Max Days	Time Frame	Bed-Hold Policies		ERD	Baseline NH		Other NH		Non NH			
				Occupancy Requirement	ERD		N	%	N	%	N	%		
103.30	64.8	16.6	N/A	N/A	N/A	16.0	All States	77,955	72,158	92.5	4,783	5.4	1,614	2.1
101.00	0	0	NA	NA	NA	0.0	Nevada	140	117	83.6				
118.91	0	0	NA	NA	NA	0.0	New Hampshire	178	170	95.5			17	0.4
127.63	90	10	Hosp	No	No	10.8	New Jersey	4,618	4,516	97.8	85	1.8		
92.96	50	12	Cal	Yes	Yes	6.0	New Mexico	82	75	91.5				
160.66	100	20	Hosp	Yes	Yes	24.0	New York	7,026	6,763	96.3	212	3	51	0.7
122.14	0	0	NA	NA	NA	0.0	North Carolina	1,467	1,362	92.8	89	6.1	16	1.1
104.94	100	15	Hosp	Yes	Yes	18.0	North Dakota	115	108	93.9				
121.76	50	30	Cal	No	No	15.0	Ohio	5,956	5,657	95	225	3.8	74	1.2
66.57	50	5	Cal	No	No	2.5	Oklahoma	883	694	78.6	143	16.2	46	5.2
95.43	0	0	NA	NA	NA	0.0	Oregon	114	108	94.7				
122.91	33	15	Hosp	No	No	6.0	Pennsylvania	6,058	5,872	96.9	111	1.8	75	1.2
116.03	0	0	NA	NA	NA	0.0	Rhode Island	626	577	92.2				
94.36	100	10	Hosp	No	No	12.0	South Carolina	727	699	96.1				
79.60	100	5	Hosp	No	No	6.0	South Dakota	153	143	93.5				
81.76	100	15	Hosp	Yes	Yes	18.0	Tennessee	1,666	1,556	93.4	89	5.3	21	1.3
83.53	0	0	NA	NA	NA	0.0	Texas	4,533	3,806	84	454	10	273	6
90.19	0	0	NA	NA	NA	0.0	Utah	107	90	84.1				
112.56	80	6	Hosp	Yes	Yes	5.8	Vermont	44	43	97.7				
99.99	0	0	NA	NA	NA	0.0	Virginia	1,264	1,190	94.1	58	4.6	16	1.3
121.79	0	0	NA	NA	NA	0.0	Washington	654	603	92.2				
113.68	100	12	Cal	Yes	Yes	12.0	West Virginia	411	403	98.1				
98.77	85	15	Hosp	Yes	Yes	15.3	Wisconsin	1,364	1,276	93.5	70	5.1	18	1.3
97.89	100	14	Cal	No	No	14.0	Wyoming	43	41	95.3				

Some cells have been suppressed to meet the Centers for Medicare and Medicaid Services privacy guidelines. ERD, Equivalent Reimbursement Days is calculated as the product of proportion of daily coverage and maximum number of days covered (converted to annual when necessary and multiplied by the average number of annual hospitalizations, 1.2).

model, the unit of analysis is a resident who was hospitalized from the original nursing home, and who was discharged from the hospital alive to an institution.

The model postulates that residents in the same nursing home would have unmeasured errors that are more similar than residents in other nursing homes, and likewise nursing homes in the same county, and counties in the same state. Thus, this method partitions the overall model error variance to components, accounting for the variance at each level of nesting. Denoted by Y_{ijk_r} , the location of hospital discharge for resident i in nursing home j in county k in state r , where discharge location is denoted 0 for return to original nursing home, 1 to another nursing home, and 2 for an institution other than a nursing home, we model the logit of the probability of hospital discharge to another nursing home or to a nonnursing home institution compared with discharge to original nursing home, i.e.

