

## The primary divisions of life: a phylogenomic approach employing composition-heterogeneous methods

Peter G. Foster<sup>1</sup>, Cymon J. Cox<sup>2</sup> and T. Martin Embley<sup>3,\*</sup>

<sup>1</sup>Department of Zoology, Natural History Museum, Cromwell Road, London SW7 5BD, UK <sup>2</sup>Centro de Ciências do Mar, Universidade do Algarve, Campus de Gambelas, 8005-139 Faro, Portugal <sup>3</sup>Institute for Cell and Molecular Biosciences, Newcastle University, Newcastle NE2 4HH, UK

The three-domains tree, which depicts eukaryotes and archaebacteria as monophyletic sister groups, is the dominant model for early eukaryotic evolution. By contrast, the 'eocyte hypothesis', where eukaryotes are proposed to have originated from within the archaebacteria as sister to the Crenarchaeota (also called the eocytes), has been largely neglected in the literature. We have investigated support for these two competing hypotheses from molecular sequence data using methods that attempt to accommodate the across-site compositional heterogeneity and across-tree compositional and rate matrix heterogeneity that are manifest features of these data. When ribosomal RNA genes were analysed using standard methods that do not adequately model these kinds of heterogeneity, the three-domains tree was supported. However, this support was eroded or lost when composition-heterogeneous models were used, with concomitant increase in support for the eocyte tree for eukaryotic origins. Analysis of combined amino acid sequences from 41 protein-coding genes supported the eocyte tree, whether or not composition-heterogeneous models were used. The possible effects of substitutional saturation of our data were examined using simulation; these results suggested that saturation is delayed by among-site rate variation in the sequences, and that phylogenetic signal for ancient relationships is plausibly present in these data.

Keywords: universal tree of life; eukaryote origins; archaebacteria; eocyte; heterogeneous phylogenetic models

### **1. INTRODUCTION**

Phylogenetic reconstruction of the earliest diverging lineages in the tree of life is one of the most difficult, but important, problems in evolutionary biology. At present there are two main hypotheses concerning the primary divisions in the tree of life based upon different analyses of molecular sequence data (figure 1). The 'three-domains hypothesis' (Woese et al. 1990) posits that eubacteria (or Bacteria), archaebacteria (Archaea) and eukaryotes (Eukarya) form separate monophyletic groups (domains). The three-domains tree has provided support for theories of eukaryotic origins that have eukaryotes as old as archaebacteria and derived from a common ancestor, sometimes called a neomuran, shared with that group (Cavalier-Smith 2002; Pace 2006). By contrast, the 'eocyte hypothesis' (Rivera & Lake 1992) posits that essential components of eukaryotes branch from within the archaebacteria, sharing common ancestry with a specific group of archaebacteria called the Crenarchaeota (Woese et al. 1990) or eocytes. Current

\* Author for correspondence (martin.embley@ncl.ac.uk).

Electronic supplementary material is available at http://dx.doi.org/ 10.1098/rstb.2009.0034 or via http://rstb.royalsocietypublishing.org. versions of both hypotheses hold that the root of the tree of life is either on the branch separating the eubacteria from the archaebacteria and eukaryotes, in line with rooting studies using ancient paralogous genes (e.g. Baldauf *et al.* 1996; Zhaxybayeva *et al.* 2005), or it lies within the eubacteria based on the polarization of cladistic characters or indels (Cavalier-Smith 2006; Skophammer *et al.* 2007). For the purpose of this paper we also follow the convention of a eubacterial root, while recognizing that there is still a healthy debate about its reliability (Philippe & Forterre 1999; Zhaxybayeva *et al.* 2005; Lake *et al.* 2008, and references therein).

The main evidence for the two competing hypotheses comes from analyses of molecular sequences, often the same ones, at different times and using different methods (Lake 1988; Yang & Roberts 1995; Baldauf *et al.* 1996; Barns *et al.* 1996; Tourasse & Gouy 1999; Brown *et al.* 2001; Katoh *et al.* 2001; Harris *et al.* 2003). It has been suggested that the strongest support for archaebacterial monophyly, and hence the three-domains tree, comes from the simplest methods (Tourasse & Gouy 1999; Katoh *et al.* 2001); the inference being that archaebacterial monophyly might be a phylogenetic artefact of model mis-specification. With this in mind, we recently re-investigated the support for the three-domains tree and the eocyte tree from the small number of genes, typically

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Figure 1. Two views of the tree of life. The root of the tree is often considered to be on the branch leading to the eubacteria (e.g. Baldauf *et al.* 1996) or within the eubacteria (Cavalier-Smith 2006; Skophammer *et al.* 2007). Under any of those rootings, the three-domains tree has a monophyletic archaebacteria, where Euryarchaeota group with the Crenarchaeota/eocytes. By contrast, the eocyte hypothesis groups the eukaryotes with eocytes making the archaebacteria paraphyletic.

