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Larger Body Mass Index and Waist Are Associated with Lower Mortality in the Chinese Long-Term Care Facility Residents

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Abstract

Objectives—We aimed to investigate the association between BMI and WC and all-cause mortality of Chinese residents in long-term care facilities in Taiwan.

Design—Prospective cohort study.

Setting—Eight long-term care facilities in Taiwan.

Participants—Three hundred and fifty-four residents aged 60 and above (men=156, women=198; median: 78.4 years, range: 60–101 years) were recruited during the study period.

Measurements—Anthropometrics and metabolic parameters were measured at baseline. The range of BMI and WC in this study was from 11.6 to 35.3 kg/m² (mean±SD=21.7±4.2) and from 55.0 to 124.0 cm (mean±SD=82.4±10.9), respectively. Mortality data was from the Department of Health in Taiwan.

Results—During the 5 years of follow-up, there were 219 deaths. After adjusting for age, gender, albumin, Karnofsky performance status scale, hypertension, and diabetes, subjects in the highest quartile of BMI (mean±SD =27.3±2.8 kg/m²) and WC (mean±SD =96.7±7.4cm) had a significantly lower mortality rate than did subjects in the lowest quartile (BMI, mean±SD

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=16.7 \pm 1.7kg/m²; WC, mean \pm SD =69.6 \pm 4.2cm). After further stratification by central obesity status, the subjects in the 2 highest BMI quartiles had a lower mortality rate than those in the lowest BMI quartile, but only in the central obesity group (\geq 90cm in men or \geq 80cm in women). The adjusted relative risk for all-cause mortality in the highest vs. lowest BMI quartile was 0.17 (95% CI: 0.05–0.57).

Conclusion—BMI and WC were negative predictors for all-cause mortality in Chinese elderly living in long-term care facilities. Subjects with higher WC and BMI had lower all-cause mortality rate.

Keywords

body mass index; waist circumference; elderly; mortality; institutionalized

INTRODUCTION

According to the World Health Organization (WHO), 1.6 billion adults were overweight and at least 400 million adults were obese globally in 2005^1 . Prevalence of obesity increases with the increase in age. In the United States, 73.9 % of men and 68.4 % of women age 60 and above were either overweight or obese (defined as body mass index(BMI) $\geq 25 \text{kg/m}^2$). In China, the prevalence of obesity (BMI $\geq 25 \text{kg/m}^2$) among individuals age 20–29 is 21.3%, but it increases to 46.2% among age 60–69. The increased prevalence of obesity represents a great health burden globally.

Many studies have found a U- or a J-shaped association between BMI and mortality among adults 3^{-7} . The relationship, however, remains controversial in older populations. Studies found that the influence of obesity on mortality is reduced with increasing age and the optimal weight and waist circumference (WC) for older populations may be higher $^{8-10}$. In 2007, in a review of 28 studies done by Janssen et al indicated that overweight individuals were not at increased risk for mortality, while obese individuals had only a modest increase in mortality risk 11. Moreover, recent studies reported that BMI was actually inversely related to mortality in the older subjects 12^{-13} . Baumgartner et al proposed a potential explanation for this paradox, i.e. that total mortality was increased at high body fat mass (FM) and at low fat-free mass (FFM) $^{14-15}$.

Most studies investigating the relationship between BMI and WC and mortality have been carried out in community-dwelling elderly populations, not in long-term care facilities. The number of people living in long-term care facilities has been rapidly growing in many countries, including Taiwan16. The prevalence of obesity may also increase in long-term care settings. There are only a limited number of studies assessing the relationship between BMI and WC and mortality in residents living in long-term care facilities17⁻¹⁸. Therefore, the aim of this study was to prospectively evaluate the relationship between BMI and all-cause mortality in older Chinese adults living in long-term care facilities in Taiwan. We also used WC (reflecting the degree of central obesity) or calculated resting energy expenditure, REE (reflecting the degree of lean body mass (LBM)) to substitute for BMI in our analyses.

