Serum 25-hydroxyvitamin D concentrations in relation to cardiometabolic risk factors and metabolic syndrome in postmenopausal women $1-3$

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ABSTRACT

Background: Low concentrations of serum 25-hydroxyvitamin D [25(OH)D] may be associated with cardiometabolic disorders; however, little is known about their relation to intermediate metabolic and lipid markers.

Objective: We investigated the relation of serum 25(OH)D concentrations to fasting insulin, glucose, dyslipidemia, adiposity, and prevalent metabolic syndrome.

Design: We conducted this cross-sectional analysis in 292 postmenopausal women aged 50–79 y in the Women's Health Initiative Calcium–Vitamin D (WHI-CaD) trial. Data were collected from 3 nested case-control studies that measured baseline serum 25(OH)D concentrations. Inverse probability weighting was used to approximate parameter estimates for the WHI-CaD population.

Results: In weighted linear regression models adjusted for age, raceethnicity, month of blood draw, region, case-control status, smoking, alcohol, physical activity, and history of cardiometabolic risk factors, there was an inverse association of serum 25(OH)D with adiposity [body mass index (BMI): $\beta = -1.12 \pm 0.30$, $P = 0.0002$; waist circumference: $\beta = -3.57 \pm 0.49$, $P < 0.0001$; waist-hip ratio: β = -0.01 \pm 0.002, P < 0.0001], triglycerides (β = -0.10 \pm 0.02, $P < 0.0001$), and triglyceride:HDL-cholesterol ratio ($\beta = -0.11 \pm$ 0.03, $P = 0.0003$). The multivariable-adjusted odds ratio for metabolic syndrome for the highest $(\geq 52 \text{ nmol/L})$ compared with the lowest (<35 nmol/L) tertile of serum 25(OH)D concentrations was 0.28 (95% CI: 0.14, 0.56). Significant associations remained after adjustment for BMI. We observed no significant associations with LDL cholesterol, HDL cholesterol, insulin, glucose, homeostatic model assessment of insulin resistance (HOMA-IR), or homeostatic model assessment of β cell function (HOMA- β).

Conclusion: Higher serum 25(OH)D concentrations may be inversely associated with adiposity, triglycerides, triglyceride:HDL-cholesterol ratio, and metabolic syndrome but are not associated with LDL and HDL cholesterol, insulin, glucose, HOMA-IR, or HOMA- β in postmenopausal women. This trial was registered at clinicaltrials.gov as NCT00000611. Am J Clin Nutr 2011;94:209–17.

INTRODUCTION

Vitamin D deficiency is an increasingly recognized health concern related to skeletal and nonskeletal outcomes. Although accumulating evidence suggests that low concentrations of serum 25-hydroxyvitamin D [25(OH)D] may be associated with increased risk of cardiometabolic disorders including type 2 diabetes (1) and

cardiovascular disease (2–4), biological mechanisms that underlie these relations remain poorly understood.

Vitamin D receptors are present on pancreatic β cells and insulinsensitive tissues including skeletal muscle tissue (5), and vitamin D repletion improves insulin and glucose homeostasis in animal models of vitamin D deficiency (6, 7). However, findings from crosssectional and prospective cohort studies that examined the relation of serum 25(OH)D to fasting insulin (8, 9), fasting glucose (8, 10), insulin resistance (8–13), and β cell dysfunction (10, 11, 13) in observational settings have been inconsistent. Low serum 25(OH)D concentrations may also be associated with dyslipidemia (14–16), but data to support this relation are sparse. Furthermore, the role of adiposity remains unclear. There have been few studies of the relation of serum 25(OH)D concentrations to these intermediate metabolic and lipid markers. Additional research in this

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area may provide insight into possible intermediate pathways for complex cardiometabolic diseases.

To investigate the hypothesis that higher serum 25(OH)D concentrations may be protective for cardiometabolic disease through beneficial effects on intermediate metabolic biomarkers and adiposity, we conducted a cross-sectional analysis in postmenopausal women enrolled in the Women's Health Initiative Calcium–Vitamin D (WHI-CaD) trial. In particular, we examined concentrations of serum 25(OH)D in relation to metabolic biomarkers including total cholesterol, LDL cholesterol, HDL cholesterol, triglycerides, the triglyceride:HDL cholesterol ratio, insulin, glucose, and insulin resistance and β cell dysfunction as measured by homeostatic model assessment, measures of adiposity, including body mass index (BMI; in $kg/m²$), waist circumference, and waist-hip ratio, and prevalent metabolic syndrome.

