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Supervised machine learning and logistic regression identifies novel epistatic risk factors with PTPN22 for rheumatoid arthritis

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Abstract

Investigating genetic interactions (epistasis) has proven difficult despite the recent advances of both laboratory methods and statistical developments. With no ‘best’ statistical approach available, combining several analytical methods may be optimal for detecting epistatic interactions. Using a multi-stage analysis that incorporated supervised machine learning and methods of association testing, we investigated epistatic interactions with a well-established genetic factor (PTPN22 1858T) in a complex autoimmune disease (rheumatoid arthritis (RA)). Our analysis consisted of four principal stages: Stage I (data reduction)—identifying candidate chromosomal regions in 292 affected sibling pairs, by predicting PTPN22 concordance using multipoint identity-by-descent probabilities and a supervised machine learning algorithm (Random Forests); Stage II (extension analysis)—testing detailed genetic data within candidate chromosomal regions for epistasis with PTPN22 1858T in 677 cases and 750 controls using logistic regression; Stage III (replication analysis)—confirmation of epistatic interactions in 947 cases and 1756 controls; Stage IV (combined analysis)—a pooled analysis including all 1624 RA cases and 2506 control subjects for final estimates of effect size. A total of seven replicating epistatic interactions were identified. SNP variants within CDH13, MYO3A, CEP72 and near WFDC1 showed significant evidence for interaction with PTPN22, affecting susceptibility to RA.

Keywords

epistasis; rheumatoid arthritis; PTPN22; Random Forests

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Conflict of interest

The authors declare no conflict of interest.

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Introduction

Genome-wide association studies, which provide the ability to simultaneously investigate hundreds of thousands of genetic markers in large numbers of individuals, have successfully led to the discovery of genetic risk factors with modest effects in several complex diseases, including autoimmune diseases.[1,2] Nevertheless, it is apparent that current approaches to genetic analysis, which include almost exclusively, marginal associations using a univariate approach, are not able to identify a substantial fraction of the genetic burden. This may reflect the involvement of rare variants, copy number variation, *gene* × *gene* interactions, *gene* × *environment* interactions and/or epigenetic mechanisms. Currently, there is no consensus regarding appropriate approaches for evaluating these components of complex diseases.

Investigating genetic or *gene* × *gene* interactions (also known as 'epistasis', where the action of one gene is modified by one or several other genes) has proven difficult, despite recent advances of both laboratory methods and statistical developments. For example, the 15th biennial Genetic Analysis Workshop (GAW15) investigated genetic interactions in rheumatoid arthritis (RA (MIM 180300)) using several data sets. A variety of statistical approaches were used to investigate epistasis in RA in both family and population-based data sets with varying genetic marker density; results varied greatly, showing that robust and comprehensive approaches not restricted by a small sample size or sparse data are necessary.[3–8] With no 'best' statistical approach available, combining several analytical methods may be optimal for detecting epistatic interactions.[9,10] Here, we performed a comprehensive multi-stage genetic investigation with a replication analysis to reveal epistatic relationships, involving the well-established *PTPN22* (protein tyrosine phosphatase non-receptor 22; GeneID 26191) risk variant, that alter susceptibility to RA.

RA is a chronic multi-system autoimmune disease, resulting from persistent inflammatory synovitis and subsequent erosion of the joint architecture. It is considered a complex disease with a multi-factorial etiology influenced by both genetic and environmental risk factors. Genetic predisposition to RA is suspected to involve multiple genes with incomplete penetrance, interacting in a variety of biological pathways.[11] A large number of earlier studies have investigated the role of genetic factors in RA. The *HLA-DRB1* (GeneID 3123) shared epitope (SE; *HLA-DRB1* alleles: 0101, 0102, 0401, 0404, 0405, 0408, 0413, 1001, and 1402) and *PTPN22* 1858T alleles have been consistently associated with greater risk[12–14] Similar to *HLA-DRB1*, *PTPN22* is a strong biological candidate as it regulates T-cell receptor signaling and confers risk for multiple autoimmune diseases, though there is limited knowledge regarding the exact etiological mechanism(s).[15,16]

Previous genome-wide linkage scans in RA identified evidence for additional loci, however uncovering the susceptibility genes within these regions has been difficult.[17,18] Recent candidate gene and genome-wide association studies have reported strong evidence for several genes with modest contributions to RA susceptibility, including *CTLA4*, *PADI4*, *REL*, *STAT4* and *TRAF1-C5*, as well as loci within chromosomal regions 6q23, 10p15, 12q13 and 22q13.[19–26] Multiplex interactions between genetic and environmental risk factors (*gene* × *gene*, *gene* × *environment*, *gene(n)* × *environment*) have been implicated in RA, however the scope of earlier studies has been restricted primarily to candidate risk factors.[14,27–30] To understand the etiologic mechanisms involved in RA, a clear understanding of the intricate genetic architecture underlying susceptibility, including the complete identification of epistatic relationships between genes, particularly those that do not show independent, or marginal, disease associations, is necessary. Using a multi-stage strategy framed within a genome-wide perspective, we identified evidence for novel epistatic risk factors interacting with *PTPN22* to influence RA susceptibility.

