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Evaluation of disorder predictions in CASP9

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Abstract

Lack of stable three-dimensional structure, or intrinsic disorder, is a common phenomenon in proteins. Naturally unstructured regions are proven to be essential for carrying function by many proteins and therefore identification of such regions is an important issue. CASP has been assessing the state of the art in predicting disorder regions from amino acid sequence since 2002. Here we present the results of the evaluation of the disorder predictions submitted to CASP9. The assessment is based on the evaluation measures and procedures used in previous CASPs. The balanced accuracy and the Matthews correlation coefficient were chosen as basic measures for evaluating the correctness of binary classifications. The area under the receiving operating characteristic curve was the measure of choice for evaluating probability-based predictions of disorder. The CASP9 methods are shown to perform slightly better than the CASP7 methods but not better than the methods in CASP8. It was also shown that capability of most CASP9 methods to predict disorder decreases with increasing minimum disorder segment length.

Keywords

CASP; intrinsically disordered proteins; unstructured proteins; rediction of disordered regions; assessment of disorder prediction

INTRODUCTION

It has been widely accepted that that the ability of proteins to perform specific cellular functions is directly associated with their unique spatial structure¹. Numerous experiments have shown that proteins lose their activity upon loss of ordered structure due to exposure to non-physiological environments such as high temperature, urea or acid. It was also shown that a denatured protein can regain practically all of its original activity by recovering its structure upon restoration of physiological conditions². Based on these observations, the concept that proteins achieve their biological function upon folding into unique structural conformations became widely accepted. In the latest two decades ample information has been collected in evidence of proteins that do not follow this general rule^{3–6}. These so-called naturally unstructured or intrinsically disordered proteins (IDPs) lack stable structures under physiological conditions but are nevertheless biologically active. Many other proteins contain structured regions alongside extended intrinsically disordered regions (IDRs) that often play an important functional role (e.g., BRCA1, a breast cancer susceptibility protein

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damaged \overline{DNA}^7). The IDPs and proteins including IDRs are highly abundant in both eukaryotes and prokaryotes $8-10$ and tend to be enriched for regulatory functions related to molecular recognition and signal transduction^{11–14}. The number of experimentally verified IDPs and IDRs is rapidly rising¹⁵: the DisProt database¹⁶ currently contains annotations for more than 640 proteins with disordered regions, and recent reviews on this topic^{6,11,12} cite hundreds of papers. The association of several IDPs with human disease, such as cancer, cardiovascular disease, amyloidoses, diabetes, neurodegenerative diseases, and others, has triggered additional research on the subject^{17–19}.

With the high level of interest in disordered proteins, a substantial effort was placed to develop experimental and computational methods to study this phenomenon^{12,20}. Computational methods have quickly become a particularly valuable tool, in part because of their ability to keep pace with the large-scale genome sequencing projects. These techniques are based on the premise that the amino acid sequence encodes protein non-folding similarly to protein folding. Indeed, comparison of composition and complexity of protein sequences in ordered and disordered regions shows that they are statistically different^{12,21}. Based on this observation, many methods were built to predict the IDRs through recognition of amino acid motifs characteristic of disorder.

The first formal method for computational protein disorder prediction was published in 1997²², and, since then, more than fifty methods to identify disorder have been developed^{23–25}. In a recent review, He et al²⁴ provide a historical perspective of progress in this field, pointing out also the important role that the CASP experiments have played in these advancements since 2002²⁶.

The present paper analyzes the results obtained by the thirty-two disorder prediction groups participating in CASP9. While the initial round of disorder prediction (2002) was assessed by the organizers, the following three rounds were evaluated by independent assessors (2004, 2006, 2008). Since the methods to evaluate disorder prediction in CASP have developed to the point where assessments are relatively straightforward and can be made fully automatically, in CASP9 the evaluations were again performed by the organizers.

