Smoking as a Risk Factor for Prostate Cancer: A Meta-Analysis of 24 Prospective Cohort Studies

Michael Huncharek, MD, MPH, K. Sue Haddock, PhD, Rodney Reid, MD, and Bruce Kupelnick, BA

Prostate cancer is the most common solid tumor diagnosed among men in the United States, with an estimate of more than 186 000 new cases and 28 000 deaths in 2008. Unfortunately, few risk factors have been identified, other than advanced age and family history.² That environmental factors may play a role in its etiology is suggested by data demonstrating wide international variation in incidence. Migrant studies document increased occurrence among those moving from low- to high-incidence countries, as was observed among Japanese men who immigrated to the United States.³ Autopsy studies have documented a consistent prevalence (15%-30%) of histologic or latent prostate cancer across populations.4 These findings suggest that initiating events for prostate cancer may differ from those contributing to progression and the occurrence of clinically evident disease.

Despite the demonstrated links between smoking and several solid tumors, the association between cigarette smoking and prostate cancer remains a matter of debate. Although this disease is not considered to be tobacco related, ⁵ cigarette smoke is known to contain multiple carcinogens, including *N*-nitroso compounds (recognized animal carcinogens). ^{6,7} An association with smoking could also have a hormonal basis: male smokers were found to have elevated levels of circulating androsterone and testosterone, which may increase prostate cancer risk or contribute to cancer progression. ⁸

Unfortunately, results from human observational studies are inconsistent across study types, with case—control analyses showing particular heterogeneity. In addition, epidemiological analyses suggest that the outcomes of studies examining the influence of smoking on prostate cancer incidence may differ from the results in studies of prostate cancer mortality. This may further complicate analysis of a causal association. Because this topic had not previously been subjected to meta-analysis and results from

Objectives. We evaluated the relationship between smoking and adenocarcinoma of the prostate.

Methods. We pooled data from 24 cohort studies enrolling 21579 prostate cancer case participants for a general variance-based meta-analysis. Summary relative risks (RRs) and 95% confidence intervals (Cls) were calculated separately for mortality and incidence studies. We tested the robustness of effect measures and evaluated statistical heterogeneity with sensitivity analyses.

Results. In the pooled data, current smokers had no increased risk of incident prostate cancer (RR=1.04; 95% Cl=0.87, 1.24), but in data stratified by amount smoked they had statistically significant elevated risk (cigarettes per day or years: RR=1.22; 95% Cl=1.01, 1.46; pack years of smoking: RR=1.11; 95% Cl=1.01, 1.22). Former smokers had an increased risk (RR=1.09; 95% Cl=1.02, 1.16). Current smokers had an increased risk of fatal prostate cancer (RR=1.14; 95% Cl=1.06, 1.19). The heaviest smokers had a 24% to 30% greater risk of death from prostate cancer than did nonsmokers.

Conclusions. Observational cohort studies show an association of smoking with prostate cancer incidence and mortality. III-defined exposure categories in many cohort studies suggest that pooled data underestimate risk. (*Am J Public Health*. 2010;100:693–701. doi:10.2105/AJPH.2008.150508)

epidemiological studies were inconsistent, we pooled data from the available cohort studies to elucidate the possible relationship between smoking and the etiology and progression of adenocarcinoma of the prostate.

METHODS

Our methods are described elsewhere. Our methods are described elsewhere. Our meta-analysis to examine the risk of prostate cancer, both incidence and mortality, associated with cigarette smoking. We prospectively determined eligibility criteria for study inclusion and the specific data elements to be extracted from each published report.

We designed a data extraction form for recording relevant information, with 2 researchers performing data extraction. Differences were resolved by consensus. Other data collected but not included in the eligibility criteria were number of patients and location for each study; length of follow-up; cohort description; type of statistical adjustments, if any, to the individual study odds ratios (ORs) or relative risks (RRs); estimates of smoking dose

(e.g., number of cigarettes smoked or pack years [packs smoked per day×number of years smoked]).