encoding components of the genetic machinery, that are conserved across all the three domains (Cox et al. 2008). These genes have been called the genealogydefining core of genes whose common history dates back to the root of the universal tree (Woese 2002), and it is widely held that their phylogeny reflects the three-domains tree (Woese 2002; Pace 2006; Yutin et al. 2008). For our analyses, we used new phylogenetic models that allow for changing compositions across data (Lartillot & Philippe 2004) and across the tree (Foster 2004), reflecting the observation that compositional heterogeneity of both types is pervasive among molecular sequences (Cox et al. 2008). Analyses using these more sophisticated methods consistently favoured the eocyte tree with important implications for theories of eukaryotic origins (Cox et al. 2008). In the present work, we have extended our previous analyses to include recently sequenced additional Crenarchaeota/eocytes, and we also introduce the node-discrete rate matrix heterogeneity (NDRH) model, which enables heterogeneous substitution rates to evolve across the tree. As part of our analyses we have also investigated model fit and model adequacy and the properties of the data with regard to substitutional saturation. Our results support the recent findings of Cox et al. (2008), that when manifest properties of the data comprising acrossdata or across-tree compositional or rate heterogeneity are taken into consideration, it is the eocyte tree rather than the three-domains tree that is favoured in phylogenetic analyses.

### 2. MATERIAL AND METHODS

Two combined datasets, one rRNA and one protein, were constructed based on the data presented in Cox *et al.* (2008). Seven additional Crenarchaeota/eocytes that had been sequenced since our original analyses

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(Cox et al. 2008) were added to the datasets, namely, Caldivirga maquilingensis, Cenarchaeum symbiosum, Hyperthermus butylicus, Ignicoccus hospitalis, Nitrosopu-Staphylothermus marinus milus maritimus, and Thermofilum pendens, and a total of 12 taxa from the eubacteria, euryarchaeota and eukaryotes were removed to reduce the computational complexity (table 1). In total, there were 35 taxa included in the combined protein dataset and 34 in the combined rRNA dataset (Phytophora ramorum rRNA sequences were unavailable). The protein dataset consisted of 41 proteins, those used by Cox et al. (2008), but excluding chaperonin TCPl subunits 1 ( $\alpha$ ), 3 ( $\gamma$ ), 4 ( $\delta$ ), 7 ( $\eta$ ), which are paralogues of chaperonincontaining TCPl subunit 5 ( $\epsilon$ ). The combined protein dataset was also recoded into Dayhoff groups (Hrdy et al. 2004). Dayhoff recoding defined the following six groups of aminoacids corresponding to the PAM matrix: 1, cysteine; 2, alanine, serine, threonine, proline, glycine; 3, asparagine, aspartic acid, glutamic acid, glutamine; 4, histidine, arginine, lysine; 5, methionine, isoleucine, leucine, valine; 6, phenylalanine, tyrosine, tryptophan. Constant sites and singletons were removed from the datasets as they do not contribute to topological resolution and their composition differs from that of the variable sites in a  $\chi^2$ -test of significance ( $p \approx 0$ ).

Maximum-parsimony (MP) and maximumlikelihood (ML) analyses were conducted in PAUP\* (v. 4.b10; Swofford 2002) and RAXML (v. 2.2.3; Stamatakis 2006), respectively. Bayesian Markov Chain Monte Carlo (MCMC) analyses were conducted in MRBAYES (v. 3.1.2; Ronquist & Huelsenbeck 2003), p4 (v. 0.86; Foster 2004) and PHYLOBAYES (v. 2.3; Lartillot & Philippe 2004). MP bootstrap analyses of the small subunit (SSU) and large subunit (LSU) combined rRNA dataset (1045 sites) were performed with the data in a single partition using 1000 heuristic search replicates with treebisection reconnection (TBR) branch-swapping. Neighbour-joining (NJ) analysis using log-determinant distances (LogDet; Lake 1994; Lockhart et al. 1994) were conducted in PAUP\* with 500 bootstrap replicates. ML bootstrap analyses (500 replicates) were conducted with each rRNA partition modelled separately by the general-time reversible (GTR) plus gamma-distributed among-site rate variation ( $\Gamma$ ) model (labelled GTRGAMMA in RAXML). Treehomogeneous MCMC analysis of the rRNA data was performed in P4 for 2000000 generations with a separate  $GTR + \Gamma$  model for each partition, and with the polytomy prior, and a free amongpartition rate parameter. Covarion model analyses were performed in MRBAYES, with a GTR +  $\Gamma$  applied to each rRNA partition, and the MCMC run for 2 000 000 generations. Bayesian MCMC analysis using the NDCH and the NDRH models was performed using P4. The NDCH (node-discrete composition heterogeneity) model allows different compositions on different branches, and the NDRH model allows different rate matrices on different branches. Ten replicate MCMC runs were performed for each configuration of NDCH and NDRH, for four to six million generations. A prior Table 1. Taxa and data provenance.