MATERIALS AND METHODS

Targets and definition of disorder

One hundred and twenty nine targets were released for modeling in CASP9 and all of them were made available for disorder prediction. The structures of twelve targets were not solved in time for prediction assessment and thus were canceled²⁷, leaving 117 targets (98 X-ray and 19 NMR structures) for assessment*. Structure data for five of these targets (T0533, T0536, T0600, T0612, T0637) were compromised in the period between the corresponding server and human prediction deadlines²⁷ and therefore only the server predictions were evaluated on these targets.

Disorder regions for each target were defined based on the best structure determination available at the time of the assessment and using the sequence released for prediction (which sometimes was slightly different than the sequence later deposited in the PDB). In cases where both the NMR and X-ray structures were available (T0551), the X-ray structure was used.

 $*$ Target T0549, which was excluded from the tertiary structure assessment as lacking big parts of structure, was retained for the disorder assessment.

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Disorder in CASP9 was defined similarly to the previous $CASPs^{28-31}$. A residue was considered to be in a disordered state if it appeared in the protein's amino acid sequence but either (1) lacked the spatial coordinates or (2) showed a high conformational variability across different X-ray chains or NMR models. We have defined "high variability" as cases where distances between positions of the same residue in any pair of models in the NMR ensemble or in any pair of X-ray chains in the asymmetric unit exceeded 3.5Å in the optimal LGA^{32} superposition†. In all other cases, the residue was assumed to be ordered. This is an oversimplification as residues may be disordered under physiological conditions but forced into the "ordered" state by crystallization. Also, long disordered regions often contain "dual personality" fragments³³ that become structured when binding to a partner. These segments are often predicted to be ordered even though they are disordered in the absence of their partner³⁴. However, such transitions are impossible to detect given only the crystal structure of the isolated protein.

Overall, 2,677 residues (or 10.2% of residues in all CASP9 targets) were classified as disordered, including 403 residues in NMR structures (or 19.9% of all residues in NMR targets) and 2,274 residues (9.4%) in X-ray targets. Thus, percent-wise, CASP9 NMR structures contain approximately twice as many disordered residues as X-ray structures. At the target level, the fraction of disordered residues is approximately the same for both types of targets, varying from 0 to 55% in X-ray structures and from 2 to 53% in NMR structures. Two targets at the high end of this range were T0603 (X-ray, 305 residues) containing 6 separate unstructured regions summing up to 55% of its length, and T0590 (NMR, 137 residues) containing two long disordered segments covering 54% of its sequence. The statistics on the number and length of the IDRs in the CASP9 targets are shown in Figure 1. Short disordered regions are much more common than the long ones. To reduce noise due to experimental uncertainty, segments consisting of less than four consecutive residues of the same order/disorder type were not considered in the assessment. After eliminating short segments, the assessment was performed on a set of 26,075 residues, including 2,417 classified as disordered. We also assessed the ability of methods to identify longer disordered regions by setting the minimum length of a disordered region to 20, 30 and 40 residues.

Participating groups and prediction format

Thirty two groups participated in prediction of disordered regions in CASP9, including 22 servers and 10 human-expert groups. These groups could submit up to five DR predictions (here called models) per target, but only models identified by the predictors as number "1" were evaluated. The overwhelming majority of groups submitted predictions on all or almost all of the targets (see Table I). The two exceptions were human-expert groups G147 and G462, which submitted predictions on 53 and 57 out of the 112 targets, respectively. We assessed the performance of these two groups but did not include them in the final rankings.

The format of the predictions in the DR category has not changed since CASP5. The predictors were asked to identify the IDRs by assigning to each residue a binary classifier of order or disorder ("O" for the ordered state and "D" for the disordered), and a probability of belonging to a disordered region (a real number in the [0;1] range). The detailed description of the DR format can be found at the Prediction Center website

<http://predictioncenter.org/casp9/index.cgi?page=format#DR>. Learning from the lessons of CASP8, in CASP9 we required that all the residues that were assigned a binary disordered/

[†]This definition slightly differs from the previous CASP definitions, in part due to the 3.5Å deviation criterion for X-ray structures, which classifies additional 246 residues (or 1.0% of all residues in CASP9 X-ray structures) as disordered.

