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# Depressive Symptoms, Health Behaviors, and Risk of Cardiovascular Events in Patients With Coronary Heart Disease

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# Abstract

**Context**—Depressive symptoms predict adverse cardiovascular outcomes in patients with coronary heart disease, but the mechanisms responsible for this association are unknown.

**Objective**—To determine why depressive symptoms are associated with an increased risk of cardiovascular events.

**Design and Participants**—The Heart and Soul Study is a prospective cohort study of 1017 outpatients with stable coronary heart disease followed up for a mean (SD) of 4.8 (1.4) years.

**Setting**—Participants were recruited between September 11, 2000, and December 20, 2002, from 12 outpatient clinics in the San Francisco Bay Area and were followed up to January 12, 2008.

**Main Outcome Measures**—Baseline depressive symptoms were assessed using the Patient Health Questionnaire (PHQ). We used proportional hazards models to evaluate the extent to which the association of depressive symptoms with subsequent cardiovascular events (heart failure, myocardial infarction, stroke, transient ischemic attack, or death) was explained by baseline disease severity and potential biological or behavioral mediators.

**Results**—A total of 341 cardiovascular events occurred during 4876 person-years of follow-up. The age-adjusted annual rate of cardiovascular events was 10.0% among the 199 participants with depressive symptoms (PHQ score  $\geq$ 10) and 6.7% among the 818 participants without depressive

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symptoms (hazard ratio [HR], 1.50; 95% confidence interval, [CI], 1.16–1.95; P=.002). After adjustment for comorbid conditions and disease severity, depressive symptoms were associated with a 31% higher rate of cardiovascular events (HR, 1.31; 95% CI, 1.00–1.71; P=.04). Additional adjustment for potential biological mediators attenuated this association (HR, 1.24; 95% CI, 0.94–1.63; P=.12). After further adjustment for potential behavioral mediators, including physical inactivity, there was no significant association (HR, 1.05; 95% CI, 0.79–1.40; P=.75).

**Conclusion**—In this sample of outpatients with coronary heart disease, the association between depressive symptoms and adverse cardiovascular events was largely explained by behavioral factors, particularly physical inactivity.

Depression has long been recognized as a risk factor for the development of cardiovascular disease in healthy patients, for recurrent events in patients with established cardiovascular disease, and for adverse outcomes after coronary bypass graft surgery.  $^{1-6}$  Depression is also a risk factor for the development of heart failure and for adverse outcomes in patients with existing heart failure.  $^{7-10}$  In a recent survey of 245 404 adults from 60 countries, patients with comorbid depression reported worse overall health than those with asthma, diabetes, arthritis, or cardiovascular disease alone.  $^{11}$  Based on the results of this study, the World Health Organization highlighted the detrimental effects of depression on medical illnesses as 1 of its 10 most important global public health statistics for 2007.  $^{12}$ 

Despite the substantial body of evidence demonstrating a strong link between depression and cardiovascular disease, the explanation for this association remains unclear. Several candidate mechanisms have been suggested as potential mediators, including smoking, lack of exercise, <sup>13</sup> medication non-adherence, <sup>14,15</sup> worse underlying cardiac disease severity, <sup>16</sup> lower heart rate variability, <sup>17</sup> antidepressant toxicity, <sup>18</sup> enhanced activity of the hypothalamic pituitary axis, <sup>19</sup> greater catecholamine levels, <sup>20</sup> dietary factors, lowomega-3 fatty acid levels, <sup>21</sup> increased serotonin and platelet activation, <sup>22</sup> and inflammatory processes. <sup>23,24</sup> However, the extent to which these proposed mechanisms explain the increased risk of cardiovascular events in depressed patients is unknown.

Understanding how depression leads to cardiovascular events is necessary for developing interventions to decrease the excess cardiovascular morbidity and mortality associated with depression. In a prospective cohort study of 1017 participants with stable coronary disease, we evaluated the extent to which the association of depressive symptoms with subsequent cardiovascular events was explained by differences in comorbid conditions, cardiac disease severity, use of antidepressant medications, and potential biological and behavioral mediators.

# METHODS

The primary goal of the Heart and Soul Study was to determine why depression is associated with an increased risk of cardiovascular events in outpatients with stable coronary heart disease. <sup>25</sup> We used administrative databases to identify outpatients with documented coronary artery disease at 2 Department of Veterans Affairs Medical Centers (San Francisco VA Medical Center and the VA Palo Alto Health Care System, California), 1 university medical center (University of California, San Francisco), and 9 public health clinics in the Community Health Network of San Francisco.

Patients were eligible to participate if they had at least 1 of the following: a history of myocardial infarction, angiographic evidence of at least 50% stenosis in 1 or more coronary vessels, prior evidence of exercise-induced ischemia by treadmill or nuclear testing, a history of coronary revascularization, or a diagnosis of coronary artery disease documented by an internist or cardiologist. Between September 11, 2000, and December 20, 2002, a total of 1024

participants were enrolled: 240 from the public health clinics, 346 from the university medical center, and 438 from the VA medical centers.