METHODS

Study subjects

The target population was residents living in 8 long-term care facilities which cooperated with the largest tertiary hospital in an urban city in Taiwan during 2002–2003. There were 447 residents living in these 8 facilities during study period. Among these 447 residents, 393 residents aged 60 years old and above were invited to participate in this study. Finally, 354

subjects (men=156, mean age \pm SD (standard deviation) = 76.4 \pm 7.6 years; women =198, 79.9 \pm 7.7 years) agreed to participate in this study were recruited. Deaths were ascertained by computer linkage to the national death registry using ID number. All deaths that occurred between study entry and December 2007 were included. Ethics approval for patient recruitment and data analyses was obtained from the Institutional Review Board.

Anthropometric index

Trained staff measured the weight, height, and WC. WC is taken at the midway point between the inferior margin of the last rib and the iliac crest in a horizontal plane. BMI is calculated as weight (kg) divided by height squared (m²). BMI and WC were divided into quartiles by genders as follows. BMI quartiles I–IV: <18.5, 18.5–20.5, 20.6–23.6, >23.6 kg/ m² in men; <19.4, 19.4–21.4, 21.5–25.0, >25.0 kg/m² in women; WC quartiles I–IV: <73.6, 73.6–81.0, 81.1–87.4, > 87.4 cm in men; <75.5, 75.5–81.5, 81.6–89.1, > 89.1 cm in women. BMI was also divided into four levels according to the obesity definition using Taiwan and WHO for Asians criteria 20 .

Biomedical markers

Blood pressure (BP) was measured by the same trained staff on the right arm using an appropriately sized cuff and a standard mercury sphygmomanometer in a seated position. A venous blood sample was taken after a 12-hr fast for determination of plasma glucose, albumin, and lipid profile using a biochemical autoanalyzer (Beckman Cou, Fullerton, CA, USA) at the Clinical Laboratory Department of China Medical University Hospital.

Resting energy expenditure (REE)

REE was calculated according to Liu's predictive equation for Chinese²¹ as shown below: REE (kcal/day) = $13.88 \times$ (weight in kg) + $4.16 \times$ (height in cm) - $3.43 \times$ (age in year) - $112.4 \times$ (gender) + 54.34, where gender is 0 for men and 1 for women.

REE were divided into quartiles by genders as follow: I–IV: <1078, 1078–1201, 1202–1381, > 1208 kcal/day in men; <835, 836–939, 940–1041, > 1041 kcal/day in women.

Questionnaire, diabetes, hypertension, and Karnofsky performance status scale (KPS)

Each participant completed a structured questionnaire. Medical records were reviewed by physicians. Smoking history was divided into 3 classes as follows: never, former, and current. Diabetes was defined as: (1) fasting glucose \geq 126 mg/dL and/or (2) diabetes history and on oral hypoglycemic agents or insulin treatment. Hypertension was defined as: (1) systolic BP \geq 140 mmHg and/or diastolic BP \geq 90 mmHg and/or (2) hypertension history and on anti-hypertensive drugs. KPS is an assessment tool used to assist physicians and caretakers in measuring a patient's ability to carry out activities of daily living. KPS runs from 0–100, where 0 is death and 100 is perfect health 22 . KPS below 60 was defined as subjects who require some assistance22.

Statistical analysis

Data are presented as the mean (SD) for continuous variables. Log transformation was used for variables with significant deviation from normal distribution, assessed by the Kolmogorov–Smirnov test before further analyses. Analysis of variance (ANOVA) test was used for comparing mean values of continuous variables across BMI quartiles. Student's t test for unpaired data was used to compare mean values between two groups. Proportions and categorical variables were presented as percentage and tested by the χ^2 test and by the two-tailed Fisher's exact method when appropriate. Cox proportional hazard regression analyses adjusted for potential confounders were used to estimate the relative risks (RRs) for

all-cause mortality. Survival curves adjusted for other covariates were drawn for different BMI quartiles23⁻²⁴. All statistical tests were 2-sided at the 0.05 significance level. These statistical analyses were performed using the PC version of SPSS statistical software (13th version, SPSS Inc., Chicago, IL, USA).

RESULTS

During the 5-years of follow-up, there were 219 deaths. The overall 5-year mortality rate was 208.6 per 1000 person-years. The baseline characteristics of the population sample according to BMI quartiles are listed in Table 1. Compared to the lowest quartile of BMI, the higher quartiles of BMI tended to have greater weight, BMI, REE, WC, fasting glucose, total cholesterol (TCHOL), triglyceride (TG), albumin, prevalence of high KPS and diabetes, but lower age and high-density-lipoprotein cholesterol (HDL-C).

