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# An Ecological Momentary Assessment Analysis of Relations among Coping, Affect, and Smoking During a Quit Attempt

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# Abstract

**Aims**—This study used Ecological Momentary Assessment (EMA) data from smokers trying to quit to assess relations among coping, positive affect, negative affect, and smoking. The effects of stress coping on affect and smoking were examined.

**Design**—Data from a randomized clinical trial of smoking cessation treatments were submitted to multilevel modeling to test the effects of coping with stressful events on subsequent affect and smoking.

**Participants**—372 adult, daily smokers who reported at least one stressful event and coping episode and provided post-quit data.

**Measurements**—Participants' smoking, coping, and affect were assessed in near real time with multiple EMA reports using electronic diaries pre-and post-quit.

**Findings**—Multilevel models indicated that a single coping episode did not predict a change in smoking risk over the next 4 or 48 hours, but coping in men was associated with concurrent reports of increased smoking. Coping predicted improved positive and negative affect reported within 4 hours of coping, but these affective gains did not predict reduced likelihood of later smoking. Pre-quit coping frequency and gender moderated post-quit stress coping relations with later positive affect. Men and those with greater pre-quit coping frequency reported greater gains in positive affect following post-quit coping.

**Conclusions**—Coping responses early in a quit attempt may help smokers trying to quit feel better but may not help them stay smoke-free.

# Keywords

smoking cessation; stress; coping; affect; ecological momentary assessment; multilevel modeling

# Introduction

Smoking remains the leading preventable cause of illness and death [1–2] and relapse remains the most common outcome of cessation attempts [3–4]. Stress and associated

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**Conflict of Interest Statement** 

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affective distress contribute to cessation failure and relapse [5–8]. Shiffman [9] found that most smokers reported that lapses were preceded by stress, a result replicated in other studies [10–11]. Similarly, within-subjects studies show that a third or more of smokers report that they lapsed during stressful events (see [12] for a review). However, Shiffman's[9] retrospective study indicated that stressful experiences may be precursors to highly tempting situations, but not necessarily to lapses or relapses. In the Shiffman study, the use of coping to deal with relapse crisis was the only predictor of temptation outcome; any coping predicted greater likelihood of successful resolution of the temptation without smoking. Later studies found additional evidence that coping is closely associated with temptation outcomes [13–15].

Smoking cessation treatment programs were quick to integrate such findings and typically aim to promote coping in response to smoking temptations and warn would-be quitters about high-risk situations [3,16]. For instance, the Public Health Service Clinical Practice Guideline *Treating Tobacco Use and Dependence* [3] recommends practical counseling that focuses on teaching problem-solving skills, such as cognitive coping strategies to regulate negative mood. Indeed, substantial evidence supports a positive relation between coping and successful resolution of temptations or relapses crises [7,13–15,17,18]. For instance, a retrospective study [13] suggested that the number of temptation-coping strategies used positively predicted abstinence. Similarly, a community intervention trial [19] demonstrated that 7-day point prevalence abstinence at the six-month follow-up was associated with the average number of temptation-coping strategies used rather than the average number of temptations reported.

Although much research supports the consensus that stressful events and coping with temptations predict distal cessation outcomes, the effects of stress coping on affect and lapse vulnerability in the short-term remain unknown. Research has focused mainly on the effects of coping with temptations to smoke rather than coping with stressful events. The conceptual framework proposed by Wills and Shiffman [20] assumes a distinction between *stress-coping*, responses intended to deal with general life stressors, and *temptation coping* responses specific to temptations for substance use, and argued that stress coping and temptation coping make independent contributions to substance use. Affect is hypothesized to be a potent motivator of drug use [21] and frequent target of coping (i.e., emotion-focused coping) [22], and is therefore a candidate mediator of coping effects on smoking. Most studies have focused on negative affect, but not positive affect. Negative reinforcement models identify anhedonia (e.g., lack of positive affect) as an aversive state that prompts drug use [21]. Sustained or increased positive affect may indicate that this aversive state has been avoided or eliminated, and in turn, predicts abstinence.

Thus, the extent to which coping effectively improves affect, by either alleviating distress or increasing positive affect, has not been explored as a possible mediator of coping effects on abstinence. Previous research on coping has mainly studied post-quit coping and no studies have yet examined the role of pre-quit coping experience in cessation efforts. Pre-quit stress coping experience may help to automate coping so that it is less demanding and more effective, much like practice can make complex behaviors such as driving nearly effortless.

