

Optimizing infectious disease interventions during an emerging epidemic

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The emergence and global impact of the novel influenza A(H1N1)v highlights the continuous threat to public health posed by a steady stream of new and unexpected infectious disease outbreaks in animals and humans. Once an emerging epidemic is detected, public health authorities will attempt to mitigate the epidemic by, among other measures, reducing further spread as much as possible. Scarce and/or costly control measures such as vaccines, anti-infective drugs, and social distancing must be allocated while epidemiological characteristics of the disease remain uncertain. Here we present first principles for allocating scarce resources with limited data. We show that under a broad class of assumptions, the simple rule of targeting intervention measures at the group with the highest risk of infection per individual will achieve the largest reduction in the transmission potential of a novel infection. For vaccination of susceptible persons, the appropriate risk measure is force of infection; for social distancing, the appropriate risk measure is incidence of infection. Unlike existing methods that rely on detailed knowledge of group-specific transmission rates, the method described here can be implemented using only data that are readily available during an epidemic, and allows ready adaptation as the epidemic progresses. The need to observe risk of infection helps to focus the ongoing planning and design of new infectious disease surveillance programs; from the presented first principles for allocating scarce resources, we can adjust the prioritization of groups for intervention when new observations on an emerging epidemic become available.

human influenza | pandemic | immunization | school closure | mathematical model

The need to plan for countering new emerging diseases is highlighted by the worldwide rise of HIV infections since its discovery in 1981 (1, 2), by the spread of the foot-and-mouth epidemic in the United Kingdom in 2002 (3), by the global impact of several outbreaks of severe acute respiratory syndrome (SARS) in 2003 (4), by the projected impact of a possible new influenza pandemic (5–7), and recently by the actual spread of a novel influenza A(H1N1)v pandemic (8, 9). Available control measures for new infections will be scarce because of supply (e.g., vaccines, drugs, masks) and logistical (e.g., distribution) constraints, and their use will be costly. As a consequence, public health bodies face the questions of how to deploy limited control measures to minimize transmission, and of which groups in the population should be targeted for infectious disease control (10, 11). The general problem is how to choose groups of the population that should receive priority in getting the intervention. Typically groups are defined by age, but for some diseases sex, occupation, or other demographic characteristics are more salient. Hereafter we refer to age groups, for brevity.

Existing approaches to allocating infection control rely either on detailed knowledge of transmission parameters (12–15) that may remain ambiguous, as standard methods for estimating transmission rates among groups suffer from indeterminacy (16–18), or on predictions for the eventual number of infections that occur in each group during the entire epidemic (16, 19–21),

which are unlikely to be available at the start of an emerging epidemic. Here we provide a robust solution to the allocation problem that applies to a general class of infectious diseases for which the objective is to minimize transmission and at-risk contacts are reciprocal, that is, for which spread of infection requires the proximity of two individuals (17). We show that there exist simple principles to find optimal allocation schemes for scarce control measures that require observation of only a few key risk measures of infection that can be observed in the initial phase of an emerging epidemic.

Results

Vaccination. Suppose that we wish to target vaccination of susceptible individuals to minimize transmission of an infection. A measure of the transmission potential is the reproduction number R , defined as the number of secondary infections caused by a typical primary case (22–25); this monotonically affects the rate of increase in the number of cases (26). We find that the marginal benefit of allocating a dose of vaccine to a given age group i is approximately proportional to the product of the incidence rate per person, denoted by x_i/n_i , and the force of infection, denoted by x_i/s_i , where force of infection is defined as incidence rate per susceptible person (22). If contact reciprocity holds and all else is equal, this implies that the greatest reduction in transmission of the infection population-wide can be achieved by vaccinating a person in the group with the highest product of incidence and force of infection. More generally, this reduction depends on the efficacy of the vaccine in each group q_i , the per contact probability of becoming infected for each group a_i , and the per contact infectiousness of each group c_i :

$$\frac{1}{R} \frac{dR}{du_i} \approx -hq_i \frac{c_i}{a_i} \frac{x_i(t)}{s_i(t)} \frac{x_i(t)}{n_i} \quad [1]$$

[relative change in transmission] \approx [constant] [vaccine efficacy] [per contact probability of transmitting infection/ per contact probability of becoming infected] [force of infection] [incidence of infection].

