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Common Genetic Influences Underlie Comorbidity of Migraine and Endometriosis

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Abstract

We examined the co-occurrence of migraine and endometriosis within the largest known collection of families containing multiple women with surgically confirmed endometriosis and in an independent sample of 815 monozygotic and 457 dizygotic female twin pairs. Within the endometriosis families, a significantly increased risk of migrainous headache was observed in women with endometriosis compared to women without endometriosis (odds ratio [OR] 1.57, 95% confidence interval [CI]: 1.12–2.21, $P = 0.009$). Bivariate heritability analyses indicated no evidence for common environmental factors influencing either migraine or endometriosis but significant genetic components for both traits, with heritability estimates of 69 and 49%, respectively. Importantly, a significant additive genetic correlation ($r_G = 0.27$, 95% CI: 0.06–0.47) and bivariate heritability ($h^2 = 0.17$, 95% CI: 0.08–0.27) was observed between migraine and endometriosis. Controlling for the personality trait neuroticism made little impact on this association. These results confirm the previously reported comorbidity between migraine and endometriosis and indicate common genetic influences completely explain their co-occurrence within individuals. Given pharmacological treatments for endometriosis typically target hormonal pathways and a number of findings provide support for a relationship between hormonal variations and migraine, hormone-related genes and pathways are highly plausible candidates for both migraine and endometriosis. Therefore, taking into account the status of both migraine and endometriosis may provide a novel opportunity to identify the genes underlying them. Finally, we propose that the analysis of such genetically correlated comorbid traits can increase power to detect genetic risk loci through the use of more specific, homogenous and heritable phenotypes.

Keywords

heritability; headache; women; hormones; pain; comorbid

INTRODUCTION

Typical migraine is a frequent, debilitating and painful disorder that normally affects people during their most productive years. Mutually confirmed prevalence data from population-based studies in Denmark [Rasmussen et al., 1991b; Russell and Olesen, 1995], France [Henry et al., 1993; Michel et al., 1996], the USA [Stewart et al., 1992, 1996] and Canada [Edmeads et al., 1993; O'Brien et al., 1994] all using International Headache Society (IHS) diagnostic criteria [Headache Classification Committee of the International Headache Society, 1988] have shown a homogenous picture regarding migraine prevalence. These studies produce a reasonable estimate of the 1-year prevalence of migraine in adults of 10–12%; 6% among men and 15–18% among women. The World Health Organization [2001] recently identified migraine among the world's top 20 leading causes of disability, with an impact that extends far beyond the suffering individual, to the family and community. Although migraine is highly prevalent in our society, its aetiology remains relatively obscure and there are no laboratory-based diagnostic tests that identify those who suffer from the disorder. Twin studies indicate that migraine has a significant genetic component, with heritability estimates of 33–65% [Gervil et al., 1999; Honkasalo et al., 1995; Larsson et al., 1995; Mulder et al., 2003; Nyholt et al., 2004; Ulrich et al., 1999].

Endometriosis is a gynaecologic disorder affecting up to 10% of women and is defined by the presence of endometrial tissue outside of the uterus, growing, bleeding cyclically and causing adhesions. Common symptoms are severe pelvic pain, painful menstrual periods (dysmenorrhea) and infertility. The main pathological processes associated with the disease are peritoneal inflammation and fibrosis, and the formation of adhesions and ovarian cysts. The disorder often recurs and has a major impact on women's health, relationships, productivity and life choices. The risk of endometriosis increases with age within the reproductive years, however its aetiology remains obscure. Some risk factors relate to oestrogen levels, including positive smoking history (which reduces ovarian production of oestrogen and reduces risk of endometriosis) and increased peripheral body fat (which increases oestrogen), plus increased exposure to menstruation (i.e., shorter cycle length, longer duration of flow and nulliparity) [Berkley et al., 2005; Eskenazi and Warner, 1997]. Genetic factors also contribute to endometriosis with a heritability of 51% and relative recurrence risk to siblings of 2.34 estimated in our Australian sample of twins [Treloar et al., 1999]; however, the value may be as high as 15 in the sisters of women with more severe disease, based upon imaging studies [Kennedy et al., 1997].