Moreover, most of the studies discussed above relied heavily on retrospective self-reports and between-subjects analyses. Stone et al. [23] cogently demonstrated that the correspondence between retrospective and EMA assessments was low; in retrospective assessments, cognitive coping was underreported while behavioral coping was overreported, relative to EMA reports. Also, when coping efforts are assessed once per subject, the results from between-subjects comparisons only indicate differences between people who use a certain coping style and those who do not. Such differences may reflect stable

individual differences rather than the effects of coping per se. Within-subject designs can disentangle individual differences and situation-specific coping effects. Such studies have yielded discrepant results [18,24].

In the current study, we attempted to examine how coping with stress was related to subsequent affect and smoking status. We predicted that stress coping-smoking relations would depend on: (1) the *efficacy* of coping (as measured by affect improvement), and (2) prior coping *experience* (as measured by frequency of pre-quit coping with stress). Coping efficacy was assessed by decreases in negative affect and increases in positive affect. We also sought to explore gender differences in coping effects as a possible explanation for lower smoking cessation rates in women than men [25,26]. Past research has shown that men and women tend to differ in terms of: coping likelihood and style, success in quitting [25,26,27], and how smoking risk is affected by stress [28]. In this study, we explored whether gender moderated coping effects on later affect and smoking.

In this study, EMA data were used to determine whether coping with stress predicted change in the likelihood of a smoking lapse over 48 hours, relative to stressful events subjects did not cope with. We predicted that coping with stress would protect against smoking lapses and would improve affect (i.e., decrease negative and increase positive affect) over 4 hours. We further hypothesized that the relation between coping (vs. no coping) and both later lapse and affect would be moderated by the frequency of stress coping pre-quit, such that those who reported more coping pre-quit would have improved affect and decreased lapse risk following coping post-quit. We also predicted that the relation between stress coping and subsequent lapses would be mediated by intervening changes in affect.

# Method

#### **Participants**

Data in the current study were collected for a double-blind, randomized, placebo-controlled clinical trial of bupropion SR and individual counseling for smoking cessation [29]. Participants were daily smokers (age 18 years and older) recruited in the Madison, Wisconsin area via mass media. Inclusion/exclusion criteria are shown in Table 1. A total of 463 participants enrolled. Of these, 372 (80.3%) attended the quit day visit and reported at least one stressful event and one stress-coping effort (Figure 1). Demographic characteristics of the subjects are summarized in Table 2. Subjects retained for analyses were equally distributed across treatment conditions  $\chi^2(3, N = 463) = 2.10, p = .552$  and those dropped from the present analyses (due to attrition or lack of stress or coping reports) did not differ from those retained in terms of age, minority status, years smoking, CO level, or number of past quit attempts (all *ps* > .05). The dropped and retained samples differed in terms of cigarettes smoked per day, gender, and post-quit stressful event frequency, as shown in Table 2. The actual sample sizes for the analyses varied from 346 to 347 due to missing data.

#### Measures

**Baseline Assessment**—Participants provided demographic information (see Table 2), baseline depressive symptoms using the Center for Epidemiologic Studies-Depression scale (CES-D) [30], smoking history, and completed the Fagerström Test of Nicotine Dependence (FTND [31]), Wisconsin Smoking Withdrawal Scale (WSWS [32]), the Positive and Negative Affect Schedule (PANAS [33]), and the Negative Emotionality Scale (NES [34]).

**Ecological Momentary Assessment**—EMA data were collected via Electronic Diaries (EDs; Palm Vx Palmtop Computer, Palm, Inc., Santa Clara, CA) programmed by in vivo

data Inc. (Pittsburgh, PA) to administer four to seven two-minute momentary reports (random reports) at pseudo-random times separated by at least 30-minute intervals daily for two weeks pre- and four weeks post-quit.

Random reports assessed the occurrence of stressful events (yes/no) since the last report, and, if so, whether participants tried to cope with stress (yes/no). Subjects were instructed to endorse the stress item when they felt subjectively stressed by an event, regardless of the magnitude of the precipitating event (i.e., there was no formal, *a priori* definition of a stressful event). The stress coping item was displayed only when stressful events were reported (which occurred on 8% of reports). In addition, the number of cigarettes smoked since the last report (0–20) was assessed and subjects rated affect and withdrawal just before the prompt on an 11-point scale ranging from 1 (No!!) to 11 (Yes!!). Two highly correlated positive affect P-PANAS items ("interested" and "enthusiastic") (r=.83 [29]) were averaged to yield a positive affect summary score. The average of the items "tense or anxious" and "sad or depressed" (r = .45) was used as a measure of momentary negative affect. Positive and negative affect scores were negligibly related ( $\beta$ =-0.01, p=0.012, df=34,123).