Literature Search

Our methods of literature retrieval are described elsewhere. We conducted a MED-LARS search of English language articles published between January 1966 and February 2007, and a review of CancerLit and the CD-ROM version of Current Contents. We searched the Cochrane database for publications between January 1966 and February 2003. Search terms were *smoking* and *prostatic neoplasms*. For series of articles, all data were retrieved from the most recent article. We also performed hand searches of bibliographies of published reports, review articles, and textbooks. Manual searches included review of studies that did not specify smoking as the primary risk factor analyzed.

The initial electronic and manual searches yielded 290 abstracts, which were screened by a physician-investigator according to our inclusion criteria. Most excluded articles did not report peer-reviewed prospective cohort

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studies, reported on studies with a sample of fewer than 50 participants, reported on in vitro or animal studies, or were literature reviews. Eligibility criteria also included publishing an observational study that enrolled participants with histologically proven adenocarcinoma of the prostate, availability of data on cigarette smoking, availability of ORs or RRs with 95% confidence intervals (CIs) for each report or availability of raw data to calculate these values, and availability of data on the outcome of interest, including incident or fatal prostate cancer.

After the initial screening, we had 34 citations. ^{12–45} Ten reports were excluded because of lack of data on smoking, ^{12–14} small sample size, ^{15–17} lack of specified outcome data (i.e., prostate cancer incidence or mortality), ¹⁸ or unavailability of 95% CIs or raw data for their calculation. ¹⁹ One report dealt only with patients treated for benign prostatic hypertrophy or stage A prostatic tumors and was excluded. ²⁰ Data for a 1980 study ²¹ were updated in 1991 ³⁴; we used data from the latter report for statistical pooling.

Statistical Analysis

Our data analysis followed meta-analytic procedures described by Greenland. 46 For each included study, we derived RRs or ORs reflecting the risk of developing or dying from prostatic adenocarcinoma associated with smoking, followed by calculation of the natural logarithm of the estimated RR for each data set as well as calculation of an estimate of the variance. When both crude and adjusted RRs were provided, we used the most fully adjusted value. We calculated the variance in each study's measure of effect from the 95% CIs. We used ORs or RRs for the highest versus lowest exposure categories. If these measures were missing, we calculated them with standard methods. 10 Whenever possible, we used adjusted outcome measures for statistical pooling.

We calculated a weight for each included report as 1 divided by the variance followed by a summation of the weights. We then calculated and summed the product of the study weight and the natural logarithm of the estimated RR. Finally, we calculated a summary RR and 95% CI.

We performed a statistical test for homogeneity (Q). This procedure tests the hypothesis that the effect sizes are equal in all of the

included studies.¹¹ If Q exceeds the upper-tail critical value of χ^2 (P < .10) at k-1 df, the observed variance in study effect sizes is greater than expected by chance if all studies share a common population effect size. If the studies are not homogeneous, they are not measuring an effect of the same size and calculation of a pooled estimate of effect may be of questionable validity. Explanations for the observed heterogeneity must be sought. Sensitivity analyses or further stratified analyses are then performed according to the magnitude of Q.

We did not examine the potential for publication bias. Publication bias occurs because published studies may not be representative of all studies that have ever been done. The funnel plot method and other statistical tools were constructed to address this issue. Unfortunately, these methods lack firm statistical theoretical support and are not generally recommended for medical applications.¹¹

RESULTS

Our meta-analysis included 24 cohort studies. ^{22–45} Table 1 provides an overview of the entire database. Seventeen reports (71%) were from the United States, ^{25,26,29–37,39–44} 3 from Norway, ^{28,38,45} 2 from Japan, ^{23,24} and 1 each from Sweden ²² and the United Kingdom. ²⁷ Prostate cancer cases totaled 21579.

