A Testable Prognostic Model of Nicotine Dependence

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Abstract

Individuals' dependence on nicotine, primarily through cigarette smoking, is a major source of morbidity and mortality worldwide. Many smokers attempt but fail to quit smoking, motivating researchers to identify the origins of this dependence. Because of the known heritability of nicotine-dependence phenotypes, considerable interest has been focused on discovering the genetic factors underpinning the trait. This goal, however, is not easily attained: no single factor is likely to explain any great proportion of dependence because nicotine dependence is thought to be a complex trait (i.e., the result of many interacting factors). Genomewide association studies are powerful tools in the search for the genomic bases of complex traits, and in this context, novel candidate genes have been identified through single nucleotide polymorphism (SNP) association analyses. Beyond association, however, genetic data can be used to generate predictive models of nicotine dependence. As expected in the context of a complex trait, individual SNPs fail to accurately predict nicotine dependence, demanding the use of multivariate models. Standard approaches, such as logistic regression, are unable to consider large numbers of SNPs given existing sample sizes. However, using Bayesian networks, one can overcome these limitations to generate a multivariate predictive model, which has markedly enhanced predictive accuracy on fitted values relative to that of individual SNPs. This approach, combined with the data being generated by genomewide association studies, promises to shed new light on the common, complex trait nicotine dependence.

Keywords

addiction; nicotine; genetics; prediction

Cigarette smoking is the most prevalent form of tobacco use and is a major contributor to worldwide morbidity and mortality (DHHS, 2004), including being the single largest preventable cause of lung cancer (Murray, 2006). Nevertheless, 21.6% of adult Americans smoke (Centers for Disease Control and Prevention, 2005). Many want to stop smoking, as evidenced by the finding that 42.5% of American smokers reported having quit for at least 1 day in the previous year (Centers for Disease Control and Prevention, 2006). Most of these attempts fail: A meta-analysis of clinical trials demonstrated that an average of only 10% of individuals achieve abstinence for 6 months or more (Hughes et al., 2007). Nicotine dependence contributes to the failure rate, with more highly dependent smokers having greater difficulty in attaining abstinence (John et al., 2004; Xian et al., 2007). It has been established that nicotine dependence and associated characteristics are highly heritable phenotypes (Agrawal et al., 2006; Broms et al., 2007; Broms et al., 2006; Haberstick et al., 2007; Lessov et al., 2004; Maes et al., 2004; Pergadia et al., 2006; True et al., 1999; Vink et al., 2005; Xian et al., 2003, 2005). Despite its clear heritability, the search for the genetic variations underlying this trait is challenging. There have been issues with consistent replication of findings of association (Lerman & Swan, 2002; Vandenbergh et al., 2002), and no individual single-nucleotide polymorphism (SNP) is a sound predictor of nicotine dependence. This paper will address the challenges to identifying accurate genomic prognostic models of nicotine dependence and will review Bayesian networks, a novel proposed solution.

NICOTINE DEPENDENCE AS A COMPLEX PHENOTYPE

Despite the established heritability of nicotine-dependence phenotypes (Lessov-Schlaggar et al., 2008; M. R. Munafo & Johnstone, 2008), identifying replicable underlying genetic variations has proven difficult. Genomewide association studies (GWASs) hold promise and, indeed, novel associations have already been identified from using this approach. However, even the most strongly associated SNPs identified in these studies fail to *predict* nicotine dependence. This is likely due to the fact that nicotine dependence is a complex and potentially multidimensional trait, that is, a result of complicated interactions of multiple genetic and environmental factors (Cardon & Bell, 2001; M. R. Munafo & Johnstone, 2008; Rice et al., 2001; Risch, 2000).

Defining Nicotine Dependence

Careful phenotype definition is essential to achieving the phenotypic clarity required to identify valid, meaningful genetic predictors of complex human disorders (Rice et al., 2001). Several measures exist for defining nicotine dependence, the predominant measures being (1) diagnostic nicotine dependence as defined by the *Diagnostic and Statistical Manual for Mental Disorders*, 3rd edition Revised and 4th edition (DSM-III-R and DSM-IV) (APA, 1987, 1994) and (2) the Fagerström Tolerance Questionnaire (FTQ) and the Fagerström Test for Nicotine Dependence (FTND). Because of its widespread use (Hatsukami et al., 2008) and pertinence to the analyses presented, this report will focus on the FTND; in-depth descriptions of the other instrument can be found in the review articles (Piper et al., 2006; Lessov-Schlaggar et al., 2008).

The FTND is a revision of the FTQ, designed to improve the psychometric properties of the scale (Heatherton et al., 1991; Payne et al., 1994). The FTND (Figure 1) is a six-question instrument, with a minimum score of 0 and a maximum score of 10. This scale focuses primarily on physiological dependence (Bierut et al., 2007), as opposed to other behavioral and psychological dimensions of dependence, for example, the subjective view of nicotine use being a "problem" in some way (Lessov-Schlaggar et al., 2008).

