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Individual Differences in Discount Rate Are Associated With Demand for Self-Administered Cocaine, But Not Sucrose

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Abstract

Substance abusers, including cocaine abusers, discount delayed rewards to a greater extent than do matched controls. In the current experiment, individual differences in discounting of delayed rewards in rats (choice of one immediate over three delayed sucrose pellets) were assessed for associations with demand for either sucrose pellets or an i.v. dose of 0.1 mg/kg/infusion cocaine. Twenty-four male Sprague Dawley rats were split into three groups based on sensitivity to delay to reinforcement. Then, demand for sucrose pellets and cocaine was determined across a range of fixed-ratio (FR) values. Delay discounting was then reassessed to determine the stability of this measure over the course of the experiment. Individual differences in impulsive choice were positively associated with elasticity of demand for cocaine, a measure of reinforcer value, indicating that rats having higher discount rates also valued cocaine more. Impulsive choice was not associated with demand for sucrose. Individual sensitivity to delay was correlated with the initial assessment when reassessed at the end of the experiment, although impulsive choice increased for this cohort of rats as a whole. These findings suggest that impulsive choice in rats is positively associated with valuation of cocaine, but not sucrose.

Keywords

Behavioral economics; cocaine demand; delay discounting; impulsive choice; impulsivity; Sprague Dawley rats

Impulsivity is a construct that is commonly split into at least two types of impulsive behavior: impulsive choice or inter-temporal choice and what is termed impulsive action or behavioral inhibition (Dalley *et al*, 2008; de Wit, 2009; Evenden, 1999; Perry and Carroll, 2008; Winstanley, Eagle, and Robbins, 2006). Impulsive choice is the tendency to choose smaller rewards delivered sooner over larger rewards delivered later, while impulsive action refers to the inability to withhold or inhibit a prepotent response. Impulsive choice is typically measured using procedures that provide choice opportunities between a smaller

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amount of a reinforcer delivered after little or no delay and larger amounts of the same reinforcer delivered after a longer delay (Ainslie, 1975). Animal and human models of impulsive choice are widely studied, and in humans extensive evidence links delay discounting to substance abuse. Substance abusers demonstrate a higher degree of impulsive choice than do matched controls across a range of substances, including cocaine (Coffey *et al*, 2003; Kirby and Petry, 2004).

Despite the robust relationship between delay discounting and substance abuse in people, relatively little research has examined the analogous relationship between delay discounting and drug self-administration in animals. Animal-subjects research offers many opportunities not available in human-subjects research, including the ability to easily measure behavioral characteristics both before and after a subject is exposed to a drug like cocaine. Delay discounting is modeled straightforwardly in animals with procedures that offer choices between varying amounts of some reinforcer available at varying delays (Evenden and Ryan, 1996; Mazur, 1987; Richards et al, 1997). Impulsive choice is increased after a period of self-administered cocaine (Mendez et al, 2010), and noncontingent exposure to cocaine has been shown to produce lasting (Simon, Mendez, and Setlow, 2007) or transient (Logue et al, 1992; Paine, Dringenberg, and Olmstead, 2003) increases in impulsive choice. Individual differences in impulsive choice also predicted acquisition of cocaine selfadministration in female rats (Perry et al, 2005), as well as level of nicotine selfadministration at high response requirements and reinstatement of extinguished nicotine self-administration in male rats (Diergaarde et al, 2008). It is not clear, however, whether animals that discount delayed rewards steeply value these drug reinforcers to a greater extent, as response-rate-based measures of drug reinforcement have many shortcomings (for recent reviews, see Bergman and Peronis, 2006; Hursh and Silberberg, 2008). When selfadministered, different doses of the same drug support different amounts of behavior, with moderate doses typically supporting more behavior than both high and low doses (Bergman and Peronis, 2006). This feature makes dose an influential variable when assessing reinforcer value, with no clear method of choosing which dose of a given drug best represents the reinforcing value of that drug. In addition, many self-administered drugs function to increase or decrease general activity, confounding the independent variable being assessed (e.g., drug A versus drug B) and the dependent measure (e.g., lever presses maintained by drug A versus drug B).

Behavioral economics, the application of economic terms, concepts, and analytical tools to the study of the behavior of individual organisms (Bickel, Green, and Vuchinich, 1995), provides a system of assessing reinforcer value that is independent of drug dose (Hursh and Silberberg, 2008). In such an analysis, total consumption of a reinforcer rather than response rate is the dependent measure, and this is determined at a variety of prices (response requirements). The rate at which consumption declines as price increases is termed the elasticity of demand, and this measure is assumed to reflect the reinforcer value. When demand of a reward is inelastic, an individual will defend its consumption of that reward even as price is increased dramatically, leading to large increases in total effort expended to consume that reward. As a metric of reinforcer value, elasticity of demand goes beyond a response-rate based measure by more fully characterizing the relationship between the price of a reinforcer and its consumption. Elasticity of demand can be used to rank order reinforcer value across different reinforcers (e.g., Hursh and Winger, 1995) or across different environmental conditions with responding maintained by the same reinforcer (e.g., Hursh, 1991). Note that elasticity of demand only depends on rate of decline in consumption, and does not depend on total level of responding, total consumption, or on the dose of the self-administered drug. Elasticity of demand therefore avoids many of the potential confounds introduced by rate of responding measures when assessing value of selfadministered drugs.