taxonomy	taxon	provenance
eubacteria		
Campylobacterales	Campylobacter jejuni	GB: NC_003912
Chlamydiae	Chlamydia trachomatis	GB: NC_000117
Firmicutes	Clostridium acetobutylicum	GB: NC_003030
Gammaproteobacteria	Escherichia coli	GB: NC_000913
Planctomycetes	Rhodopirellula baltica	GB: NC_005027
Alphaproteobacteria	Rhodopseudomonas palustris	GB: NC_005296
Cyanobacteria	Synechocystis sp. PCC6803	GB: NC_000911
Spirochaetes	Treponema pallidum	GB: NC_000919
archaebacteria	· ·	
Crenarchaeota/eocyte	Aeropyrum pernix	GB: NC_000854
Euryarchaeota	Archaeoglobus fulgidus	GB: NC_000917
Crenarchaeota/eocyte	Caldivirga maquilingensis	EMBL: CP000852
Crenarchaeota/eocyte	Cenarchaeum symbiosum	EMBL: DP000238
Crenarchaeota/eocyte	Hyperthermus butylicus	EMBL: CP000493
Crenarchaeota/eocyte	Ignicoccus hospitalis	EMBL: CP000816
Euryarchaeota	Methanococcus jannaschii	GB: NC_000909
Euryarchaeota	Methanosarcina mazei	GB: NC_003901
Euryarchaeota	Methanothermobacter thermautotrophicus	GB: NC_000916
Crenarchaeota/eocyte	Nitrosopumilus maritimus	EMBL: CP000866
Crenarchaeota/eocyte	Pyrobaculum aerophilum	GB: NC_003364
Euryarchaeota	Pyrococcus furiosus	GB: NC_003413
Crenarchaeota/eocyte	Staphylothermus marinus	EMBL: CP000575
Crenarchaeota/eocyte	Sulfolobus solfataricus	GB: NC_002754
Crenarchaeota/eocyte	Thermofilum pendens	EMBL: CP000505
Euryarchaeota	Thermoplasma acidophilum	GB: NC_002578
eukaryotes		
Viridiplantae	Arabidopsis thaliana	TIGR <sup>a</sup>
Mycetozoa	Dictyostelium discoideum	dictyBase <sup>b</sup>
Fungi	Encephalitozoon cuniculi	GB: nr <sup>c</sup>
Mycetozoa	Entamoeba histolytica	TIGR <sup>a</sup>
Diplomonadida	Giardia lamblia	GiardiaDB <sup>d</sup>
Metazoa	Homo sapiens	EMBL-EBI <sup>e</sup>
Euglenozoa	Leishmania major	WTSI <sup>f</sup>
Oomycetes	Phytophthora ramorum	$\mathrm{JGF}^{\mathrm{g}}$
Fungi	Saccharomyces cerevisiae	GB: nr <sup>c</sup>
Bacillariophyta	Thalassiosira pseudonana	$\mathrm{JGF}^{\mathrm{g}}$
Parabasalidea	Trichomonas vaginalis	TIGR <sup>a</sup>

<sup>a</sup>The Institute for Genomic Research, Center for the Study of Biological Complexity.

<sup>b</sup>http://www.dictybase.org/.

<sup>c</sup>ftp://ftp.ncbi.nih.gov/blast/—NCBI non-redundant protein BLAST database.

dhttp://gmod.mbl.edu/perl/site/giardial4.

<sup>e</sup>European Molecular Biology Laboratory—European Bioinformatics Institute. <sup>f</sup>Wellcome Trust Sanger Institute.

<sup>g</sup>Joint Genome Institute.

probability ratio on changing composition vectors or rate matrices associated with nodes was used, as described further in the supplemental materials. CAT model analyses were performed in PHYLOBAYES with a GTR rate matrix and  $\Gamma$  distribution of rates among sites. Two independent CAT model runs were conducted, each >2 000 000 cycles, to check convergence to the same posterior probability distribution.

Combined protein analyses were all performed with the entire dataset (5222 sites standard amino acid coding and 4008 sites Dayhoff-recoded data) treated as a single partition. NJ bootstrap analyses of protein LogDet distances were performed with 1000 replicates in P4. Equally weighted MP analyses were performed in PAUP\* with 500 bootstrap heuristic search replicates with TBR branch-swapping. ML bootstrap analyses (300 replicates) under a WAG (Whelan & Goldman 2001) rate matrix with gamma-distributed amongsite rate variation (PROTGAMMAWAG) were performed in RAXML. Homogeneous MCMC analyses of the combined protein dataset were conducted in MRBAYES under a WAG +  $\Gamma$ , with two independent runs each of 1 000 000 generations. Covarion MCMC analyses in MRBAYES were performed with two independent runs, each with two chains, under a WAG +  $\Gamma$  substitution model, and run for 800 000 generations. MCMC analyses in PhyloBAYES were conducted with the CAT-Poisson model with a  $\Gamma$  distribution of rates among sites. Two independent CAT-poisson MCMC runs were performed, each of  $>3\,000\,000$  cycles, to check for convergence to the

