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Striatal Volume Contributes to the Prediction of Onset of Huntington Disease in Incident Cases

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Abstract

Background—Previous neuroimaging research indicates that brain atrophy in Huntington disease (HD) begins many years before movement abnormalities become severe enough to warrant diagnosis. Most clinical trials being planned for individuals in the prediagnostic stage of HD propose to use delay of disease onset as the primary outcome measure. Although formulae have been developed, based on age and CAG repeat length, to predict when HD motor onset will occur, it would be useful to have additional measures that can improve the accuracy of prediction of disease onset.

Methods—The current study examined MRI measures of striatum and white matter volume in 85 individuals prospectively followed from pre-HD stage through diagnosable motor onset (“incident cases”) and 85 individuals individually-matched with incident cases on CAG repeat length, sex, and age, who were not diagnosed with HD during the course of the study.

Results—Volumes of striatum and white matter were significantly smaller in individuals who would be diagnosed 1 to 4 years following the initial MRI scan, compared to those who would remain in the pre-HD stage. Putamen volume was the measure that best distinguished between the two groups.

Conclusions—Results suggest that MRI volumetric measures may be helpful in selecting individuals for future clinical trials in pre-HD where HD motor onset is the primary outcome measure. In planning for multisite clinical trials in pre-HD, investigators may also want to consider using more objective measures, such as MRI volumes, in addition to onset of diagnosable movement disorder, as major outcome measures.

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Keywords

Huntington disease; MRI; onset; prediction; prospective; striatum

Introduction

Huntington disease (HD) is an autosomal dominant neurodegenerative disorder caused by a CAG repeat expansion in the HTT gene on chromosome 4 (p16.3) (1). HD diagnosis is based on the presence of extrapyramidal movement abnormalities, typically beginning in the patient's 40s or 50s, but the features of the disease also include cognitive decline and psychiatric impairment (see Ross & Tabrizi (2) for a review of disease pathogenesis and characteristics). Discovery of the HD gene in 1993 made it possible to identify individuals who have the gene mutation prior to the onset of symptoms. Previous neuroimaging studies indicate that significant brain atrophy, particularly in the striatum, occurs many years before diagnosis of HD (3,4). Thus, treatment will likely be most effective if it is administered prior to diagnosis. Most clinical trials that are currently being planned for subjects in the pre-diagnostic stage of HD propose to use delay of onset of HD as the primary outcome measure.

Because CAG repeat length is inversely correlated with age of onset (5,6), it is possible to predict approximate age of onset of HD motor impairments for individuals who have the HD mutant allele. Several formulae have been derived from retrospective data to predict when diagnosable motor signs of HD will occur (3,7,8). As the relationship between age at onset and the length of the CAG repeat expansion accounts for only 47% (8) to 73% (9) of the variance, it would be useful to identify other measurements that could be used to improve the accuracy of predicting the age at which motor impairment will become severe enough to warrant diagnosis. The current study was designed to determine whether neuroimaging measures could be useful in distinguishing individuals prospectively diagnosed with HD ("incident cases") from individuals of equivalent CAG and age who remain in the pre-diagnostic stage of disease ("pre-HD"). Improving prediction of age at diagnosis will be useful in designing clinical trials in pre-HD that rely on diagnosis of HD as the major outcome measure.

Methods and Materials

Participants

The analyses presented here are based on a sample of 170 participants from PREDICT-HD (10), an international multi-site study following a large sample of individuals who are at risk for HD by virtue of having a parent with HD. All participants for the current study tested positive for the HD gene mutation but were not diagnosed with HD at study enrollment ("pre-HD"). One half of these individuals were diagnosed with the disorder sometime during the course of the study ("incident cases"), and the other half ("non-diagnosed") were selected as individual matches for the incident cases, based on sex, age (within 2 years), and CAG repeat length (within 1). All incident cases had MRI data at least one year prior to being diagnosed; for 42 of the incident cases, the MRI was obtained two years prior to diagnosis; for 3 incident cases the MRI was obtained three years prior to diagnosis; and for another 3 the MRI was obtained four years prior to diagnosis. All non-diagnosed participants were seen at least one year following the MRI visit and none had yet received a diagnosis of HD at the time of data analysis (23 had verification of non-diagnosed status 2 or more years after the MRI; 17 had verification 3 or more years after the MRI; and 11 had verification 4 or more years after the MRI). Participants for this study came from 28 sites, with 1 to 16 participants per site. It was not possible to match the incident cases and non-

diagnosed participants on site. However, there was no systematic bias in the diagnostic status of participants coming from any given site. Table 1 provides demographic and clinical information on the two groups. All aspects of the study were approved by the Institutional Review Board at each participating institution, and all participants gave written informed consent.