In the current experiment, individual differences in discounting rate, assessed by a procedure developed by Evenden and Ryan (1996), were related to individual differences in valuation of sucrose pellets and cocaine injections, assessed using behavioral economic demand analysis. The relationship between cocaine use and steep delay discounting is consistently found in human-subjects research (Coffey *et al*, 2003; Kirby and Petry, 2004), and the demonstration of a similar relationship in experimental animals would provide a framework for determining the nature of this relationship, including whether treatments that reduce impulsive choice could also reduce the valuation of drugs of abuse. Based on the strong positive relationship between delay discounting rate and demand for cocaine in rats in the present experiment. Delay discounting measures were also assessed for stability over the course of the experimental procedure with a delay discounting reassessment after demand determination.

Method

Subjects

Twenty-four male Sprague Dawley rats served as subjects (Harlan, Indianapolis, IN). Rats were approximately 10 weeks old at the start of the experiment and approximately 46 weeks at the end of the experiment. Weights were maintained at approximately 325 g throughout all four phases of the experiment, which is approximately 85% of the mean adult weight supplied by the manufacturer for this strain. When not in session, rats were individually housed in accordance with institutional animal care and use guidelines in polycarbonate cages with fresh water continuously available. The lights in the housing colony were on from 7:00 AM to 7:00 PM, and sessions were conducted between 9:00 AM and 5:00 PM. These protocols were approved by the University of Michigan Committee on the Use and Care of Animals and conformed to the guidelines established by the NIH Guide for the Use of Laboratory Animals.

Apparatus

Sessions were conducted in rodent operant conditioning chambers (ENV-008; Med Associates Inc., St. Albans, VT). Both sides of the front panel of the chamber held a retractable lever (ENV-112CM). Between the levers was a food tray connected to a 45 mg pellet dispenser (ENV-200R1AM and ENV-203M-45). Above both of the levers and the food tray were triple stimulus lights containing a red, green, and yellow LED (ENV-222M). Centered on the opposite wall was a nose-poke response device containing a yellow LED (ENV-114BM) and a houselight near the top of the wall (ENV-215M), which was unused in the current procedure. A syringe pump was located outside the chamber for drug deliveries (PHM-107). Med-PC IV software was used to control experimental events and record data.

Procedure

See Figure 1 for a schematic of the entire experimental procedure. Rats were initially trained to respond on a mixed fixed-time 60 s FR 1 schedule of reinforcement, with the active lever alternating each session. This schedule arranged one sucrose pellet to be delivered every 60 s independent of behavior, with every lever press also producing a pellet. After four sessions, no more response-independent pellets were delivered. The FR 1 schedule continued until all subjects emitted 80 responses or more on two consecutive 20-min sessions, which took 12 sessions.

Phase 1: Initial delay discounting assessment—The sessions were then extended to 75 min and split into five components of ten discrete-choice trials each. Total trial duration was 90 s and began with one or both levers extending into the chamber. If a single response

was made within 20 s, the levers retracted and the consequence programmed for that lever was delivered. If no response was made within 20 s, that trial was recorded as an omission and the levers retracted for the remaining 70 s of that trial. The first two trials of each component were always forced-choice trials where only one lever was extended into the chamber, forcing the subject to sample the contingencies for that component. The remaining eight trials were free-choice trials where both levers were extended into the chamber, allowing the rat to respond on either. The red stimulus light above each lever was lit whenever that lever was inserted in the chamber, with the left light constant and the right light flashing at 2 Hz. The green and yellow stimulus lights above the pellet tray were lit during sucrose-pellet deliveries. Initially, the consequences for the left and right levers were immediate deliveries of either one or three 45-mg sucrose pellets, respectively. This condition was continued until all the rats chose the three-pellet option on at least 85% of free-choice trials (7 sessions total). At this point, delays were introduced between responses made on the 3-pellet lever and the delivery of the 3 pellets. The delays to the three-pellet option were 0, 10, 20, 40, or 60 s and were always presented in ascending order with one delay in effect in each of the five 10-trial components. Rats were exposed to this procedure for 48 sessions.