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Evaluation criteria

Disorder predictions in CASP9 were evaluated with the *MCC*, *Acc*, and *AUC* measures also used in previous $CASPs^{29-31}$. The S_w measure was dropped from the assessment as it was shown35 to be equivalent to the *Acc*, when calculated with the weights used in CASPs6–8. The statistical significance of the differences in group performance was assessed using procedures adopted in $CASP7^{30}$, i.e. the bootstrap confidence interval method^{36,37} and the DeLong tests³⁸.

Measures for evaluating binary order/disorder predictions—In CASP, the ability to correctly assign the order/disorder tags to residues in a target has been evaluated with several measures^{28–31}: sensitivity and specificity (used in CASP5–8), statistical accuracy $Q2$ (CASP6), Matthews correlation coefficient *MCC* (CASP6), the weighted score S_w (CASP6– 8), and the balanced accuracy *Acc* (CASP7–8).

The disorder prediction data are characterized by a large class imbalance: in the latest five CASPs ordered residues outnumbered disordered ones 9 to 1 or higher. As disordered residues are relatively rare and therefore harder to predict, their correct prediction should be rewarded more generously than the prediction of ordered residues, and vice versa – the incorrect prediction of disordered regions should be penalized less severely than the incorrect prediction of ordered residues. Not all measures are equally effective in handling these tasks. Below, we briefly discuss the relative strengths and weaknesses of the aforementioned evaluation measures for the disorder assessment.

Sensitivity and specificity

$$
Sensitivity = \frac{TP}{TP + FN} = \frac{TP}{N_d},
$$

$$
Specificity = \frac{TN}{TN + FP} = \frac{TN}{N_o}
$$

are the two statistical measures routinely used for evaluating the accuracy of a two-class binary predictor. In prediction of disorder, TP (true positives) and TN (true negatives) are the numbers of correctly predicted disordered and ordered residues, respectively; FP (false positives) and FN (false negatives) are the numbers of misclassified ordered and disordered residues, and *Nd* and *No* are the total numbers of disordered and ordered residues in all targets predicted by a particular group. Specificity determines the fraction of negative examples (ordered residues) correctly identified in a prediction. For datasets dominated by negative examples, specificity is high for practically all predictors and therefore is not a discriminative measure of prediction quality. Sensitivity represents the fraction of positive examples (disordered residues) correctly identified in a prediction and has a better discriminative power but at the same time is completely insensitive to negative examples (see the corresponding formula). Predictors can increase the sensitivity or specificity of their classifiers by deliberately predicting more residues as disordered or ordered, respectively. There is a tradeoff between these two measures and increasing one of them usually leads to decreasing the other. A prediction method can be considered to perform well only if it scores high in both sensitivity and specificity; neither of these two measures is a good estimator of

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methods' strength when used alone. The one-sided nature of sensitivity and specificity can be overcome by employing measures that use all four parameters of prediction quality (TP, FP, TN and FN).

The statistical accuracy (used under the name of *Q2* in previous CASP disorder prediction assessments) is calculated according to

$$
Q2 = \frac{TP + TN}{TP + FN + TN + FP}
$$
\n⁽¹⁾

and accounts for all four components of prediction quality. Nevertheless, it strongly favors conservative classifications (i.e. predicting more residues as ordered)^{29,30} and therefore is not well suited for disorder assessment.

The balanced accuracy *Acc*

$$
Acc = \frac{Sensitivity + Specificity}{2} = \frac{TP}{TP + FN} + \frac{TN}{TN + FP}
$$
\n(2)

is a much better measure as it does not reward over-prediction of the ordered state. On contrary, it has a desired feature of rewarding prediction of disordered state more generously than the prediction of the ordered, but it is also known to strongly favor greedy classifications (i.e. predicting more residues as disordered) 29 .