Results

We investigated epistatic interactions with the *PTPN22* 1858T risk variant (rs2476601) in RA cases (available through the North American Rheumatoid Arthritis Consortium (NARAC)) and who were predominantly positive for the presence of antibodies to cyclic citrullinated peptide (anti-CCP; Table 1) and healthy controls of European origin using four principal stages: Stage I (data reduction)—identification of relevant chromosomal regions in 292 affected sibling pairs (ASPs) using Random Forests, a supervised machine learning algorithm; Stage II (extension analysis)—testing for epistasis with *PTPN22* 1858T in 677 cases and 750 controls; Stage III (replication analysis)—confirmation of epistatic interactions in 947 cases and 1756 controls; Stage IV (combined analysis)—a combined analysis including all 1624 RA cases and 2506 control subjects to determine final estimates of effect size (Figure 1).

Stage I (data reduction)

A non-parametric approach using a supervised machine learning algorithm (Random Forests) and multipoint identity-by-descent (mIBD) probabilities identified promising regions harboring epistatic candidates for *PTPN22* in 292 ASPs from 512 multiplex RA families. This case-only analysis investigated mIBD probabilities for 379 microsatellite markers (MSMs) from across the genome (~10 cM coverage), as well as gender and *HLA-SE*, were used to predict *PTPN22* 1858T concordance in ASPs using the Random Forests algorithm (see Materials and methods). Random Forests assigns each predictor a variable importance (VI) score, with more *important* predictors having greater values. On the basis of the distribution of the VI scores, the five top-ranking predictors were considered *important* predictors of *PTPN22* concordance in the ASPs (see Supplementary Figure 1), as there appeared to be a clear distinction between the VI scores for these predictors compared with the remaining predictors. The five *important* predictors were mIBD probabilities for MSMs in the following chromosomal regions: 4q34, 5p15, 10p11, 14q23 and 16q23 (Table 2).

Stage II (extension analysis)

We investigated *important* chromosomal regions through association tests in individuals with at least 90% Northern European ancestry (677 cases; 750 healthy controls) from the collective NARAC I data set (908 cases; 1260 healthy controls) for which dense genome-wide SNP data were available (Table 1). We evaluated our power to detect *gene* × *gene* interactions (see Materials and methods) and were sufficiently powered to investigate interactions under a dominant mode of inheritance. The presence of the *PTPN22* 1858T risk variant was significantly associated with RA in Caucasians when using the full NARAC I data set (odds ratio (OR) = 2.09, 95% CI: 1.69–2.60), as described earlier.[22] Similar results were observed when the analysis was restricted to individuals with at least 90% Northern European ancestry (OR = 1.89, 95% CI: 1.45–2.46). A total of 10 589 SNPs within 5 Mb of an *important* locus satisfied quality control criteria (see Materials and methods; Table 2) and were investigated using the Breslow–Day test for homogeneity. A total of 665 SNPs significantly modified the effect of *PTPN22* on susceptibility to RA ($P\text{-value}_{BD}<0.05$). Under a dominant model, 449 SNPs tested for epistasis (logistic regression) with *PTPN22* showed significant ($P\text{-value}_{I}<0.05$; see Materials and methods) evidence for interaction in the NARAC I data set.

Stage III (replication analysis)

The replication analysis was performed in the NARAC II data set (947 cases and 1756 controls of European ancestry). The presence of the *PTPN22* 1858T risk variant was also significantly associated with RA in the replication data set as expected (OR =1.81, 95% CI: 1.49–2.20); results for the NARAC II cases have been reported earlier.[26,31] Four hundred

and forty-seven of the 449 SNPs that showed evidence for epistasis (Stage II) met quality control criteria and were formally tested for epistasis with *PTPN22* in the NARAC II data set, under a dominant model. A total of seven epistatic relationships were replicated (P-value$_I$<0.05; Table 3). Individuals were stratified as carriers or non-carriers for the minor allele of each replicating SNP to characterize the relationship of *PTPN22* 1858T with RA susceptibility; consistent associations were observed across both Stage II (NARAC I) and Stage III (NARAC II) populations (Table 3).

Stage IV (combined analysis)

To determine final estimates of epistasis, replicated findings (N =7) were investigated in the combined analysis of the data sets from Stage II (NARAC I) and III (NARAC II), and included 1624 RA cases and 2506 healthy controls; minor allele frequencies for these SNPs did not differ between data sets showing no evidence for heterogeneity (data not shown) (Table 4). The effect of *PTPN22* was significantly increased among individuals who were carriers of the minor allele for four SNP variants versus non-carriers, including: an intronic *CDH13* variant (rs1895535 AA/AG: OR =3.95, 95% CI: 2.39–6.55, $P = 1.0 \times 10^{-7}$; rs1895535 GG: OR =1.69, 95% CI: 1.44–1.98, $P<1 \times 10^{-8}$); an intronic *MYO3A* variant (rs12573019 AA/AG: OR =2.78, 95% CI: 2.05–3.76, $P<1 \times 10^{-8}$; rs12573019 GG: OR = 1.60, 95% CI: 1.34–1.90; $P = 1.7 \times 10^{-7}$); and two SNPs with no known function (rs4695888 CT/TT: OR = 2.10, 95% CI: 1.76–2.51, $P<1 \times 10^{-8}$; rs4695888 CC: OR =1.22, 95% CI: 0.91–1.65, P =0.18; and rs1168587 CT/TT: OR =2.13, 95% CI: 1.78–2.56, $P<1 \times 10^{-8}$; rs1168587 CC: OR =1.29, 95% CI: 0.98–1.70, P = 0.07). The effect of *PTPN22* was reduced in carriers of three SNP variants versus non-carriers, including an intronic *CDH13* variant (rs7200573 AA/AC: OR =1.37, 95% CI: 1.11–1.70, P =0.0034; rs7200573 CC: OR =2.47, 95% CI: 1.99–3.06, $P<1 \times 10^{-8}$); a variant 7 kb upstream *WFDC1* (rs11865624 CC/CT: OR =1.01, 95% CI: 0.66–1.55, P = 0.95; rs11865624 TT: OR =2.01, 95% CI: 1.71–2.37, $P = 1 \times 10^{-8}$); and an intronic *CEP72* variant (rs7726839 AG/GG: OR =1.46, 95% CI: 1.16–1.82, P = 0.0011; rs7726839 GG: OR =2.22, 95% CI: 1.81–2.73, $P< 1 \times 10^{-8}$) (Table 4).

