

Searching for Footprints of Positive Selection in Whole-Genome SNP Data From Nonequilibrium Populations

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ABSTRACT

A major goal of population genomics is to reconstruct the history of natural populations and to infer the neutral and selective scenarios that can explain the present-day polymorphism patterns. However, the separation between neutral and selective hypotheses has proven hard, mainly because both may predict similar patterns in the genome. This study focuses on the development of methods that can be used to distinguish neutral from selective hypotheses in equilibrium and nonequilibrium populations. These methods utilize a combination of statistics on the basis of the site frequency spectrum (SFS) and linkage disequilibrium (LD). We investigate the patterns of genetic variation along recombining chromosomes using a multitude of comparisons between neutral and selective hypotheses, such as selection or neutrality in equilibrium and nonequilibrium populations and recurrent selection models. We perform hypothesis testing using the classical *P*-value approach, but we also introduce methods from the machine-learning field. We demonstrate that the combination of SFS- and LD-based statistics increases the power to detect recent positive selection in populations that have experienced past demographic changes.

GENOMES contain information related to the history of natural populations. Past neutral and selective processes may have left footprints in the genome. Recent advances in population genetics aim to understand the patterns of genetic diversity and identify events that have led to genetic adaptations. Among them, positive selection has been a focus of many recent studies (HARR *et al.* 2002; KIM and STEPHAN 2002; GLINKA *et al.* 2003; AKEY *et al.* 2004; ORENGO and AGUADÉ 2004). Their goal is to (i) provide evidence of positive selection, (ii) estimate the strength and the rate of selection, and (iii) localize the targets of selection. These objectives form the basis of a long-term pursuit, which is the understanding of the molecular basis of adaptation of populations in a changing environment.

Positive selection can cause genetic hitchhiking when a beneficial mutation spreads in the population (MAYNARD SMITH and HAIGH 1974). When a strongly beneficial mutation occurs and spreads in a population, linked neutral or slightly deleterious variants hitchhike with it, and their frequency increases. According to Maynard Smith and Haigh's model, three patterns are generated locally around the position of the beneficial

mutation. First, the level of variability will be reduced since standing variation of the population that is not linked to the beneficial allele vanishes, and tightly linked polymorphisms may fix (KAPLAN *et al.* 1989; STEPHAN *et al.* 1992). Second, the site frequency spectrum (SFS), which describes the frequency of allelic variants, shifts from its neutral expectation toward rare and high-frequency derived variants (BRAVERMAN *et al.* 1995; FAY and WU 2000). The third signature describes the emergence of specific linkage disequilibrium (LD) patterns around the target of positive selection, such as an elevated level of LD in the early phase of the fixation process of the beneficial mutation and a decay of LD across the selected site at the end of the selective phase (KIM and NIELSEN 2004; STEPHAN *et al.* 2006).

The availability of genome-wide SNP data has made possible the scanning of genomes and the identification of loci that may have been targets of recent selective events. Several approaches have been developed within the last years that can detect the molecular signatures of positive selection (KIM and STEPHAN 2002; JENSEN *et al.* 2005; NIELSEN *et al.* 2005). While the methods of KIM and STEPHAN (2002) and JENSEN *et al.* (2005) are designed to analyze subgenomic SNP data, the approach of NIELSEN *et al.* (2005) can be applied to both subgenomic and whole-genome data (reviewed in PAVLIDIS *et al.* 2008). For this reason we concentrate here on the latter procedure. This method, called *SweepFinder*, calculates the probability $P(x)$ that a polymorphism of multiplicity x is linked to a beneficial

Supporting information available online at <http://www.genetics.org/cgi/content/full/genetics.110.116459/DC1>.

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mutation using a simple selective model and the SFS prior to the selective event. Then, for each location in the genome it compares a selective with a neutral model assuming independence between the SNPs, therefore calculating the composite likelihood ratio Λ . Thus, it identifies regions where the likelihood of the selective sweep is greater than that of the neutral model using the maximum value Λ_{MAX} of Λ .

The ω -statistic, developed by KIM and NIELSEN (2004), detects specific LD patterns caused by genetic hitchhiking (described above). In the study by KIM and NIELSEN (2004) the maximum value of the ω -statistic was used to identify the targets of selective sweeps. Later, JENSEN *et al.* (2007) studied its performance in separating demographic from selective scenarios. An important result by JENSEN *et al.* (2007) is the demonstration that for demographic parameters relevant to nonequilibrium populations (such as the cosmopolitan populations of *Drosophila melanogaster*) the ω -statistic can distinguish between neutral and selective scenarios. This article further develops *SweepFinder* and the ω -statistic such that they can eventually be applied to whole-genome SNP data sets that have been collected from nonequilibrium populations. In particular, populations undergoing population-size bottlenecks are of interest as these size changes may confound the patterns of selective sweeps (BARTON 1998). For this reason we use the following approach: first, we theoretically analyze the genealogies of bottlenecked populations under neutrality and show to what extent they resemble the genealogies of single hitchhiking (SHH) events. We also point out the importance of high-frequency-derived variants in the identification of selective sweeps. Second, we study the statistical properties of *SweepFinder* and the ω -statistic separately and in combination. As the main result, we demonstrate that the combination of these two methods (that include both SFS and LD information) increases the power for detecting recent SHH events in nonequilibrium populations, in particular when machine-learning techniques are employed. Third we analyze the performance of *SweepFinder* and the ω -statistic in the detection of recurrent hitchhiking (RHH) events.

METHODS

Modifications of the ω -statistic and *SweepFinder*: The proposed modifications aim at (i) adapting the ω -statistic for the analysis of whole-genome data and (ii) increasing the accuracy of *SweepFinder* to predict the target of selection. Instead of fixed windows, variable-size windows are used in the ω -statistic, and in the *SweepFinder* algorithm a fraction of monomorphic sites is incorporated.

The hitchhiking model by MAYNARD SMITH and HAIGH (1974) predicts that an excess of LD arises after

the completion of the selective sweep within each of the two regions flanking the selected site, but does not extend across the two regions (STEPHAN *et al.* 2006; McVEAN 2007; PFAFFELHUBER *et al.* 2008). This is due to the assumption that any observed polymorphism around the sweep has been introduced in the population prior the selective sweep and entered the beneficial genetic background through recombination. Since independent recombination events are necessary to explain polymorphisms on both sides of the selective sweep, the LD vanishes across the site of the beneficial mutation, but not within each side. This genomic footprint may be captured using the ω -statistic (KIM and NIELSEN 2004). Assume a genomic window with S segregating sites that is split into a left and right subregion with l and $S-l$ segregating sites, respectively. The ω -statistic (Equation 1) quantifies to what extent average LD is elevated on each side of the selective sweep (see the numerator of Equation 1) but not across the selected site (see the denominator of Equation 1):

$$\omega = \frac{\left(\binom{l}{2} + \binom{S-l}{2} \right)^{-1} \left(\sum_{i,j \in L} r_{ij}^2 + \sum_{i,j \in R} r_{ij}^2 \right)}{(l(S-l))^{-1} \sum_{i \in L, j \in R} r_{ij}^2}. \quad (1)$$

The ω -statistic considers the space between the left and right subregions as the center of the selective sweep. Thus, a genomic region may be scanned and scores are reported for each position. Then, using simulations, a significance threshold is determined. The maximum value ω_{MAX} predicts the target of recent positive selection. In the original version of the ω -statistic, the borders of the left and right subregions are assumed constant (KIM and NIELSEN 2004; JENSEN *et al.* 2007). This may be valid for a subgenomic analysis, when the recombination rate ρ and mutation rate θ do not fluctuate much or a single selective event may have occurred. However, in a whole-genome study these parameters that affect the extent of LD may vary dramatically. Additionally, the polymorphism patterns may have been shaped by recurrent selective sweeps. Thus, the constant-border approach implemented by KIM and NIELSEN (2004) may be limited. If the subregions are large, then ω_{MAX} tends to decrease and the signal disappears. On the other hand, short subregions might contain no SNPs and the ω -statistic cannot be calculated.

We have implemented a variable-window size ω -statistic. The borders of the left and right subregions vary and the configuration that maximizes ω is reported. This approach overcomes the aforementioned problems inherent in the constant-border approach of KIM and NIELSEN (2004). Thus, it may be suitable for scanning large genomic regions or whole chromosomes characterized by variable ρ or θ parameters and shaped by recurrent adaptive substitutions.