The Matthews correlation coefficient (*MCC*)

$$
MCC = \frac{TP \cdot TN - FP \cdot FN}{\sqrt{(TP + FP)(TP + FN)(TN + FP)(TN + FN)}}
$$
\n(3)

does not favor over-prediction of any of the prediction classes and had been recommended for handling cases with skewed class frequencies^{39,40}. MCC varies between −1 and 1 with a random prediction scoring zero. It was noticed, though, that *MCC* can yield unreasonably high scores in cases where prediction algorithms assign very few or no false positives and at the same time very few true positives⁴¹. As this situation can happen in DR prediction (overprediction of ordered residues), we have conducted a numerical experiment to estimate the scale of possible discrepancies. We have run these calculations on artificial datasets with the TP, TN, FP and FN values varying in the ranges typical of the CASP9 data. The *MCC* appeared to yield reasonable and consistent scores for all combinations of prediction characteristics, leading to a conclusion that in general it does not overinflate scores for overprediction of ordered residues in our data.

The general conclusions on the effectiveness of measures (1) – (3) hold true for the CASP9 data. First, the tendency of *Q2* to unreasonably favor conservative predictions can be illustrated by an example of two CASP9 groups: G291 and G067. Group G291 is ranked high according to all three measures used in our evaluation (Table I), while group G067 is at the very bottom of the table. Surprisingly, G067 outscores G291 0.91 to 0.87 according to *Q2* (data not shown). This result can be directly attributed to more conservative predictions submitted by G067 (only 391 residues predicted as disordered; the remaining 98.5% of residues predicted as ordered, the highest figure in CASP9). Second, both *Acc* and *MCC*, reproduce the overall trends in prediction quality fairly well but emphasize numerical contributions from TP, TN, FP and FN differently (on CASP data, the Spearman ranked correlation coefficient between these two measures is only $p=0.56$). The balanced accuracy

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(*Acc*) generously rewards correct prediction of disordered regions and mildly penalizes their incorrect prediction, encouraging development of riskier methods tuned to identify large numbers of disordered residues. The *MCC* is more balanced, it does not reward "greedy" predictions as strongly as the *Acc* does, but instead rewards classifiers with higher predictive precision

$$
precision = PPV = \frac{TP}{TP + FP}.
$$
\n(4)

The difference between the *Acc* and *MCC* can be illustrated by the example of groups G119 and G015 (Table I). Group G119 is one of the most "greedy" CASP9 classifiers. It has predicted 5115 residues as disordered, but only 1570 of these classifications had been correct (31%). Group G015 predicted only 1019 residues as disordered, of which 839 were correct (82%). The *Acc* score favors G119 as able to identify almost twice as many disordered residues as G015. At the same time the *MCC* favors G015 for obtaining a much higher level of precision, while still predicting a relatively high number of disordered residues. The decision of which of these two measures should be used in assessments is to some degree subjective and therefore we present the results of both, noting the better balance of the *MCC*.

In previous three CASPs, also the S_w score was extensively used in assessments

$$
S_w = \frac{w_d TP - w_o FP + w_o TN - w_d FN}{w_d (TP + FN) + w_o (TN + FP)}.
$$

This measure was specifically designed²⁹ to address the imbalance in the ratio of ordered and disordered residues through adjustable weights w_o and w_d . It was recently shown³⁵ that for the weights used in CASP6–8

$$
w_o = \frac{N_d}{N_o + N_d}, \ w_d = \frac{N_o}{N_o + N_d}
$$

this score is equivalent to the *Acc* as there is a linear relationship between the two:

$$
S_w=2Acc-I.
$$

Therefore we kept only one of these measures (*Acc*) in our analysis.

In addition to the scores used in previous CASPs we have tested other evaluation measures. One such measure is the *F*-score, which, similarly to the *MCC*, had been recommended to handle skewed data $42,43$. Our calculations on CASP data have shown a high correlation of this measure with the MCC (Spearman's $p=0.9$) and therefore these results are not shown.