PREDICT-HD participants are seen yearly by clinicians experienced in the evaluation of movement disorders and specifically trained for PREDICT-HD on administration of the Unified Huntington's Disease Rating Scale (UHDRS) (11). Using this standardized scale that includes a series of specific assessments of HD-related motor movements, the clinician assigns a Motor Score, ranging from 0 to 124, and then assigns a score from 0 to 4 on the HD Diagnostic Rating Scale indicating the rater's level of confidence that the motor abnormalities reflect the presence of HD. In accordance with clinical practice (12), HD diagnosis is operationally defined as a score of 4, indicating that the rater has $\geq 99\%$ certainty that the participant shows "unequivocal presence of an otherwise unexplained extrapyramidal movement disorder." Participants were excluded from the current study if they received a rating of 4 at baseline. All incident cases received a rating of 4 sometime during the course of the study; all non-diagnosed participants received scores of < 4 at all visits prior to the MRI scan and at least one year following the MRI scan.

Because of the large number of clinicians involved in this multi-site study and because it was not possible to individually match incident cases with non-diagnosed participants from the same sites, there was some concern about the consistency of criteria used for making diagnoses. A reanalysis was performed using groups that were based on UHDRS Motor score rather than on clinical diagnosis. The UHDRS Motor score of each incident case at the time of diagnosis was noted, along with the number of years between the scan and the diagnosis ("scan-to-diagnosis interval"). For each matched non-diagnosed case, we identified the UHDRS Motor score that was obtained at the same post-scan interval as the matched pair's scan-to-diagnosis interval. A review of these UHDRS Motor scores revealed some overlap between the incident cases and non-diagnosed cases in the range of 9 (the lowest score for an incident case) to 26 (the highest score for a non-diagnosed participant). The UHDRS Motor score that best discriminated the incident cases from the non-diagnosed participants was 10.5. Thus, the groups were reformed, using this cut-off rather than actual clinical diagnosis. As a result, there were 13 of the original 85 pairs in which both members of the pair now fell in the same group. For 2 pairs the incident case had a UHDRS Motor Score < 10.5 , and in 11 pairs the non-diagnosed participant had a UHDRS Motor Score > 10.5 . Group analyses were then redone omitting these 13 pairs.

MRI Acquisition and Analysis

All MRI scans were obtained using a standard protocol that included an axial 3D volumetric spoiled gradient echo series and a dual echo proton density/T2 series. Scans were processed at The University of Iowa using AutoWorkup (13), an automated procedure implemented in BRAINS (14) and artificial neural networks (15). Volume measures were determined for caudate, putamen, total striatum (caudate + putamen), and "cortical" white matter volume (excluding white matter in the cerebellum, brainstem, and subcortical region). These regions were selected based on our previous research examining which brain regions showed the greatest volumetric differences between pre-HD and gene-negative individuals (10). After completion of AutoWorkup, all scans were individually inspected for correct realignment and coregistration, tissue classification, and accuracy of brain and subcortical structures. Participants were included in this study only if they had scans that passed inspection for all measures. Intracranial volume was also calculated to allow for correction of structural volumes for overall head size. (See Supplement for details on scan acquisition and analysis.)

Statistical Analysis

The sample consisted of 1-to-1 matched pairs. For continuous outcome variables, paired *t*-tests were conducted to examine group differences in MRI volumes (corrected for intracranial volume), based on either diagnostic status (incident vs. non-diagnosed) or UHDRS-Motor group (10 and under vs. 11 and above). For binary outcome variables, conditional logistic regression was performed to determine which imaging variables (corrected for ICV) best predicted group membership. Each imaging measure was standardized to have a mean of 0 and standard deviation (SD) of 1. Each of the standardized variables was analyzed separately using conditional logistic regression. All reported confidence intervals (CI) are two-sided and have confidence levels of 95%, unless specified otherwise. A significance level of 0.05 is used for all hypothesis testing. All statistical analyses were carried out with SAS v9.1 (Cary, NC).