Discussion

We investigated a role for epistasis in RA, a complex autoimmune disease, between *PTPN22* and other variants across the genome using large, well-characterized study populations with detailed clinical and genetic information. A multi-stage analysis, with extension and replication, that combined robust non-parametric and parametric methods was used. We report evidence for novel epistatic interactions with *PTPN22* in RA for variants within *CDH13* (GeneID 1012), *MYO3A* (GeneID 53904), *CEP72* (GeneID 55722) and near *WFDC1* (GeneID 58189). Interestingly, when RA cases were compared with controls, none of the final SNP variants showed significant associations with RA (P<0.05; data not shown); that is, marginal effects in the absence of *PTPN22* were not present for any SNP.

PTPN22 is the second strongest known genetic risk factor for RA, and confers risk in several other autoimmune diseases.[15,16] Located on chromosome 1p13.3–13.1, *PTPN22* encodes a non-receptor classical class I tyrosine protein, lymphoid tyrosine phosphatase, which negatively regulates the T-cell receptor signaling by dephosphorylating several molecules (that is Src family kinases) immediately downstream of the T-cell receptor.[32,33] The 1858T risk variant replaces the amino acid at position 620 from an arginine to a tryptophan, resulting in a gain of function mutation that increases the capacity of lymphoid tyrosine phosphatase to negatively regulate T-cell receptor signaling.[34] There is also evidence suggesting that the 1858T risk variant impairs B-cell receptor signaling and subsequent proliferation.[35–37] Given the importance of *PTPN22* in RA and autoimmunity, the current evidence for statistical interactions, which appear to modify RA risk conferred

by variation within *PTPN22*, sheds new light on potential biological mechanisms for future genetic and molecular investigations. Indeed, epistatic relationships observed in this study may also be relevant to other autoimmune diseases.

We identified and replicated seven epistatic interactions with *PTPN22* under a dominant genetic model. In this analysis *CDH13*, on chromosome 16, showed the strongest epistatic relationship with *PTPN22* in RA. Two *CDH13* SNP variants ~350 kb apart (rs7200573 and rs1895535) showed significant evidence for interaction with *PTPN22*. For rs7200573, ~14 kb from exon 8, *PTPN22* conferred increased risk of RA in both carriers and non-carriers of rs7200573A; however *PTPN22* risk was significantly (P-value$_I$ = 0.00015) less in carriers (OR =1.37) relative to non-carriers (OR =2.47). Interestingly, rs1895535, ~22 kb from exon 5, showed the opposite association for *PTPN22* and RA risk; *PTPN22* risk was significantly (P-value$_I$ = 0.0016) greater in rs1895535A carriers (OR = 3.95) relative to non-carriers (OR =1.69). This variation in association suggests that these risk variants exist on separate haplotypes; r^2 between the two SNPs was <0.1. Unfortunately, SNP variants within exons 5 and 8 of *CDH13* have not been identified, and in available CEPH HapMap data (http://hapmap.org), the linkage disequilibrium between the *CDH13* epistatic risk variants and other SNP variants near the respective exons is also low (r^2<0.1). There was also evidence supporting epistasis for an intronic *MYO3A* SNP variant (rs12573019), an intronic *CEP72* SNP variant (rs7726839) and an SNP variant 7 kb upstream of *WFDC1* (rs11865624), on chromosomes 10, 5 and 16, respectively. *PTPN22*-associated risk for RA was significantly increased in carriers of the *MYO3A* minor allele, whereas *PTPN22* risk was significantly increased in non-carriers of the *CEP72* and *WFDC1* minor alleles. An investigation of the underlying haplotype block structure for the epistatic risk variants in CEPH suggests that the identified variants may be tagging functional variants within their respective genes. For example, the *MYO3A* variant, between exons 9 and 10 (~2.2 kb range), occurs within a large 150 kb haplotype block spanning 22 exons; the *CEP72* risk variant, located in the first intron, exists within a 38 kb haplotype block that extends through the entire gene; and the risk variant near *WFDC1* has modest linkage disequilibrium (r^2 =0.2–0.3) with several SNP variants within the 5′ untranslated region and first intron of *WFDC1*.