A Finnish clinical study published in 1975 of 125 patients, aged 19–44 who were in hospital for abdominal operation found of the 50 patients with the solitary diagnosis of endometriosis, headache was reported by 42 (84%) compared to only 26 (60%) of the 44 controls with no endometriosis ($\chi^2_1 = 7.26, P = 0.007$). In addition, 14 (28%) of the endometriosis patients described their headache to be of the nature of migraine, compared to only 8 (18%) of controls ($\chi^2_1 = 5.19, P = 0.023$) [Tervila and Marttila, 1975]. Also, a USA clinical trial assessing the effect of leuprolide acetate [a synthetic nonapeptide analogue of naturally occurring gonadotropin releasing hormone (GnRH or LH-RH) which is a very potent blocker of follicle-stimulating hormone (FSH) and luteinizing hormone (LH): lowering levels of oestrogen, progesterone and testosterone] and oophorectomy (surgical removal of an ovary or ovaries) on IHS menstrual migraine in 29 reproductive age female subjects, noted 11 of 29 volunteers experienced dysmenorrhea and 10 of 23 patients (with previous pelvic surgery) had surgically confirmed endometriosis. These symptoms and findings were present equally in those whose migraines responded to leuprolide and those who did not respond [Lichten et al., 1995].

More recently, two studies reported comorbidity of migraine and endometriosis. The first study of 133 Italian women with endometriosis reported a significantly higher ($P < 0.001$) prevalence

of migraine ($n = 51$, 38.3%, 95% CI: 30.1–47.2%) than in 166 Italian women without endometriosis (controls) ($n = 25$, 15.1%, 95% CI: 10.0–21.4%) [Ferrero et al., 2004]. Similarly, a USA study reported a higher prevalence of endometriosis in 50 females with migraine than in 52 age-matched women without migraine (30 versus 4%, $P < 0.001$) [Tietjen et al., 2006]. When migraine and endometriosis are comorbid, the well-being of the patient may be further impaired.

Therefore, to further investigate the comorbidity of migraine and endometriosis we collected migraine symptom data from women in our cohort of 931 Australian families containing at least two sisters with endometriosis (≥ 1 affected-sister-pair, ASP) [Treloar et al., 2005]. In addition, we used an independent sample of 815 monozygotic (MZ) and 457 dizygotic (DZ) female twin pairs for whom endometriosis status and migraine symptom data were available to assess the aetiology of migraine and endometriosis comorbidity.

Also, several studies have demonstrated a cross-sectional relation between psychiatric disorders and migraine in community samples, and, in particular, subjects with migraine report more affective and anxiety disorders and exhibit elevated rates of neuroticism and somatisation compared to non-migraine subjects [Merikangas et al., 1993]. Indeed, a recent cross-sectional study of female headache outpatients and healthy controls found anxiety was more common in migraine with endometriosis group, than in the cohort with migraine without endometriosis (OR = 2.2, 95% CI: 1.0–4.7) [Tietjen et al., 2007]. Given persons with a high degree of neuroticism are sensitive and vulnerable to various stimuli, including stress and pain, they are more likely to fulfil the disability criterion and thereby the criteria for migraine; we also investigated whether neuroticism [Eysenck and Eysenck, 1974] underlies the co-occurrence of migraine and endometriosis in our twin data.

MATERIALS AND METHODS

SAMPLES

The endometriosis ASP families were recruited for our studies aimed at identifying the molecular mechanisms underlying the disorder by looking for genomic regions co-inherited with endometriosis and represent the largest known cohort of women diagnosed (surgically confirmed) with endometriosis and their family members [Treloar et al., 2005].

Given the strong selection bias of the endometriosis ASP families (i.e., there is little variation in endometriosis status due to most women being affected), to determine whether comorbidity of migraine and endometriosis results from (i) coincidence or selection bias, (ii) one condition causing another or (iii) shared environmental or genetic risk factors, we investigated whether genetic influences on endometriosis overlap with genetic influences on migraine, utilising a community sample of twins for whom endometriosis and migraine phenotype data were available.

The twin sample participants were members of the cohort of 1,979 female twin pairs that were identified originally in 1980–1982 from the Australian National Health and Medical Research Council (NHMRC) Twin Register and followed up in 1988–1990 [Heath et al., 1994].

Approval to conduct the research was obtained from the Queensland Institute of Medical Research (QIMR) Human Research Ethics Committee and from the Australian Twin Registry.

ASSESSMENT

Analogous to our previous studies on the genetics of migraine [Nyholt et al., 2004, 2005], migraine symptom data were obtained from 961 women aged 40–70 in our endometriosis ASP families during the course of an extensive, semi-structured telephone interview that included

diagnostic assessments of migraine, based on IHS criteria (Table I) and similar to the Finnish Migraine-Specific Questionnaire [Kallela et al., 2001]. The interview performed by lay interviewers trained and experienced in migraine diagnosis, yielded diagnoses for *migraine without aura* (MO) and *migraine with aura* (MA), with the use of visual prodromal symptoms as an index of MA.