Each study used 1 of 2 endpoints: incident prostate cancer or prostate cancer mortality. Smoking status was categorized by a binary measure, current versus ever or former, or by a quantitative measure, such as number of cigarettes per day or pack years of smoking (Table 1).

Pooling homogeneous data (P=.131) from 8 cohorts in studies examining the risk of incident prostate cancer among current smokers^{22,25,28,29,38,42–44} gave a summary RR of 1.04 (95% CI=0.87, 1.24; Figure 1).

Because use of the broad exposure categorization of current smoker could mask a dose-dependent effect, if one exists, we then calculated summary RRs for several subgroups of studies that quantified smoking history. Seven cohorts contained information on number of cigarettes smoked per day versus years or pack years of smoking and risk of incident prostate cancer among current smokers. ^{22,25,30,36–38,42} Comparing the highest to the lowest exposure

category yielded an RR of 1.13 (95% CI=1.03, 1.23), consistent with a 13% greater risk of prostate cancer among the heaviest smokers. Combining data only from cohorts that were stratified specifically by number of cigarettes smoked per day^{25,37,38,42} (i.e., 4 of the 7 studies that collected these data) showed a 22% increased risk of incident prostate cancer (RR=1.22; 95% CI=1.01, 1.46).

Homogeneous data (P=.15) on incident cancer risk among current smokers measured in years or pack years of smoking (highest versus lowest) from the 5 cohort studies containing such information 22 .25,30,37,38 were also consistent with a small but significant increase in prostate cancer risk (RR=1.11; 95% CI=1.01, 1.22). As shown in Table 1, these 5 studies used varying cutoff points for their highest smoking exposure categories, ranging from more than 21 years of smoking 22 to more than 55 pack years. 25 These differences could contribute to attenuation of the resultant RRs.

Ten cohort studies also provided data on risk of incident prostate cancer among exsmokers. $^{22,25,28,33,37-40,42,43}$ The RR was consistent with our pooled estimates derived from current smokers: 1.09~(95%~CI=1.02,1.16; Figure 2). Excluding from our analysis a study that analyzed incident and fatal cases together 37 slightly attenuated the outcome (RR=1.07; 95% CI=1.01, 1.14). None of these 10 studies provided quantitative estimates of amount smoked.

Among reports in which prostate cancer mortality was the outcome of interest, 22,26,27,29,31,32,34,35 current smokers had a 17% greater risk of death from prostate cancer than did nonsmokers (RR=1.17; 95% CI=1.10, 1.23). The pooled data were heterogeneous (P=.007), with 1 study accounting for more than half of the observed heterogeneity (data not shown).31 That study was among the oldest in the database, dating from 1958, when the demographic characteristics of smokers in the United States differed from those after 1989, the period during which 22 of our 24 cohort studies were published. Excluding the data from that older study from the pooled analysis eliminated the statistical heterogeneity (RR=1.14; 95% CI=1.06, 1.19; P=.15).

These findings were further supported by summary RRs derived by pooling data from mortality studies that quantified smoking by cigarettes per day (RR=1.30; 95% CI=1.16,