#### Procedures

Study procedures are described in detail elsewhere [29]. Subjects were randomized to one of four conditions resulting from the full crossing of medication (9-weeks of bupropion SR vs. placebo) and counseling (8–10-minute counseling sessions vs. no counseling), attended a total of 13 study visits, carried EDs for 6 weeks, and received monthly follow-up phone calls through 1-year post-quit. Compensation for participation did not exceed \$200.

## **Data Analysis**

Data were analyzed using Hierarchical Linear Modeling (HLM) Version 6.04 software [35]. Models predicting smoking within 48 hours of an index episode and predicting affect within 4 hours of an index report were run. The following person-level covariates were tested in all models but were pruned if not significantly related to occasion-level coefficients: medication and counseling conditions; gender (1=female, 0=male); age; baseline FTND; P-PANAS (for positive affect models); N-PANAS and NES scores (for negative affect models); and initial cessation success vs. failure (1=quit smoking for at least 24 hours, 0=failed to quit on quit day). Pre-quit coping count was included as a moderator of coping-affect and coping-smoking relations. All models controlled for any smoking 48 hours before the index report (1=smoked at least once, 0=no smoking). As such, the models predicting smoking within 48 hours of an index report capture predictors of change in smoking status (i.e., lapse) from 48 hours before to 48 hours after the report.

# Results

### **Smoking Risk Over 48 Hours**

Results from Bernoulli models predicting smoking status 48 hours after an index report are shown in Table 3 (top panel). The trimmed model treats only the intercept reflecting the mean probability of smoking in the absence of stress and coping as random (reliability=.84) because allowing additional parameters to vary resulted in non-convergence. Smoking was modeled as a function of stress and coping at an index report (t<sub>0</sub>; treated as binary) and during the 48 hours following an index report (the same period smoking was assessed; with stress treated as a count variable and coping treated as binary). Results indicated that those taking bupropion (vs. placebo) and those who abstained on the quit day were less likely to smoke over 48-hours, whereas older subjects and women were more likely to smoke. A single episode of stress or stress coping had no relation with smoking 48 hours later overall, but the accumulation of stress and any coping over the next 48 hours were both positively

related to concurrent smoking. The positive relation between concurrent stress coping and smoking was stronger for men than for women. Smoking in the 48 hours before the index report positively predicted smoking in the next 48 hours, particularly for men.

Separate models in which earlier levels of negative and positive affect were added to the predictors listed above were run (Table 3, bottom panel). Affect at the time of the index report and four hours after the index report (after stress and coping) were included. Neither positive nor negative affect was significantly predictive of smoking status within 44–48 hours.

We also tested models in which the dependent variable was any smoking reported within four hours of the index report. These models also failed to detect relations between stress coping and later smoking. Results are presented in the supplementary material online.

#### Affect within 4 Hours

To test the efficacy of coping in improving affect, we ran the models shown in Tables 4 and 5. Separate models predicted positive and negative affect at the next report within four hours of an index report to assess whether coping was associated with changes in affect. Models controlled for the level of affect at the index report, smoking in the 48 hours preceding an index report, and the person-level covariates listed above (if significantly related to model coefficients). Models also examined relations between recent smoking, stress, and stress coping reported at the same time as affect to examine concurrent relations between these variables.

Negative affect results indicated that those who received the combination of active (vs. placebo) bupropion SR and active counseling (vs. assessment control) and those who quit successfully for 24 hours reported lower negative affect, whereas those who had higher FTND, N-PANAS, and NES scores reported higher negative affect overall, in the absence of stress, coping, and smoking. Stressful events were marginally related to increased negative affect, controlling for negative affect at the index report. Stressful events reported at the same time as negative affect were positively related to negative affect, especially for women. Concurrently reported stress coping was marginally associated with lower negative affect. Smoking in the 48 hours before an index report was significantly related to reduced negative affect four hours post-index for those who quit successfully, but not for those who failed to quit. Recent smoking was not significantly related to negative affect.