A naive implementation of the ω -statistic scanning algorithm would recalculate the LD of the positions as the center of the sweep moves along the chromosome. This is particularly critical for the variable-window size approach since the number of calculations increases. Our implementation, as illustrated in [supporting information](#), Table S1, guarantees a single calculation between any two sites that may participate in the ω calculation. Thus, it results in an algorithm that is efficient when the number of polymorphisms is large. Calculations are performed using a matrix Z (Table S1), which stores the unweighted Z_{ns} (KELLY 1997) values (not divided by the number of comparisons) for all possible windows. For a pair $(i, i + 1)$, $Z_{i,i+1}$ equals the correlation coefficient between these two positions. This value is then added to all cells $Z_{j,i+1}$, with $j < i$ to form the Z_{ns} for the region $[j, i + 1]$. With this method all possible numerators of the ω -statistic are formed. When the left and right subregions are defined by $[i, k]$ and $[k + 1, j]$, respectively, then the denominator is simply a weighted version of $Z_{i,j} - Z_{i,k} - Z_{k+1,j}$.

SweepFinder detects the shift of the SFS as a signature of hitchhiking. Demographic effects are incorporated through the neutral SFS, which is either provided by the user or calculated from the data itself. Monomorphic sites are generally excluded from the analysis (NIELSEN *et al.* 2005; SVETEC *et al.* 2009) since tests that include them may be more sensitive to assumptions regarding the mutation rate (NIELSEN *et al.* 2005). Additionally, for realistic mutation rates, the majority of the sites remain monomorphic. Thus, by including invariant sites, the data set and the computational time required for the analysis increase dramatically. On the other hand, the decrease of diversity represented by the monomorphic sites constitutes a well-known signature of the hitchhiking effect. Omitting them may decrease the power of the tests (NIELSEN *et al.* 2005) and lead to inaccurate predictions about the target of selection. Inaccuracies mainly emerge due to changes in the input site density when only polymorphic sites are included. We incorporate a fraction of the monomorphic sites into the analysis in a way that (i) generates a uniform input site density and (ii) preserves the signature of low diversity in regions of depleted variation. Additionally, since only a small fraction of monomorphic sites are used, the computational time is only increased slightly. Given a genomic region with S polymorphic sites we include Sq monomorphic sites, where $0 < q < 1$. In this study $q = 0.1$, so that the number of monomorphic sites are in the same order as the polymorphic sites. We proceed as follows. In the first step, there are $S - 1$ intervals between the S polymorphic sites. A monomorphic site is included at a random location within the largest interval. In the second step there are $S + 1$ sites and S intervals and the process is repeated. The cutoff value is defined by treating the neutral simulations in the same way. With this process the SNP density differ-

ences are reduced and monomorphic sites are embedded in regions of depleted variation.

Quantifying the effects of population bottlenecks on neutral genealogies: The ω -statistic and *SweepFinder* can scan genomes from natural populations that have experienced demographic changes and detect targets of selection. We investigated whether the neutral demographic scenarios inferred by LI and STEPHAN (2006) and THORNTON and ANDOLFATTO (2006) to describe the demography of a European population of *D. melanogaster* can result in patterns along a recombining chromosome that resemble selective sweeps. In particular, we examined which effects of population bottlenecks are responsible for the polymorphism patterns that mimic the effects of selective sweeps. We focused on the properties of genealogies that are generated by those two demographic models because genealogies reflect demographic properties more comprehensively than summary statistics.

A way to measure the effect of a bottleneck on the genealogies of a recombining genome is through the ratio $f = L_n/H_n$ of the total length to the height of the coalescent. Short, star-like genealogies have large ratios and $\max(L_n/H_n) = n$ is obtained for a n -furcated star-like tree. On the other hand, for genealogies with long internal branches the ratio takes small values and $\min(L_n/H_n) = 2$ is obtained when the genealogy is dominated by two very long internal branches. Using simulations we first calculate the percentage of n -furcated star-like genealogies (with large f values) in a region of 50 kb. Then, for each simulated instance we relate the percentage of n -furcated star-like genealogies with the resemblance to a selective sweep as this is measured using *SweepFinder* (see THEORETICAL ANALYSES).

The joint effects of population bottlenecks and selective sweeps on high-frequency derived alleles: A hallmark of selective sweeps in constant populations is the excess of high-frequency-derived variants around the target of positive selection. High-frequency-derived variants consist of mutations that were present in the population prior to the selective sweep, hitchhike with the beneficial allele, and, due to recombination, appear as polymorphisms. This signature forms the basis of a multitude of neutrality tests that are based on the SFS (FAY and WU 2000; KIM and STEPHAN 2002; NIELSEN *et al.* 2005) and contributes to the precise detection of the target of selection. However, in natural populations positive selection may occur simultaneously with demographic changes. Using simulations from the demographic models that were inferred by LI and STEPHAN (2006) and THORNTON and ANDOLFATTO (2006), we examine whether high-frequency-derived alleles occur when demographic changes occur simultaneously with positive selection.

Measuring the precision of the inferred selective sweep position: An objective of the genome-scanning studies is the precise prediction of the selective sweep locations. Usually, every position or a subset of them is

scored for a given statistic (for example, the ω -statistic or the *SweepFinder*). Thus, peaks and valleys are formed along the genomic region. Then, some of the peaks may survive a cutoff value delimiting the potential targets of selection. As illustrated in Figure S1, we determine the distance between a peak on the landscape of the statistic and the closest location where a selective sweep has occurred given a user-defined threshold. In Figure S1, two selective sweeps have occurred recently in the history of the population. The positions of the sweeps are illustrated as vertical green lines. A peak is defined as the highest point in an isolated region by the cutoff value. Thus, five peaks (*a* to *e*) have been formed in the example of Figure S1. D measures the distance between a peak and the closest selective sweep location. On the basis of this approach we can measure the accuracy of the different methods. Furthermore, we implemented a simple randomization of the peaks to evaluate the quality of the predictions. This is necessary because finite genomic regions are simulated, and therefore the distance between any location and the target of selection is bounded.

Supervised-learning techniques: We introduce supervised-learning approaches from the field of machine learning that can be useful for the classification of a genomic region as either neutral or selected. In a classification problem, the goal is to separate these classes using a function, which is inferred from the available data. Such a process is called “learning from the data” or “supervised learning” and is related to finding the optimal hyperplane that distinguishes the two classes. Typically, in a supervised-learning problem, data consist of pairs of input and output objects. Input consists of a vector of multiple entries that summarize the data and are called features. Inputs can be set arbitrarily depending on the specific problem. However, the efficiency of the algorithm increases when they are independent and capture the whole information of the data. Output can be binary, denoting the class that the object belongs to. In supervised learning the goal is to use the input to predict the value of the output, and the problem can be formulated as teaching the computer the combinations of feature values that are associated with either of the classes. In the specific problem we examine here, the output is coded as neutrality/selection. Then, using simulations of the neutral demographic model and the model with selection we train the algorithm to separate these two classes. As input for the machine-learning approach we use ω_{MAX} , Λ_{MAX} (from the original algorithms), and combinations of ω and Λ , such as the distance between the genomic positions of ω_{MAX} and Λ_{MAX} and the correlation coefficient between ω and Λ . The reasoning for this choice of inputs is as follows. First, Λ_{MAX} and ω_{MAX} capture different aspects of the data. Λ_{MAX} is affected mostly by the SFS, whereas ω_{MAX} is affected by LD. Even if SFS and LD can be correlated (KIM and NIELSEN 2004), it is

expected that this correlation is lower than that from using statistics that are based exclusively on the SFS or LD. Second, previous studies have shown that Λ_{MAX} and ω_{MAX} are relatively robust to demographic changes (but see ORENKO and AGUADÉ 2010). Third, it seems intuitively obvious that the peaks of ω and Λ profiles should point to the same genomic location if a selective sweep has occurred. Thus, using the distance between the peaks or the correlation of the profiles should increase the classification performance of the algorithm. In this study, both the distance between the peaks and the correlation between the profiles are used.

For each demographic scenario that was simulated in this study, we used a subset of simulations for training and the remaining for testing the performance. The supervised-learning approach can be employed to classify a certain genomic region as either neutral or selected. However, within a region the specific target of selection cannot be specified by the method itself. To achieve this, the features of the method (*i.e.*, the ω and Λ profiles) should be inspected. Tables 1–4 provide information about the accuracy of the features under various demographic scenarios.

Traditionally, when neutrality tests are employed to detect targets of positive selection, neutral simulations are performed and the 5% percentile is used as a threshold. This methodology assumes that neutrality tests produce significantly larger values in data with selection. This may be the case when the population size remains constant. However, in nonequilibrium models the values of the neutrality tests may overlap significantly between neutral models and models with selection, and therefore their performance decreases. Combining different statistics that capture different aspects of the data may contribute to increasing the classification performance.