Measures for evaluating probability-based predictions of disorder—The ability to identify the IDRs through assigning per residue disorder confidence scores [0;1] was assessed with the receiver operating characteristic (*ROC*) analysis. This method is frequently used to assess the accuracy of a classifier, and has been previously used in the assessment of protein disorder predictions (both in CASP and elsewhere) $29-31,44$.

In essence, a *ROC* curve illustrates the correspondence between the true positive rate of a predictor (*Sensitivity*) and its false positive rate (*FPR = FP*/(*TN + FP*) 1 − *Specificity*) for a

set of probability thresholds (from 0 to 1 in our case). For each threshold, a residue is considered as a positive example (disordered) if its predicted probability is equal to or greater than the threshold value. The area under a *ROC* curve (*AUC*) is indicative of the classifier accuracy. An *AUC* of 1 identifies a perfect predictor, while an *AUC* of 0.5 corresponds to a random classifier. We have computed the *AUC* scores using the trapezoid

Statistical significance of differences in group performance—Performance of groups as binary order/disorder classifiers was statistically compared using the re-sampling procedure. For each group, 80% of targets were randomly drawn from the list of targets predicted by that group and the evaluation scores were re-calculated on that subset. The procedure was repeated 1000 times, and a discrete distribution of the two-class classifications was learned for every group. Based on these distributions, we have calculated the 95% confidence intervals for each assessment measure using the two-tailed bootstrap percentile method $36,37$. Statistical significance of the differences in group performance was inferred based on the comparison of the confidence intervals obtained for each group⁴⁵.

integration rule with a threshold increment of 0.01.

Performance of groups as predictors of the per-residue disorder probabilities was compared using the DeLong non-parametric tests 38 , designed to assess the statistical significance of the differences between the *AUC* scores in the *ROC* analysis. The evaluation was performed using the statistical package R^{46} and the *pROC* library⁴⁷.

RESULTS

Performance of disorder prediction methods

Numerical evaluations of DR predictions for all groups participating in CASP9 are summarized in Table I and illustrated in Figure 2. Scores from the main three evaluation measures used in our assessment (*Acc*, *MCC* and *AUC*) are provided together with the ranges of the corresponding 95% confidence intervals and the group ranks. The ROC curves based on the continuous-scale disorder predictions are plotted in Figure 3. Note that the uneven distribution of the assigned probability scores can affect the smoothness of ROC curves, which is imperative for an accurate calculation of the *AUC* scores. In CASP9, all top-ranked groups have assigned sufficiently distinct probabilities to enable an accurate calculation of the *AUC* scores. The only exception is group G193, which submitted predictions yielding good scores according to the binary-classification measures *Acc* and *MCC* but poor *AUC* scores in the probability-based analysis. This was due to uniformly assigning a value of zero to all residues predicted as ordered.

Table I shows that *prdos2* is the only group to rank among the top three prediction groups according to all three evaluation measures. In addition to this group, there are three other groups (*Zhou-Spine-D, Multicom-refine and biomine_dr_pdb*) to rank among the best 10 groups according to all three measures.

Figure 2 shows that there are several groups that perform equally well according to the *Acc* measure (grey bars). However, as we have discussed in Materials and Methods, some high *Acc* scores may be an artifact due to over-prediction of disordered residues. As the scores for top groups are very close, the statistical significance of the differences between them could not be established by the comparison of the confidence intervals.

Group *DisoPred3C* has obtained a relatively low *Acc* score but, at the same time, the best and the second best *MCC* and *AUC* scores, respectively. The high *MCC* score can be attributed to the high (highest in CASP9) precision (4) of classifications submitted by this group. The low *Acc* score is most likely due to the relatively low levels of disorder

prediction. Based on the comparison of confidence intervals for the *MCC* scores, results of *DisoPred3C* are statistically better than those of all other groups, except for *biomine dr_pdb_c*, which is second best according to the *MCC*.