Table 2. Support for the three-do	omains tree and the eocyte	hypothesis from con	mbined SSU and L	SU rRNA genes.
A A		<i>v</i> .		0

	method	model <sup>a</sup>	log-marginal likelihood <sup>b</sup>	Euryarchaeota monophyletic	Eocytes monophyletic	Archaebacteria monophyletic	Eocyte hypothesis
A	MP <sup>c</sup>			94	41	96	<5
В	LogDet-NJ <sup>d</sup>			85	75	31	68
С	ML <sup>e</sup>	$\operatorname{GTR} + \Gamma$		93	80	73	27
D	Bayesian	$\text{GTR} + \Gamma^{\text{f}}$	-21717	100	100	82	18
Е		$GTR + \Gamma^{g}$	-21723	100	96	76	24
F		Covarion <sup>f</sup>	-21493	100	100	99	1
G		$NDCH(2,2)^{h}$	-21373	100	99	60	39
Η		$NDCH(4,4)^{h}$	-21 291	100	95	12	87
Ι		NDCH(2,2),NDRH(2,2) <sup>h,i</sup>	-21288	100	96	14	86
J		$NDCH(4,4), NDRH(2,2)^{h,i}$	$-21\ 221$	99	98	11	88
Κ		CAT <sup>j</sup>	-19948	1	39	0	100

Archaebacterial monophyly implies support for the three-domains tree. Bootstrap support or Bayesian posterior probability is shown as percent.

<sup>a</sup>All models except the CAT model (that is rows C–J) had separate models for the SSU and LSU partitions, and all of those except the ML analysis in row C had free partition rates. All were  $GTR + \Gamma$ -like models, and so for example the covarion model in row F was separate  $GTR + \Gamma$  in each data partition, with the covarion superimposed. All analyses using p4 used the polytomy prior.

<sup>b</sup>Marginal likelihood was estimated from posterior likelihood samples using Equation (16) in Newton & Raftery (1994).

<sup>c</sup>Maximum Parsimony bootstrap, using PAUP\*.

<sup>d</sup>Bootstrap values are from neighbour-joining trees using LogDet distances, using PAUP\*.

<sup>e</sup>Using RAXML-VI-HPC v. 2.2.3.

<sup>f</sup>Using MrBAYES v. 3.1.2.

<sup>g</sup>Tree-homogeneous model using P4.

<sup>h</sup>NDCH model where the numbers in parentheses show the number of composition vectors in the two data partitions using P4.

<sup>i</sup>NDRH model where the numbers in parentheses show the number of independent GTR rate matrices in each data partition using P4. <sup>j</sup>CAT mixture model using PHYLOBAYES v. 2.3.

same posterior probability distribution. Maximum parsimony analyses of the Dayhoff-recoded protein dataset were performed with 1000 bootstrap replicates and equally weighted characters. The Dayhoffrecoded dataset was analysed using a homogeneous  $GTR + \Gamma + NDCH(14)$ , that is, a model with a  $GTR + \Gamma$  rate matrix and 14 composition vectors, with the polytomy prior (Lewis *et al.* 2005). The MCMC was run for 2 000 000 generations. CAT +  $\Gamma$ analyses were also conducted in PHYLOBAYES with the 'dayhoff6' option, and two independent MCMC's run for >8 500 000 cycles.

### 3. RESULTS AND DISCUSSION

# (a) Support for alternative trees of life based on rRNA genes

Historically, it has mainly been conflicting phylogenetic analyses of large and small subunit ribosomal RNA sequences (Lake 1988; Yang & Roberts 1995; Barns et al. 1996; Tourasse & Gouy 1999) that have fuelled the debate over which tree, three-domains or eocyte, is better supported. We therefore analysed concatenated SSU rRNA and LSU rRNA gene sequences using a variety of methods to investigate support for the competing hypotheses. Most analyses showed high support for monophyletic Euryarchaeota and monophyletic Crenarchaeota/eocytes (table 2); these groups are generally, but not universally, considered to be monophyletic. In accord with previous studies, MP recovered high bootstrap support for the threedomains tree (electronic supplementary material, figure S1). ML or Bayesian analysis with the GTR +  $\Gamma$  model also recovered the three-domains tree, but with less support (ML: 73% bootstrap support (BS)

in electronic supplementary material, figure S2, GTR +  $\Gamma$ : 82% Bayesian posterior probability (BPP) in figure 2*a*, 76% BPP in electronic supplementary material, figure S3).

We used the GTR +  $\Gamma$  analyses (figure 2*a*, electronic supplementary material, figure S3) as a baseline and improved the models by accommodating covarion, tree-heterogeneous or data-heterogeneous processes. The improved fit of the models to the data was indicated by an improvement in the log marginal likelihood. Bayes factors are the ratio of the marginal likelihoods of two models, and can be used to compare models: a log(BF) of 5 or more is considered to be a very strong support in favour of the better fitting model (Kass & Raftery 1995). In all cases described here the log(BF) greatly exceeded 5, suggesting that we were adding important aspects to the evolutionary model (table 2).