Several methods have been developed for data classification. For example, Bayesian classifiers, rule-based classifiers, k -nearest neighbors, and linear discriminant analysis are some of the approaches that have been applied to supervised-learning problems (DUDA *et al.* 2000; HAN and KAMBER 2000; HASTIE *et al.* 2001). Here, we demonstrate the use of SVMs with a radial basis kernel, which is the most widespread kernel. In general, SVM uses a nonlinear mapping to transform the original training data into a higher-dimensional space and to search for an optimal linear hyperplane in this space. A great advantage of SVMs is that they are highly accurate and less prone to overfitting; *i.e.*, they have desirable generalization properties (HAN and KAMBER 2000).

Implementation and code availability: The C++ source code is available from <http://www.bio.lmu.de/~pavlidis>. For the ω -statistic, the user is able to choose between constant- or variable-window-size scanning modes. Additionally, besides r^2 various other measurements of LD, such as $\text{abs}(D)$ and $\text{abs}(D_\omega)$ (LANGLEY and CROW 1974), may be used in Equation 1.

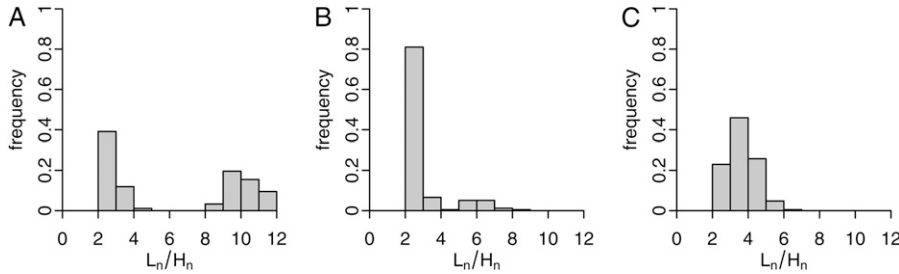


FIGURE 1.—Histogram of the ratio $f = L_n/H_n$ for the following demographic scenarios. (A) a single realization of the bottleneck scenario inferred by LI and STEPHAN (2006). Long coalescent trees that escape the bottleneck tend to produce small ratios (<4). On the other hand, genealogies that coalesce within the bottleneck period produce star-like trees because of the recent, rapid, and severe contraction of the

population. (B) A realization of the bottleneck scenario inferred by THORNTON and ANDOLFATTO (2006). In contrast to LI and STEPHAN (2006), coalescent events occur continuously. (C) The standard neutral model. For the LI and STEPHAN (2006), THORNTON and ANDOLFATTO (2006), and the neutral scenario, 12 chromosomes of 50 kb have been simulated. The recombination rate is $\rho = 0.05/\text{bp}$ and the mutation rate $\theta = 0.004/\text{bp}$. The parameter values for the LI and STEPHAN (2006) and THORNTON and ANDOLFATTO (2006) scenarios are described in the main text.

There are no specific library dependencies and the software can be installed on any Linux machine that runs the g++ compiler. Also, the modified version of *SweepFinder* that has been used here to analyze data with monomorphic sites is provided. In this version the likelihood curve of monomorphic sites has been modified so that the probability to observe a monomorphic site is high in the proximity of the sweep position but becomes negligible as distance increases (the rate of decrease is larger than in the original version). The original version of *SweepFinder* is provided by the website of Rasmus Nielsen (<http://people.binf.ku.dk/rasmus/webpage/sf.html>). Furthermore, perl scripts that have been used in the analysis are available from <http://www.bio.lmu.de/~pavlidis> or upon request from the authors.

THEORETICAL ANALYSES

The genealogies of bottlenecked populations may resemble those of SHH in constant-size populations: Past demographic changes such as bottlenecks may confound the patterns of a selective sweep (BARTON 1998). Similarly to a selective sweep, a bottleneck scenario may result in coalescent trees dominated by either external or internal branches. Short coalescent trees with long external branches are obtained when, due to a rapid, recent, and severe decrease of population size, the time of the most recent common ancestor of the sample is found within the bottleneck period. On the other hand, if some of the lineages escape the bottleneck, then long internal branches will be created. In recombining genomic regions, short and long trees may alternate, creating sweep-like patterns in the SFS (BARTON 1998).

We illustrate the effect of bottlenecks on genealogies using the demographic scenarios that have been inferred by LI and STEPHAN (2006) and THORNTON and ANDOLFATTO (2006) to describe the history of the European population of *D. melanogaster*. Scaling the time in units of $4N$ generations (where N is the present effective population size) the LI and STEPHAN (2006)

model describes a four-epoch scenario. Backward in time, the population experiences a bottleneck from 0.0367 time units until 0.0375 time units. Within this bottleneck period $N_b = 0.002N$, where N_b denotes the effective population size in the bottleneck. Then, instantly, the size of the population size changes to $7.5N$, and eventually at the time 0.1395 it becomes $1.5N$. The bottleneck phase models the founding of the European population from the ancestral population, whereas the transition from $7.5N$ to $1.5N$ models a (forward-in-time) expansion of the ancestral population. The demographic scenario inferred by THORNTON and ANDOLFATTO (2006) implements a three-epoch model. The values of the parameters depend on the ratio ρ/θ and here we use the results obtained when $\rho/\theta = 10$. The present population size N is estimated to be 2.4×10^6 , and backward in time at 0.0042 it contracts to $0.029N$. Finally, the population reaches instantly the present-day level at time 0.022.

The demographic model of LI and STEPHAN (2006) produces both star-like and long genealogies in the same genomic region of a recombining chromosome (Figure 1). The length of these trees is on average shorter than that of the standard neutral trees, thus reducing variation. The effect of the THORNTON and ANDOLFATTO (2006) demographic model is similar, but milder. On average, it creates shorter genealogies and effectively reduces the nucleotide polymorphism. However, it does not result in extreme star-like coalescent trees as often as the LI and STEPHAN (2006) model (Figure 1). This is because the population-size changes are milder, the bottleneck period is longer, and starts (backward in time) very recently in the usual coalescent time scale, allowing for a series of coalescent events.

Next we used simulations to examine the relationship between the percentage of star-like genealogies, the number of segregating sites, and Λ_{MAX} of *SweepFinder*, which can be considered a proxy for the resemblance of polymorphism patterns (based on the SFS) to a signature of a selective sweep. A 50-kb genomic region was simulated using *ms* (HUDSON 2002) for a sample of 12

chromosomes. The recombination rate $\rho = 0.05/\text{bp}$ and the mutation rate $\theta = 0.004/\text{bp}$. The demographic model describes a recent population bottleneck (as inferred by LI and STEPHAN (2006)). As illustrated in Figure 2, a small number of star-like trees create a large number of segregating sites and small Λ_{MAX} values. Similarly, when a genomic region is dominated by short, star-like genealogies, the number of segregating sites and Λ_{MAX} decrease. Even if this constitutes a polymorphism valley, the pattern does not look like a sweep because of a lack of the high-frequency derived variants (KIM and STEPHAN 2002). On the other hand, the simultaneous presence of star-like and long genealogies creates sweep-like patterns. For intermediate frequencies of star-like genealogies, Λ_{MAX} assumes large values. Since neighboring genealogies are not independent, star-like genealogies form clusters and effectively create valleys of reduced polymorphism resembling a selective sweep. These results help to interpret some of our findings below.

Selective sweeps in nonequilibrium populations may result in a loss of high-frequency-derived variants and violate the assumptions of *SweepFinder* and the ω -statistic: We examined the effects of selective sweeps on polymorphisms, when they occur within demographic bottlenecks. A 50-kb genomic fragment was simulated under the bottleneck model inferred by THORNTON and ANDOLFATTO (2006), and a selective sweep ($\alpha = 2500$) was assumed to take place within the bottleneck period (THORNTON and JENSEN 2007). First, we show that the combined action of selective sweeps and bottlenecks results in SFS that differ considerably from those generated by selective sweeps in equilibrium populations. Figure 3 compares the modifications of the average SFS around the target of selection in a constant-size demographic scenario with the model inferred by THORNTON and ANDOLFATTO (2006). It is apparent that in equilibrium demographic models there is a dramatic increase of the class $n - 1$ in the proximity of the selective sweeps (Figure 3A). Neutrality tests based on the SFS can detect the increase of the high-frequency-derived variants and therefore the accurate prediction of the target of selection is possible. In nonequilibrium scenarios, when population contraction and selective sweeps co-occur, the $n - 1$ class vanishes in a large genomic region around the target of selection (Figure 3B). The joint effect of selection and population contraction increases the probability of coalescences, resulting in short genealogies where the most recent common ancestor is located within the bottleneck phase. Consequently, the frequency of the $n - 1$ class vanishes in the present-day sample. Furthermore, the part of the genealogy that is older than the selective sweep/bottleneck phase is eliminated. Therefore the vast majority of the present-day polymorphisms are younger than the selective sweep. This violates the assumptions of *SweepFinder* and the ω -statistic and