Positive affect results indicated that older subjects and those with higher P-PANAS baseline scores had higher positive affect in the absence of stress, coping, and smoking post-quit. Stressful events at the index report were marginally predictive of reduced positive affect within four hours, whereas stress coping was predictive of improved positive affect, at least for men and those with more frequent coping prior to quitting. Those with more pre-quit stress reported less positive affect benefit from an index coping effort. Positive affect was also negatively related to concurrently reported stressful events, but positively related to concurrently reported stress coping for those with more practice coping pre-quit. Stress coping was less positively related to concurrently reported positive affect among those with more pre-quit stressful events, however. Positive affect was not significantly related to smoking in the 48 hours before the index report or in the four hours between the index report and rating of positive affect.

**Coping Strategy Analyses**—See supplementary material online for the results for the exploratory analyses of relations between specific coping strategies (i.e., cognitive, behavioral, and acceptance-based coping) and later affect and smoking.

# Discussion

The purpose of this study was to use EMA data to test the hypothesis that stress coping improves affect and decreases smoking during a quit attempt. A secondary aim was to determine whether individual differences in prior coping moderated such relations. Results indicated that stress coping did not protect against smoking during a quit attempt. Instead, any stress coping in a 48 hour period was associated with increased risk of smoking during the same 48 hour period, particularly for men. Results also showed that coping was associated with short-term improvements in affect. Negative affect was lower within four hours of coping with stress, and positive affect was higher after coping, particularly in men and those who coped more frequently prior to quitting. So, despite the fact that there was some evidence that coping was effective (i.e., it appeared to have intended affective consequences), neither coping nor these affective gains were protective against smoking over the next 44–48 hours.

The finding that contemporaneous stress coping was associated with a greater likelihood of smoking, after controlling for any smoking in the past 48 hours, runs counter to our hypothesis. It is important to note that we did not detect a link between a single, index episode of stress coping and smoking in the short term (4 hours) or longer term (48 hours), but found instead that contemporaneous reports of coping and smoking were positively related. Because coping and smoking were assessed in the same 48 hour period, it is possible that smoking preceded coping (i.e., that smoking triggered coping efforts rather than the reverse) or that smoking prompted recall of coping efforts. The gender difference observed in this coping-lapse relation may, therefore, reflect a gender difference in smoking effects on coping or recall of coping. On the other hand, the results may indicate that men experience a greater increase in the likelihood of smoking after attempting to cope with stress than do women. Perhaps this reflects gender differences in thresholds for coping (that men engage in active coping for only the most stressful events that are likely to increase smoking risk, whereas women cope with less intense risk). This may also reflect gender differences in coping efficacy, although men seemed to benefit more from coping in terms of positive affect than did women and benefitted equally in terms of negative affect. Alternatively, this gender difference may reflect differences in expectancies about what coping will do to smoking urges. Research suggests that men are more motivated to smoke by nicotine effects than women, whereas women are more motivated by non-nicotine (e.g., sensory) effects than men [25,36]. Perhaps this leads men to view stress coping as unlikely to reduce urges to smoke or withdrawal, and they are especially likely therefore to turn to smoking per se to achieve these effects.

Although stress coping did not protect against smoking, such coping predicted improved mood. Negative affect was lower within four hours of a coping episode, compared with noncoping episodes (controlling for stressful event occurrence). This effect was not moderated by cessation status on the quit day, gender, or pre-quit stress coping frequency. Positive affect improvements within four hours of coping, however, were greater for those who coped more often prior to quitting (i.e., those more practiced in coping), as anticipated, and for men. Men may experience a greater boost in positive affect following active coping because men may be socialized to use more active and instrumental coping strategies than are women; this type of coping may be more social-role congruent for men than women [27]. Stress coping was less effective in boosting positive affect among those who reported stressful events more frequently prior to quit controlling for pre-quit coping frequency, however. This suggests that those with the greatest stress burden pre-quit, perhaps due to chronic stressors or persistent hassles, benefit less from stress coping bouts post-quit. Short-term improvements in affect related to stress coping did not reduce lapse risk over the next 44–48 hours after controlling for additional stressful events encountered during that period. This is consistent with past research indicating no relation between negative affect and ad lib smoking in non-treatment seeking smokers [37,38]. The time frame used to detect affect-lapse relations may have been suboptimal, as past research suggests that negative affect influences first lapses over the course of hours rather than days [5]. This time interval may be too long to detect such effects, but it is important to note that coping did not reduce smoking risk in even a four-hour window in this study (see supplementary material). In any event, stress coping appears to improve affect more than it protects against a return to smoking during a quit attempt.