Detailed functional data are not available for these epistatic candidates; however, there is sufficient evidence to suggest a plausible biological relationship between these loci and RA. For example, *CDH13* (cadherin 13; T-cadherin (truncated); H-cadherin (heart)) encodes a unique cadherin lacking transmembrane and cytosolic domains necessary for homophilic adhesive activity of classical cadherins.[38] T-cadherin is likely involved in signal transduction and not cell–cell adhesion, as it concentrates in lipid raft domains of the plasma membrane, affects cellular migration, angiogenesis, survival under oxidative stress and contributes to the invasive potential of various cancers.[39–44] Interestingly, *CDH13* has also been associated with attention-deficit/hyperactivity disorder, blood pressure and adult height.[45–47]

Additionally, *MYO3A* (myosin IIIA) encodes a unique myosin motor protein that contains an N-terminal kinase domain and is primarily expressed within the retina and inner ear of vertebrates.[48] Mutations within the motor domain of *MYO3A* results in non-syndromic hearing loss DFNB30.[49] Both sensorineural and conductive hearing loss have increased prevalence in RA patients.[50–52] Furthermore, *MYO3A* influences sterocillia shape and length, and is capable of inducing filopodial actin protrusions in culture cells, and thus may have an unknown function in immune cell locomotion similar to *MYO2A*.[53,54]

CEP72 (centrosomal protein 72 kb) encodes a centro-mere protein that is critical for chromosomal alignment and proper tension generation between sister chromatids during

mitosis.[55] Interestingly, anti-centromere antibodies are primarily observed in patients with CREST syndrome, but have also been observed in other autoimmune diseases.[56–59]

Another candidate identified here, *WFDC1* (whey acidic protein four-disulphide core domain 1), encodes ps20 and has a highly conserved core domain.[60] WFDC1/ps20 is a multi-functional protein that facilitates endothelial cell motility and angiogenesis, inhibits cell proliferation and promotes cellular senescence.[60–62] CD4 T cells normally express ps20 after restimulation and IL2 expansion.[63] ps20 also increases CD4 T-cell permissiveness to HIV spread through CD54 integrin expression, and identifies a subset of CD4 memory T cells at an early differentiation stage (CD45RO +/CD28 +/CD57−).[63] Despite the lack of detailed experimental data, there is plausible statistical and biological evidence to support further investigation of these candidates. Our results underscore important lessons derived from recent genome-wide association studies, including findings that most replicated associations do not involve previous candidate genes, suggesting new biological hypotheses, and that many have implicated non-protein coding regions.[64]

In this analysis, we investigated epistatic interactions with *PTPN22* in RA using a multi-stage approach that combined robust non-parametric and parametric methods. There are several clear advantages to our methodology: (1) Random Forests is a model-free approach, robust to uninformative predictors and outliers and attaches a measure of importance to each predictor. The VI reported by Random Forests incorporates additional information compared to a univariate test for a predictor, as it reflects both the individual effect and the possible effect through multiplex interactions. In this analysis, we applied Random Forests to an ASP data set that was sufficiently powered to detect marginal genetic associations with OR≤0.5 and ≥2.0 for a dominant model (data not shown). (2) We used conventional logistic regression models to test for epistasis assuming multiplicative interaction, which are readily interpretable. (3) We included a replication analysis to confirm epistatic interactions. In addition, we used well-defined study populations with detailed clinical information. The principal limitation in this analysis is our interpretation of the Random Forests results in Stage I. First, there is no clear standard for identifying an *important* variable; we based our selection on the empirical distribution of the VI scores and subjectively determined that the there was a clear distinction in the VI for the five top-ranking predictors versus all others (see Supplementary Figure 1). Second, the VI potentially includes the effect of multiplex interactions between the predictors, as each variable selected at a node is essentially *important* conditional on the variable selected at the prior node; we did not explore multiplex genetic interactions in this analysis (that is three-way or other higher order interactions), which may explain why epistatic interactions were detected in four of the five *important* regions. Furthermore, the imposition of a dominant genetic model for epistatic candidates might not have been the appropriate assumption to follow up the Random Forests findings.

The *HLA-DRB1* SE is a critical genetic component of RA in Northern European Caucasians, conferring approximately 30–50% of the genetic risk, which suggests an important role for antigen presentation and subsequent T-cell activation in RA pathogenesis; recent experimental data suggests that the SE triggers pro-oxidant signaling and an innate immune response.[65–68] Given the significance of *HLA-DRB1* in RA, we specifically included *HLA-DRB1* and a dense set of MSMs across chromosome 6p21, which provided extensive coverage of the major histocompatibility complex region, as predictors in Stage I of the analysis. Interestingly, Random Forests did not identify any chromosome 6p21 marker, including the *HLA-DRB1* locus, as an *important* predictor of *PTPN22* carrier status in the ASP investigation in Stage I. Results did not change when SE status, specifically, was considered in ASPs (data not shown); therefore, we did not further investigate the relationship between *HLA-DRB1* and *PTPN22*. Our results are in strong agreement with

several recent studies,[69–71] but in contrast with others.[14,72] We do acknowledge that a large proportion of the ASPs were positive for *HLA-SE*. Additionally, the RA cases in Stages II and III were anti-CCP positive, the phenotypic subgroup for which the classical *HLA-DRB1* association has been exclusively established.[73]

In summary, results from this study show the genetic contribution to RA risk is more complex than originally considered. Here, we identified novel candidate genes (*CDH13*, *MYO3A*, *CEP72*, *WFDC1*) that modify the effect of *PTPN22*-associated risk for RA, and provide an important framework for future studies of *gene* × *gene* interactions. Genetic studies in complex disease must include application of multi-analytical strategies. Efforts to explore higher order interactions are needed and will require very large sample sizes and clearly defined phenotypes.