Heritability of Nicotine-Dependence Phenotypes

Studies of twins have shown that a substantial proportion of the phenotypic variance in nicotine dependence is heritable. Indeed, studies of the FTND score, in particular, have shown it to be markedly heritable, with estimates ranging from 40 to 75% (Broms et al., 2006; Maes et al., 2004; Vink et al., 2005). Further, component measures of the FTND are substantially heritable. The daily cigarette quantity and time to first cigarette have been estimated to be 45.0–70.0% heritable (Broms et al., 2006; Carmelli et al., 1990; Haberstick et al., 2007; Hettema et al., 1999; Lessov et al., 2004; Pergadia et al., 2006; Prescott & Kendler, 1995; Swan et al., 1996, 1997; Swan et al., 1990) and 55.0–68.0% heritable (Haberstick et al., 2007; Lessov et al., 2004), respectively. Although the nicotine-dependence phenotype has a clear genetic component, it has been difficult to identify the causative variations, or even consistently associated variations, because of the complex nature of the phenotype (Lerman & Swan, 2002; M. R. Munafo et al., 2004; M. R. Munafo & Johnstone, 2008; Vandenbergh et al., 2002).

Genomewide Association Studies of Nicotine Dependence

It is thought that complex traits are influenced by multiple loci across the genome (Cardon & Bell, 2001; Risch, 2000). Thus, to identify the set of variations that contribute to vulnerability to nicotine dependence, GWASs have been conducted.

Bierut and colleagues first performed a pooled analysis of smokers who were nicotine dependent and smokers who were not nicotine dependent at over 2.4 million SNPs and, on this basis, selected 31,960 SNPs for genotyping in individuals (Bierut et al., 2007). Both cases and controls reported having smoked at least 100 cigarettes in their lifetime. Cases were individuals who had a FTND score of 4 or more when smoking the most, which is a commonly employed definition of nicotine dependence. Controls' status was defined as having a lifetime FTND of 0. Although none of the individual findings was statistically significant after correcting for multiple tests, they identified 35 SNPs with P-values less than 10^{-4}. Some of these SNPs were in intergenic regions and the β3 nicotinic receptor (*CHRNB3*), a known candidate gene, but many SNPs were in genes not previously associated with nicotine dependence. The novel genes have been associated with the following biological processes: protein transport (*CLCA1* and *VPS13A*), protein catabolism (*RNF5*, *FBXL17*), regulation of transcription (*PBX2*), signal transduction (*GPSM3*), ion transport (*TRPC7*), lipid metabolism (*FTO*, *AGPAT1*), cell adhesion (*NRXN1*, *AGER*, *CTNNA3*), and developmental processes (*NOTCH4*) (Lessov-Schlaggar et al., 2008). In the same case-control sample, a targeted analysis of 348 known candidate genes was conducted (S. F. Saccone et al., 2007). The authors identified 39 SNPs with the greatest evidence for association with nicotine dependence, based upon the false discovery rate. One of these SNPs (*rs6474413*, *CHRNB3*) was shared with the 35 reported by the Bierut study. The majority of these SNPs were in cholinergic nicotinic receptor genes, including *CHRNB3* and *CHRNA5*, the α5 nicotinic receptor subunit gene.

Uhl and colleagues (Uhl et al., 2007) conducted a case-control study at 520,000 SNPs in pools of DNA, with a special interest in SNPs in genes that have overlap with dependence on other substances. To identify variants associated with nicotine dependence, cases were European-American active smokers (mean FTND = 6.4, average carbon monoxide = 34.7) recruited into a smoking cessation study, and controls were European-American people with no substantial lifetime histories of use of any addictive substance. The authors identified SNPs in 32 genes that were both: (1) significant at $P<0.0005$ between cases and controls in their study and (2) had been significantly associated with polysubstance abuse, alcohol dependence, and methamphetamine dependence in previous studies. Like the study by Bierut and colleagues (Bierut et al., 2007), this investigation identified genes involved in cell adhesion, signal

transduction, and transport functions, revealing convergence in the findings of these studies, despite the difference in their study populations and methods (Lessov-Schlaggar et al., 2008).

More recently, the associations reported by (Saccone et al., 2007) between nicotine dependence and variants in the CHRNA5-CHRNA3-CHRNB4 gene cluster have been replicated in several additional samples studying smoking quantity and nicotine dependence (Berrettini et al., 2008; Bierut et al., 2008; Thorgeirsson et al., 2008). Further, the nonsynonymous *CHRNA5* SNP, rs16969968, or its r^2 proxies (SNPs highly correlated with it) have now been associated with lung cancer (Amos et al., 2008; Hung et al., 2008; Thorgeirsson et al., 2008). These findings provide strong confirmation of the association between these SNPs and nicotine dependence. However, the challenge of moving beyond individual SNP associations to identify predictive models of phenotype remains.