Table 2 shows the baseline characteristics of the survivors and those who died. The survivors were younger and had greater WC and albumin.

Table 3 shows all-cause mortality rate and RRs according to different BMI, REE, and WC quartiles, in unadjusted and adjusted analyses. The subjects in the lowest quartile of BMI, WC, and REE had the highest mortality rate (Table 3). Compared to the lowest quartile of BMI, REE, and WC, using Cox proportional hazard regression analyses, the unadjusted RRs for all-cause mortality were significantly decreased in highest BMI, REE, and WC quartiles (all *p* values < 0.05, Table 3). Knowing that age, nutritional status (such as albumin), lipid profiles (such as TCHOL, TG, HDL-C), diabetes, hypertension, and performance status (such as KPS) were risk factors for all-cause mortality, we further adjusted for these potential confounders. Compared to the lowest quartile, the adjusted RRs for all-cause mortality were significantly decreased in the highest quartile of BMI, REE, or WC (all *p* values< 0.05, Table 3). Compared to the lowest quartile of BMI, REE, and WC, the adjusted RRs for the highest quartile of BMI, REE, or WC were 0.62(95% confidence interval [CI]: 0.40,0.96), 0.61(0.38,0.99), or 0.58(0.37,0.91), respectively. We also found that BMI categories using Taiwan and/or WHO for Asians criteria were inversely related to mortality (Table 3).

The correlation between BMI and WC was high (Pearson correlation=0.789, p< 0.001) and the interaction between WC and BMI quartiles was significant (p< 0.001), so we further stratified WC into 2 categories according to the central obesity definition of WHO for Asians 20 . Figure 1 shows the association between BMI quartiles and mortality in the central obesity group (\geq 90 cm in men or \geq 80cm in women). Among the central obesity group, compared to subjects in the lowest BMI group (quartile I), the adjusted RRs among subjects in BMI quartiles II, III, IV were 0.56(0.17,1.82), 0.20(0.06,0.65), 0.17(0.05,0.57), respectively. Among the non-central obesity group, the adjusted RRs were 0.59(0.37,0.93), 0.95(0.56,1.60), and 0.93(0.47,1.86), respectively. We also stratified BMI into 2 categories according to the general obesity definition by WHO for Asians (BMI \geq 25kg/m² was defined as general obesity). Among the general obesity group, compared to subjects in the lowest WC group (quartile I), the adjusted RRs among subjects in WC quartiles II, III, IV were 2.07(0.15, 28.8), 0.41(0.04,4.59), and 0.35(0.04,3.51), respectively. Among nongeneral obesity group, the adjusted RRs were 0.60(0.38,0.94), 0.93(0.61,1.42), and 0.58(0.32,1.05), respectively(Figure not shown).

DISCUSSION

We have demonstrated that BMI and WC are inversely related to all-cause mortality in older Chinese adults living in long-term care facilities in Taiwan. We also have found that

residents with central obesity and higher BMI have a lower mortality than residents with central obesity and lower BMI. Even after adjusting for potential confounders, the results are similar. As the world is aging and the number of residents living in long-term care facilities is increasing, our findings are important for health care and policy makers affecting the elderly living in long-term care facilities.

The association between BMI and mortality has been identified as U- or J- shape among younger or middle-aged adults $^{3-4}$. This association in the elderly, however, is still controversial. A meta-analysis done by Janssen et al found that BMI was inversely associated with mortality risk in elderly Asians ¹¹. Compared to middle-aged adults, previous studies found that the optimal BMI among the elderly shifted upwards ^{3, 25}. The optimal range of BMI proposed among the elderly varied among studies, but ranged from approximately 25–27 kg/m² in Chinese ³ and 25–30 kg/m² in Caucasians ²⁵. The results of our study are similar to those of most but not all studies done in Asians $^{26-28}$. For example, a study done among Korean elderly ⁴ found that elderly subjects with underweight but not obesity had higher all-cause mortality compared to subjects with normal weight. However, another study done by Gu, in China, showed that obese (BMI ≥30 kg/m²) elderly subjects had higher total mortality compared to subjects with BMI between 24–24.9 kg/m² ³. Compared to this latter study, the subjects in our study had a lower mean BMI (BMI quartile IV mean(SD) was 27.3(2.8) kg/m² in our study vs. 32.3(2.6) kg/m² in the obese group of Gu's study). Actually, elderly subjects with BMI ≥25 and <30 kg/m² did not show a significant increase in all-cause mortality compared to subjects with BMI between 24-24.9 kg/m² in Gu's study. Age may be also a factor that could explain these differences. Studies done in Caucasians found that, in relatively younger elderly (age from 60 to 75 years), increased BMI was associated with an increase of all-cause mortality, whereas in older elderly (age over 75 years) higher BMI protected from mortality ²⁷. The mean age in our study subjects was 78.4 years and our results are consistent with those of other studies of older elderly 27.