The dataset we analysed was heterogeneous in composition; a  $\chi^2$ -test for compositional homogeneity over the tree failed for both data partitions ( $p \approx 0.0$  for both SSU and LSU). Compositional heterogeneity over the tree was accommodated with the NDCH model, and the model fit was assessed by posterior predictive simulation (Bollback 2002). The  $X^2$  values for the original data were 434 for the SSU and 839 for the LSU (figure 3, arrows). (Here we distinguish the statistic  $X^2$  sensu Sokal & Rohlf (1981) from the  $\chi^2$ curve that is often used to assess its significance).  $\hat{X}^2$ values from simulations of posterior samples using the  $GTR + \Gamma$  model (table 2, row D) were mostly less than 100 for both partitions (figure 3, black bars), showing that the tree-homogeneous  $GTR + \Gamma$ model did not adequately fit the data. However, a tree-heterogeneous NDCH model using two



Figure 2. Bayesian analysis of combined SSU and LSU rRNA genes. In panel (*a*) the model is  $GTR + \Gamma$ , separate in the two data partitions, with free partition rate. This is the analysis in table 2, row D. Panel (*b*) shows an analysis with the NDCH(4,4) NDRH(2,2) tree heterogeneous model from table 2, row J.

composition vectors for each data partition (table 2, row G; electronic supplementary material, figure S5) adequately modelled the data by this test (figure 3, white bars). Furthermore, use of this model increased the likelihood by 334 log units compared with the  $GTR + \Gamma$  model; this was done by the addition of two extra composition vectors over the  $GTR + \Gamma$ model, for a total of six extra parameters. This model eroded support for the three-domains tree somewhat (60% BPP) and increased support for the eocyte tree (39% BPP). Although the NDCH(2,2) model adequately modelled the composition as shown by posterior predictive simulation, addition of two more composition parameters (table 2, row H, NDCH(4,4)model; electronic supplementary material, figure S6) improved the likelihood by a further 82 log units. This latter analysis recovered a marked increase in support for the eocyte hypothesis (87% BPP) over a monophyletic archaebacteria (12% BPP). By accommodating a covarion process in the model (table 2, row F; electronic supplementary

material, figure S4), the likelihood improved greatly by 224 log units over the GTR +  $\Gamma$  model. With this model, support for the three-domains tree also increased greatly (to 99% BPP); however, the improvement to the likelihood imparted by the covarion was not as great as that for compositional heterogeneity in these data, with contrary support.

The composition and covarion are not the only aspects of heterogeneity over the tree. The NDRH model developed here allows the rate matrix to differ over the tree in the same way that the composition can differ over the tree in the NDCH model. A tree-homogeneous model such as the GTR +  $\Gamma$  model for two data partitions can be considered a NDRH(1,1) model; if we have two rate matrices in each of two data partitions, then we have a NDRH(2,2) model. The NDCH and NDRH models can be used together, and are independent of each other.

The summary results for a NDCH(2,2) + NDRH(2,2) model are shown in row I of table 2, (electronic supplementary material, figure S7), which



Figure 3. Assessment of model composition fit by posterior predictive simulation for the rRNA analysis. The test quantity was  $X^2$  (sensu Sokal & Rohlf 1981), the statistic used in  $\chi^2$  tests. Data sets were simulated based on samples from the posterior distribution and for each the  $X^2$  was calculated. Black bars show the distribution for a tree-homogeneous model, and white bars show the distribution for a tree-heterogeneous NDCH model with two composition vectors on each data partition. Panel (a) shows distributions for the SSU partition and panel (b) shows distributions for the LSU partition. Arrows show the  $X^2$  for the original data, showing that by this test two composition vectors for each data partition are needed to adequately model the data.

gives an improvement of the likelihood of 85 log units. Combining extra composition vectors with heterogeneous rate matrices in the NDCH(4,4) + NDRH(2,2) model (figure 2b) improves the likelihood by 70 log units over the NDCH(4,4) model alone. All of these models (table 2, rows H–J) show 86–88 per cent BPP support for the eocyte hypothesis. Interestingly, the LogDet analysis (table 2, row B; electronic supplementary material, figure S8) has 68 per cent BS support for the eocyte hypothesis; LogDet distances are relatively immune to compositional differences over the tree (but see Hirt *et al.* 1999).

The CAT model accommodates heterogeneity over the data, and using this model increased the fit to the data greatly, as shown by an increase in the likelihood of 1769 log units over the GTR +  $\Gamma$  model, making it by far the best-fitting model used (electronic supplementary material, figure S9). This model had 100 per cent BPP support for the eocyte hypothesis. A notable feature of this analysis is that there was low support for the monophyly of both the euryarcheotes and crenarcheotes.

