METHODOLOGY

Non-response to baseline, non-response to follow-up and mortality in the Whitehall II cohort

Jane E Ferrie,¹* Mika Kivimäki,^{1,2} Archana Singh-Manoux,^{1,3} Alison Shortt,¹ Pekka Martikainen,^{1,4} Jenny Head,¹ Michael Marmot,¹ David Gimeno,^{1,5} Roberto De Vogli,¹ Marko Elovainio² and Martin J Shipley¹

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Background	Little is known about the associations between non-response to follow-up surveys and mortality, or differences in these associations by socioeconomic position in studies with repeat data collections.
Methods	The Whitehall II study of socioeconomic inequalities in health provided response status from five data collection surveys; Phase 1 (1985–88, $n = 10308$), Phase 5 (1997–99, $n = 6533$), and all-cause mortality to 2006. Odd-numbered phases included a medical examination in addition to a questionnaire.
Results	Non-response to baseline and to follow-up phases that included a medical examination was associated with a doubling of the mortal- ity hazard in analyses adjusted for age and sex. Compared with complete responders, responders who missed one or more phases, but completed the last possible phase before they died, had a 38% excess risk of mortality. However, those who missed one or more phases including the last possible phase before death had an excess risk of 127%. There was no evidence that these associations differed by socioeconomic position.
Conclusion	In studies with repeat data collections, non-response to follow-up is associated with the same doubling of the mortality risk as non- response to baseline; an association that is not modified by socio- economic position.
Keywords	Non-response to follow-up, partial response, socioeconomic inequal- ities, all-cause mortality, occupational cohort, white-collar

E-mail: j.ferrie@public-health.ucl.ac.uk

Most observational epidemiology studies rely on participation of the target population to obtain data on risk exposures and health outcomes. If non-response at baseline is non-random, findings from such studies may be biased. In prospective studies with repeat data collections, the problem of bias introduced by nonresponse to the baseline survey may be compounded by non-response to follow-up surveys.

Many studies have compared mortality risks between responders and non-responders to baseline, with most,^{1–14} but not all,¹⁵ providing strong evidence of higher rates of mortality among non-responders.

¹ Department of Epidemiology and Public Health, University College London, London, UK.

² Finnish Institute of Occupational Health, Helsinki, Finland.

³ INSERM, U687-IFR69, France.

⁴ Helsinki Collegium for Advanced Studies, University of Helsinki, Finland.

⁵ University of Texas School of Public Health, Health Sciences Center, Houston, San Antonio Regional Campus Texas, USA.

^{*} Corresponding author. Department of Epidemiology and Public Health, University College London, 1-19 Torrington Place, London WC1E 6BT, UK.

However, it is considered that observational epidemiologists have paid insufficient attention to response bias in cohort studies.¹⁶ Unresolved issues include associations between partial response or non-response to follow-up and mortality in studies with successive data collections, and differences in these associations by socioeconomic position. Socioeconomic position is important for two reasons: one, as a confounder in aetiological studies, and two, as an exposure in studies of social inequalities in health.¹⁷ If the association between non-response and mortality were to differ across the socioeconomic strata residual confounding might remain a problem for aetiological studies, even after adjustment for socioeconomic position, and analyses of socioeconomic inequalities might be biased.

This article uses data from the Whitehall II study, which was set up to explain socioeconomic gradients in health.^{18,19} The article has two main aims: (i) to examine associations between non-response or partial response to repeat surveys subsequent to baseline and mortality and (ii) to examine associations between non-response and mortality by socioeconomic position.

Methods

The target population for Whitehall II was all Londonbased office staff aged 35–55 working in 20 civil service departments in 1985 (n = 14121). Baseline screening (Phase 1, 1985–88) involved a medical examination and a self-administered questionnaire. Subsequent to baseline, odd-numbered data collection phases have included both a medical examination and a questionnaire, while even-numbered phases have been questionnaire only (Box 1).

Response status

Phase 1: of the target population, 10 308 people (3413 women and 6895 men) were responders to baseline. Invitations to participate in follow-up phases were extended only to these baseline responders. Phases 2 and 4: responders at these Phases completed a full questionnaire. From Phase 4, response was augmented in the last resort through a short-form telephone interview consisting of selected questions from the questionnaire; such participants were categorized as partial

Box 1 Whitehall II Study response ra	tes
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(telephone) responders. Phases 3 and 5: responders at these Phases attended the medical examination; partial (questionnaire) responders completed a questionnaire, but did not attend the medical examination and partial (telephone) responders completed the shortform telephone interview only. We created a single, time-dependent, response status summary measure using all available data from Phases 1–5. The measure has four categories, which indicate the cumulative response status at each of the five phases. Response status summary measure: (i) completed all possible phases (all five phases, or all phases prior to death or censoring if this occurred before Phase 5): (ii) responders to Phase 1 who subsequently were nonresponders at one or more phases, but responded at their last possible phase i.e. the last phase they could have attended up to that point in time: (iii) responders to Phase 1 who subsequently were non-responders at one or more phases, including their last possible phase and (iv) baseline (Phase 1) non-responders.