SUBJECTS AND METHODS

Study population

The WHI-CaD trial was designed to test the effect of calcium and vitamin D supplementation on bone fracture and colorectal cancer in postmenopausal women. A total of 36,282 participants were randomly assigned in a double-blind fashion to consume either 1000 mg elemental calcium (as calcium carbonate) and 400 IU of vitamin D_3 or a placebo. Details on the design and recruitment have been published elsewhere (17, 18). Eligibility criteria for the WHI-CaD trial included no medical condition associated with a predicted survival of \leq 3 y, no prior history of renal calculi, hypercalcemia, corticosteroid use, and calcitriol use, and no safety, adherence, or retention issues (18). All Women's Health Initiative (WHI) study procedures were approved by the institutional review board at each clinical center, and all women provided written informed consent before participating in the study.

The current study took advantage of data collected from 3 nested case-control studies examining fractures (19), breast cancer (20), and colorectal cancer (21) that measured baseline serum 25(OH)D concentrations in women enrolled in the WHI-CaD trial. Controls were free of disease for the duration of the study and were individually matched to case participants according to age, latitude of the clinical center, race-ethnic group, and date of venipuncture. The sample for the current study included women with available measurements of serum 25(OH)D from these case-control studies as well as overlapping measurements of fasting insulin, glucose, triglycerides, total cholesterol, LDL cholesterol, and HDL cholesterol collected previously in a 6% subsample of clinical trial participants. The sample in-

cluded incident cases of fracture, breast cancer, and colorectal cancer ($n = 166$ cases total) ascertained over a mean follow-up period of 7.0 y.

To account for sampling on the basis of case-control status and prior matching, we used inverse probability weighting to provide approximate parameter estimates for the entire WHI-CaD population. In addition, the case-control status and all matching variables including age, ethnicity, geographic region (proxy for latitude of clinical center), and month of blood draw were adjusted for in weighted multivariable models.

Baseline measurements

Certified WHI trained staff measured the height, weight, waist and hip circumference, and blood pressure of each subject at the baseline visit. Height (in cm) was measured with a wall-mounted stadiometer, and weight (in kg) was measured with a balancebeam scale. BMI was calculated as weight (in kg) divided by height (in m²). Waist and hip circumferences (in cm) were determined with a standardized measuring tape. Standardized questionnaires including information on age, ethnicity, education, income, occupation, medical and family histories, smoking status, alcohol use, recreational physical activity, and medication and supplement use were administered at the baseline visit. Metabolic syndrome was defined on the basis of updated guidelines proposed by the International Diabetes Federation, American Heart Association, and National Heart, Lung, and Blood Institute (22) as a presentation with ≥ 3 of the following criteria: *1*) waist circumference >88 cm for women, 2) triglyceride concentration $>$ 150 mg/dL, 3) HDL concentration $<$ 50 mg/dL for women, 4) systolic blood pressure \geq 130 mm Hg or diastolic blood pressure \geq 85 mm Hg, and 5) fasting glucose concentration \geq 100 mg/dL.

Blood collection and assessment of biomarkers

Fasting blood specimens were collected from all participants at baseline according to a standardized protocol. Participants were instructed to fast for 12 h before collection, take all regular medications except for diabetes medication, take no aspirin or nonsteroidal antiinflammatory drugs for 48 h before the visit except for those medications taken regularly, refrain from smoking for 1 h before the visit, and perform no vigorous physical activity for 12 h before the visit. Aliquots of serum, plasma, and buffy coat were frozen and shipped on dry ice to a central repository and stored at -70° C for future assays.

Serum 25(OH)D was measured with the DiaSorin Liaison 25(OH)D chemiluminescent immunoassay system at Diasorin headquarters (Stillwater, MN). Serum insulin was measured by using the stepwise sandwich enzyme-linked immunosorbent assay procedure with an ES 300 (Boehringer Mannheim Diagnostics, Indianapolis, IN). Glucose was measured in serum by using the hexokinase method with a Hitachi 747 analyzer (Boehringer Mannheim Diagnostics). Total cholesterol and triglycerides were measured by enzymatic methods with a Hitachi 747 analyzer (Boehringer Mannheim Diagnostics) as previously described (23). HDL cholesterol was isolated by using heparin-manganese chloride and measured enzymatically with a Hitachi 747 analyzer (Boehringer Mannheim Diagnostics). LDL cholesterol concentrations were calculated by using Friedewald's formula as follows (23, 24):

OmegA-3 TriaL (VITAL), which is a large-scale randomized trial of vitamin D and omega-3s in the prevention of cancer and cardiovascular disease. YS is a recipient of a VITAL ancillary study investigating diabetes prevention by vitamin D and omega-3 fatty acid supplementation (R01 DK088078) and is also supported by grant K01-DK078846 from the National Institute of Diabetes and Digestive and Kidney Diseases, National Institutes of Health.