Phase 2: Sucrose pellet demand—Demand for 45-mg sucrose pellets was then determined. Levers remained retracted throughout this procedure and the nose-poke on the back wall of the chamber was the active response device. The response was changed to a nose-poke for the demand assessments to differentiate this task from the discounting task, and to avoid confounding responses on one of the discounting levers with the demand contingencies. At the start of the 30-min sessions, the nose-poke device was lit and reinforcers were delivered on an FR schedule. The same FR schedule remained in effect for the entire session, but the FR schedule value changed between sessions. The consequence of each completed FR was a brief flash of the yellow and green stimulus lights above the pellet tray, the nose-poke light extinguishing, and a single 45-mg sucrose pellet delivered to the tray. After a 5-s timeout period, the nose-poke was illuminated and the FR schedule was again active. FR values of 1, 3, 10, 32, and 100 were examined in an ascending order. This sequence was repeated three times with an extra FR-1 session before the first sequence only, for a total of 16 sessions.

Catheter surgery—Each rat was then implanted with an indwelling femoral vein catheter for intravenous infusion of cocaine. Rats were surgically prepared under ketamine (100 mg/ kg, i.m.) and xylazine (5 mg/kg, i.m.) anesthesia. The surgical field was shaved and cleaned with betadine, and lacrilube was applied to the eyes prior to the beginning of the surgery. A small incision was made just above the femoral vein, and the overlying tissue was dissected. The wound was closed using 5-0 Ethilon suture, and the catheters were tunneled under the skin and attached to stainless steel tubing, exiting the back through a Dacron mesh tether button which was sutured to the muscle between the scapula. Rats were allowed five to seven days to recover from surgery prior to the resumption of the experiment. Catheters were flushed with 0.25 ml of saline containing heparin (100U/ml) daily to promote catheter patency.

Phase 3: Cocaine demand—Rats were initially allowed to respond for contingent infusions of 0.56 mg/kg/infusion cocaine (National Institute on Drug Abuse, Bethesda, MD) on the nose-poke device on an FR 1 schedule. This continued until rats self-administered at least 20 infusions of cocaine in a 60-min session (one subject never met this criterion). The dose of cocaine was then lowered to 0.1 mg/kg/infusion and rats were allowed to self-administer this dose for two sessions on an FR 1 schedule. The session length was then shortened to 30 min and cocaine demand was assessed as it was with sucrose pellets. The FR

sequence for cocaine demand was 1, 3, 10, 18, 32 and this sequence was repeated three times for most rats. Fewer repetitions were conducted for some rats that experienced catheter patency problems before three repetitions were complete (n = 10). Of those, three completed no full FR sequences and therefore have no cocaine self-administration data presented. The mean number of total self-administration sessions was 28.15 (SD = 3.59).

Phase 4: Delay discounting redetermination—Rats were then allowed to respond on the delay discounting procedure as described above for 43 sessions.

Data Analysis

Choice data from each rat, expressed as percent choice of three pellets at each delay to three pellets out of the total choices made (omissions were excluded), were analyzed in GraphPad Prism 5 (La Jolla, CA). To be included in the data for group analysis, delay had to significantly affect choices for the individual subject. This criterion included a significant main effect of delay on choices, determined by a one-way repeated measures analysis of variance (ANOVA) over the last five delay discounting sessions. These ANOVAs assessing the effect of delay on choices were repeated for each subject, and were not inferential in that they were not intended to indicate anything beyond the behavior patterns of individual subjects. In addition, three-pellet choice in the 60-s delay condition had to be significantly lower than in the 0-s delay condition, as measured by a planned *post hoc* comparison test. These criteria were in place due to a similar rationale to that presented in Johnson and Bickel (2008). A basic assumption of this research is that reinforcer delay is inversely related to reinforcer value, with the relevant characteristic being the degree of this inverse relationship. For rats that did not show a consistent effect of delay or showed an increased choice of delayed rewards, there is a distinct likelihood that they were not under discriminative control of the relevant contingency (delay to reinforcement). The ANOVAs with the planned *post hoc* comparisons were used to measure whether such a pattern existed for each subject. For purposes of group formation, choice data were then fit to the hyperbolic equation developed by Mazur (1987) to analyze delay discounting data:

$$V = \frac{A}{1+kD} \quad (1)$$

where V is the percent choice of three pellets at D delay, and A and k are fit parameters. A is the derived percent choice of three pellets when delayed 0 s, and k is a measure of the effect of delay on choices. A was fitted instead of set to a fixed value because, in practice, it was often less than the predicted proportion of 1, and it also provided an index of discriminability of the amount of pellets delivered.

Demand functions for sucrose pellets and cocaine infusions were plotted as reinforcers earned as a function of response requirement, known as a demand analysis (Hursh and Silberberg, 2008; Hursh and Winger, 1995). Number of reinforcers was reported as total responses divided by FR value. Data were fit with non-linear regression to the exponential equation

$$Y = \log L^* e^{-aX} \quad (2)$$

where Y is reinforcer consumption at X price and L and a are fit parameters (Winger *et al*, 2006). L represents the derived level of unconstrained consumption under the experimental conditions, and is typically nearly equal to consumption at an FR 1. The a parameter indicates the elasticity of the curve, or the rate that consumption declines with increases in price.