Another aspect represents body composition. BMI and WC are the most common anthropometric indices used for predicting effects of obesity. However, the associations between BMI and body fat and fat distribution were weaker in the elderly than in the younger or middle-aged adults ^{11, 29} and previous studies found that excess body fat may have less influence on mortality in the elderly than in the younger or middle-aged adults ¹¹. A negative association between LBM and all-cause mortality has been observed in prospective studies ^{7, 30}. We used calculated REE as a surrogate for LBM and also found that REE is negatively associated with all-cause mortality in our population. It is reasonable to speculate that the subjects in the higher BMI quartiles in our study had greater LBM which may have offered some protection against mortality. This supports the hypothesis that the preservation of LBM may have greater importance than excess fatness with regard to survival in older elderly.

Studies on the association between all-cause mortality and WC among the elderly are scanty and controversial. Janssen et al found that WC was a positive predictor of mortality ³¹, but Price and Heitmann found that WC was not associated with all-cause mortality in older persons ^{30, 32}. Ours, however, may be the first study to show that WC is a negative predictor for all-cause mortality in the elderly. In Figure 1, we also found that high BMI was protective among those with central obesity, but not among those without central obesity. In other words, the association between BMI and mortality was stronger among the centrally obese. There are some possible explanations. First, it is the selective survival theory which states that obese individuals with susceptibility to the adverse effects may have already died at a younger age 11. Thus, elderly individuals with less susceptibility to the adverse effects are left. Second, our finding may be due to differences in population's race i.e. being

specific to elderly Chinese. Body composition differs in Caucasians vs. Asians, i.e. for a given BMI, Asians have greater body fat accumulation than Caucasian adults 33. Aginginduced adipose tissue increase differs among races. Wu found that, with increasing age, Asian women have more truncal fat than African-American or Caucasian women 34. Study on dialysis patients where obese had lower mortality than normal weight patients ³⁵, the authors proposed that, in these sick individuals, the adipose tissue may secrete protective cytokines and other hormonal products. This hypothesis, that body fat may actually be protective, merits further study in elderly Chinese. Third, our finding may be due to our unique study population living in long-term care facilities, not living in the community. There are some possible differences between these two populations which may influence the association between obesity and mortality. Malnutrition (such as low serum albumin) and poor performance status (such as low KPS) increased the risk of all-cause mortality in the instituted elderly ³⁶. We adjusted for these potential confounders into models in Table 3, and the inverse association still persisted. Chronic inflammation could also play an important role in the increased risk of all-cause mortality in the elderly. If we further adjusted for Creactive protein into models in Table 3, the results were similar (data not shown). Older subjects living in long-term care facilities tend to have more chronic diseases or undiagnosed diseases than subjects living in the community. Chronic diseases (such as diabetes and hypertension) were the major causes of death in the elderly, so we adjusted for these diseases in our models. From adjusted models in Table 3, we found that the inverse association between mortality and BMI and WC still persisted after these adjustments but was weaker. Meanwhile, malnutrition and cachexia are common in institutional settings. In the context of cachexia, an "obesity paradox" has been observed, such that obese persons have a survival advantage^{37–38}. This may be related to the large amount of unintentional weight loss experienced by residents of long-term care facilities and the possibility that physiologic reserve provided by fat is important for survival. Therefore, subjects with normal weight at baseline may have been previously overweight or obese but they may have lost weight due to undiagnosed disease ³⁹. To clarify this potential influence, we excluded those who died during the first 6 months of follow-up and the results remained the same. Finally, smoking decreases body weight and increases fat accumulation, so smoking may have influenced the association ¹¹. In our study, from Table 1 and 2, we found that smoking status was not a potential confounder. When we further adjusted for smoking status in the adjusted models in Table 3, the associations between BMI/WC/REE and mortality were still not changed. Although studies done by Bigaard and Heitmann found that both high body fat and low FFM were independent predictors of all-cause mortality ^{7, 30}, in the United States, Zhu found that FM had a negative relationship with mortality ⁴⁰ which is consistent with our study results. The interplay between WC and BMI on all-cause mortality deserves further research in the elderly.