In 1993–1994, a questionnaire focused on gynaecologic conditions and hysterectomy was sent to both members of 1,570 female twin pairs, plus a further 158 individual female twins in incomplete pairs (3,298 individuals) who could still be contacted and were willing to participate in twin research [Treloar et al., 1999]. Participants were asked to complete a four-page questionnaire (“Gynaecological Health Study”), which included questions on gynaecologic problems that might predispose to hysterectomy, other surgical interventions, and medical and hormonal treatments. Twins were asked to report whether they had ever had endometriosis, their age at its onset, the nature of its investigation and/or treatment, and whether their cotwin or mother had endometriosis.

In 1993–1995, a telephone interview (by lay interviewers trained and experienced in migraine diagnosis) including a diagnostic assessment of psychiatric disorders, including alcohol use and abuse, anxiety, depression and phobias was administered to the same cohort of twins [Heath et al., 1997]. As before, twins were unselected with regard to personal or family history of alcoholism or other psychiatric or medical disorders. Participants answering “yes” to ever having “migraine or recurrent attacks of headache” (screening positive) then answered a number of questions developed by an experienced migraine researcher (K.R.M.) based on IHS criteria [Nyholt et al., 2004, 2005].

Neuroticism scores were collected via lay telephone interview (by lay interviewers trained and experienced in the assessment of psychiatric disorders) as part of a previous study in 1988 on the same set of twins for whom migraine and endometriosis data were available. The neuroticism sample and data have been described elsewhere [Birley et al., 2006; Kirk et al., 2000], and previous analyses have shown that this population-based twin sample is typical of the Australian population in many respects including the prevalence of psychiatric symptoms [Kendler et al., 1986].

STATISTICAL ANALYSIS

Cross-tabulations and kappa (K) calculations were performed using SPSS 13 (SPSS Inc.). Tetrachoric correlations (r_t) and odds ratios (ORs) were calculated using the Mstat program (<http://www.qimr.edu.au/davidD/>)

Latent class analysis (LCA) is a statistical method for finding subtypes of related cases (latent classes) from multivariate categorical data. LCA model associations (covariation) between observed variables that imperfectly measure a non-observable (latent) variable [McCutcheon, 1987; Rindskopf and Rindskopf, 1986]. By assessing the symptom profile of individual patients, LCA produces mutually exclusive groups (classes) of *patients* based upon their patterns of symptoms.

As described previously [Nyholt et al., 2004, 2005], latent class cluster models were fitted to 10 trichotomous symptom response variables (i.e., negative to screening question, screening positive but negative for symptom, and screening positive and positive for symptom) based on IHS diagnostic criteria (Table I) using the Latent GOLD 2.0 package (Statistical Innovations Inc.). For each Latent GOLD run, up to 10,000 iterations of the EM algorithm were allowed using a convergence criterion of 1×10^{-10} . Each LCA solution was restarted at least 100 times with new starting values to find the maximum likelihood estimates for the parameters.

Classification information was requested for latent class cluster models. For each symptom profile, the classification output contains the associated probabilities of belonging to each cluster. Also, bivariate residuals were obtained, which identify correlations between symptom pairs that have not been adequately explained by the model.

Although estimates of class membership and symptom endorsement probabilities ignored the twin structure of the data, which may result in the likelihood-ratio chi-square (χ^2) test overestimating the significance of adding an extra class, as in traditional factor analysis [Neale and Cardon, 1992], estimates of symptom endorsement and class membership probabilities will remain statistically unbiased [Madden et al., 1997]. Subsequently, and as recommended for large sample sizes, the comparative fits of LCA models were assessed by evaluating the Bayes information criterion (BIC) [Schwarz, 1978] where, if the BIC of a more complex model fails to decrease, the simpler model (having the lower BIC) will be selected.

Subgroups of migraine/severe headache sufferers were identified from the LCA results. Briefly, LCA identified one asymptomatic class (CL0) and three major symptomatic classes, representing (i) a mild form of recurrent non-migrainous headache (CL1), (ii) a moderately severe form of migrainous headache (CL2) typically *without* visual aura, loading (i.e., present at least 50% of the time) on all IHS MO symptoms except “unilateral location” and “nausea, vomiting or diarrhoea”, and (iii) a severe form of migraine typically *with* visual aura, loading on all IHS symptoms (CL3). Of particular importance, although “aura” was often associated with other and more severe neurological symptoms and was predominantly found in latent class 3 (CL3), almost one quarter of class-3 individuals did not report “aura”. Furthermore “aura” was present in 39.3% of individuals in class-2 (CL2). Individuals in LCA groups CL2 and CL3 are considered to suffer *migrainous headache*. We have previously shown our LCA-derived migraine phenotype provides a stable and accurate correspondence between genetic risk and migraine headache [Nyholt et al., 2004, 2005]. Importantly, genetic-model fitting indicated a greater genetic contribution to migraine using the LCA *migrainous headache* classification ($h^2 = 0.41$, 95% CI: 0.23–0.49) compared with the IHS MO/MA classification ($h^2 = 0.33$, 95% CI: 0.05–0.44) [Nyholt et al., 2004].