The relative change in transmission depends on the product of two measures for risk infection, which implies that small differences between groups in risk of infection could hint at substantial benefits for targeting specific groups.

Intuitively, the change in transmission by vaccinating one susceptible individual has two components: first, the risk of infection of this individual; and second, the number of resulting infections that this individual will cause once infected. The first

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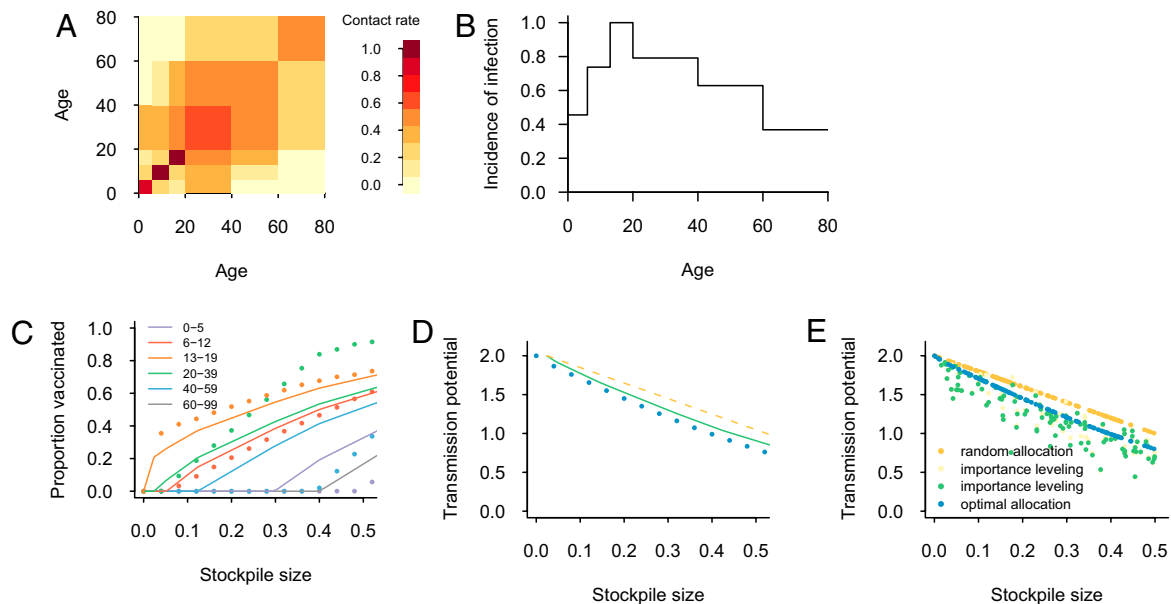


Fig. 2. A test of the importance-leveling scheme against simulated data. (A) The 21 contact parameters required to describe reciprocal contacts among six age groups, based on self-reported number of social contacts (17). (B) Incidence of infection during the initial phase of the epidemic, as simulated from the contact parameters in panel A. (C) Allocation of a perfect vaccine over six age groups for different stockpile sizes, according to the importance leveling scheme (lines) that uses only information about the incidence of infection as in B; for comparison, we show the optimal allocation that requires knowing the entire contact matrix (dots). (D) Reduction in transmission potential by importance leveling (green line) is indistinguishable from maximal reduction by optimal allocation (blue dots) for stockpile sizes up to 20% of the total population. Both are considerably better than random allocation (orange line). (E) Sensitivity analysis of reduction in transmission potential to age-specific variation in per contact probability of acquiring infection, a_i , and per contact probability of transmitting infection, c_i . Importance leveling was applied while ignoring the variation in a_i and c_i (yellow) and while accounting for this variation (green) (*Methods*). Transmission parameters are scaled such that the largest value equals 1 in A; incidence of infection is scaled such that the largest value equals 1 in B. Size of stockpile is expressed relative to total population size; transmission potential is scaled such that it equals 2.0 when stockpile size is zero in C, D, and E.

allocation for a range of vaccination coverages, reproduction numbers, for different contact patterns, and with uncertain estimates of risk of infection (*Methods* and *SI Text*).