To better isolate the elasticity parameter, normalized demand curves were also compared. Consumption data were normalized to a proportion of consumption at an FR 1 with a procedure described previously (Hursh and Winger, 1995). The price that supported the most overall responding was also computed. This value, P_{max} , is directly related to elasticity and can be derived from the *a* parameter of Equation 2 with the equation $P_{\text{max}} = 0.29/a$. Pearson correlations were conducted in Prism 5. Distributions of *k* values were log transformed to preserve a normal distribution. The best-fit parameters of demand functions were compared between groups using non-linear regression analyses in Prism 5 which are mathematically equivalent to an Analysis of Covariance (Motulsky and Christopoulos, 2003).

To address whether response rates tended to increase or decrease within cocaine and sucrose demand sessions, a whole-session index of curvature (Fry, Kelleher, and Cook, 1960) was calculated for each session. Each session was split into ten 3-min components, and the index of curvature quantifies the degree to which response rates tended to increase or decrease across components. This index ranges from +1 (maximum positive acceleration) to -1 (maximum negative acceleration), with 0 indicating a steady response rate. For each subject, index of curvature values were averaged across all FR values and compared across conditions with a paired *t* test in Prism 5.

Results

Phase 1: Initial discounting assessment

At the end of the Phase 1 initial discounting assessment, 20 of the 24 rats met the statistical criteria for inclusion in a discounting group. These 20 rats were ranked by their individual k value from Equation 1 and split into three groups based on this ranking: High (n = 7), Medium (n = 6), and Low (n = 7). The choice to use three equally sized groups was made prior to the experiment, as it allowed for more potential differentiation between the High and Low groups than a median split. The percent choice of three pellets at each delay to three pellets is shown for the last five sessions of the initial discounting assessment in Figure 2 (top panel). A two-way ANOVA revealed an overall main effect of delay on choice ($F_{4,72}$ = 61.42, p < .001) and a main effect of discounting group ($F_{2,72} = 5.72$, p = .012). The choices of the three groups were similar when the delay to three pellets was 0 s, with a difference among the groups emerging at higher delays. This pattern resulted in a significant delay by discounting group interaction ($F_{8,72} = 7.35$, p < .001). Individual discounting functions were generally well described by Equation 1, with median $r^2 = 0.828$ (interquartile range = 0.151) for subjects meeting criteria (Figure 2, bottom panel). Mean k values (+1 SD and -1 SD in parentheses, which were asymmetrical because means and SD values were based on logtransformed k values) for the High, Medium, and Low groups were 0.030 (0.042 to 0.021), 0.014 (0.017 to 0.011), and 0.0050 (0.0085 to 0.0029), respectively. Mean A values (SD in parentheses) for the High, Medium, and Low groups were 0.90 (0.15), 0.81 (0.20), and 0.96 (0.13), respectively.

Phase 2: Sucrose pellet demand

Demand for sucrose pellets was then assessed in all 24 rats, with the 20 rats that made up the three discounting groups analyzed in detail. Demand for sucrose pellets was described well by Equation 2 with median $r^2 = 0.966$ (interquartile range = 0.027). Demand did not differ among the three groups when either the *L* parameter or *a* parameter was compared with curve-fitting procedures (Table 1; Figure 3, top panel). The corresponding P_{max} values for each of the three groups were nearly identical (Table 1). Normalized demand curves were also similar among the discounting groups (Table 1; Figure S1, top panel), and Equation 2 provided a good fit to the data with median $r^2 = 0.968$ (interquartile range = 0.032). The

Phase 3: Cocaine demand

Unlike demand for sucrose pellets, demand for intravenous infusions of 0.1 mg/kg/infusion cocaine did differ as a function of group (Table 1; Figure 3, bottom panel). The High discounting group had less elastic demand than the Low or Medium group, which was reflected by a significant difference in the *a* parameter of Equation 2 fit to these data. The *L* parameter was not different between groups, however, indicating this group difference was restricted to higher FR values. The corresponding P_{max} value for the High group (6.37) was considerably higher than the Low (3.26) and Medium (3.43) groups, and Equation 2 provided a good fit to the cocaine demand data (median $r^2 = 0.961$, interquartile range = 0.077). When cocaine demand curves were normalized, demand in the High group was still less elastic than in the Low and Medium groups (Table 1; Figure S1, bottom panel), and the normalized curves were also well-approximated by Equation 2 (median $r^2 = 0.962$, interquartile range = 0.060). The mean index of curvature for cocaine self-administration sessions was -0.11, indicating a slight deceleration of responding within the sessions. This value did not differ between sucrose and cocaine sessions ($t_{19} = 0.32$, p = .751).

Phase 4: Discounting reassessment

Data from the discounting reassessment were analyzed to determine if the discounting parameters changed following sucrose and cocaine self-administration and whether k values obtained from the discounting reassessment were also associated with cocaine demand. 21 rats remained at the end of the cocaine demand determination. Of these, 18 met the statistical criteria for inclusion in a second set of discounting groups. Two of the rats that failed to meet criteria also did not meet criteria in the initial assessment. The other rat met criteria in the original assessment, but failed to meet criteria in the reassessment did meet criteria in the reassessment.

