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### Using Clinical Factors and Mammographic Breast Density to Estimate Breast Cancer Risk: Development and Validation of a New Predictive Model

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#### Abstract

**Background**—Current models for assessing breast cancer risk are complex and do not include breast density, a strong risk factor for breast cancer that is routinely reported with mammography.

**Objective**—To develop and validate an easy-to-use breast cancer risk prediction model that includes breast density.

**Design**—Empirical model based on Surveillance, Epidemiology, and End Results incidence, and relative hazards from a prospective cohort.

**Setting**—Screening mammography sites participating in the Breast Cancer Surveillance Consortium.

**Patients**—1 095 484 women undergoing mammography who had no previous diagnosis of breast cancer.

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**Measurements**—Self-reported age, race or ethnicity, family history of breast cancer, and history of breast biopsy. Community radiologists rated breast density by using 4 Breast Imaging Reporting and Data System categories.

**Results**—During 5.3 years of follow-up, invasive breast cancer was diagnosed in 14 766 women. The breast density model was well calibrated overall (expected–observed ratio, 1.03 [95% CI, 0.99 to 1.06]) and in racial and ethnic subgroups. It had modest discriminatory accuracy (concordance index, 0.66 [CI, 0.65 to 0.67]). Women with low-density mammograms had 5-year risks less than 1.67% unless they had a family history of breast cancer and were older than age 65 years.

**Limitation**—The model has only modest ability to discriminate between women who will develop breast cancer and those who will not.

**Conclusion**—A breast cancer prediction model that incorporates routinely reported measures of breast density can estimate 5-year risk for invasive breast cancer. Its accuracy needs to be further evaluated in independent populations before it can be recommended for clinical use.

In 2007, breast cancer will have been diagnosed in more than 178 000 women in the United States, and more than 40 000 women will have died of breast cancer (1). Most of these women never had their risk for breast cancer assessed, and even fewer considered chemoprevention (2–5). Providing women with an estimate of their risk for breast cancer would provide an opportunity for them to consider options to decrease their risk. Women at low short-term risk for breast cancer may experience less anxiety about their health and would be less likely to benefit from prevention efforts. Women at very high risk may warrant additional screening tests, such as breast magnetic resonance imaging (6), and might benefit from chemoprevention of breast cancer with tamoxifen or raloxifene. The standard risk assessment model available to practitioners (the Gail model) (7) identifies only a minority of women who eventually develop breast cancer being at high risk (8). Better breast cancer risk prediction tools are needed (9).

The radiographic appearance of the breast has been consistently shown to be a major risk factor for breast cancer, whether it is defined by a qualitative assessment of the parenchymal pattern or a quantitative measure of percentage of density (10–12). Women in whom more than 50% of total breast area is mammographically dense have high breast density and are at 3- to 5-fold greater risk for breast cancer than women in whom breast density is less than 25% (10,13– 16). The increased risk for breast cancer associated with breast density is due in part to the lower sensitivity of mammography in dense breasts (17–19), but the association remains strong after accounting for masking (20,21). Mammographically dense breast tissue is rich in epithelium and stroma (10), and the association could represent activation of epithelial cells or fibroblasts (22–25). Recently, several models have been published that incorporate breast density: One uses a continuous measure of breast density that is not available to clinicians and has not been validated (26), and the other predicts 1-year risk for breast cancer (27).

We previously demonstrated that a simple model based on age, ethnicity, and a categorical measure of breast density had predictive accuracy similar to that of the Gail model in a multiethnic cohort of women receiving screening mammograms in northern California (28). We expand on that work by using data from more than 1 million ethnically diverse women throughout the United States to develop and validate a risk assessment tool that incorporates breast density and therefore might improve breast cancer screening and prevention efforts.

#### Methods

#### **Study Population**

We included 1 095 484 women age 35 years or older who had had at least 1 mammogram with breast density measured by using the Breast Imaging Reporting and Data System (BI-RADS) classification system in any of the 7 mammography registries participating in the National Cancer Institute–funded Breast Cancer Surveillance Consortium (BCSC) (available at http://breastscreening.cancer.gov) (29). The BCSC is a community-based, ethnically and geographically diverse sample that broadly represents the United States (30).

We excluded women who had a diagnosis of breast cancer before their first eligible mammography examination. Because our goal was to develop a model of long-term risk for invasive breast cancer, we excluded women with cancer diagnosed in the first 6 months of follow-up to minimize the number of cases of cancer included in the model that were diagnosed on the basis of the mammogram used for risk assessment. Women were also excluded if they had breast implants. Women in whom ductal carcinoma in situ was diagnosed were censored at the time of diagnosis in the primary analysis. When women had several mammograms, we based our analysis on findings from the first mammogram.

Each registry obtains annual approval from its institutional review board for consenting processes or a waiver of consent, enrollment of participants, and ongoing data linkage for research purposes. All registries have received a Certificate of Confidentiality from the federal government that protects the identities of research participants.

#### **Measurement of Risk Factors**

Patient information was obtained primarily from self-report at the time of mammography. We selected 2 risk factors in addition to breast density for inclusion in the model on the basis of simplicity (yes or no) and a high attributable risk: history of breast cancer in a first-degree relative and history of a breast biopsy. Body mass index was later considered for addition to the model, but it was excluded to maintain parsimony and because it had minimal effect on model discrimination (the increase in the concordance statistic [c-statistic] was only 0.003). For modeling and validation, missing data for relatives with breast cancer and number of breast biopsies were set to 0. The 5-year Gail risk was computed for each woman by using the algorithms provided by the National Cancer Institute to calculate the Gail model risk for individual women (31). For Gail model calculations, missing data were coded as specified by that model (age at menarche as  $\geq$ 14 years, age at first live birth as <20 years, number of breast biopsies as 0, and number of first-degree relatives as 0). Ethnicity was coded by using the expanded race and ethnicity definition currently used in the Surveillance, Epidemiology, and End Results (SEER) database and U.S. Vital Statistics (non-Hispanic White, non-Hispanic Black, Asian or Pacific Islander, Native American/Alaskan Native, Hispanic, or other). We classified women who self-identified as mixed or other race with participants who did not report race and ethnicity.

**Breast Density**—Community radiologists at each site classified breast density on screening mammograms as part of routine clinical practice by using the American College of Radiology BI-RADS density categories (32): almost entirely fat (category 1), scattered fibroglandular densities (category 2), heterogeneously dense (category 3), and extremely dense (category 4). The BI-RADS category 2 was used as the reference group for breast density because it formed the largest group.

**Ascertainment of Breast Cancer Cases**—Breast cancer outcomes (invasive cancer and ductal carcinoma in situ) were obtained at each site through linkage with the regional population-based SEER program, state tumor registries, and pathology databases.

**Vital Status**—Vital status was obtained through linkage to SEER registries, state tumor registries, and the individual state vital statistics or the National Death Index.