Applications to Pandemic Influenza. Practical application of this approach depends on the existence of large and detectable differences between groups in risk of infection in a real emerging epidemic. Reconstruction of the time course of the 1957–1958 influenza A(H2N2) pandemic shows (Fig. 3) that both incidence and force of infection were significantly greater in persons less than 20 years of age 2 weeks before the first pandemic peak (the week of 18 September 1957), indicating that interventions in this group would have been most effective early on. After the first peak, the group-specific risks of infection become more similar. Two weeks before the second peak (the week of 29 January 1958), the age groups revealed no measurable differences in incidence, and differences in force of infection were modest and not statistically significant, suggesting that at least for social distancing and possibly for vaccination, there would have been no benefit in targeting the 0- to 19-year-old group during the second pandemic wave. Two results emerge: first, the current best target group can be identified through observing differences between group-specific risk of infection during emerging epidemics; and second, the best target group for intervention may change during the course of an epidemic.

The principles for allocating scarce control measures are immediately applicable in the influenza A(H1N1)v pandemic. Age-specific incidence has been reported for several countries, including the United States and Chile. In these two countries, the highest incidence during the initial phase occurred among the 0- to 19-year-olds, higher by a factor of at least 4 compared with that in 30- to 39-year-olds (Fig. 4) (34). From these observations

we would conclude, assuming no difference between groups in per contact probability of transmitting infection, that the highest-priority age groups for social distancing in the initial phase would have been the 0- to 19-year-olds and that differences between groups in expected reduction of transmission after targeting social distancing can be greater than 16-fold.

The observed age-specific incidence for influenza A(H1N1)v also suggests that in the initial phase the highest priority age groups for vaccination would have been the 0- to 19-year-olds, assuming no difference between groups in proportion of those susceptible to infection. A striking feature of the age-specific incidence is the relative sparing of older age groups. Older age groups are spared because they have lower levels of social contact with infectious age groups, or because of preexisting immunity to symptomatic infection. To refine estimates of the benefits of group-specific vaccination, which depend on the proportion of susceptible individuals in each group, one would need estimates of the proportion of individuals who are immune. This could be provided by serological surveys of the population before vaccination, or by keeping track of the number of infections per group. If a serological survey finds that there is indeed preexisting immunity among older age groups, this leads to a counterintuitive consequence that the score for importance for prioritization of these older age groups would increase: for a given number of cases, a higher level of immunity implies a higher value for force of infection, which in turn implies a larger expected reduction in transmission if the older groups are targeted for vaccination.

Discussion

Ongoing comparisons of incidence by age are an appropriate way to quantify how target groups change as the epidemic progresses