$$\log\left[\frac{\Pr(Y_{ijk_r} = y|Z_{ijk_r})}{\Pr(Y_{ijk_r} = 0)}\right] = \beta_{0jkr}^y + Z'_{ijk_r}\beta^y + \varepsilon_{ijk_r}^y, \quad \text{for } y = 1, 2;$$

$$\beta_{0jkr}^y = \beta_0^y + u_{0j}^y + u_{0k}^y + u_{0r}^y,$$

where β_{0jkr}^y is the facility-specific intercept and β^y are the regression parameters associated with each outcome y , in particular, they contain the regression parameters associated with any bed-hold policy, $\beta_{\text{bed-hold}}^y$ for which we test the null hypotheses that bed-hold is not related to discharge locations, i.e. that the parameters are zero, versus the alternative hypotheses that each is different from zero. The intercept is composed of a constant β_0^y and three random error terms u_{0j}^y , u_{0k}^y , and u_{0r}^y , each from a normal distribution with zero mean and specific variances, σ_j^{2y} , σ_k^{2y} , and σ_r^{2y} , which along with the overall model error's variance σ_ε^{2y} comprise the four-model variance parameters for each outcome $y = 1, 2$.

To simplify the estimation of the multinomial outcome, we estimated two multilevel logistic regression models comparing discharges to the baseline nursing home with discharges to a different nursing home and with discharges to nonnursing home settings. This method provides unbiased estimates with potentially inflated standard errors (Begg and Gray 1984).

Sensitivity Analyses

We conducted several exploratory analyses to examine the robustness of the results. We first removed the restriction of studying a single hospitalization. Next, we allowed transitions from other nursing homes including hospital-based nursing homes and studied whether the residents returned to the dis-

charging nursing home, as well as whether they returned to the baseline nursing home.

Nursing home residents in states with a bed-hold policy were more likely to be hospitalized (Intrator et al. 2007) raising a concern that residents admitted to hospital in those states were less sick and more likely to be readmitted back to the baseline nursing home. Another sensitivity analysis was conducted by estimating the models only on residents who were in the upper quartile of the posterior predicted probability of hospitalization ($> .2$) providing a perspective on the association of bed-hold and discharge destination on a more homogeneous sample of hospitalized residents who were sicker. The hospitalization model included only resident characteristics and controlled for the competing event of death, adjusting for clustering of residents in nursing homes (following Intrator et al. 2007).

RESULTS

Sample Description

The final cohort of 77,955 hospitalized nursing home residents had 100,269 hospitalizations over the 5-month window of which the first was used in the analyses. Overall, 92.5 percent of residents returned to their baseline nursing home. This varied from less than 82 percent in Oklahoma, Louisiana, and Nebraska to over 98 percent in Delaware, West-Virginia, and Montana (Table 1). The proportion of residents discharged to another nursing home averaged 5.4 percent, and the proportion of residents returning to an institution other than nursing home averaged 2.1 percent. Twelve states offered no bed-hold payment. Among the 36 states with bed-hold, most allowed bed-hold days per hospitalization (25 states), with an average ERD of 11.5 days (SD = 5.7 days). Average state-level Medicaid per-diem rates ranged between \$66.57 and \$160.66, with an average of \$103.30, (SD = \$19.60).

Table 2 presents the characteristics of the 8,652 urban free-standing nursing homes (NHs) with any hospitalized residents in this cohort. These characteristics do not differ from the characteristics of all urban free-standing NHs (Intrator et al. 2007). Of note, 19.4 percent of NHs had a dementia unit, and 3.4 percent had a rehabilitation unit. Almost 32 percent of NHs had less than 85 percent occupancy suggesting that there were plenty of beds available for returning hospitalized nursing home residents.

The average length of inpatient stay was 7 days indicating that almost 15 percent of the hospitalizations were less than the minimum stay required for