The discounting reassessment curves, in general, appeared to be steeper than those in the initial discounting assessment (see Figure S2, bottom panel). This was confirmed by a paired *t* test on the log *k* values from the individual curves in the initial assessment and reassessment (Figure 4; $t_{15} = 2.8$, p = .013). The *A* values from Equation 1 did not differ between discounting assessments (Figure 4; $t_{15} = 1.1$, p = .306). Any changes noted between the initial discounting assessment and reassessment did not depend on initial discounting group for log *k* values ($F_{2,13} = 0.78$, p = .480) or *A* values ($F_{2,13} = 0.43$, p = .662). In those rats that completed and met significance criteria for inclusion in both discounting assessments, the respective *k* and *A* values from Equation 1 were significantly correlated across assessments (log *k* correlation r = .698, n = 16, p = .003; *A* correlation r = .770, n = 16, p < .001).

If the 18 rats that met criteria in the reassessment were again split into three groups of six rats each using k from Equation 1 fit to the individual choice data (median $r^2 = 0.879$, interquartile range = 0.133), nine of the rats that met criteria both assessments stayed in the same discounting group, while seven switched groups. Since assignment to groups was based on rank order k values, a switch among groups was not necessarily indicative of an increase or decrease in k value for that subject, but rather a change in rank compared to the other subjects. Analyzed by group in these 18 subjects, there was a main effect of delay ($F_{4,60} = 24$, p < .001) and discounting group ($F_{2,60} = 48$, p < .001) on choices of three pellets (Figure S2, top panel). These differences also tended to be larger at the higher delays, leading to a significant group by delay interaction ($F_{8,60} = 3.8$, p = .001). Note that in the

discounting reassessment, the choices of the groups were not equal in the 0-s delay component.

If performance in the discounting reassessment was used as a basis for group selection, the effect of discounting group on demand for cocaine remained, and there was still no difference in sucrose demand (Table 2; Figure S3). This indicates that it did not matter if discounting was assessed prior to or following demand, the same pattern of results remained (association with cocaine demand, but not sucrose demand).

Discussion

Subjects appeared to learn the contingencies of the operative schedule of reinforcement during each phase of the current experiment, with data largely following the expected patterns. Individual differences in delay discounting did not predict level or elasticity of demand for sucrose, but did predict elasticity of demand for self-administered cocaine injections. This relationship occurred whether delay discounting was measured prior to or after demand assessments.

Sizeable individual differences in sensitivity to delay produced discounting groups that significantly differed from one another in their choices of immediate and delayed rewards (Figure 2). Individual log k and A values correlated between the initial discounting assessment and reassessment (Figure 4), But overall, discounting was steeper in the second assessment. It should be noted that Equation 1, used here to differentiate subjects based on sensitivity to delay, is not typically used to analyze data obtained from the Evenden and Ryan (1996) procedure. For k to be a true representation of discounting rate as proposed by Mazur (1987), a series of indifferent points assessed with distinct amount and delay comparisons should first be obtained. Only one amount comparison was included in the current experiment (i.e., one versus three sucrose pellets), and therefore only one indifference point could be obtained from any subject's choice data. Curve-fitting with a single datum point is of little use. Rather, this hyperbolic function was used because it happened to describe the choice data in the current experiment well, and provided a simple one-parameter assessment of sensitivity to delay (k) and amount (A). To determine if the findings of the current experiment were a byproduct of the specific equation chosen to summarize obtained data, the slope and y-intercept of linear regression analyses drawn through choice data were also obtained and used as measurements of sensitivity to delay (slope) and amount (y-intercept). While this method of grouping subjects did not result in an identical composition of the High, Medium, and Low groups, statistical conclusions of data described in this manner were not appreciably different, and did not lead to different logical conclusions (i.e., the Linear Regression High group had significantly less elastic demand for cocaine than the Linear Regression Low group, with no difference in sucrose demand).

Individual differences in discounting were associated with elasticity of demand for cocaine (Figure 3), suggesting that impulsive choice differentiates assessments of cocaine value. An analogous relationship was not found between impulsive choice and demand for sucrose (Figure 3), indicating that differences in cocaine demand were not simply due to differences in propensity to respond for reinforcers of any type or overall level of responding. Instead, a specific relationship seems to exist between cocaine demand and impulsive choice. Previous research has suggested a relationship between impulsive choice and acquisition of cocaine self-administration in female rats (Perry *et al*, 2005). The current research extends this finding by relating elasticity of cocaine demand in male rats to impulsive choice, with a sucrose demand control condition. It is notable that level of cocaine demand (L from Equation 2), which approximates responding on an FR 1, did not differ between delay discounting groups. Only elasticity of demand differed, which agrees with the finding that

delay discounting is associated with nicotine self-administration at high FR values only (Diergaarde *et al*, 2008). Analogous to the current experiment, behavioral economic demand analyses have been conducted in human nicotine-dependent adults (MacKillop and Tidey, 2011) and alcohol users (MacKillop *et al*, 2010). Among nicontine-independent adults (MacKillop and Tidey, 2011), demand elasticity was correlated with delay discounting of a small monetary reward, but the same relationship was not found among alcohol users (MacKillop *et al*, 2010). A similar study has not been conducted in cocaine users.