## Limitations

Results from this study may have been influenced by sampling bias, measurement problems, and the fact that we did not manipulate the causal variables of interest (e.g., coping). The sample retained for analyses may differ from the general population of smokers trying to quit. Although reports were randomly prompted to minimize sampling biases, missing reports may be associated with certain situations or states of interest (e.g., severe stress or demanding coping). In addition, although past research supports the validity and reliability of brief EMA affect measures, measures of coping and stress occurrence are less well validated [39]. Assessments were brief to reduce subject burden, and so did not assess subjective coping efficacy, the severity or chronicity of stressful events, or the events that triggered smoking episodes. Past research has found that negative affect predicts smoking only when the smoking is triggered by stress [40]. If we limited analyses to smoking following stress, we might similarly find a relation with negative affect. The fact that our data replicated some known relations (e.g., the relation between any past smoking and increased risk of future smoking; the relations between stressful events and affect and smoking; bupropion SR reduction of smoking risk; relations among baseline affect and momentary affect [3,9,41,42] suggests that our measures were somewhat valid and reliable, however. Finally, stress coping might, in theory, work via other means than affective change (e.g., altered self-efficacy). However, this would not change the fact that such coping was not related to smoking in the current data set.

#### Conclusions

Results of the present study of relations among coping, affect, and smoking suggest that stress coping in the early phases of a quit attempt may be more effective in improving affect than in reducing smoking risk, and may even increase smoking. Additional analyses should address the extent to which stress (vs. temptation events) moderates relations between coping and later smoking and may identify the types of coping that are effective and should be promoted in treatment.

# **Supplementary Material**

Refer to Web version on PubMed Central for supplementary material.

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#### Figure 1.

Participation flow diagram depicting the number of subjects excluded from analyses due to drop out prior to the quit day, the absence of any reports of stressful events, and the absence of any reports of stress coping. The number of people screened for enrollment was estimated assuming a 22.86% enrollment rate based on two cohorts for which data were available, as data were not available for all cohorts.

#### Inclusion/exclusion criteria.

Inclusion Criteria	Exclusion Criteria
18 years of age or greater	History of bipolar disorder or psychosis
Able to read and write English	diagnosis or treatment
At least fairly motivated to quit smoking	Current depression (CES-D score over 16)*
Willing to fulfill study requirements	Current illegal drug use
Smoking at least 10 cigarettes per day	Use of other tobacco products in last 7 days
Baseline CO level of at least 10 parts per million	Current use of stop-smoking treatments
	Participation in a study in the past 30 days
	Living with someone enrolled in the study
	Uncontrolled hypertension
	Current heavy drinking
	History of seizure
	Past negative reactions to bupropion,
	Pregnancy
	Breast feeding

\* Except when an interview with a licensed clinical psychologist suggested another cause for elevated scores (such as an anxiety disorder, sleep disorder, or pain disorder associated with elevated ratings on specific relevant symptoms).

Demographic characteristics of final sample and differences in subsample of subjects who were retained (N=372) vs. excluded (N=91) from current analyses.

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Variable	Value	Retained Cases (N=372)	Excluded Cases (N=91)	$\chi^2$	d
Sex (N=372)	Female	198 (53.2%)	35(38.5%)	6.38	.012*
Race/Ethnicity	Hispanic	4 (1.1%)	1 (1.1%)		.984
	White	332	80 (87.9%)	5.02	.285
	African-American	22 (5.9%)	4 (4.4%)		
	Asian, Pacific Islander	3 (0.8%)	1 (1.1%)		
	American Indian	1 (0.3%)	2 (2.2%)		
	Other	11 (3.0%)	4 (4.4%)		
Marital Status	Married	162 (43.5%)	36 (43.5%)	1.18	.946
	Divorced	69 (18.5%)	16 (17.6%)		
	Never married	91 (24.5%)	26 (28. 6%)		
	Cohabitating	34 (9.1%)	10 (11.0%)		
	Separated	8 (2.2%)	2 (2.2%)		
	Widowed	6 (1.6%)	1 (1.1%)		
Education	< High school	13 (3.5%)	7 (7.7%)	9.27	.055
	High school graduate	77 (20.7%)	27(29.7%)		
	Some college	184 (49.5%)	40(44.0%)		
	College degree	96 (25.8%)	17 (18.7%)		
Employment Status	Employed for wages	273 (73.4%)	50 (54.9%)	11.70	.111
	Self-employed	36 (9.7%)	8 (8.8%)		
	Unemployed <1 year	18 (4.8%)	7 (7.7%)		
	Homemaker	17 (4.6%)	1 (1.1%)		
	Student	8 (2.2%)	4 (4.4%)		
	Retired	9 (2.4%)	5 (5.5%)		
	Disabled	6 (1.6%)	4 (4.4%)		
Household Income	< \$25,000	105 (28.9%)	36 (39.6%)	9.95	.127