#### **Model Development**

We used a proportional hazards model of invasive breast cancer to estimate the hazard ratios for each BI-RADS breast density category. Women entered the model 6 months after the index mammogram and were censored at the time of death, diagnosis of ductal carcinoma in situ, or the end of follow-up. All models were adjusted for age (in 5-year intervals) and race and ethnicity. The strength of the breast density association with breast cancer was greater for women younger than age 65 years (*P* for interaction < 0.001). Thus, separate models were fitted for women younger than age 65 years and for women age 65 years or older. No other interaction terms were included in the final model. We calculated similar estimates for first-degree relatives with breast cancer (yes or no) and a personal history of breast biopsy (yes or no) from the BCSC. All predictors met the proportional hazards assumption that was assessed by log–log plots and by including interaction terms with time for each predictor variable.

We then developed an absolute risk model by using methods described in the Appendix Figure (available at www.annals.org). The model primarily estimates predicted incidence of invasive breast cancer by using age, race or ethnicity, and breast density. These estimates are then adjusted for family history and biopsy history if available. We based our estimates of breast cancer incidence on the SEER age- and ethnicity-specific risk for invasive breast cancer (1992 to 2002) (33). Age-specific incidence for each ethnic group was estimated by fitting a thirdorder polynomial model to the SEER data. Age-specific incidence rates for the Native American and Alaskan Native group were inconsistent in SEER, so we excluded this group from further analyses. We calculated the baseline risk for the model by adjusting SEER incidence for the population's attributable risk for each breast density subgroup. We estimated the age- and ethnicity-specific distribution of mammographic breast density needed for these calculations by using data from a larger set of 3 343 047 mammograms from the BCSC. The distribution of breast density varied statistically significantly by age and by race or ethnicity (P < 0.001 for each comparison). The model used these variations by age and race to distribute the 5-year risk for invasive breast cancer across the 4 breast density subgroups. We used the methods described by Gail and colleagues (7) to translate the hazard ratios and risk factor distributions into absolute risks. The age-, sex-, and ethnicity-specific competing risks for death for women were calculated by using 2002 U.S. Vital Statistics (34). Age-specific death for each ethnic group was estimated by fitting an exponential model to the 2002 U.S. Vital Statistics. To assess the effect of breast density on the model, we developed a similar model, the risk factor model, by using the same approach but excluding breast density.

#### Statistical Analysis

We developed the model by using a random sample of 60% of the women and validated it in the remaining 40%. Model calibration was assessed by calculating the ratio of expected cases of breast cancer to observed cases of breast cancer (expected–observed ratio) by age group, race or ethnicity, individual risk factor distributions, decile of predicted risk, and Gail risk. The 95% CI for this ratio was calculated by assuming that the observed breast cancer events follow a Poisson distribution; thus, we calculated the CI as follows: (expected–observed ratio)\*exp ( $\pm$  1.96\*1/sqrt [O]). For a group of women, calibration assesses how closely the number of women in whom the model predicts that breast cancer will develop matches the actual number

of women in whom breast cancer is diagnosed. An expected–observed ratio of 1.0 would indicate perfect calibration.

We summarized the discriminatory accuracy of the model by using the c-statistic (35). Standard errors used to calculate 95% CIs around the c-statistic were estimated by using the method of DeLong and colleagues (36). We calculated the age-adjusted c-statistic by using the method of Rockhill and colleagues (37). We also used a 5-fold cross-validation to confirm the internal validity of the model (38,39). The c-statistic measures the ability of the model to separate women who will develop breast cancer from those who will not by calculating the proportion of pairs of women in which the woman with breast cancer has a higher predicted risk than the woman without breast cancer. A c-statistic of 0.5 is equivalent to no discrimination, and a c-statistic of 1.0 indicates perfect discrimination between women who develop breast cancer and women who do not.

We used the derivation sample to guide model development and the validation sample to assess calibration and discrimination of the model predictions. To evaluate the contribution of breast density in addition to clinical risk factors and for comparisons with the Gail model, we limited analyses to the 629 229 women with complete follow-up information from 0.5 to 5.5 years from the index mammogram (a 5-year interval). We used the method advocated by Cook and colleagues (40,41) to compare the model based on breast density with a model based on clinical risk factors. Because data were missing, direct comparisons between the breast density model and the Gail model may be biased against the Gail model and should be interpreted with caution. We reported incidence rates per 500 woman-years in the larger data set of women with variable follow-up length to approximate 5-year risks. A 2-sided *P* value of 0.05 or less was considered statistically significant. All analyses were done by using Stata, version 9.2 (Stata, College Station, Texas).

#### **Role of the Funding Source**

The BCSC and Building Interdisciplinary Research Careers in Women's Health had no role in the design, conduct, and analysis of this study, nor did they participate in the decision to submit the manuscript for publication.

#### Results

At the time of their earliest mammogram in the BCSC, 46% of women in our study were younger than age 50 years (Table 1). The majority of women were white (71%), but more than 25 000 women represented each of the black, Asian, and Hispanic groups. During a median follow-up of 5.3 years, 14 766 women developed invasive breast cancer. As expected, older age, non-Hispanic white race or ethnicity, a family history of breast cancer, a personal history of breast biopsies, and high breast density were all associated with the development of breast cancer (Table 1).

The strength of the association between breast density and breast cancer was greater for women younger than age 65 years (the relative hazard for extremely dense versus almost entirely fat breasts decreased from 4.2 to 2.2; *P* for interaction < 0.001). The distribution of breast density also varied by age and race or ethnicity (P < 0.001). The proportion of women with extremely dense breasts was greatest among Asian women at all age ranges (Table 2) and decreased with age across all race or ethnicity groups.

The model was well calibrated in the validation sample (Table 3). Within a subset of the validation cohort with 5 years of follow-up, the observed rate of invasive breast cancer was 1.38% (3465 cases of cancer among 251 789 women). The expected rate according to the model was 1.41% (expected-observed ratio, 1.03 [CI, 0.99% to 1.06%]). Model discrimination

measured by the c-statistic in the validation set was 0.660 (CI, 0.651 to 0.669), which was statistically and possibly clinically significantly greater than that of the Gail model (c-statistic, 0.613 [CI, 0.604 to 0.622]). In addition, the average c-statistic from 5-fold cross-validation was 0.6576. As expected, age-adjusted c-statistics were lower for both the breast density model (0.622) and the Gail model (0.562).

Calibration of the model was reasonably accurate across risk factor subgroups (Table 4). The model slightly underestimated breast cancer rates in younger women (expected– observed ratio, 0.94 for women age 40 to 44 years). It also underestimated cancer rates among Asian (expected– observed ratio, 0.95) and Hispanic women (expected–observed ratio, 0.94). The model was well calibrated across other risk factor subgroups, including those defined by the Gail model.

Table 5 shows the predicted 5-year risk for women in the study, by age and breast density groups and by subgroups defined by the presence or absence of a first-degree relative with breast cancer and a history of breast biopsy. Women with the lowest mammographic breast density (almost entirely fat) had a 5-year risk greater than 1.66% only if they were at least age 65 years and had both a first-degree relative with breast cancer and a personal history of a breast biopsy. In contrast, women with extremely dense breasts had an observed risk greater than 2% by age 45 years if they had either a family history of breast cancer or a personal history of a breast biopsy and by age 50 years, regardless of whether they had had previous biopsies or a family history of breast cancer.