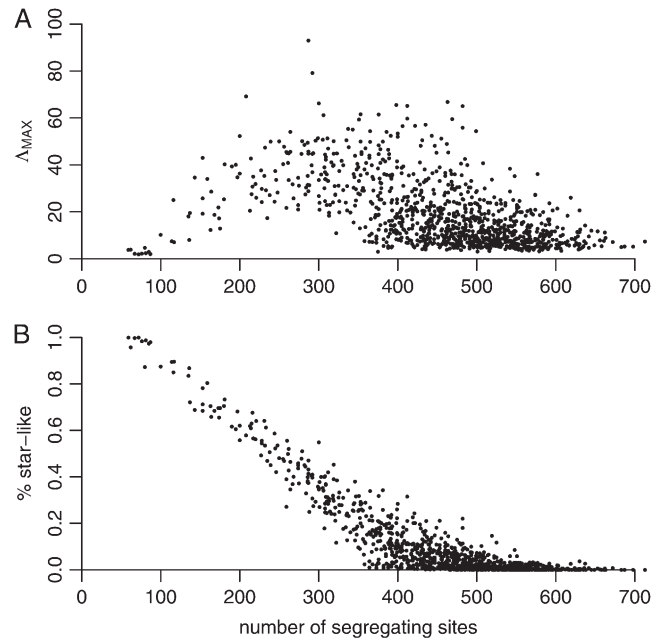
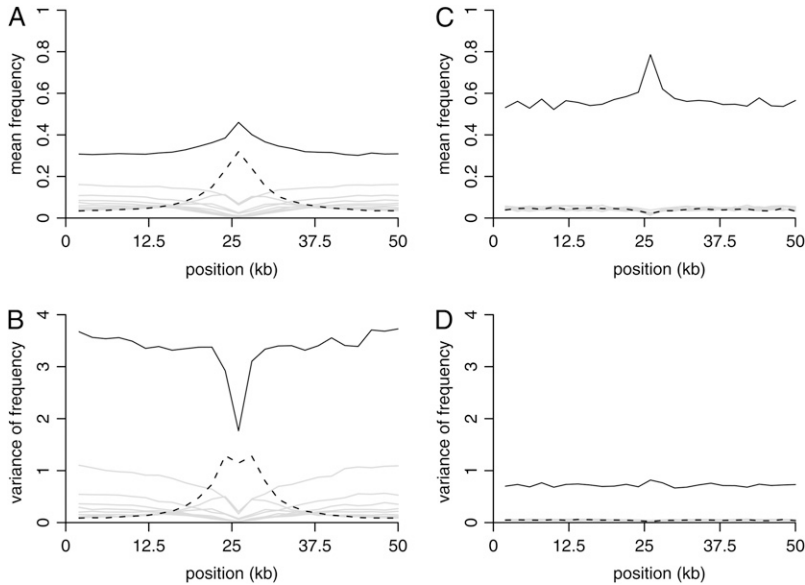


FIGURE 2.—The relation between (A) Λ_{MAX} and (B) the percentage of star-like genealogies and the number of segregating sites in the LI and STEPHAN (2006) demographic scenario. We have performed neutral simulations for 12 recombining chromosomes, assuming a length of 50 kb. The recombination rate $\rho = 0.05/\text{bp}$ and the mutation rate $\theta = 0.005/\text{bp}$. The parameter values for the demographic model inferred by LI and STEPHAN (2006) are described in the main text. The number of short genealogies in the LI and STEPHAN (2006) scenario determines both the number of segregating sites and the sweep resemblance (measured by the *SweepFinder* statistic). When a genomic region is dominated by short star-like genealogies only a few segregating sites are present. Even if this constitutes a polymorphism valley, the pattern does not look like a single sweep because of a lack of the high-frequency derived variants (KIM and STEPHAN 2002). Similarly, when the star-like trees are absent Λ_{MAX} is small. On the other hand, the simultaneous presence of star-like and long genealogies creates sweep-like patterns. This is because star-like trees tend to cluster together along the recombining chromosome, creating valleys within polymorphism islands.

may result in imprecise prediction of the target of selection.

STATISTICAL PERFORMANCE OF THE TESTS IN THE DETECTION OF SHH

In this section, the discrimination capacity of *SweepFinder* and the ω -statistic is scrutinized, and the distance between the predicted and the true target of selection is evaluated for single sweeps under the scenarios (i) selection *vs.* neutrality in equilibrium populations (*i.e.*, standard neutral populations), (ii) selection in equilibrium populations *vs.* neutrality in nonequilibrium populations (*i.e.*, populations that have



derived variants are very small due to the simultaneous action of the sweep and the bottleneck. Therefore, the observed polymorphisms (mostly singletons) are younger than the selective event and spread over the whole genomic region, obscuring the location of the selective sweep.

experienced past demographic changes), and (iii) selection *vs.* neutrality in nonequilibrium populations. The performance is assessed as follows. First, the false positive (FP) rate of the SVM is estimated. Using this false positive rate we compare the true positive (TP) rates of each test. Thus, all comparisons refer to the same false positive rate. Second, for the evaluation of the distance between the true and predicted targets we use only simulated results that survive the threshold defined by the false positive rate. Finally, for the nonequilibrium models with selection we implement a simple randomization process to assess the quality of results (see METHODS).

SHH *vs.* neutrality in equilibrium populations: We simulate a single selective sweep in the middle of a 50-kb genomic region using the *ssw* software (KIM and STEPHAN 2002). The parameter values have been chosen for their relevance to natural populations of *D. melanogaster*. Specifically, the parameter $\alpha = 2Ns$, where s is the selection coefficient of the beneficial mutation, assumes the values 500, 2500, and 5000 that are realistic for *D. melanogaster* (BEISSWANGER and STEPHAN 2008). For all data sets the mutation rate $\theta = 0.005/\text{bp}$, similar to the estimation of θ for the European population of *D. melanogaster* by LI and STEPHAN (2006). The scaled recombination rate $\rho = 0.05/\text{bp}$, so that the ratio $\rho/\theta = 10$ (THORNTON and ANDOLFATTO 2006). The standard neutral simulations were performed using the same value of ρ . We used a sample size of 12 for all simulations.

Each realization of the selective sweep was compared with those of the standard neutral model that are obtained using $\theta_{\text{NEU}} = \theta_{\text{W}} = S_n/h_n$. θ_{NEU} denotes the

FIGURE 3.—A selective sweep causes a spatial modification of the SFS. The mean and the variance of the frequency are modified when a selective sweep has occurred in the middle of a 50-kb genomic fragment. The 50-kb region is split in 2-kb nonoverlapping windows and in each one the average mean (f_i) (A and C) and the variance $\text{var}(f_i)$ (B and D) of the frequency f_i of the polymorphism class i is calculated. In A the plots refer to a selective event in equilibrium populations ($\alpha = 2500$) that has been completed recently, whereas in C, the plots refer to the nonequilibrium model of THORNTON and ANDOLFATTO (2006) ($\alpha = 2500$). The solid lines refer to the singletons, the dashed lines to the class 11, and the gray lines to the classes 2–10. The dramatic change of the high-frequency derived alleles in A contributes to the precise localization of the selective event. On the contrary, in C the high-frequency-derived SNPs are absent even in the proximity of the selective sweep. This is because the length of the branches of the coalescent tree that may generate high-frequency-

θ value used in standard neutral simulations, θ_{W} is Watterson's (1975) estimator of θ obtained using the number of segregating sites S_n of the selective sweep realization, and $h_n = \sum_{i=1}^{n-1} \frac{1}{i}$. Thus, a selective sweep is compared with the standard neutral realizations that on average create the observed number of polymorphic sites [$F\theta$ procedure (RAMOS-ONSINS *et al.* 2007)]. Alternative approaches to calculating the threshold value may use the observed number of segregating sites S_n or take into account the uncertainty on θ by considering a prior distribution of θ . In neutral equilibrium populations these approaches result in the same threshold values for the models tested in this study (Figure S2). Here, for the calculation of thresholds we use the $F\theta$ approach. Since, the null model is represented by an equilibrium standard neutral model, θ can be estimated using the estimator θ_{W} . Figure S2 shows that the cutoff value of the ω -statistic decreases as S_n increases and the opposite tendency is seen for the *SweepFinder* statistic.

Consistent with previous studies (JENSEN *et al.* 2007) a selective sweep is discriminated easily from the standard neutral model. Indeed as illustrated in Figure 4A, the ω_{MAX} and Λ_{MAX} are distributed to a large extent distinctly even for relatively small values of α (*e.g.*, 500). Results are summarized in Table 1. Next, the distance between the true target of selection and the predicted target of selection is estimated (Table 1). The ω -statistic is more accurate than the *SweepFinder* and the median distance from the target of selection is about 0.5 kb. However, the performance of *SweepFinder* in discriminating the two scenarios is higher. Combining *SweepFinder* with the ω -statistic increases the classification performance (last column in Table 1).