GENETIC-MODEL FITTING

We used genetic-model fitting techniques in Mx software [Neale et al., 2003] to obtain estimates of the genetic and environmental factors. These techniques make optimum use of the information available in the twin-co-twin covariance structure [Neale and Cardon, 1992]. Model fitting is based on the comparison of the observed and expected variance-covariance matrices in MZ and DZ twin pairs.

Genetic-model fitting used structural equation modelling (SEM) to estimate parameters of a model that included additive genetic effects (A), non-additive genetic effects (i.e., dominance and/or epistasis) (D) or shared family environment (C), and random or unique environment (E) [Neale and Cardon, 1992]. For each model, we obtained: a χ^2 goodness-of-fit statistic ($-2LL$) and calculated the sample size adjusted BIC parsimony measure which indexes the fit of the model and its parsimony. Nested models were compared using the likelihood ratio (χ^2 difference) test ($\Delta-2LL$), a significant χ^2 value indicating a deterioration in model fit, and by examining the change in BIC (better fitting models produce lower values of BIC). Likelihood-based 95% confidence intervals for estimates of genetic and environmental parameters were then computed using Mx [Neale et al., 2003].

RESULTS

A large proportion of the 961 women in our endometriosis ASP families suffer from migraine headache: 509 (53%) were diagnosed with *migrainous headache* via our LCA approach [Nyholt et al., 2004, 2005] and 345 (35.9%) satisfied the strict criteria for IHS MO and/or MA.

Analysis of the 931 women for whom both migraine and endometriosis status were available (i.e., 30 women had unknown endometriosis status) revealed a significantly increased risk of migrainous headache in women with endometriosis compared to women without endometriosis (Table II) ($\chi^2_1 = 6.86$, $P = 0.009$), with an OR of 1.57 (95% CI: 1.12–2.21). Similarly, an increased risk of IHS migraine in women with endometriosis compared to women without endometriosis was observed (Table II) (OR = 1.35, 95% CI: 0.94–1.95). As shown in Table II, fewer individuals are considered affected for the strict IHS migraine phenotype compared to the more sensitive LCA-based *migrainous headache* phenotype [Nyholt et al., 2004] (35.8 versus 52.9%), resulting in a reduction in power to detect a significant relationship between IHS migraine and endometriosis.

It is worth noting that in the endometriosis family dataset, the mean age of women *with* endometriosis (50.35 yrs) was significantly lower than women *without* endometriosis (56.88 yrs) ($P < 0.001$). Similarly, the mean age of women *with* migraine (50.81 yrs) was significantly lower than women *without* migraine (52.25 yrs) ($P = 0.002$). Also, the mean age of women *without* endometriosis was similar in the non-migraine (57.54 yrs) and migraine subgroup (56.03 yrs), as was the mean age of women *with* endometriosis in the non-migraine (50.85 yrs) and migraine subgroup (49.94 yrs). Given the prevalence of endometriosis and lifetime migraine both increase with age, age differences between women with and without endometriosis cannot underlie the observed increased prevalence of migraine in endometriosis sufferers (i.e., women with endometriosis would need to be *older* than women without endometriosis for age to underlie comorbid migraine and endometriosis). Indeed, the observation of women with comorbid migraine and endometriosis being younger than the women without migraine and endometriosis further supports an association between these two traits. As a consequence, in these data it is neither necessary nor sensible to age-adjust ORs relating the risk of migraine in women with endometriosis compared to women without endometriosis.

Migraine symptom data were available on 815 MZ and 457 DZ female twin pairs for whom endometriosis status had previously been determined [Treloar et al., 1999]. Twin-pair cross-tabulations (Table III) show an increased correlation in *both* migrainous headache and endometriosis for MZ co-twins compared to DZ co-twins, of proband twins presenting with either trait, suggesting the presence of both individual and shared genetic effects.

Except for a subgroup of women who participated in a later study which collected more detailed symptom data, the majority (84.7%) of the 2,544 female twins were not asked questions relating to having at least five migraine/episodes of headache during their lifetime, whether their average typical migraine/headache lasted between 4 and 72 h, and whether they would describe the pain associated with their headache as moderate or severe. Consequently, although IHS criteria C and D (Table I) status were available on all participants, strict IHS migraine diagnoses could be determined for only a small proportion (15.3%) of the twins and resulted in too small a sample to perform bivariate heritability analyses for IHS migraine. However, we have previously shown there to be a high correlation between our LCA-derived migrainous headache phenotype and migraine diagnosed according to strict IHS criteria [Nyholt et al., 2004,2005].