There are some limitations to our study. First, we did not measure body composition (such as FM) among our subjects, so we cannot analyze the association between FM/FFM and mortality. However, since REE represents FFM, we reasonably assumed that subjects with high REE had higher FFM which led to lower mortality. Second, the study had limited follow-up period. The results may be confounded by previous BMI misclassification. However, we tried to exclude those deaths which occurred in the first 6 months and the results remained the same and believe that this effect of potential BMI misclassification was minimal. Third, our subjects were residents in long-term care facilities and we did not assess the proportion of indigenous Taiwanese in our population; the generalization to community-dwelling population or indigenous Taiwanese or people living in the rural area should be made with caution.

We conclude that, in the elderly living in long-term care facilities, both BMI and WC are negatively associated with mortality in Taiwan. This negative effect was enhanced among

centrally obese subjects. This finding may have great public health impact as the number of instituted elderly is increasing in many developed countries. Clinicians should regularly screen these residents using both BMI and WC. Further studies regarding the association between FM and FFM in the elderly living in long-term care facilities should be done.

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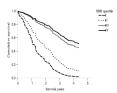


Figure 1.

The association between BMI quartile and risk of all-cause mortality in the central obesity defined by WHO for Asians using the Cox proportional hazard regression analyses after adjusting for age, gender, albumin, Karnofsky performance status scale, total cholesterol, triglycerides, high-density lipoprotein cholesterol, hypertension, and diabetes. As compared to lowest BMI group (quartile I), the adjusted relative risks were 0.56 (95% CI: 0.17, 1.82), 0.20 (0.06, 0.65), 0.17 (0.05, 0.57) among BMI quartile II, III, IV, respectively.

Table 1

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Baseline subjects' characteristics according to BMI quartiles by genders.

		BMI q	BMI quartile		Ь
Characteristic	I(n=87)	II(n=87)	III(n=92)	IV(n=88)	Value
$\operatorname{Men}\left[n\left(\%\right)\right]^*$	38 (43.7)	39 (44.8)	40 (43.5)	39 (44.3)	0.998
Age (years)†	79.6±8.2	80.0±7.6	77.5±7.7	76.4±7.4	9000
Height $(cm)^{\dagger}$, \sharp	152.4±7.5	153.3±7.5	152.6±8.7	153.4±7.5	0.743
Weight (kg)	38.8±5.1	47.1 <u>±</u> 4.2	52.8 ± 5.9	64.3 ± 8.2	<0.001
BMI $(kg/m^2, range)^{\dagger}$	16.7±1.7 (M:11.9–18.3; F:11.6–19.3)	20.0±0.8 (M:18.5–20.5; F:19.4–21.4)	22.6±1.1 (M:20.6–23.6; F:21.5–25.0)	27.3±2.8 (M:23.7–34.0; F:25.1–35.3)	<0.001
REE (kcal/day)†	891±148	1010±138	$1092{\pm}172$	1260±177	<0.001
$ m WC~(cm)^{\dagger}$	71.8±5.9	78.7±6.4	84.3±6.2	94.1 ± 9.6	<0.001
Systolic BP (mmHg) \dagger , \ddagger	125.0 ± 14.2	125.9 ± 14.5	124.7±15.1	124.1±14.4	0.869
Diastolic BP (mmHg)†,‡	75.0 ± 10.4	73.6 ± 10.0	76.7±11.5	73.6 ± 10.2	0.177
$TCHOL(mg/dL)^{\vec{\tau}}$	171.8±39.7	164.6 ± 38.7	174.0 ± 45.2	185.9±43.7	0.010
TG $(mg/dL)^{\dagger}, \sharp$	83.8 ± 70.3	98.2±66.7	103.8 ± 49.1	164.8 ± 341.3	<0.001
$\mathrm{HDL ext{-}C}(\mathrm{mg/dL})^\dagger$	55.6 ± 16.2	52.6 ± 12.8	49.9±12.1	48.2 ± 12.5	0.002
Albumin (g/dL) \dagger ; \ddagger	3.03 ± 0.46	3.16 ± 0.44	3.24 ± 0.42	3.33±0.44	<0.001
$KPS > 60\% [n (\%)]^*.$$	9 (10.8)	13 (16.3)	19 (22.4)	28 (32.6)	<0.001
Diabetes $[n\ (\%)]^*$	15 (17.2)	23 (26.4)	32 (34.8)	42 (47.7)	<0.001
Hypertension $[n\ (\%)]^*$	44 (50.6)	46 (52.9)	57 (62.0)	57 (64.8)	0.163
Smoking $[n (\%)]^*$, f					0.146
Current	0(0)	2 (2.5)	2 (2.4)	4 (4.7)	
Former	6 (7.2)	13 (16.3)	15 (17.6)	15 (17.6)	
Never	77 (92.8)	65 (81.3)	68 (80.0)	66 (77.6)	