We used reclassification tables to compare the breast density model with other models. We divided women into 4 risk categories: low (<1%), low or intermediate (1% to 1.66%), high or intermediate (1.67% to 2.5%), and high ( $\geq$ 2.5%). Using the method suggested by Cook and colleagues (40,41), we calculated the proportion of women reclassified correctly (patients with cancer reassigned to a higher-risk category and patients without cancer reassigned to a lowerrisk category) and patients reclassified incorrectly. The addition of breast density to age, race or ethnicity, family history, and history of breast biopsy correctly reclassified 22% of women and incorrectly reclassified 16% of women (Table 6). When we used a cut-point greater than 1.66% to define high risk, the true-positive rate increased slightly (from 52% to 53%) and the false-positive rate decreased slightly (from 33% to 30%). The positive predictive value thus increased slightly from 2.2% to 2.4%. Compared with the Gail model, the model we developed correctly reclassified 14% of women but incorrectly reclassified 35% of women (Table 7). Nonetheless, the true-positive rate increased from 28% to 53% and the positive predictive value increased from 2.3% to 2.4%. The false-positive rate also increased from 17% to 30%. Results were similar when we used a cut-point greater than 2% to define high risk (Appendix Tables 1 and 2, available at www.annals.org).

#### Discussion

We developed a risk prediction model to assess 5-year risk for invasive breast cancer that is based on breast density and demonstrate that the model has some features that might make it useful for risk prediction. The model is well calibrated in major race and ethnic groups in the United States. In addition, it has modest ability to discriminate between women who will develop breast cancer and those who will not.

Two other risk assessment models have incorporated mammographic breast density (26,27). The first, a refinement of the Gail model by Chen and colleagues (26), used a continuous measure of breast density in addition to age, age at first live birth, number of first-degree relatives with breast cancer, body weight, number of breast biopsies, and presence of atypical hyperplasia on biopsy to estimate a woman's future risk for breast cancer. This improves risk

discrimination compared with the Gail model, but calibration of the model in different demographic and risk factor subgroups has not yet been published. In addition, the model is based on a continuous measure of breast density that is not routinely available in clinical practice because it requires digital scanning of the mammogram and specialized software to estimate the percentage of the total breast area on the mammogram that is dense, a relatively expensive and labor-intensive process. Thus, it cannot be readily implemented among women receiving mammography today. Barlow and colleagues (27) also developed a model in a different subset of the BCSC that used BI-RADS as its measure of breast density. However, their model focused on 1-year risk for breast cancer and was validated on the basis of cases of breast cancer diagnosed by the index mammogram. This method may overestimate a woman's long-term risk for breast cancer by including incident cancers detected by the first mammogram.

Our model has advantages over these other models. It uses simple, easily obtained variables and is based on many races and ethnicities. Despite these advantages and its excellent calibration, however, its discriminatory accuracy was modest (c-statistic, 0.660). The age-adjusted c-statistic, a better measure for comparing models across different study populations because it removes the age distribution of the sample from the measure (the strongest risk factor for breast cancer), was even lower (0.622). We believe that these modest measures of discrimination must be interpreted with the understanding that extremely high relative risks (>100) are required for risk factor models to have high c-statistics (42–44) and that other risk models with modest discriminatory accuracy, such as the Framingham model (c-statistic, 0.63 to 0.83) (45), are commonly used to guide clinical decisions. Thus, in our view, the range of c-statistics we report does not exclude the model as a potential risk prediction tool.

The Gail model without breast density (7) remains the standard risk assessment tool for clinicians. In a previous study (28), we demonstrated that a model based on BI-RADS density alone had predictive accuracy similar to one based on the Gail model variables. Adding breast density to the Gail model was statistically significant (P < 0.001), although the small increase in the c-statistic was unlikely to be clinically important (28). In that analysis, we directly compared a simple proportional hazards model by using Gail model variables with the same model plus breast density. The model that we report here differs from our previous model in several ways, including the use of SEER data to estimate breast cancer incidence based on age and race or ethnicity and the use of different coefficients for BI-RADS density for younger (age <65 years) and older women (age  $\geq$ 65 years). It also more carefully accounts for differences in the age- and race- or ethnicity-specific distribution in breast density.

Compared with the Gail model, the breast density model reclassified a larger percentage of women incorrectly than correctly. However, it reclassified a higher proportion of women who developed breast cancer into higher-risk categories, and the positive predictive value and c-statistic of the breast density model were higher than that of the Gail model. This paradoxical finding can occur when 1 of the models is not well calibrated in a population. The Gail model has recently been shown to underestimate the risk for invasive breast cancer in black women by a factor of 1.5 to 3.2 (46). The racial diversity of our cohort and missing data for some Gail model risk factors produced an inappropriately low risk estimate for some women when using the Gail model (expected–observed ratio, 0.88 for Gail model). When the same women were reclassified by using the breast density model, their risk was appropriately increased so that the overall model would be well calibrated, even though this meant that some women in whom breast cancer was not diagnosed had their risk estimate increased. Further comparisons of the 2 models in additional populations will help to clarify their relative value.

The BI-RADS measure of breast density has modest reproducibility. In an earlier study (47), we found that without specific training of the radiologists, intraobserver variability for the

measure was fair ( $\kappa = 0.72$ ; 83% agreement) and interobserver variability was modest ( $\kappa = 0.59$ ; 75% agreement). Therefore, a woman's estimated risk for breast cancer might differ somewhat from radiologist to radiologist and from one examination to the next. One study has reported that 2 longitudinal measures of BI-RADS breast density may better predict a woman's risk for breast cancer than a single measure (48).

Alternatively, a more precise and reproducible measure than BI-RADS breast density that did not depend on the subjective impression of radiologists may improve the utility of breast density for risk prediction. Breast density can be measured as a continuous risk factor, and recent studies have preferred to use the percentage of mammographic density (area of density divided by total breast area) as the unit of measure (10,13,49). The extremes of BI-RADS breast density, which capture most of risk for breast cancer attributed to breast density, correlate well with quantitative measures (50,51). This implies that a categorical measurement may be sufficient for risk assessment. However, new automated approaches being pilot tested include a phantom calibrator in the mammographic image that should more precisely and accurately characterize breast tissue composition (52,53). However, none of these quantitative approaches are available for clinical use. As better measurements of breast density become available, our model can be updated by using the more precise measurement.

No single model can address all needs in breast cancer risk assessment. For example, our breast density model does not adequately capture risk in women with a very strong family history of breast cancer and other diseases associated with hereditary breast cancer syndromes, such as ovarian cancer, prostate cancer, sarcomas, and thyroid disease. These patients should be identified and referred to genetic counselors for detailed pedigree analysis and for genetic testing when appropriate. None of the general risk assessment tools, including those developed to capture detailed family history (54–56), adequately quantifies the nuances of these patients' family and personal history. For the general population, the most cost-effective and efficient approach may be stepwise, starting with a simple model and family history. Women above a certain risk level may then be referred for more detailed risk assessment by using a comprehensive model, such as that of Tyrer and colleagues (56). In women at average or low risk for breast cancer, risk assessment should be repeated periodically to capture changes in risk over time. In either scenario, risk estimates apply to populations, not individuals, and those estimates should always be combined with the potential benefits and harms of tests and treatments (along with a woman's preference for those tests and treatments) to make the best clinical decisions for individual patients.