TABLE 1—Smoking and Prostate Cancer Risk: Overview of 24 Cohort Studies

20	2368 (Incident Ex-smoker				RR (95% CI)	Dulation	RR (95% CI)	to KK
15	(00000	Ex-smoker	1.09 (0.98, 1.22) > 25 Cigarettes	>25 Cigarettes	1.00 (0.72, 1.38)	>21 y	1.03 (0.90, 1.19)	Age, BMI, marital
15	cases)			(all smokers)			(Current smokers)	status
15		Current smoker	1.11 (1.01, 1.23)					
15	709 (Fatal	Ex-smoker	1.03 (0.84, 1.33) > 25 Cigarettes	· 25 Cigarettes	1.05 (0.82, 1.35)	>41 y	1.28 (0.99, 1.65)	
15	cases)			(all smokers)			(Current smokers)	
15		Current smoker	1.26 (1.06, 1.50)					
	108 (Fatal	Ever smoker	1.1 (0.7, 1.5) >	> 35 Cigarettes	3.0 (1.0, 7.1)	:		Residence, age,
	cases) _a							observation
								period
33	196 (Incident	Ever smoker	0.80 (0.60, 1.07)	:		:		Age period, city of
	cases)							residence, radiation
								dose, education
11	71 (Incident	Current smoker	2.2 (1.2, 4.4)	≤20 Cigarettes	1.8 (0.7, 4.4)	> 55	2.0 (1.1, 3.8)	Age
	cases)					Pack		
						years ^b		
			^	· 20 Cigarettes	2.7 (1.2, 6.0)	:		
		Ex-smoker	1.2 (0.7, 2.1)					
16	826 (Fatal	Current smoker	1.31 (1.13, 1.52) ≤	≤25 Cigarettes	1.21 (1.01, 1.46)	:		Age, race, serum
	cases)							cholesterol, income,
								diabetes
			ΛI	226 Cigarettes	1.45 (1.19, 1.77)	:		
20	878 (Fatal cases)	Current smoker	0.99 (0.87, 1.34)	<u>:</u>		:		Age, study
		Ex-smoker	0.86 (0.80, 1.12)					
27	dent	Current smoker	1.1 (0.9, 1.3)	:		:		Age at start of smoking,
	cases)							of cigarette smoked, urban/rural residence,
			:					pipe/cigar smoking
I and the second se	11 16 27		7.1 (Incident cases) 826 (Fatal cases) 878 (Fatal cases) 707 (Incident cases)	rases) Ex-smoker 1.2 (0.7, 2.1) 826 (Fatal Current smoker 1.31 (1.13, 1.52) cases) RR (Fatal Current smoker 0.99 (0.87, 1.34) cases) Ex-smoker 0.86 (0.80, 1.12) 707 (Incident Current smoker 1.1 (0.9, 1.3) cases)	rases) Ex-smoker 1.2 (0.7, 2.1) 826 (Fatal Current smoker 1.31 (1.13, 1.52) cases) Rx-smoker 0.99 (0.87, 1.34) cases) Ex-smoker 0.99 (0.87, 1.34) cases) Ex-smoker 0.99 (0.87, 1.34) cases) Ex-smoker 0.90 (0.87, 1.34) cases)	cases) Ex-smoker 1.2 (0.7, 2.1) 826 (Fatal Current smoker 1.31 (1.13, 1.52) ≤ 25 Cigarettes cases) Ex-smoker 0.39 (0.87, 1.34) Ex-smoker 0.86 (0.80, 1.12) 707 (Incident Current smoker 1.1 (0.9, 1.3) cases) Ex-smoker 0.86 (0.80, 1.12) Cases)	cases) 2.0 Cigarettes 1.8 (U.f., 4.4) cases) 2.1 (1.2, 6.0) Ex-smoker 1.2 (0.7, 2.1) 826 (Fatal Current smoker 1.31 (1.13, 1.52) ≤ 25 Gigarettes 1.21 (1.01, 1.46) cases) Ex-smoker 0.99 (0.87, 1.34) Ex-smoker 0.86 (0.80, 1.12) 707 (Incident Current smoker 1.1 (0.9, 1.3) cases) Ex-smoker 0.9 (0.7, 1.1)	cases) cases) cases) cases) Aurient smoker 1.2 (1.2, 4.4) 2.0 Cigarettes 2.7 (1.2, 6.0) Ex-smoker 1.2 (0.7, 2.1) 826 (Fatal Current smoker 2.8 (0.87, 1.34) 2.9 Cigarettes 1.9 (1.01, 1.46) 2.0 Cigarettes 1.1 (1.01, 1.46) 2.2 Cigarettes 1.2 (1.01, 1.46) 2.2 Cigarettes 1.3 (1.13, 1.52) ≤ 25 Gigarettes 1.45 (1.19, 1.77) 2.6 Cigarettes 1.45 (1.19, 1.77) 2.707 (Incident Current smoker 1.1 (0.8, 1.12) 2.2 Cigarettes 1.2 (1.01, 1.46) 3 2.2 Cigarettes 1.3 (1.13, 1.52) ≤ 25 Gigarettes 1.45 (1.19, 1.77) 3 3 3 3 4 5 6 6 6 7 7 6 7