Association with versus Prediction of Nicotine Dependence

The genomewide data that have been and will be collected hold the promise for uncovering the genetic origins of nicotine dependence, and the association analyses of individual SNPs have revealed new possible suspects. In addition, these analyses help to define overall risk classes, for example, a man with the G allele at SNP rs2791480 in *CLCA1* may be at higher risk for being nicotine dependent (Bierut et al., 2007). Beyond association, the genomewide data can be used to identify predictors of complex traits such as nicotine dependence. Predictive modeling is a logical complement to the association-based approach, because predictive measures are both amenable to translation into clinical practice and are able to address some of the challenges of the analysis and interpretation of genomewide data. In particular, measures of predictive accuracy can address the issues of replication and generalizability of genetic association studies (Heidema et al., 2006; M. R. Munafo et al., 2004; M. R. Munafo & Johnstone, 2008). Even when the basis for the predictive measure is a multivariate model, the predictive accuracy can be used as a single, comprehensive measure of the extent of the generalizability of the model among different studies. Further, measuring the predictive accuracy is directly responsive to concerns that genetic tests will have a poor disease-predictive ability (Holtzman, 1992; Khoury et al., 1985). Indeed, couching results in the language of prediction poises them for translation into clinical practice, where they can be used in risk communication and counseling.

Predictive Accuracy of Individual SNPs and Demographic Factors

Because of the utility of predictive measures, the predictive accuracies of the 73 SNPs identified by Bierut (Bierut et al., 2007) and Saccone (S. F. Saccone et al., 2007), as well as the demographic factors, age and gender, were calculated. The predictive accuracy of the single-SNP model was computed as the area under the receiver-operator characteristic curve (AUROC) (DeLong et al., 1988) of the prediction of the fitted values. The probability of nicotine dependence, given the genotype of an individual subject, was calculated by using the clique algorithm (S. L. Lauritzen, & Spiegelhalter, 1988).

As shown in AUROC: area under the receiver operator characteristic curve, the highest predictive accuracy demonstrated by an individual SNP was 54.4% ($P = 0.002$), while gender and quartiles of age achieved predictive accuracies of 54.5 ($P = 0.001$) and 54.7% ($P = 0.001$), respectively. In these and the following cases, the *P*-value tests the hypothesis that the AUROC is significantly better than random (i.e., AUROC = 50%). The highest AUROC attained by an SNP in *CHRNB3*, a known candidate gene that was identified as highly associated in both the Bierut (Bierut et al., 2007) and Saccone (S. F. Saccone et al., 2007) papers, was 51.9% ($P = 0.093$), a performance not significantly better than random assignment. Because of the complex nature of nicotine dependence, the low predictive accuracies of individuals SNPs are perhaps not surprising, given that the effect of any single SNP on the trait is likely to be very small

(Lerman & Swan, 2002). This issue can be addressed only by considering the polygenic nature of nicotine dependence (Lerman & Swan, 2002). Additional benefits accrue via the multivariate approach, including responsiveness to the issues of confounding (Cardon & Bell, 2001), linkage disequilibrium (Cardon & Bell, 2001), and genetic heterogeneity (Heidema et al., 2006). It has been noted, however, that a significant obstacle to accomplishing this goal is the large sample sizes required by standard multivariate methods such as logistic regression (Heidema et al., 2006; Lerman & Swan, 2002; M. R. Munafo & Johnstone, 2008). There has, therefore, been an interest in exploring alternative analytical methods that are more efficient. One such approach with a demonstrated track record in the analysis of complex phenotypes is the Bayesian network.