In summary, we developed a risk prediction model that incorporates breast density to estimate a woman's 5-year risk for invasive breast cancer. It is well calibrated in the major race and ethnicity groups in the United States and across the age range of women who would be considered candidates for chemoprophylaxis for breast cancer. It is convenient enough that it could be incorporated into routine breast cancer screening, and primary care physicians could use it to calculate an individual woman's breast cancer risk. However, its accuracy must be further evaluated in independent populations before it can be recommended for clinical use.

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Let a = the age of a woman at the beginning of the risk prediction interval, t = the length of the interval, and r = her race or ethnicity. We initially estimated the woman's absolute risk (AR) by using the following equation:

 $\begin{array}{l} \mathsf{AR} = \mathsf{S}_{\mathsf{x}=3} \overset{a}{=} \mathsf{t}^{1} \mathsf{I}_{\mathsf{x},\mathsf{r}} * \mathsf{P}_{\mathsf{x},\mathsf{r}} \text{ in which} \\ \mathsf{I}_{\mathsf{x},\mathsf{r}} = \mathsf{the} \text{ age and race specific invasive breast cancer incidence and} \\ \mathsf{P}_{\mathsf{x},\mathsf{r}} = \mathsf{the proportion of women at risk for breast cancer at age x} \end{array}$ 

We estimated the age-specific annual incidence of breast cancer by fitting a third-order polynomial to the 1998 to 2002 SEER invasive breast cancer rates stratified by race or ethnicity. The breast cancer incidence (per 100 000) for women age X years was modeled as follows:  $I_{x,r} = aX^3 + bX^2 + cX + d$  in which the coefficients vary by race or ethnicity r as follows:

Race or Ethnicity	a	b	с	d
White	-0.000007231	0.001129372	-0.044330046	0.520553050
Black	-0.000004654	0.000689680	-0.023186530	0.218993144
Asian	-0.000003314	0.000421460	-0.009405533	0.004344518
Hispanic	-0.000003928	0.000585266	-0.020678768	0.210555571

After the first year, the number at risk for breast cancer is decreased by the number of women with a diagnosis of breast cancer and the number of women dying from causes other than breast cancer. We estimated the age-specific mortality rate by fitting an exponential curve to the 2002 U.S. Vital Statistics (total death rate minus breast cancer death) stratified by race or ethnicity. The non-breast cancer death (per 100 000) for women age X years was modeled as follows:

 $D_{x,r} = n^* e^{(m^*X)}$  in which the coefficients vary by race or ethnicity r as follows:

Race or Ethnicity	n	m
White	0.003695745	0.08844949
Black	0.011685888	0.077263338
Asian	0.01333534	0.09479173
Hispanic	0.002435331	0.091388373

Using  $I_{x,r}$  and  $D_{x,r}$  = we can now define the proportion of women at risk for breast cancer,  $P_{x,r}$ :  $P_{x,r} = 1$  for x = a

 $P_{x,r}^{x,i} = 1 - (1 - P_{x-l,r}) - I_{x-l,r} - D_{x-l,r}$  for x > a

To calculate the 5-year risk, we calculated the incidence during 5 consecutive years. This gives the average 5-year risk for a woman age X years of 1 specific race or ethnicity. Next, we adjusted the risk based on the woman's breast density and the distribution of breast density for women her age. The relative hazards were calculated by using the proportion hazards model described in the paper. The relative hazards for the final model were as follows:

Age	Relative Hazard
<65 y	
BI-RADS 1	0.48107
BI-RADS 2	1.00
BI-RADS 3	1.5506
BI-RADS 4	2.0119
≥65 y	
BI-RADS 1	0.65706
BI-RADS 2	1.00
BI-RADS 3	1.3881
DI DADC 4	4.4400
BI-RADS 4	1.4499

The calculated average 5-year incidence was then adjusted to that of the reference group, women with scattered fibroglandular densities (BI-RADS 2 density), based on the hazard ratios for each density category and the age- and race- or ethnicity-specific distribution of density. The adjustment factors were as follows:

Race or Ethnicity	Age 35–44 Years	Age 45–54 Years	Age 55–64 Years	Age 65–74 Years	Age ≥75 Years
White	0.7182	0.7565	0.8246	0.9069	0.9134
Black	0.7470	0.7810	0.8486	0.9201	0.9334
Asian	0.6341	0.6742	0.7530	0.8866	0.9127
Hispanic	0.7535	0.7950	0.8780	0.9615	0.9865

The adjusted incidence is then multiplied by the hazard ratio based on the woman's mammographic density. If family history and breast biopsy status are unknown, no further calculations are done. If a woman has no first-degree relatives with breast cancer, the estimated incidence is multiplied by 0.938; if she has at least 1 first-degree relative, it is multiplied by 1.454. Similarly, if a woman has never had a breast biopsy, the estimated incidence is multiplied by 0.906; if she has had a least 1 biopsy, her estimated incidence is multiplied by 1.495. The final product is the woman's estimated 5-year risk based on her age, race or ethnicity, breast density, family history, and breast biopsy history.

Appendix Figure. The Breast Cancer Surveillance Consortium breast density model algorithm

# BI-RADS = Breast Imaging Reporting and Data System; SEER = Surveillance, Epidemiology, and End Results.

Let a = the age of a woman at the beginning of the risk prediction interval, t = the length of the interval, and r = her race or ethnicity. We initially estimated the woman's absolute risk (AR) by using the following equation:

 $AR = S_{x=a}^{a+t} I_{x,r} * P_{x,r}$  in which

 $x_{r,r} =$  the age and race specific invasive breast cancer incidence and  $P_{x,r} =$  the proportion of women at risk for breast cancer at age x

We estimated the age-specific annual incidence of breast cancer by fitting a third-order polynomial to the 1998 to 2002 SEER invasive breast cancer rates stratified by race or ethnicity. The breast cancer incidence (per 100 000) for women age X years was modeled as follows:  $I_{x,r} = aX^3 + bX^2 + cX + d$  in which the coefficients vary by race or ethnicity r as follows:

Race or Ethnicity	a	b	с	d
White	-0.000007231	0.001129372	-0.044330046	0.520553050
Black	-0.000004654	0.000689680	-0.023186530	0.218993144
Asian	-0.000003314	0.000421460	-0.009405533	0.004344518
Hispanic	-0.000003928	0.000585266	-0.020678768	0.210555571

After the first year, the number at risk for breast cancer is decreased by the number of women with a diagnosis of breast cancer and the number of women dying from causes other than breast cancer. We estimated the age-specific mortality rate by fitting an exponential curve to the 2002 U.S. Vital Statistics (total death rate minus breast cancer death) stratified by race or ethnicity. The non-breast cancer death (per 100 000) for women age X years was modeled as follows:

 $D_{x,r} = n^* e^{(m^*X)}$  in which the coefficients vary by race or ethnicity r as follows:

Race or Ethnicity	n	m
White	0.003695745	0.08844949
Black	0.011685888	0.077263338
Asian	0.01333534	0.09479173
Hispanic	0.002435331	0.091388373

Using  $I_{x,r}$  and  $D_{x,r}$  = we can now define the proportion of women at risk for breast cancer,  $P_{x,r}$ :  $P_{x,r}$  = 1 for x=a