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LDL cholesterol ¼ total cholesterol ½HDL cholesterol þ ðtriglycerides 3 0:2Þ ð1Þ

The CVs for each analyte were 11.8% for serum 25(OH)D, 5.3–8.8% for insulin, 2.0–2.3% for glucose, 0.8–1.3% for total cholesterol, 1.5–2.2% for LDL cholesterol, 1.8–2.0% for triglycerides, and 2.4–2.6% for HDL cholesterol.

Statistical analysis

Differences in baseline characteristics of the study population across tertiles of serum 25(OH)D were compared by using analysis of variance and Pearson's chi-square test. Continuous outcomes with skewed distributions were logarithmically transformed before analysis to achieve normal distributions. The homeostatic model assessment of insulin resistance (HOMA-IR) was computed by using the formula

[Fasting plasma insulin (mU/L)

$$
\times
$$
 fasting glucose (mmol/L)]÷22.5 (2)

The homeostatic model assessment of β cell function (HOMA- β) was calculated as follows (25, 26):

$$
[20 \times \text{fasting plasma insulin (mU/L)]}
$$

$$
\div \text{fasting plasma glucose (mmol/L)} - 3.5 \qquad (3)
$$

We divided participants into clinically relevant categories of serum 25(OH)D concentrations (\leq 50, 50–75, and $>$ 75 nmol/L) (27–29) as well as tertiles for categorical analysis [to convert 25(OH)D concentrations in nmol/L to ng/mL, divide by 2.496). Tertiles are presented because of small numbers within certain clinically relevant categories $($ >75 nmol $/L$).

We used inverse probability weighting to account for prior matching in the 3 nested case-control studies to allow findings in our sampled study population to be representative of the entire WHI-CaD population ($n = 36,282$). Inverse probability weights equal to the inverse of the conditional probability of being included in the sample were estimated by fitting a logistic regression model that included outcomes from the 3 case-control studies (hip, spine, lower arm, and wrist fractures; invasive breast cancer; and colorectal cancer) and matching variables (age, race-ethnicity, month of blood draw, and geographic region) as predictor variables. All analyses were performed as weighted analyses (with the Proc GenMod procedure in SAS software; version 9.2; SAS Institute, Cary, NC).

We performed weighted multiple linear regression models to compute geometric means of biomarker concentrations across categories of serum 25(OH)D after adjusting for potential confounding variables. Geometric means were calculated by regressing the natural logarithmic values of plasma concentrations of biomarkers on serum 25(OH)D concentrations and taking the antilog of the resulting mean logarithmic value. To test for the linear trend across increasing categories of serum 25(OH)D, we used the median value of each category as a continuous variable in the model. We also calculated the corresponding changes in biomarker concentrations and measures of adiposity associated with

an increase of 25 nmol/L (10 ng/mL) of serum 25(OH)D concentration. To model the shape of the dose-response relation between serum 25(OH)D and biomarker outcomes while allowing for variation within and across categories (30), we fit restricted, weighted, quadratic spline models with knots at medians of tertiles of serum 25(OH)D concentrations (26, 43, and 70 nmol/L). The resulting curves from the adjusted spline models were plotted to provide a visual representation of the dose-response trend. In addition, we performed weighted logistic regression modelsto assessthe odds ratios (ORs) and 95% CIs of prevalent metabolic syndrome across tertiles of serum 25(OH)D concentrations. A comparison of the highest (>52 nmol) to the lowest ($<$ 35 nmol/L) tertile of serum 25 (OH)D concentrations was of particular interest because tertile boundaries approximately coincided with clinically relevant cutoffs (\leq 30 and \geq 50 nmol/L) recently proposed in the 2011 Institute of Medicine report (28, 31).