USING BAYESIAN NETWORKS TO BUILD PREDICTIVE MODELS OF COMPLEX PHENOTYPES

Bayesian networks are multivariate dependency models that account for simultaneous associations and interactions among multiple SNPs. A Bayesian network is a directed acyclic graph in which nodes represent random variables and arcs (arrows) define directed stochastic dependencies quantified by probability distributions. When an arc connects two nodes, the node at the point of the arrow is called a "child" of the "parent" node at the end of the arrow. Figure 2 depicts two Bayesian networks, starting with a simple network describing the dependency of a phenotype P on a single SNP G (Figure 2A). The graph decomposes the joint probability distribution of the two variables, P and G, into the product of the marginal distribution of G (the parent node) and the conditional distribution of P (the child node) given G. The marginal and conditional probability distributions are sufficient to define the association between P and G because their product determines the joint probability distribution. The property persists when we invert the direction of the arc in the graph, and when we expand the graphical structure to include several variables (Figure 2B): the overall association is measured by the joint probability distribution that is still defined by the product of each child-parent conditional distribution. This modular nature of a Bayesian network is due to the conditional independences among the variables encoded by the directed acyclic graph (Cowell et al., 1999): each node is independent of its predecessors given its parent nodes. This modular representation captures complex dependency models that are able to integrate associations between SNPs and phenotype; associations between SNPs due to linkage disequilibrium or evolutionary patterns (Chakravarti, 1999); and interaction processes linking SNPs, phenotypes, and modulating factors (Hoh & Ott, 2003) with a small number of parameters. Reducing the number of parameters allows us to "learn" large dependency networks from comparatively small datasets, and well-established techniques exist to develop Bayesian networks from data in an almost automated manner (Cowell et al., 1999).

The Bayesian network approach has been used to analyze several types of genomic data, including gene expression (Friedman, 2004; Friedman et al., 2000; Sebastiani et al., 2004), protein-protein interactions (Jansen et al., 2003), and pedigree analysis (S. L. Lauritzen & Sheehan, 2004). In the context of SNP data, Bayesian networks have been used to model the multigenic risk of stroke in sickle cell anemia (Sebastiani et al., 2005). While sickle cell anemia is a paradigmatic monogenic disease, the occurrence of stroke in individuals with the disorder is, like nicotine dependence, complex. This study demonstrated that Bayesian networks can generate a model of a complex trait that has a high predictive accuracy. Indeed, the stroke in the sickle cell anemia model demonstrated a predictive accuracy of 98.2% in an independent population.

APPLICATION OF BAYESIAN NETWORKS TO NICOTINE-DEPENDENCE PREDICTION

Bayesian networks are able to generate SNP-based models of complex traits by using relatively small sample sizes. Combining this method with the existing large-scale genotyping data affords a powerful, prognostic vantage on nicotine dependence.

Study Population

This study was based upon the genetic data and FTND phenotype definitions of the Collaborative Genetic Study of Nicotine Dependence (COGEND) (Bierut et al., 2007; S. F. Saccone et al., 2007). The 73 SNPs reported as associated to nicotine dependence in this population were subjected to Bayesian network analysis.

Statistical Analysis

All of the SNPs considered in the Bayesian network analysis were tested for deviation from the Hardy-Weinberg equilibrium (as described in Bierut et al., 2007; S. F. Saccone et al., 2007). The analysis was conducted by using Bayesware Discoverer (Bayesware, Boston, Massachusets, USA). The dependencies in the model of *CASE* (nicotine dependent) status were identified by calculating the Bayes Factors corresponding to the ratio of the marginal log likelihood of the statusdetermining genotype to the marginal log likelihood of status being independent of genotype. In the absence of an independent validation set, we assessed the goodness-of-fit of the model by calculating the AUROC of the prediction of the fitted values. The probability of nicotine dependence given the genotype of an individual subject was calculated by using the clique algorithm implemented in Bayesware Discoverer; the predictive accuracy of the model was estimated as the AUROC curve convex hull, using the trapezoidal rule (DeLong et al., 1988).

Predictive Accuracy of the Multivariate Model

The multivariate model generated by the Bayesian network analysis is shown in Figure 3. Out of the original 73 SNPs, 60 were incorporated into the model, along with the two demographic factors (*SEX* and *AGE*). However, to predict nicotine-dependence status, it is not always necessary to have information about all 60 SNPs. If one has information on all of the nodes that are children of *CASE*, as well as any other parents of these nodes, one does not need any information about any other nodes. This subset of nodes is called the Markov blanket of *CASE*. In this network shown in Figure 3, there are 20 nodes in the Markov blanket: 18 SNPs and both demographic factors. The SNPs directly connected to *CASE* are in genes *CHRNA3, CHRNB3, CLCA1, CLTCL1, CTNNA3, FBXL17, FTO, GABRA4, KCNJ6, NRXN1, OPRM1*, and *VPS13A*, with the remaining SNPs being in intergenic regions. Individually, the best-performing of these SNPs, rs2836823, had a predictive accuracy on fitted values of 54.4% ($P = 0.002$; Table 1). While this is statistically significantly different from randomness, such a low AUROC indicates a lack of predictive ability. By contrast, the network as a whole (Figure 4) was able to achieve a predictive accuracy on fitted values of 75.0% ($P<0.0001$), as shown in Figure 4.