All participants completed a baseline examination that included an interview, fasting blood draw, psychiatric interview, questionnaire, echocardiogram, exercise treadmill test, 24-hour ambulatory electrocardiogram, and 24-hour urine collection. Of the 1024 participants who completed the baseline examination, we were not able to contact 7 (<1%) during the follow-up period, leaving 1017 for this analysis. Our protocol was approved by the appropriate institutional review boards, and all participants provided written informed consent.

#### **Depressive Symptoms**

We assessed depressive symptoms using the 9-item Patient Health Questionnaire (PHQ),<sup>26</sup> a self-report instrument that measures the frequency of depressive symptoms corresponding to the 9 *Diagnostic and Statistical Manual of Mental Disorders Fourth Edition*  $(DSM-IV)^{27}$  criteria for depression. The PHQ has demonstrated excellent validity when compared with a structured diagnostic interview for depression in patients with coronary artery disease.<sup>28</sup> Participants indicated the frequency of experiencing each symptom during the prior 2 weeks; the items were scored as no points for not at all, 1 point for several days, 2 points for more than half the days, or 3 points for nearly every day. We evaluated depressive symptoms as a continuous variable (range, 0–27), as a categorical variable ( $\leq 4$ , >4-<10, 10-<15, 15-<20, and  $\geq 20$ ), and as a dichotomous variable using the standard cut point of 10 or higher.<sup>26</sup> There was no difference in the mean (SD) PHQ score between the 1017 participants included in the analysis (5.2 [5.5]) and the 7 participants lost to follow-up (3.6 [1.0]; *P*=.43).

We also ascertained the presence of major depressive disorder (past month, past year, or lifetime) using the Computerized Diagnostic Interview Schedule for the *DSM-IV* (C DIS-IV). <sup>29</sup> Computerized versions of the Diagnostic Interview Schedule have previously demonstrated acceptable validity and reliability.<sup>30,31</sup> Participants with a major depressive episode in the past month were informed of this diagnosis, instructed to discuss their symptoms with their primary care physician, and provided a list of local resources. Of the 2 participants who did not complete the C DIS-IV, one had a PHQ score of 10 or higher, and the other had a PHQ score of 0.

#### **Baseline Cardiac Disease Severity and Risk Factors**

All participants underwent resting echocardiography using an Acuson Sequoia Ultrasound System (Siemens, Mountain View, California). We obtained standard 2-dimensional views and performed planimetry with a computerized digitization system to determine left-ventricular ejection fraction. We categorized participants as having diastolic dysfunction if their mitral inflow ratio of peak early-to-late diastolic filling velocity was more than 0.75 and if the velocity time integral in their pulmonary vein was greater during diastole than during systole.<sup>32</sup> Fasting venous blood samples were drawn to determine low- and high-density lipoprotein cholesterol levels.

#### **Potential Biological Mediators**

Three-channel 24-hour ambulatory Holter electrocardiography was used to assess heart rate variability,<sup>33</sup> including the SD of 5-minute mean NN intervals and the natural log of very low frequency power. We collected 24-hour urine samples to measure norepinephrine and cortisol excretion. Norepinephrine was assessed by gas chromatography-mass spectrometry at the Associated Regional and University Pathologists, Inc (Salt Lake City, Utah). Cortisol was analyzed using radioimmunoassay or (due to a change at the laboratory) high-performance liquid chromatography/tandem mass spectrometry.<sup>19,20</sup>

We used high-pressure liquid chromatography with electrochemical detection to assay whole blood serotonin levels. High-sensitivity C-reactive protein levels were measured using the Roche (Indianapolis, Indiana) Integra assay in the first 229 participants and (due to a change in the laboratory) the Beckman Extended Range (Galway, Ireland) assay in the remaining samples.<sup>34</sup> Blood levels of 2 omega-3 fatty acids, docosahexaenoic acid and eicosapentaenoic acid, were measured by capillary gas chromatography as the percentage composition of total fatty acid methyl esters in the red blood cell membranes.<sup>35</sup>

#### **Potential Behavioral Mediators**

Smoking and alcohol use were determined by self-report questionnaire. To assess medication adherence, we asked, "In the past month, how often did you take your medications as the doctor prescribed?" Possible responses were all of the time (100%), nearly all of the time (90%), most of the time(75%), about half the time(50%), or less than half the time (<50%). We defined medication non-adherence as taking prescribed medications 75% or less of the time.<sup>36</sup> We chose to measure self-reported medication adherence for several reasons. First, self-reported medication adherence has been validated as are liable predictor of health outcomes, including blood pressure control,<sup>37</sup> hospitalization for heart failure,<sup>38</sup> serum drug concentrations,<sup>39</sup> and cardiovascular events.<sup>36,40</sup> Second, in a study of hypertensive patients taking hydrochlorothiazide, self-reported medication adherence was more strongly correlated with qualitative urinary hydrochlorothiazide levels, changes in serum potassium, and decreases in blood pressure than was pill count adherence.<sup>41</sup> Third, pharmacy refills are often provided for 90 days at a time, and thus pharmacy data would not have allowed us to assess the association of current (past month) adherence with depression. Fourth, Medication Event Monitoring System (MEMS) caps have their own limitations<sup>42</sup> because the number of cap openings does not necessarily reflect the number of pills ingested by the patient.