Variable	Value	Retained Cases (N=372) Excluded Cases (N=91)	Excluded Cases (N=91)	$\chi^2$	d
	\$25,00-\$34,999	57 (15.3%)	13 (14.2%)		
	\$35,000-\$49,999	71 (19.1%)	17 (18.7%)		
	\$50,000-\$74.999	76 (20.4%)	13 (14.2%)		
	>\$75.000	55 (14.5%)	11 (12.1%)		
		(SD)	M (SD)	$M(SD)$ tor $\chi^2$	d
Age		38.94 (11.94)	38.00 (13.06)663	663	.507
Cigarettes smoked per day		21.44 (10.26)	23.95 (10.79)	2.06	$.040^{*}$
Previous quit attempts		6.19 (11.45)	4.92 (4.68)	959	.338
Baseline CO level		24.54 (11.74)	24.84 (10.80)	.211	.833
<b>Baseline FTND Score</b>		4.98 (2.33)	5.63 (2.46)	2.31	.021*
Years smoking		21.19 (11.31)	20.43 (12.09)	571	.569
Pre-quit stressful event count		9.06 (9.33)	6.87 (9.72)	-1.84	.067

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Trimmed multilevel model of stress, stress coping, and affect effects on smoking risk over 48 hours.

Predictor Level-1	<b>Predictor Level-2</b>	Coefficient	Standard Error	T-ratio	Odds Ratio	95% CI	df	p-value
Coping-lapse relations								
Mean P (Smoking) 48 hrs after index <sup><i>a</i></sup>		- 0.880	0.387	- 2.276	0.415	(0.194, 0.886)	342	$0.023^{*}$
	Active bupropion SR	- 1.171	0.360	- 3.251	0.310	(0.153, 0.629)	342	$0.002^{*}$
	Age	0.047	0.015	3.173	1.048	(1.018, 1.080)	342	$0.002^{*}$
	Gender	0.980	0.361	2.716	2.664	(1.312,5.411)	342	$0.007^{*}$
	Quit at least 24 hours	- 3.141	0.374	- 8.399	0.043	(0.021, 0.090)	342	$0.000^*$
Index Stressful Event (Y/N) (48 hrs earlier)		- 0.266	0.168	- 1.578	0.767	(0.551, 1.066)	33,560	0.114
Index Stress Coping (Y/N) (48 hrs earlier)		0.078	0.208	0.373	1.081	(0.719, 1.624)	33,560	0.709
Smoking (Y/N) (48 hrs before index)		0.890	0.085	10.447	2.434	(2.060, 2.876)	33,560	$0.000^{*}$
	Gender	- 0.316	0.103	-3.055	0.729	(0.596, 0.893)	33,560	$0.003^{*}$
Stressful Events (Count; in 48 hrs following index)		0.318	0.027	11.585	1.375	(1.303, 1.451)	33,560	$0.000^*$
Stress Coping $(Y/N)$ (48 hrs after index)		0.548	0.123	4.461	1.730	(1.360,2.201)	33,560	$0.000^*$
	Gender	- 0.434	0.120	- 3.623	0.648	(0.512,0.819)	33,560	$0.001^{*}$
Affect-lapse relations								
Negative Affect (At index)		- 0.004	0.021	-0.211	0.996	(0.955, 1.038)	21,271	0.839
Negative Affect (Within 4 hrs of index)		- 0.015	0.021	-0.730	0.985	(0.945, 1.026)	21,271	0.465
Positive Affect (At index)		- 0.014	0.019	- 0.754	0.986	(0.950, 1.023)	21,271	0.451
Positive Affect (Within 4 hrs of index)		- 0.026	0.018	- 1.424	0.974	(0.940, 1.010)	21,271	0.155
aRandom coefficient, reliability = .826. All other predictors were treated as fixed to facilitate model convergence.	ctors were treated as fixed	l to facilitate m	odel convergence.					
* p<.05								

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Trimmed model of negative affect within 4 hours of an index report.