# (b) Support for alternative trees of life based on protein-coding genes

It has been demonstrated previously (Brown *et al.* 2001; Cox *et al.* 2008) that equally weighted MP analyses of concatenated proteins across the tree of life support the traditional three-domains hypothesis. With our expanded eocyte sampling, this remained the case with both standard amino acid coding (86% BS for a monophyletic archaebacteria—table 3, row A; electronic supplementary material, figure S10) and Dayhoff-recoded data (88% BS—table 3, row G;

electronic supplementary material, figure S11). However, all other methods and models we employed supported the alternative tree of life; the eocyte hypothesis. One possible explanation of these contradictory results is that they are due to the increased tendency of the MP method to suffer from the distorting effects of long-branch attraction (LBA—Felsenstein (1978)), whereby taxa are grouped by an excess of unrecognized homoplasy rather than homologous changes. That we obtain similar results (i.e. the three-domains tree) to previous studies under the MP criterion suggests that our selection of genes and site inclusion was not responsible for our obtaining a different result (i.e. the eocyte tree) from other methods. Further highlighting the importance of composition heterogeneity in the tree of life are the results of LogDet distance analysis, which clearly identifies the eocyte hypothesis in preference to a monophyletic archaebacteria (88% versus 8% BS, respectively; table 3, row B; electronic supplementary material, figure S12). ML analyses (table 3, row C; electronic supplementary material, figure S13) also resolve the eocyte tree, but with little support (68% BS). By contrast, topologies identifying a monophyletic archaebacteria are not recovered from samples of the posterior probability distributions of any Bayesian analysis; in fact, all of these analyses support the eocyte hypothesis at, or near, maximum support values (figure 4a,b, table 3, rows D–F, H–J; electronic supplementary material, figures S14-S17). It is notable that including a covarion parameter had little effect on the fit of the model, increasing the log marginal likelihood by only 2 units (table 3, row D versus E). This result suggests that covarion-like rates of lineage evolution are not an important factor in these data. Nevertheless, caution should be urged with regard to this conclusion as the result of combining loci into a single partition may have the effect of homogenizing covarion structures of lineage substitution rates particular to individual loci. Despite this caution, we note that Cox et al. (2008) were only able to identify three genes included in the current dataset where a significant covarion structure was present.

Analyses using the CAT model with both standard amino acid coding and Dayhoff-recoded data fit the data much better than homogeneous models or the tree-heterogeneous NDCH model. That is, the CAT model fits the standard amino acid coded data 26 048 log marginal likelihood units better than the homogeneous WAG +  $\Gamma$  model (table 3, row D versus F), and the same model showed an improvement of 7312 log marginal likelihood units over the homogenous  $GTR + \Gamma$  model when the data were recoded into Dayhoff groups (table 3, row H versus J). Such remarkable improvements in model fit indicate the utility of the CAT model and the importance of modelling across-data compositional heterogeneity. Nevertheless, PhyloBAYES analyses of the amino acid data did suffer from a lack of convergence between independent runs. In both cases, the standard and Dayhoff coding of the amino acid data, the two runs differed in topology with respect to the placement of Encephalitozoon cuniculi, a taxon whose placement is

Table 3. Support for the three-domains tree and the eocyte hypothesis from combined protein coding genes.

	method	model	log-marginal likelihood <sup>a</sup>	Euryarchaeota monophyletic	Eocytes monophyletic	Archaebacteria monophyletic	Eocyte hypothesis
A	$MP^{b}$			79	36	86	10
В	LogDet-NJ <sup>c</sup>			64	69	8	88
С	$ML^d$	$WAG + \Gamma$		97	58	15	66
D	Bayesian	$WAG + \Gamma^{e}$	-246692	100	86	0	99
Е		Covarion <sup>e</sup>	-246690	100	99	0	100
F		$CAT + \Gamma^{f}$	-220644	0	100	0	100
G	MP-Dayhoff <sup>b</sup>			52	57	88	6
Η	Bayesian-Dayhoff	$GTR + \Gamma^{g}$	-106068	100	100	0	98
Ι		$NDCH(14)^{h}$	-105488	100	97	0	99
J		$CAT + \Gamma^{f}$	-98756	0	100	0	100

Rows A-F used standard AA coding, and rows G-J recoded the AA data into the six Dayhoff groups.

<sup>a</sup>Calculated as described in table 2.

<sup>b</sup>Maximum Parsimony bootstrap using PAUP\*.

<sup>c</sup>Bootstrap values are from neighbour-joining trees using PAUP\* made from LogDet distances calculated using P4. <sup>d</sup>Using RAXML-VI-HPC v. 2.2.3.

<sup>e</sup>Using MrBayes v. 3.1.2. The covarion model was WAG +  $\Gamma$  + Covarion.

<sup>f</sup>CAT mixture model, using PHYLOBAYES v. 2.3.

<sup>g</sup>Tree-homogeneous  $6 \times 6$  GTR +  $\Gamma$  model using P4.

<sup>h</sup>NDCH model where the numbers in parentheses show the number of composition vectors using P4.

known to be problematic (Embley & Martin 2006), but importantly, not with respect to the status of the eocytes or archaebacterial monophyly. In the Dayhoffrecoded amino acid data analyses, the differences between runs were not well supported (i.e. <95% posterior probability). In the CAT analysis of the standard amino acid coding data, however, there was strong (>95%) support for two alternative placements of *E. cuniculi;* either at the base of the eukaryotic tree, or with the fungi. In both cases we chose to present the results from the run with the best log marginal likelihood score: Dayhoff-recoded: run 1 -98 785 versus run 2 -98 754, and standard coding: run 1 -220 776 versus run 2 -220 644.