To assess physical activity, we asked, "Which of the following statements best describes how physically active you have been during the last month, that is, done activities such as 15 to 20 minutes of brisk walking, swimming, general conditioning, or recreational sports?" Participants chose from 1 of the following 6 categories: not at all active, a little active(1–2 times per month), fairly active (3–4 times per month), quite active (1–2 times per week), very active (3–4 times per week), or extremely active( $\geq$ 5timesperweek). Self-report has been shown to be a reliable, valid, and accurate method of assessing physical activity.<sup>43–46</sup> In particular, single-response items have demonstrated excellent construct validity.<sup>45–47</sup> Participants who reported that they were not at all or a little active were considered physically inactive.

As an objective measure of physical fitness, participants also completed an exercise treadmill test according to a standard Bruce protocol at the baseline examination. Those who were unable to continue the standard Bruce protocol were switched to slower settings and encouraged to exercise for as long as possible. Exercise capacity was calculated as the total number of metabolic equivalent tasks achieved. Treadmill exercise capacity is a gold-standard measure of physical fitness and is highly correlated with self-reported physical activity.<sup>43,46,48–50</sup>

#### **Other Patient Characteristics**

Age, sex, race, educational achievement, and medical history were determined by self-report questionnaire. We measured height and weight, and calculated body mass index. Participants were instructed to bring their medication bottles to their appointment, and study personnel recorded all current medications. Medications were categorized using Epocrates Rx (San Mateo, California).

#### **Cardiovascular Events**

Between the baseline examination and the last day of follow-up on January 12, 2008, we conducted annual telephone follow-up interviews with participants (or their proxy), asking specifically about hospitalization for "heart trouble." For any reported event, medical records, electrocardiograms, death certificates, and coroner's reports were retrieved and reviewed by 2 independent blinded adjudicators. In the event of disagreement, the adjudicators conferred, reconsidered their classification, and requested consultation from a third blinded adjudicator as necessary.

Outcome events were heart failure, myocardial infarction, stroke, transient ischemic attack, or death. *Heart failure* was defined as hospitalization for a clinical syndrome involving at least 2 of the following: paroxysmal nocturnal dyspnea, orthopnea, elevated jugular venous pressure, pulmonary rales, third heart sound, and cardiomegaly or pulmonary edema on chest radiography. These signs and symptoms must have represented a clear change from the usual clinical status.<sup>32</sup> *Nonfatal myocardial infarction* was defined using standard criteria.<sup>51</sup> *Stroke* was defined as a new neurological deficit not known to be secondary to brain trauma, tumor, infection, or other cause. *Transient ischemic attack* was defined as a focal neurological deficit(in the absence of head trauma)lasting more than 30seconds and no longer than 24 hours, with rapid evolution of the symptoms to the maximal level of deficitin less than 5minutes and with subsequent complete resolution. *Death* was determined by death certificates and coroner's reports.

#### Statistical Analysis

We aimed to enroll 1000 patients, including an expected 200 with depression. Assuming a 25% cumulative event rate over 5 years among patients without depression, this provided 80% power in 2-sided tests with an  $\alpha$  of 5% to detect a relative hazard for depression of 1.56, after accounting for loss of precision due to adjustment for confounders.<sup>52</sup> Our testing was 2-sided because we were open to the possibility that depression would be associated with better outcomes, and we did not know whether adjustment for potential mediators would increase or decrease the effect of depressive symptoms on cardiovascular events; in addition, 2-sided tests require stronger evidence at any given  $\alpha$  level to reject the null hypothesis.

Baseline differences in characteristics between participants with and without depressive symptoms were compared using *t* tests and  $\chi^2$  tests. C-reactive protein was log transformed because it did not have a normal distribution. We estimated the risk of cardiovascular events associated with depressive symptoms (both as a continuous and a dichotomous variable) using Cox proportional hazards models. We also estimated the risk of cardiovascular events associated with the presence of major depressive disorder (based on C DIS-IV interview).

To evaluate whether a covariate changed the strength of association between depressive symptoms and cardiovascular events, we calculated the percentage change in the effect size (log hazard ratio [HR]) for depressive symptoms before and after adjustment for the potential confounder or mediator. To avoid any artifact due to different sample sizes between the 2 nested models, participants missing the covariate of interest were excluded. We sequentially considered demographic variables, comorbid conditions, cardiac disease severity, use of medications, potential biological mediators, and potential behavioral mediators.<sup>53</sup> All variables that resulted in a more than 5% change in the effect size (log HR) for depressive symptoms were considered confounders or potential mediators and included in the final multivariable models,<sup>54</sup> with the exception of "other antidepressants." Adjusting for use of "other antidepressants" (eg, mirtazapine, venlafaxine, or bupropion) reduced the effect size for depressive symptoms on cardiovascular events by 8.8%. Although this met our a priori definition for potential mediation, we thought the reduction in effect size most likely occurred

In model checking, we verified the log-linearity assumption for continuous variables by checking for improvement in fit after addition of quadratic and cubic terms. We verified the proportional hazards assumption of these models using log-minus-log survival plots and by checking for secular patterns in scaled Schoenfeld residuals. We used Wald tests to check for interactions of depressive symptoms with age, race, history of myocardial infarction, stroke, diabetes, left ventricular ejection fraction, smoking status, medication adherence, and physical activity in both age- and multivariable-adjusted models. Plots of estimated cumulative risk were generated using PROC PHREG followed by PROC GPLOT. Analyses were performed using SAS version 9.0 (SAS Institute Inc, Cary, North Carolina).