In the current twin data for whom IHS migraine status could be determined ($n = 389$), the tetrachoric correlation between migrainous headache and IHS migraine was estimated at 0.94

(95% CI: 0.87–0.99) with a diagnostic reliability kappa (K) of 0.47 (95% CI: 0.38–0.56, $P = 4 \times 10^{-28}$). The tetrachoric correlation between fulfilling only IHS criteria C1D and IHS migraine was estimated at 0.95 (95% CI: 0.89–1.00), $K = 0.52$ (95% CI: 0.43–0.60, $P = 4 \times 10^{-31}$), while the tetrachoric correlation between fulfilling IHS criteria C1D and being in the most severe migrainous headache class (CL3) [Nyholt et al., 2004] (IHS criteria C+D+CL3) (termed *migrainous IHS headache*) and IHS migraine was estimated at 0.90 (95% CI: 0.85–0.95), $K = 0.69$ (95% CI: 0.61–0.77, $P = 4 \times 10^{-42}$). As suggested by the high correlation between these migraine measures, similar results from bivariate heritability analyses were obtained for migrainous headache, IHS criteria C+D and IHS criteria C+D+CL3 (*migrainous IHS headache*). Given all individuals diagnosed as ‘affected’ under the strict IHS scheme are also diagnosed with migrainous headache, for simplicity we only report bivariate heritability results based on the latter, more severe and ‘clinically relevant’ migraine phenotype, *migrainous IHS headache* (see Table III for MZ/DZ cross-tabulations). Individuals classified as *migrainous IHS headache* suffer a severe form of migraine headache (i.e., suffering a greater number of IHS symptoms) and typically fulfil strict IHS criteria.

Bivariate analyses utilising the correlated liability (triangular decomposition or “Cholesky” factor) model performed using the Mx program [Neale et al., 2003], indicated significant additive genetic components for both migrainous IHS headache and endometriosis, with heritability estimates of 69% (95% CI: 60–77%) and 48.76% (95% CI: 32–63%), respectively. There was no evidence for common environmental or non-additive genetic factors influencing either migraine or endometriosis, with non-shared environmental factors (which also subsumes any errors of measurement) explaining the remaining variance in liability to migraine (31%, 95% CI: 23–40%) and endometriosis (51.24%, 95% CI: 37–68%). Also, there was no significant evidence for age, age² or age³ effects.

Importantly, a significant additive genetic correlation was observed between migraine and endometriosis ($r_G = 0.27$, 95% CI: 0.06–0.47), while the observed non-shared environmental correlation between migraine and endometriosis ($r_E = 0.08$, 95% CI: 0–0.32) was not significant. Moreover, the significant bivariate heritability estimate of 15.37% (95% CI: 4–27%) ($P = 0.010$), and non-significant non-shared environmental factors estimate of 2.99% (95% CI: 0–13%) ($P = 0.560$) indicate common genetic influences wholly underlie the observed covariance (comorbidity) between migraine and endometriosis. Figure 1 portrays the bivariate model best fitting the covariation between migraine and endometriosis (i.e., only the significant ($P \leq 0.05$) parameters are included).

The principal genetic features of this model are as follows. First, the common additive genetic factors accounting for 69% (95% CI: 60–77%) of the variance in migraine also account for 17% (95% CI: 8–27%) ($P = 0.0005$) of the variance in endometriosis and completely explain their covariance (i.e., bivariate heritability of 17%). Secondly, there are specific additive genetic factors accounting for 49% (95% CI: 32–64%) of variance in endometriosis. Individual (non-shared) environmental factors explain the remaining variance in liability to migraine and endometriosis.

Furthermore, bivariate heritability analyses utilising direction-of-causation models [Duffy and Martin, 1994; Neale and Cardon, 1992] did not support migraine as the cause of endometriosis or vice versa, with both models providing a similar fit to the data (sample size adjusted BIC parsimony measures of –8,280.13 and –8,279.44, respectively).

Finally, although results from trivariate heritability analyses indicated that neuroticism and migraine share common additive genetic factors, neuroticism and endometriosis do not. Moreover, inclusion of neuroticism in the model caused no significant change in the variance

component estimates of genetic and environmental covariance between migraine and endometriosis (data not shown).

DISCUSSION

Our results represent the two largest studies to date on the comorbidity of migraine and endometriosis and the first study on bivariate heritability of these two disorders. The results do not support migraine and endometriosis co-occurrence being due to coincidence or selection bias, shared environmental effects or one condition causing another. Instead, the results confirm migraine and endometriosis comorbidity and indicate that additive genetic effects completely explain their co-occurrence within individuals and which are unrelated to personality.