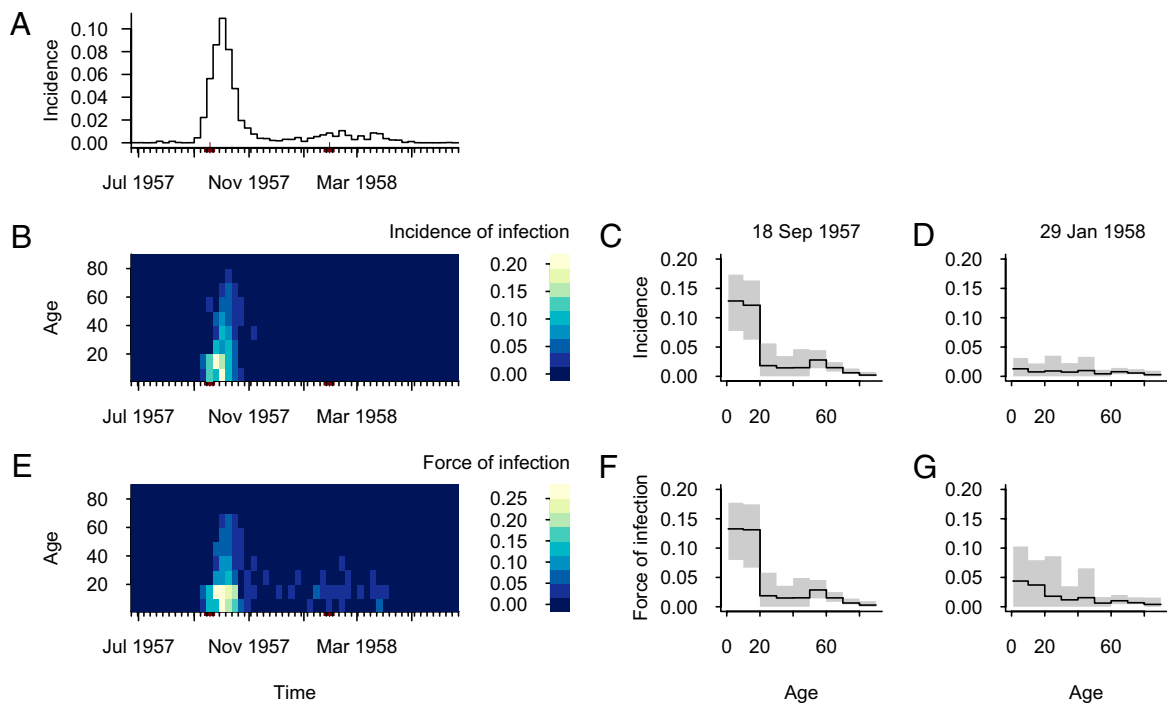


Fig. 3. Incidence and force of infection during the “Asian” 1957–1958 influenza A(H2N2) pandemic in the Netherlands, as reconstructed from age-specific records of mortality and from serological cross-sectional surveys conducted in June 1957 and June 1958. (A) Time course of overall weekly incidence of infection. (B) Time course of weekly incidence of infection by age group. (C) Age-specific incidence of infection in week of 18 September 1957, 2 weeks before peak of first pandemic wave, and (D) in week of 29 January 1958, 2 weeks before peak of second pandemic wave. (E) Time course of weekly force of infection by age groups. (F) Age-specific force of infection 2 weeks before peak of first pandemic wave, and (G) age-specific force of infection 2 weeks before peak of second pandemic wave. Gray areas in C, D, F, and G indicate 95% bootstrapped confidence intervals (*Methods*).

and more persons become immune. For the influenza A(H1N1)v pandemic, interventions targeted at school-aged children, such as school closures, should be most effective in the early stages; but, as in 1957, such interventions may be considerably less effective as the epidemic progresses. In particular, if the incidence becomes comparable among children and adults, such measures will be significantly less valuable than they would have been at the start of the pandemic.

Our results suggest that the key risk factors for infection should be used to define a population structure that is relevant for monitoring incidence of infection and for optimal targeting interventions. For newly emerging infections, this is an important result: elucidating the risk factors for infection helps in deciding whether we should partition the population by age, sex, occupation, or other demographic characteristics before we set out to

monitor risk of infection per group and identify the best target group for infection control.

The principles for identifying target groups apply immediately when the objective of control is to reduce transmission. Simulations suggest that allocation strategies that reduce transmission also tend to do better than random allocation in minimizing the peak incidence and minimizing the total number of infections (Fig. S1). Therefore, these principles may also apply to a wider range of objectives, such as the reduction of the total number of infections in the population or the reduction of the total number of severe cases. For this wider range of objectives, there is evidence that small supplies of vaccines should target high-risk groups, whereas larger supplies should target groups who are major transmitters (13–15, 19, 21, 28). The contribution of this paper is to define how to identify the major transmitters, given limited data. For the objectives other than reducing transmission, the question remains open at what level of supplies of vaccine there is a switch from targeting high-risk groups to targeting major transmitters.