In our previous analyses (Cox et al. 2008), we found only two proteins, the largest subunits of RNA polymerase I (RPA1) and III (RPC1), that resolved archaebacterial monophyly under the NDCH model. Further analyses of RNA polymerases (RPA1, RPB1, RPC1, RPA2, RPB2 and RPC2) with additional eocyte taxa, under both the NDCH and CAT models, failed to find additional support for archaebacterial monophyly. Indeed, support for a monophyletic archaebacteria from RPA1 under the NDCH(2) model was eroded from 99 per cent to 57 per cent, and NDCH(2) analyses of RPC1 failed to resolve a monophyletic archaebacteria. All analyses of RNA polymerases under the CAT model failed to resolve a monophyletic archaebacteria or strongly identify any group as most closely related to the eukaryotes.

### (c) Decay of phylogenetic information is delayed by among-site rate variation

A potential criticism of any phylogenetic study based on anciently diverged molecules is that the sequences may be saturated with superimposed mutations, masking the historical signal (Philippe & Forterre 1999; Penny *et al.* 2001; Ho & Jermiin 2004). One way to visualize saturation is by using saturation plots (Philippe & Forterre 1999). These plots can be made for either simulated data, where the observed pairwise p-distances are plotted against the simulation branch lengths, or they can be made from empirical data where the p-distances are plotted against inferred branch lengths. In simulations using the simple Jukes-Cantor model for DNA (figure 5a), as branch lengths increased, the observed pairwise distances increased but plateaued at 0.75 as the sequences became randomized by superimposed mutations. This appeared to happen at branch lengths above about three mutations per site; if evolution behaved in this way it would be difficult or impossible to make reliable phylogenetic inferences based on such diverged sequences. A similar effect is shown in simulated protein sequences in figure 5b, which used the WAG model. Here saturation appears to have occurred at branch lengths above about 6.

However, the situation changed greatly in our simulations when we allowed the process of evolution to have among-site rate variation. The simulations shown in figure 5c were performed using the WAG +  $\Gamma$  model, that is, with gamma-distributed among-site rate variation. Here, we can see that complete saturation was never reached even after an average of 50 mutations per site. When there is among-site rate variation there are both slow sites and fast sites; the fast sites will become saturated at small simulation branch lengths (even sooner than the average sites in simulations without among-site rate variation), but the slow sites will be relatively immune from saturation even at high simulation distances. The biological sequences that we used do show among-site rate variation, and together with invariant sites, these slow sites allow us to recognize that anciently diverged sequences are homologous, and allow us to align sequences from some conserved genes over the entire tree of life with confidence in positional homology.



Figure 4. Bayesian phylogenetic analyses of concatenated amino acid data. The analysis in panel (*a*) used Dayhoff-recoded data with a  $GTR + \Gamma + NDCH(14)$  tree-heterogeneous substitution model in P4. This is the analysis summarized in table 3, row I. The analysis shown in panel (*b*) used standard amino acid-coded data with a CAT-Poisson +  $\Gamma$  substitution model in PHYLOBAYES. This is the analysis summarized in table 3, row F.

Using simulations we asked whether we would expect saturation to cause problems for our methods of analysis and our data. In figure 5d, points were from simulations based on samples from the posterior distribution of the tree-heterogeneous model analysis shown in row I of table 2. Lack of a plateau showed that in evolutionary scenarios such as this we would not expect complete saturation.

Turning now to the data that we used in this study, we asked whether plots of these data show saturation. Figure 6 shows that neither the rRNA genes, nor the protein sequences, nor the grouped amino acid sequences showed complete saturation. The rRNA genes are perhaps close to saturation, and we can speculate that this contributes to the generally ambiguous results that are shown in table 2 and in the published literature (e.g. Cox *et al.* 2008). It appears that the protein and grouped amino acid data are not saturated. The achaebacteria-eukaryote pairs are the most relevant to our problem, and those distances, isolated in figure 6e-h, are clearly not saturated, as there are larger pairwise distances with larger p-distances evident in figure 6a-d.