An interesting result revealed by the topology of the network in Figure 3 is the role of the SNP, rs16969968 (labeled CHRNA5_2 in the network), which is connected to the phenotype through CHRNA5 (rs578776 labeled as CHRNA5_1). This finding is of particular note given that several recent studies have shown the associations of SNPs in the CHRNA5-CHRNA3-CHRNB4 cluster to both nicotine dependence and lung cancer (Amos et al., 2008;Berrettini et al., 2008;Bierut et al., 2008;Hung et al., 2008;Thorgeirsson et al., 2008). Further, an in-depth study in the COGEND population of the complete family of 16 nicotinic receptor subunits

identified both rs578776 and rs16969968 as statistically significantly associated with nicotine dependence even though the correlation between these SNPs is low (N. L. Saccone et al., in press).

SUMMARY AND FUTURE DIRECTIONS

Nicotine dependence has been shown to have a heritable component, and therefore, large-scale genetic studies hold tremendous promise for revealing the genetic bases for nicotine dependence. Single-SNP association analyses have yielded novel findings that shed light on the biological underpinnings of the trait. However, none of even the most highly associated SNPs alone has been shown to *predict* nicotine dependence better than at random, which is a result that might be anticipated given the complex nature of the trait. While standard analytical approaches, such as logistic regression, require prohibitively large sample sizes to consider multiple interacting SNPs, the Bayesian network approach has been shown to be well-suited to this type of analysis in the context of existing sample sizes. Indeed, when applied to nicotine dependence, the multivariate Bayesian network analysis demonstrated markedly enhanced prediction on fitted values relative to individual SNPs. This work is an important first step in discovering predictors of nicotine dependence. An accurate predictive model would give the research community a new vantage point from which to consider the causes of this trait, and such a model would have the further advantage of being framed in the language of diagnostic accuracy and risk communication.

The next steps toward the goal of generating a valid prognostic model of nicotine dependence should be to seek to improve the predictive accuracy by building a model drawing upon all SNPs in the genomewide analysis and incorporating environmental exposures. Further, the models generated should be validated in independent populations, rather than upon fitted values. In preference of a retrospective case-control design, the ideal study for such a validation would be prospective to allow for true *prediction* of nicotine dependence status, as opposed to correct classification. The distinction between prediction and correct classification is smaller in the context of SNPs, because they are fixed over a person's entire lifetime, but is of great importance when modeling the effect of mutable environmental exposures. Moreover, such an approach would allow for the prospective determination of phenotype.

In addition to identifying predictors of nicotine dependence, developing predictive models of the trait should help refine the phenotype definitions of "being a smoker" and nicotine dependence themselves. Proper phenotype definition is fundamental to any search for the origins of complex traits, and phenotyping has been identified as an issue in need of additional attention and clarification by a number of researchers (Lerman & Swan, 2002; Lessov et al., 2004; M. R. Munafo & Johnstone, 2008; Robert, 2006). The nature of the disorder will become clearer as the molecular bases of the trait subtypes become better understood. Along these lines, Cardon and Bell have envisioned an iterative process of developing an understanding of the genetic bases of complex traits, in which weak associations are successively improved upon by generating hypotheses on the basis of subgroups and assessing them in new cohorts (Cardon & Bell, 2001). Through these efforts, the nicotine-dependence research community will develop a deeper knowledge of the causes of nicotine-dependence phenotypes. This knowledge promises to define etiology-driven definitions of disease (Piper et al., 2006) and to reveal new therapeutic avenues.

Finally, the ability to develop a prognostic model of the complex trait nicotine dependence highlights the feasibility of identifying predictors of another complex trait: response to smoking cessation therapy. Currently, the untargeted assignment of existing therapies suffers from disappointing relapse rates. Through pharmacogenetic targeting, success rates could be enhanced by assigning people who want to quit smoking to the therapy that is most likely to

work for them. A small number of studies have highlighted the promise of the approach, but as with nicotine dependence, a lack of replication has plagued these investigations (M. Munafo & Lerman, 2006). An efficient multivariate solution, such as the Bayesian network approach, may help address the challenge of finding translatable pharmacogenetic solutions for the treatment of nicotine dependence.

ACKNOWLEDGEMENTS

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	Points
How soon after you wake up do you smoke your first cigarette?	
Within 5 minutes	3
6–30 minutes	2
31–60 minutes	1
After 60 minutes	0
Do you find it difficult to refrain from smoking in place where it is forbidden, e.g., in church, at the library, in cinema, etc.?	
Yes	1
No	0
Which cigarette would you hate most to give up?	
The first one in the morning	1
All others	0
How many cigarettes per day do you smoke?	
≤10	0
11–20	1
21–30	2
≥31	3
Do you smoke more frequently during the first hours after waking than you do during the rest of the day?	
Yes	1
No	0
Do you smoke if you are so ill that you are in bed most of the day?	
Yes	1
No	0

Figure 1.
The Fagerström Test for Nicotine Dependence.

a.