Although migraine diagnosis is preferably performed via direct communication between physician and patient (which is impractical for thousands of participants spread throughout Australia), our study obtained symptom information and level of detail similar to other migraine studies, which were repeatedly found to be a valid diagnostic tool for migraine [Gervil et al., 1998; Rasmussen et al., 1991a; Russell et al., 1995]. Compared to their sole use, our combined use of both questionnaire and interview increases the reliability of symptom data. For example, using a similar screening approach, Stewart et al. (1997) obtained a 92.6% positive predictive value of their telephone interview diagnosis compared with their clinical examination [Stewart et al., 1997]. Furthermore, any misclassification resulting from diagnosis based on questionnaire and/or trained lay interview is expected to result in *underestimation* of migraine prevalence compared to expert interview [Tzourio et al., 2003], will be minimised via our more sensitive LCA approach and be consistent across the endometriosis and non-endometriosis individuals, and thus have negligible effect on the results of our comorbidity analyses.

The agreement (Cohen's kappa) between LCA-derived *migrainous headache* and strict IHS migraine was 0.66 in the endometriosis family data and 0.69 between LCA-derived *migrainous IHS headache* and strict IHS migraine in the twin data. These kappas represent substantial agreement [Landis and Koch, 1977] and are very good in light of the uncertainty associated with the clinical diagnosis of migraine. For example, kappas range from 0.55 to 0.81 among different neurologists assigning headache diagnosis based on review of videotaped patient interviews [Granello et al., 1994]. These results support the validity and reliability of our LCA-based migraine phenotype; indeed the utility of our LCA approach to facilitate migraine research was highlighted in two recent reviews on migraine genetics [van den Maagdenberg et al., 2007; Wessman et al., 2007].

Similarly, although surgery is required for definitive diagnosis of endometriosis (which is obviously not possible for large population-based samples), we have previously shown our endometriosis self-report data (supplemented with information from medical, surgical and/or pathology reports where available) to have high agreement with medical report ($K = 0.52$, 95% CI: 0.42–0.61) and self-reported endometriosis adjusted by available medical and pathology data ($K = 0.82$, 95% CI: 0.78–0.86) [Treloar et al., 1999].

While twin studies indicate that migraine has a significant genetic component, importantly, these studies do not support the presence of genetic sex-specific effects (i.e., effects expressed in one sex but not the other), although this may be due to a lack of power to detect such effects [Martin et al., 1978]. In the absence of sex-specific genetic effects, it is likely that sex hormones play a pivotal role in the increased vulnerability to migraine among women. A number of findings provide support for a relationship between hormonal variations and migraine. These include evidence indicating that many women with migraine report worsening of their headaches around the time of menstruation [Edelson, 1985; Greene, 1967]. In fact, menstrual

migraine (i.e., 90% of attacks of MO occur in association with menstruation) is experienced by 24–25% of females with MO [Rasmussen, 1993].

Further support for a relationship between hormonal variations and migraine include reports of the oral contraceptive pill precipitating a first attack of migraine or worsening or improving the frequency and severity of existing attacks [Dalton, 1976; Kudrow, 1975; Ryan, 1978]. In addition, many women with migraine have fewer attacks during pregnancy [Callaghan, 1968; Nattero, 1982; Somerville, 1972], although the condition may be exacerbated in the post-partum period [Stein, 1981]. Also, some women have fewer attacks or cease to have attacks after menopause [Martin et al., 1971].

Finally, the prevalence of migraine has generally been shown to be equal in boys and girls before puberty but increases at a greater rate in girls as adolescence approaches, so that by adulthood the female to male ratio has increased to 3:1 [Lance, 1982; Welch et al., 1984]. It has been postulated that this age-dependent increase in female prevalence could be the result of hormonal fluctuations triggering a primary predisposition [Dennerstein et al., 1978; Greene, 1969].

An association has been reported between the oestrogen receptor 1 (*ESR1*) gene exon 8 G594A (rs2228480) polymorphism and migraine susceptibility in two independent Australian case-control (CC) cohorts (224CC and 260 CC) [Colson et al., 2004], while an exon 4 C325G (rs1801132) polymorphism was reported to be associated with migraine in women in a Spanish cohort (367 cases, 232 controls) [Oterino et al., 2006]. Interestingly, rs2228480 was not associated in the Spanish cohort and rs1801132 was not found to be associated in the Australian cohort [Colson et al., 2006]. More recently, a large Finnish study of 898 MA cases and 900 controls reported nominal association for 5 out of 26 *ESR1* SNPs tested (rs6557170, rs2347867, rs6557171, rs4870062, rs1801132; uncorrected *P*-values 0.007–0.034), but these results did not remain significant after taking multiple testing into account [Kaunisto et al., 2006]. Although these results indicate that neither rs2228480 nor rs1801132 are likely to be the causal variant(s), they nonetheless provide support for a role of *ESR1* in migraine susceptibility and highlight the need for more detailed analysis of variants within *ESR1*.