Table 2: Description of Cohort

	<i>N</i>	<i>%</i>	<i>Mean</i>	<i>SD</i>
State policies (<i>N</i> = 48)				
Any bed-hold policy	36	0.75		
Average per-diem Medicaid rate			103.29	19.62
County (<i>N</i> = 812)				
Area wage index adjuster to geographic variation in \$			0.96	0.12
Average number of empty beds			14.62	9.32
Average hospital occupancy			52.62	23.01
Baseline nursing home (NH) (<i>N</i> = 8,652)				
Total beds			123.72	69.48
For profit	6,364	73.56		
Chain ownership	5,151	59.54		
Average nursing case-mix index on admission assessments			0.96	0.11
Average nursing case-mix index on annual assessments			0.72	0.06
Occupancy < 85%	2,754	31.83		
Dementia unit available	1,681	19.43		
Rehabilitation unit available	298	3.44		
> 35% Residents nongovernment paid	1,991	23.01		
> 15% Medicare paid	1,338	15.46		
> 1/2 FTE MD	921	10.64		
Any NP/PA	1,914	22.12		
Ratio RNs to (RNs+LPNs)			0.34	0.20
Hospitalization event (<i>N</i> = 77,955)				
Event length of stay (LOS)			7.04	6.59
LOS < 3 days	11,334	14.53		
LOS 3–7 days	32,752	42.00		
LOS 8–14 days	27,225	34.90		
LOS > 14 days	6,684	8.57		
Primary diagnosis is ambulatory care sensitive	23,859	30.61		
Primary diagnosis is stroke	3,496	4.48		
Primary diagnosis is chronic obstructive pulmonary disease	1,944	2.49		
Primary diagnosis is congestive heart failure	4,218	5.41		
Hip fracture	4,197	5.38		
Has a psychiatric diagnosis	3,649	4.68		
Elixhauser comorbidity score			2.54	1.38
Any ICU during hospitalization event	16,286	20.89		
Resident baseline Minimum Data Set (MDS) (<i>N</i> = 77,955)				
Age			83.07	7.82
Married	13,775	17.67		
African American	10,888	13.97		
Male	21,464	27.53		
Do-not-resuscitate order	35,813	45.94		
Moderate cognitive impairment (CPS = 3, 4)	30,691	39.37		
Severe cognitive impairment (CPS = 5, 6)	16,396	21.03		
ADL < 8	16,373	21.00		
ADL 8–16	19,148	24.56		

continued

Table 2. *Continued*

	<i>N</i>	%	<i>Mean</i>	<i>SD</i>
Weight loss	9,800	12.57		
Flacker score			4.36	2.15
Diagnosis of diabetes affecting functioning	22,203	28.48		
Diagnosis of CHF affecting functioning	15,936	20.44		
Bipolar/schizophrenia without treatment	994	1.28		
Bipolar/schizophrenia with treatment	2,783	3.57		
More than nine medications in past 7 days	31,601	40.54		
Obese (BMI > 30)	11,250	14.43		
Low weight (BMI < 18)	7,024	9.01		

FTE, full-time-equivalent (% time employed); NP/PA, nurse practitioner/physician assistant; ICU, intensive care unit; CPS, cognitive performance scale; ADL, activity of daily living scale; BMI, body mass index.

Medicare SNF benefits (3 days). Another 8.6 percent of hospitalizations were longer than 2 weeks.

Multilevel Model

Table 3 presents the results of the multilevel model. Nursing home residents from states with a bed-hold policy were more likely to return to their baseline nursing home. When compared with returning to the baseline nursing home, the adjusted odds ratio (AOR) of discharge to another nursing home was 0.55 (95 percent confidence interval [CI] 0.52–0.58). Being discharged to an institution other than a nursing home was also less likely (AOR = 0.76; CI 0.44–1.31), although not statistically significant. Nursing home residents from states with higher Medicaid per-diem rates were more likely to return to their baseline nursing home. When compared with return to baseline, the AOR of returning to another nursing home with a \$10 higher Medicaid per-diem rate was 0.85 and 0.88 for discharge to another institution, both statistically significant at the 5 percent level.

Higher hospital occupancy in the local market was associated with increased odds of being discharged to another nursing home (AOR = 1.11). Higher competition for nursing home residents as measured by excess capacity (empty beds) in the market was associated with higher odds of being discharged to an institution other than a nursing home (AOR = 1.26).