# RESULTS

Of the 1017 participants, 199 (19.6%) had depressive symptoms (PHQ  $\geq$ 10). Compared with participants who did not have depressive symptoms, those with depressive symptoms were younger and less likely to be male (Table 1). They were more likely to smoke, to be less physically active, to be less adherent to medications, and to have a higher mean body mass index. Participants with depressive symptoms had more comorbid conditions, including history of myocardial infarction, heart failure, and diabetes. Depressive symptoms were also associated with greater antidepressant use, greater 24-hour urinary norepinephrine excretion and C-reactive protein levels, and lower omega-3 fatty acid levels.

A total of 341 cardiovascular events occurred during 4876 person-years with a mean (SD) of 4.8 (1.4) years' follow-up. The age-adjusted annual rate of cardiovascular events was 10.0% among the 199 participants with depressive symptoms (PHQ  $\geq$ 10) and 6.7% among the 818 participants without depressive symptoms (HR, 1.50; 95% confidence interval [CI], 1.16–1.95; *P*=.002; Table 2).

#### Potential Mediators

Several variables, including comorbid conditions, smoking status, medication adherence, physical activity, left ventricular function, and C-reactive protein level, met the criterion for potential confounding or mediation (Table 3). Most notably, adjustment for physical activity was associated with a 31.7% reduction in the strength of association between depressive symptoms and cardiovascular events. Variables that did not change the effect size for depressive symptoms by 5% or more included use of selective serotonin reuptake inhibitors or tricyclic antidepressants, heart rate variability, levels of serotonin and omega-3 fatty acids, and 24-hour excretion of norepinephrine and cortisol (Table 3).

When adjusted for comorbid conditions and cardiac disease severity, depressive symptoms remained associated with a 31% greater rate of cardiovascular events (HR, 1.31; 95% CI, 1.00– 1.71; P=.04). Additional adjustment for C-reactive protein attenuated this association (HR, 1.24; 95% CI, 0.94–1.63; P=.12; Table 4). After further adjustment for potential behavioral mediators, including physical inactivity, there was no longer a significant association between depressive symptoms and cardiovascular events (Table 4, Figure 1). Substituting exercise capacity (for self-reported physical activity) resulted in a similar reduction in the effect size for depressive symptoms (HR, 0.96; 95% CI, 0.70–1.31; = .79). In the final model, physical inactivity was associated with a 44% greater rate of cardiovascular events (HR, 1.44; 95% 1.14–1.82; P=.002), adjusting for depressive symptoms, comorbid conditions, left ventricular ejection fraction, C-reactive protein, smoking, and medication adherence (Table 5).

#### Alternative Approaches to Classifying Depression

We observed a dose-response relation between increasing depressive symptom scores and cardiovascular events (Figure 2). When adjusted for age, co-morbid conditions, and left ventricular ejection fraction, each SD (5.5-point) increase in depressive symptom score was associated with a 15% greater rate of cardiovascular events (HR, 1.15; 95% CI, 1.03–1.28; P=.01). Adjusting for log C-reactive protein minimally attenuated this relation (HR, 1.13; 95% CI, 1.02–1.26; P=.02), and further adjustment for smoking and non-adherence moderately diminished the association (HR, 1.09; 95% CI, 0.98–1.22; P=.12). Again, there was no association between depressive symptoms and cardiovascular events after adjustment for physical activity (HR, 1.03; 95% CI, 0.92–1.16; P=.53). We found no evidence that the association between depressive symptom score and cardiovascular events varied by age, race, history of myocardial infarction, stroke, diabetes, left ventricular ejection fraction, smoking, medication adherence, or physical activity (all P values for interaction  $\geq$ 0.1).

Of the 1015 participants who completed the C DIS-IV for major depressive disorder, 120 had both current depressive symptoms (PHQ $\geq$ 10) and major depressive disorder in the past month, 78 had depressive symptoms only, 102 had major depressive disorder only, and 715 had neither. In age-adjusted models, we did not observe any association between past month (HR, 1.08; 95% CI, 0.82–1.44; *P*=.56), past year (HR, 1.12; 95% CI, 0.86–1.46; *P*=.39), or lifetime (HR, 0.98; 95% CI, 0.78–1.23; *P*=.86) major depressive disorder and cardiovascular events.

# COMMENT

In this prospective cohort study of more than 1000 outpatients with stable coronary heart disease, we found that participants with baseline depressive symptoms had a 50% greater rate of subsequent cardiovascular events (heart failure, myocardial infarction, stroke, transient ischemic attack, or death) than participants without depressive symptoms. After adjustment for comorbid conditions and cardiac disease severity, depressive symptoms remained associated with a 31% increased rate of cardiovascular events. Inflammation (as measured by C-reactive protein) explained a small part of this association. However, no significant association between depressive symptoms and cardiovascular events remained after adjustment for physical activity and other health behaviors. These findings raise the hypothesis that the increased risk of cardiovascular events associated with depression could potentially be preventable with behavior modification, especially exercise.