Given pharmacological treatments for endometriosis typically target hormonal pathways and *ESR1* gene variants have also been reported to be associated with endometriosis [Georgiou et al., 1999; Hsieh et al., 2007; Kim et al., 2005], hormone-related genes and pathways are highly plausible candidates for both migraine and endometriosis. Hence, taking into account the status of both migraine and endometriosis may provide a novel opportunity to identify the genes underlying them. The identification of such genes will provide clues to the further elucidation of the complex molecular pathways of migraine and endometriosis and will help in the development of diagnostic tests and rational treatment strategies.

Most recently, a retrospective study of 258 Taiwanese women with menstrual migraine who underwent laparoscopic uterosacral nerve ablation (LUNA) for the treatment of dysmenorrhea found a significant amelioration of migraine severity after LUNA in women who had at least one grade improvement in the multidimensional scoring of dysmenorrhea at the 12 month follow-up. The results suggest LUNA may offer an approximate 25% chance of lessening migraine severity in certain women [Juang et al., 2006].

Furthermore, women with comorbid migraine and endometriosis may particularly benefit from the use of nonsteroidal anti-inflammatory drugs (NSAIDs) commonly used to treat migraine, as NSAIDs are one of the recommended therapies in women with menorrhagia [Kirtava et al., 2003; MacGregor, 2000]. Menorrhagia (excessive and/or prolonged menstrual bleeding) is a risk factor for endometriosis [James, 2005] and was also recently associated with migraine [Tietjen et al., 2006]. Similarly, given the strong oestrogen link to both migraine and

endometriosis, women comorbid for these conditions may especially benefit from therapies which suppress and/or stabilise oestrogen production. In contrast, treatment of migraine with ergotamine may exacerbate endometriosis, as ergot derivatives are known to induce fibrosis [Robert et al., 1984].

In summary, these results confirm the previously reported comorbidity between migraine and endometriosis and indicate common genetic influences completely underlie their co-occurrence within individuals. Therefore, to improve the physical health and emotional well-being of many women—via improved treatment and pain relief—we suggest the presence of migraine be investigated in women with endometriosis and vice versa. Additionally, our data indicate comorbid migraine and endometriosis represents a subset of individuals with unique genetic effects and may therefore provide a novel opportunity to identify the genes underlying them. Consequently, we propose that the analysis of such genetically correlated comorbid traits can increase power to detect genetic risk loci through the use of more specific, homogenous and heritable phenotypes. Also, the identification of genetically correlated comorbid traits has particular relevance to the currently popular genome-wide association (GWA) paradigm by providing a tangible means to substantially improve the efficient use of both existing and future GWA data. For example, in addition to the primary association analysis of GWA data to the phenotype for which it was ascertained, secondary analyses may be performed for genetically correlated comorbid traits. Furthermore, examination of comorbid/non-comorbid cases (and controls) provides a unique opportunity to knowingly stratify according to genetic homogeneity.