Residents of for-profit nursing homes were more likely to be discharged to other nursing homes (AOR = 1.22), but chain membership was not associated with discharge location. Residents hospitalized from larger nursing homes were less likely to be discharged to another nursing home (AOR = 0.87

Table 3: Model of Discharge Location Following Hospitalization, Compared with Returning to Baseline Nursing Home

<i>Characteristics/Factors</i>	<i>Other Nursing Home</i>			<i>Institution Other Than Nursing Home</i>		
	<i>AOR</i>	<i>LCI</i>	<i>UCI</i>	<i>AOR</i>	<i>LCI</i>	<i>UCI</i>
State policies						
Any bed-hold policy	0.55**	0.52	0.58	0.76	0.44	1.31
Average per-diem Medicaid rate (std to $[x - \$103.3]/\10)	0.85**	0.82	0.87	0.88*	0.78	0.99
County market						
Area wage index adjusted to geographic variation (std to $[x - 0.96]/0.12$)	0.93**	0.90	0.96	0.95	0.84	1.08
Average number of empty beds (std to $[x - 14 \text{ beds}]/10 \text{ beds}$)	1.05	0.96	1.14	1.26**	1.06	1.50
Average hospital occupancy (std to $[x - 52.6\%]/20\%$)	1.11*	1.02	1.21	1.04	0.91	1.19
Baseline nursing home						
Total beds (std to $[x - 100 \text{ beds}]/70 \text{ beds}$)	0.87**	0.83	0.92	1.06 ⁺	1.00	1.13
For-profit	1.22**	1.14	1.30	1.09	0.93	1.29
Chain ownership	1.03	0.99	1.08	1.05	0.92	1.21
Average nursing case-mix index on admission assessments (std to $[x - 0.96]/0.125$)	0.84**	0.79	0.89	1.00	0.92	1.10
Average nursing case-mix index on annual assessments (std to $[x - 0.719]/0.066$)	0.89**	0.84	0.95	0.92 ⁺	0.84	1.01
Occupancy < 85%	1.08**	1.03	1.14	0.98	0.85	1.12
Dementia unit available	0.93 ⁺	0.85	1.01	1.11	0.95	1.30
Rehabilitation unit available	1.13	0.90	1.41	1.09	0.79	1.52
> 35% Residents nongovernment paid	1.10 ⁺	0.98	1.22	1.05	0.88	1.25
> 15% Medicare paid	1.01	0.93	1.09	1.08	0.90	1.29
> 1/2 FTE MD	1.01	0.87	1.16	0.95	0.77	1.16
Any NP/PA	1.04	0.98	1.10	1.18*	1.02	1.36
Ratio RNs/(RNs+LPNs) (std to $[x - 0.3]/0.2$)	0.98	0.94	1.01	1.02	0.94	1.11
Hospitalization event						
Event length of stay (LOS) (std to $[x - 7.3 \text{ days}]/7.3 \text{ days}$)	1.41**	1.35	1.46	1.28**	1.20	1.37
LOS < 3 days	0.20**	0.17	0.25	0.99	1.04	0.94
LOS 3-7 days	0.73**	0.69	0.76	0.61**	0.53	0.70
LOS > 14 days	1.33**	1.21	1.45	1.37**	1.12	1.68
Primary diagnosis is ambulatory care sensitive	1.07 ⁺	1.00	1.15	0.67**	0.58	0.77
Primary diagnosis is stroke	1.23 ⁺	1.03	1.49	2.59**	2.13	3.16
Primary diagnosis is chronic obstructive pulmonary disease	0.71**	0.55	0.92	1.01	0.67	1.54
Primary diagnosis is congestive heart failure	0.70**	0.61	0.80	0.87	0.63	1.20
Hip fracture	1.97**	1.73	2.24	1.72**	1.36	2.16
Has a psychiatric diagnosis	1.29**	1.17	1.42	0.31**	0.22	0.45
Elixhauser comorbidity score	1.00	0.97	1.04	1.13**	1.07	1.19