Materials and methods

Informed consent was obtained from all participants and approvals from local institutional review boards were secured at each recruitment site before enrollment. All RA cases satisfied the American College of Rheumatology (ACR) 1987 classification criteria for RA.[74]

Study population

Stage I—A total of 292 ASPs from 512 multiplex RA families recruited by the NARAC were used as described earlier (Table 1).[12] Briefly, families were eligible if ≥ two siblings satisfied the ACR 1987 criteria for RA,[74] ≥ one sibling had erosions on hand radiographs and ≥ one sibling had disease onset between the ages of 18 and 60 years. Families were excluded from participation if other diseases associated with similar articular symptoms were present. Greater than 90% of RA cases were positive for the presence of anti-CCP antibody, an antibody that is highly specific to RA;[75] 82.1% were positive for rheumatoid factor (data not shown).

Stage II—A collective data set of 908 RA cases was used in the extension analysis, which included 464 unrelated probands from the NARAC ASP families (described above; of which 127 probands overlapped with Stage I), 168 RA cases from the National Data Bank for Rheumatic Diseases, 162 RA cases from the National Inception Cohort of Rheumatoid Arthritis and 114 RA cases from the Study of New Onset Rheumatoid Arthritis and 1260 healthy control subjects from the New York Health Project, as described earlier (Table 1).[22,76–79] All RA cases were positive for the presence of anti-CCP antibody; 85.9% were positive for rheumatoid factor (data not shown). The study was then restricted to subjects with at least 90% Northern European ancestry (see Laboratory Procedures) to avoid bias because of population stratification (*N* =682 RA cases, 752 controls). Similarly, all RA cases in this group were positive for anti-CCP; 85.0% were positive for rheumatoid factor (data not shown). Subjects were also excluded if genotype data for SNP rs2476601 (*PTPN22* 1858T) was missing. The final data set for extension analysis comprised 677 RA cases and 750 healthy controls (Table 1).

Stage III—The NARAC II data set used in the replication analysis consists of 952 RA cases, which included 175 RA probands from the NARAC family studies, 332 RA cases from the Veterans Affairs Rheumatoid Arthritis Registry, 160 RA cases from the Studies of the Etiologies of Rheumatoid Arthritis cohort, 105 RA probands as members of the Multiple Autoimmune Disease Genetics Consortium, 86 RA patients from the UCSF Rheumatoid Arthritis Genetics Project and 94 RA patients from the Early Rheumatoid Arthritis Treatment Evaluation Registry and 1760 control subjects, as described earlier (Table

1).[12,22,26,31,80–82] Patients for whom anti-CCP data were available were all anti-CCP positive (data was not available for 31 Veterans Affairs Rheumatoid Arthritis Registry subjects). Control data were taken from publicly available control data sets in the Illumina iControl database (Illumina, San Diego, CA, USA; http://www.illumina.com/iControlDB) and the neurodevelopmental control group obtained from the NIH Laboratory of Neurogenetics (http://neurogenetics.nia.nih.gov/paperdata/public/). The control genotypes were selected from the entire set of European American genotypes available in these resources based on the following data filters: (1) >90% complete genotyping data; (2) >90% European continental ancestry. The European continental ancestry was determined using ancestry informative markers as described earlier.[83] Subjects were also excluded if genotype data for SNP rs2476601 (*PTPN22* 1858T) was missing. The final data set for the replication analysis comprised 947 RA cases and 1756 healthy controls (Table 1). There was no overlap between individuals in the replication data set, and those used in both Stage I and Stage II analyses.

Laboratory procedures

Stage I—RA ASPs were genotyped for *PTPN22* 1858T (rs2476601), the *HLA-DRB1* locus and 379 MSMs from the Marshfield Set 8A Combo List (http://research.marshfieldclinic.org/genetics/GeneticResearch/screeningsets.asp) with additional MSMs in specific chromosomal regions (that is the HLA complex) as described earlier.[12] The MSMs provided approximately 10 cM genome-wide coverage on average.

Stage II—The NARAC I subjects were genotyped for 545 080 SNPs at the Feinstein Institute for Medical Research. The NARAC I RA case samples and 601 control samples were genotyped with the Infinium HumanHap550 v1.0 (Illumina), 411 controls on Human-Hap550 v3.0 and 248 controls on Infinium Human-Hap300 and HumanHap240S arrays.[22] Genotypes were collected for samples across the three Illumina platforms and plate membership was assessed by the top 10 principal components (EIGENSTRAT); no systematic differences were observed.[22] Subjects were excluded if more than 5% of genotypes were missing, had non-European ancestry, had evidence of relatedness or had evidence of possible DNA contamination.[22] Subjects were evaluated for Northern European ancestry by applying the software program STRUCTURE to an ancestry informative set of 704 SNPs.[84,85]

Stage III—The NARAC II subjects were genotyped using 373 400 SNPs on the Illumina HapMap370 BeadChip at the Feinstein Institute for Medical Research, as described earlier.[26] Subjects with more than 5% missing genotype data or showing evidence of non-European ancestry were excluded. In addition, samples showing evidence of relatedness with other samples in the study population, or possible DNA contamination were also excluded. Of the SNP variants identified in Stage II (see Statistical analysis), a total of 219 SNP variants were imputed in the NARAC II data set using maximum likelihood imputation and applying the greedy algorithm as implemented in MACH v1.0.16.[86] Five Markov Chain iterations were set to obtain map estimates, which were used as conditions for finding the most likely genotype. The NARAC I healthy controls were used as the reference population.