T TOMOT TOACT TOACT T	Predictor Level-2	Coefficient	Standard Error	T-ratio	df	P-value
Mean Negative Affect Rating 4 hrs after index $a$		3.151	0.175	17.975	339	$0.000^{*}$
	Active bupropion SR	0.293	0.214	1.368	339	0.179
	Counseling	0.244	0.210	1.165	339	0.245
1	Active bupropion SR X Counseling	- 0.833	0.299	- 2.786	339	$0.006^*$
	FTND	0.077	0.032	2.395	339	$0.017^{*}$
	Negative PANAS	0.034	0.014	2.457	339	$0.015^{*}$
	NES	0.116	0.024	4.832	339	$0.000^*$
	Quit at least 24 hours	- 0.027	0.165	- 0.166	339	0.868
Index Negative Affect <sup>b</sup> (4 hrs earlier)		0.309	0.014	22.296	346	$0.000^*$
Index Stressful Event (Y/N) (4 hrs earlier)		0.135	0.074	1.827	21,522	0.067
Index Stress-Coping (Y/N) (4 hrs earlier)		- 0.250	0.089	- 2.799	21,522	$0.006^*$
Smoking (Y/N) (48 hrs before index)		0.063	0.042	1.513	21,522	0.130
	Quit at least 24 hours	- 0.170	0.058	- 2.919	21,522	$0.004^{*}$
Smoking (Y/N) (contemporaneous)		- 0.018	0.034	- 0.547	21,522	0.584
Stressful Event <sup>c</sup> $(Y/N)$ (contemporaneous)		1.009	0.136	7.437	345	$0.000^*$
	Gender	0.411	0.146	2.806	345	$0.006^*$
Stress-Coping (YN) (contemporaneous)		- 0.239	0.128	- 1.870	21,522	0.061

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b Random coefficient, Reliability = 0.644, Random coefficient, Reliability = 0.679. All other predictors were treated as fixed to facilitate model convergence.

\* p<.05

Trimmed model of positive affect within 4 hours of an index report.

Predictor Level-1	Level-2	Coefficient	Standard Error	<b>T-ratio</b>	df	P-value
Mean Positive Affect Rating 4 hrs after index $a$		7.457	0.183	40.848	343	$0.000^{*}$
	Age	0.045	0.00	5.171	343	$0.000^*$
	Positive PANAS	0.095	0.015	6.386	343	$0.000^*$
	Quit at least 24 hours	- 0.229	0.219	- 1.041	343	0.299
Index Positive Affect $^{b}\left( Y/N ight) (4$ hrs earlier)		0.295	0.015	19.298	346	$0.000^{*}$
Index Stressful Event $(Y/N)$ (4 hrs earlier)		- 0.172	060.0	- 1.923	21,528	0.054
Index Stress-Coping (Y/N) (4 hrs earlier)		0.298	0.124	2.402	21,528	$0.016^{*}$
	Gender	- 0.262	0.092	- 2.843	21,528	$0.005^{*}$
	Pre-quit Stressful Events	- 0.071	0.027	- 2.660	21,528	$0.008^{*}$
	Pre-quit Stress-Coping Efforts	0.072	0.028	2.540	21,528	$0.011^{*}$
Smoking (Y/N) (48 hrs before index)		- 0.028	0.036	- 0.777	21,528	0.437
Smoking (Y/N) (contemporaneous)		0.027	0.042	0.638	21,528	0.523
Stressful Event <sup><math>c</math></sup> (Y/N) (contemporaneous)		- 0.558	0.127	- 4.385	346	$0.000^*$
Stress-Coping (Y/N) (contemporaneous)		0.185	0.148	1.251	21,528	0.211
	Pre-quit Stressful Events	- 0.093	0.039	- 2.391	21,528	$0.017^{*}$
	Pre-quit Stress-Coping Efforts	0.117	0.043	2.705	21,528	$0.007^{*}$

 $^{a}$ Random coefficient, Reliability= .986,

 $b_{
m Reliability=.651,}$ 

 $^{c}$ Reliability= 0.607. All other predictors were treated as fixed to facilitate model convergence.

\* p<.05