continued

Table 3. *Continued*

<i>Characteristics/Factors</i>	<i>Other Nursing Home</i>			<i>Institution Other Than Nursing Home</i>		
	<i>AOR</i>	<i>LCI</i>	<i>UCI</i>	<i>AOR</i>	<i>LCI</i>	<i>UCI</i>
Any ICU during hospitalization event	0.93 ⁺	0.86	1.00	1.59**	1.41	1.80
Resident						
Age (std to [x - 83 years]/8 years)	1.00	0.94	1.06	0.93*	0.88	0.99
Married	1.00	0.94	1.06	1.15*	1.00	1.33
African American	0.86 ⁺	0.72	1.02	0.99	0.84	1.17
Male	1.12**	1.06	1.19	1.07	0.95	1.21
Do-not-resuscitate order	0.86**	0.79	0.92	0.97	0.86	1.09
Moderate cognitive impairment (CPS = 3, 4)	0.93	0.84	1.03	0.80**	0.71	0.91
Severe cognitive impairment (CPS = 5, 6)	0.87*	0.77	0.98	1.05	0.90	1.22
ADL < 8	0.72**	0.63	0.81	0.77**	0.64	0.92
ADL 8-15	0.91**	0.85	0.97	0.73**	0.63	0.83
Weight loss	1.03	0.93	1.14	0.85 ⁺	0.70	1.02
Flacker score (std)	0.98	0.94	1.02	1.23**	1.13	1.34
Diagnosis of diabetes affecting functioning	0.93 ⁺	0.85	1.01	0.87*	0.77	0.98
Diagnosis of CHF affecting functioning	1.09*	1.02	1.18	0.90	0.77	1.04
Bipolar/schizophrenia without treatment	0.63**	0.47	0.83	0.87	0.54	1.41
Bipolar/schizophrenia with treatment	0.70**	0.55	0.89	0.93	0.68	1.27
More than nine medications in past 7 days	1.01	0.93	1.09	1.05	0.94	1.18
Obese (BMI > 30)	0.93 ⁺	0.86	1.01	1.07	0.91	1.26
Low weight (BMI < 18)	1.11*	1.02	1.21	0.98	0.82	1.17

<i>Variance components</i>	<i>Var</i>	<i>SE</i>	<i>ICC (%)</i>	<i>Var</i>	<i>SE</i>	<i>ICC (%)</i>
State	0.199	0.052	4.72	0.418	0.143	8.66
County	0.216	0.014	5.13	0.6	0.099	12.45
Facility	0.514	0.011	12.19	0.516	0.09	10.70

Std, standardized; AOR, adjusted odds ratio; LCI, lower confidence interval; UCI, upper confidence interval; ICC, intra-class correlation; FTE, full-time-equivalent (% time employed); NP/PA, nurse practitioner/physician assistant; ICU, intensive care unit; CPS, cognitive performance scale; ADL, activity of daily living scale; BMI, body mass index; Var, variance; SE, standard error.

**p-value < .01.

*p-value < .05.

⁺p-value < .1.

for every 70 beds more than 100), but more likely to be discharged to another institution (AOR = 1.06 for every 70 beds more than 100).

Discharge location was not affected by payer mix at the baseline nursing home. On the other hand, residents from nursing homes with low occupancy (<85 percent) had higher odds of being discharged to another nursing home (AOR = 1.08). The availability of a rehabilitation unit in the baseline nursing home was not associated with discharge location in this cohort, and the avail-

ability of a dementia unit was only marginally associated with lower odds of discharge to another nursing home (AOR = 0.93). Finally, neither nursing home staffing level nor skill mix was related to the hospital discharge location.

The length of patients' hospitalization was a major determinant of discharge location. Overall, longer length of stay (LOS) was associated with higher odds of being discharged to locations other than the baseline nursing home. Residents hospitalized for very brief stays (<3 days) were much less likely to be discharged to another NH (AOR = 0.20), and those hospitalized for 3–7 days were also less likely to be discharged to another nursing home (AOR = 0.73). Other hospitalization characteristics such as diagnosis and use of the ICU were also associated with the discharge location.

Model Robustness

An examination of all hospitalizations, including those not emanating from the baseline nursing home, controlling for discharging nursing home characteristics, resulted in similar estimates that were somewhat tempered (particular results available upon request). Residents hospitalized from a hospital-based nursing home had 3.24 higher odds of being discharged to another nursing home (CI = 1.90–5.52).