In multivariable analyses, we first adjusted for case-control status (yes or no) and matching factors including age (continuous), race-ethnicity (white, black, Hispanic, and Asian), geographic region (Northeast, South, Midwest, West), and month of blood draw (model 1). In addition, we adjusted for smoking status (never, past, and current smokers), alcohol intake (never, past, and current drinkers), physical activity (continuous), and a composite riskfactor profile that incorporated the history of metabolic risk factors including hypertension, high cholesterol that required medication, myocardial infarction, stroke, or prior treatment of diabetes ("yes" if participant had at least one history of a risk factor and "no" if otherwise) (model 2). The composite risk-factor variable was excluded when we modeled odds of metabolic syndrome because of an overlap between conditions included in the historical profile and metabolic syndrome, the outcome of interest. Further adjustments were made for the use of supplemental vitamin D, calcium, or magnesium or multivitamins with minerals (yes or no) (model 3), and BMI (continuous) (model 4). We also adjusted for waist circumference, although because of the high correlation between BMI and waist circumference $(R = 0.84)$ we did not adjust for both variables in the same model. To further examine whether the association between serum 25(OH)D and cardiometabolic biomarker concentrations was modified by measures of adiposity including BMI and waist circumference prior history of disease, season, and vitamin D supplementation, we conducted stratified analyses by BMI (\leq 30 and \geq 30), waist circumference (\leq 88 cm and >88 cm), prior history of metabolic risk factors including hypertension, high cholesterol requiring medication, myocardial infarction, stroke, or prior treatment of diabetes (yes or no), season (winter, spring, summer, an fall), and use of vitamin D supplementation (yes or no). We also entered multiplicative interaction terms into the model for other lifestyle and demographic confounders and tested their significance by using likelihood ratio tests.

All P values were 2-tailed, and $P < 0.05$ was considered to indicate statistical significance unless otherwise specified. All statistical analyses were conducted with SAS software (version 9.2; SAS Institute).

RESULTS

As shown in Table 1, women in the highest tertile of serum 25 (OH)D concentrations were generally healthier than women with lower concentrations; these women had a lower BMI and

waist circumference, were more likely to be physically active, were less likely to have prevalent metabolic syndrome, and reported higher intakes of vitamin D, calcium, and magnesium. Metabolic profiles also varied across tertiles of serum 25(OH)D concentrations; insulin, HOMA-IR, and LDL cholesterol were lower in women with higher serum 25(OH)D concentrations, and HDL cholesterol concentrations were higher in this group.

The multivariable-adjusted geometric means of cardiometabolic biomarkers acrosstertiles of serum 25(OH)D concentrations as well as the linear regression coefficients for a corresponding increase in 25 nmol/L of the serum 25(OH)D concentration is shown in Table 2. Overall, we observed that higher serum 25(OH)D concentrations were inversely associated with insulin, HOMA-IR, HOMA- β , triglycerides, and the triglyceride: HDL-cholesterol ratio but not fasting glucose, total cholesterol, LDL cholesterol, or HDL cholesterol after controlling for age, race-ethnicity, month

of blood draw, geographic region, and case-control status in categorical analyses. Multivariable-adjusted geometric means across increasing tertiles of serum $25(OH)D$ were 11.3, 10.1, 9.9 μ IU/mL for insulin concentrations (P for linear trend < 0.0001); 2.82, 2.48, and 2.45 for HOMA-IR ($P = 0.002$); 117.9, 111.9, and 103.0 for HOMA- β (P = 0.02); 152.9, 142.4, and 117.0 mg/dL for triglyceride concentrations ($P < 0.0001$); and 2.9, 2.5, and 2.1 for the triglyceride:HDL cholesterol ratio ($P < 0.0001$) (model 1). After further adjustment for smoking status, alcohol intake, physical activity, history of cardiometabolic risk factors, use of supplements, and BMI, inverse associations with insulin, HOMA-IR, and HOMA- β were attenuated, whereas inverse associations with triglycerides and the triglyceride:HDL- cholesterol ratio remained statistically significant. With the assumption of a linear relation, an increase of 25 nmol/L of the serum 25(OH)D concentration was inversely associated with triglycerides ($\beta = -0.08 \pm 0.02$, P <

TABLE 1

Baseline characteristics according to tertiles of serum 25-hydroxyvitamin D [25(OH)D] in postmenopausal women $(n = 292)^{1}$