While the saturation curves in figure 6 show that the sequences are not saturated, we also approached the question from a different angle and asked whether we would expect there to be enough phylogenetic signal remaining in protein sequences that have evolved under a WAG +  $\Gamma$  model at the level of divergence seen in figure 6*b*. Simulations were made using a WAG +  $\Gamma$  model on a four-taxon tree with terminal branch lengths of 1.5 and an internal branch length of 0.1. The simulation sequences were 5222 characters long, the same length as the concatenated protein



Figure 5. Saturation plots from simulated data. Simulation distances are simulation branch lengths, measured in average mutations per site. Panel (*a*) is from DNA simulated under the Jukes-Cantor model. Panel (*b*) is from protein simulated under the WAG model. The simulations in panels (*a*) and (*b*) were performed without among-site rate variation. Panel (*c*) shows protein simulations under the WAG +  $\Gamma$  model, i.e. including among-site rate variation. Panel (*d*) shows DNA simulations based on samples from the posterior distribution of the analysis shown in row I of table 2. Lines were fit from the second half of the points.

dataset, and the alpha value for the gamma-distributed rate variation was set to 1.94, the posterior average. For each replicate simulation, the ML of the three possible four-taxon trees was calculated. Of 500 replicates, 483 ML trees (96.6%) were the simulation topology. If we had saturation, we would have expected that the three possible topologies would be the ML tree about one-third of the time each. This shows that even when the sequences had been hit by an average of three mutations per site between taxa pairs (more than is seen between taxa pairs in figure 6b) there was still enough phylogenetic signal remaining to allow high accuracy in a phylogenetic reconstruction.

### 4. CONCLUSIONS AND IMPLICATIONS FOR THEORIES OF EUKARYOTIC ORIGINS

The alignments that we used were conservative to ensure that our hypotheses of positional homology between domains were as robust as we could make them. In doing so, we removed many positions that could be reliably aligned within domains but not between them. This inevitably removed some signal for relationships within domains and may have contributed to the recovery of some of the controversial or unconventional relationships that we observed in our trees. For example, in most analyses the long-branched microsporidian *Encephalitozoon* was recovered near the base of the eukaryote cluster, rather than with fungi where most data would place it (Embley & Martin

2006). The CAT method is reported to be more robust to long-branch artefacts than other methods (Lartillot et al. 2007); so it was interesting to see that CAT analyses of the concatenated proteins did indeed unite Encephalitozoon with Saccharomyces (figure 4b; electronic supplementary material, figure S17). Most analyses recovered the Crenarchaeota/eocytes and the euryarchaeotes as monophyletic groups, as classically depicted in both the three-domains and eocyte trees. By contrast, the CAT analyses recovered euryarchaeotes as a paraphyletic group, with Pyrococcus as the sister to the Crenarchaeota/eocytes plus eukaryotes. It will be interesting to test how robust these relationships are to increased taxon sampling and to an expanded sequence alignment. Taken at face value, they raise the possibility that Crenarchaeotes/eocytes plus eukaryotes may have originated from within the Euryarchaeote radiation.

There is currently a debate about how far back molecular phylogenetics might be able to take us (e.g. Philippe & Forterre 1999; Penny et al. 2001; Ho & Jermiin 2004). Phylogenetic methods will generally construct trees irrespective of whether any historical signals for relationships remain in the data. The message from computer simulations is that success in recovering any ancient signal is related to the properties of the data, for example whether it contains a mixture of site rates, and how well the model fits the data (Penny et al. 2001; Ho & Jermiin 2004; Lartillot et al. 2007). Our simple simulations illustrate some of these issues and are consistent with the possibility of there being signal for historical relationships in the data that we analysed. There are unicellular microfossils (acritarchs) argued to be eukaryotic in strata of about 1.45 Gyr of age that provide one estimate for a minimal age for eukaryotes (Javaux et al. 2001). This figure is consistent with an age of between 950 and 1259 Myr for the diversification of major eukaryotic groups that has been estimated from concatenated molecular sequence data using a relaxed molecular clock (Douzery et al. 2004).

The three-domains hypothesis (Woese et al. 1990) explains the similarities in the eukaryotic and archaebacterial transcription and translation machinery (Zillig et al. 1985; Olsen & Woese 1997), as originating in a common ancestor that was not shared with eubac-This putative common ancestor teria. was subsequently called a neomuran by Cavalier-Smith (2002). By contrast, the eocyte hypothesis posits that the observed similarities reflect the origin of eukaryotes from within the archaebacteria as the sister group of a specific group of archaebacteria called the Crenarchaeota or eocytes. In the current work, we have investigated the support for these competing hypotheses from genes and proteins that largely, but not exclusively, comprise components of the salient genetic machinery (Cox et al. 2008). Our results, based upon an increased sampling of eocytes, are in agreement with those that we published earlier (Cox et al. 2008), namely, when methods are used that are designed to overcome across-tree (Foster 2004), or across-data compositional heterogeneity (Lartillot & Philippe 2004), features that are manifestly evident for these data, it is the eocyte tree that is favoured and not the three-domains tree.



Figure 6. Saturation plots of empirical data. Inferred distances are patristic distances between taxa pairs following the tree path. Panels (a)-(d) show all points; panels (e)-(h) isolate pairs where one member of the pair is a eukaryote sequence and the other is an archaebacterial sequence. Panels (a) and (e): rRNA data with the tree-heterogeneous model shown in row I in table 2. Panels (b) and (f): protein sequences analysed with the WAG +  $\Gamma$  model. Panels (c) and (g): protein sequences analysed with the CAT model. Panels (d) and (h): protein sequences recoded into the six Dayhoff groups and analysed with a GTR +  $\Gamma$ -like model.

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