Abbreviation: BMI, body mass index; REE, resting energy expenditure; BP, blood pressure; TCHOL, total cholesterol; TG, triglycerides; HDL-C, high-density-lipoprotein cholesterol; KPS, Karnofsky performance status scale.

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Pearson chi-square test was used for categorical data. Data were presented with n (%).

 $^{^{\}uparrow} ANOVA~test~was~used~for~comparing~mean~values~of~continuous~variables~between~groups.~Data~were~presented~as~mean~\pm SD.$

 $\slash\hspace{-0.4em}^{\slash\hspace{-0.4em} \uparrow}$ Statistics are tested using the log-transformed values.



Table 2
Baseline subjects' characteristics by survival status and all-cause deaths

	Survivors (n=135)	All-cause deaths (n=219)	P Value
Men [n (%)]*	56 (41.5)	100 (45.7)	0.442
Age (years) †	76.8±7.8	79.3±7.7	0.004
Height (cm) [†] , ‡	153.0±6.9	152.8±8.3	0.796
Weight $(kg)^{\dagger}$	52.2±11.4	50.0±10.8	0.065
BMI $(kg/m^2)^{\dagger}$	22.3±4.4	21.4±4.1	0.055
REE (kcal/day) †	1086±208	1050±207	0.120
WC (cm) †	84.0±12.2	81.4±9.9	0.038
Systolic BP (mmHg) [†] , ‡	124.5±14.7	125.2±14.4	0.659
Diastolic BP (mmHg) † , ‡	74.2±9.5	75.1±11.2	0.534
$TCHOL(mg/dL)^{\dagger}$	177.0±45.5	172.2±40.5	0.303
TG $(mg/dL)^{\dagger}$, \ddagger	108.8±74.9	114.8±220.1	0.428
HDL-C(mg/dL) †	51.3±13.3	51.7±14.0	0.767
Albumin $(g/dL)^{\dagger}$, \ddagger	3.26±0.43	3.15±0.46	0.028
$KPS > 60\% [n (\%)]^*, $ ¶	27 (21.6)	42 (20.1)	0.742
Diabetes $[n\ (\%)]^*$	38 (28.1)	74 (33.8)	0.268
Hypertension $[n (\%)]^*$	78 (57.8)	126 (57.5)	0.964
Smoking [<i>n</i> (%)]*,#			0.992
Current	3 (2.4)	5 (2.4)	
Former	18 (14.4)	31 (14.9)	
Never	104 (83.2)	172 (82.7)	

Abbreviation: BMI, body mass index; SBP, systolic blood pressure; DBP, diastolic blood pressure; TCHOL, total cholesterol; TG, triglycerides; HDL-C, high-density-lipoprotein cholesterol; REE, resting energy expenditure; KPS, Karnofsky performance status scale.

 $^{^{\}ast}$ Pearson chi-square test was used for categorical data. Data were presented with n (%).