	Serum 25(OH)D				
	Total (median: 47 nmol/L)		Tertile $(<35 \text{ nmol/L})$ Tertile 2 $(35-51 \text{ nmol/L})$	Tertile 3 (>52 nmol/L)	P
n	292	96	94	102	
Age (y)	63.3 ± 7.5^2	62.4 ± 7.6	64.3 ± 7.5	63.4 ± 7.3	0.48
BMI $(kg/m2)$	28.7 ± 5.6	30.9 ± 6.3	28.3 ± 5.1	27.0 ± 4.6	< 0.0001
Waist circumference (cm)	87.7 ± 12.5	91.9 ± 13.1	87.5 ± 12.6	84.0 ± 10.7	< 0.0001
Waist-hip ratio	0.8 ± 0.1	0.8 ± 0.1	0.8 ± 0.1	0.8 ± 0.1	0.13
Current smoker (%)	9	13	11	$\overline{4}$	0.21
Current drinker $(\%)$	65	62	63	68	0.70
Physical activity (METs/wk)	9.8 ± 12.0	7.9 ± 10.1	9.3 ± 9.1	12.0 ± 15.3	0.02
Metabolic syndrome (%)	33	42	31	26	< 0.0001
Diabetes ever $(\%$ yes)	8	8	9	6	0.73
Hypertension ever (% yes)	35	40	34	30	0.36
Systolic blood pressure (mm Hg)	127.8 ± 16.2	127.4 ± 15.9	128.3 ± 16.2	127.5 ± 16.7	0.98
Diastolic blood pressure (mm Hg)	74.6 \pm 8.8	75.5 ± 8.4	74.6 ± 9.4	73.7 ± 8.6	0.14
Race-ethnicity (%)					
White	64	51	71	70	0.005
Black	14	24	14	6	0.001
Hispanic	11	16	$\overline{7}$	9	0.14
Asian	7	$\overline{4}$	3	12	0.03
Geographic region $(\%)$					
Northeast	24	22	30	22	0.32
South	27	29	26	26	0.84
Midwest	24	22	22	28	0.48
West	24	27	22	24	0.73
Supplemental nutrient intake					
Vitamin $D(\mu g/d)$	4.6 ± 6.4	1.8 ± 3.9	5.8 ± 5.5	6.2 ± 7.9	< 0.0001
Calcium (mg/d)	284.6 ± 407.1	100.9 ± 245.1	301.8 ± 368.0	439.7 \pm 488.7	< 0.0001
Magnesium (mg/d)	49.7 ± 89.3	21.5 ± 56.7	53.8 ± 73.8	72.1 ± 116.6	0.0001
Metabolic biomarkers					
Insulin $(\mu I U/mL)$	11.0 ± 7.0	12.9 ± 8.5	10.5 ± 6.9	9.8 ± 5.0	0.003
Glucose (mg/dL)	100.2 ± 28.3	103.2 ± 28.0	98.7 ± 26.8	98.8 ± 30.1	0.33
HOMA-IR	2.9 ± 2.7	3.6 ± 3.1	2.7 ± 2.5	2.6 ± 2.5	0.02
$HOMA-\beta$	126.7 ± 97.5	131.9 ± 87.1	129.9 ± 124.8	118.7 ± 75.0	0.33
Total cholesterol (mg/dL)	218.2 ± 37.4	220.9 ± 38.7	219.9 ± 36.6	214.2 ± 36.8	0.19
LDL cholesterol (mg/dL)	128.0 ± 34.3	133.1 ± 35.0	128.6 ± 32.5	122.6 ± 34.7	0.03
HDL cholesterol (mg/dL)	60.0 ± 15.7	56.7 ± 15.2	60.6 ± 14.2	62.6 ± 17.1	0.01
Triglycerides (mg/dL)	151.2 ± 70.3	155.5 ± 63.2	153.4 ± 81.6	145.1 ± 65.4	0.28
Triglyceride:HDL-cholesterol ratio	2.8 ± 1.8	3.0 ± 1.7	2.8 ± 2.1	2.6 ± 1.6	0.08

¹ METs, metabolic equivalent tasks; HOMA-IR, homeostatic model assessment of insulin resistance; HOMA- β , homeostatic model assessment of β cell

function.
² Mean \pm SD (all such values).

TABLE 2

Adjusted geometric means of insulin, glucose, homeostatic model assessment of insulin resistance (HOMA-IR), homeostatic model assessment of β cell function (HOMA- β), triglycerides, total cholesterol, LDL cholesterol, HDL cholesterol, and the triglyceride:HDL ratio across tertiles of serum 25-hydroxyvitamin D [25(OH)D] and linear regression coefficients for a corresponding 25-nmol/L increase in serum 25(OH)D in a sample of postmenopausal women in the Women's Health Initiative Calcium–Vitamin D (WHI-CaD) trial ($n = 292$)¹

 $¹$ Model 1 was adjusted for matching factors (age, race-ethnicity, month of blood draw, and geographic region) and case-control status (yes or no). Model</sup> 2 was adjusted for variables in model 1 plus smoking status, alcohol intake, physical activity, and history of cardiometabolic risk factors [including hypertension, high cholesterol that required medication, myocardial infarction, stroke, or prior treatment of diabetes (yes or no)]. Model 3 was adjusted for variables in model 2 plus the use of supplemental vitamins including vitamin D, calcium, or magnesium or multivitamins with minerals (yes or no). Model 4 was adjusted for variables in model 3 plus BMI. To convert 25(OH)D concentrations in nmol/L to ng/mL, divide by 2.496. ² Adjusted geometric mean; 95% CI in parentheses (all such values).