TABLE 1—	TABLE 1—Continued									
Giovannucci et al. ²⁹ (2007)	Health Professionals Follow-Up Study, 51529 male health professionals in the United States, initiated in 1986	16	3544 (Incident cases)	3544 (Incident Current smoker cases)	0.98 (0.89, 1.07)	:		÷		Smoking history, race, family history, physical activity, BMI, height, energy consumption, Calcium, tomato sauce, and alpha-linoleic acid intake
			312 (Fatal cases)	Current smoker	1.41 (1.04, 1.91)					
Giovannucci et al. ³⁰ (1999)	Health Professionals Follow-Up Study, 51.529 male health professionals in the United States, initiated	∞	1369 (Incident NA cases)	W		÷		Total lifetime pack years	1.07 (0.85, 1.36)	Smoking history, race, family history, physical activity, BMI, height, energy consumption, Calcium, tomato sauce, and
Hammond ³¹ (1958)	in 1986 American Cancer Society cohort, 187 783 White men	ю	134 (Fatal cases)	Ever smoker	1.75 (1.37, 2.19)	:		:		aipna-linoleic acid intake Age
Hammond et al. ³²	American Cancer Society cohort follow-up	7	319 (Fatal cases)	Ex-smoker	1.02 (0.81, 1.28)	:		÷		Age
Hiatt et al. ³³ (1994)	Kaiser Permanente Medical Care Program, 43432 men in the United States, aged > 30 y, enrolled 1978–1985	-	238 (Incident cases)	Ex-smoker	1.1 (0.8, 1.5)	<1 Pack	1.0 (0.6, 1.6)	<u>:</u>		Age, race, education, alcohol
Hsing et al. ³⁴ (1991)	Veterans Administration Cohort, 293916 US veterans who served 1917-1940	26	4607 (Fatal cases)	Ex-smoker	1.13 (1.03, 1.24) > 39 Ggarettes (current smo	> 39 Gigarettes (current smokers)	1.5 (1.20, 1.90)	> 39 y (Current smokers)	1.23 (0.97, 1.56)	Age
Hsing et al. ³⁵ (1990)	Lutheran Brotherhood Cohort Study, 17 633 White men in the United States, aged > 35 v. followed 1966-1986	20	149 (Fatal cases)	Ever smoker [©]	1.18 (1.1, 2.9)	20-29 Cigarettes ^d	1.7 (0.8, 3.5)	÷		Age
LeMarchand et al. ³⁶ (1994)	Hawaii State Department of Health Cohort, 20316 men of various ethnicities (36% Caucasian), 1975-1989	14	198 (Incident cases)	:		Highest quartile ^e	1.0 (0.6, 1.6)	:		Age, ethnicity, income