Given the relatively modest effects of traditional therapies on depressive symptoms in patients with heart disease,  $^{55-58}$  there is increasing urgency to identify interventions that not only reduce depressive symptoms but also directly target the mechanisms by which depression leads to cardiovascular events.  $^{59}$  Several large-scale studies have clearly documented that increased physical activity reduces cardiovascular events,  $^{60-62}$  but whether exercise interventions can eliminate the excess risk of cardiovascular events associated with depressive symptoms has not been studied. Patients with depressive symptoms are less likely to adhere to dietary, exercise, and medication recommendations,  $^{13-15}$  and poor health behaviors can lead to cardiovascular events.  $^{36,62}$  Exercise training can improve both depressive symptoms and markers of cardiovascular risk.  $^{63,64}$  It is therefore possible that the combination of exercise therapy plus antidepressant medication may reduce the risk of cardiovascular events in patients with depression.

Our study evaluated depression and physical activity at the same point in time, and thus we cannot determine whether physical inactivity was the cause or result of depression. Indeed, the association is almost certainly bidirectional because depression leads to physical inactivity, 14,65-68 and physical inactivity exacerbates depression.69-72 This can result in a downward spiral in which depression and physical inactivity become mutually reinforcing. Likewise,

treating depression can increase physical activity, and physical activity can elevate mood.<sup>64</sup>, 73-75 To the extent that physical inactivity precedes depression, it would be acting as a confounder rather than as a mediator of the association between depression and cardiovascular events. Regardless of whether physical inactivity was the cause (a confounder) or the result (a mediator) of depressive symptoms, it appeared to account for almost half of the association between depressive symptoms and cardiovascular events in our sample. These findings raise the possibility that increased exercise may decrease the risk of cardiovascular events associated with depression.

Because we measured the potential behavioral mediators by self-report, it is possible that participants with depressive symptoms were more likely to underreport medication adherence and physical activity. We attempted to address this possibility by substituting exercise capacity (instead of self-reported physical activity) as an objective measure of physical fitness. Although exercise capacity may also have been influenced by worse underlying cardiac disease severity that was not otherwise accounted for in our models, adjusting for this gold standard measure of physical fitness resulted in the same findings as adjusting for self-reported physical activity.

Many cross-sectional studies have examined potential mediators of the association between depression and cardiovascular disease. However, only a few prospective studies have evaluated whether adjusting for any of these candidate mechanisms actually changes the effect size for depression on cardiovascular events. Carney et al<sup>17</sup> found that low heart rate variability partially mediated the effect of depression on survival among 311 depressed patients after an acute myocardial infarction. In light of the present results, an intriguing possibility is that the mediating effects of heart rate variability may have been related to a less physically active lifestyle among the depressed participants. Vaccarino and colleagues<sup>24</sup> reported that inflammatory biomarkers explained a small portion of the association between depression and cardiovascular disease incidence among 559 women with suspected coronary ischemia, but Empana et al<sup>76</sup> found that adjustment for inflammatory markers did not change the strength of association between depressive symptoms and incident cardiovascular disease in a nested case-control study of 1005 healthy middle-aged men. In the present study, we found that adjusting for C-reactive protein was associated with a decrease in the strength of association between depressive symptoms and cardiovascular events. However, it is unclear whether inflammation was functioning as a mediator between depressive symptoms and cardiovascular events or as a marker of worse cardiac disease severity.

Others have raised the possibility that antidepressant toxicity may be responsible for the adverse cardiovascular outcomes associated with depressive symptoms. Cohen et al<sup>77</sup> reported an excess risk of myocardial infarction in users of tricyclic antidepressants, and Sherwood et al<sup>18</sup> found that use of antidepressant medication was associated with an increased risk of cardiovascular hospitalization or death in patients with heart failure. In contrast, Taylor and colleagues evaluated 2481 patients after experiencing a myocardial infarction and found that users of selective serotonin re-uptake inhibitors were at lower risk of death or recurrent myocardial infarction than nonusers.<sup>78</sup> To evaluate the antidepressant toxicity hypothesis, we compared the effect size for depressive symptoms on cardiovascular events before and after adjustment for use of antidepressant medications. Although adjusting for use of selective serotonin reuptake inhibitors or tricyclic agents did not change the effect size for depressive symptoms, adjusting for use of "other antidepressants" reduced the effect size for depressive symptoms by 8.8%. One potential interpretation of these findings is that antidepressant toxicity may (at least partly) be responsible for the increased risk of cardiovascular events associated with depressive symptoms. However, a more likely explanation is that use of "other antidepressants" was a marker of more severe (or treatment-resistant) depression and that the effect size for depressive symptoms on cardiovascular events was reduced because adjusting for use of "other antidepressants" was tantamount to adjusting for depression severity. This

interpretation is supported by previous studies identifying treatment-resistant depression as a high-risk subtype for cardiovascular events.<sup>59,79</sup>

We found that major depressive disorder as measured by the C DIS-IV was not associated with cardiovascular events, whereas depressive symptoms as measured by the PHQ were. It is possible that some participants may have felt more comfortable endorsing depressive symptoms on an anonymous questionnaire than in a face-to-face interview, making the interview a less accurate measure of depression. Assuming the interview was accurate, however, this discrepancy could inform our understanding of the mechanisms of association between depression and cardiovascular disease. An interview diagnosis of major depressive disorder was based on endorsing a minimum of 5 symptoms during the past month.