³ Linear regression coefficient \pm SD; P values in parentheses (all such values).

FIGURE 1. Restricted quadratic spline plots showing the fully adjusted geometric means (solid lines) and pointwise 95% CIs (dashed lines) of triglycerides (mg/dL) (left panel) and the triglyceride:HDL-cholesterol (HDL-C) ratio (right panel) by serum vitamin D concentrations $(mnol/L)$; $n = 292$. Three knots at medians of serum 25-hydroxyvitamin D (Serum Vitamin D) concentration tertiles (26, 43, and 70 nmol/L). All models were adjusted for matching factors (age, race-ethnicity, month of blood draw, and geographic region), case-control status (yes or no), smoking status, alcohol intake, physical activity, history of cardiometabolic risk factors [including hypertension, high cholesterol that required medication, myocardial infarction, stroke, or prior treatment of diabetes (yes or no)], use of supplemental vitamins including vitamin D, calcium, or magnesium or multivitamins with minerals (yes or no), and BMI.

0.0001) and the triglyceride:HDL cholesterol ratio (β = -0.08 \pm 0.04, $P = 0.04$) (model 4). There appeared to be a suggestion of an inverse trend with total cholesterol with a small but significant decrease observed per 25 nmol/L increase in the 25(OH)D concentration ($\beta = -0.03 \pm 0.01$; P = 0.04), although the trend was not significant in categorical analyses (model 4). After further adjustment for waist circumference (in place of BMI), the inverse associations remained significant for triglycerides ($\beta = -0.06 \pm 0.02$; P = 0.0001) and borderline significant for the triglyceride:HDL ratio $(\beta = -0.04 \pm 0.02; P = 0.10)$ (data not shown).

Smoothed dose-response curves generated from restricted quadratic spline models (model 4) were consistent with these findings and showed a decrease in triglycerides and the triglyceride:HDL cholesterol ratio as serum 25(OH)D concentrations rose (Figure 1). In subgroup analyses stratified by BMI $\ll 30$ and $>$ 30) and waist circumference ($<$ 88 and $>$ 88 cm), we observed a significant positive association with HDL cholesterol in women with higher adiposity that was not present in normalweight women (HDL cholesterol: P for trend = 0.0002 in women with BMI \geq 30; P = 0.001 in women with a waist circumference ≥ 88 cm). Although multiplicative interaction terms entered into the model were not significant (BMI: P for interaction = 0.16; waist circumference: P for interaction = 0.31), the notable difference in significant trends across groups suggested that adiposity may have modified the association of serum 25(OH)D with HDL cholesterol concentrations. Similarly, inverse associations with triglycerides and the triglyceride:HDL ratio appeared more pronounced during the winter and spring (compared with during the summer and fall) (data not shown). We observed no significant interactions with a prior history of disease, use of supplements, and other lifestyle factors including alcohol, smoking, and physical activity for all metabolic outcomes.

The multivariable-adjusted associations of serum 25(OH)D concentrations with BMI, waist circumference, and the waist-hip ratio (model 2) are shown in Table 3. Serum 25(OH)D concentrations were consistently and inversely associated with all 3 measures of adiposity ($P < 0.001$ for all) even after controlling for demographic and lifestyle risk factors including physical activity and geographic region, suggesting that serum 25(OH)D concentrations may contribute to or be affected by adiposity apart from these factors (the magnitude and statistical significance of the regression coefficients did not change materially in models 1 and 3).

As shown in Table 4, women in the highest tertile of serum vitamin D concentrations (\geq 52 nmol/L) were less likely to have prevalent metabolic syndrome compared with those in the lowest tertile $(<$ 35 nmol/L) (26% compared with 42% of women, respectively; $P \leq$ 0.0001). After adjustment for matching factors (age, race-ethnicity, month of blood draw, and region), case-control status, smoking, alcohol, physical activity, and use of supplements, the multivariableadjusted OR for metabolic syndrome for the highest $(\geq 52 \text{ nmol/L})$ compared with the lowest $(<$ 35 nmol/L) tertile of serum 25(OH)D concentrations was 0.28 (95% CI: 0.14, 0.56) (model 3). After the final adjustment for BMI, the association was slightly attenuated but remained significant.

DISCUSSION

In apparently healthy postmenopausal women enrolled in the WHI-CaD trial, serum 25(OH)D concentrations were inversely associated with triglycerides and the triglyceride:HDL cholesterol ratio, measures of adiposity, and prevalent metabolic syndrome. These associations appeared to be independent of demographic characteristics and traditional risk factors for cardiometabolic disorders. We observed no significant associations between serum 25(OH)D concentrations and LDL cholesterol, HDL cholesterol, fasting insulin and glucose, or insulin resistance and β cell dysfunction as reflected by HOMA measures.