BLE 1—	TABLE 1—Continued									
Lotufo et al. ³⁷ (2000)	Physicians Health Study, 22 071 men in the United States, aged 40-84 y at baseline (92% White, 49% nonsmokers)	12.5	883 (Incident cases) 113 (fatal cases)	Ex-smoker	1.14 (1.00, 1.30) <20 Cigarettes (current smo	20 Cigarettes (current smokers)	1.10 (0.78, 1.55)	≥40 Pack years (incident cases)	1.18 (0.95, 1.46)	Age, aspirin/beta carotene assignment, BMI, height, physical activity, alcohol use
					^	>20 Cigarettes (current smokers)	1.10 (0.84, 1.44)	≥40 Pack years	0.91 (0.47, 1.75)	
Lund Nilsen et al. ³⁸ (2000)	National Health Screening Service, 22895 men in Norway, aged≥40 y, enrolled 1984-1986	6.	644 (Incident cases)	Current smoker	0.96 (0.78, 1.19) > 15 Cigarettes	-15 Cigarettes	1.27 (0.91, 1.76)	Z5 Pack years	1.22 (0.93, 1.60)	Age, cigarettes/d, alcohol consumption, leisure time physical activity index, marital status, occupation
Mills et al.	Adventist Health Study, 14000 men in the United States,	6.0	180 (Incident cases)	Ex-smoker Ex-smoker ^f	0.98 (0.80, 1.19) 1.24 (0.91, 1.67)	÷		÷		Age
Putnam et al. (2000)	Retrospective cohort study, 1572 men in lowa,	Up to 9.0	56 (Incident cases)	Ex-smoker ^g	1.4 (0.9, 2.3)	÷		÷		Age
Rodriguez et al. ⁴¹ (1997)	Cancer Prevention Study II, 450.279 men in the United States, aged≥30 y, enrolled in 1982	9.0	1748 (Fatal cases)	Ever smoker Ex-smoker Current smoker	1.02 (0.92, 1.14) 0.99 (0.87, 1.12) 1.34 (1.16, 1.56) > 20 Cigarettes	 20 Cigarettes	1.25 (1.00, 1.57)		1.26 (1.04, 1.53)	Age, race, education, family history, vasectomy, exercise, BMI, alcohol use, vegetable/meat intake
Rohrmann et al. ⁴² (2007)	102 Men from 2 private censuses in Washington County, MD, 1963 and 1975	15 and 19	498 (Incident cases)	Current smoker (1963 cohort) Ex-smoker (1963 cohort) Current smoker (1975 cohort)	1.00 (0.63, 1.59) ≥ 20 cigarettes (1963 cohort 1.33 (0.85, 2.10) (1975 cohort (1975 cohort	2 20 Cigarettes (1963 cohort) 2 20 Cigarettes (1975 cohort)	1.01 (0.65, 1.57)		foor took of the foot	Age
			424 (Fatal cases)	Ex-smoker (1975 cohort) Current smoker (1963 cohort) Ex-smoker (1963 cohort)	1.04 (0.80, 1.36) 0.93 (0.67, 1.29) ≥ 20 Cigarettes (1.01 (0.70, 1.46)	≥ 20 Cigarettes (1963 cohort)	0.95 (0.62, 1.47)			

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				Current smoker (1975 cohort)		≥ 20 Cigarettes (1975 cohort)	1.58 (0.94, 2.64)		
				Ex-smoker (1975 cohort)	1.02 (0.69, 1.50)				
Severson	7999 Men of Japanese	21	174 (Incident	174 (Incident Current smoker 0.87 (0.61, 1.23)	0.87 (0.61, 1.23)	:		:	Age
et al.	ancestry in Hawaii,		cases)						
(1989)	born 1900-1919,								
	enrolled 1965-1968								
				Ex-smoker	0.89 (0.61, 1.29)				
Thompson	Lipid Research Clinics	14	54 (Incident	Current smoker	1.06 (0.62, 1.80)	:		:	None
et al.	Prevalence Study,		cases)						
(1989)	6110 adults from								
	Rancho Bernardo, CA,								
	1972-1974								
Veierod	25 708 Men in Norway,	15	72 (Incident Ex-smoker ^h	Ex-smoker ^h	0.6 (0.3, 1.1)	≤ 10 Cigarettes	0.5 (0.3, 1.1)	:	Age
et al. ⁴⁵	aged 16-56 y, enrolled		cases)			(current smoker)			
(1997)	1977-1983 via a health								
	screening								
						≥ 11 Cigarettes	0.6 (0.3, 1.2)		
						(acilomo tacasio)			

Note. RR=relative risk; Cl = confidence interval; BMI = body mass index; NA= not applicable. Ellipses indicate data that were not included in study.