All-cause mortality

Mortality data were obtained from the National Health Services Central Registry. The analyses of non-response to follow-up and partial-response to follow-up include the 10 308 responders to baseline from all the 20 departments that make up the Whitehall II study. Of these, 99.9% (10 297) have been followed for mortality. Mortality data were available for non-responders to baseline in 14 of the 20 departments that make up the Whitehall II study. Effective tracing of non-responders to baseline in the six remaining departments was impossible due to insufficient information.

Covariates

Information on age, sex and employment grade at baseline was derived from the Phase 1 questionnaire for responders and from Civil Service department files for non-responders. Civil Service employment grade was classified into high (senior administrators), intermediate (executives, professionals and technical staff) and low (clerical and office support staff). As of August 1992 the salary range among high-grade employees was £25330–87620 (€32583–112680, US\$44428–153622) and among low-grade employees £7387–11917 (€9498–15322, US\$12949–20895).^{18,19}

Phase of data collection	Start–Finish	Content	Response rate
Phase 1	1985–88	Medical and questionnaire	(73% of the target population)
Phase 2	1989–90	Questionnaire only	(79% of responders to Phase 1)
Phase 3	1991–94	Medical and questionnaire	(86% ^a of responders to Phase 1)
Phase 4	1995–96	Questionnaire only	(84% ^a of responders to Phase 1)
Phase 5	1997–99	Medical and questionnaire	(76% ^a of responders to Phase 1)

^aMaximum response rates—include questionnaire only and short-form telephone interviews.

Statistical methods

Associations between response status and mortality are presented as mortality rates per 1000 personyears, or hazard ratios (HRs) and 95% CIs determined using Cox proportional hazard models with follow-up as the time-scale and responders as the reference category. Person-years of follow-up were computed as follows: for responders, the start of follow-up (Phase 1) was their date of screening while for subsequent phases it was the date of first participation in that phase. For non-responders, the start of follow-up (Phase 1) was taken as the date that screening began for those who worked in the same building, while for subsequent phases the midpoint date for the responders was used. Participants were followed for mortality up to 31 July 2006 or their date of death (8%), embarkation (1%) or deregistration with a health authority (1%). For all analyses that included the non-responders at Phase 1, an indicator variable was included in the models to distinguish and adjust for any mortality differences between responders in the 14 departments that provided mortality data for the non-responders and the six departments that did not.

As there was no strong evidence of sex differences, pooled estimates are presented. The cumulative effect of non-response across phases was assessed using the response status summary measure updated at each phase and fitted as a time-dependent variable. The estimates are presented as HR and 95% CIs for the total cohort and separately by employment grade category.

Ethical approval

Ethical approval for the Whitehall II study was obtained from the University College London Medical School committee on the ethics of human research.

Results

The age- and sex-adjusted, all-cause mortality hazard for Phase 1 non-responders is double that for Phase 1 responders (Table 1). Non-response to follow-up at Phases 3 and 5 is associated with an increased hazard similar to that associated with Phase 1

Table 1 All-cause mortality by response status at Phase 1 (baseline) and at follow-up Phases 2–5, women and men combined

	п	Deaths	Mortality rate ^a	Adjusted for age, sex	Adjusted for age, sex, grade
Phase 1 (Baseline-Medical and qu	estionnaire	2)			
<i>n</i> = 11 819					
Responders	10297^{b}	737	3.8	1.0	1.0
Non-responders	1522	211	7.6	2.03 (1.73-2.39)	1.83 (1.56–2.16)
Phase 2 (Questionnaire only)					
n = 10239					
Responders	8125	509	3.9	1.0	1.0
Non-responders	2114	180	5.6	1.40 (1.18-1.66)	1.34 (1.13–1.59)
Phase 3 (Medical and questionnair	e)				
n = 10165					
Responders	8089	420	3.8	1.0	1.0
Partial (questionnaire) responders	706	73	8.6	2.00 (1.55-2.58)	1.98 (1.53–2.55)
Non-responders	1370	138	8.1	2.15 (1.76-2.62)	2.00 (1.64-2.45)
Phase 4 (Questionnaire only)					
n = 10053					
Responders	8598	429	4.6	1.0	1.0
Non-responders	1455	122	8.1	1.76 (1.44-2.16)	1.64 (1.34-2.02)
Phase 5 (Medical and questionnair	e)				
<i>n</i> = 9931					
Responders	6533	225	4.2	1.0	1.0
Partial (questionnaire) responders	736	49	8.7	1.89 (1.39, 2.58)	1.83 (1.34, 2.50)
Non-responders	2662	189	9.2	2.23 (1.83, 2.71)	2.10 (1.72, 2.57)

^aCrude mortality rate per 1000 person-years.