Both delay discounting and demand analysis purport to assess different contributors to the value of a reinforcer. Delay discounting assesses how a delay to the delivery of a reinforcer influences value as measured by choice behavior, while demand analysis assesses the consumption of a reinforcer across a range of prices or response requirements, and is thought to measure the intrinsic value inherent to a specific reinforcer, sometimes called `essential value' (Hursh and Silberberg, 2008). Given that both procedures measure value, it is notable that discounting of delayed sucrose pellets is not related to demand for sucrose pellets, but is related to demand for the distinct reinforcer cocaine. This pattern of results suggest that delay discounting is measuring a separate and distinct contributor to value than is demand analysis, and the results of the current experiment are not simply a byproduct of measuring value in two different ways. Instead, the relationship between discounting of delayed sucrose pellets and elasticity of demand for cocaine seems to be more similar to the analogous pattern of results seen with human cocaine abusers who exhibit greater discounting of delayed money, a situation where the reinforcers for each measure also differ (Coffey *et al*, 2003; Kirby and Petry, 2004).

Due to the inherent limitations of human-subjects research, the causal direction, if any, of the correlation between cocaine abuse and delay discounting in humans has not been conclusively determined. Four explanations for this observed correlation are plausible: inborn variability in impulsive choice predisposes an individual to an increased likelihood of abusing cocaine, cocaine exposure and the life experiences associated with procuring and consuming cocaine increase impulsive choice, both characteristics cause the other in a positive-feedback loop, or both impulsive choice and cocaine abuse vulnerability are caused by an unknown third variable and do not otherwise interact. Limited evidence exists demonstrating that impulsivity, as measured by personality questionnaires and behavioral assessments, predicts development of drug abuse in human subjects (for review, see de Wit, 2009). Perhaps most relevant to impulsive choice is a single study finding that individual differences in delay discounting assessed at grade 10 predicts initiation of smoking within the following four years (Audrain-McGovern et al, 2009). The results of the current experiment, to the extent that the models employed simulate the human condition, suggest that such a positive relationship does occur between natural variability in impulsive choice and propensity to self-administer cocaine.

Whether the act of abusing a substance increases discounting has not been studied experimentally in humans, for obvious reasons. However, acute drug effects have been shown to affect impulsive choice (for review, see de Wit, 2009) and recent opioid intake influences level of discounting in opioid abusers (Giordano *et al*, 2001). However, ex-smokers and never smokers discount delayed rewards similarly and less than current-smokers, suggesting that if smoking alters impulsive choice, it does so temporarily (Bickel, Odum, and Madden, 1999). Delay discounting assessments conducted before and after the initiation of smoking also did not find evidence that smoking altered discounting of delayed rewards (Audrain-McGovern *et al*, 2009), despite discounting in smokers being higher (Baker, Johnson, and Bickel, 2003; Bickel *et al*, 1999; Heyman and Gibb, 2006; Johnson, Bickel, and Baker, 2007; Jones *et al*, 2009; Mitchell, 1999; Reynolds, 2006; Reynolds *et al*, 2007). These results suggest that impulsive choice may predict smoking, but not vice versa.

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The current study was not designed optimally to determine if cocaine self-administration influences discounting, but an increase in impulsive choice was measured from the initial discounting assessment to the discounting redetermination assessed after a period of cocaine self-administration. Increased impulsive choice with age is not typical, as impulsive choice typically negatively correlated with age in people (Green, Fry, and Myerson, 1994) and in rats (Simon *et al*, 2010). An appropriate control group that experienced all the behavioral and surgical components of the current experiment without cocaine self-administration was not included, however. This trend does agree with previous research showing noncontingent injections (Logue *et al*, 1992; Paine *et al*, 2003; Simon *et al*, 2007) and self-administered injections (Mendez *et al*, 2010) of cocaine leading to a subsequent increase in impulsive choice.