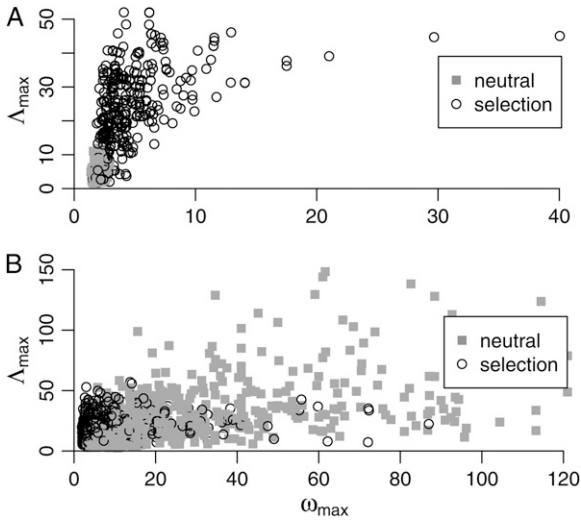


FIGURE 4.—The joint distributions of Λ_{MAX} and ω_{MAX} in scenarios with and without selection. (A) We compare the joint distribution of Λ_{MAX} and ω_{MAX} between a model with selection ($\alpha = 500$) in a constant population and a standard neutral model. The overlap between the distributions is limited and the scenarios can be discriminated by the *SweepFinder* (y -axis) and to a lesser extent by the ω -statistic (x -axis). (B) We compare a model with selection ($\alpha = 500$) with a neutral model that has experienced a bottleneck as it has been inferred by LI and STEPHAN (2006). Neither of the statistics can discriminate accurately the two scenarios (see also Table 2). Note that the scales of the statistics are different in A and B.

SHH in equilibrium populations vs. neutrality in nonequilibrium populations: Using simulations, selective sweeps have been generated as described above. For realizing past bottleneck events we used the LI and STEPHAN (2006) demographic history for the European population of *D. melanogaster*. We follow a similar approach as described in the previous section to assess the cutoff value. However, since the null hypothesis is not represented by the standard neutral model, θ_W is not an appropriate estimator of θ . Instead, we use the generalized unbiased estimator $\hat{\theta} = 2S_n/E(T_c)$, where $E(T_c)$ is the expected total length

of the coalescent of n sequences (ZIVKOVIC and WIEHE 2008). $E(T_c)$ depends only on the demographic history of the population.

For large values of α ($\alpha = 2500$) the true positive rate of the statistics ω_{MAX} and Λ_{MAX} is greater than 70% when the false positive rate is 18% (Table 2). For the same false positive rate, the true positive rate of the modified version of *SweepFinder* is above 90%. However, when smaller selection coefficients (e.g., $\alpha = 500$) define the hitchhiking effect, the selective sweep may be inseparable from bottleneck scenarios similar to that inferred by LI and STEPHAN (2006), using the original version of *SweepFinder* or the ω -statistic (TP rates < 10%, Table 2 and Figure 4B). The modified version of *SweepFinder* has a larger discrimination performance (true positive rate $\sim 40\%$). The low discrimination performance is indicated by the resemblance of genealogies between bottleneck models and selective sweeps in constant populations (see also THEORETICAL ANALYSES). The distributions of ω_{MAX} and Λ_{MAX} are largely overlapping as illustrated in Figure 4B. The SVM approach performs considerably better than any of the tests alone. The true positive rate is 75% when the false positive is 26% (Table 2). The main reason for the superior performance of the SVM approach is that it uses information about the distance of the peaks. In the scenarios with selection the target can be predicted accurately (Table 2); therefore, the distance between the peaks is considerably smaller than in the neutral scenarios.

SHH vs. neutrality in nonequilibrium populations: In this section we examine the statistical performance of the neutrality tests to detect selection in a genomic region and assess the distance between the true and the predicted targets of selection. We focus on two bottleneck scenarios. The first one describes a deep and short-lasting bottleneck (model A), whereas the second scenario describes a shallow and long-lasting bottleneck (model B). In both cases the severity (i.e., the product depth \times length) is the same ($=0.375$ in units of $4N$), and the bottleneck begins (backward in time) at 0.01. The present effective population size is assumed 10^6 , and the simulated region 50 kb. The recombination rate ρ for

TABLE 1
Equilibrium neutrality vs. selection in equilibrium populations

Model parameter	Performance	SF	SF*	ω	ω^*	SVM
$\alpha = 500$	TP (FP = 0.03)	0.85	0.97	0.13	0.14	0.9
$\alpha = 500$	Median distance (bp) from target (SD)	1728 (5597)	754 (1333)	528 (480)	540 (525)	—
$\alpha = 2500$	TP (FP = 0)	0.97	0.99	0.82	0.85	0.98
$\alpha = 2500$	Median distance (bp) from target (SD)	5383 (4509)	4582 (3905)	789 (657)	794 (680)	—

Using the SVM approach a false positive rate (FP) is estimated. For this FP rate, the true positive rates (TP) of the various neutrality tests are compared. The median distance and the standard deviation (SD) are also shown. SF, original *SweepFinder*; SF*, modified *SweepFinder*; ω , ω algorithm with constant-size windows; ω^* , ω algorithm with variable-size windows.

TABLE 2
Nonequilibrium neutrality vs. selection in equilibrium populations

Model parameter	Performance	SF	SF*	ω	ω^*	SVM
$\alpha = 500$	TP (FP = 0.26)	0.1	0.41	0.04	0.03	0.75
$\alpha = 500$	Median distance (bp) from target (SD)	899 (878)	522.982 (824)	423 (428)	603 (513)	—
$\alpha = 2500$	TP (FP = 0.18)	0.73	0.93	0.72	0.74	0.84
$\alpha = 2500$	Median distance (bp) from target (SD)	3065 (3209)	2074 (3361)	917 (1653)	956 (1629)	—

the whole region is set to 500. In the deep bottleneck scenario, the depth

$$\frac{\text{present population size}}{\text{bottlenecked population size}} = 500$$

and the length 0.00075.

In the shallow bottleneck scenario, the depth equals 20 and the length 0.01875.

Neutral simulations have been performed using Hudson's *ms* (HUDSON 2002) and simulations with selection using the *mbs* algorithm (TESHIMA and INNAN 2009). The design of simulations is as follows. In both cases we fix the number of polymorphic sites (=50) by employing broad uniform priors on θ and accepting only those instances that result in 50 segregating sites. This is justified by the dependence of the ω -statistic and *SweepFinder* on the number of segregating sites (Figure S2 and Figure S3) and the large variance on segregating

sites that neutral bottleneck scenarios generate. Furthermore, the rejection process guarantees that the total length of the tree, the posterior θ values, and the number of segregating sites are coupled. The 25th and 75th quantiles of the posterior distribution of θ are 32 and 52, respectively, for the deep-bottleneck scenario and 32 and 48 for the shallow scenario; therefore, the ratio ρ/θ is close to 10. In the simulations with selection, we examine scenarios of selective sweeps occurring recently (between the present and the bottleneck, sweep in phase 1), within the bottleneck (sweep in phase 2), and after the bottleneck (backward in time, sweep in phase 3). The parameters of the models with selection are described in Table 3 and Table 4 for the deep and shallow models, respectively. Similar to the neutral cases, a broad uniform prior on θ has been used, and we condition on observing 50 segregating sites. The posterior range of θ depends on the timing of the selective sweep; therefore, the ratio ρ/θ is close to 10

TABLE 3
Neutrality vs. selection in nonequilibrium populations (deep bottlenecks)

Model Parameter	Performance	SF	SF*	ω	ω^*	SVM
Sweep in phase 1	TP (FP = 0.51)	0.64	0.66	0.39	0.49	0.71
	Median distance (bp) from target (SD)	10813 (6768)	10497 (6832)	11986 (6595)	10239 (6186)	—
	Random target distance (SD)	11053 (6827)	11308 (6803)	11575 (6645)	11944 (6945)	—
Sweep in phase 2	TP (FP = 0.20)	0.62	0.64	0.36	0.44	0.73
	Median distance (bp) from target (SD)	9666 (6531)	10828 (6896)	11854 (6500)	10469 (6123)	—
	Random target distance (SD)	11508 (6885)	11397 (6808)	11877 (6750)	11555 (6804)	—
Sweep in phase 2*	TP (FP = 0.08)	0.72	0.78	0.63	0.12	0.97
	Median distance (bp) from target (SD)	9512 (6659)	10986 (6977)	10905 (6482)	11328 (6487)	—
	Random target distance (SD)	12067 (6983)	12265 (6920)	11647 (6950)	13236 (7213)	—
Sweep in phase 3	TP (FP = 0.56)	0.53	0.55	0.48	0.46	0.63
	Median distance (bp) from target (SD)	10377 (6831)	10845 (6833)	11342 (6662)	10624 (6541)	—
	Random target distance (SD)	12202 (6908)	11641 (6860)	12151 (6920)	12220 (6824)	—

A deep bottleneck, named model A, is examined. The ratio

$$\frac{\text{present population size}}{\text{bottlenecked population size}} = 500$$

and the length of the bottleneck is 0.00075. A beneficial mutation may appear within each phase of this three-epoch model (where time is measured backward in units of $4N$ generations): a recent sweep at time 0.01 (sweep in phase 1), a sweep within the bottleneck at time 0.0107 (sweep in phase 2), and an old sweep at 0.115 (sweep in phase 3). The selection coefficient is 0.002. Additionally, in the "sweep in phase 2*" model we describe a sweep that completes within the bottleneck ($s = 0.8$). The true positive rates of the neutrality tests are shown for each sweep model. The other rows depict the distance between the predicted and true targets and the random expectations for the distance.