G		
AA	Aa	aa
0.6	0.3	0.1

	P	
G	+	−
AA	0.7	0.3
Aa	0.5	0.5
aa	0.1	0.9

b.
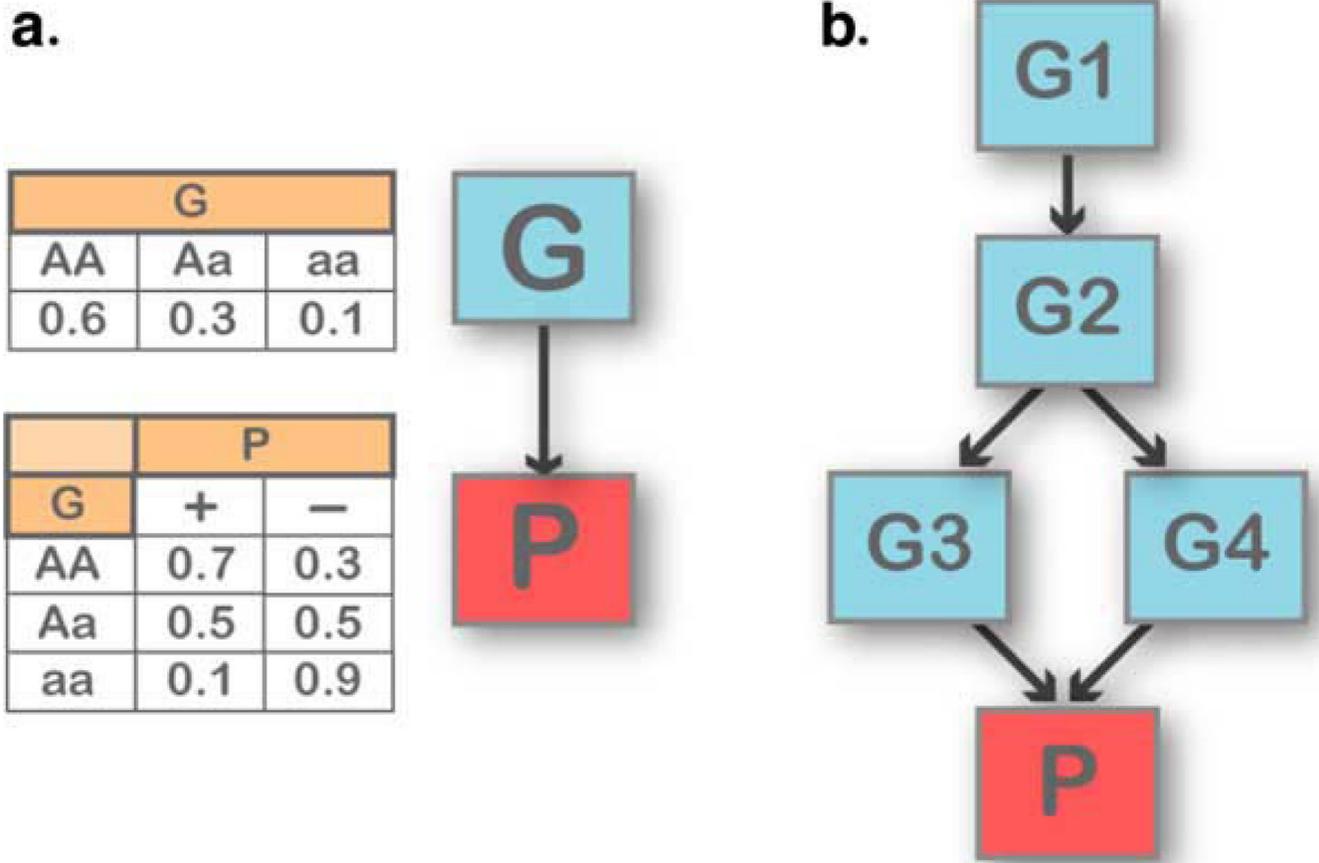
Figure 2.
Examples of Bayesian network structures. (A) A simple Bayesian network with two nodes, representing a SNP (G) and a phenotype (P). The probability distribution of G represents the genotype distribution in the population, and the conditional probability distribution of P describes the distribution of each phenotype given each genotype. The direction of the association between G and P can be reversed by using Bayes theorem. (B) A Bayesian network linking four SNPs (G1–G4) to a phenotype P. The phenotype is independent of G1 and G2 given G3 and G4. The joint probability distribution of the network is fully specified by the five distributions representing the distribution of G1 (two parameters), of G2 given G1 (six parameters), of G3 given G2 (six parameters), of G4 given G2 (six parameters), and of P given G3 and G4 (nine parameters). The full probability distribution requires $81 \times 2 - 1 = 161$ parameters, while this network requires only 29.