Procedural details in the current experiment that potentially affected results included the use of an ascending delay sequence in the delay discounting task, and an ascending series of FR values in the two demand determinations. However, it is unlikely that an ascending order of delays and FR values were influential, since ascending delays in the delay discounting task result in similar discounting rates as descending delays (Slezak and Anderson, 2009), and an ascending series of FR values result in similar demand curves as a random FR sequence (Raslear et al, 1988). An additional potential problem with measuring the `essential value' of a reinforcer involves the somewhat arbitrary limit set on session length of the demand sessions (30 min) and the potential that that marginal value of each additional reinforcer changed within demand sessions in a way that differed between cocaine and sucrose sessions, biasing results. To assess whether the marginal value of each successive reinforcer changed within a session, index of curvature values were collected that quantified the positive or negative acceleration of response rates within sessions. Mean index of curvature values for sucrose sessions (-0.10) and cocaine sessions (-0.11) did not differ, suggesting that the marginal value of each successive sucrose pellet and cocaine injection may have decreased somewhat within a session, but this decrease was very similar across reinforcers. Since the pattern of response rates within a session was similar across reinforcers, it is unlikely that changes in marginal value biased the measures of essential value that were obtained in the demand analyses.

In conclusion, individual differences in impulsive choice are associated with elasticity of cocaine demand, a measure of reinforcer value. This relationship holds if impulsive choice is measured before or after cocaine demand is determined, and sucrose demand is not differentiated by individual differences in impulsive choice. Impulsive choice was increased following cocaine self-administration, but the cause of this increase cannot be conclusively identified by the current experiment.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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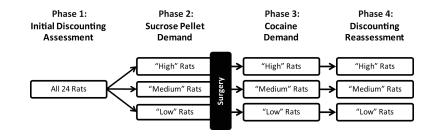
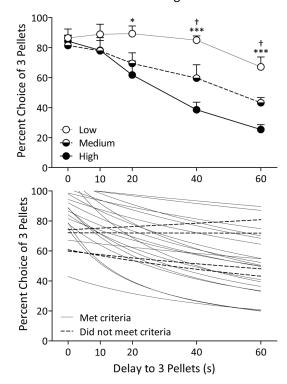


Figure 1.

A schematic of the study design. All rats experienced four phases in the order shown in the figure. Rats were split into groups after the Phase 1 discounting assessment, but all rats were treated equally and exposed to the same conditions throughout the entire experiment. See the Procedure section for a detailed description of this progression through the experiment.



Initial Discounting Assessment

Figure 2.

Choice data from the initial delay discounting assessment. Top panel: Groups of rats divided based on the *k* parameter from Equation 1 fit to the individual subject data. Data are presented as percent choice of 3 pellets as a function of the delay to 3 pellets for the High (filled symbols), Medium (half-filled symbols), and Low (open symbols) groups. Symbols near points indicate that point is significantly different from the corresponding point in the High (* p < .05, ** p < .01, *** p < .001) or the Med group († p < .05), as measured by a Bonferroni-adjusted *post hoc* test. Bottom panel: The curves obtained by fitting Equation 1 to the individual subject data. Curves in solid lines represent the 20 rats for which delay significantly affected their choice behavior, as described in the Data Analysis section. Curves in broken lines represent the four rats for which delay did not significantly reduce choice behavior. These four rats are not included in the groups in the top panel.

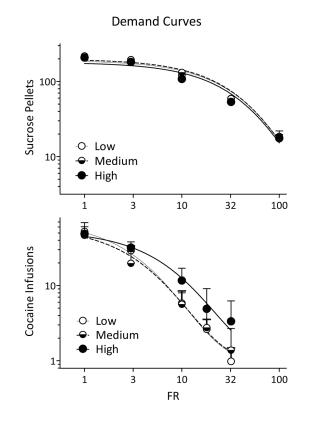


Figure 3.

Demand curves. Top panel: Demand for sucrose pellets as a function of discounting group. Data are plotted as consumption of sucrose pellets as a function of FR value. Bottom panel: Demand for cocaine injections as a function of discounting group.

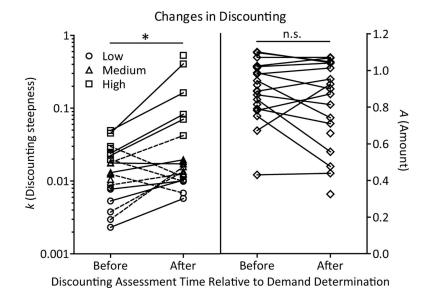


Figure 4.

Changes in discounting parameters from the initial discounting assessment to the discounting reassessment. The shapes of the symbols in the left panel indicate the group subject was placed into for the initial assessment before demand evaluation and the reassessment after demand evaluation. Solid lines connect points for a single subject when that subject stayed in the same group, while dotted lines connect points for subjects that switched groups from one assessment to the next. The results of statistical comparisons between time points are also shown (* p < .05; n.s. = not significant).

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assessment. Normalized and standard demand analyses are presented with both sucrose and cocaine reinforcers, as well as the results of F tests comparing A comparison of the demand parameters for each group of rats split on the basis of the individual k values obtained from the initial discounting these parameters across groups. Statistically significant results appear in bold.