Statistical analysis

Stage I—A graphical summary of the analytical approach for this study, including all steps, statistical methods and significance criteria is provided in Figure 1. The first stage of analysis used Random Forests, a supervised machine learning algorithm that grows recursively partitioned trees without pruning.[87] Each tree is independently grown on a bootstrapped sample of observations, and at each node in the tree, the predictor that best discriminates the outcome is selected from a random subset of predictors. Classification

accuracy of the forest is assessed for observations not included in the bootstrapped sample by comparing the predicted versus the actual outcome across all trees. Finally, the value of each predictor is randomly permuted across all trees, and a single VI score for each predictor is determined by the change in classification accuracy (misclassification). The VI scores can be used to identify *important* predictors that most contribute to outcome classification without model specification, and are robust to uninformative predictors and outliers.

Our Random Forests analysis explored the specific hypothesis that ASPs may share other genetic regions relevant to the biological mechanism(s) through which their shared *PTPN22* status confers risk for RA; therefore we used mIBD probabilities for 379 MSMs to predict *PTPN22* concordance in ASPs using Random Forests v5.1 (http://www.stat.berkeley.edu/~breiman/RandomForests/cc_software.htm). ASPs were categorized as positively or negatively concordant for *PTPN22* 1858T carrier status (both siblings having at least one variant (N = 83 pairs), or neither sibling having a variant (N =209 pairs)). All MSMs were in Hardy–Weinberg equilibrium in unrelated individuals and had inheritance consistency within families,[12] and mIBD probabilities for each MSM were generated using GENEHUNTER v2.1.[88] On the basis of the distribution of the VI scores (see Supplementary Figure 1), five *important* (top-ranking) predictors from the Random Forests analysis were chosen for further study (Table 2).

Stage II—All SNPs within 5 Mb of an *important* locus were selected for investigation in the NARAC I case–control data set (N =11 207 SNPs). SNPs were excluded if they did not meet set criteria for Hardy–Weinberg equilibrium ($P>1\times10^{-4}$), genotype call rates (>90% completeness) or minor allele frequency (>0.01). A final subset of SNPs used for this stage of analysis (N =10 589 SNPs) had an average call rate of 99.2% (Table 2).

Power to detect *gene* × *gene* interactions with *PTPN22* in Stage II was investigated, using dominant and recessive inheritance modes. We assumed a two-sided type 1 error of 5% (α = 0.05), and a frequency and effect estimate for *PTPN22* 1858T of 9.0% and OR =1.8 based on published estimates. Power to detect ratio of odds ratio (ROR_I) ranging from 0.1 to 3.0 was examined. Our analyses revealed that there was sufficient power (~70–99%) to detect *gene* × *gene* interactions with $ROR_I \leq 0.5$ and ≥ 2.0 for almost all dominant models considered, and for recessive models where minor allele frequency>35% in Stage II (data not shown). Therefore, we restricted our investigation to dominant genetic models, and all analyses were performed and results interpreted in accordance with these established criteria.

Identified SNPs were evaluated for effect modification of *PTPN22* risk in NARAC I cases and controls using the Breslow–Day test for homogeneity implemented in PLINK v1.06.[89] Significant SNPs ($P\text{-value}_{BD}<0.05$) were selected as candidates for epistasis with *PTPN22*, and were formally tested for epistasis using logistic regression models in Stata v9.2 (StataCorp. LP, College Station, TX, USA). The test for epistasis was based on the coefficient of the interaction term (where P-value of the interaction term (ROR_I) reflects the difference in the likelihood between the full model and a reduced model containing only main effects); interactions showing a significance level $P\text{-value}_I<0.05$ were considered significant.

Stage III—The SNPs that provided evidence for epistatic interaction with *PTPN22* in Stage II were further investigated. Power to detect *gene* × *gene* interactions between *PTPN22* and each final candidate SNP was investigated for the replication (NARAC II) data set using previously defined parameters. Our analyses revealed that power was sufficient (~70–99%) in Stage III to detect *gene* × *gene* interactions with $ROR_I \leq 0.6$ and ≥ 1.6 for almost all dominant models considered. Candidate SNPs that satisfied quality control criteria

mentioned earlier were formally tested for epistasis using logistic regression models in Stata v9.2. Interactions showing a significance level $P\text{-value}_I<0.05$ were considered significant.

Stage IV—A combined analysis including all 1624 RA cases and 2506 control subjects gave final estimates of effect size for replicating interactions. Allelic frequencies for all SNPs pursued in the fourth stage were first tested for heterogeneity (Stata v9.2) in each case and control group separately before combining. Only those showing homogeneity across groups were included. SNP stratified analyses (minor allele carrier versus non-carrier subjects (dominant genetic model)) were used to further characterize the epistatic relationships for *PTPN22* using logistic regression (Stata v9.2).