Results

Volumes of caudate, putamen, striatum, and white matter (corrected for intracranial volume) were significantly smaller for incident cases than for the non-diagnosed participants (see Table 1). Using conditional logistic regression, the results for caudate show that for any two matched subjects with the same age, sex, and CAG length, if the case subject's caudate was 1 SD unit less than the control (non-diagnosed) subject, the odds of HD diagnosis for the case subject had a multiplicative increase of 1.9, which was equivalent to a 47.0% increase in the odds that the case subject would be in the diagnosed group, as compared to the control subject (95% CI = 21.5%–64.3%). For putamen, a 1 SD unit lower score for a case subject was accompanied by a multiplicative increase in the odds of HD diagnosis of 3.1, which was a 68.1% increase in odds (95% CI = 44.9%–81.6%). For total striatum, the increase in the odds was 2.7 (63.5% increase) (95% CI = 44.9%–81.6%), and for white matter the increase in odds was 1.7 (41.4% increase) (95% CI = 12.7%–60.3%). Multivariate logistic regression was then performed to determine which imaging variable had the highest predictive value of diagnostic status (incident vs. non-diagnosed). When the caudate, putamen, and total white matter volume were all included in the model, multivariate conditional logistic regression showed that the putamen was the only variable which was statistically significant, $\chi^2(1) = 7.9, p = .0049$.

When the data were reanalyzed based on the UHDRS-Motor score groups (10 and under vs. 11 and above), group differences became somewhat more robust (see Table 2). Multivariate logistic regression continued to demonstrate that putamen was the only variable which was statically significant in predicting group membership, $\chi^2(1) = 9.1, p = .0025$.

Although the group analysis showed that MRI measures were significantly related to future diagnosis, it was not the case that *all* incident cases had smaller striatal volumes (corrected for ICV) 1 to 4 years prior to diagnosis as compared with age- and CAG-matched participants who remained non-diagnosed (Figure 1). For the majority of matched pairs, however, the relevant structure volumes were larger for the non-diagnosed participant than for the matched incident case. The number of matched pairs with this pattern (i.e., larger volume of any magnitude for the non-diagnosed participant) was 56 (65.88%) for caudate, 61 (71.76%) for putamen, 62 (72.94%) for total striatum, and 50 (58.82%) for white matter. For the UHDRS-Motor groupings, the number of matched pairs (out of 72) in which the member with the lower UHDRS-Motor score had a larger volume than the member with the higher score was 48 (66%) for caudate, 56 (77%) for putamen, 54 (75%) for total striatum, and 45 (63%) for white matter.

Discussion

Results from this study demonstrate that volumetric MRI measures can aid in the prediction of diagnosis of HD in individuals, 1 to 4 years prior to disease onset. All of these regions were significantly smaller in those individuals who would be diagnosed 1–4 years later, as compared to those whose signs and symptoms would not be severe enough to warrant a diagnosis, even though the groups were well-matched on age, sex, and CAG repeat length.

Among the regions studied, putamen contributed most to prediction of diagnosis. As diagnosis is based on motor signs (and not cognitive or psychiatric impairment), it is possible that this reflects a stronger association between putamen and motor signs than between caudate and motor signs, as suggested by some early research in HD (16) and lesion studies (17). Cross-sectional studies by our group and others have suggested that putamen volume reduction may be slightly greater than caudate volume reduction at all pre-HD stages (4) and in early stages of manifest HD (18,19), and that the correlation with estimated years to onset may be slightly higher for putamen volume than for caudate volume (20). Longitudinal studies, however, suggest that there is somewhat greater change *over time* in caudate than in putamen volume, both in pre-HD (21) and manifest HD (22,23) and that the rate of atrophy may become significantly greater than zero earlier in the pre-HD period for caudate than for putamen (24). Our longitudinal findings are not consistent, however, with longitudinal functional imaging studies using PET (positron emission tomography) that show a slightly greater change over time for putamen than caudate (25,26). While white matter volume did not predict group status (incident vs. nondiagnosed) as well as striatal volume, it is possible that other methods of assessing white matter, such as diffusion tensor imaging would yield different results.

Strengths of this study include the relatively large sample of individuals who have been followed prospectively from non-diagnosed to diagnosed status, as well as very close individual matching with participants who would remain non-diagnosed. Although many of the non-diagnosed participants had some motor signs and were likely approaching HD diagnosis (as reflected by confidence ratings of 3 on the HD Diagnostic Rating Scale in the non-diagnosed group at the follow-up “verification” visit), they continued to be considered to not meet criteria for motor onset at least one year following the MRI that was used for the current data analysis. Since diagnosis is based on motor signs, it is not surprising that initial motor scores were higher in participants later given the diagnosis, and we did not attempt to incorporate both motor and imaging variables in the analysis.