In contrast, the continuous PHQ-9 score captured both the number and frequency of depressive symptoms experienced in the past 2 weeks. Some patients with a major depressive episode in the past month may not have reported substantial depressive symptoms within the past 2 weeks. Likewise, some patients who did not meet interview criteria for major depressive disorder because they did not have 5 or more symptoms of depressive symptoms more frequently or more recently than patients who did meet interview criteria. This would suggest that the frequency of depressive symptoms may be a stronger risk factor for physical inactivity and cardiovascular events than whether the patient meets criteria for major depressive disorder.

Our study has several strengths, including detailed assessments of depression and cardiac disease severity, careful measurement of potential biological and behavioral mediators, and comprehensive assessment of cardiovascular events. However, several limitations must also be considered. First, most of our participants were older men, and almost half were recruited from VA medical centers. Therefore, our results may not generalize to women or to other patient populations. Second, low physical activity and exercise capacity may have been the result of greater cardiac disease severity. We attempted to address this possibility by carefully measuring and adjusting for comorbid conditions and cardiovascular disease severity. However, no observational study can completely eliminate confounding, and it remains possible that the effect of physical inactivity on cardiovascular outcomes may have been influenced by worse underlying cardiac disease severity that was not otherwise accounted for in our models. Third, we did not assess dietary factors other than blood levels of omega-3 fatty acids. Finally, our study was limited to outpatients with stable coronary heart disease, and thus we cannot comment on the mechanisms of association between depression and cardiovascular outcomes in healthy populations or in patients following acute coronary syndrome.

In summary, we found that the association between depressive symptoms and cardiovascular events was largely explained by health behaviors, especially physical inactivity. These results suggest that the relationship between depression and cardiovascular events may be modifiable with behavioral interventions. The ongoing Understanding Prognostic Benefits of Exercise and Antidepressant Therapy (UPBEAT) study is comparing the effects of exercise vs antidepressive symptoms and coronary heart disease.<sup>80</sup> The longer-term goal is to identify an intervention that will improve both depression and cardiovascular disease outcomes. Regardless of whether physical inactivity causes depressive symptoms or vice-versa, increased activity has the potential to reduce the excess risk of cardiovascular events associated with depressive symptoms in patients with stable coronary heart disease.

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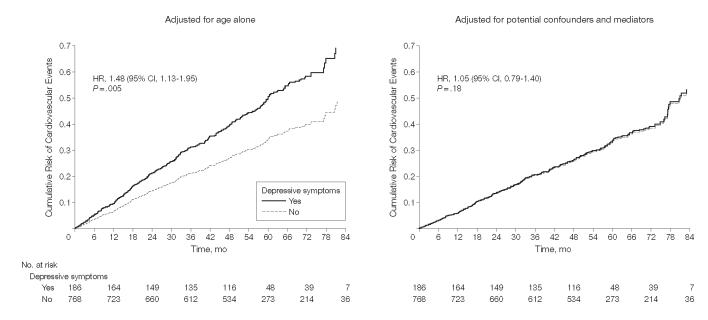
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#### Figure 1. Cumulative Risk of Cardiovascular Events

Data are stratified by depressive symptoms before and after adjustment for health behaviors in 954 participants with complete data. The adjusted hazard ratio (HR) differs slightly from Table 4 because 63 patients with incomplete data were excluded from the analysis. CI indicates confidence interval.

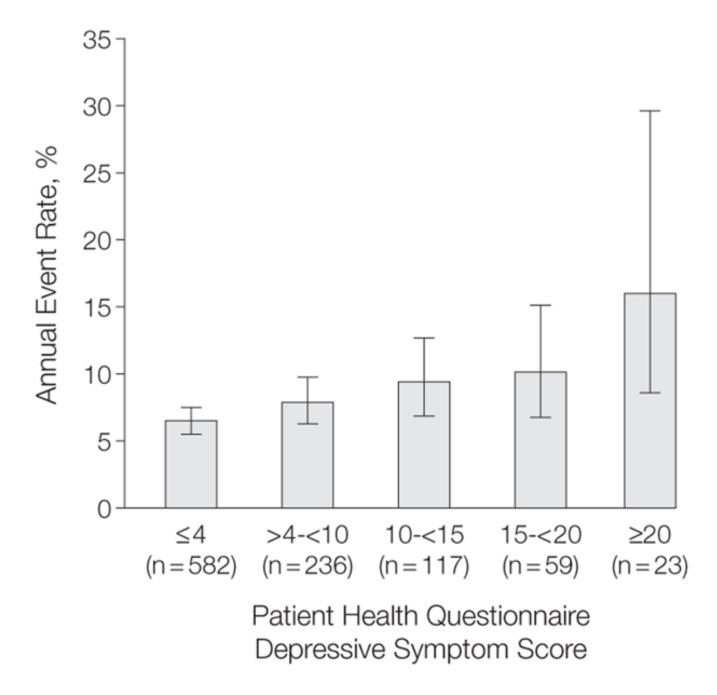


Figure 2. Age-Adjusted Annual Rate of Cardiovascular Events by Depressive Symptom Score in 1017 Participants

Events include heart failure, myocardial infarction, stroke, transient ischemic attack or death.  $\chi^2 P$  for trend<.001. Error bars indicate 95% confidence intervals.