Among residents more likely to be hospitalized ($N = 27,524$), 1,678 (6.1 percent) were discharged to another nursing home and 661 (2.4 percent) were discharged to an institution other than a nursing home. Model results revealed that Medicaid per-diem rate was associated with lower odds of discharge to another nursing home (AOR = 0.88; 95 percent CI 0.82–0.95); however, the effect of bed-hold policy was much smaller and not statistically significant (AOR = 0.81; 95 percent CI 0.58–1.12).

DISCUSSION

Previously reported nursing home to nursing home transfer was 2–3 percent annually (Mor et al. 1997; Hirth et al. 2000), lower than the rate observed in this cohort of long-stay residents following hospitalization (5.4 percent). Whereas the cohorts and unit of analysis were different in those earlier studies, this suggests that the majority of nursing home to nursing home transfers occur via an intervening hospitalization. Moreover, residents who stayed longer in the hospital were more likely to be transferred to other nursing homes. These results suggest that a hospitalization is a catharsis for re-evaluation of a relative's goals of care and needs.

The study confirmed the hypothesis that bed-hold policies were associated with a lower rate of transfer to other nursing homes. Interestingly, this result did not vary by the generosity of the bed-hold policy. Moreover, higher Medicaid reimbursement was associated with a greater likelihood of return to the original nursing home, continuing a line of research that suggests that higher Medicaid payments for nursing home resident care are associated with better nursing home quality and resident outcomes, possibly due to financial stability allowing investment in infrastructure and staffing (Grabowski 2001, 2004; Grabowski et al. 2004; Intrator and Mor 2004; Intrator et al. 2005, 2007; Feng et al. 2008). These findings indicate the need for a comprehensive analysis of the relative merits of bed-hold policies compared to their price and to other policies that might achieve comparable merits.

Nursing home transfer may be a consequence of potential quality differences between the baseline and new nursing home. A previous study showed that nursing home to nursing home transfer, though infrequent, was more common from lower quality facilities (Hirth et al. 2000). However, the relative quality of the new facility was not investigated. Among residents transferred to another nursing home in this study, 48.8 percent subsequently returned to their baseline nursing home within 92 days, indicating that quality might not have been the reason for those transitions. Moreover, lower occupancy rate at the baseline nursing home did not appear to guarantee a return to that nursing home. Indeed, contrary to our expectations, lower occupancy was associated with more transfers to another nursing home, suggesting that occupancy per se might be a surrogate measure of nursing home quality with lower occupancy serving as a marker of poorer quality. Other literature suggests the relationship between lower occupancy and poor nursing home quality (Mor et al. 2004; Smith et al. 2007).

Medicare policies allow SNF care following a hospitalization of 3 days or more. Thus, it was hypothesized that one reason for hospitalization would be to establish resident eligibility for Medicare SNF-level care, which is generally compensated at higher rates than Medicaid (MEDPAC 2005). Among residents returning to their baseline nursing home, only 20,387 (29.1 percent) returned with Medicare SNF level of care, while among residents transferred to another nursing home, 53.7 percent received Medicare SNF level of care from their new nursing home. Even though most nursing homes in the country were dually certified to provide Medicaid and Medicare covered services, it is likely that the type of SNF care required following a hospitalization was not available in residents' baseline nursing home requiring the temporary transition.

There has been a growing concern regarding unnecessary hospitalizations (Saliba et al. 2000; Miller et al. 2003; Intrator et al. 2004; Porell and Carter 2005), another indicator of poor quality of care in nursing homes. Indeed, among the hospitalizations in this cohort, 30.6 percent were for an Ambulatory Care Sensitive condition, indicating the hospitalization could potentially have been avoided (Intrator et al. 2004). Moreover, the fact that many hospitalizations did not result in SNF care increases doubts regarding the circumstances of the hospitalizations. As bed-hold policies have been associated with increased hospitalizations (albeit not more potentially preventable than otherwise), the number of potentially preventable hospitalizations would necessarily have been higher than without the policy.