An unavoidable weakness of the study is the potential lack of consistency among raters regarding the criteria they use to make a diagnosis. Because of the rarity of HD, it would not be possible to conduct this study at a single site, and it was not possible to individually match subjects from the same sites. Because criteria for making a diagnosis are subjective (the rater must have $\geq 99\%$ certainty that the participant shows “unequivocal presence of an otherwise unexplained extrapyramidal movement disorder”), some variability is expected in raters’ judgment of the severity of motor signs necessary to warrant a diagnosis. More robust group differences were observed for the reanalysis involving groupings based on UHDRS-Motor score cutoffs, whereby cases were omitted when the diagnosis was based on a relatively low UHDRS-Motor score (10 or less) and (more often), when a diagnosis had not been made despite a relatively high score (11 or greater). Of course, the assignment of the UHDRS-Motor score, like diagnosis, involves some subjective judgment by the clinician, and necessarily influences the decision to make a diagnosis. It must be noted as well that our cut-off of 10.5 for the UHDRS-Motor groupings is based only on data from this sample and should not be considered as a recommendation for a diagnostic threshold. Because it was not

possible to recruit the matched pairs from the same site, another limitation of the study was that the members of each pair were not scanned on the same scanner.

Results of this study suggest that structural MRI volumes, especially of putamen, can be useful in identifying those individuals who can be expected to be diagnosed within a relatively short timeframe. However, because of the lack of 100% predictive validity, it is not recommended that these measures be used clinically to predict for individual patients when diagnosable symptoms will occur. These measures could, however, be quite useful in selecting participants for clinical trials in which diagnosis of HD is a primary outcome measure. In planning for multisite clinical trials in pre-HD, investigators may also want to consider using more objective measures, such as MRI volumes, in addition to onset of diagnosable movement disorder, as a major outcome measure.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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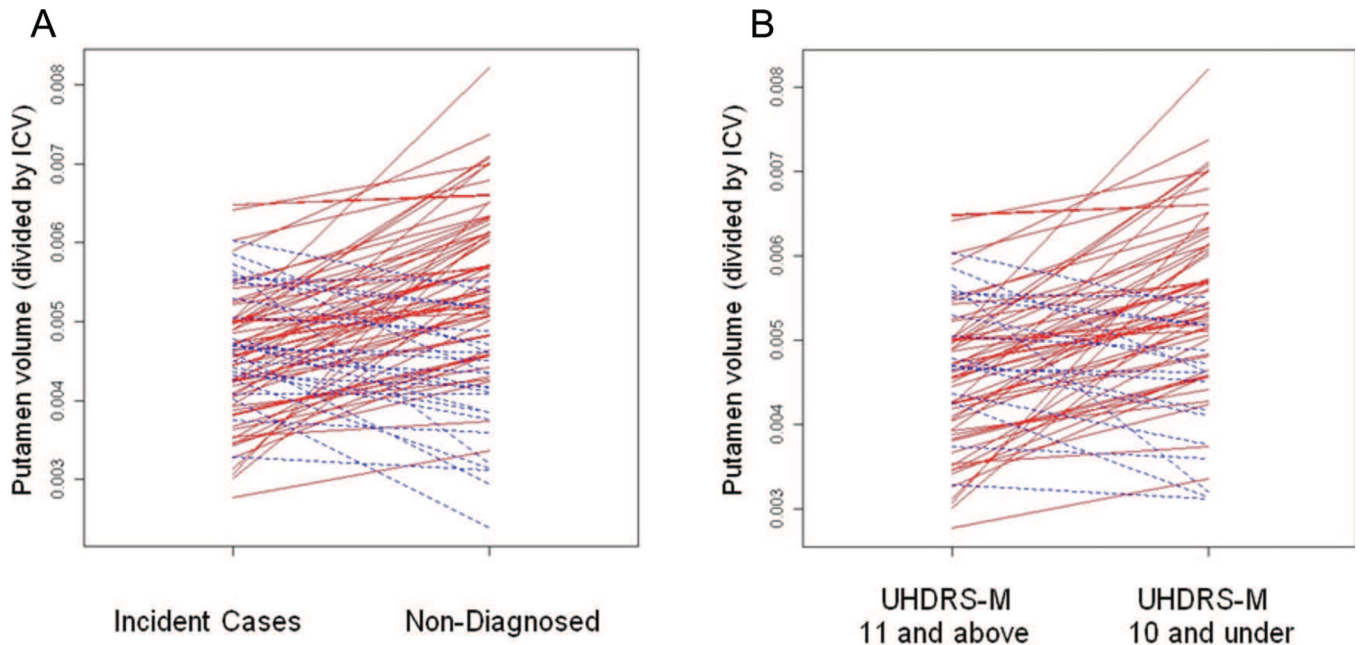


Figure 1.