Before targeting interventions at the groups with the highest risk of infection, the underlying assumptions and conditions of this simple rule should be checked and verified. In absence of data on incidence of infection by location and household size, we have assumed here that transmission occurs according to mass-action-type dynamics, thus ignoring the subtlety of network interactions, such as the intense and repeated nature of contacts within households. The analysis assumes that at-risk contacts for transmission are reciprocal. Reciprocal contacts arise, for example, whenever spatial and temporal proximity of two individuals is required for transmission to occur. This condition is not strongly violated by a broad class of diseases such as influenza. The analysis requires furthermore that the distribution of risk of infection over groups should not change much before or

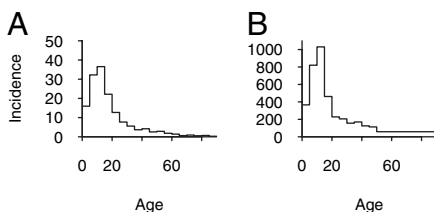


Fig. 4. Incidence by age during first phase of outbreak of A(H1N1)v, as cases per million. (A) Incidence in United States up to 13 May 2009. A total of 3,369 confirmed and probable cases have been reported. Incidence was calculated using 3,097 case patients for which age was reported or could be calculated using date of birth and who did not report a recent history of travel from Mexico. (B) Incidence in Chile up to 21 June 2009. A total of 5,186 confirmed cases have been reported. Incidence was calculated using 5,085 confirmed cases for which age was known (34).

during the observation interval, and the number of infections should be large enough to obtain reliable estimates of risk of infection. Our stochastic simulations have shown that there is only a small risk that stochastic variation in the epidemic will mislead the analyst. A more salient problem is bias in group-specific incidence, when detection may be heavily biased toward severe cases (29). It is important for users of our approach to base estimates of incidence and force of infection on the best possible proxy for infection, and to consider biases introduced by the surveillance system.

In conclusion, we have described a robust strategy to target infectious disease interventions in the face of uncertainty about the precise epidemiological characteristics of a newly emerging infection. If all else is equal, priority for vaccination goes to groups with the highest product of incidence and force of infection, and priority for social distancing goes to groups with the highest incidence of infection. Real-time monitoring of group-specific risk differences during an emerging epidemic will be crucial to determine the optimal targets of an intervention with scarce resources.

Methods

The 1957–1958 “Asian” Influenza A(H2N2) Epidemic in The Netherlands. We obtained data on mortality with influenza as a primary cause in the Netherlands from 26 June 1957 to 25 June 1958 from Statistics Netherlands. The number of deaths was reported weekly and stratified by 10 year age cohorts (1–9 years, 10–19 years, 20–29 years, . . . , 70–79 years, 80 years and older). The number of persons in 1957 in these age groups was also obtained from Statistics Netherlands. The age-specific proportions of susceptible individuals were estimated from a cross-sectional serological study conducted in June 1957, before the first pandemic wave arrived (30), and from a follow-up study conducted in June 1958, after the second pandemic wave had passed (31). The age-specific risk of death upon infection is obtained by dividing the number of deaths due to influenza by the number of influenza infections. As the time lag between moment of death and moment of infection was short for all age groups (32), we obtain the number of infections for any week in each age group by dividing the number of deaths in that week by the risk of death upon infection, and we obtain the number of susceptible individuals for any week by subtracting the cumulative number of infections until that week from the initial number of susceptible individuals as observed in June 1957. The weekly incidence is calculated as the number of infections divided by population size, and the weekly force of infection is calculated as the number of infections divided by the number of susceptible individuals. To assess the uncertainty in the outcome, we constructed 100 bootstrapped data sets for the serological surveys, and repeated the calculations for each data set to obtain 95% bootstrap intervals.

Outline of Derivation of Eqs. 1 and 2. To specify the precise relationship between the expected change in the reproduction number R and the magnitude of a targeted intervention, we study the transmission of an infection in a host population that is stratified into m groups (key to notation in Table S1). The transmission of infection from group j to group i is quantified by k_{ij} , the mean number of individuals in group i that are infected by a single individual in group j during its entire infectious period. The matrix with elements k_{ij} is the reproduction matrix \mathbf{K} . A targeted intervention, such as vaccination or social distancing, will change the values of one or more elements of this reproduction matrix \mathbf{K} ; this change in the reproduction matrix is denoted by $d\mathbf{K}$. The reproduction number R is defined as the top eigenvalue of the reproduction matrix \mathbf{K} , and it gives the number of secondary infections produced by a “typical” infective in the stratified population (23). In general, it is not possible to compute the change in the top eigenvalue dR from a change in the reproduction matrix $d\mathbf{K}$ when the precise values of all matrix elements k_{ij} are unknown. However, for a broad