This study had 12 357 fatal cases from all cancers.

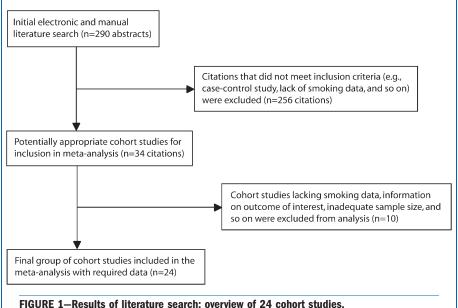
Classified smokers as "ever used tobacco."

^oUpdate published in 2007 for same cohort; 1999 study was included for total pack years risk data.

⁴Highest intake category of tobacco by cigarettes per day had only 3 prostate cancer deaths, making risk estimates questionable. RRs shown are for second highest intake, 20 to 29 cigarettes per day not specified in article; data are presented for highest quartile.

⁶Unly data on former smokers was used because there were only 3 cases of prostate cancer among current smokers.

⁴Rumber of cases for current smokers of fewer than 20 cigarettes per day and 20 or more cigarettes per day were both fewer than 10 participants; therefore only data on former smokers were included in the meta-analysis. ¹⁹Based on data from 11 patients. ¹⁹Based on data from 11 patients. ¹⁹Based on data from 14 patients.



1.46: P=.20)^{22,23,26,34,35,37,41,42} or pack years (RR=1.24; 95% CI=1.09, 1.40; $P=.83)^{22,34,37,41}$: these data were homogeneous. The available data on ex-smokers and prostate cancer mortality were heterogeneous (P<.001) and could not be statistically pooled.

DISCUSSION

Pooled data from 24 cohort studies enrolling more than 26000 participants with prostate cancer showed a modest, although consistent, 9% to 30% increase in both incident and fatal prostate cancer associated with smoking (Table 2). Former smokers had the smallest increase in prostate cancer risk: their risk of incident tumors was 9% higher than that of nonsmokers. Because former or ex-smoker is such a broad, nonquantitative measure of exposure, it is reasonable to assume that this RR represents an underestimate of the strength of the true underlying association. This may also be true of the summary estimates of effect reflecting the effect of current smoking on incident and fatal prostate tumors. The higher RRs calculated from incidence or mortality data for current smokers in studies that used quantitative measures of exposure support this contention.

Previous published data suggested no clear evidence of a causal relationship between

smoking and prostate cancer development.⁵ Studies that used cancer incidence as the endpoint had inconsistent outcomes. Mortality studies also provided little evidence of an association.⁵ By contrast, our meta-analysis revealed a statistically significant and consistent increase in prostate cancer incidence as well as an increased risk of death from this disease with increased smoking.

TABLE 2-Meta-Analysis Overview: 24 **Cohort Studies**

Risk Category	Studies, N	lo. RR (95% CI)
Incident prostate canc	er	
Current smoker	8	1.04 (0.87, 1.24)
Current smoker ^a	4	1.22 (1.01, 1.46)
Current smoker ^b	5	1.11 (1.01, 1.22)
Ex-smoker	10	1.09 (1.02, 1.16)
Fatal prostate cancer		
Current smoker	7	1.14 (1.06, 1.19)
Current smoker ^a	8	1.30 (1.16, 1.46)
Current smoker ^c	4	1.24 (1.09, 1.40)
Ex-smoker ^d	6	

Note. RRs = summary relative risk; CI = confidence interval. Ellipsis indicates RR not calculated. ^aCigarettes per day, highest versus lowest. ^bYears, pack years, highest versus lowest. ^cPack years, highest versus lowest.

^dOnly study that was not homogeneous.