TABLE 4
Neutrality vs. selection in nonequilibrium populations (shallow bottlenecks)

Model Parameter	Performance	SF	SF*	ω	ω^*	SVM
Sweep in phase 1	TP (FP = 0.27)	0.46	0.49	0.22	0.25	0.5
	Median distance (bp) from target (SD)	10116 (6872)	10691 (7001)	10268 (6658)	10868 (6670)	—
	Random target distance (SD)	11604 (6862)	11452 (6835)	10744 (6895)	11192 (7115)	—
Sweep in phase 2	TP (FP = 0.22)	0.58	0.56	0.27	0.32	0.6
	Median distance (bp) from target (SD)	10233 (6866)	11059 (6807)	11659 (6721)	11531 (6643)	—
	Random target distance (SD)	11725 (6889)	11375 (6855)	10846 (6829)	11245 (6882)	—
Sweep in phase 2*	TP (FP = 0.35)	0.67	0.74	0.65	0.4	0.67
	Median distance (bp) from target (SD)	9610 (6814)	10148 (6962)	11356 (6683)	10680 (6539)	—
	Random target distance (SD)	11906 (6889)	12102 (6846)	12432 (6894)	11583 (7079)	—
Sweep in phase 3	TP (FP = 0.25)	0.4	0.38	0.23	0.27	0.46
	Median distance (bp) from target (SD)	10232 (6710)	10447 (6744)	11693 (6965)	10829 (6625)	—
	Random target distance (SD)	11372 (6906)	11574 (6857)	11666 (6817)	13068 (6914)	—

A shallow bottleneck, named model B, is examined. The ratio

$$\frac{\text{present population size}}{\text{bottlenecked population size}} = 20$$

and the length of the bottleneck is 0.01875. A recent sweep at time 0.01 (sweep in phase 1), a sweep within the bottleneck at time 0.0107 (sweep in phase 2), and an old sweep at 0.115 (sweep in phase 3) are described. The selection coefficient in the model “sweep in phase 2*” is 0.1.

when the sweep is either recent or old, but it decreases when the selective sweep occurs within the bottleneck phase.

First, we examined the performance of the ω -statistic and *SweepFinder* to detect whether a genomic region has been shaped by positive selection. Results are presented in Table 3 and Table 4. For all comparisons, we used the false positive rate that is reported by the SVM. Then, we compare the TP rates between the various tests; the performance of a test is better when the TP rate is higher. The combination of *SweepFinder* and ω -statistic performs better than each test (SVM column in Table 3 and Table 4). Also, *SweepFinder* outperforms the ω -statistic. In model A (deep bottleneck), when the sweep is either recent or old, the discrimination between neutral and selective models becomes problematic; when the false positive rate is about 50%, the true positive is as low as 70 and 63%, respectively, for the SVM approach. For the separate tests, the performance is even lower. This result suggests that recent or old selection in populations that have experienced deep bottlenecks cannot be discriminated from neutrality. However, when selection has occurred within the bottleneck phase, the false positive rate decreases to 20% and the true positive rate is 73% for the SVM and about 10% lower for the *SweepFinder* (Table 3, sweep phase 2). Higher discrimination performance is achieved when the sweep completes within the bottleneck (Table 3, sweep phase 2*), but this requires unrealistically high values of s .

In model B (shallow bottleneck), the discrimination performance is slightly better than that of model A. However, again the most challenging scenarios are either recent or old sweeps and the performance increases when the sweep occurs within the bottleneck phase

(Table 4). Finally, the distances between the true target and the predicted target of selection are estimated. For both models A and B the distance is large and close to random expectations (Table 4).

Distinguishing RHH from neutrality in equilibrium populations: In contrast to single selective sweep (SHH) models, recurrent selected substitutions occur randomly along a chromosome according to a time-homogeneous Poisson process at a rate ν per generation (KAPLAN *et al.* 1989; WIEHE and STEPHAN 1993; STEPHAN 1995). Well-known patterns of SHH models are modified under RHH. As an example, the SFS is skewed toward the rare variants; however, the excess of high-frequency-derived alleles decreases (KIM 2006; JENSEN *et al.* 2008). Previously, JENSEN *et al.* (2007) have shown that it is difficult to separate RHH models from neutrality on the basis of ω_{MAX} -values or site frequency spectrum statistics. We explore the same problem with our new versions of the ω -statistic and the *SweepFinder* algorithm. Using the software developed by JENSEN *et al.* (2008) we simulated 100-kb genomic regions for a given reduction of heterozygosity (WIEHE and STEPHAN 1993), namely $H_{\text{RHH}}/H_{\text{NEU}} = 0.05, 0.25, 0.5, 0.75, \text{ or } 0.95$. $H_{\text{RHH}}/H_{\text{NEU}}$ denotes the ratio of heterozygosity in the RHH model to the heterozygosity in the absence of selective sweeps. The selection coefficient $s = 0.0001$ or 0.01. The null hypothesis is represented by the standard neutral model.

The null model used for the *SweepFinder* calculations and represented by the SFS of the population prior to the selective sweep in the SHH cases (n-SFS) cannot be described precisely by the standard neutral model. The population size is assumed to be constant. However, since adaptive mutations occur according to a time-homogeneous Poisson process it remains obscure what

the “prior to the sweep” SFS should be. Here, we follow two approaches. First, we assume that the n-SFS is derived from the standard neutral model and second that the n-SFS is obtained from the genomic region itself. Clearly, both approaches are approximations. On one hand, using the standard neutral model we increase the sensitivity of the *SweepFinder*. On the other hand, the nucleotide polymorphism patterns of the genomic region under investigation have been shaped by selective sweeps, so the n-SFS forms a conservative null model with small sensitivity. However, if real data are consistent with the RHH model, the standard neutral model cannot be supported as a null model since the whole genome will be affected by recurrent sweeps.

When the n-SFS is derived from the data itself then the power of the *SweepFinder* is greater for small values (*e.g.*, 0.0001) than for large values (*e.g.*, 0.01) of the selection coefficient s (Figure S4). Even if this appears to be counterintuitive, it is reasonable because when s is small the footprints of the selective sweep are local, and a large fraction of the genome remains neutral. On the other hand, for large values of s almost the entire genomic region may be affected by RHH, contradicting the assumption of the *SweepFinder* test that only a small fraction of the genome has been shaped by positive selection (Figure S4).

Under RHH models selective sweeps occur in different genomic locations during the evolution of the population following a time-homogeneous Poisson process (WIEHE and STEPHAN 1993). When subgenomic data are analyzed it is possible that the target of selection is either inside or outside of the sequenced genomic region. Furthermore, since selective events occur with a certain probability per generation (WIEHE and STEPHAN 1993), patterns of polymorphism are shaped by both old and new selective events. However, the ω -statistic and *SweepFinder* are based on the assumption that a single selective sweep has just been completed. Thus, it is important to test whether the algorithms are able to predict the correct position of the adaptive events.

Incorporating a fraction of monomorphic sites into *SweepFinder* analysis increases the precision of the algorithm (Figure S5). Similarly, the variable-size sliding window approach appears more accurate than the constant-size sliding window method for high cutoff values. When $H_{\text{RHH}}/H_{\text{NEU}} = 0.25$, *SweepFinder* and the ω -statistic predict that a target of selection is within a 5-kb distance from a true selective sweep position in about 40% of the cases. However, this fraction becomes smaller for higher values of $H_{\text{RHH}}/H_{\text{NEU}}$ (Figure S5).