class of infectious diseases, the reproduction matrix \mathbf{K} has a special structure that does allow for an explicit derivation of the change in the top eigenvalue dR , even when the values of k_{ij} are unknown. The derivation proceeds in four steps. First, when contacts are reciprocal, the reproduction matrix \mathbf{K} can be factorized into a product of symmetric matrices. Specifically, we can write $k_{ij} = s_i a_i b_{ij} c_j$. From right to left, c_j reflects the per contact probability of transmitting infection; $b_{ij} = b_{ji}$ reflects the contact parameter (defined here as the proportion of individuals in group i contacted by an individual in group j during its entire infectious period); a_i reflects the per contact probability of becoming infected; s_i reflects the number of susceptible individuals. Second, because of this structure, the normalized top left eigenvector \mathbf{v} of the reproduction matrix is directly related to the normalized top right eigenvector \mathbf{w} of the reproduction matrix. Specifically, the i th element of the top left eigenvector \mathbf{v} can be related to the i th element of the top right eigenvector \mathbf{w} as $v_i \sim c_i w_i / s_i a_i$. Third, the normalized top right eigenvector can be related to the group-specific number of new infections $\mathbf{x}(t)$: the i th element of the top right eigenvector \mathbf{w} is proportional to the group-specific number of new infections in group i as observed around time t : $w_i \sim x_i(t)$. Fourth, the change in top eigenvalue dR is given by the matrix product of the normalized top left eigenvector \mathbf{v} , the change in the reproduction matrix $d\mathbf{K}$, and the normalized top right eigenvector \mathbf{w} (33). By multiplying the left eigenvector, the change in the reproduction matrix (expressed in terms of either $R q_i d u_i / n_i$ when vaccination is targeted at group i , or $a_i c_i s_i d p_{ij} / n_i^2$ when social distancing is targeted at group i), and the right eigenvector we obtain the Eqs. 1 and 2 for the change in reproduction number R (full derivation and conditions in SI Text).

Sensitivity Analysis. We explored the sensitivity of reduction in transmission potential R to 4-fold variation in per contact probability of acquiring infection, a , and per contact probability of transmitting infection, c . 100 parameter sets were generated using a contact matrix as shown in Fig. 2A and Table S2. For each parameter set, the per contact probabilities of acquiring infection, a_i , and the per contact probabilities of transmitting infection, c_i , were drawn independently for each group from a uniform distribution on the interval 0.25–1. In each parameter set, a fixed quantity of vaccine was assumed to be available, randomly chosen between 0 and 0.5 of the population size. The results are shown in Fig. 2E. Initial incidences of infection were simulated, and the importance leveling algorithm was used to calculate the allocation of the vaccines over the age groups, as if we were ignorant of the existing variation in a_i and c_i (Fig. 2E, yellow dots). Subsequently, the importance leveling algorithm was applied again with knowledge of the age-specific variation in a_i and c_i (Fig. 2E, green dots). The difference in the reduction of transmission potential between both approaches is very small. For comparison, the outcome is shown for random allocation (Fig. 2E, orange dots) and the optimal allocation that minimizes the transmission potential, using simulated annealing and taking advantage of the full information on the entire transmission matrix and parameter set (Fig. 2E, blue dots). These results show that the performance of the importance leveling scheme is robust to ignoring existing variation in the per contact probability of acquiring infection and the per contact probability of transmitting infection.

In addition, we compared the performance of the importance leveling scheme against random allocation and optimal allocation in a simulation study for a range of vaccine coverages, for different values of reproduction numbers, and for different contact patterns while explicitly allowing for a delay between observing risk of infection and implementing control measures (Fig. S1). We also addressed the uncertainty about group-specific risk of infection (Fig. S2).

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