Our cross-sectional analysis has several limitations. We identified associations between policy and discharge destination, but could not ascertain causality. Whereas we have controlled for occupancy with annually updated OSCAR data, we were unable to consider the day-to-day variations in occupancy that could disallow bed-hold payments. Therefore, some hospitalizations in states, which require minimum nursing home occupancy (e.g., 90 percent) in order to activate a bed-hold policy, may not have resulted in bed-hold payments. Furthermore, our cohort includes nongovernment paid nursing home residents for whom Medicaid bed-hold payments would not apply. However, both limitations serve to bias our results towards the null hypothesis of no bed-hold policy effect on discharge destination.

We conducted a rough cost estimate based on the model presented in Table 3 and assuming 17.4 percent hospitalization rate without bed-hold and 22.3 percent with bed-hold (assuming AOR = 1.36, based on prior literature). Using these figures, we estimated that bed-hold policies were associated with 9.5 fewer annual relocations for every 1,000 residents, with an associated 77.9 more hospitalizations. It is therefore instructive to examine the associated budgetary implications, especially in light of arguments made by the nursing home industry that bed-hold payments were integral to the financial viability of nursing homes (Florida 1999; Massachusetts 2003; Rotstein 2006). A simple calculation shows that with a generous bed-hold policy such as in Massachusetts and New York, and other states that pay for bed-hold at 100 percent for 12 calendar days or more, the average Medicaid payment for the bed-hold would be \$201.31 per resident per year.² With a less generous bed-hold policy such as in Oklahoma, with a maximum of 5 days reimbursed annually at 50 percent, the extra revenue to nursing homes from the bed-hold policy would be \$1.18 per resident year.³ A simple calculation shows that by increasing Medicaid daily reimbursement rate by 55 cents, the most generous bed-hold

policies would be offset. Assuming that the causal effect of increased Medicaid rate on limiting hospitalizations and increasing return to baseline nursing home is unchanged, this substitution would encourage better quality care, and ultimately cost Medicare less by saving payments for the additional hospitalizations. These rough calculations, coupled with the lower levels of nursing home occupancy in today's markets, appear to indicate that bed-hold policies may have reached the end of their life.

In making this estimation, we acknowledge that we do not have all the inputs to conduct a full cost-benefit analysis regarding bed-hold policy. Specifically, any benefit or harm associated with relocation has yet to be quantified. For example, when a state has no bed-hold policy, some residents may refuse hospitalization to avoid loss of their bed (Nohlgren 2004). In particular, if ill patients refuse hospitalization, mortality may increase. Providing care to more acutely ill patients in nursing homes ill-prepared to do so may take time away from other residents, increasing the potential for adverse events and litigation (Stevenson and Studdert 2003). Absence of a bed-hold policy may prolong hospitalization due to difficulty in determining a locale for hospital discharge (Nohlgren 2004). Verification of these potential effects will require further research.

A full-scale cost analysis examining all cost implications of removing or adding a bed-hold policy should be considered to further elucidate the issues of financial viability. A cost-benefit study could be designed to examine the costs relative to the social and personal benefits of such policies.

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NOTES

1. Case-mix measures were aggregated from all admission MDS assessments and separately for all annual MDS assessments in 2000 using the nursing case-mix index, an index of the ratio of the average number of nursing staff minutes required to care for residents at each of 44 Resource Utilization Groups to the overall average of staff minutes.
2. This calculation is based on the average number of hospitalization days, 11.2 days, multiplied by 100 percent bed-hold rate, and per-diem of \$103.30, applied to 17.4 percent residents estimated to be hospitalized. The additional hospitalized resident days due to the increased hospitalization rates would not incur additional cost since they would have been paid for had the resident not been hospitalized. Under this scenario, the total cost to Medicaid would be \$201.31 per resident year. In making the calculations, we assumed that the effect of bed-hold policy on the hospitalization rate does not vary by the generosity of the policy.
3. Given bed-hold payment at 50 percent of rate for a maximum of 5 days would reimburse NHs \$44.97 for hospitalizations that would have occurred if the hospitalization rate had remained at 17.4 percent. However, Medicaid would save \$43.71 for the extra residents hospitalized due to the higher hospitalization rate by paying the partial bed-hold rate on days that the resident would have otherwise been in the nursing home. This brings the total extra payment due to bed-hold policy to \$1.18 per resident year.

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