 $P_{x,r} = 1 - (1 - P_{x-l,r}) - I_{x-l,r} - D_{x-l,r}$  for x > a

To calculate the 5-year risk, we calculated the incidence during 5 consecutive years. This gives the average 5-year risk for a woman age X years of 1 specific race or ethnicity. Next, we adjusted the risk based on the woman's breast density and the distribution of breast density for women her age. The relative hazards were calculated by using the proportion hazards model described in the paper. The relative hazards for the final model were as follows:

Age	Relative Hazard
<65 y	
BI-RADS 1	0.48107
BI-RADS 2	1.00
BI-RADS 3	1.5506
BI-RADS 4	2.0119
≥65 y	
BI-RADS 1	0.65706
BI-RADS 2	1.00
BI-RADS 3	1.3881
BI-RADS 4	1.4499

The calculated average 5-year incidence was then adjusted to that of the reference group, women with scattered fibroglandular densities (BI-RADS 2 density), based on the hazard ratios for each density category and the age- and race- or ethnicity-specific distribution of density. The adjustment factors were as follows:

Race or Ethnicity	Age 35-44 Years	Age 45–54 Years	Age 55-64 Years	Age 65–74 Years	Age ≥75 Years
White	0.7182	0.7565	0.8246	0.9069	0.9134
Black	0.7470	0.7810	0.8486	0.9201	0.9334
Asian	0.6341	0.6742	0.7530	0.8866	0.9127
Hispanic	0.7535	0.7950	0.8780	0.9615	0.9865

The adjusted incidence is then multiplied by the hazard ratio based on the woman's mammographic density. If family history and breast biopsy status are unknown, no further calculations are done. If a woman has no first-degree relatives with breast cancer, the estimated incidence is multiplied by 0.938; if she has at least 1 first-degree relative, it is multiplied by 1.454. Similarly, if a woman has never had a breast biopsy, the estimated incidence is multiplied by 0.906; if she has ad a least 1 biopsy, her estimated incidence is multiplied by 1.459. The final product is the woman's estimated 5-year risk based on her age, race or ethnicity, breast density, family history, and breast biopsy history.

Appendix Figure. The Breast Cancer Surveillance Consortium breast density model algorithm BI-RADS = Breast Imaging Reporting and Data System; SEER = Surveillance, Epidemiology, and End Results.

#### Appendix Table 1 Change in Risk Categorization by Using the Breast Cancer Surveillance Consortium Breast Density Model Compared with the Risk Factor Model<sup>\*</sup>

5-Year Risk in Risk	5-Year Ri	isk in BCSC	Breast Density	Model	Row Totals	Row Totals         Events (95% CI), n (%)			Accuracy of Model
Model	0 to <1%	1% to 2%	2% to 2.99%	≥3%		Correctly Reclassified $\dot{f}$	Incorrectly Reclassified $^{\ddagger}$	without Breast Density for Women with Risk ≥2%	with Breast Density for Women with Risk ≥2%
Total, <i>n</i>	249 959	253 656	90 201	35 413	629 229	118 943 (19)	95 401 (15)	_	_
Events	1761	3649	2129	1245	8784				
Nonevents	248 198	250 007	88 072	34 168	620 445				
0 to <1%, <i>n</i>	176 831	38 571	0	0	215 402	415 (0.2)	38 156 (18)	-	-
Events	1161	415	0	0	1576				
Nonevents	175 670	38 156	0	0	213 826				
1% to 2%, <i>n</i>	71 012	184 299	43 767	476	299 554	71 446 (24)	43 809 (15)	-	-
Events	585	2749	1002	17	4353				
Nonevents	70 427	181 550	42 765	459	295 201				
2% to 2.99%, n	2116	26 990	31 805	12 987	73 898	29 109 (39)	12 984 (18)	True positive: 2855 (32%)	True positive: 3374 (38%)
Events	15	422	738	440	1615				
Nonevents	2101	26 568	31 067	12 547	72 283			False positive: 111 418 (18%)	False positive: 122 240 (20%)
≥3%, <i>n</i>	0	3796	14 629	21 950	40 375	17 973 (44)	452 (1.1)	PPV: 2.50%	PPV: 2.69%
Events	0	63	389	788	1240				
Nonevents	0	3733	14 240	21 162	39 135				

The risk factor model uses the same algorithm as the breast density model, but breast density was removed from the calculations. BCSC = Breast Cancer Surveillance Consortium; PPV = positive predictive value.

<sup>†</sup>Events were reassigned to higher-risk categories and nonevents were reassigned to lower-risk categories.

<sup>#</sup>Events were reassigned to lower-risk categories and nonevents were reassigned to higher-risk categories.

Appendix Table 1

Change in Risk Categorization by Using the Breast Cancer Surveillance Consortium Breast Density Model Compared with the Risk Factor Model<sup>\*</sup>

5-Year Risk in Risk Factor Model	5-Year Ri	5-Year Risk in BCSC Breast Density Model				Events (95	% CI), n (%)	Accuracy Accuracy of Model of Model	
	0 to <1%	1% to 2%	2% to 2.99%	≥3%		Correctly Reclassified $^{\dagger}$	Incorrectly Reclassified <sup><math>\ddagger</math></sup>	Without Breast Density for Women with Risk ≥2%	With Breast Density for Women with Risk ≥2%
Total, <i>n</i>	249 959	253 656	90 201	35 413	629 229	118 943 (19)	95 401 (15)	_	_
Events	1761	3649	2129	1245	8784				
Nonevents	248 198	250 007	88 072	34 168	620 445				
0 to <1%, <i>n</i>	176 831	38 571	0	0	215 402	415 (0.2)	38 156 (18)	-	-
Events	1161	415	0	0	1576				
Nonevents	175 670	38 156	0	0	213 826				

5-Year Risk in Risk	5-Year Risk in BCSC Breast Density Model			Row Totals	Row Totals         Events (95% CI), n (%)			Accuracy of Model	
Model	0 to <1%	1% to 2%	2% to 2.99%	≥3%		Correctly Reclassified $^{\dagger}$	Incorrectly Reclassified $\stackrel{+}{\neq}$	Without Breast Density for Women with Risk ≥2%	with Breast Density for Women with Risk ≥2%
1% to 2%, <i>n</i>	71 012	184 299	43 767	476	299 554	71 446 (24)	43 809 (15)	_	_
Events	585	2749	1002	17	4353				
Nonevents	70 427	181 550	42 765	459	295 201				
2% to 2.99%, n	2116	26 990	31 805	12 987	73 898	29 109 (39)	12 984 (18)	True positive: 2855 (32%)	True positive: 3374 (38%)
Events	15	422	738	440	1615				
Nonevents	2101	26 568	31 067	12 547	72 283			False positive: 111 418 (18%)	False positive: 122 240 (20%)
≥3%, <i>n</i>	0	3796	14 629	21 950	40 375	17 973 (44)	452 (1.1)	PPV: 2.50%	PPV: 2.69%
Events	0	63	389	788	1240				
Nonevents	0	3733	14 240	21 162	39 135				

The risk factor model uses the same algorithm as the breast density model, but breast density was removed from the calculations. BCSC = Breast Cancer Surveillance Consortium; PPV = positive predictive value.

<sup>†</sup>Events were reassigned to higher-risk categories and nonevents were reassigned to lower-risk categories.

Events were reassigned to lower-risk categories and nonevents were reassigned to higher-risk categories.