Groups *prdos2*, *DisoPred3C* and *Multicom* are the best performing groups according to the probability-based assessment. These groups are statistically indistinguishable from each other by the *AUC* score and better than all other groups according to the results of the DeLong tests (see Table II). This conclusion is also confirmed by the comparison of the *AUC* confidence intervals (Table I).

Evaluation of results for longer disorder regions

As noted earlier, short disordered regions prevail in the CASP data set (see Figure 1). While such regions may sometimes consist of chain termini or short loops without any obvious functional role, they are often of functional importance (for example, flaps over enzyme active sites, pieces of chain that order into DNA grooves, or loops that become ordered in protein-protein interfaces), so their inclusion in the methods testing is important. At the same time, long disorder regions require separate attention as they are found in abundance in the human disease-associated proteins 48 and their properties and functional roles are likely different from those of short disorder regions (for example, ordering of complete domains upon complex formation). The issue of the different length of disordered regions has been taken into consideration in several disorder predictions methods^{49–53}. To address this issue in the assessment, we additionally evaluated the predictions taking into account only segments longer than a specified length cutoff.

Figure 4 compares results of CASP9 methods for four minimum length thresholds: 4, 20, 30 and 40 residues. The "average group" splines ('AVG', thicker line) in all three panels of the graph show that the discriminatory power of the methods tends to decrease with the increase of the minimum disorder segment length. The average drop in performance is moderate according to the *Acc* and *AUC* scores and more pronounced according to the *MCC* score. The *MCC* panel suggests that an average CASP9 method can identify 40+ residue long disorder segments just slightly better than a random predictor (*MCC*=0). It should be mentioned, though, that the results for disorder regions spanning 40 residues or more should be interpreted with caution as there were only four qualified segments constituting to only 0.8% of all residues in CASP9 targets.

Curves for the vast majority of participating groups follow the average trend to decrease, resulting in high correlation (0.85 – 0.98) between the scores of the same evaluation measure at neighboring length thresholds. The lowest (even though still high in absolute value: 0.85) correlation between the 4+ and 20+ *Acc* score sets reflects the fact that this score was the most prone to the shifting of ranks. While the majority of groups performed worse in identifying 20+ residue long disorder segments (compared to 4+ segments), there were five groups that performed somewhat better, with two of them - *DisoPred3C* (G015) and *GSMetaDisorder3D* (G421) - improving their *Acc* scores significantly (by more than 6%) and consequently raising their ranks by 13 positions (to #8 and #7). *DisoPred3C* is the only group that demonstrated an ability to better discriminate 20+ residue long disorder regions according to all three evaluation scores, and is the best group in this length range according to the *MCC* and *AUC* scores. This group also has quite high scores for the 30+ residue-long regions, comparable to those they obtained for the 4+ ones. *GSMetaDisorder3D* also proved to be successful in identifying longer disorder segments, consistently placing in the best three according to the *MCC* and *AUC* scores at all longer disorder length levels.

Comparison between recent CASPs

To compare the accuracy of disorder predictions across all CASPs, we have re-evaluated predictions in previous CASPs using exactly the same disorder definitions and evaluation measures as in CASP9. Even so, it is hard to ensure full objectivity of such a comparison as the targets, methods and databases change in time.

Figure 5 shows the results of a comparison of the *MCC* scores for the twelve best performing groups in the latest three CASPs. The CASP9 scores are higher than those in CASP7 but lower than in CASP8. This tendency holds when these methods are compared using other scores (see Figure S1 in Supplementary Material). As the majority of the top performing CASP8 methods were among the best also in CASP9, the drop in performance is most likely due to a greater difficulty of the CASP9 targets⁵⁴.