TABLE 3

Multivariable-adjusted relations of serum 25-hydroxyvitamin D [25(OH)D] with measures of adiposity in a sample of postmenopausal women in the Women's Health Initiative Calcium–Vitamin D (WHI-CaD) trial $(n = 292)^{1}$

	Change in metabolic risk factor for each 25-nmol/L increase		
	in serum $25(OH)D$	P	
BMI $(kg/m2)$	-1.12 ± 0.30	0.0002	
Waist circumference (cm)	-3.57 ± 0.49	< 0.0001	
Waist-hip ratio	-0.01 ± 0.002	< 0.0001	

¹ All values are linear regression coefficients \pm SDs. Models were adjusted for matching factors (age, race-ethnicity, month of blood draw, and geographic region), case-control status (yes or no), smoking status, alcohol intake, physical activity, and history of cardiometabolic risk factors [including hypertension, high cholesterol that required medication, myocardial infarction, stroke, or prior treatment of diabetes (yes or no)].

TABLE 4

Multivariable-adjusted odds ratios (95% CIs) of metabolic syndrome across tertiles of serum 25-hydroxyvitamin D [25(OH)D] concentrations in a sample of postmenopausal women in the Women's Health Initiative Calcium–Vitamin D (WHI-CaD) trial $(n = 292)^T$

	Tertile 1 $(<35 \text{ nmol/L})$	Tertile 2 $(35-51 \text{ nmol/L})$	Tertile 3 $($ >52 nmol/L)	P for linear trend
Median (nmol/L)	25.7	43.0	69.6	
Metabolic syndrome				
Unadjusted prevalence $(\%)$	42	31	26	< 0.0001
Model 1	1.00	0.37(0.21, 0.67)	0.35(0.25, 0.48)	< 0.0001
Model 2	1.00	0.33(0.20, 0.52)	0.31(0.21, 0.47)	< 0.0001
Model 3	1.00	0.30(0.15, 0.61)	0.28(0.14, 0.56)	0.0002
Model 4	1.00	0.43(0.20, 0.93)	0.38(0.16, 0.91)	0.03

 $¹$ Odds ratios and 95% CIs were derived from weighted logistic regression models. Model 1 was adjusted for matching</sup> factors (age, race-ethnicity, month of blood draw, and geographic region) and case-control status (yes or no). Model 2 was adjusted for variables in model 1 plus smoking status, alcohol intake, and physical activity. Model 3 was adjusted for variables in model 2 plus the use of supplemental vitamins including vitamin D, calcium, or magnesium or multivitamins with minerals (yes or no). Model 4 was adjusted for variables in model 3 plus BMI.

Abnormalities in concentrations of triglycerides, which are a primary source of fat storage in the blood, and HDL cholesterol, and the lipoprotein responsible for the transport of cholesterol back to the liver for excretion are 2 major criteria for metabolic syndrome. The ratio of triglycerides to HDL cholesterol is also a marker for the atherogenic effect of circulating lipids (32). Although substantial variability has been observed in the association between serum 25(OH)D and dyslipidemia (11, 15, 33– 37), the inverse association with triglycerides has been reported fairly consistently in both cross-sectional and prospective cohort studies in diverse populations (8, 11, 14–16, 38–42). In line with these findings, we observed an inverse association between serum 25(OH)D and triglycerides as well as the triglyceride:HDL ratio. The potential biological mechanism underlying this relation is still not completely understood but may be mediated, in part, by the effects of dietary calcium. Higher serum 25(OH)D concentrations increase the absorption of intestinal calcium (43), which may bind to fatty and bile acids and form insoluble lipid-calcium complexes, thereby inhibiting the absorption of cholesterol and increasing fecal excretion (44, 45). Alternatively, the association may be mediated by reductions in the hepatic triglyceride formation or secretion in response to increased hepatocellular calcium amounts (46). Excess concentrations of parathyroid hormone associated with low serum 25(OH)D concentrations (47) may also drive this association; the decreased peripheral removal of triglycerides and hypertriglyceridemia has been observed in states of hyperparathyroidism (48). Some (49) but not all (50, 51) randomized trials that tested the effects of varying doses of vitamin D supplementation on metabolic outcomes have reported decreases in triglycerides in response to vitamin D supplementation, although the dosage in the null studies (50) may have been too low to achieve beneficial serum 25(OH)D concentrations. Thus, additional investigation of the relation of serum 25(OH)D to triglycerides seems warranted in large randomized settings of vitamin D supplementation.