Lines connect corrected putamen volumes (putamen divided by intracranial volume) of matched pairs: (A) Volume for incident case on the left connected to volume for matched non-diagnosed pair member on the right ($N = 85$ pairs); (B) Volume for member with higher UHDRS-Motor score (11 and above) on the left connected to volume for matched pair member with the lower score on the right (10 and under). Red lines indicate pairs for which the incident case (or member with the higher Motor score) had a smaller caudate volume than the matched non-diagnosed participant (or member with the lower Motor score); blue lines indicate pairs for which the reverse was true.

Table 1

Demographic Information and Structural Volumes for Incident Cases and Non-diagnosed Participants (Based on Clinician Judgment)

	Incident	Non-diagnosed	Statistic* (p-value)
N	85	85	
Gender	61 F 24 M	61 F 24 M	
Mean CAG repeat length (s.d.)	43.1 (2.4)	43.1 (2.3)	1.0 ($p = .32$)
Mean Age at MRI (s.d.)	44.9 (9.2)	44.9 (9.1)	0.98 ($p = .33$)
Years between MRI visit and "verification" visit**	1.75 (0.77)	2.46 (1.40)	
Mean UHDRS Motor Score at MRI Visit (s.d.)	11.4 (6.6)	4.9 (5.5)	7.00 ($p < .0001$)
Caudate Volume			
Raw (cc) \pm s.d.	4.61 (1.38)	5.17(1.28)	
Corrected*** \pm s.d.	.35 (.10)	.40 (.10)	5.39 ($p < .0001$)
Putamen Volume			
Raw (cc) \pm s.d.	6.01 (1.24)	5.17(1.28)	
Corrected*** \pm s.d.	.46 (.08)	.40 (.10)	3.55 ($p = .0006$)
Total Striatal Volume			
Raw (cc) \pm s.d.	10.63 (2.47)	11.95 (2.58)	
Corrected*** \pm s.d.	.81 (.17)	.92 (.20)	5.01 ($p < .0001$)
White Matter Volume			
Raw (cc) \pm s.d.	342.39 (61.29)	357.86 (63.48)	
Corrected*** \pm s.d	26.06 (3.31)	27.42 (3.57)	2.89 ($p = .0049$)

* matched sample *t*-tests, based on ICV-corrected volumes (region divided by ICV) for structural measures.

** the verification visit, which occurred 1 to 4 years after the MRI scan, is when a diagnosis of HD was made for the incident cases and the most recent visit at which the non-diagnosed participants were verified to still be free of diagnosable signs and symptoms

*** Corrected volumes = (structure volume/intracranial volume) * 100

Table 2

Demographic Information and Structural Volumes for Participants Categorized by UHDRS Motor Score

	UHDRS Motor Score 11 and over	UHDRS Motor Score 10 and under	Statistic* (<i>p</i> -value)
N	72	72	
Gender	49 F 23 M	49 F 23 M	
Mean CAG repeat length (s.d.)	43.1 (2.4)	43.0 (2.2)	.70 (<i>p</i> = .48)
Mean Age at MRI (s.d.)	44.4 (8.7)	44.4 (8.6)	.77 (<i>p</i> = .44)
Years between MRI visit and "verification" visit**	1.7 (.73)	2.5 (1.4)	
Mean UHDRS Motor Score at MRI Visit (s.d.)	11.8 (6.4)	3.6 (3.5)	9.49 (<i>p</i> < .0001)
Caudate Volume			
Raw (cc) ± s.d.	4.7 (1.4)	5.3 (1.3)	
Corrected*** ± s.d.	.35 (.10)	.41 (.10)	3.71 (<i>p</i> = .0004)
Putamen Volume			
Raw (cc) ± s.d.	6.1 (1.3)	7.0 (1.4)	
Corrected*** ± s.d.	.46 (.09)	.54 (.11)	6.25 (<i>p</i> < .0001)
Total Striatal Volume			
Raw (cc) ± s.d.	10.8 (2.6)	12.3 (2.5)	
Corrected*** ± s.d.	.81 (.17)	.94 (.19)	5.59 (<i>p</i> < .0001)
White Matter Volume			
Raw (cc) ± s.d.	348.3 (62.4)	364.3 (63.2)	
Corrected*** ± s.d.	26.10 (3.2)	27.87 (3.3)	3.48 (<i>p</i> = .0009)

* matched sample *t*-tests, based on ICV-corrected volumes (region divided by ICV) for structural measures.

** the verification visit, which occurred 1 to 4 years after the MRI scan, is when a diagnosis of HD was made for the cases with higher UHDRS Motor Scores and the most recent visit at which those with lower UHDRS Motor Scores were verified to still be free of diagnosable signs and symptoms

*** Corrected volumes = (structure volume/intracranial volume) * 100