#### Appendix Table 2 Change in Risk Categorization by Using the Breast Cancer Surveillance Consortium Breast Density Model Compared with the Gail Model<sup>\*</sup>

5-Year Risk in	5-Year Ri	5-Year Risk in BCSC Breast Density Model R		Row Totals	Events (95	Events (95% CI), n (%)			
	0 to <1%	1% to 2%	2% to 2.99%	≥3%		Correctly Reclassified $\dagger$	Incorrectly Reclassified $\ddagger$	Model for Women with Risk ≥2%	Women with Risk ≥2%
Total, <i>n</i>	249 959	253 656	90 201	35 413	629 229	81 786 (13)	197 741 (31)	_	_
Events	1761	3649	2129	1245	8784				
Nonevents	248 198	250 007	88 072	34 168	620 445				
0 to <1%, n	201 037	111 008	14 552	884	327 481	1879 (0.6)	124 565 (38)	_	_
Events	1350	1552	306	21	3229				
Nonevents	199 687	109 456	14 246	863	324 252				
1% to 2%, n	45 598	121 807	51 984	12 912	232 301	46 865 (20)	63 629 (27)	_	-
Events	379	1760	1223	423	3785				
Nonevents	45 219	120 047	50 761	12 489	228 516				
2% to 2.99%, n	2361	14 506	14 497	9256	40 620	16 943 (42)	9180 (23)	True positive: 1770 (20%)	True positive: 3374 (38%)
Events	19	230	353	325	927				
Nonevents	2342	14 276	14 144	8931	39 693			False positive: 67 677 (11%)	False positive: 122 240 (20%)

5-Year Risk in	5-Year Ri	isk in BCSC	Breast Density	Model	Row Totals	Events (95	% CI), n (%)	Accuracy of Gail	Accuracy of for
	0 to <1%	1% to 2%	2% to 2.99%	≥3%		Correctly Reclassified <sup><math>\dagger</math></sup>	Incorrectly Reclassified <sup><math>\ddagger</math></sup>	Model for Women with Risk ≥2%	women with Risk ≥2%
≥3%, <i>n</i>	963	6335	9168	12 361	28 827	16 099 (56)	367 (1.3)		
Events	13	107	247	476	843			PPV: 2.55%	PPV: 2.69%
Nonevents	950	6228	8921	11 885	27 984				

BCSC = Breast Cancer Surveillance Consortium; PPV = positive predictive value.

<sup>†</sup>Events were reassigned to higher-risk categories and nonevents were reassigned to lower-risk categories.

<sup>#</sup>Events were reassigned to lower-risk categories and nonevents were reassigned to higher-risk categories

Appendix Table 2

Change in Risk Categorization by Using the Breast Cancer Surveillance Consortium Breast Density Model Compared with the Gail Model<sup>\*</sup>

5-Year Risk in	5-Year Ri	sk in BCSC	Breast Density	Model	Row Totals	Events (95	% CI), n (%)	Accuracy of Gail	Accuracy of for
	0 to <1%	1% to 2%	2% to 2.99%	≥3%		Correctly Reclassified $^{\dot{\tau}}$	Incorrectly Reclassified $\ddagger$	Model for Women with Risk ≥2%	women with Risk ≥2%
Total, n	249 959	253 656	90 201	35 413	629 229	81 786 (13)	197 741 (31)	-	_
Events	1761	3649	2129	1245	8784				
Nonevents	248 198	250 007	88 072	34 168	620 445				
0 to <1%, <i>n</i>	201 037	111 008	14 552	884	327 481	1879 (0.6)	124 565 (38)	_	_
Events	1350	1552	306	21	3229				
Nonevents	199 687	109 456	14 246	863	324 252				
1% to 2%, n	45 598	121 807	51 984	12 912	232 301	46 865 (20)	63 629 (27)	_	_
Events	379	1760	1223	423	3785				
Nonevents	45 219	120 047	50 761	12 489	228 516				
2% to 2.99%, n	2361	14 506	14 497	9256	40 620	16 943 (42)	9180 (23)	True positive: 1770 (20%)	True positive: 3374 (38%)
Events	19	230	353	325	927				
Nonevents	2342	14 276	14 144	8931	39 693			False positive: 67 677 (11%)	False positive: 122 240 (20%)
≥3%, <i>n</i>	963	6335	9168	12 361	28 827	16 099 (56)	367 (1.3)		
Events	13	107	247	476	843			PPV: 2.55%	PPV: 2.69%
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BCSC = Breast Cancer Surveillance Consortium; PPV = positive predictive value.

<sup>†</sup>Events were reassigned to higher-risk categories and nonevents were reassigned to lower-risk categories.

 $\stackrel{\neq}{}$ Events were reassigned to lower-risk categories and nonevents were reassigned to higher-risk categories

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#### Table 1

#### **Baseline Patient Characteristics**\*

Breast , n (%)	All Patier

Risk Factor	Patients without Breast Cancer, n (%)	Patients with Breast Cancer, n (%)	All Patients, n (%)
Age			
35–39 у	102 884 (10)	518 (4)	103 402 (9)
40–44 y	213 318 (20)	1557 (11)	214 875 (20)
45–49 y	182 151 (17)	2038 (14)	184 189 (17)
50–54 у	162 150 (15)	2284 (15)	164 434 (15)
55–59 у	113 155 (10)	2049 (14)	115 204 (11)
60–64 y	89 792 (8)	1753 (12)	91 545 (8)
65–69 y	79 944 (7)	1718 (12)	81 662 (7)
70–74 у	66 482 (6)	1543 (10)	68 025 (6)
75–79 у	47 031 (4)	904 (6)	47 935 (4)
80–84 y	23 811 (2)	402 (3)	24 213 (2)
Race or ethnicity			
White, non-Hispanic	764 265 (71)	11 181 (76)	775 446 (71)
Black, non-Hispanic	72 995 (7)	853 (6)	73 848 (7)
Asian or Pacific Islander	28 931 (3)	249 (2)	29 180 (3)
Native American	8746 (1)	59 (0.4)	8805 (1)
Hispanic	81 866 (8)	694 (5)	82 560 (8)
Other, mixed, unknown	123 915 (11)	1730 (12)	125 645 (11)
First-degree relatives with breast cancer			
0	891 022 (82)	11 246 (76)	902 268 (82)
1	90 144 (8)	1821 (12)	91 965 (8)
$\geq 2$	35 626 (3)	951 (6)	36 577 (3)
Missing	69 926 (6)	748 (5)	64 674 (6)
Age at menarche			
≥14 y	80 559 (7)	1212 (8)	81 771 (7)
12 or 13 y	98 980 (9)	1557 (11)	100 537 (9)
<12 y	147 535 (14)	2417 (16)	149 952 (14)
Missing	753 644 (70)	9580 (65)	763 224 (70)
Age at first birth			
<20 y	91 284 (8)	979 (7)	92 263 (8)
20–24 у	149 142 (14)	2089 (14)	151 231 (14)
25-29 y or nulliparous	183 772 (17)	2744 (19)	186 516 (17)
≥30 y	74 317 (7)	1132 (8)	75 449 (7)
Missing	582 203 (54)	7822 (53)	590 025 (54)
Breast biopsies			
0	671 587 (62)	7184 (49)	678 771 (62)
≥1	168 784 (16)	4212 (28)	172 996 (16)
Missing	240 347 (22)	3370 (23)	243 717 (22)
<b>BI-RADS</b> breast density classification			
Almost entirely fat (category 1)	100 502 (9)	831 (6)	101 333 (9)

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Risk Factor	Patients without Breast Cancer, n (%)	Patients with Breast Cancer, n (%)	All Patients, n (%)
Scattered fibroglandular densities (category 2)	474 282 (44)	5968 (40)	480 250 (44)
Heterogeneously dense (category 3)	407 050 (38)	6281 (43)	413 331 (38)
Extremely dense (category 4)	98 884 (9)	1686 (11)	100 570 (9)

\*BI-RADS — Breast Imaging Reporting and Data System.