^bOf the 10308 participants at baseline, 99.9% (10297) have been followed for mortality.

					HR (95% CI)	
Employment grade	Response status summary measure	Deaths	Person- years	Mortality rate ^a	Adjusted for age and sex	Adjusted for age, sex and grade
All	Completed all possible phases ^b	356	136 701	2.6	1.0	1.0
	Missed ≥ 1 phases and completed last phase ^c	88	22 124	4.0	1.38 (1.09–1.74)	1.34 (1.06–1.70)
	Missed ≥ 1 phases and missed last phase ^d	293	34 421	8.5	2.27 (1.94-2.67)	2.14 (1.82-2.51)
	Phase 1 non-responders	211	27 802	7.6	2.77 (2.32-3.31)	2.52 (2.10-3.02)
High	Completed all possible phases	106	11 858	2.4	1.0	1.0
	Missed ≥ 1 phases and completed last phase	21	1639	3.4	1.28 (0.80-2.04)	
	Missed ≥ 1 phases and missed last phase	58	1469	8.0	2.53 (1.82-3.50)	
	Phase 1 non-responders	28	1336	5.5	2.25 (1.46-3.45)	
Middle	Completed all possible phases	164	18451	2.4	1.0	1.0
	Missed ≥ 1 phases and completed last phase	37	2833	3.6	1.29 (0.90-1.84)	
	Missed ≥ 1 phases and missed last phase	119	3040	7.8	2.20 (1.73-2.80)	
	Phase 1 non-responders	92	3421	7.1	2.81 (2.16-3.67)	
Low	Completed all possible phases	86	7342	3.3	1.0	1.0
	Missed ≥ 1 phases and completed last phase	30	1633	5.4	1.46 (0.96–2.21)	
	Missed ≥ 1 phases and missed last phase	116	2420	9.8	1.85 (1.39–2.48)	
	Phase 1 non-responders	91	2559	9.5	2.14 (1.56-2.95)	

Table 2 All-cause mortality by summary response status stratified by employment grade, women and men combined

^aCrude mortality rate per 1000 person-years.

^bPhases 1–5 or every phase prior to death or censoring.

^cPhase 1 responders who missed ≥ 1 phases after Phase 1, but completed their last possible phase.

^dPhase 1 responders who missed ≥ 1 phases after Phase 1 including their last possible phase.

non-response, but non-responders to Phases 2 and 4 (questionnaire-only phases) have lower HRs.

Partial response

There was no evidence that the HR for mortality among partial (questionnaire) responders at Phase 3 (medical screening phase) was different from that for Phase 3 non-responders. At Phase 4 (questionnaire only), there was no evidence that the HR for mortality among partial (telephone) responders, HR (95% CIs) 1.15 (0.79–1.67), was different to that for Phase 4 responders. At Phase 5 (medical screening phase), the HR for mortality among partial (telephone) responders [2.08 (1.45-2.98)] did not differ from that for partial (questionnaire) responders [1.89 (1.38–2.58)], and both were no different to that for Phase 5 non-responders [2.23 (1.83-2.71)]. Adjustment of the HRs for non-responders and partial responders for employment grade gave estimates only marginally below those adjusted for age and sex.

Non-response

Table 2 presents associations between the response status summary measure and mortality. Due to similar HRs for mortality between partial responders and non-responders at medical examination phases, and between partial (telephone) responders and responders at questionnaire phases, partial responders and non-responders are combined at Phases 3 and 5, while partial responders and responders are combined at Phase 4. Compared with participants who completed all possible follow-up surveys, the HR for mortality among non-responders at Phase 1 was 2.77 (2.32–3.31)—Table 2, upper section. The HR was 2.27 (1.94-2.67) among those who were nonresponders at one or more subsequent phases and non-responders at their last possible phase, but 1.38 (1.09–1.74) among those who were non-responders at subsequent surveys but responders at their last possible phase. Adjustment for employment grade had little effect on these findings and analyses of associations between the response status summary measure and mortality within each of the three employment grade categories provided no evidence that associations differ by socioeconomic position-Table 2, lower section.