Baseline Characteristics of 1017 Participants With Stable Coronary Heart Disease, by Depressive Symptoms (Patient Health Questionnaire Score of  $\geq 10$  vs  $< 10)^a$ 

Variable	Depressive Symptoms (n = 199)	No Depressive Symptoms (n = 818)	<i>P</i> Value
Demographic characteristics Age, mean (SD), y	63 (12)	68 (10)	<.001
Male sex, No. (%)	152 (76)	682 (83)	.02
White, No. (%)	110 (55)	500 (61)	.13
High school graduate, No. (%)	166 (83)	718 (88)	.08
Body mass index, mean (SD)	29.2 (5.7)	28.2 (5.2)	.02
Comorbid conditions, No. (%) Hypertension	151 (76)	569 (70)	.07
Myocardial infarction	121 (62)	423 (52)	.01
Stroke	33 (17)	115 (14)	.34
Revascularization	109 (55)	492 (60)	.21
Congestive heart failure	49 (25)	130 (16)	.004
Diabetes mellitus	68 (34)	197 (24)	.003
Cardiac disease severity and risk factors Resting left ventricular ejection fraction, mean (SD), %	61 (10)	62 (10)	.06
Diastolic dysfunction, No. (%)	22 (13)	94 (13)	.96
Cholesterol, mean (SD), mg/dL Low-density lipoprotein	105 (36)	104 (33)	.58
High-density lipoprotein	44 (13)	46 (14)	.02
Medication use, No. (%) Aspirin	151 (76)	638 (78)	.52
β-Blocker	119 (60)	472 (58)	.59
Renin-angiotensin system inhibitor	104 (52)	420 (51)	.82
Tricyclic antidepressant	15 (8)	29 (4)	.01
SSRI	38 (19)	59 (7)	<.001
Other antidepressant	35 (18)	44 (5)	<.001
Potential biological mediators, mean (SD) Heart rate variability, SDANN, ms	106 (34)	110 (37)	.29
Heart rate variability, ln VLF, ms <sup>2</sup>	6.28 (0.81)	6.32 (0.88)	.67
Serotonin among non-SSRI users, ng/mL	120 (74)	120 (67)	>.99
Cortisol, µg/d	36.0 (25.3)	39.1 (30.3)	.23
Norepinephrine, µg/d	56 (33)	51 (25)	.02
Log C-reactive protein, mg/L	0.89 (1.33)	0.67 (1.30)	.04
Omega-3 fatty acid levels, % DHA 3 EPA	3.9 (2.2)	4.2 (2.0)	.03
Potential behavioral mediators Regular alcohol use, No. (%)	55 (28)	237 (29)	.69
Smoking status, No. (%) Current	67 (34)	131 (16)	<.001
Past	74 (38)	428 (53)	
Never	54 (28)	252 (31)	
Medication nonadherence	29 (15)	54 (7)	<.001
Self-reported physical activity, No. (%)	59 (30)	130 (16)	<.001

Variable	Depressive Symptoms (n = 199)	No Depressive Symptoms (n = 818)	P Value
Not at all active			
A little active, 1–2 times/mo	55 (28)	127 (15)	
Fairly active, 3-4 times/mo	33 (17)	122 (15)	
Quite active, 1–2 times/wk	13 (7)	141 (17)	
Very active, 3-4 times/wk	24 (12)	193 (24)	
Extremely active, ≥5 times/wk	14 (7)	103 (13)	
Exercise capacity, MET	6.5 (3.2)	7.5 (3.3)	<.001

Abbreviations: DHA, docosahexaenoic acid; EPA, eicosapentaenoic acid; In VLF, natural log of very low frequency; MET, metabolic equivalent task; SDANN, SD of 5-minute mean NN intervals; SSRI, selective serotonin reuptake inhibitor; body mass index is calculated as weight in kilograms divided by height in meters squared.

SI conversion factors: To convert high-density and low-density lipoprotein to mmol/L, multiply by 0.0259.

 $^{a}$ Totals may not sum to their respective totals due to missing data, and percentages may not sum to 100 due to rounding.

# Cardiovascular Events During Follow-up, by Depressive Symptoms

	Events (Age-Adjusted Annual Rate), No. (%)		(ó)	
Event	Depressive Symptoms (n = 199)	No Depressive Symptoms (n = 818)	Age-Adjusted Hazard Ratio (95% Confidence Interval)	P Value
Heart failure	36 (4.5)	103 (2.4)	1.87 (1.27–2.74)	.001
Myocardial infarction	23 (2.9)	80 (1.9)	1.45 (0.91–2.33)	.12
Stroke or transient ischemic attack	14 (1.7)	33 (0.8)	2.11 (1.12–3.99)	.02
All-cause mortality	48 (5.5)	190 (4.1)	1.36 (0.99–1.87)	.06
Any of above outcomes	73 (10.0)	268 (6.7)	1.50 (1.16–1.95)	.002

Change in the Strength of Association Between Depressive Symptoms and Cardiovascular Events (Expressed as the Percent Change in the Age-Adjusted Log Hazard Ratio) After Adjustment for Potential Confounders and Mediators