DISCUSSION

The demography of natural populations: A major challenge of population genomics studies is to identify the loci that driven by positive selection contribute to the adaptation of natural populations

and to localize the beneficial mutation accurately (KIM and STEPHAN 2002; SABETI *et al.* 2002; JENSEN *et al.* 2005; NIELSEN *et al.* 2005; AKEY 2009; NIELSEN *et al.* 2009; PICKRELL *et al.* 2009). To address these questions, it is important to consider the demographic history of the population, as this neutral nonequilibrium model represents the null (LI and STEPHAN 2006; THORNTON and ANDOLFATTO 2006). Since the standard neutral model does not reflect accurately the demography of most natural populations, neutrality tests should not be performed using the standard neutral scenario as the null model. In this study, we examined two bottleneck scenarios that are relevant to the demographic history of the European population of *D. melanogaster* (LI and STEPHAN 2006; THORNTON and ANDOLFATTO 2006). The properties of the coalescent trees that underlie these demographic models differ considerably. In a recombining genomic region, the model inferred by LI and STEPHAN (2006) produces both star-like short coalescent trees and genealogies with long internal branches. Star-like genealogies are generated less frequently by the THORNTON and ANDOLFATTO (2006) model (Figures 1 and 2). As a consequence, the null distributions of the neutrality statistics may differ. Thus, inferring the demographic history of a population is a prerequisite for performing genomic scans for selective sweeps, which has been shown to be a challenging task (MYERS *et al.* 2008).

Separating single selective sweeps from neutral models: When the value of the selection intensity α is large, the joint distribution of Λ and ω overlaps only partially between a model of selection in an equilibrium population and the bottleneck model inferred by LI and STEPHAN (2006). However, for smaller values of α the two distributions overlap greatly. A useful approach for classifying an observation as either a neutral or selective model is by combining the Λ and ω profiles. Here, we use the distance between the peaks and the correlation of ω and Λ . These features can be used in a classifier (*e.g.*, SVM). Training requires that there are known instances of both neutral and selective models. For simple selective and neutral models this is currently possible, using coalescent-based programs. However, it remains challenging for more complicated scenarios. Forward simulations provide greater flexibility when selective events occur in nonequilibrium populations and they can be used efficiently when the population size is relatively small (*i.e.*, on the order of thousands) or diffusion scaling applies (HOGGART *et al.* 2007; CHADEAU-HYAM *et al.* 2008; HERNANDEZ 2008).

The rationale for employing combinations of Λ and ω is that under a selective model the two statistics assume high values close to the target of selection. This implies that the target of selection can be localized accurately. Under selection models in equilibrium populations this assumption is met even for small α values. Modifying

SweepFinder to include a fraction of nonpolymorphic sites in the analysis increased the accuracy of the algorithm and the performance in separating neutral scenarios from scenarios with selection. Furthermore, both versions of the ω -statistic, the constant- and the variable-size sliding window approach, are very accurate for selection models in equilibrium populations.

However, in severe nonequilibrium scenarios (*e.g.*, the estimated bottlenecks of LI and STEPHAN 2006 and THORNTON and ANDOLFATTO 2006), when selection and past demographic changes occur within the same model, the target of selection cannot be predicted, neither by *SweepFinder* nor by the ω -statistic. The accuracy of the target prediction when a selective sweep has occurred within the bottleneck period is comparable to that of randomized experiments. The reason is that polymorphism valleys and short coalescent trees may extend over large genomic regions, and the often used sweep signature of an excess of high-frequency-derived alleles vanishes. This result should be taken into account when regions of strong and recent positive selection are identified in genome scans. Since natural populations can be described by equilibrium demographic models only rarely, the true target of selection may be tens of kilobases away from the predicted target.

In the case of a severe bottleneck, such as model A, recombinants (carrying the selected mutation and the derived neutral allele) are most likely formed in the early period of the selective phase (forward in time), but they will be lost with high probability due to drift after the population size crashes. Therefore, high-frequency-derived variants may not be observed. In contrast, the frequency of rare variants (singletons) will dramatically increase. Therefore, on the basis of site frequency spectrum it is possible to discriminate, to some extent, neutral from nonneutral scenarios (Table 3).

The analysis of the likelihood curves of *SweepFinder* can provide further insights into the technical reasons that, in the cases of selection in nonequilibrium populations, make the prediction of the target of selection challenging. *SweepFinder* implements a model of selective sweep, which assumes that each observed SNP existed prior to the sweep. It uses the compound parameter $\gamma = (r/s)\log(2N)$ (named α in NIELSEN *et al.* 2005) and the position x where the selective event occurred. (Here r denotes the recombination rate per basepair.) As Figure S6 illustrates, low- and high-frequency SNPs affect the likelihood in a similar way by contributing high values in the proximity of the sweep. Examining how the SFS changes over a genomic region under an equilibrium demographic model with selection and the THORNTON and ANDOLFATTO (2006) model with selection ($\alpha = 2500$), it is apparent that there is a dramatic increase of the class $n - 1$ in the proximity of the selective sweep in the equilibrium model (Figure 3), but a very slight change of singletons in the nonequilibrium model. In the equilibrium-model case the precise

localization of the sweep is possible, due to the spatial patterns of the rare and high-frequency-derived variants. However, in the THORNTON and ANDOLFATTO (2006) model with selection this pattern vanishes, the high-frequency-derived variants disappear and the singletons spread over the whole genomic region. Thus, the target of selection cannot be estimated accurately.

It should be noted, however, that the poor performance of *SweepFinder* and the ω -statistic under the nonequilibrium models (bottlenecked populations with selection) does not imply that the performance of the tests is poor under any nonequilibrium model with selection. These models represent extreme cases that violate major assumptions of the algorithms. The slightly improved performance of the machine-learning approach is due to the fact that it uses information from the sweep scenarios and, furthermore, it combines information from both the ω -statistic and *SweepFinder*.

Studying a scenario where a selective event took place in a bottleneck period is of great biological importance. Often, population bottlenecks are associated with a major migration event. For example, the bottleneck inferred by LI and STEPHAN (2006) for the European population of *D. melanogaster* describes the colonization of Europe from the African ancestral population. Therefore, positive selection may have occurred in the new habitat that contributed to the adaptation of flies to the environmental conditions of Europe. As Tables 3 and 4 show, the performance of the tests (especially the SVM, and to a lesser extent, the *SweepFinder*) is high when the sweep occurs within the bottleneck. This suggests that the approaches tested in this study can be used for the detection of selective sweeps in populations that have recently migrated to new environments. Furthermore, Tables 3 and 4 suggest that the power of SFS-based tests is higher than LD-based tests.

A difficulty that arises from using simulations with selection to train the algorithms is that the parameters of the scenarios with selection are unknown, *i.e.*, the selection intensity α , the position of the sweep x , and the time at which the sweep occurred. In the models that we presented it was assumed that these parameters are known. However, when real data are analyzed these parameters are generally unknown, and moreover no methods that can estimate them in scenarios with past demographic changes are available. Thus, heuristic approaches have to be used. First, the position x can be assumed to be in the center of the fragment. Then, in real-data analysis overlapping windows should be used so that windows where x is located near their center will exist. The time of the sweep should be recent ($< 0.1N$). In the classical approach this parameter is also implicitly specified by assuming that the sweep has just been complete. Finally, the selection intensity can be drawn from a prior uniform distribution. In this case the training set is composed of a mixture of models with various selection intensities.

Recurrent selective sweep analysis: Recurrent selective sweeps invalidate the assumption that a single hitchhiking event has just been completed. In agreement with JENSEN *et al.* (2007), we find that for greater rates v of selective events per generation the power of the tests increases for a given $H_{\text{RHH}}/H_{\text{NEU}}$. One possible explanation is that for smaller v a few strong selective sweeps that affect a large portion of the genome and shift the SFS of large genomic regions have occurred. Thus, the local characteristic of the signature of a selective event is lost. Another possible explanation is that for smaller v the selective events are old on average and the signature of selective sweep has faded away (JENSEN *et al.* 2007).

The variable-size sliding window approach increases the accuracy of the ω -statistic to predict the target of selection. However, the performance is still poor. In $\sim 20\%$ of the peaks above a certain threshold found in a scan of a given genomic region, the real position of the sweep is located within a 5-kb distance. The performance of the constant-size sliding window is about half that of the variable-size approach and comparable to the randomization experiments. A similar improvement has been achieved with the modified *SweepFinder* algorithm. RHH models imply that adaptive substitutions occur at a time-homogeneous rate, *i.e.*, uniformly in the history of the population. This assumption may be violated in domesticated populations or in populations that experienced environmental changes. Thus, an increase of the performance of the tests (lower false positive rate, greater accuracy in target prediction) may result when RHH models are incorporated within the *SweepFinder* or the ω -statistic algorithms.