Non-response to last survey

As indicated above, missing one or more follow-up surveys was associated with an excess mortality risk of only 38% (P=0.007) provided the last possible survey had been completed. To quantify the additional mortality risk associated with non-response to the last possible survey, we undertook further analyses in which responders who missed one or more phases including the last possible phase were

compared with those who either completed all their possible phases or who missed one or more phases but completed their last possible phase. This comparison showed that the additional mortality risk associated with missing the last possible phase was >100%, HR 2.02 (1.73–2.36), P < 0.001.

Discussion

The key finding of this article is that non-response to clinical follow-up was associated with the same doubling of the mortality risk as non-response to baseline. Non-response to one or more follow-up surveys was associated with a 1.4-fold excess risk among those who responded to their last possible survey, but a 2.3-fold excess risk among those who missed their last possible survey. We also found that excess mortality risk is negligible among partial responders to questionnaire-only surveys, but similar to that for non-responders for participants who miss a medical examination but respond partially through questionnaire completion. In addition, the study examined non-response by socioeconomic position and found no evidence that socioeconomic position modifies the association between non-response and mortality. Our findings that non-responders to baseline had a mortality hazard double that for responders is in line with findings from previous studies.^{2,5,7–14,20–26}

Although a number of studies have reported somewhat higher levels of non-response in low socioeco-nomic groups,^{5,6,23,27–33} few studies have examined whether the effects of non-response differ across the socioeconomic strata.^{29,30,33} In the present study, mortality risks among non-responders were little attenuated by adjustment for socioeconomic position measured by employment grade. Similarly, analyses within grade provided no evidence of differences in the associations between non-response and mortality. Recent findings from the population-based FINRISK surveys similarly observed little attenuation of the association between non-response to baseline and mortality on adjustment for socioeconomic position, and a 2-fold higher mortality risk in non-responders to baseline in every socioeconomic category.¹⁴ Studies that have examined morbidity have similarly provided little evidence that the bias introduced by non-response to baseline differs across socioeconomic strata.²

It is implausible that there is a direct association between non-response and mortality, rather it is an association driven by a variety of common causes, for example ill-health, caring and accessibility. Some of these, such as measureable ill-health at baseline, were available for participants in the Whitehall II study, but not for non-responders to baseline, and many others were not measured. Consequently, we have chosen to present our analyses adjusted only for age, sex and grade, which were available for everyone included in the analyses. Participants in the Whitehall II study were white-collar civil servants at study entry and covered a wide range of grades with a salary difference of >10-fold between the top and bottom of the hierarchy. In this respect they are representative of the majority of employees in Western countries. Nonetheless, this work needs to be repeated in general population samples as the generalizability of our findings will be limited by the healthy worker effect at baseline, which may produce an underestimation of associations between non-response and mortality.

Our findings from a study with repeat data collections show the excess mortality associated with nonresponse to previous follow-up surveys is \sim 1.4-fold if the most recent possible survey is completed, but >2-fold if it is missed. When collecting questionnaire data only, the health-related selection introduced by the inclusion of partial responders who only complete a short telephone interview is similar to that for those who complete a full questionnaire. This means their inclusion in analyses, while increasing the power of the study, will not reduce the bias associated with complete non-response to follow-up at that phase. However, when data collection includes both a medical and a questionnaire, partial responders who miss the medical examination are similar in health status to those who are complete non-responders. Inclusion in analyses of questionnaire or telephone interview data for these partial responders would thus both increase power and reduce the health-selection bias associated with non-response to follow-up. Our findings regarding partial responders are plausible. Attendance at a medical examination requires much more time and effort than completing a questionnaire or short-form telephone interview at home. Thus, nonresponse to surveys that include a medical examination outside the home is likely to capture a wider range of the unmeasured common causes of nonresponse and mortality. Additionally, any association between non-response and mortality may be strengthened further if non-responders miss a diagnosis or treatment that might have been brought to attention by the medical examination.

With respect to studies designed to examine inequalities in health, good news is provided by evidence that socioeconomic position does not modify the association between non-response and mortality.

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Conflict of interest: None declared.

KEY MESSAGES

- In common with non-response to baseline, mortality among non-responders to clinical follow-up is double that of responders. Non-response to follow-up surveys prior to the most recent survey is associated with a lower excess mortality risk.
- Excess mortality risk is negligible among partial responders to questionnaire only surveys, but similar to that for non-responders for participants who miss a medical examination (but have partial response via questionnaire).
- Socioeconomic position does not modify the association between non-response and mortality.

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