It is possible that the sample sizes in many previous studies were too small to detect an effect. By contrast, our meta-analysis combined information on more than 21000 prostate cancer cases. The validity of our findings was further supported by the elevated risk of prostate cancer we observed among former smokers, which was lower than for current smokers and higher than for nonsmokers. It is likely that the use of ill-defined smoking categories, as well as failure to update smoking status over time. 28,36 contributed to attenuating. and therefore masked the modest association we observed in our pooled data.

The studies we analyzed varied in their outcome measures, so it is important to interpret their results in context. For example, cancer incidence and mortality address somewhat different issues. Incidence data likely reflect an effect of smoking on disease etiology or initiation; observational studies that take prostate cancer mortality as an endpoint only indirectly provide insight into this relationship.

Mortality studies may reflect the effect of smoking on tumor progression by various proposed mechanisms, for example, smoking may increase serum estrogen metabolites that have been postulated to induce a more aggressive tumor phenotype and thereby increase prostate cancer death. 42 Other investigators suggest that smoking may cause mutation of the p53 tumor suppressor gene, creating another pathway to an aggressive tumor phenotype and increased mortality. 30 Whether smokers differ from nonsmokers in the type of therapy received, which in turn could influence survival, is also unknown.9

We found that former smokers have increased risk for prostate cancer, which does not support the suggestion that smoking is related to poorer survival during treatment.³⁴ The increasing gradient in risk across higher exposure categories suggests a biological relationship. Although the exact mechanism underlying the positive association between smoking and prostate cancer death is unknown, it is clear that data from incidence and mortality studies should be considered separately.

Despite these caveats, the results of our analysis of pooled data were quite consistent across endpoints. Although the calculated summary estimates of effect showed a modest (weak) effect of smoking on prostate cancer, it is

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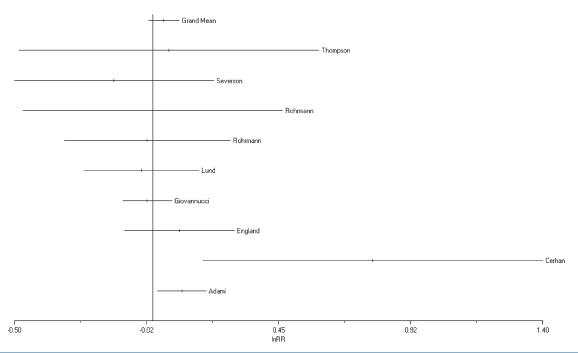


FIGURE 2—Forest plot of pooled results reflecting risk of incident prostate cancer among subjects stratified as current smokers: overview of 24 cohort studies.

likely that several factors contributed to attenuating the association. These included ill-defined smoking status, lack of repeat assessments of smoking status in many cohorts, and the possibility of a screening effect. The effects of screening for prostate cancer could attenuate an association: previous work suggested lower rates of such screening among smokers than nonsmokers. Although stage stratification could help address this issue, we did not find such data in our literature search.

As suggested by Hickey et al., ⁹ future studies should collect data on stage and grade of tumor and on smoking history, including quantity smoked and details of smoking cessation. Although additional work is needed to clarify the relationship between smoking and prostate cancer, our results support a causal association.

About the Authors

Michael Huncharek and Bruce Kupelnick are with the Meta-Analysis Research Group, Columbia, SC. Michael Huncharek is also with the Division of Radiation Oncology, St. Louis VA Medical Center, MO. K. Sue Haddock and Rodney Reid are with the Dorn VA Medical Center, Columbia.

Correspondence should be sent to Michael Huncharek MD, MPH, Meta-Analysis Research Group, 10 Sasanqua Circle, Columbia, SC 29209 (e-mail: info@metaresearchgroup. org; metaresearch@hotmail.com). Reprints can ordered at http://www.ajph.org by clicking the "Reprints/Eprints" link

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Contributors

M. Huncharek originated the study and collected and analyzed the data. K. S. Haddock and R. Reid contributed to the design of the study. B. Kupelnick designed the electronic search strategies and assisted with manual literature searches. All authors contributed to writing the article.

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Human Participant Protection

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