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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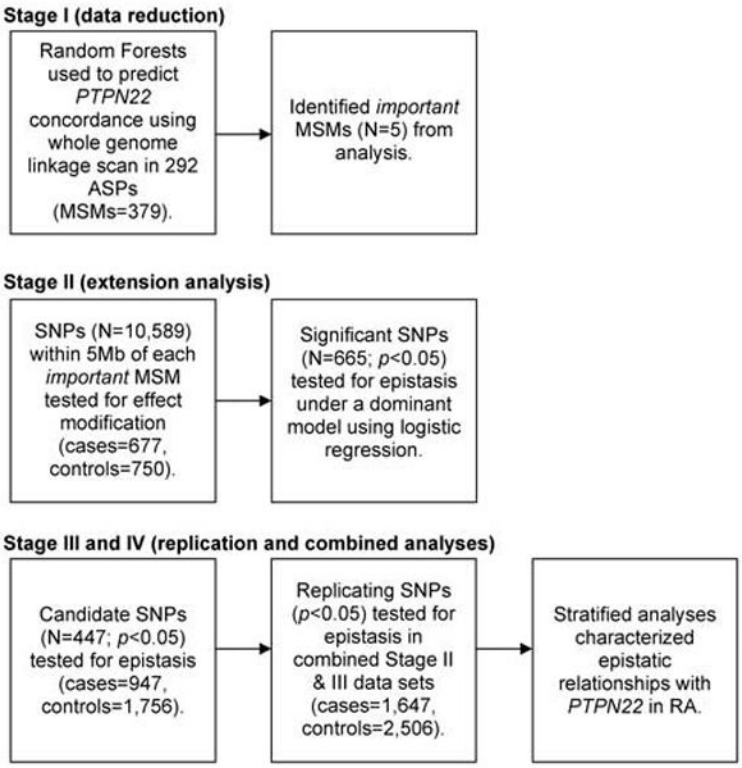

Figure 1.
Summary of analytical approach to identify evidence for epistatic relationships with *PTPN22* in RA.

Table 1

Clinical and demographic features of the RA family and case–control data sets

Study population	N	Female:male ratio	Age of onset (years)	PTPN22 1858T carrier (%)[a]	Shared epitope carrier (%)[b]	Anti-CCP status (%)
Stage I						
RA cases	530	3.3:1	39.2	29.1% (*n* = 154)	83.8% (*n* = 444)	90.8% (*n* = 481)
ASPs	292	—	—	83	227	259
Stage II (NARAC I)						
All RA cases	908	2.8:1	45.5 (*n* = 884)	27.8% (*n* = 898)	97.9% (*n* = 865)	100% (*n* = 907)
All controls	1260	2.5:1	NA	15.6% (*n* = 1253)	44.6% (*n* = 1159)	NA
NE[c] RA cases	682	2.8:1	45.7 (*n* = 670)	27.6% (*n* = 677)	97.7% (*n* = 641)	100% (*n* = 681)
NE[c] controls	752	2.3:1	NA	16.8% (*n* = 750)	46.7% (*n* = 752)	NA
Stage III (NARAC II)						
All RA cases	952	1.2:1[d]	47.2 (*n* = 917)	27.4% (*n* = 947)	NA	100% (*n* = 921)
All controls	1760	NA	NA	17.2% (*n* = 1756)	NA	NA

Abbreviations: ASP, affected sibling pairs; NARAC, North American Rheumatoid Arthritis Consortium; RA, rheumatoid arthritis.

[a] Individuals were *PTPN22* 1858T carriers if they had one or more 1858T alleles. In Stage II, the *PTPN22* 1858T allele was associated with RA using all individuals (OR = 2.09, 95% CI: 1.69–2.60) and when restricted by Northern European ancestry (OR = 1.89, 95% CI: 1.45–2.46). In Stage III, the *PTPN22* 1858T allele was associated with RA (OR = 1.81, 95% CI: 1.49–2.20).

[b] Individuals were shared epitope carriers if they had one or more of the following *HLA-DRB1* alleles: 0101, 0102, 0401, 0404, 0405, 0408, 0413, 1001 and 1402. However, four-digit typing was not available for all subjects, thus cases with *HLA-DRB1**04/01/10 were assumed to be carriers for shared epitope.

[c] Subjects with at least 90% Northern European ancestry.

[d] Lower female:male ratio reflects overrepresentation of male RA cases from the Veterans Affairs Rheumatoid Arthritis Registry.

Table 2

Top genomic regions identified by the Random Forests analysis in Stage I (292 ASPs) and number of SNPs to be investigated in Stage II (NARAC I: 677 cases, 750 controls)

Chr	MSM	Band	No. SNPs[a]
4	D4S2431	4q34.1	1580
5	D5S2488	5p15.33	1167
10	D10S1426	10p11.23	2016
14	D14S592	14q23.1	2006
16	D16S402	16q23.3	3820

Abbreviations: ASP, affected sibling pairs; GWA, genome-wide association; MSM, microsatellite marker; NARAC, North American Rheumatoid Arthritis Consortium.

[a]SNP variants were selected if present in the NARAC I GWA data set, satisfied quality control criteria, and were within 5 Mb of an *important* MSM identified by Random Forests (see Supplementary Figure 1).

Table 3

Replicating results for tests of interaction between *PTPN22* and candidate SNPs in two populations of European origins (Stages II and III), and the effect of *PTPN22* when stratified by SNP of interest