Recurrent selective sweep parameters such as the rate v of adaptive substitutions and the decrease of heterozygosity have been estimated recently. JENSEN *et al.* (2008) and LI and STEPHAN (2006) have estimated that heterozygosity has decreased in genomic regions of normal recombination by 50% whereas the estimate of MACPHERSON *et al.* (2007) and ANDOLFATTO (2007) is about 20% (*i.e.*, $H_{\text{RHH}}/H_{\text{NEU}} = 0.8$). We examined the performance of the *SweepFinder* and the ω -statistic for various levels of heterozygosity reduction, $H_{\text{RHH}}/H_{\text{NEU}} = 0.25, 0.5, 0.75,$ and 0.95 , and selection coefficients $s = 10^{-2}$ and 10^{-4} (Figure 5). The power of *SweepFinder* is greater for the LI and STEPHAN (2006) and JENSEN *et al.* (2008) estimations than for that of MACPHERSON *et al.* (2007) and ANDOLFATTO (2007), given that selection is strong ($s = 10^{-2}$). For $s = 10^{-4}$ the differences in the performance of *SweepFinder* for various levels of $H_{\text{RHH}}/H_{\text{NEU}}$ are small. The reason is that for $s = 10^{-4}$ the diversity is similar for values of $H_{\text{RHH}}/H_{\text{NEU}}$ between 0.05 and 0.95. This may be due to inaccuracies of the RHH theory when s is small or due to the stochastic trajectory of the beneficial mutation (COOP and GRIFFITHS 2004; SPENCER and COOP 2004).

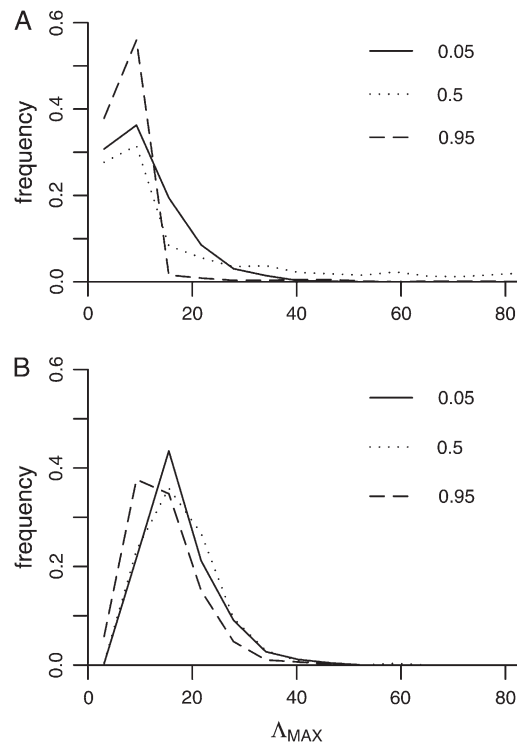


FIGURE 5.—The distributions of Λ_{MAX} for various levels of the decrease of heterozygosity and $s = 10^{-2}$. Each distribution is discrete and the size of each bin has been set to 6. (A) For $H_{\text{RHH}}/H_{\text{NEU}} = 0.05, 0.5,$ and 0.95 the cutoff values (95th percentile) are 5.7, 9.7, and 11.9, respectively, and the sensitivities of the test (percentage of true positives) given the cutoff values are 0.74, 0.48, and 0.07. The power of *SweepFinder* is greater for the LI and STEPHAN (2006) and JENSEN *et al.* (2008) estimations than those of MACPHERSON *et al.* (2007) and ANDOLFATTO (2007) because selection is strong ($s = 10^{-2}$). (B) When $s = 10^{-4}$ the amount of diversity is similar for $H_{\text{RHH}}/H_{\text{NEU}} = 0.05, 0.5,$ and 0.95 . Therefore, the performance of *SweepFinder* is relatively independent of the $H_{\text{RHH}}/H_{\text{NEU}}$.

Time of the selective sweep: For SHH models (in demographic equilibrium) we assume that the selected mutation has reached fixation very recently. The selective model that underlies the *SweepFinder* algorithm assumes a recent and strong selective sweep. Therefore, the power of *SweepFinder* is expected to be higher for recently completed hitchhiking effects. Indeed, simulations have shown that the power decreases exponentially after the selective sweep (P. PAVLIDIS, unpublished results). It should be mentioned that the demographic scenario that follows the selective sweep (*i.e.*, between the time of completion of the selective sweep and the time of sampling) affects the performance of *SweepFinder*. Simulations have shown that if the completion of a selective sweep is followed by population expansion, the performance of the likelihood ratio test implemented in *SweepFinder* remains high even after the completion of the selective sweep (P. PAVLIDIS, unpublished results). The rationale behind this is that a population expansion decreases the coalescent rate; therefore, the return to

the equilibrium SFS is slower and the signature of the selective sweep is preserved for a longer period. In contrast to *SweepFinder*, which is based on the low- and high-derived variants, the ω -statistic is more sensitive to the time since the completion of the selective sweep. Indeed, the LD pattern captured by the ω -statistic vanishes rapidly (JENSEN *et al.* 2007), comparable to the fixation rate of the high-frequency-derived alleles (KIM and STEPHAN 2000; PRZEWORSKI 2002; JENSEN *et al.* 2007).

Overlapping selective sweeps: In this study we focused on nonoverlapping selective sweeps. The RHH model we have used describes successive and nonoverlapping selective events. CHEVIN *et al.* (2008) have shown that two interfering selective sweeps may modify the pattern of linked neutral variation. A related process, when the targets of selection are located closely to each other in the genome, causes trafficking (KIRBY and STEPHAN 1996; KIM and STEPHAN 2003). A most extreme scenario, which describes the appearance of beneficial mutations at the same site, is described as “soft” sweep (HERMISSON and PENNINGIS 2005). Soft sweeps may emerge during the evolution of organisms (*e.g.*, *Plasmodium*) with high mutation rates (NAIR *et al.* 2007). Conversely, they may be of limited importance in the evolution of *D. melanogaster* or *Homo sapiens*, for instance. The patterns of neutral variation under these selective scenarios are different from those of single selective events. For example, the skew of Tajima’s *D* toward negative values vanishes in the interference scenarios described by CHEVIN *et al.* (2008) and can even be positive between the selected sites. In general, SFS-based approaches may not work under overlapping selective sweeps because the frequency of the class of polymorphisms in intermediate frequency may be quite large. In such cases, LD-based statistics can be useful because a multitude of extended haplotypes may exist on the left and right sides of the selected region (SABETI *et al.* 2002; VOIGHT *et al.* 2006; TANG *et al.* 2007).

Machine-learning approaches in population genetics: Machine-learning approaches are widely used in a variety of applications from image processing to classification of microarrays. Here, we are interested in the subfield of machine learning that is related to supervised learning or classification. Typically, in a classification problem a training set teaches the algorithm to predict the class label of an input object (DUDA *et al.* 2000; HASTIE *et al.* 2001). The goal is to decide between a selective and a neutral model. However, classifying a data set as either neutral or selective is challenging because the parameters of the neutral and selective models are unknown. Therefore, parameter estimation is required prior to the classification. In the cases that an equilibrium model with selection is employed, the selection intensity α can be estimated using the *clsW* software (KIM and STEPHAN 2002) or the *SweepFinder* algorithm (given that ρ is known). To our knowledge,

currently the only method able to estimate α given a nonequilibrium (stepwise) model with selection has been developed by LI and STEPHAN (2006). On the other hand, several approaches exist for the estimation of parameters in a neutral demographic model (NIELSEN 2000; EXCOFFIER *et al.* 2005; LI and STEPHAN 2006; HEY and NIELSEN 2007). Usually, these approaches require multiple loci to infer the demographic parameters of a population. The next step in a classification problem is feature selection, which aims at using a subset of the features available from the data. Here, Λ_{MAX} , ω_{MAX} , and their combinations (distance between peaks and correlation of ω and Λ) have been used. Combining ω and Λ is powerful in comparisons between equilibrium models with selection and neutral nonequilibrium models when the selection intensity is small (Table 2). Alternatively, various summary statistics, such as TAJIMA (1989)’s *D*, FAY and WU (2000)’s *H*, or Z_{nS} (KELLY 1997) can be used. Our choice is based on the fact that *SweepFinder* uses SFS information whereas the ω -statistic is based on LD. The choice of the classification technique is important and depends on the problem and the nature of the data. Here, we demonstrate an application using the SVM classifier (with the radial kernel), as it is implemented in the e1071 package of the R-project. To our knowledge, there are no studies on separating neutral from selective scenarios that use supervised-learning approaches. Future work will provide insight into the feature selection problem and will also evaluate the performance of the supervised-learning approaches.

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