Figure 3.
Bayesian network of nicotine dependence. The Bayesian network describing the joint association of 60 SNPs (blue) and two demographic factors (pink) with nicotine dependence. CASE represents nicotine-dependence status. The RS numbers corresponding to the codes in the nodes are shown in Table 1.
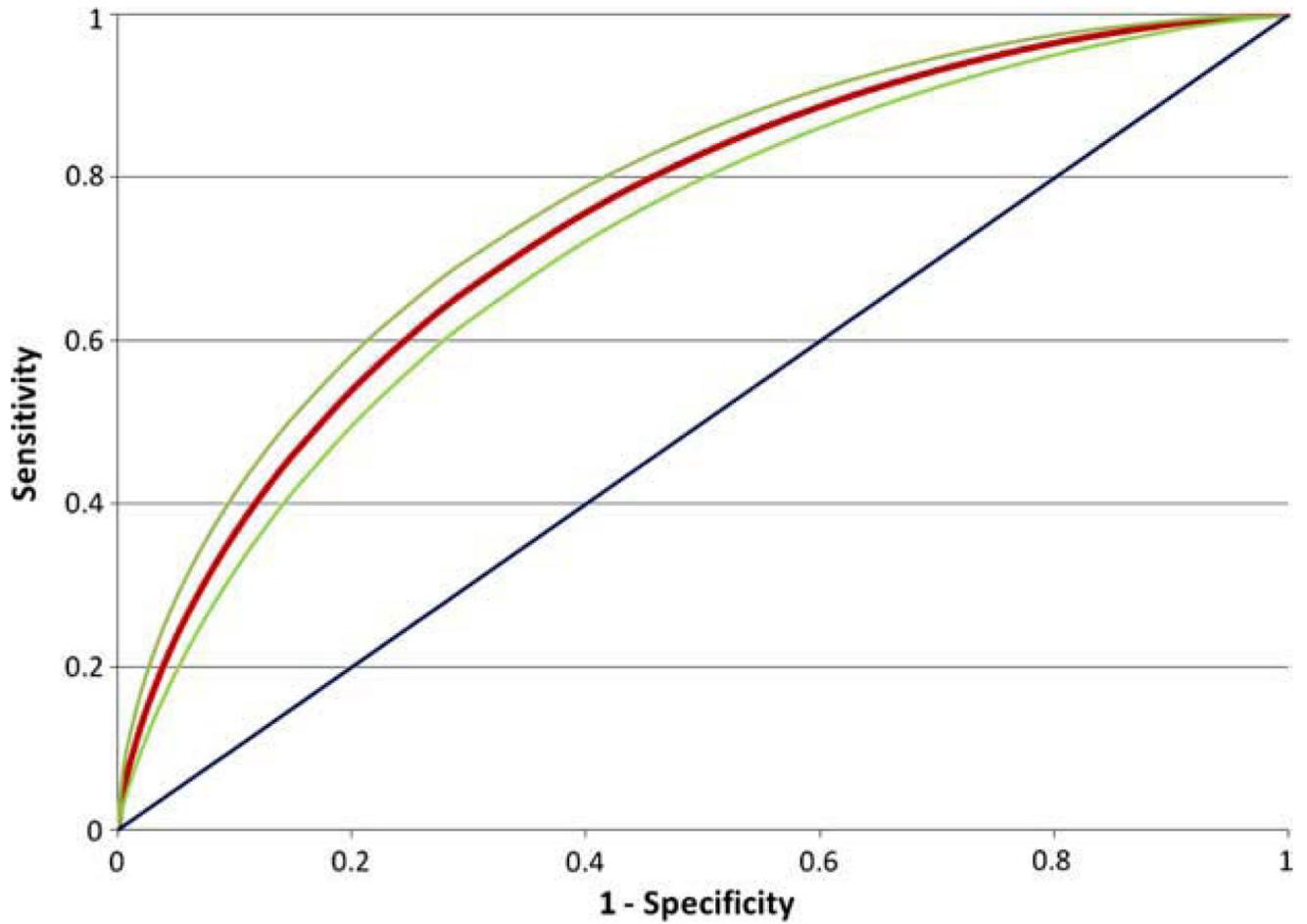
Figure 4.
ROC curve representing the predictive accuracy of the Bayesian network in Figure 3 on the fitted values. The curve emanating from the origin has an area under the ROC curve (AURCO) of 50%, which represents random classification. The red ROC curve corresponds to the Bayesian network (AUROC = 75.0%), the 95% confidence interval of which is bounded by the green curves.

Table 1

Predictive Accuracies (AUROC) of Individual SNPs and Demographic Factors on the Fitted Values.