Koffarnus and Woods

			dno to Summoorta manut	dinoin			
H	Parameter	Low	Medium	High	Ч	df	d
	Г	191	198	179	0.588	2,94	.557
Sucrose: Not Normalized	а	0.00629	0.00640	0.00644	0.0364	2,94	.964
1	$P_{ m max}$	46.1	45.3	45.0			
	а	0.00391	0.00387	0.00416	0.535	2,97	.588
Sucrose: Normalized	$P_{ m max}$	74.2	75.0	69.6			
	Т	72.8	61.3	53.8	0.214	2,73	.808
Cocaine: Not Normalized	а	0.0889	0.0846	0.0455	3.53	2,73	.034
1	$P_{ m max}$	3.26	3.43	6.37			
	а	0.118	0.108	0.0660	5.23	2,76	.007
Cocarne: Normalized	$P_{ m max}$	2.47	2.68	4.39			

 $P_{\text{max}} = Price$ supporting the maximum level of responding.

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A comparison of the demand parameters for each group of rats split on the basis of the individual k values obtained from the discounting reassessment. Normalized and standard demand analyses are presented with both sucrose and cocaine reinforcers, as well as the results of Ftests comparing these parameters across groups. Statistically significant results appear in bold.

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			Discountin	Discounting Reassessment Group	ent Group			
L 187 194 202 0.312 a 0.00615 0.00613 0.00526 0.0271 P_{max} 47.2 47.3 46.3 0.0271 a 0.000175 0.000245 0.115 p_{max} 75.2 77.3 78.1 P_{max} 75.2 77.3 78.1 P_{max} 75.2 77.3 78.1 P_{max} 75.2 60.1 0.0144 a 0.0953 0.0545 0.0144 a 0.0953 0.0545 0.0440 3.19 P_{max} 3.04 5.32 6.01 3.19 P_{max} 0.121 0.0720 0.0608 5.82		Parameter	Low	Medium	High	Ł	df	Ч
a 0.00615 0.00626 0.0271 P_{max} 47.2 47.3 46.3 0.0211 a 0.000175 0.000217 0.000245 0.115 P_{max} 75.2 77.3 78.1 0.0144 P_{max} 75.2 77.3 78.1 0.0144 P_{max} 75.2 77.3 78.1 0.0144 P_{max} 3.04 6.3.5 60.1 0.0144 a 0.0953 0.0545 0.0440 3.19 P_{max} 3.04 5.32 6.60 3.19 P_{max} 0.121 0.0720 0.0608 5.82		Т	187	194	202	0.312	2,84	.733
$P_{\rm max}$ 47.2 47.3 46.3 a 0.000175 0.000217 0.000245 0.115 $P_{\rm max}$ 75.2 77.3 78.1 1 L 65.6 63.5 60.1 0.0144 ixed a 0.0953 0.0545 0.0144 $red a 0.0953 0.0545 0.0144 red a 0.0953 0.0545 0.0440 3.19 red a a 0.0720 0.0608 5.82 red a a a a a a a $	Sucrose: Not Normalized	в	0.00615	0.00613	0.00626	0.0271	2,84	.973
$\begin{array}{c c c c c c c c c c c c c c c c c c c $		$P_{ m max}$	47.2	47.3	46.3			
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	-	в	0.000175	0.000217	0.000245	0.115	2,87	.892
$\begin{array}{cccccccc} L & 65.6 & 63.5 & 60.1 & 0.0144 \\ \hline a & 0.0953 & 0.0545 & 0.0440 & \textbf{3.19} \\ \hline P_{\max} & 3.04 & 5.32 & 6.60 \\ a & 0.121 & 0.0720 & 0.0608 & \textbf{5.82} \\ \hline P_{\max} & 2.40 & 4.03 & 4.77 \end{array}$	Sucrose: Normanzed	$P_{ m max}$	75.2	77.3	78.1			
ized a 0.0953 0.0545 0.0440 3.19 $P_{\rm max}$ 3.04 5.32 6.60 5.82 a 0.121 0.0720 0.0608 5.82 $P_{\rm max}$ 2.40 4.03 4.77		Г	65.6	63.5	60.1	0.0144	2,69	986.
$\begin{array}{ccccccc} P_{\rm max} & 3.04 & 5.32 & 6.60 \\ & & & & \\ a & 0.121 & 0.0720 & 0.0608 & \textbf{5.82} \\ P_{\rm max} & 2.40 & 4.03 & 4.77 \end{array}$	Cocaine: Not Normalized	в	0.0953	0.0545	0.0440	3.19	2,69	.047
a 0.121 0.0720 0.0608 5.82 $P_{\rm max}$ 2.40 4.03 4.77		$P_{ m max}$	3.04	5.32	6.60			
$P_{\rm max}$ 2.40 4.03	:	в	0.121	0.0720	0.0608	5.82	2,72	.005
	Cocane: Normalized	$P_{ m max}$	2.40	4.03	4.77			
	a = Demand elasticity parameter from Equation 2.	eter from Equa	tion 2.					

 P_{max} = Price supporting the maximum level of responding.