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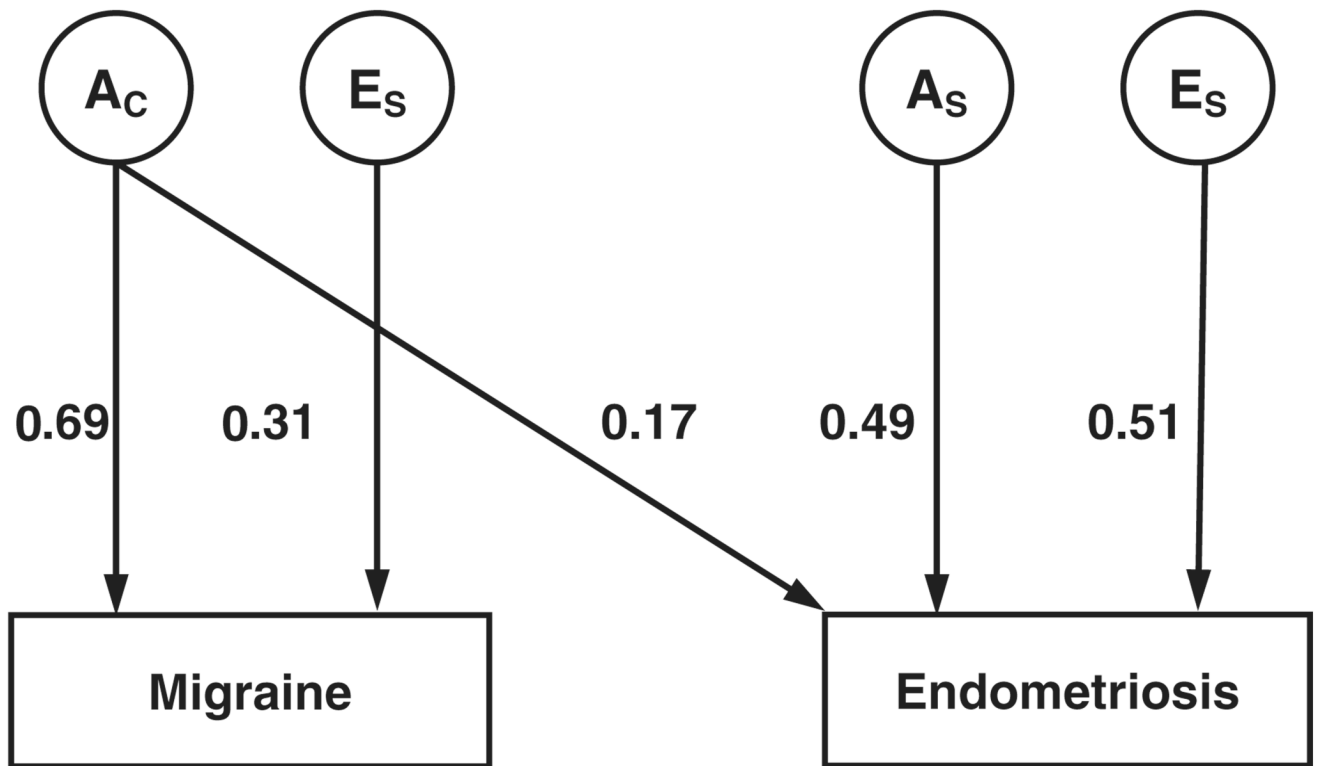


Fig. 1. Path diagram of the best fitting bivariate correlated liability (triangular decomposition or “Cholesky” factor) model. Latent variables are represented by circles and phenotypes are represented by boxes. A_C represents additive genetic influences common to migraine and endometriosis; A_S and E_S are specific additive genetic and non-shared environmental influences. Parameter values represent the standardised proportion of variance in liability accounted for by the latent variables.

TABLE I
Diagnostic criteria for migraine without aura, excerpt from International Headache Society (IHS) classification of headache

Migraine without aura (MO)

- A. At least five attacks fulfilling B through D
 - B. Headache lasting 4–72 h (untreated or unsuccessfully treated)
 - C. Headache has at least two of the following characteristics:
 - 1. Unilateral location
 - 2. Pulsating quality
 - 3. Moderate or severe intensity (inhibits or prohibits daily activities)
 - 4. Aggravation by walking stairs or similar routine physical activity
 - D. During headache at least one of the following:
 - 1. Nausea and/or vomiting
 - 2. Photophobia and phonophobia
-

TABLE II

Prevalence of migraine conditional on endometriosis in ASPs

	Migrainous headache	<i>N</i>	IHS migraine	<i>N</i>
Non-endometriosis	Unaffected	92 (0.56)	Unaffected	114 (0.70)
	Affected	72 (0.44)	Affected	50 (0.30)
Endometriosis	Unaffected	344 (0.45)	Unaffected	482 (0.63)
	Affected	423 (0.55)	Affected	285 (0.37)

TABLE III

Twin-pair cross-tabulations^a

Zygoty Group	Twin 1	Twin 2	r _t (95% CI)
MZ	Endometriosis	–	+
		713	38
		47	17
		410	22
DZ	+	23	2
<i>Migrainous IHS headache</i>			
MZ	Endometriosis	–	+
		444.5	311
		28	31.5
		236	196.5
DZ	–	14.5	10
MZ	Migrainous headache	348	121
		128	218
		155	91
		100	111
DZ	–	155	91
MZ	Endometriosis	–	+
		619.5	136
		43.5	16
		330	102.5
DZ	–	19.5	5
MZ	Migrainous IHS headache	594	67
		71	83
		280	70
		69	38
DZ	–	280	70
MZ	Endometriosis	–	+
		619.5	136
		43.5	16
		330	102.5
DZ	–	19.5	5
MZ	Migrainous IHS headache	594	67
		71	83
		280	70
		69	38
DZ	–	280	70
MZ	Endometriosis	–	+
		619.5	136
		43.5	16
		330	102.5
DZ	–	19.5	5
MZ	Migrainous IHS headache	594	67
		71	83
		280	70
		69	38
DZ	–	280	70

^aCounts may not be a whole number for endometriosis × migraine cross-tabulations, as tables were averaged over using either Twin 1 or Twin 2 as proband.