Vitamin D receptors have been identified on pancreatic β cells (52), and the active metabolite of vitamin D, 1,25-dihydroxyvitamin D, is thought to be required for normal glucose-stimulated insulin release from β cells (53). Although cross-sectional and pro-

spective inverse associations with insulin resistance $(8, 10-13, 10)$ 33), β cell function (10, 11, 13), and glycemia (8, 10) have been reported, we observed no consistent significant associations between serum 25(OH)D and insulin, glucose, HOMA-IR, and $HOMA-\beta$. This is an interesting finding because of the biological evidence and prior cross-sectional and prospective cohort findings that reported an association. Our findings may differ, in part, because we controlled for the history of cardiometabolic risk factors in all multivariable-adjusted models, unlike in some previous studies (10, 11), which more adequately controlled for the confounding effects of adiposity. Our findings are consistent with evidence from randomized settings (49, 54–61) in which confounding by adiposity would generally not have been an issue because of randomization; a recent meta-analysis reported that 5 of 8 randomized trials of vitamin D supplementation observed no effect on fasting plasma glucose or incident diabetes (62).

Increased adiposity has been consistently associated with reduced serum 25(OH)D concentrations and adverse cardiometabolic outcomes, although the mechanism underlying the relation to serum 25(OH)D is not clear. It may be that overweight individuals at increased risk of cardiometabolic disorders are more likely to have low serum 25(OH)D concentrations because of the high lipid-solubility of serum 25(OH)D and sequestration in excess adipose tissue that result in reduced bioavailability (63). An alternate explanation is that adiposity confounds the relation because overweight individuals have less exposure to ultraviolet light because of lower levels of outdoor physical activity, which results in lower serum 25(OH)D concentrations. In the current analysis, we accounted for BMI and waist circumference as potential confounders and effect modifiers of the relation. Our primary findings were unchanged after controlling for BMI and waist circumference, which suggested that serum 25(OH)D may be related to triglycerides and the triglyceride:HDL ratio independently of adiposity. We also directly examined the relation of serum 25(OH)D to BMI, waist-circumference, and the waist-hip ratio to determine whether serum 25(OH)D was associated with adiposity independent of physical activity. Although we cannot confirm causality in this study, our findings lend support to the hypothesis that lower serum 25(OH)D concentrations may be physiologically associated with increased adiposity apart from the effects of physical activity and sunlight exposure.

In line with our findings regarding triglycerides and measures of adiposity, we observed a significant inverse association between serum 25(OH)D concentrations and metabolic syndrome. Although 3 cross-sectional studies reported no association with metabolic syndrome (64–66), the observed OR of 0.38 in the current study was approximately consistent in magnitude with 6 cross-sectional studies that reported ORs that ranged from 0.26 to 0.59 (16, 35, 42, 67–69). Inconsistencies in the literature regarding this association may be due to differences in baseline concentrations of serum 25(OH)D, with higher mean concentrations potentially making it difficult to detect an association (64). Varying control for measures of adiposity may also contribute; 2 (65, 66) of 3 null studies controlled for adiposity, whereas other studies that reported an association (16, 35, 41, 42, 69) did not control for adiposity. Our findings are of note because we observed a significant inverse association with metabolic syndrome even after controlling for BMI.

Principal limitations to this study included the relatively small sample size that limited our statistical power and the crosssectional design. To account for sample-size limitations, we compared results including serum 25(OH)D in the model continuously and categorically and plotted the adjusted results by using a restricted quadratic spline model to more realistically model the relation and allow for variation within categories. Our findings were generally consistent across methods. The crosssectional design precluded us from making causal conclusions about the effect of serum 25(OH)D concentrations on cardiometabolic outcomes; we could not rule out the possibility that metabolic disturbances that were already present led to lower serum 25(OH)D concentrations because of excess adipose tissue or other unknown mechanisms.

In conclusion, we observed no relation between serum 25(OH)D concentrations and LDL and HDL cholesterol, insulin, glucose, HOMA-IR, and HOMA- β after adjustment for demographic and lifestyle factors. In contrast, we observed a consistent inverse association between serum 25(OH)D concentrations and triglycerides and the triglyceride:HDL cholesterol ratio as well as measures of adiposity and prevalent metabolic syndrome in this population of postmenopausal women. These findings support the need for future, large-scale longitudinal studies to quantify the potential beneficial effects of increasing serum 25(OH)D on the components of metabolic syndrome, especially in obese populations. Randomized clinical trials may ultimately be necessary to confirm these findings.

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