Chr	SNP	*P*-value$_{BD}$[a]	Stage II[b]				Stage III[c]			
			P-value$_I$	ROR$_I$[d] (95% CI)	SNP carriers[e]	SNP non-carriers[f]	*P*-value$_I$	ROR$_I$[d] (95% CI)	SNP carriers[e]	SNP non-carriers[f]
					OR$_{PTPN22}$(95% CI)	OR$_{PTPN22}$ (95% CI)			OR$_{PTPN22}$ (95% CI)	OR$_{PTPN22}$ (95% CI)
4	rs4695888T	0.041	0.016	2.02 (1.14–3.60)	2.26 (1.66–3.06)	1.11 (0.68–1.81)	0.038	1.59 (1.03–2.45)	2.04 (1.64–2.54)	1.29 (0.88–1.87)
5	rs7726839G	0.025	0.021	0.55 (0.33–0.91)	1.39 (0.97–2.00)	2.55 (1.77–3.66)	0.046	0.68 (0.46–0.99)	1.45 (1.09–1.94)	2.15 (1.67–2.76)
10	rs12573019A	0.029	0.022	1.99 (1.11–3.57)	3.13 (1.89–5.19)	1.58 (1.17–2.12)	0.048	1.56 (1.00–2.42)	2.53 (1.73–3.71)	1.63 (1.31–2.03)
10	rs1168587T	0.019	0.018	1.96 (1.12–3.44)	2.30 (1.69–3.13)	1.17 (0.73–1.87)	0.037	1.55 (1.03–2.35)	2.07 (1.65–2.60)	1.33 (0.94–1.89)
16	rs1895535A	0.018	0.038	2.94 (1.06–8.12)	5.03 (1.89–13.43)	1.71 (1.31–2.24)	0.0076	2.42 (1.26–4.62)	4.02 (2.17–7.45)	1.67 (1.36–2.03)
16	rs7200573A	0.026	0.021	0.54 (0.32–0.91)	1.42 (1.00–2.02)	2.61 (1.79–3.81)	0.0015	0.54 (0.37–0.79)	1.33 (1.01–1.74)	2.46 (1.89–3.21)
16	rs11865624C	0.028	0.050	0.48 (0.23–1.00)	1.01 (0.51–2.01)	2.12 (1.60–2.79)	0.020	0.51 (0.28–0.90)	1.00 (0.59–1.71)	1.98 (1.62–2.42)

Abbreviations: NARAC, North American Rheumatoid Arthritis Consortium; RA, rheumatoid arthritis.

[a] Breslow–Day test was used as a test of effect modification, evaluating the homogeneity of association (OR) between an SNP variant and the risk for RA, across each strata of *PTPN22* 1858T carrier status in NARAC I data (677 RA cases and 750 controls). Only those SNPs with *P*-value$_{BD}$<0.05 were tested for epistasis.

[b] Using NARAC I data: 677 RA cases and 750 controls.

[c] Using NARAC II data: 947 RA cases and 1756 controls.

[d] The test of epistasis is based on the interaction term (reported ratio of odds ratios (ROR$_I$) and 95% CI) between the dominant genetic model of *PTPN22* 1858T (CT/TT versus CC) and the dominant genetic model for the SNP variant.

[e] Using RA cases and controls who carried the minor allele of the SNP variant of interest.

[f] Using RA cases and controls who did not carry the minor allele of the SNP variant of interest.

Table 4

Results for tests for interaction between *PTPN22* and replicating SNPs in the combined data set, and results for the effect of *PTPN22* 1858T when stratified by SNP variants of interest under a dominant model (Stage IV)[a]

Chr	SNP	MAF[b]	Stage IV		SNP carrier[d]		SNP non-carrier[e]		Gene
			ROR_I[c] (95% CI)	*P*-value$_I$	OR (95% CI)	*P*-value	OR (95% CI)	*P*-value	
4	rs4695888T	0.50	1.71 (1.22–2.43)	0.0021	2.10 (1.76–2.51)	$<1 \times 10^{-8}$	1.22 (0.91–1.65)	0.184	
5	rs7726839G	0.26	0.65 (0.48–0.89)	0.0061	1.46 (1.16–1.82)	0.0011	2.22 (1.81–2.73)	$<1 \times 10^{-8}$	*CEP72*
10	rs12573019A	0.13	1.74 (1.23–2.47)	0.0019	2.78 (2.05–3.76)	$<1 \times 10^{-8}$	1.60 (1.34–1.90)	1.7×10^{-7}	*MYO3A*
10	rs1168587T	0.43	1.66 (1.19–2.31)	0.0028	2.13 (1.78–2.56)	$<1 \times 10^{-8}$	1.29 (0.98–1.70)	0.073	
16	rs1895535A	0.06	2.34 (1.38–3.97)	0.0016	3.95 (2.39–6.55)	1.0×10^{-7}	1.69 (1.44–1.98)	$<1 \times 10^{-8}$	*CDH13*
16	rs7200573A	0.29	0.56 (0.41–0.75)	0.00015	1.37 (1.11–1.70)	0.0034	2.47 (1.99–3.06)	$<1 \times 10^{-8}$	*CDH13*
16	rs11865624C	0.07	0.50 (0.32–0.79)	0.0031	1.01 (0.66–1.55)	0.946	2.01 (1.71–2.37)	$<1 \times 10^{-8}$	5′ *WFDC1*

Abbreviations: NARAC, North American Rheumatoid Arthritis Consortium; RA, rheumatoid arthritis.

[a] Using the combined data from NARAC I and NARAC II: 1647 RA cases and 2506 controls.

[b] Minor allele frequencies (MAF) were determined using data from controls.

[c] The test of epistasis is based on the interaction term (reported ratio of odds ratios (ROR_I) and 95% CI) between the dominant genetic model of *PTPN22* 1858T (CT/TT versus CC) and the dominant genetic model for the SNP variant.

[d] Using RA cases and controls who carried the minor allele of the SNP variant of interest.

[e] Using RA cases and controls who did not carry the minor allele of the SNP variant of interest.