Covariate	Change in Effect Size After Adjustment, % <sup><i>a</i></sup>
Demographic characteristics Male sex	3.1
White	0.5
High school graduate	-2.5
Body mass index	0.7
Comorbid conditions Hypertension	-1.1
Myocardial infarction	$-7.8^{a}$
Stroke	$-5.4^{a}$
Revascularization	-0.4
Congestive heart failure	$-15.7^{a}$
Diabetes mellitus	$-8.4^{a}$
Cardiac disease severity and risk factors Left ventricular ejection fraction	$-19.0^{a}$
Diastolic dysfunction	-2.7
Low-density lipoprotein cholesterol level	-0.8
High-density lipoprotein cholesterol level	-4.4
Medication use Aspirin	-2.7
β-Blocker	-0.1
Renin-angiotensin system inhibitor	-0.7
Tricyclic antidepressant	-3.1
SSRI	-1.3
Other antidepressant	$-8.8^{a}$
Potential biological mediators Heart rate variability, SDANN	4.5
Heart rate variability, ln VLF	4.1
Serotonin level among nonusers of SSRIs	-1.5
Cortisol excretion	0.5
Norepinephrine excretion	-0.5
Log C-reactive protein level	$-11.3^{a}$
Omega-3 fatty acid levels	-1.5
Potential behavioral mediators Regular alcohol use	-0.5
Smoking	$-10.9^{a}$
Medication nonadherence	$-5.3^{a}$
Self-reported physical activity	$-31.7^{a}$

Abbreviations: In VLF, natural log of very low frequency; SDANN, SD of 5-minute mean NN intervals; SSRI, selective serotonin reuptake inhibitor.

<sup>a</sup>Changes of more than 5% were considered as biologically or clinically important, and (with the exception of other antidepressant use) included in subsequent multivariate regression models.

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**Table 4** Association Between Baseline Depressive Symptoms (Patient Health Questionnaire Score ≥10 vs >10) and Subsequent Cardiovascular Events, With Sequential Adjustment for Potential Confounders and Mediators

		Hazard I	Hazard Ratio (95% Confidence Interval)	(T	
Adjusted Variables <sup>a</sup>	Heart Failure	Myocardial Infarction	Stroke or Transient Ischemic Attack	Death	Any Event
Age per 10-y increase	1.87 (1.27–2.74)	1.45 (0.91–2.33)	2.11 (1.12–3.99)	1.36 (0.99–1.87)	1.50 (1.16–1.95)
History of myocardial infarction, stroke, diabetes, and heart failure	1.63 (1.10–2.42)	1.32 (0.82–2.12)	1.89 (1.00–3.59)	1.23 (0.89–1.70)	1.37 (1.05–1.79)
Left ventricular ejection fraction per 10% increase	1.57 (1.06–2.33)	1.26 (0.78–2.02)	1.80 (0.95–3.42)	1.15 (0.83–1.59)	1.31 (1.00–1.71)
Inflammation, log C-reactive protein per SD increase	1.48 (0.99–2.20)	1.17 (0.72–1.92)	1.55 (0.78–3.07)	1.05 (0.75–1.47)	1.24 (0.94–1.63)
Smoking status	1.46 (0.98–2.17)	1.09 (0.66–1.80)	1.43 (0.71–2.90)	1.01 (0.72–1.42)	1.20 (0.91–1.58)
Medication adherence	1.45 (0.97–2.16)	1.07 (0.65–1.78)	1.31 (0.64–2.67)	1.00 (0.71–1.41)	1.18 (0.89–1.56)
Physical activity	1.18 (0.78–1.80)	0.98 (0.58–1.64)	1.47 (0.70–3.11)	0.87 (0.61–1.23)	1.05 (0.79–1.40)
$\frac{a}{2}$ Each model includes the variables from the precedit	he preceding row so that the	ng row so that the final model includes all the variables listed in this table.	listed in this table.		

Fully Adjusted Model of the Association Between Baseline Depressive Symptoms and Subsequent Cardiovascular Events (Heart Failure, Myocardial Infarction, Stroke, Transient Ischemic Attack, or Death) in 954 Participants With Complete Data on Covariates

Variable	HR (95%Confidence Interval)	P Value
Depressive symptoms <sup>a</sup>	1.05 (0.79–1.40)	.75
Age per 10-y increase	1.72 (1.52–1.94)	<.001
Medical history Myocardial infarction	1.06 (0.83–1.34)	.64
Diabetes	1.76 (1.38–2.25)	<.001
Stroke	1.31 (0.98–1.75)	.07
Heart failure	1.27 (0.96–1.67)	.09
Left ventricular ejection fraction per 10% decrease	1.40 (1.26–1.56)	<.001
Log C-reactive protein per SD increase	1.27 (1.13–1.43)	<.001
Current smoking	1.39 (1.17–1.66)	<.001
Medication nonadherence	1.90 (1.31–2.76)	<.001
Physical inactivity <sup>b</sup>	1.44 (1.14–1.82)	.002

 $^{a}$ A score of 10 or higher on the Patient Health Questionnaire.

 $b_{\rm Not}$  at all or a little vs fairly, quite, very, or extremely active.