## Table 2 Variation in the Distribution of Mammographic Density, by Age and Race or Ethnicity\*

BI-RADS Density Category		Race or Ethnic	ity, n (%)	
	White	Black	Asian	Hispanic
Age <50 y				
Almost entirely fat	30 969 (4)	4579 (6)	349 (1)	6133 (7)
Scattered fibroglandular densities	265 277 (35)	31 741 (38)	4852 (19)	33 418 (38)
Heterogeneously dense	354 545 (47)	38 619 (47)	12 978 (52)	38 071 (44)
Extremely dense	103 397 (14)	8085 (10)	6752 (27)	9320 (11)
Age 50–64 y				
Almost entirely fat	66 005 (7)	6671 (9)	1188 (4)	11 426 (12)
Scattered fibroglandular densities	413 961 (47)	37 267 (49)	10 450 (35)	45 999 (49)
Heterogeneously dense	351 884 (40)	28 956 (38)	14 438 (48)	30 833 (33)
Extremely dense	56 931 (6)	3427 (4)	3945 (13)	5027 (5)
Age ≥65 y				
Almost entirely fat	66 184 (11)	5807 (12)	1889 (13)	9697 (20)
Scattered fibroglandular densities	320 899 (54)	26 411 (56)	6825 (46)	25 593 (54)
Heterogeneously dense	187 657 (32)	13 600 (29)	5049 (34)	10 984 (23)
Extremely dense	20 535 (3)	1205 (3)	982 (7)	1348 (3)

\*BI-RADS — Breast Imaging Reporting and Data System.

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# Overall Calibration and Discrimination of the Breast Cancer Surveillance Consortium Breast Density Model among Women Table 3 with 5 Years of Follow-up

Sample	Women, <i>n</i>	Expected 5-Year Rate	Observed 5-Year Rate	Expected-Observed Rate Ratio (95% CI) <sup>*</sup>	Concordance Statistic (95% CI)
Development	377 440	1.41	1.41	1.00 (0.98–1.03)	0.657 (0.65–0.67)
Validation	251 789	1.41	1.38	1.03 (0.99–1.06)	0.660 (0.65–0.66)
Total	629 229	1.41	1.40	1.01 (0.99–1.03)	0.658 (0.65–0.66)

\* Expected rate divided by the observed rate. The observed 5-year rate is the actual rate observed in the individual subcohorts. The expected rate is the average of the Breast Cancer Surveillance Consortium breast density model predicted risk for each woman in the subcohort. No additional adjustments were done.

# Table 4 Calibration of the Breast Cancer Surveillance Consortium Breast Density Model in Risk Factor Subgroups\*

Risk Factor Group	Expected 5-Year Rate	Observed 5-Year Rate	Expected– Observed Rate Ratio (95% CI)	Concordance Statistic
Total cohort	1.41	1.40	1.01 (0.99–1.03)	0.66
Age				
40–44 y	0.68	0.73	0.94 (0.89–1.00)	0.63
45–49 y	1.06	1.07	0.99 (0.94–1.04)	0.61
50–54 y	1.33	1.38	0.96 (0.92–1.01)	0.62
55–59 y	1.72	1.77	0.97 (0.92–1.02)	0.64
60–64 y	1.94	1.87	1.04 (0.98–1.10)	0.63
65–69 y	2.23	1.97	1.13 (1.07–1.20)	0.60
70–74 y	2.34	2.17	1.08 (1.02–1.15)	0.61
Race or ethnicity				
White, non-Hispanic	1.48	1.46	1.02 (0.99–1.04)	0.66
Black, non-Hispanic	1.17	1.18	1.00 (0.91–1.09)	0.63
Asian	0.95	0.99	0.95 (0.81-1.12)	0.66
Hispanic	0.95	1.01	0.94 (0.85–1.04)	0.67
Other, mixed, unknown	1.45	1.41	1.03 (0.97–1.09)	0.65
BI-RADS breast density classification				
Almost entirely fat (category 1)	0.78	0.75	1.04 (0.95–1.14)	0.67
Scattered fibroglandular densities (category 2)	1.27	1.24	1.02 (0.99–1.06)	0.64
Heterogeneously dense (category 3)	1.66	1.65	1.00 (0.97–1.04)	0.65
Extremely dense (category 4)	1.80	1.81	1.00 (0.94–1.06)	0.64
First-degree relatives with breast cancer				
0	1.30	1.32	0.99 (0.96–1.01)	0.65
1	2.11	1.98	1.07 (1.01–1.13)	0.64
≥2	2.52	2.44	1.04 (0.95–1.13)	0.62
Breast biopsies				
0	1.19	1.16	1.03 (0.99–1.06)	0.64
≥1	2.35	2.37	0.99 (0.96-1.03)	0.62
Decile of predicted risk				
1st	0.40	0.39	1.04 (0.92–1.18)	0.64
2nd	0.61	0.66	0.93 (0.84–1.02)	0.56
3rd	0.76	0.84	0.91 (0.83-0.99)	0.57
4th	0.92	0.95	0.97 (0.89–1.05)	0.58
5th	1.12	1.07	1.04 (0.97–1.12)	0.57
6th	1.30	1.37	0.95 (0.89–1.02)	0.58
7th	1.53	1.47	1.04 (0.98–1.11)	0.58

Risk Factor Group	Expected 5-Year Rate	Observed 5-Year Rate	Expected– Observed Rate Ratio (95% CI)	Concordance Statistic
8th	1.85	1.87	0.99 (0.93–1.05)	0.57
9th	2.26	2.27	0.99 (0.94–1.05)	0.55
10th	3.40	3.10	1.10 (1.05–1.15)	0.58
Quintile of Gail risk				
1st	0.75	0.76	0.99 (0.93–1.05)	0.62
2nd	1.04	1.05	0.99 (0.94–1.04)	0.64
3rd	1.30	1.29	1.01 (0.96–1.06)	0.62
4th	1.72	1.69	1.02 (0.98–1.06)	0.62
5th	2.26	2.20	1.03 (0.99–1.07)	0.61
Gail high risk				
No (<1.67%)	1.22	1.21	1.00 (0.98–1.03)	0.65
Yes (≥1.67%)	2.37	2.29	1.03 (0.99–1.07)	0.61

\* Among the 629 229 women with complete 5-year follow-up. The observed 5-year rate is the actual rate observed in the individual subcohorts. The expected rate is the average of the Breast Cancer Surveillance Consortium breast density model predicted risk for each woman in the subcohort. No additional adjustments were done. BI-RADS = Breast Imaging Reporting and Data System.