 $^{^{\}dagger}$ Student's *t*-test for unpaired data was used for the comparison of mean values between groups. Data were presented as mean \pm SD.

[‡]Statistics are tested using the log-transformed values.

^{¶&}lt;sub>No.=334.</sub>

[#]No.=333

Table 3

Unadjusted and adjusted relative risks of all-cause mortality by baseline BMI, WC, and REE quartiles (district models) using the Cox proportional hazards regression analyses.

Variables, quartile (No.)	Deaths (No.)	5-year mortality rate per 1000 persons- years	Unadjusted all-cause mortality relative risk (95% CI)	Adjusted all-cause mortality relative risk (95% CI)*, \dagger †
BMI_quartiles [†]				
I (87)	58	256.6	1.00(Reference)	1.00(Reference)
II (87)	53	210.9	0.82(0.57-1.19)	0.73(0.49-1.08)
III (92)	59	213.6	0.83(0.57-1.19)	0.79(0.53-1.17)
IV (88)	49	165.5	0.64(0.44-0.93)**	0.62(0.40-0.96)**
BMI_Asia [‡]				
I (75)	52	274.0	1.00(Reference)	1.00(Reference)
II (158)	101	224.8	0.82(0.59-1.14)	0.75(0.53-1.07)
III (44)	22	141.3	0.51(0.31–0.84)***	0.54(0.32-0.92)**
IV (77)	44	172.7	0.63(0.42–0.94)**	0.60(0.38-0.95)**
BMI_TW [§]				
I (75)	52	274.0	1.00(Reference)	1.00(Reference)
II (181)	114	216.6	0.79(0.57-1.09)	0.74(0.53-1.05)
III (62)	35	168.5	0.61(0.40-0.94)***	0.60(0.37-0.96)**
IV (36)	18	143.0	0.52(0.30–0.89)***	0.50(0.28-0.89)**
ree [¶]				
I (88)	61	276.1	1.00(Reference)	1.00(Reference)
II (89)	56	222.8	0.80(0.56-1.16)	0.78(0.53-1.17)
III (89)	54	189.7	0.68(0.47-0.98)**	0.73(0.49-1.10)
IV (88)	48	164.1	0.59(0.40–0.86)***	0.61(0.38-0.99)**
WC [#]				
I (83)	57	258.9	1.00(Reference)	1.00(Reference)
II (88)	50	192.8	0.74(0.51–1.09)	0.69(0.45-1.06)
III (83)	58	238.5	0.91(0.63-1.32)	0.93(0.62-1.39)
IV (86)	46	159.1	0.61(0.42-0.90)**	0.58(0.37-0.91)**

Abbreviation: CI, confidence interval; BMI, body mass index; REE, resting energy expenditure; KPS, Karnofsky performance status scale; TCHOL, total cholesterol; TG, triglycerides; HDL-C, high-density-lipoprotein cholesterol.

^{*} Analyses adjusting for age, gender, albumin, KPS, TCHOL, TG, HDL-C, hypertension, and diabetes

 $^{^{\}dagger}$ BMI was divided into quartiles by genders: <18.5, 18.5–20.5, 20.6–23.6, >23.6 kg/m 2 in men; <19.4, 19.4–21.4, 21.5–25.0, >25.0 kg/m 2 in women

 $^{^{\}ddagger}$ BMI was divided using WHO for Asians criteria as below: I (underweight): <18.5, II (normal weight): 18.5–22.9, III (Overweight): 23.0–24.9, IV (Obesity): \geq 25.0 kg/m²

 \S BMI was divided using Taiwan's criteria as below: I (underweight): <18.5, II (normal weight): 18.5–23.9, III (Overweight): 24.0–26.9, IV (Obesity): \ge 27.0 kg/m²

 \P REE was divided into quartiles by genders: <1078, 1078–1201, 1202–1381, > 1208 kcal/day in men; <835, 836–939, 940–1041, > 1041 kcal/day in women

 $^{\#}$ WC was divided into quartiles by genders: <73.6, 73.6–81.0, 81.1–87.4, > 87.4 cm in men; <75.5, 75.5–81.5, 81.6–89.1, > 89.1 cm in women

^{**} *P* value < 0.05.

^{***} *P* value < 0.01.

^{††&}lt;sub>No.=334</sub>