Factor	AUROC	p-value
Age[*]	54.7%	0.001
Gender[*]	54.5%	0.001

Gene	Label	SNP	AUROC	p-value
AVPR1A	AVPR1A_1	3021529	47.10%	0.98
CHRNA5/3	**CHRNA5_1**[*]	578776	54.00%	0.01
CHRNA5	CHRNA5_2[*]	16969968	50.30%	0.40
	CHRNA5_3[*]	1051730	49.60%	0.62
	CHRNA5_4[*]	637137	48.90%	0.79
	CHRNA5_5[*]	684513	49.20%	0.72
	CHRNA5_6[*]	3743078	48.40%	0.89
CHRNA6	CHRNA6_1[*]	2304297	47.10%	0.98
CHRNB3	**CHRNB3_1**[*]	13277254	51.90%	0.09
	CHRNB3_2[*]	6474413	48.10%	0.91
	CHRNB3_3[*]	4953	48.80%	0.82
	CHRNB3_5[*]	4952	47.90%	0.93
CHRNB4	CHRNB4_1[*]	3813567	47.70%	0.95
CHRND	CHRNG_1	2767	47.00%	0.98
	CHRNG_2	3791729	47.60%	0.96
CLCA1	**CLCA1_1**[*]	2791480	54.20%	0.00
CLTCL1	**CLTCL1_1**[*]	1206549	52.20%	0.06
CTNNA3	**CTNNA3_1**[*]	4142041	53.10%	0.02
	CTNNA3_2[*]	9332406	49.00%	0.75
CYP2B6	CYP2B6_1[*]	4802100	47.90%	0.93
	CYP2B6_2[*]	3760657	48.50%	0.86
DAO	DAO_1	17041074	47.10%	0.98
DBH	DBH_1[*]	4531	49.30%	0.68
	DBH_2[*]	3025382	47.10%	0.98
DOCK3(GRM2)	DOCK3_1[*]	6772197	44.80%	1.00
CHRNG	EIF4E2_1	6749955	48.80%	0.82
FBXL17	**FBXL17_1**[*]	10793832	52.50%	0.04
FMO1	FMO1_1[*]	7877	48.30%	0.89
	FMO1_2[*]	742350	47.10%	0.98
	FMO1_3[*]	7517376	47.20%	0.98
	FMO4_1[*]	16864387	48.10%	0.92
FTO	**FTO_1**[*]	2302673	52.60%	0.03

Gene	Label	SNP	AUROC	p-value
GABRA4	**GABRA4_1**[*]	3762611	50.70%	0.30
	GABRA4_2	3762607	48.30%	0.89
GPSM3,	AGER_1[*]	999	47.50%	0.97
AGPAT1,				
NOTCH4,				
RNF5,				
AGER,				
PBX2,				
AGER				
HTR5A	HTR5A_1	6320	47.90%	0.93
	HTR5A_2	1657273	48.60%	0.85
KCNJ6	**KCNJ6_1**[*]	6517442	54.10%	0.01
NRXN1	**NRXN1_1**[*]	12623467	51.20%	0.20
	NRXN1_2[*]	10490162	47.60%	0.95
	NRXN1_3[*]	12467557	47.10%	0.98
OPRM1	**OPRM1_1**[*]	510769	53.60%	0.01
PDYN	PDYN_1	6045733	48.50%	0.86
PIP5K2A	PIP5K2A_1[*]	10508649	46.00%	1.00
TRPC7	TRPC7_1	2673931	47.70%	0.95
	TRPC7_2	2546657	47.10%	0.98
VPS13A	**VPS13A_1**[*]	12380218	53.10%	0.01
	VPS13A_2[*]	2022443	47.00%	0.98
	VPS13A_3[*]	11145381	48.30%	0.88
Intergenic		10049135	48.00%	0.93
		1031006[*]	52.20%	0.06
		10958726[*]	47.40%	0.97
		11157219[*]	46.20%	1.00
		11694463[*]	51.70%	0.12
		1612945[*]	48.50%	0.86
		17602038[*]	47.40%	0.97
		17633211[*]	47.40%	0.97
		17633258[*]	45.80%	1.00
		17706299[*]	46.90%	0.99
		17706334[*]	49.50%	0.65
		1782134[*]	47.10%	0.98
		1782141[*]	47.00%	0.98
		1782144[*]	47.90%	0.93
		1782145[*]	47.90%	0.93
		1782159[*]	52.80%	0.03
		1782182[*]	47.10%	0.98

Gene	Label	SNP	AUROC	p-value
		2276560	45.90%	1.00
		2798983[*]	48.60%	0.84
		2836823[*]	54.40%	0.00
		4142603[*]	48.20%	0.90
		4245150[*]	48.70%	0.83
		4859365[*]	53.60%	0.01

AUROC: area under the receiver operator characteristic curve

Label: label used in Figure 3

[*] included in the model shown in Figure 3

bold: in the Markov blanket of CASE in the network shown in Figure 3