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BI-RADS Density Category	5-Year Risk for Breast Cancer $\overset{\gamma}{r}$	5-Year Risk with N	o Family History <sup>†</sup>	5-Year Risk with	Family History <sup>†</sup>
		No Breast Biopsy	History of Breast Biopsy	No Breast Biopsy	History of Breast Biopsy
Age 40–44 y					
Almost entirely fat	0.2	0.2	0.3	0.3	0.5
Scattered fibroglandular densities	0.5	0.4	0.7	0.7	1.1
Heterogeneously dense	0.7	0.7	1.1	1.0	1.7
Extremely dense	1.0	0.8	1.4	1.3	2.2
Age 45–49 y					
Almost entirely fat	0.4	0.3	0.5	0.5	0.8
Scattered fibroglandular densities	0.8	0.7	1.0	1.0	1.7
Heterogeneously dense	1.2	1.0	1.7	1.6	2.6
Extremely dense	1.6	1.3	2.1	2.1	3.3
Age 50–54 y					
Almost entirely fat	0.5	0.4	0.7	0.7	1.1
Scattered fibroglandular densities	1.0	0.9	1.4	1.4	2.2
Heterogeneously dense	1.6	1.3	2.2	2.1	3.4
Extremely dense	2.1	1.7	2.8	2.7	4.4
Age 55–59 y					
Almost entirely fat	0.7	0.6	0.9	6.0	1.5
Scattered fibroglandular densities	1.4	1.2	1.9	1.9	3.0
Heterogeneously dense	2.2	1.8	3.0	2.8	4.6
Extremely dense	3.0	2.3	3.8	3.6	5.9
Age 60–64 y					
Almost entirely fat	0.8	0.7	1.1	1.1	1.7
Scattered fibroglandular densities	1.7	1.4	2.3	2.2	3.5
Heterogeneously dense	2.6	2.1	3.5	3.3	5.4
Extremely dense	3.4	2.7	4.4	4.2	6.9
Age 65–69 y					

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DI-KADS DEnsity Category	5-Year Risk for Breast Cancer <sup>†</sup>	5-Year Risk with N	do Family History $^{\hat{T}}$	5-Year Risk with	Family History $\dot{r}$
		No Breast Biopsy	History of Breast Biopsy	No Breast Biopsy	History of Breast Biopsy
Almost entirely fat	1.3	1.1	1.8	1.8	2.9
Scattered fibroglandular densities	2.0	1.7	2.8	2.6	4.3
Heterogeneously dense	2.9	2.3	3.8	3.6	5.9
Extremely dense	3.0	2.4	4.0	3.8	6.2
Age 70–74 y					
Almost entirely fat	1.4	1.2	1.9	1.8	3.0
Scattered fibroglandular densities	2.1	1.8	2.9	2.8	4.5
Heterogeneously dense	3.1	2.5	4.0	3.8	6.2
Extremely dense	3.3	2.5	4.2	3.9	6.4

BI-RADS = Breast Imaging Reporting and Data System.

fMean predicted risk among the 629 229 women with complete 5-year follow-up. A woman was considered to have a family history of breast cancer if invasive breast cancer was diagnosed in  $\geq$ 1 firstdegree relative.

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ear Risk in ¢ Factor		5-Year Risk in BCSC B	<b>3reast Density Model</b>		Row Totals	Events (95%	% CI), n (%)	Accuracy of Model without	Accuracy of Model with
	0 to <1%	1% to 1.66%	1.67% to 2.49%	≥2.5%		Correctly Reclassified $^{\dagger}$	Incorrectly Reclassified ${}^{\sharp}$	Breast Density for Women with Risk ≥1.67%	Breast Density for Women with Risk ≥1.67%
ll, n	249 959	186 106	124 420	68 744	629 229	137 404 (22)	61) 777 (16)		
vents	1761	2390	2513	2120	8784			Ι	Ι
onevents	248 198	183 716	121 907	66 624	620 445				
<1%, <i>n</i>	176 831	38 500	71	0	215 402	415 (0.2)	38 156 (18)		
ents	1161	415	0	0	1576			Ι	Ι
onevents	175 670	38 085	71	0	213 826				
o 1.66%, n	64 297	99 456	37 149	1025	201 927	64 557 (32)	37 914 (19)		
ents	526	1328	754	32	2640			Ι	Ι
nevents	63 771	98 128	36 395	993	199 287				
6 to 2.49%,	8741	45 478	71 309	23 267	148 795	54 157 (36)	23 329 (16)	True positive: 4568 (52%)	True positive: 4633 (53%)
ents	74	609	1419	621	2723				
nevents	8667	44 869	69 890	22 646	146 072			False positive: 207 332 (33%)	False positive: 188 531 (30%)
%, n	90	2672	15 891	44 452	63 105	18 275 (29)	378 (0.6)		
ents	0	38	340	1467	1845			PPV: 2.16%	PPV: 2.40%
nevents	90	2634	15 551	42 985	61 260				

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tEvents were reassigned to higher-risk categories and nonevents were reassigned to lower-risk categories. tEvents were reassigned to lower-risk categories and nonevents were reassigned to higher-risk categories.

5-Year Risk in Gail Model		5-Year Risk in BCSC l	Breast Density Model		Row Totals	Events (95 <sup>e</sup>	% CI), n (%)	Accuracy of Gail Model for	Accuracy of for Women
	0 to <1%	1% to 1.66%	1.67% to 2.49%	≥2.5%		Correctly Reclassified $^{\hat{r}}$		Women with Risk ≥1.67%	with Kisk ≥1.67%
Total, <i>n</i>	249 959	186 106	124 420	68 744	629 229	85 262 (14)	222 672 (35)	1	I
Events	1761	2390	2513	2120	8784				
Nonevents	248 198	183 716	121 907	66 624	620 445				
0 to <1%, <i>n</i>	201 037	94 843	27 582	4019	327 481	1879 (0.6)	124 565 (38)	I	Ι
Events	1350	1192	582	105	3229				
Nonevents	199 687	93 651	27 000	3914	324 252				
1% to 1.66%, <i>n</i>	42 444	69 051	59 604	24 355	195 454	44 000 (23)	82 403 (42)	I	Ι
Events	341	875	1188	602	3113				
Nonevents	42 103	68 176	58 416	23 646	192 341				
1.67% to 2.49%, <i>n</i>	4833	17 117	26 355	15 518	63 823	22 083 (35)	15 385 (24)	True positive: 2442 (28%)	True positive: 4633 (53%)
Events	51	253	513	437	1254				
Nonevents	4782	16 864	25 842	15 081	62 569			False positive: 103 852 (17%)	False positive: 188 531 (30%)
≥2.5%, n	1645	5095	10 879	24 852	42 471	17 300 (41)	319 (0.8)		
Events	19	70	230	869	1188			PPV: 2.30%	PPV: 2.40%
Nonevents	1626	5025	10 649	23 983	41 283				

 ${\cal F}$ Events were reassigned to higher-risk categories and nonevents were reassigned to lower-risk categories.

 ${m \sharp}$  Events were reassigned to lower-risk categories and nonevents were reassigned to higher-risk categories.

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Table 7

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