CONCLUSIONS

The number of disorder prediction methods published in the literature is now well over fifty²⁴ and continues to grow as new methods continue to appear^{55–58}. This growth correlates well with the increase in the number of disorder prediction groups participating in CASP experiments. However, the increased number of participating groups does not seem to result in a better performance. Rather, our analysis show that the scores obtained in CASP9 have slightly decreased in comparison with those in CASP8 according to all three measures used in the CASP9 evaluation. As we discussed in this paper, this might be related to a higher difficulty of the CASP9 targets but perhaps also to the lack of conceptually new methods. New meta-predictors or slight modifications of established methods were not sufficient to achieve substantial progress in the field. By analyzing the submitted abstracts we could identify only one group (*Zhou-Spine-D*), claiming development of a conceptually new method based on neural networks. This method was assessed to be among top 10 according to all scores in CASP9, but did not outperform other already established CASP performers. A brief description of the best performing automatic methods participated in CASP9 is provided in Table III. It seems that performance of disorder prediction methods in CASP has reached a plateau and new breakthroughs are needed.

Besides more effective disorder prediction methods, we also need better target sets in CASP, since the vast majority of targets are solved by X-ray crystallography and therefore typically contain only short disorder regions. This type of data likely does not fully represent the type of disorder observed in functionally relevant long disordered segments. Thus, test sets containing more targets with extended disordered regions are required for more comprehensive testing of disorder prediction methods.

For the first time, we have analyzed differences in the capability of methods to recognize disorder regions of different length. The surprising result is that, independent of the exact evaluation metrics, there is a rather dramatic fall-off in performance with disorder length increase. Perhaps this reflects a tendency for the methods used for CASP to be trained on the short disorder segments typical of the targets. Nevertheless, it is a disturbing result.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

Abbreviations

3D three-dimensional

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Figure 1. Length distribution of disordered regions in CASP9 target proteins.

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Figure 2.

Performance of DR groups according to three evaluation scores: *AUC* (black bars), *Acc* (grey bars) and *MCC* (light grey bars). The groups are sorted according to decreasing *AUC* score. The error bars on the plot indicate boundaries of the 95% confidence intervals for each measure.

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Figure 3.

ROC curves of disordered region predictions for all CASP9 groups. Legends are shown for the best 12 groups according to the *AUC*. There are four non-regular ROCs corresponding to poorly performing groups, two of which misinterpreted DR format (G193 used only a single value for ordered residues and G114 did not use continuous scale but rather 5 different numbers).

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Figure 4.

Comparison of prediction performance across four different minimum disorder segment length thresholds. Different panels show scores for different evaluation measures (*Acc, MCC* and *AUC*). Each group is marked with a different color; groups in the legend are sorted according to the *AUC* score (across and then down); the artificial average group ('AVG', black thicker line) is added to the graph as a point of reference.

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\bullet CASP7 \blacksquare CASP8 \blacktriangle CASP8* \circ CASP9

Figure 5.

Comparison of the performance of the best 12 groups in the latest 3 CASPs. Groups in each CASP are sorted according to the MCC score. CASP8 results are evaluated for both the full set of targets and the set without target T0500, a long, completely unfolded protein considerably influencing the scores. The reduced target set is marked with an asterisk in the legend. The scores in CASP9 are higher than in CASP7 but lower than in CASP8. The CASP8–CASP9 drop in scores may be attributed to the greater difficulty of targets in CASP9.

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The groups are sorted according to the AUC score. The two groups at the bottom of the table (in grey) are not included in the rankings as they submitted predictions on less than half of the targets. The groups are sorted according to the *AUC* score. The two groups at the bottom of the table (in grey) are not included in the rankings as they submitted predictions on less than half of the targets.

Table II

Statistical comparison of the best 12 groups according to the AUC score Statistical comparison of the best 12 groups according to the *AUC* score

Results of the non-parametric DeLong tests on the *AUC* scores. Shaded cells highlight pairs of groups that are statistically indistinguishable at the 0.05 significance level.

Results of the non-parametric DeLong tests on the AUC scores. Shaded cells highlight pairs of groups that are statistically indistinguishable at the 0.05 significance level.

Table III

Methods description for the best CASP9 DR servers

